



Prognostic Role of the Platelet-Lymphocyte Ratio in Acute Ischemic Stroke Patients Undergoing Reperfusion Therapy: A Meta-Analysis

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ABSTRACT

BACKGROUND: Both inflammation and thrombotic/hemostatic mechanisms may play a role in acute ischemic stroke (AIS) pathogenesis, and a biomarker, such as the platelet-to-lymphocyte ratio (PLR), considering both mechanisms may be of clinical utility.

OBJECTIVES: This meta-analysis sought to examine the effect of PLR on functional outcomes, early neurological changes, bleeding complications, mortality, and adverse outcomes in AIS patients treated with reperfusion therapy (RT).

DESIGN: Systematic Review and Meta-Analysis

DATA SOURCES AND METHODS: Individual studies were retrieved from the PubMed/Medline, EMBASE and Cochrane databases. References thereof were also consulted. Data were extracted using a standardised data sheet, and systematic reviews and meta-analyses on the association of admission (pre-RT) or delayed (post-RT) PLR with defined clinical and safety outcomes were conducted. In the case of multiple delayed PLR timepoints, the timepoint closest to 24 hours was selected.

RESULTS: Eighteen studies (n=4878) were identified for the systematic review, of which 14 (n=4413) were included in the meta-analyses. PLR collected at admission was significantly negatively associated with 90-day good functional outcomes (SMD=-.32; 95% CI = -.58 to -.05; P=.020; z=-2.328), as was PLR collected at delayed timepoints (SMD=-.43; 95% CI = -.54 to -.32; P<.0001; z=-7.454). PLR at delayed timepoints was also significantly negatively associated with ENI (SMD=-.18; 95% CI = -.29 to -.08; P=.001). Conversely, the study suggested that a higher PLR at delayed timepoints may be associated with radiological bleeding and mortality. The results varied based on the type of RT administered.

CONCLUSIONS: A higher PLR is associated with worse outcomes after stroke in terms of morbidity, mortality, and safety outcomes after stroke.

KEYWORDS: stroke, endovascular therapy, meta-analysis, platelet-lymphocyte ratio, reperfusion therapy

TYPE: Meta-analysis

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SUPPLEMENTAL MATERIAL: Supplemental material for this article is available online.

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Background

Acute ischemic stroke (AIS) forms the vast majority of stroke and has a large burden of disease and death, with sizeable case-fatality and disability rates.^{1,2} As such, the identification of a prognostic biomarker that could inform treatment decision making in AIS has attracted great research interest.³ Blood-based biomarkers are of particular clinical interest because they are easier and cheaper to obtain than imaging-based biomarkers.⁴ The platelet-lymphocyte ratio (PLR) is such a biomarker that has shown utility in emergency medicine and trauma settings,⁵ acute illnesses such as acute coronary

syndrome,⁶ and cardiovascular reperfusion.^{7,8} It is particularly promising as it could potentially provide insight into both inflammation and thrombotic/hemostatic mechanisms thought to play a role in AIS pathogenesis, whereas other biomarkers shown to have prognostic value such as C-Reactive Protein, platelet count and the neutrophil-lymphocyte ratio encompass only one of these mechanisms.^{4,9-11} While there has been some evidence showing benefit of PLR in predicting clinical outcomes,¹²⁻¹⁹ mortality^{12,16,17,19-21} and bleeding risk^{12,17,18,22,23} in AIS patients treated with reperfusion therapy (RT), this is yet to be clearly elucidated; to the best of our knowledge there is no



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systematic review or meta-analysis currently in the literature evaluating this. As such, this study aims to investigate the association of PLR at admission and PLR collected at delayed timepoints through a systematic review and meta-analysis of published literature, and to gauge the clinical utility thereof in prognostication and clinical decision-making.

Our underlying research questions are as follows; in AIS patients receiving RT:

- Is lower baseline or delayed PLR associated with (1) good functional outcomes; (2) modified Rankin scale (mRS) 0-1; (3) favourable recanalisation outcomes; (4) early neurological improvement (ENI), and (5) dramatic ENI?
- Is higher baseline or delayed PLR associated with: (1) mortality; (2) intracerebral haemorrhage (ICH) (3) symptomatic ICH (sICH); (4) early neurological deterioration (END) and (5) stroke associated infection (SAI)?

Methods

The study was performed as per the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA)²⁴ flowchart (Figure 1) and the Meta-analysis of Observational Studies in Epidemiology (MOOSE) Checklist (Supplemental Table 3).²⁵ Ethics approval was not required as this study was a systematic review and meta-analysis of previously published studies.

Literature Search: Identification and Selection of Studies

The following databases were searched: Embase, PubMed/Medline, and Cochrane Library until 30 October 2021. Keywords used included a combination of terms including “acute stroke”, “cerebrovascular accident”, “brain ischemia”, “reperfusion”, “endovascular therapy”, “endovascular thrombectomy”, “thrombolysis”, “PLR” and “platelet-lymphocyte ratio.” Full search strategies and a complete list of keywords are provided in the **Supplementary Information (search strategy)**. In addition, references of related articles were also examined to retrieve studies relevant to our analysis.

Inclusion and Exclusion Criteria

Following Inclusion criteria were applied: (1) patients aged 18 years or above; (2) patients diagnosed with AIS; (3) patients who received RT; and (4) studies with good methodological design (including sufficient sample size, defined as >20 patients). The exclusion criteria were: (1) animal/preclinical studies; (2) duplicated publications; (3) studies with smaller sample sizes or shorter study periods, where multiple studies from overlapping centres with varying study periods reporting similar outcomes were present; (4) full-text articles not available; (5) systematic reviews or meta-analyses, conference abstracts, letters and case reports or series; and (6) studies presented as abstracts, with relevant PLR or outcome data not reported.

Data Extraction

Titles and abstracts were first reviewed on Endnote to exclude articles mismatched to eligibility criteria. The remaining articles underwent thorough full text examination to determine if they were to be included in the systematic review or meta-analysis as per the eligibility criteria. Reviews, former systematic reviews and meta-analyses and opinion articles were kept separately for discussion in the manuscript. Two authors conducted the screening independently, and any disagreements were discussed until a consensus was made. Data from each study/trial were extracted independently using a standardised data extraction sheet to obtain the following information: (1) baseline demographics: author, country, and year of publication; (2) study population: age of patients, sample size, characteristics of acute stroke patients, and RT type (EVT/IVT); (3) PLR; (4) time of blood collection (preintervention vs postintervention; for delayed timepoints, the timepoint closest to 24 hours was included); (5) outcome measures: primary and secondary outcomes, including clinical outcomes, angiographic outcomes, and mortality; and (6) adverse effects/safety outcomes. The primary outcome was in terms of morbidity: 90-day good functional outcomes, defined as mRS 0-2 across all studies. One study looked at mRS 0-1 and was considered separately. Mortality was defined at 90 days in all studies. Regarding angiographic outcomes, successful recanalisation was defined as mTICI_{2b} across all included studies, and where applicable, the first pass effect (FPE) as complete recanalisation (mTICI 3) achieved with a single pass. In all studies, ENI was defined in terms of improvements in NIHSS score, with this generally being 4 points in 24 hours, unless otherwise indicated (Table 2). Dramatic ENI was defined as improvement in NIHSS score by 8 points across all studies reporting this outcome. END was conversely defined as NIHSS score worsening across all studies, with any specifications on this indicated (Table 2). sICH was determined by neurological decline along with imaging confirmation across all studies, with criteria such as European Cooperative Acute Stroke Study-I (ECASS-I) and Safe Implementation of Thrombolysis in Stroke-Monitoring Study (SITS-MOST) used. The radiological bleed outcome was defined as any radiological evidence of bleeding, with CT being the most common imaging modality used to ascertain this, and some studies using additional MRI.

Quality Assessment of Included Studies

The methodological quality of each study was assessed using the modified Jadad scale by two researchers independently.²⁶ The scale evaluates study quality based on the following evaluation criteria: randomisation, blinding, withdrawals, dropouts, inclusion/exclusion criteria, adverse effects, and statistical analysis. A double-blind study received 1 score; a single blind study received .5 scores. The total score for each study ranged from 0 to 8 points, and studies were divided into low-quality (0-3 points) and high-quality (4-8 points) levels.

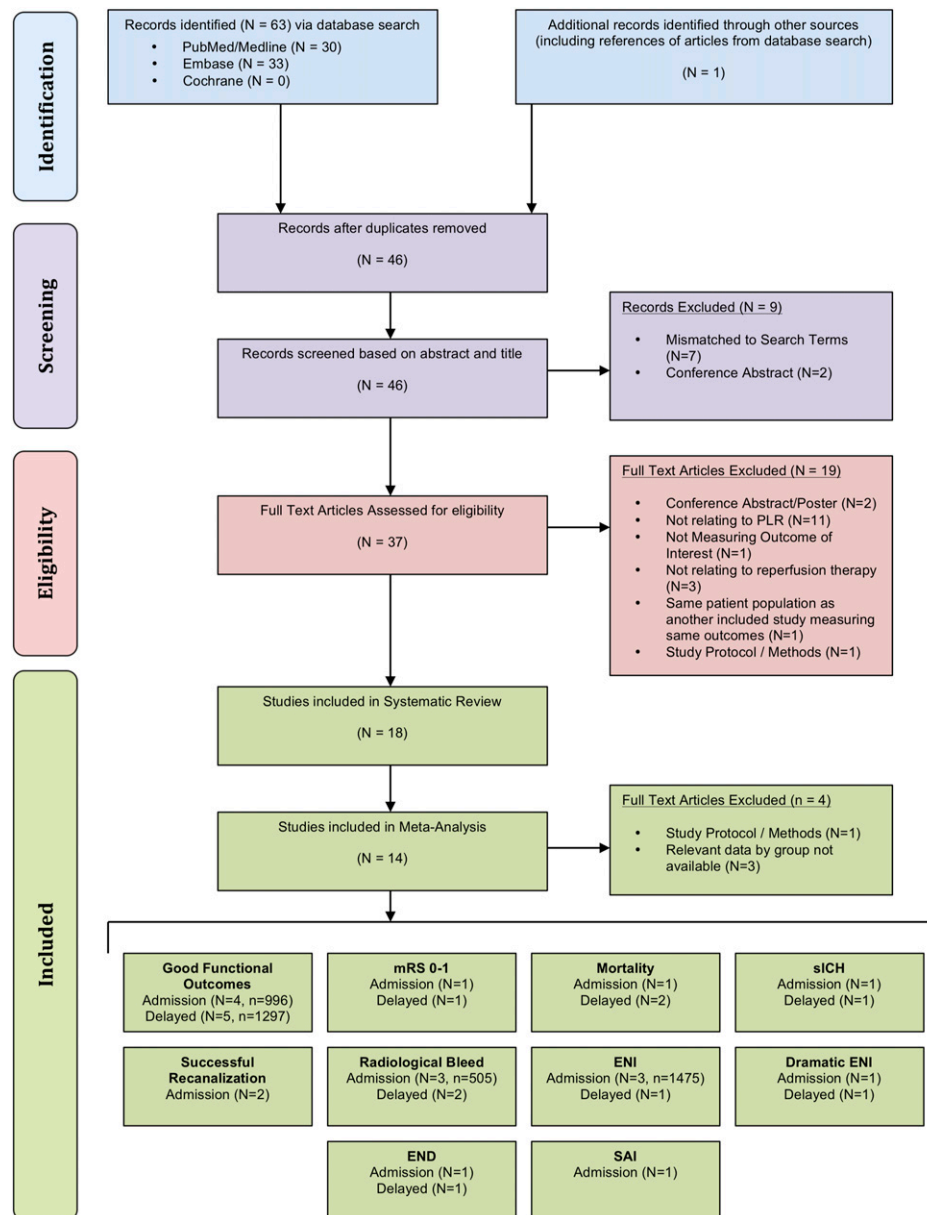


Figure 1. PRISMA Diagram. Note: The PRISMA flowchart shows the main characteristics of the included studies. Outcomes for which a meta-analysis could successfully be carried out also have the number of patients shown. Abbreviations: mRS=Modified Rankin Scale; SAI=Stroke Associated Infection; SAP=Stroke Associated Pneumonia; sICH=symptomatic intracerebral hemorrhage; PH=Parenchymal Hematoma; PLR=Platelet-Lymphocyte Ratio; N=Number of Included Studies; n=number of patients.

The risk of funding bias in the included studies was evaluated independently from the quality assessment by using the scoring test developed by Saunders et al²⁷ (2017), which analyses the declaration of funding sources and conflicts of interest. A score of 1-2 was considered to indicate moderate potential for bias. The absence of industry funding was not considered to signify an absence of bias, but the presence of industry funding or conflicts of interest was assumed to indicate bias.

Statistical Analysis

All statistical analyses were performed using STATA (Version 13.0, StataCorp LLC, College Station, Texas, USA). Forest plots were generated to present the standard mean difference

(SMD), P values, 95% confidence intervals (CI), percentage weight and heterogeneity between studies included in the meta-analysis. Meta-analyses were split by admission PLR and delayed PLR (pre- and postintervention, respectively). In cases where there were multiple delayed PLR timepoints, the timepoint closest to 24 hours was taken (Table 2). I^2 statistics and P values were used to assess heterogeneity between studies, with <40%, 30-60%, 50-90% and 75-100% representing low, moderate, substantial, and considerable heterogeneity, respectively.²⁸ Random-effects modelling was used across all subgroup analyses. Subgroup analyses stratified by treatment group^{xxxx} were performed for all outcomes, with any adjunct treatment indicated by '±'. Baseline characteristics of patient populations were synthesized from all included studies (Table 1). Where

Table 1. Summary of combined clinical characteristics, risk factors and stroke etiologies across all included studies.

FACTOR	NUMBER OF PATIENTS FOR WHOM DATA WAS AVAILABLE	NUMBER OF PATIENTS WITH FACTOR	% OR MEAN (\pm SD)
Age (yrs.)	4788	N/A	69.67 \pm 13.25
Male gender	4713	2784	59.07
Baseline NIHSS	4348	N/A	11.51 \pm 7.67
Baseline PLR	4553 <i>Excluding Inanc & Inanc:* 4497</i>		91.23 \pm 168.51 <i>Excluding Inanc & Inanc:* 91.99 \pm 169.29</i>
Delayed PLR	1111		166.66 \pm 106.70
BSBP	2936		149.60 \pm 24.62
Etiology			
LAA	3533	1134	32.10
CE	3533	1260	35.66
SVO	2721	440	16.17
Other and/or undetermined	3533 (as reported) 2384 (excluding studies not reporting SVO to avoid overlap)	571 (as reported) 384 (excluding studies not reporting SVO to avoid overlap)	16.16 (as reported) 16.11 (excluding studies not reporting SVO to avoid overlap)
Risk factors			
CAD	3476	758	21.81
AF	4306	1270	29.49
HTN	4361	2869	65.79
DM	4713	1173	24.89
HL/DL	3757	1255	33.40
Smoking	3111	916	29.44
PS/TIA	3877	757	19.53

*Inanc & Inanc was excluded, and the analysis was repeated to ensure that this did not skew the data because it was an outlier compared to other studies. Abbreviations: LAA=Large Artery Atherosclerosis; CE=Cardioembolic; SVO=Small Vessel Occlusion; CAD=coronary artery disease; AF=Atrial Fibrillation; HTN=Hypertension; DM=Diabetes Mellitus; HL=Hyperlipidemia; DL=Dislipidemia; PS=Previous Stroke; TIA=Transient Ischemic Event; BSBP=Baseline Systolic Blood Pressure; PLR=Platelet-Lymphocyte Ratio.

applicable, median, and interquartile ranges were converted to mean and standard deviation using Wan et al.'s (2012) method, and median and ranges were converted to mean and standard deviation using the methods described by Luo et al (2018)²⁹ and Wan et al (2014),³⁰ respectively. For studies where SD was not available, we used the method proposed by Walter and Yao³¹ (2007) to determine the SD, assuming data were normally distributed. Combined means were calculated where applicable. A (Begg's) funnel plot was used to visually detect the presence of publication bias in the meta-analysis; asymmetry was indicative of publication bias. This was confirmed using Egger's test of effect sizes for publication bias. The command "metaling" was used in STATA to determine the impact of individual studies on the overall meta-analysis (Supplemental Figure 1). P values <.05 were considered statistically significant.

Results

Description of Included Studies

The total number of patients considered in the meta-analysis was 4413. An additional four studies, with 465 patients, were

included in the systematic review. The mean age was 69.67 (SD 13.25) years. The clinical characteristics of all studies included in the meta-analysis are shown in Table 1, details about outcomes in all studies in Table 2, and details about PLR stratified by outcome in Table 3. The results of the methodological quality and funding bias assessment are provided in Supplemental Table 1.

Association of PLR With 90-day Good Functional Outcomes

Four studies (n=996) reported admission PLR levels for this outcome, and 5 studies (n=1297) reported PLR collected at delayed timepoints. All studies defined this outcome as mRS 0-2.

Admission PLR. The meta-analysis showed a significant negative association of admission PLR with 90-day good functional outcomes (SMD=-.32; 95% CI = -.58 to -.05; P=.020; z=-2.328; Figure 2). However, considering treatment provided, a nonsignificant effect was observed in both subgroups,

including IVT-treated patients (SMD=-.19; 95% CI = -.38 to .01; P=.063; z=-1.860) and patients who received EVT±IVT (SMD=-.47; 95% CI = -1.12 to .18; P=.155; z=-1.423). There was nonsignificant heterogeneity between groups (P=.410) and substantial to considerable overall heterogeneity (I²=75.0%, P=.007). No evidence of publication bias was observed by visual inspection of the funnel plot (Figure 4), but this was not consistent with Egger's test (Supplemental Table 2 and Figure 2). Notably, omitting the study by Ozgen et al¹⁶ markedly reduced the magnitude of the trend, especially in comparison to removing other studies (SMD=-.17; 95% CI = -.31 to -.03), although significance was maintained (Supplemental Figure 1).

Delayed PLR. A significant negative association was observed between PLR collected at delayed timepoints and 90-day good functional outcomes (SMD=-.43; 95% CI = -.54 to -.32; P<.0001; z=-7.454; Figure 2). This was consistent in the group comprising patients receiving IVT only (SMD=-.47; 95% CI = -.63 to -.32; P<.0001; z=-5.911). There was only one study in both the IVT±EVT (SMD=-.33; 95% CI = -.57 to -.09; P=.006; z=-2.746) and all treatment combination groups (SMD=-.42; 95% CI = -.64 to -.20; P<.0001; z=-3.748), both of which provided a significant effect. There was no significant heterogeneity between groups (P=.615) and nonsignificantly low overall heterogeneity (I²=0.0%, P=.680). Some evidence of publication bias was observed by visual inspection of the funnel plot (Figure 4), and this was confirmed by Egger's test (Supplemental Table 2 and Figure 2).

Association of PLR With 90-day Mortality

One study (n = 150) reported admission PLR values for this outcome, and 2 studies (n = 566) reported PLR values collected at delayed timepoints. A meta-analysis could not be carried out for either due to the limited number of studies.

Admission PLR. A meta-analysis could not be performed. However, the systematic review indicated that a higher admission PLR could be associated with 90-day mortality. However, this was a significant difference in only the results of Ozgen et al and not Inanc & Inanc.^{16,21} Altintas et al¹² did not provide PLR values by outcome but carried out an analysis stratified by an optimal PLR value determined from receiver operating curve analysis, where mortality was significantly higher in the group with a higher PLR than in the group with a lower PLR.

Delayed PLR. A meta-analysis could not be performed. However, a systematic review indicated that a higher admission PLR could be associated with 90-day mortality, with both Chen, Ren et al²⁰ and Xu et al¹⁹ reporting this association. The difference in delayed PLR between groups was significant in both studies.

Association of PLR With Radiological Bleed

There were 3 studies (n=505) reporting admission PLR values for this outcome and 2 studies (n=490) reporting PLR collected at delayed timepoints. A meta-analysis was not carried out for the latter due to a lack of studies.

Admission PLR. The meta-analysis showed that although there was a positive association of admission PLR with radiological bleeding, this was not significant (SMD=.27; 95% CI = -.15 to .70; P=.209; z=1.256; Figure 3). Considering the treatment provided, a nonsignificant effect was also observed in IVT-treated patients (SMD=.09; 95% CI = -.25 to .42; P=.614; z=.505). There was only one study reporting on patients who received EVT±IVT (SMD=.70; 95% CI = .26 to 1.14; P=.002; z=3.138), which showed a significant effect. There was significant heterogeneity between groups (P=.029) and significant substantial to considerable overall heterogeneity (I²=71.6%, P=.030). No major evidence of publication bias was observed by visual inspection of the funnel plot (Figure 4), and this was confirmed by Egger's test (Supplemental Table 2 and Figure 2). Notably, omitting the study by Eren et al¹⁴ provided a result not crossing the line of no effect (SMD=.45; 95% CI = .01 to .90).

Delayed PLR. A meta-analysis could not be performed. However, a systematic review indicated that a higher admission PLR could be associated with radiological bleeding, with both Ferro et al²³ and Topcuoglu et al¹⁸ reporting this. Ferro et al.'s results, split into grades of radiological bleeding, while significant in univariate analyses, did not provide a significant result in multivariate modelling. Topcuoglu et al showed a significant difference in delayed PLR between groups.

Association of PLR With Early Neurological Improvement (ENI)

There were 3 studies (n=1475) reporting admission PLR values and 1 study (n=165) reporting relevant data for PLR collected at delayed timepoints. A meta-analysis was not carried out for the latter group.

Admission PLR. A significant decrease in admission PLR was associated with ENI (SMD=-.18; 95% CI = -.29 to -.08; P=.001; Figure 2). This significant effect was observed in both patients receiving IVT only (SMD=-.25; 95% CI = -.44 to -.06; P=.012) and bridging therapy (SMD=-.16; 95% CI = -.28 to -.03; P=.014), although the latter group had only one study. There was no significant heterogeneity between groups (P=.426) and nonsignificantly low overall heterogeneity (I²=0.0%, P=.656). No major evidence of publication bias was observed by visual inspection of the funnel plot (Figure 4), and this was confirmed by Egger's test (Supplemental Table 2 and Figure 2).

Delayed PLR. A meta-analysis could not be performed. However, the systematic review indicated mixed results. A lower delayed PLR was observed in patients with ENI in both studies included in the systematic review, but this reached statistical significance only in the results of Topcuoglu et al and not Inanc & Inanc.^{18,21}

Association of PLR With Early Neurological Deterioration (END)

Only one study each for both admission (n=1060) and delayed (n=325) PLR was available with relevant data for consideration in the meta-analysis.

Admission PLR. A meta-analysis was not possible. However, the systematic review provided mixed results, with Inanc & Inanc reporting nonsignificantly higher admission PLR values in patients with END than in those without END, but Gong et al reporting a significantly higher admission PLR in the END group than in the ENI and neither ENI nor END groups.^{15,21}

Delayed PLR. Only one study was available for this outcome for PLR collected at delayed timepoints, and thus, neither a meta-analysis nor systematic review was possible.

Association of PLR With Dramatic Early Neurological Improvement (DENI)

For this outcome, only one study (n=165) reported relevant data for consideration in meta-analyses, providing both admission and delayed PLR data.

Admission PLR. A meta-analysis was not possible. The systematic review indicated a possible association, with Inanc & Inanc reporting significantly lower admission PLR values in patients with DENI compared to those without, and Topcuoglu et al reporting the same trend but no statistical significance.^{18,21}

Delayed PLR. Only one study was available for this outcome for PLR collected at delayed timepoints, and thus, neither a meta-analysis nor systematic review was possible.

Association of PLR With 90-day mRS 0-1

Only one study (n=165) was available with relevant data for consideration in the meta-analyses for this outcome, providing both admission and delayed PLR data.

Admission PLR. A meta-analysis was not possible. However, the systematic review indicated that admission PLR was not associated with 90-day mRS 0-1, with both Topcuoglu et al and Inanc & Inanc reporting nonsignificant differences between groups, although PLR was lower in both the mRS 0-1 group in both studies.^{18,21}

Delayed PLR. Only one study was available for this outcome for PLR collected at delayed timepoints, and thus, neither a meta-analysis nor systematic review was possible.

Association of PLR With Symptomatic Intracerebral Hemorrhage (sICH)

Only one study (n=165) was available for consideration in the meta-analyses for this outcome, providing both admission and delayed PLR data.

Admission PLR. A meta-analysis was not possible. However, the systematic review provided unclear results; although a higher PLR was observed in patients with sICH by both Topcuoglu et al and Inanc & Inana, this was only significant in the former study.^{18,21} Additionally, Diestro et al³² reported that admission PLR was not significantly associated with sICH in adjusted analyses.

Delayed PLR. Only one study was available for this outcome for PLR collected at delayed timepoints, and thus, neither a meta-analysis nor systematic review was possible.

Association of PLR With Successful Recanalisation

Only two studies provided relevant admission PLR data for consideration of the meta-analysis for this outcome (n=373), with the same definition of successful recanalisation (mTICI 2b-3).^{33,34} The systematic review indicated that admission PLR may be associated with recanalisation outcomes, as the aforementioned studies both reported lower admission PLR in patients with a favorable recanalisation outcome, as did Inanc & Inanc,²¹ who defined this in terms of complete recanalisation or percentages thereof using the Thrombolysis in Brain Ischemia (TIBI) scale. Of these three studies, the differences between groups were significant only in Lee et al.³³'s work. Another study, reporting on the first pass effect (FPE), where complete recanalisation (mTICI 3) was achieved with a single pass, showed a significantly higher PLR in the non-FPE group,³⁵ further supporting an association of admission PLR with recanalisation outcomes.

Only one study was available for this outcome for PLR collected at delayed timepoints, and thus, neither a meta-analysis nor systematic review was possible.

Association of PLR With Stroke Associated Infection (SAI)

Only one study was available for this outcome, providing only admission PLR data, and thus no systematic review or meta-analysis was possible.

Discussion

This study investigated the association of PLR with outcomes after AIS in patients receiving RT and is, to the best of our

Table 3. Platelet-Lymphocyte Values Stratified by Outcome.

ID	AUTHOR	PLR TIME-POINT	OUTCOME GROUP	PATIENTS (N)	PLR VALUE (MEAN (SD))
1	Yi et al ⁶⁰	Admission	GFOs	245	119.2 (108.5)
			No GFOs	195	137.4 (127.2)
2	Chen, Lin et al ¹³	Admission	GFOs	136	134.9 (60.6)
			No GFOs	105	146.7 (73.3)
			GFOs	136	148.4 (78.9)
			No GFOs	105	187.5 (88.6)
3	Ferro et al ²³	Delayed	GFOs	147	163.67 (104.82)
			No GFOs	178	215 (133.79)
			END	85	230.67 (157.62)
			No END	240	229.33 (152.89)
			RB*	168	214.32 (143.79)
			No RB*	160	168.00 (109.97)
4	Ozgen et al ¹⁶	Admission	GFOs	92	123.54 (44.13)
			No GFOs	58	183.5 (104.2)
			Mortality	33	178.71 (79.60)
			No mortality	117	144.76 (103.77)

(Continued)

Table 3. Continued.

ID	AUTHOR	PLR TIME-POINT	OUTCOME GROUP	PATIENTS (N)	PLR VALUE (MEAN (SD))			
5	Topcuoglu et al ¹⁸	Admission	mRS 0-1**	54	131 (79)			
			No mRS 0-1**	110	142 (89)			
			GFOs	81	130 (79)			
			No GFOs	84	147 (92)			
			RB	43	155 (105)			
			No RB	122	134 (78)			
			sICH	11	215 (112)			
			No sICH	154	134 (81)			
			ENI	86	126 (82)			
			No ENI	79	152 (89)			
			DENI	47	133 (85)			
			No DENI	118	141 (87)			
			Delayed			mRS 0-1	54	141 (62)
						No mRS 0-1	110	207 (146)
						GFOs	81	146 (66)
						No GFOs	84	223 (162)
						RB	43	246 (185)
No RB	122	163 (95)						
sICH	11	427 (271)						
No sICH	154	168 (93)						
ENI	86	157 (96)						
No ENI	79	215 (154)						
6	Xu et al ¹⁹	Delayed	DENI	47	147 (75)			
			No DENI	118	199 (145)			
			GFOs	166	143.83 (73.66)			
			No GFOs	120	172.27 (100.33)			
			Mortality	38	202 (124.42)			
No mortality	151.83	87.62						

(Continued)

Table 3. Continued.

ID	AUTHOR	PLR TIME-POINT	OUTCOME GROUP	PATIENTS (N)	PLR VALUE (MEAN (SD))
7	Feng et al ²²	Admission	RB	34	260.92 (181.35)
			No RB	56	170.25 (83.35)
8	Sengeze et al ³⁴	Admission	SR	43	144.77 (80.30)
			No SR	48	167.03 (110.59)
9	Sarioglu et al ³⁵	Admission	FPE	32	103.17 (37.06)
			No FPE	51	195.35 (101.49)
10	Gong et al ¹⁵	Admission	END	193	184.1 (96.35)
			No END	867	130.94 (46.41)
			ENI	398	134.6 (48.88)
			NO ENI	662	144.24 (68.73)
11	Deng et al ⁴⁹	Admission (IVT only)	SAI	141	168 (89.88)
			No SAI	196	127.91 (52.52)
		Admission (EVT±IVT)	SAI	219	208 (117.91)
			No SAI	114	174.33 (93.11)
12	Chen, ren et al ²⁰	Delayed	GFOS	194	125.27 (53.28)
			No GFOS	86	146.97 (5)
			Mortality	27	165.5 (111.81)
			No mortality	253	127.91 (52.52)
13	Eren et al ¹⁴	Admission	ENI	136	129.03 (70.58)
			No ENI	114	145.61 (86.23)
			RB	27	129.6 (61.98)
			No RB	223	137.45 (80.20)
14	Lee et al ³³	Admission	SR	224	120.97 (65.59)
			No SR	58	161.37 (90.39)

(Continued)

Table 3. Continued.

ID	AUTHOR	PLR TIME-POINT	OUTCOME GROUP	PATIENTS (N)	PLR VALUE (MEAN (SD))
15*	Inanc & inanc ²¹	Admission	Mortality	23	72.68 (107.67)
			No mortality	33	22.78 (19.17)
			RB	24	24.10 (20.51)
			No RB	32	61.56 (100.45)
16*	Altintas et al ¹²	Admission			
17*	Diestro et al ³²	Admission			
18*	Chen, Li et al ⁵¹	Admission			
		Delayed			

All values were provided to 2 decimal places where rounding was required. Outcome groups where data were available included only. Definitions of outcomes were as outlined in the text and the caption of Figure 2.

*Total numbers for this outcome were not congruent with other outcomes. However, numbers provided in data tables were used for relevant analyses.

**Data missing for one patient.

**Only included in the systematic review.

Abbreviations: n=Number of patients in each outcome group; ENI=Early Neurological Improvement; DENI=Dramatic ENI; END=Early Neurological Decline; GFOs=Good Functional Outcomes; PLR=Platelet-Lymphocyte Ratio; sICH=symptomatic intracerebral haemorrhage; RB=Radiological Bleed; SR=Successful Recanalisation; SAI=Stroke Associated Infection; SAP=Stroke Associated Pneumonia; FPE=First Pass Effect.

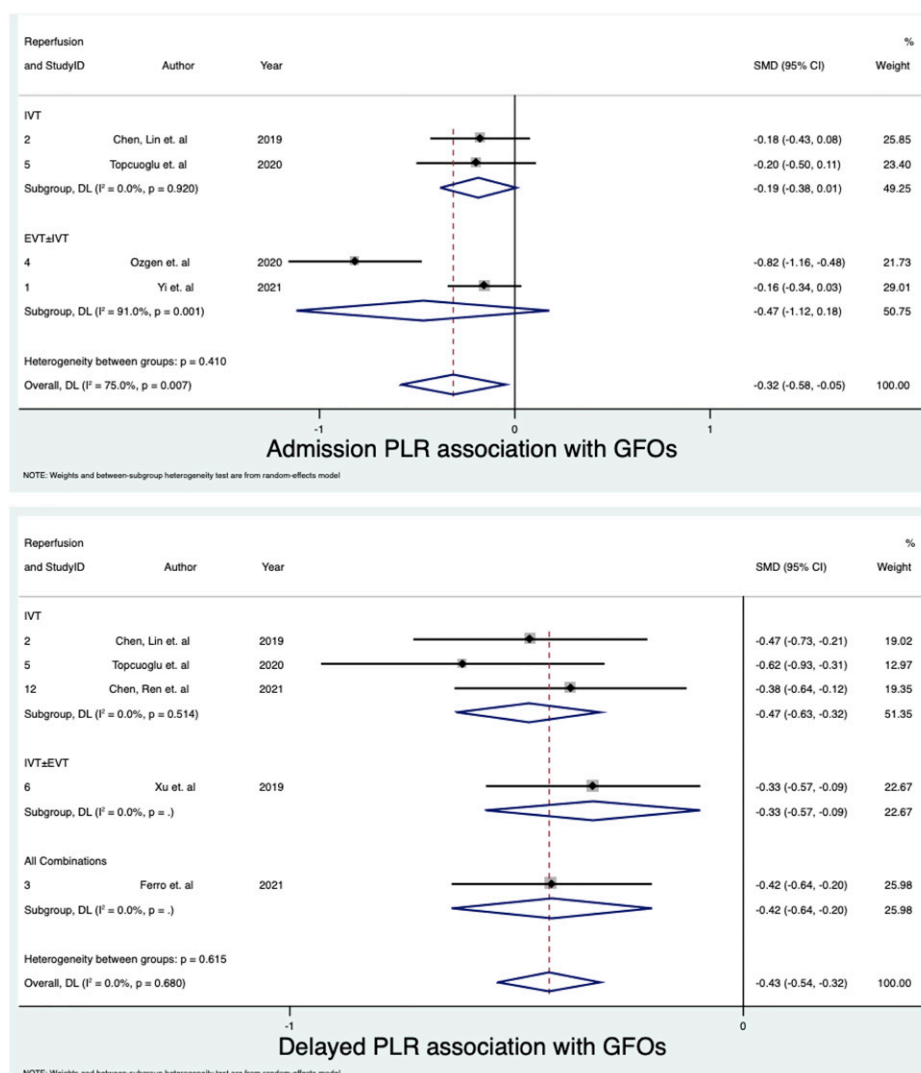


Figure 2. Forest plots showing association of platelet lymphocyte ratio (PLR) with good functional outcomes. Abbreviations: GFOs=Good Functional Outcomes; PLR=Platelet-Lymphocyte Ratio; IVT=Intravenous Thrombolysis; EVT=Endovascular Thrombectomy.

knowledge, the first of its kind. We demonstrate that a lower admission PLR is significantly associated with ENI, and a lower PLR collected at both admission and delayed timepoints is significantly associated with 90-day good functional outcomes. Additionally, our meta-analysis suggests that higher PLR collected at delayed timepoints may be associated with 90-day mortality and radiological bleeding. As such, there is a role for PLR in predicting outcomes and informing treatment decisions.

PLR is a beneficial biomarker owing to its low cost and availability from standard blood panels,^{4,10} and its ability to provide insight into both the hemostatic/thrombotic and inflammatory pathways underlying AIS pathogenesis.²² It is of particular interest in patients receiving RT due to the association of platelets with increased recruitment and activation of immune cells, including neutrophils, which have a role in ischaemia-reperfusion injury (IRI), a phenomenon where stroke continues to progress despite resolution of the occlusion,

along with the no-reflow phenomenon, whereby blood flow is not restored despite removal of the occlusion. The role of platelets in formation of neutrophil extracellular traps may contribute to this.^{4,36-38} Indeed, several studies pertaining to cardiovascular reperfusion show a strong association of elevated PLR with the no-reflow phenomenon.^{7,8} This may explain our observation that lower PLR levels collected at delayed timepoints were associated with good functional outcomes and higher delayed PLR levels with mortality in studies included in the systematic review, but admission PLR showed nonsignificant results in each of the treatment subgroups, as all included patients received RT and hence these phenomena may have occurred. These results should be interpreted with caution due to the publication bias detected in the delayed timepoints analysis. Similar trends were seen in the systematic review component of our study looking at association of admission PLR with mRS score 0-1. Although both included studies showed nonsignificant results, this should be interpreted in light

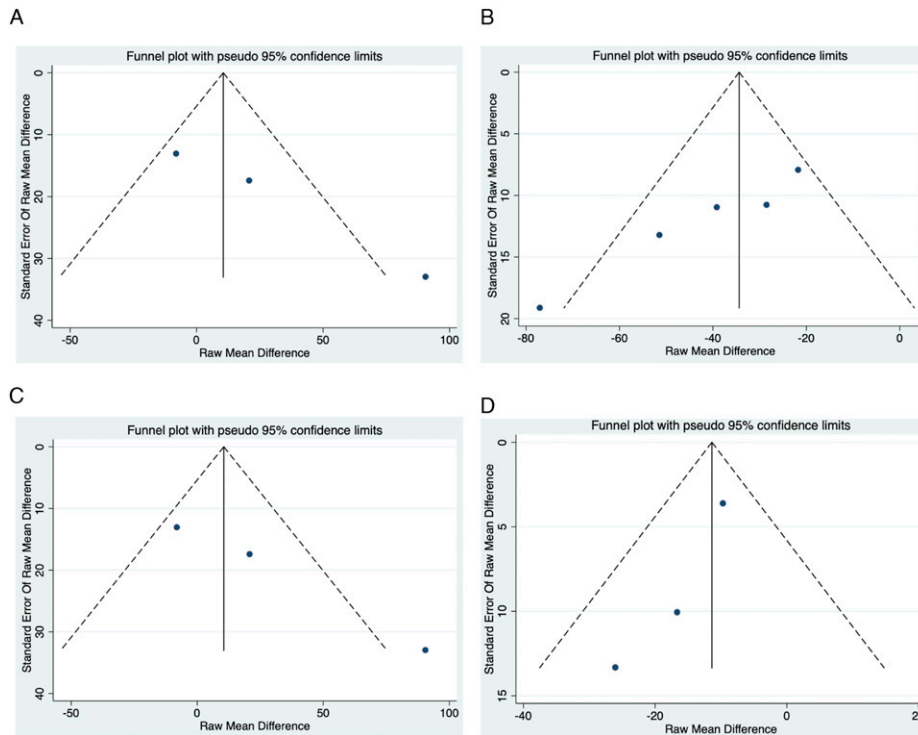


Figure 3. Forest plots showing association of platelet lymphocyte ratio (PLR) with radiological bleed and early neurological improvement (ENI). Abbreviations: PLR=Platelet-Lymphocyte Ratio; IVT=Intravenous Thrombolysis; EVT=Endovascular Thrombectomy; ENI=Early Neurological Improvement.

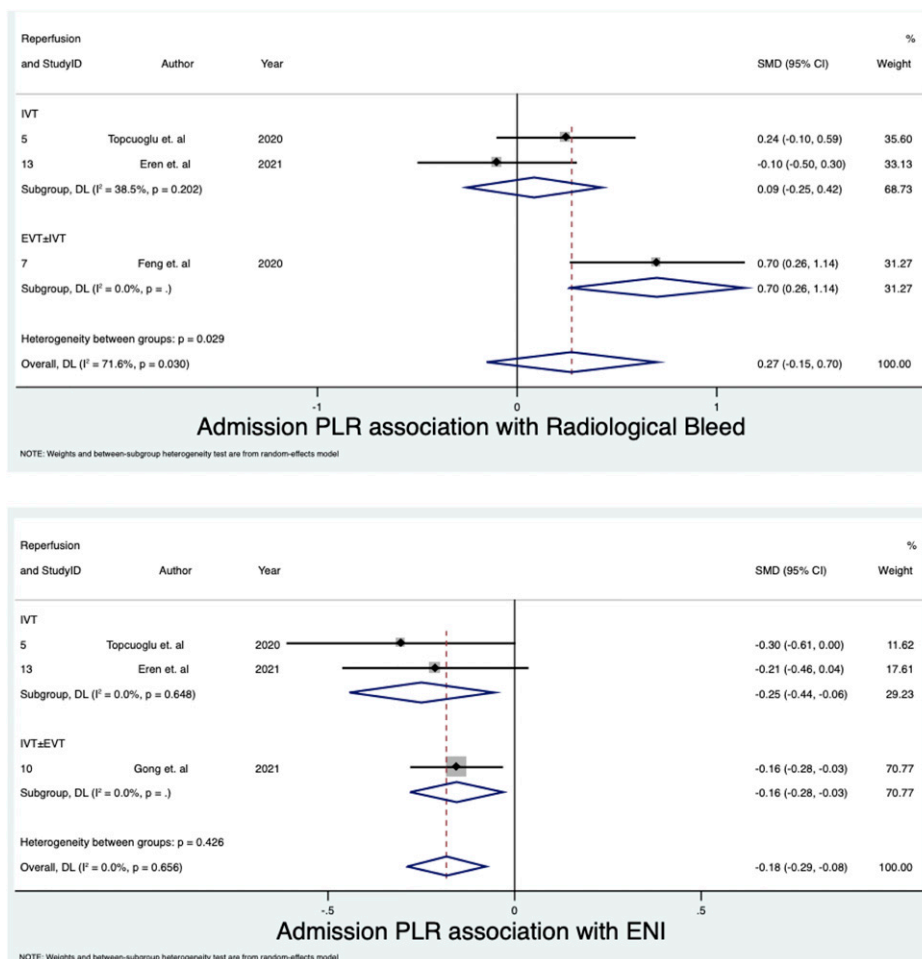


Figure 4. Funnel Plots for each meta-analysis. Note: Funnel plots for each meta-analysis. A: Admission PLR association with Good Functional Outcomes; B: Delayed PLR association with Good Functional Outcomes; C: Admission PLR association with Radiological Bleed; D: Admission PLR association with ENI.

of the small sample size and retrospective study design. There were no studies considering PLR collected at delayed timepoints. Drawing conclusions about utility of PLR in predicting recanalisation outcomes was restricted by similar issues for both admission and delayed PLR timepoints. As such, there is a pressing need for further prospective high-powered studies in this space, as this may help determine which patients would benefit from recanalisation, and could assist in monitoring patients at high risk of IRI and no-reflow injury more closely, and inform long-term follow up of high risk patients.^{4,10} Considering recent data showing a sizeable recurrence rate in AIS, the latter is of enormous significance for early intervention and initiating aggressive preventive measures.¹ There is also a pressing need for further studies with regards to bridging therapy or EVT only, as most studies included in these analyses had cohorts where patients received IVT only; it is well established that treatment with EVT has the potential to improve outcomes,³⁹ and hence, understanding the role of PLR in these patients is critical to determining its putative role in clinical practice and treatment decision-making.

We also showed that patients with ENI had lower admission PLR levels. This may be related to previous hypotheses that higher PLR levels may reflect a higher burden of high-risk plaques and atherosclerotic risk and thus may be associated with worse prognosis due to plaque instability.¹² Additionally, it may also reflect less immune-mediated inflammation, which has a notable role in early neurological changes.^{4,13} It has also recently been reported that PLR may reflect initial stroke severity, and this may have been a contributing factor to our results,⁴⁰ Though only 2 studies included in the meta-analysis for ENI provided stroke severity data, and future research into this area is necessary. Most studies did not consider delayed PLR values for early neurological change outcomes, likely owing to this outcome generally being measured at 24 hours, which is when most delayed PLR measurements are taken.^{15,18} However, a meta-analysis could not be carried out to determine the role of admission PLR in predicting other early neurological outcomes, END and DENI, and further primary research is necessary for this. Interestingly, the results reported by Gong et al (n=1060) indicated that END was statistically significantly associated with admission PLR, which is consistent with an understanding that platelets can drive early stroke inflammation via cross-talk mechanisms with deleterious immune cells.^{4,15,38,41} However, the other study in the systematic review for this outcome had only 56 patients, and thus, further research is critical to validate these results. PLR has also been shown to have utility in the coronavirus disease 2019 (COVID-19) setting, with higher PLR associated with increased clinical deterioration and thrombosis risk, and thus, considering reports of increased stroke risk in the milieu of COVID-19, along with involvement of the clotting cascade, PLR could have a putative role in predicting short-term neurological outcomes.⁴²⁻⁴⁵ As we have reported previously, such biomarkers may also help facilitate treatment owing to the increasing reliance on telemedicine to

keep up with increased healthcare demands in the COVID-19 era.^{4,10,42,46}

There has been a particular interest in using PLR for predicting bleeding outcomes considering previous reports that thrombocytopenic patients have a higher risk of sICH than patients with a normal platelet count.⁴⁷ However, our meta-analysis showed somewhat diverging results, with radiological bleeding seemingly indicated by a higher PLR collected at delayed timepoints. This could be attributed to previous observations that platelets can, in conjunction with neutrophils and fibrinogen, cause blood brain barrier (BBB) damage after initial occlusion and thus may predispose patients to hemorrhagic transformation.²² We observed no statistically significant association of admission PLR with radiological bleeding and that removing Eren et al caused the effect to no longer cross the line of no effect. This was surprising, as this was the study with the longest follow-up CT scanning, extending up to 72 hours, more than twice the next highest included study.¹⁴ This could be related to stroke etiology, as PLR, due to it being a marker of high-risk plaques, has been purported as being more relevant in LAA stroke,²³ which formed only 32.10% of our included patient population, with relevant etiology data only available for one of the 3 studies included in the admission PLR association with radiological bleed meta-analysis. Again, there is a pressing need for further studies considering bridging therapy and EVT-treated patients to further validate the utility of PLR in AIS patients in informing treatment decisions. A meta-analysis could not be carried out for sICH or SAI, and the systematic review for the former was limited by small underpowered studies. Neither a systematic review nor a meta-analysis could be conducted for SAI, but this area that warrants further attention, especially with evidence suggesting SAI may impact immune cell counts and the thromboinflammatory role of platelets.^{4,48}

The major strength of our study is that, to the best of our knowledge, it is the first of its kind. Additionally, our use of SMD to account for the continuous nature of PLR allows us to overcome the issue of varying thresholds in the published literature. We also considered admission and delayed PLR separately, using pre- and postintervention definitions, thereby accounting for the role of platelets in IRI and no-reflow injury in clinical and recanalisation outcomes and in potential BBB damage leading to HT,^{22,38} as well as the dynamicity of immune cells in AIS pathogenesis.^{4,10} By considering treatment modalities separately, we also clearly show that there are differences in treatment modalities, which can be used not only to determine the utility of PLR in different patient subgroups but also to guide further research.

Limitations of our study include that most studies retrospective in design; we sought to minimise the ensuing heterogeneity through random-effects modelling. Additionally, some outcomes lack a breadth of published literature, with many low powered studies; hence, conclusions could not be drawn about the role of PLR in prognosticating these outcomes. Additionally, we were limited from making conclusions about

bridging therapy and IVT because there were very few studies with cohorts of patients treated as such. The role of racial and ethnic differences in PLR should also not be discounted, but we were unable to consider this in our analysis due to minimal reporting of data stratified as such.⁴⁹ Notably, our findings relate to AIS patients receiving RT and thus, should be interpreted in this scope. Very few studies considered the association of dynamic PLR changes with outcomes, and thus, this could not be analysed. Finally, not all studies^{12,13,17,22,34,35} excluded patients with active infection, malignancy, or chronic disease, which can influence lymphocyte count via systemic inflammation.^{4,10}

Conclusion

Our study clearly shows that there is a putative role for PLR in the management of AIS patients treated with RT, particularly regarding predicting long-term outcomes in these patients, which may be related to IRI and no-reflow injury and for predicting bleeding outcomes. As such, PLR could be beneficial in informing management decisions and predicting higher-risk patients who might need closer monitoring post RT. Implementing systems level changes for such monitoring could help address the AIS global disease burden. Given the rapid access and low costs, PLR values could be implemented on a systems level to stratify secondary prevention and improve ongoing surveillance as well as follow-up strategies for AIS patients at high-risk of poor morbidity and mortality. However, there is a pressing need for further prospective high-powered studies considering bridging therapy and EVT-only patients. Additionally, further experimental studies are required to better elucidate pathophysiological mechanisms underlying the role of PLR in AIS, as this may mediate clinical outcomes and bleeding complications.

Author Contributions

SMMB conceptualised the study, contributed to the planning, drafting and revision of the manuscript, and supervised the student. SMMB encouraged DS to investigate and supervise the findings of this work. DS and SMMB wrote the first draft of this paper and were involved in data extraction and analyses. All authors contributed to the revision of the manuscript. All authors approved the final draft of the manuscript.

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