

# Symptoms of anxiety and depression and risk of acute myocardial infarction: the HUNT 2 study

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## Aims

The nature of the association of depression and anxiety with risk for acute myocardial infarction (AMI) remains unclear. We aimed to study the prospective association of single and recurrent self-reported symptoms of anxiety and depression with a risk of AMI in a large Norwegian population based cohort.

## Methods and results

In the second wave of the Nord-Trøndelag Health Study (HUNT2, 1995–97) baseline data on anxiety and depression symptoms, sociodemographic variables, health status including cardiovascular risk factors and common chronic disorders were registered for 57 953 adult men and women free of cardiovascular disease. The cohort was followed up during a mean (SD) 11.4 (2.9) years for a first AMI from baseline through 2008. A total of 2111 incident AMIs occurred, either identified at hospitals or by the National Cause of Death Registry. The multi-adjusted hazard ratios were 1.31 (95% CI 1.03–1.66) for symptoms of depression and 1.25 (CI 0.99–1.57) for anxiety. Two episodes of mixed symptoms of anxiety and depression (MSAD), reported 10 years apart, increased the risk for AMI by 52% (11–108%). After exclusion of the first 5 years of follow-up, the association of depression symptoms with AMI risk was attenuated. Relative risk for AMI with anxiety symptoms and MSAD weakened when participants with chronic disorders were excluded.

## Conclusion

Self-reported symptoms of depression and anxiety, especially if recurrent, were moderately associated with the risk of incident AMI. We had some indications that these associations might partly reflect reverse causation or confounding from common chronic diseases.

## Keywords

Depression • Anxiety • Prospective • Risk • Acute myocardial infarction • Epidemiology

## Introduction

Depression and anxiety, both highly prevalent conditions in the general population, are associated with elevated risk for acute myocardial infarction (AMI).<sup>1–7</sup> However, the nature of this association remains controversial.

One of the greatest challenges in the research on the prospective association of anxiety and depression with AMI is that atherosclerosis, one of the main underlying pathophysiological mechanisms of AMI, is known to develop during decades before the first clinical symptoms.<sup>8</sup> However, almost all studies of depression included middle aged or older adults, often with a short follow-up time.<sup>9</sup> Thus, individuals free from clinical heart disease in the aforementioned prospective

studies may not be free from atherosclerosis, which may facilitate depressive symptoms even before generating ischaemia.<sup>8,10</sup>

Furthermore, in most previous reports, the well-known cardiovascular risk factors were inadequately controlled for, and it is not clear to what extent these factors explain the observed association between depression and AMI.<sup>1,9,11</sup> Among the most overlooked confounders are comorbid physical disorders.<sup>12</sup> Depressive symptoms are highly correlated with the overall disease burden, and several common chronic disorders are also risk factors for AMI. Moreover, depression is known to have an episodic nature,<sup>13</sup> yet in the great majority of previous studies, depression was assessed only once. Some studies, however, suggest that recurrent depression has considerably stronger association with AMI than single depressive episodes.<sup>14</sup>

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Further, anxiety and depression are often considered separate psychopathological conditions, yet they share common symptoms and often overlap.<sup>15</sup> Still, relatively little data are available on the prospective association between anxiety and risk for cardiovascular disorders (CAD), and similar methodological concerns as those presented for depression apply to these studies as well.<sup>16</sup> Therefore, the aim of this study was to investigate the prospective association of single and recurrent self-reported symptoms of anxiety and depression with the risk of AMI controlled for potential confounders, including comorbidities in a large population-based cohort.

## Methods

### Study population and setting

All adult citizens in Nord-Trøndelag County, Norway, received a postal invitation to participate in the second wave of the Nord-Trøndelag Health Study (HUNT 2, 1995–97, <http://www.ntnu.edu/hunt>, last accessed 28 July 2013). In total, 94 187 individuals were invited and 65 215 (69%) participated in the HUNT 2 study. The participants filled in questionnaires (<http://www.ntnu.edu/hunt/data/que>, last accessed 28 July 2013) and attended a baseline clinical examination. Details about the study have been published elsewhere.<sup>17,18</sup> Of the participants in HUNT 2, 47 316 also attended the first wave of the HUNT Study (HUNT 1, 1984–86).<sup>19</sup>

This study was approved by the regional ethics committee for research, by the National Directorate of Health, the Norwegian Data Inspectorate, and the HUNT data access committee. All participants gave written consent.

### Exposure: depression and anxiety symptoms

The HUNT self-report questionnaire included a Norwegian version of the Hospital Anxiety and Depression Scale (HADS),<sup>20</sup> which assesses core psychological symptoms of anxiety (HADS-A) and depression (HADS-D) during the last week. The seven anxiety questions mirror symptoms of worry and tension, while the seven depression items mirror mainly symptoms of anhedonia and loss of interest. All questions have a 4-point Likert scale response option from 0 (no symptom) to 3 points (highest symptom level). The HADS has shown good psychometric properties across various patient samples and settings.<sup>21</sup>

In total, 62 685 (96.1%) of the HUNT 2 participants had a valid HADS response, defined as response on 5 or more items on one subscale. Missing response among those who filled in five or six items were replaced based on the sum of completed items multiplied by 7/5 or 6/5, respectively. Both subscales range from 0 (no symptoms) to 21 points.

We categorized the participants according to their HADS-D and HADS-A scores, respectively, as having no symptoms of depression or anxiety (score < 8), having mild to moderate symptoms (score 8–10), and having severe symptoms (score ≥ 11).<sup>22</sup> A combined total score (HADS-T) for assessing the impact of mixed symptoms of anxiety/depression (MSAD) was calculated by summing up valid HADS-A and HADS-D scores. We categorized the participants into three groups: no MSAD (score < 15), mild to moderate MSAD (score between 15 and 18), and severe MSAD (score ≥ 19).<sup>21</sup>

In HUNT 1 (1984–86), the Anxiety and Depression Index (ADI-4) was included. ADI-4 is an acceptable indicator for MSAD (sensitivity 0.51, specificity 0.93) and showed a high correlation (0.83) with the HADS-T score.<sup>23</sup> ADI has been used as a crude baseline measure of MSAD in HUNT 1.<sup>23</sup> ADI consists of four questions, two tapping into general anxiety and two into depression. Two of the questions have four response categories: nervousness ranging from never (1) to almost all the time (4) and calmness which range from almost all the time (1) to

never (4). The two other questions had seven response categories; mood, ranging from very downhearted (7) to very happy (1) and vitality ranging from very strong and fit (1) to very tired and worn out (7).

In total, 32 963 of the participants in the present study had valid MSAD scores both in HUNT 1 (ADI-4) and HUNT 2 (HADS-T). A variable to indicate the combined burden of MSAD in HUNT 1 and HUNT 2 was created in categories of never, one, and two episodes. MSAD in HUNT 2 was defined as cut off ≥ 19 on the HADS-T scale. We then defined previous episode of MSAD in HUNT 1 by scoring in the upper quartile on ADI, cut-off at ≥ 14 points. 'Never MSAD' was as defined as a scoring < 14 on ADI in HUNT 1 and < 19 on HADS in HUNT 2, 'MSAD once' was defined as scoring above cut-off in one of the studies, while 'MSAD twice' included those scoring above cut-off in both studies.

### Follow-up and outcome ascertainment

A total of 4734 participants were excluded at baseline because they reported a history of previous myocardial infarction ( $n = 1944$ ), angina pectoris ( $n = 2978$ ), or stroke ( $n = 1120$ ). The remaining 57 953 participants were included in this study, and *Figure 1* summarizes the recruitment of the participants.

Participants were followed-up for a first AMI from enrolment in HUNT 2 study until 31st of December 2008, either identified at the two hospitals in Nord-Trøndelag County or by the National Cause of Death Registry.<sup>24</sup>

Acute myocardial infarction was diagnosed according to the European Society of Cardiology/American College of Cardiology consensus guidelines.<sup>5</sup> Criteria for AMI included: (i) specific clinical symptoms according to case history information, (ii) changes in blood levels of cardiac enzymes, and (iii) specified ECG changes. Fatal AMI incidents that never reached the emergency units at the hospitals were identified in the National Cause of Death Registry by the International Classification of Diseases codes (code no. 410 in the 9th revision and codes I21 and I22 in the 10th revision). During mean (SD) 11.4 (2.9) years of follow-up, 231 participants who left the county, and 6409 participants who died from other causes than AMI, were censored at the time of event (emigration or death) in the statistical analyses (see *Figure 1*, flow chart).

### Covariates

#### Demographic factors

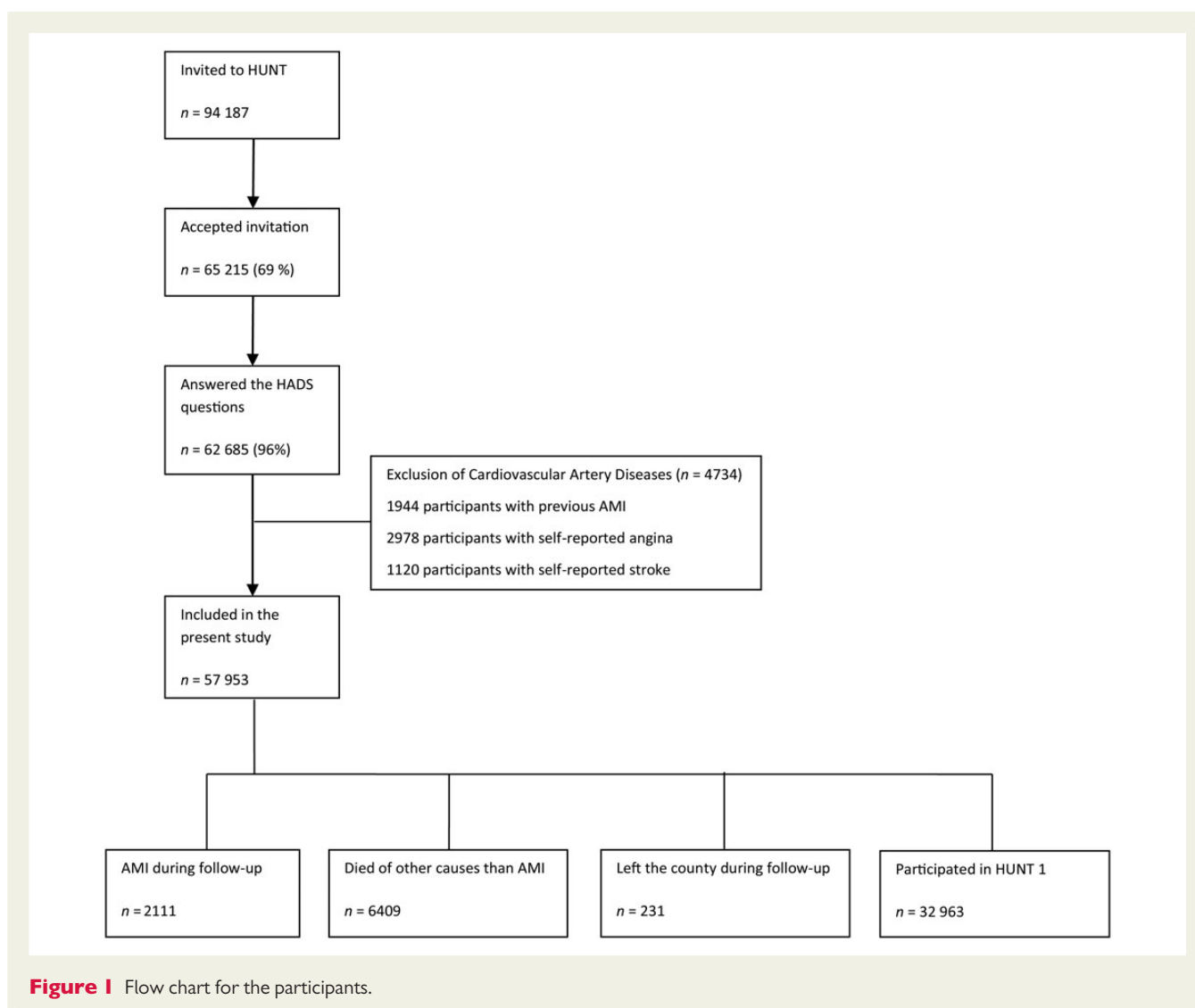
Cohabitation status was dichotomized as living with a partner or alone. Education was categorized as low (≤ 9 years), medium (between 10 and 12 years), or high (over 12 years).

#### Life style factors

Smoking habits were self-reported and were categorized as current, previous, or never smoking. Physical activity was defined as light if it did not involve sweating or feeling of breathlessness, while hard physical activity was defined by sweating or breathlessness. The participants were categorized as inactive (less than 1 h of hard activity and less than 3 h of light physical activity per week), moderate active if they reported 1–3 h of hard activity or > 3 h of light activity per week, and as physically active if they reported > 3 h of hard physical activity per week.

#### Chronic somatic disorders

Participants self-reported if they had (or had ever had) cancer, asthma, diabetes mellitus, other endocrine disorders (hypothyreosis, hyperthyreosis, goitre, and thyroiditis), musculoskeletal disorders (osteoporosis, fibromyalgia, and arthrosis/arthritis), autoimmune disease (Bechterew disease, rheumatoid arthritis), epilepsy, or other chronic disorders (replying yes to the question 'do you have or ever had any other chronic disease' and no to the questions regarding previous specified diagnosis of chronic disease).



### Clinical examination

Blood pressure, heart rate, weight, height, waist, and hip circumference were measured by specially trained nurses after a standardised protocol. Heart rate, and systolic and diastolic blood pressure were measured after the participant had been seated for at least 2 min with the cuff on, and the size of the cuff was adjusted for arm circumference.<sup>17</sup> All blood pressure and heart rate recordings were performed using a Dinamap 845XT (Criticon) based on oscillometry and involved automatically three recordings at 1 min intervals. The first heart rate recording was in average 2 bpm lower than the second and third measurement and was thus used in order to mirror resting heart rate. For blood pressure, the value decreased for each recording, thus we used the average of the second and third measurements in our analysis. Height was measured without shoes to the nearest 1.0 cm and weight with light clothing to the nearest 0.5 kg.<sup>17</sup> Body mass index (BMI) was computed as weight (in kg) divided by the squared value of height (in metres).

### Blood sampling and laboratory measurements

A non-fasting whole blood sample was drawn from each participant, and the time between last meal and the venepuncture was recorded. Serum was separated by centrifugation at the screening site and immediately

placed in a refrigerator (4°C). The samples were sent (same day or following Monday for samples drawn on Friday) for analyses at the central laboratory at Levanger Hospital, Norway, where they used a Hitachi 911 Autoanalyzer, applying reagents from Boehringer Mannheim. Serum total cholesterol was measured by an enzymatic colorimetric cholesterol esterase method. The day-to-day coefficient of variation for total cholesterol was 1.3–1.9%.<sup>17</sup>

### Statistical analysis

Comparisons of continuous variables among those who developed AMI during follow-up and those who did not were made by two-sided t-test, and the  $\chi^2$  test was used to compare categorical data. We used Cox proportional hazard models to examine the association of depression and anxiety symptoms and subsequent risk for AMI. We calculated hazard ratios (HRs) and 95% confidence intervals (CIs). For test of trends, we assigned a numeric value of 0 to 2 to the HADS categories, with 0 having no, 1 having moderate, and 2 having severe anxiety or depression symptoms, treating the categories as a continuous variable. In a separate analysis, we calculated the risk for AMI associated with the presence of MSAD in HUNT 1 and HUNT 2 when those without symptoms of MSAD in both of the surveys constituted the reference group.

We adjusted all our models for sex and age as a continuous variable. In the second adjustment model, we added education and marital status as potentially confounding factors in our models. Established cardiovascular risk factors such as resting heart rate, high blood pressure, low physical activity, high BMI, smoking, dyslipidemia, and diabetes mellitus may act both as confounding and mediating factors for the association of depression and anxiety with the AMI risk. We therefore analysed the data both with and without the factors included in the analyses. We conducted several stratified analyses to assess whether the association of anxiety and depression and the risk for AMI could be modified by other factors.

We investigated the potential effect modification by sex, age (dichotomized at age 50 and age 65), BMI (dichotomized at  $\geq 35$  kg/m<sup>2</sup>), total cholesterol (dichotomized at 6.5 mmol/L), education (dichotomized at 12 years), blood pressure (systolic blood pressure dichotomized at >140 mm Hg, diastolic at >90 mm Hg), and smoking status (current vs. no current smoking). We also formally tested the homogeneity of stratum-specific relative risks. For these tests of interaction, we used the trend variable as defined above.

We present event-free survival curves using the Kaplan–Meier method according to symptom categories.

We performed several sensitivity analyses to assess the robustness of our findings. First, we excluded participants with cancer, asthma, diabetes mellitus, other endocrine disorders, musculoskeletal disorders,

and autoimmune diseases or other reported chronic disorders. We also restricted the analysis to AMI cases that were confirmed at the hospital, thus cases whose AMI diagnosis were based on death certificates alone were excluded from the analyses. Finally, in order to address the possibility of reverse causation as an explanation for possible associations, we excluded the first 5 years of follow-up and repeated the analyses. All sensitivity analyses were run with adjustments for socioeconomic status and traditional cardiovascular risk factors (Model 3).

We tested the proportionality of hazard using log–log curves and formal tests of interaction with time or log-time. There was no evidence against the proportionality assumption (all  $P > 0.10$ ).

Statistical analyses were performed in Stata IC/12.1 for windows (© Stata Corp LP).

## Results

Characteristics of the study population, by AMI status, are shown in *Table 1*. Among the 59 953 participants, 2111 (3.5%) had a first AMI during follow-up. Of these, 1632 cases were diagnosed at a hospital and 479 cases were registered by the National Cause of Death Registry alone.

**Table 1** Baseline characteristics of participants in the total population and according to AMI vs. no AMI during follow-up

Variable	No. of subjects	Total population	AMI during follow-up	No AMI during follow-up	P-value
Total % (n)	57 953	%	3.6 (2111)	96.7 (55 854)	
Variables, % (n)					
Male sex	26 551	45.8	62.5 (1312)	45.2 (25 239)	<0.0001
Diabetes mellitus	1292	2.2	8.3 (176)	2.0 (1 116)	<0.0001
Smoking					<0.0001
Never	26 669	46.2	34.0 (714)	46.6 (25 955)	
Former	13 748	23.8	28.0 (608)	23.6 (13 140)	
Current	17 334	30.0	36.0 (775)	29.8 (16 569)	
Physical activity					<0.0001
Inactive	20 333	38.1	46.9 (796)	37.8 (19 537)	
Moderately active	27 574	51.7	46.5 (788)	51.9 (26 786)	
Physically active	5434	10.2	6.6 (112)	10.3 (5322)	
Living alone	23 009	39.8	35.9 (757)	39.9 (22 252)	<0.0001
Education					<0.0001
≤9 years	18 930	33.7	63.2 (1468)	32.8 (17 763)	
10–12 years	25 311	45.1	27.3 (632)	45.7 (24 765)	
>12 years	11 857	21.1	9.0 (133)	21.5 (11 661)	
Mean (SD)					
Age, years	57 953	47.7 (16.3)	64.9 (12.7)	47.0 (16.0)	<0.0001
Body mass index, kg/m <sup>2</sup>	57 498	26.2 (4.1)	27.4 (4.0)	26.2 (4.1)	<0.0001
Systolic blood pressure, mmHg	57 586	123.5 (21.1)	153.9 (23.5)	135.8 (20.7)	<0.0001
Diastolic blood pressure, mmHg	57 586	79.9 (12.0)	87.1 (12.7)	79.7 (11.9)	<0.0001
Resting heart rate, bpm	57 673	71.6 (13.0)	73.4 (13.5)	71.6 (13.0)	<0.0001
Total cholesterol, mmol/L	57 644	5.8 (1.2)	6.7 (1.2)	5.8 (1.2)	<0.0001
HADS—depression score	57 819	3.4 (3.0)	4.1 (3.0)	3.4 (3.0)	<0.0001
HADS—anxiety score	57 023	4.2 (3.3)	3.9 (3.4)	4.3 (3.3)	<0.0001
HADS—total score	56 889	7.6 (5.6)	8.7 (5.8)	8.6 (5.4)	0.3555
ADI score (HUNT 1)	32 963	13.5 (1.2)	13.6 (1.3)	13.5 (1.2)	0.0036

Using the highest cut off ( $\geq 11$ ), the prevalence rates of HADS-defined anxiety and depression were 5.2 and 3.0, respectively. These estimates were comparable with the corresponding values (6.4 and 4.2%) found in the general European population.<sup>25</sup>

In our study, women reported higher symptom levels of anxiety and MSAD, while depression symptom levels were the same in both genders.

A total of 2111 participants had a first AMI during a mean (SD) follow-up of 11.4 (2.9) years. The prevalence of current smoking, diabetes, physical inactivity, and low education was higher among those who developed AMI during follow-up.

Table 2 presents the HRs for AMI in relation to symptoms of anxiety, depression, and MSAD. Risk for AMI was moderately elevated for participants reporting symptoms of depression, anxiety, and MSAD. The point estimates were fairly robust in multivariable models for socioeconomic and cardiovascular risk factors. Additional adjustment for resting heart rate in model 3 did not alter these results considerably (data not shown). Symptoms of anxiety were associated with higher risks for non-fatal AMI (HR 1.33, 95% CI 1.03–1.71) than for fatal AMI (HR 0.93, 95% CI 0.70–1.24), and this pattern was similar regarding symptoms of depression and MSAD (data not shown).

Figures 2–5 show Kaplan–Meier curves for event-free survival according to symptom categories and number of subjects at risk for each 2.5 years of follow-up.

Table 3 displays the HRs for AMI according to the presence of MSAD combining data from HUNT 1 and 2. In the study population, 48.4% ( $n = 15\ 657$ ) reported no episodes of MSAD, 48.9% ( $n = 15\ 770$ ) reported one MSAD episode, and 2.7% ( $n = 884$ ) experienced two MSAD episodes. There was a trend towards higher HRs in those with two episodes of MSAD compared with those who did

not report MSAD at HUNT 1 or HUNT 2 and HRs were 30% higher than for those reporting MSAD only in HUNT 2.

## Sensitivity analyses

Table 4 presents the multivariable adjusted HRs for AMI in relation to depression, anxiety, and MSAD in different sensitivity analyses.

There were 1582 AMI cases among 49 330 participants that completed the questionnaire on common chronic somatic disorders. Of these, 13 515 persons (27.4%) reported to have at least one chronic disease. The frequency of chronic disorders was: diabetes mellitus  $n = 902$ , other endocrine disorders  $n = 1998$ , autoimmune diseases  $n = 1330$ , muscle/skeletal disorders  $n = 4184$ , asthma  $n = 3898$ , cancer  $n = 1323$ , epilepsy  $n = 652$ , and 'other chronic disorders'  $n = 1793$ . After exclusion of common chronic disorders, the effect sizes were generally attenuated (Table 4). No single group of chronic disorders appeared to be responsible for this effect.

When we restricted follow-up to AMI cases that were confirmed in hospital (1632) events, we obtained similar or even higher hazard ratios for anxiety, depression, and MSAD than in the main analyses.

After the fifth year of follow-up, we observed 1443 AMI events. Exclusion of the first 5 years substantially weakened the observed association between depression and AMI risk. Risk for AMI associated with anxiety symptoms and MSAD remained fairly stable (Table 4).

## Discussion

In this large population-based cohort study, participants with symptoms of anxiety, depression and mixed symptoms (MSAD) had a 20–30% increased risk for first AMI. Participants reporting two episodes of MSAD, 10 years apart, had >50% increased risk for AMI. All effects were robust to extensive adjustments for socioeconomic and

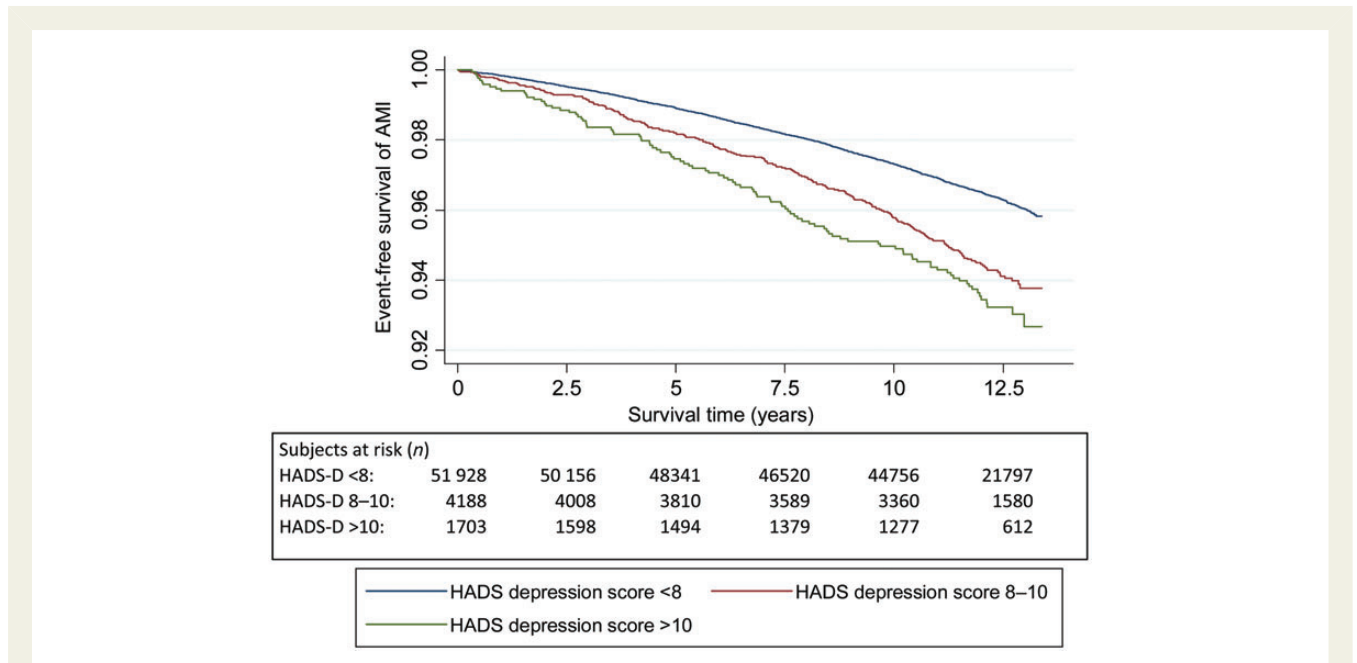
**Table 2 Hazard ratios (95% confidence intervals) for AMI during follow up according to symptoms of depression (HADS-D) and anxiety (HADS-A) and mixed anxiety/depression symptoms (MSAD measured by HADS-T) in HUNT 2 (1995–97)**

Variable		Model 1	Model 2	Model 3
HADS-D	Events/person-time	2096/656 582	1896/636 445	1561/590 365
0–7		Reference	Reference	Reference
8–10		1.08 (0.94–1.24)	1.03 (0.88–1.19)	0.91 (0.77–1.09)
$\geq 11$		1.32 (1.07–1.61)	1.25 (1.01–1.56)	1.31 (1.03–1.66)
<i>P</i> for trend		0.008	0.074	0.253
HADS-A	Events/person-time	2012/648 777	1847/630 642	1533/586 848
0–7		Reference	Reference	Reference
8–10		1.02 (0.87–1.20)	1.00 (0.84–1.18)	0.91 (0.76–1.00)
$\geq 11$		1.30 (1.06–1.60)	1.30 (1.06–1.61)	1.25 (0.99–1.57)
<i>P</i> for trend		0.027	0.049	0.309
HADS-T	Events/person-time	1997/647 414	1836/629 490	1528/586 132
0–12		Reference	Reference	Reference
15–19		1.02 (0.85–1.21)	1.00 (0.83–1.20)	0.87 (0.70–1.07)
$\geq 19$		1.26 (1.04–1.51)	1.24 (1.02–1.51)	1.22 (0.98–1.51)
<i>P</i> for trend		0.032	0.066	0.345

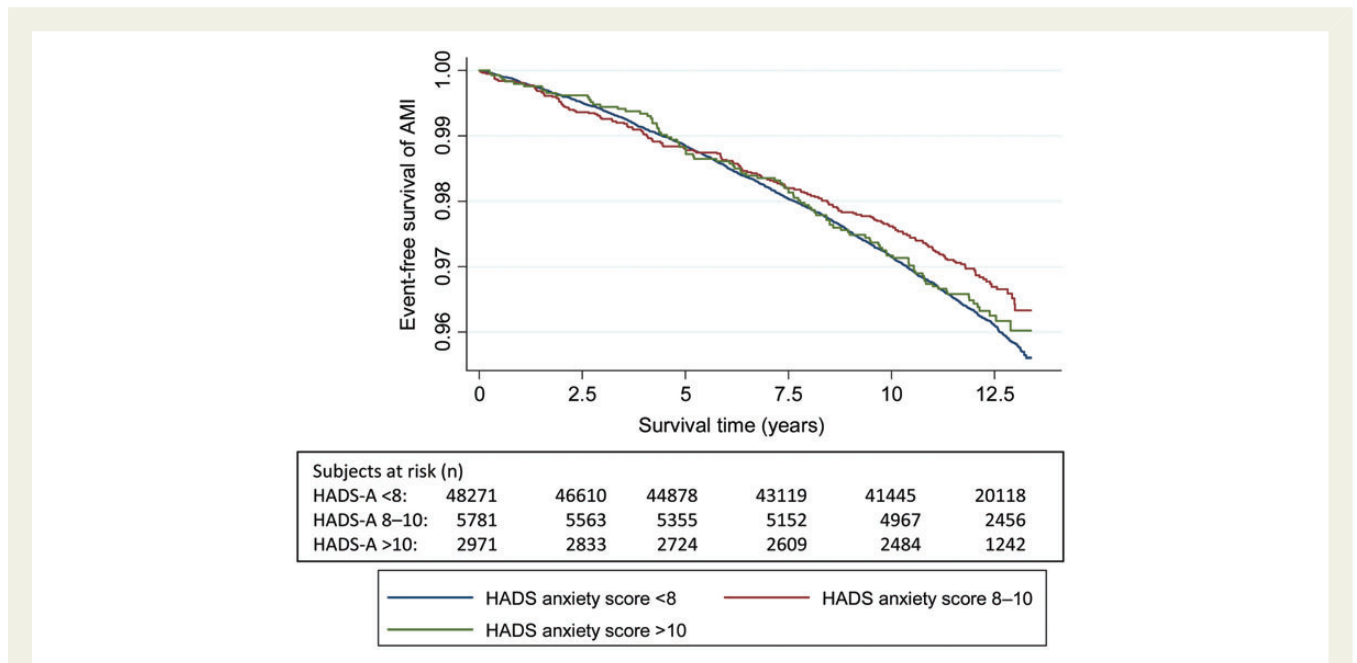
HR indicates hazard ratio; CI, confidence interval; AMI, acute myocardial infarction.

Model 1: adjusted for calendar age and sex. Model 2: Model 1 + marital status, education.

Model 3: Model 2 + smoking, physical activity, body mass index, total cholesterol, diabetes mellitus, systolic blood pressure.



**Figure 2** The Kaplan–Meier curve of incident acute myocardial infarction and subjects at risk during follow-up according to HADS-Depression symptom categories.

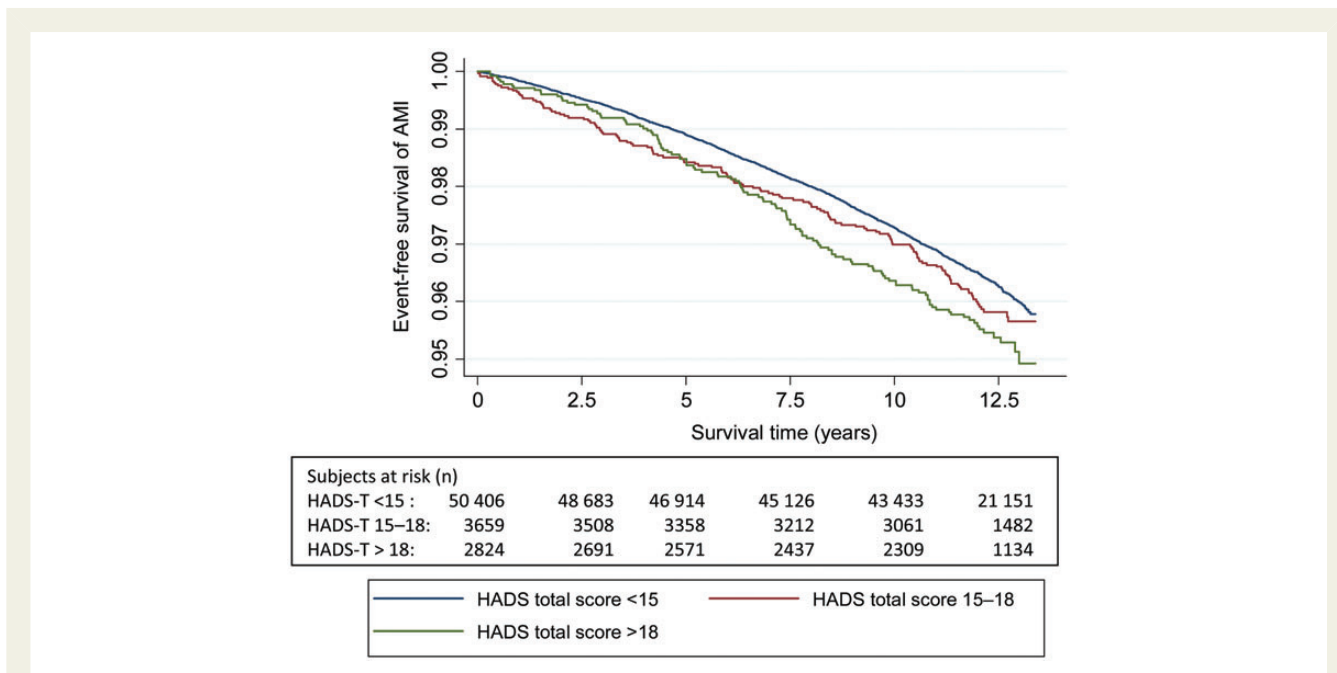


**Figure 3** The Kaplan–Meier curve of incident acute myocardial infarction and number of persons at risk during follow-up according to HADS-Anxiety symptom categories.

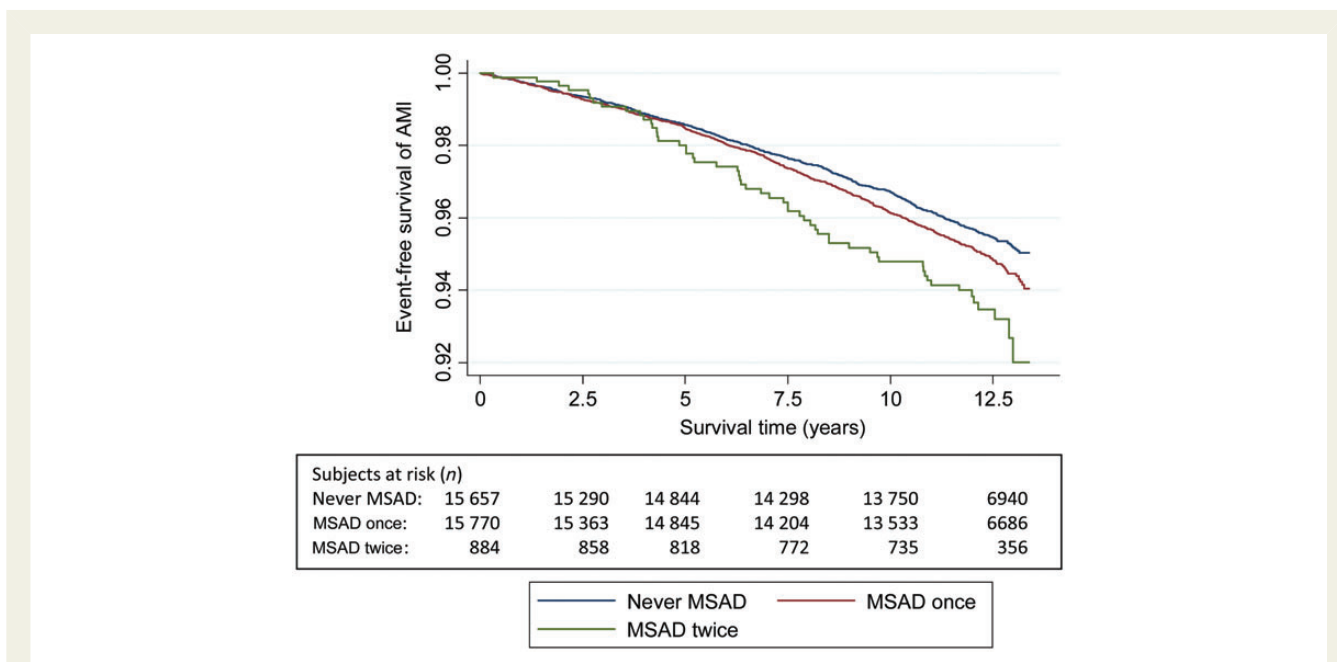
traditional cardiovascular risk factors. For depression symptoms, the effect was only evident the first 5 years of follow-up, which raises a possibility for a reverse causation. For anxiety symptoms and MSAD, AMI risk attenuated with exclusion of participants with co-existent chronic physical illnesses at baseline, which probably reflects confounding.

### Comparisons with previous studies

Although anxiety and depression symptoms tend to overlap, these conditions have to date mostly been investigated separately in terms of risk for AMI.<sup>10</sup> A recent meta-analysis of 23 prospective studies including a wide range of anxiety measures, found a pooled effect-size associated with incident cardiac heart disease of 1.26



**Figure 4** The Kaplan–Meier curve of incident acute myocardial infarction and number of persons at risk during follow-up according to HADS-Total symptom categories.



**Figure 5** The Kaplan–Meier curve of incident acute myocardial infarction and number of persons at risk during follow-up according to mixed anxiety and depression symptom categories in HUNT 1 and 2.

(95% 1.15–1.38).<sup>26</sup> Only 22% of the anxiety studies included in the meta-analysis consisted of both men and women.<sup>26</sup> In our study, the relative risks for AMI in relation to anxiety and depression symptoms were similar in the two sexes and the overall effect was comparable to previous studies.<sup>26</sup>

For depression, another meta-analysis including 21 prospective studies reported a pooled effect of 1.95 (95% CI 1.51–2.51) on CAD death and AMI,<sup>9</sup> considerably higher than the point estimates found in our study. We can think of several explanations for this discrepancy; first, studies using self-report instruments have generally

**Table 3** The association between episodes of mixed symptoms of anxiety and depression (MSAD) in HUNT 1 and 2, and risk of subsequent acute myocardial infarction

Number of episodes MSAD Events/person years	Model 1 1491/372 505	Model 2 1368/361 051	Model 3 1145/333 405
Never MSAD ( <i>n</i> = 15 657)	Reference	Reference	Reference
MSAD once ( <i>n</i> = 15 770)	1.07 (0.97–1.19)	1.06 (0.96–1.19)	1.02 (0.99–1.14)
MSAD twice ( <i>n</i> = 884)	1.63 (1.24–2.19)	1.65 (1.24–2.20)	1.52 (1.11–2.08)
<i>P</i> for trend	0.009	0.016	0.184

Results presented as HR's (95% CIs).

Model 1: adjusted for calendar age and sex.

Model 2: Model 1 + marital status, education.

Model 3: Model 2 + smoking, physical activity, body mass index, total cholesterol, diabetes mellitus+ systolic blood pressure.

found smaller effects than studies relying on clinical diagnosis or psychiatric interviews.<sup>9</sup> Second, adjustment for cardiovascular risk factors and comorbid physical illness(es) was limited in most previous studies. Third, large population-based studies tended to report lower point estimates than smaller studies probably due to publication bias.<sup>9</sup> Finally, we shall emphasize that previous studies used diagnostic interviews or self report instruments that included a wide depressive symptom spectrum, such as sleep problems, lack of energy and disturbed appetite. These symptoms greatly overlap with symptoms of common physical illnesses, including subclinical coronary artery disease.<sup>13</sup> As a result, the genuine contribution of core psychological and cognitive symptoms of anxiety and depression to cardiovascular morbidity has been uncertain. Importantly, the tool used in the current study, i.e. HADS, was originally designed for detection of anxiety and depression in patients with physical illness, and systematically excludes somatic symptoms which can mimic heart disease.<sup>22</sup> Hence, our study is one of the first to link core psychological and cognitive symptoms of anxiety and depression to a moderately increased AMI risk. Nevertheless, by using this approach, some participants with predominantly somatic symptoms of depression might go undetected. As a consequence, we are likely to have underestimated, rather than overestimated, the association between depression and AMI in our study.

Most studies confirm that comorbid anxiety and depression are associated with more severe illness, worse outcome, and increased mortality compared with single conditions.<sup>27</sup> This was not confirmed in our study, as we found no excess risk for AMI associated with MSAD in HUNT 2.

Limited previous evidence suggested that information on anxiety or depressive symptoms provided by repeated measures enable a considerably better risk prediction when compared with the information derived from a single measurement.<sup>1,3,28,29</sup> Our study supports these earlier findings and suggests that persistent or recurrent symptoms might be associated with increased cardiovascular risk. One can hypothesize that a single measurement could reflect a normal reaction to stressful events, while multiple episodes may better capture anxiety and depressive disorders,<sup>30</sup> such as bipolar disorders, known for poorer adherence with lifestyle and medical advice.<sup>31</sup>

Our sensitivity analysis regarding depressive symptoms, but not anxiety or mixed symptoms, indicate a higher risk for AMI in the first 5 years of follow-up. This might reflect a reverse causation.

The possibility that both depressive symptoms and subsequent AMI are caused by subclinical manifestations of atherosclerosis is among the greatest challenges for research on the prospective association between depression and AMI.<sup>12,13</sup> Individuals free from clinical heart disease included in prospective studies may not be free from atherosclerosis, which is known to develop during decades before first symptoms.<sup>8,10</sup> Thus, individuals free from clinical heart disease included in prospective studies may not be free from atherosclerosis. Atherosclerosis may facilitate depressive symptoms even before generating cardiac ischaemia creating a spurious association between depression and atherosclerosis. One possible pathway is the depressogenic actions of the increased inflammatory activity.<sup>32,33</sup> Moreover, neuroimaging and neuropathology studies suggest that late onset depression often has a cerebrovascular origin.<sup>34</sup> Cerebral atherosclerosis may thus facilitate depressive symptoms even before coronary atherosclerosis generates cardiac ischaemia.

Several epidemiological findings supported the reverse causality hypothesis. Studies with a shorter follow-up generally show stronger association with AMI.<sup>9</sup> In the PRIME study, depressive symptoms were associated with coronary heart disease only in the first 5 years of follow-up,<sup>35</sup> a finding which is confirmed in our study. There were some studies with null findings which were conducted in young populations, where vascular depression and subclinical CAD are less likely. In a study of healthy US army personnel 39–45 years of age,<sup>36</sup> no correlation was found between depression and subclinical coronary artery disease identified by electron-beam computed tomography. Further, in a large prospective cohort study of men who were 19–21 years old at baseline, only anxiety but not depression was associated with future CAD risk.<sup>37</sup>

While adjustment for established cardiovascular risk factors was extensive in several studies regarding anxiety,<sup>26</sup> many prospective studies of depression and AMI risk failed to adjust even for well-known cardiovascular risk factors such as smoking and physical activity.<sup>9</sup> Furthermore, comorbid physical disorders are among the most overlooked confounders in previous prospective studies of anxiety or depression and subsequent heart disease.<sup>12</sup> Depressive symptoms are highly correlated with the overall disease burden, and several common chronic disorders are risk factors for AMI.<sup>13</sup> In the present study, AMI risk in participants with symptoms of anxiety and MSAD at baseline was attenuated considerably after excluding all participants with comorbid chronic diseases from the analysis. In



**Table 4** Risk for acute myocardial infarction associated with symptoms of anxiety and depression in sensitivity analyses

	Exclusion of all chronic diseases <sup>a</sup>	Restriction to hospital verified AMI	Exclusion of first 5 years of follow-up
<b>HADS—depression</b>			
Events/person years	768/384 975	1228/51 442	1082/341 485
Cut off 8–10	0.95 (0.74–1.23)	0.94 (0.44–1.15)	0.96 (0.78–1.18)
Cut off ≥ 11	1.25 (0.81–1.91)	1.33 (1.02–1.75)	1.14 (0.83–1.57)
<i>P</i> for trend	0.632	0.203	0.673
<b>HADS—anxiety</b>			
Events/person years	762/384 068	1209/586 848	1060/345 716
Cut off 8–10	1.08 (0.82–1.43)	0.97 (0.79–1.19)	0.87 (0.69–1.10)
Cut off ≥ 11	1.14 (0.74–1.75)	1.33 (1.03–1.71)	1.32 (1.00–1.73)
<i>P</i> for trend	0.439	0.099	0.321
<b>HADS—total</b>			
Events/person years	761/383 732	1205/586 132	1057/339 216
Cut off 15–19	0.75 (0.53–1.05)	0.89 (0.70–1.31)	0.80 (0.61–1.05)
Cut off ≥ 19	1.07 (0.72–1.58)	1.27 (1.00–1.61)	1.23 (0.95–1.59)
<i>P</i> for trend	0.559	0.202	0.563
<b>MSAD (HUNT 1 and 2)</b>			
Events/person years	587/204 678	908/333 405	787/27 215
One episode MSAD	0.93 (0.79–1.09)	1.07 (0.93–1.22)	1.06 (0.92–1.23)
Two episode MSAD	1.38 (0.79–2.40)	1.81 (1.29–2.52)	1.75 (1.21–2.53)
<i>P</i> for trend	0.688	0.022	0.053

MSAD, mixed symptoms of anxiety and depression.

<sup>a</sup>Joint effect of persons that reported asthma, diabetes, other endocrine disorders, musculoskeletal disorders, autoimmune diseases, epilepsy, cancer, and other chronic diseases.

contrast, the excess AMI risk in those with depression symptoms and recurrent MSAD was more robust throughout these analyses.

## Strengths and limitations

Compared with prior studies, we had an ample statistical power to address the prospective association between symptoms of depression and anxiety, and the risk for AMI. The population-based nature of the study, the high stability (less than 0.3% net migration/year), and relatively high genetic and socioeconomic homogeneity of the population<sup>17</sup> along with reliable hospital and register information covering the entire county, ensured close to complete follow-up and minimized the possibility for the misclassification of endpoints or for a selection bias.<sup>24</sup> Moreover, in contrast to many previous studies, we could extensively control for potentially confounding factors.

Yet, the present work has some important limitations. Attendance in HUNT 2 was generally high, but considerably lower in the age group of 70 years and older.<sup>19</sup> Further, we had no information on treatment for depression. SSRIs, the most commonly used medication for anxiety and depression, are known to reduce platelet adhesion<sup>38</sup> and increase BMI<sup>39</sup> and therefore are potential mediators for the observed associations. However, in the present study, it was not possible to separate the effect of depression and its eventual treatment.

## Conclusion

Our findings indicate that self-reported core psychological symptoms of depression and anxiety, especially if recurrent, are moderately associated with the AMI risk. We had some indication that these associations might partly reflect reverse causation (depression) or confounding from common chronic diseases (anxiety and mixed anxiety and depression).

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