# **BMJ Open** Identification of acute respiratory distress syndrome subphenotypes de novo using routine clinical data: a retrospective analysis of ARDS clinical trials

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## ABSTRACT

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Correspondence to Dr Abhijit Duggal; duggala2@ccf.org **Objectives** The acute respiratory distress syndrome (ARDS) is a heterogeneous condition, and identification of subphenotypes may help in better risk stratification. Our study objective is to identify ARDS subphenotypes using new simpler methodology and readily available clinical variables.

**Setting** This is a retrospective Cohort Study of ARDS trials. Data from the US ARDSNet trials and from the international ART trial.

**Participants** 3763 patients from ARDSNet data sets and 1010 patients from the ART data set.

Primary and secondary outcome measures The primary outcome was 60-day or 28-day mortality, depending on what was reported in the original trial. K-means cluster analysis was performed to identify subgroups. Sets of candidate variables were tested to assess their ability to produce different probabilities for mortality in each cluster. Clusters were compared with biomarker data, allowing identification of subphenotypes. Results Data from 4773 patients were analysed. Two subphenotypes (A and B) resulted in optimal separation in the final model, which included nine routinely collected clinical variables, namely heart rate, mean arterial pressure, respiratory rate, bilirubin, bicarbonate, creatinine, PaO<sub>a</sub>, arterial pH and FiO<sub>a</sub>. Participants in subphenotype B showed increased levels of proinflammatory markers, had consistently higher mortality, lower number of ventilatorfree days at day 28 and longer duration of ventilation compared with patients in the subphenotype A. Conclusions Routinely available clinical data can successfully identify two distinct subphenotypes in adult ARDS patients. This work may facilitate implementation of precision therapy in ARDS clinical trials.

#### INTRODUCTION

The Berlin definition of acute respiratory distress syndrome (ARDS) encompasses acute hypoxemic respiratory failure due to a wide variety of etiologies.<sup>1</sup> Due to this inclusion

## Strengths and limitations of this study

- Largest cohort of patients used to identify subphenotypes of acute respiratory distress syndrome (ARDS) patients.
- Subphenotypes were validated in the population of a large international ARDS randomised controlled trial.
- Subphenotypes were identified by using only routinely collected clinical data.
- Our use of data exclusively from randomised controlled trials does not prove generalisability to unselected ARDS populations.
- The clinical utility of the subphenotypes has to be validated in a prospective study.

of heterogeneous conditions within the syndrome, there are significant clinical and biological differences that make ARDS challenging to treat.<sup>2 3</sup> These differences among ARDS patients are associated with variation in risk of disease development and progression,<sup>3 4</sup> potentially generating differential responses to treatments and interventions.<sup>5–10</sup> Despite evidence, clinical risk stratification of ARDS patients still solely depends on  $PaO_2/FiO_2$  ratios,<sup>11 12</sup> possibly misleading the interpretation of results in clinical trials and clinicians when evaluating treatment options for patients.<sup>13</sup>

Therefore, identifying groups of patients who have similar clinical, physiologic or biomarker traits becomes relevant<sup>6 14</sup> as it can help with stratification of patients producing better targeted therapies and interventions.<sup>15</sup> These different groups can be defined as ARDS subphenotypes.<sup>4 14</sup> Two ARDS subphenotypes have been consistently identified in previous studies.<sup>6-10 16-18</sup> However, these models are complex, and significant barriers exist in their implementation and use in clinical practice. Existing models use up to 40 predictor variables, including biomarkers and other variables that are not readily available at the bedside.<sup>6-10 16-18</sup> These limitations explain the current status quo of ARDS care, where clinicians must depend on the limited prognostic value of PaO<sub>2</sub>/FiO<sub>2</sub> ratios instead of biologically distinct subphenotypes.

We hypothesised that the use of a simpler methodology and a small number of easily available clinical variables could identify new ARDS subphenotypes and thus provide the means to allow future implementation of bedside stratification.

#### **METHODS**

#### **Data source and participants**

We performed a retrospective study using a deidentified data set pooling data from six randomised clinical trials in patients with ARDS, namely ARMA, ALVEOLI, FACTT, EDEN, SAILS and ART.<sup>19-24</sup> Patients in ARMA, ALVEOLI, FACTT, EDENand SAILS trials were eligible if they met the American-European consensus for ARDS, including patients with a PaO<sub>9</sub>/FiO<sub>9</sub> ratio <300 up to 48 hours before enrolment. From 1996 to 2013, these trials enrolled 902, 549, 1000, 1000 and 745 patients, respectively, and tested a variety of interventions.<sup>19–23</sup> Between 2011 and 2017, the international ART study enrolled 1010 adult patients diagnosed with moderate to severe ARDS according to the Berlin definition  $(PaO_9/FiO_9 ratio < 200)$  for less than 72 hours of duration and assessed two different ventilatory strategies.<sup>24</sup> To avoid biases due to high mortality in the high tidal volume group of the ARMA study,<sup>19</sup> which has not been standard of care since the beginning of 2000, only 473 patients receiving low tidal volume in that study were included.

## Predictors

Six clinical trials were assessed to identify a set of clinical variables recorded closest to time of randomisation which were most commonly available across all data sets. The list of potential candidates was then further refined to include only those that are frequently observed in the routine care of ARDS patients at the time of its diagnosis according to judgement provided by intensive care unit physicians who participated in this study. To develop a clustering algorithm for potential rapid translation into clinical use, elements which would not be commonly found in the electronic health records (EHR) at the time of ARDS diagnosis, such as biomarker levels, ARDS risk factors, organ support apart from mechanical ventilation settings and severity scores, were excluded from model development. The treatment assignment in the original trials, and clinical outcomes were not considered in the model development.

After all assessment, 16 variables that are routinely collected as part of the usual care and which were uniformly present in all the trials were considered, including age, gender, arterial pH, PaO<sub>2</sub>, PaCO<sub>2</sub>, bicarbonate, creatinine, bilirubin, platelets, heart rate, respiratory pressure (PEEP), plateau pressure, FiO<sub>2</sub> and tidal volume adjusted for predicted body weight (mL/kg PBW). The PBW was calculated as equal to 50+0.91 (centimetres of height, 152.4) in males and 45.5+0.91 (centimetres of height, 152.4) in females.<sup>18</sup> These variables were grouped into five domains named demographics, arterial blood gases, laboratory values, vital signs and ventilatory variables. Plateau pressure was excluded due to a high rate of missingness across the trials included in the training set. Amount of missing data in the training data sets is reported in online supplemental eTable 1.

## **Outcomes**

The primary outcome was 60-day mortality for all ARDSnet trials, and 28-day mortality for ART trial. Secondary outcomes included 90-day mortality, number of ventilator free days at day  $28^{25}$  and the duration of mechanical ventilation in survivors within the first 28 days postenrolment.

## **Data preparation**

Data preprocessing was performed before modelling, and the pooled data set was assessed for completeness and consistency. Patients with values out of the plausible physiological range for a specific variable were excluded from the final analysis (described in online supplemental eTable 2). The training data set was constructed using data from the two largest ARDSnet trials, EDEN and FACTT. The validation data set was sourced from the four remaining trials: ALVEOLI, ARMA, SAILS and ART. Means and SD for z-scoring variables were calculated from the training data set and subsequently applied to the validation data.

#### Statistical analysis

Baseline and outcome data were presented according to the assigned cluster. Continuous variables were presented as medians with their IQRs and categorical variables as total number and percentage. Proportions were compared using Fisher exact tests and continuous variables were compared using the Wilcoxon rank-sum test. Study outcomes were further compared using the median and mean absolute differences for continuous and categorical values, respectively.

### Model development and validation

For the model development, the K-means clustering algorithm was used. K-means is one of the simplest and most used classes of clustering algorithms. In critical care research, unsupervised machine learning techniques have already been used in several studies, attempting to find homogeneous subgroups within a broad heterogeneous population.<sup>26</sup> This specific algorithm identifies a K number of clusters in a data set by finding K centroids within the n-dimensional space of clinical features.<sup>26</sup>

For feature selection, different sets of candidate variables were tested to assess their ability to produce

significantly different mortality probabilities in each cluster using the minimum amount of readily available clinical data. For each set of candidate variables, the optimal number of clusters was determined by comparing models with between 2 and 5 clusters, using the Elbow method<sup>27</sup> and the Calinski-Harabasz index.<sup>28</sup> Information about the methods for selecting the number of clusters are provided in the online supplemental material.

The following steps were performed for the final model selection: (1) all predictors were assessed for correlation (online supplemental eTable 3); and (2) 10 different combinations of the proposed variables were investigated. These combinations were developed based on the perceived clinical importance of each variable and its combinations. All 10 models were tested for the optimal number of clusters based on both the Elbow method and the Calinski-Harabasz index, as described above. The models were then compared, aiming for the minimum set of variables with high 60-day mortality separation. The description of each model is shown in online supplemental eTable 4.

Biological and clinical characteristics of the clusters were evaluated using clinical, laboratory and (when available) biomarker data to establish subphenotypes.<sup>4</sup> All iterations in model development were done on the training set and the generalisability of the final model was assessed using the validation data set. K-means clustering analysis is structured to ignore cases with missing data. No assumption was made for missingness, and we therefore conducted a complete case analysis. Model development and evaluation was performed using Python V.3.8 and scikit-learn 0.23.1.

#### Patient and public involvement

There was no patient involvement in this study.

#### **Data availability**

Data from the ARDSnet studies (EDEN, FACTT, ARMA, ALVEOLI and SAILS) are publicly available from the NHLBI ARDS Network and data from the ART trial can be requested from study authors.

## RESULTS

## **Participants**

Data from 4777 clinical trial patients were considered for inclusion. In total, four patients were excluded for having clinical measurements outside plausible range. The remaining 1998 patients from EDEN and FACTT trials were included in the training set, while the 2775 patients from ARMA, ALVEOLI, SAILS and ART were included in the validation cohort.

Baseline characteristics of the patients in the training and validation sets are presented in table 1. Pneumonia was the prevailing aetiology followed by sepsis and aspiration in all trials. Between 29.3% and 72.7% of the patients were receiving vasopressors at the time of randomisation. At randomisation,  $PaO_9/FiO_9$  ratio ranged from 112 (75–158) to 134 (96–185) mm Hg, and PEEP from 8 (5–10) to 12 (10–14) cmH<sub>2</sub>O across trials. Mortality at 60 days for the ARDSnet trials ranged from 22.7% to 30.1%, while in the ART trial mortality at 28 days was 58.8%.

## Predictor variables and model selection

The correlation between the 15 variables selected for clustering is shown in online supplemental eTable 3. The strongest correlation was between PEEP and  $FiO_2$  (r=0.49). The comparison of the 10 models regarding the optimal number of clusters based on both the Elbow method and the Calinski-Harabasz index is shown in online supplemental eFigure 1. In all models and methods, two clusters were a better fit than a higher number of clusters.

Across the 10 models, absolute mortality difference between cluster 1 and cluster 2 ranged from 3.9% to 13.1% for the FACTT study and between 0.1% and 8.1% for EDEN (online supplemental eTable 4). The models with the highest 60-day absolute mortality separation between the clusters for each of the two trials in the training set were then further evaluated. Models 6, 5 and 8 were consistently among the models with highest separation (online supplemental eTable 4). Model 8 was selected for further investigation, as it had the fewest variables (online supplemental eTable 5).

### **Clinical characteristics of each cluster**

Based on model 8, only nine clinical and laboratory variables were needed to identify the two distinct clusters in ARDS patients, namely heart rate, mean arterial pressure, respiratory rate, bilirubin, bicarbonate, creatinine, PaO<sub>a</sub>, arterial pH and FiO<sub>9</sub>. For each variable in the model, opposing measurements could be observed for each cluster (figure 1 and online supplemental eFigure 2). For the ARDSnet trials, the incidence of cluster 1 patients varied from 57.8% (EDEN) to 73.6% (ARMA), and 41.5% of ART patients were part of cluster 1. Across all trials, patients in cluster 2 had higher severity of illness, rate of vasopressor, heart rate, respiratory rate, creatinine and bilirubin, as well as lower platelets, pH, blood urea nitrogen and bicarbonate compared with patients in cluster 1 (table 2, online supplemental eTables 6 and 7). In addition, 28-day, 60-day and 90-day mortality rate was higher in patients in cluster 2 in all trials (table 3). Likewise, for each trial, the number of ventilator-free days at day 28 was lower in patients in cluster 2 compared with cluster 1, and duration of ventilation in survivors was longer in cluster 1.

#### Identification of subphenotypes

After comparing the clinical characteristics of the clusters, each cluster was assigned to represent a distinct subphenotype of ARDS, with patients in cluster 1 assigned to subphenotype A, and patients in cluster 2 assigned to subphenotype B. Using blood biomarker information available for a subset of patients from both ARMA and ALVEOLI, subphenotype B showed increased levels of proinflammatory markers when compared with

Table 1 Baseline characteristics	s and clinical outcomes i	n the included trials				
	Training set (n=1998	()	Validation set (n=27	75)		
	EDEN (n=1000)	FACTT (n=998)	ALVEOLI (n=549)	ARMA (n=472)	ART (n=1010)	SAILS (n=744)
Age, year	52.0 (42.0–63.0)	49.0 (38.0–60.8)	50.0 (39.0–65.0)	50.0 (37.8–65.0)	52.0 (36.0-64.0)	55.0 (42.0-66.0)
Male gender, no. (%)	510 (51.0)	533 (53.4)	302 (55.0)	285 (60.4)	631 (62.5)	365 (49.0)
Aetiology, no. (%)						
Pneumonia	650 (65.0)	471 (47.2)	221 (40.3)	145 (30.7)	555 (55.0)	526 (70.7)
Sepsis	147 (14.7)	231 (23.1)	120 (21.9)	125 (26.5)	196 (19.4)	147 (19.8)
Aspiration	96 (9.6)	149 (14.9)	84 (15.3)	72 (15.3)	58 (5.7)	49 (6.6)
Trauma	36 (3.6)	74 (7.4)	45 (8.2)	59 (12.5)	31 (3.1)	6 (0.8)
Other	71 (7.1)	73 (7.3)	79 (14.4)	71 (15.0)	170 (16.8)	16 (2.2)
Severity of illness*	73.0 (59.0–89.0)	78.0 (62.0–94.0)	78.0 (64.0–93.0)	83.0 (70.0–97.0)	63.0 (50.2–75.0)	76.0 (61.0–92.0)
Vasopressors, no. (%)	489 (48.9)	397 (40.5)	156 (29.3)	147 (31.3)	734 (72.7)	395 (54.2)
Laboratory tests						
White cell count, 10 <sup>9</sup> /L	12.0 (7.8–16.7)	11.8 (7.2–17.1)	11.6 (7.7–15.7)	11.5 (7.5–16.2)	I	13.9 (8.7–20.0)
Platelets, 10 <sup>9</sup> /L	169 (108–241)	183 (106–258)	157 (83–247)	135 (80–211)	175 (106–263)	167 (96–247)
Creatinine, mg/dL	1.2 (0.8–2.0)	1.0 (0.7–1.5)	1.0 (0.7–1.7)	1.1 (0.8–1.7)	1.3 (0.8–2.2)	1.0 (0.7–1.7)
Bilirubin, mg/dL	0.8 (0.5–1.4)	0.8 (0.5–1.6)	0.8 (0.5–1.5)	1.0 (0.6–2.1)	0.8 (0.4–1.5)	0.8 (0.5–1.4)
Arterial blood gas						
*Hd	7.36 (7.30–7.42)	7.37 (7.30–7.43)	7.40 (7.34–7.44)	7.41 (7.35–7.45)	7.28 (7.19–7.36)	7.37 (7.31–7.42)
PaO <sub>2</sub> , mm Hg	83 (68–108)	79 (67–100)	77 (67–93)	76.5 (67–93)	112 (81–155)	83 (69–103)
PaO <sub>2</sub> /FiO <sub>2</sub>	125 (86–178)	118 (80–163)	134 (96–185)	112 (75–158)	112 (81–155)	133 (89–178)
PaCO <sub>2</sub> , mm Hg	38 (34–45)	39 (34–45)	38 (33–43)	36 (31–41)	50 (42–62)	39 (34–45)
Bicarbonate, mmol/L	21.0 (18.0–25.0)	21.0 (17.4–25.0)	22.0 (18.0–26.0)	22.0 (18.0–25.0)	22.9 (19.4–26.3)	22.0 (18.0–25.0)
Ventilatory variables						
Tidal volume, mL	410 (360–470)	450 (400–510)	500 (420–600)	700 (600–750)	350 (308–400)	400 (350–460)
Per PBW, mL/kg PBW	6.3 (6.0–7.3)	7.1 (6.1–8.1)	7.9 (6.6–9.4)	10.2 (9.0–11.3)	5.9 (5.1–6.1)	6.2 (6.0–7.1)
Plateau pressure, cmH <sub>2</sub> O	24.0 (20.0–27.0)	26.0 (22.0–30.0)	26.0 (22.0–31.0)	29.0 (24.8–34.0)	26.0 (22.0–29.0)	24.0 (19.0–28.0)
PEEP, cmH <sub>2</sub> O	10 (5–12)	10 (5–12)	10 (5–12)	8 (5–10)	12 (10–14)	10 (5–11)
FiO2	0.60 (0.50–0.80)	0.60 (0.50–0.80)	0.60 (0.50–0.80)	0.60 (0.50–0.74)	0.70 (0.60–1.00)	0.60 (0.40–0.70)
Clinical outcomes						
28-day mortality, no. (%)	I	I	I	I	594 (58.8)	I
60-day mortality, no. (%)	227 (22.7)	268 (26.9)	144 (26.2)	141 (30.1)	I	199 (26.7)
						Continued

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Range of variable averages between studies



**Figure 1** Differences of the variables included in the cluster algorithm among clusters. Square symbols represent the study with the highest mean z score for each phenotype; circles represent the study with the lowest mean z score for each phenotype. The coloured bands are exclusively to help visualise the opposite trends of the variables on the different clusters; Art.pH, arterial pH; Bicarb, bicarbonate; MAP, mean arterial pressure; Creat, creatinine; Resp.Rate, respiratory rate.

subphenotype A (figure 2 and online supplemental eTables 8 and 9).

## DISCUSSION

This study successfully demonstrated that nine easily obtainable clinical variables: arterial pH, partial  $O_2$  pressure, creatinine, bilirubin, bicarbonate, mean arterial pressure, heart rate, respiratory rate and FiO<sub>2</sub> at the time of study enrolment can identify two distinct ARDS subphenotypes with different clinical and biologic characteristics as well as outcomes across the test and validation cohorts. There was good generalisability among diverse populations from multiple validation data sets with temporal and geographical differences.

It is understandable that researchers feel compelled to use as much information as possible to build robust models. This is supportable for two main reasons: (1) the well-known heterogeneity of complex syndromes such as ARDS and (2) the abundance of highly granular clinical data generated by EHRs. It is anticipated that analysing this vast amount of data will provide new knowledge regarding disease mechanisms by enabling researchers to find plausible hidden patterns within the data.<sup>29</sup> However, this data-heavy approach has the potential drawback of using predictors which are not generally obtained in a time window prior to intervention, or worse yet, using variables that are not part of the routine standard of care for patients. The rationale of using fewer and easy to collect clinical variables is not new in the field of critical care. Prognostic models have already shown that it is indeed feasible to create meaningful models using fewer predictors.<sup>30 31</sup>

Unfortunately, unlike supervised algorithms (eg, regression analyses), unsupervised algorithms such as K-means clustering do not provide one straightforward and

Table 1 Continued						
	Training set (n=199	8)	Validation set (n=27	75)		
	EDEN (n=1000)	FACTT (n=998)	ALVEOLI (n=549)	ARMA (n=472)	ART (n=1010)	SAILS (n=744)
90-day mortality, no. (%)	233 (23.3)	283 (28.6)	148 (27.5)	143 (30.8)	I	204 (27.4)
Ventilator-free days, day 28	20.0 (0.0–24.0)	17.0 (0.0–23.0)	18.0 (0.0–24.0)	13.0 (0.0–23.0)	0.0 (0.0–13.0)	20.0 (0.0–25.0)
Ventilator days in survivors	7.0 (4.0–13.0)	8.0 (5.0–16.0)	8.0 (4.0–14.0)	8.0 (4.0–15.0)	13.0 (8.0–20.0)	6.0 (4.0–11.0)
Data are median (quartile 25th-quartil           *Except for ART. that uses SAPS-3. all	ile 75th) or N (%). all studies use APACHE-IV.					

APACHE, Acute Physiology and Chronic Health Evaluation; PBW, predicted body weight; PEEP, positive end-expiratory pressure.

Table 2         Baseline character	eristics and clinical	outcomes accordir	ng to the c	lusters and trials in th	e training set	
	FACTT			EDEN		
	Cluster 1 (n=407)	Cluster 2 (n=294)	P value	Cluster 1 (n=449)	Cluster 2 (n=328)	P value
Age, year*	50.0 (40.0–63.0)	47.0 (36.0–58.0)	0.002	53.0 (44.0–63.0)	51.0 (41.0–62.2)	0.183
Male gender, no. (%)	223 (54.8)	151 (51.4)	0.411	233 (51.9)	168 (51.2)	0.910
Body mass index, kg/m <sup>2</sup>	27.5 (23.3–32.1)	27.4 (23.0–32.7)	0.938	29.1 (24.6–34.5)	28.5 (23.4–35.1)	0.476
Caucasian, no. (%)	269 (66.1)	177 (60.2)	0.129	349 (81.5)	237 (75.7)	0.067
Aetiology, no. (%)			<0.001			0.003
Pneumonia	201 (49.4)	139 (47.3)		296 (65.9)	217 (66.2)	
Sepsis	78 (19.2)	101 (34.4)		50 (11.1)	60 (18.3)	
Aspiration	67 (16.5)	30 (10.2)		45 (10.0)	27 (8.2)	
Trauma	24 (5.9)	8 (2.7)		24 (5.3)	5 (1.5)	
Other	37 (9.1)	16 (5.4)		34 (7.6)	19 (5.8)	
Prognostic scores						
APACHE III	69.0 (56.0–84.0)	91 (76.0–105.0)	<0.001	66.0 (54.0–79.0)	84.0 (71.0–100.2)	<0.001
Use of vasopressor, no. (%)	118 (29.5)	189 (64.9)	<0.001	187 (41.6)	209 (63.7)	<0.001
Vital signs						
Temperature, °C	37.5 (36.8–38.2)	37.6 (37.0–38.4)	0.371	37.3 (36.8–37.8)	37.3 (36.7–38.1)	0.212
Heart rate, bpm	95.0 (81.0–110.0)	114 (102–126)	<0.001	89 (77–102)	101 (89–116)	<0.001
Mean arterial pressure, mm Hg	76.0 (68.0–88.0)	71.0 (65.0–80.8)	<0.001	77.0 (68.0–84.0)	71.0 (66.0–80.0)	<0.001
SpO <sub>2</sub> , %	96 (93–98)	95 (92–97)	<0.001	96 (94–98)	95 (92–98)	0.032
Urine output in 24 hours, mL	1785 (1192–2853)	1370 (842–2446)	<0.001	1505 (977–2250)	1165 (566–1816)	<0.001
Laboratory tests						
Haematocrit, %	30.0 (26.0–33.0)	30.0 (24.2–35.0)	0.272	30.0 (26.0–34.0)	30.0 (26.0–35.0)	0.919
White cell count, 10 <sup>9</sup> /L	11.6 (7.3–16.3)	11.7 (5.6–17.9)	0.972	11.4 (7.7–15.5)	12.7 (7.7–19.0)	0.019
Platelets, 10 <sup>9</sup> /L	195 (118.5–268)	158 (87–237)	<0.001	163 (108–241)	164 (103–227)	0.552
Creatinine, mg/dL	0.9 (0.7–1.3)	1.4 (1.0–2.0)	<0.001	1.0 (0.7–1.5)	1.6 (1.0–2.8)	<0.001
Bilirubin, mg/dL	0.7 (0.5–1.3)	0.9 (0.5–2.0)	0.003	0.8 (0.5–1.3)	0.8 (0.5–1.7)	0.128
Arterial blood gas						
pH*	7.41 (7.36–7.45)	7.29 (7.23–7.35)	<0.001	7.40 (7.35–7.44)	7.30 (7.24–7.35)	<0.001
PaO <sub>2</sub> , mm Hg	78 (68–100)	78 (65–99)	0.240	83 (70–107)	81 (67–107)	0.416
PaO <sub>2</sub> /FiO <sub>2</sub>	132 (92–173)	89 (65–126)	<0.001	133 (98–193)	101 (73–162)	<0.001
PaCO <sub>2</sub> , mm Hg	39 (34–44)	38.5 (33–47.9)	0.877	38 (34–44)	38 (33–46)	0.55
Bicarbonate, mmol/L	24.0 (21.0–27.0)	17.0 (14.0–20.0)	<0.001	23.0 (21.0–26.0)	18.5 (15.0–21.0)	<0.001
Ventilatory variables						
Tidal volume, mL	450 (400–530)	450 (382–500)	0.009	420 (356–487)	400 (350–450)	0.032
Per PBW, mL/kg PBW	7.1 (6.3–8.4)	7.0 (6.0, 8.0)	0.058	6.3 (6.0–7.5)	6.1 (6.0–7.3)	0.079
Plateau pressure, cmH <sub>2</sub> O	25.0 (20.0–29.0)	28.0 (24.0–32.0)	<0.001	23.0 (19.0–27.0)	24.0 (21.0–28.0)	0.004
PEEP, cmH <sub>2</sub> O	8 (5–10)	10 (8–14)	<0.001	10 (5–10)	10 (8–14)	<0.001
Respiratory rate, breaths/ min	22 (18–27)	30 (24–35)	<0.001	22 (19–26)	30 (25–35)	<0.001
FiO <sub>2</sub>	0.50 (0.40–0.70)	0.80 (0.60–1.00)	<0.001	0.60 (0.45–0.70)	0.80 (0.60–1.00)	<0.001

Data are mean±SD, median (quartile 25th–quartile 75th) or N (%).

APACHE, Acute Physiology and Chronic Health Evaluation; PEEP, positive end-expiratory pressure; VT/PBW, tidal volume per predicted body weight.

Table 3 Clinical outcomes according to clust	ters in each trial			
	Cluster 1	Cluster 2	Difference (95% CI)	P value
Training set				
FACTT	n=407	n=294		
60-day mortality, no. (%)	94 (23.1)	102 (34.7)	11.6% (4.9% to 18.3%)	0.001
90-day mortality, no. (%)	103 (25.4)	106 (36.3)	10.9% (4.1% to 17.8%)	0.002
Ventilator-free days at day 28	19.0 (0.0–24.0)	10.0 (0.0–21.0)	–9.0 (–11.9 to –6.1)	<0.001
Duration of ventilation in survivors, days	8.0 (4.0–13.0)	10.0 (7.0–19.0)	2.0 (0.5 to 3.5)	<0.001
EDEN	n=449	n=328		
60-day mortality, no. (%)	87 (19.4)	90 (27.4)	8.1% (2.1% to 14.0%)	0.010
90-day mortality, no. (%)	90 (20.0)	93 (28.4)	8.3% (2.3% to 14.3%)	0.009
Ventilator-free days at day 28	21.0 (0.0–25.0)	15.0 (0.0–22.2)	–6.0 (–8.1 to –3.9)	<0.001
Duration of ventilation in survivors, days	6.0 (4.0–11.0)	8.0 (6.0–18.0)	2.0 (0.9 to 3.1)	<0.001
Validation set				
ALVEOLI	n=336	n=157		
60-day mortality, no. (%)	59 (17.6)	68 (43.3)	25.8% (17.7% to 33.8%)	<0.001
90-day mortality, no. (%)	60 (18.1)	70 (45.5)	27.3% (19.2% to 35.5%)	<0.001
Ventilator-free days at day 28	21.0 (4.8–25.0)	2.0 (0.0–19.0)	–19.0 (–20.8 to –17.2)	<0.001
Duration of ventilation in survivors, days	7.0(4.0,13.0)	11.0 (6.0–22.2)	4.0 (2.1 to 5.9)	<0.001
ARMA	n=279	n=100		
60-day mortality, no. (%)	69 (24.8)	42 (42.0)	17.2% (6.9% to 27.5%)	0.002
90-day mortality, no. (%)	70 (25.5)	42 (42.0)	16.5% (6.0% to 26.9%)	0.003
Ventilator-free days at day 28	17.0 (0.0–24.0)	2.0 (0.0–19.0)	–15.0 (–18.6 to –11.4)	<0.001
Duration of ventilation in survivors, days	7.0 (4.0–13.8)	11.0 (5.0–18.0)	4.0 (1.5 to 6.5)	0.018
SAILS	n=319	n=188		
60-day mortality, no. (%)	80 (25.1)	60 (31.9)	6.8% (–1.2% to 14.9%)	0.119
90-day mortality, no. (%)	81 (25.4)	63 (33.5)	8.1% (0.0% to 16.3%)	0.063
Ventilator-free days at day 28	21.0 (0.0–25.0)	16.0 (0.0–23.0)	–5.0 (–7.3 to –2.7)	<0.001
Duration of ventilation in survivors, days	6.0 (3.0–10.0)	8.0 (5.0–14.0)	2.0 (0.7 to 3.3)	<0.001
ART	n=211	n=298		
28-day mortality, no. (%)	81 (38.4)	180 (60.4)	22.0% (13.4% to 30.7%)	<0.001
Ventilator-free days at day 28	0.0 (0.0–17.0)	0.0 (0.0–7.8)	-0.0 (-1.0 to 1.0)	<0.001
Duration of ventilation in survivors, days	12.0 (8.0–20.0)	13.5 (8.0–20.0)	2.0 (-0.3 to 4.2)	0.570

Data are median (quartile 25th–quartile 75th) or N (%). Difference is mean difference with (95% CI) for binomial variables and median difference with (95% CI) for continuous variables.

established metric to describe feature importance. In that sense, our approach of testing multiple sets of variables was also meant to select features that were most likely to be relevant, serving as surrogate for the feature selection step normally employed in supervised algorithms. While each individual variable by itself may not be significantly different across subphenotypes, their interaction in the nine-dimensional space of our model may be relevant.

Our initial choices to define variables commonly found in the EHR at ARDS diagnosis was inspired by a recent report from the WHO which showed an enormous discrepancy of medical devices availability in a survey across 135 countries.<sup>29</sup> Recognising this inconsistency is essential for widespread implementation of machine learning models regardless of varying availability of resources across countries and health systems.<sup>29</sup> The aim is to provide clinically relevant information within a defined and short period that might impact the delivery of effective interventions to the right patient population and to as many patients as possible.<sup>29</sup>

Recently, Sinha *et al* developed supervised-learning gradient boosted classifier models trained using 24 or 14 readily available clinical data elements to reproduce biomarker-derived subphenotypes which were previously identified by Calfee *et al.*<sup>17</sup> Unlike Sinha *et al*, who predicted previously identified subphenotypes, our study has identified two subphenotypes de novo using a small set of clinical variables.

Although the subphenotypes that we have identified and those that have been previously published look



**Figure 2** Heat map of the biomarkers available for the ARMA and ALVEOLI trials. For better visualisation and due to difference in scales, the values were log-normalised and z-scored. Subphenotypes A and B are shown separately to highlight their differences.

similar, our work is distinct from previous studies in several ways. We employed different training and validation data sets as well as a different and well-established unsupervised learning technique. Moreover, we utilised a process for selecting predictors which is not comparable to previous studies. Acknowledging these differences is crucial. It would not be unexpected to assume that these deviations would be relevant enough to produce different subphenotypes.<sup>32</sup> However, the clinical, laboratory characteristics and the clinical outcomes of our subphenotypes show that they are remarkably similar to subphenotypes found in previous papers, regardless of methodological differences.

At this point it is not possible to go beyond this comparative analysis, as there is no gold standard definition of ARDS subphenotypes.<sup>32</sup> Nonetheless, our work does provide robust evidence that ARDS does indeed have two subphenotypes that can be systematically identified, despite major differences in population assessed and methodological approach used compared with previous studies. It also reinforces that we should continue to explore the underlying biological pathways of such subphenotypes to find responders to new or previously tested therapies.

Our study has several strengths. First, it is the largest cohort of patients that has been studied to develop distinct subphenotypes of ARDS patients. Moreover, our validation cohort included patients from the ART trial, allowing us to validate our model in the contemporaneous population of a large international randomised clinical trial in addition to the ARDSnet studies used in other subphenotyping studies. Second, our subphenotyping model was developed exclusively on the training set and then validated across multiple separate data sets. Nevertheless, similar separation in mortality was seen between the two subphenotypes across all trials. Third, we used the K-means algorithm to identify our subphenotypes, and the results obtained with this technique can be easily interpreted by clinicians and implemented in clinical practice. Finally, this is the first phenotyping study

that has used easily available clinical variables to identify ARDS phenotypes de novo, which allows for early identification of these patients in the clinical care at the bedside. Using this algorithm with a small number of routinely collected variables could enable our model to be applied in trials that either retrospectively or prospectively assess interventions targeted to each subphenotype.

This study also has limitations. First, we have developed our models exclusively on patients enrolled in clinical trials. Due to the strict inclusion and exclusion criteria of these clinical trials, the generalisability of these results needs to be evaluated in unselected ARDS populations. Although there are clear clinical and biomarker differences between the identified subphenotypes, the model's clinical utility needs to be prospectively validated and further investigated. Additionally, our biomarker analysis is limited to those patients in which the data were made publicly available by the study authors, but future collection of biomarker data in a prospective study will allow more robust understanding of the underlying biology and validation of the subphenotype model. Also, K-means clustering does not handle missing data, and no approach was used to impute missing values. However, the extremely low rate of missingness in our study makes this issue less relevant. Finally, future work should analyse previous trials to identify possible differential treatment responses for the subphenotypes of ARDS patients identified in this study.

## **CONCLUSIONS**

This study confirms the existence of two distinct subphenotypes in ARDS patients using a novel clustering model on routinely collected clinical data. This work may allow for easier identification of ARDS subphenotypes to facilitate implementation of precision clinical trial enrolment and development of targeted therapies in a variety of settings without the added burdens of biomarker evaluation.

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## <u>d</u>

## Patient consent for publication Not applicable.

Ethics approval All patients provided informed consent in the original trials. This secondary analysis study was exempt from IRB review because it does not meet the definition of human subject as defined in 45 CFR 46.102. WIRB Work Order #1-1228617-1.

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Data availability statement Data are available upon reasonable request. All data relevant to the study are included in the article or uploaded as supplementary information. Data from the ARDSnet studies (EDEN, FACTT, ARMA, ALVEOLI) may be requested from the NIH-NHLBI BioLINCC repository. These are available by application through the BioLINCC website (https://biolincc.nhlbi.nih.gov/resource\_overview/), subject to established review processes for their distribution to qualified investigators and data from the ART trial can be requested from study authors.

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#### REFERENCES

- ARDS Definition Task Force, Ranieri VM, Rubenfeld GD, et al. Acute respiratory distress syndrome: the Berlin definition. JAMA 2012;307:2526–33.
- 2 Thille AW, Esteban A, Fernández-Segoviano P, et al. Comparison of the Berlin definition for acute respiratory distress syndrome with autopsy. Am J Respir Crit Care Med 2013;187:761–7.
- 3 Reilly J, Calfee C, Christie J. Acute respiratory distress syndrome phenotypes. *Semin Respir Crit Care Med* 2019;40:019–30.
- 4 Reddy K, Sinha P, O'Kane CM, et al. Subphenotypes in critical care: translation into clinical practice. Lancet Respir Med 2020;8:631–43.
- Shankar-Hari M, Fan E, Ferguson ND. Acute respiratory distress syndrome (ARDS) phenotyping. *Intensive Care Med* 2019;45:516–9.
   Calfee CS, Delucchi K, Parsons PE, et al. Subphenotypes in acute
- respiratory distress syndrome: latent class analysis of data from two randomised controlled trials. *Lancet Respir Med* 2014;2:611–20.
  7 Famous KR, Delucchi K, Ware LB, *et al.* Acute respiratory distress
- syndrome subphenotypes respond differently to randomized fluid management strategy. *Am J Respir Crit Care Med* 2017;195:331–8.
- 8 Calfee CS, Delucchi KL, Sinha P, et al. Acute respiratory distress syndrome subphenotypes and differential response to simvastatin: secondary analysis of a randomised controlled trial. Lancet Respir Med 2018;6:691–8.
- 9 Sinha P, Delucchi KL, Thompson BT, et al. Latent class analysis of ARDS subphenotypes: a secondary analysis of the statins for acutely injured lungs from sepsis (sails) study. *Intensive Care Med* 2018;44:1859–69.
- 10 Bos LD, Schouten LR, van Vught LA, et al. Identification and validation of distinct biological phenotypes in patients with acute respiratory distress syndrome by cluster analysis. *Thorax* 2017;72:876–83.

- 11 Ferguson ND, Fan E, Camporota L, *et al.* The Berlin definition of ARDS: an expanded rationale, justification, and supplementary material. *Intensive Care Med* 2012;38:1573–82.
- 12 Bellani G, Laffey JG, Pham T, et al. Epidemiology, patterns of care, and mortality for patients with acute respiratory distress syndrome in intensive care units in 50 countries. JAMA 2016;315:788.
- 13 Gattinoni L, Vassalli F, Romitti F. Benefits and risks of the P/F approach. Intensive Care Med 2018;44:2245–7.
- 14 Matthay MA, Arabi YM, Siegel ER, *et al.* Phenotypes and personalized medicine in the acute respiratory distress syndrome. *Intensive Care Med* 2020;46:2136–52.
- 15 Shankar-Hari M, Rubenfeld GD. Population enrichment for critical care trials: phenotypes and differential outcomes. *Curr Opin Crit Care* 2019;25:489–97.
- 16 Sinha P, Delucchi KL, McAuley DF, et al. Development and validation of parsimonious algorithms to classify acute respiratory distress syndrome phenotypes: a secondary analysis of randomised controlled trials. *Lancet Respir Med* 2020;8:247–57.
- 17 Sinha P, Churpek MM, Calfee CS. Machine learning classifier models can identify acute respiratory distress syndrome phenotypes using readily available clinical data. *Am J Respir Crit Care Med* 2020;202:996–1004.
- 18 Kitsios GD, Yang L, Manatakis DV, et al. Host-Response subphenotypes offer prognostic enrichment in patients with or at risk for acute respiratory distress syndrome. *Crit Care Med* 2019;47:1724–34.
- 19 Acute Respiratory Distress Syndrome Network, Brower RG, Matthay MA, et al. Ventilation with lower tidal volumes as compared with traditional tidal volumes for acute lung injury and the acute respiratory distress syndrome. N Engl J Med 2000;342:1301–8.
- 20 Brower RG, Lanken PN, MacIntyre N, et al. Higher versus lower positive end-expiratory pressures in patients with the acute respiratory distress syndrome. N Engl J Med 2004;351:327–36.
- 21 National Heart, Lung, and Blood Institute Acute Respiratory Distress Syndrome (ARDS) Clinical Trials Network, Wheeler AP, Bernard GR, et al. Pulmonary-Artery versus central venous catheter to guide treatment of acute lung injury. N Engl J Med 2006;354:2213–24.
- 22 National Heart, Lung, and Blood Institute Acute Respiratory Distress Syndrome (ARDS) Clinical Trials Network, Rice TW, Wheeler AP, *et al.* Initial trophic vs full enteral feeding in patients with acute lung injury: the EDEN randomized trial. *JAMA* 2012;307:795–803.
- 23 National Heart, Lung, and Blood Institute ARDS Clinical Trials Network, Truwit JD, Bernard GR, *et al.* Rosuvastatin for sepsisassociated acute respiratory distress syndrome. *N Engl J Med* 2014;370:2191–200.
- 24 Writing Group for the Alveolar Recruitment for Acute Respiratory Distress Syndrome Trial (ART) Investigators, Cavalcanti AB, Suzumura Érica Aranha, et al. Effect of lung recruitment and titrated positive end-expiratory pressure (PEEP) vs low PEEP on mortality in patients with acute respiratory distress syndrome: a randomized clinical trial. JAMA 2017;318:1335.
- 25 Yehya N, Harhay MO, Curley MAQ, et al. Reappraisal of Ventilator-Free days in critical care research. Am J Respir Crit Care Med 2019;200:828–36.
- 26 Castela Forte J, Perner A, van der Horst ICC. The use of clustering algorithms in critical care research to unravel patient heterogeneity. *Intensive Care Med* 2019;45:1025–8.
- 27 KETCHEN Jr. DJ, Shook CL. The application of cluster analysis in strategic management research: an analysis and critique. *Strateg Manag J* 1996;17:441–58.
- 28 Caliński T, Harabasz J. A dendrite method for cluster analysis. *Communications in Statistics* 1974;3:1–27.
- 29 Bulgarelli L, Deliberato RO, Johnson AEW. Prediction on critically ill patients: The role of "big data". *J Crit Care* 2020;60:64–8.
- 30 Johnson AEW, Kramer AA, Clifford GD. A new severity of illness scale using a subset of acute physiology and chronic health evaluation data elements shows comparable predictive accuracy. *Crit Care Med* 2013;41:1711–8.
- 31 Deliberato RO, Escudero GG, Bulgarelli L, et al. SEVERITAS: an externally validated mortality prediction for critically ill patients in low and middle-income countries. *Int J Med Inform* 2019;131:103959.
- 32 DeMerle KM, Angus DC, Baillie JK, *et al.* Sepsis subclasses: a framework for development and interpretation. *Crit Care Med* 2021;49:748–59.