Common Variants of Hepatocyte Nuclear Factor 1 Are Associated With Type 2 Diabetes in a Chinese Population

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OBJECTIVE—Hepatocyte nuclear factor 1 β (HNF1 β) is a transcription factor that is critical for pancreatic cell formation and glucose homeostasis. Previous studies have reported that common variants of *HNF1\beta* were associated with type 2 diabetes in Caucasians and West Africans. However, analysis in the subjects from the Botnia study and Malmö Preventive Project produced conflicting results, and the role for *HNF1\beta* in type 2 diabetes susceptibility remains unclear. We therefore investigated common variants across the *HNF1\beta* gene in a Chinese population.

RESEARCH DESIGN AND METHODS—Fifteen tagging single nucleotide polymorphisms (SNPs) were analyzed for association with type 2 diabetes in subjects with type 2 diabetes (n = 1,859) and normal glucose regulation (n = 1,785).

RESULTS—Consistent with the initial study, we observed evidence that the risk G allele of rs4430796 in intron 2 was significantly associated with type 2 diabetes (odds ratio 1.16 [95% CI 1.05–1.29], P = 0.0035, empirical P = 0.0475). Furthermore, the at-risk G allele was associated with earlier age at diagnosis in the type 2 diabetic subjects (P = 0.0228).

CONCLUSIONS—The result of this study provides evidence that variants in the $HNF1\beta$ region contribute to susceptibility to type 2 diabetes in the Chinese population. *Diabetes* **58:1023–1027, 2009**

epatocyte nuclear factor 1 β (HNF1 β , also known as transcription factor [TCF]2), is a homeodomain-containing transcription factor that functions as a homodimer or heterodimer with its homologous partner HNF1 α (1). *HNF1\beta* is widely expressed in a number of tissues where it takes a vital role in embryonic development and pancreatic cell formation and is involved in the β -cell transcription factor network (2,3). Heterozygous mutation of the *HNF1\beta* gene was identified in an autosomal dominant, early-onset subtype of diabetes known as maturity-onset diabetes of the young type 5 (MODY5) in Caucasian, Japanese, and Chinese populations (4-6). It is well recognized that common variants of genes that cause monogenic forms of diabetes may also have a role in common type 2 diabetes susceptibility, such as P12A polymorphism in PPARG, E23K polymorphism in KCNJ11, and common variants in the promoter region of HNF4A (7-9). Recently, single nucleotide polymorphisms (SNPs) on $HNF1\beta$ were identified to be associated with type 2 diabetes in Finland, Sweden, and Canada populations (10,11). The association signals, which were separated by recombination hot spots, were located on introns 2 (rs757210) and 6-8(rs1008284 and rs3110641). Moreover, a genome-wide association study on prostate cancer identified that the G allele of rs7501939 and A allele of rs4430796 (both in the intron 1–2 region) were associated with increased risk of prostate cancer in Caucasians and reduced risk of type 2 diabetes in Caucasians and Africans but not Hong Kong Chinese (12). However, the study by Holmkvist et al. (13) produced the conflicting results that common variants in $HNF1\beta$ did not predict future risk of type 2 diabetes in two prospective studies. Thus, extensive study, especially in other ethnic groups, will improve our understanding of the role of $HNF1\beta$ in type 2 diabetes predisposition. We therefore selected tagging SNPs spanning the $HNF1\beta$ gene and performed a casecontrol study in a Chinese Han population.

RESEARCH DESIGN AND METHODS

All subjects were of eastern Chinese Han ancestry and resided in Shanghai. All case subjects were from the 2001–2005 inpatient database of the Shanghai Diabetes Institute and were diagnosed with type 2 diabetes by the 1999 World Health Organization criteria (14). Type 1 diabetes and mitochondrial diabetes were excluded (15). The control subjects were selected from the Shanghai Diabetes Studies, a community-based epidemiological survey for diabetes (16). The inclusion criteria for the control subjects were I) age >40 years, 2) normal glucose regulation confirmed by a standard 75-g oral glucose tolerance test (fasting plasma glucose <6.1 mmol/l and 2-h plasma glucose <7.8 mmol/l), and 3) absence of family history of diabetes stated on a standard questionnaire. In the end, the study groups comprised 1,859 subjects with type 2 diabetes and 1,785 subjects with normal glucose regulation. The phenotypic characteristics of the study group are described in Table 1.

This study was approved by the institutional review board of Shanghai Jiao Tong University Affiliated Sixth People's Hospital. Written informed consent was obtained from each participant.

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Anthropometric parameters such as height, weight, waist and hip circumference, and blood pressure were measured as previously described (16,17). For the case subjects, fasting and postprandial plasma glucose and serum insulin were measured. For the control subjects, blood samples were obtained at 0 and 120 min during the oral glucose tolerance test, and plasma glucose and serum insulin levels were examined. Insulin resistance

TABLE	1			
Clinical	characteristics	of the	study	subjects

	Case subjects	Control subjects
\overline{n}	1,859	1,785
Sex (male/female)	973/886	739/1,046
Ages (years)	61.16 ± 12.63	57.37 ± 12.33
$BMI (kg/m^2)$	24.08 ± 3.53	23.63 ± 4.23
Fasting plasma glucose (mmol/l)	13.04 ± 5.19	4.99 ± 0.50
Age at diagnosis (years)	54.10 ± 11.85	

Data are n or means \pm SD.

and pancreatic β -cell function were assessed by homeostasis model assessment (18).

Genotyping. We initially selected 20 SNPs spanning the $HNF1\beta$ region, including both 1) tagging SNPs selected from the HapMap Phase II Chinese Han database that captured all common variants (minor allele frequency > 0.1) from 3 kb upstream to 1 kb downstream of the gene based on an r^2 of ≥ 0.7 and 2) SNPs previously reported to be associated with type 2 diabetes. SNPs were genotyped by matrix-assisted laser desorption ionization-time of flight mass spectroscopy (MassARRAY Compact Analyzer, Sequenom, San Diego, CA). Genotyping data underwent a series of quality-control checks, and cleaned data were used in the further association analyses. Fifteen SNPs had a call rate >95%, whereas five SNPs were excluded because of low genotyping call rate (rs1016990, rs7407025, rs1008284, rs3110640, and rs11263755). SNPs with Hardy-Weinberg equilibrium test P values <0.05 in the case or control subjects were excluded. No SNP was excluded because of inconsistency with Hardy-Weinberg expectations. The average reproducibility rate among 1,500 duplicate genotype pairs was 100%. The overall call rate for the remaining 15 SNPs was 99.1%

Statistical methods. The Hardy-Weinberg equilibrium test was performed in the case and control groups separately for each variant using χ^2 test. Pairwise linkage disequilibrium was computed from the combined data of case and control subjects by calculating D' and r^2 in Haploview (version 4.1) (19). Sliding windows consisting of two or three adjacent SNPs were generated for haplotype analyses using PLINK (version 0.99) (20). Allele, genotype, and haplotype frequencies for case and control subjects were compared using the χ^2 test or Fisher's exact test. Odds ratios (ORs) with 95% CIs were presented. Correction of multiple testing was performed in Haploview through 10,000 permutations. Empirical *P* values of logistic regression analysis were calculated through 500 permutations that generated simulation data by randomly shuffling case and control status. In the control group, continuous diabetes-related traits were analyzed, and quantitative traits were transformed to reduce skewness. All statistical analyses were performed by SAS (version 8.0; SAS Institute, Cary, NC), unless otherwise specified. A two-tailed *P* value

 ${<}0.05$ was considered significant. Considering a risk allele frequency of 20% and an additive model, we had ${-}80\%$ power to detect an OR of 1.15 at the 0.05 level.

RESULTS

We examined 15 tagging SNPs spanning the $HNF1\beta$ region. Figure 1 shows the pairwise linkage disequilibrium and haplotype block structure among these SNPs. The association of single SNPs with type 2 diabetes risk is shown in Table 2. Six of the 15 SNPs, including three previously reported SNPs, had a nominal P < 0.05. Three previously reported SNPs (rs4430796, rs757210 in intron 2, and rs3110641 in intron 8) were significantly associated with type 2 diabetes susceptibility (OR 1.13– 1.16, P = 0.0035 - 0.0178). Three additional SNPs (two SNPs in intron 8 and one in the 5' end of the gene region) also showed nominal significance (1.15-1.25, P = 0.0046 - 0.0178). Logistic regression analysis of type 2 diabetes with these six SNPs showed that the G allele of rs4430796 and A allele of rs11656817 tended to confer independent risks of type 2 diabetes (adjusted OR 1.135 and 1.211, P = 0.0162 and 0.0203, respectively). However, after adjustment for multiple testing by permutation, only the risk G allele of rs4430796 (in intron 2) showed significantly associated increased risk of type 2 diabetes (OR 1.16 [95% CI 1.05–1.29], P = 0.0035, empirical P = 0.0475). Logistic regression analysis adjusting for age, sex, and BMI also indicated that the association with SNP rs4430796 was statistically significant (1.18 [1.06–1.31], P = 0.0025, empirical P = 0.034). Sliding window haplotype analysis found that windows containing rs4430796 showed the strongest association with type 2 diabetes, which merely confirmed the results of single-SNP association analysis (Figure S1 available in an online appendix [http://diabetes.diabetesjournals. org/cgi/content/full/db08-1064/DC1]).

Because rs4430796 exhibited strong association with diabetes in our population, further analyses focused on this SNP for diabetes-related traits. We first treated age at diagnosis as a quantitative trait and analyzed its association with rs4430796 in the type 2 diabetic subjects. There

TABLE 2 Allele frequencies and association results for SNPs in the $HNF1\beta$ gene

						allele iency				
	Position (bp)	<i>HNF1β</i> region	Major/minor allele	Risk allele	Case subjects	Control subjects	OR (95% CI)	$P_{\rm SNP}$	$P_{\rm gene}$	$\begin{array}{c} \text{Adjusted} \\ P_{\text{SNP}} ^{*} \end{array}$
rs17626423	-3,492	5'	T/C	Т	0.886	0.885	1.01 (0.88-1.17)	0.8804	1	0.4552
rs3760511	-1,438	5'	C/A	А	0.368	0.336	1.15 (1.04–1.27)	0.0046	0.0624	0.0117
rs4430796	6,836	Intron 2	A/G	G	0.312	0.280	1.16 (1.05-1.29)	0.0035	0.0475	0.0025
rs757210	8,361	Intron 2	G/A	А	0.289	0.264	1.13 (1.02-1.26)	0.0178	0.2129	0.6318
rs3744763	13,991	Intron 4	C/T	С	0.557	0.545	1.05 (0.96-1.15)	0.3065	0.9874	0.5032
rs3786127	17,002	Intron 4	C/G	С	0.837	0.834	1.02(0.90-1.15)	0.7708	1	0.7296
rs2107131	18,187	Intron 4	C/T	Т	0.375	0.364	1.05 (0.95-1.15)	0.3403	0.9934	0.3711
rs12450628	22,455	Intron 4	C/T	С	0.651	0.646	1.02 (0.93-1.12)	0.6989	1	0.8704
rs11649743	29,897	Intron 4	G/A	А	0.328	0.323	1.02(0.93-1.13)	0.6373	1	0.8719
rs2158254	39,381	Intron 5	G/A	А	0.266	0.257	1.05 (0.94–1.16)	0.4075	0.9981	0.2062
rs2189303	44,771	Intron 7	C/T	Т	0.397	0.390	1.03 (0.94–1.13)	0.5362	1	0.9497
rs11656817	47,977	Intron 8	A/G	А	0.916	0.898	1.25 (1.06-1.46)	0.0067	0.0873	0.0210
rs11868513	52,184	Intron 8	G/A	G	0.721	0.719	1.01 (0.91-1.12)	0.8353	1	0.8193
rs1859211	53,504	Intron 8	A/G	А	0.911	0.894	1.21 (1.03–1.41)	0.0178	0.2138	0.0344
rs3110641	57,459	Intron 8	C/T	С	0.770	0.745	1.14 (1.03–1.27)	0.0139	0.1742	0.0307

*Adjusted for age, sex, and BMI.

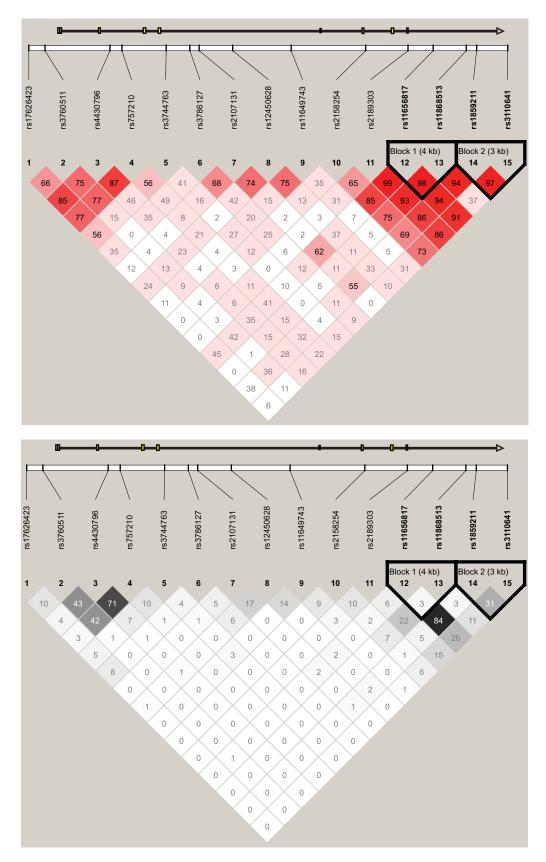


FIG. 1. Linkage disequilibrium patterns of the 15 typed SNPs in the Chinese population.

was evidence that the G allele of rs4430796 was associated with younger age at diagnosis (52.40 ± 11.18 years for G/G, 53.94 ± 11.85 years for G/A, and 54.61 ± 11.97 years for A/A, P = 0.0228). When treating age at diagnosis as a

dichotomous trait to stratify the type 2 diabetic patients and comparing the allele frequencies of rs4430796 with control subjects, we also found that the association signal was slightly stronger in the early-onset (age of diagnosis

TABLE 3

Allelic association analysis of rs4430796 among type 2 diabetic and control subjects stratified according to age at diagnosis of type 2 diabetes

		Allele frequencies			P vs. control	
	n	А	G	OR (95% CI)	subjects	
Type 2 diabetic						
patients	1,856	0.688	0.312	1.16(1.05-1.29)	0.0035	
Early onset	395	0.670	0.330	1.27 (1.07–1.49)	0.0051	
Late onset	1,461	0.693	0.307	1.14 (1.02–1.26)	0.0211	
Control subjects	1,781	0.720	0.280	1.0	—	

Early-onset type 2 diabetes age of diagnosis \leq 45 years; late-onset type 2 diabetes age of diagnosis \geq 45 years.

<45 years) subgroups (OR 1.27 [95% CI 1.07–1.49], P = 0.0051, Table 3).

In subjects with normal glucose regulation, this SNP was not associated with plasma glucose levels; adiposity-related anthropometrics such as BMI, waist circumference, and waist-to-hip ratio; lipid profile; insulin sensitivity; or insulin secretion (supplementary Table S1).

DISCUSSION

We replicated previous findings that a SNP (rs4430796) in intron 2 of $HNF1\beta$ was significantly associated with type 2 diabetes in our Chinese population. Our results are consistent with those reported by Gudmundsson et al. (12), and the effect size of the at-risk G allele for rs4430796 in our Chinese subjects (OR 1.16) is quite similar to that determined by the DECODE findings (for the G allele, overall OR 1.10). Notably, the allele frequency for the SNP rs4430796 is guite different between different ethnic groups. The at-risk G allele for rs4430796 had a frequency of 0.25-0.30 in the Chinese population but a frequency of 0.47-0.54 and 0.70-0.73 in Caucasians and West Africans, respectively (11,12). Although the minor G allele identified in the Chinese and Caucasian populations is the major allele in West Africans, this allele showed the same direction of its effect for type 2 diabetes. rs4430796 was not significantly associated with type 2 diabetes in Hong Kong Chinese, although the effect sizes between the Shanghai and Hong Kong samples were similar (for the risk G allele, OR 1.16 and 1.12, respectively) and in the same direction. One potential explanation for the discrepancy between the current findings and the results in the Hong Kong population is the difference in the control ascertainment. The Shanghai control subjects were >40 years of age and had measurements for both fasting plasma glucose and 2-h plasma glucose. In contrast to the individuals of our study, the Hong Kong control subjects were only tested for fasting plasma glucose. In addition, more than one-third of these control subjects were adolescents. All of these considerations raise the possibility of misclassification in the Hong Kong control samples; consequently, the effect size may be underestimated.

In the study by Winckler et al. (11), the authors observed that other SNPs in intron 2 (rs757210) and intron 8 (rs3110641) were associated with type 2 diabetes in a large Caucasian sample (combined OR 1.12 and 1.10, $P = 5 \times 10^{-6}$ and 0.0006, respectively). SNP rs757210 is correlated with rs4430796 (D' = 0.87 and $r^2 = 0.71$ in our popula-

tion). In the present study, although neither of the two SNPs reached significance after a correction for multiple testing, these two SNPs showed a similar trend in the same direction in our samples. In the study by Holmkvist et al. (13), the three previously reported SNPs (including rs1008284, rs3110641, and rs757210) did not predict future risk. However, our study suggests that the strongest effect for type 2 diabetes susceptibility was at rs4430796, and further study to follow up and replicate the association is still necessary.

Our study is the first to demonstrate that the at-risk G allele for rs4430796 is also associated with earlier onset of type 2 diabetes. Interestingly, previous studies have provided evidence that common variants in the MODY1 gene (*HNF4A*) also associated with type 2 diabetes and age at diagnosis (21,22). Nevertheless, rs4430796 was in the intron region of *HNF1* β , with no obvious functional effect reported; thus, we cannot exclude the possibility that other, as yet unknown or untyped, SNPs in linkage disequilibrium with rs4430796 may be the causal variants.

There are several limitations of this study. First, although the age of diagnosis of diabetes was determined by physician-documented medical records, we cannot exclude the possibility that the case subjects had an earlier age of onset. Second, the SNPs that we successfully genotyped can only capture 60% of common SNPs in the $HNF1\beta$ gene region under the threshold of $r^2 \ge 0.7$ based on phase 3 HapMap Chinese data. It may be inevitable that we miss some common variants in this region that contribute to diabetes.

In summary, our data replicated the previous finding by Gudmundsson et al. (12) and support the notion that genetic variation in the $HNF1\beta$ is associated with type 2 diabetes. Further studies in other populations and of more SNPs at this gene will be required to further elucidate the role of $HNF1\beta$ variation in the pathogenesis of type 2 diabetes.

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