

# Meta-analysis of studies using metformin as a reducer for liver cancer risk in diabetic patients

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## Abstract

Metformin has garnered more interest as a chemo-preventive agent given the increased liver cancer risk in diabetic patients. This work was undertaken to better understand the effect of metformin use on liver cancer risk in diabetic patients.

A comprehensive literature search was performed in PubMed, Embase, BIOSIS Previews, Web of Science, and Cochrane Library through July 30, 2016. Meta-analyses were performed using Stata version 12.0, with odds ratio (ORs) and 95% confidence intervals (CIs) as effect measures.

Twenty-three studies were included. Meta-analysis of 19 studies involving 550,882 diabetic subjects suggested that metformin use reduced the ratio of liver cancer by 48% (OR=0.52; 95% CI, 0.40–0.68) compared with nonusers. The protective effect was validated in all the exploratory subgroup analyses, except that pooled result of post hoc analyses of 2 randomized controlled trials found no significant difference between subjects with metformin and those without, with OR being 0.84 (95% CI, 0.10–6.83). After adjusting for hepatitis B/C virus infection, cirrhosis, obesity, behavioral factors, and time-related bias, the association was stable, pooled OR ranged from 0.42 to 0.75.

A protective effect for liver cancer was found in diabetic metformin users. However, more randomized clinical evidence is still needed to verify the results.

**Abbreviations:** ADM = antidiabetic medication, AMPK = adenosine monophosphate-activated protein kinase, CI = confidence interval, DM = diabetes mellitus, HBV = hepatitis B virus, HCC = hepatocellular carcinoma, HCV = hepatitis C virus, HR = hazard ratio, OR = odds ratio, RCT = randomized controlled trial, RR = relative risk.

**Keywords:** diabetes mellitus, liver cancer, meta-analysis, metformin

## 1. Introduction

Liver cancer is the fifth most common cancer worldwide and the third-leading cause of cancer related-death.<sup>[1]</sup> Diabetes mellitus (DM) is becoming an established independent risk factor for liver cancer as reported in multiple observational studies and subsequent meta-analyses.<sup>[2,3]</sup> In these existing studies, DM has been reported to confer a 2- to 4-fold risk of liver cancer, and the risk increases with DM severity and duration. However, this risk may be mitigated by antidiabetic medications (ADMs). Metformin, a widely used ADM, has recently attracted great attention for antitumor effect in a wide range of malignancies including liver cancer, through both insulin-dependent and insulin-independent mechanisms.<sup>[4]</sup> However, the evidence for a cancer

preventive effect for metformin has not been consistently demonstrated.

Association of metformin and risk of liver cancer are mainly studied in animal and observational human studies. A meta-analysis demonstrated that metformin appeared to have a direct antihepatocellular carcinoma (HCC) effect in animal models.<sup>[5]</sup> Accumulating epidemiologic studies comparing the risk of liver cancer between those using metformin with those using other ADMs have shown somewhat variable results,<sup>[6–9]</sup> and it was also endorsed that confounders were not well addressed in most studies.<sup>[2]</sup> Several meta-analyses have been published to determine if a consistent effect of metformin use on liver cancer incidence was evident.<sup>[10–13]</sup> Except for the incomplete included studies, meta-analyses in previous reviews were rough, and the heterogeneity was not explored in detail.

In our opinion, the differences in estimates and the heterogeneity between studies could largely be explained by differences in study designs, quality, population, the comparators used, estimation of the exposure to metformin (duration and dosage) and adjusted factors, as the inability to account for these factors may result in certain degrees of bias. To better understand the association of metformin and risk of liver cancer, we embarked on a systematic review and meta-analysis with integrated overall, subgroup and sensitivity analyses.

## 2. Methods

### 2.1. Inclusion criteria

Either observational studies (cohort and case-control studies) or post hoc analyses of randomized controlled trials (RCTs) were included if they evaluated and defined exposure to metformin or

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biguanide, reported liver cancer incidence or related outcomes of diabetic patients, provided effective comparison groups, and reported hazard ratio (HR)/relative risk (RR)/odds ratio (OR) and corresponding 95% confidence intervals (CIs), or provided sufficient data for their estimations. Inclusion was not restricted by language, study size, or publication type. The most recent or most comprehensive report was given precedence if there were multiple publications (regardless of study design) from the same population, while the others might be included in subgroup analysis according to the concrete conditions.

## 2.2. Literature search

A comprehensive literature search was performed in PubMed, Embase, BIOSIS Previews, Web of Science, Cochrane Library, National Institutes of Health database, EU clinical trials register from the earliest date available through July 30, 2016, without any restrictions. In order to include more potential literature, our overall search strategy only included terms for metformin and liver cancer. The comprehensive literature search was conducted as follow: ((liver cancer) OR (liver carcinoma) OR (liver neoplasm) OR (liver tumor) OR (hepatoma) OR (hepatocellular carcinoma) OR (HCC) OR (hepatic cancer) OR (hepatic neoplasm) OR (hepatic tumor) OR (cholangiocarcinoma)) AND ((metformin) OR biguanide). We screened bibliographies of selected original studies, review articles, and relevant conference abstracts. Attempts were made to contact the corresponding authors for additional data.

## 2.3. Data extraction

Citations were merged together in Endnote, version X7 to facilitate management. Two authors independently applied the inclusion criteria to all retrieved articles in an unblinded standardized manner, evaluated by title, abstract, and full text. For each of eligible study, information of first author, publication year, location, study design, data source, study period, mean follow-up, characteristics of study population (mean age, sex ratio), definition of exposure and control, dose and duration of exposure (if reported), comparison groups, risk estimates (included HR, RR, OR), and 95% CIs with and without adjustment for confounding factors were selectively extracted onto piloted structured forms independently by 2 authors. As subjects in most studies used combination therapy, the final analysis on exposure used the dichotomous categorical variable of “with” or “without” use of metformin.

Adjusted factors were extracted, and some of them were selected for further analysis: infected with hepatitis B virus (HBV) or hepatitis C virus (HCV), cirrhosis, obesity (including body mass index and obesity), behavioral factors (including alcohol abuse and cigarette smoking), use of statins, and time-related bias (including DM duration, duration of exposure, duration of follow up, enrollment date, time of first ADM prescription, calendar time, and other time-dependent factors). Any disagreements during study selection or data collection were resolved by discussion, referring back to the original article.

## 2.4. Quality assessment

Newcastle–Ottawa Scale<sup>[14]</sup> was used to assess the quality of observational studies and post hoc analyses of RCTs, which were treated as cohort studies to achieve quality assessment. In this scale, studies were scored across 3 categories: selection of subjects

(4 stars), comparability of study groups (2 stars), and assessment of outcome/exposure (3 stars). Star rating system was used to indicate the quality, with a maximum of 9 stars: 0 to 5 stars as low quality and 6 to 9 stars as high quality.

## 2.5. Statistical analysis

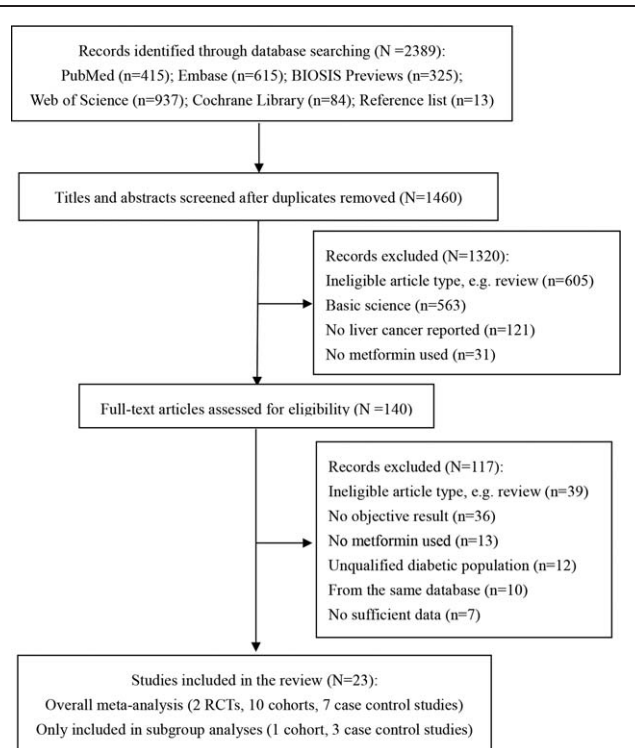
Adjusted estimate was mainly used for quantitative analysis. Crude estimate served as an alternative in case of no adjusted estimate was available. When estimates or 95% CIs were missing or incomplete, appropriate summary statistics or Kaplan–Meier curves were used to calculate based on published methods.<sup>[15]</sup> OR was employed as a common measure of the association between metformin use and liver cancer risk due to the enrollment of case–control studies in most analyses. Between-study heterogeneity was qualitatively assessed by using Cochrane Q test with a significance level of  $P \leq .1$ , and quantified by estimated  $I^2$  ( $I^2 < 50\%$  representing low heterogeneity,  $50\% \leq I^2 \leq 75\%$  representing moderate heterogeneity,  $I^2 > 75\%$  representing substantial heterogeneity).<sup>[16]</sup> An inverse variance fixed-effects model was used to calculate when the test for heterogeneity was not statistically significant, otherwise the DerSimonian–Laird random-effects model was employed.<sup>[17]</sup>

Sensitivity analyses were performed to assess the robustness of results. Between-study sources of heterogeneity were further investigated using subgroup analyses by stratifying original estimates according to study characteristics (study design, setting, and quality), controlled ADM, and adjustment. Analyses of adjusted estimates were emphasized on studies controlling for HBV/HCV infection, cirrhosis, obesity, behavioral factors, use of statins, and time-related bias, given their modifying effects on metformin’s activity on DM and liver cancer risk.<sup>[18–20]</sup> Publication bias was detected for overall analysis using Begg test and Egger test (publication bias considered present if  $P \leq .1$ ).<sup>[21,22]</sup> All the statistical analyses were 2-sided and performed using Stata version 12.0 (StataCorp, College Station, TX).

## 3. Results

### 3.1. Description of included studies

Searches identified 2389 potentially relevant studies. The selection process is shown in Fig. 1. Twenty-three studies fulfilled the inclusion criteria and were included in the meta-analysis (2 RCTs, 11 cohort studies, 10 case–control studies). These 23 studies cumulatively reported more than 35,000 cases of liver cancer in 663,335 diabetic subjects. Only 19 studies were included in the overall analysis,<sup>[6–9,23–36]</sup> the remaining 4 studies<sup>[37–40]</sup> were only included in subgroup analyses for specific conditions, as they were multiple publications from the same populations. In fact, more multiple publications were found during the study selection. Eleven studies<sup>[7,27,37–39,41–46]</sup> (6 cohort and 5 case–control) were conducted in Taiwan, China using the National Health Insurance data, and hence only 2<sup>[7,27]</sup> of them with different time period (ignoring a coinciding year) were included in our overall analysis, and 3<sup>[37–39]</sup> were just included in different subgroup analyses. Likewise, 3 Italian case–control studies<sup>[32,47,48]</sup> were from a same cohort, and only one<sup>[32]</sup> of them was included. Moreover, data of 1 cohort<sup>[31]</sup> and 1 case–control study<sup>[40]</sup> were both from the United Kingdom Clinical Practice Research Datalink, only the cohort with larger sample size and higher quality was included in our overall analysis.



**Figure 1.** PRISMA flow diagram of study selection. RCT=randomized controlled trials.

The characteristics of included studies are shown in Table 1. Twelve studies<sup>[8,23,25,26,29,30,32–36]</sup> were population-based studies, and the remainder<sup>[6,7,9,24,27,28,31]</sup> were hospital-based studies. Overall methodological quality of included studies was high. Treatment comparators were sulfonylureas,<sup>[28,31,32,37]</sup> insulin<sup>[25,26,29,32,37,40]</sup> or nonuse of any ADMs.<sup>[25,26,40]</sup> Type of liver cancer was clearly informed to be HCC in most studies. In addition to age and sex, most studies adjusted for HBV/HCV infection,<sup>[7,24,25,27,30,32,33,40]</sup> obesity,<sup>[7,25,31,32]</sup> behavioral factors,<sup>[8,24,25,31–33]</sup> use of statins,<sup>[6,31,38]</sup> and time-related bias.<sup>[6,7,9,25,28,31,32]</sup>

### 3.2. Overall analysis

On the basis of 19 studies<sup>[6–9,23–36]</sup> involving 550,882 diabetic patients, compared with metformin nonusers, metformin use reduced the ratio of liver cancer by 48% (OR=0.52; 95% CI, 0.40–0.68;  $P < .001$ ), with substantial heterogeneity ( $I^2 = 83.7%$ ) (Fig. 2). Sensitivity analysis using leave-one-out method found that the pooled result was robust when omitting any one study alone, heterogeneity kept substantial except when omitting the study<sup>[9]</sup> with maximum weight from overall analysis,  $I^2$  dropped to 33.5%, with the summary OR being 0.53 (95% CI, 0.44–0.63;  $P < .001$ ). Significant publication bias was found for overall analysis by Begg test ( $P = .069$ ) and Egger test ( $P < .001$ ).

### 3.3. Subgroup analysis

Subgroup analyses were conducted to further validate the result from overall analysis, and to explore potential sources of heterogeneity among studies (Table 2). Hierarchies of study setting, quality, controlled drugs, and adjustment did not change

over the significant reduction in ratio of liver cancer in metformin users. Pooled result of post hoc analyses of 2 RCTs<sup>[23]</sup> found no significant difference between subjects with metformin and those without, with OR being 0.84 (95% CI, 0.10–6.83;  $P = .871$ ) (Fig. 2). Subgroup analyses of hospital-based studies with relatively small sample size (OR=0.32; 95% CI, 0.24–0.44) or studies with low quality (OR=0.29; 95% CI, 0.18–0.49) showed an exaggeration in metformin's effect. Metformin showed higher protective effect of liver cancer when compared with insulin (OR=0.36; 95% CI, 0.25–0.51), other than sulfonylurea (OR=0.65; 95% CI, 0.55–0.78) and nonuser of any ADM (OR=0.62; 95% CI, 0.40–0.98).

Furthermore, use of adjusted estimates caused numerical increases on pooled OR and heterogeneity. Compared to the analysis of all adjusted estimates (OR=0.57; 95% CI, 0.42–0.76;  $P < .001$ ), numerical increases in the ratio of liver cancer in metformin users were found when the estimates were adjusted for HBV/HCV infection, cirrhosis, obesity, and behavioral factors (pooled ORs ranged from 0.42 to 0.51), while the ratio reduction decreased in studies adjusted for use of statins (OR=0.75; 95% CI, 0.68–0.83;  $P < .001$ ) and time-related bias (OR=0.65; 95% CI, 0.48–0.89;  $P = .006$ ).

Heterogeneity was significant in most subgroups, with  $I^2 (>50%)$  ranging from 50.5% to 89.2% (Table 2). Nevertheless, no heterogeneity was found in the subgroup analyses of hospital-based studies ( $I^2 = 0%$ ) and studies with low quality ( $I^2 = 0%$ ). Moreover, for subgroup analyses of RCTs, studies adjusted for use of statins, and when the controlled drugs were definitely restricted to insulin, sulfonylurea, or nonuse of any ADM, heterogeneity was limited ( $I^2$  ranged from 14.0% to 39.5%).

## 4. Discussion

This systematic review synthesized evidence on association between use of metformin and risk of liver cancer in diabetic patients from 23 studies. We used systematic strategy and broad search terms in multiple databases to identify as many studies as possible. Rigorous methods were used to extract and appraise the data. Multiple publications from the same population were checked in any analysis. Considering the potential confounding factors for liver cancer, adjusted estimates were used instead of the unadjusted ones as much as possible to make the summary results more precise and plausible.

Overall meta-analysis of 19 studies involving 550,882 diabetic subjects found that, relative to nonuse, use of metformin reduced the ratio of liver cancer by 48% (OR=0.52; 95% CI, 0.40–0.68;  $P < .001$ ), with substantial heterogeneity ( $I^2 = 83.7%$ ). Sensitivity analysis found that the heterogeneity was mainly from a high-quality population-based study,<sup>[9]</sup> with maximum weight, but without reporting the exact number of liver cancer cases. After omitting this study, heterogeneity of overall analysis was significantly decreased ( $I^2$  dropped from 83.7% to 33.5%), so was the heterogeneity in subgroup analyses of high-quality studies ( $I^2$  dropped from 86.1% to 33.7%) and population-based studies ( $I^2$  dropped from 89.2% to 7.3%). The beneficial effect of metformin was validated in observational studies, with a diminution in cohort studies (OR=0.64) and a rise in case-control studies (OR=0.50), but it lost significance in RCTs (OR=0.84; 95% CI, 0.10–6.83;  $P = .871$ ), which might be largely limited by the fewer available studies ( $n = 2$ ). Actually, most RCTs are not designed or sufficiently powered to examine cancer outcomes due to the short follow-up periods and very few cancer events.<sup>[31]</sup>

**Table 1**

**Characteristics of included studies in the meta-analysis.**

Study (year)	Design	NOS score	Location	Study population	Total subject	Mean age, y	Sex, % male	Liver cancer cases	Data source	Time period	Mean follow-up, y	Definition of exposure and control	Adjusting factors
Home 2010 (ADOPT) <sup>[23]</sup>	RCT	6	USA, Canada, Europe	Diabetic subjects	4351	56.8	57.7	4	Post hoc analysis	April 2000–June 2006	4.0	Ex: metformin ≤ 2 g/d; Con: rosiglitazone ≤ 8 mg/d; glyburide/glibenclamide ≤ 15 mg/d	Study exposure
Home 2010 (RECORD) <sup>[23]</sup>	RCT	6	Europe, Australia	Diabetic subjects	4447	58.4	51.6	4	Post hoc analysis	April 2001–December 2008	5.5	Ex: metformin ≤ 2.55 g/d and sulfonylurea/rosiglitazone; Con: sulfonylurea and rosiglitazone	Study exposure
Oliveria 2008 <sup>[24]</sup>	Cohort	9	USA	Diabetic subjects	191,223	56.0	51.0	11	Insured population	January 2000–December 2004	3.9	Ex: with metformin (filled ≥ 1 prescription); Con: without metformin (did not fill any prescriptions)	Age, sex, hepatitis B/hepatitis C, cirrhosis, alcoholism
Nkontchou 2011 <sup>[25]</sup>	Cohort	7	France	Diabetics with HCV cirrhosis	100	61.0	53.0	39	Screening program	January 1988–January 2007	5.0	Ex: metformin monotherapy or metformin plus insulin; Con: diet alone; insulin secretagogues; insulin therapy	Age, sex, BMI, DM duration, alcohol abuse, steatosis, AFP, HbA1c, other biochemical index
Ampuero 2012 <sup>[26]</sup>	Cohort	2	Spain	Diabetic subjects	82	57.0	71.0	9	NR	NR	NR	Ex: metformin with or without pioglitazone; Con: insulin; dietetic treatment	NA
Hsieh* 2012 <sup>[27]</sup>	Cohort	7	China	Diabetic subjects	10,786	61.2	52.2	220	Taiwan's NHI	2000–2008	NR	Ex: metformin monotherapy; Con: sulfonylurea monotherapy; insulin monotherapy	Age, sex
Lai 2012 <sup>[27]</sup>	Cohort	8	China	Diabetic subjects	19,349	55.5	55.8	224	Taiwan's NHI	2000–2005	4.9	Ex: with metformin; Con: without metformin	Age, sex, cirrhosis, alcoholic liver damage, hepatitis B, hepatitis C
Ruiter 2012 <sup>[28]</sup>	Cohort	7	The Netherlands	Diabetic subjects†	85,289	63.3	47.1	31	PHARMO RLS	January 1998–December 2008	3.5	Ex: metformin; Con: sulfonylurea derivatives	Sex, age at first OGLD prescription, year in first OGLD prescription was dispensed, numbers of hospitalizations and drugs
Aydin 2013 <sup>[28]</sup>	Cohort	5	Turkey	Diabetic subjects	655	57.2	42.0	2	Duzce University	January 2010–January 2011	6.8	Ex: with metformin; metformin monotherapy; Con: without metformin; insulin monotherapy	NA
Hsu 2014 <sup>[30]</sup>	Cohort	4	China	Diabetics received NUC for CHB with liver cirrhosis	49	52.8	73.3	13	E-Da hospital	September 2007–March 2013	2.1	Ex: with metformin; Con: without metformin	NA
Tsilidis 2014 <sup>[31]</sup>	Cohort	9	UK	Diabetic subjects	95,820	62.1	56.5	102	CFRD	January 1987–December 2010	5.1	Ex: with metformin; metformin monotherapy; Con: without metformin; sulfonylureas monotherapy	Age, sex, smoking status, alcohol consumption, aspirin/NSAIDs, statins, DM duration, BMI, year of first ADM prescription
Lin 2015 <sup>[7]</sup>	Cohort	9	China	Diabetic subjects	34,823	54.3	50.7	285	Taiwan's NHI	2005–2010	2.9	Ex: with metformin; Con: without metformin	Age, sex, obesity, cirrhosis, hypertension, hepatitis B and C, dyslipidemia, gout, duration of ADM exposure
Valent 2015 <sup>[9]</sup>	Cohort	7	Italy	Diabetic subjects	109,255	NR	NR	NR	Regional health information system	January 2002–December 2014	NR	Ex: with metformin; Con: without metformin	Age at start of observation, sex, time-dependent variable, other drugs
Donadon 2010 <sup>[32]</sup>	Case-control	5	Italy	Diabetic HCC, liver cirrhosis, and controls (general population)	595	68.3	80.0	190	Pordenone General Hospital	January 1994–December 2008	NA	Ex: metformin; Con: sulfonylureas; insulin	Age, sex, BMI, HBV and HCV infection, alcohol abuse, ALT level, triglycerides, cholesterol and DM duration

(continued)

**Table 1**  
**(continued).**

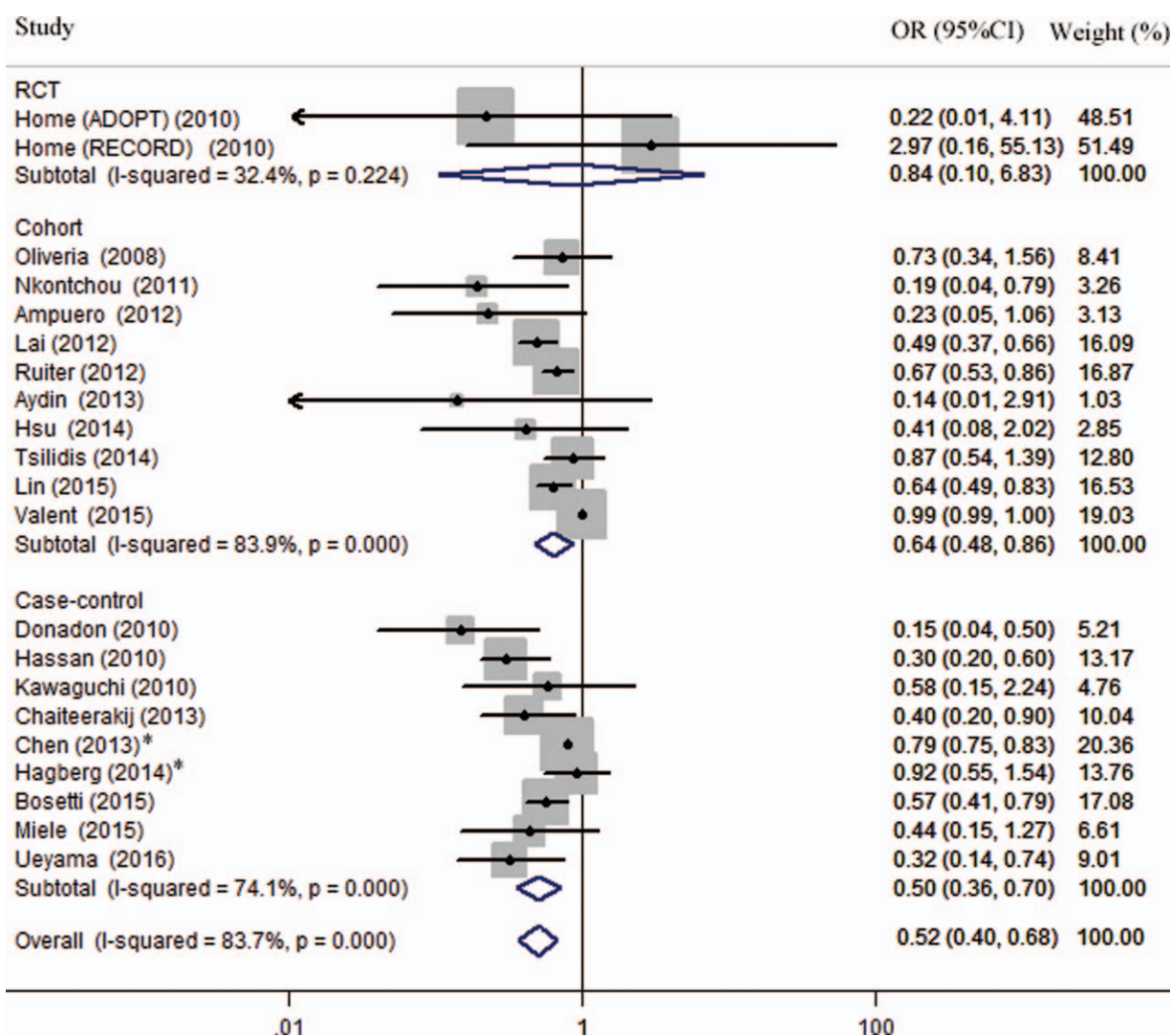
Study (year)	Design	NOS score	Location	Study population	Total subject	Mean age, y	Sex, % male	Liver cancer cases	Data source	Time period	Mean follow-up, y	Definition of exposure and control	Adjusting factors
Hassan 2010 <sup>(33)</sup>	Case-control	7	USA	Diabetic HCC and controls (healthy spouses or in-laws of specific cancer patients)	208	60.8	74.5	122	M. D. Anderson Cancer Center	January 2000–July 2008	NA	Ex: with biguanide; Con: without biguanide	Age, sex, race, educational level, smoking, alcohol drinking, HCV, HBV, family history of cancer
Kawaguchi 2010 <sup>(34)</sup>	Case-control	6	Japan	Diabetic HCC and non-HCC (both were hepatitis C patients)	241	67.0	67.6	138	Kurume Nagata and Chikugo Hospital	January 2004–December 2008	NR	Ex: with biguanide; Con: without biguanide	Age, sex
Chang* 2012 <sup>(38)</sup>	Case-control	7	China	Diabetic liver cancer and controls (a risk-set sampling)	52,588	66.7	63.9	10,741	Taiwan's NHI	January 2000–December 2007	7.9	Ex: with metformin; Con: without metformin	Age, sex, chronic liver disease, other ADMs, statins, aspirin, other drugs, and diseases
Chaitteerakij 2013 <sup>(35)</sup>	Case-control	7	USA	Diabetic ICC and controls (bio-bank participants)	113	61.4	49.7	79	Mayo Clinic	January 2000–May 2010	NA	Ex: with metformin; Con: without metformin	Match: age (±5 y), sex, ethnicity, residence
Chen* 2013 <sup>(39)</sup>	Case-control	7	China	Diabetic HCC and controls (randomly sampled subjects)	47,820	62.5	72.8	22,047	Taiwan's NHI	January 1997–December 2008	NA	Ex: with metformin; Con: without metformin	Age, sex, hepatitis B and C, liver cirrhosis, end-stage renal disease, DM duration, DM control and other ADMs
Hagberg* 2014 <sup>(40)</sup>	Case-control	8	U.K.	Diabetic liver cancer and controls (same study population)	1069	69.6	84.1	234	CRPD	1988–2011	NR	Ex: with metformin; metformin monotherapy; Con: without metformin; nonuse of ADMs; insulin monotherapy	Age, sex, calendar time, general practice, BMI, smoking, alcohol-related disorders, HBV/HCV, statins, rare metabolic disorder, DM duration
Bosetti 2015 <sup>(6)</sup>	Case-control	7	Italy	Diabetic HCC <sup>†</sup> and controls (alive and at risk of developing HCC)	3962	65.3	75.2	190	Healthcare utilization databases	January 2005–December 2007	6.0	Ex: with metformin; Con: without metformin	Age, sex, date at cohort entry, statins, duration of follow-up, Charlson comorbidity index, other ADMs and drugs
Miele 2015 <sup>(6)</sup>	Case-control	5	Italy	Diabetic HCC and controls (out-patients, patients undergoing surgical interventions)	121	NR	NR	69	Agostino Gemelli Hospital	January 2005–July 2012	NA	Ex: with metformin; Con: without metformin	Age, sex, tobacco smoking, alcohol drinking
Ueyama 2016 <sup>(56)</sup>	Case-control	3	Japan	Diabetic HCC and controls (included patients with liver cirrhosis)	389	70.9	51.7	59	Five hospitals	NR	NA	Ex: with metformin; Con: without metformin	Age, sex, tobacco smoking, alcohol drinking

Exposure: with metformin, patients who ever had metformin, including metformin-combined therapy; without metformin, patients who had no metformin, but might have no medication, or have used diet alone or other ADMs.

ADM = antidiabetic medication, ADOPT = A Diabetes Outcome Progression Trial, AFP =  $\alpha$ -fetoprotein, ALT = alanine aminotransferase, BMI = body mass index, CHB = chronic hepatitis B, Con = control, CPRD = Clinical Practice Research Datalink, DM = diabetes mellitus, Ex = exposure, HbA1c = glycated hemoglobin, HBV = hepatitis B virus, HCC = hepatocellular cancer, HCV = hepatitis C virus, ICC = intrahepatic cholangiocarcinoma, NA = not applied, NHI = National Health Insurance, NOS = Newcastle-Ottawa Scale, NR = not reported, NSAIDs = nonsteroidal anti-inflammatory drugs, NUC = nucleoside analogue, OGLD = oral glucose-lowering drug, RECORD = Rosiglitazone Evaluated for Cardiovascular Outcomes and Regulation of Glycaemia in Diabetes, RLS = record linkage system.

\* Studies were only included in subgroup analyses.

<sup>†</sup> Subjects with at least 1 antidiabetic drug prescription.



**Figure 2.** Forest plot of the association between use of metformin and liver cancer risk in diabetic patients. RCT=randomized controlled trials. \*Studies were multiple publications and they were only included in corresponding subgroup analysis.

Metformin is one of the most commonly prescribed medications in the treatment of DM. However, DM treatment is a dynamic process, ADMs might be changed continuously and often used in combination,<sup>[29]</sup> which made definition of exposure using dichotomous categorical variable of “with” or “without” use of metformin be somewhat less convincing. Thus we further subanalyzed supplemented comparisons between monotherapy of ADMs. Results showed that metformin had higher protective effect of liver cancer when compared with insulin, other than sulfonylurea and nonuser of any ADM. Beyond the plausible finding that use of insulin increased risk of liver cancer,<sup>[44]</sup> another explanation is that metformin is a first-line ADM prescribed in less severe or shorter duration of DM, while insulin is usually prescribed to patients with longer duration and more advanced DM, which in turn may be associated with higher risk of liver cancer.<sup>[49]</sup> However, when compared to nonuser of any ADM (mild or newly diagnosed DM patients), monotherapy use of metformin achieved a 38% (OR=0.62; 95% CI, 0.40–0.98) reduction in ratio of liver cancer, probably reflecting the real world scenario.

Lots of confounders may have modifying effect on association between metformin and liver cancer risk in diabetic patients.

Presence of DM in patients with cirrhosis is an independent factor for the progression to liver cancer.<sup>[50]</sup> Moreover, metformin may be specifically sensitive to certain etiological types of liver cancer.<sup>[10]</sup> After adjusting for HBV/HCV infection, cirrhosis, obesity, and behavioral factors, the beneficial effects on the ratio of liver cancer for metformin use were significant and larger (pooled OR ranged from 0.42 to 0.51), which might be the true link between metformin use and liver cancer risk in diabetic patients. Recent reviews underscored the prevalence of time-related bias in observational studies, potentially leading to inflated estimates of metformin’s protective effect.<sup>[19]</sup> Time-related bias includes immortal-time bias, time-window bias, and time-lag bias.<sup>[19]</sup> Of note, exclusion of time-biased studies from our analysis resulted in a numerical decline on the ratio reduction (OR=0.65; 95% CI, 0.48–0.89; P=.006). Thus further studies should take these biases into account in the study design and analysis.

Statins was previously found to be associated with a reduced risk of liver cancer.<sup>[51]</sup> Most of included studies did not take the concomitant use of statins into account to adjust for potential confounding. Subgroup analysis of studies adjusted for the use of

**Table 2****Summary results of subgroup analyses.**

Subgroup	No. of studies	Total subject	Summary result		I <sup>2</sup> , %
			OR (95% CI)	P	
Design					
RCT	2	8798	0.84 (0.10–6.83)	.871	32.4
Cohort	10	536,645	0.64 (0.48–0.86)	.003	83.9
Case–control	9	54,328	0.50 (0.36–0.70)	<.001	74.1
Setting					
Hospital-based	12	11,161	0.32 (0.24–0.44)	<.001	0
Population-based	7	539,721	0.69 (0.52–0.91)	.009	89.2
Quality					
Low	6	1701	0.29 (0.18–0.49)	<.001	0
High	13	549,181	0.60 (0.45–0.79)	<.001	86.1
Controlled drugs					
Insulin	6	11,100	0.36 (0.25–0.51)	<.001	33.6
Sulfonylurea	4	160,115	0.65 (0.55–0.78)	<.001	14.0
Nonuse of any ADM	3	785	0.62 (0.40–0.98)	.039	37.2
Adjusted REs used or not					
Unadjusted REs	17	406,804	0.42 (0.36–0.49)	<.001	0
Adjusted REs	11	540,555	0.57 (0.42–0.76)	<.001	88.8
Adjustment					
Infected with HBV/HCV	8	247,226	0.50 (0.36–0.69)	<.001	59.0
Cirrhosis	6	245,878	0.49 (0.34–0.69)	<.001	50.5
Obesity	4	131,338	0.51 (0.29–0.90)	.020	67.5
Behavioral factors	6	288,067	0.42 (0.24–0.75)	.003	66.0
Use of statins	3	152,370	0.75 (0.68–0.83)	<.001	39.5
Time-related bias	7	329,844	0.65 (0.48–0.89)	.006	86.8

ADM = antidiabetic medication, CI = confidence interval, HBV = hepatitis B virus, HCV = hepatitis C virus, OR = odds ratio, RCT = randomized controlled trial, REs = risk estimates.

statins caused a numerical decline on the ratio reduction (OR = 0.75; 95% CI, 0.68–0.83;  $P < .001$ ), which might suggest a synergistic effect of metformin and statins for liver cancer, in addition to their dose-dependent protective effects.<sup>[46,52]</sup> Given the rising disease burden of liver cancer, looking for chemopreventive strategy is necessary, especially for cheap nonetiology-specific medications, like metformin and statins, still with favorable safety profile.<sup>[52]</sup> However, further researches are needed to establish definitive role of metformin and statins on the prevention of liver cancer in diabetic patients.

The observational nature allows only an association to be established. Plenty of experimental studies have added evidence to metformin's protective effect on malignancies. Although the exact mechanism is not fully understood, several biologically plausible mechanisms have shown that metformin might have direct antiliver cancer activity by inhibiting proliferation and colony formation ability through adenosine monophosphate-activated protein kinase (AMPK) in HCC cells<sup>[53]</sup>; suppressing HCC cell growth through induction of cell cycle G1/G0 phase arrest, p21CIP and p27KIP expression, and down-regulation of cyclin D1<sup>[54]</sup>; inducing apoptosis in HCC cells via signaling pathways, including AMPK and p38 mitogen-activated protein kinase<sup>[55]</sup>; and suppressing xenograft tumor growth in mouse models.<sup>[56]</sup> Moreover, as an antihyperglycemic agent and insulin sensitizer, metformin treatment inhibits hepatic gluconeogenesis,<sup>[57]</sup> reduces serum concentrations of insulin and insulin growth factor I,<sup>[58]</sup> improves glycemic control, and decreases inflammatory response,<sup>[59]</sup> thus leading to less aggressive behavior of cancer cells. However, given that not all in vitro and in vivo work with animal models could be successfully translated into clinical outcomes in humans, well-designed RCTs are still needed to provide authentic evidence.

Several limitations of this study needed to be addressed and merited further discussion. First, significant heterogeneity was presented between studies in some of our analyses. However, sensitivity analyses found that the heterogeneity could be mostly interpreted by 1 same article.<sup>[9]</sup> Except for the contribution of heterogeneity, omitting this article would not change over the initial results. Second, information on treatment was obtained through prescriptions contained in patients' medical records, therefore a gap between prescribed and actual dose could bias the results. Third, adjustments of included studies might be incomplete and inconsistent. Although we performed subgroup analyses of adjusted estimates controlled for several important factors. Some other confounders were failed to control, such as information like details of DM (severity and duration) and metformin use (dose and duration) were absent in most studies, which would have been important to adjust for residual confounding.<sup>[9]</sup> Fourth, significant publication bias was found for overall analysis. However, this might probably be the small-study effect rather than true publication bias, especially in the presence of significant heterogeneity among studies.<sup>[60]</sup>

## 5. Conclusion

In conclusion, a protective effect in the risk of liver cancer was found in diabetic metformin users, and the protective effect was validated in most of our exploratory analyses. However, the conclusion should be interpreted with caution given the possibility of residual confounding. Simultaneously, limited by the observational study design, a conclusion of causality cannot be drawn. Clinical trials are needed to determine if the observations in diabetic subjects can be expanded to a wider range of population, and then to reveal the true scenario in real life.

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