Ther Adv Urol

2021, Vol. 13: 1-18

DOI: 10.1177/ 17562872211022296

© The Author(s), 2021. Article reuse guidelines: sagepub.com/journalspermissions

# Pokket Sirisreetreerux<sup>1</sup>, Rujira Wattanayingcharoenchai, Sasivimol Rattanasiri, Oraluck Pattanaprateep, Pawin Numthavaj<sup>1</sup> and Ammarin Thakkinstian

network meta-analysis and risk-benefit

Medical and non-medical interventions for

post-operative urinary retention prevention:

# Abstract

analysis

**Aims:** To assess the efficacy in lowering post-operative urinary retention, urinary tract infection and lower urinary tract symptoms and the incidence of adverse events among 12 interventions and to perform risk-benefit analysis.

**Methods:** Previous randomized controlled trials were identified from *MEDLINE*, *Scopus* and *CENTRAL* database up to January 2020. The interventions of interest included early ambulation, fluid adjustment, neuromodulation, acupuncture, cholinergic drugs, benzodiazepine, antispasmodic agents, opioid antagonist agents, alpha-adrenergic antagonists, non-steroidal anti-inflammatory drugs (NSAIDs) and combination of any interventions. The comparators were placebo or standard care or any of these interventions. Network meta-analysis was performed. The probability of being the best intervention was estimated and ranked using rankogram and surface under the cumulative ranking curve. Risk-benefit analysis was done. Incremental risk-benefit ratio (IRBR) was calculated and risk-benefit acceptability curve was constructed.

**Results:** A total of 45 randomized controlled trials with 5387 patients was included in the study. Network meta-analysis showed that early ambulation, acupuncture, alpha-blockers and NSAIDs significantly reduced the post-operative urinary retention. Regarding urinary tract infection and lower urinary tract symptoms, no statistical significance was found among interventions. Regarding the side effects, only alpha-adrenergic antagonists significantly increased the adverse events compared with acupuncture and opioid antagonist agents from the indirect comparison. According to the cluster ranking plot, acupuncture and early ambulation were considered high efficacy with low adverse events, corresponding to the IRBR. **Conclusion:** Early ambulation, acupuncture, opioid antagonist agents, alpha-adrenergic antagonists and NSAIDs significantly reduce the incidence of post-operative urinary retention with no difference in adverse events. Regarding the risk-benefit analysis of the medical treatment, alpha-adrenergic antagonists have the highest probability of net benefit at the acceptable threshold of side effect of 15%, followed by opioid antagonist agents, NSAIDs and cholinergic drugs.

Keywords: meta-analysis, post-operative urinary retention, prevention, risk-benefit analysis

Received: 22 July 2020; revised manuscript accepted: 5 January 2021.

#### Background

Urinary retention is the inability to pass urine despite persistent effort, which can be acute and

chronic urinary retention. Generally, acute urinary retention is defined as a painful, palpable or percussible bladder and the patient is unable to Correspondence to: Pawin Numthavaj Department of Clinical Epidemiology and Biostatistics, Faculty of Medicine Ramathibodi Hospital, Mahidol University, 270 Rama VI Road, Ratchathewi, Bangkok, 10400, Thailand. pawin.num@mahidol.ac.th

Pokket Sirisreetreerux

Department of Clinical Epidemiology and Biostatistics, Faculty of Medicine Ramathibodi Hospital, Mahidol University, Bangkok, Thailand

Division of Urology, Department of Surgery, Faculty of Medicine Ramathibodi Hospital, Mahidol University, Bangkok, Thailand

#### Rujira

Thailand

Wattanayingcharoenchai Department of Clinical Epidemiology and Biostatistics, Faculty of Medicine Ramathibodi Hospital, Mahidol University, Bangkok,

Department of Obstetrics and Gynecology, Faculty of Medicine Ramathibodi Hospital, Mahidol University, Bangkok, Thailand

Sasivimol Rattanasiri Oraluck Pattanaprateep

Ammarin Thakkinstian Department of Clinical Epidemiology and Biostatistics, Faculty of Medicine Ramathibodi Hospital, Mahidol University, Bangkok, Thailand

journals.sagepub.com/home/tau



Creative Commons Non Commercial CC BY-NC: This article is distributed under the terms of the Creative Commons Attribution-NonCommercial 4.0 License (https://creativecommons.org/licenses/by-nc/4.0/) which permits non-commercial use, reproduction and distribution of the work without further permission provided the original work is attributed as specified on the SAGE and Open Access pages (https://us.sagepub.com/en-us/nam/open-access-at-sage).

pass the urine, which is commonly seen after surgery with long operation time. This is known as post-operative urinary retention (POUR), which results in urinary complications including urinary tract infection (UTI), patients' discomfort and distress, prolonged hospital stay and increased cost of the treatment. The POUR incidence varied from 2.1% to 80% according to surgery types and specific study factors (e.g. age, aggressiveness of the bladder dissection, the use of opiates and underlying comorbidities).<sup>1–4</sup>

The POUR may be caused by several mechanisms,<sup>5</sup> including surgery (e.g. voiding reflex interruption, perioperative medication,<sup>6</sup> sensation of bladder fullness impairment from anesthesia, the imbalance of sympathetic and parasympathetic systems7) and other factors (e.g. immobilization, voiding in supine position, or perineal and lower abdominal pain) that could inhibit the perineal relaxation that is necessary for voiding.8 Various non-medical and medical strategies were introduced aiming to prevent POUR, for example, early ambulation (AMB), neuromodulation (NEU), antispasmodic agents (ASPs), non-steroidal anti-inflammatory drugs suppository (NSD), alpha-adrenergic antagonist (ALP), et cetera. Currently, no standard guideline has been approved for preventing this condition. According to previous meta-analysis, the effect of POUR prevention was pooled and demonstrated good efficacy in ALP, AMB and ASP compared with placebo.9 However, only a few studies have pooled the adverse events (AEs) and other outcomes such as UTI, post-void residual urine (PVRU), and lower urinary tract symptoms (LUTSs). In addition, no network meta-analysis has been reported and compared among all interventions. Our study aimed to evaluate the efficacy of all interventions, both non-medical and medical, in prophylactic POUR, UTI, PVRU, LUTS and AE, along with performing risk-benefit analysis.

## Methods

#### Protocol

The systematic review and network meta-analysis protocol was developed following the guidelines in the PRISMA extension of network meta-analysis (NMA). The review protocol was registered with PROSPERO (CRD 42019145653).

#### Search strategy and criteria for study inclusion

The studies were identified primarily from *PubMed, Scopus* and *CENTRAL* up to January 2020 and updated monthly. Additionally, *Thai-Journal Citation Index*, WHO International Clinical Trials Registry Platform (www.clinicaltrials.gov) and references of selected articles and previous systematic reviews were used to identify the eligible studies. The search terms were constructed and followed the interested population, medication name and procedures of each intervention and outcome, which were POUR, UTI, PVRU, LUTS and AE. The full search strategies are available in the Supplemental Material Table 1 online.

This study included only the randomized controlled trials (RCTs) without language restriction with the following criteria: conducted in adults aged 18 years or older who underwent any type of surgery; compared between any of the following interventions: AMB, perioperative fluid adjustment (FLU), NEU, acupuncture (ACU), cholinergic drug (CHO), benzodiazepine (BENZ), ASP, intravenous opioid antagonist related (OPI), ALP, NSD, combined interventions (COMBs), and placebo/standard care (PLA); and having at least one of the following outcomes: POUR, UTI, PVRU, LUTS and AE. The studies with insufficient data for pooling after three attempts of contacting the author every 2 weeks and the studies published in languages which reviewers could not translate were excluded.

#### Interventions and outcomes

There were 11 interventions with placebo included in our study. Four non-medical interventions were AMB, FLU, NEU and ACU. AMB was defined as the patients having no absolute bed rest and were mobilized to the toilet. FLU intended to restrict oral and intravenous fluid during perioperative and early postoperative periods. ACU included acupuncture procedure with or without electrical stimulation. CHO included bethanechol, neostigmine or distigmine by both oral and intravenous routes. BENZ comprised all oral or intravenous medication in this group. ASP contained drotaverine administered perioperatively. OPI consisted of any dosage of methylnaltrexone or naloxone given along with opioid. ALP included all medications and dosage of this drug group. NSD included any form and dosage of non-steroidal anti-inflammatory drug. PLA could be placebo, or no intervention or standard care.

The primary outcome was the incidence of POUR, which was diagnosed according to original studies. Briefly, most studies diagnosed by clinical symptoms and signs<sup>8</sup> (e.g. inability to urinate within 7 days after the surgery, after urethral catheter removal, or the feeling of strong desire to void or palpable suprapubic mass) with or without confirmation by urine volume measurement by urethral catheterization or ultrasonography. The secondary outcomes were symptomatic UTI with/without laboratory confirmation, and PVRU measured in millimeters by catheterization or ultrasonography. LUTS was reported in scores from any questionnaires and AEs.

# Data extraction and study quality assessment

Data extraction was done independently by at least two of three reviewers (PS, PN, RW) using a data extraction form. Disagreement between reviewers was discussed and solved by a team. The extracted data included general characteristics of article, baseline patients' characteristics (e.g. number of patients, mean age, type of operations, operative time, type of anesthesia), outcome measurement, interventions and data for pooling, that is, a total number of patients and events for dichotomous outcome, and mean and SD for continuous outcome.

Risk of bias assessment was done independently by at least two of three reviewers (PS, PN, RW). The assessment of risk of bias was performed using the revised Cochrane risk-of-bias tool for randomized trials (RoB 2),<sup>10</sup> which comprised five domains (randomization, deviations from the intended intervention, missing outcome data, measurement of the outcome, selection of the reported result) along with overall risk of bias. Each domain and overall risk of bias was rated as low, some concerns, or high.

# Data synthesis and statistical analyses

Direct meta-analysis (DMA). DMA was performed if there were at least three studies for each comparison. Risk ratio (RR) for dichotomous outcomes and mean difference (MD) for continuous outcome were estimated. These were then pooled across studies using a fixed-effect model if heterogeneity was not present, otherwise a randomeffect model was applied. Heterogeneity was assessed using Cochran's Q and Higgins's  $I^2$  statistic,<sup>11</sup> and considered present if *p*-value was less than or equal to 0.1 or  $I^2$  more than 25%. Source of heterogeneity was explored by fitting each of the co-variables (e.g. age, type of operation, type of anesthesia) in a meta-regression model. If the included variable could decrease  $I^2$  more than 50%, subgroup analysis and sensitivity analysis were performed accordingly. Publication bias was assessed using funnel plot and Egger test. If asymmetric funnel plot presented, contour enhanced funnel plot was then applied to distinguish the cause of asymmetry.<sup>12</sup>

NMA. The interventions of interest were coded as 0, 1, 2, 3, 4, 5, 6, 7, 8, 9, 10 and 11 for PLA, AMB, FLU, NEU, ACU, CHO, BENZ, ASP, OPI, ALP, NSD and COMB, respectively. The network of all treatments was mapped. Two-stage NMA was applied<sup>13</sup> as follows: first, a linear regression was applied to estimate relative treatment effects (i.e. In(RR), MDs). Second, relative treatment effects were pooled across the studies using multivariate meta-analysis with random-effect model. The probability of being the best intervention in each outcome was estimated and ranked using rankogram and surface under the cumulative ranking curve (SUCRA). Inconsistency assumption was assessed by applying the design-by-treatment interaction model. Publication bias was evaluated using comparison-adjusted funnel plot.

*Cluster rank and risk-benefit analysis.* A cluster ranking plot was created simultaneously considering SUCRAs of benefit (i.e. lowering POUR, UTI, PVRU, LUTS) and the risk of AE. The higher SUCRA of the outcomes and AEs represented the better efficacy and fewer AEs. The plot was divided into four quadrants, in which the right upper quadrant reflected the intervention with good efficacy and fewer AEs, whereas the left lower quadrant represented the intervention with low efficacy and high AEs.

Incremental risk–benefit ratio (IRBR) of POUR prevention was calculated as a ratio of an increment in risk ( $\Delta$ R) divided by an increment in benefit of POUR prevention ( $\Delta$ B) estimated by NMA. Monte Carlo method was used to simulate  $\Delta$ R and  $\Delta$ B with 1000 replications assuming normal distribution for both. The risk–benefit plane was constructed with the *Y*-axis representing  $\Delta$ R and the *X*-axis representing  $\Delta$ B along with risk– benefit acceptability threshold,<sup>14</sup> that is, the willingness to accept additional AEs for one additional benefit. This plot was divided into four quadrants, in which the right lower and right upper quadrants implied good benefit with fewer AEs; and good benefit, but higher AEs, respectively. For the latter quadrant, the intervention would be favored if the IRBR fell under the line of the risk– benefit acceptability threshold. In addition, risk– benefit acceptability curve (RBAC) was created based on the risk–benefit plane with any value of threshold.<sup>15</sup> This curve indicated the probability of the intervention to give the net benefit for all thresholds.

In terms of net benefit framework, net clinical benefit (NCB) was calculated by multiplying the acceptable risk threshold with  $\Delta B$  then minus with  $\Delta R.^{16}$  We also generated the NCB probability curve with NCB in the Y-axis and any value of acceptability threshold in the X-axis.<sup>15</sup> This curve represented the number of additional AEs that occured per 100 patients treated at a given threshold.

All analyses were performed using STATA® version 16.0 (StataCorp LP, College Station, TX, USA) and the simulations were done using Microsoft® Excel 2019 (Microsoft Corp., Seattle, WA, USA).

# Results

# Study selection and characteristics of eligible studies

A total of 5425 studies was identified, but only 45 RCTs were included with a total of 5387 patients; see Figure 1. Most studies had two arms with included interventions of three studies for AMB, four for FLU, two for NEU, three for ACU, five for CHO, one for BENZ, one for ASP, three for OPI, 22 for ALP, two for NSD and one for COMB of CHO and BENZ. Five outcomes were considered, including POUR (N=44), PVRU (N=10), UTI (N=7), LUTS (N=4) and AE (N=23). The characteristics of included RCTs are shown in Table 1 and Supplemental Table 2.

## Risk of bias assessment

From 45 RCTs, only three (6.7%) studies were identified as low overall risk of bias (RoB), 25 studies (55.6%) were considered as some concern and 17 (37.8%) studies were high RoB in the overall module. When each module was considered individually, there were six (13.3%) studies in randomization process, 38 (84.4%) in deviations from the intended intervention, 42 (93.3%) in missing outcome data, 33 (73.3%) in measurement of the

outcome and six (13.3%) in selection of the reported result that were graded as low RoB. The RoB is shown in Supplemental Figure 1.

# DMA

DMA was performed according to five outcomes, that is, POUR (39 studies, n=4,697), PVRU (five studies, n = 490), UTI (four studies, n=326), LUTS (three studies, n=437) and AE (19 studies, n=1949). For POUR, 39 studies were included with 4/6 pairs of interventions<sup>17,18,20,21,23,26,28,30–32,35–56,60–63</sup> which were significant in the incidence of POUR relative to PLA, that is, AMB (n=398), ACU (n=325), OPI (n=429) and ALP (n=5206) with the pooled RR [95% confidence interval (CI)] of 0.44 (0.32, 0.61), 0.32 (0.11, 0.94), 0.62 (0.40, 0.95) and 0.55 (0.40, 0.76), respectively (Supplemental Figure 2). The heterogeneity was varied with the  $I^2$  ranged from 63.54 to 73.34. Subgroup analysis was performed in ALP versus PLA according to clinical significance including gender and operative type. The results indicated that the prophylactic effect of ALP was statistically significant only in male patients (pooled RR 0.45, 95% CI 0.29, 0.69). Considering the type of operation, we divided the operation into neurourological procedure (i.e. spinal cord surgery, prostate surgery and anti-incontinence surgery) and non-neurourological procedure (i.e. craniofacial surgery, cataract, thoracic surgery or limb surgery). Subgroup analysis was performed according to the type of operation in each gender and found that only non-neurourological operation had significant prophylactic effect from ALP in both genders (Supplemental Figure 3).

Only ALP versus PLA was pooled for UTI (four studies, n = 326), 45, 48, 50, 56 PVRU (five studies, n = 490)<sup>42,45,46,48,54</sup> and LUTS (three studies, n=437),<sup>42,46,50</sup> but their effects were not significant in lowering these outcomes (Supplemental Figures 4–6). In terms of AE, 23 studies reported AE but only 19 studies of three comparisons which had at least three studies in each comparison were included in the analysis.<sup>17,18,30-32,61,62</sup> Among these, overall and individual AEs were reported in 10 and nine studies (Supplemental Table 2).<sup>38,40,41,44,48–52,54–56,60</sup> For those reported individual AEs, we used the maximum incidence of individual AE as the composite outcome. ALP significantly increased the AE relative to placebo, as shown in Supplemental Figure 7 with RR of  $1.72 (1.07, 2.78, I^2 0).$ 





**Figure 1.** Flow diagram showing selection of articles for review. AE, adverse event; LUTS, lower urinary tract symptom; PICO, patient-intervention-comparison-outcome; POUR, postoperative urinary retention; RU, residual urine; UTI, urinary tract infection.

Publication bias was assessed using the Egger's test, which suggested a symmetrical plot in all comparisons, corresponding to the funnel plot except POUR in ALP *versus* PLA. Contour-enhanced funnel plots suggested that the asymmetrical plots resulted from heterogeneity, not from publication bias (Supplemental Figure 8).

#### NMA

*POUR.* Forty-four<sup>17,18,20,21,23–26,28,30–53,55,56,58–63</sup> RCTs with 12 interventions and 5338 patients were included; see the network map in Figure 2. Although all interventions reduced POUR, only AMB, ACU, ALP and NSD were statistically significant compared with placebo, but none of the

# Therapeutic Advances in Urology 13

# Table 1. Characteristics of included studies.

Author	Country	Intervention	Total number of patients	Mean age (years)	Male (%)	Operation type	Mean operative time (min)	GA(%)	Reported outcome
AMB versus PLA									
Hansen <i>et al.</i> <sup>17</sup>	Denmark	AMB <i>versus</i> PLA	138	49.00	51.45	Disc herniation repair	60.51	100	POUR, AE
Kim et al. <sup>18</sup>	Korea	AMB <i>versus</i> PLA	153	35.45	61.44	Anorectal surgery	34.41	0	POUR, AE
Kim et al. <sup>19</sup>	Korea	AMB <i>versus</i> PLA	107	34.99	63.55	Anorectal surgery	N/A	0	POUR, AE
FLU versus PLA									
Bailey et al. <sup>20</sup>	USA	FLU <i>versus</i> PLA	496	47.00	23.19	Anorectal surgery	N/A	88.51	POUR
Kozol <i>et al.</i> <sup>21</sup>	USA	FLU <i>versus</i> PLA	113	54	100	Herniorrhaphy	N/A	16.81	POUR
Young et al. <sup>22</sup>	Korea	FLU <i>versus</i> PLA	80	40.53	46.25	Anorectal surgery	15.38	0	POUR
Orbey et al. <sup>23</sup>	Turkey	FLU <i>versus</i> PLA	38	42	57.89	Benign anorectal disease	N/A	0	POUR
NEU versus PLA									
Butwick et al. <sup>24</sup>	UK	NEU <i>versus</i> PLA	43	70.47	74.42	Knee replacement	N/A	0	POUR, LUTS
Li <i>et al.</i> <sup>25</sup>	China	NEU <i>versus</i> PLA	91	N/A	0	Extensive panhysterectomy and pelvic lymph node dissection	N/A	100	POUR, PVRU
ACU versus PLA									
Gao et al. <sup>26</sup>	China	ACU <i>versus</i> PLA	61	54.02	31.15	Arthroscopic knee surgery	84.43	0	POUR, AE
He et al. <sup>27</sup>	China	ACU <i>versus</i> PLA	154	46.44	54.55	Not report	59.62	0	POUR
Yi et al. <sup>28</sup>	China	ACU <i>versus</i> PLA	110	46.90	0	Extensive hysterectomy	N/A	0	POUR
CHO versus PLA									
Uy et al.29	Philippines	CHO <i>versus</i> PLA	106	N/A	74.53	Anorectal surgery	0	0	POUR, AE
Bowers <i>et al</i> . <sup>30</sup>	USA	CHO <i>versus</i> PLA	108	N/A	N/A	Anorectal surgery	N/A	36.11	POUR, AE
Manchana <i>et al.</i> <sup>31</sup>	Thailand	CHO <i>versus</i> PLA	62	48.50	0	Hysterectomy	180	100	POUR, PVRU, UTI, AE
Walsh <i>et al.</i> <sup>32</sup>	USA	CHO <i>versus</i> PLA	100	N/A	0.00	Vaginal hysterectomy with pelvic floor repair	N/A	0	POUR, UTI, AE

(Continued)

# Table 1. (continued)

Author	Country	Intervention	Total number of patients	Mean age (years)	Male (%)	Operation type	Mean operative time (min)	GA(%)	Reported outcome
BENZ versus PLA									
Hershberger <i>et al.</i> <sup>33</sup>	USA	BENZ <i>versus</i> PLA	92	35.52	0	Abdominal surgery, vulva procedure	84.27	96.67	POUR
ASP versus PLA									
Tomaszewski <i>et al.</i> <sup>34</sup>	Poland	ASP <i>versus</i> PLA	226	28.50	64.18	Orthopedic surgery (hip, knee)	N/A	0	POUR
OPI versus PLA									
Cepeda <i>et al</i> . <sup>35</sup>	USA	OPI <i>versus</i> PLA	265	42.13	23.77	Abdominal surgery, orthopedic surgery, thoracic surgery, craniofacial surgery	102.15	100	POUR, AE
Gallo <i>et al.</i> <sup>36</sup>	USA	OPI <i>versus</i> PLA	97	N/A	61.86	Orthopedic surgery (hip, knee, shoulder)	N/A	0	POUR, PVRU
Zand <i>et al.</i> <sup>37</sup>	Iran	OPI <i>versus</i> PLA	67	35.54	100	Fracture lower limbs	121.52	7.46	POUR, AE
ALP versus PLA									
Akkoc <i>et al.</i> <sup>38</sup>	Turkey	ALP <i>versus</i> PLA	180	35.95	N/A	Inguinal, penile, scrotal, perineal surgery	51.49		POUR, AE
Basheer <i>et al.</i> <sup>39</sup>	USA	ALP <i>versus</i> PLA	95	57.36	100	Spine surgery	221.81	0	POUR
Bazzazi <i>et al.</i> 40	Iran	ALP <i>versus</i> PLA	67	69.84	100	Cataract	N/A	100	POUR, AE
Cataldo <i>et al.</i> <sup>41</sup>	USA	ALP <i>versus</i> PLA	49	N/A	N/A	Anorectal surgery	N/A	0	POUR, AE
Chung et al. <sup>42</sup>	Korea	ALP <i>versus</i> PLA	88	66.30	100	Transrectal ultrasound guided prostate biopsy	N/A	0	POUR, PVRU, LUTS
Goldman <i>et al.</i> 43	Israel	ALP <i>versus</i> PLA	102	N/A	N/A	Hernia repair	N/A	16.67	POUR
Gonullu <i>et al.</i> 44	Turkey	ALP <i>versus</i> PLA	156	37.62	100	Hernia repair	N/A	93.59	POUR, AE
Jang et al. <sup>45</sup>	Korea	ALP <i>versus</i> PLA	94	56.50	54.26	Rectal cancer surgery	N/A	100	POUR, PVRU, UTI
Jeong <i>et al.</i> 46	Korea	ALP <i>versus</i> PLA	218	63.50	100	Robotic assisted radical prostatectomy	N/A	100	POUR, PVRU, LUTS
Livne <i>et al.</i> 47	Israel	ALP <i>versus</i> PLA	155	N/A	0	Abdominal and vaginal hysterectomy	N/A	98	POUR

(Continued)

## Table 1. (continued)

Author	Country	Intervention	Total number of patients	Mean age (years)	Male (%)	Operation type	Mean operative time (min)	GA(%)	Reported outcome
Lose <i>et al.</i> <sup>48</sup>	Denmark	ALP <i>versus</i> PLA	41	65.51	0	Anti-continence surgery	N/A	100	POUR, PVRU, UTI, AE
Madani <i>et al.</i> 49	Iran	ALP <i>versus</i> PLA	232	27.64	100	Inguinal, anorectal surgery	51.91	0	POUR, AE
Schubert <i>et al.</i> <sup>50</sup>	USA	ALP <i>versus</i> PLA	131	60.97	100	Total hip arthroplasty, total knee arthroplasty	N/A	18.32	POUR, UTI, LUTS, AE
Mohammadi-Fallah <i>et al.</i> 51	Iran	ALP <i>versus</i> PLA	80	86.33	100	Herniorrhaphy	N/A	18.75	POUR, AE
Peterson <i>et al.</i> <sup>52</sup>	USA	ALP <i>versus</i> PLA	60	65.35	100	Total hip arthroplasty, total knee arthroplasty	N/A	90	POUR, AE
Shaw et al.53	India	ALP <i>versus</i> PLA	48	N/A	100	Herniorrhaphy	N/A	0	POUR
Watson <i>et al.</i> <sup>54</sup>	UK	ALP <i>versus</i> PLA	49	N/A	0	Anti-continence surgery	N/A	0	PVRU, AE
Woo et al. <sup>55</sup>	Australia	ALP <i>versus</i> PLA	70	64.75	100	Herniorrhaphy	N/A	0	POUR, AE
Evron <i>et al.</i> <sup>56</sup>	Israel	ALP <i>versus</i> PLA	60	N/A	0	Elective C section	N/A	0	POUR, UTI, AE
Jianggao <i>et al.</i> 57	China	ALP <i>versus</i> PLA	60	N/A	N/A	Elective abdominal surgery	N/A	0	POUR
NSD versus PLA									
Khan <i>et al.</i> <sup>58</sup>	Pakistan	NSD <i>versus</i> PLA	186	35.83	82.26	Hemorrhoidectomy	N/A	0	POUR
Galán <i>et al.</i> <sup>59</sup>	Spain	NSD <i>versus</i> PLA	157	53.39	33.76	Hemorrhoidectomy	N/A	0	POUR, UTI, AE
ALP versus COMB versus PLA									
Burger <i>et al.<sup>60</sup></i>	Netherlands	ALP <i>versus</i> COMB <i>versus</i> PLA	249	59.64	62.65	Not reported	N/A	63.45	POUR
CHO versus ALP versu	s PLA								
Savona-Ventura <i>et al.</i> 61	USA	CHO <i>versus</i> ALP <i>versus</i> PLA	73	N/A	0	Vaginal hysterectomy, anterior repair	N/A	0	POUR

ACU, acupuncture; AE, adverse event; ALP, alpha-adrenergic antagonist; AMB, ambulation; ASP, anti-spasmodic agent; BENZ, benzodiazepine; CHO, cholinergic drug; COMB, combined intervention; FLU, fluid adjustment; GA, general anesthesia; LUTS, lower urinary tract symptom; NEU, neuromodulation; NSD, non-steroidal anti-inflammatory drug suppository; OPI, opioid antagonist related; PLA, placebo/standard care; POUR, postoperative urinary retention; PVRU, post-void residual urine; UTI, urinary tract infection.



Figure 2. Network maps of (a) POUR, (b) UTI, (c) PVRU, (d) LUTS and (e) AE outcomes.

ACU, acupuncture; AE, adverse event; ALP, alpha-adrenergic antagonist; AMB, ambulation; ASP, anti-spasmodic agent; BENZ, benzodiazepine; CHO, cholinergic drug; COMB, combined intervention; FLU, fluid adjustment; LUTS, lower urinary tract symptom; NEU, neuromodulation; NSD, non-steroidal anti-inflammatory drug suppository; OPI, opioid antagonist related; PLA, placebo/standard care; POUR, post-operative urinary retention; PVRU, post-void residual urine; UTI, urinary tract infection. indirect comparisons was significant; see Table 2. The probability of being the best intervention in lowering the POUR using rankogram (Supplemental Figure 9) and SUCRA demonstrated that NSD was the best prophylaxis (SUCRA 81.6%). The ranking along with SUCRA is shown in Table 3.

No evidence of inconsistency was found (Chisquare = 6.71, p = 0.0816). Sensitivity analysis for the studies with variance greater than 80th percentile showed similar results. Adjusted funnel plots were performed and showed asymmetry, and heterogeneity was suspected (Supplemental Figure 10).

UTI. Data from seven RCTs with four interventions (i.e. PLA, CHO, ALP and NSD) and 652 patients were used; see Figure 2. NSD and ALP were about 0.38 (0.02, 9.15) and 0.57 (0.25, 1.27) times better in lowering UTI whereas CHO was about 1.16 (0.25, 1.27) times higher risk of UTI relative to PLA. However, none of them was statistically significant (Supplemental Table 3). The probability of being the best intervention in UTI prevention was NSD, followed by ALP, PLA and CHO (Table 3 and Supplemental Figure 11). No evidence of inconsistency was found (Chisquare = 0.36, p=0.5498). Adjusted funnel plots were performed and showed symmetry of the plot (Supplemental Figure 12).

PVRU. A network map was constructed using data of 10 RCTs (n=900) with six interventions, that is, PLA, NEU, ACU, CHO, OPI and ALP; see Figure 2. All interventions decreased PVRU by about 8.6–20.2 ml, but none of them was statistically significant; see Supplemental Table 3. The probability of being the best intervention in decreasing PVRU was CHO, followed by NEU, ALP, ACU, OPI and PLA (Table 3 and Supplemental Figure 13). No evidence of inconsistency was found (Chi-square = 0.99, p = 0.3189). Adjusted funnel plots were performed and showed asymmetry of the plot, suggesting possible publication bias (Supplemental Figure 14).

LUTS. Data from four RCTs with three interventions (i.e. PLA, NEU and ALP) and 523 patients were analyzed. A network map was created and no closed loop was found; see Figure 2. LUTS was reported in International Prostate Symptom Score (IPSS) or American Urological Association (AUA) symptoms index. Both NEU and ALP decreased LUTS about 2–0.03, but these were not statistically significant (Supplemental Table 3). The probability of being the best intervention in lowering LUTS was NEU, followed by ALP and PLA (Table 3 and Supplemental Figure 15). Inconsistency assumption was forced analyzed and found no inconsistency (Chi-square = 0.00, p = 0.9766). Adjusted funnel plots were performed and showed symmetry of the plot (Supplemental Figure 16).

AE. The analysis comprised data from 23 RCTs with seven interventions (i.e. PLA, AMB, ACU, CHO, OPI, ALP and NSD) and 2596 patients. A network map was constructed; see Figure 2. For direct comparison, AMB, ACU and OPI were found to decrease the incidence of AEs compared with PLA with the RRs of 0.74 (0.25, 2.17), 0.32 (0.07, 1.49) and 0.72 (0.46, 1.14), respectively. On the other hand, CHO, ALP and NSD increased the AE with RRs of 2.47 (0.57, 10.70), 1.66 (0.99, 2.78) and 1.15 (0.23, 57.40), respectively. However, no statistical significances were found. Regarding indirect comparison, statistically significant increased AE was found in ALP versus ACU and ALP versus OPI (Table 2). The probability of being the least AE occurrence was ACU, followed by OPI, AMB, PLA, NSD, ALP and CHO (Table 3 and Supplemental Figure 17). Inconsistency assumption was forced analyzed and no inconsistency was found (Chi-square = 0.00, p = 0.9446). Adjusted funnel plots were performed and showed symmetry of the plot (Supplemental Figure 18).

Cluster ranking plot and risk-benefit analysis. A cluster rank plot was created according to the SUCRA of benefit and the risks from each intervention pair; see Figure 3. ACU and AMB were in the right upper quadrant, implying that ACU was the best in high efficacy of POUR prevention and PVRU reduction with the least AE, and AMB was the second best for POUR prevention.

In terms of risk-benefit analysis, the risk-benefit plane was constructed and is shown in Supplemental Figure 19. A scatter plot of ACU fell in the right lower quadrant, implying dominant risk-benefit over the placebo, although the shape of the plot was sparse, which represented the low precision of the outcome. Supposing the risk-benefit acceptability threshold was 0.2, which meant two patients with additional AE were acceptable for the 10 additional POUR

Table 2. Relativ	e treatment eff	fects for loweri	ing POUR and	AEs.							
POUR											
AE PLA	0.33 (0.09, 1.22)	0.58 (0.28, 1.20)	0.61 (0.25, 1.51)	0.27 (0.08, 0.86)	0.58 (0.30, 1.12)	0.64 [0.26, 1.62]	0.77 (0.44, 1.33)	0.46 (0.04, 5.92)	0.44 (0.23, 0.84)	0.54 (0.40, 0.73)	0.34 (0.15, 0.80)
I	ASP	1.77 [0.39, 7.99]	1.86 (0.37, 9.20)	0.81 [0.14, 4.76]	1.76 [0.40, 7.70]	1.97 [0.39, 9.82]	2.35 [0.56, 9.78]	1.40 (0.08, 24.87)	1.34 (0.31, 5.82)	1.65 (0.43, 6.38)	1.04 (0.22, 4.98)
1.38 [0.88, 2.17]	I	IdO	1.05 (0.33, 3.36)	0.46 (0.11, 1.83)	0.99 (0.37, 2.66)	1.11 [0.34, 3.59]	1.32 (0.53, 3.30)	0.79 (0.06, 11.29)	0.76 (0.28, 2.00)	0.93 (0.42, 2.05)	0.59 (0.19, 1.79)
I	I	I	NEU	0.44 [0.10, 1.94]	0.95 (0.31, 2.92)	1.06 [0.29, 3.86]	1.26 [0.44, 3.65]	0.75 (0.05, 11.38)	0.72 (0.24, 2.20)	0.89 (0.34, 2.31)	0.56 [0.16, 1.94]
0.87 [0.02, 43.50]	1	0.63 (0.01, 32.35)	I	NSD	2.16 [0.56, 8.36]	2.42 [0.54, 10.78]	2.89 (0.79, 10.59)	1.72 (0.10, 28.79)	1.65 (0.43, 6.32)	2.03 (0.60, 6.84)	1.28 (0.30, 5.45)
I	I	I	I	I	FLU	1.12 [0.36, 3.48]	1.33 (0.56, 3.17)	0.80 (0.06, 11.19)	0.76 (0.30, 1.93)	0.94 (0.45, 1.94)	0.59 (0.20, 1.73)
I	I	I	I	I	1	COMB	1.19 [0.41, 3.46]	0.71 (0.05, 10.78)	0.68 (0.22, 2.10)	0.84 (0.33, 2.11)	0.53 (0.15, 1.84)
0.40 (0.09, 1.75)	I	0.29 (0.06, 1.35)	1	0.47 (0.01, 30.33)	1	1	сно	0.60 (0.04, 8.17)	0.57 (0.24, 1.34)	0.70 (0.38, 1.29)	0.44 (0.16, 1.22)
I	I	I	I	I	I	I	I	BENZ	0.96 (0.07, 13.40)	1.18 (0.09, 15.52)	0.74 (0.05, 11.00)
1.36 [0.46, 4.00]	1	0.98 (0.30, 3.18)	I	1.56 (0.03, 90.21)	1	1	3.35 (0.54, 20.73)	I	AMB	1.23 (0.60, 2.53)	0.78 (0.27, 2.26)
0.60 [0.36, 1.01]	I	0.44 (0.24, 0.81)	I	0.69 (0.01, 35.88)	I	I	1.49 (0.32, 7.05)	ı	0.45 (0.13, 1.48)	ALP	0.63 (0.26, 1.54)
3.10 [0.67, 14.31	I	2.25 [0.46, 11.07]	1	3.56 [0.05, 237.43]	I	I	7.65 (0.92, 63.70)	I	2.28 (0.35, 14.87)	5.13 (1.02, 25.79)	ACU
Results were reg statistical signifi ACU, acupunctur restriction; NSD,	oorted as risk rat cance. Results fc e; AE, adverse ev non-steroidal ar	tios (95% confide or POUR are pres vent; ALP, alpha nti-inflammatory	ence intervals) f sented above in I-adrenergic an' / drug suppositi	rom network m tervention nam tagonist; AMB, ory; OPI, opioid	eta-analysis be es and results 1 ambulation; AS antagonist age	tween each pai or AE are belov P, antispasmod nt; POUR, post-	r of treatments. v intervention n. lic agents; BEN. operative urina	Comparisons a ames. Z, benzodiazepii ry retention.	are read from le ne; CHO, cholin	ft to right. Bold ergic agent; FL	font indicates U, fluid

P Sirisreetreerux, R Wattanayingcharoenchai *et al.* 

## Therapeutic Advances in Urology 13

		5,			5 , ,	,		5		5,5
Rank	POUR		UTI		PVRU		LUTSs		AEs	
	Intervention	SUCRA								
1	NSD	81.6	ALP	75.8	СНО	65.1	NEU	76.9	ACU	88
2	ACU	73.9	NSD	69.6	NEU	57.3	ALP	37.3	OPI	69
3	ASPs	70.3	PLA	35.9	ALP	56	PLA	35.8	AMB	63.7
4	AMB	62.7	СНО	18.7	ACU	51.3			PLA	47.3
5	BENZ	54.8			OPI	43.9			NSD	45.7
6	ALP	49.8			PLA	26.4			ALP	20.7
7	OPI	45.7							СНО	15.6
8	FLU	45.7								
9	NEU	43.6								
10	СОМВ	37.9								
11	CHOs	25.6								
12	PLA	8.4								

Table 3. Treatment ranking by SUCRA from NMA in lowering POUR, UTI, PVRU, LUTS and minimizing AE outcomes accordingly.

ACU, acupuncture; AE, adverse event; ALP, alpha-adrenergic antagonist; AMB, ambulation; ASP, antispasmodic agent; BENZ, benzodiazepine; CHO, cholinergic agent; COMB, combined intervention; FLU, fluid restriction; LUTS, lower urinary tract symptom; NEU, neuromodulation; NMA, network meta-analysis; NSD, non-steroidal anti-inflammatory drug suppository; OPI, opioid antagonist agent; PLA, placebo/standard care; POUR, postoperative urinary retention; PVRU, post-void residual urine; SUCRA, surface under the cumulative ranking curve; UTI, urinary tract infection.



**Figure 3.** A cluster rank plot of benefit in lowering POUR *versus* risk of AE outcomes. Clustered ranking plot showing SUCRA values for highest efficacy of POUR prevention *versus* AE. Right upper quadrant represents the highest efficacy with minimal AE. Each symbol represents a group of treatments in each cluster. ACU, acupuncture; AE, adverse event; ALP, alpha-adrenergic antagonist; AMB, ambulation; CHO, cholinergic drug; LUTS, lower urinary tract symptom; NSD, non-steroidal anti-inflammatory drug suppository; OPI, opioid antagonist related; PLA, placebo/standard care; POUR, post-operative urinary retention; PVRU, post-void residual urine; SUCRA, surface under the cumulative ranking curve; UTI, urinary tract infection.



**Figure 4.** Risk-benefit acceptability curve in lowering post-operative urinary retention at each distinct threshold among six intervention pairs.

The reference values are shown as vertical dashed lines with varying thresholds of 0.10, 0.20 and 0.30.

ACU, acupuncture; ALP, alpha-adrenergic antagonist; AMB, ambulation; CHO, cholinergic drug; NSD, non-steroidal anti-

inflammatory drug suppository; OPI, opioid antagonist related; PLA, placebo/standard care.

preventions, AMB, ACU and ALP were beneficial as the plots were under the threshold line. IRBRs in POUR prevention compared with PLA were 0.0118, 0.7372, -0.1291, 1.4131, -0.0608 and -0.0092 for AMB, ACU, CHO, OPI, ALP and NSD, respectively (Supplemental Table 4). RBACs of each intervention are presented in Supplemental Figure 20 and the estimations of percent chance that the IRBR would be less than risk-benefit acceptability thresholds are summarized in Supplemental Table 5. For clinical implication, RBACs among the interventions were compared in any threshold and these aided in decision making (Figure 4). Given the low acceptable threshold of 0.2, referring to only two acceptable AEs for 10 beneficial outcomes, AMB was the most favorable followed by ACU and ALP, then OPI and NSD. CHO was the least useful intervention. Another example, with the acceptable threshold moved up to 0.3, showed that AMB and ALP were the best with similar benefit, then ACU, NSD, OPI and CHO. In terms of net benefit framework, NCB acceptability curve is shown in Supplemental Figure 21, representing the probability of being the best intervention at distinct threshold.

#### Discussion

POUR is one of the most common complications after surgeries, which results in urinary complications including UTI, patients' discomfort and distress, prolonged hospital stay and increased cost of treatment.8 Currently, there are many interventions reported aiming to prevent POUR. However, there is no standard protocol for preventing this condition because of the uncertainty of the efficacy and the concern about AEs. From the previous studies, there were five meta-analyses published about intervention for POUR prevention, two studies analyzed about methods and timing of catheterization.<sup>64,65</sup> Jackson et al.9 reported not only medical but also non-medical interventions for POUR prevention, including both RCTs and non-RCTs. Clancy et al.66 assessed ALP from five RCTs, particularly in men who underwent inguinal hernia repair. Ghuman et al.67 described 15 RCTs and analyzed the effect of ALP in patients who underwent various kinds of operation. The results of our study were in a similar direction compared with the previous meta-analyses, but we added a higher number of included studies and the network technique allowed for indirect comparison of all treatments available and ranking which treatment should be best in POUR prevention.

In our study, we reported the comprehensive review, first NMA and risk-benefit analysis of all available RCTs in preventing POUR. Our interventions of interest included 11 non-medical and medical interventions with five outcomes. The direct comparisons from NMA showed that AMB, ACU, OPI and ALP significantly reduced POUR. Nevertheless, due to the limited number of the studies and the missing data of covariables, further analysis including heterogeneity, meta-regression and subgroup analysis was not possible in some comparisons. The possible source of heterogeneity might be from the diversity of the studies in terms of populations, the schedule and dosage of the intervention, the definition of outcomes, time of measurement and variation in operation type.

For NMA, our results showed that the best intervention for POUR prevention was NSD followed by ACU, ASP, AMB, BENZ, ALP, OPI, FLU, NEU, COMB, CHO and PLA. However, only NSD, ACU, AMB and ALP were found statistically significant in our analysis compared with PLA. No statistical significance was observed among other interventions from both direct and indirect comparisons. For PVRU, the best intervention was CHO followed by NEU, ALP, ACU, OPI and PLA. Regarding UTI, NSD was considered the best, followed by ALP, PLA and CHO. The ranking of probability of being the best among these outcomes was sparse and not in the same direction among the outcomes. We hypothesized that the reason behind this was probably the number of involved studies in our secondary outcomes, which was quite small, and because of the heterogeneity of the studies. Concerning the AEs, the intervention with the least AEs was ACU, followed by OPI, AMB, PLA, NSD, ALP and CHO. Nevertheless, only ALP versus OPI and ALP versus ACU were statistically significant. Most AEs were reported as mild to moderate, such as nausea, vomiting, pruritus et cetera. One of the obstacles in AE analysis was the difference in reported data. Some studies reported on overall AEs while some studies described each AE individually. Therefore, we chose to analyze as the composite outcome, which counted the most common reported incidence to include into the analysis. As a result, the incidence of AE might be underestimated. Another concern, regarding RoB, was that most of the studies were considered to be high concern and some concern for RoB, so interpretation should be done with caution.

According to the cluster ranking, ACU was found to be the best in POUR and PVRU reduction, followed by AMB. However, the decision for choosing also depended on the willingness to accept any complication, which was different in each threshold. Aiming to simplify for clinical application, we constructed the RBAC (Figure 4), which compared among six interventions. This curve was valuable for intervention selection based upon acceptable threshold. Based on our findings, we recommend patients should have a routine non-medical treatment of early AMB as soon as possible after surgery. Medical intervention, such as alpha-adrenergic antagonists might be considered in male patients undergoing nonneurourological procedure, if the patients and clinicians accept that risk of AEs might occur around 10-15%.

# Strengths and limitations

Our study has several strengths. To our knowledge, this is the first NMA reporting about POUR prevention. In addition, we comprehensively analyzed with a large number of population and included only RCTs which assumed the appropriate study design to evaluate the efficacy of the interventions. Finally, we reported the risk-benefit analysis with the RBAC, which is easy to understand and simple to implement and intuitively provides for the clinician to incorporate the information for decision making. There were some limitations in our study. First, the number of the studies in some comparisons were small and heterogeneous, so the analysis was limited to only some outcomes and also precluded further analysis such as subgroup or sensitivity analysis. It should be noted that the incidence of AEs might be underestimated due to the heterogeneity of outcome reporting and the analysis as the composite outcome. Another concern was regarding RoB, because most of the studies were considered high concern and some concern for RoB, so interpretation should be done with caution. Further well-designed study may be needed to reach further conclusions.

## Conclusions

AMB, ACU, OPI, ALP and NSD might be useful in prevention of POUR. However, when both

efficacy and AEs were considered with the acceptable threshold of side effect of 15%, ALP might be the best, followed by OPI, NSD and CHO.

#### Acknowledgements

We thank Stephen Pinder, a native-Englishspeaking advisor to the Department of Clinical Epidemiology and Biostatistics, Faculty of Medicine Ramathibodi Hospital, Mahidol University, for proofreading our manuscript.

#### **Author contributions**

PS did the data extraction, analyzed and interpreted the data and was a major contributor in writing the manuscript. RW extracted the data. SR analyzed and helped with interpretation of data. OP analyzed and helped with interpretation of data. PN has made substantial contributions to the conception, design of the work and did the data extraction. AT has made substantial contributions to the conception and revised the manuscript. All authors read and approved the final manuscript.

#### **Conflict of interest statement**

The authors declare that there is no conflict of interest.

## Funding

This research received no specific grant from any funding agency in the public, commercial, or notfor-profit sectors.

#### **ORCID** iDs

Pokket Sirisreetreerux (D) https://orcid.org/0000

-0002-6900-8356 Pawin Numthavaj 0002-1369-2945

https://orcid.org/0000-

## Supplemental material

Supplemental material for this article is available online.

#### References

- 1. Hansen BS, Soreide E, Warland AM, *et al.* Risk factors of post-operative urinary retention in hospitalised patients. *Acta Anaesthesiol Scand* 2011; 55: 545–548.
- Wu AK, Auerbach AD and Aaronson DS. National incidence and outcomes of postoperative urinary retention in the Surgical Care Improvement Project. *Am J Surg* 2012; 204: 167–171.

- Cha YH, Lee YK, Won SH, *et al.* Urinary retention after total joint arthroplasty of hip and knee: systematic review. *J Orthop Surg* 2020; 28: 2309499020905134.
- Chong C, Kim HS, Suh DH, *et al.* Risk factors for urinary retention after vaginal hysterectomy for pelvic organ prolapse. *Obstet Gynecol Sci* 2016; 59: 137–143.
- Darrah DM, Griebling TL and Silverstein JH. Postoperative urinary retention. *Anesthesiol Clin* 2009; 27: 465–484, Table of Contents.
- Choi S, Mahon P and Awad IT. Neuraxial anesthesia and bladder dysfunction in the perioperative period: a systematic review. *Can J Anaesth* 2012; 59: 681–703.
- Kamphuis ET, Ionescu TI, Kuipers PW, et al. Recovery of storage and emptying functions of the urinary bladder after spinal anesthesia with lidocaine and with bupivacaine in men. *Anesthesiology* 1998; 88: 310–316.
- Baldini G, Bagry H, Aprikian A, et al. Postoperative urinary retention: anesthetic and perioperative considerations. *Anesthesiology* 2009; 110: 1139–1157.
- Jackson J, Davies P, Leggett N, et al. Systematic review of interventions for the prevention and treatment of postoperative urinary retention. BJS Open 2019; 3: 11–23.
- Sterne JAC, Savovic J, Page MJ, et al. RoB
   2: a revised tool for assessing risk of bias in randomised trials. *BMJ* 2019; 366: 14898.
- Higgins JP, Thompson SG, Deeks JJ, et al. Measuring inconsistency in meta-analyses. BMJ 2003; 327: 557–560.
- Peters JL, Sutton AJ, Jones DR, et al. Contourenhanced meta-analysis funnel plots help distinguish publication bias from other causes of asymmetry. J Clin Epidemiol 2008; 61: 991–996.
- White IR. Network meta-analysis. Stata Journal 2015; 15: 951–985.
- Black WC. The CE plane: a graphic representation of cost-effectiveness. *Med Decis Making* 1990; 10: 212–214.
- 15. Lazo-Langner A, Rodger MA, Barrowman NJ, *et al.* Comparing multiple competing interventions in the absence of randomized trials using clinical risk–benefit analysis. *BMC Med Res Methodol* 2012; 12: 3.
- Stinnett AA and Mullahy J. Net health benefits: a new framework for the analysis of uncertainty in cost-effectiveness analysis. *Med Decis Making* 1998; 18(2 Suppl.): S68–S80.

- Hansen AB and Olsen KS. The number of in-out catheterisations is reduced by mobilising the postoperative patient with bladder needs to the toilet in the recovery room: a randomised clinical trial. *Eur J Anaesthesiol* 2015; 32: 486–492.
- Kim SB, Lee IO, Kong MH, et al. Early ambulation reduces the incidence of urinary retention after spinal anesthesia for benign anorectal surgery. Korean J Anesthesiol 1999; 37: 1001–1006.
- Kim SH, Lee IO, Kim DH, et al. The effects of early ambulation on urinary retention and postdural puncture headache following anal operations: a prospective randomized study. *J Korean Soc Coloproctol* 1999; 15(3): 179–185.
- Bailey HR and Ferguson JA. Prevention of urinary retention by fluid restriction following anorectal operations. *Dis Colon Rectum* 1976; 19: 250–252.
- Kozol RA, Mason K and McGee K. Postherniorrhaphy urinary retention: a randomized prospective study. J Surg Res 1992; 52: 111–122.
- Young LC, Cheol KH and Hee LD. A prospective study on the relationship between postoperative urinary retention and amount of infused fluid during surgery of benign anal diseases under spinal anesthesia. *J Korean Soc Coloproctol* 1999; 15: 357–361.
- Orbey BC, Alanoglu Z, Yilmaz AA, et al. Do we still need to restrict preoperative fluid administration in ambulatory anorectal surgery under spinal anaesthesia? *Tech Coloproctol* 2009; 13: 35–40.
- 24. Butwick A, Carter P and Dolin SJ. A pilot study of the effect of the Queen's Square external bladder stimulator on urinary retention after knee replacement surgery. *Anaesthesia* 2003; 58: 587–591.
- 25. Li H, Zhou CK, Song J, *et al.* Curative efficacy of low frequency electrical stimulation in preventing urinary retention after cervical cancer operation. *World J Surg Oncol* 2019; 17: 141.
- Gao Y, Zhou X, Dong X, *et al.* Electroacupuncture for bladder function recovery in patients undergoing spinal anesthesia. *Evid Based Complement Alternat Med* 2014; 2014: 892619.
- He B, Pan F and Zhang S. [Efficacy of acupuncture intervention on urinary retention after spinal anesthesia]. *Zhongguo Zhen Jiu* 2015; 35: 209–211.

- Yi WM, Li JJ, Lu XM, et al. [Effects of electroacupuncture on urinary bladder function after radical hysterectomy]. *Zhongguo Zhen Jiu* 2008; 28: 653–655.
- Uy BG, Yu BM, Torillo ML, et al. Bethanechol chloride for the prevention of postoperative urinary retention after anal surgery under spinal anesthesia. *Phil J Surg* Spec 2011; 66(2): 68–73.
- Bowers FJ, Hartmann R, Khanduja KS, *et al.* Urecholine prophylaxis for urinary retention in anorectal surgery. *Dis Colon Rectum* 1987; 30: 41–42.
- Manchana T and Prasartsakulchai C. Bethanechol chloride for the prevention of bladder dysfunction after radical hysterectomy in gynecologic cancer patients: a randomized controlled trial study. *Int J Gynecol Cancer* 2011; 21: 730–736.
- Walsh J and Bonnar J. Distigmine bromide (Ubretid) for the prevention of urinary retention after vaginal hysterectomy. *J Obstet Gynaecol Br Commonw* 1972; 79: 377–378.
- Hershberger JM and Milad MP. A randomized clinical trial of lorazepam for the reduction of postoperative urinary retention. *Obstet Gynecol* 2003; 102: 311–316.
- 34. Tomaszewski D and Balkota M. Intramuscular administration of drotaverine hydrochloride decreases both incidence of urinary retention and time to micturition in orthopedic patients under spinal anesthesia: a single blinded randomized study. *Biomed Res Int* 2015; 2015: 926953.
- Cepeda MS, Alvarez H, Morales O, *et al.* Addition of ultralow dose naloxone to postoperative morphine PCA: unchanged analgesia and opioid requirement but decreased incidence of opioid side effects. *Pain* 2004; 107: 41–46.
- Gallo S, DuRand J and Pshon N. A study of naloxone effect on urinary retention in the patient receiving morphine patient-controlled analgesia. *Orthop Nurs* 2008; 27: 111–115.
- Zand F, Amini A, Asadi S, *et al.* The effect of methylnaltrexone on the side effects of intrathecal morphine after orthopedic surgery under spinal anesthesia. *Pain Pract* 2015; 15: 348–354.
- Akkoc A, Aydin C, Topaktas R, et al. Prophylactic effects of alpha-blockers, Tamsulosin and Alfuzosin, on postoperative urinary retention in male patients undergoing urologic surgery under spinal anaesthesia. Int Braz J Urol 2016; 42: 578–584.

- Basheer A, Alsaidi M, Schultz L, *et al.* Preventive effect of tamsulosin on postoperative urinary retention in neurosurgical patients. *Surg Neurol Int* 2017; 8: 75.
- Bazzazi N, Bahar HM, Asadi H, et al. Prophylactic tamsulosin in cataract surgery under general anesthesia for preventing urinary retention: a randomized clinical trial. *Int Eye Sci* 2014; 14: 1–3.
- Cataldo PA and Senagore AJ. Does alpha sympathetic blockade prevent urinary retention following anorectal surgery? *Dis Colon Rectum* 1991; 34: 1113–1116.
- 42. Chung SJ, Jung SI, Ryu JW, *et al.* The preventive effect of tamsulosin on voiding dysfunction after prostate biopsy: a prospective, open-label, observational study. *Int Urol Nephrol* 2015; 47: 711–715.
- 43. Goldman G, Leviav A, Mazor A, *et al.* Alphaadrenergic blocker for posthernioplasty urinary retention. Prevention and treatment. *Arch Surg* 1988; 123: 35–36.
- Gonullu NN, Dulger M, Utkan NZ, et al. Prevention of postherniorrhaphy urinary retention with prazosin. Am Surg 1999; 65: 55–58.
- Jang JH, Kang SB, Lee SM, *et al.* Randomized controlled trial of tamsulosin for prevention of acute voiding difficulty after rectal cancer surgery. *World J Surg* 2012; 36: 2730–2737.
- 46. Jeong IG, You D, Yoon JH, et al. Impact of tamsulosin on urinary retention following early catheter removal after robot-assisted laparoscopic radical prostatectomy: a prospective randomized controlled trial. Int J Urol 2014; 21: 164–168.
- Livne PM, Kaplan B, Ovadia Y, et al. Prevention of post-hysterectomy urinary retention by alphaadrenergic blocker. Acta Obstet Gynecol Scand 1983; 62: 337–340.
- Lose G and Lindholm P. Prophylactic phenoxybenzamine in the prevention of postoperative retention of urine after vaginal repair: a prospective randomized double-blind trial. *Int J Gynaecol Obstet* 1985; 23: 315–320.
- Madani AH, Aval HB, Mokhtari G, et al. Effectiveness of tamsulosin in prevention of post-operative urinary retention: a randomized double-blind placebo-controlled study. Int Braz J Urol 2014; 40: 30–36.
- 50. Schubert MF, Thomas JR, Gagnier JJ, *et al.* The AAHKS clinical research award: prophylactic tamsulosin does not reduce the risk of urinary retention following lower extremity arthroplasty:

a double-blinded randomized controlled trial.  $\mathcal{J}$ Arthroplasty 2019; 34(7S): S17–S23.

- Mohammadi-Fallah M, Hamedanchi S and Tayyebi-Azar A. Preventive effect of tamsulosin on postoperative urinary retention. *Korean J Urol* 2012; 53: 419–423.
- 52. Petersen MS, Collins DN, Selakovich WG, et al. Postoperative urinary retention associated with total hip and total knee arthroplasties. a publication of the association of bone and joint surgeons<sup>®</sup>. CORR<sup>®</sup>. 1991; 269: 102–108.
- 53. Shaw MK and Pahari H. The role of perioperative use of alpha-blocker in preventing lower urinary tract symptoms in high risk patients of urinary retention undergoing inguinal hernia repair in males above 50 years. *J Indian Med Assoc* 2014; 112: 13–14, 6.
- 54. Watson AJ, Currie I and Jarvis GJ. A prospective placebo controlled double blind randomised study to investigate the use of indoramin to prevent post-operative voiding disorders after surgical treatment for genuine stress incontinence. Br J Obstet Gynaecol 1999; 106: 270–272.
- 55. Woo HH and Carmalt HL. A placebo controlled double blind study using perioperative prazosin in the prevention of urinary retention following inguinal hernia repair. *Int Urol Nephrol* 1995; 27: 557–562.
- 56. Evron S, Magora F and Sadovsky E. Prevention of urinary retention with phenoxybenzamine during epidural morphine. Br Med J 1984; 288: 190.
- Jianggao T, Qing X and Tiehu Y. Prevention and release of epidural-morphine-induced urinary retention with phenoxybenzamine and neostigmine. *Acta Acad Med Sin* 2000; 22(6): 595–596.
- Khan M, Saeed S, Ali A, *et al.* Use of diclofenac sodium post Milligan Morgan hemorrhoidectomy reduces the risk of post operative urinary retention. *Rawal Med J* 2015; 40: 84–87.
- Galán CP, Andrés Mujika J, Luis Elósegui J, et al. Una sola dosis de diclofenaco intrarrectal reduce la retención urinaria tras la cirugía de las hemorroides. Resultados de un estudio clínico controlado y aleatorizado. *Cirugía Española* 2008; 83: 301–305.
- Burger DH, Kappetein AP, Boutkan H, *et al.* Prevention of urinary retention after general surgery: a controlled trial of carbachol/diazepam versus alfusozine. *J Am Coll Surg* 1997; 185: 234–236.

- Savona-Ventura C, Grech ES and Saliba I. Pharmacological measures to prevent postoperative urinary retention; a prospective randomized study. *Eur J Obstet Gynecol Reprod Biol* 1991; 41: 225–229.
- Hahn KS, Ok LI and Hee KD. The effects of early ambulation on urinary retention and postdural puncture headache following anal operations: a prospective randomized study. *J Korean Soc Coloproctol* 1999; 15: 179–185.
- Tang J, Xu Q and Ye T. Prevention and release of epidural-morphine-induced urinary retention with phenoxybenzamine and neostigmine. *Zhongguo Yi Xue Ke Xue Yuan Xue Bao* 2000; 22: 595–596.

Visit SAGE journals online journals.sagepub.com/ home/tau

SAGE journals

64. Menshawy A, Ghanem E, Menshawy E, *et al.* Early versus delayed removal of indwelling urinary catheter after elective caesarean delivery: systematic review and meta-analysis of randomized controlled trials. *J Matern Fetal Neonatal Med* 2020; 33: 2818–2825.

- 65. Zhang W, Liu A, Hu D, *et al.* Indwelling versus intermittent urinary catheterization following total joint arthroplasty: a systematic review and meta-analysis. *PLoS One* 2015; 10: e0130636.
- 66. Clancy C, Coffey JC, O'Riordain MG, et al. A meta-analysis of the efficacy of prophylactic alpha-blockade for the prevention of urinary retention following primary unilateral inguinal hernia repair. Am J Surg 2018; 216: 337–341.
- 67. Ghuman A, de Jonge SW, Dryden SD, *et al.* Prophylactic use of alpha-1 adrenergic blocking agents for prevention of postoperative urinary retention: a review & meta-analysis of randomized clinical trials. *Am J Surg* 2018; 215: 973–979.