

Research

# Analysis of strain and regional variation in gene expression in mouse brain

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Published: 27 September 2001

*Genome Biology* 2001, **2**(10):research0042.1-0042.15

The electronic version of this article is the complete one and can be found online at <http://genomebiology.com/2001/2/10/research/0042>

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(Print ISSN 1465-6906; Online ISSN 1465-6914)

Received: 20 June 2001

Revised: 31 July 2001

Accepted: 24 August 2001

## Abstract

**Background:** We performed a statistical analysis of a previously published set of gene expression microarray data from six different brain regions in two mouse strains. In the previous analysis, 24 genes showing expression differences between the strains and about 240 genes with regional differences in expression were identified. Like many gene expression studies, that analysis relied primarily on *ad hoc* 'fold change' and 'absent/present' criteria to select genes. To determine whether statistically motivated methods would give a more sensitive and selective analysis of gene expression patterns in the brain, we decided to use analysis of variance (ANOVA) and feature selection methods designed to select genes showing strain- or region-dependent patterns of expression.

**Results:** Our analysis revealed many additional genes that might be involved in behavioral differences between the two mouse strains and functional differences between the six brain regions. Using conservative statistical criteria, we identified at least 63 genes showing strain variation and approximately 600 genes showing regional variation. Unlike *ad hoc* methods, ours have the additional benefit of ranking the genes by statistical score, permitting further analysis to focus on the most significant. Comparison of our results to the previous studies and to published reports on individual genes show that we achieved high sensitivity while preserving selectivity.

**Conclusions:** Our results indicate that molecular differences between the strains and regions studied are larger than indicated previously. We conclude that for large complex datasets, ANOVA and feature selection, alone or in combination, are more powerful than methods based on fold-change thresholds and other *ad hoc* selection criteria.

## Background

Genome-wide expression data such as that obtained with microarrays presents a significant challenge for analysis. A frequent goal of expression analysis is to identify genes whose expression is altered by an experimental condition. In general, automated methods are needed to accomplish this task owing to the large amount of data involved. A good

example of an expression dataset requiring detailed, complex analysis comes from the work of Sandberg *et al.* [1]. In their study, the expression of 13,067 genes and expressed sequence tags (ESTs) was assayed using oligonucleotide arrays [2] for each of six brain regions in two different inbred strains of mice, with each condition measured twice. This data is a potentially rich source of targets for investigating behavioral

and neurophysiological differences between the mouse strains (C57Bl/6 and 129SvEv), as well as structural and functional differences among brain regions.

Both regional and strain variations in gene expression are relevant to questions in neuroscience. Inbred mouse strains are used in many studies relating to human neurological and neuropsychiatric disorders such as stroke and alcoholism; however, it has long been recognized that different strains vary significantly in the results obtained in such studies [3]. Expression analysis is one approach to exploring the underlying molecular causes of these differences. Similarly, expression analysis can provide insight into one potential source of functional differences between brain regions. Traditional expression analysis (for example, northern blots and *in situ* hybridization) can be used to uncover such expression patterns one gene at a time, but the high-throughput nature of array technology allows researchers to take a much wider view.

Given the data of Sandberg *et al.*, and the questions one can address with it, the issue becomes how to proceed with the analysis. Many researchers use an ‘*ad hoc*’ approach: for example, the ‘fold change’ for each gene is considered, where fold change is defined as the ratio of gene expression in a test condition to that in a control condition. Changes in expression above a certain fold-change threshold are deemed to be significant. The problem with such methods is that they typically do not take into account the variability of the measurements being considered. When using the Microarray Analysis Suite (Affymetrix), one also has the option of using the ‘absent/present’ calls, determined by the software, to make discriminations as to whether a gene is expressed in one set of samples and not others. The trouble with this method is that the absent/present threshold is essentially arbitrary, and there is no easy way to estimate the numbers of false positives and false negatives obtained. Typically in such studies, genes are identified as ‘changed’ or ‘not changed’, often without a quantitative estimate of the statistical confidence in that conclusion.

Because of the shortcomings of *ad hoc* methods, some researchers have tackled the problem of how to apply more robust statistical methods to the problem of identifying changed genes in a dataset [4-6]. Most such methods deal with the case of a two-way comparison, for example between a wild-type organism and one that is mutant in a gene of interest. But, many experimenters are studying complex systems and generate large datasets with more than one experimental variable. Such is the case for Sandberg *et al.* [1], where both strain and brain region were varied in the same experiment. Here, we investigate methods appropriate for analyzing this type of dataset and determine how the methods affect the interpretation of the data of Sandberg *et al.*

There are several desirable properties of a method for analysis of microarray data. First, it should provide an estimate of

the confidence that the gene-expression pattern observed would occur by chance, that is, a  $p$ -value. Failing the ability to generate accurate  $p$ -values, the method should at least automatically rank the genes in order of interest. Second, the method should be flexible enough to allow for complex experimental designs. Methods that permit us to compare only two conditions are insufficient. Finally, an ideal method is computationally simple and fast. Mathematical simplicity and familiarity are also advantages when the methods must be used and interpreted by biologists with little statistical background.

A standard technique for analyzing such multivariate datasets, which fulfills the above criteria, is analysis of variance (ANOVA) [7]. For the data of Sandberg *et al.*, the appropriate method is a two-way ANOVA because two variables - strain and region - were involved. ANOVA is a well understood and powerful method, but has two drawbacks that can make its application difficult in practice. First, when more than two categories are present for a variable, as in the case of the work of Sandberg *et al.*, where six brain regions were studied, additional work must be done to determine which categories underlie the effects observed. Another drawback of ANOVA is that in its standard form, the number of replicates of each condition should be equal. Adjustments can be made when small differences in the number of replicates occur, but large deviations are difficult to correct for.

We note that some workers have previously applied one-way ANOVA to expression data [8], and that the use of  $t$ -tests, which is one-way ANOVA for the limiting case of two replicated conditions, is becoming a more common method for analyzing expression data [6]. Kerr *et al.* [9] applied ANOVA to two-color cDNA arrays in an experiment involving a single biological variable (tissue type). Their model encompassed all the genes on the array, as well as experimental factors such as dye and array effects. Their complex model differs from the gene-specific application for experiments with more than one biological variable that we discuss here.

An alternative way of approaching the analysis problem is to view it as one of feature selection. In machine learning, feature selection seeks to identify the data points (here, expression profiles) that are most informative when trying to learn a classification problem. In the context of the data of Sandberg *et al.*, we are seeking genes that differentiate between the genotypes and/or brain regions in our samples. While learning to predict this classification is not relevant here, the underlying task is effectively one of feature selection. The features in this situation are the individual samples (arrays). A recent application of feature selection to microarray data was described by Golub *et al.* [10]. They sought to identify genes that best differentiated between two types of tumor. Because feature selection methods can be quite simple, can be used to differentiate between any

number of categories, and permit ranking genes in accordance with how well they differentiate between regions or strains, feature selection fulfills all three criteria for a suitable analysis method. A similar method was used by Chu *et al.* [11] to perform clustering: they created a set of idealized expression patterns based on the data from hand-selected genes and identified additional genes that were similar to the desired patterns.

Sandberg *et al.* used criteria based on absent/present calls and fold-change thresholds for their analysis, and applied multiple *t*-tests to each gene to detect additional regional variations in expression [1,12]. With these methods they identified 24 genes showing strain variation, and about 240 genes showing some type of regional variation. An examination of the data and experimental design suggests that methods such as ANOVA and feature selection could be profitably applied. Here we describe the identification of additional genes showing possible strain- and region-dependent patterns of expression from the data of Sandberg *et al.* We find that ANOVA and feature selection are complementary, and can be combined to form a powerful method for detecting interesting genes in a complex dataset.

## Results

The dataset of Sandberg *et al.* [1] consists of duplicate analysis of six brain regions (amygdala, cerebellum, cortex, entorhinal cortex, hippocampus and midbrain) in two strains of mice (129SvEv and C57BL/6), for 13,067 genes and ESTs. We performed two-factor ANOVA and feature selection to look for strain- and/or region-specific variation in gene expression in this data. Our feature-selection strategy, which we call 'template matching', is depicted schematically in Figure 1.

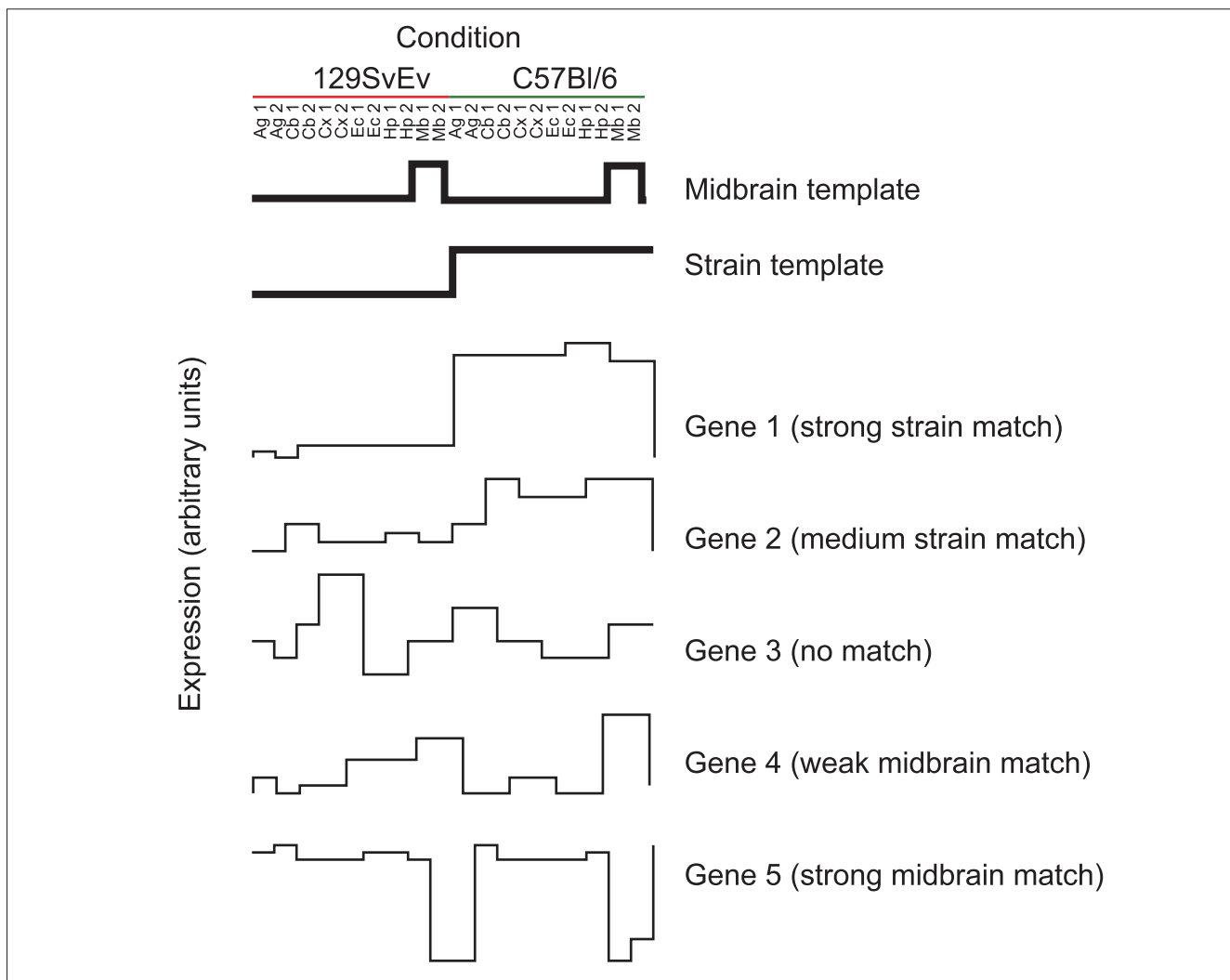
An overview of the results for both the ANOVA and template-matching methods is given in Table 1, which shows the number of genes identified by each method at three different *p*-value cut-offs. These cut-offs were selected to represent varying levels of false-positive risk (described in the Materials and methods section). The least stringent cut-off ( $p < 0.001$ ) is almost sure to yield significant numbers of false positives, whereas the most stringent ( $p < 10^{-6}$ ) will most probably yield very few false positives. The middle threshold of  $p < 10^{-5}$  is a value we use as a compromise for the purpose of comparing our results to those of the previous studies. Table 1 also shows the number of genes identified in the work of Sandberg *et al.* A more detailed view of part of the data is provided in Figure 2, which enumerates some genes identified by both methods, ranked by template-match *p*-value. Some (but not all) genes identified by Sandberg *et al.* [1] and Sandberg [12] are also given high scores (low *p*-values) by our methods, as indicated in Figure 2. As indicated in Table 1 and Figure 2, even at conservative *p*-value cut-offs we identified additional genes not noted in [1].

One goal of our study was to compare ANOVA and template matching, and their combination. In many cases the results from the two methods are comparable, and both methods give low *p*-values for many of the genes listed in Figure 2. An overall comparison of the strain-specific analysis using ANOVA and template matching is shown in Figure 3, where we plotted the *p*-values calculated for each gene by ANOVA against those calculated for the strain-specific template. Most points lie near the diagonal, showing that in general, the *p*-values agree well, though ANOVA has a tendency to give more low *p*-values and there are some significant outliers. In agreement with the trend illustrated in Figure 3, the number of genes that make each cut-off in Table 1 are very similar for strain specificity. The template-matching method, however, tends to give low *p*-values to many more genes on the basis of region-specificity. For all tissue effects, at a *p*-value of  $10^{-5}$ , the template method identifies more than twice as many genes as ANOVA, and at less conservative cut-offs the difference is even greater (Table 1). An example of a discrepancy is synuclein (GenBank C79089), which had a *p*-value of  $8.6 \times 10^{-7}$  based on a template match to the 'midbrain' pattern, but only 0.0011 for a region affect using ANOVA. The relatively poor ANOVA *p*-value appears to be due to a lower value for one of the four midbrain samples (Figure 2, last gene shown for 'Midbrain enriched'). Thus under some conditions ANOVA may be more conservative than template matching.

The combined data shown in Figure 2 takes advantage of both methods. The genes shown were selected on the basis of both ANOVA and template-match *p*-values (see Materials and methods). By listing both *p*-values, we can choose to pay the most attention to genes in which we have good confidence on the basis of two different methods. For example, the genes listed for 'cortex enriched or deficient' in Figure 2 have fairly unremarkable template-match *p*-values; the *p*-values from the generally more conservative ANOVA method are substantially worse. This leads us to suggest that few, if any, of these genes are likely to have 'real' expression patterns that are in good agreement with our expectation for a cortex-specific gene. This is borne out by the visualizations shown in Figure 2. Compared to the genes shown for other regions such as cerebellum or hippocampus, the 'cortex enriched or deficient' genes have less striking expression patterns. Thus the numerical ranking appears to be a reasonable guide to making judgments about the significance of the expression patterns.

### Strain-specific expression

At a *p*-value cut-off of  $10^{-5}$ , we found about 65 probe sets that show strain-specific variation in expression across all areas, compared to 24 genes showing overall differences between strains identified by Sandberg *et al.* [1]. The genes near the top of our ranking include many of those chosen by Sandberg [12], and all 24 genes identified by Sandberg *et al.* have *p*-values less than  $10^{-3}$  by either ANOVA, template



**Figure 1**

Schematic of the template match method. Hypothetical templates and genes are depicted as graphs of arrays (horizontal axis) vs expression (arbitrary units, vertical axis). The replicates for each condition are listed as 1 and 2 (for example, 129SvEv - Ag1 and 129SvEv - Ag2 are the two amygdala samples from 129SvEv). Two example templates are depicted schematically, and represent the profiles of idealized genes showing the patterns of interest. The top template is an idealized profile designed to select for genes showing midbrain enrichment or depletion, and the lower one is designed to select for genes showing strain differences across all regions. Each gene in the dataset is then ranked by its correlation with the template. In the hypothetical examples shown, gene 1 most closely resembles the strain template, with gene 2 giving the next best match. Gene 4 is a fairly weak match to the midbrain template. Gene 5 is a strong match to the reverse of the midbrain template; the use of the absolute value of the correlation coefficient allows the midbrain template to select both genes 4 and 5. Gene 3 resembled neither template shown here. In the actual implementation, all 13,067 genes are compared to a given template and ranked by the quality of the match. Ag, amygdala; Cb, cerebellum; Cx, cortex; Ec, entorhinal cortex; Hp, hippocampus; Mb, midbrain.

match, or both. Two examples of strain-specific genes not identified in the previous work are shown in more detail Figure 4a. The first is *Sparc/osteonectin* (testican), which was detected at lower levels in the C57Bl6 strain. The second is phosphatase *ACP1/ACP2*, with the opposite expression pattern. We rank *ACP1/ACP2* and *Sparc/osteonectin* as having the 9th and 23rd strongest strain differences, respectively (both *ACP1/ACP2* and *Sparc/osteonectin* are also shown in Figure 3). For comparison, Figure 4b shows

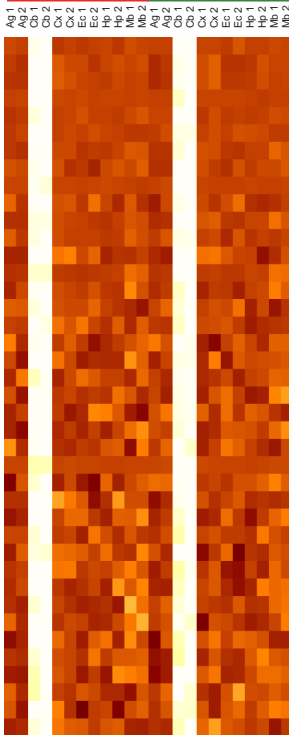
$\beta$ -globin, which was identified in [1] and was ranked 67th by template match. Overall, we rank 46 previously unrecognized genes as having stronger strain differences than  $\beta$ -globin.

Sandberg *et al.* [1] also identified an additional 49 genes that varied between strains for particular regions. We have not specifically searched for such genes; the ANOVA interaction effects would correspond most closely to such a category, but very few genes were found to have significant interaction

### Cerebellum-enriched

-2.00 2.00

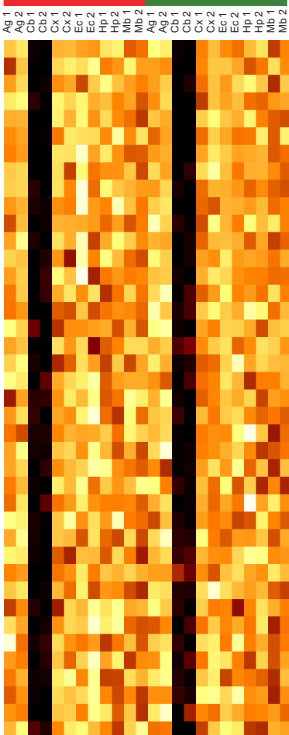
129SvEv C57Bl/6



Accession	Unigene	ANOVA	Template	Sb	Description
D83262	Mm.6257	9.24E-13	3.15E-33	•	Solute carrier family 1, member 6 (neuronal glutamate transporter EAAT4)
X70398	Mm.4919	2.59E-12	5.79E-27	•	DNA segment, human D4S114 <i>M.musculus</i> P311 mRNA.
X70398	Mm.4919	1.29E-11	2.72E-25	•	DNA segment, human D4S114 <i>M.musculus</i> P311 mRNA.
X51986	Mm.4915	2.80E-12	1.51E-24	•	Gamma-aminobutyric acid (GABA-A) receptor, subunit alpha 6
X61448	Mm.4880	6.96E-10	1.94E-24	•	Cerebellin 1 precursor protein
U28068	Mm.4636	9.55E-12	6.45E-23	•	Neurogenic differentiation 1
X59382	Mm.2766	1.68E-12	7.15E-23	•	Parvalbumin
AA289572	Mm.21099	3.31E-10	1.19E-22	•	ESTs
M21532	Mm.41456	6.61E-11	3.20E-22	•	Purkinje cell protein 2 (L7)
L16846	Mm.16596	5.33E-10	8.41E-21	•	B-cell translocation gene 1, anti-proliferative
X67141	Mm.2766	6.94E-11	6.72E-19	•	Parvalbumin
L12705	Mm.4298	2.21E-08	3.33E-18	•	Engrailed 2
M60596	Mm.41296	2.96E-12	5.23E-17	•	Gamma-aminobutyric acid (GABA-A) receptor, subunit delta (exon 9)
D32167	Mm.2719	9.73E-18	6.09E-17	•	Zinc finger protein of the cerebellum 1
AA444931	Mm.30079	8.04E-09	9.63E-17	•	Unclassifiable transcript (RIKEN cDNA 2010012F07 gene)
M21531	Mm.354	2.07E-08	1.88E-16	•	Calbindin-28K
L16846	Mm.16596	4.77E-08	3.33E-16	•	B-cell translocation gene 1, anti-proliferative
U33630	Mm.4734	8.89E-09	8.70E-16	•	Myeloid ecotropic viral integration site 1
M69068	Mm.33263	1.71E-08	4.63E-15	•	Histocompatibility 2, D region locus 1
AA472865	Mm.24719	9.31E-08	5.71E-15	•	ADP-ribosylation-like factor 6 interacting protein 2
AA473309	Mm.29910	1.05E-08	7.05E-15	•	Ribosomal protein S6 kinase polypeptide 1
X98014	Mm.5173	1.60E-08	9.65E-15	•	Sialyltransferase 8 (alpha 2, 8 sialyltransferase) E
AA008502	Mm.7414	6.61E-10	1.01E-14	•	Neuron specific gene family member 1
U91483	Mm.15343	2.80E-07	1.61E-14	•	Calsequestrin 2
X61397	Mm.156583	2.04E-09	2.62E-14	•	Carbonic anhydrase-like sequence 1
D13266	Mm.57057	6.90E-08	2.65E-14	•	Glutamate receptor, ionotropic, delta 2
Z38118	Mm.4564	1.60E-08	3.28E-14	•	Synaptonemal complex protein 1
AA059527	Mm.7414	1.67E-08	5.11E-14	•	Neuron specific gene family member 1
W41032	Mm.181894	1.97E-06	1.21E-13	•	hypothetical protein (RIKEN cDNA 2900092E17 gene) major vault protein
AA390043	Mm.28765	4.11E-07	1.41E-13	•	homolog to TRAM protein. (RIKEN cDNA 1810049E02 gene)
X69063	Mm.4789	2.69E-09	1.61E-13	•	Ankyrin 1, erythroid
X13605	Mm.18516	5.84E-07	2.07E-13	•	H3 histone, family 3B
L22144	Mm.829	2.36E-08	2.16E-13	•	S100 protein, beta polypeptide, neural
AA270913	Mm.3116	2.00E-08	2.17E-13	•	ESTs
R74735	Mm.752	3.40E-07	2.20E-13	•	ESTs
W45964	Mm.28524	1.46E-06	3.97E-13	•	DNA segment, Chr 6, Wayne State University 116, expressed
X51438	Mm.7	4.92E-08	4.70E-13	•	Vimentin
L02241	Mm.1499	1.66E-06	5.09E-13	•	Protein kinase inhibitor beta, cAMP dependent, testis specific
M90365	Mm.21990	5.88E-08	6.17E-13	•	Junction plakoglobin
X15373	Mm.2726	3.42E-07	8.48E-13	•	Inositol 1,4,5-triphosphate receptor 1

### Cerebellum-deficient

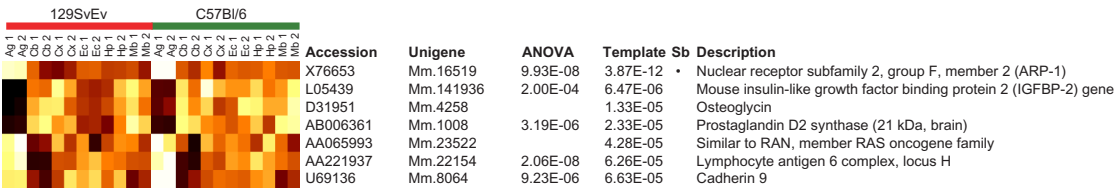
129SvEv C57Bl/6



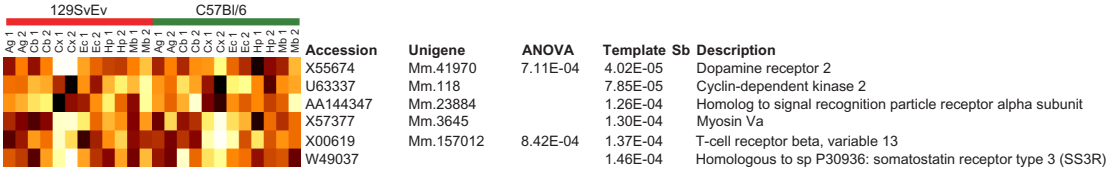
Accession	Unigene	ANOVA	Template	Sb	Description
AA220788	Mm.2388	1.15E-09	1.99E-10	•	ESTs
AA259350	Mm.21606	1.28E-06	2.32E-10	•	<i>M. musculus</i> reticulon 3 (Rtn3) mRNA, complete cds
U23184	Mm.31395	1.41E-05	2.52E-10	•	Carboxypeptidase E
L04538	Mm.2381	7.64E-07	3.66E-10	•	Amyloid beta (A4) precursor-like protein 1
AA031158	Mm.29586	4.32E-07	3.89E-10	•	Homologous to sp P13983: extensin precursor (cell wall hyd)
W57404		7.32E-05	4.46E-10	•	Homologous to sp P16870: carboxypeptidase H precursor
Z31269	Mm.29586	6.82E-07	9.61E-10	•	Homolog to brain acid soluble protein 1 (BASP1 protein)
X03151	Mm.3951	3.26E-05	1.96E-09	•	Thymus cell antigen 1, theta
X04663		4.41E-05	5.80E-09	•	Mouse mRNA for beta-tubulin (isotype Mbeta 5).
AA059763	Mm.31239	2.24E-05	1.01E-08	•	Similar to tubulin, beta 4
AA238056	Mm.25154	5.77E-05	1.04E-08	•	Wilms' tumour 1-associating protein
X04663	Mm.1703	1.42E-04	1.11E-08	•	Tubulin, beta 5
AA050852	Mm.1260	9.82E-06	1.18E-08	•	Expressed in non-metastatic cells 1, protein (NM23A)
W49178	Mm.1703	3.66E-04	4.48E-08	•	Tubulin, beta 5
AA271360	Mm.26680	7.58E-05	5.03E-08	•	ESTs
W40709	Mm.29650	1.44E-06	5.32E-08	•	Similar to mitochondrial carrier homolog 1 isoform B.
AA002366	Mm.2381	1.19E-05	5.63E-08	•	Amyloid beta (A4) precursor-like protein 1
W11020	Mm.29455	8.66E-05	6.06E-08	•	Homolog to HSPC035 protein
AA155529	Mm.2411	3.12E-06	6.45E-08	•	Ras-GTPase-activating protein (GAP<120>) SH3-domain-binding protein 2
AA271327	Mm.28422	1.07E-05	6.82E-08	•	ESTs, Moderately similar to autosomal highly conserved protein [ <i>H.sapiens</i> ]
AA259350		2.65E-05	8.47E-08	•	
L42463	Mm.1383	4.85E-06	8.68E-08	•	Rho GDP dissociation inhibitor (GDI) gamma
J04695	Mm.181021	1.01E-05	9.17E-08	•	Procollagen, type IV, alpha 2
L04538		1.36E-05	9.61E-08	•	Mouse amyloid precursor-like protein mRNA, complete cds.
D67016	Mm.34828	7.46E-05	1.02E-07	•	Heat shock protein, 105 kDa
AA217585	Mm.22025	6.62E-04	1.05E-07	•	Homolog to KIAA0290 (fragment)
AA032557	Mm.21251	2.91E-04	1.09E-07	•	Deleted in polyposis 1
M20632	Mm.1129	6.77E-05	1.51E-07	•	Repeat family 3 gene
L04280	Mm.70127	1.10E-04	1.55E-07	•	Ribosomal protein L12
M19380	Mm.116941	2.04E-06	2.06E-07	•	Calmodulin 3
AA167917	Mm.181860	1.10E-05	2.12E-07	•	Similar to tubulin, beta 4
W76777	Mm.24488	2.28E-06	2.56E-07	•	PH domain containing protein
AA408475	Mm.13020	1.21E-04	2.78E-07	•	Ribosomal protein L13a
AA689048	Mm.27923	1.09E-05	2.81E-07	•	ESTs, weakly similar to 3BP1 mouse SH3-binding protein 3BP-1
C77776	Mm.5650	2.45E-04	2.84E-07	•	ESTs
Z31298	Mm.27897	3.83E-04	2.89E-07	•	DnaJ (Hsp40) homolog, subfamily A, member 1
U28217	Mm.4654	4.18E-06	3.05E-07	•	Protein tyrosine phosphatase, non-receptor type 5
D83206	Mm.4766	3.59E-07	3.62E-07	•	Vesicular membran protein p24
C77386	Mm.24953	2.26E-04	4.04E-07	•	ESTs, weakly similar to DNA polymerase zeta catalytic subunit
U58887	Mm.736	2.86E-06	4.14E-07	•	SH3 domain protein 2 C1

Figure 2 (see the legend on the page after next)

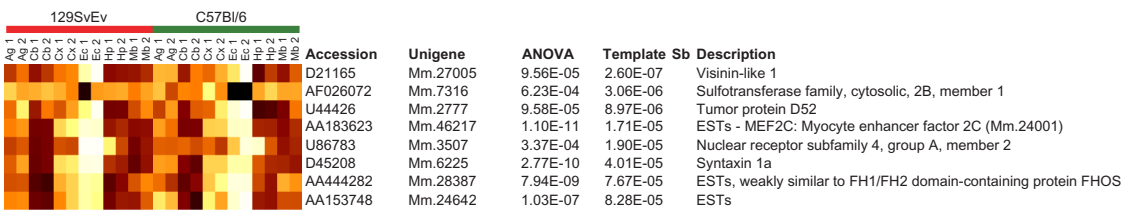
### Amygdala-enriched or deficient



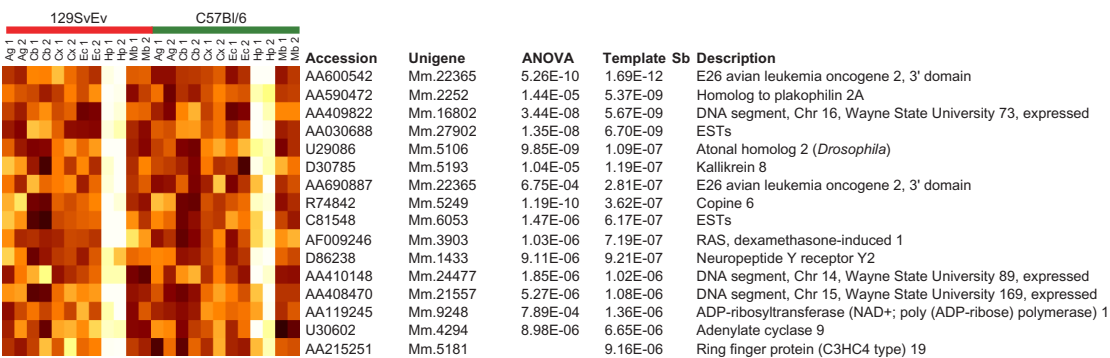
### Cortex-enriched or deficient



### Entorhinal cortex-enriched or deficient



### Hippocampus-enriched or deficient



### Midbrain-enriched

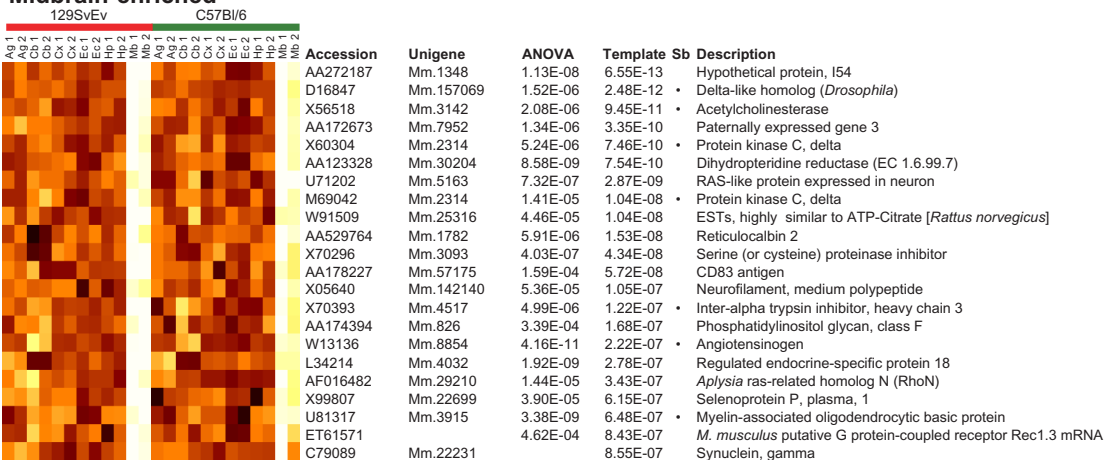
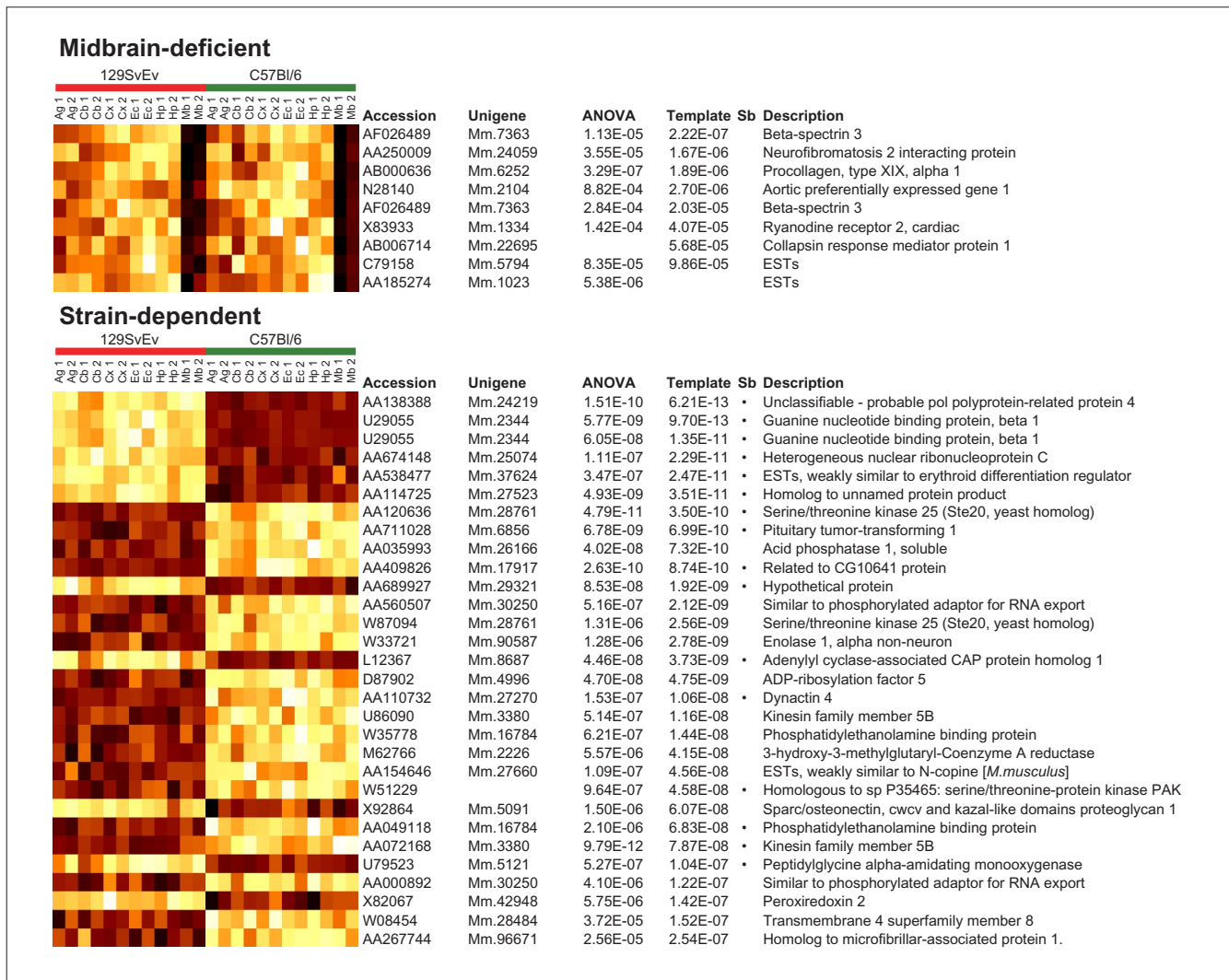


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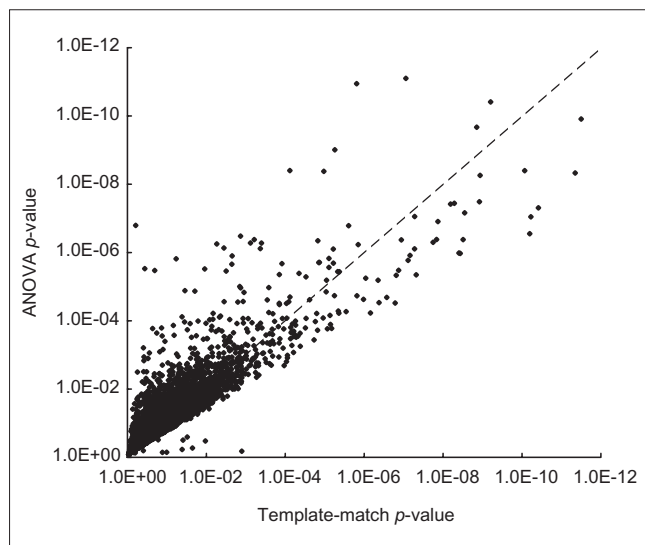


**Figure 2** (continued from the previous pages)

Visualizations and detailed information about some genes identified in this study. For each region or for the strain distinction, up to 40 genes are shown (for space considerations), all of which have template-match  $p$ -values  $< 10^{-4}$ . The genes are ranked by template-match  $p$ -value. Other rankings are possible on the basis of our data. The  $p$ -value cut-off ( $10^{-4}$ ) was chosen so that some genes would be listed for all categories; note that the 'best' genes in some categories are only comparable in specificity to the 'worst' genes in others. For two categories (cerebellum and midbrain), genes showing enrichment and depletion expression in the area of interest are separated. The results have been filtered to remove genes having primarily negative average difference values as described in Materials and methods. Besides the ANOVA and template-match  $p$ -values ( $p$ -values greater than  $10^{-4}$  are not displayed), for each gene, the GenBank accession number of the sequence used to derive the probes is shown, along with the UniGene cluster number (if applicable), and annotations drawn primarily from UniGene and FANTOM databases [29]. UniGene cluster identity and annotations were based on the current information rather than the annotations originally provided by Affymetrix and may differ from those described in Sandberg *et al.* [1]. A bullet appears in the column Sb (Sandberg) if the gene was identified in Sandberg *et al.* [1] or Sandberg [12]. The need to limit the number of genes shown may give the misleading impression that relatively few new genes were identified, especially for the cerebellum-enriched genes. On the basis of accession number, UniGene ID, and annotation, some genes appear more than once on the same list. This duplication reflects multiple probe sets targeting the same sequence, or multiple probe sets targeting different sequences in the same UniGene cluster. The visualizations were prepared as described in Materials and methods, and we emphasize that the scale shown is normalized and thus arbitrary for each gene. Abbreviations are the same as for Figure 1. The complete lists of genes with annotations and  $p$ -values can be obtained from our website [32].

effects (Table 1). Preliminary results (see Materials and methods: Availability of data and software) using templates that select specifically for genes expressed in one region in

one strain suggest that this strategy will be useful for identifying more complex expression patterns as well as the simple ones we describe in detail here.



**Figure 3**  
Comparison between template match and ANOVA  $p$ -values for strain-specific changes. The  $p$ -values for all 13,067 genes are plotted, with template-match  $p$ -values on the  $x$ -axis and ANOVA  $p$ -values on the  $y$ -axis. For most genes, the two methods are in good agreement.

### Region-specific expression

As with the strain-specific genes, we identify a number of putative region-specific genes in addition to those identified by Sandberg *et al.* [1]. The two methods lead to an estimate that, at a  $p$ -value of  $10^{-5}$ , up to approximately 600 genes show variability of expression with respect to brain region (Table 1), or around 6% of the total, compared to the 242 region-specific genes identified by Sandberg [12] and Sandberg *et al.* [1]. As for the strain-specific genes, some genes identified in [1] do not receive particularly high scores in our assays (Figure 2, and see Additional data files). For example, both hippocampal genes identified in [1] have  $p$ -values greater than 0.001 and do not appear in Figure 3. In general, however, most of the genes selected in [1] appear fairly high on our lists, although additional genes are often given even higher scores by our assays.

In agreement with Sandberg *et al.* [1], we find that the cerebellum is by far the most distinct of the brain regions studied. The bulk of the genes identified as showing region-specific effects appear to be due to cerebellum enrichment or

depletion (339 out of 595 at template-match  $p < 10^{-5}$ ). The midbrain was the second most distinct single region (Table 1). In contrast, the cortex had the least distinct expression pattern, being very similar to other forebrain regions, as pointed out by Sandberg *et al.* When considered together, the forebrain (represented by cortex, amygdala, hippocampus and entorhinal cortex) has many genes that distinguish it from the midbrain and cerebellum (Table 1). Taken separately, the cortex, amygdala, hippocampus and entorhinal cortex had relatively few genes with distinct patterns (Table 1, Figure 2). The single amygdala-specific gene (ARP-1) identified by Sandberg *et al.* appears at the top of our list of amygdala-specific genes, and although we identify a few other high-scoring genes in that category, none approaches the specificity of ARP-1 (Figure 2). Details of a few examples of high-scoring genes we identified for hippocampus and midbrain are shown in Figure 4c and d, respectively.

In many cases, the expression patterns we report seem to be novel, but in others supporting evidence is found in the literature. We did not attempt a comprehensive literature survey for each gene, but support was readily found for the hippocampus enrichment of copine 6 [13], neuropeptide Y receptor Y2 [14], Atonal 2 [15], and kallikrein 8 (neuropsin) [16]. The same can be said for the midbrain enrichment of Rin [17], Peg3 [18], and nexin [19]. Some cerebellar-enriched genes for which we found supporting evidence include calsequestrin [20,21] and aldolase 3 (zebrin) [22].

### Discussion

Our most important finding is that by using ANOVA and feature selection at even quite conservative cut-offs, we identify genes with strain- and region-dependent variation in gene expression that were not identified in the original analysis of the data [1]. These newly identified genes expand the range of genes that can be considered as candidates for underlying behavioral differences between the two strains and functional differences among brain regions. We were able to confirm the expression patterns of some of the genes we identified by examining the published literature. Although this examination was not exhaustive, we readily found supporting evidence for the tissue-specific expression patterns of a number of genes which were not identified in the previous surveys of the data. These results also show that the new patterns we identify with high confidence

**Figure 4** (see figure on the next page)

Unadjusted graphs of expression for examples of genes identified in this study. Arrays are plotted along the  $x$ -axis, and expression (average difference) is plotted on the  $y$ -axis. Below the  $x$ -axis of each graph is a schematic of the corresponding template. **(a)** Two typical examples of high-scoring genes showing strain-specific variation. **(b)** A gene identified by Sandberg *et al.* as having strain-specific variation, for comparison with (a). We gave this gene a much lower ranking than the genes shown in (a). **(c)** Three high-scoring genes showing hippocampus enrichment. **(d)** Four high-scoring genes showing midbrain enrichment.



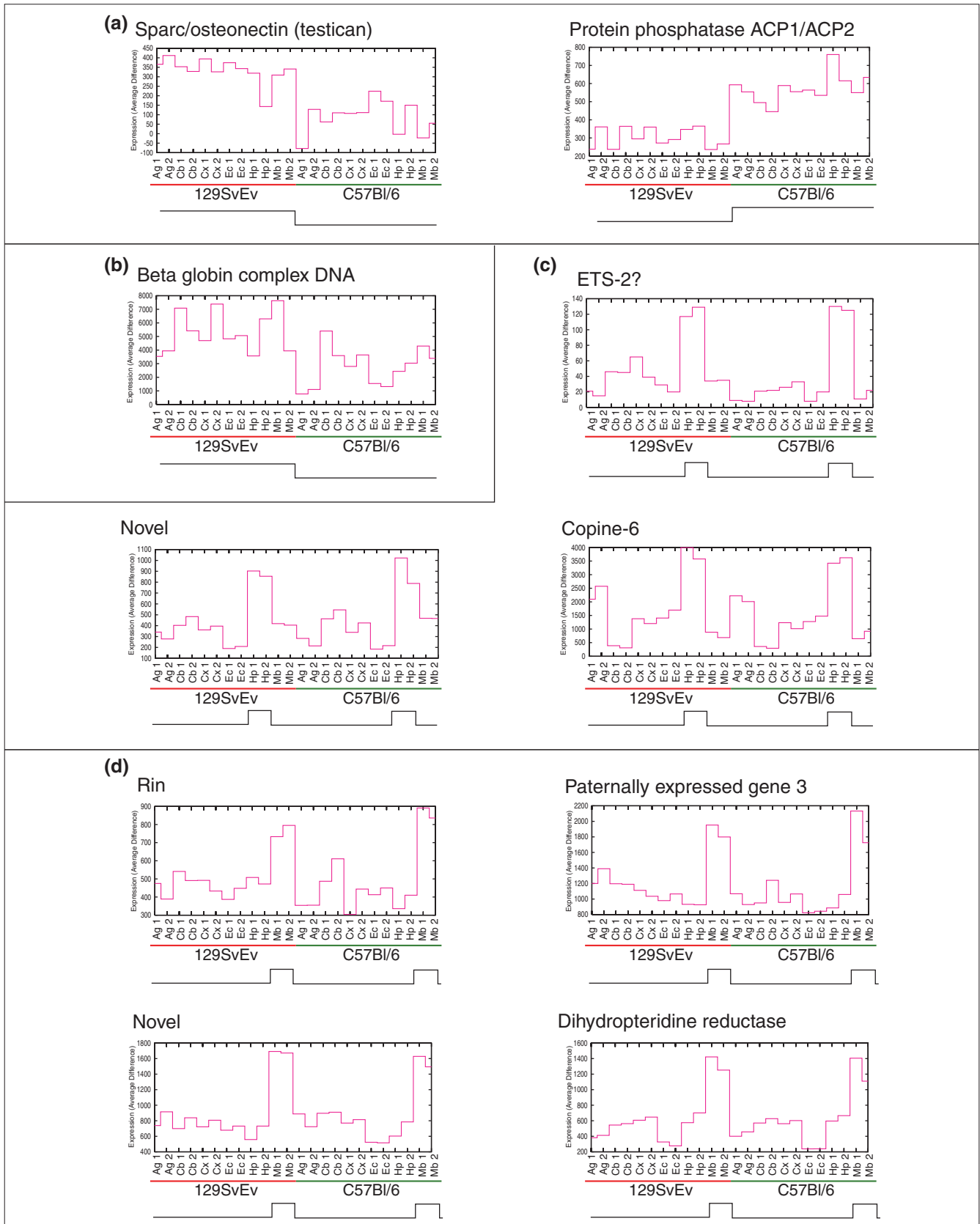


Figure 4 (see legend on the previous page)

**Table 1****Comparison of numbers of genes identified as showing strain and/or regional variation in gene expression for different methods**

Condition	Method	Number of genes selected at each <i>p</i> -value		
		<i>p</i> < 0.001	<i>p</i> < 10 <sup>-5</sup>	<i>p</i> < 10 <sup>-6</sup>
Strain	Sandberg	24	24	24
	ANOVA	213	65	36
	Template	150	56	34
Strain x region	Sandberg	49	49	49
	ANOVA	17	1	0
All strain effects	Sandberg	73	73	73
	ANOVA	230	66	36
	Template	150	56	34
Amygdala	Sandberg	1(3)	1(3)	1(3)
	Template	74	2	0
Hippocampus	Sandberg	2(10)	2(10)	2(10)
	Template	63	16	11
Cortex*	Sandberg	0	0	0
	Template	35	0	0
Forebrain*	Sandberg	58	58	58
	Template	771	191	112
Midbrain	Sandberg	13	13	13
	Template	181	44	23
Cerebellum	Sandberg	150	150	150
	Template	890	339	227
Entorhinal cortex	Sandberg	2(10)	2(10)	2(10)
	Template	37	3	1
All tissue effects†	Sandberg	242	242	242
	ANOVA	874	289	168
	Template	2051	595	374

The number of genes identified by each method, and by Sandberg *et al.* [1], for each condition or set of conditions is listed if appropriate. For ANOVA and template match, values for three different *p*-value thresholds are shown. For the various brain regions, the template-match counts include both matches (enriched genes) and anti-matches (genes expressed at lower levels in the region listed). None of the values given is corrected for genes dominated by negative average difference values. After filtering such genes (see Materials and methods) values would be about 5-10% lower. The values in the 'Sandberg' rows are the same in each column because Sandberg *et al.* did not generate confidence measures for the genes they selected. The totals for 'Sandberg' include genes listed both in Sandberg *et al.* [1] and Sandberg [12], and group together genes that were 'restricted', 'enriched', 'absent' and 'decreased'. Numbers in parentheses are the totals obtained if one includes genes mentioned in the text of [1] but not listed by name. \*Genes specific to the cortex alone. Because Sandberg *et al.* do not appear to have specifically sought genes expressed only in the cortex (instead they compared cortex only to cerebellum and midbrain [12]), the number of genes in this category for Sandberg *et al.* are listed as zero. 'Forebrain' is the combination of cortex, hippocampus, entorhinal cortex and amygdala, which most closely matches the 'Cortex' group chosen in [1,12]. It is included here to allow a fair comparison to their methods. †This category does not include strain x region effects.

are unlikely to be spurious statistical artifacts, and give an indication that the sensitivity (the ability to avoid missing significant results) of our methods is higher than the *ad hoc* approach. We also note that our criteria are not necessarily less stringent than those of Sandberg *et al.* as many genes they selected do not rank particularly high in our results. Furthermore, in some cases, such as for the entorhinal cortex at more conservative *p*-value cut-offs, we select fewer genes than Sandberg *et al.* even though at the same cut-off for other regions we select more (Table 1). Thus, our increase in sensitivity does not come at the cost of selectivity (the ability to avoid spurious positive findings) compared to the *ad hoc* approach. In the remaining discussion, we evaluate our methodology in more detail, and then consider the biological significance of the results.

**Evaluation of methodology**

The difference in the genes we selected compared to Sandberg *et al.* [1] seems to depend on two factors: ignoring the absent/present calls, and using statistical methods that provide a score for each gene, rather than simply choosing to assign genes to the group of 'changed' genes. Sandberg *et al.* appear to have used criteria that were conservative given the measures used (that is, three out of four absent/present calls must be 'present' to consider a gene expressed). But, the fact that the standard ANOVA often does not support the evaluation based on absent/present calls supports the idea that absent/present calls, even if used conservatively, are neither specific nor sensitive enough.

There are two particular drawbacks to using absent/present calls for this type of analysis. Most important, each absent/present call is based on only a single array. If one has only a single array, such a guideline for the ability to detect a given gene may be useful. When multiple measurements are made, however, as in the dataset considered here, there is no set method for evaluating the data. Is a gene 'present' in the experiment when it is 'present' on three of four arrays, or are two out of four arrays sufficient? Instead of such *ad hoc* criteria, we used statistical comparisons of the samples. This means that low-magnitude signals, which might have been called 'absent', can turn out to be reliable and meaningful when several observations are made. A second problem with the absent/present calls is that they are based on essentially arbitrary thresholds, with no simple estimate of the risk of false positives or false negatives. The use of methods that provide *p*-values allows a much finer-grained analysis. For these reasons, we feel that absent/present calls should generally not be used to analyze experiments where replicates were performed.

Despite the advantages of ignoring absent/present calls, the 'absolute difference' expression value provided by the Affymetrix software is not ideal. In particular, large negative values occur at a much higher rate than expected (data not shown). These negative values are probably due to problems

with mismatch probes, as described in the Materials and methods section. A more thorough analysis would remove, or otherwise correct for, probe pairs containing such problematic probes from the raw data, which we did not have access to here. At least two groups have described alternative methods that start with the 'perfect match/mismatch' pair data [23,24], which may be better at identifying and eliminating such troublesome data points.

The region-specific genes listed in Figure 2 reflect a combination of the two methods we used. This combination takes advantage of desirable features of both methods: ANOVA can detect any effect of region on expression, whereas template matching looks specifically for one pattern at a time. By sorting ANOVA hits into regions based on template-match *p*-values, we identify some genes that were missed by template matching. In general, this is caused by the ability of ANOVA to detect significant expression patterns that do not fit any particular template very well, but when examined manually reveal interesting patterns related to the template. We note that other methods of *post hoc* testing ANOVA data exist [7]. A benefit of the template-match method is that it is very simple to apply, and we believe we found some patterns that would be fairly difficult to find using other methods. Thus we consider the template-match method a useful complement to standard *post hoc* tests.

In the search for region-specific genes, ANOVA tended to be more conservative than template matching. This is probably because ANOVA retains information about the experimental design; for the example of synuclein given in the results, it 'knows' that there is a large disagreement between two matched replicates in the midbrain samples, whereas the template match based on the correlation coefficient considers all variations to be essentially equivalent. The fact that there are only four samples for each tissue, while there are twelve for each strain, means that tests for tissue-specific genes are even more susceptible to the effect of noise in the data. This helps explain why the discrepancies between template match and ANOVA was larger for the region-specific genes than for strain-specific genes.

The methods we describe have some limitations. ANOVA loses power when the number of replicates is not the same for all the conditions, and the experiment must also follow a complete block design, where all possible combinations of factors are tested. Even in cases where these conditions are not adhered to, multiple *t*-tests (which are generally proscribed in multivariate statistics in favor of ANOVA) would be preferable to resorting to *ad hoc* methods for assessing statistical significance of the results. The template-match method has the potential drawback that if used alone, one would miss potentially interesting patterns that do not resemble the template(s) being tested. This may be acceptable if there is a clear idea of what is being sought in the experiment, but is possibly detrimental in studies

relying on exploration and serendipity to find patterns. The ANOVA/template-match combination ameliorates this problem as one can choose to be guided primarily by the ANOVA *p*-value (which is not pattern specific) while using the template to reduce the complexity of the analysis by partitioning the data into groups based on interesting patterns.

### Biological significance

Our results resemble those of Sandberg *et al.* [1] in that many of the genes given high scores have previously defined roles in the function of the nervous system. At the same time, many of the high scoring genes have 'housekeeping' roles or more general cellular functions that are not specific to the nervous system. Currently this conclusion is based on a manual examination of the annotations and the literature associated with top-scoring genes. In the future, as annotations improve, we hope to be able to perform such assessments quantitatively and objectively. Clearly, the ranked list of genes we provide can only act as an input to further studies of nervous system function, and does not provide any ready answers to questions about behavioral or functional differences.

As already indicated, another general conclusion from our study is that the variation between regions and strains appears to be greater than originally indicated by the analysis of Sandberg *et al.* In a commentary on the original study, Geschwind [25] sought explanations for the apparent paucity of region-specific and strain-specific gene expression. Although the sensitivity of the microarray assay itself may be at issue, our results clearly show that the use of different computational analysis methods has a dramatic effect on the interpretation of the data. Geschwind [25] also expressed surprise that the cerebellum, which is histologically more homogeneous than the other brain regions assayed, was discovered to have the largest number of uniquely expressed genes. This interpretation of the data assumes that the other regions such as cortex and hippocampus, being more complex and heterogeneous, should have more differences between each other as well as from the cerebellum. However, because hippocampus, neocortex and entorhinal cortex are all cortical structures with many cell types in common, it is not surprising that they have similar gene expression profiles. In contrast, the cerebellum has a distinct developmental origin, as well as a unique cellular composition, compared to all the other regions assayed. Thus it is not at all surprising that the cerebellum is the most distinct region. In agreement with this line of logic, the midbrain, which is histologically and developmentally distinct from the cortex and cerebellum, was the next most distinct region. We also note that, contrary to the suggestion of Geschwind, the fact that the cerebellum had the largest number of 'uniquely expressed genes' is not equivalent to stating that the cerebellum has the highest molecular complexity [25]. To assess this type of complexity, the total number of mRNA species expressed in the tissue would have to be measured, but such information is not readily



ones and zeroes in the template to be reversed for any particular comparison. We note that templates that contain continuous values instead of binary ones could be applied in the same way, allowing one to search for specific patterns of greater complexity, but binary vectors were sufficient for the purposes of this study. High correlation values correspond to a good match. The  $p$ -values were calculated using a  $t$  test on the correlation coefficient [30]. We prepared templates corresponding to each of the regions assayed by Sandberg *et al.*, as well as using the strain-distinguishing template shown in the example above.

### Analysis of variance (ANOVA) and combining ANOVA with feature selection

A standard linear two-way analysis of variance [7,30] was performed independently on the data for each gene. The effects examined for each gene were strain, region, and interactions between strain and region. This is expressed by the following linear equation:

$$E_{ijk} = \mu + R_i + S_j + (R \cdot S)_{ij} + \varepsilon_{ijk} \quad \begin{cases} i = 1,2\dots6 \\ j = 1,2 \\ k = 1,2 \end{cases}$$

where  $E_{ijk}$  is the measured expression level of the gene in replicate  $k$  from one of six regions  $i$  in one of two strains  $j$ ,  $\mu$  is the overall mean expression level of the gene, and  $R$  and  $S$  are effects due to region and strain, respectively. The  $R \cdot S$  term represents interactions between strain and region, which might be detected if a gene is expressed only in specific regions in only one strain, for example. The error term ( $\varepsilon$ ) accounts for random variation from replicate to replicate; that is, any effect not accounted for by the other terms in the model. Thus this model assumes that the measured expression level of a gene can be described as a linear combination of effects due to measurement error, strain, region, and interactions between strain and region. The null hypothesis is that a given factor does not have an effect on expression, meaning that the corresponding term in the equation is zero. The extent to which the fitted parameters  $R$ ,  $S$  and  $R \cdot S$  for a gene are non-zero provides the basis of the ANOVA. We generated  $p$ -values for each of the three effects for each gene using a standard F-test [7].

As discussed, ANOVA does not allow one to directly distinguish between effects due to particular levels of one parameter. Specifically, ANOVA might indicate that there was a significant effect of region on expression, but does not directly reveal which regions show different expression from any others. Thus, one needs to apply a *post hoc* test to the ANOVA results to determine which regions, if any, show differences from others. Typically this type of multiple comparison is done using statistical tests such as the Tukey or Scheffé methods [7]. In our study, each region must be compared to each other region, so 15 comparisons must be made in total for each gene. In practice, however, interpretation of

these results is more complicated: an additional stage of analysis must be performed to generate, for example, genes expressed at higher levels in the cerebellum than in all other regions, as opposed to any one other region (as is given by Tukey or Scheffé methods). For our study we used a simple strategy suggested by the idea of template matching: we choose genes fitting a pattern in which we are interested from a subset of genes with low ANOVA  $p$ -values for region effects. Specifically, we first selected all genes which had region-effect ANOVA  $p$ -values  $< 0.001$ , and then applied template matching to this set to select genes which fit a particular pattern as described above, at a  $p$ -value of 0.0001 or less. These arbitrary, and fairly lax, thresholds were used because they proved to give reasonable numbers of genes in each partition. The resulting groups of genes can then be ranked according to either their ANOVA  $p$ -value or by the correlation with the template.

### Interpretation of $p$ -values

An important issue in interpreting the  $p$ -values generated by tests on single genes is the effect of multiple tests. Specifically, a test performed with a chosen  $p$ -value cut-off of 0.01, repeated 13,000 times, will on average generate 130 false positives if each test is independent. Corrections for this multiple testing problem are well known in statistics but are somewhat controversial [31]. Methods used by Tusher *et al.* [4] and Callow *et al.* [5] attempt to provide appropriate corrections for microarray experiments, but there is, as yet, no consensus on which methods to apply. For the purposes of presenting our data here we report results at three different thresholds, which would, by the Bonferroni criteria, yield on average about 13, 0.13, or 0.013 experiment-wide false positives under the assumptions of our tests (that is,  $p$ -values of 0.001,  $10^{-5}$  and  $10^{-6}$ , respectively). Because these cut-offs assume that each gene is independent, which is clearly not the case, these false positive estimates are probably conservative with respect to the multiple testing issue. Because we assume here that the standard F and  $t$  distributions are adequate for describing the statistics we use, however, using a conservative cut-off provides some additional protection from type I errors due to potential inaccuracy of the uncorrected  $p$ -values.

### Filtering selected genes

All the analysis described above was performed without adjusting or filtering the raw average difference values. But, a difficulty one commonly encounters with Affymetrix data is that many probe sets yield large negative expression (average difference) measures. Large negative expression values occur when one or more mismatch primers in a probe set have much higher intensities than their corresponding perfect match primers. This may occur because the mismatch primer is binding to some other nucleotide in the sample which does not represent the gene for which a measurement is intended. Because the Affymetrix software determines the presence or absence of expression of a

gene in part on the basis of how many probe pairs give high intensities, not just the average across the probe pairs, some genes may be flagged as present despite having negative average differences. Obviously this is not desirable; one would like to correct for problematic mismatch probes before calculating the average difference. A stop-gap measure, which we used here as we did not have access to the original raw data files, is to remove from consideration entire probe sets which have this problem. Specifically, genes with more than 6 values out of 24 that were less than -20 were rejected; in practice this method is quite conservative, removing only the most problematic genes. We note that a few genes (about 20 of approximately 240) selected by Sandberg *et al.* would be removed from consideration by our filter.

### Comparison with the results of Sandberg *et al.* [1]

A list of the genes identified in the previous work on this dataset were obtained from Tables 1 and 2 of Sandberg *et al.* [1] and Table 4 of Sandberg [12]. A few groups of genes referred to in these studies are not specifically identified in the publications: the selection of eight hippocampus-enriched, two amygdala-decreased, five entorhinal cortex-enriched and three entorhinal cortex-decreased genes is mentioned, but the genes are not specifically identified in [1]. From our other comparisons, it is fairly likely that some of these genes were given high scores in our analysis, so our comparison to the results of Sandberg *et al.* for those areas should be regarded as incomplete.

### Visualization of the data

In several of our figures, microarray data is displayed as a matrix of rectangles, each corresponding to an individual measurement of one probe set/gene (rows) in one sample (columns). The color, using a 'black-body' scale (from low values represented by black to dark red to orange to yellow to high values represented by white) indicates the relative level of expression of the gene across the different conditions. The values for each gene have been adjusted to a mean of zero and a variance of one to facilitate comparisons between genes, and values > 1 (or < -1) have been adjusted to 1 (or -1), increasing the contrast of the images.

### Availability of software

The software used to create the visualizations of the data and to perform template matching and ANOVA is available on request from the authors.

### Additional data files

The following text files containing the results of our analyses are available with this article online:

Anova results (the first column is the probe name and the final results are listed as "ff" (first factor) "sf" (second factor) and fxsf (interaction); the first factor is strain, the second is

region. The numbers in those columns are F values; the other columns are intermediate values used to calculate F).

Template match results (each file contains three columns. The first column is the probe name, the second is the correlation coefficient with the template. The third column is the *p*-value): cerebellum; hippocampus; midbrain; cortex; fore-brain; entorhinal cortex; amygdala; strain.

All data and results, including detailed browseable tables of the selected genes, and supplementary data and information, are also available from our website [32].

### Acknowledgements

We thank Anthony Ferrante for comments on an early version of the manuscript. W.S.N. is funded by an Award in Bioinformatics from the PhRMA Foundation, and by National Science Foundation grants DBI-0078523 and IIS-0093302.

### References

- Sandberg R, Yasuda R, Pankratz DG, Carter TA, Del Rio JA, Wodicka L, Mayford M, Lockhart DJ, Barlow C: **Regional and strain-specific gene expression mapping in the adult mouse brain.** *Proc Natl Acad Sci USA* 2000, **97**:11038-11043. [<http://ftp.gnf.org/pub/papers/brainstrain/>]
- Lipshutz RJ, Fodor SP, Gingeras TR, Lockhart DJ: **High density synthetic oligonucleotide arrays.** *Nat Genet* 1999, **21**(Suppl):20-24.
- Crawley JN, Belknap JK, Collins A, Crabbe JC, Frankel W, Henderson N, Hitzemann RJ, Maxson SC, Miner LL, Silva AJ, *et al.*: **Behavioral phenotypes of inbred mouse strains: implications and recommendations for molecular studies.** *Psychopharmacology (Berl)* 1997, **132**:107-124.
- Tusher VG, Tibshirani R, Chu G: **Significance analysis of microarrays applied to the ionizing radiation response.** *Proc Natl Acad Sci USA* 2001, **98**:5116-5121.
- Callow MJ, Dudoit S, Gong EL, Speed TP, Rubin EM: **Microarray expression profiling identifies genes with altered expression in HDL-deficient mice.** *Genome Res* 2000, **10**:2022-2029.
- Long AD, Mangalam HJ, Chan BY, Tolleri L, Hatfield GW, Baldi P: **Improved statistical inference from DNA microarray data using analysis of variance and a Bayesian statistical framework. Analysis of global gene expression profiling in *Escherichia coli* K12.** *J Biol Chem* 2001, **276**:19937-19944.
- Sahai H, Agell M: *The Analysis of Variance*. Boston: Birkhauser, 2000.
- Hill AA, Hunter CP, Tsung BT, Tucker-Kellogg G, Brown EL: **Genomic analysis of gene expression in *C. elegans*.** *Science* 2000, **290**:809-812.
- Kerr MK, Martin M, Churchill GA: **Analysis of variance for gene expression microarray data.** *J Comput Biol* 2000, **7**:819-837.
- Golub TR, Slonim DK, Tamayo P, Huard C, Gaasenbeek M, Mesirov JP, Coller H, Loh ML, Downing JR, Caligiuri MA, *et al.*: **Molecular classification of cancer: class discovery and class prediction by gene expression monitoring.** *Science* 1999, **286**:531-537.
- Chu S, DeRisi J, Eisen M, Mulholland J, Botstein D, Brown PO, Herskowitz I: **The transcriptional program of sporulation in budding yeast.** *Science* 1998, **282**:699-705.
- Sandberg R: *Gene expression profiling of brain regions in inbred mouse strains reveals candidate genes for phenotypic variation.* Masters Thesis in Biomedicine. Stockholm: Karolinska Institute; 2000.
- Nakayama T, Yaoi T, Kuwajima G: **Localization and subcellular distribution of N-copine in mouse brain.** *J Neurochem* 1999, **72**:373-379.
- Gackenhaimer SL, Schober DA, Gehlert DR: **Characterization of neuropeptide Y Y1-like and Y2-like receptor subtypes in the mouse brain.** *Peptides* 2001, **22**:335-341.
- Schwab MH, Druffel-Augustin S, Gass P, Jung M, Klugmann M, Bartholomae A, Rossner MJ, Nave KA: **Neuronal basic helix-**

- loop-helix proteins (NEX, neuroD, NDRF): spatiotemporal expression and targeted disruption of the NEX gene in transgenic mice.** *J Neurosci* 1998, **18**:1408-1418.
16. Chen ZL, Yoshida S, Kato K, Momota Y, Suzuki J, Tanaka T, Ito J, Nishino H, Aimoto S, Kiyama H, et al.: **Expression and activity-dependent changes of a novel limbic-serine protease gene in the hippocampus.** *J Neurosci* 1995, **15**:5088-5097.
  17. Lee CH, Della NG, Chew CE, Zack DJ: **Rin, a neuron-specific and calmodulin-binding small G-protein, and Rit define a novel subfamily of ras proteins.** *J Neurosci* 1996, **16**:6784-6794.
  18. Li L, Keverne EB, Aparicio SA, Ishino F, Barton SC, Surani MA: **Regulation of maternal behavior and offspring growth by paternally expressed Peg3.** *Science* 1999, **284**:330-333.
  19. Kury P, Schaeren-Wiemers N, Monard D: **Protease nexin-1 is expressed at the mouse met-/mesencephalic junction and FGF signaling regulates its promoter activity in primary met-/mesencephalic cells.** *Development* 1997, **124**:1251-1262.
  20. Volpe P, Gorza L, Brini M, Sacchetto R, Ausoni S, Clegg DO: **Expression of the calsequestrin gene in chicken cerebellum Purkinje neurons.** *Biochem J* 1993, **294**:487-490.
  21. Takei K, Stukenbrok H, Metcalf A, Mignery GA, Sudhof TC, Volpe P, De Camilli P: **Ca<sup>2+</sup> stores in Purkinje neurons: endoplasmic reticulum subcompartments demonstrated by the heterogeneous distribution of the InsP3 receptor, Ca<sup>(2+)</sup>-ATPase, and calsequestrin.** *J Neurosci* 1992, **12**:489-505.
  22. Hawkes R, Herrup K: **Aldolase C/zebrin II and the regionalization of the cerebellum.** *J Mol Neurosci* 1995, **6**:147-158.
  23. Schadt EE, Li C, Su C, Wong WH: **Analyzing high-density oligonucleotide gene expression array data.** *J Cell Biochem* 2000, **80**:192-202.
  24. Efron B, Tibshirani R, Goss V, Chu G: *Microarrays and their use in a comparative experiment.* Stanford: Stanford University Department of Statistics, 2000.
  25. Geschwind DH: **Mice, microarrays, and the genetic diversity of the brain.** *Proc Natl Acad Sci USA* 2000, **97**:10676-10678.
  26. Zirlinger M, Kreiman G, Anderson DJ: **Amygdala-enriched genes identified by microarray technology are restricted to specific amygdaloid subnuclei.** *Proc Natl Acad Sci USA* 2001, **98**:5270-5275.
  27. **NCBI-UniGene** [<http://www.ncbi.nlm.nih.gov/UniGene/>]
  28. **FANTOM: Functional Annotation of Mouse** [<http://www.gsc.riken.go.jp/e/FANTOM/>]
  29. Kawai J, Shinagawa A, Shibata K, Yoshino M, Itoh M, Ishii Y, Arakawa T, Hara A, Fukunishi Y, Konno H, et al.: **Functional annotation of a full-length mouse cDNA collection.** *Nature* 2001, **409**:685-690.
  30. Zar JH: *Biostatistical Analysis.* 4th edn. Upper Saddle River, NJ: Prentice Hall, 1999.
  31. Westfall PH, Young SS: *Resampling-based Multiple Testing.* New York: John Wiley & Sons, 1993.
  32. **Mouse brain data** [<http://www.cs.columbia.edu/compbio/brainexp/index.html>]