## REVIEW

# The mental health of young people who are not in education, employment, or training: a systematic review and meta-analysis 

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#### Abstract

Purpose There are increasing concerns about the intersection between NEET (not in education, employment, or training) status and youth mental ill-health and substance use. However, findings are inconsistent and differ across types of problems. This is the first systematic review and meta-analysis (PROSPERO-CRD42018087446) on the association between NEET status and youth mental health and substance use problems. Methods We searched Medline, EMBASE, Web of Science, ERIC, PsycINFO, and ProQuest Dissertations and Theses (1999-2020). Two reviewers extracted data and appraised study quality using a modified Newcastle-Ottawa Scale. We ran robust variance estimation random-effects models for associations between NEET and aggregate groups of mental ill-health and substance use measures; conventional random-effects models for associations with individual mental/substance use problems; and subgroup analyses to explore heterogeneity. Results We identified 24 studies from 6,120 references. NEET status was associated with aggregate groups of mental ill-health (OR 1.28, CI 1.06-1.54), substance use problems (OR 1.43, CI 1.08-1.89), and combined mental ill-health and substance use measures (OR 1.38, CI 1.15-1.64). Each disaggregated measure was associated with NEET status [mood (OR 1.43 , CI 1.21-1.70), anxiety (OR 1.55, CI 1.07-2.24), behaviour problems (OR 1.49, CI 1.21-1.85), alcohol use (OR 1.28 , CI 1.24-1.46), cannabis use (OR 1.62, CI 1.07-2.46), drug use (OR 1.99, CI 1.19-3.31), suicidality (OR 2.84, CI 2.04-3.95); and psychological distress (OR 1.10, CI 1.01-1.21)]. Longitudinal data indicated that aggregate measures of mental health problems and of mental health and substance use problems (combined) predicted being NEET later, while evidence for the inverse relationship was equivocal and sparse. Conclusion Our review provides evidence for meaningful, significant associations between youth mental health and substance use problems and being NEET. We, therefore, advocate for mental ill-health prevention and early intervention and integrating vocational supports in youth mental healthcare.


Keywords Education or employment • NEET • Youth mental health • Substance use • Systematic review • Meta-analysis

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## Introduction

Transitioning from education into work is a milestone of emerging adulthood that about one in seven young people in economically developed countries struggle to attain [1], falling into the category of NEET-not in education,

[^0]employment, or training. Concerns over these youth are growing worldwide [2,3]. The term NEET was coined in a 1999 report called "Bridging the Gap" from the United Kingdom [4]. By 2019, between 5.6\% (Luxembourg) and $28.8 \%$ (Turkey) of 15 to 29-year-olds in Organisation for Economic Co-operation and Development countries were NEET [1]. Economic fallouts of the COVID-19 pandemic are expected to swell these numbers [3], as even early on in the pandemic, data showed that NEET rates were higher in the second quarter of 2020 than the previous year in 45 out of 50 countries [5]. Youth who are NEET are considered vulnerable as they face social exclusion and disempowerment, and disproportionately come from disadvantaged backgrounds [6, 7]. Being outside school and the workforce limits their ability to gain skills and experience that could improve their prospects [8-10].

Being NEET is intertwined with mental health and substance use problems in young people. Studies have linked being NEET with the emergence of symptoms of depression, anxiety, substance use, and suicidality [11-15]. Conversely, mental health and substance use problems can deplete the drive and energy needed to enter the workforce or continue education/training and increase the risk of becoming NEET. However, the link between being NEET and poor mental health is unclear. Cross-sectional relationships are not always supported by longitudinal data $[16,17]$ and there are indications that the relationship differs by type of mental health or substance use problem [14, 18]. Furthermore, the association between being NEET and mental health problems may also differ in strength and significance across gender, depending on mental health problem [19, 20]. For instance, Henderson [19] found that the association between internalizing disorders and being NEET was significant in only men. For externalizing disorders, however, the association with NEET status was significant for both men and women.

Previous reviews have reported on the association between mental health problems and youth unemployment [21, 22] and school disengagement [23, 24], but none investigated youth disengaged from both work and school. One narrative review examined the correlates of being NEET [6], but with little in-depth information on mental health. A synthesis of the literature is needed to inform the discussion on the growing youth population who are NEET and its intersection with youth mental health and substance use problems. This information is also needed to develop intervention studies and effective strategies to promote youth engagement in employment, education, and training.

Our primary objective was therefore to systematically review and synthesize via meta-analysis the literature on the associations between being NEET and mental health and substance use problems among youth. We expected NEET status to relate to mental ill-health measures, substance use
measures, and all measures of mental ill-health and substance use combined; and the associations to vary across mental health and substance use problems. Our review thus extends the literature by focussing on youth disengagement from both education and employment and by examining the strength and consistency of associations across types of mental health and substance-use problems. Our secondary objectives were to investigate the directionality of the associations from longitudinal data and to examine subgroup differences by gender, age, and population-based versus clinical samples. We expected the association between NEET status and mental health and substance use problems to be bidirectional and to differ in strength by gender. We were agnostic as to differences by age and sample type.

## Methods

## Search strategy

We followed MOOSE reporting guidelines [25]. We searched Medline, EMBASE, ISI Web of Science, ERIC, PsycINFO, and ProQuest Dissertations and Theses Online from January 1, 1999 to May 2020, imposing no language restriction (see MEDLINE search strategy in Supplementary Appendix 1). We limited searches to the last 20 years because our objective was to synthesize contemporary knowledge of policy and practice relevance. The study is registered through PROSPERO (CRD42018087446) [26].

## Selection criteria

We included observational studies with individual-level data that estimated the association between being NEET and mental and/or substance use symptoms or disorders among persons aged 15-34 years. We chose this age range to accommodate internationally diverse definitions of youth [27]. Studies had to identify NEET status by explicitly querying work and education or training status. Measures included those for any specific or general mental or substance use disorder; psychological or behavioural problems; psychological distress or well-being; or suicidality, measured on a dichotomous or continuous scale of symptoms, severity, or score. We excluded neurodevelopmental disorders and disabilities typically diagnosed in childhood (e.g., autism, intellectual disability) since we expected developmental and learning problems to have a unique association with becoming NEET. We excluded abstracts but considered unpublished studies if information was available for data extraction and quality assessment. We searched references of primary studies and review articles for additional studies. All references were uploaded to Covidence software [28].

The screening of titles and abstracts followed by screening of full texts for inclusion and exclusion criteria; data extraction from eligible studies (see data extraction form in Supplementary Appendix 2); and quality assessment of studies using a modified Newcastle-Ottawa Scale [29] (see description in Supplementary Appendix 3) were done independently by two reviewers, including first author/epidemiologist GG and either a psychiatry graduate student or a research assistant with a Master's in mental health epidemiology (SD). Disagreements were resolved by consensus or by author SI. We emailed study authors for further information where necessary.

## Data analysis

We conducted meta-analysis to quantitatively synthesize the literature. We ran robust variance estimation (RVE) randomeffects models to obtain associations between NEET status and three aggregate groups; namely, mental ill-health (comprised of psychological distress, mood disorders, anxiety disorders, and behavioural disorders); substance use problems (comprised of alcohol, cannabis disorder, and drug use disorders); and all measures of mental ill-health and substance use problems taken together (comprised of the measures included in the mental ill-health and substance use groups, any other disorder, and suicidality). RVE allowed us to pool statistically dependent estimates (i.e., multiple estimates that are correlated because they arise from the same participant samples) into estimates incorporating all relevant measures for these aggregated groups without having to know or specify their covariance structures. Additionally, our analyses benefit from small-sample correction, which has been argued as necessary to implement when using RVE hypothesis testing [30]. It is important to note that, among small samples, hypothesis testing using RVE requires degrees of freedom be greater or equal to four to be accurate. Below four degrees of freedom, the $t$ distribution approximation on which testing is based no longer holds, and the type I error will be greater than indicated by the $p$ value being used [31]. We also conducted conventional random-effects models to obtain associations between NEET status and individual mental health and substance use problems. Forest plots were generated to display all main metaanalyses described above.

Our pooled results should be interpreted cautiously, given the highly heterogeneous study methodologies. We used the odds ratio (OR) as a summary measure since most studies with available quantitative data reported ORs. When multiple studies used the same dataset, we only included the study with the largest sample size to avoid double counting. For studies that only provided gender-stratified results, we combined the results using a fixed-effects model to include in the main meta-analysis. For studies that only reported
a $p$ value $<0.001$, we calculated a confidence interval (CI) assuming a conservative $p$ value of 0.001 . For studies reporting only a $p$ value $>0.05$, we assumed a conservative $p$ value of 0.10 . All intervals reported are $95 \%$ CIs.

To explore sources of heterogeneity, we conducted subgroup analyses by gender, age group ( $<18 \mathrm{vs} \geq 18$ years old), and sample type (population-based vs clinical). Moreover, we investigated the potential directionality of associations from longitudinal studies that examined NEET status as a predictor of later mental health and substance use problems, and studies that examined the inverse relationship. We used fixed-effects models for subgroup analyses because there were too few studies by subgroup to estimate betweenstudy variance with precision [32]. Study heterogeneity was evaluated using the $I^{2}$ index. We did not assess publication bias using quantitative methods because these are not recommended under conditions of high heterogeneity [33]. Meta-analyses were conducted in $R$ ( v 3.6 .1 ) using the meta, metafor, and robumeta packages [30, 34, 35].

## Results

From 6120 identified references, we included 24 studies (see PRISMA flow diagram in Fig. 1), which represented 548,862 unique individuals from the UK ( $k=6$ studies); Australia ( $k=4$, two using the same sample); Mexico ( $k=3$, all using the same sample); Sweden ( $k=3$ ); Italy ( $k=2$ ) [13, 36]; Canada ( $k=2$ ); Brazil ( $k=1$ ); Norway ( $k=1$ ); Ireland ( $k=1$ ); Switzerland ( $k=1$ ); and Greece ( $k=1$ ). Study characteristics, detailed study findings, and quality assessment ratings appear in Table 1 and Supplementary Appendices 4 and 5, respectively. For 11 studies, the age range of the sample at baseline was under 18 years; three studies only included samples above the age of 19 ; and 10 studies included samples that were both below and above 18 years of age (e.g., 15-25). Cohort studies were most common ( $k=13$ ), followed by cross-sectional $(k=10)$ and case-control $(k=1)$ studies. Being NEET was associated with at least one measure of mental health or substance use problems in $75 \%$ of studies (18/24). Study quality was low in five studies; moderate in nine studies; and high in 10 studies. Measures of mental health and substance use problems included problems or symptoms of $\operatorname{mood}(k=12)$, anxiety $(k=11)$, behaviour ( $k=8$ ), alcohol use ( $k=9$ ), cannabis use ( $k=6$ ), and drug use $(k=8)$ disorders, general psychological distress ( $k=9$ ), suicidal behaviours ( $k=7$ ), and any psychiatric disorder ( $k=5$ ).

4 out of the 24 studies were excluded from meta-analyses of odds ratios because two reported chi-squared statistics $[13,36]$ and 2 reported beta coefficients from linear regressions [11, 37]. Further, two pairs of studies used the same data to calculate the same estimates $[15,16,38,39]$. We

Fig. 1 PRISMA flow diagram

only included one study from each pair in the RVE metaanalyses [38, 39].

The meta-analyses found significant associations between NEET and all three aggregated groups, i.e., mental health problems ( $k=15$ studies; $n=25$ effect sizes, OR 1.28, CI 1.06-1.54; see Supplementary Appendix 6 for forest plot); substance use problems ( $k=11, n=16$, OR 1.43; CI 1.08-1.89; see Supplementary Appendix 7 for forest plot); and all measures of mental health problems, substance use problems, and suicidality combined $(k=18, n=48$, OR 1.38 ; CI 1.15-1.64).

Table 2 presents summaries of findings by type of mental health or substance use problem and Fig. 2 presents forest plot by each type of problem. The evidence most consistently pointed to an association between NEET and symptoms of mood disorders $[12,14,17,38,40,41]$ ( $k=6$ nonoverlapping studies; OR 1.43, CI 1.21-1.70); behavioural disorders [12, 14, 19, 42-44] ( $k=6$; OR 1.49, CI 1.21-1.85); cannabis use problems $[14,17,38,40,44,45](k=6$; OR 1.62, CI 1.07-2.46); drug use problems [12, 14, 19, 40, 46] ( $k=5$; OR 1.99, CI 1.19-3.31); any psychiatric disorder [12,

18, 44] ( $k=3$; OR 1.72, CI 1.37-2.16); and suicidal behaviours $[12,14,18,40](\mathrm{k}=4$; OR 2.84, CI 2.04-3.95).

The evidence was more mixed-with fewer than $50 \%$ of non-overlapping studies reporting a significant finding-for NEET status being associated with anxiety disorders [12, $14,18,40,46]$ ( $k=5$ non-overlapping studies; OR 1.55, CI 1.07-2.24); alcohol use problems ( $k=5$; OR 1.28, CI 1.12-1.46); and psychological distress ( $k=7$; OR 1.10, CI $1.01-1.21$ ). Results were similar when we excluded low quality studies.

## Sub-group analyses

For each of the three aggregate groups, subgroup analyses were conducted for directionality, age, and gender. There was evidence of mental health problems (first aggregated group; $k=8$ studies; $n=12$ effect sizes, OR 1.33 , CI 1.01-1.74) and all measures combined (third aggregated group; $k=9 ; n=19$, OR 1.39 , CI 1.03-1.86) being associated with subsequent NEET status. Other subgroup analyses conducted within the three aggregated groups had too few
Table 1 Characteristics of the selected studies

| References | Country, baseline year | Design | $N$ | Sample | NEET measure | Mental health and substance use measures | Direction(s) of association $^{\text {a }}$ | Statistical Methods | Covariates ${ }^{\text {b }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Baggio et al. [17] | Switzerland, 2010-2012 | Cohort | 4758 | Young men in their early 20's at baseline | Current status, excluding those in the military or civic service | General mental health, depressive symptoms, alcohol use, and cannabis use disorder | $\begin{aligned} & \text { MH <--> NEET } \\ & \text { NEET-> MH } \\ & \text { MH-> NEET } \end{aligned}$ | Logistic regression | Language, age, family SES |
| Bania et al. [42] | $\begin{aligned} & \text { Norway, 2003- } \\ & 2005 \end{aligned}$ | Cohort | 3987 | Youth 15-16 years old at baseline | Did not complete post-secondary school and unemployed for $1+$ year, or received $6+$ months of sickness benefits during the 9 -year study period (2003-2012) | Conduct and emotional problems | MH-> NEET | Logistic regression | Gender, residency, ethnicity, parental education |
| Basta et al. [46] | Greece, 2016 | Cross-sectional | 2771 | Youth 15-24 years old | Current status | Anxiety and depressive symptoms, drug use | MH < - N NEET | Logistic regression | Gender, insurance, income, living with parents, being married, financial support by others |
| Benjet et al. [12] | Mexico, 2005 | Cross-sectional | 3005 | Youth 12-17 years old | Current status | Mood, anxiety, substance use and behavioural disorders, and suicidal behaviours | MH < - N NEET | Logistic regression | Age, gender, SES, marital status, having children, living with parents, parental SES |
| Bynner and Parsons [20] | United Kingdom, 1987 | Cohort | 930 | Youth 16-18 years old at baseline not in full-time education | NEET for $6+$ months between 16-18 years old | Psychological distress | NEET - > MH | Logistic regression | Birth weight, cognitive abilities, hobbies, and family circumstances in childhood |
| Cairns et al. [37] | Australia, 2013 | Cross-sectional | 226 | Youth 15-25 years old seeking services | Current status, excludes those in carer role | Psychological distress, history of mental health diagnosis, and history of illicit drug use | MH <-> NEET | Multinomial logistic regression | Age, gender, secondary school dropout, psychological distress, history of mental health diagnosis, history of illicit drug use |

Table 1 (continued)

| References | Country, baseline year | Design | $N$ | Sample | NEET measure | Mental health and substance use measures | Direction(s) of association ${ }^{\text {a }}$ | Statistical Methods | Covariates ${ }^{\text {b }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Gariépy and Iyer [40] | Canada, 2014 | Cross-sectional | 5622 | Youth 15-29 years old | Current status | Depression, bipolar, generalized anxiety, alcohol use, cannabis use, and drug use disorders, and suicidal behaviours | MH $<->$ NEET | Logistic regression | Age, gender, SES, ethnicity, immigrant status, living arrangement, health conditions, location |
| Goldman-Mellor et al. [14] | United Kingdom, 2012-2013 | Cohort | 2232 | Twins 17-18 years old at baseline | Current status, excludes parents | Depressive, generalized anxiety, alcohol, cannabis, and conduct disorders (age 18), and childhood depressive, anxiety, substance use disorders, and suicidal behaviours (age 12) | $\begin{aligned} & \text { MH-> NEET } \\ & \text { MH <-> NEET } \end{aligned}$ | Modified Poisson and logistic regression | Gender, cognitive ability, family SES, neighborhood, childhood mental health problems |
| Gutierrez-Garcia et al. [15] | Mexico, 2005 | Cohort | 1071 | Youth 12-17 years old at baseline | Current status | Mood, anxiety, alcohol use, substance use and behavioural disorders, and suicidal behaviours | NEET-> MH | Logistic regression | Gender, age, marital status, family SES |
| Gutierrez-Garcia et al. [39] | Mexico, 2013 | Cross-sectional | 1071 | Youth 19-26 years old | Current status | Mood, anxiety, behavioural, and substance use disorders; suicidal behaviours | MH <-> NEET | Logistic regression | Gender, marital status, has children, education, living with family of origin |
| Hale and Viner [49] | United Kingdom, 2004 | Cohort | 8682 | Youth 13 years old at baseline | Current status | Psychological distress | MH-> NEET | Logistic regression | SES, ethnicity and educational attainment |

Table 1 (continued)

| References | Country, baseline year | Design | $N$ | Sample | NEET measure | Mental health and substance use measures | Direction(s) of association ${ }^{\text {a }}$ | Statistical Methods | Covariates ${ }^{\text {b }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Hammerton et al. [43] | Brazil, 2004-2005 United Kingdom, 2002-2003 | Cohort | $\begin{aligned} & \text { Brazil: } \\ & 3939 ; \\ & \text { UK: } 5079 \end{aligned}$ | Youth age 11 years old at baseline | Current status | Conduct problem and oppositional problem | MH-> NEET | Logistic regression | Sex, parental separation, fear of the neighborhood at age 11, maternal risk factors (SES, smoking, depression, unplanned pregnancy, alcohol use, urinary infection during pregnancy), birth factors (intrauterine growth restriction, gestational age, premature birth) |
| Henderson et al. [19] | $\begin{aligned} & \text { Canada, 2009- } \\ & 2013 \end{aligned}$ | Cross-sectional | 2576 | Youth 12-24 years old seeking services | Current status | Internalizing, externalizing, and substance use disorders | MH $<->$ NEET | Chi-square tests | Age |
| López-López et al. [41] | United Kingdom, 2002-2003 | Cohort | 4501 | Youth age 11 years old at baseline | Current status | Trajectories of depressive symptoms | MH-> NEET | Logistic regression | Sex, IQ, maternal postnatal depression, maternal education, attitude towards school and academic results at age 11 |
| Manhica et al. [48] | Sweden, 2005-2009 | Cohort | 485,839 | Youth 19-24 years old with secondary education | NEET indicator of labour market attachment in past year | Alcohol use disorder | NEET $->$ M ${ }^{\text {P }}$ | Cox regression | Sex, age, domicile and origin |
| Nardi et al. [13] | Italy, 2010-2011 | Case-control | 228 | NEET youth from juvenile court services or the community and non-NEET youth from a technical institute, $16-23$ years old | Not stated | Symptoms of nervousness, mood swings, or thoughts of suicide | MH <-> NEET | Chi-square tests | None |

Table 1 (continued)

| References | Country, baseline year | Design | $N$ | Sample | NEET measure | Mental health and substance use measures | Direction(s) of association ${ }^{\text {a }}$ | Statistical Methods | Covariates ${ }^{\text {b }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Nardi et al. [36] | Italy, 2010-2011 | Cross-sectional | 143 | Youth involved in criminal proceedings, 16-19 years old | Current status | Psychiatric disorders diagnosed by psychiatric services | MH $<->$ NEET | Chi-square tests | None |
| O'Dea et al. [38] | Australia, 2011-2012 | Cross-sectional | 676 | Youth seeking services $15-25$ years old | Current status | Symptoms of mood, anxiety, alcohol use, and cannabis use disorders | MH $<->$ NEET | Logistic regression | Age, gender, criminal charges, economic hardship, self-rated disability, clinical stage |
| O'Dea et al. [16] | Australia, 2011-2012 | Cohort | 448 | Youth seeking services 15-25 years old at baseline | Past-month status | Depressive and generalized anxiety disorders | $\begin{aligned} & \text { MH }->\text { NEET } \\ & \text { NEET }->\text { MH } \end{aligned}$ | Multinomial logistic regression and chisquare tests | Age, gender, location, immigrant background |
| Power et al. [18] | $\begin{aligned} & \text { Ireland, 2000- } \\ & 2002 \end{aligned}$ | Cohort | 168 | Youth 15-25 years old | Not stated | Mood, anxiety, alcohol, substance use disorders, and suicidal behaviours | $\begin{aligned} & \text { MH < - > NEET } \\ & \text { MH-> NEET } \end{aligned}$ | Logistic regression | Gender and SES |
| Rodwell et al. [57] | Australia, 1992-1993 | Cohort | 1938 | Youth 14-15 years old at baseline | Current status | Common mental disorders; disruptive, alcohol use, and cannabis use disorders | MH-> NEET | Logistic regression | Gender, marital status, location, parental education, year |
| Stea et al. [50] | Sweden | Cross-sectional | 480 | NEET youth from vocational services and non-NEET youth attending high school, 16-21 years old | Current status | Psychological distress | MH $<->$ NEET | Logistic regression | Sex, parental education |
| Stea et al. [45] | Sweden | Cross-sectional | 480 | NEET youth from vocational services and non-NEET youth attending high school, 16-21 years old | Current status | Cannabis use | MH $<->$ NEET | Logistic regression | Sex, age, parental education |

Table 1 (continued)

| References | Country, baseline year | Design | $N$ | Sample | NEET measure | Mental health and substance use measures | Direction(s) of association ${ }^{\text {a }}$ | Statistical Methods | Covariates ${ }^{\text {b }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Symonds et al. [11] | United Kingdom, 2005 | Cohort | 11,082 | Youth 14-15 years old at baseline | Current status, excludes those caring full-time for others | Symptoms of depression, anxiety, and positive mental functioning | MH-> NEET | Linear regression | Gender, ethnicity, SES, childhood achievement |

[^1]degrees of freedom $(d f<4)$ to be reliable, so results should be considered exploratory (Supplementary Appendix 8).

## Directionality of association

Longitudinal data provided some evidence for bidirectional associations (see Supplementary Appendix 9 for summary of significant findings of studies by directionality of association and type of mental or substance use disorder or symptoms). Ten studies measured symptoms of mental ill-health and/ or substance use problems before the emergence of NEET status. Symptoms of mood disorders [14, 16, 17, 41] $k=4$ non-overlapping studies; OR 1.12, CI 1.07-1.18); behavioural problems [42-44] ( $k=3$; OR 1.25, CI 1.19-1.32); cannabis use problems [17, 44] ( $k=2$; OR 1.10, CI 1.04-1.15); drug use problems [14] ( $k=1$; OR 1.89, CI 1.29-2.77); any mental disorder [18, 44] ( $k=2$; OR 1.83, CI 1.27-2.62); and suicidal behaviours [14] ( $k=1$; OR 3.30, CI 2.07-5.27) were associated with later NEET status. Alcohol use disorder [17, 44] was not so associated ( $k=2$; OR 0.80 , CI $0.48-1.34$ ), and evidence was equivocal for anxiety symptoms/disorders [14, 16] ( $k=2$; OR 1.38, CI $0.81-2.36$ ) and psychological distress [11, 17, 42, 47] ( $k=4$; OR 1.04, CI 1.00-1.08).

Five studies examined NEET status prior to mental health and substance use outcomes. NEET status predicted later suicidal behaviours in a single study [15] ( $k=1$; OR 2.40, CI 1.32-4.31); symptoms of mood disorder in one study [15] ( $k=1$; OR 1.67, CI 1.12-1.90), but not another [16] ( $k=1$; OR 1.94, CI 0.17-21.60); and alcohol use disorder in two studies $[15,48](k=2$; OR 1.22, CI 1.12-1.32), but not another [17] ( $k=1 ; p>0.05$, values unavailable). Being NEET did not predict later symptoms of anxiety [15, 16] ( $k=1$; OR 0.40, CI 0.10-1.65); behavioural problems [15] ( $k=1$; OR 0.83, CI 0.45-1.50); cannabis use [17] ( $k=1$; $p>0.05$, values unavailable); drug use problems [15] ( $k=1$; OR 1.03, CI 0.71-1.50); or psychological distress [9, 17] ( $k=2$; OR 1.78, CI 0.93-3.42).

## Associations by gender

Six studies conducted gender-stratified analysis [19, 20, $42,46,49,50]$. In Bania et al. [42], conduct problems at age $15-16$ predicted becoming NEET 9 years later among men (OR 1.17, CI 1.07-1.28) and women (OR 1.25, CI 1.17-1.33), while emotional problems were associated with lower odds of becoming NEET in men (OR 0.88 , CI $0.81-0.97$ ), but not women (OR 1.04, CI 0.97-1.11).

Hale and Viner [49] found that psychological distress at age 13 predicted being NEET at age 19 among men (OR 1.72, CI 1.24-2.41) and women (OR 1.49, CI 1.11-1.99). Bynner and Parsons [20] found that being NEET at age $16-18$ did not predict psychological distress at age 21 in
Table 2 Summary of findings by type of mental health and substance use disorder

| Type | All studies |  | Studies reporting odds ratios |  |  |  | Studies reporting beta coefficients |  | Studies reporting Chi-square statistics |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Number of studies | \% studies reporting a statistically significant association | Number of non-overlapping studies | Number of excluded studies ${ }^{\text {a }}$ | Pooled OR (95\% CI) | $I^{2}(\%)$ | Number of studies | Beta coefficient (95\% CI) | Number of studies | $p$ value |
| Mood | 12 | 75\% (9/12) | 6 | 3 | 1.43 (1.21, 1.70) | 88.0 | 1 | $\beta 0.10(p<0.05)^{\dagger}$ | 1 | $p>0.05^{\text {* }}$ |
| Anxiety | 10 | 40\% (4/10) | 5 | 3 | 1.55 (1.07, 2.24) | 79.5 | 1 | $\beta 0.10(p<0.05)^{\dagger}$ | 1 | $p>0.05^{*}$ |
| Behavioral | 8 | 63\% (5/8) | 6 | 2 | 1.49 (1.21, 1.85) | 87.0 | 0 | - | 0 | - |
| Alcohol use | 9 | 33\% (3/9) | 6 | 3 | 1.28 (1.12, 1.46) | 21.7 | 0 | - | 0 | - |
| Cannabis use | 6 | 66\% (4/6) | 6 | 0 | 1.62 (1.07, 2.46) | 83.4 | 0 | - | 0 | - |
| Drug use | 8 | 75\% (6/8) | 5 | 3 | 1.99 (1.19, 3.31) | 89.4 | 0 | - | 0 |  |
| Any disorder | 5 | 90\% (4/5) | 3 | 1 | 1.72 (1.37, 2.16) | 0.0 | 0 | - | 1 | $p>0.05^{\text {* }}$ |
| Suicidal behaviors | 7 | 71\% (5/7) | 4 | 2 | 2.84 (2.04, 3.95) | 15.6 | 0 | - | 1 | $p>0.05^{\text {* }}$ |
| Psychological distress | 9 | 33\% (4/12) | 7 | 1 | 1.04 (0.96, 1.14) | 82.7 | 1 | $\beta 0.06(p<0.05)^{\dagger}$ | 0 | - |

${ }^{\text {a }}$ Studies with overlapping data or with specific OR not reported ${ }^{\dagger} 95 \% \mathrm{CI}$ or specific $p$ value not reported

| Study |  | Odds Ratio [95\% CI] |
| :---: | :---: | :---: |
| Mood |  |  |
| Baggio et al 2015 | $\square$ | 1.11 [1.04; 1.18] |
| Benjet et al 2012 | - | 2.70 [1.77; 4.12] |
| Gariépy \& lyer 2019 | - | 1.70 [1.11; 2.61] |
| Goldman-Mellor et al 2016 | - | 1.68 [1.33; 2.13] |
| López-López et al 2019 |  | 5.17 [1.95; 13.70] |
| O'Dea et al 2014 |  | 1.10 [1.05; 1.15] |
| Random effects model | $\diamond$ | 1.43 [1.21; 1.70] |
| Heterogeneity: $I^{2}=88 \%, p<0.01$ |  |  |
| Anxiety |  |  |
| Basta et al 2019 |  | 1.03 [1.00; 1.07] |
| Benjet et al 2012 | - | 1.30 [0.92; 1.84] |
| Gariépy \& lyer 2019 | - | 2.60 [1.36; 4.96] |
| Goldman-Mellor et al 2016 | $\square$ | 1.84 [1.19; 2.85] |
| Power et al 2015 |  | 2.40 [0.99; 5.83] |
| Random effects model | $\infty$ | 1.55 [1.07; 2.24] |
| Heterogeneity: $l^{2}=80 \%, p<0.01$ |  |  |
| Behavioural |  |  |
| Bania et al 2019 | $\square$ | 1.22 [1.16; 1.29] |
| Benjet et al 2012 | - | 2.10 [1.54; 2.87] |
| Goldman-Mellor et al 2016 |  | 1.46 [1.12; 1.90] |
| Hammerton et al 2019 | - | 1.38 [1.13; 1.69] |
| Hammerton et al 2019 | - | 3.04 [1.99; 4.65] |
| Henderson et al 2017 |  | 0.90 [0.75; 1.07] |
| Rodwell et al 2018 | $\cdots$ | 1.71 [1.15; 2.55] |
| Random effects model | $\diamond$ | 1.49 [1.21; 1.85] |
| Heterogeneity: $I^{2}=87 \%, p<0.01$ |  |  |
| Alcohol use |  |  |
| Gariépy \& lyer 2019 |  | 1.00 [0.59; 1.68] |
| Goldman-Mellor et al 2016 | - | 1.42 [1.01; 2.00] |
| Gutierrez-Garcia et al 2018 | 픔 | 1.38 [1.16; 1.65] |
| Manhica et al 2019 | $\square$ | 1.30 [1.11; 1.52] |
| Rodwell et al 2018 | - | 0.80 [0.48; 1.34] |
| Random effects model | $\diamond$ | 1.28 [1.12; 1.46] |
| Heterogeneity: $l^{2}=22 \%, p=0.28$ |  |  |
| Cannabis use |  |  |
| Baggio et al 2015 |  | 1.09 [1.04; 1.15] |
| Gariépy \& lyer 2019 |  | 1.00 [0.50; 2.00] |
| Goldman-Mellor et al 2016 | - | 3.93 [2.31; 6.68] |
| O'Dea et al 2014 |  | 1.22 [0.72; 2.06] |
| Rodwell et al 2018 | - | 1.74 [1.10; 2.75] |
| Stea, de Ridder et al 2019 | $\square$ | 2.20 [1.11; 4.35] |
| Random effects model | $\infty$ | 1.62 [1.07; 2.46] |
| Heterogeneity: $I^{2}=83 \%, p<0.01$ |  |  |
| Drug use |  |  |
| Basta et al 2019 | - | 0.80 [0.58; 1.10] |
| Benjet et al 2012 | - | 4.40 [2.57; 7.52] |
| Gariépy \& lyer 2019 |  | 3.20 [1.31; 7.84] |
| Goldman-Mellor et al 2016 | -- | 1.89 [1.29; 2.77] |
| Henderson et al 2017 |  | 1.94 [1.63; 2.31] |
| Random effects model | $\bigcirc$ | 1.99 [1.19; 3.31] |
| Heterogeneity: $I^{2}=89 \%, p<0.01$ |  |  |
| Any disorder |  |  |
| Benjet et al 2012 |  | 1.70 [1.28; 2.26$]$ |
| Power et al 2015 | , | 2.90 [1.13; 7.42] |
| Rodwell et al 2018 | $\square$ | 1.60 [1.07; 2.40] |
| Random effects model | $\Delta$ | 1.72 [1.37; 2.16] |
| Heterogeneity: $I^{2}=0 \%, p=0.52$ |  |  |
| Suicidal behaviours |  |  |
| Benjet et al 2012 - 3.60 [2.01; 6.44] |  |  |
| Gariépy \& lyer 2019 - 1.80 [1.02; 3.17] |  |  |
| Goldman-Mellor et al 2016 - 3.30 [2.07; 5.27] |  |  |
| Power et al 2015 - 3.10 [1.02; 9.45] |  |  |
| Random effects model <br> Heterogeneity: $I^{2}=16 \%, p=0.31$ $\infty$ |  |  |
|  |  |  |
| Psychological distress |  |  |
| Baggio et al 2015 | $\pm$ | 1.11 [1.04; 1.18] |
| Bania et al 2019 |  | 0.97 [0.92; 1.03] |
| Basta et al 2019 |  | 0.98 [0.96; 1.01] |
| Bynner \& Parsons 2002 - 1.78[0.93; 3.42] |  |  |
| Hale \& Viner 2018 - 1.57[1.25; 1.98] |  |  |
| Henderson et al 2017 - 1.15 [0.95; 1.39] |  |  |
| Stea, Abildsnes et al 2019 |  | 2.40 [1.05; 5.47] |
|  |  | 1.10 [1.01; 1.21] |
| Heterogeneity: $I^{2}=85 \%, p<0.01$ |  |  |
| 0.1 | $\begin{array}{llll}0.5 & 1 & 2 & 10\end{array}$ |  |

Fig. 2 Forest plot from the meta-analysis disaggregated by mental-ill health, substance use problems and suicidality measures
men (OR 2.20, $p>0.05$, CI unavailable) and women (OR $1.69, p>0.05)$.

In their cross-sectional study, Stea et al. [50] found an association between NEET and psychological distress among women (OR 2.40, CI 1.00-5.20), but not men (estimate unavailable), whereas Basta et al. [46] found no association with distress among women (OR 0.98 , CI 0.95-1.02) and men (OR 0.99, CI 0.96-1.03).

Basta et al. [46] found an association between anxiety problems and NEET among women (OR 1.05, CI $1.01-1.10$ ), but not men (OR 1.01, CI 0.96-1.03). In a crosssectional study, Henderson et al. [19] found that being NEET was associated with substance misuse among men (OR 1.83, CI 1.43-2.34) and women (OR 2.05, CI 1.58-2.66); externalizing disorders among both men (OR 0.93 , CI $0.73-1.19$ ) and women (OR 0.87, CI 0.67-1.12); and internalizing symptoms among men (OR 1.39, CI 1.08-1.78), but not women (OR 1.08, CI 0.80-1.45).

## Associations by age

We compared findings for participants who were $<18$ years ( $k=4$ studies) [12, 19, 20, 42] and $\geq 18$ years old $(k=7)$ [14, $17-19,39,44,49]$. The association was consistent among younger ( $<18$ years) youth between NEET status and mood problems [11, 12] (beta coefficient $0.0710, p<0.05$ in one study; OR 2.70, CI 1.77-4.12 in the other study); behavioural problems [12, 19, 42] ( $k=3$; OR 1.24, CI 1.18-1.30); and drug use problems ( $k=2$; OR 1.69, CI 1.38-2.07). Results were weaker for psychological distress [19, 20, 42] ( $k=3$; OR 0.97 , CI $0.92-1.03$ ) and anxiety problems [11, 12] ( $k=2$; OR 1.30 , CI $0.92-1.84$ in one study; beta coefficient $=0.07, p>0.05$ in other study).

In youth $\geq 18$ years old, there was an association with anxiety disorders $[14,18,39](k=3$; OR 1.59 , CI 1.12-2.26), behavioural disorders [14, 19, 39, 44] $(k=4$; OR 1.32, CI 1.12-1.55); cannabis use problems ( $k=3$; OR 1.11, CI 1.05-1.16); any disorder [18, 44] ( $k=2$; OR 1.76, CI 1.21-2.54); and general psychological distress ( $k=3$; OR 1.15, CI 1.08-1.22).

In youth $\geq 18$ years old, evidence was mixed for symptoms of mood disorder $[14,17,18,39](k=4$; OR $(n=3)$ 1.14, CI 1.07-1.21; and missing OR with $p>0.05$, CI unavailable in other study); drug use disorders [14, 19, 39, 44] ( $k=4$; OR ( $n=3$ ) 1.99, CI 1.61-2.45; and missing OR with $p>0.05$, CI unavailable in other study); and alcohol use problems $[14,17,18,39,44](k=5$; OR $(n=3) 1.32$, CI 1.14-1.54; missing OR with $p>0.05$, CI unavailable in other studies).

## Associations by sample type

Similar patterns of association emerged between studies using clinical $[16,19,37,38]$ ( $k=4$ studies) and populationbased samples [11, 12, 15, 17, 18, 20, 39-44, 46, 48, 49] ( $k=18$ ).

In clinical studies, service-seeking youth were more likely to be NEET if they presented with mood disorders [16, 38]; current [19] or past [37] drug use disorders; or co-occurring mental health problems [19]; but not if they had problems with alcohol or cannabis use [16, 38], anxiety disorders [16, 38], psychological distress [37], externalizing problems [19], or a history of any mental health diagnosis [37].

## Discussion

This is the first comprehensive systematic review and metaanalysis on the association between NEET status and mental health and substance use problems in youth. Being NEET was associated with mental health problems, substance use problems, and all measures combined in aggregate analyses. When disaggregated, NEET was most consistently associated with suicidal behaviours, drug use problems, any psychiatric disorders, cannabis use problems, behavioural problems, and mood problems. Findings for the association between NEET and anxiety problems, alcohol use, and psychological distress were mixed. Results were generally consistent across clinical and population-based samples but mixed across gender. These associations were particularly consistent among younger youth ( $<18$ years old). Longitudinal data indicated that mental health problems in early youth predicted a later NEET status, while evidence for the inverse relationship was equivocal and sparse. Together, these results point to early youth as a sensitive period for mental health and substance use problems becoming related with being NEET and increasing the vulnerability to later becoming NEET.

The aggregate analyses showed meaningful and significant associations between mental health and substance use problems in youth and being NEET. Although the overall evidence is based on a relatively limited and heterogeneous body of literature, the studies were generally of moderate to high quality. These results align with previous reviews on youth unemployment [21, 22] and school dropout [23, 24] that report a close connection between vocational disengagement and poor mental health. Our review extends this literature by focussing on youth disengagement from both education and employment and by revealing that the strength and consistency of associations vary across types of mental health and substance-use problems.

In our study, meta-analytical evidence from longitudinal data suggested that mental health problems and all measures
of mental health and substance use (combined) predicted becoming NEET later. This evidence for their increased risk of becoming NEET aligns with the well-documented [24] drain of mental and substance use disorders on youths' ability to perform at school and work. Disengagement from school and work may further disadvantage those with mental health problems, widening the gap between them and peers who follow more engaged developmental trajectories. Disengagement may also further heighten feelings of shame, hopelessness, and social exclusion [20, 51]. Analyses for NEET status predicting the later occurrence of mental health problems aggregated, substance use problems aggregated, and all measures combined were not conclusive. Nonetheless, there was evidence for NEET status predicting individual mental health/substance use problems, suggesting that being out of school and work, especially in early youth, could lead to mental health and substance use problems. Regardless of the directionality, school and work can provide crucial structures and experiences that enhance feelings of belonging, productivity, and hope for the future [52].

Contrary to our hypothesis, we discerned no clear genderbased pattern in the link between mental health problems and being NEET, although the evidence base was limited and most gender-stratified studies focussed on psychological distress. Nonetheless, there is evidence that the experience of being NEET could vary by gender. For instance, young women who are NEET are more likely to be stay-at-home parents or caretakers [53, 54]. Further, the consequences of being NEET may differ by gender. A British longitudinal study [20] found that, for young men, being NEET mainly impacted their job prospects, while for young women, it further affected their psychological well-being. To develop tailored strategies to prevent youth from becoming NEET or developing mental health problems when NEET, further research is needed into the intersections between gender and other subgroupings of vulnerability, NEET status, and mental illness.

Our review was constrained by the heterogeneity of mental health measures used in the reviewed studies, ranging from specific disorder subtypes (e.g., generalized anxiety disorder) to broad categories (e.g., any anxiety disorder) and general symptom scales. Many mental disorders like psychosis, eating disorders, and personality disorders, were not represented and few studies reported on comorbidity [19]. Divergent definitions of NEET status also limited comparability. Non-paid work like parenting counted as employment in some studies [14, 44], but not others. Most studies measured current NEET status, but some used timeframes from 1 month [16] to 9 years [42]. By assessing NEET status but not its duration, almost all studies captured the association of both short- and long-term vocational disengagement with mental health and substance use problems. To formulate more effective interventions and policies, research into the
duration of NEET status and its association with mental and substance disorders is therefore needed. These definitional, methodological, and cultural challenges of measuring NEET status and the heterogeneity of its circumstances have also been previously noted [27, 54, 55].

We could not examine contextual/cultural influences on being NEET and mental health problems because the reviewed studies were from a limited number of specific geographical and political backgrounds. All the studies from this review were from Europe, North America, or Australia, with the exception of one study that included data from South America, limiting the generalizability of the evidence to other contexts like low- and middle-income countries. Furthermore, global and country-specific economic shifts may exacerbate associations between NEET status and mental-ill health and deserve exploration in the future. Evidence is from observational data thereby limiting direct causal inference. While we examined longitudinal studies to assess the potential directionality of association, none of the studies used specific panel regressions models and may therefore be biased by unobserved heterogeneity. Like other reviews, our findings may be affected by publication bias. While we could review papers in English, French, and Spanish, only English studies met our search criteria. We excluded studies that focussed on neurodevelopmental disorders or disabilities that are typically diagnosed in childhood. We recognize that these disorders could co-occur with mental and substance use disorders and contribute to being NEET and may even differ in their relationship with NEET compared to other mental disorders, and therefore should be examined in future work.

Notwithstanding these limitations, the studies provided data from diverse contexts and on a range of mental health and substance use outcomes, with generally consistent results despite methodological differences. Our review carefully assessed the association between being NEET and mental health and substance use problems, an emerging topic with important clinical and public health implications. We used rigorous methodology to search, systematically assess, and analyse current literature to explain our findings. Using RVE, we appropriately pooled multiple mental health measure estimates that were correlated because they came from the same participant samples. This allowed us to capture associations between NEET and overarching groups of mental health and substance use problems that reflect a generalized relationship between youth engagement and mental-ill health. In addition, we used subgroup analysis to investigate heterogeneity, directionality of association, and vulnerable subgroups.

We identified significant knowledge gaps in NEET and mental health research. First, the association between being NEET and mental health and substance use problems is likely context-sensitive and broadening the geographic ambit
of studies is strongly recommended. Second, the association is likely marked by gender differences that bear teasing out. Third, rigorous research on the temporal relationship of mental disorders and being NEET is needed because the question of directionality remains unresolved. Moreover, the associations between duration and recurrence of NEET status and mental ill-health have yet to be systematically explored. Fourth, information about some mental disorders (e.g., psychosis) and comorbid disorders and their associations with being NEET is lacking. Finally, future research should include intervention studies to identify whether and for whom vocational and mental health supports are useful in averting and ending NEET status.

Realization of the loss to productivity and the wealth of nations from unaddressed youth mental health problems is increasing [56]. Although more longitudinal research is needed, our review found clear evidence for NEET status being a consequence of mental health problems and substance misuse. Efforts to prevent young people from becoming or remaining vocationally and socially disengaged should therefore include provisions for the prevention of and early intervention for mental health problems. Furthermore, because there is also evidence for a bidirectional relationship between NEET status and mental ill-health and because problems with vocational functioning are well-documented among youth with mental health problems [17, 57], youth mental health services should integrate educational and employment supports and services to address vocational needs and promote recovery.

The connectedness of vocational disengagement and mental health problems among young people underlines the need for consistent, widespread policy support for broader-spectrum integrated youth-focussed services [58, 59]. Our review also highlights the importance of schools, universities, and employers developing the will and capacity to address the needs of youth experiencing mental health problems. The socioeconomic disruptions and mental health implications of the ongoing pandemic make these needs ever more urgent. Our comprehensive synthesis can serve as a useful pre-pandemic reference point for future research on the associations between youth employment/education and mental health and substance use over the course of or after the COVID-19 pandemic.

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## Declarations

Conflict of interest The authors declare that they have no conflict of interest.

Availability of data and material/code availability Not applicable.
Ethics approval and consent to participate The manuscript does not contain clinical studies or patient data.

Consent for publication Not applicable.

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[^1]:    ${ }^{3}$ NEET: not in education, employment, or training; MH: mental health

