

# Positive temporal dependence of the biological clock implies hyperbolic discounting

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Temporal preferences of animals and humans often exhibit inconsistencies, whereby an earlier, smaller reward may be preferred when it occurs immediately but not when it is delayed. Such choices reflect hyperbolic discounting of future rewards, rather than the exponential discounting required for temporal consistency. Simultaneously, however, evidence has emerged that suggests that animals and humans have an internal representation of time that often differs from the calendar time used in detection of temporal inconsistencies. Here, we prove that temporal inconsistencies emerge if fixed durations in calendar time are experienced as positively related (positive guadrant dependent). Hence, what are time-consistent choices within the time framework of the decision maker appear as time-inconsistent to an outsider who analyzes choices in calendar time. As the biological clock becomes more variable, the fit of the hyperbolic discounting model improves. A recent alternative explanation for temporal choice inconsistencies builds on persistent under-estimation of the length of distant time intervals. By increasing the expected speed of our stochastic biological clock for time farther into the future, we can emulate this explanation. Ours is therefore an encompassing theoretical framework that predicts context-dependent degrees of intertemporal choice inconsistencies, to the extent that context can generate changes in autocorrelation, variability, and expected speed of the biological clock. Our finding should lead to novel experiments that will clarify the role of time perception in impulsivity, with critical implications for, among others, our understanding of aging, drug abuse, and pathological gambling.

Keywords: biological clock, impulsivity, preference reversal, hyperbolic discounting, exponential discounting, random time change

# **1 INTRODUCTION**

Suppose we are asked to choose between \$10 now or \$11 tomorrow. We may prefer the \$10 immediately rather than the \$11 received after a day. However, if we are offered to choose between \$10 to be received after 364 days, or \$11 after 365 days, we often prefer to wait the additional day for the extra dollar. After waiting 364 days, the latter choice becomes one between an immediate \$10 or \$11 tomorrow. Now we would want to reverse it, asking for the \$10 immediately, rather than waiting the extra day we seemed to have been willing to accept in the past. This time inconsistency can be modeled using hyperbolic discounting of future rewards (Laibson, 1997). To avoid inconsistencies, rewards should be discounted exponentially over time (Sutton and Barto, 1998).

Here, we conjecture that hyperbolic discounting has a rational explanation, based on the generally accepted principal that the animal and human biological clocks tick at a different rate from the calendar clock. Animals and humans are indeed known to maintain an internal representation of time that differs from standard calendar time and whose properties change with time scale, from microseconds up to years, with representations of (calendar) time at larger scales showing the highest variability (Buonomano, 2007). In humans, drugs (Meck, 1998; Wittmann et al., 2007), and age (Mischel et al., 1989), among other things, are known to influence time perception. The neurobiological mechanism behind the biological clock has recently become a topic of intense study (Buhusi and Meck, 2005).

We start from the preposition that discounting is exponential, as required for time-consistent choice. At the same time, we posit that the internal representation of time (under which discounting occurs) varies stochastically (randomly) from calendar time. This is illustrated in **Figure 1**: two events at times 0 and  $\Delta$  in calendar are experienced to occur at  $t_0 = F(0)$  and  $t_{\Delta} = F(\Delta)$  in biological time, where *F* is some stochastic (random) transformation *F*.

Now consider two events at a later time *s* and *s* +  $\Delta$ , separated by the same amount of (calendar time) delay  $\Delta$ . The times *s* and *s* +  $\Delta$  are transformed to  $t_s = F(s)$  and  $t_{s+\Delta} = F(s + \Delta)$  in biological time. Although the delay between the pair of events is the same in calendar time, the corresponding delay between the events in biological time,  $t_{\Delta} - t_0$  and  $t_{s+\Delta} - t_s$  are generally different.

We assume that the biological clock is positive quadrant dependent (Esary et al., 1967). Intuitively, this means that if  $t_s$  is small, the chance increases that the subsequent interval  $t_{\Delta} - t_s$  is small too. It implies positive autocorrelation, i.e.,  $cov(t_s, t_{s+\Delta} - t_s) > 0$ . And crucially, it implies that discount factors are positive autocorrelated:  $cov(exp(-t_s), exp(-(t_{s+\Delta} - t_s))) > 0$ . We also assume that increments in biological time are stationary: the distribution of time changes



does not depend on when they occur; the distribution of  $t_{\Delta} - t_0$  is the same as that of  $t_{s+\Delta} - t_s$ . The following theorem states the main result.

# THEOREM: POSITIVE DEPENDENCE IN BIOLOGICAL TIME IMPLIES TEMPORAL CHOICE INCONSISTENCY

The rigorous proof of the theorem is provided in the Section "Appendix," that the reader is encouraged to peruse. Here we aim to provide the intuition with a simple numerical example. We envisage delivery of \$1 now or *K* dollars at  $\Delta$  (a point in calendar time). *K* is to be chosen so that there is indifference. We then compare delivery of \$1 at *s* = 2 with delivery of *K* dollars at *s* = 2 +  $\Delta$ .

The biological clock is as follows. Either *s* feels like it takes 1 unit of (biological) time, or it takes 3 units, with an average of 2 units.  $\Delta$  is on average half the length; it feels either short (0.5 units) or long (1.5), with an average of 1 unit.

The biological clock is positive quadrant dependent. We will make the dependence perfect, to simplify the argument. When *s* feels like it takes 1 unit, the subsequent  $\Delta$  will take 0.5 unit; when *s* feels long (2 units), the subsequent  $\Delta$  is long as well, at 1.5 units.

As mentioned, *K* is set so that there is indifference between getting \$1 now and *K* dollars at  $\Delta$ . Today's value of *K* is  $0.5^* exp(-0.5) + 0.5^* exp(-1.5)^* K$ , because with 50% probability, delivery is felt like taking place in 0.5 units of (biological) time, and with 50% chance, it feels like it takes place 1.5 time units in the future. We set *K* so that today's value equals \$1. So, K = 2/(exp(-0.5) + exp(-1.5)) = 2.411.

Essentially, *K* is set so that the gain of having to wait a short time (only 0.5 units of biological time) is offset by the loss in value for having to wait a long time (1.5 units). The former gain (relative to today's \$1) is  $2.411^*exp(-0.5) - 1 = 0.462$ , or 46.2%; the latter loss equals  $1 - 2.411^*exp(-1.5) = (1 - 0.538) = 46.2\%$ . The gains and losses offset.

As for delivery of *K* at  $s + \Delta$ , notice that, while the gain and loss from waiting an extra  $\Delta$  beyond *s* have equal probability of occurring, *they are discounted differently*. The gain occurs when s = 2arrives early under the biological clock (1 unit); this is weighted more heavily, because it is discounted with only exp(-1); its weighted value is  $exp(-1)^*1.462 = 0.538$ . The loss when s = 2 arrives late (3 units under the biological clock) is weighted less, because it occurs farther in time; it is discounted with exp(-3); the weighted value in that case is  $exp(-3)^*0.538 = 0.027$ . In expectation, the value of receiving *K* at  $s + \Delta$  equals (0.538 + 0.027)/2 = 0.283.

Compare this to getting \$1 at s = 2. There is no loss or gain (one always gets \$1). With 50% probability, the dollar arrives early (1 unit in biological time), and its discounted value is exp(-1) = 0.368, and with 50% chance the dollar arrives late (3 units under the biological clock), and its discounted value is exp(-3) = 0.050. In expectation, the value equals (0.368 + 0.050)/2 = 0.209. This is strictly less than the value of getting *K* delayed (0.283).

Hence, while *K* was set to be indifferent between receiving \$1 now and *K* dollars at  $\Delta$ , the promise of *K* at *s* +  $\Delta$  is worth more than getting \$1 at *s*.

The astute reader will have noticed that the positive dependence of the biological clock is not really needed to get time inconsistencies. They occur also with negative dependence. However, with positive dependence, the decision maker will always prefer to *wait* when comparing options in the future for which she is indifferent now. That is, she looks more patient when deciding about payoffs in the future. This is the classical time inconsistency that has led to modeling of intertemporal choice using hyperbolic discounting. With negative dependence, the decision maker looks more patient in the immediate future than when considering options in the far future.

# **2 SIMULATIONS**

We now illustrate that hyperbolic discounting provides a good fit to the choices resulting from positive temporal dependence of biological time. To model biological time, we choose the lognormal distribution (Jaynes, 2003), which has a continuous positive support. Under the biological clock, any two time increments, such as  $t_{\Delta} - t_0$ and  $t_{s+\Delta} - t_s$  are jointly lognormal. When the time increments are positively correlated, they will also be positive quadrant dependent (this follows from Pitt, 1983), and hence, the corresponding discount factors will be positively correlated as well.

To make the example concrete, consider the choice between a payoff of 1 at time 0, and 1 + *K* at a delay (in calendar time) of  $\Delta = 0.5$ . We compare this to a pair of later options equally distanced in calendar time, at s = 2 and at  $s + \Delta = 2.5$ . In biological time, these events are at  $t_0 = 0$  and  $t_{\Delta}$  for the first pair,  $t_s$  and  $t_{s+\Delta}$ , for the second pair. The time of payoff delivery under the biological clock is a random variable.

We obtain the values of the options by Monte Carlo sampling. To generate the samples, we consider *n* increments in biological time,  $\tau_1, \ldots, \tau_n$ , that correspond to time increments of 0.5 units in calendar time. The *n* increments in biological time are drawn from a multivariate lognormal distribution with common mean 0.5 and unit variance. The correlations between the increments are positive, but decrease exponentially as they are farther apart in time. We encode the correlation structure as a covariance matrix with diagonals equal to 1, and covariances equal to  $\rho$ ,  $\rho^2, \ldots, \rho^{n-1}$  in the off-diagonal spots (see Appendix). We obtain instances in biological time by adding increments:  $t_{\Delta} = \tau_1$ ,  $t_s = \sum_{i=1}^4 \tau_i$ , and  $t_{s+\Delta} = \sum_{i=1}^5 \tau_i$ . These formulae reflect the fact that the biological expiration time of the delayed early option occurs after one increment, while the two later options mature after four and five increments, respectively.

Time perception and hyperbolic discounting

For the first option pair, the value of the immediate option is 1, and the option with payoff at (calendar time)  $\Delta = 0.5$ , is valued at  $E[e^{-t_{\Delta}}(1+K)]$ , assuming a unit discount rate (in biological time). Monte Carlo analysis based on  $N(=10^6)$  samples of  $t_{\Delta}$  estimates this to be  $(1/N)\sum_{i=1}^{N} e^{-t_{\Delta}^i}(1+K)$ . For K = 0.78, we find that  $E[e^{-t_{\Delta}}(1+K)] \approx 1$ , i.e., the decision maker is approximately indifferent between immediate delivery of \$1 and a payoff of \$1.78 after a delay of  $\Delta = 0.5$  units of calendar time.

For the pair of options at the more distant future, the values of the early and later options are estimated to remain approximately equal when time increments are independent, i.e.,  $\rho = 0$   $(E[e^{-t_s}(1)] \approx E[e^{-t_{s+\Delta}}(1+K)] \approx 0.056)$ . When time increments are positively correlated, i.e.,  $\rho > 0$ , the later option is valued more highly, in accordance with our Theorem. For example, when  $\rho = 0.5$ , the early option has value  $E[e^{-t_s}(1)] \approx 0.091$ , and the later option has value  $E[e^{-t_{s+\Delta}}(1+K)] \approx 0.104$ . So the decision maker prefers to wait to receive \$1.78 later, while he was indifferent between an immediate \$1 and \$1.78 after an equally long delay of  $\Delta = 0.5$ . We thus have obtained a temporal inconsistency.

We can obtain a discounting curve by evaluating payoffs at various delays, as in the previous example. We generated n(=10) increments,  $\tau_1, \ldots, \tau_n$ , of length  $\Delta = 0.5$ . The time indicated by the biological clock at calendar time *s* is given by  $t_s = \sum_{i=1}^{s} \tau_i$ . The value of a payoff of \$1 at time 0 is 1, and at (calendar) time *s* is  $E[e^{-t_i}(1)]$  (we continue to use a unit discount rate.) The values obtained for each point in calendar time can then be compared to valuation with hyperbolic discounting (in calendar time), assuming a discount factor is 1/(1 + ks). We find the best-fitting value of *k* by minimizing the squared error between the theoretical values under the hyperbolic function, and that generated by our Monte Carlo procedure. Similarly, we can also obtain a comparison with valuation (in calendar time) assuming exponential discounting, where the discount factor equals  $e^{-\delta s}$ . The discount rate  $\delta$  is also obtained by minimizing the squared error.

We are interested, in particular, in the effect of the correlation parameter  $\rho$  of the biological clock on the shape of the discounting curve. When autocorrelation equals zero ( $\rho = 0$ ), the discounting curve is pretty much exponential, as shown in **Figure 2** (Top). The best-fitting exponential curve has  $\delta = 0.3$ ; this differs from the true discount rate (1) because the latter only applies to biological time. The hyperbolic curve has k = 2.75, but its fit is worse. **Figure 2** (Bottom) illustrates how, when the autocorrelations are very high ( $\rho = 0.97$ ), hyperbolic discounting provides the better fit. The bestfitting exponential discount rate equals  $\delta = 0.45$ , and the hyperbolic discount rate is estimated at k = 1.75.

Variability in the mapping from calendar to biological time also plays a role. In the limit, when the biological clock is accurate (i.e., the mapping is deterministic, and, because we assume a constant speed for the biological clock, linear), we of course obtain exponential discounting in calendar time. Section A3 in Appendix shows how variability induces increased convexity in the discounting curve.

# **3 DISCUSSION**

Time inconsistencies, like the ones that led to modeling time preferences with hyperbolic discounting, arise when the biological clock advances randomly in calendar time, and increments in biological



FIGURE 2 | (Top) No autocorrelation of biological time ( $\rho = 0$ ). The discounting curve in biological time is exponential with discount rate equal to 1. It generates the dotted discounting curve in calendar time ("ZeroCorr"). The best exponential ("Exp") fit in calendar time produces a discount rate of 0.30, and the best hyperbolic fit ("Hyp") has a discount rate equal to 2.75. (Bottom) Very high autocorrelation of biological time ( $\rho = 0.97$ ). The discounting curve in biological time ( $\rho = 0.97$ ). The discounting curve in biological time (r = 1.1 to generate the dotted discounting curve in calendar time ("HighCorr"). The best exponential ("Exp") fit in calendar time ("HighCorr"). The best exponential ("Exp") fit in calendar time produces a discount rate of 0.45, and the best hyperbolic fit ("Hyp") in calendar time has discount rate equal to 1.75.

time are positively dependent. When measured in calendar time, discounting becomes increasingly hyperbolic as the biological clock becomes more highly correlated and more variable.

Prior to our result, hyperbolic discounting emerged in a normative (i.e., fully rational) model because discount rates were assumed to be stochastic (Farmer and Geneakopolos, 2009). Our rational explanation of hyperbolic discounting does not rely on random discounting, but on randomness in the transformation between calendar time (which determines payoff times) and biological time (which is relevant for decision making). The two explanations can be shown to be related mathematically, but they are biologically very different. Specifically, stochastic time perception is biologically plausible, while stochastic discounting is rarely considered outside the arcane world of mathematical finance. An exception is Skog (1997). Other rational explanations of hyperbolic discounting focus on specific forms of uncertainty about the ability of the payer to deliver the future payment (because he is bankrupted) or of the payee to receive it (because she may have deceased beforehand). When the hazard rate is stochastic, the apparent discount rate can be shown to become stochastic (see Sozou, 1998; Azfar, 1999; Sozou and Seymour, 2003). However, payment uncertainty cannot provide a comprehensive explanation. In particular, it fails to explain hyperbolic discounting in experiments where design precludes bankruptcy and where the time horizon is too short for significant effects from sudden inability of the payee to take delivery (e.g., Kable and Glimcher, 2007).

Our explanation of hyperbolic discounting assumes that discounting is exponential in biological time. Consistent with this, temporal discounting has empirically been shown to have an exponential form when subjective estimates of time elapsed are taken into account (Zauberman et al., 2009). Other work has shown that discounting is hyperbolic if subjects perceive realizations of future events to be uncertain (Dasgupta and Maskin, 2005).

The importance of perceived time in discounting has been pointed out before (Kim and Zauberman, 2009; Nakahara and Kaveri, 2010), but because random time changes were never considered, some type of misperception had to be invoked to generate hyperbolic discounting and the associated choice inconsistencies. Specifically, in Kim and Zauberman (2009), Nakahara and Kaveri (2010), the mapping from calendar to biological time is concave, so that increments farther in calendar time are expected (under the biological clock) to become shorter, inconsistent with the actual experience once the future arrives. In contrast, in our case, the speed of the biological clock is in tune with the calendar time, on average. Our approach relies on variability in the estimates.

Still, we can emulate the misperception of Kim and Zauberman (2009), Nakahara and Kaveri (2010) by increasing the expected speed of the biological clock for time farther into the future, or equivalently, decreasing the drift in the mapping from calendar to biological time. This generates concavity in the (random) mapping from calendar to biological time, and hyperbolic discounting adequately captures the resulting intertemporal choices; see Section A4 in Appendix. Positive dependence in the biological clock is no longer needed; nor is variability. As such, our framework encompasses explanations that rely on concavity in the mapping from calendar to biological time.

Stochasticity in time perception has long been accepted in psychology. Gibbon et al. (1984), e.g., uses a random clock process to explain response accuracy in animal timing tasks. Consideration

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of temporal dependence of the biological clock is novel, however, and may elucidate timing anomalies that an independent clock cannot explain (Machado and Keen, 1999). Positive temporal correlation in the internal clock is neurobiologically plausible; it may be supported by the positive autocorrelations recently discovered in human brain activity oscillations, displaying slow decreases even over thousands of cycles (Linkenkaer-Hansen et al., 2001, 2004), not unlike those generated by fractional Brownian motions (Mandelbrot and Van Ness, 1968). It is unknown to what extent this generalizes to longer time horizons, however.

Our theoretical framework provides a potentially unifying account for recorded time preferences. This is fortunate, because hyperbolic discounting is known to not be universal, with the shape (and level) of discounting changing with context (Scholten and Read, 2010). Context-dependence is consistent with our theory, which implies exponential discounting when the speed of the internal clock is expected to be constant, and the relation between calendar time and biological time accurate, or increments in biological time uncorrelated. Hyperbolic discounting emerges when the biological clock exhibits temporal dependence, or when its speed is expected to decrease in the more distant future. Future research should clarify which features of the biological clock can account for the observed context-dependence of discounting. In Scholten and Read (2010), intertemporal preferences were observed to be different depending on whether a time interval is divided up, or time is extended by adding intervals. In principle, our theory could accommodate such differences, but it may require the biological clock to not be self-similar (Mandelbrot and Van Ness, 1968); that is, its temporal properties may have to change as one moves from coarser to finer sub-divisions of time.

Our theorem provides a new, unifying framework to study time perception and how it relates to impulsivity in temporal decision making (Wittmann and Paulus, 2008). Our linking the phenomena of biological time and intertemporal discounting should lead to novel studies of the symptoms and causes of many disorders involving anomalous time perception, such as attention-deficit hyperactivity syndrome, borderline personality disorder, anxiety disorder, and schizophrenia.

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# **APPENDIX**

# A1. PROOF OF THE MAIN THEOREM

### Theorem: positive correlation in biological time implies temporal choice inconsistency

*Proof.* To prove this theorem, we start with setting the exponential discount rate (in biological time) equal to 1. This is without loss of generality; any other discount rate would work. Now pick an amount *K* so that the decision maker is indifferent between an immediate (at time 0) payoff of 1 and a payoff of 1 + *K* after a delay  $\Delta(>0)$ . Again without loss of generality, we set the initial biological time  $t_0 = 0$  (although at times we will keep  $t_0$  explicit, for clarity). Let  $t_{\Delta}$  be the time that will have passed according to the biological clock by the time the calendar clock indicates  $\Delta$ . Because of the preference of our decision maker, the valuations corresponding to the immediate option (left-hand side) and to the delayed option (right-hand side) are equal:

$$1 = E[e^{-(t_{\Delta} - t_0)}(1 + K)]$$
(1)

Now consider the valuation of 1 at some later time *s*, as well as that of 1 + K at the same time *s* plus the delay  $\Delta$ . The corresponding times according to the biological clock are  $t_s$  and  $t_{s+\Delta}$  respectively. The increment from *s* to  $s + \Delta$  equals  $t_{s+\Delta} - t_s$  in biological time. The (time 0) value of the payoff of 1 at (calendar time) *s* equals  $E[e^{-t_s}]$ , and the value of the payoff of 1 + K at  $s + \Delta$  equals

$$E[e^{-t_{s+\Delta}}(1+K)] = E[e^{-t_s}e^{-(t_{s+\Delta}-t_s)}(1+K)]$$
(2)

We assume that calendar time increments are perceived to be positive quadrant dependent. Hence,  $cov(e^{-t_s}, e^{-(t_{s+\Delta}-t_s)}) > 0$ , or, applying the definition of covariance,

$$cov(e^{-t_s}, e^{-(t_{s+\Delta}-t_s)}) = E[e^{-t_s}e^{-(t_{s+\Delta}-t_s)}] - E[e^{-t_s}]E[e^{-(t_{s+\Delta}-t_s)}] > 0.$$
(3)

We can use the latter inequality to obtain a lower bound for the value of the later option:

$$E[e^{-t_s}e^{-(t_{s+\Delta}-t_s)}(1+K)] > E[e^{-t_s}]E[e^{-(t_{s+\Delta}-t_s)}(1+K)].$$
(4)

Our assumption that time increments are stationary implies, in particular, that

$$E[e^{-(t_{s+\Delta}-t_s)}(1+K)] = E[e^{-(t_{\Delta}-t_0)}(1+K)].$$
(5)

But we picked K so that the latter equals 1. Combining this with the above, we conclude that the later option is worth more than:

$$E[e^{-t_s}e^{-(t_{s+\Delta}-t_s)}(1+K)] > E[e^{-t_s}].$$
(6)

But the right-hand side is the value of the earlier option. As a result, the decision maker is no longer indifferent between the earlier and later options as she was when the earlier option was immediate; she now strictly prefers the later option, which is a time inconsistency.  $\Box$ 

# A2. GENERATING TIME INCREMENTS

Biological time increments  $\tau_1, \tau_2, ..., \tau_n$  are generated according to a multivariate lognormal distribution. The mean of the time intervals is fixed at 0.5. The covariance matrix encodes first-order autocorrelation, with correlation parameter  $\rho$ .

$$\begin{bmatrix} \tau_{1} \\ \tau_{2} \\ \vdots \\ \tau_{3} \end{bmatrix} = logN \begin{pmatrix} 0.5 \\ 0.5 \\ \vdots \\ 0.5 \end{bmatrix}, \begin{bmatrix} 1 & \rho & \cdots & \rho^{n-1} \\ \rho & 1 & \cdots & \rho^{n-2} \\ \vdots & & \ddots & \vdots \\ \rho^{n-1} & \rho^{n-2} & \cdots & 1 \end{bmatrix} \end{pmatrix}$$
(7)

The biological time after *s* increments is given by:  $t_s = \sum_{i=1}^{i=s} \tau_i$ .

# A3. INCREASED RANDOMNESS IN THE MAPPING FROM CALENDAR TO BIOLOGICAL TIME

Simulations are performed as for **Figure 2**. We set the autocorrelation of the biological clock ( $\rho$ ) equal to 0.3, and increase variability (variances of the multivariate lognormal distribution of biological time intervals) from 0.50 to 1 and 4 (the middle case is the value used to generate **Figure 2**). Variability generates increased convexity in the discounting curve in calendar time. See **Figure A1**.

# A4. CONCAVE MAPPING FROM CALENDAR TIME TO BIOLOGICAL TIME

The time increments  $\tau_1, \tau_2, ..., \tau_n$  are generated according to a multivariate lognormal distribution. To generate concavity (on average) in the (stochastic) mapping from calendar to biological time, we let the drift in the mapping from calendar to biological time decrease with time. The expected length of the *k*th interval in biological time equals:  $E[\tau_k] = (0.5k)^{\gamma} - (0.5(k-1))^{\gamma}$ , where  $\gamma < 1$  (e.g.,  $\gamma = 0.5$ ). The covariance matrix is identity, thus assuming no correlation across increments. As before, the biological time after *s* increments is given by:  $t_s = \sum_{i=1}^{i=s} \tau_i$ .

$$\begin{bmatrix} \tau_1 \\ \tau_2 \\ \vdots \\ \tau_n \end{bmatrix} = logN \left( \begin{bmatrix} 0.5^{\gamma} \\ 1^{\gamma} - 0.5^{\gamma} \\ \vdots \\ (0.5n)^{\gamma} - (0.5(n-1))^{\gamma} \end{bmatrix}, \begin{bmatrix} 1 & 0 & \cdots & 0 \\ 0 & 1 & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & 1 \end{bmatrix} \right)$$
(8)

We set  $\gamma = 0.5$ . In **Figure A2** we find that the discounting curve generated by the concavity in the relation between calendar and biological time is best modeled as hyperbolic.



FIGURE A1 | Impact of increased variability onto shape of discounting under mild autocorrelation of the biological clock ( $\rho$  = 0.3). As variability ("Var") increased from 0.5 over 1 to 4, the discounting function becomes more convex.

