

SYSTEMATIC REVIEW

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Brain natriuretic peptide as a predictive marker of mortality in sepsis: an updated systematic review and meta-analysis

Jian-li Song^{1†}, Bin Fan^{1†}, Li-quan Qiu^{1†}, Qiang Li¹ and Guan-yu Chen^{1*}

Abstract

Introduction Early identification of patients with sepsis at high risk of death remains a challenge, and whether brain natriuretic peptide (BNP) or N-terminal pro-B-type natriuretic peptide (NT-proBNP) has a prognostic effect on patients with sepsis is controversial. Here, we clarified the prognostic value of BNP and NT-proBNP and sought to establish suitable cutoff values and intervals.

Methods We searched five databases to identify studies that met the inclusion criteria. The primary outcomes were the pooled sensitivity, specificity, positive likelihood ratio (PLR), negative likelihood ratio (NLR), diagnostic odds ratio (DOR), area under the curve (AUC), and corresponding 95% credible interval (95% CI) of BNP and NT-proBNP. The secondary outcomes were the sensitivity and specificity of BNP or NT-proBNP in subgroup analyses.

Results Forty-seven studies were included in our meta-analysis. The pooled sensitivity of NT-proBNP (0.77 [0.68, 0.84]) was weaker than that of BNP (0.82 [0.76, 0.87]), the pooled specificity of NT-proBNP (0.70 [0.60, 0.77]) was less than that of BNP (0.77 [0.71, 0.82]), and the AUC of BNP (0.87 [0.83–0.89]) was greater than that of NT-proBNP (0.80 [0.76–0.83]). The results of the subgroup analysis showed that the cutoff range of 400–800 pg/mL for BNP had high sensitivity (0.86 [0.74–0.98]) and specificity (0.87 [0.81–0.93]) and was probably the most appropriate cutoff range.

Conclusions Elevated levels of BNP and NT-proBNP were significantly related to the mortality of patients with sepsis and had a moderate prognostic value in predicting the mortality of patients with sepsis. In addition, our meta-analysis preliminarily established appropriate cutoff values for BNP and NT-proBNP.

Keywords Brain natriuretic peptide (BNP), N-terminal pro-B-type natriuretic peptide, Sepsis, Mortality, Meta-analysis

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Background

Sepsis is defined as life-threatening organ dysfunction caused by a dysregulated host response to infection [1]. Over the years, great progress has been made in understanding the complex pathophysiology of sepsis; however, it remains the main cause of morbidity and mortality worldwide. Indeed, it is estimated that more than 30 million people worldwide are diagnosed with sepsis annually, resulting in approximately 6 million deaths [2].

Early identification of patients with sepsis at high risk of death remains a challenge. Several severity scoring systems have been developed, including Acute Physiology and Chronic Health Assessment (APACHE) and its revisions (APACHE II, III, and IV), Simplified Acute Physiology Score (SAPS) and SAP II, and Mortality Probability Model (MPM). In addition, the third international consensus definition of sepsis and septic shock (Sepsis-3) suggests using the SOFA score to predict the in-hospital mortality of patients with sepsis [3]. However, these scoring systems are often complicated and contain too many evaluation parameters, which leads to an untimely evaluation. Therefore, it is essential to identify reliable biomarkers as a valuable tool to predict the prognosis of patients with sepsis in a timely manner. Biomarkers can also assist with monitoring the progress of the disease and identify patients with an increased risk of complications, thereby representing important prognostic indicators for patients with sepsis [4]. At present, C-reactive protein (CRP), calcitonin (PCT), and other inflammatory markers (e.g., white blood cells) are widely used to aid in the diagnosis of sepsis and predict its progress. However, although these markers have certain clinical diagnostic value for sepsis, their prognostic ability is relatively limited [4, 5]. Recently, brain natriuretic peptide (BNP), a cardiac neurohormone synthesized by ventricular myocytes, has been suggested as a more useful laboratory parameter in aiding in the prognosis of sepsis. The N-terminal pro-brain natriuretic peptide precursor (NT-pro-BNP) is an inactive polypeptide of pro-hormone BNP [6]; both are synthesized in myocardial cells to respond to hemodynamic pressure or inflammatory state [7], and as prognostic markers of inflammatory state in critical patients [8–10], they have diagnostic value for patients with heart failure [11]. Two meta-analyses have been conducted on the efficacy of BNP and NT-proBNP in the prognosis of sepsis [12, 13]. However, only a few studies were included, and no potential confounding factors that might affect the prognostic value of BNP were investigated, limiting the universality of the results. In addition, neither of the two published meta-analyses conducted subgroup analysis of BNP and NT-proBNP according to the severity of sepsis, but simply considered the utility of these biomarkers in the whole spectrum of sepsis. Therefore, the efficacy of BNP and NT-proBNP in predicting

mortality may be different in patients with sepsis with different severities. In addition, the best cutoff values of two biomarkers have not yet been proposed.

Given that several related studies have been published recently, we aim to provide an updated meta-analysis to further understand the predictive value of BNP and NT-proBNP in sepsis-related mortality.

Methods

This study was performed according to the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) statement. The protocol for this meta-analysis is available in PROSPERO (CRD42022357140).

Search for trials

We searched PubMed, Web of Science, Cochrane Library, Embase and China National Knowledge Infrastructure (up to 1 January 2024) using the keywords “NT-proBNP,” “BNP,” “Septic Shock,” and “Sepsis” to identify studies that met the inclusion criteria. There were no restrictions on language. The detailed search strategy is presented in Supplement file 1.

Selection criteria

Two authors (JLS and LQQ) independently determined the eligibility of all studies identified in the initial research. The inclusion criteria were as follows: (1) adult patients with sepsis; (2) outcome, the association between NTproBNP or BNP and risk of mortality, and the prognostic value of NT-proBNP or BNP in mortality; and (3) studies with odds ratio (ORs) data;

Data extraction

Two researchers (JLS and BF) independently extracted the following information from each study: author, region, optimal timing, tested method, outcome, design, sepsis criteria, population, sample size (n); cutoff; and outcome data (sensitivity, specificity, nonsurvivors; survivors;). The OR data were also extracted. In cases where values from multivariate analyses were unavailable, those from univariate analysis were used. Discrepancies were resolved by consensus.

Quality of evidence

The quality of evidence for the included studies was assessed independently by the two researchers based on the Quality Assessment of Diagnostic Accuracy Studies (QUADAS-2).

Statistical analysis

Threshold effects were calculated by testing the Spearman correlation using STATA 14.0 software, with values >0.05 indicating no significant threshold effects. If there was no evidence of a threshold effect, then pooled

sensitivity, specificity, positive likelihood ratio (PLR), negative likelihood ratio (NLR), diagnostic odds ratio (DOR), area under the curve (AUC), and corresponding 95% credible interval (CI) were calculated using a bivariate regression model. I^2 and a bivariate boxplot were used to measure the heterogeneity caused by non-threshold effects. If the I^2 value was $\geq 50\%$ and the P-value was ≤ 0.05 , then meta-regression analysis was performed to identify the sources of heterogeneity. For meta-regression models, covariates were manipulated as mean-centered continuous or as dichotomous (yes=1, no=0) fixed effects. The effect of each covariate on sensitivity was estimated separately from that on specificity. Deek's funnel plot was used to detect publication bias, with $P < 0.05$ indicating publication bias. The following guidelines have been suggested for interpretation of intermediate AUROC values: low ($0.5 \geq \text{AUC} \leq 0.7$), moderate ($0.7 \geq \text{AUC} \leq 0.9$), or high ($0.9 \geq \text{AUC} \leq 1$) accuracy.

Results

Retrieved studies and their characteristics

The database search identified 645 records that potentially qualified for inclusion. The titles and abstracts of these records were then filtered, and after screening the abstracts, 320 articles were excluded because they were irrelevant to the current meta-analysis.

Full texts of 125 records were screened, and 47 met the inclusion criteria. Eventually, 47 studies were included in the meta-analysis (Fig. 1), of which 22 studies [14–34] reported NT-proBNP, 24 reported BNP [35–58], and 1 [59] reported both BNP and NT-proBNP. A total of 36 studies could construct a 2×2 table of results, while the remaining 11 studies only reported ORs.

Additional file 2: Table S1 lists the main characteristics of the 47 studies included in the meta-analysis. In terms of region, 30 (63.9%) trials recruited patients from Asia, 10 (21.3%) from Europe, 5 (10.6%) from North America, and 2 (4.2%) from Oceania. In terms of the subject population, 11 included patients with sepsis and 36 with severe sepsis or septic shock. In terms of trial design, 35 studies were prospective cohort studies, and 12 were retrospective cohort studies.

Quality of evidence

Figure 2 present the findings of the risk of bias assessment. Among the 47 studies analyzed in our meta-analysis, 14 studies demonstrated low bias in patient selection. Furthermore, 41 studies were deemed to have low bias in the administration of index tests, while all 47 studies were identified as having low bias in terms of reference standards. Additionally, 14 studies were found to have low bias in flow and timing. Regarding applicability concerns, 17 studies exhibited low bias in patient selection,

30 studies were rated as having low bias in relation to index tests, and all 47 studies were considered to have low bias in relation to reference standards.

Association between NT-proBNP, BNP, and mortality

For NT-proBNP, pooled analysis showed that an elevated NT-proBNP level was significantly associated with patient mortality (OR [95% CI]: 10.28 [3.30, 32.04], $P = 0.003$, $I^2 = 72.8\%$) (adjusted OR [95% CI]: 1.36 [1.20, 1.54], $P < 0.001$, $I^2 = 92.9\%$) (Figs. 3A and 4A). For BNP, pooled analysis showed that an elevated NT-proBNP level was significantly associated with patient mortality (OR [95% CI]: 8.58 [3.39, 21.71], $P < 0.001$, $I^2 = 86.8\%$) (adjusted OR [95% CI]: 1.0088 [1.0004, 1.0174], $P < 0.001$, $I^2 = 89.2\%$) (Figs. 3B and 4B).

Threshold effect and heterogeneity

The Spearman correlation coefficient and P-value for NT-proBNP and BNP were 0.07 and 0.08, respectively, which indicated that there was no significant threshold effect. We used I^2 and a bivariate boxplot to measure the heterogeneity caused by non-threshold effects (Fig. 5). For NT-proBNP and BNP, the I^2 values were 97% and 96%, respectively.

Forest plot and area under the summary ROC (SROC) curve

Forest plots of sensitivity and specificity are shown in Fig. 6. The pooled sensitivity, specificity, PLR, NLR, DOR, AUC, and corresponding 95% CI (95% CI) of NT-proBNP and BNP were 0.77 (0.68, 0.84), 0.82 (0.76, 0.87); 0.70 (0.60, 0.77), 0.77 (0.71, 0.82); 2.5 (1.9, 3.3); 3.6 (2.7, 4.6); 0.33 (0.24, 0.47), 0.23 (0.17, 0.32); 8 (4, 13), 15 (9, 26); and 0.80 (0.76–0.83), 0.87 (0.83–0.89), respectively. Supplement Fig. 1 shows the SROC curve for the prognosis of sepsis.

Likelihood ratio scattergram

For NT-proBNP and BNP, the summary LRP and LRN for index testing were on the right lower quadrant (RLQ), indicating that NT-proBNP or BNP was unable to identify patients with sepsis at high risk of dying (Supplement Fig. 2).

Publication bias

Supplementary Fig. 3 shows the assessment of publication bias. Based on the P-values of NT-proBNP and BNP (0.55 and 0.08, respectively) and the corresponding Deek's funnel plot, no significant publication bias was observed.

Pair-wise comparisons

Additional file 2: Table S2 shows the results of pair-wise comparisons between statistical indicators of sensitivity, specificity, and AUC. The pooled sensitivity



PRISMA 2009 Flow Diagram

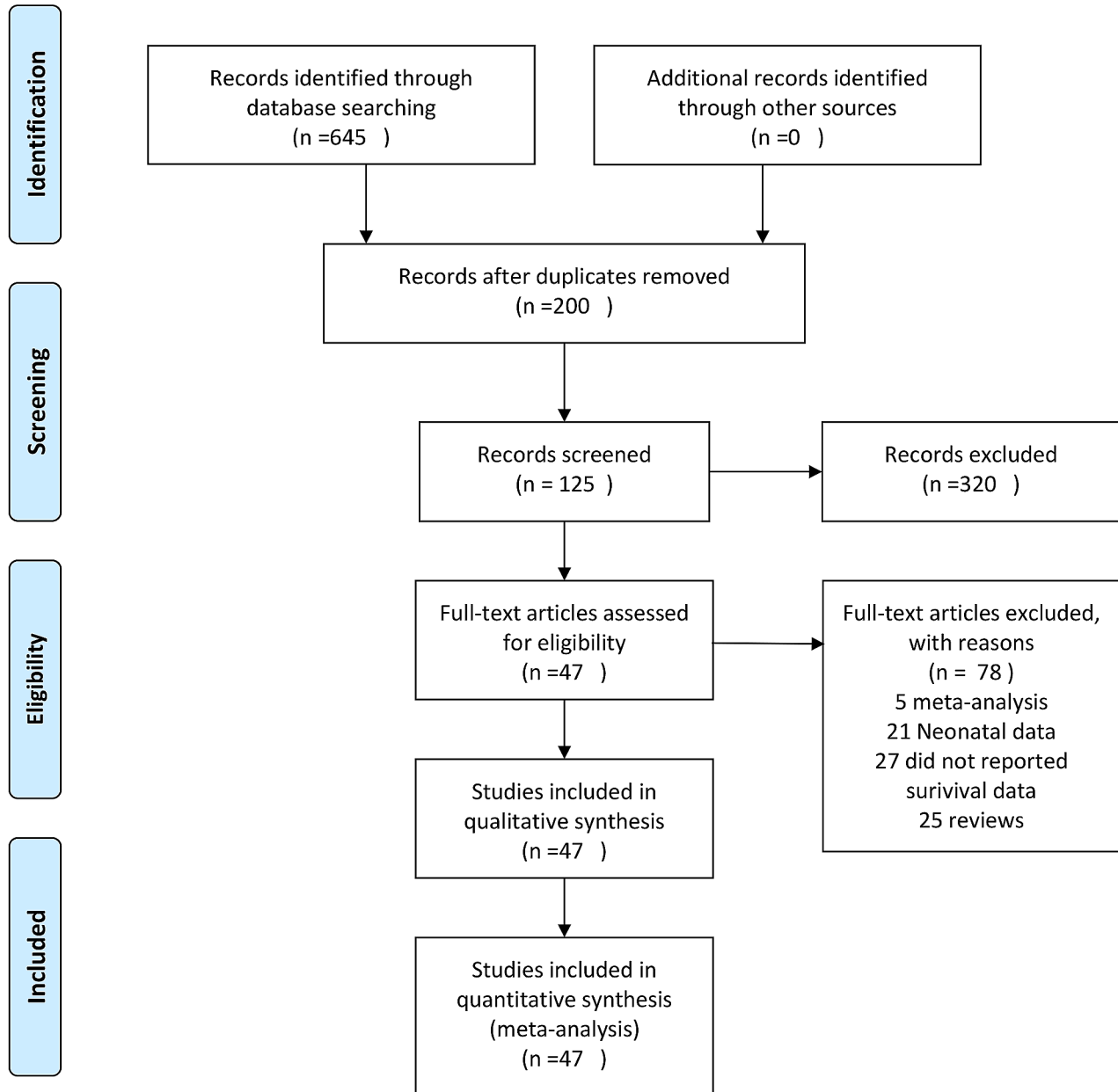


Fig. 1 Literature search and screening process

of NT-proBNP (0.77 [0.68, 0.84]) was weaker than that of BNP (0.82 [0.76, 0.87]); the pooled specificity of NT-proBNP (0.70 [0.60, 0.77]) was less than that of BNP (0.77 [0.71, 0.82]); and the AUC of BNP (0.87 [0.83–0.89]) was greater than that of NT-proBNP (0.80 [0.76–0.83]).

Meta-regression analysis

Meta-regression analysis of sensitivity, specificity, and joint models was performed to identify potential sources of heterogeneity (Additional file 2: Tables S3–S4 and Supplementary Fig. 4). According to the results of meta-regression analysis, we specified subgroups based on

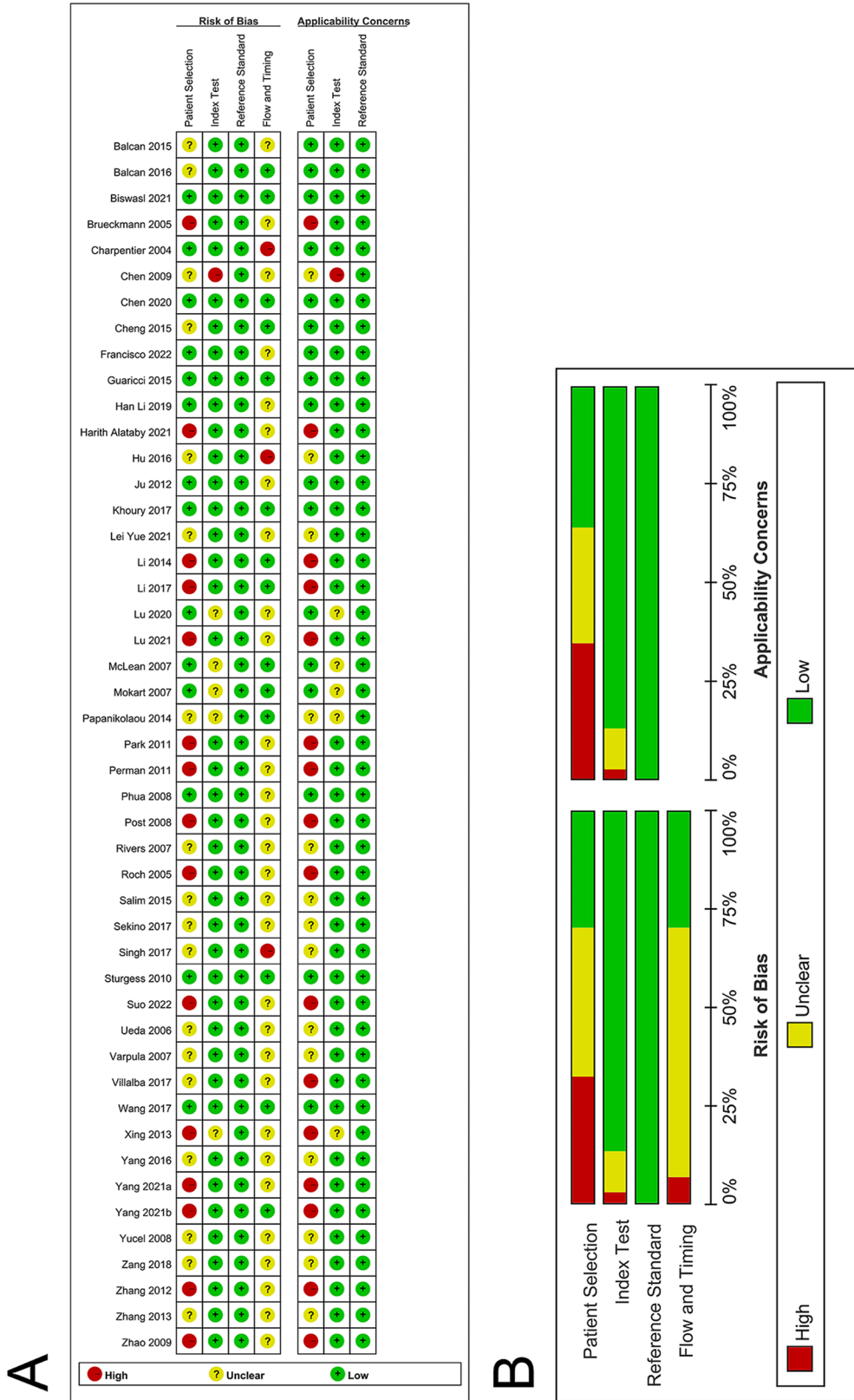


Fig. 2 Risk of bias assessment. **A:** Risk of bias summary; **B:** Risk of bias for individual studies

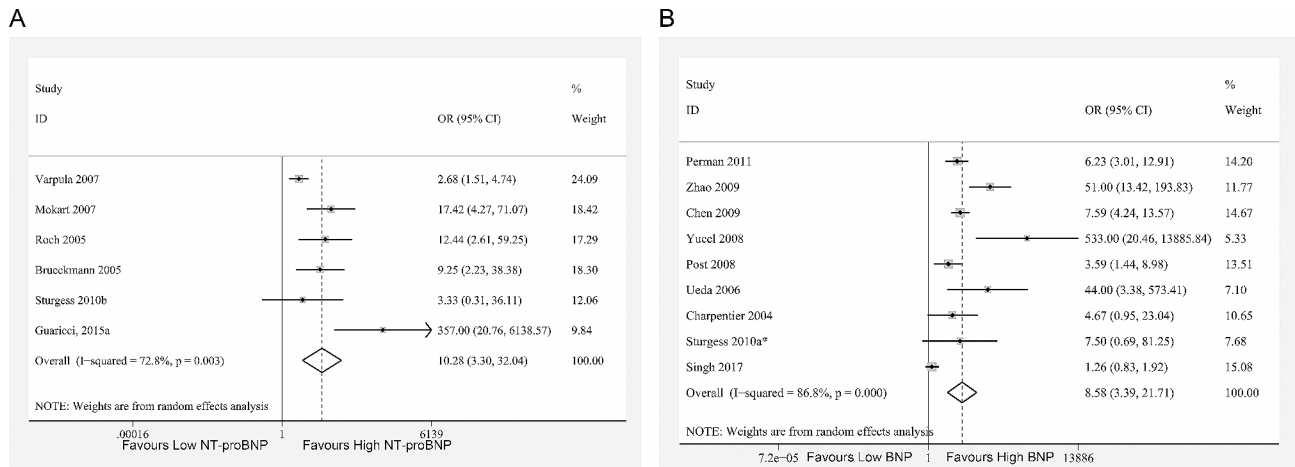


Fig. 3 Forest plot of the association between NT-proBNP or BNP and mortality in patients with sepsis; **A:** The association between NT-proBNP and mortality in patients with sepsis; **B:** The association between BNP and mortality in patients with sepsis

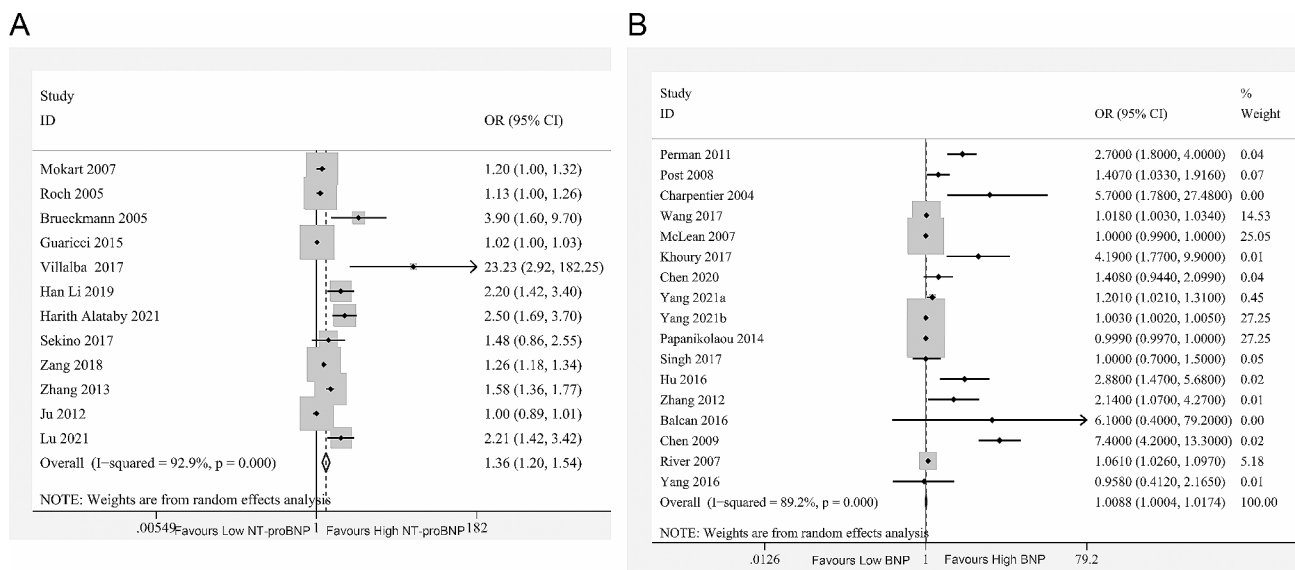


Fig. 4 Forest plot of the association between NT-proBNP or BNP and mortality in patients with sepsis adjusted for multivariate factors; **A:** The association between NT-proBNP and mortality in patients with sepsis adjusted for multivariate factors; **B:** The association between BNP and mortality in patients with sepsis adjusted for multivariate factors

population, study design, outcome, region, method, test time, and cutoff value.

Subgroup analysis

The results of the subgroup analysis are shown in Additional file 2: Tables S5, S6 and S7.

For NT-proBNP, the specificity of NT-proBNP in Europe (0.82 [0.71–0.92]) was significantly higher than that in Asia (0.65 [0.56–0.75]). In terms of study design, the sensitivity of retrospective cohort studies (0.81 [0.70–0.92]) was significantly higher than that of prospective cohort studies (0.74 [0.64–0.84]). For the cutoff, the sensitivity of NT-proBNP obtained at a cutoff interval of 3000–6000 pg/mL (0.78 [0.67–0.88]) was higher than

its sensitivity at a cutoff interval of >6000 pg/mL (0.70 [0.60–0.81]). For sepsis criteria, the specificity of Sepsis-1.0 criteria (0.78 [0.68–0.88]) was significantly higher than the Sepsis-2.0 criteria (0.70 [0.60–0.79]) and Sepsis-3.0 criteria (0.59 [0.43–0.75]). For the subject population, NT-proBNP had high sensitivity (0.85 [0.69–1.00]) and specificity (0.87 [0.81–0.94]) in patients with severe sepsis. In addition, there were no statistically significant differences in the sensitivity and specificity of outcomes, cutoff values, test time, and method.

In terms of regions, BNP had high sensitivity (0.84 [0.79–0.90]) and specificity (0.81 [0.76–0.87]) in Asia than in Europe (0.71 [0.54–0.88]) and 0.61 [0.43–0.79]) and North America (0.73 [0.56–0.90]) and 0.64

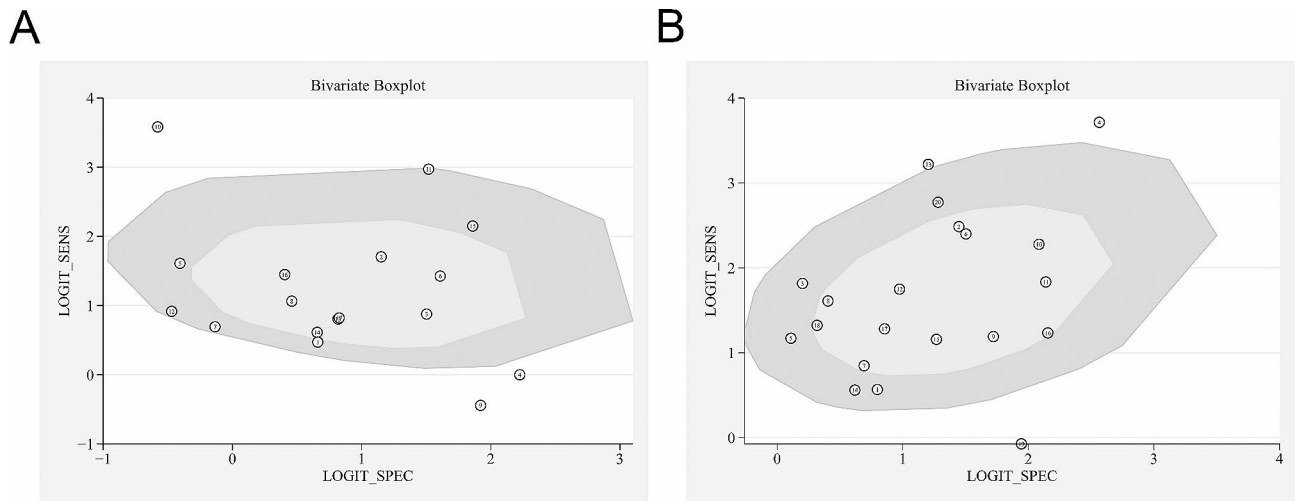


Fig. 5 Bivariate boxplots. **A:** Bivariate boxplots of NT-proBNP; **B:** Bivariate boxplots of BNP

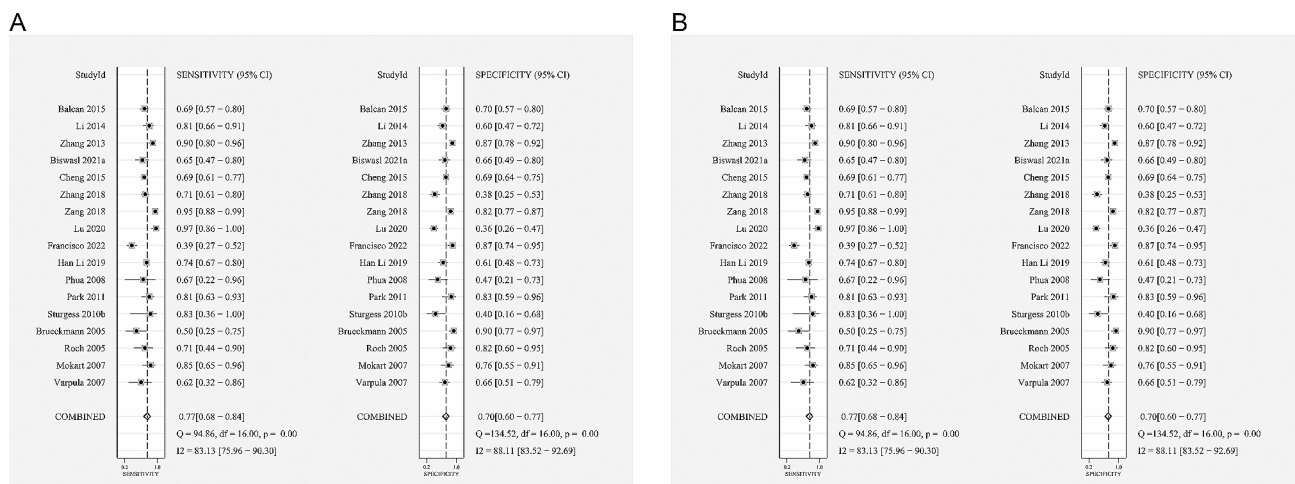


Fig. 6 Forest plots of sensitivity and specificity of NT-proBNP or BNP. **A:** Forest plots of sensitivity and specificity of NT-proBNP; **B:** Forest plots of sensitivity and specificity of BNP. Point estimates for sensitivity and 95% confidence intervals are shown with pooled estimates. Q=Cochran Q statistic

[0.46–0.83]). Regarding the method, the immunoradiometric assay was significantly more sensitive and specific (0.84 [0.79–0.90]) and (0.84 [0.79–0.90]) than the immunofluorescence assay (0.79 [0.73–0.85]) and 0.73 [0.67–0.79]). In terms of study design, the sensitivity of retrospective cohort studies (0.85 [0.77–0.91]) was significantly higher than that of prospective cohort studies (0.73 [0.67–0.80]). For BNP, a cutoff interval of 400–800 pg/L had high sensitivity (0.90 [0.83–0.94]) and specificity (0.87 [0.82–0.91]). Regarding the outcome, the sensitivity of BNP in 28-day mortality (0.83 [0.78–0.88]) was significantly higher than that of in-hospital mortality (0.57 [0.31–0.83]). For Sepsis criteria, the specificity of the Sepsis-3.0 criteria (0.83 [0.73–0.93]) was significantly higher than that of Sepsis-1.0 criteria (0.71 [0.59–0.82]). In addition, there were no statistically significant differences in the sensitivity and specificity of test time and population.

Multiple subgroup analyses

The results of multiple subgroup analyses are shown in Additional file 2: Tables S5, S6 and S7.

In terms of method and region, for NT-proBNP, the sensitivity of ECLI in Asia (0.82 [0.69–0.92]) was significantly higher than that in Europe (0.65 [0.45–0.85]). For BNP, the immunofluorescence assay had high sensitivity (0.82 [0.75–0.88]) and specificity (0.78 [0.71–0.85]) in Asia than in Europe (0.71 [0.53–0.90]) and 0.59 [0.37–0.80], respectively) and North America (0.73 [0.58–0.88]) and 0.64 [0.45–0.82], respectively).

In terms of cutoff and region, in Asia, the sensitivity of NT-proBNP obtained at a cutoff interval of <3000 pg/mL (0.90 [0.81–0.99]) was higher than its sensitivity at a cutoff interval of 3000–6000 pg/mL (0.78 [0.65–0.91]) and >6000 pg/mL (0.73 [0.67–0.80]). For BNP, the cutoff range of 400–800 pg/mL had high sensitivity (0.86 [0.74–0.98]) and specificity (0.87 [0.81–0.93]) and was probably

the most appropriate cutoff range in Asia. Moreover, the sensitivity of BNP obtained at a cutoff interval of 400–800 pg/mL (0.88 [0.82–0.93]) was higher than its sensitivity at a cutoff value <400pg/mL (0.83 [0.78–0.88]) and >800 pg/mL (0.76 [0.61–0.90]), while its specificity obtained at a cutoff interval of 400–800 pg/mL (0.87 [0.82–0.92]) was higher than its specificity at a cutoff value of <400 pg/mL (0.71 [0.62–0.80]).

In terms of the Sepsis criteria and population, for Sepsis-3.0 criteria, the sensitivity of NT-proBNP in patients with sepsis (0.84 [0.71–0.96]) was significantly higher than that in patients with sepsis and septic shock (0.58 [0.41–0.75]). In the Sepsis-1.0/2.0 criteria, the specificity of NT-proBNP in patients with severe sepsis (0.87 [0.81–0.93]) was significantly higher than that in patients with septic shock (0.68 [0.57–0.79]). For the Sepsis-3.0 criteria, the specificity of NT-proBNP (0.59 [0.43–0.76]) in predicting mortality in patients with all subtypes of sepsis was significantly lower than that of BNP (0.83 [0.71–0.95]), but no significant difference between these two markers was found in the Sepsis-1.0/2.0 criteria. The specificity of NT-proBNP (0.53 [0.33–0.72]) in predicting mortality in patients with sepsis was significantly lower than that of BNP (0.77 [0.62–0.91]), but there was no statistical difference between them in predicting mortality in patients with septic shock and severe sepsis. The above results show that the ability of BNP in predicting the mortality of all subtypes of sepsis in the Sepsis-3.0 criteria was higher than that of NT-proBNP, but this was only reflected in predicting ordinary sepsis; for patients with severe sepsis and septic shock, there was no statistical difference between the two markers in predicting the mortality of patients.

Sensitivity analysis

Sensitivity analysis was performed by sequential exclusion of each study. Additional file 1: Tables S8 and S9 show the combined DOR and 95% CI calculated after deleting each study. The combined DOR after removal did not change significantly, suggesting that the results were robust.

Discussion

The meta-analysis showed that elevated levels of BNP and NT-proBNP were significantly related to the mortality of patients with sepsis. In addition, SROC curve analysis showed that NT-proBNP and BNP had moderate prognostic value in predicting the mortality of patients with sepsis. Finally, our meta-analysis preliminarily established appropriate cutoff values and intervals of BNP and NT-proBNP. The most appropriate cutoff range of BNP and NT-proBNP was 400–800 and <3000 pg/mL, respectively. However, the heterogeneity of our results limits the strength of these conclusions.

Although the concept of sepsis is widely used, the standard definition of this common disease has not yet been established, which leads to differences in the criteria for diagnosis. Therefore, the term sepsis probably includes diseases and severities that differ among the studies included in this meta-analysis, which may also explain the high levels of heterogeneity observed in our analysis. In addition, we performed meta-regression analysis based on population, study design, outcome, region, method, test time, and cutoff value to identify potential sources of heterogeneity. The results of subgroup analysis showed significant differences in the sensitivity and specificity of NT-proBNP in terms of cutoff value, study design, regions, Sepsis criteria, and population. For BNP, there were significant differences in outcome, region, study design, cutoff value, detection method, and Sepsis criteria.

It has been shown that BNP values are significantly different between patients with sepsis and septic shock [60, 61] and between patients with severe sepsis and septic shock [43, 62]; however, no clear cutoff value has been proposed for this purpose to distinguish patients in these stages in clinical practice. The results of subgroup analysis showed that the sensitivity and specificity of NT-proBNP in predicting the mortality of severe sepsis was higher than that of septic shock and sepsis. For BNP and NT-proBNP, the specificity of NT-proBNP in predicting mortality in patients with sepsis was significantly lower than that of BNP, but there was no statistical difference between them in predicting mortality in patients with septic shock and severe sepsis. However, because the studies included in this meta-analysis did not provide cutoff value information for different stages of sepsis syndrome, we could not further evaluate the ability of markers to distinguish septic shock from sepsis or severe sepsis.

In addition to BNP and NT-proBNP, which were identified in this study as predictors of mortality in sepsis patients, emerging biomarkers such as cell free DNA, Pentraxin-3, and Neutrophil-to-lymphocyte ratio have demonstrated potential prognostic value in predicting mortality in sepsis [63–65]. Recent meta-analysis findings suggest that the AUC values of these novel biomarkers ranged from 0.73 to 0.80, slightly lower than those of BNP and NT-proBNP in the current study [63–65]. However, there is a lack of comparative studies evaluating BNP and NT-proBNP against these biomarkers for prognostic purposes in sepsis, thus hindering the assessment of their relative effectiveness. Moreover, the clinical utility of these novel biomarkers is limited by various factors, including infrequent measurement in routine clinical settings, small sample sizes in studies, and incomplete understanding of the causal relationship between these biomarkers and sepsis outcomes. Consequently, current

guidelines recommend against the use of biomarkers for prognostic evaluation in sepsis [66]. Although previous studies showed that BNP and NT-proBNP were more sensitive than SOFA [39, 45, 49], our meta-analysis shows that the use of BNP or NT-proBNP alone cannot predict the mortality of patients with sepsis. Considering the low sensitivity and high specificity of the clinical severity score, future studies should investigate whether single or multiple biomarkers, when combined with the Clinical Severity Score, can provide a more accurate assessment of the prognosis of sepsis.

Our meta-analysis has several limitations. First, there is a high degree of heterogeneity among the studies, although we conducted regression analysis and subgroup analysis on the factors in the team. Secondly, renal failure and ventricular dysfunction resulting from sepsis, along with interventions like catecholamine administration and volume resuscitation, contribute to an increase in BNP/NT-proBNP levels [67]. Conversely, certain sepsis treatments such as positive inotropic agents (levosimendan and dobutamine), intra-aortic balloon pump (IABP) insertion, and continuous renal replacement therapy (CRRT) have been shown to lower BNP/NT-proBNP levels [68–70]. In evaluating the prognostic value of BNP in sepsis, the abnormal renal function of patients with sepsis is still a major confounding factor, because the included studies exclude pre-existing chronic kidney diseases to varying degrees, and the adjustment of acute kidney injury is inconsistent in the analysis. In patients with sepsis, we showed contradictory results regarding the correlation between BNP and serum creatinin [29, 50]. Therefore, further research is needed to develop the clinically relevant BNP critical value and stratify patients with sepsis according to their renal function to determine the BNP range of these patients more effectively. Moreover, most studies exclude patients with pre-existing heart disease and do not systematically evaluate cardiac dysfunction. As these studies did not provide the aforementioned details (e.g., fluid balance, cardiac dysfunction, renal function, positive inotropic agents, IABP and CRRT), these factors were not systematically assessed in the individual studies included in our analyses, which limits the generalizability of our findings.

However, despite the above limitations, our results showed that elevated levels of BNP and NT-proBNP were significantly related to the mortality of patients with sepsis and had moderate prognostic value in predicting the mortality of patients with sepsis. In addition, our meta-analysis preliminarily established appropriate cutoff values for BNP and NT-proBNP.

Abbreviations

BNP	Brain Natriuretic Peptide
NT-proBNP	N-terminal pro-B-type natriuretic peptide
APACHE	Acute Physiology and Chronic Health Assessment

SAPS	Simplified Acute Physiology Score
CRP	C-Reactive Protein
PLR	Positive Likelihood Ratio
NLR	Negative Likelihood Ratio
DOR	Diagnostic Odds Ratio
AUC	Area Under the Curve
SOFA	Sequential Organ Failure Assessment
CI	Credible Interval
LVEF	Left Ventricular Ejection Fraction

Supplementary Information

The online version contains supplementary material available at <https://doi.org/10.1186/s12871-024-02661-z>.

Supplement File 1. Summary receiver-operating characteristic (SROC) curves for the prognosis of sepsis. A:SROC of NT-proBNP; B:SROC of BNP
AUC = area under the curve

Supplement File 2. Likelihood ratio scattergrams. A:Scattergrams evaluating the positive likelihood ratios in the prognosis of sepsis for NT-proBNP; or Scattergrams evaluating the positive likelihood ratios in the prognosis of sepsis for BNP.

Supplement File 3. Deek's funnel plots. A:Funnel plots evaluating publication bias of NT-proBNP;B:Funnel plots evaluating publication bias of BNP.

Supplement File 4. Univariable meta-regression analysis of sensitivity and specificity. A:Univariable meta-regression analysis of sensitivity and specificity of NT-proBNP;B:Univariable meta-regression analysis of sensitivity and specificity of BNP.

Supplement File 5: The Search strategy for pubmed.

Supplement File 6: Supplement tables 1-9 Supplement. table 1: Literature search and characteristics of the included studies, Supplement table 2: Pair-wise comparisons between modalities for sensitivity, Specificity, and AUC. Supplement table 3: The result of meta-regression and subgroup analysis for NT-proBNP. Supplement table 4: The result of meta-regression and subgroup analysis for BNP. Supplement table 5: Subgroup analysis of region and detection method for NT-proBNP and BNP. Supplement table 6: Subgroup analysis of region and cutoff level method for NT-proBNP and BNP. Supplement table 7: Subgroup analysis of sepsis criteria and population for NT-proBNP and BNP. Supplement table 8: Sensitivity analyses of NT-proBNP. Supplement table 9: Sensitivity analyses of BNP.

Supplement File 7 : Prisma 2020 checklist.

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Author contributions

GYC made substantial contributions conception and design of the study; JLS and LQQ searched and screened literature; JLS, BF, QL and BL extracted data from the collected literature and analyzed the data; JLS, LQQ and GYC wrote the manuscript; QL, BF and BL revised the manuscript. All the authors approved the final version of manuscript.

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Data availability

No datasets were generated or analysed during the current study.

Declarations

Ethics approval and consent to participate

This article is meta-analysis and does not require ethics committee approval or a consent statement.

Consent for publication

Not applicable. The manuscript does not contain any personal data in any form (including personal details, pictures, or videos).

Competing interests

The authors declare no competing interests.

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