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Abstract: Vitamin D has been implicated in the regulation of glucose metabolism and insulin resistance. We designed this study to provide evidence that insulin resistance is dependent on the concentration of vitamin D in the body. Forty observational studies of both type 2 diabetes mellitus patients and healthy subjects were included in this meta-analysis. Related articles were searched from Embase, PubMed, and Medline through January 2021. Filters for search were used to obtain more focused results. We used Comprehensive Meta-Analysis Version 3 for the construction of forest plots. RevMan software version 5.3 was used to build the risk of bias tables and summary plots. The observational studies included in this systematic review and meta-analysis showed an inverse relationship of insulin resistance with the status of vitamin D both in non-diabetic (r = -0.141 to -0.234; p = 0.000) and diabetic (r = -0.255; 95% CI = -0.392 to -0.107, p = 0.001) populations. From the meta-analysis we concluded that hypovitaminosis D is related to increased levels of insulin resistance in both type 2 diabetes patients and the healthy population all over the world.

Keywords: hypovitaminosis D; insulin resistance; fasting plasma insulin; type 2 diabetes; body mass index; vitamin D

1. Introduction

Insulin resistance and type 2 diabetes mellitus (T2D) are among the greatest challenges of this time. Obesity is one of the major risk factors for the spread of these diseases [1]. Insulin, the glucose lowering hormone, has an important role in the adipose tissues, liver, and skeletal muscles. After binding to its receptors in the cell membrane, the insulin starts metabolic reactions, e.g., it stores glucose in the skeletal muscles and liver, initiates glucose use in the skeletal muscles, and is involved in the regulation of genes related to lipid synthesis and glucose transport. Insulin also functions to suppress lipolysis in the liver, reducing the concentration of acetyl-CoA, thus decreasing pyruvate carboxylase activity. A decrease in pyruvate carboxylase and glycerol production helps insulin reduce gluconeogenesis [2,3]. A higher insulin level in the blood to maintain a normal status of glucose defines insulin resistance. Insulin resistance is found to be the culprit for a number of diseases such as pre-diabetes, non-alcoholic fatty liver (NAFL), and polycystic ovaries [4–6]. Continuous high insulin requirement exhausts the beta cells of the islets of Langerhans, resulting in the obvious progression of type 2 diabetes. Hypovitaminosis D is considered to be related to the development of T2D, as evident from a number of epidemiological studies [7-9]. Deficiency of vitamin D is also potentially linked with non-alcoholic fatty liver disease, cardiovascular disease, and overall mortality risk [10–12]. Vitamin D is a fat-soluble prohormone steroid that has endocrine, paracrine, and autocrine functions [13]. Studies showed that deficiency in vitamin D develops insulin resistance, which in turn promotes obesity and type 2 diabetes [14]. The 1α -hydroxylase enzyme required for the conversion of 25 (OH) vitamin D into its functionally active form 1,25 (OH)2 vitamin D and vitamin D receptor (VDR) are found in the beta cells, showing its



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Copyright: © 2021 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (https:// creativecommons.org/licenses/by/ 4.0/). role in the homeostasis of insulin production [15]. The progression of diabetes is slowed with vitamin D supplementation in animal models of diabetes. Moreover, a high risk of type 2 diabetes and an intensive hyperglycemia have been observed for carbohydrate consumption under hypovitaminosis D conditions [16,17]. A strong link has been found between vitamin D status and insulin response in tissues in non-diabetic subjects [18]. The evidence of vitamin D correlation with insulin resistance is continuously increasing all over the world, showing an inverse relationship between them, which is consistent with our hypothesis [19,20].

The goal of this review was to reveal the relationship of vitamin D status and fasting plasma insulin as a measure of insulin resistance in previous diabetic and non-diabetic observational studies. The prospective relationship of vitamin D levels and insulin resistance was examined in this study using a forest plot. Vitamin D status is also affected by sun; therefore, latitude can have an effect on this relationship. Other factors that can affect this association are the method of vitamin D determination and BMI of the selected population. To identify the influence of these parameters on the relationship between vitamin D and insulin resistance, we performed meta-regression analysis.

2. Materials and Methods

Three databases (Embase, Medline, and PubMed) were searched for this review article to find appropriate observational studies through to January 2021. The keywords used were: "cholecalciferol", "25 (OH) vitamin D", "25 (OH) D", "vitamin D3", "vitamin D", in combination with "fasting plasma insulin", "HBA1C", "homeostasis model assessment of insulin resistance", "fasting plasma glucose", "type 2 diabetes", "T2D", "adiposity", and "abdominal obesity". The search for the keywords was performed both as free keywords and in combination with EMTREE in Embase, and Medical Subject Heading (MeSH) in PubMed. The studies selected showed the relationship between vitamin D (25-hydroxy vitamin D) and fasting plasma insulin. The selection criteria included studies conducted on human beings of more than 18 years of age, written in English. Editorials, commentaries, and reports were not included in this study. The articles were also searched by other sources in addition to systematic search for more references. If the articles lacked necessary information on moderators or estimates, the authors were contacted.

Statistical Analysis and Outcome Measures

The aggregate effect measure was extracted and pooled for meta-analysis as a correlation coefficient. We used the random effect model to compute the forest plot as the summary measure for the outcome. Studies were collected from a range of populations in different regions of the world with different ethnicities, cultures, and customs, since the biological effect of vitamin D varies with location. The estimates of consistency and reliability were tested by I² and τ^2 , respectively, where I² defines total heterogeneity as percentage among included studies.

Grades of Recommendation Assessment Development and Evaluation (GRADE) was used for quality assessment of the articles. The factors that determined the quality of the study were: 1. indirectness (compromised generalizability of results); 2. inconsistency (unexplained heterogeneity between studies); 3. publication bias (small number of participants); 4. imprecision (confidence intervals too long). Comprehensive Meta-Analysis Version 3 (Biostat, Inc., Englewood, NJ, USA) was used to perform meta-analysis. Meta-regression was performed (Comprehensive Meta-Analysis Version 3, Biostat, Inc., Englewood, NJ, USA) to determine the sources of bias. The risk of bias (ROB) analysis was performed using Review Manager 5.3.

3. Results

A total of 2023 studies were identified electronically (Pubmed, Embase, and Medline). Nineteen references were recognized by other means. Endnote software was used to screen the duplicate entries and 998 entries were discarded. A total of 749 studies were excluded

on the basis of title. The rest underwent abstract and full text evaluation. The abstract evaluation discarded 200 articles and systematic assessment of full text rejected 55 articles. Forty articles fulfilled the inclusion criteria and were finally selected as eligible to be used in the meta-analysis (Figure 1).



Figure 1. Flow chart for the literature review and selection of studies.

3.1. Excluded Studies on the Basis of Full Text Evaluation

Eleven articles were excluded as data were not compatible with our outcome measure of correlation coefficient [14,21–30]. Twenty-one studies were selected for exclusion because their study design did not match our study design [31–51]. Sixteen articles were rejected because the outcome measure was calculated for a mixed population, i.e., for both diabetic and non-diabetic subjects [52–67]. One study was excluded because the number of subjects in each vitamin D quartile was not mentioned [68]. For seven references, full-length articles were not accessible [69–75].

3.2. Included Studies

3.2.1. Meta-Analysis and Meta-Regression for Non-Diabetes Patient Studies

Thirty-five studies included in this meta-analysis were collected through to January 2021. The participants of all studies were at least 18 years old. Twelve studies determined vitamin D concentration by radioimmunoassay (RIA), five by enzyme-linked immunosorbent assay, eight by chemiluminescence assay (CLIA), three by electrochemiluminescence assay (ECLIA), four by liquid chromatography-mass spectrometry (LC-MS), one by high-performance liquid chromatography (HPLC), and two studies did not mention the method of determination. The articles selected were from all over the world and from different ethnicities.

Because of the large amount of variability due to the above-mentioned sources, we used the random effect model for this meta-analysis. An inverse relationship (r = -0.188, 95% CI = -0.141 to -0.234, p = 0.000) was seen between fasting plasma insulin and vitamin D concentrations in the blood for all thirty-five non-diabetic subject studies (Figure 2). The correlation of all studies lies between r = -0.041 and r = -0.397. The meta-regression analysis showed R² to be zero for both latitude and method of determination of vitamin D, meaning the relationship between vitamin D concentration and fasting plasma insulin

is independent of these two variables (Figures 3 and 4). The summary of the GRADE assessments is presented in Figures 5 and 6. The subgroup analysis for different quartiles of BMI depicts an overall increasing strength of correlation between fasting plasma insulin and vitamin D status from lower to higher BMI quartile. For example, the correlation was r = -0.152, 95% = -0.206 to -0.097, p = 0.000 in the lowest quartile (BMI < 25) (Figure 7); r = -0.153, 95% = -0.206 to -0.099, p = 0.000 in the medium quartile; and r = -0.229, 95% = -0.322 to -0.131 (BMI = 25–30) (Figure 8), p = 0.000 in the highest quartile (BMI > 25) (Figure 9). The correlation was almost the same in the first two quartiles; however, it was significantly higher in the third quartile compared to the first two quartiles.



Non-diabetic Subject Studies

Figure 2. Forest plot for non-diabetic subject studies showing the correlation between status of vitamin D and fasting plasma insulin. The random effect model was used for the determination of correlation and 95% CI [19,20,57,76–106].





Figure 3. Meta-regression analysis for non-diabetic subject studies. Latitude R^2 shows the effect of the moderator on the heterogeneity of correlation. (**a**) total variance in true effects; (**b**) not explained by model; (**c**) explained by model.



R² for Model 1, Random effects (MM), Z-Distribution, Fisher's Z

Figure 4. Meta-regression analysis for non-diabetic subject studies. Method of determination of vitamin D R² shows the effect of the moderator on the heterogeneity of correlation. (**a**) total variance in true effects; (**b**) not explained by model; (**c**) explained by model.



Figure 5. Risk of bias assessment for non-diabetic subject studies (plus sign shows low risk, minus sign shows high risk, and question mark shows unknown bias) [19,20,36,57,76–106].



Figure 6. Summary of risk of bias assessment for non-diabetic subject studies, data shown are in percentages.



Non-diabetic Subject Studies (BMI < 25) Subgroup-analysis

Figure 7. Forest plot for the lowest BMI quartile (<25) of non-diabetic subject studies showing the correlation between the status of vitamin D and fasting plasma insulin. The random effects model was used for the determination of correlation and 95% CI [85,87,90–92,100,101,103,104,106].

Non-diabetic Subject Studies (BMI 25-30) Subgroup-analysis

Study name	Statistics for each study						Correlation and 95% Cl					
	Correlation	Lower limit	Upper limit	Z-Value	p-Value							
Abbasi 2015 [76]	-0.128	-0.214	-0.040	-2.834	0.005	1						
Al-Daghri 2013 [77]	-0.041	-0.224	0.145	-0.430	0.667							
Bhatt 2014 [79]	0.093	-0.076	0.257	1.080	0.280			_+∎_	-			
Bindal 2011 [80]	-0.397	-0.603	-0.142	-2.970	0.003			-				
Chacko 2011 [83]	-0.216	-0.345	-0.079	-3.065	0.002		-	-				
Friedman 2012 [86]	-0.220	-0.384	-0.042	-2.419	0.016							
Grimnes 2011 [88]	-0.112	-0.263	0.044	-1.409	0.159		- I ·					
Grineva 2013 [89]	-0.320	-0.431	-0.199	-5.019	0.000		-₩	-				
Hurskainen 2012 [57]	-0.070	-0.127	-0.013	-2.395	0.017							
Li 2011 [94]	-0.336	-0.635	0.050	-1.713	0.087			-				
Liu 2009 [95]	-0.201	-0.280	-0.118	-4.717	0.000							
Moore 2015 [97]	-0.200	-0.276	-0.122	-4.933	0.000		11	-				
Muscogiuri 2010 [98]	-0.130	-0.428	0.193	-0.784	0.433							
O'Hartaigh 2013 [99]	-0.047	-0.105	0.011	-1.587	0.113							
Sheth 2015 [102]	-0.197	-0.281	-0.110	-4.373	0.000							
	-0.153	-0.206	-0.099	-5.515	0.000			◆				
						-1.00	-0.50	0.00	0.50	1.00		
						Lo	w 25(OH)D Hiç	gh 25(O⊦	I) D		

Figure 8. Forest plot for the medium BMI quartile (25–30) of non-diabetic subject studies showing the correlation between the status of vitamin D and fasting plasma insulin. The random effects model was used for the determination of correlation and 95% CI [57,76,77,79,80,83,86,88,89,94,95,97–99,102].

Study name	Statistics for each study					Correlation and 95% Cl						
	Correlation	Lower limit	Upper limit	Z-Value	p-Value							
Boonchaya-anat 2014 [81]	-0.077	-0.217	0.066	-1.058	0.290							
Botella-Carretero 2007 [82]	-0.240	-0.501	0.061	-1.567	0.117			⊷				
Bril 2015 [19]	-0.301	-0.434	-0.155	-3.954	0.000			-				
De-Pergola 2013 [84]	-0.350	-0.546	-0.118	-2.901	0.004		+	- 1				
Del-Gobbo 2011 [36]	-0.140	-0.224	-0.054	-3.173	0.002							
Kayaniyil 2011 [93]	-0.310	-0.378	-0.239	-8.179	0.000							
	-0.229	-0.322	-0.131	-4.521	0.000		- ◀					
						-1.00	-0.50	0.00	0.50	1.00		
						Lo	w 25(OH)D Hig	h 25(OH) D		

Non-diabetic Subject Studies (BMI >30) Subgroup-analysis

Figure 9. Forest plot for the highest BMI quartile (>30) of non-diabetic subject studies showing the correlation between the status of vitamin D and fasting plasma insulin. The random effects model was used for the determination of correlation and 95% CI [19,36,81,82,84,93].

3.2.2. Meta-Analysis and Meta-Regression for Diabetes Patient Studies

Seven studies fulfilled the criteria to be included in this meta-analysis for the relationship of vitamin D with fasting plasma insulin in diabetic patients. In this meta-analysis, we found an inverse association (r = -0.255, 95% CI = -0.392 to -0.107, p = 0.001) between fasting plasma insulin and vitamin D levels (Figure 10). The range of correlation in all studies was -0.045 to -0.25, except for one study from Southern Spain [107], which showed an increased correlation (r = -0.882). The effect of the moderator (latitude) on the correlation of vitamin D status and fasting plasma insulin was determined by meta-regression analysis. The results showed that the latitude ($R^2 = 0.000\%$, p = 0.000) did not contribute to heterogeneity in this correlation (Figure 11). The summary of GRADE assessment is presented in Figures 12 and 13.

Study name	Statistics for each study						Correlation and 95% Cl					
	Correlation	Lower limit	Upper limit	Z-Value	p-Value							
Al-Daghri 2013 [77]	-0.230	-0.375	-0.074	-2.868	0.004		-					
Cai 2014 [108]	-0.054	-0.106	-0.002	-2.026	0.043							
Calvo-Romero 2015 [107]	-0.820	-0.882	-0.731	-10.018	0.000		-					
Dalgard 2011 [53]	-0.100	-0.175	-0.024	-2.587	0.010							
Gao 2015 <mark>[</mark> 109]	-0.045	-0.235	0.148	-0.455	0.649							
Gulseth 2010 [18]	-0.140	-0.230	-0.048	-2.966	0.003							
Sheth 2015 [102]	-0.254	-0.340	-0.163	-5.360	0.000			┣				
	-0.255	-0.392	-0.107	-3.329	0.001							
						-1.00	-0.50	0.00	0.50	1.00		
						Lo	w 25 (25))D Hig	Jh 25 (O⊦	I) D		

Diabetic Subject Studies

Figure 10. Forest plot for diabetic subject studies showing the correlation between status of vitamin D and fasting plasma insulin. The random effects model was used for the determination of correlation and 95% CI [18,53,77,107–109].



R² for Model 1, Random effects (MM), Z-Distribution, Fisher's Z

Figure 11. Meta regression analysis for diabetic subject studies. Latitude R² shows the effect of the moderator on the heterogeneity of correlation. (**a**) total variance in true effects; (**b**) not explained by model; (**c**) explained by model.



Figure 12. Risk of bias assessment for diabetic subject studies (plus sign shows low risk, minus sign shows high risk, and question mark shows unknown bias) [18,53,77,102,107–109].



Figure 13. Summary of risk of bias assessment for diabetic subject studies: data shown are in percentages.

4. Discussion

It is evident from this meta-analysis that the levels of vitamin D in the body are inversely related to insulin resistance both in diabetic and non-diabetic populations. However, the correlation is stronger in the diabetic population (r = -0.255, 95% CI = -0.392 to

-0.107, p = 0.001) (Figure 2) compared with the non-diabetic population (r = -0.188, 95% CI = -0.141 to -0.234, p = 0.000) (Figure 10).

The status of vitamin D is inversely related to insulin resistance independent of age and sex. The active form of vitamin D (1,25-hydroxy vitamin D) has been detected in the pancreas [110]; therefore, there is a possibility that vitamin D plays a role in the evolutionary development of metabolic systems such as beta cell function. Hypovitaminosis D is associated with reduced calcium status in the blood circulation, which ultimately controls insulin synthesis and insulin secretion by beta cells [111]. Vitamin D supplementation increases plasma calcium levels, which in turn increase the synthesis and secretion of calcium from the beta cells, ultimately improving glucose homeostasis [21,73]. Hypovitaminosis D therefore plays a role in the development of insulin resistance by affecting insulin synthesis and secretion from beta cells and by regulating circulating serum calcium.

The subgroup analysis on the basis of BMI showed an increasingly strong inverse relationship between vitamin D status and insulin resistance with increasing BMI in nondiabetic subject studies. The strength of correlation is stronger (r = -0.229, 95% = -0.322to -0.131) in the highest BMI quartile, and almost the same in the first (r = -0.152, 95% = -0.206 to -0.097, p = 0.000) and second (r = -0.153, 95% = -0.206 to -0.099, p = 0.000) BMI quartiles. According to previous studies, a synergy exists between hypovitaminosis D and obesity in developing insulin resistance [14,20,42]. The expression of vitamin D receptors is more pronounced in obese compared with lean subjects, and vitamin D deficiency has an independent inverse relationship with BMI [112]. The anti-insulin resistance mechanism of vitamin D might act through its anti-inflammatory mechanism in overweight subjects. A decrease in inflammatory cytokines after vitamin D treatment has been observed in many previous studies and might have a role in promoting insulin sensitivity [21]. The cycle works via insulin-stimulated fat synthesis and adipose tissue initiating the synthesis of inflammatory markers, which then lead to augmented insulin resistance. Vitamin D interrupts this cycle at the level of adipogenesis by hindering it and at the level of inflammatory marker production by lowering their synthesis [113].

The underlying cause of obesity-related insulin resistance is inflammation induced by obesity. Vitamin D is well-known for its anti-inflammatory functions as it lowers the concentration of different inflammatory indicators (C-reactive protein (CRP), tumor necrosis factor-a (TNF-alpha), and interleukin-6 (IL-6)) [114]. Numerous studies have shown the effect of insulin resistance on the risk of cardiovascular disease, which is doubled in insulin-resistant compared with normal populations. Considerable similarities in the biochemical profile of insulin resistance and inflammation have been observed in diabetic and cardiovascular patients recently. A recent study even showed a role for insulin resistance in the development of ischemic heart disease under normal glucose tolerance [115].

Vitamin D receptor (VDR) is required for the functioning of vitamin D in different tissues. However, the requirements for the expression of VDR vary in different tissues, e.g., in some tissues, it requires calcium and vitamin D for its expression, and in others, it needs neither. It has been reported that vitamin D induces insulin secretion in the beta cells of the pancreas and increases insulin sensitivity in target cells, i.e., muscle, adipose tissue, and liver [116–118]. Hypovitaminosis D has been shown to be related to hyperglycemia and insulin resistance earlier [29,119].

The epigenetic effect of vitamin D has been observed at the level of transcription for many genes. Insulin receptor substrate (IRS-1) is a protein that plays an important role in promoting insulin sensitivity. The expression of IRS protein was observed to be increased by 2.4 times in high-fat-treated mouse muscle tissue after treatment with vitamin D. The anti-insulin resistance mechanism of vitamin D appears to involve insulin-mediated intracellular functions through IRS-1 [120]. The photosynthetic production of vitamin D in the skin depends on the radiation (UV-B) from sunlight. Therefore, latitude can explain the status of vitamin D geographically, but the meta-regression analysis presented in this study does not show any variability in the correlation because of latitude ($R^2 = 0.000$, p = 0.000). This might be because many factors in the modern world have reduced the impact of these radiations on the production of vitamin D. For example, concrete buildings absorb more radiation, and the gases emitted by industry and vehicles reduce the irradiance of ultraviolet B radiation from the sun [121,122]. These and other factors, such as diet, clothing styles, industrialization, reduced time for sun exposure, and skin pigmentation, have confounded the effect of latitude on the strength of correlation between vitamin D status and insulin resistance.

The meta-regression analysis for the effect of method of determination of vitamin D also showed no heterogeneity of the correlation. However, we observed an overall increased strength of inverse correlation between vitamin D status and insulin resistance when the CLIA method was used for the determination of vitamin D; the most pronounced example being the study of Calvo-Romero [107] from Southern Spain, which reported the highest correlation of -0.82.

Vitamin D is directly related to the progression of the diabetic complications such as diabetic neuropathy, diabetic nephropathy, and diabetic retinopathy. Vitamin D regulates neurotrophin and calcium homeostasis related to nerve action, and the deficiency of vitamin D exerts diverse effects on the complication of diabetic neuropathy [123]. It was observed previously that vitamin D is inversely related to diabetic neuropathy, and this relationship does not depend on the duration of diabetes disease. Chronic nephropathy developed during type 2 diabetes was also linked to diabetic neuropathy [124]. The role of vitamin D in the functioning of neurons has been established in the last couple of decades. Many neuronal diseases have been proven to be associated with hypovitaminosis D. For example, treating multiple sclerosis patients with vitamin D can slow the progression of disability [125–127]. Nerve growth factor is important for the growth and development of neurons, and myelination of Schwann cells in case of injury. Vitamin D increases the production of nerve growth factor in glial cells after crossing the blood–brain barrier and entering the glial cells [128–130]. There was progress in the treatment of diabetic foot healing when the diabetic foot was topically treated with nerve growth factor [131].

Strengths and Weaknesses

The systematic search used for the mining of research articles is one of the major strengths of this meta-analysis. The gold standard international methodology was applied, and observational studies were evaluated by the Grading of Recommendations Assessment, Development, and Evaluation (GRADE). The meta-analysis did not reveal very wide 95% confidence intervals, which shows the dependence of insulin resistance on the status of vitamin D. Although the total number of subjects was high in this meta-analysis, the studies were observational; therefore, the chances for residual confounding cannot be ruled out, which is a limitation. Potentially confounding factors include the age, ethnicity, and lifestyle of the participants. The intake of vitamin D and sun exposure has not been mentioned in all of the studies, which may be an additional source of confounding. Observational studies have the drawback of not being blinded and randomized, which is a limitation of this study. We consider this evidence to be moderate on the basis of the strengths and weaknesses of the studies included.

5. Conclusions

Diabetic hypovitaminosis D is at the pandemic level worldwide. The present systematic review and meta-analysis suggest a role of vitamin D in the regulation of insulin production and release from the beta cells of Langerhans. However, this association is not purely independent, and strongly depends on BMI as observed in the subgroup-analysis. The inverse correlation between vitamin D status and fasting insulin strengthens with increasing BMI. The meta-regression analysis did not show any effect of latitude or the method of determination of vitamin D on the overall relationship of vitamin D levels in the body and fasting insulin in the blood. There is a significant need for high-quality, long-term, randomized controlled trials to be conducted using different doses of vitamin D to see its effect on fasting plasma insulin levels.

Author Contributions: S.R. and P.B.J. worked together to extract related articles from Embase, PubMed, and Medline through to January 2021. S.R. and P.B.J. evaluated full text articles for inclusion. S.R. worked on the mining of data from the included articles, performed the meta-analysis and wrote the manuscript. P.B.J. provided feedback and approved the final version. All authors have read and agreed to the published version of the manuscript.

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