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One-year mortality after recovery from critical illness: A retrospective cohort study

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Abstract

Rationale

Factors associated with one-year mortality after recovery from critical illness are not well understood. Clinicians generally lack information regarding post-hospital discharge outcomes of patients from the intensive care unit, which may be important when counseling patients and families.

Objective

We sought to determine which factors among patients who survived for at least 30 days post-ICU admission are associated with one-year mortality.

Methods

Single-center, longitudinal retrospective cohort study of all ICU patients admitted to a tertiary-care academic medical center from 2001–2012 who survived \geq 30 days from ICU admission. Cox's proportional hazards model was used to identify the variables that are associated with one-year mortality. The primary outcome was one-year mortality.

Results

32,420 patients met the inclusion criteria and were included in the study. Among patients who survived to \geq 30 days, 28,583 (88.2%) survived for greater than one year, whereas 3,837 (11.8%) did not. Variables associated with decreased one-year survival include: increased age, malignancy, number of hospital admissions within the prior year, duration of mechanical ventilation and vasoactive agent use, sepsis, history of congestive heart failure, end-stage renal disease, cirrhosis, chronic obstructive pulmonary disease, and the need for



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Abbreviations: BIDMC, Beth Israel Deaconess Medical Center; COPD, Chronic obstructive pulmonary disease; DRG, Diagnosis-related group; ESRD, End-stage renal disease; LOS, Length of stay; ICD-9, International Classification of Diseases 9th Revision; ICU, Intensive care unit; SOFA, Sequential organ failure assessment; MICU, Medical intensive care unit; MIMIC-III, Medical Information Mart for Intensive Care; MV, Mechanical Ventilation; SOFA, Sequential Organ Failure Assessment. renal replacement therapy. Numerous effect modifications between these factors were found.

Conclusion

Among survivors of critical illness, a significant number survive less than one year. More research is needed to help clinicians accurately identify those patients who, despite surviving their acute illness, are likely to suffer one-year mortality, and thereby to improve the quality of the decisions and care that impact this outcome.

Introduction

As early recognition, effective resuscitation, and technological advancements in the ongoing support of critically ill patients have improved mortality outcomes over recent decades[1–3], there has been growing interest in mortality prediction tools [4–8] in the intensive care unit (ICU) as well as an increasing focus on post-discharge outcomes[9]. One potentially important use for such tools is to reduce, and even avoid, unnecessary and undesirable interventions at the end of life. Among patients who do not survive their ICU admissions, greater than 50% have potentially life-sustaining therapies withheld or withdrawn during that final stay [10, 11]. However, intensivists frequently lack long-term follow up on patients who were "successfully" discharged from the intensive care unit.

Despite clinicians' efforts to determine which patients will have poor outcomes in order to limit futile care, a significant number of patients experience long, complicated clinical courses that are not captured by the current ability of providers to predict one-year mortality. Studies suggest that ICU survivors who die within one year of discharge have a poor quality of life prior to death[12, 13]. Prior studies have demonstrated that those with prolonged mechanical ventilation spend an average of 74% of all days alive post-discharge in a hospital, post-acute care facility or at home but requiring skilled care[12]. Additionally, numerous factors are associated with long term mortality, including ICU length of stay and severe sepsis[13–15]. Rather than dying at home, many deaths occurred either in an acute-care hospital, skilled nursing facility, or long-term care facility with each subsequent hospitalizations within the last year of life being associated with progressive disability and poor quality of life[16].

We therefore sought to determine which factors are associated with one-year mortality among patients who survived critical illness for at least 30 days. Note that our intention is not to categorize a group of patients who may be statically defined as 'chronically critically ill' based on ICU length of stay and certain clinical characteristics, but rather to describe the dynamic course of critically ill patients who survive their initial critical illness and then suffer one-year mortality, which prior data suggest is associated with poor quality of life and functional status during the remaining time alive[17]. While there will be some overlap between these two populations, we believe it is essential to point out and clarify this fundamental difference. The 'chronically critically ill' as defined by Kahn et al., represent a cohort of patients with particular conditions who have endured a defined period of critical care-the definition does not include further, longer term outcomes. In contrast, the very definition of our short-term survivors rests on two outcomes that combine (live) hospital discharge with one year mortality. We then proceed to attempt to identify the specific conditions and treatments associated with patients eventually included in this group.

Materials and methods

We conducted a longitudinal, single center, retrospective study of patients who survive at least 30 days from the time of ICU admission to determine which factors are associated with one-year mortality.

Data for this study was collected from the Medical Information Mart for Intensive Care (MIMIC-III) database[18]. All patient data were anonymized prior to extraction and data analysis. The creation, maintenance, and use of the MIMIC-III database was approved by the institutional review boards of the Massachusetts Institute of Technology (MIT: 0403000206) and Beth Israel Deaconess Medical Center (BIDMC: 2001-P-001699). The database contains records from 38,597 adult ICU patients admitted to Boston's Beth Israel Deaconess Medical Center between 2001–2012. The database includes hourly physiologic readings from bedside monitors, validated by ICU nurses. The database also contains records of demographic information, laboratory results including interpretation of imaging, clinician notes, medications, fluid balance, as well as International Classification of Diseases 9th Revision (ICD-9) and Diagnosis-Related Group (DRG) codes for each hospitalization. Sepsis was defined according to the same ICD-9 criteria as those presented in Angus et al.[19] because MIMIC-III antedates the introduction of sepsis-specific codes into the ICD-9 nomenclature. Most relevant to this project, post-discharge survival is captured using data from the Social Security Death Records, which are updated annually.

Definition of the cohort

Eligibility was limited to those patients who were over 15 years of age at ICU discharge and survived at least 30 days from the date of ICU admission.

Outcome measurement

The primary outcome was one-year mortality, measured from the patient's admission date. For patients with multiple ICU admissions, the most recent admission was included in the analysis. A post-hoc sensitivity analysis was performed using five-year mortality from admission date as the primary outcome.

Statistical analysis

The baseline characteristics of the population were examined using bivariate analyses for continuous and categorical variables using t-tests and chi-square tests, respectively. Non-parametric methods including the Mann-Whitney U test were used for continuous variables that were not normally distributed, as assessed by Q-Q plots, Kolmogorov–Smirnov tests, and visual inspection of the distributions.

A multivariate Cox's proportional hazards model was used to assess survival. For our study, survival was censored at one year, and the assumption of proportional hazards was assessed for each variable. No variable failed to meet the assumption of proportional hazards, which was assessed using log-log plots, Kaplan Meier curves, time-dependent covariate assessments, and Schoenfeld residuals. See the online data supplement <u>S1 Text</u> and <u>S1-S23</u> Figs for descriptions and relevant plots for the included variables.

Candidate variables, which were selected based on previous literature, clinician input, and biologic plausibility, included patient demographics, co-morbidities as captured by ICD-9 diagnoses, interventions received (e.g. mechanical ventilation, use of vasoactive agents, renal replacement therapy, tracheostomy), severity and trajectory of illness as captured by Sequential Organ Failure Assessment (SOFA) scores[20], presence or absence of sepsis, development of

acute kidney injury as defined by the Acute Kidney Injury Network [21], hospitalization/ICU admission within the previous year and ICU type. The change in SOFA score was calculated by subtracting the SOFA score from days 3 and 2 from the SOFA score on day 1. Days on ventilation and age were modeled as continuous variables, and the assumption of proportional hazards was validated using Schoenfeld's residuals in a smoothing cubic spines model. Vasoactive agent use–including norepinephrine, epinephrine, phenylephrine, isoproterenol, dopamine, dobutamine, vasopressin and milrinone–was initially considered as a binary variable, but this did not meet the assumption of proportional hazards. We instead included vasoactive exposure as a continuous variable (i.e. days on vasoactive agents. Urgency of ICU admission (elective, urgent/emergent) was included as a categorical variable. Wald tests were used to assess the impact of two-way interaction terms on a reasonable set of pre-specified, clinically relevant interaction terms. The set of variables that created the interaction terms included ventilation, sepsis, RRT, liver disease, and ESRD.

All reported p-values were rounded to four decimal places. Data were analyzed using SAS version 9.3. All statistical tests and/or confidence intervals, as appropriate, were performed at $\alpha = 0.05$ (2-sided).

Results

As noted in Fig 1, 32,420 patients who survived at least 30 days after ICU admission were included in the cohort. Of these, 3,837 patients (11.8%) survived less than one year. In univariate analyses, patients with one-year mortality had higher co-morbid rates of congestive heart failure (38.96% vs. 20.16%, p<0.0001), chronic obstructive pulmonary disease (COPD) (20.80% vs. 10.04%, p<0.0001), and end-stage renal disease (ESRD) (4.90% vs. 1.89%, p<0.0001). Those with one-year mortality were more likely to require initiation of hemodialysis (7.51% vs. 3.06%, p<0.0001) and more likely to be diagnosed with sepsis (39.95% vs. 22.36%, p<0.0001). Mechanical ventilation was more common among those who survived greater than one year (43.75% vs. 35.73%, p<0.0001), however the duration of mechanical ventilation was shorter (median duration 1 day vs. 2 days, p<0.0001). See Table 1 for details.

Using a Cox proportional hazards model, numerous variables were associated with oneyear mortality. The variables most strongly associated with one-year mortality were malignancy (HR = 2.53, p<0.0001), need for renal replacement therapy (RRT) (HR = 1.79, p<0.0001), and urgent/emergent ICU admission (HR = 1.64, p<0.0001). Other factors associated with one-year mortality are ESRD, Cirrhosis, frequency of admissions in the prior year, ICU type, Sepsis, COPD, diabetes, cardiac dysrhythmias, coagulopathy, and the duration of mechanical ventilation. The interaction term RRT*Cirrhosis (HR = 0.51, p = 0.0063) was statistically significantly associated with one-year mortality. Please see online data supplement and Table 2 for further details of the Cox proportional hazards model.

We additionally performed a post-hoc analysis using five-year mortality as the outcome of a Cox proportional hazards model. RRT (HR = 2.13, p<0.0001), malignancy (HR = 1.95, p<0.0001), ESRD (HR = 1.62, p<0.0001), COPD (HR = 1.53, p<0.0001) and Cirrhosis (HR = 1.52, p<0.0001) were most highly associated with five-year mortality. Please see the <u>S1</u> Table for details.

Discussion

Our study identified numerous acute diagnoses and comorbidities associated with one-year mortality after initial survival of critical illness. The presence of malignancy, need for renal replacement therapy and urgent/emergent admission are most strongly associated with one-year mortality among 30-day survivors. Additionally, ESRD, sepsis, COPD, Cirrhosis,



Fig 1. Flow diagram showing initial selection of cohort and excluded patients.

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Diabetes, prior hospital admissions, and duration of mechanical ventilation are all associated with one-year mortality among 30-day survivors. As expected, many of the variables associated with one-year mortality are likewise associated with five-year mortality in our proportional hazards model.

Additionally, those admitted to a MICU had lower one-year survival as compared to patients admitted to other ICU types (CCU, CSRU, or SICU). We postulate that patients with numerous co-morbidities and higher level of illness are less likely to be placed in specialty-specific intensive care units. In addition, due to their underlying medical problems, these patients

Table 1. Comparison based on length of survival $^{\ast}\dagger.$

	One-year mortality	Survival > One-year	P-value
N	3837 (11.8)	28583 (88.2)	
Age, years	72.14	61.22	< 0.0001
Male gender	2074 (54.05)	16540 (57.87)	< 0.0001
Co-Morbid Conditions			
Congestive Heart Failure	1495 (38.96)	5763 (20.16)	< 0.0001
Cirrhosis	285 (7.43)	1609 (5.63)	< 0.0001
Cerebrovascular Accident	179 (4.67)	1424 (4.98)	0.428
COPD	798 (20.80)	2871 (10.04)	< 0.0001
ESRD	188 (4.90)	540 (1.89)	< 0.0001
Obesity	121 (3.15)	1718 (6.01)	< 0.0001
Hypertension	1388 (36.17)	12961(45.35)	< 0.0001
Diabetes	1084 (28.25)	6992 (24.46)	< 0.0001
Length of Stay			
ICU LOS, days	2.28 [1.32-4.34]	1.97 [1.15–3.48]	< 0.0001
Admission within 1 year prior‡			
Hospital	0.30±0.72	0.12±0.43	< 0.0001
ICU	0.39±0.85	0.18±0.54	< 0.0001
SOFA Score			
Day 1 score	2.47±2.50	2.74±2.71	0.0003
Day 2 score—Day 1 score ^ð	-0.53±1.94	-0.98±2.27	< 0.0001
Day 3 score—Day 1 score $^{\Psi}$	-0.84±2.39	-1.23±2.76	< 0.0001
ICU Type			
MICU	1,848 (48.16)	8,901 (31.14)	< 0.0001
CCU	626 (16.31)	3,932 (13.76)	< 0.0001
CSRU	366 (9.54)	7,226 (25.28)	< 0.0001
SICU	677 (17.64)	4,714 (16.49)	0.075
Comfort Care Order	694 (18.09)	5,536 (19.37)	0.059
Mechanical Ventilation			
Mechanical Ventilation Use	1,371 (35.73)	12,505 (43.75)	< 0.0001
Duration of Mechanical Ventilation, days $^{\Theta}$	2 [1-6]	1 [1-3]	< 0.0001
Vasoactive Agent Use			
Vasoactive Agent Use	674 (17.57)	4,147 (14.51)	< 0.0001
Duration of Vasoactive Agent Use, days $^{\Omega}$	2 [1-3]	1 [1-3]	< 0.0001
Sepsis	1,533 (39.95)	6,391 (22.36)	< 0.0001
Cardiac Dysrhythmia	1,500 (39.09)	7,604 (26.60)	< 0.0001
Coagulopathy	489 (12.74)	2,335 (8.17)	< 0.0001
Dialysis	288 (7.51)	876 (3.06)	< 0.0001
Tracheostomy	272 (7.09)	1,999 (6.99)	0.814

*All continuous variables are expressed as medians [inter quartile range] or mean±SD depending on normality of data. Categorical data is expressed as n (frequency). †Abbreviations: COPD: Chronic obstructive pulmonary disease; ESRD: End-stage renal disease; LOS: Length of stay; ICU: Intensive care unit; SOFA: Sequential organ failure assessment; MICU: Medical intensive care unit; CCU: Cardiovascular intensive care unit; CSRU: Cardiac surgery intensive care unit; SICU: Surgical intensive care unit

‡Admission (hospital and/or ICU) within 1 year prior to index ICU stay

δ 16,231 observations

 Ψ 10,304 observations

 Θ 13,876 observations

 Ω 4,821 observations

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Variable	Parameter Estimate	Standard Error	Hazard Ratio	95% Confidence Interval	p value
Malignancy	0.929	0.032	2.53	(2.38, 2.69)	<.0001
RRT†	0.582	0.129	1.79	(1.39, 2.31)	<.0001
Admission Type	0.495	0.055	1.64	(1.47, 1.83)	<.0001
ESRD	0.391	0.168	1.48	(1.06, 2.05)	0.0196
Cirrhosis	0.391	0.082	1.48	(1.26, 1.74)	<.0001
COPD	0.375	0.037	1.45	(1.35, 1.57)	<.0001
Hospital Readmission§	0.347	0.043	1.41	(1.30, 1.54)	<.0001
Sepsis	0.304	0.036	1.35	(1.26, 1.45)	<.0001
ICU Type	0.249	0.017	1.28	(1.24, 1.33)	<.0001
Diabetes	0.243	0.059	1.28	(1.13, 1.43)	<.0001
Coagulopathy	0.161	0.046	1.17	(1.07, 1.29)	0.0005
Cardiac Dysrhythmia	0.127	0.033	1.14	(1.07, 1.21)	0.0001
ESRD*Cirrhosis‡	0.064	0.338	1.07	(0.55, 2.07)	0.8487
ICU Readmissions§	0.037	0.027	1.04	(0.98, 1.09)	0.18
Age	0.034	0.001	1.03	(1.03, 1.04)	<.0001
Cerebrovascular Accident	0.026	0.069	1.03	(0.90, 1.17)	0.7035
Duration of Mechanical Ventilation	0.021	0.003	1.02	(1.02, 1.03)	<.0001
Duration of Vasoactive Agent Use	0.007	0.008	1.01	(0.99, 1.02)	0.3911
Male Gender	0.007	0.031	1.01	(0.95, 1.07)	0.8303
Day 1 SOFA Score	-0.021	0.007	0.98	(0.97, 0.99)	0.0012
RRT*Sepsis	-0.155	0.150	0.86	(0.64, 1.15)	0.2989
Sepsis*Cirrhosis‡	-0.176	0.116	0.84	(0.67, 1.05)	0.1294
Comfort Care Order	-0.218	0.041	0.8	(0.74, 0.87)	<.0001
ESRD*Sepsis	-0.226	0.177	0.8	(0.56, 1.13)	0.2029
ESRD*RRT‡	-0.356	0.183	0.7	(0.49, 1.00)	0.0519
Hypertension	-0.409	0.033	0.66	(0.62, 0.71)	<.0001
RRT*Cirrhosis‡	-0.676	0.247	0.51	(0.31, 0.83)	0.0063

Table 2. Cox proportional nazarus model for outcome. One-year mortanty	Table 2.	Cox proportional	l hazards model fo	r outcome: One-	vear mortality
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*Abbreviations: COPD: Chronic obstructive pulmonary disease; ESRD: End-stage renal disease; LOS: Length of stay; ICU: Intensive care unit; SOFA: Sequential organ failure assessment; MICU: Medical intensive care unit; MV: Mechanical Ventilation

†New initiation of Renal Replacement Therapy in the ICU

\$Admission within 1 year of index ICU stay

‡Interaction terms

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may have worse prognoses independent of ICU admission, and/or subsequently be more physiologically compromised by the impact of such admissions.

As noted in previous studies, many patients have care withdrawn or withheld, and physicians use numerous variables in their determinations regarding whether to limit therapy[22]. Given the significant association with certain ICU interventions (renal replacement therapy, mechanical ventilation duration, vasopressor duration) as well as co-morbid illnesses (cirrhosis, ESRD) with one-year mortality, we endorse that discussions should address the possibility of one-year mortality even if the patient were to overcome critical illness, with Cox and Curtis describing methods with which this can be integrated into the electronic health record[23].

In addition to establishing numerous comorbidities and disease states associated with oneyear mortality, we also identified effect-modification between the use of renal replacement therapy and cirrhosis. Effect measure modification is the biological phenomenon in which a certain exposure has a different magnitude of effect in the presence of another exposure. Interaction between the effects of numerous patient-, disease- and treatment-related factors is challenging for clinicians to quantify objectively without computational assistance.

As the care of patients becomes more complex, physicians struggle to determine which patients are likely to suffer a poor quality of life and/or death after surviving critical illness, and little follow up is routinely available to ICU clinicians. Current ICU prediction tools rarely apply to patients after the current hospital stay and do not address longer clinical courses such as the one-year interval that we examined [24, 25]. Our findings should alert clinicians to the relevance of events beyond the immediate scope of the ICU stay and to consider post-ICU outcomes in day-to-day decision-making. This awareness should lead to an ability to conduct more informative and comprehensive discussions with patients and their families regarding the goals of care and the ongoing use or additional implementation of potentially futile medical interventions. Our findings are consistent with those found by Garland, et al., providing more evidence that acute ICU interventions and diagnoses along with co-morbid illnesses are associated with long-term survival among survivors of critical illness[9].

There are several limitations to our study. Our study is limited by its retrospective nature as well as by being from a single, tertiary-care academic center, and our results may therefore not be widely generalizable. Furthermore, potential unmeasured confounders not included in our analysis may also be present. Disease associations based on ICD-9 codes are limited by the reporting of the treating physician(s). Our inclusion of admissions within the prior year only capture those at our single institution, and thus may miss hospitalizations at other institutions. Additionally, we did not adjust for any treatment-related or technological advancements that may have occurred between 2001 and 2012. Furthermore, we chose to focus on post-discharge mortality among 30-day survivors rather than all patients admitted to the intensive care unit. We focused our post-discharge mortality predictions specifically on 30-day survivors rather than all ICU patients for 2 reasons. First, this approach excludes in-hospital mortality within that time window. Given that some individuals might have lived if care had not been limited or terminated, including these individuals could have led to classification error. The 30-day survivor cohort is also roughly equivalent to that of hospital survivors while providing a long post hospital discharge time window to observe mortality rates until the defined limit of one year. Second, one-year mortality among patients who are discharged after an ICU admission may reflect, to some degree, poor prognostication on the part of the clinician: these patients were thought to have potentially positive post-discharge outcomes, otherwise care would presumably have been limited or withdrawn.

Lastly, our main outcome was one-year mortality. There are significant outcomes of value that were not measured, including hospital re-admission, time spent in a facility, and post-discharge utilization of resources. We hope to integrate the Massachusetts All Payer Claims Database[26] into future versions of MIMIC to better characterize post-discharge utilization among critically-ill patients.

Conclusion

Among survivors of critical illness, a significant proportion survive less than one year. Numerous factors are associated with one-year mortality. More research is needed to help clinicians accurately predict which patients who, despite surviving their acute illness, are likely to suffer one-year mortality. Identification of these factors will aid in the development of clinical decision support tools to assist in the complex, real time therapeutic decisions required in this difficult context and provide feedback to clinicians regarding the long-term outcomes of patients who survive their ICU stay.

Supporting information

S1 Table. Cox proportional hazards model for outcome: Five-year mortality. (DOCX)

S1 Text. Supplementary information. (DOCX)

S1 Fig. Log-log plot of survival as a function of hospital readmissions. (PNG)

S2 Fig. Kaplan Meier plot of survival as a function of hospital readmissions. (PNG)

S3 Fig. Log-log plot of survival as a function of sepsis. (PNG)

S4 Fig. Kaplan Meier plot of survival as a function of sepsis. (PNG)

S5 Fig. Log-log plot of survival as a function of arrhythmia. (PNG)

S6 Fig. Kaplan Meier plot of survival as a function of arrhythmia. (PNG)

S7 Fig. Log-log plot of survival as a function of cirrhosis. (PNG)

S8 Fig. Kaplan Meier plot of survival as a function of cirrhosis. (PNG)

S9 Fig. Log-log plot of survival as a function of COPD. (PNG)

S10 Fig. Kaplan Meier plot of survival as a function of COPD. (PNG)

S11 Fig. Log-log plot of survival as a function of diabetes. (PNG)

S12 Fig. Kaplan Meier plot of survival as a function of diabetes. (PNG)

S13 Fig. Log-log plot of survival as a function of renal replacement therapy. (PNG)

S14 Fig. Kaplan Meier plot of survival as a function of renal replacement therapy. (PNG)

S15 Fig. Log-log plot of survival as a function of end stage renal disease. (PNG)

S16 Fig. Kaplan Meier plot of survival as a function of end stage renal disease. (PNG)

S17 Fig. Log-log plot of survival as a function of ICU type. (PNG)

S18 Fig. Kaplan Meier plot of survival as a function of ICU type. (PNG)

S19 Fig. Log-log plot of survival as a function of malignancy. (PNG)

S20 Fig. Kaplan Meier plot of survival as a function of malignancy. (PNG)

S21 Fig. Schoenfeld's residual plot of duration of mechanical ventilation. (PNG)

S22 Fig. Schoenfeld's residual plot of duration of mechanical ventilation using cubic splines and smoothing.

(PNG)

S23 Fig. Schoenfeld's residual plot of duration of mechanical ventilation using cubic splines without smoothing. (PNG)

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References

- Erickson SE, Martin GS, Davis JL, Matthay MA, Eisner MD, Network NNA. Recent trends in acute lung injury mortality: 1996–2005. Critical care medicine. 2009; 37(5):1574–9. Epub 2009/03/28. <u>https://doi. org/10.1097/CCM.0b013e31819fefdf</u> PMID: 19325464; PubMed Central PMCID: PMC2696257.
- Friedman G, Silva E, Vincent JL. Has the mortality of septic shock changed with time. Critical care medicine. 1998; 26(12):2078–86. Epub 1999/01/06. PMID: 9875924.
- Pro CI, Yealy DM, Kellum JA, Huang DT, Barnato AE, Weissfeld LA, et al. A randomized trial of protocol-based care for early septic shock. N Engl J Med. 2014; 370(18):1683–93. Epub 2014/03/19. https:// doi.org/10.1056/NEJMoa1401602 PMID: 24635773; PubMed Central PMCID: PMC4101700.

- Carson SS, Kahn JM, Hough CL, Seeley EJ, White DB, Douglas IS, et al. A multicenter mortality prediction model for patients receiving prolonged mechanical ventilation. Critical care medicine. 2012; 40 (4):1171–6. Epub 2011/11/15. https://doi.org/10.1097/CCM.0b013e3182387d43 PMID: 22080643; PubMed Central PMCID: PMC3395423.
- Paul E, Bailey M, Pilcher D. Risk prediction of hospital mortality for adult patients admitted to Australian and New Zealand intensive care units: development and validation of the Australian and New Zealand Risk of Death model. J Crit Care. 2013; 28(6):935–41. Epub 2013/10/01. https://doi.org/10.1016/j.jcrc. 2013.07.058 PMID: 24074958.
- Zimmerman JE, Kramer AA, McNair DS, Malila FM. Acute Physiology and Chronic Health Evaluation (APACHE) IV: hospital mortality assessment for today's critically ill patients. Critical care medicine. 2006; 34(5):1297–310. Epub 2006/03/17. https://doi.org/10.1097/01.CCM.0000215112.84523.F0 PMID: 16540951.
- Moreno RP, Metnitz PG, Almeida E, Jordan B, Bauer P, Campos RA, et al. SAPS 3—From evaluation of the patient to evaluation of the intensive care unit. Part 2: Development of a prognostic model for hospital mortality at ICU admission. Intensive care medicine. 2005; 31(10):1345–55. Epub 2005/09/01. https://doi.org/10.1007/s00134-005-2763-5 PMID: 16132892; PubMed Central PMCID: PMC1315315.
- Higgins TL, Teres D, Copes WS, Nathanson BH, Stark M, Kramer AA. Assessing contemporary intensive care unit outcome: an updated Mortality Probability Admission Model (MPM0-III). Critical care medicine. 2007; 35(3):827–35. Epub 2007/01/27. <u>https://doi.org/10.1097/01.CCM.0000257337.63529.9F</u> PMID: 17255863.
- Garland A, Olafson K, Ramsey CD, Yogendran M, Fransoo R. Distinct determinants of long-term and short-term survival in critical illness. Intensive care medicine. 2014; 40(8):1097–105. <u>https://doi.org/10.1007/s00134-014-3348-y PMID: 25011575</u>.
- Ferrand E, Robert R, Ingrand P, Lemaire F, French LG. Withholding and withdrawal of life support in intensive-care units in France: a prospective survey. French LATAREA Group. Lancet. 2001; 357 (9249):9–14. Epub 2001/02/24. PMID: 11197395.
- Sprung CL, Cohen SL, Sjokvist P, Baras M, Bulow HH, Hovilehto S, et al. End-of-life practices in European intensive care units: the Ethicus Study. JAMA : the journal of the American Medical Association. 2003; 290(6):790–7. Epub 2003/08/14. https://doi.org/10.1001/jama.290.6.790 PMID: 12915432.
- Unroe M, Kahn JM, Carson SS, Govert JA, Martinu T, Sathy SJ, et al. One-year trajectories of care and resource utilization for recipients of prolonged mechanical ventilation: a cohort study. Ann Intern Med. 2010; 153(3):167–75. Epub 2010/08/04. https://doi.org/10.7326/0003-4819-153-3-201008030-00007 PMID: 20679561; PubMed Central PMCID: PMC2941154.
- 13. Prescott HC, Langa KM, Liu V, Escobar GJ, Iwashyna TJ. Increased 1-year healthcare use in survivors of severe sepsis. Am J Respir Crit Care Med. 2014; 190(1):62–9. Epub 2014/05/30. https://doi.org/10. 1164/rccm.201403-0471OC PMID: 24872085; PubMed Central PMCID: PMC4226030.
- Prescott HC, Osterholzer JJ, Langa KM, Angus DC, Iwashyna TJ. Late mortality after sepsis: propensity matched cohort study. BMJ. 2016; 353:i2375. <u>https://doi.org/10.1136/bmj.i2375</u> PMID: 27189000; PubMed Central PMCID: PMCPMC4869794.
- Moitra VK, Guerra C, Linde-Zwirble WT, Wunsch H. Relationship Between ICU Length of Stay and Long-Term Mortality for Elderly ICU Survivors. Critical care medicine. 2016; 44(4):655–62. https://doi. org/10.1097/CCM.00000000001480 PMID: 26571190; PubMed Central PMCID: PMCPMC4792682.
- Gill TM, Gahbauer EA, Han L, Allore HG. The role of intervening hospital admissions on trajectories of disability in the last year of life: prospective cohort study of older people. BMJ. 2015; 350:h2361. Epub 2015/05/23. https://doi.org/10.1136/bmj.h2361 PMID: 25995357; PubMed Central PMCID: PMC4443433.
- Kahn JM, Le T, Angus DC, Cox CE, Hough CL, White DB, et al. The epidemiology of chronic critical illness in the United States*. Critical care medicine. 2015; 43(2):282–7. Epub 2014/11/08. <u>https://doi.org/ 10.1097/CCM.00000000000710 PMID: 25377018</u>.
- Johnson AE, Pollard TJ, Shen L, Lehman LW, Feng M, Ghassemi M, et al. MIMIC-III, a freely accessible critical care database. Sci Data. 2016; 3:160035. https://doi.org/10.1038/sdata.2016.35 PMID: 27219127; PubMed Central PMCID: PMCPMC4878278.
- Angus DC, Linde-Zwirble WT, Lidicker J, Clermont G, Carcillo J, Pinsky MR. Epidemiology of severe sepsis in the United States: analysis of incidence, outcome, and associated costs of care. Critical care medicine. 2001; 29(7):1303–10. Epub 2001/07/11. PMID: 11445675.
- Vincent JL, Moreno R, Takala J, Willatts S, De Mendonça A, Bruining H, et al. The SOFA (Sepsisrelated Organ Failure Assessment) score to describe organ dysfunction/failure. Intensive care medicine. 1996; 22(7):707–10. https://doi.org/10.1007/bf01709751 PMID: 8844239

- Mehta RL, Kellum JA, Shah SV, Molitoris BA, Ronco C, Warnock DG, et al. Acute Kidney Injury Network: report of an initiative to improve outcomes in acute kidney injury. Crit Care. 2007; 11(2):R31. https://doi.org/10.1186/cc5713 PMID: 17331245; PubMed Central PMCID: PMCPMC2206446.
- Kranidiotis G, Gerovasili V, Tasoulis A, Tripodaki E, Vasileiadis I, Magira E, et al. End-of-life decisions in Greek intensive care units: a multicenter cohort study. Crit Care. 2010; 14(6):R228. Epub 2010/12/ 22. https://doi.org/10.1186/cc9380 PMID: 21172003; PubMed Central PMCID: PMC3219993.
- 23. Cox CE, Curtis JR. Using Technology to Create a More Humanistic Approach to Integrating Palliative Care into the Intensive Care Unit. Am J Respir Crit Care Med. 2016; 193(3):242–50. https://doi.org/10. 1164/rccm.201508-1628CP PMID: 26599829.
- Badawi O, Breslow MJ. Readmissions and death after ICU discharge: development and validation of two predictive models. PloS one. 2012; 7(11):e48758. Epub 2012/11/13. https://doi.org/10.1371/ journal.pone.0048758 PMID: 23144958; PubMed Central PMCID: PMC3492441.
- Ouanes I, Schwebel C, Francais A, Bruel C, Philippart F, Vesin A, et al. A model to predict short-term death or readmission after intensive care unit discharge. J Crit Care. 2012; 27(4):422 e1–9. Epub 2011/ 12/17. https://doi.org/10.1016/j.jcrc.2011.08.003 PMID: 22172798.
- Massachusetts All Payer Claims Database: Center for Health Information and Analysis; 2015. Available from: http://www.chiamass.gov/ma-apcd/.