

# 

**Citation:** Duan H, Deng T, Chen Y, Zhao Z, Wen Y, Chen Y, et al. (2018) Association between vasectomy and risk of testicular cancer: A systematic review and meta-analysis. PLoS ONE 13(3): e0194606. https://doi.org/10.1371/journal. pone.0194606

Editor: Sabine Rohrmann, University of Zurich, SWITZERLAND

Received: November 5, 2017

Accepted: March 6, 2018

Published: March 22, 2018

**Copyright:** © 2018 Duan et al. This is an open access article distributed under the terms of the Creative Commons Attribution License, which permits unrestricted use, distribution, and reproduction in any medium, provided the original author and source are credited.

**Data Availability Statement:** All relevant data are within the paper.

**Funding:** This work was financed by grants from National Natural Science Foundation of China (NO. 81370804 and NO. 81670643) (GZ), Guangzhou Science Technology and Innovation Commission (NO. 201604020001 and NO. 201704020193) (GZ), and the China Postdoctoral Science Foundation (NO. 2017M612636) (TD).

**Competing interests:** The authors have declared that no competing interests exist.

**RESEARCH ARTICLE** 

# Association between vasectomy and risk of testicular cancer: A systematic review and meta-analysis

Haifeng Duan<sup>1,2°</sup>, Tuo Deng<sup>1,2°</sup>, Yiwen Chen<sup>1,2°</sup>, Zhijian Zhao<sup>1,2</sup>, Yaoan Wen<sup>1,2</sup>, Yeda Chen<sup>1,2</sup>, Xiaohang Li<sup>1,2</sup>, Guohua Zeng<sup>1,2</sup>\*

1 Department of Urology, Minimally Invasive Surgery Center, The First Affiliated Hospital of Guangzhou Medical University, Guangzhou, Guangdong, China, 2 Guangdong Key Laboratory of Urology, Guangzhou, Guangdong, China

So These authors contributed equally to this work.

\* gzgyzgh@vip.sina.com

# Abstract

# Objectives

A number of researchers have reported that vasectomy is a risk factor for testicular cancer. However, this conclusion is inconsistent with a number of other published articles. Hence, we conducted this meta-analysis to assess whether vasectomy increases the risk of testicular cancer.

# Materials and methods

We identified all related studies by searching the PubMed, Embase, and Cochrane Library database from January 01, 1980 to June 01, 2017. The Newcastle-Ottawa Scale (NOS) checklist was used to assess all included non-randomized studies. Summarized odds ratios (ORs) and 95% confidence intervals (CIs) were used to assess the difference in outcomes between case and control groups. Subgroup analyses were performed according to the study design and country.

# Results

A total of eight studies (2176 testicular cancer patients) were included in this systematic review and meta-analysis. Six articles were case-control studies, and two were cohort studies. The pooled estimate of the OR was 1.10 (95% CI: 0.93-1.30) based on the eight studies in a fixed effects model. Two subgroup analyses were performed according to the study design and country. The results were consistent with the overall findings. Publication bias was detected by Begg's test and Egger's test and *p* values > 0.05, respectively.

# Conclusions

Our meta-analysis suggested that there was no association between vasectomy and the development of testicular cancer. More high-quality studies are warranted to further explore the association between vasectomy and risk of testicular cancer.

# Introduction

Testicular cancer accounts for 1% of male tumors[1]. This cancer is the most common cancer for men aged 20 to 39 years[2]. Testicular cancer's incidence and mortality rate have been gradually increasing. Testicular cancer caused 8300 deaths worldwide in 2003 and 7000 deaths in 1990, an increase of 19%[3]. Ghazarian et al[4] showed that the incidence of testicular cancer continues to increase in the United States. Vasectomy is a simple surgical procedure that can be performed with local lidocaine anesthesia in the outpatient. Vasectomy is the most effective permanent form of contraception available to men. It is reported that the prevalence of vasectomy was approximately 8% in China[5]. New Zealand, Canada and the United Kingdom have a higher prevalence of vasectomy: 25% of the married men and 57% in the age group of 40–49[6,7].

However, several studies showed that vasectomy was associated with testicular cancer[8,9] and prostate cancer[10,11]. This association especially caught the public's attention because of the extensive use of vasectomy. Recently, Liu et al[12] published a meta-analysis showing that vasectomy may not contribute to the risk of prostate cancer. The relation between vasectomy and testicular cancer has not been thoroughly elucidated to date. Several studies have reported a positive association between vasectomy and testicular cancer[8,9]. Other studies found the opposite results[13,14,15,16,17,18,19]. However, the number of patients in these studies are small, and the level of evidence is not high.

Therefore, we conducted this systematic review and meta-analysis to assess the association between vasectomy and risk of testicular cancer, aiming to provide further evidence and guidelines for the general public.

## Materials and methods

#### Search strategy

A comprehensive electronic literature search of Pubmed, Embase, and Cochrane Library databases was conducted by two members to identify studies which assessed the association between vasectomy and the risk of testicular cancer. Separate searches were performed with the following search terms: ('vasectomy' or 'deferentectomy' or 'vasoligation' or 'vasoligature') AND ('testicular neoplasms' or 'testicular cancer'). The detailed retrieval process is shown in S2 File. References of all included studies were also checked for potential papers. This search was repeated until no additional articles were found. The full date range for the search is from January 01, 1980 to June 01, 2017.

#### Study selection, inclusion and exclusion criteria

We performed an initial screening based on the titles and abstracts. If uncertain, a subsequent full-text assessment was conducted. The second screening was a full-text review. Populations for this review and meta-analysis were broadly inclusive, involving any country and race. Studies were included if they met the following criteria: (i) the study design was a cohort, case-control study or randomized controlled trial; (ii) contained vasectomy and testicular cancer measures; (iii) studies that reported the odds ratio (OR) or relative risk (RR) relating vasectomy to testicular cancer outcome and the corresponding 95% confidence interval (CI) (or sufficient data to calculate them); and (iv) the publication language was confined to English. Exclusion criteria were (i) duplicate literature; (ii) reviews, meeting summaries, and editorials comments; and (iii) studies that had no control group.

#### Study quality assessment and data extraction

Two researchers (Duan and Chen) independently extracted data from all eligible studies. Any differences were discussed or decided by the third reviewers (Deng). Data extraction was performed using a standardized data collection form. The following data were collected from each of the selected articles: the name of the first author, data source, study type, age, study period, publication year, length of follow-up, the number of patients and participants, OR, RR, 95% CIs and statistical adjustments for confounding factors.

The Newcastle-Ottawa Scale (NOS) was applied to assess the quality of all studies. The NOS checklist contains three parameters of quality: (i) selected population, (ii) comparability of groups, and (iii) assessment of either the exposure or outcome of interest for case-control or cohort studies. Each study was assigned a score of 0-9. The studies scored greater than or equal to 7 were considered to be high quality articles.

#### **Statistics analysis**

In the meta-analysis, RevMan analytical software package (Version 5.3, Cochrane Collaboration, Oxford, UK) was used to combine the extracted data. The pooled OR or RR and the corresponding 95% CIs were calculated to assess the relationship between vasectomy and risk of testicular cancer. Heterogeneity was assessed using the chi-square test based Q- and  $I^2$ - statistic. If heterogeneity was not present (P > 0.10,  $I^2 < 50\%$ ), a fixed-effect model was used to calculate the combined OR values. Otherwise, a random-effect model was used. Subgroup analysis was performed according to the study design and country. All results in this analysis were considered as significant only with a two-tailed P < 0.05. Both the Begg's test and the Egger's test were performed using Stata 12.1 (Stata Corp., College Station, TX) to determine whether publication bias existed.

# Results

After screening the manuscripts, we included eight articles (Strader et at., 1988; Forman et at., 1994; Moss et al., 1986; Brown et at., 1987; Rosenberg et at., 1994; Swerdlow et at., 1987; Nienhuis et at., 1992; Eisenberg et at., 2015). The detailed retrieval process is shown in Fig 1. The characteristics of qualified studies are listed in Table 1. In total, 11,141 participants in case-control studies and 908,927 in cohort studies (a total of 2176 testicular cancer patients were included in this systematic review and meta-analysis. Six articles were case-control studies, and two were cohort studies. Five studies were conducted in the United States of America (USA), and three were in England.

As shown in Table 2, according to the NOS checklist, six studies with scores  $\geq$  7 stars were considered high quality, and the remaining two studies were medium quality for 5 and 6 stars, respectively.

## Outcomes

Overall, the pooled estimate of the OR was 1.10 (95% CI: 0.93–1.30, P = 0.28) based on eight studies in a fixed effects model with no significant heterogeneity ( $I^2 = 33\%$ , P = 0.15) (Fig 2). One of our subgroup analyses was carried out according to research types.

The pooled estimate of the OR was 1.04 (95% CI: 0.84–1.27, P = 0.69) based on six case-control studies in a fixed effects model, without heterogeneity ( $I^2 = 0\%$ , P = 0.51) (Fig 3). The pooled estimate of the OR was 0.83 (95% CI: 0.24–2.84, P = 0.77) based on two cohort studies in a random effects model with significant heterogeneity ( $I^2 = 80\%$ , P = 0.03) (Fig 4)





https://doi.org/10.1371/journal.pone.0194606.g001

The other subgroup analysis was carried out based on countries. The pooled estimate of the OR was 1.15 (95% CI: 0.93–1.42, P = 0.21), based on five USA studies in a fixed effects model without significant heterogeneity ( $I^2 = 40\%$ , P = 0.151) (Fig 5). The pooled estimate of the OR was 1.02 (95% CI: 0.78–1.34, P = 0.88), based on three English studies in a fixed effects model with mild heterogeneity ( $I^2 = 40\%$ , P = 0.19) (Fig 5).

# **Publication bias**

Both the Begg's test and Egger's test were conducted to assess publication bias, indicating no significant publication bias among all the included studies (Fig 6).

# Discussion

1. To the best of our knowledge, this report presents the first systematic review and meta-analysis assessing the association between vasectomy and the risk of testicular cancer. A total of 11,141 participants in case-control studies and 908,927 in cohort studies (a total of 2176 testicular cancer patients) were included in our systematic review and meta-analysis. Our results showed that there was no significant association between vasectomy and the development of testicular cancer. Vasectomy, as a common method of permanent birth control, should not be forbidden unless there are more high quality papers reporting the positive association between

#### Table 1. The characteristics of qualified researches.

| First author                 | Publication<br>time | Country   | Study<br>design  | Age<br>(years) | Number of cases                | Number of controls                 | OR   | 95%<br>CI     | Diagnosis                        | Variable adjustment   | Quality<br>scores |
|------------------------------|---------------------|-----------|------------------|----------------|--------------------------------|------------------------------------|------|---------------|----------------------------------|---|-------------------|
| D Forman<br>[ <u>13]</u>     | 1994                | England   | Case-<br>control | 15-49          | 794                            | 794                                | 1.1  | 0.8–<br>1.5   | testicular<br>germ cell<br>tumor | Same practitioner   | 7                 |
| Moss AR<br>[14]              | 1985                | Americans | Case-<br>control | >18            | 173                            | 212                                | 0.6  | 0.3-<br>1.2   | testicular<br>germ cell<br>tumor | Friends, race, age  | 8                 |
| Brown LM<br>[ <u>15]</u>     | 1987                | American  | Case-<br>control | 18-42          | 266                            | 254                                | 1    | 0.3–<br>3.3   | testicular<br>cancer             | same hospital, other<br>malignancy, age, race,<br>vital status. | 7                 |
| Rosenberg L<br>[16]          | 1994                | American  | Case-<br>control | <70            | 132                            | 7027                               | 0.8  | 0.4–<br>1.9   | testicular<br>cancer             | no history of cancer  | 7                 |
| Strader CH<br>[8]            | 1987                | American  | Case-<br>control | 20–69          | 228                            | 513                                | 1.5  | 1.0–<br>2.2   | testicular<br>germ cell<br>tumor | Residence time,<br>education, religion.                         | 7                 |
| Swerdlow AJ<br>[17]          | 1986                | England   | Case-<br>control | >10            | 259                            | 489                                | 1.1  | 0.63-<br>2.04 | testicular<br>cancer             | town, age   | 7                 |
| First author                 | Publication<br>time | Country   | Study<br>design  | Age<br>(years) | Number of<br>exposed<br>groups | Number of<br>non exposed<br>groups | RR   | 95%<br>CI     | Diagnosis                        | Variable adjustment   | Quality<br>scores |
| Nienhuis H<br>[ <u>18]</u>   | 1992                | England   | Cohort<br>study  | 25-49          | 13246                          | 22196                              | 0.46 | 0.1–<br>1.4   | testicular<br>cancer             | age   | 6                 |
| Eisenberg<br>ML [ <u>19]</u> | 2014                | American  | Cohort<br>study  | 18-50          | 112655                         | 760830                             | 1.27 | 0.94–<br>1.73 | testicular<br>cancer             | age   | 5                 |

https://doi.org/10.1371/journal.pone.0194606.t001

#### Table 2. The Newcastle-Ottawa Scale (NOS).

| First<br>author              | Quality<br>evaluation | Case definition                      | Representativeness                | Selection of<br>Controls     | Definition<br>of Controls              | Comparability | Ascertainment<br>of exposure | Same<br>method?             | Non-<br>Response<br>rate |
|------------------------------|-----------------------|--------------------------------------|-----------------------------------|------------------------------|--|---------------|------------------------------|-----------------------------|--------------------------|
| D Forman<br>[ <u>13]</u>     | 7                     | 1                                    | 1                                 | 1                            | 1                                      | 1             | 1                            | 1                           | 0                        |
| Moss AR<br>[14]              | 8                     | 1                                    | 1                                 | 1                            | 1                                      | 1             | 1                            | 1                           | 1                        |
| Brown LM<br>[ <u>15</u> ]    | 7                     | 1                                    | 1                                 | 0                            | 1                                      | 1             | 1                            | 1                           | 1                        |
| Rosenberg<br>L [16]          | 7                     | 1                                    | 1                                 | 0                            | 1                                      | 1             | 1                            | 1                           | 1                        |
| Strader CH<br>[8]            | 7                     | 1                                    | 1                                 | 1                            | 1                                      | 1             | 1                            | 1                           | 0                        |
| Swerdlow<br>AJ [ <u>17</u> ] | 7                     | 1                                    | 1                                 | 0                            | 1                                      | 1             | 1                            | 1                           | 0                        |
| First<br>author              | Quality<br>evaluation | Representativeness of exposed cohort | Selection ofnon<br>exposed cohort | Ascertainment<br>of exposure | outcome not<br>present<br>before study | Comparability | Assessment of outcome        | follow-up<br>long<br>enough | Non-<br>Response<br>rate |
| Nienhuis H<br>[ <u>18]</u>   | 6                     | 1                                    | 1                                 | 1                            | 1                                      | 1             | 1                            | 0                           | 0                        |
| Eisenberg<br>ML [19]         | 5                     | 1                                    | 0                                 | 1                            | 1                                      | 1             | 1                            | 0                           | 0                        |

https://doi.org/10.1371/journal.pone.0194606.t002



|  | Vasect        | tomy   | No vase | ctomy  |        | Odds Ratio         | latio |                   | Ratio         |    |  |
|--|---------------|--------|---------|--------|--------|--------------------|-------|-------------------|---------------|----|--|
| Study or Subgroup  | <b>Events</b> | Total  | Events  | Total  | Weight | M-H, Fixed, 95% Cl |       | <u>M-H, Fixec</u> | <u>1, 95%</u> | CI |  |
| Brown LM [15]  | 5             | 10     | 261     | 510    | 2.0%   | 0.95 [0.27, 3.34]  | -     |                   |               | -  |  |
| D Forman[13]   | 81            | 156    | 713     | 1432   | 26.4%  | 1.09 [0.78, 1.52]  |       |                   | <u> </u>      |    |  |
| Eisenberg ML [19]  | 52            | 112655 | 251     | 760830 | 25.3%  | 1.40 [1.04, 1.89]  |       | -                 | -             |    |  |
| Moss AR [14]   | 15            | 45     | 158     | 340    | 9.6%   | 0.58 [0.30, 1.11]  | -     |                   |               |    |  |
| Nienhuis H [18]  | 4             | 13246  | 17      | 22196  | 5.0%   | 0.39 [0.13, 1.17]  |       | +                 |               |    |  |
| Rosenberg L [16]   | 7             | 438    | 125     | 6721   | 5.9%   | 0.86 [0.40, 1.85]  |       |                   |               |    |  |
| Strader CH [8]   | 46            | 134    | 182     | 607    | 16.9%  | 1.22 [0.82, 1.81]  |       | +                 |               |    |  |
| Swerdlow AJ [17]   | 22            | 58     | 237     | 690    | 8.9%   | 1.17 [0.67, 2.03]  |       | -                 |               |    |  |
| Total (95% Cl)   |               | 126742 |         | 793326 | 100.0% | 1.10 [0.93, 1.30]  |       | •                 | •             |    |  |
| Total events   | 232           |        | 1944    |        |        |                    |       |                   |               |    |  |
| Heterogeneity: Chi² = 10.43, df = 7 (P = 0.17); l² = 33% |               |        |         |        |        |                    |       |                   | <u> </u>      |    |  |
| Test for overall effect: 2                               | 0.1 0.2       | 0.5 1  | Z       | 5      | 10     |                    |       |                   |               |    |  |

#### Fig 2. Forest plot of all included studies.

https://doi.org/10.1371/journal.pone.0194606.g002

|                                     | Testicular cancer |          | Control |       | Odds Ratio |                   |     | Odds Ratio    |         |          |         |
|-------------------------------------|-------------------|----------|---------|-------|------------|-------------------|-----|---------------|---------|----------|---------|
| Study or Subgroup                   | Events            | Total    | Events  | Total | Weight     | M-H, Fixed, 95% C |     | <u>M-H, F</u> | ixed, 9 | 5% CI    |         |
| Brown LM [15]                       | 5                 | 266      | 5       | 254   | 2.8%       | 0.95 [0.27, 3.34] |     |               | -       |          | _       |
| D Forman[13]                        | 81                | 794      | 75      | 794   | 37.8%      | 1.09 [0.78, 1.52] |     |               |         | -        |         |
| Moss AR [14]                        | 15                | 173      | 30      | 212   | 13.8%      | 0.58 [0.30, 1.11] | -   | -             | +       |          |         |
| Rosenberg L [16]                    | 7                 | 132      | 431     | 7027  | 8.5%       | 0.86 [0.40, 1.85] |     |               | •       |          |         |
| Strader CH [8]                      | 46                | 228      | 88      | 513   | 24.3%      | 1.22 [0.82, 1.81] |     |               | -+      |          |         |
| Swerdlow AJ [17]                    | 22                | 259      | 36      | 489   | 12.8%      | 1.17 [0.67, 2.03] |     | _             |         |          |         |
| Total (95% CI)                      |                   | 1852     |         | 9289  | 100.0%     | 1.04 [0.84, 1.27] |     |               | •       |          |         |
| Total events                        | 176               |          | 665     |       |            |                   |     |               |         |          |         |
| Heterogeneity: Chi <sup>2</sup> = 4 | 4.26, df = 5 (P   | = 0.51); | l² = 0% |       |            |                   | +   |               |         | <u> </u> | <u></u> |
| Test for overall effect:            | Z = 0.34 (P = 0   | 0.73)    |         |       |            |                   | 0.2 | 0.5           | 1       | 2        | 5       |

#### Fig 3. Subgroup analysis of case-control studies.

https://doi.org/10.1371/journal.pone.0194606.g003

vasectomy and testicular cancer. More studies are required to further explore the association between vasectomy and risk of testicular cancer.

2. There were two studies reporting positive association between vasectomy and testicular cancer[8,9]. Strader et al[8] reported that OR was 1.5 and 95 CI was 1.0–2.2. The authors

|  | Vasectomy | y No vase    | No vasectomy |        | <b>Risk Ratio</b>   | Risk Ratio |                  |    |  |
|--|-----------|--------------|--------------|--------|---------------------|------------|------------------|----|--|
| Study or Subgroup  | Events T  | Total Events | Total        | Weight | M-H, Random, 95% CI | M-H, Ra    | <u>ndom, 95%</u> | CI |  |
| Eisenberg ML [19]  | 52 112    | 2655 251     | 760830       | 58.8%  | 1.40 [1.04, 1.89]   |            |                  |    |  |
| Nienhuis H [18]  | 4 13      | 3246 17      | 22196        | 41.2%  | 0.39 [0.13, 1.17]   |            | +                |    |  |
| Total (95% CI)<br>Total events                             | 56        | 5901<br>268  | 783026       | 100.0% | 0.83 [0.24, 2.84]   |            |                  |    |  |
| Heterogeneity: 1au <sup>2</sup> = Test for overall effect: | 0.2 0.5   | 1 2          | 5            |        |                     |            |                  |    |  |

#### Fig 4. Subgroup analysis of cohort studies.

https://doi.org/10.1371/journal.pone.0194606.g004

|  | Vasectomy   |            | No vasectomy            |        |        | Odds Ratio         | Odds Ratio         |  |  |  |
|--|-------------|------------|-------------------------|--------|--------|--------------------|--------------------|--|--|--|
| Study or Subgroup  | Events      | Total      | Events                  | Total  | Weight | M-H, Fixed, 95% CI | M-H, Fixed, 95% Cl |  |  |  |
| 1.5.1 American   |             |            |                         |        |        |                    |                    |  |  |  |
| Brown LM [15]  | 5           | 10         | 261                     | 510    | 2.0%   | 0.95 [0.27, 3.34]  |                    |  |  |  |
| Eisenberg ML [19]  | 52          | 112655     | 251                     | 760830 | 25.3%  | 1.40 [1.04, 1.89]  |                    |  |  |  |
| Moss AR [14]   | 15          | 45         | 158                     | 340    | 9.6%   | 0.58 [0.30, 1.11]  |                    |  |  |  |
| Rosenberg L [16]   | 7           | 438        | 125                     | 6721   | 5.9%   | 0.86 [0.40, 1.85]  |                    |  |  |  |
| Strader CH [8]   | 46          | 134        | 182                     | 607    | 16.9%  | 1.22 [0.82, 1.81]  | +                  |  |  |  |
| Subtotal (95% CI)  |             | 113282     |                         | 769008 | 59.7%  | 1.15 [0.93, 1.42]  | •                  |  |  |  |
| Total events   | 125         |            | 977                     |        |        |                    |                    |  |  |  |
| Heterogeneity: Chi <sup>2</sup> = 6  | 6.68, df =  | 4 (P = 0.1 | 5); l <sup>2</sup> = 40 | )%     |        |                    |                    |  |  |  |
| Test for overall effect: 2   | Z = 1.27 (  | P = 0.21)  |                         |        |        |                    |                    |  |  |  |
| 1.5.2 England  |             |            |                         |        |        |                    |                    |  |  |  |
| D Forman[13]   | 81          | 156        | 713                     | 1432   | 26.4%  | 1.09 [0.78, 1.52]  |                    |  |  |  |
| Nienhuis H [18]  | 4           | 13246      | 17                      | 22196  | 5.0%   | 0.39 [0.13, 1.17]  |                    |  |  |  |
| Swerdlow AJ [17]   | 22          | 58         | 237                     | 690    | 8.9%   | 1.17 [0.67, 2.03]  |                    |  |  |  |
| Subtotal (95% CI)  |             | 13460      |                         | 24318  | 40.3%  | 1.02 [0.78, 1.34]  | <b>•</b>           |  |  |  |
| Total events   | 107         |            | 967                     |        |        |                    |                    |  |  |  |
| Heterogeneity: Chi <sup>2</sup> = 3  | 3.31, df =  | 2 (P = 0.1 | 9); l <sup>2</sup> = 40 | )%     |        |                    |                    |  |  |  |
| Test for overall effect: 2   | Z = 0.15 (  | P = 0.88)  |                         |        |        |                    |                    |  |  |  |
| Total (95% CI)   |             | 126742     |                         | 793326 | 100.0% | 1.10 [0.93, 1.30]  | •                  |  |  |  |
| Total events   | 232         |            | 1944                    |        |        |                    |                    |  |  |  |
| Heterogeneity: Chi <sup>2</sup> = 1  | 10.43, df = | 7 (P = 0.  | 17); l <sup>2</sup> = 3 | 33%    |        |                    |                    |  |  |  |
| Test for overall effect: Z = 1.08 (P = 0.28) 0.1 0.2 0.5 1 2 5 10                      |             |            |                         |        |        |                    |                    |  |  |  |
| Test for subgroup differences: Chi <sup>2</sup> = 0.44, df = 1 (P = 0.51), $I^2 = 0\%$ |             |            |                         |        |        |                    |                    |  |  |  |
| Fig 5. Subgroup analysis based on countries.   |             |            |                         |        |        |                    |                    |  |  |  |

https://doi.org/10.1371/journal.pone.0194606.g005

only found this association in Catholic men. A history of vasectomy was reported with approximately equal frequency by Catholics and non-Catholics in the case group. However, a great difference was reported by Catholics and non-Catholics in the control group, 6.3% VS 19.7%. The authors explained that was mainly owing to selective underreporting by



https://doi.org/10.1371/journal.pone.0194606.g006

Catholic men. In this study, there had been a extended period time between vasectomy and the study. Information bias should be controlled by querying physicians and checking the medical records, not just based on the patient;'s self—report. Cale et al[9] suggested that vasectomy could accelerate the development of testicular cancer. However, it has limited cases(37 testicular cancer patients), no well-controlled group and it remains unclear whether testicular cancer existed before vasectomy. So it could not provide good evidence. A welldesigned research is required to further explore whether vasectomy accelerates the development of testicular cancer. The largest of these case-control studies (794 testicular cancer patients) were conducted by face-to-face interviewing and showed no association between vasectomy and risk of testicular cancer[13]. The median time between case diagnosis and interview was only 10 months, and patients' self-reported medical histories were confirmed by general practitioners. That could well reduce the information bias, giving more convincing results. To get more reliable results, two cohort studies were performed [18,19]. The first was a retrospective cohort study [18], performed in 1992. A total of 13,246 men undergoing vasectomy and 22,196 comparison subjects were included, showing there was no evidence of an increase associated with vasectomy in the incidence of testicular cancer. The information was recorded according to medical records. The research should record the diseases of cryptorchidism, infertility and testicular injury, which increase the risk of testicular cancer. The other cohort study, which analyzed US claims data, was performed in 2014[19] and was the most recent of the included studies. A total of 112,655 vasectomized men and 760,830 control men were included. Few or no infertile men were included in the vasectomized men group. The research showed that vasectomy was not a risk factor for testicular cancer. In our opinion, we cannot conclude that vasectomy increased risk of testicular cancer based on the current literature.

3. Several human immune system effects caused by vasectomy were reported in a number of studies. For example, 50% of the men with vasectomy were found to have circulating spermatozoal antibodies[20]. The level of testosterone was confirmed to be unchanged after vasectomy[21]. However, all studies to date have lacked information on the mechanism by which vasectomy increases the risk of testicular and prostate cancer. The World Health Organization (WHO) meeting in 1991 concluded that there was no biological mechanism to account for any association between vasectomy and prostate cancer.

4. Potential bias was unavoidable. First, men who had a vasectomy were more proactive in seeking medical care if they sensed an abnormality in their bodies[22]. Thus, the detection of testicular cancer would be more possible for men with a vasectomy surgery compared with men who have no vasectomy. Second, cryptorchidism, infertility, exposure to organochlorine pesticides and some unreported factors could increase the risk of testicular cancer[15,23]. However, it is impossible to ensure that the risk factors for testicular cancer are evenly distributed to two groups. To reduce the bias, subgroup analysis was conducted according to research types. Neither case-control group nor cohort-study group showed positive association between vasectomy and the risk of testicular cancer. Third, the development of the testicular cancer may take a long time. The follow-up time for each study was different. If the follow-up time was not long enough, it would produce false negative results.

5. Our meta-analysis had several potential limitations. First, only British and American articles were included. This would limit the results of our research to other populations. Second, due to the limited data available in the original article, we could not perform more comprehensive and detailed subgroup analysis, such as subgroups based on the staging of testicular cancer, follow-up time, and age. Third, most of the included articles were not recent. More newly published studies are required.

# Conclusions

Our meta-analysis suggested that there is no association between vasectomy and the development of testicular cancer. More studies are required to further explore the association between vasectomy and the risk of testicular cancer.

# **Supporting information**

**S1 File. PRISMA checklist.** (DOC)

**S2** File. Detailed search criteria in PubMed. (DOCX)

**S1 Table.** The 19 references which were excluded by reading the full texts. (XLSX)

# Acknowledgments

We are thankful for the data provided by the authors of included studies.

# **Author Contributions**

Data curation: Haifeng Duan, Yiwen Chen, Yeda Chen, Xiaohang Li.

Formal analysis: Tuo Deng.

Methodology: Yaoan Wen, Yeda Chen.

Project administration: Guohua Zeng.

Resources: Haifeng Duan.

Writing - original draft: Haifeng Duan, Yiwen Chen.

Writing - review & editing: Tuo Deng, Zhijian Zhao, Yaoan Wen, Guohua Zeng.

#### References

- 1. Rosen A, Jayram G, Drazer M, Eggener SE. Global trends in testicular cancer incidence and mortality. Eur Urol. 2011; 60: 374–379. https://doi.org/10.1016/j.eururo.2011.05.004 PMID: 21612857
- Hayes-Lattin B, Nichols CR. Testicular cancer: a prototypic tumor of young adults. Semin Oncol. 2009; 36: 432–438. https://doi.org/10.1053/j.seminoncol.2009.07.006 PMID: 19835738
- GBD 2013 Mortality and Causes of Death Collaborators. Global, regional, and national age-sex specific all-cause and cause-specific mortality for 240 causes of death, 1990–2013: a systematic analysis for the Global Burden of Disease Study 2013. Lancet. 2015; 385: 117–71. https://doi.org/10.1016/S0140-6736(14)61682-2 PMID: 25530442
- Ghazarian AA, Trabert B, Devesa SS, McGlynn KA. Recent trends in the incidence of testicular germ cell tumors in the United States. Andrology. 2015; 3: 13–18. https://doi.org/10.1111/andr.288 PMID: 25331158
- 5. Hsing AW, Wang RT, Gu FL, Lee M, Wang T, Leng TJ, et al. Vasectomy and prostate cancer risk in China. Cancer Epidemiol Biomarkers Prev. 1994; 3: 285–288. PMID: 8061575
- Sneyd MJ, Cox B, Paul C, Skegg DC. High prevalence of vasectomy in New Zealand. Contraception. 2001; 64: 155–159. PMID: <u>11704094</u>
- Pile JM, Barone MA. Demographics of vasectomy—USA and international. Urol Clin North Am. 2009 36: 295–305. https://doi.org/10.1016/j.ucl.2009.05.006 PMID: 19643232
- Strader CH, Weiss NS, Daling JR. Vasectomy and the incidence of testicular cancer. Am J Epidemiol. 1988; 128: 56–63. PMID: 2837898

- Cale AR, Farouk M, Prescott RJ, Wallace IW. Does vasectomy accelerate testicular tumour? Importance of testicular examinations before and after vasectomy. BMJ. 1990; 300: 370–370. PMID: 2106990
- Mettlin C, Natarajan N, Huben R. Vasectomy and prostate cancer risk. Am J Epidemiol. 1990; 132: 1056–1061. PMID: 2260537
- Rohrmann S, Paltoo DN, Platz EA, Hoffman SC, Comstock GW, Helzlsouer KJ. Association of vasectomy and prostate cancer among men in a Maryland cohort. Cancer Causes Control. 2005; 16: 1189– 1194. https://doi.org/10.1007/s10552-005-0304-8 PMID: 16215869
- Liu LH, Kang R, He J, Zhao SK, Li FT, Wan SP, et al. Vasectomy and risk of prostate cancer: a systematic review and meta-analysis of cohort studies. Andrology. 2015; 3: 643–649. https://doi.org/10.1111/ andr.12040 PMID: 26041315
- United Kingdom Testicular Cancer Study Group. Aetiology of testicular cancer: association with congenital abnormalities, age at puberty, infertility, and exercise. BMJ. 1994; 308: 1393–1399. PMID: 7912596
- Moss AR, Osmond D, Bacchetti P, Torti FM, Gurgin V. Hormonal risk factors in testicular cancer. A case-control study. Am J Epidemiol. 1986; 124: 39–52. PMID: 2872797
- Brown LM, Pottern LM, Hoover RN. Testicular cancer in young men: the search for causes of the epidemic increase in the United States. J Epidemiol Community Health. 1987; 41: 349–354. PMID: 2901454
- Rosenberg L, Palmer JR, Zauber AG, Warshauer ME, Strom BL, Harlap S, et al. The relation of vasectomy to the risk of cancer. Am J Epidemiol. 1994; 140: 431–438. PMID: 8067335
- Swerdlow AJ, Huttly SR, Smith PG. Testicular cancer and antecedent diseases. Br J Cancer. 1987; 55: 97–103. PMID: 2880604
- Nienhuis H, Goldacre M, Seagroatt V, Gill L, Vessey M. Incidence of disease after vasectomy: a record linkage retrospective cohort study. BMJ. 1992; 304: 743–746. PMID: 1571679
- Eisenberg ML, Li S, Brooks JD, Cullen MR, Baker LC. Increased risk of cancer in infertile men: analysis of U.S. claims data. J Urol. 2015; 193: 1596–1601. https://doi.org/10.1016/j.juro.2014.11.080 PMID: 25463997
- Shulman S, Zappi E, Ahmed U, Davis JE. Immunologic consequences of vasectomy. Contraception. 1972; 5: 269–278. PMID: 4650652
- Smith KD, Tcholakian RK, Chowdhury M, Steinberger E. An investigation of plasma hormone levels before and after vasectomy. Fertil Steril. 1976; 27: 144–51. PMID: 1248661
- Cox B, Sneyd MJ, Paul C, Delahunt B, Skegg DC. Vasectomy and risk of prostate cancer. JAMA. 2002; 287: 3110–3115. PMID: 12069674
- 23. McGlynn KA, Trabert B. Adolescent and adult risk factors for testicular cancer. Nat Rev Urol. 2012; 9: 339–349. https://doi.org/10.1038/nrurol.2012.61 PMID: 22508459