### REVIEW

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# Timing for cranioplasty to improve neurological outcome: A systematic review

Maria C. De Cola | Francesco Corallo | Deborah Pria | Viviana Lo Buono | Rocco S. Calabrò

IRCCS Centro Neurolesi "Bonino Pulejo", Messina, Italy

#### Correspondence

Rocco S. Calabrò, IRCCS Centro Neurolesi "Bonino-Pulejo", Messina, Italy. Email: salbro77@tiscali.it

#### Abstract

**Introduction**: Cranioplasty is a surgical technique applied for the reconstruction of the skullcap removed during decompressive craniectomy (DC). Cranioplasty improves rehabilitation from a motor and cognitive perspective. However, it may increase the possibility of postoperative complications, such as seizures and infections. Timing of cranioplasty is therefore crucial even though literature is controversial. In this study, we compared motor and cognitive effects of early cranioplasty after DC and assess the optimal timing to perform it.

**Methods**: A literature research was conducted in PubMed, Web of Science, and Cochrane Library databases. We selected studies including at least one of the following test: Mini-Mental State Examination, Rey Auditory Verbal Learning Test immediate and 30-min delayed recall, Digit Span Test, Glasgow Coma Scale, Glasgow Outcome Scale, Coma Recovery Scale-Revised, Level of Cognitive Functioning Scale, Functional Independence Measure, and Barthel Index.

**Results**: Six articles and two systematic reviews were included in the present study. Analysis of changes in pre- and postcranioplasty scores showed that an early procedure (within 90 days from decompressive craniectomy) is more effective in improving motor functions (standardized mean difference [SMD] = 0.51 [0.05; 0.97], *p*-value = 0.03), whereas an early procedure did not significantly improve neither MMSE score (SMD = 0.06 [-0.49; 0.61], *p*-value = 0.83) nor memory functions (SMD = -0.63 [-0.97; -0.28], *p*-value < 0.001). No statistical significance emerged when we compared studies according to the timing from DC.

**Conclusions**: It is believed that cranioplasty performed from 3 to 6 months after DC may significantly improve both motor and cognitive recovery.

#### KEYWORDS

cognitive outcomes, cranioplasty, motor recovery, neurorehabilitation

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#### 1 | INTRODUCTION

Decompressive craniectomy (DC), consisting in the partial removal of the skullcap, is widely used in the management of neurological emergencies as it allows a decrease in brain swelling and intractable intracranial hypertension (Hofmeijer et al., 2009). DC is performed for a variety of reasons, but the most common are tumor removal and the reduction in increased intracranial pressure due to malignant ischemic or hemorrhagic stroke (Hofmeijer et al., 2009; Vahedi et al., 2007). Cranioplasty (CP) is a neurosurgical procedure aimed to repair the skull defect following craniectomy.

The search for materials and strategies to provide more comfortable and reliable surgical procedures is a challenging topic, both in clinical and in economical terms. However, none of the currently available materials meets the criteria required for an ideal implant (Zanotti et al., 2016).

Besides a purely aesthetic reason, CP helps the individual's rehabilitation from different points of view. The possible advantages of CP have been discussed extensively in literature, as increased cerebral blood flow (Coelho et al., 2014; Erdogan et al., 2003; Maekawa, Awaya, & Teramoto, 1999), change in cerebrospinal fluid hydrodynamics (Juul, Morris, Marshall, & Marshall, 2000; Mah & Kass, 2016; Winkler, Stummer, Linke, Krishnan, & Tatsch, 2000), and reduction in epileptic seizures (Nalbach, Ropper, Dunn, & Gormley, 2012). Recently, promising results following this procedure in both motor and cognitive outcomes have been reported. Thus, this link between the repair of the cranial defect and the changes in cerebrovascular and cerebrospinal fluid hydrodynamics seems to have positive effects on neurological functions (Bijlenga, Zumofen, Yilmaz, & Creisson, 2007).

If on one hand CP may lead to notable improvements (Sancisi et al., 2009; Stiver, Wintermark, & Manley, 2008), on the other hand it may increase the possibility of infections, the risk of hydrocephalus (especially when performed later), and the possibility of developing the "trephined" syndrome (Stiver, Wintermark, & Manley, 2008), especially when operation time exceeding 90 min (Cho & Kang, 2017). Indeed, although the mortality rate after cranioplasty is rather low, research suggests that 1 out of 3 people has overall complications (Zanaty et al., 2015), especially seizures and infection (Honeybul & Ho, 2016). Timing of cranioplasty is therefore crucial even though the literature is divided. According to several studies, it should be performed from 3 to 12 months following DC, based on the presence of infections or postoperative complications. Indeed, in order to prevent the development of devitalized autograft or allograft infections it is recommended to wait from 3 to 6 months before reconstructive surgery, even one year if there is an infected area (Aydin, Kucukyuruk, Abuzayed, Aydin, & Sanus, 2011). On the contrary, an early intervention (i.e., within 3 months) seems to reduce neurological complications, especially in patients with severe acquired brain injury, since a lesion in the postacute period might be negative for motor and cognitive recovery (Huang, Lee, Yang, & Liao, 2013). Although the timing to perform cranioplasty largely depends on personal clinical experience rather than

evidence-based data, it could be useful to estimate a suitable threshold to perform cranioplasty.

In this article, we want to review current literature on motor and cognitive effects of an early cranioplasty after decompressive craniectomy, also focusing on the optimal timing to perform it.

#### 2 | METHODS

#### 2.1 | Data sources and keywords

A systematic review and meta-analysis in line with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines were performed.

Articles published up to July 2017 were searched on the PubMed, Web of Science, and Cochrane Library databases, without language restrictions. A follow-up search was done in January 2018. Databases were queried using key words, and their combinations as follows: "Recovery AND Cranioplasty"; "Rehabilitation AND Cranioplasty"; "Timing AND Cognitive AND Cranioplasty"; "Timing AND Motor AND Cranioplasty"; "Early AND Cognitive AND Cranioplasty"; "Early AND Motor AND Cranioplasty"; "Cognitive recovery AND Cranioplasty"; "Motor recovery AND Cranioplasty."

#### 2.2 | Study selection and search strategy

All studies reporting motor and/or cognitive recovery after cranioplasty for the patients with cranial defects after DC were included. Systematic reviews that investigated the effects of cranioplasty timing on motor and cognitive recovery in patients underwent cranioplasty were also included. Reports of less than ten subjects, comments, letters, editorial articles, and studies included mainly patients <18 years old were excluded.

At first, search results were summarized and duplicate citations were deleted, together with non-English articles. Then, titles were screened for relevance to motor and cognitive recovery after cranioplasty. Next, abstracts of the remaining articles were read and those not meeting the eligibility criteria were excluded. The full text of all potential articles was evaluated in depth. In case of uncertainty, or when the abstract was not available, the entire article was read. Two reviewers performed independently the selection of the articles included in this systemic review. The Cohen's kappa score for inter-rater agreement in study selection was computed (Sands & Murphy, 1996). Discrepancy was resolved through discussion.

#### 2.3 | Data extraction and outcomes

Data from the studies were collected in an electronic sheet including age, gender, pathology, craniectomy to cranioplasty time interval, surgical site, and pre- and postcranioplasty assessment. Concerning the latter, given that our primary outcome was to compare effects of early and late cranioplasty on the cognitive and motor recovery,

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we selected studies including as assessment tools Mini-Mental State Examination (MMSE; Folstein, Folstein, & McHugh, 1975), Rey Auditory Verbal Learning Test immediate (RAVLT) and 30-min delayed recall (RAVLT-DR; McMinn, Wiens, & Crossen, 1988), and Digit Span Test (DST; Schroeder, Twumasi-Ankrah, Baade, & Marshall, 2012). We considered patients with disorders of consciousness separately, including studies reporting data from the Glasgow Coma Scale (GCS; Doyle, 1989), the Glasgow Outcome Scale (GOS; Wilson, Pettigrew, & Teasdale, 1998), the Coma Recovery Scale-Revised (CRS-R; Giacino, Kalmar, & Whyte, 2004), and the Level of Cognitive Functioning Scale (LCF; Sander, 2012). Concerning the motor recovery, we selected studies including Functional Independence Measure (FIM; Keith, 1987) or Barthel Index (BI; Collin, Wade, Davies, & Horne, 1988).

In absence of at least one of the aforementioned assessments, administered both at baseline and at follow-up, we excluded the article from the meta-analysis for inadequate study design.

#### 2.4 | Data analysis

The meta-analysis was performed using the metafor package of R (version 3.4.0; the R Foundation for Statistical Computing, Vienna, Austria), setting at  $\alpha$  = 0.05 the statistical significance. Statistical averages and relative percentages of all patient characteristics were combined, when and if appropriated. The main analysis concerned the effects of early versus late cranioplasty on motor and cognitive recovery, assessed by comparing the changes in pre- and

postcranioplasty scores. For studies reporting multiple test assessment, only the primary outcome was included in the analysis. For studies reported multiple evaluation times before CP, we considered as pre-CP evaluation the one closest to the date of the procedure.

We also performed a subgroup analysis by subdividing the studies according to the time interval from DC to CP: within 3 months and within 6 months. Where the article included both the early and the late cranioplasty groups, we considered the patient's subdivisions of the original study. Otherwise, we subdivided the patients into two groups choosing a threshold according to the median time interval between DC and CP.

Since many studies used different outcome scales, as well as had different sample dimensions, the treatment effect of an intervention was estimated by pooling the standardized mean difference (SMD) with 95% confidence interval (CI). Heterogeneity was quantified by the estimated between-study variance  $\tau^2$ ,  $I^2$ . When the level of heterogeneity was higher than 75%, we considered the results obtained by the application of the random-effects model. Risk of bias, at outcome level, was graphically investigated by funnel plot.

#### 3 | RESULTS

#### 3.1 | Study selection

Figure 1 shows our study selection process. A total of 444 records were identified: 243 articles from PubMed database, 198 articles from Web of Science database, and three articles from the Cochrane



**FIGURE 1** PRISMA flow diagram describing the study selection process

	-									
					Time	Complications after				
Study	Age (years)	Gender	Etiology	Surgical site	DC- Cranioplasty	Decompressive Craniectomy	Cranioplasty	Motor outcome	Cognitive outcome	Time between assessment
Alibhai, Balasundaram, Bridle, and Holmes (2013) <sup>a</sup>	79	Male	Tumor	Unilateral (R)	I	I	Seizure	I	Improved	8 weeks
Corallo, Calabro, Leo, and Bramanti (2015)	55	Male	Vascular	Unilateral (R)	8 months	VS, Partial seizure	I	Improved	Improved	6 months
Ratnasingam, Lovick, Weber, Buonocore, and Williams (2015)	21	Male	TBI	Bifrontal	6 months	-	Seizure, Bells' palsy	I	Improved	26 months
Jeyaraj (2015)	52	Male	TBI	Unilateral (L)	3 months	Hemiparesis DX, Hydrocephalus, Syndrome of the trephined		Improved	Improved	11 days
Nguyen, Doan, Gelsomino, Shabani, and Mueller (2016)	37	Male	Vascular	Unilateral (L)	3 months	I	I	Improved	Improved	I
Corallo, Marra, Bramanti, and Calabrò (2014)	30	Male	Vascular	Unilateral (R)	50 days	Hemiparesis SX with dysesthesia, depression, anhedonia, irritability, sleep alterations	1	Improved	Improved	3 months
Castaño-Leon et al. (2017)	36	Male	TBI	Bilateral	7 months	Hydrocephalus, cephalea, dizziness, vomiting, diplopia	None	Improved	Improved	I
Segal, Oppenheim, and Murovic (1994)	35	Male	TBI	Bilateral	6 months	Blind, left leg paretic, right leg plegic, left hand plegic.	None	Improved	1	7 days
"We reported 1 case out of 2 pati	ents TBI·Traur	matic hrain ii	niurv							

 TABLE 1
 Overview of case report excluded in this review

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					Early	СР			Late	СЪ			Mator		
Study	Type	Etiology	Surgical site	Early cutoff	z	Age (years)	Male	DC-CP time interval (d)	z	Age (years)	Male	DC-CP time interval (d)	out- come	<b>Cognitive</b> outcome	Follow- up
Honeybul et al. (2016)	Retrospective	TBI (72%) Vascular (20%) Tumor (6%)	Bifrontal (28.0%) Unilateral R (34.0%) Unilateral L (38.0%)	3 months	20	45.5 ± 16.6	16	64.0 ± 15.2	30	37.2 ± 16.0	22	157.0 ± 125.5	Σ Ξ	DST	<3 days
Stefano et al. (2016)	Prospective	TBI (72.5%) Vascular (27.5%)	Bilateral (24.2%) Unilateral R (31.0%) Unilateral L (37.9%) NA (6.9%)	6 months	15	<b>39.1 ± 15.6</b>	12	127.9 ± 31.9	14	41.0 ± 10.9	10	399.9 ± 85.9	1	RAVLT, RAVLI-D, DST	1 month, 6 months <sup>a</sup>
Corallo et al. (2017)	Prospective	TBI (40%) Vascular (60%)	Bilateral (20.0%) Unilateral R (53.3%) Unilateral L (26.7%)	6 months	15	51.5 ± 15.5	Ŷ	4.5±3.0	15	51.1 ± 12.8	11	7.9 ± 3.5	1	MMSE, GCS, DST, RAVLI RAVL-DR, CRS-R, LCF	1 month, 1 year <sup>a</sup>
Songara et al. (2016)	Prospective	TBI (100%)	Bilateral (6.2%) Unilateral R (31.3%) Unilateral L (62.5%)	3 months	Ŷ	34.5 ± 14.6	4	63.7 ± 16.4	10	38.7 ± 12.0	10	195.8 ± 104.9	1	MMSE, GCS, GOS	1 week <sup>a</sup> , 4 weeks
Kim et al. (2017)	Retrospective	TBI (50%) Vascular (50%)	NA	3 months	12	58.7 ± 15.5	2	74.0 ± 14.5	12	<b>51.4 ± 13.1</b>	α	219.0 ± 131.3	FIM, K-MBI <sup>a</sup>	K-MMSE	<4 weeks
Jasey et al. (2018)	Retrospective	TBI (69%) Vascular (31%)	Bilateral (7.7%) Unilateral R (61.5%) Unilateral L (30.8%)	3 months	Ś	40.8 ± 17.8	ო	75.4 ± 19.4	ω	45.5 ± 19.2	Ŷ	135.5 ± 33.7	Σ	1	۲ Z
Notes. MMSI Span Test; G Barthel Inde» <sup>a</sup> Not conside	<ul> <li>Mini-Mental St</li> <li>S: Glasgow Com:</li> <li>TBI: traumatic k</li> </ul>	ate Examinatio a Scale; GOS: G rain injury; NA sis	n; K-MMSE: Ko lasgow Outcom .: Not Available	rean Mini-Mo ie Scale; CRS-	ental ( -R: Co	State Examina ma Recovery S	tion; RA\ Scale-Rev	/LT: Rey Audito /ised; LCF: Leve	ory Ver I of Co	bal Learning Te gnitive Functio	est imm ning Sc	ıediate; RAVLT-D ale; FIM: Functio	R: and 30-n nal Independ	nin delayed recall: dence Measure; K	DST: Digit -BI: Korean

**TABLE 2** Characteristics of included studies

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TABLE 3 CF	naracteristics of f	ull-text excluded	d studies							
Study	Type	Etiology	Surgical site	Patients	Age, years	Male	Time interval from DC to CP	Motor outcomes	<b>Cognitive</b> outcomes	Follow-up
Honeybul et al. (2013)	. Retrospective	TBI (76.0%) Vascular (20.0%) Tumor (4%)	Bifrontal (40%) Unilateral R (32%) Unilateral L (28%)	25	40 (25-59)	19 (76%)	100 ± 128 days	Σ E	DST	<3 days
Bender et al. (2013)	Retrospective	TBI (46.2%) Vascular (51.8%)	Bifrontal (5%) Unilateral R (50%) Unilateral L (45%)	147	48.3 ± 16.8	95 (64.6%)	86.4 ± 129.7 days	BI, FIM	CRS-R	161.7 ± 68.3 days
Stelling et al. (2011)	Retrospective	TBI (65.0%)	NA	23	Mean 37 16-64 ranged	16 (69.6%)	12 days to 35 months	I	GCS, GOS	<15 months
Shahid et al. (2018)	Prospective	TBI (100.0%)	Bifrontal (2.9%) Unilateral R (50%) Unilateral L (47.1%)	34	31.53 ± 10.08	30 (88.2%)	Mean 5 months 3-29 ranged	I	GCS (pre), GOS, RAVLT, RAVLT-DR, DST	3 months
Huang et al. (2013)	Retrospective	TBI (100.0%)	NA	105	41.94 ± 19.73	71 (67.6%)	78.84 ± 49.04 days	I	GCS (pre), GOS (post)	25.96 ± 15.61 months
Liang et al. (2007)	Retrospective	TBI (100.0%)	Unilateral (86.9%) Bilateral (13.1%)	23	28.6 (16-41)	18 (78.3%)	5-8 weeks	I	GCS (pre), GOS (post)	1 month
Su et al. (2017)	Retrospective	TBI (100.0%)	Bilateral (31.2%) Unilateral R (37.5%) Unilateral L (31.3%)	16	42.4 ± 15.8	12 (75.0%)	AN	В	MMSE	31.2 ± 7.5 days

Note. MMSE: Mini-Mental State Examination; K-MMSE: Korean Mini-Mental State Examination; RAVLT: Rey Auditory Verbal Learning Test immediate; RAVLT-DR: and 30-min delayed recall; DST: Digit Span Test; GCS: Glasgow Coma Scale; GOS: Glasgow Outcome Scale; CRS-R: Coma Recovery Scale-Revised; LCF: Level of Cognitive Functioning Scale; FIM: Functional Independence Measure; K-BI: Korean Barthel Index; TBI: traumatic brain injury; NA: not available; pre: measured only at baseline; post: measure only at follow-up.

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Library database. After removing 239 duplicates and 13 non-English articles, 192 articles were identified. Later, 103 records were excluded by reading titles and 74 articles by reading abstract, including all case-report and case-control studies for unsustainability of results. The case reports and/or case series studies excluded were summarized for the purpose of this systematic review in Table 1.

The 15 remaining studies were full-text-screened. After reading them, one article was excluded because reported duplicate patients (Honeybul, Janzen, Kruger, & Ho, 2013), who were included in a more recent and larger simple size study (Honeybul, Janzen, Kruger, & Ho, 2016), whereas two articles were excluded for inadequate study design (Huang et al., 2013; Liang et al., 2007). Although we contacted eight authors for further information regarding missing data (Bender et al., 2013; Di Stefano et al., 2016; Honeybul et al., 2016; Jasey, Ward, Lequerica, & Chiaravalloti, 2018; Shahid, Mohanty, Singla, Mittal, & Gupta, 2018; Songara, Gupta, Jain, Rege, & Masand, 2016; Stelling, Graham, & Mitchell, 2011; Su et al., 2017), only four of them were able to provide original individual data useful for our meta-analysis (Di Stefano et al., 2016; Honeybul et al., 2016; Jasey et al., 2018; Songara et al., 2016). At the end of selection, six articles and two systematic reviews (Malcolm et al., 2018; Xu et al., 2015) have been included in the present study.

The kappa score for inter-rater agreement in study selection was 0.88 indicating an "almost perfect agreement," (Landis & Koch, 1977) with a percentage of agreement between the two reviewers of 99.3%.

#### 3.2 | Study characteristics

In Table 2 are reported the six studies included in quantitative analysis, with a total of 162 patients (70.99% males and 29.01% females), whereas in Table 3 are reported the seven excluded studies.

Four of selected articles included both early and late cranioplasty groups: Two studies (Kim, Kim, & Hyun, 2017; Songara et al., 2016) used 90 days as threshold for dividing patients into early and late groups, whereas two studies (Corallo et al., 2017; Di Stefano et al., 2016) used a threshold of 180 days. For two studies (Honeybul et

Study	Total	Early	CP	Total	Late	CP		Standardi	sed Mean	l	SWD	04	%_CI
Study	Total	Mean	30	Total	Weall	30		Dille	ence		SIND	50	/0-01
DIGIT-SPAN													
Corallo, 2016	8	1.63	0.92	4	2.25	2.06	100		20		-0.42	[-1.64;	0.80]
Di Stefano, 2016	7	-1.14	2.04	10	1.20	1.69	-				-1.21	[-2.28;	-0.14]
Honeybul, 2016	19	-0.05	1.08	22	0.40	1.67		<del></del> +			-0.31	[-0.93;	0.31]
Fixed effect model	_34			36				$\sim$			-0.51	[-1.00;	-0.02]
Heterogeneity: $I^2 = 3\%$	6, τ <sup>2</sup> = (	0.007, p	<b>b</b> = 0.3	36									
MMSE													
Corallo, 2016	9	5.11	2.15	8	6.63	4.50					-0.42	[-1.38;	0.55]
Kim, 2017	12	4.50	7.49	12	-1.08	3.65					0.91	[ 0.07;	1.76]
Songara, 2016	6	7.17	2.71	8	10.25	4.37					-0.77	[-1.88;	0.34]
<b>Fixed effect model</b>	27			28				$\sim$			0.06	[-0.49;	0.61]
Heterogeneity: $I^2 = 71^{\circ}$	%, τ <sup>2</sup> =	0.6057	, p = (	0.03									
RAVLT													
Corallo, 2016	8	6.25	4.39	5	7.75	3.98					-0.33	[-1.46;	0.80]
Di Stefano, 2016	11	-4.40	8.24	13	4.85	5.21		+			-1.32	[-2.22;	-0.42]
Fixed effect model	19			18			-	$\sim$			-0.93	[-1.64;	-0.23]
Heterogeneity: $I^2 = 45^{\circ}$	%, τ <sup>2</sup> =	0.2218	, p = (	0.18									
RAVLT-DR													
Corallo, 2016	8	3.39	0.62	4	4.10	1.23			20		-0.77	[-2.03;	0.49]
Di Stefano, 2016	10	-0.97	2.72	13	0.19	2.31					-0.45	[-1.28;	0.39]
Fixed effect model	្18			17					-		-0.55	[-1.24;	0.15]
Heterogeneity: $I^2 = 0\%$	b, τ <sup>2</sup> = (	D, p = 0	.68					1					
							2		· ·	2			
					Eavo	re lato	-Z	-i ( oplactv	J I Eavore	Z	ranion	lactu	
					Favo	rs late	cranie	oplasty	Favors	early c	raniop	lasty	

**FIGURE 2** Comparison of early cranioplasty (early CP) versus late cranioplasty (late CP) on pre- and postcognitive scores. Number of participants, with mean and standard deviation of changes in MMSE score, is presented for each study in any group. The point estimate and the overall effect, with 95% confidence intervals, are indicated by a diamond in the forest plots

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Study	Total	Early Mean	CP SD	Total	Late Mean	CP SD	Standardis Differ	sed Mean ence	SMD	95%-CI
CRS-R Corallo, 2016 Fixed effect model Heterogeneity: not app	5 <b>5</b> licable	2.20	3.90	7 7	4.14	4.63			-0.41 <b>-0.41</b>	[-1.58; 0.75] <b>[-1.58; 0.75]</b>
GCS Corallo, 2016 Songara, 2016 Fixed effect model Heterogeneity: $I^2 = 0\%$	5 6 <b>11</b> , τ <sup>2</sup> = (	1.20 1.00 ), p = 0	1.64 0.63 .90	7 10 <b>17</b>	2.43 1.30	2.76 0.82			-0.48 -0.37 <b>-0.42</b>	[-1.65; 0.69] [-1.40; 0.65] <b>[-1.19; 0.35]</b>
GOS Songara, 2016 Fixed effect model Heterogeneity: not app	6 <b>6</b> licable	1.17	0.75	10 <b>10</b>	0.90	0.57			0.40 <b>0.40</b>	[-0.63; 1.42] <b>[-0.63; 1.42]</b>
LCF Corallo, 2016 Fixed effect model Heterogeneity: not app	5 <b>5</b> licable	0.40	0.55	7 7	0.57 Favo	0.98 rs late	-1.5 -1 -0.5 0	0.5 1 1 Favors early	-0.19 <b>-0.19</b> 1 .5 7 craniop	[-1.34; 0.96] <b>[-1.34; 0.96]</b> lasty

FIGURE 3 Comparison of early cranioplasty (early CP) versus late cranioplasty (late CP) on pre- and postcognitive test scores for postcoma patients. Number of participants, with mean and standard deviation of changes in test score, is presented for each study in any group. The point estimate and the overall effect, with 95% confidence intervals, are indicated by a diamond in the forest plots

al., 2016; Jasey et al., 2018), we set at 3 months the threshold to split the patients in two groups, as the median timings were 99 days and 105 days, respectively. Notably, the study population by Jasey et al. (2018) consisted in 26 subjects with a decompressive craniectomy, but only 13 underwent also cranioplasty, who were the only participants included in our analysis.

Half of selected studies were prospective (Di Stefano et al., 2016; Kim et al., 2017; Songara et al., 2016). For 65.44% of patients, the cause of DC was trauma, followed by a cerebrovascular disease (30.86%) and other causes (3.70%). One study included only traumatic brain injury (TBI) patients (Songara et al., 2016). The percentage of trauma was rather homogeneous between early and late patients, 46.40% and 53.60%, respectively.

Cranial procedures locations, when specified, included unilateral, bilateral and bifrontal.

The mean ± SD time interval from DC to CP was 146.76 ± 108.46 days, and it was significantly longer in late (195.79 ± 122.62 days)than in early patients (86.08 ± 33.65 days).

The mean age of participants at baseline was 44.62 ± 15.96 years, with no statistically different between patients submitted to early (46.30 ± 16.42 years) or late (43.22 ± 15.51 years) CP.

Only two studies reported complications after cranioplasty (Corallo et al., 2017; Songara et al., 2016), and in both cases, they were observed in patients belonging to the late group.

For the primary outcome, the studies used different assessment tools to evaluate the functional recovery: FIM was used in three studies, MMSE in three studies, DST in three studies, RAVLT and RAVLT-DR in two studies, and GCS in two studies, whereas LCF, GOS, and CRS-R in only one study. The interval between assessments after cranioplasty, as well as the number of evaluations, was different among studies: One study performed one follow-up within 3 days (Honeybul et al., 2016), one study a follow-up within 4 weeks (Kim et al., 2017), and four studies performed two followup (Corallo et al., 2017; Di Stefano et al., 2016; Jasey et al., 2018; Songara et al., 2016). However, except for one study (Jasey et al., 2018), all authors reported one follow-up after 1 month from cranioplasty, which was the postcranioplasty evaluation considered in our analysis.

#### 3.3 | Timing effects in cognitive domain

Given that there was no significant heterogeneity for any analyses, a fixed-effects analysis was used. Figure 2 shows meta-analyses of early CP versus late CP, subdivided by type of cognitive outcome. There were 55 participants undergoing MMSE in three studies (27 early, 28 late). Here, the estimates of heterogeneity ( $\tau^2$  = 0.61 and  $I^2$  = 71.3% [2.4%; 91.5%]) indicated a moderate statistical heterogeneity probably due to a bias, as hinted by the corresponding



**FIGURE 4** Comparison of early cranioplasty (early CP) versus late cranioplasty (late CP) on pre- and postcognitive test scores according to the timing from decompressive craniectomy. Number of participants, with mean and standard deviation of changes in test score, is presented for each study in any group. The point estimate and the overall effect, with 95% confidence intervals, are indicated by a diamond in the forest plots



**FIGURE 5** Comparison of early cranioplasty (early CP) versus late cranioplasty (late CP) on pre- and postmotor test scores. Number of participants, with mean and standard deviation of changes in test score, is presented for each study in any group. The point estimate and the overall effect, with 95% confidence intervals, are indicated by a diamond in the forest plots

funnel plot. The fixed-effects model showed that an early procedure did not significantly improve MMSE score (SMD = 0.06 [-0.49; 0.61], *p*-value = 0.83). Concerning the remaining cognitive tests, instead, 70 participants underwent the DS test in three studies (34 early, 36 late) with a very low statistical heterogeneity ( $\tau^2 = 0.007$  and  $I^2 = 3\%$ ); 37 participants underwent the RAVLT in two studies (19 early, 18 late) with a moderate statistical heterogeneity ( $\tau^2 = 0.22$  and  $I^2 = 45\%$ ); and 35 participants underwent the RAVLT in two studies (18 early, 17 late) with an absent statistical heterogeneity ( $\tau^2 = 0$  and  $I^2 = 0\%$ ). The corresponding fixed-effects models showed that a late procedure was more effective in improving memory functions (SMD = -0.63 [-0.97; -0.28], *p*-value < 0.001).

Figure 3 shows meta-analyses of early versus late CP, subdivided by type of cognitive outcome used in postcoma patients. There were 28 participants undergoing GCS in two studies (11 early, 17 late) with an absent statistical heterogeneity ( $\tau^2 = 0$  and  $l^2 = 0$ %). We found that a late procedure was significantly effective in improving the clinical condition compared to an early procedure, although it did not reach the statistical significance (SMD = -0.42 [-1.19; 0.35], *p*-value = 0.29). All the remaining scales were used in only one study; therefore, no consistent results emerged.

Overall, these results show very strong evidence of the positive effects of cranioplasty on cognitive functions, but independently from the timing (SMD = -0.19 [-0.68; 0.31], *p*-value = 0.50).

#### 3.4 | Subgroup analysis between 3 and 6 months

We subdivided the studies according to the time interval from DC to CP. Three studies set at 3 months the threshold between early and late CP, including 79 participants (37 in the early and 42 in the late group). Results of the meta-analysis (Figure 4) showed a moderate heterogeneity across studies ( $\tau^2 = 0.48$  and  $l^2 = 72\%$ ). The

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fixed-effects model indicated a nonsignificant difference between early and late CP groups (SMD = -0.03 [-0.49; 0.42], *p*-value = 0.89). On the contrary, two studies set at 6 months the threshold between early and late CP, including 29 participants (15 in early and 14 in late CP). Here, we found that a late procedure was more effective in improving memory functions (SMD = -0.86 [-1.67; -0.06], *p*value = 0.03). However, no statistical significance emerged in the overall model (SMD = -0.24 [-0.63; 0.16], *p*-value = 0.25).

#### 3.5 | Timing effects in motor domain

Three studies assess the motor recovery by means of the FIM scale for 77 participants (35 early, 42 late) with an absent statistical heterogeneity ( $\tau^2 = 0$  and  $l^2 = 0$ %). We found that an early procedure was significantly effective in improving the motor functions compared to a late procedure (SMD = 0.51 [0.05; 0.97], *p*-value = 0.03), as showed in Figure 5.

As all studies set at 3 months the threshold between early and late CP, we did not perform the subgroup analysis.

#### 4 | DISCUSSION

The optimal cranioplasty timing is a controversial matter. This choice mainly depends on the presence of complications, as well as the time needed for the recovery.

Several studies define "early cranioplasty" as a cranioplasty performed within 91 days from decompressive craniectomy (Malcolm et al., 2018, 2016; Xu et al., 2015). Notably, Xu et al. (2015) sustain that early cranioplasty may reduce the duration of surgery by reducing difficulties in dissecting the scalp flap and fitting the bone flap. Nonetheless, this early procedure cannot reduce the complications and may even increase the risk of hydrocephalus. Indeed, Tasiou et al. reported that delayed cranioplasty should be preferred to minimize the risk of infection that may be caused by intervening in a still contaminated wound (Tasiou et al., 2014). Malcolm et al. (2016) showed that early cranioplasty, with almost certain hydrocephalus management, has similar complication rates to late cranioplasty.

In the last few years, researcher interest is moving toward the association of cranioplasty with the recovery of consciousness and cognitive function as well as the timing of performing cranioplasty (Huang et al., 2013; Shahid et al., 2018; Songara et al., 2016). Rish et al. (1979) reported that cranioplasty performed within 6 months after DC is associated with poor outcomes, Huang et al. (2013) sustained that the timing of cranioplasty is not related to the neurological outcomes of TBI, and Corallo et al. found that the neurological recovery is independent from timing and patient's clinical status. (Shamay-Tsoory, Tomer, Goldsher, Berger, & Aharon-Peretz, 2004) However, Malcom et al., in a more recent meta-analysis, including three motor outcomes and a tool specific for postcoma patients, confirmed the positive effect of cranioplasty on neurological function and claimed that an early procedure may enhance this effect (Malcolm et al., 2018). Similarly, many recent studies recommend early cranioplasty because of its association with clinical improvement (Bender et al., 2013; Chibbaro et al., 2011; Liang et al., 2007; Quah et al., 2016), which can be performed as early as 2 weeks postcraniectomy (and in any case not later than 6 months) to lower the overall cost of care by eliminating the need for additional hospital admissions (Beauchamp et al., 2010). Indeed, it would seem that the majority of neurocognitive changes tend to be at their maximum initially and then decline gradually (Di Stefano et al., 2016), given that ipsilateral low cerebral blood flows increased and reached normal levels after CP (Erdogan et al., 2003), raising the recovery of motor and cognitive functioning (Su et al., 2017).

These contradictory results may be attributed to several factors. First of all the heterogeneity of the population studied, but also the study design features, the choice of surgical approach and operational factors (Sancisi et al., 2009). Thus, our review was aimed at shedding some light on the ongoing debate concerning the right timing to perform cranioplasty and to observe positive effects on cognitive and motor functions. The main question was whether it is reasonable to suggest performing cranioplasty within 90 days from craniectomy to improve the neurological recovery.

Our results showed that such timing is "optimal" only when considering motor outcomes. Indeed, in all studies included in this work, we observed greater positive effects on motor function in the early than late cranioplasty group. On the contrary, to observe a significant cognitive recovery CP should be performed later, although Kim et al. (2017) reported a strong evidence of effects on cognitive functions within 90 days. However, its retrospective study design may lead to a minor reliability since data collected and the measured outcomes are not planned before the study began. Indeed, the follow-up assessment was performed not "after" but "within" 4 weeks; hence, the recovery could be not evident in all patients, especially by using the MMSE test. Although it is one of the most popular tests in clinical and research settings, this tool is not sensitive enough to detect cognitive recovery, as it is rather a screening test. In our opinion, patients undergoing CP should be evaluated by means of a detailed neurocognitive battery, without lingering on their global recovery often assessed through short evaluations. Indeed, it is necessary to standardize common guideline on what kind of tests should be administered to patients following CP, since the assessment is not homogenous, often because of the different patient's etiology and clinical conditions. In this study, about 65% of patient's disease was trauma, whereas about 30% was cerebrovascular disease. Although the difference in rehabilitation approaches between vascular versus TBI is little (Shamay-Tsoory et al., 2004), the pathology may affect the timing of performing CP to manage the risk of complications. Even if this issue is not so important in motor functions, it is fundamental in the cognitive domain. To this aim, the assessment should be specifically based on the site and side of lesion, as some brain areas are more strictly related to specific cognitive functions than others (Redolfi et al., 2017).

With regard to the memory tests, the findings suggest that late CP leads to better overall effects. Notably, when focusing on the Digit Span test results, two studies (Corallo et al., 2017; Di Stefano et al., 2016) showed a more significant recovery after 6 months from CP, whereas one study (Honeybul et al., 2016), which has the highest weight, after only 3 months. Thus, we could suppose that a CP performed between 3 and 6 months leads to more significant cognitive recovery, maybe by the restoration of physiological cerebrospinal fluid circulation that, in turn, allows an efficient restoration of blood circulation and, consequently, of the large-scale neuronal networks responsible for cognition (Corallo et al., 2017; Rish et al., 1979). Indeed, before CP, most of the cognitive abnormalities may be due to changes in cerebrovascular and cerebrospinal fluid hydrodynamics, as per the "sinking skin flap syndrome." (Coelho et al., 2014; Erdogan et al., 2003; Juul et al., 2000; Maekawa et al., 1999; Mah & Kass, 2016; Winkler et al., 2000). However, it is possible that the early group of Honeybul (Liang et al., 2007) had an improved outcome after CP in a shorter time due to less severely injured patients than those reported in Corallo et al. (2017) and Di Stefano et al. (2016). Moreover, we have to underline that in Honeybul et al. (2016) the follow-up assessment was performed within 3 days from CP, as opposed to Corallo et al. (2017) and Di Stefano et al. (2016) who performed it after 1 month, thus explaining the substantial difference that might influence the test scores. Indeed, the difference in follow-up assessment times after CP is another important issue to discuss, since it can affect the measurement. After all, in many studies the greatest improvements were evident many months after cranioplasty and most of the clinical improvement due to cranioplasty is secondary to prolonged effects on brain physiology, rather than immediate changes (Jasey et al., 2018). However, neurorehabilitation programs (if performed) might affect outcomes after longer times (Jolliffe, Lannin, Cadilhac, & Hoffmann, 2018), reinforcing cranioplasty effects on spontaneous cognitive recovery. Su et al. (2017) observed synergetic effects of cranioplasty on TBI patients with rehabilitation training, both in the motor and in the cognitive domains. Moreover, it is well known that an early neuropsychological rehabilitation that has been performed for an adequate time can affect the outcomes in both severe brain injured and patients with disorder of consciousness (Sancisi et al., 2009).

It is noteworthy to highlight that in postcoma patients, results showed very strong evidence of effects of cranioplasty on cognitive functions, but independently from the timing. Unfortunately, our meta-analysis included only two studies; thus, the findings might not correctly reflect reality. After all, the current literature is poor of studies investigating cranioplasty effects on cognitive functions by means of specific neuropsychological assessment, and, to the best of our knowledge, this is the first attempt to do such analysis.

To summarize, cranioplasty performed within 30 days after initial craniectomy may minimize infection, seizure, and bone flap resorption, whereas waiting >90 days may minimize hydrocephalus but may increase the risk of seizure (Morton et al., 2018; Thavarajah, Lacy, Hussien, & Sugar, 2012). Moreover, at 6-month follow-up patients with severe brain injury got better functional outcomes after early than late CP (Yang, Song, Yoon, & Seo, 2018).

A limitation of the study consists in the fact that we did not include the key word "complications" in our database search, although it has been reported that postsurgical complications after Brain and Behavior

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cranioplasty may influence the motor and cognitive recovery and the outcome. Thus, further research is needed to address this important issue.

#### 5 | CONCLUSIONS

Despite the limitations of this meta-analysis, findings confirm that cranioplasty may improve cognitive and motor recovery. Although 6 months is considered the minimum time to reduce complications, cranioplasty performed within 3 months from decompressive craniectomy may lead to greater effects on motor functions, while for the cognitive domain that the best choice seems to be from three to 6 months, especially if the patient underwent neuropsychological rehabilitation. Future prospective larger sample studies are needed to standardize the best timing of performing CP in patients with different disorders, also by using specific psychometric approaches in order to improve functional recovery and thus patient's quality of life.

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#### CONFLICT OF INTEREST

The authors declare that they have no financial or other conflict of interests in relation to this research and its publication.

#### ORCID

Rocco S. Calabrò (D) http://orcid.org/0000-0002-8566-3166

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