abstract

Persistent Disparity: Socioeconomic Deprivation and Cancer Outcomes in Patients Treated in Clinical Trials

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PURPOSE Patients with cancer living in socioeconomically disadvantaged areas have worse cancer outcomes. The association between socioeconomic deprivation and outcomes among patients with cancer participating in clinical trials has not been systematically examined.

METHODS We examined survival outcomes for patients enrolled in phase III and large phase II clinical trials for major cancers conducted by the SWOG Cancer Research Network from 1985 to 2012. Socioeconomic deprivation was measured using trial participants' residential zip codes linked to the Area Deprivation Index (ADI). Five-year overall survival, progression-free survival, and cancer-specific survival were examined using Cox regression frailty models, adjusting for age, sex, and race, and separately for insurance status, prognostic risk, and rural or urban residency.

RESULTS We examined 41,109 patients from 55 trials comprising 24 cancer histology and stage-specific cohorts. Compared with trial participants in the most affluent areas (ADI, 0%-20%), trial participants from areas with the highest socioeconomic deprivation (ADI, 80%-100%) had worse overall (hazard ratio [HR] = 1.28, 95% CI, 1.20 to 1.37, P < .001), progression-free (HR = 1.20, 95% CI, 1.13 to 1.28, P < .001), and cancer-specific survival (HR = 1.27, 95% CI, 1.18 to 1.37, P < .001). The results were similar after adjusting for insurance status, prognostic risk, and rural or urban residency. There was a continuous increase in risk of all outcomes as the ADI quintile increased.

CONCLUSION In patients with cancer with access to protocol-directed care in clinical trials, high area-level socioeconomic deprivation was associated with worse survival. Future research should examine whether the etiology of this residual disparity is related to reduced access to supportive care or postprotocol therapy and/or to differences in health status not reflected by protocol selection criteria.

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INTRODUCTION

Poverty is detrimental to health. A key underlying reason is that individuals from socioeconomically disadvantaged areas have limited access to vital healthcare resources, which may become a chronic risk factor for worse disease outcomes.^{1,2} Indeed, the association between low socioeconomic status and poor health outcomes has been well-documented in the literature for many diseases.^{3,4}

Patients with cancer from socioeconomically deprived areas have limited access to screening and treatment services and tend to have more advanced disease at presentation.⁵⁻⁸ Studies using cancer population data have shown that these patients have worse cancer outcomes.⁹⁻¹¹ Evidence suggests that these disparities persist even after accounting for individual-level socioeconomic variables such as insurance status.^{12,13} An important and unresolved question is whether these disparate outcomes remain after accounting for access to quality cancer care. We used data from patients participating in cancer treatment clinical trials to systematically examine whether patients from poor areas experienced worse outcomes. Patients in trials have uniform access to protocol-guided care and are uniformly staged to limit potential differences in underlying health status. These advantages enable investigators to account for the confounding influences of inconsistent pretreatment evaluation, care, and post-treatment surveillance. A better understanding of whether patients from socioeconomically deprived areas experience worse cancer outcomes can provide valuable insight into whether policies to improve outcomes should address not just individuals but also communities.

METHODS

Patients

In this retrospective observational study of pooled clinical trial participants, we used data from the SWOG

ASSOCIATED CONTENT

Data Supplement

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CONTEXT

Key Objective

Poverty is detrimental to health. A key underlying reason is that individuals from socioeconomically deprived areas have limited access to healthcare resources. In cancer, an unresolved question is whether socioeconomic disparities in outcomes remain after accounting for access to quality cancer care. We examined the relationship between socio-economic deprivation and outcomes for 41,109 patients enrolled in clinical trials, who have uniform access to protocol-directed treatment.

Knowledge Generated

Compared with trial patients in the most affluent areas, patients from the highest socioeconomically deprived areas had a 28% increased risk of death (P < .001). The findings were consistently observed within major patient groups.

Relevance

These findings suggest that initial access to quality cancer care as represented by treatment in a clinical trial is insufficient to eliminate the disparate outcomes related to socioeconomic deprivation. Policies to mitigate socioeconomic differences in cancer outcomes should emphasize access to cancer care services beyond initial therapy.

Cancer Research Network, a member of the National Cancer Institute's funded National Clinical Trials Network and Community Oncology Research Program. We systematically reviewed the SWOG trial database to identify phase III trials or large phase II trials completed beginning in 1985 (when zip code data began to be routinely collected) for which the primary analysis was previously published. Data were pooled across trials of similar histology and stage to increase power to identify potential differences in outcomes between patients from areas with different levels of socioeconomic deprivation. A cancer-specific cohort was included in the analysis if—based on the observed sample size and number of events—the power to detect a hazard ratio (HR) of 1.5 or greater between groups was 80% or greater (Data Supplement, online only).¹⁴

Each trial included in this analysis was previously approved by an institutional review board; informed consent was previously obtained from all patients for each study included. This research was approved by the Fred Hutchinson Cancer Research Center's Institutional Review Board.

Area Deprivation Index and Covariates

We measured socioeconomic deprivation using trial participants' residence zip codes linked to the Area Deprivation Index (ADI) (ADI; University of Wisconsin School of Medicine and Public Health, 2019; 2013 version). The ADI is a comprehensive index composed of 17 indicators reflecting a diverse set of socioeconomic variables including neighborhoodlevel measures of education (two variables), employment (two), housing quality (five), and poverty (eight; Data Supplement).¹⁵ The ADI has been rigorously tested, is inclusive of all US neighborhoods, and is regularly updated. The ADI measure is scored from 0 to 100, with higher scores indicating greater socioeconomic deprivation. Based on recent studies, we split the ADI measure into quintiles; thus, for analysis purposes, the most deprived patients were defined as those in the highest ADI quintile (81%-100%) and the most affluent were patients in the lowest ADI quintile (0%-20%).¹⁶⁻¹⁹

Adjustment covariates included the following demographic variables: age (< 30, 30-34, ..., 75-79, \geq 80), race (Black *v* Others, by self-report), and sex. Patient-level insurance status was available for a subset of patients, categorized as using Medicaid, Medicare plus Medicaid, or self-pay versus using private insurance, Medicare, or military and/or Veteran's Administration insurance. To account for potential differences in clinical risk, for each cancer cohort, we identified key clinical risk factors, typically reflected by their inclusion as stratification variables in the trials (Data Supplement). For each patient, we summed the number of adverse clinical risk factors, creating a composite prognostic score, standardized to a 0-100 scale.¹⁴ Additionally, rural or urban residency was coded according to the US Department of Agriculture rural-urban continuum codes (1-3 = urban; 4-9 = rural).²⁰

Outcomes

The primary end point was overall survival, measured as days from study registration to death by any cause or, for patients still alive, as days to last contact (censored). Secondarily, we examined progression-free survival, defined as time from trial registration to the date of death by any cause, evidence of protocol-defined relapse or progressive disease, or date of last contact for those alive and progression free at last contact (censored). We also examined cancer-specific survival, defined as time from study registration to date of cancer-specific death, death by another cause (censored), or last contact for those alive at last contact (censored). Detailed cause-of-death information was available for 28% of patient deaths; for remaining patients, we considered any death preceded by documented relapse or progression a cancer-specific death.

Statistical Analysis

We evaluated outcomes by quintile of the ADI score from highest quintile (most deprived) to lowest quintile (most affluent). We limited analyses to the first 5 years after registration to focus on cancer-related and treatment-related survival. Patients with last contact date or death date > 5 years postregistration were censored at 5 years of follow-up. For each cancer-specific cohort, we used multivariable Cox regression to assess the association of ADI quintile score and survival outcomes, with covariate adjustment.²¹ We used multivariable Cox regression frailty models to generate overall estimates of the association of ADI quintile score and survival outcomes aggregated across cancer cohorts, including each cancer cohort as a random effect to account for differences in period of enrollment, cancer-specific staging, and prognosis.^{21,22} Multiple models were conducted using a nested approach, to evaluate the extent to which additional adjustment for important covariates potentially modified the results. The primary model included covariate adjustments for age, sex, and race. Subsequent models also adjusted, separately, for insurance status (including an indicator variable for whether insurance status was available, yes v no), the prognostic score, rural or urban residency, and all three covariates.

We used Wald chi-square statistics derived from multivariable Cox regression models for statistical testing. Tests for statistical significance were two-sided, $\alpha = .05$. We examined data through April 24, 2020.

Additional Analyses

Analyses were also conducted in the subsets of patients defined by age (< 65 $v \ge$ 65 years), sex, race (White *v* non-White), insurance status (Medicaid or no insurance *v*

	ADI Quintile (Q)												
Patient Characteristic		ADI Q1: 0%- 20% (n = 4,727)		ADI Q2: 21%- 40% (n = 8,775)		ADI Q3: 41%- 60% (n = 11,878)		ADI Q4: 61%- 80% (n = 11,935)		ADI Q5: 81%- 100% (n = 3,794)		Total (N = 41,109)	
		%	No.	%	No.	%	No.	%	No.	%	No.	%	
Age													
\geq 65 years	1,194	25.3	2,423	27.6	3,494	29.4	3,635	30.5	1,162	30.6	11,908	29.0	
< 65 years	3,533	74.7	6,352	72.4	8,384	70.6	8,300	69.5	2,632	69.4	29,201	71.0	
Sex													
Female	3,208	67.9	5,500	62.7	7,300	61.5	7,130	59.7	2,135	56.3	25,273	61.5	
Male	1,519	32.1	3,275	37.3	4,578	38.5	4,805	40.3	1,659	43.7	15,836	38.5	
Race or ethnicity													
Any minority	811	17.2	1,273	14.5	1,447	12.2	1,902	15.9	1,521	40.1	6,954	16.9	
Black	204	4.3	490	5.6	829	7.0	1,454	12.2	1,246	32.8	4,223	10.3	
Hispanic	172	3.6	405	4.6	405	3.4	344	2.9	235	6.2	1,561	3.8	
Not minority	3,916	82.8	7,502	85.5	10,431	87.8	10,033	84.1	2,273	59.9	34,155	83.1	
Insurance ^a													
Medicaid, Medicaid and Medicare, and self-pay	247	6.5	532	7.9	771	8.9	1,022	12.3	549	22.8	3,121	10.4	
Medicare, private, military, and VA	3,543	93.5	6,242	92.1	7,880	91.1	7,286	87.7	1,855	77.2	26,806	89.6	
Unknown	937		2,001		3,227		3,627		1,390		11,182		
Cancer histology													
Breast	2,481	52.5	4,048	46.1	5,080	42.8	4,574	38.3	1,327	35.0	17,510	42.6	
Lung	189	4.0	443	5.0	789	6.6	985	8.3	325	8.6	2,731	6.6	
Colorectal	412	8.7	789	9.0	1,313	11.1	1,400	11.7	435	11.5	4,349	10.6	
Prostate	597	12.6	1,120	12.8	1,365	11.5	1,336	11.2	507	13.4	4,925	12.0	
Others	1,048	22.2	2,375	27.1	3,331	28.0	3,640	30.5	1,200	31.6	11,594	28.2	
Stage													
Advanced	1,500	31.7	3,075	35.0	4,369	36.8	4,669	39.1	1,513	39.9	15,126	36.8	
Not advanced	3,227	68.3	5,700	65.0	7,509	63.2	7,266	60.9	2,281	60.1	25,983	63.2	

TABLE 1. Patient Characteristics by ADI Quintile

Abbreviations: ADI, Area Deprivation Index; VA, Veterans Affairs.

^aPercentages calculated among patients with known data.

Medicare or private), disease stage (advanced v not advanced), time period of enrollment in approximate 10-year intervals (1985-1994 v 1995-2004 v 2005-2012), and geographic location (rural v urban). Furthermore, for each outcome, we examined whether the findings were sensitive to the influence of individual cancer cohorts by iteratively excluding each individual cohort and regenerating the aggregate frailty model regression estimates.

RESULTS

In total, 41,109 patients from 55 trials comprising 24 cancerspecific cohorts were examined (Table 1; Data Supplement). Overall, 29.0% of patients were \geq 65 years, 61.5% were female, 10.3% were Black, and 10.4% (among those with known data) had no insurance or were enrolled in Medicaid. The four most common cancers (breast, colorectal, lung, and prostate) comprised 71.8% of the sample, and 36.8% had advanced-stage disease. There were clear trends toward higher rates of older patients, males, Black patients, and patients with Medicaid or no insurance as the level of deprivation increased (Table 1). One third (32.8%) of patients from the most deprived areas were Black.

Association of Socioeconomic Deprivation and Survival Outcomes

Despite cohort-level heterogeneity, a trend toward worse overall survival for patients from the most deprived areas was apparent, even after accounting for the demographic covariates (Data Supplement). Among the 24 cancer cohorts we analyzed, the risk of a poorer outcome was statistically significantly greater for patients in the most deprived quintiles in eight instances for overall survival, in seven instances for progression-free survival, and in eight instances for cancer-specific survival. In contrast, in no instance was the risk of a poorer outcome statistically significantly greater for patients in the most affluent group.

These trends are reflected in aggregate models. In a multivariate Cox regression frailty model with the cancer cohorts treated as random effects, trial participants from the highest deprivation areas had worse overall survival on average (HR = 1.28, 95% CI, 1.20 to 1.37, P < .001) compared with patients from the most affluent areas (Fig 1; Table 2). The findings were of similar magnitude and direction for both progression-free survival and cancer-specific survival.

The addition of other socioeconomic, clinical, and geographic variables to the regression models only modestly attenuated the association (Table 2). For instance, the aggregate estimate of the increased risk of death comparing patients in the most deprived versus the most affluent areas of 28% (HR = 1.28, 95% Cl, 1.20 to 1.37, P < .001) was reduced to 22% (HR = 1.22, 95% Cl, 1.14 to 1.31, P < .001) when insurance status, baseline clinical risk, and rural or urban residency were included in the model.

There was a clear trend toward consistently increasing risk of death as the level of deprivation increased (Fig 2). This trend was also apparent for progression-free and cancer-specific survival and for both the primary and secondary models.

Additional Analyses

Compared with patients in the most affluent areas, patients in the most deprived areas had statistically significantly greater risk of overall, progression-free, and cancer-specific survival among those who were < 65 years, \geq 65 years, female, male, White, and non-White, using Medicare or private insurance, with adjuvant disease, with advanced



FIG 1. Forest plot of hazard ratio comparing highest deprivation quintile versus lowest deprivation quintile. The results are from primary modeling approach, adjusting for age, race, and sex in frailty model with cancer type as random effect. The boxes in the forest plot represent the hazard ratios, and the horizontal lines are the 95% CIs; the vertical line is the line of equal hazard. HR, hazard ratio.

TABLE 2. Changes in HR Comparing Lowest Versus Highest Area Deprivation Index Quintiles

End Point and Statistic	Primary Model ^a	Secondary Model 1: Add Insurance Status	Secondary Model 2: Add Baseline Clinical Risk	Secondary Model 3: Add Insurance Status and Baseline Clinical Risk	Secondary Model 4: Add Rural or Urban Residency	Secondary Model 5: Add Insurance status and Baseline Clinical Risk and Rural or Urban Residency
Overall survival						
HR	1.28	1.24	1.26	1.22	1.28	1.22
95% CI	1.20 to 1.37	1.16 to 1.33	1.17 to 1.35	1.14 to 1.31	1.19 to 1.37	1.14 to 1.31
Р	< .001	< .001	< .001	< .001	< .001	< .001
Reduction in HR ^b	—	14.3%	7.1%	21.4%	0%	21.4%
Progression-free survival						
HR	1.20	1.18	1.18	1.16	1.20	1.17
95% CI	1.13 to 1.28	1.11 to 1.25	1.11 to 1.26	1.09 to 1.24	1.13 to 1.28	1.10 to 1.24
Р	< .001	< .001	< .001	< .001	< .001	< .001
Reduction in HR ^b	—	10.0%	10.0%	20.0%	0%	15.0%
Cancer-specific survival						
HR	1.27	1.24	1.24	1.21	1.27	1.21
95% CI	1.18 to 1.37	1.15 to 1.33	1.16 to 1.34	1.13 to 1.31	1.18 to 1.37	1.12 to 1.30
Р	< .001	< .001	< .001	< .001	< .001	< .001
Reduction in HR ^b		11.1%	11.1%	22.2%	0%	22.2%

Abbreviation: HR, hazard ratio.

^aAdjusting for age, race, and sex in frailty model with cancer type as random effect.

^bCalculated as, 1—([1-secondary model HR]/[1-primary model HR]).

disease, enrolled from 1985 to 1994, 1995 to 2004, or 2005 to 2012, and in rural or urban areas (Fig 3). In contrast, in the small set (10.4%) of patients with Medicaid or no insurance, there was no observed difference in risk of outcomes between patients from the most deprived compared with the most affluent areas.

The exclusion of individual cancer cohorts changed the overall hazard ratio from 1.260 (-7.1% change) to 1.311 (+11.1% change; Fig 4). Similarly, only modest changes were also observed for progression-free and cancer-specific survival. The findings were also largely insensitive to the exclusion of individual studies (Data Supplement).

DISCUSSION

The examination of the relationship between socioeconomic deprivation and cancer outcomes in clinical trial patients with cancer is crucial for understanding whether socioeconomic disparities persist even after accounting for access to quality cancer care. This comprehensive study combined data from more than 40,000 clinical trial patients. Participants were from a wide variety of cancer types and stages, were uniformly staged, prospectively treated, and followed under protocol-directed care. Patients from the most socioeconomically deprived areas were older and more likely to be male, Black, and to have Medicaid or no insurance. Even after accounting for these factors, patients from the most deprived areas had a greater risk of death. The findings were consistent for progression-free and cancer-specific survival, suggesting an association with both cancer and noncancer outcomes. Moreover, the risk of death increased consistently as the level of deprivation increased, irrespective of the strategy for covariate adjustment, representing a dose response to the increasing burden of deprivation. The findings were uniformly observed for all subgroups of patients, with the exception of those with Medicaid or no insurance. Thus, socioeconomic disparities in cancer outcomes persist for almost all patients after accounting for initial access to protocol-directed, quality cancer care in clinical trials.

This study complements and extends previous research linking area-level attributes and outcomes among patients with cancer.^{11,13,23-27} First, studies using cancer population data cannot account for access to care, an essential consideration when examining cancer health disparities. By using trial data, we implicitly accounted for initial access to uniform, high quality care in the design. Second, we relied on the recently developed ADI, a comprehensive socioeconomic deprivation index.¹⁵ One common limitation of clinical trial data is lack of patient-level socioeconomic information.

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FIG 2. Percentage increase in HR by end point, model, and ADI quintile. (A) Overall survival, (B) progression-free survival, and (C) cancer-specific survival. The results are shown by quintile, defined as Q2, ADI 21%-40%; Q3, ADI 41%-60%; Q4, ADI 61%-80%; and Q5, ADI-81%-100%, most disadvantaged. The *P* value below each bar represents the comparison for the given quintile of area level deprivation with the lowest area-level deprivation quintile (Q1, ADI 0%-20%, most affluent). Analyses were conducted using Cox regression frailty models; in each model, the cancer type (Data Supplement) was considered a random effect. The vertical lines indicate the 95% Cls. ADI, Area Deprivation Index; HR, hazard ratio; PM, Primary Model.

Instead, the ADI is an area- or community-level attribute, although it likely correlates with individual-level measures, so is useful in discriminating outcomes between different individuals in numerous settings.²⁸⁻³⁰ Third, detailed trial records allowed us to account for uniformly collected clinical risk variables that were unavailable in previous studies. Finally, we included a broad range of different cancers, reinforcing the generalizability of the results.

	Overall	Survival		Progression	-Free Surv	vival	Cancer-Specific Survival				
	Hazard Ratio (95% CI)		Р	Hazard Ratio (95%	CI)	Р	Hazard Ratio (95% (Р			
Age < 65 years	1.33 (1.22 to 1.45)	⊢ ∎-1	< .001	1.26 (1.16 to 1.36)	H -H	< .001	1.32 (1.20 to 1.45)	1 -1	< .001		
Age ≥ 65 years	1.21 (1.09 to 1.35)	⊢ ∎–⊣	< .001	1.12 (1.01 to 1.24)		.027	1.21 (1.07 to 1.35)	H---1	.002		
Female	1.38 (1.20 to 1.59)		< .001	1.37 (1.20 to 1.56)		< .001	1.46 (1.25 to 1.69)	⊢→	< .001		
Male	1.22 (1.09 to 1.37)	F=-1	< .001	1.11 (1.00 to 1.23)	⊢ ⊷i	.051	1.22 (1.08 to 1.38)	H -1	.001		
Non-White	1.28 (1.11 to 1.46)	⊢ ∎⊸1	< .001	1.21 (1.07 to 1.37)	н -н	.002	1.27 (1.10 to 1.47)		.001		
White	1.34 (1.24 to 1.45)	⊢ ∎-4	< .001	1.23 (1.14 to 1.32)	нн	< .001	1.32 (1.21 to 1.43)	H -H	< .001		
Medicaid or no insurance	1.03 (0.81 to 1.30) -		.810	1.08 (0.87 to 1.35) ⊢		.460	1.09 (0.84 to 1.40) 🛏		.512		
Medicare or private insurance	1.27 (1.16 to 1.39)	H=-1	< .001	1.16 (1.07 to 1.26)	H=4	< .001	1.22 (1.11 to 1.34)	H=4	< .001		
Adjuvant disease	1.42 (1.27 to 1.59)	1	< .001	1.28 (1.16 to 1.40)	H=1	< .001	1.41 (1.25 to 1.60)		< .001		
Advanced disease	1.20 (1.10 to 1.31)	H= -1	< .001	1.14 (1.05 to 1.24)	H=H	.001	1.20 (1.09 to 1.31)	H=-1	< .001		
1985–1994	1.17 (1.04 to 1.31)	H -4	.006	1.16 (1.05 to 1.29)	H=-1	.004	1.22 (1.08 to 1.38)	HHH	.002		
1995–2004	1.25 (1.12 to 1.40)	⊢ ∎−4	< .001	1.11 (1.00 to 1.23)	⊨I	.040	1.23 (1.09 to 1.38)	H--1	< .001		
2005–2012	1.21 (1.04 to 1.40)		.015	1.17 (1.02 to 1.33)		.022	1.16 (0.99 to 1.36)		.071		
Rural	1.63 (1.17 to 2.27)	\mapsto	.004	1.55 (1.15 to 2.09)	→	.004	1.69 (1.17 to 2.45)	\mapsto	.005		
Urban	1.22 (1.13 to 1.31)	H=-1	< .001	1.16 (1.09 to 1.25)	нн	< .001	1.21 (1.11 to 1.31)	нн	< .001		
					+			+ + + + +			
	0.8	1.2 1.6		0.8	1.2 1.0	6 	0.8	1.2 1.6	; 		

FIG 3. Forest plot of hazard ratio comparing highest deprivation quintile versus lowest deprivation quintile within patient groups. The results are from primary modeling approach, adjusting for age, race, and sex in frailty model with cancer type as random effect. The boxes in the forest plot represent the hazard ratios, and the horizontal lines are the 95% CIs; the vertical line is the line of equal hazard. Findings are statistically significant for all groups except patients with Medicaid or no insurance.

Previous studies have used census-based area-level poverty or education measures and showed highly variable increases (18%-164%) in risk of mortality for patients in the lowest socioeconomic categories.^{23,25,26} Only a few studies have applied a multiattribute index to examine this relationship. Singh and Jemal categorized counties according to 11 US Census-based indicators using registry data from the 1990s.¹¹ They found that patients with cancer from the most deprived areas had a 56% increased mortality risk. Hastert et al¹³ applied a socioeconomic index based on six indicators to a cohort of patients of age 50 to 76 years recruited from 2000 to 2002 and showed a 68% increased mortality risk for those in the most deprived areas. Taken together, these studies showed about a 60% increased risk of mortality, notably greater than the 28% increased risk in our study. If these studies accurately reflect the outsized risk of poor outcomes for patients with cancer living in the most socioeconomically deprived areas, the comparison suggests that access to protocol-guided care may mitigate half of the negative effects of living in socioeconomically deprived areas.

The magnitude of the socioeconomic disparity observed in our study is equivalent to the treatment effect for contemporary large treatment trials.^{31,32} Our finding is consistent with a previous study demonstrating the enormous potential reductions in cancer-related mortality because of the elimination of socioeconomic disparities.²⁵ The observation that socioeconomic disparities in cancer outcomes persist despite access to treatment is consistent with the idea that low socioeconomic status represents a risk factor for poor health outcomes because of poor access to resources (including medical, technical, and financial resources).¹ In this setting, the receipt of protocol-guided therapy may provide beneficial initial access to guidelinebased cancer care, but will not otherwise materially alter the tendency toward poor access to healthcare resources over the long term. Other research has suggested the potential effects of chronic stress exposure on health outcomes through mechanisms of altered immunological and neuroendocrine function, mechanisms which would endure despite access to guideline-based care.^{33,34} Additionally, living in a lower socioeconomic area is associated with more risky health behaviors, including smoking, obesity, physical inactivity, alcohol use, and poor diet.^{35,36} For vulnerable populations, a diagnosis of cancer may amplify high-risk behaviors.³⁷ Finally, although trial patients are uniformly staged and treated, residual differences in prognosis between patients from socioeconomically deprived and

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FIG 4. Change in overall frailty regression model results due to excluding individual cancer cohorts (eg, leave one out analysis). For overall survival, progression-free survival, and cancer-specific survival, the difference between the primary model adjusted hazard ratio and the adjusted hazard ratio excluding the specified cancer cohort is shown by the bars. The percentage difference is also shown. Adj, adjuvant; Adv, advanced; AML, acute myeloid leukemia; CR, castration-resistant; ER, estrogen-receptor; GIST, GI stromal tumor; NHL, non-Hodgkin's lymphoma; NSCLC, non–small-cell lung cancer; PR, progesterone receptor; SCLC, small-cell lung cancer.

affluent areas may be associated with differences in clinical outcomes.

These findings have policy implications for patients with cancer both in trials and outside trials. Oncology care resources are sparse in certain areas, which will adversely affect those lacking the financial resources to travel for care.³⁸ Thus, even patients able to access initial guidelinebased care might not have access to follow-up care, including supportive care and postprogression care. Telehealth models may be vital in such situations, providing the convenience that allows patients to more easily overcome the economic constraints of travel, having to take time off work, or childcare—issues that may be particularly burdensome for individuals in more deprived areas that lack an adequate oncology clinical care infrastructure.³⁹ One notable consequence of the recent COVID-19 pandemic has been the accelerated adoption of telehealth by many oncology care centers.⁴⁰ This development—if sustained—is likely to benefit both patients and physicians over the long term, especially those living and working in more deprived areas. More generally, these findings suggest that mechanisms that enable patients with cancer to maintain access to quality care—such as adequate health insurance—are required. Of note, we observed no gradient in clinical outcomes among patients with Medicaid or no insurance, suggesting that residence in affluent areas provides no added benefit for these patients. State-level Medicaid programs provide needed access to healthcare services for socioeconomically deprived individuals, but have historically varied widely in eligibility and the provision of services, with reduced and interrupted access related to adverse cancer outcomes.⁴¹ The Affordable Care Act not only provided more ready access to health insurance for all individuals but also strengthened cancer coverage for states adopting the Medicaid expansion,⁴² with subsequent improvements in coverage for cancer survivors.⁴³ But not all states have participated in the Medicaid expansion, and this legislation is yet to be modified or discontinued. Additionally, Congress recently passed legislation to require state Medicaid programs to cover routine patient costs for

cancer clinical trial participation.⁴⁴ This legislation could provide coverage for routine supportive care that patients with Medicaid may be lacking.

There are limitations to the current study. Although the trials included in the analysis routinely capture protocol treatment data, other data on supportive care administered during therapy, the management of acute complications, and treatments administered after progression are lacking. Details about the types of resources relevant to cancer survivorship were also not available. Additionally, the findings relate to enrollments from a single National Cancer Institute network group, so the results may not be generalizable to other groups. The absence of an observed ADI gradient among patients with Medicaid or no insurance is limited by the relatively few patients in this subgroup, potentially limiting power to identify trends over ADI quartiles. Also, the quality of care for patients enrolled over multiple decades to multiple trials may vary, although since treatment was guided by protocol, the risk of poor quality care was likely minimized. Finally, since the individual cohort-specific effects were considered random, the

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interpretation about which cancers are primarily subject to socioeconomic adversity is limited.

Our findings suggest that *initial* access to protocol-guided therapy in a clinical trial is not-by itself-sufficient to eliminate the disparate outcomes related to deprivation. Future research should examine whether the etiology of this residual disparity is related to reduced access to supportive care or postprotocol therapy and/or to differences in health status-determined using linked data sources such as Medicare claims—not reflected by protocol staging criteria. Policies to mitigate socioeconomic differences in cancer outcomes should emphasize access to cancer care services beyond initial therapy. Moreover, understanding whether there are differences in outcomes for patients from poor areas receiving clinical trial care is important for researchers designing and interpreting clinical trials. This is especially true given the increasingly recognized importance of ensuring that trial cohorts adequately represent the cancer population for which the experimental therapy is targeted.

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AUTHORS' DISCLOSURES OF POTENTIAL CONFLICTS OF INTEREST

Persistent Disparity: Socioeconomic Deprivation and Cancer Outcomes in Patients Treated in Clinical Trials

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Open Payments is a public database containing information reported by companies about payments made to US-licensed physicians (Open Payments).

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