# Socio-economic and lifestyle factors associated with the risk of prostate cancer 

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#### Abstract

Summary International and interethnic differences in prostate cancer incidence suggest an environmental aetiology, and lifestyle and socioeconomic factors have been studied, but with divergent results. Information on a cohort of 22895 Norwegian men aged 40 years and more was obtained from a health examination and two self-administered questionnaires. Information on incident cases of prostate cancer was made available from the Cancer Registry. We used the Cox proportional hazards model to calculate incidence rate ratios as estimates of the relative risk (RR) with $95 \%$ confidence interval (CI). Reported $P$-values are two-sided. During a mean follow-up of 9.3 years, 644 cases were diagnosed. Risk was elevated among men in occupations of high compared to low socio-economic status ( $\mathrm{RR}=1.30 ; 95 \% \mathrm{Cl} 1.05-1.61$ ), and among men with high education compared to the least educated ( $\mathrm{RR}=1.56$; 95\% $\mathrm{Cl} 1.11-2.19$ ). A RR of 1.56 ( $95 \% \mathrm{Cl} 0.97-2.44$ ) suggests a higher risk among divorced or separated men, compared with married men. We also found indications of a weak negative association with leisure-time physical activity ( $\mathrm{RR}=0.80 ; 95 \% \mathrm{Cl} 0.62-1.03$ for high vs low activity), a weak positive association with increasing number of cigarettes $(P=0.046)$, while alcohol consumption was not related to the risk of prostate cancer. These results show that high socio-economic status is associated with increased risk of prostate cancer, and that divorced or separated men might be at higher risk than married men. Data from this study also indicate that high levels of physical activity may reduce prostate cancer risk. © 2000 Cancer Research Campaign


Keywords: prostate cancer; risk; socio-economic factors; lifestyle

Prostate cancer is a common malignancy in Western societies, and a leading cause of male cancer deaths (Cancer Registry of Norway, 1998; Landis et al, 1998), but the current knowledge of its aetiology is meagre. Nonetheless, the international variation (Muir et al, 1991; Parker et al, 1998), and especially that risk is modified by migration (Haenszel and Kurihara, 1968), indicates environmental mechanisms of causation.

Cigarette smoking is generally not considered a risk factor for prostate cancer (Colditz, 1996), although a few studies have reported a positive association (Hsing et al, 1991; Cerhan et al, 1997). Similarly, there are conflicting findings on the association between alcohol consumption and the risk of prostate cancer (Hiatt et al, 1994; Andersson et al, 1996; Hayes et al, 1996; Lumey et al, 1998). Available results for physical activity and fitness are inconsistent (Le Marchand et al, 1991; Lee et al, 1992; Thune and Lund, 1994; Hartman et al, 1998), as are those considering marital status and socio-economic factors (Talamini et al, 1986; Severson et al, 1989; Hayes et al, 1992; Andersson et al, 1996; Harvei and Kravdal, 1997). A possible relation between diabetes and prostate cancer has been investigated under the hypothesis that an altered endocrine milieu in men with this disease protects against prostate cancer, but findings are inconclusive (Steenland et al, 1995; Giovannucci et al, 1998).

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Using prospective data from a large health screening survey, we examined the association between several lifestyle and socio-economic factors and the development of prostate cancer in a cohort of Norwegian men.

## MATERIALS AND METHODS

## The cohort

From 1984 to 1986, the National Health Screening Service in Norway conducted a large health survey in the county of NordTrøndelag, where all residents aged 20 years or more by 31 December 1983 were invited to participate. Among 85100 eligible persons, 77310 ( $90.8 \%$ ) filled in the questionnaire that was mailed with the invitation. The health examination included measurements of height and weight, capillary blood glucose level, blood pressure and pulse rate. A second questionnaire was handed out together with a pre-stamped envelope that the participants were asked to fill in and return from home. Among other items, this second questionnaire included detailed questions on smoking history, alcohol consumption, physical activity, educational attainment and occupation. A more comprehensive description of the participants, questionnaires and screening procedures is previously given by Holmen and Midthjell (1990).

We restricted the study to men who were at risk of developing the disease, and included all 22895 men aged 40 years or more at baseline who had no history of any cancer at study entry. Each participant contributed person-years from the date of study entry until the date of a cancer diagnosis (of all sites), death, emigration, or the cut-off date of 1 January 1996, whichever occurred first.

## Follow-up

Every citizen in Norway is given a unique 11-digit identity number, and Norwegian law strictly regulates the use of this number. The Norwegian Data Inspectorate, the Norwegian Board of Health, and the Regional Committee for Ethics in Medical Research approved our study protocol. Being attached to each participant's record, the identity number enabled linkage to the Norwegian Cancer Registry in order to identify incident cases of prostate cancer that occurred in the cohort during follow-up.

Mandatory reporting of cancer is regulated by Norwegian law, and the Cancer Registry has since 1953 registered all incident cases of cancer in Norway. To achieve a high degree of completeness and high data quality, the material of the Cancer Registry is matched against the Register of Deaths at Statistics Norway. For all cases registered since $1953,84.7 \%$ are histologically verified, and only $1.7 \%$ of the diagnoses are based on death certificate alone (Cancer Registry of Norway, 1998). Evaluation of the quality of prostate cancer data in the Cancer Registry has shown a completeness of nearly $100 \%$ (Harvei et al, 1996).

## Study factors

Exposure data was collected from three sources. Marital status was added to the health survey database from Statistics Norway, clinical data were obtained at the physical examination included in the screening programme, and the questionnaires provided information on lifestyle and demographic characteristics. We classified smoking status in three categories, where individuals who had never smoked cigarettes daily were considered non-smokers, and men who reported previous or present daily smoking were classified as former and current smokers respectively. We also analysed the association with number of cigarettes smoked per day (in quartiles) and pack-years of smoking (in quartiles), using non-smokers as the reference. For alcohol consumption, the reference category included men not drinking alcohol the past 2 weeks, whereas men who reported drinking 1-4 times and more than 4 times the past 2 weeks comprised category two and three. Moreover, men who reported to be teetotallers were compared to drinking men, and men who reported periods of excessive drinking were compared to men without such periods.

In the health survey questionnaire on leisure-time physical activity, the participants where asked, 'how often do you exercise?', 'how hard do you exercise?', and "for how long do you carry on?', with five, three and four response options respectively. Regarding frequency of leisure-time physical activity, we considered men who exercised less than once a week as inactive, whereas individuals exercising $1-3$ times per week and more than 3 times per week were classified as moderately active and highly active. In addition, we utilized the information on frequency, intensity and duration to calculate a summary measure (index) of physical activity, categorized into three equal groups labelled low, medium and high. Occupational physical activity was classified according to how often they reported being physically worn out after a day's work (low activity when reporting almost never or infrequently worn out and high activity when reporting often or nearly always worn out). For the association with cardiovascular fitness, we analysed pulse rate and systolic blood pressure in quartiles, however, the results are presented as dichotomized variables.

History of diabetes mellitus (both insulin-dependent and non-insulin-dependent) was assessed from the baseline questionnaire, whereas levels of non-fasting blood glucose among nondiabetics was dichotomized according to the cut point for a 'positive screening' used in the health survey ( $\geq 8.0 \mathrm{mmol} / \mathrm{l}$ ) (Holmen and Midthjell, 1990), which is in conformity with the WHO criteria of 1980 (WHO, 1980).

Marital status was classified as married, unmarried, widower and divorced/separated. Educational attainment was classified as primary and lower secondary school ( $0-9$ years), upper secondary school ( $10-12$ years) and college or university ( $>12$ years). Finally, different occupational categories with respect to socioeconomic status were classified as follows: unemployed men, unskilled manual workers, fishermen and subordinate staff comprised the reference category, while category two included men occupied in farming, agriculture, or forestry. The third category (i.e. occupations of highest socio-economic status) included skilled manual workers, men in professional or management positions and self-employed individuals.

## Statistical analysis

The Cox proportional hazards model (Kleinbaum, 1995) was used to calculate relative risk estimates for each of the variables under study. Because prostate cancer is highly age-dependent, we included age (across 2.5 -year categories) and the relevant exposure variable as independent variables, and individual number of person-years as the dependent variable. This produced ageadjusted incidence rate ratios as estimates of the relative risk (RR) with $95 \%$ confidence interval (CI). When appropriate, a two-sided test for trend across exposure categories was calculated by treating the categories as ordinal variables in the proportional hazards model.

On the basis of the results thus obtained, we evaluated potential confounding between each exposure variable and other available factors. Thus, the multivariate analysis included cigarettes per day, alcohol consumption, leisure-time physical activity index, marital status, educational attainment and occupation. Due to inter-correlation between educational attainment and occupation, these variables were not treated simultaneously in the multivariate analysis.

In a separate analysis, we used 1 January 1993 as cut-off date for follow-up, in order to avoid potential bias due to differential testing with prostate-specific antigen (PSA) between categories of marital status, education and occupation. We also conducted a separate analysis for metastatic prostate cancer, since the risk of localized and invasive prostate cancer may vary with respect to the variables under study. Due to small numbers with metastatic prostate cancer we analysed some of the variables using fewer categories than in the previous analysis, and adjustment for age was made using 5 -year categories.

All statistical analyses were performed using the statistical software SPSS for Windows (Release 8.0.0, Copyright © SPSS Inc. 1989-1997).

## RESULTS

A total of 644 men were diagnosed with prostate cancer (mean follow-up $=9.3$ years), and 212720 person-years were observed altogether. Mean age at diagnosis was 75.5 years (range 49.2-96.2

Table 1 Follow-up study of 22895 Norwegian men participating in a large health survey. Number of individuals included between 1984 and 1986 and accumulated person years during 12 years of follow-up, ${ }^{\text {a }}$ incident cases and incidence rate of prostate cancer by age at entry

| Age at entry <br> (years) | No. of men <br> (\% of total) | No. of person years <br> (\% of total) | No. of cases <br> (\% within age-group) | Incidence rate <br> (cases/10 000 py) |
| :--- | :---: | :---: | :---: | ---: |
| $40-49$ | $6133(26.8)$ | $65432(30.8)$ | $7(0.1)$ | 1.1 |
| $50-59$ | $5540(24.2)$ | $56882(26.7)$ | $78(1.4)$ | 13.7 |
| $60-69$ | $6034(26.3)$ | $55086(25.9)$ | $229(3.8)$ | 41.6 |
| $70-79$ | $3818(16.7)$ | $28328(13.3)$ | $257(6.7)$ | 90.7 |
| $\geq 80$ | $1370(6.0)$ | $6992(3.3)$ | $73(5.3)$ | 104.4 |

${ }^{\text {a }}$ Censored by death, cancer diagnosis (all sites) and migration.

Table 2 Age-adjusted relative risks (RRs) and 95\% confidence intervals (Cls) of incident prostate cancer associated with smoking habits and alcohol consumption in a cohort of 22895 Norwegian men

| Variable ${ }^{\text {a }}$ | No. of cases | No. of men | No. of person-years | RR | 95\% CI | $P$ for trend ${ }^{\text {b }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Smoking status |  |  |  |  |  |  |
| Never | 222 | 6316 | 58155 | 1.00 | Reference |  |
| Former | 183 | 6097 | 57748 | 0.98 | 0.80-1.19 |  |
| Current | 153 | 6627 | 62162 | 0.96 | 0.78-1.19 | 0.693 |
| Cigarettes/day |  |  |  |  |  |  |
| 0 | 222 | 6316 | 58155 | 1.00 | Reference |  |
| 1-8 | 73 | 2754 | 25574 | 0.84 | 0.64-1.10 |  |
| 9-10 | 67 | 2689 | 25748 | 1.05 | 0.79-1.39 |  |
| 11-15 | 51 | 2079 | 20344 | 1.37 | 1.00-1.88 |  |
| >15 | 45 | 2255 | 22289 | 1.27 | 0.91-1.76 | 0.046 |
| Pack-years (20 cigarettes/pack) |  |  |  |  |  |  |
| 0 | 222 | 6316 | 58155 | 1.00 | Reference |  |
| 1-10 | 50 | 2520 | 25064 | 0.95 | 0.70-1.30 |  |
| 11-17 | 34 | 2021 | 19850 | 0.84 | 0.58-1.21 |  |
| 18-25 | 58 | 2107 | 20334 | 1.24 | 0.92-1.67 |  |
| >25 | 73 | 2241 | 20593 | 1.22 | 0.93-1.60 | 0.102 |
| Alcohol consumption the past 2 weeks |  |  |  |  |  |  |
| None (but not teetotaller) | 281 | 8447 | 76302 | 1.00 | Reference |  |
| 1-4 times | 148 | 6844 | 67939 | 1.15 | 0.94-1.41 |  |
| >4 times | 40 | 1721 | 16756 | 0.90 | 0.64-1.25 | 0.862 |
| Teetotaller |  |  |  |  |  |  |
| No | 469 | 17012 | 160998 | 1.00 | Reference |  |
| Yes | 80 | 1700 | 14445 | 1.22 | 0.96-1.55 |  |
| Excessive drinking during some period(s) of life |  |  |  |  |  |  |
| No | 319 | 11368 | 107573 | 1.00 | Reference |  |
| Possibly/yes | 140 | 5291 | 50278 | 1.09 | 0.90-1.34 |  |

${ }^{\text {a }}$ Information on each variable was not available on all participants. ${ }^{\text {b }}$ Two-sided $P$-values for trend by proportional hazards model when variables were treated as ordinal variables.
years). The age-specific incidence rates (Table 1) correspond to those of the total Norwegian population (Cancer Registry of Norway, 1998), and the lifetime risk (up to 75 years) of prostate cancer would be $9.7 \%$.

Table 2 shows the age-adjusted RR of prostate cancer associated with smoking habits and alcohol consumption. Smoking status and pack-years of smoking were not associated with prostate cancer risk, but there was a slightly increased risk with increasing number of cigarettes $(P=0.046)$. Frequency of alcohol consumption the past 2 weeks was not significantly associated with prostate cancer. However, comparing teetotallers to drinking men, we found a slightly increased risk among teetotallers ( $\mathrm{RR}=1.22$; $95 \% \mathrm{CI}$ $0.96-1.55)$. The associations with cigarette smoking and alcohol consumption were not materially changed after adjustment for potential confounding with each other or with marital status,
education and occupation. Adjustment for physical activity changed the estimates only modestly (data not shown).
The results for physical activity, cardiorespiratory fitness, diabetes and blood glucose are presented in Table 3. Overall, there was no strong association, but the most physically active men had a suggestive $20 \%$ reduction in prostate cancer risk compared to the least active ( $R R=0.80 ; 95 \%$ CI $0.62-1.03$ ), which did not change after adjustment for potentially confounding variables (data not shown). Additionally, men with known diabetes mellitus had a relative risk of 1.31 compared to men without the disease ( $95 \%$ CI $0.93-1.82$ ).

Table 4 shows that divorced or separated men might be at higher risk of prostate cancer than married men $(R R=1.56 ; 95 \%$ CI 0.97-2.44). Increasing years of education was positively associated with prostate cancer risk ( $P=0.007$ ), and men in occupations of high socio-economic status had higher risk than men in

Table 3 Age-adjusted relative risks (RRs) and 95\% confidence intervals (CIs) of incident prostate cancer associated with measures of physical activity, cardiovascular fitness, diabetes, and blood glucose levels in a cohort of 22895 Norwegian men

| Variable ${ }^{\text {a }}$ | No. of cases | No. of men | No. of person-years | RR | 95\% CI | $P$ for trend ${ }^{b}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Leisure-time physical activity |  |  |  |  |  |  |
| Inactive | 178 | 7385 | 68384 | 1.00 | Reference |  |
| Moderately active | 217 | 8053 | 78205 | 1.05 | 0.86-1.28 |  |
| Highly active | 136 | 3187 | 28042 | 1.01 | 0.81-1.27 | 0.882 |
| Frequency-intensity-duration-index of leisure-time physical activity |  |  |  |  |  |  |
| Low | 141 | 4163 | 37459 | 1.00 | Reference |  |
| Medium | 136 | 4348 | 42125 | 1.08 | 0.85-1.37 |  |
| High | 107 | 4259 | 41450 | 0.80 | 0.62-1.03 | 0.099 |
| Occupational physical activity |  |  |  |  |  |  |
| Low | 148 | 7180 | 71669 | 1.00 | Reference |  |
| High | 116 | 6172 | 62166 | 1.04 | 0.82-1.32 |  |
| Pulse rate (beats/min) |  |  |  |  |  |  |
| $\leq 72$ | 322 | 12103 | 114608 | 1.00 | Reference |  |
| >72 | 307 | 10334 | 95104 | 1.05 | 0.89-1.23 |  |
| Systolic blood pressure ( mmHg ) |  |  |  |  |  |  |
| $\leq 141$ | 227 | 11366 | 111257 | 1.00 | Reference |  |
| >141 | 402 | 11068 | 98447 | 0.98 | 0.82-1.16 |  |
| History of diabetes |  |  |  |  |  |  |
| No | 604 | 21921 | 206337 | 1.00 | Reference |  |
| Yes | 37 | 855 | 5749 | 1.31 | 0.93-1.82 |  |
| Blood glucose levels among non-diabetics (mmol/l) |  |  |  |  |  |  |
| <8.0 | 556 | 19071 | 178418 | 1.00 | Reference |  |
| $\geq 8.0$ | 37 | 967 | 8626 | 1.15 | 0.82-1.61 |  |

${ }^{\text {a }}$ Information on each variable was not available on all participants. ${ }^{\mathrm{b}}$ Two-sided $P$-values for trend by proportional hazards model when variables were treated as ordinal variables.

Table 4 Age-adjusted relative risks (RRs) and $95 \%$ confidence intervals (CIs) of incident prostate cancer associated with marital status, education and occupation in a cohort of 22895 Norwegian men

| Variable ${ }^{\text {a }}$ | No. of cases | No. of men | No. of person-years | RR | 95\% CI | $P$ for trend ${ }^{\text {b }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Marital status |  |  |  |  |  |  |
| Married | 487 | 18070 | 172230 | 1.00 | Reference |  |
| Unmarried | 70 | 2631 | 23485 | 0.94 | 0.73-1.21 |  |
| Widower | 68 | 1380 | 9364 | 0.98 | 0.75-1.29 |  |
| Divorced/separated | 19 | 760 | 7127 | 1.56 | 0.97-2.44 |  |
| Educational attainment |  |  |  |  |  |  |
| Primary or lower secondary (0-9 years) | 435 | 13505 | 122994 | 1.00 | Reference |  |
| Upper secondary (10-12 years) | 80 | 3355 | 33555 | 1.18 | 0.92-1.50 |  |
| College or university (13+ years) | 37 | 1602 | 16216 | 1.56 | 1.11-2.19 | 0.007 |
| Occupation |  |  |  |  |  |  |
| Unemployed, unskilled, fishermen, subordinate staff | 137 | 5021 | 46516 | 1.00 | Reference |  |
| Farmers | 150 | 4517 | 41694 | 1.09 | 0.86-1.37 |  |
| Skilled, professional/management, self-employed | 218 | 8112 | 79021 | 1.30 | 1.05-1.61 |  |

 ordinal variables.
occupations of lower status ( $\mathrm{RR}=1.30 ; 95 \%$ CI 1.05-1.61). These results were not altered after adjustment for potentially confounding variables (data not shown).

To explore whether the results of the study could be biased due to differential testing with PSA between different categories of exposure, we used 1 January 1993 as cut-off date for follow-up. These analyses included 460 cases of prostate cancer diagnosed during a mean follow-up of 7.1 years, and mean age at diagnosis was 75.6 years. However, the estimates of relative risks were similar to those obtained with full follow-up (data not shown).

We also restricted the analysis to metastatic prostate cancer (data not shown), and found a similar increase in risk associated with socio-economic status, but confidence limits were wide due to fewer cases in each category. Moreover, the indications of a negative association with leisure-time physical activity persisted in the analysis of metastatic disease. The RR for highly active men was 0.65 ( $95 \%$ CI $0.40-1.06$ ) compared with the least active men, and the RR for men with a high frequency/intensity/duration-index was 0.65 ( $95 \%$ CI $0.39-1.09$ ) compared with men with a low index value.

## DISCUSSION

In this 12-year follow-up of 22895 Norwegian men 40 years and older, we found that men with high socio-economic status (i.e. long education and/or employed in occupations of high socioeconomic status) had an elevated risk of prostate cancer. This is in agreement with some studies (Rimpelä and Pukkala, 1987; Yu et al, 1988; Williams et al, 1991; Harvei and Kravdal, 1997), contradictory to others (Oishi et al, 1989; Fincham et al, 1990), while a few have reported no association (Talamini et al, 1986; Severson et al, 1989). We also found that divorced or separated men might have a higher risk of prostate cancer than married men. Epidemiological data on the association between marital status and prostate cancer are inconsistent (Newell et al, 1987; Yu et al, 1988; Severson et al, 1989; Hayes et al, 1992; La Veccia et al, 1993; Harvei and Kravdal, 1997).

It is conceivable that some of these results could be biased by differential PSA testing between categories of the cohort. However, when we restricted our analysis to the period before PSA screening became prevalent in Norway, using 1 January 1993 as cut-off for follow-up, there was no material change in the estimates of relative risk. Another possibility that we cannot rule out is a higher incidence among men with high socio-economic status due to greater medical attention. The frequency of TURPs (transurethral resection of the prostate) might also be higher in this group, and thus influence the prostate cancer detection rate. However, neither a strong positive association with age, nor a pronounced international and racial variation is likely to disturb the results of our study. First, our estimates were adjusted for age (across 2.5-year categories), and second, inhabitants in this part of Norway constitute a population of homogenous ethnicity; hence, very few people are non-Caucasian.

Sexual factors have been suggested to have bearing on the development of prostate cancer (Rotkin, 1977; Ross et al, 1987; Oishi et al, 1990, La Veccia et al, 1993), partly because sexual activity might mirror the endocrine milieu of androgens, and partly because men with an active sexual life have more chances to be exposed to transmittable oncogenic agents. Although marital status probably is not a good indicator of sexual activity, the increased risk seen among divorced and separated men in our study could be explained by such mechanisms. Moreover, the higher educated may have more sexual partners than those with less education (Binson et al, 1993). A higher risk of prostate cancer has been found among men in occupations that involve travelling (Pearce et al, 1987), and their sexual activity may also include several partners (National Institute of Public Health, 1993).

We found a weak positive association with number of cigarettes, which is in agreement with some studies (Hsing et al, 1991; Cerhan et al, 1997), while others have reported no association (Adami et al, 1996, Lumey et al, 1997). It has been reported that certain nitrosamines may induce prostate cancer in laboratory rats (Pour, 1983), but an association could also be mediated through the endocrine pathway, since smoking appears to increase levels of circulating androgens in men (Dai et al, 1988).

Available epidemiological data do not indicate a causal role of alcohol (Breslow and Weed, 1998), and our findings are consistent with this. However, our data suggest a slightly increased risk among teetotallers compared to drinking men, but biological mechanisms explaining such an association are not evident,
although animal studies show that alcohol may decrease plasma testosterone (Badr and Bartke, 1974).

We also found suggestive evidence of a negative association between physical exercise and risk of prostate cancer, and this association persisted in the analysis of metastatic disease. Physical activity may reduce levels of circulating testosterone (Aakvaag et al, 1978; Hackney et al, 1988), but previous findings have been inconsistent (Le Marchand et al, 1991; Lee et al, 1992; Thune and Lund, 1994; Hartman et al, 1998). Our measures of pulse rate and blood pressure could, in addition to mirroring levels of physical activity, serve as surrogate markers of sympathetic nervous activity, which may influence prostatic growth (Wang et al, 1991). We found, however, no association with these physiological variables, although previous epidemiological data indicate a relation (Gann et al, 1995).
Previously, a lower risk of prostate cancer has been suggested for men with diabetes mellitus (Thompson et al, 1989; Giovannucci et al, 1998). However, we found a non-significantly higher risk of prostate cancer among men with diabetes in our study, a finding that is supported by Steenland et al (1995). In addition, we hypothesized that high levels of blood glucose among men without diabetes could identify men with lower risk of prostate cancer, but found no association between levels of nonfasting blood glucose and prostate cancer among non-diabetic men.
In summary, our study suggests that men of higher socioeconomic status have increased risk of prostate cancer, and that divorced men might be at higher risk of developing the disease. Furthermore, number of cigarettes smoked per day showed a weak positive association with prostate cancer, while men reporting high physical activity may have had a reduced risk of prostate cancer.

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