#### ORIGINAL RESEARCH



# Cefazolin Versus Anti-Staphylococcal Penicillins for the Treatment of Patients with Methicillin-Susceptible *Staphylococcus aureus* Infection: A Meta-Analysis with Trial Sequential Analysis

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### ABSTRACT

*Introduction*: Methicillin-susceptible *Staphylococcus aureus* (MSSA) is a common cause of infection in humans. Beta-lactam antibiotics are the preferred agents, with anti-staphylococcal penicillins (ASPs) or the first-generation cephalosporin, cefazolin, favored by clinicians. Recent studies comparing the two strategies suggest similar outcomes between the agents. The purpose of this meta-analysis was to explore

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differences between cefazolin and ASPs for the treatment of MSSA infections.

*Methods*: We performed a meta-analysis with trial sequential analysis (TSA) of observational or cohort studies using a random-effects model. Two blinded reviewers independently assessed studies for inclusion, risk of bias, and data extraction. The primary outcome was all-cause mortality. Secondary outcomes included clinical failure, infection recurrence, and antibiotic discontinuation due to adverse events. Subgroup analyses were conducted for the primary outcome by type of ASP, studies with a high percentage of deep-seated infections, and studies of low to moderate risk of bias.

**Results**: After performing a comprehensive search of the literature, and screening for study inclusion, 19 studies (13,390 patients) were included in the final meta-analysis. Fifteen of the 19 studies (79%) were judged as having a

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A. F. Pomputius Biomedical and Health Information Services, University of Florida Health Science Center Libraries, Gainesville, FL, USA low or moderate risk of bias. Use of cefazolin was associated with lower all-cause mortality [odds ratio (OR) 0.71, 95% confidence interval (CI) 0.56–0.91, p = 0.006,  $I^2 = 28\%$ ], clinical failure (OR 0.55, 95% CI 0.41–0.74, p < 0.001,  $I^2 = 0\%$ ), and antibiotic discontinuation due to adverse events (OR 0.25, 95% CI 0.16–0.39, p < 0.001,  $I^2 = 23\%$ ). Infection recurrence was higher in the cefazolin patients (OR 1.41, 95% CI 1.04–1.93, p = 0.03,  $I^2 = 0\%$ ).

*Conclusion*: This meta-analysis demonstrated that the use of cefazolin was associated with significant reductions in all-cause mortality, clinical failure, and discontinuation due to adverse events, but was associated with an increased risk of infection recurrence.

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**Keywords:** Bacteremia; Cefazolin; Infection; Meta-analysis; Penicillins; *Staphylococcus aureus* 

## **INTRODUCTION**

Staphylococcus aureus (S. aureus) is a gram-positive bacterium that is a part of normal human skin flora. S. aureus is the causative pathogen for a wide array of clinically significant infections, ranging from skin and soft tissue infections to more severe infections including endocarditis, bacteremia, pneumonia, and osteomyelitis. S. aureus infections are typically broken down into two categories based on their susceptibility methicillin-susceptible profile, (MSSA) or methicillin-resistant (MRSA), with MRSA being associated with higher mortality rates [1] and fewer treatment options. Despite these data, MSSA is still a clinical concern as it still accounts for > 50% of all S. aureus bacteremias, with reported mortality rates of 15-30%, depending on patient comorbidities [2]. Selection of an appropriate treatment regimen is an optimal component of ensuring optimal outcomes in patients with MSSA infections.

Historically, beta-lactams have been the preferred agents for the treatment of MSSA

infections, owing to being more efficacious than other agents such as vancomvcin. reducing overall mortality, and preventing infection recurrence [3]. Within the beta-lactam drug class, anti-staphylococcal penicillins (ASPs) such as nafcillin and oxacillin are often utilized as first-line agents to treat MSSA infections. Concerns with the frequency of dosing, relatively high acquisition cost, and patient tolerance associated with ASPs have led to an interest in suitable alternatives. Cefazolin is a first-generation cephalosporin that has in vitro activity against MSSA and has a favorable pharmacokinetic profile allowing for less frequent dosing compared with ASPs. However, this enthusiasm is tempered by in vitro data that suggest that an inoculum effect can occur, resulting in higher rates of antibiotic failure compared with ASPs for the treatment of MSSA infections with high bacterial inoculum [4]. However, despite these in vitro data, the impact of the inoculum effect on clinically relevant outcomes is controversial [5]. Additionally, the data are unclear on whether there is a difference between ASPs and cefazolin for the treatment of MSSA infections. The purpose of this metaanalysis is to determine which agent is more effective at reducing mortality and infection recurrence. Additionally, the study will evaluate whether there is any difference related to adverse events between the two treatment strategies.

## METHODS

This meta-analysis was performed based on the methodology recommended by the Cochrane Collaboration and adhered to recommendations according to the Preferred Reporting Items for Systemic Reviews and Meta-analyses statement [6]. The study protocol was submitted and registered with the PROSPERO International Prospective Register of Systematic Reviews (study ID: CRD42018106442) before the start of the literature search. No specific funding sources were utilized to complete this study. This article is based on previously conducted studies and does not contain any studies with human participants or animals performed by any of the authors. Our hypothesis was that, compared with ASPs, cefazolin would be associated with reduced mortality and have superior safety outcomes [7].

#### **Eligibility Criteria**

We sought to include all randomized controlled trials utilizing a parallel study design, or observational studies, including cohort and casecontrol studies comparing the use of ASPs with cefazolin. We sought to include only studies in which patients received MSSA treatment with cefazolin or ASP for at least 72 h. No restrictions were imposed based on publication status. We excluded studies conducted in pediatric patients, studies that did not report mortality as an outcome, studies that used antibiotics other than cefazolin or ASPs, studies that included bacteria other than MSSA, and animal and in vitro studies.

#### Search Strategy

A comprehensive population, intervention, comparator, and outcome question was created using applicable medical subject headings (MeSH) and individual substances to answer the question, "Does cefazolin reduce all-cause mortality compared with antistaphylococcal penicillins for the management of MSSA infections?" The initial search was conducted by a medical librarian (AFP) across PubMed/Medline, Web of Science, BIOSIS, and Cochrane CENTRAL Register of Controlled Trials. Gray literature was searched using ClinicalTrials.gov, Agency for Healthcare Research and Quality (AHRQ), and infectious diseases professional organizations. In each database, search strategies used controlled vocabulary, if applicable, as well as a variation of keywords related to cefazolin, staphylococcal infections, and antistaphylococcal penicillin. Additionally, relevant conference proceedings and abstracts identified in the search were included if no resulting peerreviewed article was found. A secondary search of Google Scholar was also performed to identify other relevant articles. All references, recent systematic reviews, and guidelines were handsearched for trials and citations not included from the initial search. Search results were not limited by language or date. The literature search was initially conducted in September 2018 and re-run in June 2019. Full search strategies are available in the supplementary appendix.

#### **Study Selection**

Two investigators (JMA, AMC) independently screened each article for eligibility and inclusion using an electronic spreadsheet (Microsoft Corp., Redmond, WA). Disagreements were resolved by adjudication between the investigators until consensus was achieved.

#### **Data Extraction**

Electronic data collection forms were developed to assess primary and secondary outcomes. Data were also collected on inclusion/exclusion criteria for each study, patient demographics (i.e., age, gender, deep-seated infection, ICU admission at baseline, history of IV drug abuse, baseline need for hemodialysis), year of publication, and type of intervention (drug, dose, duration). Deep-seated infections were defined as osteomyelitis, endocarditis, or pneumonia. Using an electronic data extraction tool, two investigators (JMA, LB) independently extracted data from each study. Disagreements were resolved by adjudication until consensus was achieved.

#### **Risk of Bias Assessment**

Individual risk of bias at the study level was assessed independently by two investigators (JMA, LB) using the Risk of Bias in Non-randomized Studies of Interventions (ROBINS-I) for non-randomized studies [8]. Disagreements were resolved by adjudication between the investigators until consensus was achieved. A priori sensitivity analyses were performed to test the effect of only including studies with low or moderate risk of bias.

#### **Statistical Analysis**

The primary study outcome was all-cause mortality, reported as odds ratio (OR), and 95% confidence intervals (95% CI). Secondary outcomes included clinical failure, infection recurrence, and antibiotic discontinuation due to adverse events. Subgroup analyses were conducted for the primary outcome by type of ASP used, studies of low to moderate risk of bias, studies utilizing at least 4 g/day of cefazolin, and studies with a high percentage of deepseated infections. Studies with a high percentage of deep-seated infections were defined as having > 25% of cefazolin patients with deepseated infections (i.e., pneumonia, osteomyelitis, and endocarditis) [9]. Due to the anticipated variability in patient populations, risk factors, and interventions across included studies, all outcomes were analyzed using the Mantel-Haenszel random-effects model. Heterogeneity among trials was quantified by using the inconsistency factor  $(I^2)$ . All p values  $\leq 0.05$ were considered statistically significant. Statistical analysis was conducted using the Cochrane Review Manager v.5.3 (Cochrane, London, UK) and the Trial Sequential Analysis (TSA) software published by the Copenhagen Trial Unit, Centre for Clinical Intervention Research (available at http://www.ctu.dk/tsa/ index.html). TSA is a methodology that combines an information size calculation (i.e., cumulated sample sizes of all included trials) for a meta-analysis with the threshold of statistical significance. A sensitivity analysis using TSA was performed to confirm the findings of the meta-analysis. Furthermore, before the required information size is reached, TSA constructs monitoring boundaries to determine when an estimated effect is so convincingly large (or small) that the conclusions are unlikely to change with more evidence, which is analogous to an interim analysis of a single randomized controlled trial. The use of TSA to validate metaanalysis results has previously been described [10]. An additional sensitivity analysis to assess the robustness of the results compared with unmeasured confounding using the E value methodology was performed [11, 12]. The E value estimates what the relative risk would have to be for any unmeasured confounder to overcome the observed association of cefazolin with mortality in this study.

## RESULTS

#### **Study Characteristics**

The initial study search yielded 2211 studies. After de-duplication, 979 studies were included in the initial screening process. After a review of titles and abstracts, and subsequent full-text review, 961 studies were excluded. A subsequent study search yielded an additional study that was not included in the original search, leaving 19 studies in the final analysis (13,390 patients) [13–31]. There was one prospective, observational study included in the final analysis, while the remaining studies were retrospective study designs. The study inclusion flow diagram is included in the study supplemental information. Among the included studies, four had a serious risk of bias (21%), seven had a moderate risk of bias (37%), and eight had a low risk of bias (42%). The risk of bias for each included study is described in the supplemental study information. Of the included studies. 12 utilized nafcillin as the comparator ASP. Other ASPs included in the analysis were cloxacillin (n = 3), oxacillin (n = 2), floxacillin (n = 1), and nafcillin/oxacillin (n = 1). Nineteen of the studies included only patients with positive blood cultures for MSSA. The most-reported follow-up period for mortality was 30 day, with 11 studies reporting this as a study outcome. Other periods reported include 90-day (n = 4), in hospital (n = 3), and 28-day (n = 1). Among all studies in which specific sources of infection were reported, 3343/12,649 (26.4%) patients a deep-seated infection (cefazolin, had 660/2819, 23.4%; ASPs, 2683/9830, 27.3%). A similar number of patients in both groups required ICU admission at study baseline (cefazolin, 395/2454, 16.1%; ASPs, 1407/8902, 15.8%). Patients with a history of IV drug abuse were similar between cefazolin and ASPs (cefazolin, 158/1472, 10.7%; ASPs, 288/2336, 12.3%). Other specific study characteristics are summarized in Table 1.

Study name (n =)	Study design, country	Baseline severity of illness	Baseline ICU admission 12, (%)	Renal dysfunction, n, (%)	IVDA n, (%)	Deep-seated infection <i>n</i> , (%)	Mortality n, (%)	Clinical failure Infection n, (%) recurrenc (%)	Infection recurrence <i>n</i> , (%)	Antibiotic discontinuation n, (%)	Risk of bias
Algrim (2017) [13] (n = 176)	Retrospective cohort; US	NR	NR	NR	NR	NR	In hospital Cefi 6 (6.8) Naf: 11 (12.5)	In hospital Cef. 11 (12.5) Naf: 17 (19.3)	NR	NR	High
Bai (2015) [14] (n = 354)	Retrospective cohort, Canada	NR	Cef. 10 (10) Clox: 45 (18)	CKD Cef: 29 (28) Clox: 53 (21) Hemodialysis at baseline Cef: 8 (8) Clox: 23 (9)	ХХ	Endocarditis Cef: 2 (2) Clox: 30 (12)* Osteomyelitis Cef: 15 (14) Clox: 28 (11) Pneumonia Cef: 17 (16) Clox: 44 (18)	90 day Cefi 21 (20) Clox: 75 (30)*	NR	90 day Cef: 6 (6) Clox: 4 (2)	NR	Low
Burrelli (2019) [15] $(n = 157)$	Retrospective cohort; US	APACHE II, median (IQR) Cef: 17 (10–20) Naf: 10 (6–14)*	Cef: 15 (36.6) Naf: 30 (25.9)	Renal dysfunction Cef: 20 (48.8) Naf: 12 (10.3)* Hemodialysis prior to admission Cef: 15 (36.6) Naf: 5 (4.3)*	Cefi 4 (9.8) Naf: 11 (9.5)	Endocarditis Cef: 4 (9.8) Naf: 14 (12.1) Osteomyelitis Cef: 8 (19.5) Naf: 31 (26.7) Pneumonia Cef: 1 (2.4) Naf: 1 (1.7)	In hospital Cefi 0 Nafi 4 (3.4)	Clinician documented failure Cef: 13 (31.7) Naf: 33 (28.4)	Microbiologic recurrence Cef: 0 Naf: 2 (1.7)	ЯК	роМ

Study name (n =)	Study design, country	Baseline severity of illness	Baseline ICU admission <i>n</i> , (%)	Renal dysfunction, n, (%)	IVDA n, (%)	Deep-seated infection <i>n</i> , (%)	Mortality n, (%)	Clinical failure Infection 1, (%) recurrenc (%)	Infection recurrence <i>n</i> , (%)	Antibiotic discontinuation <i>n</i> , (%)	Risk of bias
Davis (2018) [16] (n = 7312)	Retrospective cohort, Australia	NR	Cef: 110 (13.9) Flox: 878 (13.5)	Cef. 149 (20.8) Flox: 519 (8.9)	NR	Endocarditis Cef: 47 (5.9) Flox: 524 (8.0)* Osteomyelitis Cef: 128 (16.2) Flox: 1195 (18.3)	30 day Cefi 83 (10.5) Flox: 731 (11.2) 7 day Cefi 29 (3.7) Flox: 294 (4.5)	NR	NR	NR	PoM
[17] $[17]$ $(n = 149)$	Retrospective cohort; US	Pitt bacteremia > 4, <i>n</i> (%) Cef: 7 (9.4) Naf: 13 (15.9)	NN	NR	Cefi 7 (9.3) Naf: 16 (19.5)	Endocarditis Cef: 11 (16.2) Naf: 22 (27.2) Osteomyelitis Cef: 13 (19.1) Naf: 20 (24.7)	30 day Cefi: 4 (6) Nafi: 4 (5)	30 day Cefi 1 (1.5) Nafi 2 (2.5)	NN	NR	роМ
Lee $(2011)$ [18] (n = 82)	Retrospective case- control; South Korea	McCabe fatal classification, n (%) Cef: 27 (66) Naf: 30 (73)	NR	End-stage renal disease Cef: 2 (5) Naf: 6 (15)	NR	Endocarditis Cefi 1 (2) Nafi 1 (2) Osteomyelitis Cefi 10 (24) Nafi $7$ (17) Pneumonia Cefi 3 (7) Nafi 4 (10)	4 week Cefi 2 (4) Nafi 2 (4) 12-week SAB- related Cefi 1 (2) Nafi 5 (12)	12 week Cefi 6 (15) Nafi 6 (15)	12 week Cef: 1 (2) Naf: 1 (2)	Cef. 0 Naf: 7 (17)*	Low

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Study name (n =)	Study design, country	Baseline severity of illness	Baseline ICU admission <i>n</i> , (%)	Renal dysfunction, n, (%)	IVDA n, (%)	Deep-seated infection <i>n</i> , (%)	Mortality n, (%)	Clinical failure Infection n, (%) recurrenc (%)	Infection recurrence <i>n</i> , (%)	Antibiotic discontinuation n, (%)	Risk of bias
Lee $(2018)$ [19] (n = 242)	Prospective, observational cohort; South Korea	SOFA = 2, n (%) Cef: 36 (45.6) Naf: 111 (68.1)*	NR	CKD, Stage 4–5 Cef: 14 (17.7) Naf: 32 (19.6)	NR	Endocarditis Cef: 4 (5.1) Naf: 24 (14.7)* Ostcomyclitis Cef: 28 (35.4) Naf: 61 (37.4) Pneumonia Cef: 3 (3.8) Naf: 16 (9.8)	30 day Cefi 2 (2.5) Nafi 13 (8%) 90 day Cefi 2 (2.5) Nafi 24 (14.7) *	30 day Cef: 24 (30.4) Naf: 82 (50.3)	Cefi 2 (2.5) Nafi 6 (3.7)	Cef. 10 (12.7) Naf. 48 (29.4)*	Low
Li $(20]$ (n = 93)	Retrospective cohort; US	Pitt Bacteremia score, median (IQR) Cefi 0 (0–1) Oxa: 0 (0–1)	Cefi 4 (7) Oxa: 6 (18)	End-stage renal disease Cef: 15 (25) Oxa: 0*	Cefi 19 (32) (32) (30)	Endocarditis Cef: 14 (25) Oxa: 3 (12) Osteomyelitis Cef: 18 (31) Oxa: 20 (59) * Pneumonia Cef: 2 (4) Oxa: 2 (6)	30 day Cef: 0 Oxa: 1 (3) 90 day Cef: 0 Oxa 1 (3)	90 day Cef: 14 (24) Oxa: 16 (47)	90 day Cefi 4 (6.8) Oxa: 3 (8.8)	Cef. 2 (3) Oxa: 7 (21)	роМ

Study name (n =)	Study design, country	Baseline severity of illness	Baseline ICU admission 2,	Renal dysfunction, n, (%)	IVDA n, (%)	Deep-seated infection 11, (%)	Mortality n, (%)	Clinical failure n, (%)	Infection recurrence <i>n</i> , (%)	Antibiotic discontinuation n, (%)	Risk of bias
Mcdanel (2017) [21] (n = 3167)	Retrospective cohort; US	APACHE III ≥ 34, n (%) Cef: 651 (56) Naf/Oxa: 1040 (52)	(19)* Cefi 178 (15) Naf/Oxa: 378 (19)*	Dialysis at baseline Cef: 176 (15) Naf/Oxa: 168 (8)*	Cef: 110 (9) Naf/ Oxa: (11) (11)	Endocarditis Cef: 52 (4) Naf/Oxa: 145 (7) Osteomyelitis Cef: 138 (12) Naf/Oxa: 267 (13)	30 day Cef: 113 (10) Naf/Oxa: 307 (15)* 90 day Cef: 231 (20) Naf/Oxa: 502	X	45–365 day Cef: 58 (5) Naf/Oxa: 72 (3.6)	N	poM
Miller (2018) [22] (n = 90)	Retrospective cohort; US	Charlson comorbidity index, median (IQR) Cef: 6 (3–8) Naf: 3 (0–6)*	Cefi 18 (36) Nafi 20 (50)	End-stage renal disease Cef: 17 (34) Naf: 2 (5)*	NR	Endocarditis Cef: 8 (16) Naf: 8 (20) Osteomyelitis Cef: 13 (26) Naf: 12 (30) Pneumonia Cef: 3 (6) Naf: 2 (5)	(10) 30 day Cef: 3 (6) Naf: 4 (10)	30 day Cef: 3 (6) Naf: 7 (17)	ХX	Cef: 16/51 (31.4) Naf: 41/79 (51.9)*	Low
Monogue (2018) [23] (n = 142)	Retrospective cohort; US	Pitt bacteremia score, mean (SD) Cef: 0.76 (0.89) Naf: 1.1 (1.7)	Cef: 8 (11.4) Naf: 23 (32.4)*	End-stage renal disease Cefi 22 (31) Naf: 10 (14.1)	Cefi 3 (4.2) Nafi 11 (15.5)*	Endocarditis Cef: 3 (4.2) Naf: 6 (7) Osteomyelitis Cef: 6 (8.5) Naf: 10 (14.1) Pneumonia Cef: 3 (4.2) Naf: 4 (5.6)	30 day Cefi 0 Nafi 3 (4.2)	90 day Cefi 6 (8.5) Nafi 8 (11.3)	90 day Cefi 4 (5.6) Nafi 1 (1.4)	Cef. 0 Nafi 2 (2.8)	PoM

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Study name (n =)	Study design, country	Baseline severity of illness	Baseline ICU admission 12, (%)	Renal dysfunction, n, (%)	IVDA n, (%)	Deep-seated infection <i>n</i> , (%)	Mortality n, (%)	Clinical failure n, (%)	Infection recurrence <i>n</i> , (%)	Antibiotic discontinuation n, (%)	Risk of bias
Paul (2011) [24] (n = 353)	Retrospective cohort; Israel	NR	NR	NR	NR	NR	90 day Cef: 29 (40.3) Clox: 91 (32.4)	NR	NR	NR	High
Phe (2015) [25] (n = 78)	Retrospective cohort; US	NR	NR	NR	NR	NR	30 day Cef: 0 Naf: 2 (7.7)*	30 day Cef: 8 (15.4) Naf: 8 (30.8)	NR	NR	High
Pollett (2016 [26] (n = 100)	Retrospective cohort; US	Charlson comorbidity index, median (IQR) Cefi 4 (2–7) Nafi 3 (1–7)	Cef. 9 (13) Naf: 8 (27)	Dialysis at baseline Cefi 43 (61) Naf: 16 (53)	Cefi 12 (17) Nafi 8 (27)	Endocarditis Cef: 10 (14) Naf: 5 (17) Osteomyelitis Cef: 9 (13) Naf: 5 (17) Pneumonia Cef: 1 (1) Naf: 4 (13)	90 day Cefi 5 (7.1) Nafi 5 (16.7)	NR	X	XR	Low
Rao (2015) [27] (n = 161)	Retrospective cohort; US	Modified APACHE II, mean (SD) Cef: 13 (6.3) Oxa: 10.3 (5.8)*	Cef: 43 (41.8) Oxa: 19 (32.8)	Baseline or acute renal dysfunction Cefi 51 (49.5) Oxa: 17 (29.3)*	NR	Endocarditis Cef: 17 (16.5) Oxa: 12 (20.7) Oxteomyelitis Cef: 15 (14.6) Oxy: 8 (13.8)	In hospital Cefi 1 (1) Oxa: 3 (5.2)	In hospital Cef: 6 (5.8) Oxa: 7 (12.1)	In hospital Cef: 5 (4.9) Oxa: 3 (5.2)	Cef. 21 (20.4) Naf: 25 (43.1)*	Low

Study name $(n =)$	Study design, country	Baseline severity of illness	Baseline ICU admission <i>n</i> , (%)	Renal dysfunction, n, (%)	IVDA n, (%)	Deep-seated infection <i>n</i> , (%)	Mortality n, (%)	Clinical failure Infection n, (%) recurrenc (%)	Infection recurrence <i>n</i> , (%)	Antibiotic discontinuation n, (%)	Risk of bias
Renaud (2011) [28] (n = 27)	Retrospective cohort; Singapore	Charlson comorbidity age-adjusted score, mean (SD) Cef: 5.4 (1.9) Clox: 5.8 (3.1)	XX	Baseline hemodialysis Cef: 14 (100) Clox: 13 (100)	NR	Endocarditis Cef: 0 Clox: 1 (7.7) Osteomyelitis Cef: 1 (7.1) Clox: 1 (7.7)	30 day Cef: 1 (7.1) Clox: 2 (15.4)	NR	NR	NR	High
Schweizer (2011) [29] (n = 142)	Retrospective cohort; US	NR	NR	NR	NR	NR	In hospital Cef: 5 (8.5) Naf: 6 (7.2)	NR	NR	NR	Low
Shah (2018) [30] (n = 80)	Retrospective cohort; US	NR	NR	NR	NR	NR	30 day Cef: 1 (3) Naf: 6 (13)	30 day Cefi 3 (9) Nafi 6 (13)	90 day Cefi 1 (3) Naf: 0	Cef: 0 Naf: 11 (48)*	High
Youngster (2014) [31] (n = 485)	Retrospective cohort; US	NR	Cefi 0 Nafi 0	Chronic renal disease Cef: 12 (10) Naf: 29 (7.1)	NR	Endocarditis Cef: 7 (5.8) Naf: 28 (7.6) Osteonyelitis Cef: 39 (32.7) Naf: 112 (30.6)	Overall Cef: 1 (0.8) Naf: 4 (1.1)	X		Cef: 8 (6.7) Naf: 124 (33.9)*	Low

#### **Primary Outcome**

When all studies were included, the all-cause mortality was 12.8% (394/3081) in the cefazolin group and 14.4% (1483/10309) in the ASP group. The use of cefazolin was associated with lower all-cause mortality compared with ASPs (OR 0.71; 95% CI 0.56–0.91; p = 0.006;  $I^2 = 28\%$ ). Among a priori subgroup analysis, cefazolin was also associated with lower allcause mortality compared with ASPs in studies with low/moderate risk of bias (OR 0.75; 95% CI 0.59–0.96; p = 0.02;  $I^2 = 32\%$ ), studies that used nafcillin (OR 0.46; 95% CI 0.29–0.74; *p* = 0.001;  $I^2 = 0\%$ ), and studies with a high percentage of deep-seated infections (OR 0.48; 95% CI 0.32–0.73; p = 0.0005;  $I^2 = 0\%$ ). In studies that utilized at least 4 g/day of cefazolin, there were no differences observed in all-cause mortality between cefazolin and ASPs (OR 0.55; 95% CI 0.24–1.24; p = 0.15;  $I^2 = 0\%$ ). The results were unaffected when sensitivity analysis was performed evaluating the impact of the two largest included trials. Figure 1 summarizes all-cause mortality for all included studies and multiple a priori subgroup analyses. Sensitivity analysis of the primary outcome was also performed based on studies with low risk of bias (Supplement). To assess the validity of the meta-analysis results, TSA was performed for all-cause mortality, using various power and zero-event handling assumptions. TSA confirmed the meta-analysis results for most assumed event rates, powering, and zero-event analyses (Supplement). The E value (relative risk) for the point estimate for all-cause mortality was 2.17.

#### **Secondary Outcomes**

Eleven studies reported results on clinical failure with cefazolin use being associated with less clinical failure compared with ASPs (OR 0.55; 95% CI 0.41–0.71; p < 0.0001;  $I^2 = 0\%$ ) (Supplement). Upon completion of TSA, significance was maintained (OR 0.55; 95% CI 0.32–0.97). Nine studies had extractable data on infection recurrence, with cefazolin-treated patients having high risk of infection recurrence compared with ASPs (OR 1.41; 95% CI 1.04–1.93; p = 0.03;  $I^2 = 0\%$ ) (Supplement), and TSA confirmed these results (OR 1.41; 95% CI 1.02–1.96). Eight studies reported on antibiotic discontinuation due to adverse events occurring less frequently in cefazolin-treated patients compared with ASPs (OR 0.25; 95% CI 0.16–0.39; p < 0.001;  $I^2 = 23\%$ ) (Supplement). TSA confirmed the meta-analysis results (OR 0.26; 95% CI 0.11–0.62).

### DISCUSSION

Our study results suggest that cefazolin is associated with a small benefit in all-cause mortality compared with ASPs. Additionally, in our a priori subgroup analysis of all-cause mortality, cefazolin continued to be associated with less mortality across multiple subgroups, including studies with low/moderate risk of bias, studies that used nafcillin as the ASP comparator, and studies with high rates of deep-seated infections. One particular subgroup analysis that warrants additional discussion is deep-seated infections. Deep-seated infections such as endocarditis, osteomyelitis, and pneumonia can be notoriously difficult to treat, often require prolonged courses of therapy, and are associated with a higher degree of treatment failure. Previous literature questioned the utility of cefazolin for deep-seated MSSA infections because of an inoculum effect [32]. In our a priori subgroup analysis of studies with a high percentage of deep-seated infections, cefazolin outperformed ASPs in all-cause mortality, suggesting that it could be a reasonable therapeutic option in this patient population. TSA also confirmed the meta-analysis result for all-cause mortality across multiple power and zero-constant assumptions. Cefazolin also was associated with reduced risk of clinical failure and antibiotic discontinuation due to adverse drug events compared with ASPs. The substantially lower risk of clinical failure and antibiotic discontinuation due to adverse drug events with cefazolin compared with ASPs indicated that patients with MSSA infections were more likely to complete therapy on cefazolin than to switch to another agent. Combined with the observed mortality benefit, these findings appear to imply that cefazolin may be the

	Cefazolir Events T	otal	Anti-staphylococo Events		Weight	Odds Ratio M-H, Random, 95% Cl	Odds Ratio M-H, Random, 95% Cl
.1.1 All studies- All-ca							
ligrim, 2017	6	88 105	11	88	4.5%	0.51 [0.18, 1.45]	
3ai 2015 Burrelli 2018	21 0	41	75	249 116	11.8% 0.6%	0.58 [0.33, 1.00] 0.30 [0.02, 5.72]	
Davis 2018		792	731	6520	23.7%	0.93 [0.73, 1.18]	-
lynt 2017	4	68	4	81	2.6%	1.20 [0.29, 5.00]	
.ee 2011	1	41	5	41	1.1%	0.18 [0.02, 1.61]	
.ee, 2018	2	79	24	163	2.4%	0.15 [0.03, 0.65]	
.i 2014	0	59	1	34	0.5%	0.19 [0.01, 4.74]	•
IcDanel 2017		163 50	502 4	2004	26.6%	0.74 [0.62, 0.88]	
filler, 2018 fonogue 2018	3 0	71	4	40 71	2.2% 0.6%	0.57 [0.12, 2.73] 0.14 [0.01, 2.70]	• • • • • • • • • • • • • • • • • • • •
aul 2011	29	72	91	281	12.2%	1.41 [0.83, 2.40]	
he, 2015	0	52	2	26	0.6%	0.09 [0.00, 2.02]	· · · · · · · · · · · · · · · · · · ·
ollett 2016	5	70	5	30	3.0%	0.38 [0.10, 1.44]	
Rao 2015		103	3	58	1.1%	0.18 [0.02, 1.77]	
Renaud 2011	1	14	2	13	0.9%	0.42 [0.03, 5.32]	
Schweizer, 2011	5	59	6	83	3.3%	1.19 [0.34, 4.09]	
Shah, 2018 Youngster 2014	1 1	35 119	6 4	45 366	1.2% 1.1%	0.19 [0.02, 1.67] 0.77 [0.08, 6.93]	
Subtotal (95% CI)		081	4	10309	100.0%	0.71 [0.56, 0.91]	•
otal events	394		1483	10000	1001010		•
leterogeneity: Tau <sup>2</sup> = 0.		= 25.1		2): I <sup>2</sup> = 28%	6		
est for overall effect: Z				.,,,			
.1.2 Low/Moderate Ris			-				
3ai 2015		105	75	249	12.6%	0.58 [0.33, 1.00]	
Burrelli 2018	0	41	4	166	0.7%	0.44 [0.02, 8.24]	
Davis 2018		792	731	6520	25.7%	0.93 [0.73, 1.18]	
iynt 2017 .ee 2011	4 1	68 41	4	81 41	2.7%	1.20 [0.29, 5.00] 0.18 [0.02, 1.61]	
.ee, 2018	2	79	24	163	2.6%	0.15 [0.02, 1.61]	
.i 2014	0	59	1	34	0.6%	0.19 [0.01, 4.74]	·
IcDanel 2017	231 1	163	502	2004	29.0%	0.74 [0.62, 0.88]	+
filler, 2018	3	50	4	40	2.3%	0.57 [0.12, 2.73]	
lonogue 2018	0	71	3	71	0.7%	0.14 [0.01, 2.70]	· · · · · · · · · · · · · · · · · · ·
2011 Paul 2011	29	72	91	281	13.1%	1.41 [0.83, 2.40]	+ <b>-</b> -
Pollett 2016	5	70	5	30	3.1%	0.38 [0.10, 1.44]	
Rao 2015		103	3	58	1.1%	0.18 [0.02, 1.77]	
Schweizer, 2011 Youngster 2014	5 1	59 119	6 4	83 366	3.5% 1.2%	1.19 [0.34, 4.09] 0.77 [0.08, 6.93]	
Subtotal (95% CI)		892	+	10187	100.0%	0.75 [0.59, 0.96]	•
otal events	386		1462				
Heterogeneity: Tau <sup>2</sup> = 0.	05; Chi <sup>2</sup> =	= 20.5	5, df = 14 (P = 0.11	); I <sup>2</sup> = 32%	6		
est for overall effect: Z							
.1.3 Nafcillin- All-caus		-		0.0	20.00	0.54 10 40 4 17	
ligrim, 2017	6	88	11	88	20.6%	0.51 [0.18, 1.45]	
Burrelli 2018 Swot 2017	0 4	41	4	116	2.6%	0.30 [0.02, 5.72]	
lynt 2017 .ee 2011	4	68 41	4	81 41	11.0% 4.6%	1.20 [0.29, 5.00] 0.18 [0.02, 1.61]	
.ee, 2018	2	79	24	163	10.4%	0.15 [0.03, 0.65]	
filler, 2018	3	50	4	40	9.2%	0.57 [0.12, 2.73]	
Ionogue 2018	0	71	3	71	2.5%	0.14 [0.01, 2.70]	·
he, 2015	0	52	2	26	2.4%	0.09 [0.00, 2.02]	• • • • • •
Pollett 2016	5	70	5	30	12.8%	0.38 [0.10, 1.44]	
Schweizer, 2011	5	59	6	83	14.6%	1.19 [0.34, 4.09]	
Shah, 2018	1	35	6	45	4.8%	0.19 [0.02, 1.67]	
oungster 2014		119 773	4	366 1150	4.6% 100.0%	0.77 [0.08, 6.93]	
Subtotal (95% CI) Total events	28	115	78	1150	100.0%	0.46 [0.29, 0.74]	•
leterogeneity: Tau <sup>2</sup> = 0.		- 0 06		IZ - 0%			
est for overall effect: Z:				. = 0.0			
.1.4 High deep-seated							_
3ai 2015		105	75	249	54.6%	0.58 [0.33, 1.00]	
Burrelli 2018	0	41	4	116	1.9%	0.30 [0.02, 5.72]	· · · · ·
lynt 2017	4	68	4	81	8.1%	1.20 [0.29, 5.00]	
.ee 2011	1	41 79	5 24	41	3.4%	0.18 [0.02, 1.61]	
.ee, 2018 .i 2014	2 0	79 59	24	163 34	7.6% 1.6%	0.15 [0.03, 0.65] 0.19 [0.01, 4.74]	•
Ailler, 2018	3	59 50	4	40	6.8%	0.19 [0.01, 4.74]	
Pollett 2016	5	70	4	30	9.4%	0.38 [0.10, 1.44]	
Rao 2015		103	3	58	3.2%	0.18 [0.02, 1.77]	· · · · · · · · · · · · · · · · · · ·
oungster 2014	1	119	4	366	3.4%	0.77 [0.08, 6.93]	
		735			100.0%	0.48 [0.32, 0.73]	◆
Subtotal (95% CI)	38		129				
otal events	00 Chiž-			<sup>2</sup> = 0%			
otal events leterogeneity: Tau² = 0.		= 0.00	JU5)				
otal events			use Mortality				
otal events leterogeneity: Tau² = 0. est for overall effect: Z	= 3.50 (P	all.ce	and moreality	81	32.6%	1.20 [0.29, 5.00]	
otal events leterogeneity: Tau <sup>2</sup> = 0. est for overall effect: Z . <b>1.5 Cefazolin &gt; 4 gra</b> n	= 3.50 (P ns/daily- /		4	01	32.0% 6.4%	0.19 [0.01, 4.74]	·
otal events Heterogeneity: Tau <sup>2</sup> = 0. Test for overall effect: Z :1.5 Cefazolin > 4 gram Hynt 2017	= 3.50 (P ns/daily- / 4	68	4	34		0.10[0.01, 4.74]	
otal events Heterogeneity: Tau <sup>2</sup> = 0. Test for overall effect: Z <b>.1.5 Cefazolin &gt; 4 gram</b> lynt 2017 ji 2014	= 3.50 (P ns/daily- /		4 1 4	34 40		0.57 (0.12, 2.73)	
otal events Heterogeneity: Tau <sup>2</sup> = 0. Test for overall effect: Z :1.5 Cefazolin > 4 gram Hynt 2017	= 3.50 (P ns/daily- / 4 0	68 59	1		27.3% 7.4%	0.57 [0.12, 2.73] 0.14 [0.01, 2.70]	· · · · · · · · · · · · · · · · · · ·
otal events leterogeneity: Tau <sup>a</sup> = 0. est for overall effect: Z .1.5 Cefazolin > 4 gran lynt 2017 j 2014 diller, 2018	= 3.50 (P ns/daily- / 4 0 3 0	68 59 50	1 4	40	27.3%	0.57 [0.12, 2.73] 0.14 [0.01, 2.70] 0.18 [0.02, 1.77]	· _ · · · · · · · · · · · · · · · · · ·
iotal events teterogeneity: Tau <sup>*</sup> = 0 iest for overall effect: Z: .1.5 Cefazolin > 4 gran i/ynt 2017 j 2014 diller, 2018 donogue 2018 Rao 2015 Oungster 2014	= 3.50 (P <b>ns/daily</b> - <i>I</i> 4 0 3 0 1 1 1	68 59 50 71 103 119	1 4 3	40 71 58 366	27.3% 7.4% 12.7% 13.7%	0.14 [0.01, 2.70] 0.18 [0.02, 1.77] 0.77 [0.08, 6.93]	·
otal events teterogeneity. Tau <sup>2</sup> = 0. cest for overall effect. Z: .1.5 Cefazolin > 4 gran lynt 2017 .2 2014 diller, 2018 donogue 2018 donogue 2018 donogue 2015 'oungster 2014 uibtotal (95% CI)	= 3.50 (P ns/daily- / 4 0 3 0 1 1 1	68 59 50 71 103	1 4 3 3 4	40 71 58 366	27.3% 7.4% 12.7%	0.14 [0.01, 2.70] 0.18 [0.02, 1.77]	· · · · · · · · · · · · · · · · · · ·
otal events           teterogeneity: Tau" = 0;           set for overall effect. Z:           1.5. Cefazolin > 4 gran           typt 2017           1.2014           tiller, 2018           donogue 2018           ao 2015           'oungster 2014           aubtotal (95% CI)           'otal events	= 3.50 (P <b>ns/daily</b> - <i>J</i> 4 0 3 0 1 1 9	68 59 71 103 119 <b>470</b>	1 4 3 3 4 19	40 71 58 366 <b>650</b>	27.3% 7.4% 12.7% 13.7%	0.14 [0.01, 2.70] 0.18 [0.02, 1.77] 0.77 [0.08, 6.93]	<
Total events Teterogeneity, Tau <sup>2</sup> = 0. Test for overall effect Z 1.5 Cefazolin > 4 gran Iynt 2017 1.2014 Miller, 2018 Monogue 2018 Monogue 2018 Monogue 2018 Monogue 2015 Youngster 2014 Mubtotal (95% Cl) Total events Teterogeneity, Tau <sup>2</sup> = 0.	= 3.50 (P ns/daily- / 4 0 3 0 1 1 9 .00; Chi <sup>2</sup> =	68 59 70 71 103 119 <b>470</b> = 3.46	1 4 3 4 4 i, df = 5 (P = 0.63); I	40 71 58 366 <b>650</b>	27.3% 7.4% 12.7% 13.7%	0.14 [0.01, 2.70] 0.18 [0.02, 1.77] 0.77 [0.08, 6.93]	
otal events leterogeneity. Tau <sup>2</sup> = 0. est for overall effect. Z : 1.5 Cefazolin > 4 gran lynt 2017 12014 11012 12014 11012 1012	= 3.50 (P ns/daily- / 4 0 3 0 1 1 9 .00; Chi <sup>2</sup> =	68 59 70 71 103 119 <b>470</b> = 3.46	1 4 3 4 4 i, df = 5 (P = 0.63); I	40 71 58 366 <b>650</b>	27.3% 7.4% 12.7% 13.7%	0.14 [0.01, 2.70] 0.18 [0.02, 1.77] 0.77 [0.08, 6.93]	
otal events teterogeneity: Tau <sup>*</sup> = 0. est for overall effect: Z 1.5 Cefazolin > 4 gran ynt 2017 i 2014 iller, 2018 ionogue 2018 ao 2015 oungster 2014 ubtotal (95% CI) otal events teterogeneity: Tau <sup>*</sup> = 0.	= 3.50 (P ns/daily- / 4 0 3 0 1 1 9 .00; Chi <sup>2</sup> =	68 59 70 71 103 119 <b>470</b> = 3.46	1 4 3 4 4 i, df = 5 (P = 0.63); I	40 71 58 366 <b>650</b>	27.3% 7.4% 12.7% 13.7%	0.14 [0.01, 2.70] 0.18 [0.02, 1.77] 0.77 [0.08, 6.93]	
otal events leterogeneity. Tau <sup>2</sup> = 0. est for overall effect. Z : .1.5 Cefazolin > 4 gran lynt 2017 i 2014 liller, 2018 lonogue 2018 lonogue 2018 lonogue 2018 uongster 2014 ubtotal (95% CI) otal events leterogeneity. Tau <sup>2</sup> = 0.	= 3.50 (P ns/daily-/ 4 0 3 0 1 1 9 .00; Chi <sup>2</sup> =	68 59 70 71 103 119 <b>470</b> = 3.46	1 4 3 4 4 i, df = 5 (P = 0.63); I	40 71 58 366 <b>650</b>	27.3% 7.4% 12.7% 13.7%	0.14 [0.01, 2.70] 0.18 [0.02, 1.77] 0.77 [0.08, 6.93]	0.01 0.1 10 10 Favors cefazolin Favors anti-staph PCN

Fig. 1 All-cause mortality

preferred option for the treatment of MSSA infections. However, these results should be interpreted with trepidation because of the lack of randomization and retrospective study design of all of the studies included in our analysis. Additionally, 18 of the 19 studies included only patients with MSSA infection complicated by bacteremia. The inclusion of positive MSSA blood cultures as an inclusion criterion reduces potential heterogeneity in the patient population compared with MSSA infections without concurrent bacteremia. However, this limited our ability to detect any differences in treatment effects among patients with MSSA infections complicated with bacteremia compared with MSSA infections without concurrent bacteremia. It is unclear whether cefazolin has benefits compared with ASPs in patients without concurrent bacteremia.

Our study results are consistent with previous meta-analyses completed on this topic. Weis and colleagues completed a meta-analysis of 14 studies (n = 12,859), which showed cefazolin use was associated with lower 30-day mortality and less nephrotoxicity compared with ASPs. No difference was noted in treatment failure or relapse rates [33]. Bidell and colleagues completed a meta-analysis of seven studies (n = 4391)patients) comparing cefazolin with ASP, which suggested cefazolin was associated with a lower risk of 90-day mortality risk compared with ASPs, but no difference in clinical failure [34]. Shi and colleagues completed a meta-analysis of ten studies (n = 4779 patients), which showed cefazolin use was associated with reduced all-cause mortality and less risk of antibiotic discontinuation due to adverse drug events compared with ASPs [35]. Lee and colleagues published a metaanalysis of nine studies (n = 4442 patients) in acutely ill patients that showed reduced mortality and lower risk of treatment failure with cefazolin compared with ASPs [36]. Our metaanalysis of 19 studies encompassing over 13,000 patients is the largest meta-analysis to date comparing cefazolin to ASPs for the treatment of MSSA infection and is concordant with the previous studies. Inclusion of a larger study [16] published in 2018, which has not previously been included in multiple meta-analyses [34–36], and inclusion of gray literature explains the increased number of patients in our study. Additionally, our study is the first to apply TSA to confirm the meta-analysis results for cefazolin versus ASPs. Most meta-analyses do not have sufficient statistical power to detect or refute intervention effects. When the number of participants and the corresponding number of trials in a meta-analysis are insufficient, the use of the traditional 95% confidence interval or the 5% statistical significance threshold will lead to false conclusions (both type I and type II errors) [9]. Application of TSA to meta-analysis results allows for better interpretation of meta-analysis with small numbers of participants and included studies. In our study, TSA confirmed all of the observed results from the traditional meta-analvsis. The finding that TSA confirmed all of our significant study results suggests that, based on the available evidence, cefazolin use is associated with better outcomes compared with ASPs. Additionally, we performed a sensitivity analysis for confounding variables for the primary outcome using the *E* value methodology. The calculated *E* value was 2.17, meaning that residual confounding could explain the observed association if an unmeasured covariate exists with a relative risk association at least as large as 2.17 with both ASP use and mortality. However, our meta-analysis is not without limitations. As previously noted, all of the included studies had an observational, retrospective study design, which alone introduces the risk of bias, compared with a prospective, randomized trial design. Additionally, 11 of the 19 included studies had a moderate or serious risk of bias according to the ROBINS-I assessment tool used to identify risk of bias in non-randomized trials. However, sensitivity analysis including trials solely with a low risk of bias, and those with low/moderate risk of bias studies did not affect the results of the primary outcome. An additional limitation is that none of the included studies evaluated the potential presence or influence of the cefazolin inoculum effect among MSSA isolates. Inconsistent study definitions across studies also were observed. Future studies on this topic should have a prospective study design, evaluate for the presence of a cefazolin inoculum effect, and account for flaws commonly noted in retrospective and

observational studies by including proper randomization of study participants.

### CONCLUSION

The results of our meta-analysis, the largest that has been completed on this topic to date, suggest that cefazolin use is associated with a small benefit in all-cause mortality, including in studies with a high percentage of deep-seated infections such as endocarditis, osteomyelitis, and pneumonia. Cefazolin was also associated with reduced clinical failure and was better tolerated compared with ASPs for the treatment of patients with MSSA infections. These findings were derived from retrospective, observational studies, which weakens the study findings. However, TSA confirmed the benefits of cefazolin compared with ASPs. Further investigation should be undertaken to evaluate whether cefazolin is associated with an increase in infection recurrence in patients with MSSA infections compared with ASPs. Pragmatically, it would appear reasonable to avoid using cefazolin in patient scenarios with a high-risk of infection recurrence until further evidence is available (e.g., MSSA endocarditis without the possibility of surgical management, prosthetic joint infection without the possibility of hardware removal). Due to its favorable pharmacokinetic profile requiring less frequent dosing, and associated positive clinical outcomes, cefazolin appears to be a reasonable alternative to ASPs for the management of MSSA infections. Further research is needed to evaluate for patient-specific factors to guide antibiotic selection in patients with MSSA infections.

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*Authorship.* All named authors meet the International Committee of Medical Journal

Editors (ICMJE) criteria for authorship for this article, take responsibility for the integrity of the work as a whole, and have given their approval for this version to be published.

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*Compliance with Ethics Guidelines.* This article is based on previously conducted studies and does not contain any studies with human participants or animals performed by any of the authors

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