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Population structure of chum salmon and selection on the markers collected for stock identification

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Abstract

Genetic stock identification (GSI) is a major management tool of Pacific salmon (Oncorhynchus Spp.) that has provided rich genetic baseline data of allozymes, microsatellites, and single-nucleotide polymorphisms (SNPs) across the Pacific Rim. Here, we analyzed published data sets for adult chum salmon (Oncorhynchus keta), namely 10 microsatellites, 53 SNPs, and a mitochondrial DNA locus (mtDNA3, control region, and NADH-3 combined) in samples from 495 locations in the same distribution range (n = 61,813). TreeMix analysis of the microsatellite loci identified the greatest convergence toward Japanese/Korean populations and suggested two admixture events from Japan/Korea to Russia and the Alaskan Peninsula. The SNPs had been purposively collected from rapidly evolving genes to increase the power of GSI. The largest expected heterozygosity was observed in Japanese/Korean populations for microsatellites, whereas it was largest in Western Alaskan populations for SNPs, reflecting the SNP discovery process. A regression of SNP population structures on those of microsatellites indicated the selection of the SNP loci according to deviations from the predicted structures. Specifically, we matched the sampling locations of the SNPs with those of the microsatellites and performed regression analyses of SNP allele frequencies on a 2-dimensional scaling (MDS) of matched locations obtained from microsatellite pairwise F_{ST} values. The MDS first axis indicated a latitudinal cline in American and Russian populations, whereas the second axis showed differentiation of Japanese/Korean populations. The top five outlier SNPs included mtDNA3, U502241 (unknown), GnRH373, ras1362, and TCP178, which were identified by principal component analysis. We summarized the functions of 53 nuclear genes surrounding SNPs and the mtDNA3 locus by referring to a gene database system and propose how they may influence the fitness of chum salmon.

KEYWORDS

demographic history, hatchery, microsatellite, mtDNA, population structure, SNPs

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1 | INTRODUCTION

Chum salmon (*Oncorhynchus keta*) has a wide distribution range from the Arctic to California in North America and from Siberia to Japan and Korea in Asia (Salo, 1991). The ability to adapt to the varied environments of spawning rivers in the past and at present may have expanded its distribution range. Chum salmon also comprise the world's largest hatchery release program (Amoroso et al., 2017; Flagg, 2015; Kitada, 2018; Naish et al., 2007), and more than 3 billion hatchery-reared chum salmon are released annually into the North Pacific. Therefore, chum salmon is one of the best species to study the structure and history of populations, which could be underpinned by environmental adaptation and anthropogenic selection pressures.

The genetic population structure of chum salmon has traditionally been studied in the context of genetic stock identification (GSI). GSI has been a major management tool of Pacific salmon (Oncorhynchus Spp.) since the early 1980s (Beacham, Candy, Wallace, et al., 2009; Beacham, Sato, et al., 2008; Beacham, Varnavskaya, et al., 2008; Beacham, Wetklo, et al., 2008; Beacham, Winther, et al., 2008; Beacham et al., 2020; Grant et al., 1980; Larson et al., 2014; Shaklee et al., 1999; Utter, 1991), and it has contributed to understanding high-seas and coastal migration patterns (Myers et al., 2007; Seeb et al., 2004). In Pacific salmon, samples from mixed-stock fisheries and forensic studies have been analyzed to provide optimal resolution of proportions of mixed stocks at a reasonable cost (Beacham et al., 2020). GSI studies have provided genetic baseline data for salmon populations across the Pacific Rim, and these data have contributed to studies into population structure, mixed-stock fisheries, and genetic interactions between hatchery and wild salmon (Waples et al., 2020).

Genetic markers for GSI have progressed from allozymes to microsatellites and single-nucleotide polymorphisms (SNPs) (Beacham et al., 2020; Bernatchez et al., 2017). Allozyme loci often have a small number of alleles. To improve the power of GSI resolution for the high gene flow salmonids, microsatellites were developed because the number of alleles is generally much larger than that of allozymes, and much more information can be included. However, standardizing hundreds of microsatellite alleles across sampling points in different countries is difficult (Seeb et al., 2011). To avoid the standardization problem, genotyping of microsatellites of salmon species was generally performed by a single laboratory (Beacham, Candy, Le, et al., 2009; Beacham, Candy, Wallace, et al., 2009; Beacham, Sato, et al., 2008; Beacham, Varnavskaya, et al., 2008; Seeb et al., 2011). In contrast, calibrating SNP genotyping is more straightforward because genotype data can be stored in a unified format and can be accessed by different laboratories on different continents (Waples et al., 2020).

Populations of chum salmon have been widely surveyed for genetic variability and show large allele frequency differences in three studies (Elfstrom et al., 2007; Smith, Baker, et al., 2005; Smith et al., 2005). The markers selected for SNP typing were originally identified as rapidly evolving genes (Elfstrom et al., 2007; Seeb et al., 2011) that also showed positive selection in humans and chimpanzees (Nielsen et al., 2005). They included genes associated with fatty acid synthesis, testis-specific expression, olfactory receptors, immune responses, and cell growth and differentiation (Elfstrom et al., 2007; Smith, Baker, et al., 2005; Smith, Elfstrom, et al., 2005). The population structure determined using the SNPs selected for the GSI was influenced not only by natural selection on the genes but also by the SNP discovery process. Specifically, the three studies were focused on Western Alaska, which was the area of the authors' interest (Seeb et al., 2011). As a result, the SNP allelic richness and heterozygosity are high in Alaskan populations.

The use of neutral and adaptive markers in various combinations can be useful in establishing optimal management strategies (Funk et al., 2012). Population structures inferred using neutral markers reflect gene flow and genetic drift (Waples & Gaggiotti, 2006), which affect within and among population variations and can lead to adaptive divergence in the genome (Funk et al., 2012). To integrate adaptive markers into the definition of conservation units, Funk and colleagues proposed a framework of comparing population structures inferred from putatively neutral and adaptive loci. The inclusion of information for loci putatively under selection can help to understand mechanisms of local adaptation and is useful for conservation and management of the species (Moore et al., 2014).

Here, we analyzed the published data sets of microsatellites and the SNPs genotyped for chum salmon GSI. First, we inferred the chum salmon population structure and its demographic history using the microsatellite data in a distribution range. Then, we matched the sampling locations of the SNP genotyping studies with those of the microsatellite data. By regressing the SNP population structure on the microsatellite population structure, we estimated the selection on the SNPs as deviations from the predicted structure.

2 | MATERIAL AND METHODS

2.1 | Screening of population genetics data for chum salmon

We screened population genetics studies of chum salmon in the literature published after 1990 using the Google Scholar search system with keyword searches of "mixed-stock fisheries," "population structure," "salmon," and "stock identification." We also added studies that were known to us. The population genetics of chum salmon in the North Pacific has been studied extensively over the last four decades. International sampling of chum salmon has been conducted cooperatively across its distribution range to establish baseline genotype data for effective GSI (Beacham, Candy, Le, et al., 2009; Beacham, Candy, Wallace, et al., 2009; Beacham, Sato, et al., 2008; Beacham, Varnavskaya, et al., 2008; Seeb & Crane, 1999; Seeb et al., 1995; Urawa et al., 2005, 2009; Wilmot et al., 1994). The population structure of chum salmon has also been studied extensively using various markers, such as isozymes (Kijima & Fujio, 1979; Okazaki, 1982; Sato & Urawa, 2015; Seeb & Crane, 1999; Seeb VILEY_Ecology and Evolution

et al., 1995; Wilmot et al., 1994; Winans et al., 1994), mitochondrial DNA (mtDNA) (Garvin et al., 2010; Park et al., 1993; Sato et al., 2004; Yoon et al., 2008), minisatellites (Taylor et al., 1994), and microsatellites (Beacham, Sato, et al., 2008; Beacham, Varnavskaya, et al., 2008; Beacham, Candy, Le, et al., 2009; Beacham, Candy, Wallace, et al., 2009; Olsen et al., 2008. More recently, SNPs were developed to improve accuracy of mixed-stock identification and have also been used for population genetics (Garvin et al., 2013; Petrou et al., 2014; Saito et al., 2020; Sato et al., 2014; Seeb et al., 2011; Small et al., 2015). A high-throughput panel was developed for chum salmon, and patterns of linkage disequilibrium have been examined (McKinney et al., 2020).

Through the data screening, we found several publicly available data sets that covered the distribution range of chum salmon, including Japan. These data sets comprised microsatellite allele frequencies (Beacham, Candy, Le, et al., 2009), nuclear SNP genotypes with a combined mtDNA locus (Seeb et al., 2011), isozyme allele frequencies (Seeb et al., 1995; Winans et al., 1994), and mtDNA control regions (Sato et al., 2004). In this study, we analyzed the microsatellite (Beacham, Candy, Le, et al., 2009) and SNP (Seeb et al., 2011) data sets. All Japanese samples were from hatcheries and weirs in hatchery-enhanced rivers and were therefore hatchery-reared fish and/or hatchery descendants.

2.2 | Microsatellite and SNP data sets analyzed in this study

Microsatellite allele frequencies of chum salmon were retrieved from the Molecular Genetics Lab Online Data, Fisheries, and Oceans Canada website (https://www.pac.dfo-mpo.gc.ca/science/facilities -installations/pbs-sbp/mgl-lgm/data-donnees/index-eng.html, accessed 01 August 2020). The data consisted of the allele frequencies at 14 loci from chum salmon populations at 381 localities in a distribution range (n = 51,355) (Table S1; Figure S1) (Beacham, Candy, Le, et al., 2009). These data were used to infer the population structure and demographic history of chum salmon.

SNP genotypes were retrieved from the Dryad data repository (Seeb et al., 2011). The data consisted of 58 SNPs collected from 114 locations in a distribution range (Table S2; Figure S2) (Seeb et al., 2011). We excluded four loci, *Oke_U401-220*, *Oke_GHII-2943*, *Oke_IL8r-272*, and *Oke_U507-87*, to avoid pseudo-replication following the original study. A total of 53 nuclear SNPs and a combined mtDNA locus (mtDNA3) were included in our analysis (n = 10,458). The mtDNA3 locus included three gene loci, *Oke-Cr30* and *Oke-Cr386* (mtDNA control region), and *Oke-ND3-69* (mtDNA NADH-3).

2.3 | Accounting for ascertainment bias in the data sets

Significant deviations from Hardy-Weinberg equilibrium at four loci (Oke3, Ots103, One111, and OtsG68) from a 14-microsatellite

locus panel were found in Japanese, Russian, and representative North American populations (Beacham, Sato, et al., 2008; Beacham, Varnavskaya, et al., 2008). These deviations were likely due to ascertainment bias (Seeb et al., 2011), and therefore, we excluded these four loci and used the remaining 10 loci in our analysis.

For the SNP data, the highest levels of allelic richness detected in parts of the Alaskan region may be affected by ascertainment bias (Seeb et al., 2011). Allele/site frequency spectra have been used to examine potential ascertainment bias in SNPs (Nielsen et al., 2004). We calculated allele frequency spectra in six geographical areas according to the major lineage used in the original study (Seeb et al., 2011), namely Western Alaska, Yukon/ Kuskokwim upper, Alaskan Peninsula, Southeast Alaska (SEA)/ British Columbia (BC)/Washington (WA), Russia, and Japan/ Korea.

2.4 | Inferring population structure and demographic history

We used microsatellite allele frequencies (Beacham, Candy, Le, et al., 2009) to infer the population structure of chum salmon. In previous coalescent simulations, we found that the bias-corrected G_{ST} moment estimator (Nei & Chesser, 1983) performed better than other F_{ST} estimators when estimating pairwise F_{ST} values (Kitada et al., 2017). Therefore, we used G_{ST} (NC83) to estimate pairwise F_{ST} between populations. Because only allele frequency data could be accessed, we used the "read.frequency" function and computed pairwise F_{ST} values averaged over loci using the "GstNC" function in the R package FinePop v1.5.1 on CRAN.

A multidimensional scaling (MDS) analysis was performed on the pairwise F_{ST} distance matrix using the "cmdscale" function in R. We used the cumulative contribution ratio up to the *k*th axis (j = 1, ..., k, ..., K) as the explained variation measure, which was computed in the R function as $C_k = \sum_{j=1}^k \lambda_j / \sum_{j=1}^K \lambda_j$, where λ_j is the eigenvalue and $\lambda_j = 0$ if $\lambda_j < 0$. Following the original study (Beacham, Candy, Le, et al., 2009), populations were classified into eight geographical areas: Japan, Korea, Russia, Alaskan Peninsula, Western Alaska, Yukon/Tanana/Upper Alaska, SEA/BC, and WA (Table S1). A neighbor-joining (NJ) tree (Saitou & Nei, 1987) was constructed using the pairwise F_{ST} distance matrix using the "nj" function in the R package "ape."

To infer demographic history with admixture, we applied TreeMix (Pickrell &Pritchard, 2012) to the microsatellite allele frequencies (Beacham, Candy, Le, et al., 2009). The analyses were performed using six regional genetic groups, where Japan/Korea and SEA/BC/WA were combined to reduce the number of parameters to be estimated under the limited number of loci. On the basis of the allele frequencies and the lengths (numbers of repeats), we computed the mean and variance in length at each microsatellite locus and used them to run TreeMix v1.13. We tested up to five migration events and found the best tree using the Akaike Information Criterion (AIC; Akaike, 1973); AIC = $-2\text{Log}\hat{L} + 2 \times s$, where \hat{L} is the maximum composite likelihood and *s* is the number of parameters to be estimated, which, in this study, was three per migration event, namely, migration edge, branch length, and migration weight. For example, for the model with two migration events, the number of parameters was six. TreeMix maximizes composite likelihood and the 2 × log likelihood ratio is expected to follow an approximate χ^2 distribution with three degrees of freedom. Because TreeMix calculates composite likelihoods, the penalty in AIC needs to account for the correlation in the likelihoods. To avoid unconscious bias toward selection of parameter-rich models, we kept all the candidate models except those where the AIC values were decisively different.

2.5 | Gene flow and genetic diversity

We recorded approximate longitudes and latitudes of the sampling sites according to the names of rivers and/or areas and maps from the original studies using Google Maps. Sampling locations were plotted on the map using the "sf" package in R. Sampling points with pairwise F_{ST} values lower than a given threshold were connected by yellow lines to visualize the population connectivity. Under the assumption of Wright's island model, $F_{ST} \approx \frac{1}{4N_em+1}$, where N_e is the effective population size and *m* is the average migration rate between populations (Slatkin, 1987), a threshold of $F_{ST} = 0.01$ corresponds to $4N_em \approx 99$ migrants per generation (see, Waples & Gaggiotti, 2006; Whitlock & Mccauley, 1999).

We computed expected heterozygosity (H_e) values for the data sets. For microsatellites, we computed H_e for each population at each locus based on the allele frequencies. We then averaged H_e values over loci. For the 53 nuclear SNPs, we converted the original genotype data to genepop format (Rousset, 2008) and read the data using the "read.GENEPOP" function in the R package FinePop2 v0.4 on CRAN, where H_e values were computed automatically. The mtDNA3 locus had five alleles, and H_e was computed using the allele frequencies. We also computed NC83 pairwise F_{ST} values using 53 nuclear SNPs using the "pop_pairwiseFST" function in FinePop2.

Sampling points were visualized by a color gradient of populationspecific H_e values. The color of population *i* was rendered as RGB $(1-H_{e0,i}, 0, H_{e0,i})$, where $H_{e0,i} = (H_{e,i}-\min H_e)/(\max H_e-\min H_e)$. This conversion represents the standardized magnitude of an H_e value at the sampling point, with a continuous color gradient ranging from blue to red for the lowest and highest H_e values, respectively. Population-specific H_e values of microsatellites and SNPs were grouped by seven geographical areas: Japan/Korea, Russia, Alaskan Peninsula, Western Alaska, Yukon/Tanana/Kuskokwim, SEA/BC, and WA.

Differences in average H_e values between all pairs of the groups were tested by one-way ANOVA, where *p*-values were adjusted for the multiple comparisons using the "TukeyHSD" function in R.

2.6 | Identifying highly differentiated SNPs beyond the neutral population structure

To compare the genetic differentiation of SNP allele frequencies with the population structure inferred using the microsatellite markers, we analyzed 10 microsatellite loci and 53 nuclear SNPs loci plus the combined mtDNA3 locus. The mtDNA3 locus had five alleles, and the major allele (second one) was used for the meta-analysis. The functions of all 54 analyzed gene loci were confirmed by referring to the GeneCards database system (https://www.genecards. org/) and published literature.

First, we matched the 114 sampling locations of the SNPs (Figure S2) to the nearest location of the 381 sampling points of the microsatellites (Figure S1). The same locations were not necessarily used for genotyping. When we did not find the same location names, we assigned the closest locations of the microsatellites to those of SNPs using the longitudes and latitudes of the sampling locations. Among the 114 sampling locations for SNPs, 56 locations also had samples for microsatellites to the other 58 locations for SNPs. Then, we identified loci with allele frequencies that were significantly differentiated from the microsatellite population structure in the pairwise $F_{\rm ST}$. Specifically, we performed regression analyses of allele frequencies (axes) of the MDS of the pairwise $F_{\rm ST}$ estimated using the microsatellite allele frequencies as:

$$p_j^{(i)} = \beta_0^{(i)} + \beta_1^{(i)} \mathsf{mds1}_j + \beta_2^{(i)} \mathsf{mds2}_j + \varepsilon_j^{(i)}, \varepsilon_j^{(i)} \sim N\left(0, \sigma^2\right)$$

where $p_j^{(i)}$ is the allele frequency of SNP *i* (=1,...,*l*) in population *j* (=1,...,*l*). The explanatory variables mds1_{*j*} and mds2_{*j*} are the 1st and 2nd MDS coordinates of pairwise F_{ST} values estimated from the microsatellite allele frequencies for the matched populations *j* (=1,...,*J*). These explanatory variables are population-specific (*J* = 114). The regression analysis was performed for each marker (*i* = 1, ..., 54), and 54 sets of regression coefficients ($\beta_0^{(i)}, \beta_1^{(i)}, \beta_2^{(i)}$) were obtained. The *p*-values for the coefficients were adjusted for the multiple comparison (*q*-values) by the method of Benjamini and Hochberg (1995) using the "p.adjust" function in R. Given the population structure predicted from the neutral microsatellites, the outlier loci from the set of 53 nuclear SNPs and mtDNA3 locus that deviated largely from the predicted population structure are likely under natural selection and/or anthropogenic selection.

To identify genes that best characterize geographical areas, we conducted a principal components analysis (PCA) for the allele frequencies of the 53 nuclear SNPs and mtDNA3 locus using the "prcomp" function in R. Finally, we visualized the geographical distributions of the identified highly differentiated genes using boxplots and geographical maps of the allele frequencies and H_e values. Sampling locations were mapped with a color gradient of allele frequencies for population *j* rendered as RGB (p_{j0} , 0, 1– p_{j0}), where $p_{j0} = (p_j - \min p)/(\max p - \min p)$. This conversion represents the standardized allele frequency values at the sampling point, with a continuous

color gradient ranging from blue to red for the lowest and highest allele frequencies, respectively.

3 RESULTS

(a)

3.1 | Population structure and demographic history of chum salmon

Pairwise F_{ST} values based on the microsatellite allele frequencies averaged 0.019 \pm 0.010 (SD, standard deviation). The MDS of pairwise F_{sT} values (Figure 1a) indicated a latitudinal cline in the American and Russian populations and a separation of Alaskan populations from the others along the first axis (mds1) explaining 29% of the variance. The second axis (mds2, 16% of variance) showed a differentiation of Japanese/Korean populations from the others, although they were closely related to southern Russian populations from Sakhalin, Amur, and Primorye. Plots of mds1 versus mds3, and mds1 versus mds4 produced similar results, showing most differentiation was driven by divergence from a latitudinal cline (Figure S3). The mds1 to mds4 explained 60% variation in the pairwise F_{sT} distant matrix (Figure S4). In the unrooted NJ tree generated from this data set (Figure S5a), five large regional population groupings were inferred: (a) Alaskan Peninsula, (b) Western Alaska/Yukon (Canada), (c) Russia/Japan/Korea, (d) SEA/Northern BC, and (e) Southern BC/WA. The TreeMix analysis indicated two periods of admixture (migration events), from Japan/Korea to Russia and the Alaskan Peninsula (Figure 1b, Figure S6). Japanese/Korean populations had the greatest convergence among all the groups, whereas admixture was smallest in the Alaskan Peninsula and SEA/BC/WA populations.

3.2 Ascertainment bias in the SNP data

The allele frequency spectra (Figure S7) showed a similar pattern in all areas, except in Western Alaska, where much smaller occurrence was found for the lowest allele frequencies in Western Alaska.

3.3 Gene flow and genetic diversity

Using the criterion of pairwise $F_{ST} < 0.01$ (4 $N_{p}m \approx 99$), substantial gene flow between American and Asian populations was identified for the microsatellites (Figure 2a). Japanese populations were found to be connected to Russian and Korean populations. For the SNPs, using the same criterion, we detected substantial gene flow within Alaska, but the SEA and WA populations appeared to be isolated with little gene flow between American and Russian populations (Figure 2b). Japanese and Korean populations were isolated from all the others.

For the microsatellites, H_{ρ} values were generally large in almost all the populations (Figure 2a, Figure S8a). The average H_{a} value over loci was 0.89 \pm 0.02 (range 0.81–0.93). Large H_e values were found in Japanese, Russian, and Alaskan populations, whereas H_{a} values were slightly smaller in SEA/BC, WA, and Canadian Yukon populations (Figure S5b). Japanese populations had the highest H_{a}



(b)

Alaskan.Pen

(MDS) plots (mds1 versus mds2) of the population structure of chum salmon based on pairwise F_{sT} values estimated from 10 microsatellite loci in 381 populations. (b) Admixture graph inferred using TreeMix with two migration events. The scale bar (drift parameter) is 10 times the average standard error of the sample covariance matrix of allele frequencies between populations (Pickrell & Pritchard, 2012)

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FIGURE 2 Visualization of genetic diversity and population connectivity of chum salmon based on microsatellite and SNP markers. (a) Map obtained using 10 microsatellite loci of 381 populations (n = 51,355) (Beacham, Candy, Le, et al., 2009). (b) Map obtained using 53 SNPs of 114 populations (n = 10,458) (Seeb et al., 2011). The color of each population reflects the magnitude of the expected heterozygosity (H_e) values, with a continuous color gradient ranging from blue to red for the lowest and highest H_e values, respectively. The yellow lines connect populations with pairwise $F_{ST} < 0.01$

values, but the mean was similar to those in Russian and Western Alaskan populations, smaller in Alaskan Peninsular, SEA/BC, and WA populations, and extremely small in the Yukon population (Figure S8a).

In contrast, the largest SNP H_e values were found in Western Alaska populations, followed by Yukon, Alaskan Peninsula, Russia, and SEA populations (Figure 2b, Figure S8b). The average H_e value over loci was 0.27 \pm 0.02 (range 0.20–0.30). Japan and WA populations had the smallest H_e values. The H_e values were similar in populations from Western Alaska, Yukon, Alaskan Peninsula, and Russia (Figure S8b), and, although Russian populations had larger H_e values than the SEA/BC and WA populations, the differences between these populations were not significant. Japanese populations had the smallest H_e values with no significant difference between Japanese and WA populations.

3.4 | Differentiated SNPs beyond the neutral population structure

We summarized 53 nuclear SNPs and the mtDNA3 locus with functions (if known), locus names, and raw data locus names used in our analysis (Table S3). Five of the outlier SNPs had previously been identified as potential candidates for selection (Seeb et al., 2011). We found 31 SNPs in functional genes, and 23 SNPs were unknown.

The results of our regression analysis are summarized in Table 1. Most of the SNPs had large coefficients for both mds1 and

mds2 with highly significant q-values, showing that most of SNPs were differentiated beyond the neutral population structure. The top five outlier loci (Table 1) with the highest regression coefficients and q-values were mtDNA3, U502241, GnRH373, ras1362, and TCP178, which characterized mds2 of the pairwise F_{sT} . The scatter plot of mds1 and mds2 (Figure 3a) identified the outlier SNPs that diverged beyond the scale of mds1. The PCA (Figure 3b) identified differences between Japanese/Korean and other populations as a primary component (PC1, 43% variance), whereas the second component (PC2, 20% variance) corresponded to the latitudinal cline among the Russian/American populations. Biplots of PCA1 versus PCA3 and PCA4 described outliers (Susitna River, and Sturgeon River on Kodiak Island), but most differentiation was caused by divergence of the Japanese/Korean populations (Figure S9). PCA1 to PCA4 explained 76% of variance of the SNP allele frequencies (Figure S10). The PCA (Figure 3b) confirmed that differentiation of Japanese/Korean populations was characterized mainly by the top five outlier SNPs detected by the regression analysis (Figure 3a).

The allele frequencies of the top five outlier loci were distinct in Japanese/Korean populations in the distribution range (Figure 4). The major allele of mtDNA3 was fixed at close to 1.0 in American and Russian populations, whereas the allele frequencies were significantly lower (0.21 ± 0.11) in Japanese/ Korean populations. The U502241 allele frequencies were much larger in Japanese/Korean and Washington populations than they were in the other populations. GnRH373, ras1362, and TCP178 had similar geographical distributions of allele frequencies, and -WILEY-Ecology and Evolution

TABLE 1 Regression analysis of SNP allele frequencies on abs(mds1) and abs(mds2) of the microsatellite pairwise F_{ST} values

		abs(mds1)		abs(mds2)	
Locus raw name	Locus name	Estimate ^a	-(log ₁₀ q) ^b	Estimate ^a	-(log ₁₀ q) ^b
mtDNA3	mtDNA3 ^c	9.83	6.13	41.92	14.92
U502241 ^d	Oke_U502-241	7.04	2.68	34.78	18.61
GnRH373	Oke_GnRH-373	0.43	0.12	26.25	25.07
ras1362	Oke_ras1-249	3.84	2.26	18.84	15.88
TCP178	Oke_TCP1-78	3.50	3.52	18.44	25.07
U22	Oke_u200-385	3.62	1.56	17.69	11.43
Tf278 ^d	Oke_Tf-278	13.27	13.06	17.32	11.66
U506110 ^d	Oke_U506-110	13.07	14.42	16.06	11.87
IL8r406	Oke_IL8r-406	3.90	4.76	15.62	22.51
DM20548	Oke_DM20-548	0.27	0.09	15.34	13.59
U217	Oke_u217-172	9.31	13.02	14.63	15.02
GHII3129	Oke_GHII-3129	13.05	22.84	14.36	16.23
U505112	Oke_U505-112	3.94	3.85	12.93	14.46
u1519	Oke_u1-519	9.95	10.38	12.44	8.49
GPDH	Oke_GPDH-191	7.97	11.28	11.93	12.21
AHR178	Oke_AhR1-78	0.75	0.31	11.89	12.10
U212	Oke_u212-87	0.88	0.89	11.31	25.07
CKS389	Oke_CKS-389	2.79	1.19	10.00	5.15
IL1RA	Oke_il-1racp-67	2.11	1.38	9.67	9.21
RFC2 ^d	Oke_RFC2-618	15.49	23.68	8.75	6.26
U504228	Oke_U504-228	9.09	8.52	8.26	4.08
COPA ^d	Oke_copa-211	0.45	0.11	8.05	3.35
ZAN132	Oke_U305-130	2.50	1.35	8.01	4.95
U507286	Oke_U507-286	4.61	2.66	7.98	3.65
serpin	Oke_serpin-140	8.49	7.99	7.33	3.54
CTS1627	Oke_PP2A-635	6.62	4.07	7.16	2.62
TSHA1	Oke_Tsha1-196	2.53	1.30	6.89	3.63
U514150	Oke_U514-150	1.47	1.02	6.63	6.42
RH1OP	Oke_RH1op-245	7.08	15.95	6.21	7.85
U216	Oke_u216-222	1.91	3.23	6.01	11.65
CCT3143	Oke_U401-143	2.97	5.92	5.95	10.35
HP182	Oke_HP-182	3.77	2.92	5.94	3.47
GPH105	Oke_GPH-105	2.03	0.70	5.31	1.69
U511271	Oke_U511-271	3.10	3.83	5.05	4.78
U302195	Oke_U302-195	5.48	3.96	4.91	1.85
ctgf105	Oke_ctgf-105	3.47	6.46	4.17	4.93
hnRNPL239	Oke_hnRNPL-239	2.07	2.44	3.93	3.90
MOESIN	Oke_Moesin-160	3.55	11.04	3.72	6.77
U509219	Oke_U509-219	4.66	2.88	3.32	0.99
MARCKS362	Oke_MARCKS-362	13.58	11.23	3.09	0.66
GnRH527	Oke_GnRH-527	4.47	4.19	3.03	1.30
ARF	Oke_arf-319	4.30	5.07	2.81	1.42
GPH78	Oke_GPH-78	0.76	0.31	2.53	1.06
U510204	Oke_U510-204	2.30	2.19	2.43	1.35

TABLE 1 (Continued)

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		abs(mds1)		abs(mds2)	
Locus raw name	Locus name	Estimate ^a	-(log ₁₀ q) ^b	Estimate ^a	$-(\log_{10} q)^{\rm b}$
FARSLA242	Oke_FARSLA-242	5.23	9.87	1.86	1.07
EIF4EB	Oke_eif4ebp2-64	5.64	13.79	1.85	1.35
MAPK1135	Oke_MAPK1-135	3.70	5.41	1.78	0.99
ZAN311	Oke_U305-307	4.70	9.02	1.52	0.88
IGFI1	Oke_hsc71-199	1.42	3.25	1.22	1.42
U503272	Oke_U503-272	2.76	13.06	1.16	1.84
U202	Oke_u202-131	5.24	12.89	1.07	0.65
KPNA287	Oke_KPNA2-87	2.36	0.98	0.75	0.15
AhR1278	Oke_AhR1-278	0.26	0.26	0.61	0.53
HSP90BA299	Oke_HSP90BA-299	0.70	2.25	0.39	0.57

Note: The markers are ordered by the regression coefficients for mds2, which characterized the differentiation of Japanese/Korean populations. ^aAbsolute values of the regression coefficients.

^bp-values adjusted by the Benjamini & Hochberg's method.

^cOke-Cr30, Oke-Cr386, Oke-ND3-69 were combined (Seeb et al., 2011).

^dFive outlier loci found in the original study (Seeb et al., 2011).

in Japanese/Korean populations, they were differentiated from those in the other populations. The average genetic diversity (H_e) was significantly larger in Japanese/Korean populations than it was in Western Alaska for mtDNA3 (Welch two sample *t* test, t = 30.5, df = 17.5, $p = 2.2 \times 10^{-16}$) and TCP178 loci (t = 36.8, df = 51.3, $p = 2.2 \times 10^{-16}$), whereas it was similar in Western Alaska for U502241 (t = 0.9, df = 19.7, p = 0.37), and ras1362 (t = 0.6, df = 35.7, p = 0.57). GnRH373 had smaller H_e values in Japanese/Korean populations (0.33) than in Western Alaskan populations (0.40) (t = -3.0, df = 21.2, p = 0.006).

4 | DISCUSSION

The inferred population structure based on microsatellites was consistent with the one obtained by the TreeMix analysis (Figure 1). The mds1 of pairwise $F_{\rm ST}$ identified the north-south cline in American and Russian populations, and mds2 clearly showed divergence of Japanese/Korean populations, in agreement with the original study (Beacham, Candy, Le, et al., 2009). TreeMix indicated that modern chum salmon populations may have been recolonized from source populations between the Alaskan Peninsula and SEA/BC/WA, which expanded to the north and to Asia. However, a higher level of divergence of the Japanese/Korean populations was clear relative to the Russian and American populations, suggesting that the Japanese/Korean populations were the most diverged from the ancestral population and that they diverged between Russia and the Alaskan Peninsula. Two inferred admixture events may reflect the straying of Japanese/Korean chum salmon into Russian and Alaskan rivers during long migration in the North Pacific.

The SNP allele frequency spectra (Figure S7) showed a similar pattern in all areas except Western Alaska. Because the entire range was represented for the SNP selection from Japan to WA (Seeb et al., 2011), ascertainment bias caused by the elimination of SNP with minor alleles may have occurred in Western Alaska, as indicated in Figure S7. The use of SNPs with intermediate allele frequencies produced larger genetic diversities in the Western Alaska samples relative to other samples (Figure S8b), although the effects of the ascertainment bias were expected to be minimal within Alaska (Seeb et al., 2011).

Our regression analysis identified the top five outlier gene loci (Table 1, Figures 3 and 4) that were significantly differentiated beyond the neutral population structure in Japanese/Korean populations, namely, mtDNA3, U502241, GnRH373, ras1362, and TCP178. The scatter plot of mds1 and mds2 indicated that most of the SNPs were differentiated beyond the neutral population structure. Selection can be induced by natural and anthropogenic pressures. The mds1 described the north-south cline of population expansion (Figure 1). The maximum regression coefficient of mds1 was 15.5 for RFC2 (Table 1). In contrast, the top five outlier gene loci had mds2 coefficients >18.4, which suggested that the allele frequencies of the top five gene loci were the result of anthropogenic selection pressures in Japanese/Korean populations.

mtDNA3 (a combined locus of the control region and NADH-3, Smith, Baker, et al., 2005) was the most differentiated gene locus (Figures 3 and 4). mtDNA encodes some of the proteins in the oxidative phosphorylation enzymatic complex and plays a key role in aerobic ATP production, by contributing to the ability to respond to endurance exercise (reviewed by Stefàno et al., 2019). The mtDNA control region functions in dNTP metabolism (Nicholls & Minczuk, 2014) and oxygen consumption (Kong et al., 2020). Subunits



FIGURE 3 SNPs that characterized the distinctiveness of Japanese/Korean chum salmon populations. (a) Absolute values of the regression coefficients of mds1 versus those of mds2 on the same scale. Red indicates the outlier SNPs that highly diverged beyond the neutral population structure. (b) Principal component analysis (PCA) biplot, based on 53 nuclear SNPs and the combined mtDNA3 locus

of the NADH dehydrogenase complex encoded by mtDNA are involved in the pathway from NADH to oxygen (Rivera et al., 1998). A significant reduction in mtDNA3 allele frequencies might reduce the efficiency of aerobic exercise ability and endurance, energy metabolism, and oxygen consumption. These results are consistent with the lower swimming endurance of Japanese hatchery-born chum salmon fry measured in a stamina measurement experimental device in the Japanese oldest Chitose Hatchery, Hokkaido (Kobayashi & Ohkuma, 1983). Wild chum salmon fry had a 1.4-fold higher swimming ability (56.6 \pm 11.1 cm/s) than hatchery-reared fry (41.4 \pm 12.3) (t = 2.45, df = 8.7, p = .038) in the early 1980s.

U502241 allele frequencies were significantly larger in Japanese/ Korean and Washington populations (Figure 4). Elevated levels of U502241 allele frequencies in Japanese/Korean and Washington populations were similar (t = 1.3, df = 1.1, p = .39) and may indicate a common effect. The original study (Seeb et al., 2011) found this phenomenon and detected U502241 as an outlier SNP. No function has been associated with the chum U502241 locus so far (Elfstrom et al., 2007), but the Atlantic salmon immunoglobulin IgH locus B genomic sequence was the most likely match for U502241 (Seeb et al., 2011). Immunoglobulins initiate immune responses (Gene Cards).

GnRH (gonadotropin-releasing hormone) is a key regulator of vertebrate reproduction, including that of salmonids (Khakoo et al., 1994). TCP1 (T-Complex 1) is related to sperm-zona-pellucida interaction and spermatozoan fertilization ability (Dun et al., 2011). The *ras* gene has a central role in cell growth (Rotchell et al., 2001). These outliers related to reproduction and growth are key parameters in hatchery production. In chum salmon, GnRH is involved in gonadal maturation during the early and final phases of upstream migration (Kudo et al., 1996). GnRH also improved stream odor discrimination in adult fish (Ueda, 2019), and expression of GnRH increased in the brain during homing migration (Ueda et al., 2016). The differentiation of GnRH373 (Figure 4) might influence homing migration timing, potentially leading to the current introgression from Japanese into Russian populations, as suggested by our TreeMix analysis.



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FIGURE 4 Geographical distribution of the top five outlier SNPs that characterized the distinctiveness of Japanese/Korean chum salmon populations. The points on the maps show allele frequencies (center) and expected heterozygosity (H_e) (right) at sampling locations, with a continuous color gradient ranging from blue to red for the lowest and highest allele frequencies and H_e values, respectively

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Apparently different distributions in H_e values were observed in microsatellites and SNPs in the Japanese/Korean populations (Figure S8). In the Japanese populations, the largest microsatellite H_e values may have been influenced by the transplantation history in Japanese river populations (Beacham, Sato, et al., 2008), whereas the smallest SNP H_e values may be a consequence of artificial selection in the hatcheries. The H_e values in Japanese/Korean and Western Alaskan populations were similar for U502241 and ras1362, whereas the H_e values for GnRH373 were smaller in the Japanese/Korean populations than in the Western Alaskan populations as shown in Figure 4. In contrast, the H_e values in Japanese/ Korean populations were significantly larger than those in Western Alaskan populations for mtDNA3 and TCP178, suggesting increases in heterozygosity. These results suggested genotype-specific effects on fitness.

The analyses in this study relied on limited data sets. Clearly, further genomic studies are needed to obtain more precise views of the demographic history, and environmental and anthropogenic selection of chum salmon. Population genomics studies can also provide finer resolution of GSI and will be useful for management and conservation of chum salmon.

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CONFLICT OF INTEREST

None declared.

AUTHOR CONTRIBUTION

Shuichi Kitada: Conceptualization (equal); Data curation (lead); Formal analysis (equal); Funding acquisition (equal); Investigation (lead); Methodology (equal); Project administration (lead); Resources (equal); Software (equal); Supervision (lead); Validation (equal); Visualization (equal); Writing-original draft (equal); Writingreview & editing (equal). Hirohisa Kishino: Conceptualization (equal); Data curation (supporting); Formal analysis (equal); Funding acquisition (equal); Investigation (equal); Methodology (equal); Project administration (supporting); Resources (equal); Software (equal); Supervision (supporting); Validation (equal); Visualization (equal); Writing-original draft (equal); Writing-review & editing (equal).

DATA AVAILABILITY STATEMENT

The authors affirm that all data necessary for confirming the conclusions of this article are present within the article, figures, and supporting information.

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SUPPORTING INFORMATION

Additional supporting information may be found online in the Supporting Information section.

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