Seasonality of gestational diabetes mellitus and maternal blood glucose levels

Evidence from Taiwan

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Abstract

Previous studies have indicated that the prevalence of gestational diabetes mellitus (GDM) was related to the season. However, there was no relevant information in Asia. The aim of this study was to determine whether there was seasonality of GDM and maternal blood glucose level in Taiwanese women.

A total of 6396 pregnancies were enrolled between 2012 and 2014 in this retrospective study. A 2-step approach according to the Carpenter-Coustan criteria was used for GDM diagnosis. A generalized linear mixed model was used to estimate the effect of season on GDM diagnosis by adjusting for age, prepregnancy body mass index, parity, history of GDM, fetal sex, and the rate of weight gain.

During the study period, 418 (6.5%) pregnancies were diagnosed as GDM. The model demonstrated an increased prevalence of GDM in spring and summer (odds ratio: 1.59, 95% confidence interval: 1.13–2.24; odds ratio: 1.59, 95% confidence interval: 1.14–2.23, respectively) compared to winter. For the glucose level variation, the model demonstrated an increase of 2.56 mg/dL glucose in the 50-g glucose challenge test in summer compared to winter. In glucose challenge test-positive pregnancies, the season also had an effect on the results of the 100-g 1-h, 2-h, and 3-h oral glucose tolerance tests, but no effect on the 100-g fasting oral glucose tolerance tests.

GDM prevalence in Taiwan presents seasonal variation, with the highest risk during spring and summer due to post-glucose load level variations. These findings could serve as reference data for countries in Southeast Asia or areas with a similar climate.

Abbreviations: BMI = body mass index, CI = confidence interval, DMF-CYCH = Ditmanson Medical Foundation Chia-Yi Christian Hospital, GCT = glucose challenge test, GDM = gestational diabetes mellitus, OGTT = oral glucose tolerance test, OR = odds ratio.

Keywords: Asia, gestational diabetes mellitus, glucose levels, season

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PW and C-SW contributed equally to this work.

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The data that support the findings of this study are available from a third party, but restrictions apply to the availability of these data, which were used under license for the current study, and so are not publicly available. Data are available from the authors upon reasonable request and with permission of the third party.

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1. Introduction

Gestational diabetes mellitus (GDM) is defined as diabetes diagnosed in the second or third trimester of pregnancy that is not clearly overt diabetes that was present prior to gestation.^[1] GDM is an important medical and public health issue and is associated with adverse pregnancy outcomes. About 2.7% to 13.2% of pregnant women develop GDM during their pregnancies and the prevalence has increased considerably during the last decade.^[2-4] Even the prevalence of GDM in high-risk populations reaches 25.7% in the second trimester and more than 30% in the third trimester.^[5,6] If left untreated or undiscovered, perinatal morbidity, and mortality risk of mothers increase. Maternal diabetes during the pregnancy is strongly linked to adverse pregnancy outcomes, also at higher risk of subsequent development of type 2 diabetes.^[7,8] There is evidence showing that exposure to maternal diabetes in utero has long-term adverse effects on the offspring, which likely occurs due to epigenetic modifications of the fetal genome, and could be averted by therapies applied early during pregnancy.^[9] Therefore, it is important to realize the relevant factors associated with GDM, especially for the high-risk populations.

Medicine

Differences in screening programs and diagnostic criteria may have resulted in different incidence and prevalence of GDM among various populations. Nevertheless, ethnicity has been proven to be an independent risk factor for GDM.^[10] Women from Asia are a heterogeneous group by genetic background, culture, and lifestyle, such as diet. Previous studies showed that Southeast Asian women were more likely to have GDM, even with a lower (BMI).^[2,11–14] From a pregnancy-outcome point of view, Southeast Asian women have lower rates of macrosomia compared with non-Hispanic White women, but preterm delivery with preeclampsia occurs more often compared with Japanese, and non-Hispanic White women.^[15] Additionally, Southeast Asian women with GDM have a higher risk of developing type 2 diabetes compared with Caucasians.^[16] A recent review reported a pooled GDM prevalence of 10% in Eastern and Southeast Asia. However, a difference in GDM prevalence can still be observed across countries.^[17] Therefore, establishing appropriate diagnostic criteria of GDM and proper management strategies for Asian women with GDM is necessary.

Apart from the well-known risk factors, other parameters are also discussed in the literature, such as Vitamin D and temperature. A few interventional studies indicated that supplementation optimizes maternal vitamin D status or improves maternal glucose metabolism.^[18,19] Vitamin D is the product of sun rays, which might explain the relationship between high temperature and GDM. However, observational studies about the association between maternal vitamin D status and risk of GDM are conflicting.^[20,21] In the 1990s, Schmidt et al reported that ambient temperature may affect venous glucose concentrations after glucose tolerance tests.^[22] However, subsequent research has shown that GDM is not seasonal.^[23,24] Until recently, studies from different countries, including Sweden, Australia, Italy, and Greece, indicated that the prevalence of GDM was related to the season, with higher risk during summer.^[25–28]

The prevalence of GDM varies among races and ethnicities, and climates vary around the world. The GDM risk factors have been explored in many research reports, and season is a factor that has been proposed in recent years to be associated with GDM. However, there has been no study on the relationship between GDM and season in Asian women. Therefore, the aim of this study is to determine whether there is a seasonal variation in GDM diagnosis and maternal blood glucose levels. The hypothesis was that the prevalence of GDM is related to the season in Taiwanese women.

2. Materials and methods

2.1. Materials

This retrospective study involved collecting the laboratory data and medical records of pregnant women without overt diabetes who underwent a 50-g 1-h glucose challenge test (GCT) at 24 to 30 weeks of gestation between January 2012 and December 2014 and delivered at the Ditmanson Medical Foundation Chia-Yi Christian Hospital (DMF-CYCH). Women with multifetal pregnancies, preexisting diabetes, and chronic/preexisting or unspecified hypertension were excluded. This study was approved by the Institutional Review Board of the DMF-CYCH (CYCH IRB No: 100006). We collected data from medical records; thus, the committee agreed that the informed consent of each participant was not necessary. Plasma glucose levels were measured using a Hitachi 7170 automatic analyzer (Hitachi Co., Tokyo, Japan) at the DMF-CYCH central laboratory according to a standard clinical protocol.

2.2. GDM diagnosis

The included women underwent the 2-step approach for GDM screening. The 50-g 1-h GCT was considered positive if the screening value was \geq 140 mg/dL. If the GCT was positive, the women were referred for a 100-g 3-h oral glucose tolerance test (OGTT) to confirm whether they had GDM. Women were

diagnosed with GDM when 2 or more plasma glucose level results exceeded or equaled 95, 180, 155, or 140 mg/dL in the fasting, 1-, 2- or 3-h plasma glucose tests, respectively (the Carpenter-Coustan criteria).^[29]

2.3. Statistical analyses

Normality was assessed by using the Shapiro-Wilk test. The differences between the season groups and the characteristics of the pregnancies were analyzed using Kruskal-Wallis tests for the continuous variables because of nonnormal distributions and are presented as medians (interquartile ranges). The GDM rate between the season groups and the characteristics of the pregnancies were analyzed using Wilcoxon rank sum tests for the continuous variables because of nonnormal distributions and are presented as medians (interquartile ranges). Continuous variables with normal distributions were compared by Student t test and are presented as means \pm standard deviations. The Chisquared test was used for the categorical variables, which are presented as counts (percentages). Pregnant women with more than 1 pregnancy during the study period were not excluded. Therefore, we applied generalized linear mixed models to estimate the effects of different seasons on GDM prevalence and the glucose levels by adjusting for maternal age at GCT, prepregnancy BMI, parity, history of GDM,^[10,30-32] fetal sex,^[33] and the rate of weight gain from prepregnancy to GDM screening (50-g GCT).^[34-36] In the mixed models, all covariates were included as fixed effects. We considered correlation of data from the same individual as a random effect that incorporated an unstructured variance-covariance matrix. The associations were described in terms of adjusted odds ratios (ORs) with 95% confidence intervals (CIs) for binary outcomes and estimates with 95% CIs for continuous glucose values. We managed to consider "Maternal age at GCT" and "Prepregnancy BMI" as continuous variables or categorical variables in different models in analyzing the relationship between season and GDM. The months of March to May were defined as spring; June to August were defined as summer; September to November were defined as autumn; and December to February were defined as winter. BMI was calculated as the ratio of body weight (in kilograms) divided by the square of height (in meters). According to the Bureau of Health Promotion, Department of Health, Taiwan, a BMI of <18.5 kg/m² denotes underweight, 18.5 to <24 kg/m² normal weight, 24 to $<27 \text{ kg/m}^2$ overweight, and $\geq 27 \text{ kg/m}^2$ obesity (https://www.hpa.gov.tw/Pages/Detail.aspx?nodeid=542&pid= 9734). The rate of weight gain was calculated by dividing weight gain from prepregnancy to GDM screening by the corresponding number of weeks. The average monthly temperature during the study period was obtained from the nearby fixed-site Chia-Yi monitoring station, operated by the Taiwan Environmental Protection Administration, which measured the weather data hourly throughout Taiwan (https://tagm.epa.gov.tw/tagm/tw/ default.aspx). We calculated the daily average temperature and matched these data with the women's screening dates. A 2-sided P value < .05 was considered statistically significant. All data were merged and analyzed using SAS 9.4 (SAS Institute, Cary, NC).

3. Results

A total of 6809 pregnancies occurred in patients without overt diabetes who underwent a 50-g 1-h GCT at 24 to 30 weeks of gestation between January 2012 and December 2014 and





delivered at the DMF-CYCH; 287 pregnancies in patients with multifetal pregnancies, preexisting diabetes, or chronic/preexisting or unspecified hypertension were excluded. A total of 126 subjects did not complete the OGTTs and were excluded. Finally, 5923 women with 6396 pregnancies were enrolled in the study. Among them, 468 (7.9%) women had multigravida during the study period, and the pregnancies of these women were analyzed separately. During this period, 418 (6.5%) pregnancies were associated with GDM (Fig. 1).

Table 1 shows the associations between the season and the characteristics of the pregnancies. The pregnancies screened during the winter and spring were associated with an increased rate of weight gain (P < .001). There were no significant correlations between the season and the other characteristics.

Table 2 shows the characteristics of the GDM and non-GDM groups. In univariate analysis, the overall GDM prevalence was 6.5%. The highest rate of GDM occurred in summer, and the lowest rate occurred in winter (7.4% vs 4.8%, P = .02). Women with increasing age, increasing prepregnancy BMI, or a history of GDM all had a significant higher rate of being diagnosed with GDM (all P < .001). Multiparous women also were more likely to develop GDM (7.3% vs 5.7%, P = .01). A male fetus tended to cause GDM more often than female (7.1% vs 6.0%, P = .07). The rate of weight gain had no difference between GDM and non-GDM pregnancies, both having a median of 0.29 kg/wk. In multivariate analysis, after adjusting for maternal age (continuous), BMI (continuous), parity, history of GDM, fetal sex, and rate of weight gain, the model demonstrated an increased prevalence of pregnancies associated with GDM in spring and summer (OR: 1.59, 95% CI: 1.13-2.24; OR: 1.59, 95% CI: 1.14-2.23, respectively) compared to winter in model 1. In model 2, when maternal age and prepregnancy BMI were classified as categorical variables, the relationship between GDM, and season was stronger. The rate of weight gain carried a significant risk for GDM; a weight gain increase of 1 kg/wk made a GDM diagnosis 2.97 and 2.57 more likely in model 1 and model 2, respectively.

Figure 2 shows the trend of GDM rate and temperature by month. As the temperature rose and as the standard deviation of temperature grew smaller, the GDM rate was higher. In contrast, as the temperature got lower and the standard deviation rose, the GDM rate was lower. The highest prevalence of GDM was noted in August (8.1%), and the lowest one was in February (4.7%).

Table 3 shows the abnormal glucose rates according to season. The mean \pm standard deviation of temperature in summer was 29.0 \pm 1.3°C, and in winter it was 18.3 \pm 2.7°C. The highest positive rate of 50-g GCT occurred in summer, and the lowest rate occurred in winter (27.3% vs 23.0%, P = .06). After adjusting for confounding factors, the model demonstrated an increased prevalence of pregnancies associated with abnormal 50-g GCT in spring and summer (OR: 1.22, 95% CI: 1.02–1.47; OR: 1.27, 95% CI: 1.06–1.53, respectively) compared to winter. In GCT-positive pregnancies, the highest abnormal rates and adjusted ORs of the 100-g 1-h, 2-h and 3-h OGTTs occurred in spring (27.4%, P = .08; OR: 1.55, 95% CI: 1.04–2.31), spring (32.8%, P = .07; OR: 1.70, 95% CI: 1.170–2.48), and autumn (19.2%, P = .01; OR: 2.00, 95% CI: 1.23-3.24), respectively, but the results of the 100-g fasting OGTT showed no difference according to season (P = .98).

Table 4 shows the association between season and oral glucose level. Compared to other seasons, winter had a lower median glucose in the 50-g GCT (P = .01), 100-g OGTT at 2 h (P = .04), and 100-g OGTT at 3 h (P = .02) but a higher median in the 100-g OGTT at fasting (P = .003) than other seasons. After adjusting for confounding factors, the model demonstrated an increase of 2.56 mg/dL glucose in the 50-g GCT in summer compared to winter. In GCT-positive pregnancies, the season had no effect on

Table 1

The association between season and the characteristics of the pregnancies.

			Sea	ason		
Variable	Total	Spring	Summer	Autumn	Winter	Р
Total number	6396	1613	1779	1643	1361	
Maternal age at GCT (yr)	6396	30.2 (27.1–33.3)	30.2 (27.2-33.4)	30.5 (27.1–33.8)	30.3 (27.3-33.4)	.39
<25	859	220 (25.6)	229 (26.7)	224 (26.1)	186 (21.7)	.26
25–30	2193	564 (25.7)	639 (29.1)	528 (24.1)	462 (21.1)	
30–35	2311	587 (25.4)	631 (27.3)	591 (25.6)	502 (21.7)	
≥35	1033	242 (23.4)	280 (27.1)	300 (29.0)	211 (20.4)	
Prepregnancy BMI (kg/m ²)	6136	21.2 (19.2–23.4)	21.1 (19.2–23.8)	21.0 (19.2–23.3)	20.9 (19.2–23.5)	.57
<18.5	970	246 (25.4)	278 (28.7)	263 (27.1)	183 (18.9)	.83
18.5–24	3822	995 (26.0)	1081 (28.3)	1026 (26.8)	720 (18.8)	
24–27	791	199 (25.2)	247 (31.2)	192 (24.3)	153 (19.3)	
≥27	553	133 (24.1)	161 (29.1)	153 (27.7)	106 (19.2)	
Parity						
Nulliparous	3039	770 (25.3)	871 (28.7)	780 (25.7)	618 (20.3)	.27
Multiparous	3357	843 (25.1)	908 (27.1)	863 (25.7)	743 (22.1)	
50-g GCT (weeks)	6396	26.0 (25.1-27.3)	26.0 (25.0-27.1)	26.1 (25.0-27.4)	26.1 (25.0-27.3)	.12
100-g OGTT (wk)	1635	27.9 (26.7-29.3)	27.9 (26.7-29.4)	28.0 (26.7-29.3)	28.0 (26.7-29.4)	.98
History of GDM						
Yes	90	21 (23.3)	21 (23.3)	23 (25.6)	25 (27.8)	.46
No	6306	1592 (25.3)	1758 (27.9)	1620 (25.7)	1336 (21.2)	
Fetal sex						
Boy	3343	846 (25.3)	929 (27.8)	868 (26.0)	700 (20.9)	.93
Girl	3041	766 (25.2)	844 (27.8)	775 (25.5)	656 (21.6)	
Missing	12					
Rate of weight gain (kg/wk)*	6115	0.30 (0.22-0.38)	0.28 (0.20-0.36)	0.28 (0.21-0.36)	0.31 (0.23-0.39)	<.001
Missing	281					

Categorical data are presented as n (%) and were compared by the Chi-squared test; continuous data are presented as median (interquartile range) and were compared by the Kruskal–Wallis test. BMI=body mass index, GCT=glucose challenge test, GDM=gestational diabetes mellitus, OGTT=oral glucose tolerance test.

[®] Rate of weight gain was calculated by dividing weight gain from prepregnancy to gestational diabetes mellitus screening by the corresponding number of weeks.

Table 2

The correlates of gestational diabetes mellitus in univariate and multivariate analysis.

			Univariate			Multiva	riate	
		GDM di	agnosis		Model 1		Model 2	
Variable	Total	Yes (n=418)	No (n=5978)	Р	OR (95% CI)	Р	OR (95% CI)	Р
Intercept					0.0004 (0.0004-0.001)	<.001	0.06 (0.03-0.13)	<.001
Season								
Spring	1613	115 (7.1)	1498 (92.9)	.018	1.59 (1.13-2.24)	.008	1.63 (1.16-2.29)	.005
Summer	1779	132 (7.4)	1647 (92.6)		1.59 (1.14-2.23)	.007	1.64 (1.18-2.29)	.004
Autumn	1643	106 (6.5)	1537 (93.6)		1.38 (0.98-1.95)	.07	1.40 (0.994-1.98)	.054
Winter	1361	65 (4.8)	1296 (95.2)		1		1	
Maternal age at GCT (yr)	6396	33.0 ± 4.6	30.1 ± 4.7	<.001	1.13 (1.11–1.16)	<.001		
<25	859	20 (2.3)	839 (97.7)	<.001			1	
25–30	2193	93 (4.2)	2100 (95.8)				1.72 (1.05-2.84)	.032
30–35	2311	159 (6.9)	2152 (93.1)				2.88 (1.78-4.68)	<.001
≥35	1033	146 (14.1)	887 (85.9)				5.94 (3.63–9.73)	<.001
Prepregnancy BMI (kg/m ²)	6136	23.1 (20.9-25.7)	20.9 (19.2-23.3)	<.001	1.13 (1.10–1.15)	<.001	· · · · ·	
<18.5	970	30 (3.1)	940 (96.9)	<.001			1	
18.5–24	3822	218 (5.7)	3604 (94.3)				1.57 (1.06-2.34)	.025
24–27	791	86 (10.9)	705 (89.1)				3.06 (1.97-4.74)	<.001
≥27	553	74 (13.4)	479 (86.6)				4.18 (2.64-6.61)	<.001
Parity								
Nulliparous	3039	174 (5.7)	2865 (94.3)	.013	1.26 (1.01-1.58)	.042	1.21 (0.97-1.51)	.099
Multiparous	3357	244 (7.3)	3113 (92.7)				· · · · ·	
History of GDM				<.001				
Yes	90	40 (44.4)	50 (55.6)		1.24 (1.001-1.53)	.05	1.24 (1.01-1.54)	.045
No	6306	378 (6.0)	5928 (94.0)		1		1	
Fetal sex		. ,						
Boy	3343	236 (7.1)	3107 (92.9)	.074	0.10 (0.06-0.16)	<.001	0.09 (0.06-0.15)	<.001
Girl	3041	181 (6.0)	2860 (94.1)		1		1	
Rate of weight gain $(kg/wk)^*$	6115	0.29 (0.21-0.37)	0.29 (0.21–0.37)	.869	2.97 (1.39-6.34)	.005	2.57 (1.20-5.49)	.015

Categorical data are presented as n (%) and were compared by the Chi-squared test; continuous data with a normal distribution are presented as mean ± standard deviation and were compared by Student *t* test; data without a normal distribution are presented as median (interquartile range) and were compared by the Wilcoxon rank sum test.

BMI = body mass index, CI = confidence interval, GCT = glucose challenge test, GDM = gestational diabetes mellitus, OR = odds ratio.

[®] Rate of weight gain was calculated by dividing weight gain from prepregnancy to gestational diabetes mellitus screening by the corresponding number of weeks.



the results of the 100-g fasting OGTT, but it did have an effect on the results of the 100-g 1-h, 2-h, and 3-h OGTTs.

4. Discussion

During the study period, 6.5% of the pregnancies were associated with a diagnosis of GDM. In this retrospective study, we observed clear seasonality in the prevalence of GDM by using a 2-step approach for GDM diagnosis: an increased GDM prevalence occurred among pregnancies in spring and summer compared to winter due to post-glucose load level variations after adjusting for maternal age, BMI, parity, history of GDM, fetal sex, and rate of weight gain. The GDM prevalence peak occurred in spring and summer, and the lowest prevalence was during winter. The degree of variation in the glucose level in the 50-g GCT in summer was significantly higher than in winter. In GCT-positive pregnancies (either abnormal glucose rates or glucose levels), the season had an effect on the results of the 100-g 1-h, 2-h, and 3h OGTTs, but it did not have effect on the results of the 100-g fasting OGTT.

Advanced maternal age is an independent risk factor for GDM.^[37] Our study has the concordant result. After controlling for other risk factors (such as BMI and parity), women older than 35 years old had 6-time odds ratio to be diagnosed with GDM compared with those who were less than 25 years old. Recently, a meta-analysis demonstrates that the risk of GDM increases linearly with successive age-groups.^[38] Subgroup analyses also indicated that from the age of 25, Asian women had a significantly higher risk of developing GDM than Europid women. On the other hand, the average age of childbearing over the past 4 decades has been increasing.^[39] This may be the reason for the incidence of GDM increased year by year.^[2-4] Therefore, it is important to realize the relevant factors associated with GDM, especially for the high-risk populations.

Recent studies have indicated that the prevalence of GDM is related to the season^[25-27] and that GDM prevalence increases</sup>during the summer compared to winter. The diagnosis of GDM in this study was in line with those results. A large multicenter study including 11,538 pregnant women in southern Sweden reported that summer was associated with an increased frequency of GDM compared to all other months (OR 1.51, 95% CI 1.24-1.83) and that there was seasonal variation in the 2-h glucose concentration in the OGTT (P < .001); the fasting glucose levels were unaffected by temperature.^[26] In the same year, another

	Temperature, $^\circ extsf{C}$	5	50-g GCT				100-	3 OGTT			
		ΛI	140 mg/dL	Fastinç	g, ≥ 95 mg/dL	1-h, ∠	- 180 mg/dL	2-h, 2	2 155 mg/dL	3-h,	≥ 140 mg/dL
Season	mean±SD	(%) u	OR (95% CI)	(%) u	OR (95% CI)	(%) u	OR (95% CI)	(%) u	OR (95% CI)	(%) u	OR (95% CI)
Spring	24.5 ± 3.5	418 (25.9)	1.22 (1.02–1.47)*	42 (10.2)	0.95 (0.55–1.66)	113 (27.4)	1.55 (1.04–2.31)*	135 (32.8)	1.70 (1.17–2.48)	64 (15.5)	1.69 (1.03–2.79)*
Summer	29.0 ± 1.3	485 (27.3)	1.27 (1.06–1.53)**	50 (10.3)	0.94 (0.55-1.62)	129 (26.4)	1.54 (1.04–2.28)*	150 (30.7)	1.56 (1.08–2.26)*	85 (17.4)	1.86 (1.15–3.01)*
Autumn	25.9 ± 3.0	419 (25.5)	1.17 (0.97–1.41)	42 (9.9)	0.88 (0.51-1.52)	108 (25.5)	1.36 (0.91–2.03)	129 (30.5)	1.45 (1.00–2.11)	81 (19.2)	2.00 (1.23–3.24)
Winter	18.3 ± 2.7	313 (23.0)		34 (10.9)		61 (19.6)	-	75 (24.0)	-	32 (10.3)	-
Ρ		.06		98.		.08		.07		.01	

CI = confidence interval, GCT = glucose challenge test, OGTT = oral glucose tolerance test, OR = ords ratio, SD = standard deviation. and rate of weight gain.

P<.01. P<.001 P<.05.

Table 3

		50-0	GCT				UOF	-u OCTT		
		6-0C	100				001			
	Noi	nfasting	Ľ	asting		1-h		2-h		3-h
Season	Median (IQR)	Estimate (95% CI)	Median (IQR)	Estimate (95% CI)	Median (IQR)	Estimate (95% CI)	Median (IQR)	Estimate (95% CI)	Median (IQR)	Estimate (95% CI)
Spring	121 (104–140)	1.21 (-0.79, 3.22)	84 (79–89)	-0.61 (-1.95, 0.74)	161 (141–182)	4.22 (-0.49, 8.92)	141 (121–162)	6.94 (2.17, 11.70)**	113 (97–130)	5.11 (0.72, 9.51)*
Summer	123 (106–142)	2.56 (0.60, 4.52)*	83 (79–88)	-0.86 (-2.15, 0.43)	161 (141–182)	4.62 (0.08, 9.16)*	142 (124–161)	7.30 (2.70, 11.90)**	115 (97–133)	7.02 (2.79, 11.25)**
Autumn	121 (104–140)	1.54 (-0.45, 3.52)	82 (78–88)	-1.29 (-2.61, 0.03)	159 (140–177)	3.34 (-1.29, 7.98)	138 (121-158)	5.68 (0.98, 10.39)*	110 (96–130)	6.23 (1.91, 10.55)**
Winter	119 (103–138)	0	85 (80–89)	0	161 (139–177)	0	135 (120–154)	0	110 (93–125)	0
Ρ	.01		.003		.48		.04		.02	

Data are presented as median (interquartile range) and were compared by the Kruskal–Vallis test. Estimates for glucose levels were adjusted for maternal age at 50-g glucose challenge test (continuous), prepregnancy body mass index (continuous), prepresented as the test of the test of test (continuous), prepresented as te and rate of weight gair mellitus, fetal sex,

ratio CI = confidence interval, GCT = glucose challenge test, IQR = interquartile range, OGTT = oral glucose tolerance test, OR = odds

P<.05. *P*<.01. *P*<.001.

study in a coastal city of Australia found that the prevalence of GDM according to the 1-h or 2-h diagnostic test was higher in summer and lower in winter than the overall prevalence.^[25] Recently, Chiefari et al^[27] found that the prevalence of GDM, according to the Italian Health Ministry guidelines, in Italian patients was also significantly higher in summer than in spring, autumn and winter. Regarding glucose levels, no differences were observed between the median fasting glycemia values among the 4 seasons; rather, serum glucose levels according to the 1-h and 2h OGTTs in summer were higher than those in other seasons.^[27] These results indicated that seasonal variation affected the GDM diagnosis rate due to post-glucose load level variation and not to the fasting OGTT measurement. This study had similar findings on the post-glucose load data in the 50-g 1-h GCT and 100-g 1-h, 2-h, and 3-h OGTT.

A 50-g 1-h GCT is the first step in the 2-step GDM diagnosis approach. Only 1 previous study conducted in Israel has investigated the seasonality of GCT results.^[15] Wainstock et al found that the GCT results in winter were independently associated with the lowest risk of pathological GCT values compared to all other seasons. In our study, winter was also associated with the lowest median GCT value and lowest risk for abnormal rates of GCT. Taiwan is in a subtropical region, where the temperature varies less between the 4 seasons than in Israel, with a Mediterranean climate, but there is still seasonal variation (Fig. 2), and the diagnostic GCT-positive rates in spring and summer were significantly higher than that in winter (ORs: 1.22 and 1.27, respectively). Considering only the glucose level, a GCT during summer had significantly different effect compared to winter.

Taiwan is an island on the border of 2 climatic zones. The northern part of the island has a subtropical climate with wet summers and dry, cool winters. The southern part has a monsoon tropical climate, which experiences an extraordinarily rainy wet season (summer) and pronounced dry season (winter). Our study was conducted in the monsoon climate in southern Taiwan, where the average monthly temperature is higher (17.4°C to 29.6°C, see Fig. 2) than that of other countries where similar research topics have been studied (e.g., UK,^[24] Australia,^[25] Sweden,^[26] Greece).^[28] In our study area, for every degree increase in temperature, the likelihood of a diagnosis of GDM increased by 3% (OR 1.03, 95% CI 1.01-1.06). To further explore the reasons for the seasonal variation in GDM, we further conducted a stratified analysis to investigate the correlation between temperature changes and GDM. After adjusting for other factors, temperatures within seasons had no influence on GDM. We speculate that the temperature variation within seasons is too small to affect GDM. Although the temperature is higher and does not vary much within a year in the study area, there is still the phenomenon of seasonality. And we still found that GDM was affected by the season after controlling for the confounding factors. Therefore, season may be considered one of the parameters of GDM diagnosis in Taiwanese women.

We further explored the mechanism of GDM seasonality. The effect of season on the diagnosis rate of GDM was due to postglucose load level variation.^[25-27] In 1997, the effects of changes in temperature on 1-h and 2-h OGTT results were investigated in healthy males in a climate-controlled chamber.^[40] The 2-h glucose result increased nonlinearly with each 5°C increase, and the rise was most apparent between 25°C and 30°C. A recent report also found a significantly higher glucose level in the heat (43°C) compared with both cold (7.2°C) and neutral (22°C) environments from the 15- to the 120-minute time point (P < .002).^[41] The pancreas anticipates the increase in insulin resistance and thus elevated glucose that occur late in pregnancy and therefore increases β -cell number and function early in pregnancy.^[42] A recent study conducted by Retnakaran et al, in pregnant women, showed that a rising mean daily temperature in the 3 to 4 weeks prior to the OGTT may have an adverse effect on maternal β -cell function and thereby increase the risk of GDM.^[43] Another possible mechanism might be the activation of brown adipose tissue upon cold exposure. Orava et al reported that cold exposure resulted in a 12-fold increase in glucose disposal in brown adipose tissue.^[44] In addition, its activation increases whole-body glucose disposal and insulin sensitivity in humans.^[45]

A strength of this study is that it is the first to investigate the effect of seasonality on the diagnosis of GDM in a large clinical sample of Asian women. These findings could serve as reference data for countries in Southeast Asia or areas with a similar climate. Some mothers had more than 1 pregnancy during the study period; therefore, we used generalized linear mixed models to address the repeated gravidity data and adjust for risk factors of GDM to retain complete gravidity information and increase the precision of the analysis.

A limitation of this study was that we used only the variables that had been recorded in the hospital information system, so we lacked information on some GDM risks, such as diet and physical activity, that might change according to the weather; thus, these risks were not evaluated. Notably, seasonal dietary changes were noted in a meta-analysis of the literature.^[46] The winter or the postharvest season is associated with increased energy intake. In the winter in Taiwan, because of the Chinese New Year, individuals consume more food and engage in less physical activity than in other seasons. Our research bears out these trends. The pregnant patients who underwent GDM screening during winter presented a faster rate of weight gain than those who underwent screening in other seasons; however, they had a lower prevalence of GDM, and the model was adjusted for this factor. Second, to further explore the reasons for the seasonal variation of GDM, we conducted a stratified analysis to investigate the correlation between temperature variation and GDM. The results did not show a significant correlation between temperature and GDM within a season. Temperature seems a part of the explanation, but not all. To determine whether this was because stratified analysis reduced the number of study samples or there was interference from seasonal-related factors, further research is needed.

5. Conclusions

Asian women were more likely to have GDM. GDM prevalence in Taiwan presents seasonal variation, with increased risk during spring and summer due to post-glucose load level variations. These findings could serve as reference data for countries that still use the 2-step approach for GDM screening and can also serve as a reference for countries in Southeast Asia or areas with a similar climate. Season may be considered as one of the factors that poses influence of GDM diagnosis in Southeast Asian women, especially for the women who were at high risk of GDM. More research is warranted to determine the seasonality of GDM in the general population from the other parts of Asia.

Author contributions

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