

RESEARCH

Open Access



Healthcare utilisation among low-income individuals affected by violence in Rio de Janeiro, Brazil: a retrospective cohort study

Sophia Medeiros^{1*}, Julia Guerra^{2,3}, Vinicius Peçanha², Joana Monteiro⁴, Rudi Rocha^{2,3}, Christopher Millett^{1,5} and Thomas Hone^{1,2}

Abstract

Background Violence is associated with adverse health outcomes and affects healthcare utilisation, but there is limited robust evidence on this relationship in the context of violence in urban environments. This study examines subsequent patterns of healthcare utilisation by individuals who accessed healthcare as a result of experiencing violence in Rio de Janeiro, Brazil.

Methods A retrospective cohort analysis was conducted on 529,219 low-income individuals from January 2010 to December 2016. Electronic medical records were screened for ICD-10 and primary care codes, and clinical free text identifying individuals who used healthcare due to violence. Control individuals (who did not use healthcare due to violence) were identified via coarsened exact matching. Logistic panel regression assessed the association between healthcare use due to violence and subsequent healthcare utilisation.

Results 2,038 individuals used healthcare due to violence. This group were 11% more likely (OR:1.11;95%CI:1.00–1.24; $p < 0.05$) to use PHC in the subsequent 1–3 months after initial healthcare use compared to those not using healthcare due to violence. They were 79% less likely (OR:0.21;95%CI:0.05–0.93; $p < 0.05$) to be Hospitalised at 7–12 months. Female users were 4.25 times (OR:4.25;95%CI:2.14–8.44; $p < 0.001$) and 2.57 times (OR:2.57;95%CI:1.20–5.53; $p < 0.05$) more likely to use PHC for poor pregnancy outcomes in the first 1–3 months and 7–12 months, respectively. PHC use for mental health increased by 50% (OR:1.50;95%CI:1.27–1.77; $p < 0.005$) in the subsequent 1–3 months after initial healthcare use.

Conclusions Violence was associated with short-term increases in healthcare utilisation among low-income individuals in Rio de Janeiro, Brazil.

Keywords Violence, Health outcomes, Primary healthcare, Healthcare utilisation, Hospitalisation, Rio de Janeiro, Brazil, Urban, LMIC, Coarsened exact matching

*Correspondence:

Sophia Medeiros
sophia.medeiros21@imperial.ac.uk

¹Public Health Policy Evaluation Unit, Department of Primary Care and Public Health, School of Public Health, Imperial College London, London, UK

²Instituto de Estudos para Políticas de Saúde (IEPS), São Paulo, Brazil

³Escola de Administração de Empresas de São Paulo da Fundação Getúlio Vargas (FGV EAESP), São Paulo, Brazil

⁴Escola Brasileira de Administração Pública e de Empresas da Fundação Getúlio Vargas (FGV EBAPE), Rio de Janeiro, Brazil

⁵Public Health Research Center, Comprehensive Health Research Center, NOVA National School of Public Health, NOVA University Lisbon, Lisbon, Portugal



Background

Violence is a major threat to global and human development. Violence is responsible for over 500,000 deaths each year, which is estimated to increase to 610,000 by 2030 [1, 2]. Violence is a persistent public health problem in Brazil [3]. Among Latin American countries, a region with one of the highest Homicide rates globally, Brazil ranks among the highest with a Homicide rate of 30.76 per 100,000 population in 2020 – three times the global average [3, 4] whilst violence is the sixth leading cause of hospital admissions [5].

Experiencing violence, including being directly harmed, affected by, or witnessing violence, has been associated with poorer physical and psychological health in numerous studies [6–14]. The most visible, immediate health impacts of violence are injuries and homicides. However, there are long-term effects on both physical and mental health. Existing evidence from the United States and conflict-affected regions including Palestine, Libya, and Afghanistan, suggests that violence increases the risk of chronic illnesses like asthma, cardiovascular disease, cancer, and stroke as well as poor psychological health and birth outcomes [6–9]. Individuals fear accessing healthcare services in violence-afflicted settings which discourages health seeking behaviours and service utilisation [15–18]. On the other hand, some evidence suggests that worsening health status and stress in violent settings may increase health service utilisation. For example, injuries may lead to an increase in emergency room visits and violence-related stress may result in long-term mental health needs, increasing short- and long-term utilisation [8, 19, 20].

Violence affects health through multiple pathways that increases risk factors leading to adverse health outcomes.

Considering Scott-Storey's [21] and Guha-Sapir & van Panhuis' [22] frameworks on violence and health, four key pathways through which violence increases health risks can be identified: biological and nervous system responses, health risk behaviours as coping mechanisms, chronic stress, and injury. These pathways contribute to various health conditions, including adverse pregnancy and birth outcomes, cardiovascular disease and stroke, diabetes, obesity, poor mental health, infections and infectious diseases, and disability (Fig. 1; Appendix 1).

Despite the understanding that violence affects both short- and long-term health, there is little empirical evidence on patterns of healthcare utilisation in those exposed to violence [23–25]. Studies in low- and middle-income countries (LMICs) tend to focus on intimate partner violence and utilisation of maternal care services [26–30]. Most evidence originates from high income countries [31, 32], focusses on violence exposure in childhood and adolescence [11, 12, 15, 17], or is based on self-reported violence and health status [10, 12, 15], with limited investigation on patterns of healthcare utilisation.

The city of Rio de Janeiro is an important setting for this study. An estimated one fifth of Rio de Janeiro's population live in slums (favelas) where violence can be endemic [33–35]. Individuals living in slums across Latin America frequently experience high levels of violence as a result of drug trafficking, organised crime, gang activity, and police operations. While existing evidence demonstrates associations between violence and unsafe communities, reduced healthcare access, and reduced investment into public services [36, 37], no study has comprehensively explored the impact of violence on healthcare utilisation.

This study quantifies patterns of healthcare utilisation following an initial healthcare consultation due to

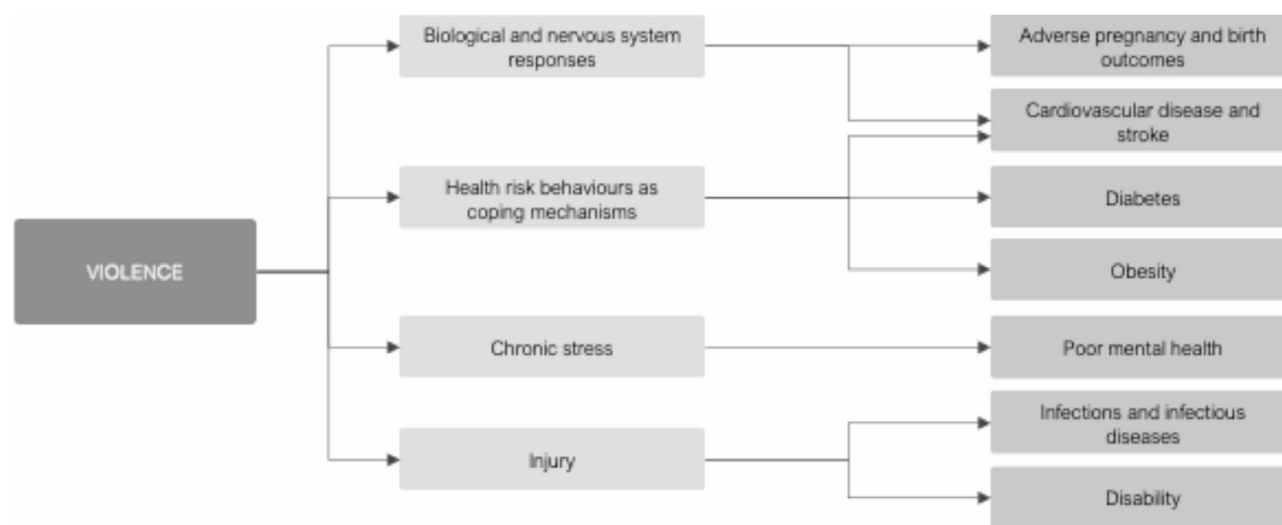


Fig. 1 Conceptual framework of the pathways through which violence increases health risk factors. Sources: Scott-Storey [21], Guha-Sapir & van Panhuis [22], and author's own work. See Appendix 1 for more detail

violence among low-income individuals in Rio de Janeiro. We combine individual electronic health records and hospitalisation records with socioeconomic and demographic data. This study employs a novel approach by screening clinical free text records—an approach that has been used in high-income countries and shown to offer advantages over structured data alone. It enables the identification of individuals who report or discuss violence, even when it is not recorded as a diagnosis or procedure code, thereby capturing more detailed information and improving research quality and accuracy [38]. In this study we focus on violence that would expectedly occur outside the home and violence which relates to the city and urban environment (i.e. in the community/neighbourhood rather than the domestic environment). Rape and sexual abuse/assault in Brazil is largely interpretable as domestic violence and warrants separate investigation given the distinctly different policy responses, interventions, and advocacy efforts [39, 40]. Therefore, we exclude violence related to rape and sexual abuse/assault from our analyses. We aim to understand the effect of violence, such as physical assault, crime, and gang activity, on the risk of further healthcare use over time by quantifying primary healthcare utilisation and hospitalisation for violence-related health conditions previously shown to be adversely impacted by violence exposure.

Methods

Study design

This is a retrospective cohort analysis of primary healthcare-registered users exposed to violence and their matched controls with linked welfare, primary healthcare (PHC), and hospitalisation records.

Study population

Our study population is derived from a larger cohort of 3,148,383 individuals registered with PHC clinics in Rio de Janeiro. We obtained information on individuals from the *Ficha de Cadastro Domiciliar e Territorial* (Household and Territorial Registration Form; Ficha A) from the Brazilian Ministry of Health [41]. These PHC registration records were linked to PHC electronic medical records from the Primary Care Assessment Tool (PCAT) dataset to identify PHC users and public hospitalisation records (*Sistema de Informações Hospitalares*; SIH), all covering the period from 1 January 2010 to 31 December 2016. The datasets were linked via a combination of deterministic and probabilistic approaches, with details published elsewhere [42].

We identified PHC users exposed to violence, and using coarsened exact matching identified control individuals not exposed to violence, resulting in a final study

population of 529,219 individuals. PHC services in Rio de Janeiro are concentrated in low-income areas, meaning the study population largely consisted of lower-income individuals living in poorer neighbourhoods at greater risk of violence.

Matching

We exploited one-to-many coarsened exact matching (CEM) with no repeats to reduce potential unobserved confounding and ensure covariate balance. Covariate balance was assessed using the linear 1 (L1) statistic [43]. Exposed individuals who did not have matches were dropped from the dataset ($n=11$). Exposure status was determined prior to conducting CEM (on the original 3.1 million sample) to prevent bias in group classification and ensure that the matching process was based solely on confounders. Nine covariates were used for matching: micro area code (a geographical administrative unit responsible for local health services to account for potential differences in healthcare access), sex, race/skin colour, education, age group, household income, private healthcare insurance, whether an individual's family claimed welfare, and whether an individual used healthcare (PHC or hospitalisation) in the study period. Information on the sample size and covariate balancing pre- and post-matching can be found in [Appendix 4](#).

Exposure

Individuals were considered exposed to violence if they used PHC or were hospitalised due to violence. This was classified according to International Classification of Disease (ICD-10) codes and the International Classification of Primary Care (ICPC) code Z25 (assault/harmful event problem) (Table 1) – for either primary or secondary causes. We also searched clinical free text in PHC consultation notes, detailed in [Appendix 2](#). For each individual, we obtained the month-year of their first healthcare use for violence and used this to construct our exposure variable.

Outcomes

Two outcomes were analysed: PHC consultations and hospital admissions. These outcomes were considered for conditions previously shown to be adversely impacted by violence exposure (see conceptual framework; Fig. 1). Health conditions were grouped into five categories according to ICD-10 and ICPC codes: adverse pregnancy and birth outcomes; cardiovascular disease and stroke, diabetes, and obesity; poor mental health; infections and infectious diseases; and disability ([Appendix 3](#)) based on clinical presentation. Both primary and secondary causes of PHC utilisation and hospitalisations were used to identify outcomes.

Table 1 International classification of disease (ICD-10) and international classification of primary care (ICPC) codes to identify cases of violence in electronic medical records

VIOLENCE SUB-TYPES	CODES
Perpetrated with object	X99, Y00
Firearms	X93-X95, W32-W34, Y22-Y24
Drugs, chemicals, explosives	X85-X90, X96-X98
Suffocation	X91-X92
Intentional action (pushing, crashing)	Y01-Y03
Bodily force and physical violence	Y04, R45.6
Other maltreatment by unknown person	Y07.2, Y07.3
Assault by other means	Y08-Y09
Other intentional injuries	Y35
Sequelae of assault and other intentional injuries	Y87.1, Y89
Assault/harmful event problem	Z25*
Examination and observation following other inflicted injury	Z04.5
Victim of crime and terrorism – victim of torture	Z65.4
Exposure to disaster, war and other hostilities	Z65.5

Sources: World Health Organization [44], Romeo et al., [45] Clery et al., [46] Reynolds [47], and author's own work. *ICPC code. All other codes refer to ICD-10

Covariates

Individual-level covariates, obtained from PHC registration and utilisation records, included race (based on self-reported skin colour in Brazil – Black, White, Pardo [mixed race], Amarelo [Asian], Indigenous, other or unknown [missing]); sex (male, female); highest educational attainment (none/pre-school/literacy class, elementary school, high school or higher education, none reported [missing]); and age group in years (0–18, 19–34, 35–54, 55–64, 65+). Household-level covariates, obtained from Ficha A, included household income (some salary reported to less than full minimum salary, more than one minimum salary but less than two minimum salaries, two minimum salaries or more, no salary reported [unknown or missing]); private healthcare insurance (yes, no); and whether the family receives Bolsa Família—a conditional cash transfer programme for families earning below US\$70 per day (yes, no).

Covariates were selected to control for any potential confounders between healthcare use due to violence (exposure) and future healthcare use (outcomes) in sensitivity analyses. Individuals of lower-income, lower educational attainment, and Black/Pardo Brazilians often have limited access to healthcare, influencing health outcomes [48–50]. Additionally, these individuals are more likely to reside in the neighbourhoods with high levels of poverty and crime, increasing the risk of experiencing or being affected by violence [51–54].

Dataset

An individual-level panel was constructed at the month-year level. Outcomes (PHC consultations or hospitalisations for selected health conditions) were observed for individuals from 1 January 2010 to 31 December 2016. Individuals' start time was defined as either (i) the start of the observation period (1 January 2010) if they were born before 1 January 2010 and PHC registered before 1 January 2010 or (ii) their registration date if they were born after 1 January 2010 and PHC registered before 31 December 2016. Individuals' end date was defined as either (i) their date of death or (ii) the end of the observation period (31 December 2016).

Statistical analyses

The total number of PHC consultations and hospitalisations for conditions previously shown to be adversely impacted by violence exposure were reported, overall and by health condition for individuals who used healthcare due to violence and their matched controls. Cohort characteristics, person-years and crude rates of PHC utilisation and hospitalisation of the study population were reported.

Firstly, separate logistic panel regressions were conducted for PHC utilisation and hospitalisation outcomes with a categorical exposure variable representing time since using healthcare due to violence (0 months or did not use healthcare due to violence, 1–3 months, 4–6 months, 7–12 months, over twelve months). For models on hospitalisations, we additionally adjusted for PHC utilisation. Secondly, separate panel logistic regression models were conducted stratified by specific health conditions to understand which health conditions may be driving utilisation. Only PHC outcome regression results were analysed by condition due to low numbers in hospital admissions. Additional two-way interactions between the categorical violence exposure variable and sex, race/skin colour, education level, and household income, were tested to explore socioeconomic inequalities in the associations between violence and PHC utilisation. This analysis aims to identify whether certain groups are disproportionately affected by violence in terms of their healthcare utilisation [48–54].

To address the potential for unobserved heterogeneity in individual characteristics (e.g., baseline health status or socioeconomic factors) to bias our estimates, we conducted a fixed-effects analysis. A dummy variable was included for each individual, effectively controlling for all time-invariant characteristics and accounting for unobservable factors that do not change over time. All regression models were clustered at the individual level, and incorporated individual, month, and year fixed effects. Regression models assessing healthcare use for adverse pregnancy and birth outcomes were restricted to females.

As all variables were expressed categorically, effect estimates were interpreted as the ratio of the category of interest and the baseline category. Adjusted odds ratios (aORs) or odds ratios (ORs) with 95% confidence intervals (CIs) and p-values were reported. All analyses were performed in Stata V.18.1.

Sensitivity analyses

To evaluate the robustness of our findings, we conducted sensitivity analyses using alternative model specifications. First, to test the assumption that there may be unobserved heterogeneity across individuals, we used a random-effects logistic regression model which adjusted for differences between individuals and changes within individuals over time. It also allowed us to include individuals who did not have any PHC use or hospitalisation. Second, to test the assumption that fixed-effects logistic regression might not fully address the potential for robust standard errors, we employed a fixed-effects Poisson regression model. This model controlled for all time-invariant individual characteristics while enabling the calculation of robust standard errors. Finally, to test the assumption that correlations within repeated measures of individuals over time could affect the results, we used generalized estimating equations (GEE) with a binomial family and logit link to estimate population-averaged associations. GEE accounted for these correlations, providing a broader perspective on the observed relationships.

Results

We identified 2,038 individuals who used healthcare due to violence during the study period and selected 527,181 matched controls, representing 1,468,225 person-years of observation (Table 2). The mean number of person-years of observation was 2.77. Over 70% of the study population were female (377,087; 71.3%). By race/skin colour, most of the sample were Pardo (mixed race) (324,583; 61.3%) or White (157,222; 29.7%). Most individuals had either elementary school (216,733; 41.0%) or high school or higher education (199,433; 37.7%). The study population was mostly aged between 35 and 54 years of age (175,108; 33.1%).

The 2,038 individuals who used healthcare due to violence had a total of 10,029 PHC consultations and 87 Hospital admissions during the study period. Matched controls used PHC 1,024,982 times and were Hospitalised 9,941 times (Appendix 5). Individuals who used healthcare due to violence had over two times the rate of PHC utilisation compared to individuals who did not (1,535.64 compared to 701.23 per 1,000 person-years, respectively) with a similar pattern observed in rates of hospitalisation (13.32 compared to 6.80 per 1,000 person-years, respectively) (Figs. 2 and 3, Appendix 6).

Patterns of healthcare utilisation over time

In Fixed-effects logistic regression models, individuals who used healthcare due to violence were 11% more likely (OR: 1.11; 95% CI: 1.00–1.24; $p < 0.05$) to use PHC in the first 1–3 months after their initial healthcare contact for violence compared to their period before and to individuals who did not use healthcare due to violence (Table 3). In the longer-term, individuals who used healthcare due to violence were 15% (OR: 0.85; 95% CI: 0.76–0.94; $p < 0.01$) and 22% (OR: 0.78; 95% CI: 0.71–0.85; $p < 0.005$) less likely to use PHC for the selected conditions in the subsequent 7–12 months and >12 months following their initial healthcare contact for violence, respectively. For Hospitalisation, individuals who used healthcare due to violence were 79% less likely (OR: 0.21; 95% CI: 0.05–0.93; $p < 0.05$) to be Hospitalised 7–12 months after their initial healthcare contact for violence compared to individuals who did not use healthcare due to violence. When adjusting for PHC utilisation in the hospitalisation model, individuals who used PHC for the same conditions were more than twice (aOR: 2.08; 95% CI: 1.92–2.26; $p < 0.005$) as likely to be hospitalised for these conditions compared to individuals who did not use PHC for the same conditions (Appendix 7), irrespective of violence exposure.

In regression analyses by health condition (Figure 4, see Appendix 8 for regression output), females who used healthcare due to violence were 4.25 times more likely (aOR: 4.25; 95% CI: 2.14–8.44; $p < 0.001$) and 2.57 times more likely (aOR: 2.57; 95% CI: 1.20–5.53; $p < 0.05$) to use PHC for poor pregnancy and birth outcomes 1–3 months and 7–12 months after using healthcare due to violence, respectively. Individuals who used healthcare due to violence were 18% (aOR: 0.82; 95% CI: 0.73–0.93; $p < 0.01$) and 24% (aOR: 0.76; 95% CI: 0.69–0.85; $p < 0.005$) less likely to use PHC for cardiovascular disease, diabetes, and obesity 7–12 months and over 12 months after using healthcare due to violence. PHC utilisation for mental health in individuals who used healthcare due to violence increased by 50% (aOR: 1.50; 95% CI: 1.27–1.77; $p < 0.005$) in the first 1–3 months after seeking care due to violence, but decreased by 33% (aOR: 0.67; 95% CI: 0.57–0.77; $p < 0.005$) after 12 months. After 12 months, individuals who used healthcare due to violence were almost 60% (aOR: 0.41; 95% CI: 0.27–0.61; $p < 0.005$) less likely to use PHC for infections and infectious diseases.

Due to low numbers, analyses for the hospitalisation outcome measure were not conducted in the condition-disaggregated subgroups. Additionally, all interactions between key socioeconomic variables (sex, race/skin colour, education level, and household income) and months since using PHC due to violence were not significant in any model specification (Appendices 9, 10, 11, 12).

Table 2 Cohort characteristics

Characteristics	Original Cohort	Matched Sample*		
		Controls	Used Health-care due to Violence	Total
Sex				
Male	1,386,067 (44.0%)	151,481 (28.7%)	651 (31.9%)	152,132 (28.7%)
Female	1,762,316 (56.0%)	375,700 (71.3%)	1,387 (68.1%)	377,087 (71.3%)
Race/Skin Colour				
White	1,085,190 (34.5%)	156,635 (29.7%)	587 (28.8%)	157,222 (29.7%)
Black	365,452 (11.6%)	42,511 (8.1%)	364 (17.9%)	42,875 (8.1%)
Pardo (Mixed Race)	1,574,381 (50.0%)	323,545 (61.4%)	1,038 (50.9%)	324,583 (61.3%)
Amarelo (Asian), Indigenous, Other or Unknown (Missing)	123,360 (3.9%)	4,490 (0.9%)	49 (2.4%)	4,539 (0.9%)
Education				
None/pre-school/literacy class	1,296,861 (41.2%)	47,085 (8.9%)	274 (13.4%)	47,359 (8.9%)
Elementary school	789,131 (25.1%)	215,774 (40.9%)	959 (47.1%)	216,733 (41.0%)
High school or higher education	633,051 (20.1%)	198,900 (37.7%)	533 (26.2%)	199,433 (37.7%)
None reported (missing)	429,340 (13.6%)	65,422 (12.4%)	272 (13.3%)	65,694 (12.4%)
Age Group (years)				
0–18	762,075 (24.2%)	76,794 (14.6%)	240 (11.8%)	77,034 (14.6%)
19–34	768,211 (24.4%)	140,193 (26.6%)	615 (30.2%)	140,808 (26.6%)
35–54	855,136 (27.2%)	174,378 (33.1%)	730 (35.8%)	175,108 (33.1%)
55–64	362,962 (11.5%)	80,333 (15.2%)	266 (13.1%)	80,599 (15.2%)
65+	399,999 (12.7%)	55,483 (10.5%)	187 (9.2%)	55,670 (10.5%)
Household Income				
Some salary reported to less than full minimum salary	314,183 (10.0%)	29,517 (5.6%)	322 (15.8%)	29,839 (5.6%)
More than one minimum salary but less than two minimum salaries	634,445 (20.2%)	95,684 (18.2%)	427 (21.0%)	96,111 (18.2%)
Two minimum salaries or more	226,460 (7.2%)	5,613 (1.1%)	91 (4.5%)	5,704 (1.1%)
No salary reported (unknown or missing)	1,973,295 (62.7%)	396,367 (75.2%)	1,198 (58.8%)	397,565 (75.1%)
Registered with Private Health Insurance				
No	2,858,881 (90.8%)	523,641 (99.3%)	1,986 (97.4%)	525,627 (99.3%)
Yes	289,502 (9.2%)	3,540 (0.7%)	52 (2.6%)	3,592 (0.7%)
Bolsa Família-Claiming Family				
No	2,796,963 (88.8%)	498,256 (94.5%)	1,686 (82.7%)	499,942 (94.5%)
Yes	351,420 (11.2%)	28,925 (5.5%)	352 (17.3%)	29,277 (5.5%)
N	3,148,383 (100.0%)	527,181 (99.6%)	2,038 (0.4%)	529,219 (100.0%)

* Cohort used for analysis; one exposed to multiple controls.

Sensitivity analyses

Sensitivity analyses were consistent with the findings from our primary models. Coefficients from the random-effects logistic regression were comparable to the fixed-effects models, indicating no substantial influence of unobserved heterogeneity across individuals (Appendix 13). Similarly, the fixed-effects Poisson regression, which incorporated robust standard errors, yielded estimates that aligned closely with the primary analysis (Appendix 14). Finally, the generalised estimating equations (GEE) approach confirmed the robustness of our results, with population-averaged associations showing no meaningful deviations from the individual-level findings (Appendix 15). Across all alternative model specifications, the direction, magnitude, and significance of associations

remained stable, reinforcing the robustness and reliability of our findings.

Discussion

Low-income individuals in Rio de Janeiro who experienced violence had increased subsequent healthcare use in the short term, notably with increased PHC use in the first six months. This effect was mainly driven by consultations for conditions relating to adverse pregnancy and birth outcomes, mental health, infectious diseases and disability associated outcomes. This effect did not appear to vary across socioeconomic groups, but in the longer-term, individuals who experienced violence had reduced healthcare use, mainly for CVD and stroke, diabetes and obesity, mental health, and infections and infectious diseases.

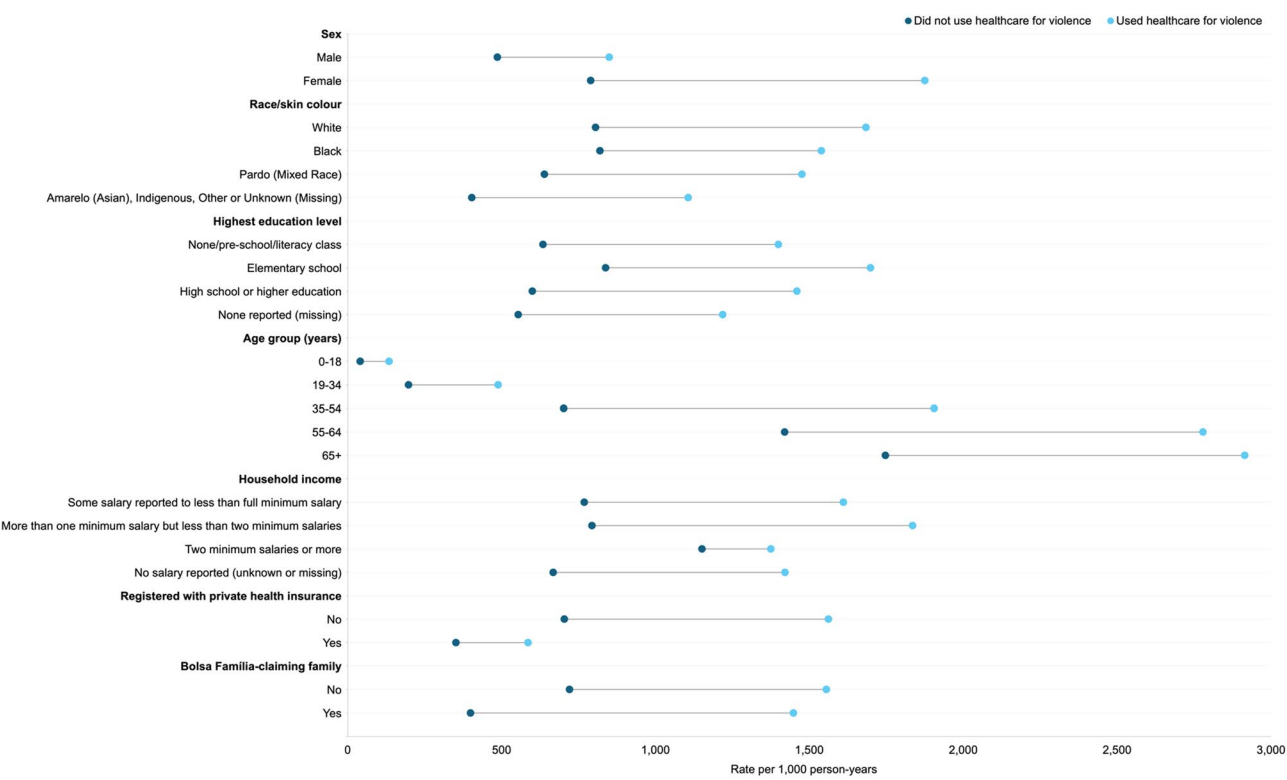


Fig. 2 Crude rates of primary healthcare utilisation by sociodemographic variables per 1,000 person-years. See [Appendix 6](#) for full tabulated results

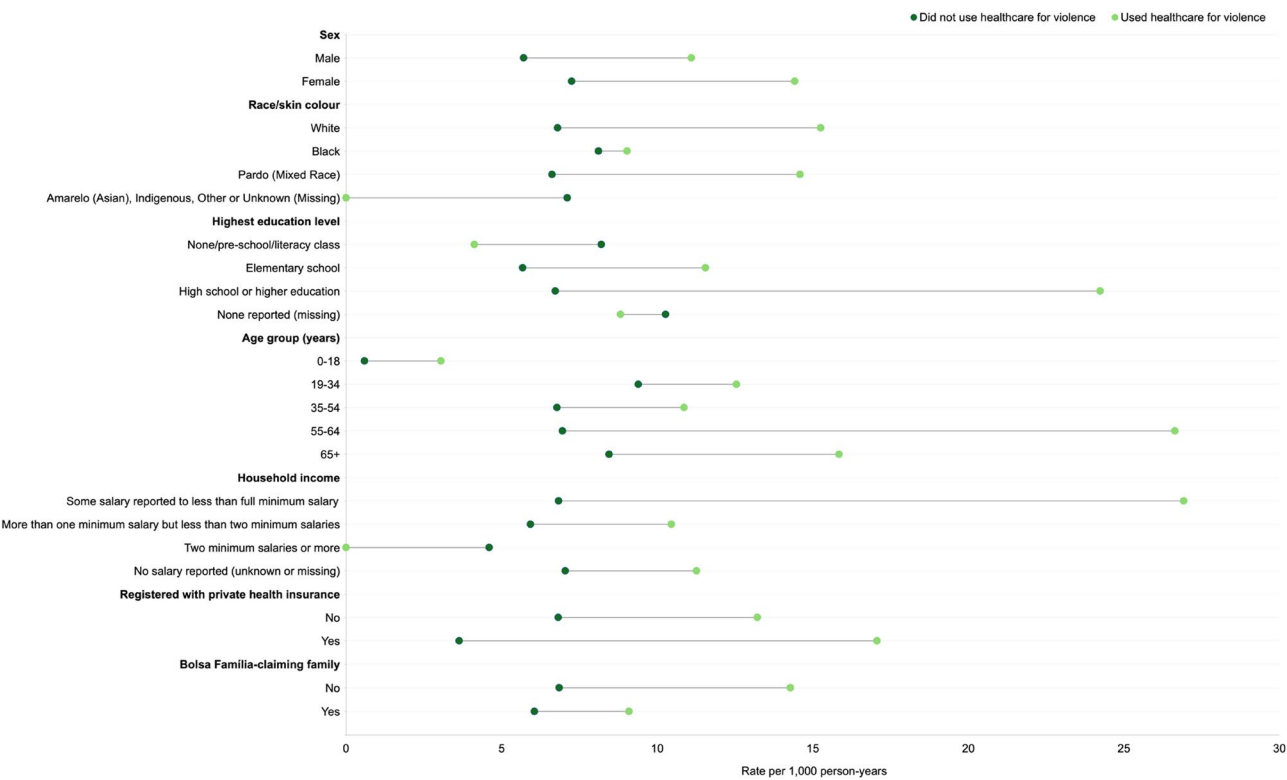


Fig. 3 Crude rates of Hospitalisation by sociodemographic variables and whether individuals used healthcare due to violence per 1,000 person-years. See [Appendix 6](#) for full tabulated results

Table 3 Odds ratios from fixed effects logistic regression models of conditions with risk factors associated with using healthcare due to violence for PHC utilisation and hospitalisation by all sociodemographic variables

Characteristics	PHC Utilisation		Hospitalisation	
	OR	95% CI	aOR	95% CI
Months after using healthcare due to violence				
0, or did not use healthcare due to violence	1 (ref)	–	1 (ref)	–
1–3	1.11 *	(1.00–1.24)	1.15	(0.48–2.77)
4–6	0.98	(0.87–1.11)	0.78	(0.26–2.35)
7–12	0.85 **	(0.76–0.94)	0.21 *	(0.05–0.93)
Over 12	0.78 ***	(0.71–0.85)	0.57	(0.25–1.33)
Total Observations (N)	6,802,113		290,393	
Total Groups (N)	193,275		7,004	

Separate fixed effects logistic regressions per outcome (PHC utilisation and hospitalisation), adjusted for year and month. PHC Primary Healthcare, OR Odds Ratios, 95% CI 95% Confidence Intervals. * $p < 0.05$; ** $p < 0.01$; *** $p < 0.001$

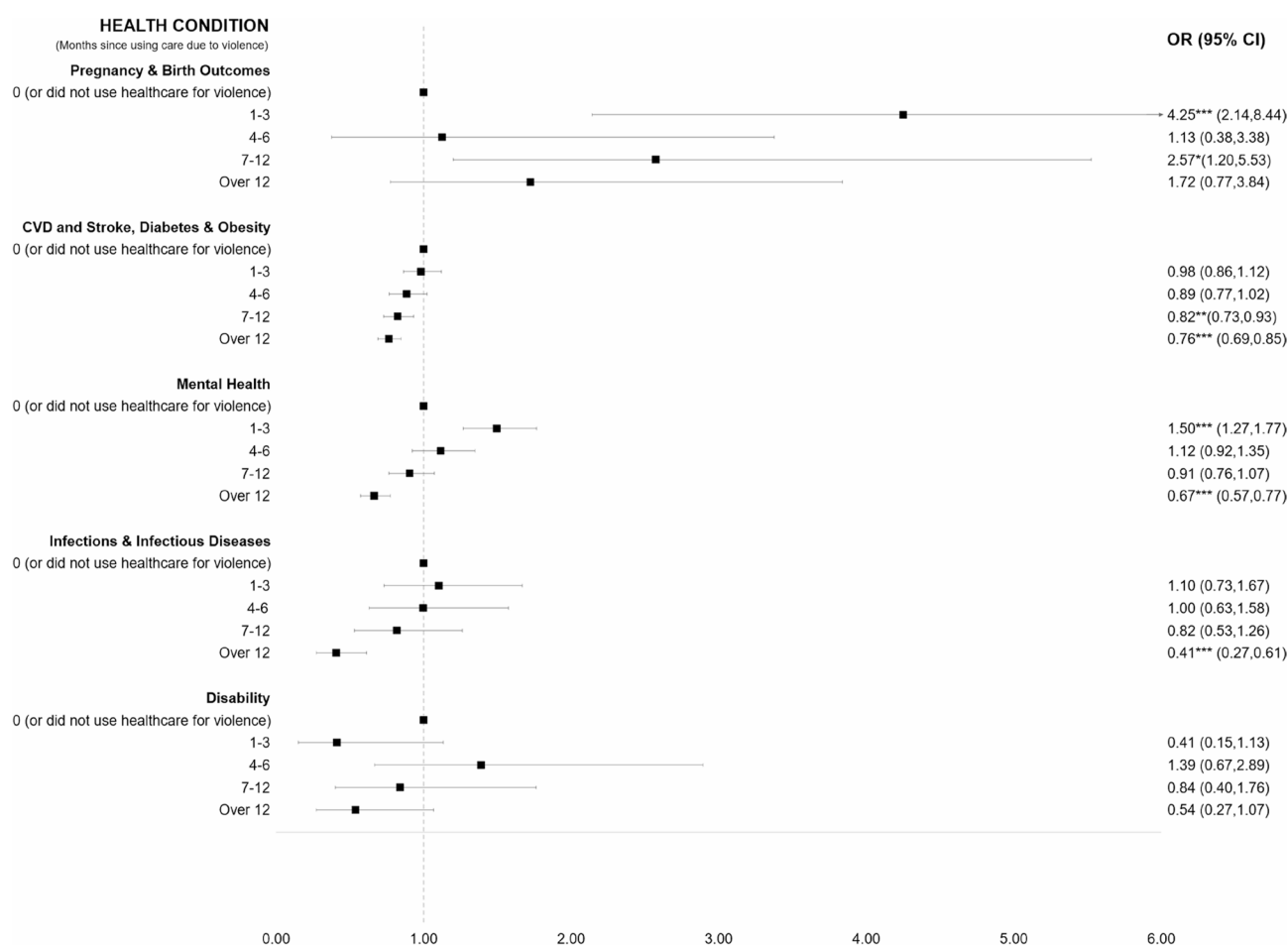


Fig. 4 Odds of PHC utilisation months after using care for violence by health condition. Each group of plotted coefficients (per health condition) is from a separate fully adjusted fixed effects logistic regression model. Robust standard errors. * $p < 0.05$; ** $p < 0.01$; *** $p < 0.001$. 95% CI – 95% Confidence Intervals. Pregnancy & Birth Outcomes model was restricted to females

The finding of a short-term increase in PHC utilisation after experiencing violence is concordant with the wider literature.[8, 19, 20] It may reflect increased healthcare needs due to direct harms or for stress-induced health concerns. The dissipation of this effect over time suggests healthcare needs may have been addressed, and there is

some evidence of lower PHC utilisation after 12 months for certain conditions.[55] In contexts of violence, both patients and healthcare providers may deprioritise chronic conditions or long-term care, contributing to a decline in healthcare utilisation over time. This longer-term decline may also be explained by barriers to access,

reduced healthcare availability, or shifts in healthcare priorities, as observed in this study.

In this study, PHC utilisation for mental health was one of conditions associated with violence, aligning with evidence that violence exposure is associated with poor mental health [56, 57]. Studies in Brazil and other Latin American countries have demonstrated associations between residing in high-violence areas and risk of having common mental disorders [57–60]. Violence can trigger acute mental health crises such as post-traumatic stress disorder and anxiety, prompting health-seeking. Over time, the association between violence and higher PHC use for mental health may diminish due to symptom relief, increased recovery self-efficacy, or individuals discontinuing care as their symptoms ease. A similar pattern was observed for other conditions, including cardiovascular disease (CVD), diabetes, obesity, and infections. For CVD, relief from chest pain or shortness of breath may lead to fewer PHC visits. In diabetes and obesity, better management of blood sugar levels and weight loss can reduce the need for visits. Similarly, in infections, the resolution of fever or pain often leads to discontinuing treatment [61, 62].

PHC utilisation for pregnancy and birth outcomes rises sharply in the short term after using healthcare due to violence in this study. Violence is strongly associated to adverse pregnancy and birth outcomes in high-income countries and LMICs alike, including Brazil [32, 63–66]. Violence is linked to adverse pregnancy outcomes globally, associated with injuries, stress, and high-risk behaviors (e.g., alcohol, smoking), which contribute to low birth weight and preterm birth [3, 65, 67]. Concern for the unborn or newborn child likely drives increased PHC visits which can lead to poor psychological health and may explain increased PHC use for mental health as well [68]. Similarly, PHC utilisation for infections, infectious diseases, and disability-related outcomes also rises in the months following violence-related care. Evidence shows that violence, notably injury, can lead to increased infection rates [69–71] and disability [72, 73]. Care-seeking in the 4–6 months after seeking care for violence may result from the gradual onset of chronic symptoms like pain and musculoskeletal injuries, and as infections become symptomatic, prompting individuals to seek additional care [70].

Our study found that individuals who used PHC for the selected conditions after experiencing violence, were more than twice as likely to be hospitalised for those same conditions, compared to those who did not use PHC for these conditions. Additionally, we did not find any statistically significant interactions between sex, race/skin colour, education level and income and post-violence PHC utilisation. One plausible explanation is that this study was underpowered to detect inequalities

in this association. It may also be the case that increases in PHC use after violence is similar across socioeconomic groups. However, this is unlikely to be the case given well-established differences in exposure to violence, PHC utilisation, and inequalities in violence-related mortality by sex [74–77]. This stems from greater morbidity in women, gender-based social hierarchies, and differing health perceptions both internationally [78–82] and in Brazil [83–85]. Victims of violence are often Black or Pardo, male, and from lower-income or lower-educational backgrounds [75, 76, 86, 87], suggesting further evidence is needed to explore socioeconomic inequalities in the healthcare impacts of violence.

A key strength of this study was the use of individual-level data with linked health and socioeconomic data facilitating comprehensive analyses. The screening of clinical free text records to identify individuals who experienced violence is another strength, and novel for a study from a LMIC. The CEM achieved multiple exact matches for 2,038 of the 2,049 exposed individuals improving covariate balance, reducing the potential selection bias, and retaining a large proportion of exposed individuals.

Nevertheless, there are limitations. CEM reduced our study population, but sensitivity analyses showed the matched sample was representative of the original 3.1 million cohort. Sociodemographic data was self-reported with potential for misclassification, but is unlikely that biases were sufficiently large to invalidate the findings or correlated with exposure or outcomes. The study population consisted of individuals that used PHC, suggesting the true healthcare impacts of violence are underreported. As our study population was already using healthcare, increases in utilisation after using healthcare due to violence may seem smaller, potentially underestimating violence's true effect on healthcare use. Additionally, since women are more likely to use PHC services, they are potentially overrepresented in the sample, reflecting broader healthcare-seeking patterns rather than a specific bias toward certain types of violence. Socioeconomically disadvantaged populations experience greater barriers in access to healthcare, suggesting that underestimations are greater in these populations. Given the low-income focus of the study population, the findings may not be generalisable to the entire city or other settings. Another limitation is that violence exposure was defined by PHC consultations or hospitalisations following a violent event and therefore excludes individuals that were exposed to violence and did not use care and those that may have been exposed to extreme violence which led to death or unreported emergency care. This is likely substantially under-estimating the true healthcare impacts of violence. To mitigate selection bias from defining exposure through healthcare use, we limited our analysis to healthcare users. However, this could

introduce bias by potentially overestimating the effects of violence, as those who do not seek care were excluded.

Additionally, we used broad inclusion criteria for healthcare visits related to violence, encompassing any mention of violence regardless of timing. This may reduce specificity in distinguishing recent versus past, or direct versus indirect exposures. Finally, it is possible that some cases of domestic violence including rape and sexual abuse/assault were inadvertently included despite efforts to exclude them, for example by relying on ICD-10 and ICPC codes. Although free-text screening was conducted in an attempt to both identify further cases of violence not coded by ICD-10 and ICPC and to minimise the inclusion of domestic violence cases, complete exclusion is challenging given data limitations.

The findings of this study underscore the need for comprehensive and multifaceted policy interventions to mitigate the adverse health effects associated with violence in Rio de Janeiro. Given the associations between violence and healthcare utilisation were predominantly in the short-term, it is important to strengthen health services to effectively respond to violence-related health issues. Immediate actions could include an increase in mental health and maternal care services in high-risk areas and ensuring efficient referral pathways to specialised services. Enhanced training for healthcare providers in PHC clinics which are often Brazilians' first point of contact with the health system, especially in low-income areas, is important [88]. Improving access to healthcare and ensuring health facilities are adequately equipped and staffed to handle violence-related cases, whilst ensuring the safety of healthcare workers is essential. Long-term strategies should address the root causes of violence, investing in social and economic development initiatives and public services to disrupt the cycles of poverty, socioeconomic inequality, and consequently violence. Overall, a collaborative approach involving local government, public, and social services, is essential to create a safer and consequently, healthier urban environment.

Conclusion

There were short-term increases in PHC utilisation in the first six months for individuals using healthcare due to violence in Rio de Janeiro, Brazil. This increase in PHC utilisation is largely driven by a higher number of PHC consultations for poor pregnancy and birth outcomes, poor mental health infections, and disability associated outcomes.

Appendix 1 Development of a conceptual framework of the pathways through which violence increases health risk factors

Applying and expanding Scott-Storey's [21] and Guha-Sapir & van Panhuis' [22] frameworks on violence and health, we developed a conceptual framework illustrat-

ing four key pathways through which violence in general increases health risk factors (Fig. 1). This conceptual framework was then used to inform the health conditions to consider as PHC and hospitalisation outcomes in the study.

First, chronic exposure to violence triggers neuroendocrine, metabolic, and immunological changes, resulting in chronic inflammation and elevated blood pressure, key markers of cardiovascular risk. Research in Sweden has shown that violence increases coronary heart disease by 75% and 39% in females and males, respectively [88]. Second, coping strategies including smoking, substance abuse, and overeating, adopted in response to stress from violence, persist even after abuse ends, exacerbating cardiovascular risk factors such as high cholesterol, obesity, and diabetes. Evidence from high income countries including Australia and the United States has demonstrated associations between neighbourhood violence and increased body mass index and type 2 diabetes, respectively [31, 89]. Similarly, in Brazil, women living in the least socially cohesive neighbourhoods and those considered the most violent had a 25% increased risk of obesity [90].

Third, violence is strongly associated with poor mental health [91, 92]. Violence's toll on mental health, specifically depression, contributes to cardiovascular disease through elevated cortisol levels and inflammatory markers. Violence amplifies chronic stress and associated health risks. Elevated cortisol levels from stress, exacerbated by experiencing violence, can disrupt women's reproductive health, leading to adverse pregnancy outcomes as evidenced in the United States [32]. In low- and middle-income countries, especially in Latin America, violence and poor neighbourhood socioeconomic conditions hampers youth access to sexual and reproductive healthcare, contributing to higher rates of adolescent births [17, 93, 94]. Limited access to healthcare due to violence also heightens the risk of sexually transmitted infections, as seen in cases of intimate partner violence [95, 96]. Lastly, violence-related injuries, including those from weapons, increase infection rates [69–71] and may result in long-term disabilities like paralysis and chronic pain [72, 73].

Appendix 2 Methods for violence free text search of PHC consultation records

The free text search covered terms (in Portuguese) such as 'bullet', 'violence', 'aggression', 'assault', 'crime', 'robbery', 'trafficker', 'assassinated' and excluded terms related to domestic violence including 'domestic', 'partner', 'husband', 'wife', 'home', 'residence', 'suspected', and 'negligence' – to ensure the focus remained on violence outside the home and perpetrated by strangers. Records were included if there was any mention of violence in the consultation record, regardless of the time since occurrence and whether exposure was direct (e.g. victim, injury) or indi-

rect (e.g. loss of a loved one, witnessing urban violence). Spelling errors and typographical errors were accounted for by including both the noun and verb forms of associated terms and using truncations where appropriate.

To validate the search, a random sample of 100 records per violence sub-category (weapon & firearms; physical assault & injury; crime & gang activity) were hand-screened, totalling 300 records, to ensure only relevant cases of violence were identified. Sub-categories of violence were used to facilitate the search and screening process, but were combined for the analysis. If incorrect terms were identified in the screening process, the search was modified, and screening was repeated. Sample screening was repeated until a plateau was reached, defined as minimal changes in precision ($\text{true positives} / (\text{true positives} + \text{false positives})$) and the number of records identified by the search, indicating that further screening would not significantly improve results. In total, we performed eight rounds of sample screening. In the final round of screening all 2,734 records identified were manually screened (weapon & firearms, 823; physical assault & injury, 1,443; crime & gang activity, 468).

Table A2 Final* violence free text search per violence category

Search Strategy

Weapon & Firearm	(\bTIRO\b.*\^bTIRO\b.*\bTIROTEIO\b\^bARMA\b.*\b(ARMA DO A)DO\b.* PISTOL.* ESFAQUEA.* FACAD.* BALEAD.* FERID.* FACA) AND NOT (NAMORAD.* PARCEIR.* MARIDO MULTHER ESPOS.* FAM.* CONJUGE COMPANHEIR.* CONSORTE DOMESTIC.* CASA RESIDENCIA FAMILIA.* ESTUPR.* CONJUGAL NEGULIG.* SUSPEIT.* SEXUAL NOIV.*) AND NOT (TIROID.* LEVOTIROXINA TIRO DUVIDAS REFACA FARMACIA RETIROU DIALEVOTIROXINA FARMA.* TUMEFAC.* FARMACEUTICO FARMACOLOGICO ALARMA TIROU RETIRO AFERIDA RETIRADA HIPERTIROIDISMO LECOTIROXINA ARMARIO CORRIMENTO \bDOR TIPO FACAD.*\b \bCOMO UMA FACADA\b \bDOR EM FACADA\b \bTIPO FACAD.*\b \bEM FACADA\b \bTIRO A.*\b \bTIRO G.*\b \bTIRO D.*\b \bTIRO O DIAG.*\b \bCOMO SE FOSSE FACAD.*\b \bCOMO SE FOSSE FACAD.*\b \bCOMO UMA FACAD.*\b \bTIRO S.*\b \bEM FACAD.*\b ALTERA.*\b SENSA.* DE FACAD.*\b \bBESTOU MEIO BALEAD.*\b \bTIRO OS PONTO.*\b PISTOLA PNEUMATICA CORRIMENTO \bEM FACADA.*\b NAO TIRO ELE NUNCA TIRO PISTOLA DE PINO \bTIRO O DIANG.*\b GRUPO DE ENTREGA MATOU A M.* CORRIEM.* AUTIAGRESSAO QUE FACA EXAME SENSACAO DE FACADAS COMO SE FOSSE UMA FACADA)
---------------------	--

Table A2 Final* violence free text search per violence category

	Search Strategy
Assault & Injury	(\bTIRO\b.* \^bTIRO\b.* \bTIROTEIO\b \^bARMA\b.* \b(ARMA DO A)DO\b.* PISTOL.* ESFAQUEA.* FACAD.* BALEAD.* FERID.* FACA) AND NOT (NAMORAD.* PARCEIR.* MARIDO MULHER ESPOS.* FAM.* CONJUGE COMPANHEIR.* CONSORTE DOMESTIC.* CASA RESIDENCIA FAMILIA.* ESTUPR.* CONJUGAL NEGLEGIG.* SUSPEIT.* SEXUAL NOIV.*) AND NOT (COMPANHEIR.*\sQUE\STEM\sQUAD-ROS\sDE\sAGRESSIVIDADE AGRESSIVA/AGRESSIVO VIOLETA AGRESSAO\sFISICA\sDE\sNAMORADO NEGA\sAGRESSAO PACIENTE\sAPRESENTA/COM\sAGRESSIVIDADE CRIANCA\sESTA\sAGRESSIVA VIOLENCIA\sSEXUAL ERITEMATOVIOLACEAS DOR\sVIOLENTA AGRESSAO\sFISICA\sFACADA\sPOR\sPARTE\sD.* APRESENTA-SE\sAGRESSIVA EVIOLUCAO AGRESSIVO AGRESSIVA VIOLACE.* RESSALTA AGRESSIVIDA.* \bCOR VIOL.* \b VIOLA.* VIOLINO VIOLÃ AUTO AGRESSAO AUTO-AGRESSAO PESSOA CONHECIDA ESCOLA COMPARTAMENTO HETEROAGRESSAO AGRESSAO POR PARTE \bAGRESSAO D.*\b CACHORRO AGRESSAO SOLAR EPISODIOS DE AGRESSIV.* VOLUMEN \bPELO SEU COMPANHEIR.* O FILHO A MALTRATA EPISODIOS DE AGRESSAO CRIANCA AGITADA AGRESS.* IN FANTIL OBSTETRICA MAGRESSA AGRESS.* A.* FILH.* HETEROAGRESS.* AGRESS.* CONTRA A FILHA AGRESS.* CONTRA O FILHO.* NEGA VIOL.* NEGA AGRESS.* ELE E VIOL.* MAE A VIOLENTO.* IMPULSOS VIOLENTOS)
Crime & Gang	(\bTIRO\b.* \^bTIRO\b.* \bTIROTEIO\b \^bARMA\b.* \b(ARMA DO A)DO\b.* PISTOL.* ESFAQUEA.* FACAD.* BALEAD.* FERID.* FACA) AND NOT (NAMORAD.* PARCEIR.* MARIDO MULHER ESPOS.* FAM.* CONJUGE COMPANHEIR.* CONSORTE DOMESTIC.* CASA RESIDENCIA FAMILIA.* ESTUPR.* CONJUGAL NEGLEGIG.* SUSPEIT.* SEXUAL NOIV.*) AND NOT (LACRIMEJA LACRIMEJAMENTO GANGLIO.* GANGLIONAR LACRIMAL LACRIMEJANDO LACRIMA INFACCAO PARA\sEL.*\sN.*\sROUBAR.*\sAS\sCOISAS ENDOC.* LAVCRIME.* LACRIMEO GANGRE.* GANGK.* GANGIO.* GANGATE GANGATA LACRIME.*\bASSINAD.*\b AL ECRIM DISCRIM.* TUMEFAC.* GANGO.* GARGANGTA GANGANGTA GANGLOS GANGLONAR GANGLI GANGLIAO MOSTROUB HIPOCRIMICA ACRIMEJAMENTO BACRIM LACRIMRJAMENTO CRIMENTO LACRIMMASA HIPERCRIMICAS LACRIMIANDO LINGANGITES GANGANTA GARGANGANTA GANGLIO.* RUBEFACCAO GANGGLIO GANGRANA DARACRIM LACRIMONASAL REFACCAO FUROUB HASSASSIN.* LECRIMEJ.* PROVUROUB.* MASSASSINAL LCRIMEJAMENT.* CRIMA ASSINTOMATICO ACRIMIO GANGRIO GANGARTA LACRIMOGNEO FILHA A ROUBA FILHO A AGRIDE ASSASSINARLA VOU ROUBAR LACRIM.* SENTE IMPULSO DE ROUBAR IRM.* ROUBOUJA ROUBOU VARIAS COISAS)

Sources: Nakamura et al. [97] and authors' own work

*Prior to manual screening of records identified by search

Appendix 3 International classification of disease (ICD-10) and international classification of primary care (ICPC) codes for analysis of health conditions with risk factors associated with experiencing violence

HEALTH CONDITION (ANALYSIS GROUPS)	ICD-10 CATEGORIES	CODES	ICPC CATEGORIES	CODES
Adverse Pregnancy and Birth Outcomes	Outcome of delivery (including stillbirth)	Z37.1, Z37.3, Z37.4, Z37.6, Z37.7	Bleeding	W03, W17
	Premature rupture of membranes	O42	Pregnancy complications	W27, W29, W71, W75, W80, W81
	Preeclampsia	O11, O14, O15	Post-partum symptom/complaint other	W18
	Gestational hypertension	O13, O16	Unwanted pregnancy	W79
	Gestational diabetes	O24.4	Outcome of delivery	W92, W93, W99, A93
	Poor foetal growth	O36.5, O36.8, O36.9	Abortion	W82, W83
	Placental abruption	O45, O46	Gestational diabetes	W85
	Spontaneous preterm labour	O60.1		
	Other conditions complicating pregnancy	O99		
Cardiovascular Dis- ease & Stroke	Hypertensive diseases	I10-I15	Pain, complaints, and risk factors for cardiovascular disease	K01-K03, K22, K29, A11
	Ischaemic heart diseases	I20-I25	Abnormal heart rhythms	K04, K05, K78-K81
	Heart failure	I50	Ischaemic heart disease	K74-K76, K89-K94
	Cardiac arrest	I46	Heart disease and heart failure	K77, K82-K84, K99
	Cerebrovascular diseases	I60-I69	Hypertensive diseases	K85-K88
	Abnormal heart rhythms	I47-I49		
Diabetes	Type 2 diabetes mellitus	E11-E11.9	Diabetes non-insulin dependent	T90
	Abnormal glucose tolerance test (prediabetes)	R73		
Obesity	Obesity	E66-E66.9	Obesity	T82
	BMI for overweight and obese	Z68.25-Z68.45	Overweight	T83
			Weight gain	T07
			Excessive appetite	T02
Mental Health	Substance abuse/dependence disorders	F10-F19	Substance abuse/dependence disorders	P15, P17-P19
	Psychotic syndromes	F20-F25, F28, F29	Psychotic syndromes	P29, P71-P73, P98, P99
	mood affective disorders	F30-F34, F38, F39	mood affective disorders	P03, P76
	Neurotic, stress-related and somatoform disorders (including anxiety)	F40-F45, F48	Neurotic, stress-related and somatoform disorders (including anxiety)	P01, P02, P74, P79, P82
	Personality and behaviour disorders	F50, F51, F60-F63, F91-F94	Personality and behaviour disorders	P80, P86
	Suicide/associated outcomes	X60-X84, R45.851	Suicide/associated outcomes	P77
Infections & Infec- tious Diseases	Post-traumatic wound infection, not elsewhere classified	T79.3	Infectious disease other/NOS	A78
	Additional codes (B95-B98) to identify infectious agent:	B95-B98	Infections of circulatory system	K70
	• Streptococcus and staphylococcus as the cause of diseases classified to other chapters		Blood infections	B70, B71
	• Other specified bacterial agents as the cause of diseases classified to other chapters		Digestive infections	D70-D73
	• Viral agents as the cause of diseases classified to other chapters			
	• Other specified infectious agents as the cause of diseases classified to other chapters			

HEALTH CONDITION (ANALYSIS GROUPS)	ICD-10 CATEGORIES	CODES	ICPC CATEGORIES	CODES
Disability	Infections with a predominantly sexual mode of transmission	A50-A64	HIV-infection/aids	B90
	Contact with and exposure to infections with a predominantly sexual mode of transmission	Z20.2	Female sexually transmitted infections	X70-X73, X90-X92
	Special screening examination for infections with a predominantly sexual mode of transmission and human immunodeficiency virus [HIV]	Z11.3, Z11.4	Male sexually transmitted infections	Y70-Y76
	Other infections with a predominantly sexual mode of transmission complicating pregnancy, childbirth and the puerperium	O98.3		
	Chronic pain	R52.1, R52.2	Pain general/multiple sites	A01
	Amputation	Y83.5, T05, T13.6, T11.6	Paralysis/weakness	N18
	Paralysis	T14.4		

Sources: Wonca International Classification Committee [98], World Health Organization [44], Sun et al. [99], de Fatima Marinho de Souza et al. [100], Dugan & Shubrook [101], Suissa et al. [102], Ho et al. [103], Glasgow et al. [104]

Appendix 4 Coarsened exact matching output

(a) Covariate Balance

Multivariate L1 distance: 0.8333353

Univariate imbalance:

	L1	Mean	Min	25%	50%	75%	Max
microarea_code	0.02817	-0.03641	1	0	0	0	-3
sex	0.11968	0.11968	0	0	0	0	0
race_skinc	0.07057	0.03622	0	0	0	0	0
age_group	0.15984	0.35934	0	1	0	0	0
education_highest	0.281	0.43787	0	1	1	0	0
hh_income	0.06818	-0.22771	0	0	0	0	0
insurance	0.06516	-0.06516	0	0	0	0	0
bf	0.06363	0.06363	0	0	0	0	0

(b) Matching Summary

Number of strata: 63,004

Number of matched strata: 1,550

	0	1
All	3146334	2049
Matched	527181	2038
Unmatched	2619153	11

Multivariate L1 distance: 5.309e-14

Univariate imbalance:

	L1	Mean	Min	25%	50%	75%	Max
microarea_code	6.90E-14	1.30E-12	0	0	0	0	0
sex	1.50E-14	-1.80E-13	0	0	0	0	0
race_skinc	2.30E-14	1.10E-12	0	0	0	0	0
age_group	6.10E-14	5.50E-13	0	0	0	0	0
education_highest	7.90E-14	1.70E-13	0	0	0	0	0
hh_income	3.00E-14	1.20E-12	0	0	0	0	0
insurance	1.20E-14	2.10E-15	0	0	0	0	0
bf	1.30E-14	2.40E-14	0	0	0	0	0

Appendix 5 Summary statistics of PHC utilisation and hospitalisation by sociodemographic variables and violence exposure

Characteristics	Counts of PHC Consultations			Counts of Hospitalisations		
	Did Not Use Healthcare due to Violence	Used Healthcare due to Violence	Total	Did Not Use Healthcare due to Violence	Used Healthcare due to Violence	Total
Sex						
Male	206,646	1,837	208,483	2,426	24	2,450
Female	818,336	8,192	826,528	7,515	63	7,578
Race/skin colour						
White	351,867	3,090	354,957	2,971	28	2,999
Black	99,274	1,875	101,149	983	11	994
Pardo (Mixed Race)	567,777	4,857	572,634	5,880	48	5,928
Amarelo (Asian), Indigenous, Other or Unknown (Missing)	6,064	207	6,271	107	0	107
Highest education level						
None/pre-school/literacy class	73,383	1,357	74,740	949	4	953
Elementary school	535,416	5,296	540,712	3,628	36	3,664
High school or higher education	321,059	2,409	323,468	3,600	40	3,640
None reported (missing)	95,124	967	96,091	1,764	7	1,771
Age group (years)						
0–18	8,258	88	8,346	121	2	123
19–34	75,168	935	76,103	3,576	24	3,600
35–54	346,152	4,559	350,711	3,344	26	3,370
55–64	321,357	2,608	323,965	1,574	25	1,599
65+	274,047	1,839	275,886	1,326	10	1,336
Household income						
Some salary reported to less than full minimum salary	69,741	1,735	71,476	620	29	649
More than one minimum salary but less than two	222,564	2,458	225,022	1,663	14	1,677
Two minimum salaries or more	11,015	287	11,302	44	0	44
No salary reported (unknown or missing)	721,662	5,549	727,211	7,614	44	7,658
Registered with Private Health Insurance						
No	1,021,410	9,926	1,031,336	9,904	84	9,988
Yes	3,572	103	3,675	37	3	40
Bolsa Família-Claiming Family						
No	989,742	8,277	998,019	9,406	76	9,482
Yes	35,240	1,752	36,992	535	11	546
N	1,024,982	10,029	1,035,011	9,941	87	10,028

PHC primary healthcare

Appendix 6 Person-years and crude rates of primary healthcare utilisation and hospitalisation (per 1,000 person-years) overall and by sociodemographic variables and whether individuals used healthcare due to violence

Characteristics	Person-Years			Rates of PHC Utilisation			Rates of Hospitalisation		
	Did Not Use Healthcare due to Violence	Used Healthcare due to Violence	Total	Did Not Use Healthcare due to Violence	Used Healthcare due to Violence	Total	Did Not Use Healthcare due to Violence	Used Healthcare due to Violence	Total
Sex									
Male	425,229.50	2,162.11	427,391.60	485.96	849.63	487.80	5.71	11.10	5.73
Female	1,036,465.00	4,368.73	1,040,833.00	789.55	1,875.15	794.10	7.25	14.42	7.28
Race/Skin Colour									
White	436,959.70	1,835.18	438,794.90	805.26	1,683.76	808.94	6.80	15.26	6.83
Black	121,165.20	1,218.13	122,383.30	819.33	1,539.25	826.49	8.11	9.03	8.12
Pardo (Mixed Race)	888,518.60	3,290.43	891,809.00	639.02	1,476.10	642.10	6.62	14.59	6.65
Amarelo (Asian), Indigenous, Other or Unknown (Missing)	15,050.72	187.10	15,237.82	402.90	1,106.35	411.54	7.11	0.00	7.02
Education									
None/pre-school/literacy class	115,684.30	969.63	116,654.00	634.34	1,399.50	640.70	8.20	4.13	8.17
Elementary school	639,046.00	3,117.25	642,163.20	837.84	1,698.93	842.02	5.68	11.55	5.71
High school or higher education	535,251.50	1,650.27	536,901.80	599.83	1,459.76	602.47	6.73	24.24	6.78
None reported (missing)	171,712.40	793.69	172,506.10	553.97	1,218.36	557.03	10.27	8.82	10.27
Age Group (years)									
0–18	204,788.20	655.50	205,443.70	40.32	134.25	40.62	0.59	3.05	0.60
19–34	380,602.90	1,912.88	382,515.80	197.50	488.79	198.95	9.40	12.55	9.41
35–54	493,147.60	2,393.04	495,540.60	701.92	1,905.11	707.73	6.78	10.86	6.80
55–64	226,310.20	938.45	227,248.60	1,419.98	2,779.05	1,425.60	6.96	26.64	7.04
65+	156,845.40	630.97	157,476.40	1,747.24	2,914.54	1,751.92	8.45	15.85	8.48
Household Income									
Some salary reported to less than full minimum salary	90,770.79	1,077.23	91,848.02	768.32	1,610.61	778.20	6.83	26.92	7.07
More than one minimum salary but less than two minimum salaries	280,508.80	1,339.11	281,847.90	793.43	1,835.55	798.38	5.93	10.45	5.95
Two minimum salaries or more	9,570.12	208.72	9,778.84	1,150.98	1,375.03	1,155.76	4.60	0.00	4.50
No salary reported (unknown or missing)	1,080,845.00	3,905.78	1,084,750.00	667.68	1,420.72	670.40	7.04	11.27	7.06
Registered with Private Health Insurance									
No	1,451,525.00	6,355.04	1,457,880.00	703.68	1,561.91	707.42	6.82	13.22	6.85
Yes	10,168.88	175.80	10,344.68	351.27	585.90	355.26	3.64	17.07	3.87
Bolsa Família-Claiming Family									
No	1,373,346.00	5,321.11	1,378,668.00	720.68	1,555.50	723.90	6.85	14.28	6.88
Yes	88,347.85	1,209.73	89,557.57	398.88	1,448.26	413.05	6.06	9.09	6.10
Total	1,461,694.00	6,530.84	1,468,225.00	701.23	1,535.64	704.94	6.80	13.32	6.83

PHC Primary Healthcare

Appendix 7 Odds ratios from a fixed effects logistic regression model of conditions with risk factors associated with using healthcare due to violence for hospitalisation, adjusted for PHC utilisation for the same conditions

Characteristics	OR	95% CI
Months after violence exposure		
0 months or did not seek healthcare for violence	1 (ref)	–
1–3	1.15	(0.48–2.77)
4–6	0.80	(0.27–2.40)
7–12	0.22 *	(0.05–0.94)
Over 12	0.56	(0.24–1.31)
PHC Utilisation for selected conditions		
No	1 (ref)	–
Yes	2.08 ***	(1.92–2.26)
Total Observations (N)	290,393	
Total Groups (N)	7,004	

Fixed effects logistic regression, adjusted for year and month. OR Odds Ratios, 95% CI 95% Confidence Intervals. * $p < 0.05$; ** $p < 0.01$; *** $p < 0.001$

Appendix 8 Odds ratios from fixed effects logistic regression models of conditions with risk factors associated with using healthcare for violence by months following healthcare use for PHC utilisation by all sociodemographic variables by health condition

Characteristics	Pregnancy & Birth Outcomes ^a		CVD and Stroke, Diabetes & Obesity		Mental Health ^b		Infections & Infectious Diseases		Disability	
	OR	95% CI	OR	95% CI	OR	95% CI	OR	95% CI	OR	95% CI
Months after using healthcare for violence										
0, or did not use healthcare for violence	1 (ref)	–	1 (ref)	–	1 (ref)	–	1 (ref)	–	1 (ref)	–
1–3	4.25	(2.14–8.44)	0.98	(0.86–1.12)	1.50	(1.27–1.77)	1.10	(0.73–1.67)	0.41	(0.15–1.13)
4–6	1.13	(0.38–3.38)	0.89	(0.77–1.02)	1.12	(0.92–1.35)	1.00	(0.63–1.58)	1.39	(0.67–2.89)
7–12	2.57	(1.20–5.53)	0.82	(0.73–0.93)	0.91	(0.76–1.07)	0.82	(0.53–1.26)	0.84	(0.40–1.76)
Over 12	1.72	(0.77–3.84)	0.76	(0.69–0.85)	0.67	(0.57–0.77)	0.41	(0.27–0.61)	0.54	(0.27–1.07)
Total Observations (N)	142,196		5,597,694		1,160,452		452,943		349,260	
Total Groups (N)	3,758		159,151		32,320		12,837		7,560	

Separate fixed effects logistic regressions; adjusted for year and month. PHC Primary Healthcare, OR Odds Ratios, 95% CI 95% Confidence Intervals, CVD cardiovascular disease. * $p < 0.05$; ** $p < 0.01$; *** $p < 0.001$. ^aPregnancy & Birth Outcomes model was restricted to females

Appendix 9 Odds ratios from a fixed effects logistic regression model of PHC utilisation for conditions with risk factors associated with using healthcare for violence from interactions between by months following healthcare use and sex

Characteristics	OR	95% CI
Months after using healthcare for violence		
0 months or did not seek healthcare for violence	1 (ref)	–
1–3	1.33 *	(1.05–1.69)
4–6	1.02	(0.78–1.34)
7–12	0.78	(0.61–1.00)
Over 12	0.91	(0.75–1.12)
Months after violence exposure × Sex		
1–3 × Female	0.80	(0.61–1.05)
4–6 × Female	0.95	(0.70–1.29)
7–12 × Female	1.10	(0.84–1.45)
Over 12 × Female	0.82	(0.65–1.03)
Total Observations (N)	6,802,113	
Total Groups (N)	193,275	

Fixed effects logistic regression, adjusted for year and month. OR Odds Ratios, 95% CI 95% Confidence Intervals. * $p < 0.05$; ** $p < 0.01$; *** $p < 0.001$

Appendix 10 Odds ratios from a fixed effects logistic regression model of PHC utilisation for conditions with risk factors associated with using healthcare for violence from interactions between by months following healthcare use and race/skin colour

Characteristics	OR	95% CI
Months after using healthcare for violence		
0 months or did not seek healthcare for violence	1 (ref)	–
1–3	1.10	(0.90–1.34)
4–6	0.87	(0.69–1.09)
7–12	0.96	(0.79–1.16)
Over 12	0.87	(0.74–1.03)
Months after violence exposure × Race/skin colour		
1–3 × Black	0.97	(0.71–1.32)
1–3 × Pardo (Mixed Race)	1.04	(0.81–1.34)
1–3 × Amarelo (Asian), Indigenous, Other or Unknown (Missing)	1.03	(0.47–2.26)
4–6 × Black	1.24	(0.88–1.75)
4–6 × Pardo (Mixed Race)	1.20	(0.91–1.60)
4–6 × Amarelo (Asian), Indigenous, Other or Unknown (Missing)	0.59	(0.24–1.43)
7–12 × Black	0.86	(0.63–1.16)
7–12 × Pardo (Mixed Race)	0.85	(0.66–1.09)
7–12 × Amarelo (Asian), Indigenous, Other or Unknown (Missing)	0.58	(0.30–1.14)
Over 12 × Black	0.86	(0.66–1.11)
Over 12 × Pardo (Mixed Race)	0.88	(0.72–1.09)
Over 12 × Amarelo (Asian), Indigenous, Other or Unknown (Missing)	0.38 **	(0.21–0.69)
Total Observations (N)	6,802,113	
Total Groups (N)	193,275	

Fixed effects logistic regression, adjusted for year and month. OR Odds Ratios, 95% CI 95% Confidence Intervals. * $p < 0.05$; ** $p < 0.01$. *** $p < 0.001$

Appendix 11 Odds ratios from a fixed effects logistic regression model of PHC utilisation for conditions with risk factors associated with using healthcare for violence from interactions between by months following healthcare use and highest education level

Characteristics	OR	95% CI
Months after violence exposure		
0 months or did not seek healthcare for violence	1 (ref)	–
1–3	0.92	(0.67–1.27)
4–6	0.94	(0.66–1.33)
7–12	0.64 **	(0.47–0.87)
Over 12	0.61 ***	(0.48–0.78)
Months after violence exposure × Highest education level		
1–3 × Elementary school	1.22	(0.86–1.73)
1–3 × High school or higher education	1.06	(0.72–1.56)
1–3 × None reported (missing)	1.80 **	(1.16–2.79)
4–6 × Elementary school	1.04	(0.71–1.54)
4–6 × High school or higher education	1.05	(0.68–1.62)
4–6 × None reported (missing)	1.10	(0.67–1.82)
7–12 × Elementary school	1.31	(0.94–1.84)
7–12 × High school or higher education	1.38	(0.94–2.01)
7–12 × None reported (missing)	1.71 *	(1.11–2.62)
Over 12 × Elementary school	1.22	(0.93–1.60)
Over 12 × High school or higher education	1.47 *	(1.08–2.00)
Over 12 × None reported (missing)	1.61 *	(1.12–2.32)
Total Observations (N)	6,802,113	
Total Groups (N)	193,275	

Fixed effects logistic regression, adjusted for year and month. OR Odds Ratios, 95% CI 95% Confidence Intervals. * $p < 0.05$; ** $p < 0.01$; *** $p < 0.001$

Appendix 12 Odds ratios from a fixed effects logistic regression model of PHC utilisation for conditions with risk factors associated with using healthcare for violence from interactions between by months following healthcare use and household income

Characteristics	OR	95% CI
Months after violence exposure		
0 months or did not seek healthcare for violence	1 (ref)	–
1–3	0.93	(0.71–1.20)
4–6	1.04	(0.79–1.38)
7–12	0.92	(0.71–1.19)
Over 12	0.79 *	(0.64–0.98)
Months after violence exposure × Household income		
1–3 × More than one minimum salary but less than two minimum salaries	1.34	(0.95–1.88)
1–3 × Two minimum salaries or more	1.21	(0.64–2.29)
1–3 × No salary reported (unknown or missing)	1.22	(0.91–1.64)
4–6 × More than one minimum salary but less than two minimum salaries	0.84	(0.57–1.24)
4–6 × Two minimum salaries or more	0.96	(0.44–2.08)
4–6 × No salary reported (unknown or missing)	0.96	(0.70–1.33)
7–12 × More than one minimum salary but less than two minimum salaries	1.12	(0.80–1.57)
7–12 × Two minimum salaries or more	0.71	(0.33–1.52)
7–12 × No salary reported (unknown or missing)	0.84	(0.63–1.13)
Over 12 × More than one minimum salary but less than two minimum salaries	0.99	(0.75–1.32)
Over 12 × Two minimum salaries or more	1.53	(0.74–3.15)
Over 12 × No salary reported (unknown or missing)	0.96	(0.75–1.23)
Total Observations (N)	6,802,113	
Total Groups (N)	193,275	

Fixed effects logistic regression, adjusted for year and month. OR Odds Ratios, 95% CI 95% Confidence Intervals. * $p < 0.05$; ** $p < 0.01$; *** $p < 0.001$

Appendix 13 Adjusted odds ratios from random effects logistic regression models of conditions with risk factors associated with using healthcare due to violence for PHC utilisation and hospitalisation by all sociodemographic variables

Characteristics	PHC Utilisation		Hospitalisation	
	aOR	95% CI	aOR	95% CI
Months after using healthcare due to violence				
0, or did not use healthcare due to violence	1 (ref)	–	1 (ref)	–
1–3	1.36 ***	(1.20–1.55)	1.95	(0.85–4.52)
4–6	1.15	(1.00–1.33)	1.33	(0.51–3.45)
7–12	0.97	(0.85–1.11)	0.36	(0.11–1.19)
Over 12	0.87	(0.76–1.01)	0.69	(0.35–1.36)
Sex				
Male	1 (ref)	–	1 (ref)	–
Female	1.34	(1.32–1.37)	1.43	(1.34–1.53)
Race/skin colour				
White	1 (ref)	–	1 (ref)	–
Black	1.32 ***	(1.29–1.36)	1.15 **	(1.04–1.26)
Pardo (Mixed Race)	1.03 ***	(1.01–1.05)	1.02	(0.96–1.08)
Amarelo (Asian), Indigenous, Other or Unknown (Missing)	0.58 ***	(0.53–0.64)	0.86	(0.65–1.12)
Highest education level				
None/pre-school/literacy class	1 (ref)	–	1 (ref)	–
Elementary school	1.33 ***	(1.29–1.38)	0.45 ***	(0.41–0.50)
High school or higher education	1.05 **	(1.02–1.08)	0.42 ***	(0.38–0.47)
None reported (missing)	0.93 ***	(0.90–0.97)	0.65 ***	(0.58–0.74)
Age group (years)				
0–18	1 (ref)	–	1 (ref)	–
19–34	5.47	(5.23–5.72)	20.20	(16.42–24.84)
35–54	30.66 ***	(29.40–31.97)	10.92	(8.87–13.43)
55–64	101.84 ***	(97.65–106.21)	10.92 ***	(8.81–13.54)
65+	156.50 ***	(150.07–163.21)	16.28 ***	(13.17–20.11)
Household income				
Some salary reported to less than full minimum salary	1 (ref)	–	1 (ref)	–
More than one minimum salary but less than two	1.03	(1.00–1.07)	0.85 *	(0.75–0.96)
Two minimum salaries or more	1.34 ***	(1.25–1.43)	0.70 ***	(0.48–1.01)
No salary reported (unknown or missing)	0.90 ***	(0.87–0.93)	0.92 ***	(0.82–1.03)
Registered with Private Health Insurance				
No	1 (ref)	–	1 (ref)	–
Yes	0.64 ***	(0.57–0.70)	0.71	(0.49–1.02)
Bolsa Família-Claiming Family				
No	1 (ref)	–	1 (ref)	–
Yes	0.96 *	(0.92–1.00)	1.20 **	(1.07–1.35)
Total Observations (N)	17,875,018		17,875,018	
Total Groups (N)	529,219		529,219	

Separate fully adjusted random effects logistic regressions per outcome (PHC utilisation and hospitalisation); adjusted for sex, race/skin colour, education level, age group, household income, private health insurance, Bolsa Família-receiving family, year, and month. Robust standard errors. PHC Primary Healthcare, aOR Adjusted Odds Ratios, 95% CI 95% Confidence Intervals. * $p < 0.05$; ** $p < 0.01$; *** $p < 0.001$

Appendix 14 Adjusted rates ratios from Poisson regression models of conditions with risk factors associated with using healthcare due to violence for PHC utilisation and hospitalisation by all sociodemographic variables

Characteristics	PHC Utilisation		Hospitalisation	
	aRR	95% CI	aOR	95% CI
Months after using healthcare due to violence				
0, or did not use healthcare due to violence	1 (ref)	–	1 (ref)	–
1–3	1.03	(0.95–1.12)	1.14	(0.47–2.79)
4–6	0.95	(0.86–1.04)	0.79	(0.27–2.28)
7–12	0.85 ***	(0.77–0.93)	0.22 *	(0.06–0.85)
Over 12	0.77 ***	(0.69–0.85)	0.59	(0.29–1.17)
Total Observations (N)	6,805,893		290,393	
Total Groups (N)	194,389		7,004	

Separate fully adjusted random effects logistic regressions per outcome (PHC utilisation and hospitalisation); adjusted for sex, race/skin colour, education level, age group, household income, private health insurance, Bolsa Família-receiving family, year, and month. Robust standard errors. PHC Primary Healthcare, aRR Adjusted Rate Ratios, 95% CI 95% Confidence Intervals. * $p < 0.05$; ** $p < 0.01$; *** $p < 0.001$

Appendix 15 Adjusted rates ratios from logit-based generalised estimating equation regression models of conditions with risk factors associated with using healthcare due to violence for PHC utilisation and hospitalisation by all sociodemographic variables

Characteristics	PHC Utilisation		Hospitalisation	
	aOR	95% CI	aOR	95% CI
Months after using healthcare due to violence				
0, or did not use healthcare due to violence	1 (ref)	–	1 (ref)	–
1–3	1.54	(1.39–1.69)	1.63	(0.69–3.85)
4–6	1.33	(1.19–1.49)	1.08	(0.34–3.46)
7–12	1.16	(1.05–1.28)	0.23	(0.03–1.83)
Over 12	1.08 *	(1.00–1.17)	0.59	(0.23–1.55)
Sex				
Male	1 (ref)	–	1 (ref)	–
Female	1.27 **	(1.26–1.29)	1.25	(1.15–1.37)
Race/skin colour				
White	1 (ref)	–	1 (ref)	–
Black	1.11 ***	(1.09–1.13)	1.08	(0.94–1.23)
Pardo (Mixed Race)	0.99 **	(0.98–1.00)	0.98	(0.90–1.07)
Amarelo (Asian), Indigenous, Other or Unknown (Missing)	0.67 ***	(0.63–0.72)	0.82	(0.56–1.20)
Highest education level				
None/pre-school/literacy class	1 (ref)	–	1 (ref)	–
Elementary school	1.09 ***	(1.07–1.11)	0.48 ***	(0.42–0.55)
High school or higher education	0.93 ***	(0.91–0.95)	0.46 ***	(0.40–0.53)
None reported (missing)	0.83 ***	(0.81–0.85)	0.65 ***	(0.56–0.76)
Age group (years)				
0–18	1 (ref)	–	1 (ref)	–
19–34	4.58	(4.37–4.79)	16.57	(12.27–22.38)
35–54	15.72 ***	(15.04–16.43)	9.56	(7.07–12.92)
55–64	31.11 ***	(29.77–32.52)	10.92 ***	a
65+	38.26 ***	(36.61–39.99)	13.45 ***	(9.89–18.29)

Characteristics	PHC Utilisation		Hospitalisation	
	aOR	95% CI	aOR	95% CI
Household income				
Some salary reported to less than full minimum salary	1 (ref)	–	1 (ref)	–
More than one minimum salary but less than two	1.01	(0.99–1.03)	0.88	(0.74–1.04)
Two minimum salaries or more	1.19 ***	(1.14–1.24)	0.72 ***	(0.44–1.17)
No salary reported (unknown or missing)	0.91 ***	(0.89–0.93)	0.95 ***	(0.82–1.11)
Registered with Private Health Insurance				
No	1 (ref)	–	1 (ref)	–
Yes	0.71 ***	(0.66–0.77)	0.70	(0.41–1.19)
Bolsa Família-Claiming Family				
No	1 (ref)	–	1 (ref)	–
Yes	1.02	(0.99–1.04)	1.08	(0.91–1.28)
Total Observations (N)	17,075,018		17,075,018	
Total Groups (N)	529,219		529,219	

Separate fully adjusted random effects logistic regressions per outcome (PHC utilisation and hospitalisation); adjusted for sex, race/skin colour, education level, age group, household income, private health insurance, Bolsa Família-receiving family, year, and month. Robust standard errors. PHC Primary Healthcare, aOR Adjusted Odds Ratios, 95% CI 95% Confidence Intervals. * $p < 0.05$; ** $p < 0.01$; *** $p < 0.001$

Supplementary Information

The online version contains supplementary material available at <https://doi.org/10.1186/s12889-025-24567-2>.

Supplementary Material 1

Acknowledgements

This work was funded by the Foreign, Commonwealth & Development Office (FCDO), the Medical Research Council (MRC) and Wellcome Trust (grant number MR/V014420/1) and by the NIHR (NIHR150067 and NIHR133252) using UK international development funding from the UK Government to support global health research.

The views expressed in this publication are those of the authors and not necessarily those of the NIHR, the UK government or other funders. The funders had no role in the study design, in the collection, analysis, and interpretation of data, in the writing of the report, or in the decision to submit the paper for publication.

Authors' contributions

SM was responsible for conceptualisation, data curation, formal data analysis, investigation, methodology, project administration, and writing—original draft preparation. TH contributed to conceptualisation, data curation, investigation, methodology, project administration, supervision, and writing—review and editing. JG, VP, CM, RR, and JM contributed to conceptualisation, methodology, and writing—review and editing. SM and TH have directly accessed and verified the data reported in the manuscript and have final responsibility for the decision to submit the study for publication. All authors critically reviewed and edited the manuscript and approved its final version.

Data availability

The datasets generated and analysed during the current study are not publicly available due to confidentiality of the linked data. They are available from the corresponding author on reasonable request and following approval from CONEP and CEP.

Declarations

Ethics approval and consent to participate

Approval for the data used in this study was acquired from the *Comissão Nacional de Ética em Pesquisa* (Brazilian National Commission for Ethics in Research; CONEP) (protocol numbers 2.689.528 and 6.468.159), and from the *Comitês de Ética em Pesquisa* (Municipal Research Ethics Committee; CEP) (protocol number 6.210.448). The dataset used was fully anonymised following secure linkage. The authors had full access to all anonymised databases employed in this analysis.

Competing interests

The authors declare no competing interests.

Received: 21 March 2025 / Accepted: 22 August 2025

Published online: 16 October 2025

References

- McEvoy C, Hideg G, Survey SA, editors. *Global Violent Deaths 2017: Time to Decide*. Graduate Institute of International and Development Studies. 2017. Available from: <https://www.smallarmssurvey.org/sites/default/files/resource/s/SAS-Report-GVD2017.pdf> [Accessed: 27 November 2022].
- World Health Organization, Habitat UN. *Global report on urban health: equitable healthier cities for sustainable development*. 2016. Available from: <https://apps.who.int/iris/handle/10665/204715> [Accessed: 8 January 2023].
- Krug E, Dahlberg LL, Mercy JA. *The world report on violence and health*. World Health Organization. 2002. Available from: https://apps.who.int/iris/bitstream/handle/10665/42495/9241545615_eng.pdf [Accessed: 6 Jan 2023].
- Ceccato V, Kahn T, Vazquez LC. Lethal violence in Brazil: a systematic review of Portuguese-language literature from 2000 to 2020. *Crim Justice Rev*. 2021;46(4):404–34.
- Reichenheim ME, de Souza ER, Moraes CL, de Mello Jorge MHP, da Silva CMFP, de Souza Minayo MC. Violence and injuries in Brazil: the effect, progress made, and challenges ahead. *Lancet*. 2011;377(9781):1962–75.

6. Rivara F, Adhia A, Lyons V, Massey A, Mills B, Morgan E, et al. The effects of violence on health. *Health Aff.* 2019;38(10):1622–9.
7. Suglia SF, Sapra KJ, Koenen KC. Violence and cardiovascular health: a systematic review. *Am J Prev Med.* 2015;48(2):205–12.
8. Apter AJ, Garcia LA, Boyd RC, Wang X, Bogen DK, Ten Have T. Exposure to community violence is associated with asthma hospitalizations and emergency department visits. *J Allergy Clin Immunol.* 2010;126(3):552–7.
9. Keasley J, Blickwedel J, Quenby S. Adverse effects of exposure to armed conflict on pregnancy: a systematic review. *BMJ Glob Health.* 2017;2(4):e000377.
10. Chanlongbutra A, Singh GK, Mueller CD. Adverse childhood experiences, health-related quality of life, and chronic disease risks in rural areas of the United States. *J Environ Public Health.* 2018;2018: 7151297.
11. Soares ALG, Howe LD, Matijasevich A, Wehrmeister FC, Menezes AMB, Gonçalves H. Adverse childhood experiences: prevalence and related factors in adolescents of a Brazilian birth cohort. *Child Abuse Negl.* 2016;51:21–30.
12. Olofsson N, Lindqvist K, Shaw BA, Danielsson I. Long-term health consequences of violence exposure in adolescence: a 26-year prospective study. *BMC Public Health.* 2012;12(1):411.
13. John HP, Rosemary C, Patricia PW, Edward AR, Dawn L, Karen EW, et al. Violent victimization of women and men: physical and psychiatric symptoms. *J Am Board Family Pract.* 2003;16(1):32.
14. Campbell JC. Health consequences of intimate partner violence. *Lancet.* 2002;359(9314):1331–6.
15. Mmari K, Marshall B, Hsu T, Shon JW, Eguavoen A. A mixed methods study to examine the influence of the neighborhood social context on adolescent health service utilization. *BMC Health Serv Res.* 2016;16(1):433.
16. Vargas L, Vélez-Grau C, Camacho D, Richmond TS, Meisel ZF. The permeating effects of violence on health services and health in Mexico. *J Interpers Violence.* 2022;37(13–14):Np10883–911.
17. Green G, Swartz A, Tembo D, Cooper D, George A, Matzopoulos R, et al. A scoping review of how exposure to urban violence impacts youth access to sexual, reproductive and trauma health care in LMICs. *Glob Public Health.* 2022;18:1–18.
18. Barbar AE, Delorme C, Fokeladeh HS, Perazzi S. *Violence against healthcare survey report.* 2021. Available from: <https://www.icn.ch/system/files/2022-07/Violence%20against%20healthcare%20survey%20report.pdf> [Accessed: 16 Jun 2023].
19. Ahern J, Matthay EC, Goin DE, Farkas K, Rudolph KE. Acute changes in community violence and increases in hospital visits and deaths from stress-responsive diseases. *Epidemiol (Cambridge Mass).* 2018;29(5):684.
20. Gates S, Hegre H, Nygård H, Strand H. Consequences of armed conflict. *World Dev.* 2012;40:1713–22.
21. Scott-Storey KA. Abuse as a gendered risk factor for cardiovascular disease: a conceptual model. *J Cardiovasc Nurs.* 2013;28(6):E1–8.
22. Guha-Sapir D, van Panhuis WG. *Armed conflict and public health: a report on knowledge and knowledge gaps.* CRED Brussels. 2002. Available from: https://dial.uclouvain.be/download/downloader.php?pid=boreal:179723&datastream=PDF_01
23. Lo Fo Wong S, Wester F, Mol S, Römkens R, Lagro-Janssen T. Utilisation of health care by women who have suffered abuse: a descriptive study on medical records in family practice. *Br J Gen Pract.* 2007;57(538):396–400.
24. Rivara FP, Anderson ML, Fishman P, Bonomi AE, Reid RJ, Carrell D, et al. Healthcare utilization and costs for women with a history of intimate partner violence. *Am J Prev Med.* 2007;32(2):89–96.
25. Ulrich YC, Cain KC, Sugg NK, Rivara FP, Rubanowicz DM, Thompson RS. Medical care utilization patterns in women with diagnosed domestic violence. *Am J Prev Med.* 2003;24(1):9–15.
26. Daoud N, Kraun L, Sergienko R, Batat N, Shoham-Vardi I, Davidovitch N, et al. Patterns of healthcare services utilization associated with intimate partner violence (IPV): effects of IPV screening and receiving information on support services in a cohort of perinatal women. *PLoS One.* 2020;15(1):e0228088.
27. Leight J, Wilson N. Intimate partner violence and maternal health services utilization: evidence from 36 national household surveys. *BMC Public Health.* 2021;21(1):405.
28. Anguzu R, Walker RJ, Babikako HM, Beyer KMM, Dickson-Gomez J, Zhou Y, et al. Intimate partner violence and antenatal care utilization predictors in Uganda: an analysis applying andersen's behavioral model of healthcare utilization. *BMC Public Health.* 2023;23(1):2276.
29. Torche F, Villarreal A. Prenatal exposure to violence and birth weight in Mexico. *Am Sociol Rev.* 2014;79(5):966–92.
30. Vargas LX, Richmond TS, Allen HL, Meisel ZF. A longitudinal analysis of violence and healthcare service utilization in Mexico. *Int J Equity Health.* 2021;20(1):75.
31. Hanigan M, Heisler M, Choi H. Relationship between county-level crime and diabetes: mediating effect of physical inactivity. *Prev Med Rep.* 2020;20:101220.
32. Mayne SL, Pool LR, Grobman WA, Kershaw KN. Associations of neighbourhood crime with adverse pregnancy outcomes among women in Chicago: analysis of electronic health records from 2009 to 2013. *J Epidemiol Community Health.* 2018;72(3):230.
33. FÓRUM BRASILEIRO DE SEGURANÇA PÚBLICA. 14 Anuário Brasileiro de Segurança Pública. 2020. Available from: <https://forumseguranca.org.br/wp-content/uploads/2020/10/anuario-14-2020-v1-interativo.pdf> [Accessed: February 18 2024].
34. Anuário Brasileiro de Segurança Pública. Brazilian yearbook of public security. 12ª Edição Fórum Brasileiro de Segurança Pública, São Paulo. 2018.
35. FÓRUM BRASILEIRO DE SEGURANÇA PÚBLICA. 17º Anuário Brasileiro de Segurança Pública. 2023. Available from: <https://forumseguranca.org.br/wp-content/uploads/2023/07/anuario-2023.pdf> [Accessed: February 18 2024].
36. Bortz M, Kano M, Ramroth H, Barcellos C, Weaver SR, Rothenberg R, et al. Disaggregating health inequalities within Rio de Janeiro, Brazil, 2002–2010, by applying an urban health inequality index. *Cad Saude Publica.* 2015;31(suppl 1):107–19.
37. Fahlberg A, Vicino TJ, Fernandes R, Potiguara V. Confronting chronic shocks: social resilience in Rio de Janeiro's poor neighborhoods. *Cities.* 2020;99:102623.
38. Ford E, Curlewis K, Squires E, Griffiths LJ, Stewart R, Jones KH. The potential of research drawing on clinical free text to bring benefits to patients in the United Kingdom: a systematic review of the literature. *Front Digit Health.* 2021;3:606599.
39. Forum Brasileiro de Segurança Pública. *Feminicídios em 2023.* 2024. Available from: <https://apidspace.universilab.com.br/server/api/core/bitstreams/eca3a94f-2981-488c-af29-572a73c8a9bf/content> [Accessed: 23 Apr 2024].
40. Forum Brasileiro de Segurança Pública. *Violência contra meninas e mulheres no 1º semestre de 2023.* 2023. Available from: <https://apidspace.universilab.com.br/server/api/core/bitstreams/1dad654e-1682-4ddb-93b2-68f7583d60f2/content> [Accessed: 23 Apr 2024].
41. Prefeitura da Cidade do Rio de Janeiro. Available from: <https://prefeitura.rio/>
42. Coeli CM, Saraceni V, Medeiros PM, da Silva Santos HP, Guillen LCT, Alves LGSB, et al. Record linkage under suboptimal conditions for data-intensive evaluation of primary care in Rio de Janeiro, Brazil. *BMC Med Inform Decis Mak.* 2021;21(1):190.
43. Blackwell M, Iacus S, King G, Porro G. Cem: coarsened exact matching in Stata. *The Stata Journal: Promoting communications on statistics and Stata.* 2009;9(4):524–46.
44. World Health Organization. International Statistical Classification of Diseases and Related Health Problems. 1994. Available from: <https://icd.who.int/brows/e10/2019/en>
45. Romeo ACDCB, Cardoso PLP, Correia-Jr GB, Joaquim-de-Carvalho MEA, Santos FM, Serafim DF, et al. Guerra civil não-declarada? Um recorte do status Da violência urbana Em Uma capital no Brasil. *Rev Colégio Bras Cir.* 2020. <http://doi.org/10.1590/0100-6991e-20202506>.
46. Clery MJ, Hudson PJ, Moore JC, Kollar LMM, Wu DT. Exploring injury intentionality and mechanism via ICD-10-CM injury codes and self-reported injury in a large, urban emergency department. *Inj Prev.* 2021;27(Suppl 1):i62–5.
47. Reynolds SA. Do health sector measures of violence against women at different levels of severity correlate? Evidence from Brazil. *BMC Womens Health.* 2022;22(1):226.
48. Coube M, Nikoloski Z, Mrejen M, Mossialos E. Persistent inequalities in health care services utilisation in Brazil (1998–2019). *Int J Equity Health.* 2023;22(1):25.
49. Barraza-Lloréns M, Panopoulou G, Díaz BY. Income-related inequalities and inequities in health and health care utilization in Mexico, 2000–2006. *Rev Panam Salud Publica.* 2013;33:122–30.
50. Constante HM, Marinho GL, Bastos JL. The door is open, but not everyone may enter: racial inequities in healthcare access across three Brazilian surveys. *Cienc Saude Coletiva.* 2021;26(9):3981–90.
51. Fix RL, Vest N, Thompson KR. Evidencing the need to screen for social determinants of health among boys entering a juvenile prison: a latent profile analysis. *Youth Violence Juv Justice.* 2022;20(3):187–205.

52. Sheats KJ, Irving SM, Mercy JA, Simon TR, Crosby AE, Ford DC, et al. Violence-related disparities experienced by black youth and young adults: opportunities for prevention. *Am J Prev Med*. 2018;55(4):462–9.
53. Curto BM, Paula CS, Do Nascimento R, Murray J, Bordin IA. Environmental factors associated with adolescent antisocial behavior in a poor urban community in Brazil. *Soc Psychiatry Psychiatr Epidemiol*. 2011;46(12):1221–31.
54. Armstead TL, Wilkins N, Nation M. Structural and social determinants of inequities in violence risk: a review of indicators. *J Community Psychol*. 2021;49(4):878–906.
55. Turnbull J, Pope C, Prichard J, McKenna G, Rogers A. A conceptual model of urgent care sense-making and help-seeking: a qualitative interview study of urgent care users in England. *BMC Health Serv Res*. 2019;19(1):481.
56. Borg BA, Rabinak CA, Marusak HA. Violence exposure and mental health consequences among urban youth. *Curr Psychol*. 2023;42(10):8176–85.
57. Miliauskas CR, Faus DP, da Cruz VL, do Nascimento Vallaperde JGR, Junger W, Lopes CS. Community violence and internalizing mental health symptoms in adolescents: a systematic review. *BMC Psychiatry*. 2022;22(1):253.
58. Projects People'sP. Redes da Maré. *Bulletin Building the Barricades*. 2021. Available from: <https://peoplespalaceprojects.org.uk/wp-content/uploads/2021/1/boletim-const-pontes-ingles-online.pdf> [Accessed: 23 May 2023].
59. Cuartas J, Roy AL. The latent threat of community violence: indirect exposure to local homicides and adolescents' mental health in Colombia. *Am J Community Psychol*. 2019;64(1–2):219–31.
60. Velez-Gomez P, Restrepo-Ochoa DA, Berbesi-Fernandez D, Trejos-Castillo E. Depression and neighborhood violence among children and early adolescents in Medellín, Colombia. *Span J Psychol*. 2013;16:E64.
61. Dilgul M, McNamee P, Orfanos S, Carr CE, Priebe S. Why do psychiatric patients attend or not attend treatment groups in the community: a qualitative study. *PLoS One*. 2018;13(12):e0208448.
62. Harvey J, Lawson V. The importance of health belief models in determining self-care behaviour in diabetes. *Diabet Med*. 2009;26(1):5–13.
63. Buitrago G, Moreno-Serra R. Conflict violence reduction and pregnancy outcomes: a regression discontinuity design in Colombia. *PLoS Med*. 2021;18(7):e1003684.
64. Foureaux Koppensteiner M, Manacorda M. Violence and birth outcomes: evidence from homicides in Brazil. *J Dev Econ*. 2016;119:16–33.
65. Dos Santos AC, Brentani A, Fink G. Associations between neighborhood violence during pregnancy and birth outcomes: evidence from São Paulo's western region birth cohort. *BMC Public Health*. 2021. <https://doi.org/10.1186/s12889-021-10900-y>.
66. Goin DE, Gomez M, Farkas A, Zimmerman K, Matthay SC, Ahern EC. Exposure to community homicide during pregnancy and adverse birth outcomes: a within-community matched design. *Epidemiology*. 2019;30(5):713–22.
67. Mazza M, Marano G, Del Gonzalez A, Chieffo D, Albano G, Biondi-Zoccai G, et al. Interpersonal violence: serious sequelae for heart disease in women. *World J Cardiol*. 2021;13(9):438–45.
68. Perera E, Chou S, Cousins N, Mota N, Reynolds K. Women's experiences of trauma, the psychosocial impact and health service needs during the perinatal period. *BMC Pregnancy Childbirth*. 2023;23(1):197.
69. Acosta Pedemonte NB, Rocchetti NS, Villalba J, Lerman Tenenbaum D, Settecase CJ, Bagilet DH, et al. *Bacillus cereus* bacteremia in a patient with an abdominal stab wound. *Rev Argent Microbiol*. 2020;52(2):115–7.
70. Bodur H, Onguru P, Akinci E, Kanyılmaz D, Yetkin A, Ozdemir B, et al. Two cases of acute hepatitis after a serial knife stabbing. *Diagn Microbiol Infect Dis*. 2012;74(3):311–2.
71. Baum GR, Baum JT, Hayward D, MacKay BJ. Gunshot wounds: ballistics, pathology, and treatment recommendations, with a focus on retained bullets. *Orthop Res Reviews*. 2022;14:293–317.
72. Kuchyn I, Horoshko V. Chronic pain in patients with gunshot wounds. *BMC Anesthesiol*. 2023;23(1):47.
73. Patil R, Jaiswal G, Gupta TK. Gunshot wound causing complete spinal cord injury without mechanical violation of spinal axis: case report with review of literature. *J Craniovertebr Junction Spine*. 2015;6(4):149–57.
74. de Ribera OS, Trajtenberg N, Shenderovich Y, Murray J. Correlates of youth violence in low- and middle-income countries: a meta-analysis. *Aggress Violent Behav*. 2019;49:101306.
75. Ministério da Economia Instituto Brasileiro de Geografia e Estatística. IBGE, editor. Pesquisa nacional de saúde: 2019 : acidentes, violências, doenças transmissíveis, atividade sexual, características do trabalho e apoio social. 2019. Available from: <https://biblioteca.ibge.gov.br/visualizacao/livros/liv101800.pdf>
76. Murray J, Cerqueira DR, Kahn T. Crime and violence in Brazil: systematic review of time trends, prevalence rates and risk factors. *Aggress Violent Behav*. 2013;18(5):471–83.
77. Pan American Health Organization. *Homicide in Brazil: A Gender and Diversity Analysis*. n.d. Available from: https://www3.paho.org/hq/dmdocuments/2010/Homicide_Brazil_gender_case_study.pdf
78. Redondo-Sendino Á, Guallar-Castillón P, Banegas JR, Rodríguez-Artalejo F. Gender differences in the utilization of health-care services among the older adult population of Spain. *BMC Public Health*. 2006;6(1):155.
79. Mutran E, Ferraro KF. Medical need and use of services among older men and women. *Journals Gerontol*. 1988;43(5):S162–171.
80. Ladwig KH, Marten-Mittag B, Formanek B, Dammann G. Gender differences of symptom reporting and medical health care utilization in the German population. *Eur J Epidemiol*. 2000;16(6):511–8.
81. Krasnik A, Hansen E, Keiding N, Savitz L. Determinants of general practice utilization in Denmark. *Dan Med Bull*. 1997;44(5):542–6.
82. Bertakis KD, Azari R, Helms LJ, Callahan EJ, Robbins JA. Gender differences in the utilization of health care services. *J Fam Pract*. 2000;49(2):147–52.
83. Cobo B, Cruz C, Dick PC. Desigualdades de Gênero e raciais no Acesso e Uso Dos serviços de Atenção primária à Saúde no Brasil. *Cienc Saude Coletiva*. 2021;26(9):4021–32.
84. Travassos C, Viacava F, Pinheiro R, Brito A. Utilização Dos serviços de Saúde no Brasil: gênero, características familiares e Condição social. *Rev Panam Salud Publica*. 2002;11(5–6):365–73.
85. Minayo MCS, Mariz RSA. Profile of Rio de Janeiro's violent lethality perpetrators, Brazil (2015). *Cien Saude Colet*. 2021;26(suppl 3):5023–32.
86. Malta DC, Soares Filho AM, Pinto IV, de Souza Minayo MC, Lima CM, Machado ÍE, et al. Association between firearms and mortality in Brazil, 1990 to 2017: a global burden of disease Brazil study. *Popul Health Metrics*. 2020;18(1):19.
87. Instituto Brasileiro de Geografia e Estatística. Pesquisa Nacional de Saúde 2013: Acesso e Utilização dos Serviços de Saúde, Acidentes e Violências: Brasil, Grandes Regiões e Unidades da Federação. 2013. Available from: <http://biblioteca.ibge.gov.br/visualizacao/livros/liv94074.pdf> [Accessed: 23 Jun 2022].
88. Sundquist K, Theobald H, Yang M, Li X, Johansson S-E, Sundquist J. Neighborhood violent crime and unemployment increase the risk of coronary heart disease: a multilevel study in an urban setting. *Soc Sci Med*. 2006;62(8):2061–71.
89. Christian H, Giles-Corti B, Knuiman M, Timperio A, Foster S. The influence of the built environment, social environment and health behaviors on body mass index. Results from RESIDE. *Prev Med*. 2011;53(1):57–60.
90. Chaparro MP, Pina MF, de Oliveira Cardoso L, Santos SM, Barreto SM, Giatti Gonçalves L, et al. The association between the neighbourhood social environment and obesity in Brazil: a cross-sectional analysis of the ELSA-Brasil study. *BMJ Open*. 2019;9(9):e026800.
91. Fowler PJ, Tompsett CJ, Braciszewski JM, Jacques-Tiura AJ, Baltes BB. Community violence: a meta-analysis on the effect of exposure and mental health outcomes of children and adolescents. *Dev Psychopathol*. 2009;21(1):227–59.
92. Lund C, Brooke-Sumner C, Baingana F, Baron EC, Breuer E, Chandra P, et al. Social determinants of mental disorders and the sustainable development goals: a systematic review of reviews. *Lancet Psychiatry*. 2018;5(4):357–69.
93. United Nations Population Fund. State of world population 2013: motherhood in Childhood-Facing the challenge of adolescent pregnancy. UN; 2013.
94. Braverman-Bronstein A, Vidana-Pérez D, Ortigoza AF, Baldovino-Chiquillo L, Diez-Canseco F, Maslowsky J, et al. Adolescent birth rates and the urban social environment in 363 Latin American cities. *BMJ Glob Health*. 2022. <https://doi.org/10.1136/bmjgh-2022-009737>.
95. World Health Organization. Violence against women and HIV/AIDS: critical intersections—intimate partner violence and HIV/AIDS. *Inform Bull Ser*. 2004;1:1–9.
96. Wheeler J, Anfinson K, Valvert D, Lungu S. Is violence associated with increased risk behavior among MSM? Evidence from a population-based survey conducted across nine cities in central America. *Glob Health Action*. 2014;7:24814.
97. Nakamura IB, Silva MT, Garcia LP, Galvao TF. Prevalence of physical violence against Brazilian women: systematic review and meta-analysis. *Trauma Violence Abuse*. 2023;24(2):329–39.
98. Wonca International Classification Committee (WICC). International Classification of Primary Care – 2nd Edition. 2003. Available from: <https://www.who.int/standards/classifications/other-classifications/international-classification-of-primary-care>

99. Sun S, Savitz DA, Wellenius GA. Changes in adverse pregnancy outcomes associated with the COVID-19 pandemic in the United States. *JAMA Netw Open*. 2021;4(10):e2129560.
100. de Fatima Marinho de Souza M, Gawryszewski VP, Orduñez P, Sanhueza A, Espinal MA. Cardiovascular disease mortality in the Americas: current trends and disparities. *Heart*. 2012;98(16):1207–12.
101. Dugan J, Shubrook J. International classification of diseases, 10th revision, coding for diabetes. *Clin Diabetes*. 2017;35(4):232–8.
102. Suissa K, Schneeweiss S, Lin KJ, Brill G, Kim SC, Patorno E. Validation of obesity-related diagnosis codes in claims data. *Diabetes Obes Metab*. 2021;23(12):2623–31.
103. Ho YA, Rahurkar S, Tao G, Patel CG, Arno JN, Wang J, et al. Validation of international classification of diseases, tenth revision, clinical modification codes for identifying cases of chlamydia and gonorrhea. *Sex Transm Dis*. 2021;48(5):335–40.
104. Glasgow AE, Wilder J, Caskey R, Munoz G, Van Voorhees B, Kim S. Mental health diagnoses among children and adolescents with chronic medical conditions in a large urban cohort. *J Behav Health Serv Res*. 2020;9(4):1–8.

Publisher's Note

Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.