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## A rapid response magnitude scale for timely assessment of the high frequency seismic radiation

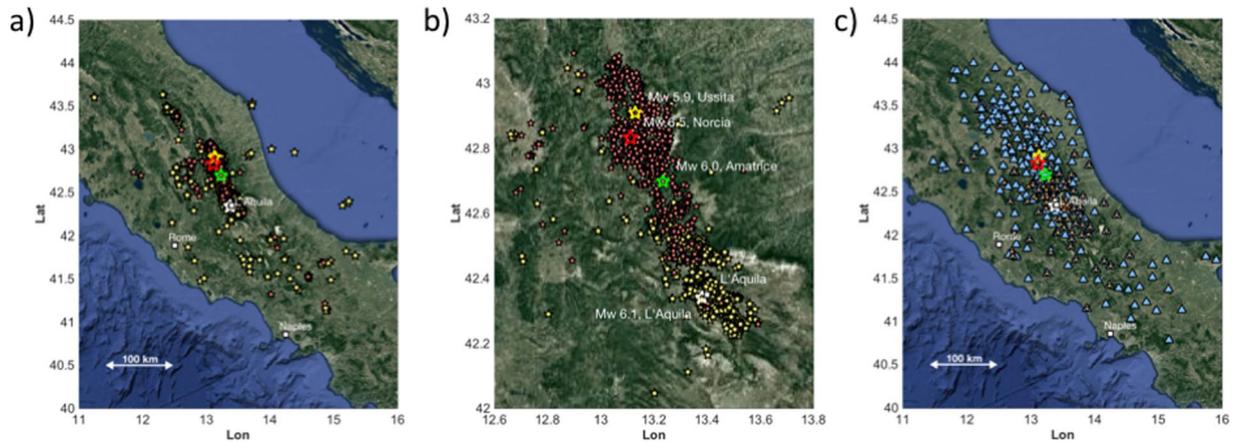
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In this work the scaling of seismic moment ( $M_0$ ) and radiated energy ( $E_r$ ) is investigated for almost 800 earthquakes of the 2016–17 Amatrice-Norcia sequences in Italy, ranging in moment magnitude ( $M_w$ ) from 2.5 to 6.5. The analysis of the  $M_0$ -to- $E_r$  scaling highlights a breaking of the source self-similarity, with higher stress drops for larger events. Our results show the limitation of using  $M_{or}$  and in turn  $M_{wr}$  to capture the variability of the high frequency ground motion. Since the observed seismicity does not agree with the assumptions on stress drop in the definition of  $M_{wr}$ , we exploit the availability of both  $E_r$  and  $M_0$  to modify the definition of  $M_w$  and introduce a rapid response magnitude ( $M_r$ ), which accounts for the dynamic properties of rupture. The new  $M_r$  scale allows us to improve the prediction of the earthquake shaking potential, as shown by the reduction of the between-event residuals computed for the peak ground velocity. The procedure we propose is therefore a significant step towards a quick assessment of earthquakes damage potential and timely implementation of emergency plans.

In the aftermath of an earthquake, the rapid assessment of both location and extension of potentially damaging ground shaking is a primary task for seismological agencies supporting emergency managers. In this context, shaking maps<sup>1</sup> become a de-facto standard for a timely dissemination of the ground shaking experienced in the area struck by an earthquake. To improve the spatial resolution of such maps for rapid response actions, the rapid determination of an earthquake size and location supports the information provided by the actual ground motion measurements (if available) and predicted ones. The moment magnitude  $M_w$ <sup>2,3</sup> is used by the seismological community as the primary measure of the earthquake size. Since  $M_w$  is based on an estimate of the seismic moment  $M_0$ <sup>4,5</sup>, it provides fault-averaged, low-frequency information on source processes but relatively less information about the small-wavelength high-frequency rupture details<sup>6</sup>. For example, earthquakes with similar  $M_w$  but different stress drop ( $\Delta\sigma$ ) can generate different ground motion levels<sup>7</sup>, suggesting that a rapid assessment of  $\Delta\sigma$  could allow more reliable predictions of the earthquake-induced ground motion severity to engineering structures (hereinafter, shaking potential). Moreover, recent studies<sup>8–10</sup> showed that the  $\Delta\sigma$  variability is a key parameter for explaining the between-event residuals at short periods. Since different magnitude scales provide different information about the static and dynamic features of the earthquake rupture, magnitudes other than  $M_w$  could better characterize the earthquake size in terms of high-frequency energy release<sup>11–13</sup>. For example,  $M_w$  can be complemented with a magnitude scale based on the high-frequency level of the Fourier spectrum<sup>14</sup>. Using teleseismic broadband P-wave recordings, a magnitude scale ( $M_c$ ) was introduced<sup>15</sup> based on measurements of the radiated energy ( $E_r$ ) and revising the Gutenberg and Richter relationship<sup>16</sup> between  $E_r$  and the surface-wave magnitude  $M_s$ . Since  $M_c$  is directly linked to the source dynamics, it is more sensitive to high-frequency source details such as variations of the slip and/or stress conditions, and the dynamic friction at the fault surface during the rupture process. Indeed, high energy-to-moment ratios indicate that the intensity of radiated energy at high frequencies is large relative to the size (measured by moment) of an earthquake, with significant implications on hazard assessment<sup>17</sup>. Although automatic procedures for the rapid estimation of  $M_c$  using P-wave recordings have been proposed both at teleseismic<sup>18</sup> and local<sup>19</sup> distances, a strategy to combine the information provided by  $M_w$  and  $M_c$  for a rapid assessment of the damage potential of an earthquake has not been proposed yet.

In this study, we present a new procedure to measure the earthquake size using rapid assessments of both  $E_r$  and  $M_0$ , considering S-wave recordings within 100 km from the epicenter. The methodology is applied to Central

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**Figure 1.** (a) Locations of the earthquakes considered in this study, occurring between 2009 and 2015 (yellow) and between 2016 and 2017 (red); (b) zoom in of the epicentral area, with the locations of the largest earthquakes shown by stars. (c) Location of the seismic stations relevant to either the calibration (orange reverse triangles) or application (cyan triangles) datasets. Maps were made using MATLAB (R2016b; 9.1.0.441655; <https://it.mathworks.com/>, last accessed January 2018).

Italy, used as an example region where the seismic hazard for residential buildings is dominated by close-distance earthquakes of low-to-moderate magnitude<sup>20</sup> (i.e., from 4.5 to 6.5). We first calibrate empirical attenuation relationships between the integral of squared velocity ( $IV2_s$ ) and the peak displacement ( $PD_s$ ) with respect to  $E_r$  and  $M_0$ , respectively, considering 229 earthquakes ranging from  $M_w$  2.4 to 6.1, mostly belonging to the L'Aquila (2009) seismic sequence<sup>21</sup>. Then, the procedure is applied to about 775 earthquakes of the 2016–2017 Central Italy sequence<sup>22</sup> with  $M_w$  in the range 2.5–6.5. Innovatively, we use the estimates of  $E_r$  and  $M_0$  (i.e., through the slowness parameter<sup>23</sup>) to introduce a rapid response magnitude ( $M_r$ ). The new  $M_r$  magnitude scale is tied to the non-saturating  $M_w$ , but it includes a term which accounts for the difference between the earthquake dynamic conditions and those assumed by Kanamori for the original definition of  $M_w$ . Finally, the impact of the new magnitude scale is discussed in terms of assessment of the earthquake shaking potential quantified through the peak ground velocity (PGV) and peak ground acceleration (PGA).

## Background

The moment magnitude  $M_w$  is defined as<sup>2</sup>

$$M_w = (\log M_0 - 9.1)/1.5 = (\log M_0 - 4.3 - 4.8)/1.5, \quad (1)$$

where the slope value of 1.5 and the constant  $-4.8$  are inherited from the relationship between  $\log(E_r)$  and the surface-wave magnitude  $M_s$ <sup>16</sup>, whilst the term  $-4.3$  derives from the following assumptions: (a) the energy required for fracturing is negligible; (b) the final average stress and the average stress during faulting are equal (also known as “complete stress drop”  $\Delta\sigma$  or “Orowan’s model”<sup>24</sup>); (c) the average rigidity in the source area ( $\mu$ ) is ranging from 3 to  $6 \times 10^4$  MPa under average crust-upper mantle conditions; (d)  $\Delta\sigma$  is nearly constant for very large earthquakes, with values between 2 and 6 MPa.

The above assumptions can be also formalized as:

$$E_r/M_0 = \Delta\sigma/(2\mu) = 5 \times 10^{-5} \quad (2)$$

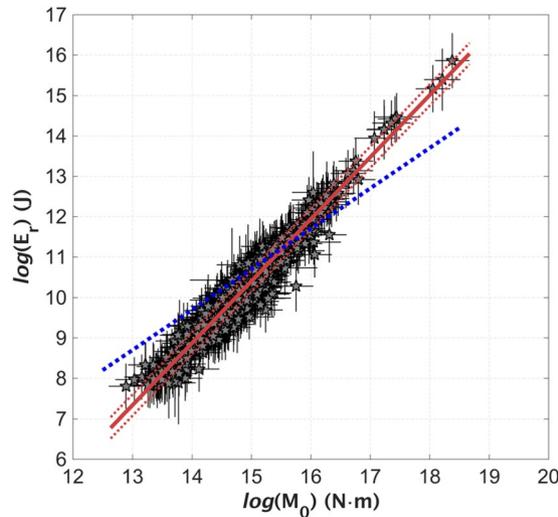
Introducing the slowness parameter term  $\theta = \log(E_r/M_0)$ <sup>23</sup>, equation (2) corresponds to:

$$\theta_K = -4.3 \quad (3)$$

that is,  $\theta$  assumes the value  $-4.3$  under the Kanamori assumptions<sup>2</sup> on  $\Delta\sigma$  and  $\mu$ . However, it is well-known that rupture processes may deviate from the complete stress drop, ranging between the “partial stress drop”<sup>25</sup> and “frictional overshoot”<sup>26</sup> models. Global datasets of earthquakes with  $M_w$  varying between 5.5 and 9 suggest that  $\theta$  varies between  $-4.7$  and  $-4.9$ <sup>15,27,28</sup>, which would correspond to an average global stress drop a factor 3 to 4 smaller than that assumed by Kanamori<sup>2</sup>. This difference in the average stress drop values is also reflected by the expected over estimation of about 0.27 m.u. of  $M_e$  with respect to  $M_w$  when an earthquake satisfies the condition  $\theta = \theta_K$ <sup>29</sup>.

**Dataset.** In this study, we analyze about 1000 earthquakes with  $M_w$  between 2.5 and 6.5, occurred in Central Italy between July 2008 and September 2017 (see Text Sa). This area of the Apennines is characterized by southwest-northeast extension (i.e., active NE- and SW-dipping normal and normal-oblique faults) since the Middle Pliocene<sup>30</sup>.

The dataset includes the main seismic sequences that have occurred over the past 10 years in Central Italy: the 2009 L'Aquila sequence<sup>21</sup>, and the 2016–2017 Central Italy sequence<sup>31</sup> (Fig. 1a). The four largest analyzed



**Figure 2.**  $E_r$  versus  $M_0$  computed for the 2016–2017 sequence (grey stars, with  $\pm 1$  standard deviations shown as vertical and horizontal bars); the best-fit model (mean  $\pm 1$  standard deviation are shown as red lines) is compared to the energy-to-moment ratio assumed by Kanamori<sup>2</sup> (blue dashed line).

earthquakes ( $M_w \geq 5.9$ ) are: 1) the April 6, 2009  $M_w$  6.3, which occurred near the town of L'Aquila and caused about 300 fatalities, 2) the August 24, 2016  $M_w$  6.0, Amatrice earthquake, 3) the October 26, 2016  $M_w$  5.9 Visso earthquake, and 3) the October 30, 2015  $M_w$  6.5 Norcia earthquake, which ruptured a  $\sim 40$  km long fault, bridging the seismic gap left from the previous two earthquakes<sup>22</sup>.

The data were recorded by the Italian strong motion network (RAN) and by the permanent and temporary stations of other networks (see Text Sa and Fig. 1). To reproduce a rapid response procedure, the recordings are pre-processed and the squared-velocity parameter ( $IV2_s$ ) and the peak displacement ( $PD_s$ ) are both measured on direct S-waves following an analysis scheme suitable for real-time operations<sup>32</sup>, (see Text Sb). From the whole compiled dataset, which consists of more than 1400 earthquakes, we extracted a subset considering the following selections: hypocentral distance smaller than 100 km; events recorded by a minimum of 8 stations and stations having at least 8 records; the sum of SNR for the three components  $\geq 200$ . The selected dataset is composed by 1004 earthquakes recorded at 340 stations. This dataset is split in two parts. The first part consists of 229 earthquakes, which occurred between July 2008 and January 2016 (hereinafter referred to as dataset D1), including the 2009 L'Aquila  $M_w$  6.3 sequence. Data set D1 is used to calibrate the attenuation models used in this study. The second dataset (hereinafter referred to as D2) is composed of 775 earthquakes occurred in the study area since January 2016 and it is used to exemplify the application of the proposed methodologies. For D2, the number of good quality recordings per event varies from 10 to 50 (Figure S1).

## Results

**Rapid response assessment of radiated energy and seismic moment.** The rapid assessment of  $E_r$  and  $M_0$  is based on extracting, from the time histories, proxies for these two source parameters, and correcting them for attenuation along the path. In particular, the radiated seismic energy is estimated from the squared velocity integrated over the S-wave time window<sup>33</sup> ( $IV2_s$ ), whereas the seismic moment is derived from the S-wave peak-displacement<sup>34</sup> ( $PD_s$ ). The parameters  $IV2_s$  and  $PD_s$  are linked to  $E_r$  and  $M_0$  through empirical attenuation models derived using data set D1, as detailed in the section Method (eqs 8 and 9). Since linear non-parametric regressions<sup>19</sup> are applied, the attenuation models are tables listing the attenuation coefficients for the discretized distance range. The attenuation models are reported in Table S1.

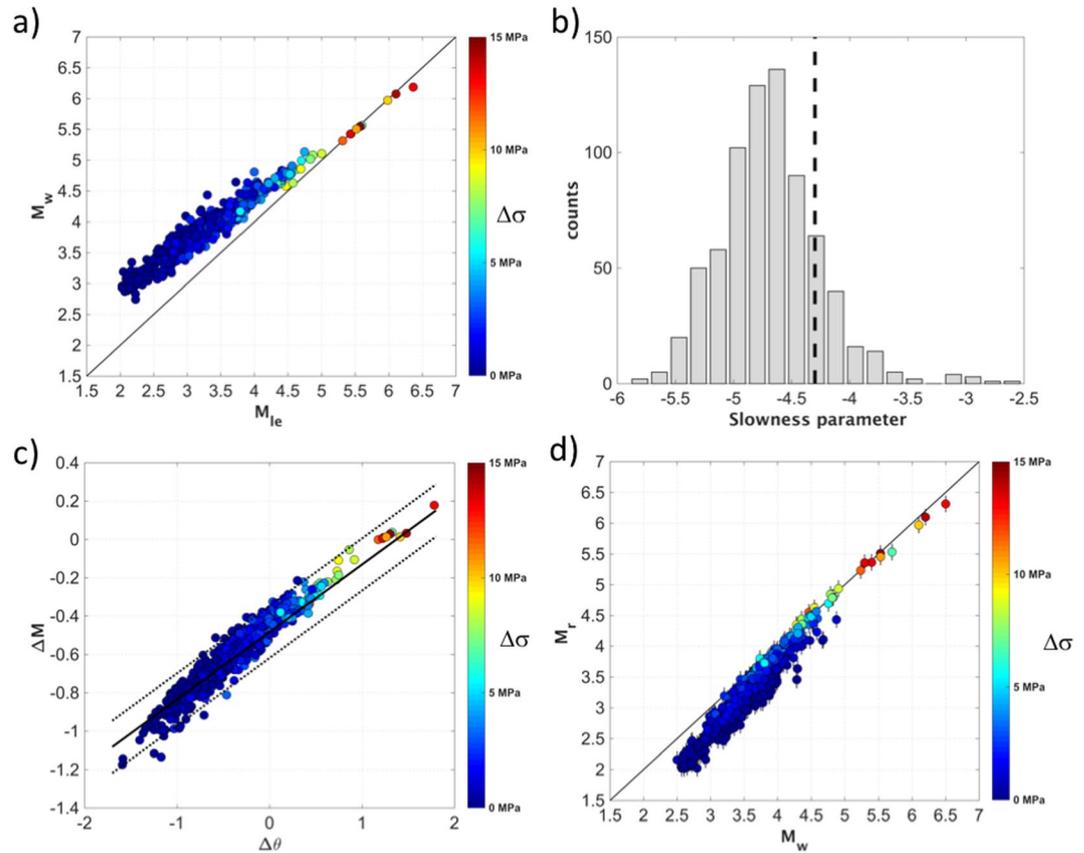
The attenuation models (8) and (9) calibrated over data set D1 are used to compute  $E_r$  and  $M_0$  of earthquakes included in D2. The  $E_r$  and  $M_0$  values are obtained by averaging over several recording stations (see Text Sc and Figure S2 for details).

Figure (2) shows that a linear scaling between  $E_r$  and  $M_0$  holds over 7 orders of magnitude for seismic moment. The observed  $M_0$ -to- $E_r$  scaling well compares to that of other tectonic areas (Figure S3). The best-fit model is:

$$\log(E_r) = (1.53 \pm 0.02)\log(M_0) - (12.59 \pm 0.26) \quad (4)$$

with a coefficient of determination ( $R^2$ ) equal to 0.94 and a standard deviation equal to 0.26.

The model of Eq. (4) deviates from the  $M_0$ -to- $E_r$  scaling related to the Kanamori's condition (i.e.,  $\log(E_r) = \log(M_0) + \theta_K$ ), shown in Fig. (2) as blue dashed line. The obtained slope larger than 1 (equation 4) implies a breaking of the source self-similarity, with stress drops larger for larger events. Moreover, the value of 12.59 obtained for the intercept implies that the condition  $\theta = \theta_K$  is met by the analyzed data set for  $M_0 = 10^{15.58}$  [Nm], corresponding to  $M_w = 4.3$ , as shown in Fig. (2).



**Figure 3.** (a) Moment magnitude ( $M_w$ ) versus energy-based local magnitude ( $M_{le}$ ), colored per  $\Delta\sigma$ . (b) Distribution of the slowness ratio  $\theta$  (grey) and  $\theta_K = -4.3$  (black dashed line). (c) Distribution of  $\Delta\theta = \theta - \theta_K$  versus  $\Delta M = (M_{le} - M_w)$  (circles) compared to the best-fit model given in Eq. 8 (the mean and the mean  $\pm 1$  standard deviation are the continuous and dashed black lines, respectively). (d) Rapid response magnitude ( $M_r$ ) versus  $M_w$ , circles, with  $\pm 1$  standard deviation bars; the 1:1 relationship (black) is shown for reference.

**Energy to moment scaling and event specific adjustment to  $M_w$ .** A straightforward application for rapid response purposes of the  $E_r$  and  $M_0$  estimates derived for the 2016–2017 dataset would consist of converting them into energy and moment magnitudes. In the case of  $M_0$ , we adopt the original definition proposed by Kanamori<sup>2</sup> (i.e., Eq. 1). By contrast, in the case of the energy magnitude, we use the calibration dataset to parameterize a linear relationship between  $E_r$  and the local magnitude ( $M_L$ )<sup>35</sup>. The best-fit model obtained by an iteratively reweighted least squares<sup>36</sup> and with 200 bootstrap replications<sup>37</sup> considered to assess the uncertainties (Figure S4) is:

$$M_L = (0.568 \pm 0.01)\log(E_r) - (2.54 \pm 0.1); \quad (5)$$

with  $R^2$  equal to 0.93 and the standard deviation of the residuals is equal to 0.13. Applying Eq. (5), the derived magnitudes agree with  $M_r$ , but being tied to  $E_r$ , can be extended toward larger earthquakes to avoid magnitude saturation effects<sup>38</sup>. The magnitudes obtained by the 2016–2017  $E_r$  estimates are referred to as energy-based local magnitude<sup>19</sup> ( $M_{le}$ ).

Figure (3a) shows that, while  $M_w$  and  $M_{le}$  well agree for  $M_w > 5$ , for smaller events  $M_w$  tends to progressively overestimate  $M_{le}$ , and the  $(M_w - M_{le})$  difference increases with decreasing stress-drop ( $\Delta\sigma$ )<sup>39</sup>.

The  $M_0$ -to- $E_r$  scaling (Fig. 2) and the differences between the energy and moment magnitudes in relation to  $\Delta\sigma$  (Fig. 3a) reflect a regional behavior of the source scaling which deviates from the Kanamori's assumptions  $\theta_K = -4.3$  (Eq. 3). As shown in Fig. (2), the ratio  $\log(E_r/M_0)$  is expected to be smaller than the values implied by  $\theta_K = -4.3$  for  $M_w < 4.3$ . Considering the magnitude distribution of the analyzed earthquakes, most of the empirical  $\theta$  values are smaller than  $\theta_K$  Fig. 3b, in agreement with  $\Delta\sigma$  smaller than the value assumed by Kanamori<sup>2</sup>. Therefore, these earthquakes can potentially generate a ground shaking variability at high frequency larger than the variability expected by assuming  $\theta$  constant as in Eq. (1).

In the context of early warning and rapid response applications where event-specific analysis is of concern, it is worth considering a new magnitude scale that accounts for the differences in the events dynamic conditions with respect to those assumed by Kanamori for the original definition of  $M_w$ .

Figure (3c) shows that  $\Delta M = (M_{le} - M_w)$  scales with  $\Delta\theta = (\theta - \theta_K)$  and correlate with  $\Delta\sigma$ . While  $\Delta\theta$  provides a mean for inferring, in the rapid response timeframe, how much the occurring earthquake deviates from the

Kanamori's condition (i.e., if the earthquake is as energetic as it is expected from the  $\theta_K$  condition), an empirical linear model between  $\Delta\theta$  and  $\Delta M$  can be used for modifying the original definition of  $M_w$ :

$$\Delta M = (0.34 \pm 0.04)\Delta\theta - (0.48 \pm 0.04); \quad (6)$$

with  $R^2$  equal to 0.88 and a standard deviation, assessed by a bootstrap approach<sup>37</sup>, equal to 0.16.

Therefore, given Eq. (6),  $\Delta M$  can be added to  $M_w$  for deriving a new rapid response magnitude scale:

$$M_r = \Delta M + (\log M_0 - 9.1)/1.5 = \Delta M + M_w \quad (7)$$

where  $M_r$  is used in the remainder of this work to indicate the new magnitude scale introduced to take into account the event-specific deviation of  $\theta$  from  $\theta_K$ . Figure (3d) compares  $M_r$  with  $M_w$  estimated by applying a non-parametric spectral inversion approach<sup>39</sup>. It is worth noting that for earthquakes characterized by  $\Delta\sigma$  values from  $\sim 5$  to  $\sim 15$  MPa,  $\Delta M$  varies between  $-0.2$  to  $0.2$  magnitude units. Therefore, the difference between  $M_r$  and  $M_w$  is negligible for the largest events of the sequence (i.e. for  $M_w > 5$ ) which fulfill Kanamori's condition. However, for smaller earthquakes characterized by  $\Delta\sigma$  below 2 MPa, the application of the  $\Delta M$  factor leads to  $M_r$  being significantly smaller than  $M_w$  (Fig. 3d).

To quantify the impact of  $M_r$  on predicting the ground shaking potential, we follow the earthquake engineering approach of analyzing the event-specific deviations from the expected mean regional attenuation trend. We refer to PGV as a measure of earthquake shaking potential, since it is controlled by the seismic energy at intermediate frequencies and, hence, better suited than peak ground acceleration (PGA) for statistical analysis correlating ground motion with damage<sup>40</sup>. We calibrate a ground motion prediction equation (GMPE) to model the PGV scaling with distance and magnitude in the target area (see Text Sc). The calibration is repeated twice, considering either  $M_w$  or  $M_r$ . Then, the between-event residuals<sup>41</sup> ( $\delta Be$ ) are computed as the average difference, for any given event, between the PGV measured at different stations and the corresponding values predicted by the GMPE. However, considering that PGA can be anyway of interest for engineering<sup>42</sup> or seismological<sup>10,43,44</sup> applications, the same analyses are repeated for PGA, (see Text Sc).

Figure (4a) shows that the standard deviation of  $\delta Be$  computed considering PGV reduces from 0.18 to 0.08 when  $M_w$  is replaced with  $M_r$ , meaning that  $M_r$  better accounts for those variations in the rupture process which can introduce systematic event-dependent deviations from the mean regional PGV scaling. As shown in Fig. (4b), the reduced  $\delta Be$  variability is mostly related to the small magnitude events (i.e.  $M_w < 4.5$ ) which show the largest deviation from the  $\theta = \theta_K$  condition. A similar result is obtained also for PGA (Fig. 4c and d), for which the standard deviation of  $\delta Be$  reduces from 0.22 to 0.13 when  $M_w$  is replaced by  $M_r$ .

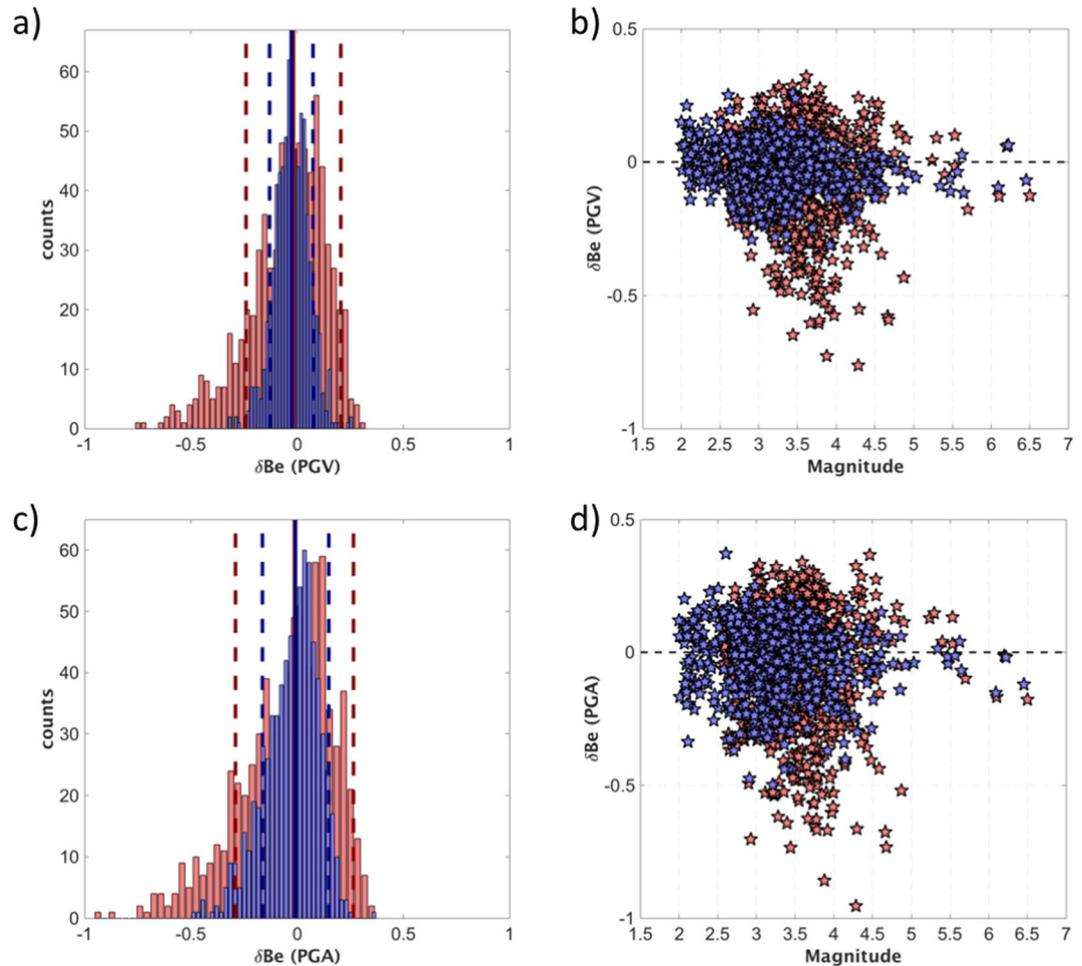
## Discussion

This study was motivated by the need of developing a procedure for the rapid assessment of the earthquake shaking potential. To accomplish such a task, we worked on two topics: first, we introduced an approach for the rapid assessment of the seismic moment and the radiated energy; second, we used these two source parameters to define a measure of the earthquake size which is not only informative for the average tectonic effects but it accounts for the efficiency of the earthquake to radiate seismic energy at high-frequency.

Concerning the first topic, standard approaches adopted by seismic observatories for estimating  $M_0$  from moment tensor analysis<sup>45,46</sup> are delayed by the requirement of acquiring full waveforms at regional-teleseismic distances. For instance, in Italy the moment tensor solutions are provided by INGV generally between 6 to 9 minutes after the earthquake occurrence (<http://info.terremoti.ingv.it/en/help#TDMT>) and ShakeMaps are released to the Civil Protection after this time span (<http://shakemap.rm.ingv.it/shake/about.html>). In this work, we have proposed an approach for quick and robust estimations of  $M_0$  and  $E_r$  based on proxies measured on S-waves recorded by a local network. The application of the attenuation models calibrated for both  $M_0$  and  $E_r$  to 775 earthquakes of the 2016–2017 Central Italy sequence confirmed the suitability of the approach for rapid-response applications.

Figure (S8) shows the comparison between the  $E_r$  estimates obtained by the procedure proposed in this study, which follows Picozzi *et al.*<sup>19</sup>, and the  $E_r$  obtained for the same dataset by a non-parametric inversion approach<sup>39</sup>, which follows Kanamori *et al.*<sup>38</sup>. As shown within Figure (S8), the two considered procedures provide consistent results, within a factor 1.5/2 (i.e., the  $\log(E_r)$  differences are within  $0.2/0.3 \log_{10}$  units). We observe that the rapid estimates of  $E_r$  tend to be slightly larger than those computed by Bindi *et al.*<sup>39</sup>. At least part of this overestimation could be ascribed to local site amplification effects, which have been considered in Bindi *et al.*<sup>39</sup> through site correction terms, not included in equation (8), an issue of interest for future studies. However, considering that  $E_r$  estimates on the same earthquakes provided by different investigators often differ up to one order of magnitude<sup>47</sup>, the result supports the reliability of the rapid estimates of  $E_r$  from IV2<sub>s</sub>. The availability of a dense network in the area under study was certainly an important condition to average out radiation pattern and rupture directivity effects on  $E_r$  and  $M_0$  estimates. Future examination will explore the bias on  $E_r$  and  $M_0$  estimates for network configurations less dense than the one considered.

Regarding the earthquake damage potential assessment, the analyzed data set represented a suitable test-case to discuss the limitation of using  $M_0$  to capture the variability of the high frequency ground motion. The topic is so important that it has been recently debated by different studies<sup>7–10</sup>. In this study, we have considered the same scientific topic but from the point of view of an observatory which must provide real-time information to support civil protection decisions and seismic risk mitigation actions. It is important to note that early-warning and rapid-response systems often deal with seismic sequences occurring in a specific target region. Since the within-sequence stress drop variability is expected to be larger than the between-mainshock stress drop variability<sup>48</sup>, the  $E_r$ -to-  $M_0$  scaling can deviate from the assumptions behind the definition of  $M_w$ <sup>2</sup>. For example,  $\Delta\sigma$  of



**Figure 4.** (a) Histograms of the between-event residuals ( $\delta Be$ ) computed for PGV considering either  $M_w$  (red) or  $M_r$  (blue); the  $\pm 1$  standard deviations are red and blue dashed lines for  $M_w$  and  $M_r$ , respectively. (b)  $\delta Be$  residuals for PGV considering either  $M_w$  (red) or  $M_r$  (blue) versus  $M_w$ ; the zero-bias value (dashed line) is shown for reference. (c) Same as a), but  $\delta Be$  computed for PGA. (d) Same as b), but  $\delta Be$  computed for PGA.

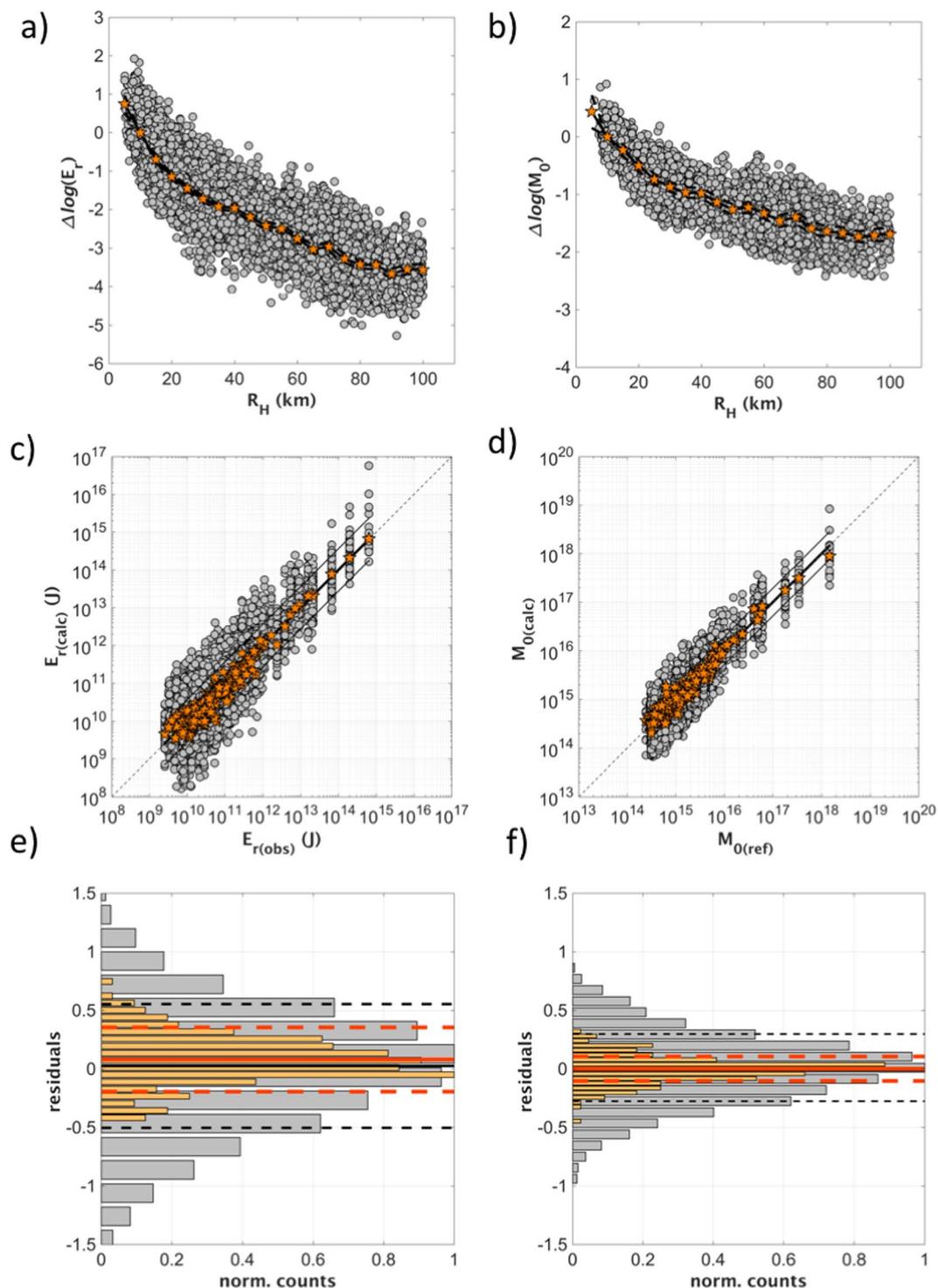
the earthquakes analyzed in this study varies from about 0.1 to 15 MPa, leading to slowness ratios significantly different from  $\theta_k$  (Fig. 3b).

The rapid assessment of  $M_0$  and  $E_r$  can thus be exploited to define a new rapid response magnitude scale ( $M_r$ ) tied to  $M_w$  but including a term that accounts for the difference between the actual energy-to-moment ratio and the value used by Kanamori<sup>2</sup>.

To demonstrate the potential value of the rapid response magnitude  $M_r$  in better capturing the event-to-event shaking potential variability, we quantified the reduction of the between-event standard deviation for PGV and PGA, which is controlled by the  $\Delta\sigma$  variability. Considering  $\Delta\sigma$  values<sup>37</sup> and the scheme proposed by Choy<sup>49</sup> over the energy-to-moment ratio, we observe that earthquakes with  $M_w > 4$  (showing  $\Delta\sigma$  from 5 to 15 MPa) are moderately or highly energetic, whilst small magnitude events (characterized by  $\Delta\sigma < 1$  MPa) are enervated (Figure S5). In our opinion, such large variability in the source properties justifies the introduction of the new magnitude scale for rapid response procedures, which is capable of better predicting the ground motion over a frequency range of engineering interest.

The proposed approach for computing  $E_r$ ,  $M_0$  and  $M_r$  can be implemented into algorithms for real-time operations<sup>50,51</sup>, so that  $M_r$  estimates can be available as soon as the S-waves are recorded by a minimum number of stations that allow stable  $E_r$  and  $M_0$  estimates to be obtained (i.e., within a few tens of seconds, depending on the seismic network density and telemetry efficiency), and can be exported to other areas for which the attenuation model for  $E_r$  and  $M_0$  can be calibrated. The application of the procedure for magnitudes larger than those in our dataset might be influenced by the saturation limit of the proxies used for the estimation of  $E_r$  and  $M_0$ . We expect that extending the observation scale from local to regional networks and considering other seismic phases than direct S-waves would allow for estimates of  $M_r$  to be extended to earthquakes with magnitude larger than  $M_w$  6.5. The saturation limit for the rapid-response procedure here proposed will be investigated in future studies.

The implications of our results are for a broad scientific audience (e.g., seismological, engineering seismology, disaster manager communities). Indeed, the implementation of our procedure for computing the seismic moment ( $M_0$ ), the radiated energy ( $E_r$ ) and  $M_r$  in real-time operations can allow civil protection operators to



**Figure 5.** Results of the calibrations between IV2<sub>s</sub> and  $E_r$  (Eq. 8) and between PD<sub>s</sub> and  $M_0$  (Eq. 9). **(a)** The coefficients  $C_j$  of Eq. 8 (orange) are compared with the residuals  $\Delta \log(E_r) = \log[IV2_s(R_H)] - A - B \log(E_r)$  (grey circles); **(b)** the coefficients  $G_j$  in Eq. 9 (orange circles) are compared with the residuals  $\Delta \log(M_0) = \log[PD_s(R_H)] - D + F \log(M_0)$  (grey circles); **(c)** the IV2<sub>s</sub> values corrected for  $C_j$  are compared with the energy  $E_{r(\text{obs})}$  (the corrected values for each recording are in grey, the average for each earthquake in orange); **(d)** PD<sub>s</sub> values corrected for  $G_j$  are compared with  $M_{0(\text{obs})}$  (the corrected values for each recording are in grey, the average for each earthquake in orange); **(e)** histograms of the residual distribution computed for Eq. (8) considering either single recording (grey, with  $\pm 1$  standard deviation, dashed black lines) or the average computed by grouping recordings per event (orange). **(f)** The same as in panel e), but considering Eq. (9).

better assess within a rapid response timeframe (i.e., within a few tens of seconds from the earthquake's origin time) an earthquake damage potential and the timely implementation of emergency plans. Furthermore, considering that  $M_r$  reduces the variability of the between-event residuals<sup>41</sup> with respect to  $M_w$ , we foresee its possible application in the development of ground motion prediction equations for seismic hazard assessment.

## Method

**Calibration of the empirical relationships  $IV2_s$ - $E_r$  and  $PD_s$ - $M_0$ .** Following previous studies<sup>19</sup>, we have derived an empirical relationship between  $IV2_s$  and  $E_r$ . To this purpose, we considered 229 earthquakes, and we computed the theoretical seismic radiated energy following a procedure<sup>52</sup> based on the spectral integration of the squared theoretical Brune's velocity spectrum for S-waves<sup>25</sup>. These have been derived using the corner frequency ( $f_c$ ) and seismic moment ( $M_0$ ) computed in previous studies<sup>8,53</sup>.

Hence, we related the  $IV2_s$  estimated for each recording to  $E_r$  radiated by the source assuming a linear model, where the attenuation of average  $IV2_s$  amplitude as a function of distance is expressed without assuming any a-priori functional form (i.e., non-parametric, or data-driven, approach):

$$\log[IV2_s(R_H)] = A + B \log(E_r) + w_j C_j + (1 - w_j) C_{j+1} \quad (8)$$

where the hypocentral distance  $R_H$  range is discretized into  $N_{bin}$ ; the index  $j = 1, \dots, N_{bin}$  indicates the  $j$ -th node selected such that  $R_H$  is between the distances  $r_j \leq R_H < r_{j+1}$ ; the attenuation function is linearized between nodes  $r_j$  and  $r_{j+1}$  using the weights  $w_j = (r_{j+1} - R_H) / (r_{j+1} - r_j)$ .

The  $R_H$  range 5–100 km is discretized into 19 bins with equal width (i.e., 5 km) (Fig. 5a). The minimum distance is fixed to 5 km given the lack of recordings at shorter distances. The coefficients  $A$ ,  $B$ ,  $C_j$  are determined by solving the over-determined linear system (8) in a least-square sense. To fix the trade-off between  $A$  and  $C_j$ , the attenuation is constrained to zero at  $r_2 = 10$  km (i.e., at the upper boundary of the first distance bin). However, it is worth noting that changing the node constrained to zero corresponds to changes of both the  $A$  and  $C_j$  coefficients which compensate each other, so that the coefficient  $B$  in Eq. (8) is unaltered.

Similarly, the relationship between the peak displacement  $PD_s$  and  $M_0$  has been expressed in the form:

$$\log[PD_s(R_H)] = D + F \log(M_0) + w_j G_j + (1 - w_j) G_{j+1} \quad (9)$$

The coefficients  $A$ ,  $B$ ,  $C_j$  and  $D$ ,  $F$ ,  $G_j$  are reported in Table S1. The calculated regressions have a  $R^2$  equal to 0.84 and 0.87, respectively.

The scaling with distance of the derived attenuation models is shown in Fig. 5a and b, where the distance distribution of the observed  $IV2_s$  and  $PD_s$  values corrected for the source scaling  $A + B \log(E_r)$  and  $D + F \log(M_0)$  are compared with the attenuation models  $C_j$  and  $G_j$ , respectively (i.e., where the index  $j$  indicates different bins of hypocentral distances). For both cases, the good agreement with the corrected data confirms the suitability of the obtained attenuation models.

Figure 5c and d compare the observed  $IV2_s$  and  $PD_s$  values corrected for attenuation effects, as modelled through the  $C_j$  and  $G_j$  coefficients, with  $E_r$  and  $M_0$ . The source scaling (black line) defined by the coefficients  $A$ ,  $B$  and  $D$ ,  $F$  capture well the trend in the data over the entire energy and moment ranges, as indicated by the agreement with the median value obtained for each event by averaging over more than 10 recording (yellow circles in Fig. 5c and d). Finally, Fig. 5e and f summarize the residual distributions between predicted and observed  $\log(E_r)$  and  $\log(M_0)$ . Considering the values computed for each recording (gray), the residuals are unbiased with standard deviations equals to 0.56 and 0.29 for  $E_r$  and  $M_0$ , respectively; when the average values per event are considered (yellow), the standard deviations reduce to 0.27 and 0.1, respectively.

**Data and Resources.** The waveforms used in this study have been obtained from European Integrated Data Archive-EIDA (<https://www.orfeus-eu.org/data/eida/>) and from the Italian Civil Protection (DPC) repository (<http://ran.protezionecivile.it/IT/index.php>). Regarding the permanent networks, we used data from the networks with FDSN (<http://www.fdsn.org/networks/>) code: MN (<https://doi.org/10.13127/SD/fBBBtDtd6q>), IV (<https://doi.org/10.13127/SD/X0FXnH7QfY>), IT (<https://doi.org/10.7914/SN/IT>). All of the webpages were last visited in March 2018. The analysis has been performed using the MATLAB software (R2016b; 9.1.0.441655; <https://it.mathworks.com/>, last accessed January 2018). Data generated during this study are available from the corresponding author on request.

## References

1. Worden, C. B. *et al.* A revised ground-motion and intensity interpolation scheme for ShakeMap. *Bull. seism. Soc. Am.* **100**(6), 3083–3096 (2010).
2. Kanamori, H. The energy release in great earthquakes. *J. Geophys. Res.* **82**, 2981–2876, <https://doi.org/10.1029/JB082i020p02981> (1977).
3. Hanks, T. C. & Kanamori, H. A moment magnitude scale. *J. Geophys. Res.* **84**, 2348–2350, <https://doi.org/10.1029/JB084iB05p02348> (1979).
4. Aki, K. Seismic displacements near a fault. *J. Geophys. Res.* **73**, 5359–5376, <https://doi.org/10.1029/JB073i016p05359> (1968).
5. Kanamori, H. & Anderson, D. L. Theoretical basis of some empirical relations in seismology. *Bull. Seismol. Soc. Am.* **65**, 1073–1095 (1975).
6. Beresnev, I. A. The reality of the scaling law of earthquake-source spectra? *J. Seismol.* **13**, 433–436, <https://doi.org/10.1007/s10950-008-9136-9> (2009).
7. Baltay, A., Hanks, T. & Beroza, G. Stable stress-drop measurements and their variability: Implications for ground-motion predictions. *Bull. Seismol. Soc. Am.* **103**, 211–222 (2013).

8. Bindi, D., Spallarossa, D. & Pacor, F. Between-event and between-station variability observed in the Fourier and response spectra domains: comparison with seismological models. *Geophysical Journal International* **210**, 1092–1104, <https://doi.org/10.1093/gji/ggx217> (2017).
9. Baltay, A. S., Hanks, T. C. & Abrahamson, N. A. Uncertainty, Variability, and Earthquake Physics in Ground-Motion Prediction Equations. *Bull. Seismol. Soc. Am.* **107**, 1754–1772, <https://doi.org/10.1785/0120160164> (2017).
10. Oth, A., Miyake, H. & Bindi, D. On the relation of earthquake stress drop and ground motion variability. *J. Geophys. Res. Solid Earth* **122**, 5474–5492, <https://doi.org/10.1002/2017JB014026> (2017).
11. Kanamori, H. Magnitude scale and quantification of earthquakes. *Tectonophysics* **93**, 185–199. ISSN 0040-1951 (1983).
12. Atkinson, G. M. Optimal choice of magnitude scales for seismic hazard analysis. *Seismol. Res. Lett.* **66**(1), 51–55, <https://doi.org/10.1785/gssrl.66.1.51> (1995).
13. Bormann, P., Wendt, S., & Di Giacomo, D. Seismic sources and source parameters. In *New Manual of Seismological Observatory Practice 2 (NMSOP2)*, edited by Bormann, P., Potsdam, Deutsches GeoForschungsZentrum GFZ, 1–259, [https://doi.org/10.2312/GFZ.NMSOP-2\\_ch3](https://doi.org/10.2312/GFZ.NMSOP-2_ch3) (2012).
14. Atkinson, G. M. & Hanks, T. C. A high-frequency magnitude scale. *Bull. Seismol. Soc. Am.* **85**, 825–833 (1995).
15. Choy, G. L. & Boatwright, J. Global patterns of radiated seismic energy and apparent stress. *J. Geophys. Res.* **100**(18), 205–226, <https://doi.org/10.1029/95JB01969> (1995).
16. Gutenberg, B. & Richter, C. F. Magnitude and energy of earthquakes. *Ann. Geofisica* **9**, 1–15 (1956).
17. Choy, G. L. & Kirby, S. Apparent stress, fault maturity and seismic hazard for normal-fault earthquakes at subduction zones. *Geoph. Journ. Int.* **159**(3), 991–1012, <https://doi.org/10.1111/j.1365-246X.2004.02449.x> (2004).
18. Di Giacomo, D. *et al.* Suitability of rapid energy magnitude estimations for emergency response purposes. *Geophys. J. Int.* **180**, 361–374, <https://doi.org/10.1111/j.1365-246X.2009.04416.x> (2010).
19. Picozzi, M. *et al.* Rapid determination of P wave- based energy magnitude: Insights on source parameter scaling of the 2016 Central Italy earthquake sequence. *Geophys. Res. Lett.* **44**, 4036–4045, <https://doi.org/10.1002/2017GL073228> (2017).
20. Barani, S., Spallarossa, D. & Bazzurro, P. Disaggregation of probabilistic ground-motion hazard in Italy. *Bull. Seismol. Soc. Am.* **99**, 2638–2661 (2009).
21. Ameri, G. *et al.* The 6 April 2009 Mw 6.3 L'Aquila (Central Italy) earthquake: Strong-motion observations. *Seismol. Res. Lett.* **80**, 951–966, <https://doi.org/10.1785/gssrl.80.6.951> (2009).
22. Chiaraluce, L. *et al.* The 2016 Central Italy Seismic Sequence: A First Look at the Mainshocks, Aftershocks, and Source Models. *Seismological Res. Letters* **88**, 757–771, <https://doi.org/10.1785/0220160221> (2017).
23. Newman, A. V., & Okal, E. A. Teleseismic estimates of radiated seismic energy: the Er/Mo discriminant for tsunami earthquakes. *J. Geophys. Res.* **103**, 26, 885–26, 898 (1998).
24. Orowan, E. Mechanism of seismic faulting. *Geol. Soc. Am. Mem.* **79**, 323–345 (1960).
25. Brune, J. N. Tectonic stress and the spectra of seismic shear waves from earthquakes. *J. Geophys. Res.* **75**, 4997–5009, <https://doi.org/10.1029/JB075i026p04997> (1970).
26. Savage, J. C. & Wood, M. D. The relation between apparent stress and stress drop. *Bull. Seismol. Soc. Am.* **61**, 1381 (1971).
27. Choy, G. L., McGarr, A., Kirby, S. H., & Boatwright, J. An overview of the global variability in radiated energy and apparent stress. In: *Abercrombie R., McGarr, A., and Kanamori, H. (eds): Radiated energy and the physics of earthquake faulting, AGU Geophys. Monogr. Ser.* **170**, 43–57 (2006).
28. Weinstein, S. A. & Okal, E. A. The mantle magnitude  $M_m$  and the slowness parameter  $\Theta$ : Five years of real-time use in the context of tsunami warning. *Bull. Seism. Soc. Am.* **95**(3), 779–799, <https://doi.org/10.1785/0120040112> (2005).
29. Bormann, P., & Di Giacomo D. Earthquakes, Energy. In *Encyclopedia of Solid Earth Geophysics*, Ed. Springer Nature, 233–237, [https://doi.org/10.1007/978-90-481-8702-7\\_24](https://doi.org/10.1007/978-90-481-8702-7_24) (2011).
30. Pace, B., Peruzza, L., Lavecchia, G. & Boncio, P. Layered seismogenic source model and probabilistic seismic-hazard analyses in central Italy. *Bull. Seism. Soc. Am.* **96**, 107–132, <https://doi.org/10.1785/0120040231> (2006).
31. Luzi, L. *et al.* The Central Italy Seismic Sequence between August and December 2016: Analysis of Strong-Motion Observations. *Seismological Research Letters* **88**(5), 1219–1231, <https://doi.org/10.1785/0220170037> (2017).
32. Zollo, A., Lancieri, M. & Nielsen, S. Earthquake magnitude estimation from peak amplitudes of very early seismic signals on strong motion. *Geophys. Res. Lett.* **33**, L23312, <https://doi.org/10.1029/2006GL027795> (2006).
33. Brondi, P., Picozzi, M., Emolo, A., Zollo, A. & Mucciarelli, M. Predicting the macroseismic intensity from early radiated P wave energy for on-site earthquake early warning in Italy. *Journal of Geophysical Research. Solid Earth* **120**, 7174–7189, <https://doi.org/10.1002/2015JB012367> (2015).
34. Festa, G. *et al.* Performance of Earthquake Early Warning Systems during the 2016–2017 Mw 5–6.5 Central Italy Sequence. *Seismological Research Letters* **89**(1), 1–18, <https://doi.org/10.1785/0220170150> (2017).
35. Di Bona, M. A local magnitude scale for crustal earthquakes in Italy. *Bull. Seism. Soc. Am.* **106**, 242–258 (2016).
36. Street, J. O., Carroll, R. J. & Ruppert, D. A Note on Computing Robust Regression Estimates via Iteratively Reweighted Least Squares. *The American Statistician* **42**, 152–154 (1988).
37. Efron, B. Bootstrap methods: Another look at the jackknife. *Ann. Stat.* **7**(1), 1–26, <https://doi.org/10.1214/aos/1176344552> (1979).
38. Kanamori, H., Hauksson, E., Hutton, L. K. & Jones, L. M. Determination of earthquake energy release and ML using TERRASCOPE. *Bull. Seismol. Soc. Am.* **83**, 330–346 (1993).
39. Bindi D., Spallarossa, D., Picozzi, M., Scafidi, D., & Cotton, F. Impact of magnitude selection on aleatory variability associated with Ground Motion Prediction Equations: Part I – local, energy and moment magnitude calibration for Central Italy, *Bull. Seismol. Soc. Am.* <https://doi.org/10.1785/0120170356> (2018).
40. Perrault, M. & Guéguen, P. Correlation between Ground Motion And Building Response Using California Earthquake Records. *Earthquake Spectra* **31**, 2027–2046, <https://doi.org/10.1193/062413EQS.168M> (2015).
41. Al Atik, L. *et al.* The variability of ground-motion prediction models and its components. *Seismol. Res. Lett.* **81**(5), 794–801, <https://doi.org/10.1785/gssrl.81.5.794> (2010).
42. Yakut, A. & Yilmaz, H. Correlation on Deformation Demands with Ground Motion Intensity. *Journal of Structural Engineering* **134**(12), 1818–1828, [https://doi.org/10.1061/\(ASCE\)0733-9445\(2008\)134:12\(1818\)](https://doi.org/10.1061/(ASCE)0733-9445(2008)134:12(1818)) (2008).
43. Archuleta, R. J. & Ji, C. Moment rate scaling for earthquakes  $3.3 \leq M \leq 5.3$  with implications for stress drop. *Geophys. Res. Lett.* **43**, 12004–12011, <https://doi.org/10.1002/2016GL071433> (2016).
44. Cotton, F., Archuleta, R. & Causse, M. What is sigma of the stress drop? *Seismol. Res. Lett.* **84**, 42–48, <https://doi.org/10.1785/0220120087> (2013).
45. Fukuyama, E. & Dreger, D. S. Performance test for automated moment tensor determination system by using synthetic waveforms of the future Tokai earthquake. *Earth Planet. Space* **52**, 383–392 (2000).
46. Dreger, D. S. TDMT\_INV: Time Domain Seismic Moment Tensor INVersion. In: W. K. Lee, H. Kanamori, P. C. Jennings, C. Kisslinger (Eds). *International Handbook of Earthquake and Engineering Seismology* **81B**, 1627 (2003).
47. Pérez-Campos, X. & Beroza, G. C. An apparent mechanism dependence of radiated seismic energy. *Journal of Geophysical Research* **106**, 11127–11136 (2001).
48. Baltay, A. S. & Hanks, T. C. Understanding the magnitude dependence of PGA and PGV in NGA-West2 Data. *Bull. Seis. Soc. Am.* **104**, 6, <https://doi.org/10.1785/0120130283> (2014).

49. Choy G. L. Stress conditions inferable from modern magnitudes: development of a model of fault maturity. **IS3.5** in Bormann, P. (Ed.) *New Manual of Seismological Observatory Practice (NMSOP-2)*, IASPEI, GFZ German Research Centre for Geosciences, Potsdam; <http://nmsop.gfz-potsdam.de>; <https://doi.org/10.2312/GFZ.NMSOP-2> (2012).
50. Satriano, C., Elia, L., Martino, C., Lancieri, M. & Zollo, A. & Iannacone G. PRESTo, the earthquake early warning system for Southern Italy: Concepts, capabilities and future perspectives. *Soil Dyn. Earthq. Eng.* **31**(2), 137–153, <https://doi.org/10.1016/j.soildyn.2010.06.008> (2011).
51. Picozzi, M. *et al.* Exploring the feasibility of a nationwide earthquake early warning system in Italy. *J. Geophys. Res. Solid Earth* **120**, 2446–2465, <https://doi.org/10.1002/2014JB011669> (2015).
52. Izutani, Y. & Kanamori, H. Scale-dependence of seismic energy-to-moment ratio for strike-slip earthquakes in Japan. *Geophys. Res. Lett.* **28**, 4007–4010, <https://doi.org/10.1029/2001GL013402> (2001).
53. Pacor, F. *et al.* Spectral models for ground motion prediction in the LAquila region (central Italy): evidence for stress-drop dependence on magnitude and depth. *Geophysical Journal International* **204**(2), 697–718, <https://doi.org/10.1093/gji/ggv448> (2016).

## Acknowledgements

We thank the Editor H. Tkalcic and two anonymous reviewers for their comments which helped improving the manuscript. This study has been partially funded by the H2020-INFRAIA project SERA (Seismology and Earthquake Engineering Research Infrastructure Alliance for Europe). We thank F. Pacor, J. Anderson, K. Fleming and E. Rivalta for useful discussions. We thank Lonni Brown (ISC) who has kindly improved the English. Analysis and plots were made using MATLAB (R2016b; 9.1.0.441655; <https://it.mathworks.com/>, last accessed January 2018).

## Author Contributions

M.P. designed the study and performed the analysis; M.P. and D.B. wrote the manuscript and, together with D.S., contributed to the data analysis; D.S. prepared and processed the dataset. All authors contributed to the interpretation of the results and approved the final version of the manuscript.

## Additional Information

**Supplementary information** accompanies this paper at <https://doi.org/10.1038/s41598-018-26938-9>.

**Competing Interests:** The authors declare no competing interests.

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