

# Incidence and risk of fatigue in cancer patients treated with MET inhibitors

A systematic review and meta-analysis

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#### Abstract

**Background:** The N-methyl-N'-nitroso-guanidine human osteosarcoma transforming gene (MET) inhibitors show a surprising survival benefit in the treatment of numerous tumors especially in MET-high tumor. Besides their impressive efficacy, fatigue reduced by MET inhibitors is still the safety issue during treatment. Thus, an understanding of this risk in the context of expanding MET-inhibitors use is an important cost and patient safety issue.

**Methods:** We searched PubMed, Embase, and the Cochrane Library databases for relevant studies up to October 2017. Eligibility criteria included phase II/III trials of MET inhibitors that reported adequate safety profiles of fatigue. The principal summary measures were incidence and relative risk (RR) of all-grade (grade 1–4) and high-grade (grade 3–4) fatigue, respectively. Random-effects model was applied to consider within-study and between-study variation.

**Results:** A total of 5028 patients from 17 clinical trials were identified. The results revealed that the incidences of MET inhibitorsassociated all-grade and high-grade fatigue were 41.9% and 9.6%, respectively. The RR of high-grade fatigue was (RR = 1.37; 95% confidence interval, 1.14–1.66; P = .0009), whereas the RR of all-grade fatigue was (RR = 1.02; 95% confidence interval, 0.91–1.15; P = .71).

**Conclusion:** Our meta-analysis has demonstrated that MET inhibitors-based treatment is associated with an increased risk of high-grade fatigue compared with control.

**Abbreviations:** EGFR = epidermal growth factor receptor, ESMO = European Society of Medical Oncology Conference, mAbs = monoclonal antibodies, MET = the N-methyl-N'-nitroso-guanidine human osteosarcoma transforming gene, TKI = tyrosine kinase inhibitor, VEGF = vascular endothelial growth factor.

Keywords: fatigue, MET inhibitors, meta-analysis, systematic review

## 1. Introduction

Fatigue is classically reported as a more frequently symptom than any others in patients with underlying malignancy, which causes

The involvement of MET in solid tumor has been demonstrated in preclinical studies and clinical trials suggest that MET inhibitors are likely to be critical to future combination-based therapies. Ideally, biomarkers of MET detected and the publication of more randomized trials report these toxicities will provide the basis for personalized treatment regimens.

HT and YZ equally contributed to this work.

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the most interference with function during survivorship phases.<sup>[1]</sup> Cancer-related fatigue has been defined as a distressing, persistent, subjective sense of tiredness or exhaustion that is not proportional to recent activity and exertion.<sup>[2]</sup> These toxicities of mild-to-moderate severity are frequently long-lasting, debilitating and thus, impact the functional status and decrease health-related quality of life.<sup>[3,4]</sup>

The N-methyl-N'-nitroso-guanidine human osteosarcoma transforming gene (MET) has emerged as a critical intermediate in cell-signaling pathways commonly deregulated in solid tumors. MET tyrosine kinase is the cell-surface receptor for hepatocyte growth factor (also known as scatter factor) that is involved in regulating tumorigenic growth, metastasis, and therapeutic resistance.<sup>[5]</sup> MET activity is normally detected in defined stages of embryogenesis and organogenesis.<sup>[5]</sup> Furthermore, MET pathway engages in crosstalk with the epidermal growth factor receptor (EGFR), vascular endothelial growth factor (VEGF), and Kirsten rat sarcoma viral oncogene signaling pathway, which in part explained MET amplification in various primary tumors, including medulloblastoma, colorectal cancer, hepatocellular carcinoma, gastric cancer, and NSCLC with resistance to EGFR inhibitors.<sup>[6-10]</sup> Based on this scientific rationale, targeting approaches to the Met signaling axis is mostly comprised of monoclonal antibodies (mAbs) and small-molecule tyrosine kinase inhibitors (TKIs), such as crizotinib, cabozantinib, tivantinib, onartuzumab, and rilotumumab that have been developed as MET inhibitors.<sup>[5]</sup> Presently, cabozantinib has been approved for treatment of renal cell carcinoma,

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medullary thyroid cancer, and lung adenocarcinoma, whereas crizotinib is approved for medullary thyroid carcinoma and lung cancer.<sup>[11–14]</sup>

According to several comprehensive global phase III study and a meta-analysis, patients with MET-high tumor show better survival benefit when a MET inhibitor was added into an EGFR TKI.<sup>[15–17]</sup> Besides their impressive efficacy, there have been concerns regarding the safety of MET inhibitors. The onset of severe cancer-related fatigue has been the major dose-limiting toxicity of TKIs.<sup>[18]</sup> Thus, an understanding of this risk in the context of expanding MET-inhibitors use is an important cost and patient safety issue. In this study, we reported a meta-analysis of trials of MET inhibitors in patients with cancer and compared the incidence of fatigue among the cohorts with different tumors types and between the groups treated with various types of MET inhibitors.

#### 2. Methods

#### 2.1. Data sources and study identification

A literature search utilizing PubMed, Embase, and the Cochrane Library databases was conducted from inception to October 2017, using Boolean logic that consisted of the key terms "c-Met," "MET," "cabozantinib," "onartuzumab," "rilotumumab," and "tivantinib" being searched. We did not include "crizotinib" because it was approved only for the treatment of anaplastic lymphoma kinase-positive NSCLC. Medical subject headings in Pubmed and EMTREE terms in Embase were used to identify synonyms. Meeting abstracts from the American Society of Clinical Oncology and not published full-text original articles of the European Society of Medical Oncology Conference were not eligible. Updated manufacturer's package inserts, Clinical Trials Registry Platform, and ClinicalTrials.gov was also searched to identify relevant information. Finally, references in reports of all included publications will be tracked to retrieve relevant citations. This meta-analysis was conducted according to the Preferred Reporting Items for Systematic Reviews and Meta-Analyses Statement recommendation.<sup>[19]</sup> Ethical approval for this study was unnecessary because it was a review of existing literature and did not involve any handling of individual patient data. The primary goal of the analysis was to establish any association between fatigue and exposure to MET inhibitors; therefore, trials with hormonal agents and corticosteroids use were excluded as they can modulate fatigue. Eligible trials that fulfill the following concepts were included: prospective phase II and III trials involving adults with malignancy; random assignment to MET-inhibitors-based treatment or control treatments (placebo or concurrently chemotherapeutic or best supportive care); and available data reporting safety and sample size for treatment-emergent, nondisease-related, all-grade and high-grade fatigue.

## 2.2. Date extraction and clinical end points

Data extraction was conducted independently by 2 reviewers (HT and YZ) using a predesigned, pilot-tested extraction form. For each study, the following baseline demographic and clinical characteristics were extracted from each study: the first author's name, year of publication, study design, trial phase, underlying malignancy, population size, treatment and dosing regimens, median age, median treatment duration, masking, number of

patients available for analysis, and number of relevant adverse events. Count data for fatigue were defined and recorded according to the common toxicity criteria of adverse events version 3.0 or 4.0, which are widely used in grading the relevant adverse events. Discrepancies between the reviewers were settled by consensus or revisiting original trials. Whether the toxicities attributed to treatment related or disease related depended on the decision of investigators in the original articles. Data were extracted from intention-to-treat analyses wherever possible. We also initially aimed to collect the median follow-up duration of all-grade and high-grade fatigue, only to find that this was rarely reported. If there were arms that exposed patients to different doses and schedules of MET inhibitors/control, those arms were pooled together. Study quality was assessed by using the Jadad 7-item scale including randomization, double-blinding, and withdrawals,<sup>[20]</sup> with disagreements adjudicated by a senior reviewer (YL).

#### 2.3. Statistical analysis

All statistical analyses were carried out by using RevMan (version 5.3; the Nordic Cochrane Centre, Copenhagen), STATA (version 12.0; StataCorp, College Station, TX), and the "R" statistical software version 2.15.2. The principal summary measures were incidence and relative risk (RR) of allgrade (grade 1-4) and high-grade (grade 3-4) fatigue, respectively. We first calculated the incidence of fatigue by using the total number of patients evaluable for the safety profiles and the number of patients experiencing fatigue in selected randomized clinical trials (RCTs). For RR analysis, data were extracted from prospective clinical trials, comparing the incidence of outcomes in patients assigned to MET inhibitors vs control arm in the same trial. Due to substantial clinical heterogeneity between studies such as diagnosis, random assignment, and in exposure to prior treatments, a random-effects model (DerSimonian and Laird) that considered within-study and between-study variation was applied.<sup>[21]</sup> Cochrane Q test  $(\chi^2)$  was used to qualitatively describe the heterogeneity and I<sup>2</sup> was used to quantitate the inconsistency.<sup>[22]</sup> When substantial heterogeneity was observed, prespecified exploratory subgroup analysis was conducted and stratified to examine the potential risk factors. The subgroups were only chosen for their clinical relevance and importance. A continuity correction was applied in RCTs with zero events. Presence of potential publication bias was assessed using funnel plots with Begg and Egger tests (if more than 10 studies were available).<sup>[23,24]</sup> All statistical tests were set as 2-sided, and P < .05 was considered statistically significant.

## 3. Results

#### 3.1. Literature search results

Following study inclusion (Fig. 1), our search identified 523 potentially relevant articles, describing the use of MET inhibitors from the initial databases (Tables S1–S3, http://links.lww.com/MD/D3). After eligibility assessment, 17 trials were considered to be most relevant for the meta-analysis,<sup>[3,16,25–40]</sup> (including 13 phase II trials (n=1823) and 4 phase III trials (n=3205). Literature review articles, irrelevant topics, case reports, and preclinical studies were excluded. The characteristics of trials and patients are listed in Table 1.



## 3.2. Population characteristics

A total of 5028 cancer patients were available for the metaanalysis, of which 1832 patients had NSCLC (36%), 455 (9%) had colorectal cancer, 803 (16%) had renal cancer, 1161 (23%) had prostate cancer, and 777 15%) were treated for other malignancies. All studies were well randomized with adequate follow-up. All selected trials included patients with an Eastern Cooperative Oncology Group performance status of 2 or less and were required to have an adequate renal, hepatic, and hematologic function. For the RR analysis, 2784 patients were treated with MET inhibitors (1169 patients treated with cabozantinib, 660 with onartuzumab, 737 with tivantinib, and 218 with rilotumumab), and the rest with control/placebo treatment. In tivantinib studies, all patients were given 360 mg twice daily. Onartuzumab was administered at a dose of 10 and 15 mg/kg in 2 or 3-week cycles, and rilotumumab was administered 15/10/7.5 mg/kg daily continuously in 3 trials. Cabozantinib was given 60 or 40 mg daily continuously in studies. The baselines characteristics of included studies are presented in Table 1. We graded the quality of each study reported by Jadad et al.<sup>[20]</sup>

### 3.3. Quality of studies

Randomized treatment allocation sequences were generated in all trials. The Jadad scores of the included 7 randomized, controlled,

# Table 1

Baseline characteristics of the prospective randomized trials included in the meta-analysis.

Authors/year/ phase	Histology	Patients enrolled	Treatment regimens	Median PFS/ TTP (mo)	No. for analysis	High-grade fatigue	
1 NSCLC studies							
Spigel/2016/III*	NSCLC	492	Erlotinib 150 mg qd+onartuzumab 15 mg/kg a3w	2.7	248	67	11
			Erlotinib 150 mg gd+placebo	2.6	244	74	12
Neal/2016/II <sup>+,‡</sup>	NSCLC	119	Erlotinib 150 mg+cabozantinib 40 mg gd	4.7	39	33	6
			Cabozantinib 60 mg gd	4.3	40	28	6
			Erlotinib 150 mg gd	1.8	40	28	5
Scauliotti/2015/III*	NSCI C	1037	Friotinib 150 mg gd+tivantinib 360 mg bid	3.6	520	226	47
0000g/1010/ 20 10/ 11	10020	1001	Friotinib 150 mg gd+placebo	1.9	517	197	41
Spigel/2013/II*	NSCLC	136	Erlotinib 150 mg qd+onartuzumab 15 mg/kg q3w	2.2	69	22	6
			Erlotinib 150 mg qd+placebo	2.6	67	24	2
Sequist/2011/II*	NSCLC	167	Erlotinib 150 mg qd+tivantinib 360 mg bid	3.8	84	28	4
			Erlotinib 150 mg qd+placebo	2.3	83	31	5
2. Colorectal cancer							
Bendell/2017/II*	Metastatic colorectal cancer	192	mFOLFOX-6+bevacizumab 5 mg/kg +onartuzumab 10 mg/kg q3w	11.0	99	NR	23
			mFOLFOX-6+bevacizumab 5 mg/kg q3w +placebo	10.3	93	NR	8
Eng/2016/II <sup>†</sup>	Metastatic colorectal cancer	121	Cetuximab 500 mg/m <sup>2</sup> +tivantinib 360 mg bid	8.3	62	24	3
			Cetuximab 500 mg/m <sup>2</sup> +placebo	7.3	59	20	2
Cutsem/2014/II*	Metastatic colorectal cancer	142	Rilotumumab 10 mg/kg+panitumumab 6 mg/ kg q2w	5.2	48	18	4
			Panitumumab 6 mg/kg+ganitumab 12 mg/kg g2w	5.3	46	36	2
			Panitumumab 6 mg/kg+placebo g2w	3.7	48	30	6
3. Prostate cancer							
Smith/2016/III <sup>†</sup>	Metastatic prostate cancer	1023	Cabozantinib 60 mg qd	5.6	681	347	119
			Prednisone 5 mg bid	2.8	342	120	30
Ryan/2013/II*	Prostate cancer	138	Mitoxantrone 12 mg/m <sup>2</sup> qw+prednisone 5 mg bid+rilotumumab 15/7.5 mg/kg	3.0	89	37	6
			Mitoxantrone 12 mg/m <sup>2</sup> qw+prednisone 5 mg bid+placebo	2.9	49	22	5
4. Renal cell carcinoma	1						
Choueiri/2017/II*	RCC	150	Cabozantinib 60 mg qd	8.2	78	67	5
			Sunitinib 50 mg qd	5.6	72	59	11
Choueiri/2016/II <sup>*,‡</sup>	RCC	653	Cabozantinib 60 mg qd	7.4	331	195	36
			Everolimus 10 mg qd	3.9	322	154	24
5. Adenocarcinoma of t	the stomach						
Shah/2016/II*	Adenocarcinoma of the Stomach	120	mFOLFOX-6+onartuzumab 10 mg/kg q3w	6.77	60	26	NR
			mFOLFOX-6 q3w+placebo	6.97	60	33	NR
lveson/2014/ll*	Adenocarcinoma of the Stomach	120	Rilotumumab 15/7.5 mg/kg+ECX q3w	5.7	81	29	11
			ECX q3w+placebo	4.2	39	13	6
6. Other cancer							
Cloughesy/2016/II*	Glioblastoma	129	Onartuzumab 15 mg/kg q2w+bevacizumab 15 mg/kg q2w	3.9	64	10	NR
			Bevacizumab 15 mg/kg q2w+placebo	2.9	65	13	NR
Diéras/ 2015/II <sup>†</sup>	Breast cancer	182	Paclitaxel 90 mg/m <sup>2</sup> qw+onartuzumab 10 mg/kg q2w+bevacizumab 10 mg/kg q2w	7.3	62	29	NR
			Paclitaxel 90 mg/m <sup>2</sup> qw+onartuzumab 10 mg/kg q2w+placebo	5.4	58	24	NR
			Paclitaxel 90 mg/m <sup>2</sup> qw+bevacizumab 10 mg/kg q2w+placebo	7.2	62	30	NR
Santoro/2013/II <sup>†</sup>	Hepatocellular carcinoma	107	Tivantinib 360 mg bid	1.5	71	7	3
			Placebo	1.4	36	5	1

bid = twice every day, ECX = (epirubicin 50 mg/m<sup>2</sup> intravenously on day 1, cisplatin 60 mg/m<sup>2</sup> intravenously on day 1, capecitabine 625 mg/m<sup>2</sup> twice a day orally on days 1–21, respectively), administered every 3 weeks, mF0LF0X-6 = (oxaliplatin [85 mg/m<sup>2</sup> IV], 5-FU [400 mg/m<sup>2</sup> IV bolus], and LV [400 mg/m<sup>2</sup> IV]) and bevacizumab (5 mg/kg IV) every 2 weeks, NR = not reported, NSCLC = nonsmall cell lung cancer, PFS=progression-free survival, q2w=once every 2 weeks, q3w=once every 3 weeks, qd=once every day, RCC=renal cell carcinoma, TTP=time to tumor progression.

Randomized trial open label. <sup>†</sup> Randomized controlled trials (RCTs).

<sup>‡</sup> Open label.

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and double-blinded trials ranged from 4 to 5,<sup>[27,30,32,37–40]</sup> whereas there were 2 phase II, randomized, open-labeled studies that were assigned a Jadad score of 3.<sup>[26,35]</sup> Therefore, quality of these studies was fair and acceptable.

#### 3.4. Overall incidence of all-grade and high-grade fatigue

For the incidence analysis, we considered arms receiving MET inhibitors or MET inhibitor-based combination (due to the limited use of MET inhibitors monotherapy treatment and potential events associated with concomitant therapy). Thus, a total of 2685 patients from 16 trials were available from the included trials. The summary incidences of all-grade fatigue ranged between 10% and 86%, with

the highest incidence seen in a phase III trial among patients with colorectal cancer.<sup>[34]</sup> Using a random-effects model, the calculated overall incidence of all-grade fatigue among patients receiving MET inhibitors was 41.9% (95% confidence interval, CI, 34.6–49.6) (I<sup>2</sup> = 92.1%, P < .01), whereas the overall incidences of high-grade fatigue ranged between 4% and 23%, and the highest incidence was observed in a trial of patients with colorectal cancer.<sup>[29]</sup> The calculated overall incidence of high-grade fatigue among patients receiving MET inhibitors was 9.6% (95% CI, 7.1–12.8) (I<sup>2</sup> = 79.5%, P < .01) (Fig. 2). Further exploratory analysis was conducted to evaluate the incidence of fatigue based on underlying malignancy and type of MET inhibitors. Data stratified by tumor type showed that the incidence of all-grade fatigue was higher in

Study	Events	Total		Proportion	95%-CI	(fixed)	(randor
Sequist(2011)	28	84	<del></del> ;	0.33	[0.23; 0.44]	3.1%	6.3
Ryan(2013)	37	89		0.42	[0.31; 0.53]	3.5%	6.4
Santoro(2013)	7	71	II	0.10	[0.04; 0.19]	1.0%	5.0
Spigel(2013)	22	69		0.32	[0.21: 0.44]	2.5%	6.19
lveson(2014)	29	81	<u> </u>	0.36	[0.25: 0.47]	3.1%	6.3
Cutsem(2014)	18	48		0.38	[0.24; 0.53]	1.8%	5.8
Diéras(2015)	53	120	<u> </u>	0.44	[0.35; 0.54]	4.9%	6.6
Scagliotti(2015)	226	520		0.43	[0.39: 0.48]	21.0%	7.19
Spigel(2016)	67	248	- <u></u>	0.27	[0.22: 0.33]	8.0%	6.9
Cloughesy(2016)	10	64	<u>→</u>	0.16	[0.08: 0.27]	1.4%	5.4
Shah(2016)	26	60	<u> </u>	0.43	10.31: 0.571	2.4%	6.19
Eng(2016)	24	62	<b>*</b>	0.39	[0.27: 0.52]	2.4%	6.19
Neal(2016)	61	79	<b>→</b>	0.77	[0.66; 0.86]	2.3%	6.0
Smith(2016)	347	681	1	0.51	[0.47: 0.55]	27.9%	7.2
Choueiri(2016)	195	331	- <u></u>	0.59	[0.53: 0.64]	13.2%	7.0
Choueiri(2017)	67	78		- 0.86	[0.76; 0.93]	1.6%	5.6
			i.	0.46	[0.44; 0.48]	100.0%	
Fixed effect model Random effects model Heterogeneity: $I^2 = 92\%$ , $\tau^2$	<sup>2</sup> = 0.3410	2085 , p < 0.	01 0.2 0.4 0.6 0.8	0.42	[0.35; 0.50]		100.09
Fixed effect model Random effects model Heterogeneity: $I^2 = 92\%$ , $\tau^2$	<sup>2</sup> = 0.3410	2085 , p < 0.	01 0.2 0.4 0.6 0.8	0.42	[0.35; 0.50]	 Weight	100.09 Weigl
Fixed effect model Random effects model Heterogeneity: I <sup>2</sup> = 92%, τ <sup>2</sup>	<sup>2</sup> = 0.3410 Events	2085 , p < 0. Total	01 0.2 0.4 0.6 0.8	0.42 0.42 Proportion	[0.35; 0.50] 95%-CI	 Weight (fixed)	100.09 Weigl (randon
Fixed effect model Random effects model Heterogeneity: / <sup>2</sup> = 92%, τ <sup>2</sup> Study Sequist(2011)	<sup>2</sup> = 0.3410 Events 4	2085 , p < 0. Total 84	01 0.2 0.4 0.6 0.8	0.42 Proportion	[0.35; 0.50] 95%-CI [0.01; 0.12]	 Weight (fixed) 1.5%	100.09 Weigl (randon 5.3
Fixed effect model Random effects model Heterogeneity: $I^2 = 92\%$ , $\tau^2$ Study Sequist(2011) Ryan(2013)	<sup>2</sup> = 0.3410 Events 4 6	2085 , <i>p</i> < 0. Total 84 89		0.42 Proportion 0.05 0.07	<b>95%-CI</b> [0.01; 0.12] [0.03; 0.14]	 Weight (fixed) 1.5% 2.2%	100.09 Weigl (randon 5.3 6.3
Fixed effect model Random effects model Heterogeneity: $I^2 = 92\%$ , $\tau^2$ Study Sequist(2011) Ryan(2013) Santoro(2013)	<sup>2</sup> = 0.3410 Events 4 6 3	2085 , <i>p</i> < 0. Total 84 89 71		0.42 Proportion 0.05 0.07 0.04	<b>95%-CI</b> [0.01; 0.12] [0.03; 0.14] [0.01; 0.12]		100.09 Weigh (randon 5.3 6.3 4.6
Fixed effect model Random effects model Heterogeneity: $I^2 = 92\%$ , $\tau^2$ Study Sequist(2011) Ryan(2013) Santoro(2013) Spigel(2013)	<sup>2</sup> = 0.3410 Events 4 6 3 6	2085 , <i>p</i> < 0. Total 84 89 71 69		0.42 Proportion 0.05 0.07 0.04 0.09	<b>95%-Cl</b> [0.01; 0.12] [0.01; 0.12] [0.01; 0.12] [0.03; 0.18]		100.09 Weigh (random 5.3° 6.3° 4.6° 6.3°
Fixed effect model Random effects model Heterogeneity: $I^2 = 92\%$ , $\tau^2$ Study Sequist(2011) Ryan(2013) Santoro(2013) Spigel(2013) Iveson(2014)	<sup>2</sup> = 0.3410 Events 4 6 3 6 11	2085 , <i>p</i> < 0. Total 84 89 71 69 81		0.42 Proportion 0.05 0.07 0.04 0.09 0.14	<b>95%-Cl</b> [0.01; 0.12] [0.03; 0.14] [0.03; 0.18] [0.07; 0.23]		100.09 Weigh (randon 5.3° 6.3° 4.6° 6.3° 7.6°
Fixed effect model Random effects model Heterogeneity: $I^2 = 92\%$ , $\tau^2$ Study Sequist(2011) Ryan(2013) Santoro(2013) Spigel(2013) lveson(2014) Cutsem(2014)	<sup>2</sup> = 0.3410 Events 4 6 3 6 11 4	2085 , p < 0. Total 84 89 71 69 81 48		0.42 Proportion 0.05 0.07 0.04 0.09 0.14 0.08	<b>95%-Cl</b> [0.01; 0.12] [0.01; 0.12] [0.03; 0.14] [0.01; 0.12] [0.03; 0.18] [0.07; 0.23] [0.02; 0.20]	 Weight (fixed) 1.5% 2.2% 1.1% 2.2% 3.8% 1.5%	100.09 Weigh (random 5.3° 6.3° 4.6° 6.3° 7.6° 5.2°
Fixed effect model Random effects model Heterogeneity: $I^2 = 92\%$ , $\tau^2$ Study Sequist(2011) Ryan(2013) Santoro(2013) Spigel(2013) lveson(2014) Cutsem(2014) Scagliotti(2015)	<sup>2</sup> = 0.3410 Events 4 6 3 6 11 4 4 47	<b>Total</b> 84 89 71 69 81 48 520		0.42 Proportion 0.05 0.07 0.04 0.09 0.14 0.08 0.09	<b>95%-Cl</b> [0.01; 0.12] [0.01; 0.12] [0.03; 0.14] [0.01; 0.12] [0.03; 0.18] [0.07; 0.23] [0.02; 0.20] [0.07; 0.12]		100.09 Weigh (random 5.3° 6.3° 4.6° 6.3° 7.6° 5.2° 9.9°
Fixed effect model Random effects model Heterogeneity: $I^2 = 92\%$ , $\tau^2$ Study Sequist(2011) Ryan(2013) Santoro(2013) Spigel(2013) lveson(2014) Cutsem(2014) Scagliotti(2015) Spigel(2016)	<sup>2</sup> = 0.3410 Events 4 6 3 6 11 4 47 11	<b>Total</b> 84 89 71 69 81 48 520 248		0.42 Proportion 0.05 0.07 0.04 0.09 0.14 0.08 0.09 0.04	<b>95%-Cl</b> [0.35; 0.50] [0.01; 0.12] [0.03; 0.14] [0.01; 0.12] [0.03; 0.18] [0.07; 0.23] [0.02; 0.20] [0.07; 0.12] [0.02; 0.08]		100.09 Weigl (random 5.3 6.3 4.6 6.3 7.6 5.2 9.9 7.8
Fixed effect model Random effects model Heterogeneity: $I^2 = 92\%$ , $\tau^2$ Study Sequist(2011) Ryan(2013) Santoro(2013) Spigel(2013) lveson(2014) Cutsem(2014) Scagliotti(2015) Spigel(2016) Eng(2016)	<sup>2</sup> = 0.3410 <b>Events</b> 4 6 3 6 11 4 47 11 3	<b>Total</b> 84 89 71 69 81 48 520 248 62		0.42 Proportion 0.05 0.07 0.04 0.09 0.14 0.08 0.09 0.04 0.09 0.04 0.05	<b>95%-Cl</b> [0.35; 0.50] [0.01; 0.12] [0.03; 0.14] [0.01; 0.12] [0.03; 0.18] [0.07; 0.23] [0.02; 0.20] [0.07; 0.12] [0.02; 0.08] [0.01; 0.13]		100.09 Weigh (random 5.3' 6.3' 4.6' 6.3' 7.6' 5.2' 9.9' 7.8' 4.5'
Fixed effect model Random effects model Heterogeneity: $I^2 = 92\%$ , $\tau^2$ Study Sequist(2011) Ryan(2013) Santoro(2013) Spigel(2013) lveson(2014) Cutsem(2014) Scagliotti(2015) Spigel(2016) Eng(2016) Neal(2016)	<sup>2</sup> = 0.3410 <b>Events</b> 4 6 3 6 11 4 47 11 3 12	<b>Total</b> 84 89 71 69 81 48 520 248 62 79		0.42 Proportion 0.05 0.07 0.04 0.09 0.14 0.08 0.09 0.04 0.05 0.05 0.15	<b>95%-Cl</b> [0.35; 0.50] [0.01; 0.12] [0.03; 0.14] [0.01; 0.12] [0.03; 0.18] [0.07; 0.23] [0.02; 0.20] [0.07; 0.12] [0.02; 0.08] [0.01; 0.13] [0.08; 0.25]		100.09 Weigh (random 5.3 6.3 4.6 6.3 7.6 5.2 9.9 7.8 4.5 7.8
Fixed effect model Random effects model Heterogeneity: $I^2 = 92\%$ , $\tau^2$ Study Sequist(2011) Ryan(2013) Santoro(2013) Spigel(2013) lveson(2014) Cutsem(2014) Scagliotti(2015) Spigel(2016) Eng(2016) Smith(2016)	<sup>2</sup> = 0.3410 <b>Events</b> 4 6 3 6 11 4 47 11 3 12 119	<b>Total</b> 84 89 71 69 81 48 520 248 62 79 681		0.42 Proportion 0.05 0.07 0.04 0.09 0.14 0.08 0.09 0.04 0.05 0.15 0.17	<b>95%-Cl</b> [0.35; 0.50] [0.01; 0.12] [0.03; 0.14] [0.01; 0.12] [0.03; 0.18] [0.07; 0.23] [0.02; 0.20] [0.07; 0.12] [0.02; 0.08] [0.01; 0.13] [0.08; 0.25] [0.15; 0.21]		100.09 Weigh (random 5.3' 6.3' 4.6' 6.3' 7.6' 5.2' 9.9' 7.8' 4.5' 7.8' 10.3'
Fixed effect model Random effects model Heterogeneity: $I^2 = 92\%$ , $\tau^2$ Study Sequist(2011) Ryan(2013) Santoro(2013) Spigel(2013) lveson(2014) Cutsem(2014) Scagliotti(2015) Spigel(2016) Eng(2016) Smith(2016) Choueiri(2016)	<sup>2</sup> = 0.3410 <b>Events</b> 4 6 3 6 11 4 47 11 3 12 119 36	<b>Total</b> 84 89 71 69 81 48 520 248 62 79 681 331		0.42 Proportion 0.05 0.07 0.04 0.09 0.14 0.08 0.09 0.04 0.05 0.15 0.17 0.11	<b>95%-Cl</b> [0.35; 0.50] [0.01; 0.12] [0.03; 0.14] [0.01; 0.12] [0.03; 0.18] [0.07; 0.23] [0.02; 0.20] [0.07; 0.12] [0.02; 0.08] [0.01; 0.13] [0.08; 0.25] [0.15; 0.21] [0.08; 0.15]		100.09 Weigh (random 5.3' 6.3' 4.6' 6.3' 7.6' 5.2' 9.9' 7.8' 4.5' 7.8' 10.3' 9.6'
Fixed effect model Random effects model Heterogeneity: $I^2 = 92\%$ , $\tau^2$ Study Sequist(2011) Ryan(2013) Santoro(2013) Spigel(2013) Meson(2014) Cutsem(2014) Scagliotti(2015) Spigel(2016) Eng(2016) Neal(2016) Smith(2016) Choueiri(2017)	<sup>2</sup> = 0.3410 <b>Events</b> 4 6 3 6 11 4 47 11 3 12 119 36 5	<b>Total</b> 84 89 71 69 81 48 520 248 62 79 681 331 78		0.42 Proportion 0.05 0.07 0.04 0.09 0.14 0.08 0.09 0.04 0.05 0.15 0.17 0.11 0.06	<b>95%-Cl</b> [0.35; 0.50] [0.01; 0.12] [0.03; 0.14] [0.01; 0.12] [0.03; 0.18] [0.07; 0.23] [0.02; 0.20] [0.07; 0.12] [0.02; 0.08] [0.01; 0.13] [0.08; 0.25] [0.15; 0.21] [0.08; 0.15] [0.02; 0.14]		100.0 Weigl (randor 5.3 6.3 4.6 6.3 7.6 5.2 9.9 7.8 4.5 7.8 10.3 9.6 5.9
Fixed effect model Random effects model Heterogeneity: $I^2 = 92\%$ , $\tau^2$ Study Sequist(2011) Ryan(2013) Santoro(2013) Spigel(2013) lveson(2014) Cutsem(2014) Scagliotti(2015) Spigel(2016) Eng(2016) Smith(2016) Choueiri(2017) Bendell	<sup>2</sup> = 0.3410 <b>Events</b> 4 6 3 6 11 4 47 11 3 12 119 36 5 23	<b>Total</b> 84 89 71 69 81 48 520 248 62 79 681 331 78 99		0.42 Proportion 0.05 0.07 0.04 0.09 0.14 0.08 0.09 0.04 0.05 0.15 0.17 0.11 0.06 - 0.23	<b>95%-Cl</b> [0.35; 0.50] [0.01; 0.12] [0.03; 0.14] [0.01; 0.12] [0.03; 0.18] [0.07; 0.23] [0.02; 0.20] [0.07; 0.12] [0.02; 0.08] [0.01; 0.13] [0.08; 0.25] [0.15; 0.21] [0.08; 0.15] [0.02; 0.14] [0.15; 0.33]	 Weight (fixed) 1.5% 2.2% 3.8% 1.5% 17.1% 4.2% 1.1% 4.2% 1.1% 39.3% 12.8% 1.9% 7.1%	100.04 Weigl (random 5.3 6.3 4.6 6.3 7.6 5.2 9.9 7.8 4.5 7.8 10.3 9.6 5.9 8.8

Figure 2. Incidence of all-grade and high-grade fatigue.

RCC patients (74%), whereas high-grade fatigue was higher in gastric or esophagogastric junction adenocarcinoma (14%). When stratified by MET inhibitors, the incidence of all-grade and high-grade fatigue was higher in patients using cabozantinib (68% and 13%, respectively). The incidence of severity adverse events varied significantly with cancer type (P < .001) and type of MET inhibitors (P < .001) (Figs. S1–S3, http://links.lww.com/MD/D3).

## 3.5. RR of all-grade and high-grade fatigue

To evaluate the specific contribution of MET inhibitors to the development of fatigue, 16 and 14 randomized trials were available

to calculate the RR of all-grade and high-grade fatigue in patients who received MET inhibitors or controls (concurrent chemotherapy or placebo or targeted therapy). The summary of RR of developing all-grade and high-grade fatigue with the administration of MET inhibitors vs controls was 1.02 (95% CI, 0.91–1.15; P=.71) and 1.37 (95% CI, 1.14–1.66; P=.0009), respectively (Figs. 3–6). This estimate was obtained by using random-effects model and fixed-effects model as the heterogeneity tested by I<sup>2</sup> statistics in the increased risk of all-grade fatigue ( $\chi^2$ =39.94, P=.0005, I<sup>2</sup>=62%) and high-grade fatigue ( $\chi^2$ =18.58, P=.14, I<sup>2</sup>=30%). Considering the confounding factors, a sensitivity

	Experim	iental	Contr	ol		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight M-H	, Random, 95% Cl	M-H, Random, 95% Cl
3.1.1 NSCLC							
Neal 2017	61	79	23	40	7.2%	1.34 [1.00, 1.80]	
Scagliotti 2015	226	520	197	517	10.6%	1.14 [0.98, 1.32]	-
Sequist 2011	28	84	31	83	5.0%	0.89 [0.59, 1.35]	
Spigel 2013	22	69	24	67	4.2%	0.89 [0.56, 1.43]	2.4 17
Spigel 2016	67	248	74	244	7.5%	0.89 [0.67, 1.18]	-
Subtotal (95% CI)		1000		951	34.5%	1.06 [0.91, 1.24]	•
Total events	404		349				
Heterogeneity: Tau <sup>2</sup> =	0.01; Chi	= 6.06,	df = 4 (P)	= 0.19)	; I= 34%		
Fest for overall effect	Z= 0.76 (	P = 0.45	)	C	a valore car		
3.1.2 Prostate cance	F						
Ryan 2013	37	89	22	49	5.3%	0.93 [0.62, 1.38]	
Smith 2016	347	681	120	342	10.2%	1.45 [1.24, 1.71]	+
Subtotal (95% CI)		770		391	15.5%	1.20 [0.78, 1.86]	*
Fotal events	384		142			200 00 00	
Heterogeneity: Tau <sup>2</sup> =	0.08: Chi	<sup>2</sup> = 4.26.	df = 1 (P)	= 0.04)	$ ^{2} = 77\%$		
Test for overall effect	7 = 0.84 (	P = 0.40	)				
	2-0.010	0.10	/				
3.1.3 Gastrointestina	l cancer						
Cutsem 2014	18	48	66	94	5 4 %	0.53 (0.36, 0.79)	
=ng 2016	24	62	20	59	4 2%	1 14 [0 71 1 84]	
vacon 2014	29	01	12	20	2.6%	1 07 0 62 1 92	
2630112014 2636 2016	25	60	22	80	5.0%	0.79 [0.63, 1.63]	
Subtotal (95% CI)	20	251	33	252	19.0%	0.22 [0.59, 1.14]	•
Total quanta	07	231	100	LJL	10.370	0.02 [0.30, 1.10]	
I utar events	9/	2-754	132	- 0.00	17- 000		
Heterogeneity. Tau-=	7 4 4 0 /	T= 7.51,		= 0.06)	1, 1- = 00%		
rest for overall ellect	2=1.10(	P=0.27	,				
1 4 DCC							
Dhavairi 2016	105	224	101	222	40.000	4 22 14 07 4 421	+
chouein 2016	195	331	154	322	10.0%	1.23 [1.07, 1.42]	L.
Choueiri 2017	67	18	59	12	10.7%	1.05 [0.91, 1.21]	
Subtotal (95% CI)	000	409	240	394	21.470	1.14 [0.95, 1.55]	
l otal events	262		213				
Heterogeneity: Tau*=	0.01; Chr	= 3.05,	df = 1 (P	= 0.08)	; 1*= 67%		
l est for overall effect	Z=1.41 (	P = 0.16	)				
A E Other maliana	ala a						
5.1.5 Other malignar	icies		10				
Cloughesy 2016	10	64	13	65	2.1%	0.78 [0.37, 1.65]	
Dieras 2015	53	120	30	62	6.5%	0.91 [0.66, 1.27]	
Santoro 2012	7	71	5	36	1.1%	0.71 [0.24, 2.08]	
subtotal (95% CI)	10000	255		163	9.7%	0.88 [0.66, 1.17]	
Fotal events	70		48				
Heterogeneity: Tau <sup>2</sup> =	= 0.00; Chi	<sup>2</sup> = 0.31,	df = 2 (P	= 0.86)	; I² = 0%		
Fest for overall effect	Z= 0.90 (	P = 0.37	)				
		Carries -		2225	1990/200	1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1	
		2685		2151	100.0%	1.02 [0.91, 1.15]	
fotal (95% CI)			004				
Fotal (95% CI) Fotal events	1217		884				
Total (95% CI) Total events Heterogeneity: Tau² :	1217 0.03; Chi <sup>a</sup>	<b>*</b> = 39.94	884 , df = 15	(P = 0.1	0005); F= 629	6	
Total (95% CI) Total events Heterogeneity: Tau <sup>2</sup> = Fest for overall effect	1217 = 0.03; Chi <sup>a</sup> Z = 0.38 (1	² = 39.94 P = 0.71	884 l,df=15 )	(P = 0.1	0005); I² = 629	6	0.01 0.1 1 10 100 Eavours (experimental) Eavours (control)

	Experim	ental	Contr	ol		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% Cl
5.1.1 NSCLC							
Neal 2017	12	79	5	40	3.8%	1.22 [0.46, 3.21]	
Scagliotti 2015	47	520	41	517	23.6%	1.14 [0.76, 1.70]	
Sequist 2011	4	84	5	83	2.9%	0.79 [0.22, 2.84]	
Spigel 2013	6	69	2	67	1.2%	2.91 [0.61, 13.93]	
Spigel 2016	11	248	12	244	6.9%	0.90 [0.41, 2.00]	
Subtotal (95% CI)		1000		951	38.4%	1.13 [0.83, 1.55]	•
Total events	80		65				
Heterogeneity: Chi <sup>2</sup> =	2.04, df=	4 (P = 0	.73);  = 1	0%			
Test for overall effect:	Z= 0.77 (	P = 0.44	)				
5.1.2 Prostate cance	r						
Ryan 2013	6	89	5	49	3.7%	0.66 [0.21, 2.05]	
Smith 2016	119	681	30	342	22.9%	1.99 [1.36, 2.91]	+
Subtotal (95% CI)		770		391	26.6%	1.81 [1.27, 2.58]	•
Total events	125		35				
Heterogeneity: Chi <sup>2</sup> =	3.28, df=	1 (P = 0)	.07); <b>F</b> = I	69%			
Test for overall effect:	Z= 3.26 (	P = 0.00	11)				
5.1.3 Gastrointestina	I cancer						
Bendell 2017	23	99	8	93	4.7%	2.70 [1.27, 5.74]	
Cutsem 2014	4	48	8	94	3.1%	0.98 [0.31, 3.09]	
Eng 2016	3	62	2	59	1.2%	1.43 [0.25, 8.24]	
lveson 2014	11	81	6	39	4.6%	0.88 [0.35, 2.21]	
Subtotal (95% CI)		290		285	13.7%	1.58 [0.98, 2.56]	◆
Total events	41		24				
Heterogeneity: Chi <sup>2</sup> =	4.17, df=	3 (P = 0	.24); F= :	28%			
Test for overall effect:	Z=1.86 (	P = 0.06	)				
5.1.4 RCC							
Choueiri 2016	36	331	24	322	14.0%	1.46 [0.89, 2.39]	
Choueiri 2017	5	78	11	72	6.6%	0.42 [0.15, 1.15]	
Subtotal (95% CI)		409		394	20.5%	1.13 [0.73, 1.73]	<b>+</b>
Total events	41		35				
Heterogeneity: Chi <sup>2</sup> =	4.75, df=	1 (P = 0	.03); P= 1	79%			
Test for overall effect:	Z= 0.54 (	P = 0.59	))				
5.1.5 Other malignan	cies						
Santoro 2012	3	71	1	36	0.8%	1.52 [0.16, 14.11]	
Subtotal (95% CI)		71		36	0.8%	1.52 [0.16, 14.11]	
Total events	3		1				
Heterogeneity: Not ap	plicable						
Test for overall effect:	Z= 0.37 (	P = 0.71	)				
Total (95% CI)		2540		2057	100.0%	1.37 [1.14, 1.66]	•
Total events	290		160				
Heterogeneity: Chi <sup>2</sup> =	18.58, df =	= 13 (P =	= 0.14); P	= 30%			
Test for overall effect:	Z= 3.31 (	P = 0.00	09)				Favours [experimental] Favours [control]
Test for subgroup diff	erences: (	$Chi^2 = 4.$	87, df = 4	(P = 0	.30), I <sup>2</sup> = 1	7.9%	terrare federational terrare featured
	-						

Figure 4. Forest plots for subgroup for the risk of high-grade fatigue stratified by tumor type.

analysis was performed to examine the stability and reliability of pooled RRs. The result indicated that the pooled RRs were not materially altered by changing the pooling models and by omitting any single study from the sequence (Fig. S4, http://links.lww.com/ MD/D3). Prespecified subgroup analysis was then conducted to explore the potential cause of heterogeneity.

# 3.6. Subgroup analyses

To examine whether the observed RRs of developing all-grade and high-grade fatigue were the result of confounding bias, prespecified subgroup analyses were performed according to the underlying malignancy, type of MET inhibitors, and controlled therapy. As suggested in the Cochrane Handbook, we tried to limit the number of subgroup analyses.

#### 3.7. Influence of tumor type

Different tumor biological behavior and associated treatment might affect the risk of fatigue. We thus evaluated RRs of all-grade and high-grade fatigue with MET inhibitors according to tumor types (Figs. 3 and 4). For all-grade fatigue, there was no significantly increased risk in the included trials, 1.02 (95% CI, 0.91–1.15; P=.71). Interestingly, the effect size was significantly greater in the

	Experime	ental	Contr	ol		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H, Random, 95% Cl
4.1.1 rilotumumab							
Cutsem 2014	18	48	66	94	5.4%	0.53 [0.36, 0.79]	
lveson 2014	29	81	13	39	3.6%	1.07 [0.63, 1.83]	
Rvan 2013	37	89	22	49	5.3%	0.93 [0.62, 1.38]	
Subtotal (95% CI)	0.20	218	1.756	182	14.3%	0.79 [0.52, 1.21]	-
Total events	84		101				124
Heterogeneity: Tau <sup>2</sup> =	= 0.09 Chi2	= 5 75	df = 2 P	= 0.06	P = 65%		
Test for overall effect:	Z= 1.07 (F	P = 0.28	i)	0.00	n 00 /		
4.1.2 onartuzumab							
Cloughesv 2016	10	64	13	65	21%	0.78 (0.37, 1.65)	
Dieras 2015	53	120	30	62	6.5%	0.91 [0.66 1.27]	
Shoh 2016	26	60	22	60	5 7%	0 70 [0.54 1 14]	
Shan 2010 Spigol 2012	20	60	24	67	4 206	0.00 [0.54, 1.14]	
Opigel 2015	22	240	24	244	4.270	0.09 [0.00, 1.43]	
Subtotal (05% Ch	0/	248	14	244	26.0%	0.89 [0.07, 1.18]	•
Subtoral (95% CI)	470	201		490	20.0%	0.87 [0.75, 1.05]	
Total events	1/8		1/4				
Heterogeneity: Tau*=	= 0.00; Chi*	= 0.48,	df = 4 (P	= 0.98	);  * = 0%		
Test for overall effect:	: Z= 1.66 (F	<sup>2</sup> = 0.10	0				
4.1.3 tivantinib							
Eng 2016	24	62	20	59	4.2%	1.14 [0.71, 1.84]	
Santoro 2012	7	71	5	36	1.1%	0.71 [0.24, 2.08]	
Scagliotti 2015	226	520	197	517	10.6%	1.14 [0.98, 1.32]	
Sequist 2011	28	84	31	83	5.0%	0.89 (0.59, 1.35)	
Subtotal (95% CI)	: 20	737	222	695	21.0%	1.10 [0.97, 1.26]	*
Total events	285		253			100 12 100	
Heterogeneity Tau <sup>2</sup> =	= 0.00° Chi <sup>2</sup>	= 1.89	df = 3 (P)	= 0.60	$1^{\circ} I^{2} = 0.96$		
Test for overall effect:	Z = 1.47 (F	P = 0.14	() ()	- 0.00			
4.1.4 cabozantinih							
Choueiri 2016	195	331	154	322	10.6%	1 23 [1 07 1 42]	
Choueiri 2017	67	79	59	72	10.7%	1 05 [0 91 1 21]	
Neal 2017	61	70	22	40	7 2%	1 34 [1 00 1 00]	
Smith 2016	347	691	120	340	10.2%	1 45 [1 24 1 74]	
Subtotal (05% CI)	347	1160	120	776	38 8%	1 25 [1 06 1 49]	•
Total evente	670	1109	250	110	30.070	1.25 [ 1.00, 1.40]	
Hotorogonoity Tou?-	0.02: 01:2	- 11 -	300 df= 2.4	- 0.0	11-12- 70	04	
Test for overall effect:	Z = 2.59 (F	P = 0.01	o, ui – s (r 0)	0.0	1),1 = 73	70	
Total (95% CI)		2685		2151	100.0%	1 02 [0 01 1 15]	•
Total evente	1017	2005	004	2131	100.070	102 [0.01, 1.10]	T
Hotorogonolity Tau?	- 0.02: OK	- 20.0	004	/D = 0	0005-12	- 600	
Heterogeneity: Tau*=	- 0.03; Chi*	= 39.94	4, ui = 15	(r=0.	0005); 1*=	- 02%	0.2 0.5 1 2 5
Test for overall effect:	∠= 0.38 (F	= 0.71	100 11	2.00	0.041	70.40	Favours [experimental] Favours [control]
lest for subgroup dif	terences: C	$n_{1} = 1$	1.26, df =	3 (P =	0.01), 1*=	73.4%	
Fi	oure 5 Fo	rest nlc	ts for sub	aroun	for the ris	sk of all-grade fatigue acc	cording to the type of MET inhibitors

RR analysis of high-grade fatigue considering only the trials in prostate cancer than other cancers, with RRs of 1.81 (95% CI, 1.27-2.58; P=.001) and 1.22 (95% CI, 0.97-1.52; P=.08).

## 3.8. Influence of MET inhibitors

The RRs of all-grade and high-grade fatigue events might be related to MET inhibitors (Figs. 5 and 6). Our results demonstrated that the use of cabozantinib significantly increased the risk of all-grade fatigue (RR=1.25; 95% CI, 1.06–1.48; P=.01). Additionally, there was significant increased risk of high-grade fatigue among patients who received onartuzumab (RR= 1.75; 95% CI, 1.06–2.88; P=.03) and cabozantinib (RR=1.55; 95% CI, 1.19–2.04; P=.001). The use of other MET inhibitors did not increase the risk of all-grade fatigue (RR=0.91; 95% CI, 0.79–1.04; P=0.17) and high-grade fatigue (RR=1.04; 95% CI, 0.76–1.43; P=.41).

#### 3.9. Influence of controlled therapy

As an exploratory analysis, patients were stratified according to the type of treatment (MET inhibitors monotherapy vs METbased combination; Fig. S3, http://links.lww.com/MD/D3). We found no statistically significant difference between the 2 groups regarding all-grade fatigue (P=.05). With regard to RRs of highgrade events, we found a significant difference between MET inhibitor monotherapy vs control (RR=1.58; 95% CI, 1.20– 2.09; P=.001), whereas no differences between RR of MET inhibitor combination vs control were found (P=.14).

#### 3.10. Publication bias

For the meta-analysis of all-grade fatigue, visual examination of Begg funnel plot indicated a degree of asymmetry, confirmed by Egger test (P=.025). No evidence of publication bias was observed in any category of the toxicities of high-grade fatigue.

	Experim	ental	Contr	ol		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% Cl
6.1.1 rilotumumab							
Cutsem 2014	4	48	8	94	3.1%	0.98 [0.31, 3.09]	
lveson 2014	11	81	6	39	4.6%	0.88 [0.35, 2.21]	
Ryan 2013	6	89	5	49	3.7%	0.66 [0.21, 2.05]	
Subtotal (95% CI)		218		182	11.5%	0.84 [0.46, 1.53]	+
Total events	21		19				
Heterogeneity: Chi <sup>2</sup> =	0.25, df=	2 (P = 0	.88);  2 =	0%			
Test for overall effect:	Z= 0.58 (I	P = 0.56	)				
6 4 2 enerturnumah							
0.1.2 onartuzuman	22	00		00	1 700	2 70 14 27 6 741	Second Second
Bendell 2017	23	99	8	93	4.7%	2.70[1.27, 5.74]	
Spiger 2013	0	09	42	244	1.2%	2.91 [0.61, 13.93]	
Spiger 2016 Subtotal (05% CD	11	248	12	244	13.9%	0.90 [0.41, 2.00]	
Subtotal (95% CI)	40	410	22	404	12.070	1.75[1.00, 2.00]	•
I otar events	40	2 /0 - 0	113.12-1	5 4 04			
Test for everall effect:	4.33, 01=	2(P=0	.11), [*=:	0470			
restior overall ellect.	2 = 2.19 (1	r = 0.03	9				
6.1.3 tivantinib							
Eng 2016	3	62	2	59	1.2%	1.43 [0.25, 8.24]	
Santoro 2012	3	71	1	36	0.8%	1.52 [0.16, 14.11]	
Scagliotti 2015	47	520	41	517	23.6%	1.14 [0.76, 1.70]	
Sequist 2011	4	84	5	83	2.9%	0.79 [0.22, 2.84]	
Subtotal (95% CI)		737		695	28.4%	1.13 [0.78, 1.63]	<b>•</b>
Total events	57		49				
Heterogeneity: Chi <sup>2</sup> =	0.44, df=	3 (P = 0	.93);  = 1	0%			
Test for overall effect:	Z=0.63 (	P = 0.53	)				
6 1 4 cabozantinih							
Choueiri 2016	36	331	24	322	14 0%	1 46 10 90 2 301	
Choueiri 2017	5	78	11	72	6.6%	0.42 [0.05, 2.35]	
Neal 2017	12	79	5	40	3.8%	1 22 [0.15, 1.15]	n
Smith 2016	119	681	30	342	22.9%	1 99 [1 36 2 91]	
Subtotal (95% CI)	110	1169		776	47.3%	1.55 [1.19, 2.04]	<ul> <li>له</li> </ul>
Total events	172		70				
Heterogeneity: Chi <sup>2</sup> =	8.45. df=	3 (P = 0	$(04):  ^2 = 1$	64%			
Test for overall effect:	Z= 3.19 (	P = 0.00	1)				
							120
Total (95% CI)		2540		2057	100.0%	1.37 [1.14, 1.66]	•
Total events	290		160				
Heterogeneity: Chi <sup>2</sup> =	18.58, df=	= 13 (P =	= 0.14); l <sup>2</sup>	= 30%			
Test for overall effect:	Z = 3.31 (I	P = 0.00	09)				Eavours [experimental] Eavours [control]
Test for subaroup diff	ferences: (	Chi² = 5.	38. df = 3	(P = 0)	15),   <sup>2</sup> = 4	14.2%	rate and Texperimental Trate and Texanol
Fig	ure 6. Fore	est plots	for subg	roup fo	r the risk	of high-grade fatigue	according to the type of MET inhibitors.

The *P* value of Egger test for the association between MET inhibitors and high-grade fatigue was 0.24, thus suggesting a minor influence of bias in the identification of high-grade fatigue. Funnel plots are presented in Figure S5 (http://links.lww.com/MD/D3).

## 4. Discussion

MET transmembrane glycoprotein has been demonstrated to be overexpressed in a wide variety of tumor, and it correlates with more advanced disease, poor survival, and the presence of metastases.<sup>[41]</sup> Meanwhile, the signaling pathway is also associated with chemotherapy, radiotherapy, and targeted therapy including EGFR and VEGF inhibitors.<sup>[42,43]</sup> Therefore, inhibition of the MET signaling pathway could be an attractive therapeutic target for cancer therapy. MET inhibitor can enhance the cytotoxic effects of chemotherapy or chemoradiotherapy in MET expressing cell carcinoma.<sup>[43]</sup> Several drugs that target MET signaling pathway, including antibodies and small molecule inhibitors, have reached clinical evaluation and showed promise in animal models.<sup>[44]</sup> Results from the clinical studies indicated that those patients who exhibit *c-Met* gene amplification and mutations respond more successfully to targeted therapy and achieve maximum therapeutic outcome.<sup>[45]</sup> Furthermore, the most promising clinical data comes from combination therapies in which MET agent is believed to participate in mechanisms of drug resistance.<sup>[46]</sup>

Fatigue represents a fundamental and essential issue in oncology practice since it correlates with poor life quality in patients with underlying malignancies. Cancer treatment-related fatigue is a multifactorial process, and the precise underlying pathophysiology remains unclear.<sup>[18]</sup> Fatigue can arise as a result of cancer induced or as an adverse event of drugs. Fatigue occurs through putative mechanisms that include cytokine dysregula-

tion, dysregulation of the hypothalamic-pituitary-adrenal axis, alterations in the autonomic nervous system, anemia, neurotransmitter dysregulation, as well as patient-related factors such as psychosocial states and demographic and medical factors.<sup>[47]</sup> Furthermore, fatigue might be a consequence of TKI-induced anemia or endocrine disorders, such as adrenal dysfunctions, thyroid alterations, mineral, gonadal, and other metabolic alterations.<sup>[48]</sup> Additionally, TKI inhibitors can also lead to hypothyroidism, cardiotoxicity, skeletal muscle atrophy, and pneumonitis, which indirectly cause fatigue. In our metaanalysis, fatigue was reported in a substantial proportion of patients in the control arms, which provides a benchmark for future research attempting to mitigate fatigue. According to the National Comprehensive Cancer Network guidelines, initial evaluation of cancer patients suffering from fatigue should include focused disease status as well as fatigue assessment. Moreover, endocrinologic evaluation, hematological function, and hepatic and renal function are always warranted.<sup>[49]</sup> The guidelines published by the American Society of Clinical Oncology recommended that fatigue assessment should be conducted for all cancer survivors from the point of diagnosis onwards.<sup>[50]</sup>

Fatigue in relation to the use of MET inhibitors was reported in the overwhelming majority of completed clinical trials. These trials have highlighted how fatigue and people reported adverse events emerged as fatal when weighing the efficacy and safety of these inhibitors. However, both subjective outcome scales of fatigue assessment and long-term effect of cancer treatment may limit the accurate quantification of the events for fatigue in cancer patients. Additionally, concurrent therapies might exacerbate fatigue. For this reason, it is crucial for both clinicians and patients to properly understand the risk of MET inhibitorsrelated fatigue to promptly allow early management and take the appropriate measures to face these events.

Previously, the prevalence of fatigue in cancer survivors has been examined in reviews, with the use of immune checkpoint inhibitors, mammalian target of rapamycin (mTOR), and VEGF receptor.<sup>[51,52]</sup> The results demonstrated that the use of mTOR and VEGF receptor inhibitors is associated with a significantly increased risk of all-grade fatigue, but not for high-grade fatigue. However, to date, only a narrative review examined the safety results of MET inhibitors, which was not estimated with a metaanalysis.<sup>[5]</sup> In this article, we tried to assess the risk of fatigue, as evaluated by treating investigators using the common toxicity criteria of adverse events version 2.0 or 3.0, treated by several developed MET TKI and mAbs. This meta-analysis confirms that cancer patients treated with a MET inhibitor-containing regimen are at higher risk of high-grade fatigue relative to control treatment. It is noteworthy that incidence of all-grade and highgrade fatigue varied among patients with different types of tumors. Similarly, we found higher incidence values in the cabozantinib and tivantinib subgroup compared to other groups, by 22% (P < .001). This data analysis also showed that MET inhibitors were linked to a higher risk of high-grade fatigue compared with the control groups, whereas all-grade fatigue showed no significant differences. Nevertheless, the above results should be interpreted with cautious while some potential heterogeneity associated with this analysis. To overcome the potential bias, sensitive analysis and subgroup analyses have been performed.

In line with a meta-analysis of clinical studies, our sensitivity analysis excluded patients receiving crizotinib, which was approved as an anaplastic lymphoma kinase inhibitor.<sup>[53]</sup> In the subgroup analysis, numerical higher RRs of both all-grade and high-grade fatigue were found in trials of prostate cancer. This may be because these patients are more prone to develop drug-related renal toxic effects and endocrine-induced fatigue. Stratified by type of MET inhibitors, the different risk of fatigue among TKI and mAbs in the current analysis highlights that cabozantinib and tivantinib have a deep impact on their toxicity of fatigue. The different risk of hepatotoxicity among MET inhibitors in the current analysis highlights the importance of not dealing with MET inhibitor as a single entity, but rather classifying them according to their mechanism of action and the combinations used, which would have a profound impact on their efficacy and toxicity. Furthermore, subanalyses did not identify any differential impacts in all-grade fatigue based on the line of therapy, age, and median duration of treatment (data not shown). Unlike a small increase in a critical toxicity such as death, these findings are not likely to affect the therapy given by an individual physician. However, high-grade fatigue may always contribute to dose reductions, treatment interruptions, or treatment discontinuation. As no prior review has applied meta-analytic methods to assess the MET-inhibitors-associated fatigue, these results of potential interactions may have an indirect impact on the level of management of these toxicities and consequently the optimal continuation of cancer treatment. Psychosocial and pharmacologic interventions could be offered, and the reported outcomes may help to inform treatment choice in certain tumors.

#### 4.1. Our meta-analysis weakness

This study has several potential limitations. First, there was apparent heterogeneity between studies. This was likely to be derived from different underlying malignancies, concomitant therapies, control regimens, phase of trials, and schedule doses. Moreover, comparing fatigue among trials is inherently difficult as the diverse criteria were used to define measurement instrument, clinical time points, and heterogeneity in tumor type and disease stage.<sup>[54]</sup> The influence was minimized by using prespecified subgroup analyses to explore the possible reasons for the heterogeneity. Second, safety profile is usually not the primary outcome measure in clinical trials; all-grade and high-grade fatigue reported to be suboptimal and variable. Third, this is a meta-analysis at study level, and confounding variables at the patient level could not be incorporated into the analysis. Fourth, it is uncertain if these findings can be generalized to drugs-related side effects because fatigue is the well-known cancer-related symptom. It is also noteworthy that the stringent eligibility criteria of RCTs may exclude patients with comorbidities. Practicing oncologists need to be aware of this risk and provide continuous monitoring to patients.

#### 5. Conclusion

Our meta-analysis indicates that the RR of high-grade fatigue is higher in patients receiving MET inhibitors than the control groups. Early detection and effective management of high-grade fatigue that can occur in cancer patients with MET inhibitors are crucial for the safer use of these drugs. The findings of this metaanalysis should be considered when assessing the benefits and risks of MET inhibitors in clinical practice. Data curation: Hongxuan Tong, Yutian Zhu, Yihua Liu.

**Investigation:** Yutian Zhu. **Resources:** Hongxuan Tong.

Software: Yutian Zhu.

Visualization: Yihua Liu.

Writing – original draft: Hongxuan Tong, Yihua Liu.

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