Re-evaluating the relationship between female social bonds and

infant survival in wild baboons

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1 ABSTRACT

2 Over the past few decades studies have provided strong evidence that the robust links between the 3 social environment, health, and survival found in humans also extend to non-human social animals. 4 A number of these studies emphasize the early life origins of these effects. For example, in several 5 social mammals, more socially engaged mothers have infants with higher rates of survival 6 compared to less socially engaged mothers, suggesting that positive maternal social relationships 7 causally improve offspring survival. Here we show that the relationship between infant survival 8 and maternal sociality is confounded by previously underappreciated variation in female social 9 behavior linked to changes in reproductive state and the presence of a live infant. Using data from 10 a population of wild baboons living in the Amboseli basin of Kenya – a population where high 11 levels of maternal sociality have previously been linked to improved infant survival - we find that 12 infant- and reproductive state-dependent changes in female social behavior drive a statistically 13 significant relationship between maternal sociality and infant survival. After accounting for these 14 state-dependent changes in social behavior, maternal sociality is no longer positively associated 15 with infant survival in this population. Our results emphasize the importance of considering 16 multiple explanatory pathways-including third-variable effects-when studying the social 17 determinants of health in natural populations.

18

19 **INTRODUCTION**

20 In social species, social interactions play a major role in determining an animal's access to 21 resources (Wenzel & Pickering, 1991; Ranta et al., 1993; Wright et al., 2001; Silk, 2007), exposure 22 to pathogens (Kappeler et al., 2015), protection against predators (Sterck et al., 1997; Wright et 23 al., 2001; Krause & Ruxton, 2002; Silk, 2007), and adoption of novel behaviors (Burkart, 2017). 24 Consistent with this idea, the social environment has been linked to health and survival outcomes 25 in a wide variety of social mammals, including humans (Silk et al., 2010; Archie et al., 2014; Brent 26 et al., 2017; Ellis et al., 2019; Campos et al., 2020; Snyder-Mackler et al., 2020). For the dependent 27 young of many species, the social environment is primarily determined by the social bonds of their 28 mothers. For this reason, maternal social traits have been proposed to influence offspring survival, 29 with evidence supporting this link presented in non-human primates (Silk et al., 2003a; Silk et al., 30 2009; Kalbitzer et al., 2017; McFarland et al., 2017; Schneider-Crease et al., 2022; Blersch et al.,

2023), dolphins (Frère et al., 2010), sheep (Vander Wal et al., 2015), and horses (Cameron et al.,
2009).

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34 Because of their long period of maternal dependence and close evolutionary relationship to 35 humans, non-human primates have been a particular focus of studies linking maternal sociality to 36 offspring survival. In the Amboseli baboon population of southern Kenya (an admixed population 37 of yellow, Papio cynocephalus, and anubis, P. anubis, baboons) females that have the strongest 38 social bonds show the highest relative infant survival over their lifetime (Silk et al., 2003a). In a 39 long-term study of chacma baboons in Botswana (P. ursinus), offspring of females with stronger 40 social bonds also live significantly longer lives (Silk et al., 2009). Similarly, in the De Hoop 41 chacma baboons of South Africa, baboon infants whose mothers have many weak social bonds are 42 also more likely to survive the first 12-months of life than infants whose mothers have fewer weak 43 social bonds (McFarland et al., 2017) and evidence in vervet monkeys (*Chlorocebus pygerythrus*) 44 suggests infant survival increases with the number of maternal spatial partners (Blersch et al., 45 2023). These studies suggest a positive role of maternal sociality in bolstering offspring survival. 46 However, Schneider-Crease et al. (2022) recently revealed no significant relationship between 47 maternal sociality and offspring survival in kinda baboons (P. kindae) and in white-faced 48 capuchins (*Cebus imitator*), the offspring of highly social females exhibit higher survivorship than 49 those of less social females during socially stable periods, but lower survivorship during the less 50 stable periods surrounding alpha male turnover (Kalbitzer et al., 2017). Therefore, associations 51 between maternal sociality and offspring survival within primates appear to vary as a function of 52 species, population, and/or prevailing demographic conditions.

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54 Notably, all analyses of the relationship between maternal social behavior and infant survival 55 face a key challenge: the amount of time a mother has a living infant may itself drive patterns of 56 female sociality. For example, in a number of primate species, conspecific adult females are 57 attracted to young infants (Seyfarth, 1976; Altmann, 1980; Small, 1982; Silk et al., 2003b; Tiddi 58 et al., 2010; Dunayer & Berman, 2018). This effect could cause mothers with surviving infants to 59 appear more social than mothers whose infants die, simply because they have a socially attractive 60 infant for a longer time period (Barrett et al., 2007). Under this scenario, infant survival drives 61 estimates of female sociality, instead of female sociality driving infant survival, providing an

62 explanation that is plausible regardless of whether maternal sociality is measured over the lifetime 63 (e.g., Silk et al., 2003a) or over fixed yearly intervals (e.g., Kalbitzer et al., 2017, Silk et al., 2009; 64 McFarland et al., 2017; Schneider-Crease et al., 2022). Similarly, the relationships females have 65 with males may depend on patterns in infant survival. In many primates the death of an infant is 66 followed by the rapid resumption of sexual cycling (the state during which grooming and 67 proximity to males peak) and males may also be socially attracted to young neonates (e.g., Nguyen 68 et al., 2009; Baniel et al., 2016). Some approaches to avoiding infant-dependent variation in female 69 social relationships in such analyses include discarding maternal social interactions that occur 70 when infants are very young (e.g., less than 100 days: Silk et al., 2009) or measuring maternal 71 sociality before an infant is born (Blersch et al., 2023). However, these approach have limitations. 72 The first approach does not elimate infant-dependent variation in maternal sociality that occurs 73 after the discarded time window (e.g., cycling resumption following an infant's death). Moreover, 74 both approaches by design eliminate the social interactions that are likely to be the most 75 consequential for infant survival, possibly obscuring the true causal effects of maternal social 76 relationships on offspring survival.

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78 Here, we directly address the potential influence of infant-driven variation in maternal 79 social behavior by measuring and adjusting for the magnitude of this potential confound in a study 80 of the baboons living in the Amboseli basin of southern Kenya – a population where higher levels 81 of maternal lifetime sociality have previously been linked to improved infant survival (Silk et al., 82 2003a). We use a previously established measure of maternal social behavior (maternal social 83 connectedness, or SCI; Archie et al., 2014, see Methods) to demonstrate that a female's social 84 interactions with both adult females and adult males are strongly dependent upon her reproductive 85 state along with the age and survival status of any current infant. Next, we demonstrate how the 86 relationships between a female's reproductive state, infant age/presence, and her rates of social 87 interactions produce statistically significant associations between maternal SCI and infant survival 88 that can change in both direction and magnitude depending upon the time interval over which SCI 89 is measured. After accounting for infant- and reproductive state-driven variation in maternal 90 sociality, we find no compelling evidence that sociality positively predicts infant survival in the 91 present analysis. Finally, we demonstrate that the same confounds attenuate the original effect

92 reported in Silk et al. (2003a) and present some post-hoc analyses assessing the eco-evolutionary

93 consequences of a trade-off between maternal sociality's effect on adult versus offspring survival.

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96 <u>RESULTS</u>

97 <u>MATERNAL SOCIAL BEHAVIOR STRONGLY DEPENDS ON REPRODUCTIVE STATE</u> 98 <u>AND INFANT PRESENCE</u>

99 Social connectedness of adult female baboons to other females (SCI-F, a non-dyadic measure of 100 the quantity/overall rate of grooming, the primary affiliative social behavior in baboons) tends to 101 peak for mothers with infants in the first one to six months of life (Fig. 1A; Table S1). This pattern 102 is consistent with the idea that having a young living infant increases a mother's social interactions 103 with females e.g., by attracting other females to groom with her. Females do not experience this 104 peak in grooming interactions if their infants die in the first year of life (Fig. S1A).

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106 Social connectedness to males (SCI-M) peaks during the pre-pregnancy period when 107 females are cycling, and steadily declines from pregnancy through the first year of a surviving 108 infant's life (Fig. 1B; Table S2). If an infant dies, the mother's connectedness to males returns to 109 pre-pregnancy levels as she resumes cycling (Fig. S1B).

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111 Associations between the presence of a live infant, reproductive state, and maternal SCI 112 could confound an apparent effect of maternal sociality on infant survival. The same concern 113 affects analyses using alternative measures of social integration such as the dyadic sociality index 114 (DSI; Silk et al., 2006; Silk et al., 2013) and the composite sociality index (CSI, which combines 115 data on grooming and proximity to individuals of both sexes; Silk et al., 2003a), which are also 116 affected by reproductive state and infant presence (albeit less so for CSI; Figs. S2 and S3). These 117 patterns indicate that the association between infants, reproductive state, and maternal sociality are 118 not unique to SCI, our specific social measure of interest. 119

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121 Figure 1: Comparison of A) social connectedness to females (SCI-F) and B) social connectedness 122 to males (SCI-M) for mothers with infants who survived to one year across four six-month time 123 intervals (six months prior to pregnancy, pregnancy, one to six months following birth, and seven 124 to 12 months following birth). Black diamonds show the median for each time interval. Each series 125 of connected dots represents a single infant ID for 257 unique infant births (n=146 unique mothers) 126 where behavioral data for the mother was available for all four time intervals and the mother did 127 not have a young infant (less than six months old) from a previous pregnancy during the pre-128 pregnancy period.

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130 <u>THE RELATIONSHIP BETWEEN MATERNAL SOCIALITY AND INFANT SURVIVAL</u> 131 <u>VARIES DEPENDING ON WHEN MATERNAL SOCIALITY IS MEASURED</u>

132 After confirming that mothers experience patterns of social behavior that depend on infants and 133 reproductive state, we investigated the effects of this relationship on the association between 134 maternal social behavior and infant survival. To do so, we analyzed the link between the 135 probability of infant mortality within the first year of life and, in separate models, social 136 connectedness to females (SCI-F) and social connectedness to males (SCI-M) measured during 137 four different time intervals. These time intervals included (i) the fixed six-month window after an 138 infant's conception (i.e., the pregnancy period in baboons, where gestation lasts an average of six 139 months; Altmann et al., 1977), (ii) the fixed six-month window beginning with the infant's live 140 birth, (iii) the fixed six-month window beginning seven months after an infant's birth, and (iv) a

141 'shifting time window' that represents the six months prior to each infant's death or – if it survived 142 – its first birthday. The shifting time window was designed to exclude any maternal social behavior 143 occurring after (and possibly as a result of) an infant's death while also capturing key parts of the 144 infant's life which were likely critical to its survival. The other time intervals occur over the same 145 fixed time period relative to birth for all mothers, regardless of infant outcome, and thus more 146 closely parallel approaches taken in previous studies connecting maternal sociality to infant 147 survival (e.g. Silk et al., 2009; McFarland et al., 2017; Blersch et al., 2023).

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149 We used binomial Generalized Linear Models (GLMs) implemented in R to test the 150 relationship between SCI-F and SCI-M measured over each of the four time intervals and infant 151 survival, controlling for other variables that could influence infant survival (maternal social rank, 152 age, parity, group size, and sexual receptivity; see Methods). In these models the outcome was 153 scored as 1 (if the infant died within one year) or 0 (if the infant survived). The total number of 154 infants included in these analyses varied between 824 and 923 depending on the time interval of 155 interest. This variation in sample sizes was a result of gaps in behavioral data for some mothers 156 during some time intervals when behavioral sampling was constrained (e.g., if groups ranged 157 outside of the study area or demographic events decreased sampling opportunities; see sample 158 sizes reported in Tables S3 to S10). In all data sets, approximately 20% of infants died before their 159 first birthday (range 19 - 24%). For example, in the shifting time window case, 873 infants were 160 included in the analysis, 203 of whom died before reaching age one.

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162 The relationship between SCI-F and infant survival varied dramatically in direction and 163 magnitude depending when SCI-F was measured. SCI-F showed no statistically significant 164 relationship with infant survival when measured over the pregnancy period (coefficient=-0.125; 165 odds ratio (OR)=0.882; p=0.109; Fig. 2; Table S3) or seven to 12 months after an infant's birth 166 (coef=0.109; OR=1.116; p=0.240; Fig. 2; Table S5). However, when SCI-F was measured over 167 the fixed six-month window directly following a live birth, SCI-F positively predicted survival: 168 infants born to mothers with higher SCI-F experienced lower mortality rates compared to infants 169 born to mothers with a lower SCI-F (coef=-0.208; OR=0.812; p=0.013; Fig. 2; Table S4). In 170 contrast, when measured over the shifting time window, which excludes all time periods after an 171 infant's death, higher SCI-F negatively predicted survival: infants born to mothers with higher

SCI-F experienced higher mortality rates compared to infants born to mothers with a lower SCI-F
(coef=0.374; OR=1.454; p=<0.001; Fig. 2; Table S6).

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175 These contrasting results are consistent with what should be expected if adult females' 176 attraction to infants, and hence their mothers (Fig. 1A) explain observed associations between 177 infant survival and maternal sociality. Specifically, measuring SCI-F during the fixed six-month 178 time window after birth means that females whose infants survive the entire six-month period 179 appear more social than females whose infants die during that period. Using a shifting time 180 window, however, would cause females whose infants die to appear more social than females 181 whose infants survive because in cases of infant death the mother's SCI-F tends to be measured 182 earlier in the infant's life, when infants are most likely to attract social attention.

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184 Similar to SCI-F, the link between SCI-M and infant mortality depended upon the time 185 interval over which SCI-M was measured (Fig. 2). Specifically, SCI-M did not predict infant 186 survival when measured during pregnancy (coef=0.005; OR=1.005; p=0.952; Fig. 2; Table S7). 187 When measured during the six months following birth, higher SCI-M was associated with 188 somewhat lower infant mortality (coef=-0.229; OR=0.795; p=0.048; Fig. 2; Table S8). On the 189 other hand, when measured during the seven to 12 months following birth or using the shifting 190 time window, higher SCI-M was associated with considerably higher infant mortality (coef=0.656; 191 OR=1.928; p=<0.001; Fig. 2; Table S9 and coef=0.607; OR=1.834; p=<0.001; Fig. 2; Table S10, 192 respectively).

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194 As with the SCI-F results, results with SCI-M were consistent with the idea that the 195 reproductive state-dependent nature of female relationships with males affects the apparent 196 relationship between SCI-M and infant survival. Specifically, high levels of sociality with males 197 may be negatively associated with infant survival when measured seven to 12 months after birth 198 because mothers with surviving infants of this age typically have not yet resumed sexual cycling 199 (Gesquiere et al., 2018) and are less social with adult males than females in any other reproductive 200 state (Fig. 1B). In contrast, mothers whose infants die resume sexual cycling and return to their 201 pre-pregnancy levels of social interactions with adult males soon after infant death, resulting in 202 high levels of social interactions with males seven to 12 months after their (non-surviving) infants'

- 203 births (Fig. S1B). Similarly, when using the shifting time window, SCI-M for mothers whose
- 204 infants survive is measured when those mothers are least likely to be interacting with adult males
- 205 (i.e., when their infants are older but the mothers are not yet cycling).
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208Figure 2: Infant survival as a function of maternal SCI-F and SCI-M measured over different time209intervals relative to infant birth; a coefficient (i.e., the natural logarithm of the odds ratio) = 0210(shown by vertical dashed line) indicates no effect of the social measure of interest on infant211survival. See Tables S3 to S10 for full model results and sample sizes. $+p \le 0.1$; * p < 0.05; ** p < 0.01212213INFANT-DEPENDENT AND REPRODUCTIVE STATE-DEPENDENT TRENDS IN

214 MATERNAL SOCIAL BEHAVIOR EXPLAIN THE RELATIONSHIP BETWEEN INFANT

215 <u>SURVIVAL AND MATERNAL SOCIALITY MEASURED OVER MULTIPLE TIME</u>

- 216 <u>INTERVALS</u>
- 217

218 To determine whether infant-dependent and reproductive state-dependent changes in 219 female social behavior could completely account for the associations between maternal sociality 220 and infant survival in our observed data sets, we used a data randomization procedure to force 221 independence between maternal sociality and infant survival. In this data randomization (hereafter 222 the 'Randomization of SCI Values'), maternal SCI-F and SCI-M values were assigned randomly 223 to the infant outcomes used in our actual analyses, removing any possible influence of maternal 224 sociality on infant survival. These randomly assigned SCI values were then changed systematically 225 over time, as a function of time since infant's birth, depending only on whether and when an infant 226 died (see further explanation below). If the confounds we described above were sufficient to 227 account for the observed links between maternal social behavior and infant survival, then the 228 randomized SCI data should drive effects comparable to those obtained with the observed data.

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230 To randomize SCI-F values, we randomly sampled SCI-F trajectories, with replacement, 231 from the set of trajectories shown in Figure 1A (where infant death never occurred), and matched 232 them one-by-one to infant-mother pairs in the true data set (randomization procedure visualized in 233 Fig. S4). For all infants in the true data set we directly substituted the randomly matched SCI-F 234 value during pregnancy for the real SCI-F value during pregnancy. For cases in the real data set 235 where the infant survived to 12 months of age, we also directly substituted the randomly matched 236 SCI-F value in the first six-month interval and the seven to twelve month interval following birth. 237 For cases in the real data where the infant died, we followed the same procedure as for surviving 238 infants with one exception: for all months following the infant's death we assigned the maternal 239 SCI-F value for pre-pregnancy from the randomly sampled SCI-F trajectory (recall that a female's 240 pre-pregnancy SCI-F values closely match her SCI-F values after the loss of her infant, Fig. S1). 241 We repeated this procedure across all observed infant outcomes 1,000 times, resulting in 1,000 242 data sets with randomized SCI-F values.

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We randomized SCI-M values as described above for SCI-F, but with an additional control for patterns in sexual cycling (since sexual cycling attracts social attention from males; see Fig. S5) using data from Fig. S6 (randomization procedure visualized in Fig. S7). For all months after birth where a female had a live infant and had not resumed cycling, we substituted the matching SCI-M value from the randomly sampled female trajectory (i.e., representing the period when the

249 randomly sampled female's infant was alive and its mother had not yet resumed cycling). For all 250 months after birth where a female had a live infant and had resumed sexual cycling, we substituted 251 the SCI-M values from the randomly sampled trajectory during pre-pregnancy. For cases in which 252 an infant died, we assigned the mother the randomized trajectory SCI-M value in pre-pregnancy, 253 starting two months after the infant's death and lasting for four months (reflecting rapid resumption 254 of cycling in baboons after infant death and mean cycling length before the next conception: Zipple 255 et al., 2017). After four months we assigned the pregnancy SCI-M value from the randomized 256 trajectory.

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We averaged the randomized monthly SCI values described above to get a mean SCI-F and SCI-M value for each infant's mother during pregnancy, six months following birth, seven to 12 months following birth, and the shifting time window. We then ran parallel binomial GLMs on the randomized data sets to estimate the association between SCI measured over each time interval and infant survival, controlling for the same fixed effects as in analyses with observed data. The set of infant outcomes for each time interval matched the sample of outcomes for complementary analyses with real data in Tables S3 to S10.

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266 After repeating this process across 1,000 randomized data sets, we found that in many cases 267 the distribution of effect sizes generated from analyses with randomized data did not center on 268 zero, indicating that infant-dependent and reproductive state-dependent changes in maternal social 269 behavior indeed present a detectable confound over these time intervals (Fig. 3; Table S11). 270 Furthermore, for both SCI-F and SCI-M, we found that all of the observed effect sizes fell well 271 within the distribution of effect sizes generated from randomized data sets, indicating the observed 272 effect sizes could be entirely explained by infant-dependent and reproductive state-dependent 273 changes in maternal social behavior (Fig. 3; Table S11). In several cases the observed effects fell 274 within the tails of the distributions (Fig. 3A, 3B, 3C, 3F, and 3G) making these effects less likely 275 to be entirely explained by the confound than cases where effects fell within the center of the 276 distribution. Among cases where the observed coefficients fell within one tail of a distribution, 277 effects generated from SCI-F and SCI-M measured seven to 12 months after birth were the least 278 likely to be entirely explained by the confound: less than one percent of coefficients from 279 randomized data were larger than the observed coefficient (Fig. 3C and 3G; Table S11), in the

direction of higher maternal sociality in this time interval being associated with *higher* infant mortality. Overall, these results suggest that infant-dependent and reproductive state-dependent patterns in maternal behavior are sufficient to produce non-zero associations between maternal sociality and infant survival when they are not taken into account, and that any evidence for maternal sociality having a consistent positive effect on infant survival is weak at best.

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286 As a complementary test of this hypothesis, we conducted a second data randomization 287 (hereafter 'Randomization of Time Intervals') that was designed to avoid (rather than quantify) 288 any potential confound. Here, we randomly sampled, with replacement, the age at death for infants 289 who died within one year, and assigned that age as the last day of the time interval used for SCI 290 calculation for infants who survived. This approach avoids both the problem of sampling maternal 291 behavior that occurs after an infant has died and the problem of measuring SCI at later ages for 292 infants who survive. We then calculated maternal SCI-F and SCI-M over the six months prior to 293 either the true death date (for dead infants) or the randomly sampled death date (for surviving 294 infants). We analyzed data from 100 of these randomized data sets each with between 766 and 793 295 infant outcomes, depending on the number of infants whose mothers had complete behavioral data 296 available for the sampled six month time interval, using GLMs that paralleled those applied above. 297 These analyses provided no evidence that SCI-F or SCI-M predict infant survival (mean coef-298 0.025; mean OR=0.975; 0% of p-values < 0.05 and mean coef=-0.029; mean OR=0.971; 0% of p-299 values < 0.05, respectively; distribution of coefficients shown in Fig. S8A and S8B). 300





302 <u>Figure 3:</u> Results of the "Randomization of SCI Values" analysis. Histograms show the 303 distribution of coefficients (i.e., the natural logarithm of the odds ratio) from binomial GLMs using 304 randomized SCI data; vertical dashed lines show coefficients from GLMs with observed data (see 305 Fig. 2). Left column shows SCI-F estimated over A) pregnancy, B) the six months following birth, 306 C) the seven to 12 months following birth, and D) a shifting data window (six months prior to

307 infant death or first birthday depending on survival outcome). Right column shows SCI-M 308 estimated over E) pregnancy, F) the six months following birth, G) the seven to 12 months 309 following birth, and H) a shifting data window. Negative coefficients mean high SCI is associated 310 with low risk of mortality and thus high survival, while positive coefficients mean high SCI is 311 associated with high risk of mortality and thus low survival. If an observed coefficient (dashed 312 line) is to the *left* of the distribution of permuted coefficients, then high maternal sociality is 313 associated with high infant survival after accounting for the confounding effect of infant and 314 reproductive variation in maternal social behavior. If an observed coefficient is to the right of the 315 distribution, then high maternal sociality is associated with high infant mortality after accounting 316 for the confound. See Table S11 for median coefficient values and variances associated with 317 distributions of coefficients from analyses on randomized data sets.

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319 <u>REPRODUCTIVE STATE AND INFANT-DEPENDENT TRENDS IN MATERNAL SOCIAL</u> 320 <u>BEHAVIOR EXPLAIN THE EFFECT OF LIFETIME MATERNAL SOCIALITY ON INFANT</u> 321 <u>SURVIVAL</u>

322 The present analysis was motivated in part by previous results from Silk et al. (2003a), who 323 showed that more socially integrated mothers experience higher infant survivorship. Silk et al. 324 (2003a) used the individual mother (i.e., a female's lifetime success at producing surviving infants) 325 as the unit of analysis while the analyses reported above used the infant as the unit of analysis. 326 Furthermore, Silk et al. (2003a) used a lifetime estimate of the "composite sociality index" (CSI), 327 which combines data on female social relationships with both males and females (captured by 328 grooming and proximity to others during focal points; see Supplementary Methods), as the 329 measure of maternal sociality. This approach is in contrast to the measures we used above, which 330 were summarized over short time intervals, sex-specific (SCI-F and SCI-M), and used only 331 grooming data. These differences could mitigate the influence of the confound described above. 332 To test this possibility, we recreated the analysis reported in Silk et al. (2003a) using an expanded 333 data set that included an additional 22 years of behavioral and demographic data that have since 334 accumulated. A total of 295 adult females were included in our recreation, compared to 108 in Silk 335 et al. (2003a). See the Supplementary Methods for more details about our recreation.

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We conducted three analyses on the expanded data set. First, following Silk et al. (2003a), we analyzed the relationship between a mother's lifetime relative infant survival and her lifetime CSI (an estimate of her tendency to groom with and be in proximity to other adults). In this analysis, we recovered a similar pattern to the original paper: higher maternal sociality was associated with higher relative infant survival (B=0.095; p=0.027) (Fig. 4A), although the effect size was notably smaller than originally estimated (original paper reports B=0.321; p=0.015).

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344 Next, we repeated the analysis twice more, first removing time periods from the CSI 345 calculation in which mothers had infants less than one year old, and second removing all time 346 periods except pregnancy (when females typically do not have young infants and all share the same 347 reproductive state). These two analyses limit the measures of maternal social behavior to periods 348 when the mother's behavior is least likely to be influenced by the presence of an attractive infant 349 or by sexual cycling. In neither of these analyses did the relationship between maternal CSI and 350 offspring mortality reach statistical significance (removing time periods where mothers had infants 351 less than one year of age: B=0.031, p=0.268, n=256, Fig. 4B; restricting the analysis to periods of 352 pregnancy: B=0.025, p=0.332, n=268, Fig. 4C). These results support the idea that the original 353 result reported in Silk et al. (2003a) is primarily explained by infant-dependent and reproductive 354 state-dependent trends in maternal social behavior.







358 Figure 4: Recreation of Fig. 1 in Silk et al. (2003a) using an updated data set (1984 - 2022),

359 showing the relationship between a mother's lifetime Composite Sociality Index (CSI) score and

360 relative infant survival. Panels from left to right show estimates: A) including all data from

361 mothers' adult lives (green), B) removing periods in the CSI calculation when mothers had a young

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- infant less than one year old (blue), and C) only including data from periods of pregnancy (purple) 363 in the CSI calculation.
- 364

365 EFFECTS OF SOCIALITY ON ADULT SURVIVAL OUTWEIGH EFFECTS OF SOCIALITY 366 ON OFFSPRING SURVIVAL

367 Our analysis revealed some limited evidence that mothers who are highly social during certain 368 periods of their infant's early life may experience modestly reduced offspring survival (Fig. 3C 369 and 3G). At the same time, high SCI and DSI scores are associated with higher survival for adult 370 female baboons themselves (Silk et al., 2010; Archie et al., 2014; Campos et al., 2020; Lange et 371 al., 2023). If maternal sociality does indeed having conflicting influences on infant and maternal 372 survival, how is the resulting tradeoff between maternal and infant survival resolved? To probe 373 this question, we built a simple matrix projection model based on the life cycle of female baboons. 374 The model's parameters were chosen to roughly match the life history of our study population, 375 and infant and adult female survival rates were assumed to be a function of SCI (males were not 376 explicitly modeled; see Supplementary Methods for details).

377 Our primary goals were to determine what values of SCI maximize λ_1 —the leading 378 eigenvalue from the transition matrix representing long-term per capita growth rate (the estimate 379 of female fitness generated by the model; see Supplementary Methods)-and to understand how the relationship between SCI and λ_1 changes depending on how SCI affects infant survival versus 380 381 adult survival in our model. If λ_1 is greatest when SCI is relatively high, this result would indicate that high levels of maternal sociality confer net fitness benefits for females. If λ_1 is greatest when 382 383 SCI is relatively low, this result would indicate that low levels of maternal sociality confer net 384 fitness benefits (and conversely that high levels of maternal sociality impose net fitness costs).

385 We found that the effect of SCI on overall female fitness (λ_1) depends almost entirely on 386 how SCI affects adult female survival, regardless of the outcome for infants (Fig. 5). In Fig. 5, 387 both of the top quadrants (where SCI is *positively* associated with adult survival) are almost

388 entirely dark in color, indicating that high SCI *improves* fitness (λ_1) when high SCI is good for 389 adult survival (positive values on the y-axis), regardless of the effect of SCI on infant survival. 390 Meanwhile both of the bottom two quadrants (where SCI is negatively associated with adult 391 survival) are almost entirely light in color indicating that high SCI *reduces* fitness (λ_1) when high 392 SCI is bad for adult survival (negative values on the y-axis), regardless of the effect of SCI on 393 infant survival. Notably, exceptions occur where the effects of SCI on adult survival are quite weak 394 (i.e., the adult slope is close to zero): in these cases the infant slope becomes more influential in 395 determining how SCI effects female fitness. The fact that SCI mainly influences fitness based upon 396 its effects on adult survival likely follows from there being more adults in the population than 397 infants at any given time.

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400 <u>Figure 5:</u> Colored landscape showing the value of SCI that maximizes λ_1 (female fitness) in our 401 matrix projection model. Dark colors correspond to cases where relatively high values of SCI 402 improve female fitness, while light colors correspond to cases in which high values of SCI reduce 403 female fitness (see color legend). Axes show the slope of infant survival (x-axis) and adult survival

404 (x-axis) as a function of SCI when SCI = 0 (the population mean). More positive slopes correspond
405 to a greater beneficial influence of high SCI on survival and more negative slopes correspond to a

- 406 greater detrimental influence of high SCI on survival.
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409 **DISCUSSION**

410 MATERNAL SOCIALITY AND INFANT SURVIVAL

We found no strong evidence that maternal sociality is causally associated with improved infant survival in our infant-wise analysis. We instead found that apparent relationships between maternal sociality and infant survival are more parsimoniously explained by infant-dependent and reproductive state-dependent trends in a mother's social behavior. As a result, the direction of the relationship between maternal social behavior and infant survival changes depending on the time interval maternal sociality is captured over.

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418 Our results further suggest that infant-dependent and reproductive state-dependent 419 differences in maternal social behavior can generate a spurious positive relationship between 420 maternal social relationships and infant survival not only in the first six months of an infant's life, 421 but even in the seven to 12 months following a birth. Specifically, Fig. 3C and Fig. 3G demonstrate 422 that the distribution of effect sizes from analyses with randomized SCI-F and SCI-M data do not 423 center on zero, indicating that removing data for the first 100 days following a birth (as in Silk et 424 al., 2009) would not be enough to completely eliminate the effect of this confound in our 425 population. Furthermore, controlling for this confound not only effects conclusions from infant-426 wise analyses, but also strongly attenuates the original effect reported in Silk et al. (2003a), which 427 used the liftetime estimates of maternal sociality and infant survival (Fig. 4). Notably, when tested 428 separately as predictors of relative infant survival in Silk et al. (2003a), proportion of time being 429 groomed was the only one of the three social components of CSI that was a significant predictor 430 of survival on its own (i.e., time spent grooming and spent in proximity to others are not significant 431 predictors), suggesting that the amount of time mothers spend being groomed by others drives the 432 relationship between CSI and infant survival. This further supports the idea that infants and certain 433 reproductive states (e.g., sexual cycling) attracting social attention (expressed as directed 434 grooming) contribute to the original result.

435

436 After accounting for variation in individual behavior driven by reproductive state and 437 infants, our analyses with data from the Randomization of SCI Values analysis produced some 438 limited evidence that more social mothers may experience higher infant mortality (Fig. 3C and 439 3G). A negative association between infant survival and maternal sociality captured over certain 440 time periods could have multiple, non-mutually exclusive explanations. High levels of maternal 441 social interactions could result in increased exposure to pathogens (May, 1983; Nunn, 2012), 442 increased stress (Pearson et al., 2015), reduced time spent feeding (Altmann, 1980), or increased 443 vulnerability to fatal kidnappings or infanticide (Kleindorfer & Wasser, 2004; Shopland & 444 Altmann, 1987; Kalbitzer et al., 2017). Furthermore, a high frequency of social interactions 445 between mothers and other females can result in rough handling of infants by other females, which 446 in turn is associated with signs of distress in infants in Amboseli (Nguyen et al., 2009), and with 447 lower survival in several other nonhuman primate populations (Silk 1980; Kleindorfer & Wasser 448 2004). On the other hand, maternal sociality may be negatively associated with infant survival if 449 less healthy mothers or infants attract more social attention, for example, if kin and other close 450 social partners intensify efforts to help support mothers when risk of offspring mortality appears 451 heightened. In this case, maternal social bonds would not contribute directly to infant survival, but 452 rather unhealthy mothers and infants would receive more social attention.

453

454 Notably, a potentially harmful effect of higher maternal sociality on infant survival was 455 only supported over some time intervals used in our analyses. Analyses from our Randomization 456 of Time Intervals approach did not support an effect of SCI on infant survival in either direction. 457 Moreover, when SCI-F was measured over pregnancy and SCI-M was measured over the six 458 months following a birth, the results of our Randomization of SCI Values were suggestive in the 459 opposite direction: high SCI may be associated with slightly higher survival (although we could 460 not rule out that these effects could be explained by a confound; Fig. 3A and 3F). The direction of 461 these results with SCI-F are consistent with findings in Blersch et al. (2023) who found a positive 462 relationship between maternal social connectedness and infant survival when measuring social 463 relationships during pregnancy in an attempt to avoid the confound of infant attractiveness. 464 Importantly, such a result could also be explained by reverse causality if, for example, mothers 465 tend to socialize more when they are healthier (and thus more likely to birth a heathy infant).

466 Regardless, given that we show maternal sociality has a neutral and occasionally even negative 467 relationship with infant survival when quantified over other time periods, this suggests that 468 mothers being more social does not provide universal fitness benefits for infants overall, at least 469 in our population.

470

471 Furthermore, according to the results of our matrix projection model, any negative or 472 positive effects of maternal sociality on infant survival are unlikely to shape maternal social 473 behavior if maternal social relationships are linked to improved adult female survival. Specifically, 474 even if maternal sociality directly affects both adult and infant outcomes, the survival benefits or 475 costs that a mother experiences from being social outweigh the benefits or costs experienced by 476 her infants. In other words, selection should not act to reduce sociality in baboon females if it 477 improves their own survival, even if those social relationships pose potential risks to their infants. 478 Consistent with adult survival having an outsized effect on fitness, findings suggest longevity is 479 more important than fecundity for lifetime reproductive success in the Amboseli baboons (McLean 480 et al., 2019).

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483 **CONCLUSIONS**

484 We have shown that the manner in which a female baboon's social interactions are affected by her 485 reproductive state and whether she has a live infant can lead to the erroneous inference that 486 maternal sociality improves infant survival. Specifically, the manner in which a female's social 487 behavior changes depending upon her reproductive state and infant status generates a correlation 488 between maternal sociality and infant survival that varies in both direction and magnitude 489 depending on when sociality is measured relative to an infant's birth or death (Fig. 3). Correcting 490 for reproductive state and infant-related variation in maternal sociality can attenuate, eliminate, or 491 even reverse the direction of maternal sociality-infant survival correlations. We believe that this 492 confound is the best explanation for previous results reported from our study system, where high 493 levels of maternal sociality were interpreted as a driver of enhanced infant survival (Silk et al., 494 2003a).

495

496 Taken together our results provide no strong support for a positive effect of maternal 497 sociality on infant survival after accounting for this confound. In fact, the most compelling 498 evidence for an association between maternal sociality and offspring survival was found in the 499 seven to twelve months following birth where maternal sociality may be associated with poor 500 infant survival in the Amboseli baboon population. This evidence that more social mothers may 501 experience poor infant survival is suggestive rather than definitive: the strength of the relationship 502 between high maternal sociality and poor infant survival varies across the time periods in which 503 we measured maternal sociality (Fig. 3). Importantly, female state-dependent patterns of social 504 interaction occur in other primate species (e.g., Gumert, 2007; Proctor et al., 2011; Tiddi et al., 505 2010), suggesting that our findings may generalize to other analyses. Consequently, our results 506 mandate a closer scrutiny of the causal pathways linking maternal social behavior and infant 507 survival in other nonhuman primate populations.

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510 METHODS

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512 <u>ESTIMATING THE RELATIONSHIP BETWEEN INFANT SURVIVAL AND MATERNAL</u> 513 SOCIALITY, MEASURED OVER DIFFERENT TIME INTERVALS

514 Study system and subjects. Data for this analysis come from a long-term study of baboons (Papio 515 cynocephalus and P. anubis) in the Amboseli basin of southern Kenya. This population has been 516 under continuous observation since 1971 (Alberts & Altmann, 2012). All study subjects in our 517 data set were live infants born into study groups between July of 1983 (when behavioral data began 518 being collected systematically and consistently in our population) and December of 2020.

519

Maternal sociality. We quantified SCI (our measure of maternal sociality) in four sixmonth time intervals: (i) pregnancy, (ii) the fixed six-month window after the infant's live birth (regardless of whether and when the infant died), (iii) the fixed six-month window beginning seven months after any infant's live birth (regardless of whether and when the infant died), and (iv) a 'shifting time window' that represents the six months prior to each infant's death or – if it survived – its first birthday. Maternal SCI with females (SCI-F) and with males (SCI-M) were calculated separately for each time interval using grooming data collected during "representative interaction

527 sampling" (see Supplementary Methods) by averaging the residuals from two linear regressions: 528 (i) log grooms given by the focal individual to other groupmates, regressed on log observer effort 529 (i.e., the number of focal follows per adult female in the group divided by the number of days the 530 female of interest was present in the group; see extended description in Supplementary Methods); 531 and (ii) log grooms received by the focal individual, regressed on log observer effort (Archie et 532 al., 2014). These values were z-scored within population and year and used as our final measure 533 of SCI.

534

535 Testing the relationship between SCI and infant survival. We tested whether SCI-F or SCI-536 M predicted infant survival in separate models using binomial GLMs implemented with the glm 537 function from R. Each model included one maternal social measure (SCI-F or SCI-M) measured 538 over one of four time intervals. The outcome was scored according to whether the infant died (1) 539 or survived (0). Given that the purpose of our analysis was to interrogate the previously reported 540 linear effects of maternal sociality on infant survival and that no explicit hypotheses exist to 541 support a non-linear relationship between sociality and infant survival, we did not explore the 542 possibility of non-linear relationships in our analyses.

543

544 Several variables are already known or suspected to influence offspring survival in this 545 population (Silk et al., 2003a; Zipple et al., 2021). To control for the possible influence of these 546 variables on infant survival, we also measured them over the time intervals of interest and included 547 them as additional fixed effects in our GLMs. These additional fixed effects were i) the relative 548 proportion of adult females a given mother outranked in her group, ii) maternal age, (iii) maternal 549 age squared, (iv) whether the infant of interest was the mother's first birth (1=true; 0=false), and 550 (v) group size. In models with SCI-M we also included the absolute proportion of days over the 551 six-month time interval in which the mother was experiencing menstrual cycling as a fixed effect, 552 to control for the ephemeral social relationships that arise between adult males and females when 553 females are sexually cycling (Rowell, 1968; Seyfarth, 1978). In all models, continuous fixed 554 effects were scaled by subtracting the mean and dividing by the standard deviation to ensure model 555 convergence. We initially used mixed models including the random effects of maternal identity, 556 hydrological year, and group identity in addition to fixed effects. However, these variables 557 explained little to no variance leading to issues of singular fit and thus uninterpretable results.

- 558 These terms were therefore dropped from all final analyses. Analyses of residuals with DHARMa
- 559 (Hartig, 2017) indicated acceptable model fits.
- 560

561 DATA AVAILABILITY STATEMENT

- 562 Data and code will be made available on Zenodo upon publication.
- 563

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