

BMJ Open Predicting frequent hospital admission risk in Singapore: a retrospective cohort study to investigate the impact of comorbidities, acute illness burden and social determinants of health

Lian Leng Low,^{1,2} Nan Liu,^{3,4,5} Sijia Wang,⁶ Julian Thumboo,^{3,7} Marcus Eng Hock Ong,^{5,8} Kheng Hock Lee^{1,2}

To cite: Low LL, Liu N, Wang S, *et al.* Predicting frequent hospital admission risk in Singapore: a retrospective cohort study to investigate the impact of comorbidities, acute illness burden and social determinants of health. *BMJ Open* 2016;**6**:e012705. doi:10.1136/bmjopen-2016-012705

► Prepublication history for this paper is available online. To view these files please visit the journal online (<http://dx.doi.org/10.1136/bmjopen-2016-012705>).

Received 18 May 2016
Revised 24 September 2016
Accepted 27 September 2016



CrossMark

For numbered affiliations see end of article.

Correspondence to
Dr Lian Leng Low;
low.lian.leng@singhealth.com.sg

ABSTRACT

Objectives: To evaluate the impact of comorbidities, acute illness burden and social determinants of health on predicting the risk of frequent hospital admissions.

Design: Multivariable logistic regression was used to associate the predictive variables extracted from electronic health records and frequent hospital admission risk. The model's performance of our predictive model was evaluated using a 10-fold cross-validation.

Setting: A single tertiary hospital in Singapore.

Participants: All adult patients admitted to the hospital between 1 January 2013 and 31 May 2014 (n=25 244).

Main outcome measure: Frequent hospital admissions, defined as 3 or more inpatient admissions within 12 months of discharge. Area under the receiver operating characteristic curve (AUC) of the predictive model, and the sensitivity, specificity and positive predictive values for various cut-offs.

Results: 4322 patients (17.1%) met the primary outcome. 11 variables were observed as significant predictors and included in the final regression model. The strongest independent predictor was treatment with antidepressants in the past 1 year (adjusted OR 2.51, 95% CI 2.26 to 2.78). Other notable predictors include requiring dialysis and treatment with intravenous furosemide during the index admission. The predictive model achieved an AUC of 0.84 (95% CI 0.83 to 0.85) for predicting frequent hospital admission risk, with a sensitivity of 73.9% (95% CI 72.6% to 75.2%), specificity of 79.1% (78.5% to 79.6%) and positive predictive value of 42.2% (95% CI 41.1% to 43.3%) at the cut-off of 0.235.

Conclusions: We have identified several predictors for assessing the risk of frequent hospital admissions that achieved high discriminative model performance. Further research is necessary using an external validation cohort.

INTRODUCTION

Patients with frequent admissions consume a disproportionately large amount of

Strengths and limitations of this study

- Study identified discriminative predictors for frequent admission risk which are easily retrievable from the electronic health record.
- High discriminatory power with acceptable sensitivity and specificity.
- Predictors on functional status and patient behaviour were not available.

healthcare resources and contribute to bed shortages in hospitals.¹ Moreover, frequent admitters (FAs) experience significant psychological stress and financial burden.² In countries with a rapidly ageing population, increasing life expectancy, rising chronic disease burden and decreasing working population, are factors that contribute to straining the finite healthcare resource further. There is consensus that reducing avoidable hospital readmissions require identification of patients at highest risk for readmissions, followed by identification of potentially modifiable risk factors that increase readmission risk.³ Resource intensive discharge planning and postdischarge monitoring can then be targeted at these risk factors.³ Finally, community-based home care services, primary care and case management could be provided to such high-risk individuals to optimise their care in the community and keep them out of hospitals.

In the USA, published prediction models have focused on predicting for 30-day readmission risk, in line with avoiding Medicare penalties for excessive readmission rates under the Readmission Risk Prevention Program.⁴ However, there is a case to be made for taking a longer term view, especially in predicting high-cost users. It is well

established that a high proportion of health expenditures are consumed by a small percentage of individuals. In the USA, the top 10% of residents in terms of healthcare costs accounted for 64% of all health expenditures.⁵ Beyond the postdischarge transitional period, a longer term strategy in the form of community-based case management and care integration are important to build healthy communities and cope with the complex healthcare needs of rapidly ageing populations. The ability to predict future high-cost users in the following 12 months will allow time for intervention by targeted transitional care programmes aimed at modification of behavioural risk factors, addressing social and functional care needs, with potentially large benefit. So far, most models are derived using administrative data and perform poorly in predicting 30-day readmission risk, with only six studies reporting a C-statistic >0.7.⁶ To date, no published literature has investigated the effect of severity during acute illness, such as the treatment with high-risk medications and requirement for dialysis during the hospitalisation episode, on readmission risk. Similarly, the index discharge diagnosis would be an important determinant of readmission risk⁴ but is rarely investigated in the published literature.

Singapore is a developed city state in Southeast Asia with a multiethnic population of 5.6 million people. Its population is one of the most rapidly ageing in Asia with an increasing chronic disease burden.⁷ Healthcare expenditure is expected to triple from S\$4 billion in 2011 to S\$12 billion in 2020. Like most countries in the developed world, the main cost driver of healthcare in Singapore is inpatient care.⁸ Reducing unnecessary readmissions and providing care at the most appropriate setting is the focus of health systems reforms aimed at reining in escalating healthcare costs. In Singapore, FAs are defined as patients with three or more inpatient admissions in 12 months. FAs are high-cost patients with an average cost per patient approaching SG\$30 000.⁹ Interestingly, a cumulative direct cost of 30 000 Canadian dollars (SG\$32 260) marks the top 5% of healthcare resource users in Ontario, Canada.¹⁰ Taking a population health approach, the Ministry of Health (MOH) Singapore has created six regional health systems (RHSs) in 2011, each being responsible to integrate care for a specific geographic region. The Singapore Health Services (SingHealth) RHS is the largest healthcare cluster in Singapore providing care for the south-central part of Singapore. The Singapore General Hospital (SGH) is the flagship hospital of the SingHealth RHS and the largest tertiary care hospital in Singapore with 37 clinical specialities and 88 000 inpatient admissions each year. Each RHS will receive capitated funding to manage and reduce readmission risk of patients at high risk of frequent admissions in its catchment population. Since 2014, SingHealth RHS has trained patient navigators¹¹ to function as case managers and reorganised its transitional home care service into an integrated practice unit¹² to coordinate the care of

patients through the complex healthcare system for the 1 year after hospital discharge.

For interventions to be targeted at patients at highest risk, we sought to identify predictors of frequent hospital admission risk. These candidate predictors should be readily available in the electronic health records (EHR) for a predictive score to be automated or can be easily retrieved from patient medical records and entered into an online spreadsheet or smartphone/tablet application to facilitate clinical use.

Our primary objective was to investigate the impact of comorbidities, acute illness burden and social determinants of health on predicting frequent readmission risk (defined as three or more inpatient admissions within 12 months of discharge) and measure the performance of our prediction model using receiver operating characteristic (ROC) analysis, sensitivity, specificity and precision.

MATERIALS AND METHODS

Study design and population

This was a retrospective single-centre cohort study designed to investigate the impact of comorbidities, acute illness burden and social determinants of health on the risk of frequent hospital admissions. All admitted adult patients (≥ 21 years of age) from 1 January 2013 to 31 May 2014 at the SGH were enrolled. Only index admissions were included. Patients were excluded if they were non-residents or died during the index admission. Patients were also excluded if their admission specialty was obstetrics, emergency medicine, dentistry or ophthalmology. Emergency medicine admissions were excluded as these were observation ward admissions rather than true hospital admissions. Patients admitted to the emergency medicine observation ward are typically monitored up to 24 hours and subsequently discharged home or converted to hospital admissions according to their clinical circumstances. Patients converted to hospital admissions would be captured in our data set. We excluded patients admitted to obstetrics as these admissions are pregnancy related; and admissions to dentistry and ophthalmology which are usually elective in nature.

Outcomes

The primary outcome was frequent hospital admissions, which were defined as three or more inpatient admissions within 12 months of discharge.

Predictive variables

Candidate variables were identified a priori and according to the medical literature.^{6 13–15} A total of 44 variables were selected and extracted from the hospital's EHR, including patient demographics (age, gender and ethnicity) and prior healthcare usage (number of hospital admissions during the past year, number of emergency department visits in the past 6 months and number of specialist outpatient clinic visits in the past year). Comorbidities (major

diseases listed under the Charlson Comorbidity Index (CCI) and Elixhauser Comorbidities)^{16–18} were derived by extracting the International Classification of Diseases (ICD)-9 and ICD-10 codes dating back to 7 years. We believe that this is most comprehensive among the published literature¹⁹ and would account for potential lapses in diagnostic coding.

We included social determinants of health in our list of predictors. We used eligibility for admission to subsidised ward class and assistance by Medifund as predictors of low socioeconomic status. Singapore uses a tiered system of healthcare subsidies that is adjusted using means testing. Patients in the lowest socioeconomic tier receive the highest tier of subsidy. Such patients are characterised by their entitlement for admission to the ward class that receive the highest subsidy. In addition they are eligible to receive financial support from Medifund. Medifund is an endowment fund set up by the Singapore government to help needy Singaporeans.²⁰ Medifund is a safety net for patients who face financial difficulties with their remaining hospital bills after receiving government subsidies and drawing on other means of payments including insurance. Therefore, patients who needed financial assistance using Medifund would belong to the neediest in society.

We studied markers of acute illness burden as predictors. These included length of stay of the index admission; 'treatment with intravenous furosemide 40 mg or more'; 'treatment with second-line intravenous antibiotics' (defined as piperacillin-tazobactam, meropenem, vancomycin and moxifloxacin) and 'required dialysis'. Our study is also novel in weighting the index discharge diagnoses using the CCI, which is already a validated marker for mortality within 1 year.¹⁷

The psychological state of patients is known to be an important factor in coping with illness. The presence of a debilitating level of mood disorders was inferred from treatment with antidepressants in the past 1 year. We used treatment with antidepressants as it is a more accurate indicator of depressive mood disorders, of which many patients are treated as outpatients and would not have discharge ICD codes for depression. The other comorbidities (table 1) were identified using ICD-10 codes¹⁸ in any primary or secondary diagnosis fields dating back to 7 years preceding the index admission.

Statistical analysis

The demographics and clinical characteristics of the recruited population were described and compared among patients who were FAs and who were not. We analysed continuous variables using Student's t-test and the categorical variables using χ^2 test or Fisher's exact test when appropriate. Univariable logistic regression was implemented to evaluate the associations between each of the candidate predictors and the outcome. Multivariable logistic regression with stepwise variable selection was used to build predictive models. Statistical significance was set as $p < 0.05$.

To assess the predictive performance, 10-fold cross-validation was conducted for model evaluation. First, the whole data set was evenly divided into 10 non-overlapped subsets. Second, a predictive model was built with nine subsets of data and validated on the remaining one subset. Third, the second step was repeated another nine times so that all subsets were tested. Finally, the ROC analysis was conducted to evaluate the model performance. Further measures of diagnostic accuracy such as sensitivity, specificity, positive predictive value (PPV) and negative predictive value (NPV) were presented. Data analyses in this study were conducted in R V.3.2.3 (R Foundation, Vienna, Austria).

RESULTS

A total of 25 244 patients were included in this study, of which 4322 patients (17.1%) had three or more inpatient admissions within 12 months of discharge and were classified as FA. Table 1 describes the baseline characteristics of the study population. The mean age of FAs was 65 (SD=15) years and the mean age of non-FAs was 57 (SD=18) years. There were 13 010 male patients (51.5%). Most admissions were emergent (76.5%, $n=19\,320$). Mean length of stay was 5.59 days. On average, FAs had a mean CCI of 3.62 and non-FAs had a mean CCI of 1.06. The readmitted patients were older compared with the non-readmitted patients, had significantly longer length of stay during the index admission, higher CCI and more emergency department admissions in the preceding 6 months.

The multivariable logistic regression was used in predicting frequent hospital admissions. Eleven variables were found to be significantly associated with frequent hospital admissions, that is, three or more inpatient admissions within 12 months of discharge. As shown in table 2, these variables were age, gender, eligibility for financial assistance using Medifund, number of visits to the emergency department 6 month before index admission, number of visits to specialist clinic 1 year before index admission, number of hospital admissions 1 year before index admission, admission to a subsidised ward class during index admission, required dialysis during index admission, treatment with intravenous furosemide 40 mg and above during index admission, treatment with antidepressants in the past 1 year, and CCI. The top two strongest independent predictors were treatment with antidepressants in the past 1 year (adjusted OR 2.51, 95% CI 2.26 to 2.78) and 'required dialysis' during index admission (adjusted OR 2.33, 95% CI 1.98 to 2.74). The odds of frequent hospital admission in the following 12 months for patients who were treated with intravenous furosemide 40 mg and above were 1.49 times higher than the odds of those who were not treated with the medication (adjusted OR 1.49, 95% CI 1.308 to 1.701). Every increase of one point on the CCI score was associated with 33% increased odds of becoming a FA (adjusted OR 1.33, 95% CI 1.312 to 1.347).

Table 1 Baseline characteristics by frequent hospital admission status

	All patients (n=25 244)	Frequent admitter (n=4322)	Non-frequent admitter (n=20 922)	p Value*
<i>Patient demographics</i>				
Age, mean (SD)	58.47 (17.96)	65.06 (15.11)	57.11 (18.2)	<0.001
Gender, male (%)	13 010 (51.5%)	2294 (53.1%)	10 716 (51.2%)	0.027
Ethnicity				<0.001
Chinese (%)	18 234 (72.2%)	3231 (74.8%)	15 003 (71.7%)	
Indian (%)	2504 (9.9%)	410 (9.5%)	2094 (10%)	
Malay (%)	3017 (12%)	551 (12.7%)	2466 (11.8%)	
Others (%)	1489 (5.9%)	130 (3%)	1359 (6.5%)	
<i>Social determinants of health</i>				
Required financial assistance using Medifund (%)	365 (1.4%)	130 (3%)	235 (1.1%)	<0.001
Stayed in a subsidised ward during index admission (%)	19 324 (76.5%)	3860 (89.3%)	15 464 (73.9%)	<0.001
<i>Past healthcare use</i>				
ED visits (6 months before index admission), mean (SD)	0.37 (0.92)	0.67 (1.59)	0.31 (0.69)	<0.001
Specialist clinic visits (1 year before index admission), mean (SD)	2.21 (5.52)	4.27 (8.59)	1.78 (4.52)	<0.001
Hospital admissions (1 year before index admission), mean (SD)	0.34 (1.09)	0.99 (2.02)	0.2 (0.69)	0.022
<i>Markers of acute illness burden</i>				
Index admission was urgent (%)	19 320 (76.5%)	3513 (81.3%)	15 807 (75.6%)	<0.001
Index admission was planned (%)	4822 (19.1%)	618 (14.3%)	4204 (20.1%)	<0.001
Required second line antibiotics during index admission (%)	1252 (5%)	297 (6.9%)	955 (4.6%)	<0.001
Required inpatient dialysis during index admission (%)	841 (3.3%)	406 (9.4%)	435 (2.1%)	<0.001
Required intravenous furosemide 40 mg and above during index admission (%)	1485 (5.9%)	521 (12.1%)	964 (4.6%)	<0.001
Required isolation during index admission (%)	310 (1.2%)	78 (1.8%)	232 (1.1%)	<0.001
Length of stay of index admission, mean (SD)	5.59 (10.81)	7.61 (11.59)	5.17 (10.6)	<0.001
Charlson Comorbidity Index,† mean (SD)	1.5 (2.43)	3.62 (3.01)	1.06 (2.02)	<0.001
<i>Medical comorbidities‡</i>				
Stroke (%)	397 (1.6%)	67 (1.6%)	330 (1.6%)	0.95
Metastatic disease (%)	2752 (10.9%)	482 (11.2%)	2270 (10.8%)	0.58
Non-metastatic malignancy (%)	4490 (17.8%)	749 (17.3%)	3741 (17.9%)	0.401
Peripheral vascular disease (%)	1003 (4%)	171 (4%)	832 (4%)	0.985
Heart failure or fluid overload (%)	3094 (12.3%)	539 (12.5%)	2555 (12.2%)	0.655
Pressure ulcer (%)	804 (3.2%)	142 (3.3%)	662 (3.2%)	0.714
Thromboembolism (%)	1606 (6.4%)	313 (7.2%)	1293 (6.2%)	0.01
Spine fracture (%)	693 (2.7%)	129 (3%)	564 (2.7%)	0.314
Coronary heart disease or myocardial infarction (%)	3794 (15%)	703 (16.3%)	3091 (14.8%)	0.013
Hip fracture (%)	529 (2.1%)	87 (2%)	442 (2.1%)	0.72
Atrial fibrillation (%)	1966 (7.8%)	352 (8.1%)	1614 (7.7%)	0.353
Epilepsy (%)	266 (1.1%)	47 (1.1%)	219 (1%)	0.875
Parkinsonism (%)	395 (1.6%)	84 (1.9%)	311 (1.5%)	0.033
Anxiety (%)	247 (1%)	40 (0.9%)	207 (1%)	0.761
Bipolar disorder (%)	76 (0.3%)	17 (0.4%)	59 (0.3%)	0.287
Collagen vascular disease (%)	376 (1.5%)	55 (1.3%)	321 (1.5%)	0.221
Dementia (%)	898 (3.6%)	152 (3.5%)	746 (3.6%)	0.911
Hypothyroidism (%)	593 (2.3%)	98 (2.3%)	495 (2.4%)	0.738
Chronic kidney disease, stages 1–4 (%)	4387 (17.4%)	767 (17.7%)	3620 (17.3%)	0.497
Chronic obstructive pulmonary disease (%)	719 (2.8%)	122 (2.8%)	597 (2.9%)	0.952
Osteoarthritis (%)	2152 (8.5%)	376 (8.7%)	1776 (8.5%)	0.673
Benign prostatic hypertrophy (%)	1238 (4.9%)	198 (4.6%)	1040 (5%)	0.298
Asthma (%)	781 (3.1%)	145 (3.4%)	636 (3%)	0.298
Hyperlipidaemia (%)	7216 (28.6%)	1284 (29.7%)	5932 (28.4%)	0.076
Hypertension (%)	10 136 (40.2%)	1774 (41%)	8362 (40%)	0.194
Chronic kidney disease stage 5 or end-stage renal failure (%)	3423 (13.6%)	611 (14.1%)	2812 (13.4%)	0.233
Diabetes (%)	6303 (25%)	1098 (25.4%)	5205 (24.9%)	0.478
History of alcoholism (%)	431 (1.7%)	103 (2.4%)	328 (1.6%)	<0.001
Treatment with antidepressants§ (%)	2513 (10%)	1048 (24.2%)	1465 (7%)	<0.001

*Continuous variables were analysed using Student's t-test and categorical variables were analysed using χ^2 test or Fisher's exact test when appropriate.

†Based on ICD codes of index admission.

‡Based on ICD codes in the preceding 7 years.

§Based on discharge and outpatient prescriptions in the preceding 1 year.

ED, emergency department; ICD, International Classification of Diseases.

Table 2 Multivariable logistic regression analysis

Variable	Adjusted OR (95% CI)	p Value
<i>Patient demographics</i>		
Age	1.01 (1.01 to 1.02)	<0.001
Gender (male)	1.08 (1.00 to 1.17)	0.042
Ethnicity		
Others	Reference	
Chinese	1.08 (0.88 to 1.33)	0.486
Indian	1.26 (0.99 to 1.59)	0.057
Malay	1.21 (0.97 to 1.53)	0.1
<i>Social determinants of health</i>		
Required financial assistance using Medifund	1.64 (1.27 to 2.11)	<0.001
Stayed in a subsidised ward during index admission	1.65 (1.47 to 1.86)	<0.001
<i>Past healthcare use</i>		
ED visits (6 month before index admission)	1.11 (1.06 to 1.16)	<0.001
Specialist clinic visits (1 year before index admission)	1.01 (1.00 to 1.02)	<0.001
Hospital admissions (1 year before index admission)	1.4 (1.35 to 1.46)	<0.001
<i>Markers of acute illness burden</i>		
Index admission was urgent	1.05 (0.95 to 1.16)	0.365
Required second line antibiotics during index admission	0.97 (0.82 to 1.15)	0.749
Required inpatient dialysis during index admission	2.33 (1.98 to 2.74)	<0.001
Required intravenous furosemide 40 mg and above during index admission	1.49 (1.31 to 1.70)	<0.001
Required isolation during index admission	1.16 (0.85 to 1.59)	0.342
Length of stay of index admission	1 (0.99 to 1.00)	0.862
Charlson Comorbidity Index*	1.33 (1.31 to 1.35)	<0.001
<i>Medical comorbidities†</i>		
Stroke	0.94 (0.69 to 1.27)	0.699
Metastatic disease	1.02 (0.88 to 1.18)	0.788
Non-metastatic malignancy	1.08 (0.88 to 1.32)	0.467
Peripheral vascular disease	0.97 (0.84 to 1.12)	0.681
Heart failure or fluid overload	1.16 (0.93 to 1.44)	0.173
Pressure ulcer	1.17 (1.00 to 1.37)	0.05
Thromboembolism	0.93 (0.83 to 1.05)	0.253
Spine fracture	1.08 (0.85 to 1.36)	0.513
Coronary heart disease or myocardial infarction	1.07 (0.94 to 1.22)	0.331
Hip fracture	0.98 (0.74 to 1.27)	0.857
Atrial fibrillation	1.05 (0.90 to 1.22)	0.54
Epilepsy	0.83 (0.56 to 1.19)	0.317
Parkinsonism	1.32 (0.98 to 1.76)	0.06
Anxiety	0.74 (0.49 to 1.09)	0.142
Bipolar disorder	1.52 (0.79 to 2.80)	0.189
Collagen vascular disease	0.73 (0.52 to 1.01)	0.06
Dementia	0.95 (0.77 to 1.17)	0.637
Hypothyroidism	1.07 (0.83 to 1.38)	0.589
Chronic kidney disease, stages 1–4	0.94 (0.80 to 1.11)	0.479
Chronic obstructive pulmonary disease	1.14 (0.90 to 1.43)	0.284
Osteoarthritis	1.03 (0.89 to 1.18)	0.724
Benign prostatic hypertrophy	0.95 (0.79 to 1.13)	0.553
Asthma	1.1 (0.88 to 1.37)	0.401
Hyperlipidaemia	1.01 (0.90 to 1.14)	0.864
Hypertension	0.97 (0.87 to 1.10)	0.671
Chronic kidney disease stage 5 or end-stage renal failure	0.98 (0.82 to 1.17)	0.839
Diabetes	1.02 (0.91 to 1.14)	0.728
History of alcoholism	1.14 (0.87 to 1.48)	0.326
Treatment with antidepressants‡	2.51 (2.26 to 2.78)	<0.001

*Based on ICD codes of index admission.

†Based on ICD codes in the preceding 7 years.

‡Based on discharge and outpatient prescriptions in the preceding 1 year.
ED, emergency department; ICD, International Classification of Diseases.

Based on 10-fold cross-validation, the final logistic regression model had high discriminative ability (area under the curve (AUC) 0.84, 95% CI 0.83 to 0.84; figure 1). The optimal cut-off for the regression model was 0.235, which achieved a sensitivity of 73.9% (95% CI 72.6% to 75.2%) and a specificity of 79.1% (78.5% to 79.6%). By choosing three different cut-offs, we stratified the patients into four groups, namely high risk, moderate risk, low risk and minimal risk. Table 3 shows the sensitivity, specificity, PPV and NPV according to these cut-offs. Table 4 elaborates the number of patients stratified into four risk groups as well as the corresponding FA ratios. For example, a cut-off score of 0.442 predicted 2525 patients as FAs while 1455 (57.6%) patients were true positives.

DISCUSSION

In this retrospective cohort study of 25 244 eligible adult patients, we identified 11 predictors associated with frequent admission risk after hospital discharge in Singapore. In deriving these predictors, we have intentionally selected variables that are readily available in the EHR. This will facilitate the development of a practical predictive scoring tool that can be easily derived from readily available data. This in turn will enable automation and deployment as online spreadsheets or smartphone/tablet application to facilitate clinical use. Our study intended to identify predictors of frequent hospital admission risk that has the potential to be developed into a parsimonious risk score as future work. A risk score can function as a rapid risk stratification tool to identify patients at high risk for frequent admissions among the many admissions to our health system each day. This risk stratification must then be followed by an assessment of the high-risk patient to identify potentially

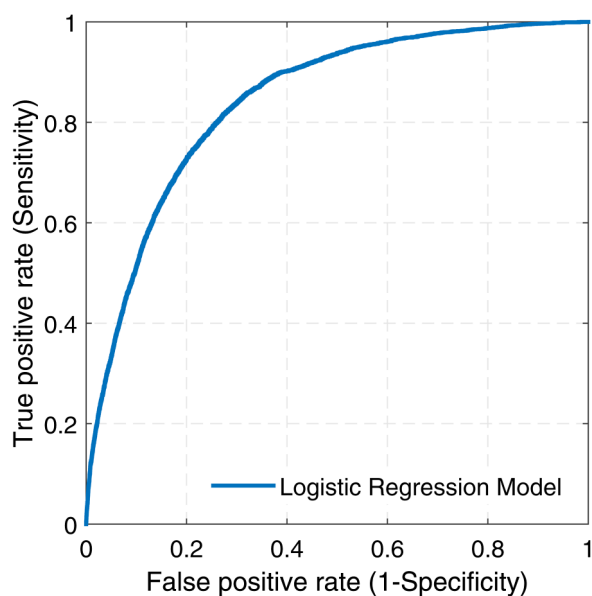


Figure 1 Area under the curve of final logistic regression model.

modifiable risk factors for intervention. Similarly, the assessment will identify patients who are going to be in rapid decline with unavoidable readmissions but can benefit from appropriate interventions that are aligned with the patients' goals of care.

Owing to the large sample size of our study, most p values in univariable analysis were very small and we conducted multivariable analysis to select significant predictors. We started with 44 variables, and we found that 11 factors explained the majority of risk associated with frequent admissions. This supports the growing evidence that simplified prediction models using a few critical predictors can achieve a high level of accuracy in its predictions. While no studies have specifically evaluated predictors for FA risk, our findings that increasing age, prior hospital usage and increased FA risk were consistent with the existing literature on 30-day, 90-day and 12-month readmission risk.^{6 13} In the final multivariable logistic regression model, chronic disease comorbidities were not independently associated with FA risk. A likely explanation is that the impact of comorbidities is confounded by usage in the past 1 year. Another explanation is that a patient's past comorbidities may have lesser significance on his current risk than recent events. We found that the effort needed to identify and retrieve individual diseases listed in the Charlson and Elixhauser comorbidity measures is not efficient in determining FA risk.

Contrary to expectations, metastatic solid tumours and tumours without metastases did not increase FA risk, in contrast to the published literature on 30-day readmission risk.^{21 22} Study population differences could have accounted for the difference. Another possible explanation could be that these patients may be in remission or had received appropriate management such as advance care planning and did not become FA. It is also possible that these patients passed away after the index hospital discharge and did not accumulate three or more hospital readmissions in the following 12 months to become a FA. Interestingly, having a pressure ulcer trended towards significance for increasing FA risk. A pressure ulcer may be a proxy marker for poor functional status and the quality of nursing care for the patient. Patients who develop a pressure ulcer tend to be bedbound and have a high level of dependency on caregivers for their activities of daily living, including regular turning in bed to prevent the development of pressure ulcers. Functional status is a known risk factor for 30-day readmission.²³ Caregiver availability is likely to be an important predictor and its inclusion is likely to improve the model performance in future studies.

We explored the use of proxy markers of severity of the acute illness episode as a predictor of FA risk. To the best of our knowledge, no previous studies have evaluated the impact of high-risk medications used during the hospitalisation although there was a study that focused on high-risk medications at discharge.²² Intravenous furosemide is used when prompt and

Table 3 Prediction results by the multivariable logistic regression model using ROC analysis

Cut-off	Sensitivity (95% CI)	Specificity (95% CI)	Positive predictive value (95% CI)	Negative predictive value (95% CI)
0.442	33.8% (32.4% to 35.2%)	94.9% (94.6% to 95.2%)	57.6% (55.6% to 59.5%)	87.4% (87.0% to 87.8%)
0.235	73.9% (72.6% to 75.2%)	79.1% (78.5% to 79.6%)	42.2% (41.1% to 43.3%)	93.6% (93.3% to 94.0%)
0.060	94.6% (93.9% to 95.3%)	47.3% (46.6% to 47.9%)	27.0% (26.3% to 27.7%)	97.7% (97.4% to 98.0%)

ROC, receiver operating characteristic.

Table 4 Patient stratification in four different risk groups

Risk group	Number of patients	Number of frequent admitters	Frequent admitter ratio (%)
High risk (0.442, 1)	2525	1455	57.6
Moderate risk (0.235, 0.442)	5048	1741	34.5
Low risk (0.06, 0.235)	7573	894	11.8
Minimal risk (0, 0.06)	10 098	232	2.3
Total	25 244	4322	17.1

effective diuresis is required and its use is an indicator of a state of severe fluid overload which required urgent reversal. Second-line antibiotics are used to treat severe hospital-acquired infections.²⁴ Its use during hospitalisation indicates the presence of a serious infection. Similarly, patients with end-stage renal failure (ESRF) requiring dialysis suggests that they are undergoing active treatment and therefore at higher risk for FA. We found that the odds of becoming FA in the following 12 months for patients with ESRF requiring inpatient dialysis and patients who were treated with intravenous furosemide 40 mg and above were 2.33 times and 1.49 times higher than the odds of patients who did not require dialysis and were not treated with the medication, respectively. It is possible that undergoing inpatient dialysis and intravenous furosemide better reflect current illness severity and FA risk compared with pre-existing comorbidity burden. CCI is a validated measure of 1 year mortality¹⁷ and we adapted its utility to weight the index admission diagnoses as a composite measure of current illness burden. We found that every increase of one point on the CCI score was associated with 33% increased risk of becoming a FA. It would be interesting to evaluate if these markers of acute illness severity also increase the risk of 30-day readmission.

In Kansagara's review of readmission risk prediction models, only 6 out of the 26 unique models used 12 months as the outcome period of interest,^{19 25–29} while the majority used a shorter time interval of 30–90 days.⁶ The current preoccupation with 30-day readmission is only of value if it determines patients who return to hospital because of bad care, and not those who would otherwise need frequent hospital admissions that our model strives to identify. Five were on general medical inpatients while one focused on patients with congestive heart failure. There are significant differences worth highlighting between these models and ours. First, five of the models predicted readmission in

the 12 months following index discharge making it difficult to compare their models with ours. The remaining model by Bottle and colleagues predicted 'high-impact users' who will have at least two or more emergency hospital admissions in the following 12 months, but the model was developed from administrative data. The model performance in terms of discriminatory ability was also poor, with five of the models having an AUC of <0.7. In comparison, our final regression model had high discriminative ability with an AUC of 0.83.

There is potential to explore other variables of interest in the future. These include functional status, health literacy, markers of social instability and caregiver availability.^{6 30} At the moment, these data are not routinely collected in most health systems although there is an intention to do so as part of a population database in our RHS. In the interim, we have focused on identifying predictors that can be easily retrieved from a patient's medical records or can potentially be automated through the EHR. Therefore, we have ignored variables that require additional collection by healthcare workers. Taha *et al*³¹ explored polypharmacy and problem medications such as anticoagulants and opioids on discharge. While these are available in our EHR system, the discharge prescription is among the finalised documents given to a patient on discharge. Delays in obtaining this information would have limited its usefulness in deriving a risk score on admission for the case managers and clinicians.

Limitations

We believe that our study had revealed many novel predictors that can be used in developing more effective and practical predictive models. However, there are limitations to our study. First, variables in our data set are restricted to those routinely collected in the EHR and administrative databases. As such, information about the social economic status of the patients is not detailed and we used tested access to financial support as a proxy.

Functional status, caregiver availability and degree of social support could have improved the performance of our model but we did not include these predictors as such data are not routinely collected in our healthcare settings and our intention was to develop an easy-to-use and generalisable risk score. Second, due to the retrospective nature of the study, we were unable to prove a causal association between the predictor variables and frequent hospital admissions. After our predictive model has identified patients at high risk for frequent hospital admissions, intervention programmes would have to identify potentially modifiable risk factors. Finally, we did not exclude patients who might have deceased after index hospital discharge. We felt that that would have biased the prediction model as these patients could have had frequent hospital admissions before death.

CONCLUSION

In this retrospective cohort study, we investigated the impact of comorbidities, acute illness burden and social determinants of health on the risk of frequent hospital admissions and evaluated the performance of these variables on predictive modelling. We identified several strong predictors such as treatment with antidepressants and requiring inpatient dialysis during index admission. Our final regression model had high discriminative ability for the prediction of frequent hospital admission risk with an AUC of 0.84. We intend to follow-up with implementation of such a predictive tool in our RSH; however, a carefully designed model and external validation are needed to ensure reproducibility.

Author affiliations

¹Department of Family Medicine & Continuing Care, Singapore General Hospital, Singapore, Singapore

²Family Medicine Program, Duke-NUS Medical School, Singapore, Singapore

³Singapore Health Services, Health Services Research Centre, Singapore, Singapore

⁴Centre for Quantitative Medicine, Duke-NUS Medical School, Singapore, Singapore

⁵Department of Emergency Medicine, Singapore General Hospital, Singapore, Singapore

⁶Integrated Health Information Systems, Singapore, Singapore

⁷Department of Rheumatology and Immunology, Singapore General Hospital, Singapore, Singapore

⁸Health Services and Systems Research, Duke-NUS Medical School, Singapore, Singapore

Contributors LLL and NL conceived and designed the study. LLL, NL and SW performed the study and analysed the data, including modelling and cross validation. LLL, NL, SW and KHL interpreted the results. LLL and NL wrote the first draft of the paper and all authors critically revised the paper and gave final approval for publication.

Funding This research received funding from Singapore Ministry of Health (MOH) Health Services Research New Investigator Grant HSRNIG14nov002 and SingHealth Foundation Health Services Research (Aging) Startup Grant SHF/HSRAg004/2015.

Competing interests None declared.

Ethics approval This study was approved by Singapore Health Services (SingHealth) Centralized Institutional Review Board CIRB 2015/2696 with a waiver of informed consent.

Provenance and peer review Not commissioned; externally peer reviewed.

Data sharing statement Details of the variables and derived predictive model are available from the corresponding authors at low.lian.leng@singhealth.com.sg or liu.nan@singhealth.com.sg.

Open Access This is an Open Access article distributed in accordance with the Creative Commons Attribution Non Commercial (CC BY-NC 4.0) license, which permits others to distribute, remix, adapt, build upon this work non-commercially, and license their derivative works on different terms, provided the original work is properly cited and the use is non-commercial. See: <http://creativecommons.org/licenses/by-nc/4.0/>

REFERENCES

- Jencks SF, Williams MV, Coleman EA. Rehospitalizations among patients in the Medicare fee-for-service program. *N Engl J Med* 2009;360:1418–28.
- Longman JM, Rolfe MI, Passey MD, *et al*. Frequent hospital admission of older people with chronic disease: a cross-sectional survey with telephone follow-up and data linkage. *BMC Health Serv Res* 2012;12:373.
- Kripalani S, Theobald CN, Anctil B, *et al*. Reducing hospital readmission rates: current strategies and future directions. *Annu Rev Med* 2014;65:471–85.
- Centers for Medicare and Medicaid Services. Readmissions Reduction Program. <https://www.cms.gov/medicare/medicare-fee-for-service-payment/acuteinpatientpps/readmissions-reduction-program.html>. (accessed 28 Mar 2016).
- Robst J. Developing models to predict persistent high-cost cases in Florida Medicaid. *Popul Health Manag* 2015;18:467–76.
- Kansagara D, Englander H, Salanitro A, *et al*. Risk prediction models for hospital readmission: a systematic review. *JAMA* 2011;306:1688–98.
- Ministry of Health Singapore. Health Facts Singapore Healthcare Institution Statistics 2012. https://www.moh.gov.sg/content/moh_web/home/statistics/healthcare_institutionstatistics.html (accessed 28 Mar 2016).
- Goetghebeur MM, Forrest S, Hay JW. Understanding the underlying drivers of inpatient cost growth: a literature review. *Am J Manag Care* 2003;9 Spec No 1:Sp3–12.
- Saxena N, You AX, Zhu Z, *et al*. Singapore's regional health systems—a data-driven perspective on frequent admittees and cross utilization of healthcare services in three systems. *Int J Health Plann Mgmt* 2015. Published Online First: 29 Jun 2015. doi:10.1002/hpm.2300
- Ronksley PE, McKay JA, Kobewka DM, *et al*. Patterns of health care use in a high-cost inpatient population in Ottawa, Ontario: a retrospective observational study. *CMAJ Open* 2015;3:E111–18.
- Lai L. Navigators steer patients through health-care system. The Straits Times. 2014. <http://www.straitstimes.com/singapore/health/singhealth-navigators-steer-patients-through-health-care-system> (accessed 1 Apr 2016).
- Low LL, Vasanwala FF, Ng LB, *et al*. Effectiveness of a transitional home care program in reducing acute hospital utilization: a quasi-experimental study. *BMC Health Serv Res* 2015;15:100.
- Low LL, Lee KH, Hock Ong ME, *et al*. Predicting 30-day readmissions: performance of the lace index compared with a regression model among general medicine patients in Singapore. *Biomed Res Int* 2015;2015:169870.
- Shadmi E, Flaks-Manov N, Hoshen M, *et al*. Predicting 30-day readmissions with preadmission electronic health record data. *Med Care* 2015;53:283–9.
- Tan SY, Low LL, Yang Y, *et al*. Applicability of a previously validated readmission predictive index in medical patients in Singapore: a retrospective study. *BMC Health Serv Res* 2013;13:366.
- Charlson ME, Pompei P, Ales KL, *et al*. A new method of classifying prognostic comorbidity in longitudinal studies: development and validation. *J Chronic Dis* 1987;40:373–83.
- Deyo RA, Cherkin DC, Ciol MA. Adapting a clinical comorbidity index for use with ICD-9-CM administrative databases. *J Clin Epidemiol* 1992;45:613–19.
- Quan H, Sundararajan V, Halfon P, *et al*. Coding algorithms for defining comorbidities in ICD-9-cm and ICD-10 administrative data. *Med Care* 2005;43:1130–9.
- Howell S, Coory M, Martin J, *et al*. Using routine inpatient data to identify patients at risk of hospital readmission. *BMC Health Serv Res* 2009;9:96.
- Ministry of Health Singapore. Medifund 2015. https://www.moh.gov.sg/content/moh_web/home/costs_and_financing/schemes_subsidies/Medifund.html (accessed 1 Apr 2016).

21. Tong L, Erdmann C, Daldalian M, *et al.* Comparison of predictive modeling approaches for 30-day all-cause non-elective readmission risk. *BMC Med Res Methodol* 2016;16:1–8.
22. Allaudeen N, Vidyarthi A, Maselli J, *et al.* Redefining readmission risk factors for general medicine patients. *J Hosp Med* 2011;6:54–60.
23. Greysen SR, Stijacic Cenzer I, Auerbach AD, *et al.* Functional impairment and hospital readmission in Medicare seniors. *JAMA Intern Med* 2015;175:559–65.
24. Peleg AY, Hooper DC. Hospital-acquired infections due to Gram-negative bacteria. *N Engl J Med* 2010;362:1804–13.
25. Billings J, Mijanovich T. Improving the management of care for high-cost Medicaid patients. *Health Aff (Millwood)* 2007;26:1643–54.
26. Billings J, Dixon J, Mijanovich T, *et al.* Case finding for patients at risk of readmission to hospital: development of algorithm to identify high risk patients. *BMJ* 2006;333:327.
27. Morrissey EFR, McElnay JC, Scott M, *et al.* Influence of drugs, demographics and medical history on hospital readmission of elderly patients. *Clin Drug Investig* 2012;23:119–28.
28. Bottle A, Aylin P, Majeed A. Identifying patients at high risk of emergency hospital admissions: a logistic regression analysis. *J R Soc Med* 2006;99:406–14.
29. Philbin EF, DiSalvo TG. Prediction of hospital readmission for heart failure: development of a simple risk score based on administrative data. *J Am Coll Cardiol* 1999;33:1560–6.
30. Amarasingham R, Moore BJ, Tabak YP, *et al.* An automated model to identify heart failure patients at risk for 30-day readmission or death using electronic medical record data. *Med Care* 2010;48:981–8.
31. Taha M, Pal A, Mahnken JD, *et al.* Derivation and validation of a formula to estimate risk for 30-day readmission in medical patients. *Int J Qual Health Care* 2014;26:271–7.