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Association of *TYK2* polymorphisms with autoimmune diseases: A comprehensive and updated systematic review with meta-analysis

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Abstract

Autoimmune diseases are characterized by the loss of self-tolerance, leading to immune-mediated tissue destruction and chronic inflammation. Tyrosine kinase 2 (TYK2) protein plays a key role in immunity and apoptosis pathways. Studies have reported associations between single nucleotide polymorphisms (SNPs) in the *TYK2* gene and autoimmune diseases; however, results are still inconclusive. Thus, we conducted a systematic review followed by meta-analysis. A literature search was performed to find studies that investigated associations between *TYK2* SNPs and autoimmune diseases (multiple sclerosis, systemic lupus erythematosus, Crohn's disease, ulcerative colitis, psoriasis, rheumatoid arthritis, type 1 diabetes, and inflammatory bowel disease). Pooled odds ratios (OR) with 95 % CI were calculated using random (REM) or fixed (FEM) effects models in the Stata 11.0 Software. Thirty-four articles were eligible for inclusion in the meta-analyses, comprising 9 different SNPs: rs280496, rs280500, rs280523, rs280519, rs2304256, rs12720270, rs12720356, rs34536443, and rs35018800. Meta-analysis results showed the minor alleles of rs2304256, rs12720270, rs12720356, rs34536443, and rs35018800 SNPs were associated with protection against autoimmune diseases. Moreover, the A allele of the rs280519 SNP was associated with risk for systemic lupus erythematosus. Our meta-analyses demonstrated that the rs2304256, rs12720270, rs12720356, rs34536443, rs35018800, and rs280519 SNPs in the *TYK2* gene are associated with different autoimmune diseases.

Keywords: Tyrosine kinase 2, autoimmunity, autoimmune disease, single nucleotide polymorphism, meta-analysis.

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Introduction

Autoimmune diseases are complex diseases triggered by multifaceted interactions between several genetic and environmental factors (Gutierrez-Roelens and Lauwerys, 2008; Rose, 2016), and are characterized by the loss of selftolerance leading to immune-mediated tissue destruction and chronic inflammation (Marrack et al., 2001; Lee and Bae, 2016; Odhams et al., 2017). These diseases share common etiological pathways, with genetic factors being considered as strong determinants of their development (Gutierrez-Roelens and Lauwerys, 2008; Lee and Bae, 2016). Regarding genetic factors, tyrosine kinase 2 (TYK2) is a candidate gene for autoimmune diseases since it encodes a member of Janus Kinase (JAK) family of tyrosine kinases, which have a central role in immune response since they mediate signaling pathways for several cytokines and type I interferon (IFN-I) (Ghoreschi et al., 2009; Strobl et al., 2011).

TYK2 is a non-receptor protein that bounds to the IFN-I receptor (IFNAR1) on the cell surface in its inactive form. After IFN- α binding to IFNAR1, TYK2 and JAK1 proteins are activated, leading to the recruitment and phosphorylation of the signal of transducers and activators of transcription

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(STAT) 1 and 2. STAT1/2 heterodimers then translocate to the nucleus, where they are major regulators of the expression of a number of IFN-stimulated genes (Yamaoka et al., 2004; Strobl et al., 2011). TYK2 is also associated with IL-6, IL-10, IL-12, and IL-23 receptors, playing a key role in the activation of these cytokine pathways (Ghoreschi et al., 2009; O'Shea and Plenge, 2012). Abnormal expression of IFN-I and other cytokines or JAK kinase members in immune cells are well known players in the pathogenesis of autoimmune diseases (Strobl et al., 2011; O'Shea and Plenge, 2012; Deng et al., 2019). Besides its role in the IFN-I and other type I and II cytokine receptor pathways, TYK2 plays a key role in other immune processes, including the activity of natural killer cells, maturation of B and Treg cells, and differentiation of Th1 and Th17 cells. Accordingly, dysregulated TYK2 expression has been associated with autoimmune diseases, specially systemic lupus erythematosus (SLE) [reviewed in (Deng et al., 2019)].

Consistent with the role of TYK2 in immune processes, several studies have suggested common single nucleotide polymorphisms (SNPs) in this gene are associated with different autoimmune diseases, including multiple sclerosis (MS) (Tao et al., 2011), SLE (Tao et al., 2011; Lee et al., 2012; Lee and Bae, 2016; Yin et al., 2018), Crohn's Disease (CD) (Lees et al., 2011; Tao et al., 2011; Ellinghaus et al., 2016), ulcerative colitis (UC) (Lees et al., 2011; Tao et al., 2011; Ellinghaus et al., 2011; Ellinghaus et al., 2011; Lee and Bae, 2016; Westra et al., 2018), type 1

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diabetes mellitus (T1DM) (Nagafuchi et al., 2015; Westra et al., 2018), and psoriasis (Pso) (Ellinghaus et al., 2016). In 2011, Tao et al. (2011) published a meta-analysis of 11 studies that investigated the association between 6 TYK2 SNPs and autoimmune and inflammatory diseases. The authors showed an association between the TYK2 rs2304256 and rs34536443 SNPs and MS, RA, SLE, CD, and UC. Lee and Bae (2016) performed a meta-analysis of 12 studies regarding the association of 7 TYK2 SNPs with SLE and RA, showing the rs2304256 and rs1270356 minor alleles were associated with protection against these rheumatic diseases. Five other SNPs (rs12720270, rs280500, rs280523, rs8108236, and rs280519) were not associated with these diseases; however, the number of studies included in their meta-analyses was small. A recent meta-analysis suggested the association of the TYK2 rs2304256 C allele with risk for SLE in Europeans (3 studies) but not in Asians (3 studies), while the rs12720270 and rs280519 SNPs were not associated with SLE (3 studies each) (Yin et al., 2018). Therefore, different SNPs in the TYK2 gene seem to be associated with autoimmune diseases, although the results on individual SNPs are still inconclusive (Tao et al., 2011; Lee et al., 2012; Ellinghaus et al., 2016; Lee and Bae, 2016; Westra et al., 2018; Yin et al., 2018) especially due to the increase in the number of studies in this field in the last few years in different ethnicities. Thus, here, we performed a comprehensive and updated meta-analysis of the related literature aiming to clarify the role of different TYK2 SNPs on susceptibility to autoimmune diseases.

Material and Methods

Search strategy and eligibility criteria

This systematic review was performed and described following PRISMA and MOOSE guidelines (Stroup et al., 2000; Moher et al., 2009), and its protocol was registered in the International Prospective Register of Systematic Reviews (PROSPERO) under the CRD42018100302 number. In order to identify studies that investigated associations between TYK2 SNPs and autoimmune diseases, we performed a literature search in Embase and PubMed resources. For this, the following MeSH terms were applied: ("TYK2 Kinase" OR "TYK2 protein, human") AND ("Autoimmune Diseases" OR "Rheumatic Diseases" OR "Lupus Erythematosus, Systemic" OR "Multiple Sclerosis" OR "Sclerosis" OR "Crohn Disease" OR "Pediatric Crohn's disease" OR "Ulcerative colitis" OR "Psoriasis" OR "Diabetes Mellitus" OR "Diabetes Mellitus, Type 1"). The search was completed on June, 2020, and was restricted to papers written in English, Spanish or Portuguese. Studies were also searched in the GWAS Catalog (https://www.ebi.ac.uk/gwas/).

Eligibility assessment was done by reviewing titles and abstracts of all articles selected, and when abstracts did not provide necessary information, the full text of the article was analyzed. This was performed independently, in a standardized manner, by two investigators (C.D. and F.M.P.), as previously described in other meta-analyses from our group (Souza *et al.*, 2013; Brondani *et al.*, 2014). Discordances were settled by debate between them and, when needed, a third investigator (D.C.) was referred. When articles had missing information,

we contacted the authors for further information. In case of duplicated data that had been published more than once, we opted to include the most complete study. In addition, reference lists from all articles fulfilling the eligibility criteria were manually searched to identify other important citations.

Studies were considered eligible if they had case-control designs and evaluated one or more *TYK2* SNPs in patients with some of the autoimmune diseases included in the MESH terms (cases) and individuals without any autoimmune condition (controls). Exclusion criteria were as follows: 1) Studies that did not have sufficient data to estimate an OR with 95 % CI; and 2) Studies where genotype distributions in control group deviated from those predicted by the Hardy-Weinberg equilibrium (HWE).

Data extraction and quality control assessment

Data were individually extracted by two researchers (C.D. and F.M.P.) using a standardized form (Souza *et al.*, 2013; Brondani *et al.*, 2014), and agreement was pursued in all extracted items. When an agreement could not be achieved, data extraction divergences were solved by referencing to the original publication or by consulting a third reviewer (D.C.). Data extracted from each study included: publication year, name of the first author, number of cases and controls, autoimmune disease, gender, age, ethnicity, genotyping technique, genotype and allele distributions in case and control samples and OR (95 % CI). We included in the meta-analysis only those SNPs investigated in at least 3 studies.

The Newcastle-Ottawa Scale (NOS) for case-control studies was used to analyze the quality of each eligible study (Wells et al., 2020). Two investigators (C.D. and F.M.P.) evaluated the 9 items of the NOS, which are categorized into 3 dimensions: selection, comparability, and exposure. Each item contains a sequence of alternative questions to be answered by the investigators. Then, a star scoring system allows the semi-quantitative analysis of article quality. In this score, the highest-quality studies receive one star for each item, excepting the comparability item that can receive two stars. Thus, the range of stars in the NOS score varies from zero to nine. The Clark-Baudouin Score (CBS) was also used to assess the quality of the studies (Clark and Baudouin, 2006). This method uses pre-defined criteria to assess each publication, highlighting quality issues in the conduction of studies and interpretation of the results. Using a 10-point scoring sheet, researchers are able to evaluate components of the articles related to reproducibility, selection of subjects, statistical analysis and genotyping methods.

Statistical analysis

Goodness-of-fit χ^2 tests were used to evaluate whether the genotype frequencies in the control groups were in agreement with those predicted by the Hardy-Weinberg Equilibrium (HWE). Gene-disease associations were measured using OR (95 % CI) calculations based on allele contrast, additive, recessive, and dominant inheritance models, which were categorized as suggested by Zintzaras *et al.* (2008). Stratifications by autoimmune disease type and/or ethnicity were performed when a disease / ethnicity

had ≥ 2 studies for each assessed SNP. Heterogeneity among studies was evaluated using a χ^2 -based Cochran's Q statistic and inconsistency was calculated using the I² metric. When P < 0.10 for the Q test and/or I² > 50 %, heterogeneity among studies was considered significant. Then, the DerSimonian and Laird random effects model (REM) was used to calculate OR (95 % CI) for each study and for the pooled effect. In the lack of significant heterogeneity, the fixed effect model (FEM) was used for this calculation (Higgins *et al.*, 2003; Melsen *et al.*, 2014).

In the case of relevant inter-study heterogeneity, sensitivity analyses were performed to identify which studies could have a considerable impact on heterogeneity. Risk of publication bias was evaluated for SNPs analyzed in ≥ 10 studies using funnel plot graphics, analyzed both visually and using the Begg and Egger statistic (Egger *et al.*, 1997). The significance of the intercept was determined by the *t* test, as proposed by Egger, with P < 0.10 being considered indicative of significant publication bias (Egger *et al.*, 1997). In case of significant publication bias, the Trim and Fill method was used for adjusting for it. This method evaluates whether publication bias is present and, then, estimates the pooled effect when biases are removed (Duval and Tweedie, 2000). The Stata 11.0 software (StataCorp, College Station, TX, USA) was used for all statistical analyses.

Results

Results from the literature search and quality of the studies

Figure 1 shows the flow diagram with the strategy used to identify and select studies for inclusion in this systematic review and meta-analysis. A total of 313 articles were retrieved after searching PubMed, Embase, and GWAS Catalog resources, and 237 of them were excluded following the review of titles and abstracts due to disagreements with our defined eligibility criteria. Seventy-six articles were therefore considered as being eligible at this point and had their full texts examined. Nevertheless, after analyzing the full texts, another 42 studies were excluded, and a total of 34 articles (Sigurdsson et al., 2005, 2007; International Consortium for Systemic Lupus Erythematosus Genetics et al., 2008; Ban et al., 2009; Hellquist et al., 2009; Kyogoku et al., 2009; Sato et al., 2009; Suarez-Gestal et al., 2009a, b; Genetic Analysis of Psoriasis et al., 2010; Jarvinen et al., 2010; Mero et al., 2010; Couturier et al., 2011; Graham et al., 2011; Li et al., 2011; Lian et al., 2013; Qiu et al., 2013; Shaiq et al., 2013; Alonso-Perez et al., 2014; Can et al., 2015; Diogo et al., 2015; Nagafuchi et al., 2015; Prieto-Perez et al., 2015; Tang et al., 2015; Ellinghaus et al., 2016; Lopez-Isac et al., 2016;

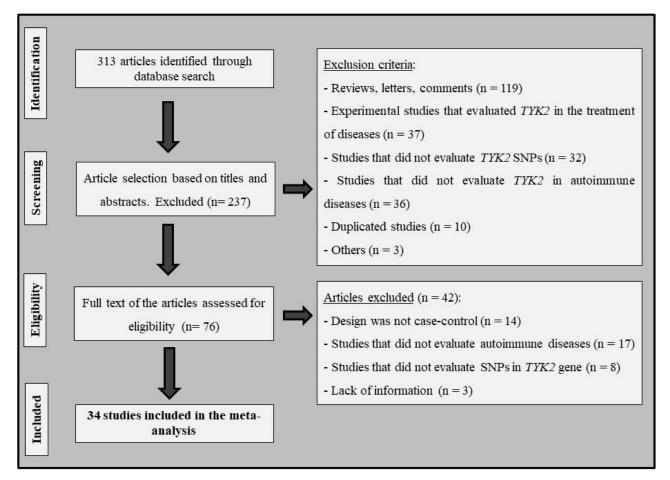


Figure 1 - Flowchart illustrating the search strategy used to identify studies of TYK2 SNPs and autoimmune diseases.

Almlof *et al.*, 2017; Langefeld *et al.*, 2017; Myrthianou *et al.*, 2017; Westra *et al.*, 2018; Zaplakhova *et al.*, 2018; Contreras-Cubas *et al.*, 2019; Graciolo *et al.*, 2019; Mohamadhosseini *et al.*, 2019) met the eligibility criteria and were included in this meta-analysis.

Nine SNPs in the *TYK2* gene were investigated in ≥ 3 studies and then were included in our meta-analyses. Table S1 shows genotype and allele frequencies of these SNPs in case and control groups from the eligible studies as well as in which autoimmune disease and ethnicity they were analyzed. Among the 34 eligible articles, 4 studies analyzed the rs280496 (g.10352804C>G) SNP (408 cases with CD or UC / 578 controls), 6 studies, the rs280500 (g.10379726A>G) SNP (2 988 cases with SLE / 6 440 controls), 14 studies, the rs280519 (g.10362257A>C) SNP (13 969 cases with CD, UC, Pso or SLE / 29 167 controls), and 6 studies investigated the rs280523 (g.10366530G>A) SNP (997 cases with CD, UC or SLE / 955 controls). Moreover, 25 studies evaluated the rs2304256 (g.10364976C>A) SNP (23 827 cases with CD, UC, SLE, MS, RA or T1DM / 35 760 controls), 9 studies, the rs12720270 (g.10365084G>A) SNP (2 792 SLE cases / 5 184 controls), and 20 studies analyzed the rs12720356 (g.10359299A>C) SNP (69 788 cases with SLE, RA, IBD, MS, CD, UC or Pso / 177 438 controls). Nineteen studies evaluated the rs34536443 (g.10352442G>C) SNP (50 011 cases with MS, RA, SLE, IBD, Pso or T1DM / 95 923 controls), while 8 studies analyzed the rs35018800 (g.10354167G>A) SNP (61 241 cases with RA, SLE, IBD, CD, Pso, UC or MS / 163 386 controls).

The quality of each individual study is shown in Table S2. As mentioned in the Material and Methods section, the highest quality articles can receive up to 9 stars for the NOS score. Most of the included studies were classified as presenting good quality since 61.8 % of the studies were awarded 6 to 8 stars and 38.2 % of the studies received the maximum of stars allowed. None of the articles scored less than 6 stars. Regarding the CBS score, most of the studies were also classified as presenting good quality since 73.5 % of them received 7 to 9 points and 26.5 % received 10 points, which is the highest score.

Meta-analyses of studies that evaluated associations between TYK2 SNPs and autoimmune diseases

Table 1 shows results of the pooled analyses for the associations between *TYK2* rs280496, rs280500, rs280519, rs280523, rs2304256, rs12720270, rs12720356, rs34536443, and rs35018800 SNPs and autoimmune diseases under different inheritance models. When the number of studies was statistically adequate, we also evaluated these associations after stratification by disease type and/or ethnicity (Table 1).

 $\textbf{Table 1} - \textbf{Pooled measures for associations between } \textit{TYK2} \texttt{rs}280496, \texttt{rs}280500, \texttt{rs}280519, \texttt{rs}280523, \texttt{rs}2304256, \texttt{rs}12720270, \texttt{rs}12720356, \texttt{rs}34536443, \\ \textbf{and rs}35018800 \texttt{SNPs} \text{ and susceptibility to autoimmune diseases.}$

| Inheritance model | n studies | n cases | n controls | I ² % | Pooled OR (95 % CI) |
|-------------------|-----------|---------|------------|------------------|----------------------|
| rs280496 | | | | | |
| Allele contrast | 4 | 408 | 578 | 0.0 | 1.15 (0.89 – 1.49)* |
| Disease | | | | | |
| CD | 2 | 143 | 289 | 0.0 | 1.19 (0.80 – 1.76)* |
| UC | 2 | 265 | 289 | 0.0 | 1.12 (0.80 – 1.58)* |
| Dominant | 3 | 294 | 378 | 0.0 | 1.21 (0.86 – 1.70)* |
| rs280500 | | | | | |
| Allele contrast | 6 | 2 988 | 6 440 | 55.7 | 1.12 (0.95 – 1.31)** |
| rs280523 | | | | | |
| Allele contrast | 6 | 997 | 955 | 0.0 | 1.11 (0.87 – 1.41)* |
| Disease | | | | | |
| CD | 2 | 143 | 289 | 0.0 | 0.90 (0.51 – 1.57)* |
| UC | 2 | 265 | 289 | 0.0 | 1.07 (0.68 – 1.70)* |
| SLE | 2 | 589 | 377 | 0.0 | 1.21 (0.87 – 1.68)* |
| Ethnicity | | | | | |
| Asian | 4 | 408 | 578 | 0.0 | 1.00 (0.70 – 1.42)* |
| Caucasian | 2 | 589 | 377 | 0.0 | 1.21 (0.87 – 1.68)* |
| Dominant | 3 | 294 | 378 | 0.0 | 0.95 (0.60 – 1.49)* |
| rs280519 | | | | | |
| Allele contrast | 14 | 13 969 | 29 167 | 87.1 | 1.07 (0.96 – 1.20)** |
| Disease | | | | | |
| CD | 3 | 223 | 389 | 58.7 | 1.08 (0.74 – 1.57)** |
| UC | 2 | 265 | 289 | 0.0 | 1.21 (0.94 – 1.55)** |
| SLE | 8 | 6 733 | 16 973 | 38.5 | 1.10 (1.04 – 1.18)** |
| Pso | 1 | 6 748 | 11 516 | _ | _ |

 $Table \ 1-Cont.$

| Inheritance model | n studies | n cases | n controls | I ² % | Pooled OR (95 % CI) |
|-------------------|-----------|---------|------------|------------------|----------------------|
| Ethnicity | | | | | |
| Asian | 8 | 1 868 | 2 538 | 34.4 | 1.08 (0.95 – 1.22)** |
| Caucasian | 6 | 12 101 | 25 245 | 94.2 | 1.07 (0.91 – 1.25)** |
| Recessive | 6 | 1 085 | 1 184 | 65.6 | 1.18 (0.80 – 1.75)** |
| Dominant | 6 | 1 085 | 1 184 | 0.0 | 0.84 (0.70 – 1.02)* |
| Additive | 6 | 618 | 629 | 41.8 | 0.89 (0.71 – 1.12)* |
| rs2304256 | | | | | |
| Allele contrast | 25 | 23 827 | 35 760 | 70.8 | 0.83 (0.77 – 0.88)** |
| Disease | | | | | |
| SLE | 12 | 7 315 | 11 736 | 72.9 | 0.77 (0.69 – 0.85)** |
| CD | 3 | 223 | 389 | 53.5 | 0.77 (0.52 – 1.15)** |
| UC | 2 | 265 | 289 | 0.0 | 0.84 (0.64 – 1.09)** |
| T1DM | 3 | 551 | 573 | 56.6 | 1.03 (0.77 – 1.38)** |
| MS | 3 | 12 312 | 20 010 | 0.0 | 0.84 (0.81 – 0.87)** |
| RA | 2 | 3 161 | 2 763 | 33.3 | 0.99 (0.89 – 1.09)** |
| Ethnicity | | | | | |
| Caucasian | 12 | 20 474 | 30 034 | 73.8 | 0.81 (0.75 – 0.87)** |
| Asian | 10 | 2 678 | 4 891 | 74.7 | 0.85 (0.71 – 1.01)** |
| Mixed Ethnicity | 3 | 675 | 835 | 31.6 | 0.87 (0.69 – 1.09)** |
| Recessive | 13 | 16 916 | 26 506 | 78.3 | 0.80 (0.65 – 0.98)** |
| Disease | | | | | |
| SLE | 4 | 3 759 | 5 545 | 88.5 | 0.62 (0.40 – 0.97)** |
| CD | 2 | 143 | 289 | 33.8 | 0.88 (0.45 – 1.75)** |
| UC | 1 | 151 | 89 | _ | _ |
| T1DM | 3 | 551 | 573 | 0.0 | 1.51 (0.97 – 2.34)** |
| MS | 3 | 12 312 | 20 010 | 0.0 | 0.79 (0.73 – 0.87)** |
| Ethnicity | | | | | |
| Caucasian | 4 | 14 794 | 24 134 | 83.7 | 0.64 (0.51 – 0.82)** |
| Asian | 7 | 1 815 | 2 053 | 80.7 | 0.87 (0.55 – 1.37)** |
| Mixed Ethnicity | 2 | 307 | 319 | 0.0 | 1.12 (0.54 – 2.30)** |
| Dominant | 14 | 16 996 | 26 606 | 45.0 | 0.78 (0.72 – 0.84)** |
| Disease | | | | | |
| SLE | 4 | 3 759 | 5 545 | 13.0 | 0.75 (0.67 – 0.83)** |
| CD | 3 | 223 | 389 | 63.9 | 0.60 (0.31 – 1.14)** |
| UC | 1 | 151 | 89 | _ | _ |
| T1DM | 3 | 551 | 573 | 49.8 | 0.93 (0.66 – 1.31)** |
| MS | 3 | 12 312 | 20 010 | 0.0 | 0.81 (0.77 – 0.85)** |
| Ethnicity | | | | | |
| Caucasian | 4 | 14 794 | 24 134 | 28.9 | 0.79 (0.75 – 0.83)** |
| Asian | 8 | 1 895 | 2 153 | 59.0 | 0.72 (0.55 – 0.95)** |
| Mixed Ethnicity | 2 | 307 | 319 | 56.0 | 0.82 (0.51 – 1.33)** |
| Additive | 13 | 10 778 | 15 955 | 72.2 | 0.68 (0.55 – 0.84)** |
| Disease | | | | | |
| SLE | 4 | 2 330 | 3 325 | 86.2 | 0.58 (0.34 – 0.99)** |
| CD | 2 | 103 | 164 | 0.0 | 0.43 (0.21 – 0.88)** |

 $Table \ 1-Cont.$

| Inheritance model | n studies | n cases | n controls | I ² % | Pooled OR (95 % CI) |
|-------------------|-----------|---------|------------|------------------|----------------------|
| UC | 1 | 113 | 54 | _ | _ |
| T1DM | 3 | 345 | 334 | 0.0 | 1.45 (0.92 – 2.28)** |
| MS | 3 | 7 887 | 12 078 | 0.0 | 0.73 (0.67 – 0.80)** |
| Ethnicity | | | | | |
| Caucasian | 4 | 9 437 | 14 493 | 84.5 | 0.59 (0.46 – 0.75)** |
| Asian | 7 | 1 136 | 1 266 | 68.3 | 0.73 (0.46 – 1.16)** |
| Mixed Ethnicity | 2 | 205 | 196 | 0.0 | 1.03 (0.49 – 2.14)** |
| rs12720356 | | | | | |
| Allele contrast | 20 | 69 788 | 177 438 | 92.1 | 0.85 (0.77 – 0.94)** |
| Disease | | | | | |
| SLE | 9 | 6 360 | 20 668 | 49.9 | 0.75 (0.65 – 0.88)** |
| RA | 2 | 6 256 | 14 544 | 11.5 | 0.91 (0.83 – 1.00)** |
| IBD | 1 | 1 346 | 13 683 | _ | _ |
| CD | 1 | 19 085 | 34 213 | _ | _ |
| Pso | 4 | 10 240 | 40 419 | 63.2 | 0.71 (0.61 – 0.83)** |
| UC | 1 | 14 413 | 34 213 | _ | _ |
| MS | 2 | 12 088 | 19 698 | 0.0 | 0.85 (0.79 – 0.90)** |
| Recessive | 4 | 13 437 | 22 606 | 1.1 | 0.85 (0.64 – 1.13)* |
| Dominant | 4 | 13 437 | 22 606 | 34.2 | 0.82 (0.77 – 0.87)* |
| Additive | 4 | 11 787 | 19 301 | 4.4 | 0.83 (0.62 – 1.10)* |
| rs34536443 | | | | | |
| Allele contrast | 19 | 50 011 | 95 923 | 75.7 | 0.68 (0.61 – 0.76)** |
| Disease | | | | | |
| MS | 9 | 21 346 | 27 989 | 38.6 | 0.75 (0.67 – 0.83)** |
| RA | 3 | 5 818 | 14 894 | 62.4 | 0.83 (0.54 – 1.29)** |
| SLE | 4 | 12 041 | 27 735 | 25.4 | 0.50 (0.43 – 0.57)** |
| IBD | 1 | 1 346 | 13 687 | _ | _ |
| Pso | 1 | 126 | 507 | _ | _ |
| T1DM | 1 | 9 334 | 11 111 | _ | _ |
| Recessive | 4 | 13 180 | 20 905 | 67.1 | 0.35 (0.03 – 3.83)** |
| Dominant | 5 | 13 306 | 21 412 | 88.0 | 0.34 (0.21 – 0.56)** |
| Additive | 4 | 12 834 | 20 478 | 67.2 | 0.35 (0.03 – 3.88)** |
| rs12720270 | | | | | |
| Allele contrast | 9 | 2 792 | 5 184 | 30.2 | 0.92 (0.84 – 1.00)* |
| Ethnicity | | | | | |
| Asian | 3 | 1 380 | 3 274 | 0.0 | 0.92 (0.87 – 1.08)* |
| Caucasian | 5 | 1 044 | 1 394 | 51.7 | 0.84 (0.72 – 0.98)* |
| Mixed Ethnicity | 1 | 368 | 516 | _ | _ |
| rs35018800 | | | | | |
| Allele contrast | 8 | 61 241 | 163 386 | 18.0 | 0.60 (0.55 – 0.65)* |

Where significant heterogeneity was detected ($I^2 > 50$ % and/or Q statistic P>0.1), the DerSimonian and Laird random effect model (REM)** was used to calculate OR (95 % CI) for each individual study and for the pooled effect; where heterogeneity was not significant, the fixed effect model (FEM)* was used for this calculation. CD: Crohn's disease; IBD: inflammatory bowel disease; MS: multiple sclerosis; Pso: psoriasis; RA: rheumatoid arthritis; SLE: systemic lupus erythematosus; T1DM: type 1 diabetes mellitus; UC: ulcerative colitis.

Our results show the *TYK2* rs280496, rs280500, and rs280523 SNPs were not associated with autoimmune diseases under an allele contrast model (Table 1 and Figures S1 and S2). Overall, the rs280519 SNP was also not associated with autoimmune diseases considering allele, dominant, recessive, and additive models (Table 1 and Figure S2). However, after stratification by disease type, the rs280519 SNP was independently associated with risk for SLE (REM OR 1.10, 95 % CI 1.04 – 1.18, allele contrast model; Table 1 and Figure S2).

The A allele of the rs2304256 SNP was associated with protection against autoimmune diseases when assuming the allele contrast model (REM OR 0.83, 95 % CI 0.77 – 0.88; Table 1 and Figure 2). This SNP was also significantly associated with autoimmune diseases under dominant, recessive, and additive models (Table 1). In addition, the C allele of the rs12720356 SNP conferred protection against different autoimmune diseases under both the allele contrast (REM OR 0.85, 95 % CI 0.77 – 0.94) and dominant models (REM OR 0.82, 95 % CI 0.77 – 0.87; Table 1 and Figure 3).

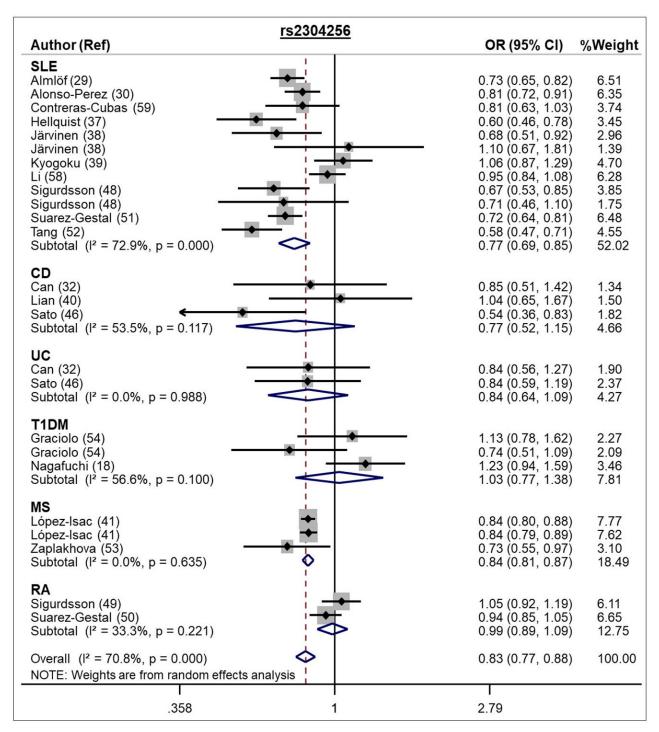


Figure 2 – Forest plot showing individual and pooled OR (95 % CI) for the association between the TYK2 rs2304256 SNP and autoimmune diseases, under an allele contrast model.

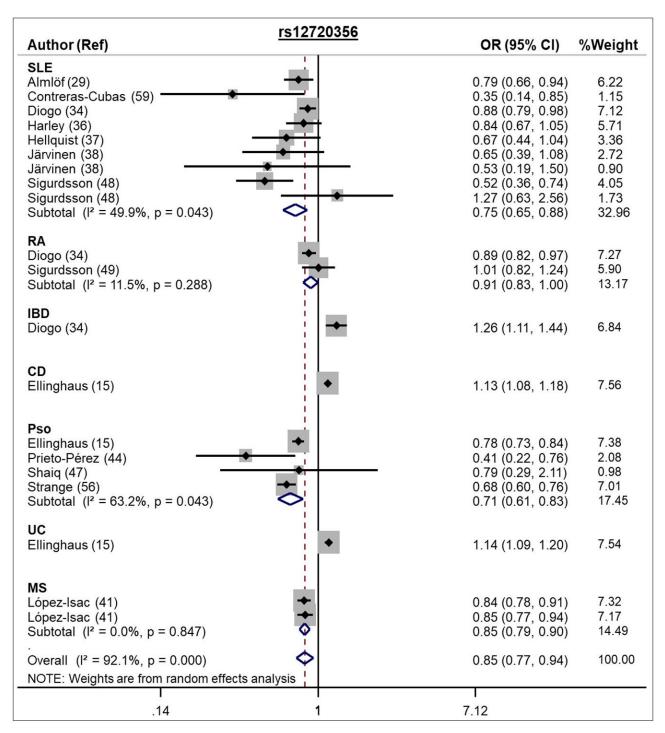


Figure 3 – Forest plot showing individual and pooled OR (95 % CI) for the associations between the TYK2 rs12720356 SNP and autoimmune diseases, under an allele contrast model.

In the same way, the rs34536443 C allele was associated with protection for autoimmune diseases under allele contrast (REM OR 0.68, 95 % CI 0.61 – 0.76) and dominant (REM OR 0.75, 95 % CI 0.58 – 0.98) models (Table 1 and Figure 4).

The A allele of the rs12720270 SNP conferred protection for SLE under the allele contrast model [FEM OR 0.92, 95 % CI 0.84-1.00 (P = 0.041); Table 1 and Figure 5]. Moreover, after stratification by ethnicity, the A allele was associated with protection for SLE in Caucasians (FEM OR 0.84, 95 % CI 0.72-0.98) but not in Asians (Table 1). Besides SLE, this SNP

was not evaluated in other autoimmune diseases. The A allele of the rs35018800 SNP also conferred protection for autoimmune diseases (FEM OR 0.60, 95 % CI 0.55 – 0.65, allele contrast model; Table 1 and Figure 5) in European populations.

Sensitivity analyses and publication bias

When significant inter-study heterogeneities were observed, sensitivity analyses were carried out in order to estimate the influence of each individual study on the meta-analysis results obtained when assuming the allele contrast

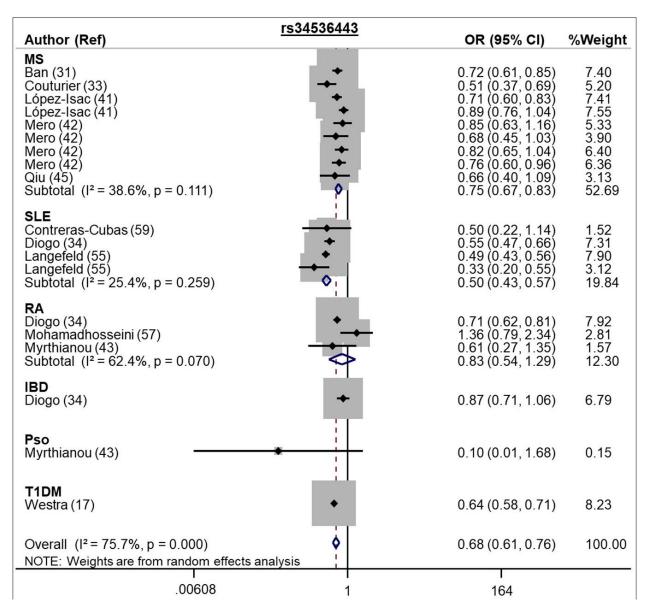


Figure 4 – Forest plot showing individual and pooled OR (95 % CI) for the associations between the TYK2 rs34536443 SNP and autoimmune diseases, under an allele contrast model.

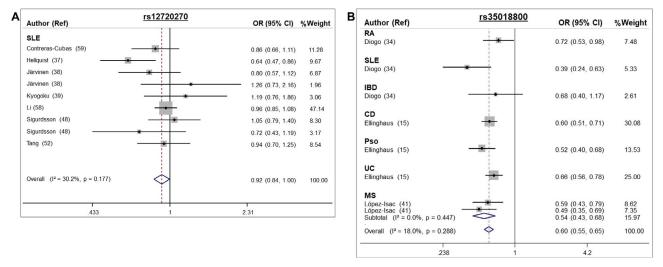


Figure 5 – Forest plots showing individual and pooled OR (95 % CI) for the associations between the TYK2 rs12720270 (A) and rs35018800 SNPs (B) and autoimmune diseases, under an allele contrast model.

model. This was performed by repeating meta-analyses excluding a different study each time. Our results showed two studies (International Consortium for Systemic Lupus Erythematosus Genetics et al., 2008; Tang et al., 2015) explained the observed heterogeneity in the meta-analysis of the rs280500 SNP since their exclusion significantly decreased the heterogeneity (all studies: I² = 42.6 %, P = 0.138, and after exclusion: I^2 0.0 %, P = 0.901). However, the exclusion of these two studies from the rs280500 metaanalysis did not change the lack of the association of this SNP with autoimmune diseases. Moreover, the exclusion of one study (Genetic Analysis of Psoriasis et al., 2010) from the rs280519 meta-analysis decreased the observed heterogeneity $(I^2 = 27.7 \%, P = 0.165)$. Importantly, the exclusion of this study significantly changed the pooled OR for this SNP, which was now associated with risk for autoimmune diseases (OR 1.11, 95 % CI 1.05 - 1.18). Of note, meta-analyses of the rs2304256, rs12720356, and rs34536443 SNPs still presented significant heterogeneity after sensitivity analyses.

Funnel plots and Egger's tests were performed to investigate the presence of possible publication bias in those meta-analyses containing at least 10 studies, which were those performed for the rs12720356, rs2304256, rs280519, and rs34536443 SNPs (Figure S3A-D). No significant publication bias was observed for the rs2304256, rs280519, and rs34536443 SNPs. However, funnel plot and Egger's test indicated a significant publication bias in the rs12720356 meta-analysis (P=0.037; Figure S3A). Trim and Fill analysis was performed to account for this bias, and the results indicated that the pooled OR obtained for this SNP did not change significantly since the adjusted effect was similar to the original effect.

Discussion

To further investigate the possible effects of the *TYK2* SNPs on susceptibility for autoimmune diseases, we performed meta-analyses of 34 published articles on the field. Our results suggest the minor alleles of rs2304256, rs12720270, rs12720356, rs34536443, and rs35018800 SNPs are associated with protection against autoimmune diseases, while the rs280519A allele is associated with risk for SLE. The rs280496, rs280500, and rs280523 SNPs do not seem to be associated with autoimmune diseases in the investigated populations.

Our meta-analysis for the rs280519 SNP included 14 studies (13 969 cases and 29 167 controls) and showed the A allele of this SNP is associated with risk for SLE. This SNP does not seem to be associated with CD, UC, and Pso. In contrast, two previous meta-analyses did not show any association of this SNP with autoimmune diseases (Tao et al., 2011; Yin et al., 2018). The discordant results may be due to the small number of studies included in the meta-analyses by Tao et al. (2011), which included 3 studies with CD and SLE, and by Yin et al. (2018), which included 4 studies with SLE. The A allele of this SNP does not cause an amino acid substitution, but it is located in a splice site of the TYK2 gene (Lopez-Rodriguez et al., 2017). No study has evaluated if this SNP has an impact on TYK2 function. Linkage disequilibrium (LD) analyses suggest the rs280519 SNP is on the same haplotype block of the rs2304256 and rs12720270 SNPs (Kyogoku *et al.*, 2009); thus, the rs280519 SNP could be a marker of the functional rs2304256 SNP.

Our meta-analysis for the rs2304256 SNP included 25 studies (23 827 cases and 35 760 controls) and showed the A allele of this SNP is associated with protection against autoimmune diseases (SLE, CD, UC, T1DM, MS, and RA) in all inheritance models analyzed. This association was confirmed in SLE and MS diseases, although the lack of individual associations with CD, UC, T1DM, and RA might be due to the small number of studies/sample sizes for each disease. In addition, after stratification by ethnicity, this SNP remained associated with autoimmune diseases in Caucasians (REM OR 0.89, 95 % CI 0.81 – 0.98, allele contrast model) but not in Asians or populations of mixed ethnicity (from Southern Brazil), which can be attributed to the fact that the studies in Asian and Brazilian populations presented a small number of subjects. Our results regarding this SNP are in agreement with the results of two previous meta-analyses (Tao et al., 2011; Lee and Bae, 2016). The meta-analysis performed by Tao et al. (2011) included only 11 studies and showed the rs2304256 A allele conferred protection for SLE, RA, UC, and CD (OR 0.78, 95 % CI 0.70 – 0.87, for the allele contrast model). In 2016, Lee and Bae (2016) published a meta-analysis of 12 studies, showing the rs2304256 A allele was associated with protection against SLE and RA in Caucasians (OR 0.82, 95 % CI 0.70 - 0.89) but not in Asians (Lee and Bae, 2016). In addition, Zuvich et al. (2010) demonstrated the rs2304256 A allele was associated with protection for MS (OR 0.90) in North American and British subjects. This study was not included in our meta-analysis due to lack of required data.

The A allele of the rs2304256 causes a substitution of valine to phenylalanine at position 362 in the JAK-homology 4 (JH4) region, which is a crucial domain for interaction of TYK2 with IFNAR1 and its function, maintaining the expression of IFNAR1 on cell membranes (Tao et al., 2011; Marroqui et al., 2015). Li et al. (2020) showed the rs2304256 A allele affects the TYK2 pre-mRNA processing since it destroys a putative exonic splicing enhancer; thus, promoting the inclusion of exon 8 in the mRNA, which is essential for TYK2 binding to cytokine receptors. Marroqui et al. (2015) isolated B lymphoblastoid cell lines (BLCLs) from subjects carrying the rs2304256 A/A genotype and demonstrated less marked IFN-α-induced STAT1 phosphorylation compared with subjects carrying the C/C genotype (3.5-fold STAT1 phosphorylation vs. 5.7-fold increase). Interestingly, TYK2 inhibition decreased cytokine-induced apoptosis and proinflammatory pathways in pancreatic beta-cells via inhibition of the IFN-I signaling and consequent decrease in STAT1/2 phosphorylation (Marroqui et al., 2015). It is well known that the initial attack in autoimmune diseases is usually followed by an inflammatory response caused by autoreactive cytotoxic cells, which then activates the release of pro-inflammatory cytokines and apoptosis via JAK-STAT pathways (Stuart and Hughes, 2002; Coomans de Brachene et al., 2020). Thus, taken together, these studies suggest the rs2304256 SNP decreases TYK2 activity and, consequently, the inflammatory response and apoptosis, explaining its association with protection against autoimmune diseases.

Our meta-analysis for the rs12720356 SNP included 20 studies (69 788 cases and 177 437 controls) and showed the C allele of this SNP provides protection for autoimmune diseases (SLE, RA, IBD, CD, Pso, UC, and MS) under allele contrast and dominant models. Among the 20 studies that evaluated this SNP, 18 were performed in Caucasian subjects. Accordingly, the meta-analysis conducted by Lee and Bae (2016) included 6 studies with rheumatic diseases and showed a similar result to ours (OR 0.81, 95 % CI 0.66 - 0.99). In contrast, another small meta-analysis, which included 4 035 cases with SLE, RA, or CD and 2 953 controls, was not able to find any association between the rs12720356 SNP and these diseases, possible because of the small number of evaluated studies (n = 4) and sample sizes (Tao et al., 2011). The C allele of the rs12720356 SNP leads to a isoleucine to serine substitution at position 684 in the pseudo-kinase region JAK-homology 2 (JH2) of TYK2. This region is required for the binding of IFN-I to IFNAR1 (Sigurdsson et al., 2005). Enerbäck et al. (2018) analyzed peripheral blood mononuclear cells (PBMCs) from patients with Pso carrying the A allele (n = 10) vs. patients with the C allele (n = 10) of this SNP. PBMCs from subjects carrying the C allele showed reduced phosphorylated (p)-STAT4 levels after induction with IL-12 compared to the A/A genotype, suggesting the rs12720356 C allele may have a functional impact on TYK2 function and, consequently, immunity (Enerback et al., 2018).

Our meta-analysis for the rs34536443 SNP included 19 studies (50 011 cases and 95 923 controls) and showed the C allele is associated with protection against autoimmune diseases (MS, RA, SLE, IBD, Pso, and T1DM) under both allele contrast and dominant models. This association was confirmed for SLE and MS; however, for IBD, Pso, and T1DM, we had a small number of studies to individually conclude about the associations with these diseases. Of note, most of the studies were performed in Caucasian subjects. Moreover, two studies (Johnson et al., 2010; International Multiple Sclerosis Genetics et al., 2013) were not included in our meta-analysis due to lack of data. Johnson et al. (2010) demonstrated the rs34536443 C allele was associated with risk for MS (OR 2.04, 95 % CI 1.01 - 4.08) in African-Americans (Johnson et al., 2010). In contrast, another study including 14 498 patients with MS and 24 091 controls of European ancestry suggested this allele conferred protection for MS (OR 0.95; P = 1.2×10^{-8}), which is in accordance to our results (International Multiple Sclerosis Genetics et al., 2013). Tao et al. (2011) also performed a meta-analysis of the rs34536443, including 9 studies with MS (10 642 MS patients / 10 620 controls), and showed the C allele was associated with protection against this disease (OR 0.76, 95 % CI 0.69 – 0.84) (Tao *et al.*, 2011).

The rs34536443 SNP is located in exon 21 and causes a change of a proline to alanine at position 1104 within the kinase domain of TYK2 (Peluso *et al.*, 2013; Gorman *et al.*, 2019). The C allele of this SNP seems to be functional since it decreased the IFN- α induced-pSTAT1 levels in PBMCs compared to cells obtained from patients carrying the G allele, thus reducing IFNAR signaling (Gorman *et al.*, 2019). The rs34536443 C allele also decreased IL-23 and IL-12 induced-p-STAT3 levels in a murine model of MS (Gorman *et al.*, 2019). Accordingly, PBMCs of patients with MS carrying the

C allele of this SNP also showed reduced IFN β induced-p-STAT2 levels compared to patients with the G allele (Couturier *et al.*, 2011).

Our meta-analysis for the rs12720270 SNP included 9 studies (2 792 cases and 5 184 controls) and showed the A allele of this SNP was associated with protection against SLE. In contrast, 3 previous meta-analyses (Tao et al., 2011; Lee and Bae, 2016; Yin et al., 2018), including only 3 to 5 studies with SLE patients, were not able to show any association between this SNP and SLE. This SNP is located in intron 7 of TYK2 gene, most specifically 36 nt upstream of the intron 7/exon 8 boundary (Li et al., 2020). The rs12720270 SNP is in strong LD with the functional rs2304256 and the rs280519 SNPs (Contreras-Cubas et al., 2019; Li et al., 2020). However, it also seems to be functional since in silico analysis and cell line experiments suggested the A allele breaks a splicingbranch point in the intron, promoting the inclusion of exon 8 in the mature TYK2 mRNA; thus, influencing TYK2 activity (Li et al., 2020).

Our meta-analysis for the rs35018800 SNP included 8 studies (61 241 cases and 163 386 controls) and demonstrated the A allele of this SNP is associated with protection against autoimmune diseases (RA, SLE, IBD, CD, Pso, UC, and MS) in Caucasian subjects. No study has evaluated this SNP in other ethnicities. We were not able to determine if this SNP provides differential protection for a given autoimmune disease since we had a small number of studies for each disease. The rs35018800 SNP causes a substitution of an alanine to valine at position 928 within the kinase domain of TYK2 (Lopez-Isac *et al.*, 2016). To date, there is no available information if this SNP has a functional significance.

Autoimmune diseases share common etiological pathways and, as a result, they may also share some similar genetic factors (Gutierrez-Roelens and Lauwerys, 2008; Luan et al., 2017). Indeed, our present meta-analysis confirms that the rs2304256, rs12720270, rs12720356, rs34536443, rs35018800, and rs280519 SNPs influence the susceptibility to different autoimmune diseases. However, the size of the effect of each individual SNP on a specific autoimmune disease might be affected by the interaction with other environmental and genetic factors involved in that disease. Thus, additional studies with larger sample sizes are required in order to clarify the effects of the rs2304256, rs12720270, rs12720356, rs34536443, rs35018800, and rs280519 SNPs on each autoimmune disease analyzed here. Moreover, since some of the analyzed SNPs (rs12720270, rs2304256, and rs280519) are in strong LD, future functional studies should evaluate which is(are) the functional(s) SNP(s) in a LD block or if they are interacting in the susceptibility for the autoimmune diseases.

Despite all the efforts, the results of the present metaanalysis should be interpreted within the context of few limitations. First, we tried to retrieve all published articles, but we cannot exclude the possibility that small negative studies could have been lost. Although we did not observe publication bias for the rs2304256, rs280519, and rs34536443 SNPs, a significant publication bias was present in the metaanalysis of the rs12720356 SNP. However, Trim and Fill analysis demonstrated that the adjusted effect did not change significantly, indicating that the number of missing studies

needed to reverse the bias is smaller than the number of missing studies needed to nullify the effect (Brondani et al., 2014). Second, we only analyzed those articles written in English, Spanish or Portuguese; hence, we could have lost few articles written in other languages. Third, we were not able to perform meta-regression analyses to explain the observed heterogeneity because of lack of data regarding age and gender. Despite of that, we performed sensibility analysis for those SNPs that showed significant heterogeneity in the respective meta-analyses. The exclusion of two studies (International Consortium for Systemic Lupus Erythematosus Genetics et al., 2008; Tang et al., 2015) from the rs280500 meta-analysis decreased its heterogeneity, but did not change the observed result. However, the exclusion of one study (Genetic Analysis of Psoriasis et al., 2010), with a large sample size, explained the heterogeneity detected in the rs280519 meta-analysis. After exclusion of this study, the rs280519 SNP was associated with risk for autoimmune diseases, suggesting that heterogeneity among studies might have influenced the results. Fourth, as already mentioned, we could not include 3 articles (Johnson et al., 2010; Zuvich et al., 2010; International Multiple Sclerosis Genetics et al., 2013) in our meta-analyses due to the lack of data. Fifth, the rs280496, rs280523, and rs35018800 SNPs were investigated by few studies, thus we were not able to stratify their meta-analyses by disease type. In addition, ethnic-specific association studies are required to confirm genetic associations in different populations (Lee et al., 2012), since we could not identify differences among ethnicities in the meta-analyses of the rs34536443, rs280500, rs280496, rs12720356, and rs35018800 SNPs due to the small number of studies evaluating different ethnicities.

In conclusion, our results suggest that the minor alleles of the rs2304256, rs12720270, rs12720356, rs34536443, and rs35018800 SNPs are involved in the protection against autoimmune diseases, and the A allele of the rs280519 SNP is associated with risk for SLE. In addition, our results indicate that the rs280496, rs280500, and rs280523 SNPs are not associated with autoimmune diseases. Additional studies with larger sample sizes are necessary to clarify the impact of each *TYK2* SNP on susceptibility for different autoimmune diseases. Functional studies are also needed to elucidate which are the *TYK2* SNPs with the highest impact on *TYK2* function and, consequently, on autoimmune diseases.

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Conflict of Interest

The authors declare that they have no conflict of interest.

Author Contributions

FMP, ACB, BMS and DC designed the study. FMP and CD carried out the data extraction, and FMP, CD and NEL performed the statistical analyses. FMP wrote the manuscript. FMP, CD, NEL, ACB, BMS and DC modified the manuscript. This manuscript has been read and approved by all authors and all authors believe that this study represents honest work.

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Supplementary material

The following online material is available for this article: Figure S1 – Forest plots showing individual and pooled OR (95 % CI) for the associations between the TYK2 rs280496 and rs280500 SNPs and autoimmune diseases, under an allele contrast model.

Figure S2 – Forest plots showing individual and pooled OR (95 % CI) for the associations between the TYK2 rs280523 and rs280519 SNPs and autoimmune diseases, under an allele contrast model. Figure S3 – Begg's funnel plots for publication bias test for TYK2 SNPs: A) rs12720356, B) rs2304256, C) rs280519, and D) rs34536443.

Table S1 – Genotype and allele distributions of TYK2 rs280496, rs280500, rs280519, rs280523, rs2304256, rs12720270, rs12720356, rs34536443, and rs35018800 SNPs in patients with autoimmune diseases and control subjects.

 $Table \ S2-New castle-Ottawa\ quality\ assessment\ scale\ for\ the\ studies\ included\ in\ the\ meta-analyses.$

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