

RESEARCH ARTICLE

Fixed-dose combination antihypertensive medications, adherence, and clinical outcomes: A population-based retrospective cohort study

Amol A. Verma^{1,2*}, Wayne Khuu³, Mina Tadrous^{3,4}, Tara Gomes^{3,4,5}, Muhammad M. Mamdani^{3,4,5}

1 Li Ka Shing Centre for Healthcare Analytics Research and Training, St. Michael's Hospital, Toronto, Ontario, Canada, **2** Eliot Phillipson Clinician-Scientist Training Program, Department of Medicine, University of Toronto, Toronto, Ontario, Canada, **3** Institute for Clinical and Evaluative Sciences, Toronto, Ontario, Canada, **4** Leslie Dan Faculty of Pharmacy, University of Toronto, Toronto, Ontario, Canada, **5** Institute of Health Policy, Management, and Evaluation, University of Toronto, Toronto, Ontario, Canada

* amol.verma@mail.utoronto.ca



OPEN ACCESS

Citation: Verma AA, Khuu W, Tadrous M, Gomes T, Mamdani MM (2018) Fixed-dose combination antihypertensive medications, adherence, and clinical outcomes: A population-based retrospective cohort study. *PLoS Med* 15(6): e1002584. <https://doi.org/10.1371/journal.pmed.1002584>

Academic Editor: Anushka Patel, The George Institute for Global Health, AUSTRALIA

Received: January 23, 2018

Accepted: May 11, 2018

Published: June 11, 2018

Copyright: © 2018 Verma et al. This is an open access article distributed under the terms of the [Creative Commons Attribution License](https://creativecommons.org/licenses/by/4.0/), which permits unrestricted use, distribution, and reproduction in any medium, provided the original author and source are credited.

Data Availability Statement: The data set from this study are held securely in coded form at ICES. Datasets provided by ICES were linked using unique encoded identifiers and analyzed at ICES. While data sharing agreements prohibit ICES from making the dataset publicly available, access may be granted to those who meet prespecified criteria for confidential access, available at www.ices.on.ca/DAS. The dataset creation plan and analytic plan have been made available with this manuscript ([S1 Dataset Creation and Analysis Plan](#)).

Abstract

Background

The majority of people with hypertension require more than one medication to achieve blood pressure control. Many patients are prescribed multipill antihypertensive regimens rather than single-pill fixed-dose combination (FDC) treatment. Although FDC use may improve medication adherence, the impact on patient outcomes is unclear. We compared clinical outcomes and medication adherence with FDC therapy versus multipill combination therapy in a real-world setting using linked clinical and administrative databases.

Methods and findings

We conducted a population-based retrospective cohort study of 13,350 individuals 66 years and older in Ontario, Canada with up to 5 years of follow-up. We included individuals who were newly initiated on one angiotensin-converting enzyme inhibitor (ACEI) or angiotensin II-receptor blocker (ARB) plus one thiazide diuretic. High-dimensional propensity score matching was used to compare individuals receiving FDC versus multipill therapy. The primary outcome was a composite of death or hospitalization for acute myocardial infarction (AMI), heart failure, or stroke. We conducted 2 analyses to examine the association between adherence and patient outcomes. First, we performed an on-treatment analysis to determine whether outcomes differed between groups while patients were on treatment, censoring patients when they first discontinued treatment, defined as not receiving medications within 150% of the previous days' supply. Second, we conducted an intention-to-treat analysis that followed individuals allowing for breaks in treatment to quantify the difference in drug adherence between groups and assess its impact on clinical outcomes. As expected, there was no significant difference in the primary outcome between groups in the on-treatment analysis (HR 1.06, 95% CI 0.86–1.31, $P = 0.60$). In the intention-to-treat analysis, the proportion of total follow-up days covered with medications was significantly greater

Funding: This study was supported by the Ontario Drug Policy Research Network (ODPRN), which is funded by a grant from the Ontario Ministry of Health and Long-Term Care (MOHLTC; grant #06673), the Health System Research Fund (HSRF), and CIHR's Strategy for Patient-Oriented Research (SPOR) Support Unit. This study was also supported by the Institute for Clinical Evaluative Sciences (ICES), which is funded by an annual grant from the MOHLTC. The funders had no role in the study design, data collection and analysis, decision to publish, or preparation of the manuscript.

Competing interests: I have read the journal's policy and the authors of this manuscript have the following competing interests: MMM has served on the following advisory boards and reports personal fees from Bristol-Myers Squibb, Eli Lilly and Company, Glaxo Smith Kline, Hoffman La Roche, Novartis, Novo Nordisk, Pfizer, and Astra Zeneca, outside of the submitted work. All other authors have no competing interests to disclose.

Abbreviations: ACEI, angiotensin-converting enzyme inhibitor; AMI, acute myocardial infarction; ARB, angiotensin II-receptor blocker; FDC, single-pill fixed-dose combination; ICES, Institute for Clinical Evaluative Sciences; TIA, transient ischemic attack.

in the FDC group (70%; IQR 19–98) than in the multipill group (42%, IQR 11–91, $P < 0.01$), and the primary outcome was less frequent in FDC recipients (3.4 versus 3.9 events per 100 person-years; HR 0.89, 95% CI 0.81–0.97, $P < 0.01$). The main limitations of this study were the lack of data regarding cause of death and blood pressure measurements and the possibility of residual confounding.

Conclusions

Among older adults initiating combination antihypertensive treatment, FDC therapy was associated with a significantly lower risk of composite clinical outcomes, which may be related to better medication adherence.

Author summary

Why was this study done?

- Most people with hypertension require more than one medication to achieve blood pressure control.
- Multiple medications can be combined into a single pill or prescribed as separate pills.
- Patients are more likely to adhere to single-pill rather than multiple-pill treatment regimens.
- It is unclear whether clinical outcomes are better in patients who receive single-pill combinations or multiple separate pills and whether this might be related to medication adherence.

What did the researchers do and find?

- We used administrative and clinical databases in Ontario, Canada to compare adults 65 years and older who were starting combination blood pressure treatment with either single-pill or multiple-pill regimens.
- We used advanced statistical methods to identify a cohort of individuals who were comparable.
- We found that people who received single-pill combinations had a significantly lower rate of the combined outcome of death or hospitalization for heart attack, heart failure, or stroke and that these differences were related to better medication adherence.

What do these findings mean?

- Approximately 675 million people globally require combination antihypertensive therapy, and up to 40% of patients with hypertension in high-income countries are treated with multiple-pill regimens.

- Our study suggests that single-pill combination treatment is associated with markedly improved medication adherence and better clinical outcomes.
- Using single-pill combinations rather than multiple pills may represent a simple and potentially low-cost intervention that could substantially reduce the global burden of death and disability related to hypertension.

Introduction

Hypertension affects an estimated 900 million adults [1] and is the leading cause of global death or disability [2]. Approximately 75% of people with hypertension require more than one medication to achieve blood pressure control [3]. Although many hypertension management guidelines recommend initiating combination treatment with either separate drugs in multipill combinations or single-pill fixed-dose combinations (FDCs) [4–6], only half of national hypertension societies recommend FDC treatment [7], which may be due to a lack of evidence about effect on clinical outcomes. Both FDC and multipill regimens are common in clinical practice. Studies in high-income countries report that 22% to 43% of patients receive multipill regimens when initiated on combination antihypertensive therapy [8,9].

The blood pressure-lowering effect of FDC therapy was similar to multipill regimens in a meta-analysis of 9 trials [10], which were all less than 1 year in duration. However, blood pressure control was found to be worse with FDC over 6 years' follow-up in the Swiss Hypertension Cohort Study [11]. It may not be possible to extrapolate the effectiveness of FDC therapy for hypertension from clinical trials to real-world settings. Although FDC use has been associated with improved medication adherence compared with multipill therapy in both clinical trials and observational settings [8,10,12,13], critics of FDC therapy argue that it makes dose titration or changing medications more difficult and that this could lead to poorer outcomes [14,15]. It is not known whether improved adherence related to FDC therapy translates into better clinical outcomes.

We examined the association between initiating FDC versus multipill antihypertensive therapy and cardiovascular events or death in a real-world setting. Addressing this question in a real-world setting is particularly important because the differences between FDC and multipill therapy arise from the way medications are used, and patterns of medication use in clinical trials may not be generalizable [16].

Methods

Setting and design

We conducted a population-based, propensity score-matched, retrospective cohort study of residents of Ontario, Canada aged 66 years or older who initiated combination antihypertensive therapy between April 1, 2004 and December 31, 2014. Individuals were followed for 5 years or until March 31, 2015. Provincial health insurance in Ontario covers physician and hospital services for all residents and prescription drugs for those over 65 years of age. This study was conducted using a prespecified analysis plan approved by the Sunnybrook Health Sciences Centre Research Ethics Board. This study is reported as per the RECORD guidelines (S1 Checklist).

Sources of data

We used the Ontario Drug Benefit claims database, which records prescription medications dispensed to all Ontarians over the age of 65, to determine exposure to combination antihypertensive therapy. Data pertaining to hospitalizations and emergency department use were obtained from the Canadian Institute for Health Information Discharge Abstract Database and National Ambulatory Care Reporting System. The Registered Persons Database was used to obtain basic demographic information and date of death for all Ontario residents. Demographic and specialty data for all physicians practicing in Ontario were obtained from the Institute for Clinical Evaluative Sciences (ICES) Physicians Database, and the Ontario Health Insurance Plan claims database was used to identify claims for all insured physician services. These databases have excellent data completeness and quality [17], and they were anonymously linked using encrypted person-level identifiers, as in previous studies [18–20].

Cohort design

We identified a cohort of new users of combination antihypertensive medication who were prescribed one angiotensin-converting enzyme inhibitor (ACEI) or angiotensin II-receptor blocker (ARB) plus one thiazide diuretic, either as FDC or as a multipill combination. These medication combinations were selected because they are common guideline-recommended options for initial antihypertensive therapy [4,5] and to avoid potential confounding when comparing combinations of various medication classes. We included only new users of antihypertensive medications to avoid selection bias based on prior medication adherence. New users were defined as receiving no prescription for any antihypertensive medication in the year prior to study enrollment.

The index date for study enrollment was defined as the date of first prescription of antihypertensive medications. To match FDC therapy, in which both medications are taken together, the multipill combination group included only individuals who were dispensed both medications on the same index date. Initiating multipill combination therapy on separate days might reflect treatment intensification in response to failure of monotherapy and select for a higher-risk population, therefore these individuals were not included.

We excluded individuals with any hospitalization for stroke, transient ischemic attack (TIA), heart failure, or myocardial infarction in the year prior to study enrollment to reduce selection bias that might arise from differential prescribing of FDC versus multipill combinations after hospitalization. We also excluded individuals with any emergency department visit for stroke or TIA in the year prior to study enrollment because the combination of ACEI and thiazide may be used for secondary stroke prevention even among nonhypertensive adults [21]. Finally, we excluded individuals who were prescribed any antihypertensive medications in addition to the initial combination therapy on the day of study enrollment. See cohort flow diagram for details (Fig 1).

To minimize selection bias, we used high-dimensional propensity score matching to identify comparable groups [22]. Each multipill combination user was matched with one FDC user based on the dose of index antihypertensive medications and the propensity score (allowing no more than a difference of 20% of the SD of the logit of the propensity score between matched pairs). The following datasets were used to develop the high-dimensional propensity score based on data from the year prior to study enrollment: prescription drug claims, diagnosis codes, and procedure codes from all hospitalizations and emergency department visits as well as insurance claims and diagnosis codes for physician services. In total, this represented 7 dimensions of data, and the 200 most prevalent codes from each dimension were retained as candidate covariates. All potential covariates were sorted in descending order by the magnitude of the log

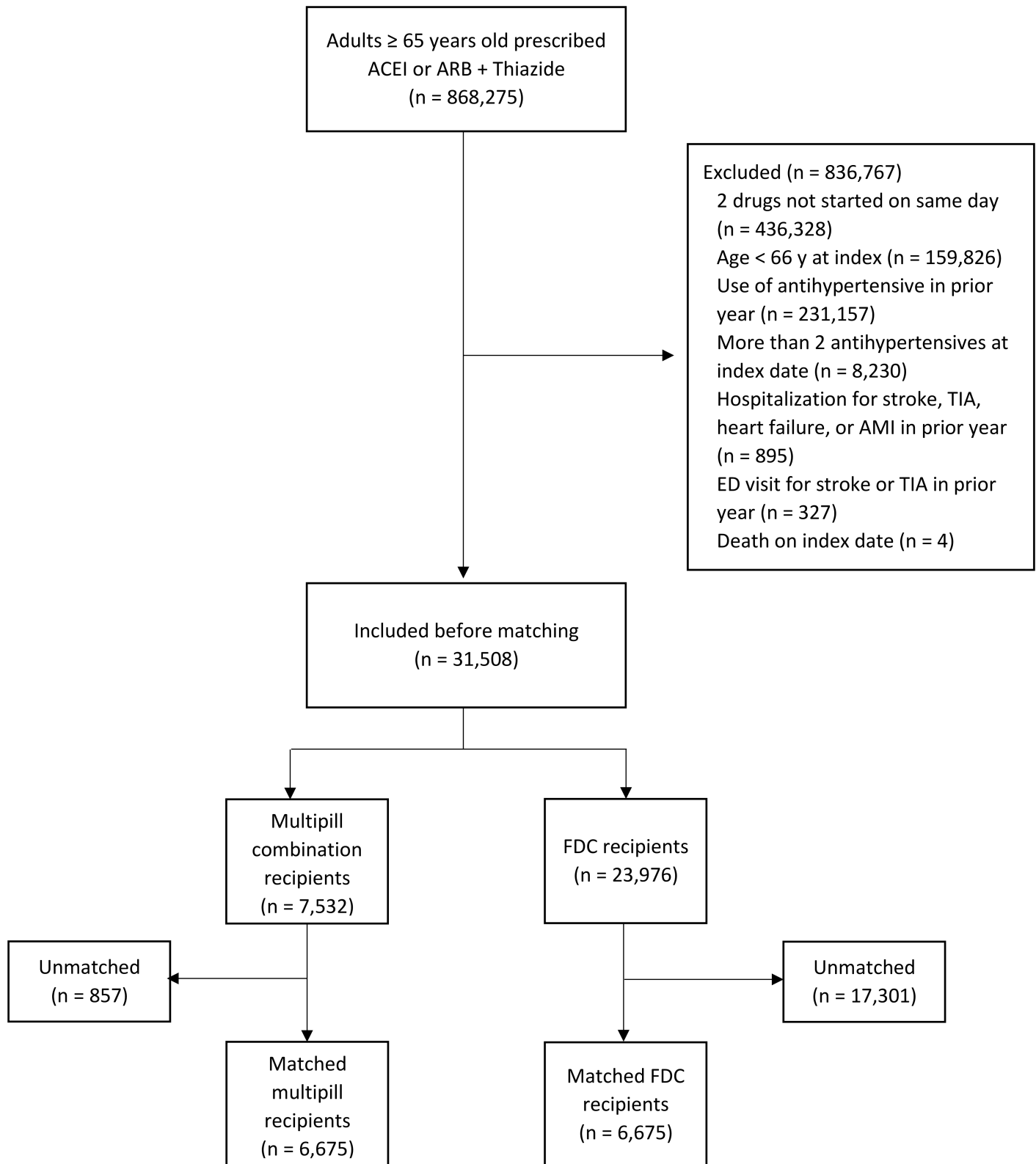


Fig 1. Cohort creation diagram. ACEI, angiotensin-converting enzyme inhibitor; AMI, acute myocardial infarction; ARB, angiotensin II-receptor blocker; ED, emergency department; FDC, single-pill fixed-dose combination; TIA, transient ischemic attack.

<https://doi.org/10.1371/journal.pmed.1002584.g001>

of the multiplicative bias term as previously described [22], and the top 500 covariates were included in the propensity score model. We also included the following prespecified covariates: age, sex, income quintile, year of the index date to account for possible changes in clinical practice over the duration of the study period, Charlson comorbidity index, the number of outpatient physician visits in the year prior to study enrollment, and any cardiology visit in the 3 months prior to study enrollment. The high-dimensional propensity score was calculated using version 2.4.4 of the freely available macro developed at Harvard University [23].

The dose of index antihypertensive medications was used as a proxy for severity of hypertension at baseline. As described in the supporting information, each medication was categorized into high- and low-dose categories based on the usual daily dose range described in clinical practice guidelines (S1 Table) [24]. Individuals were then categorized into 3 groups based on the dose of their index antihypertensives: “low” if both medications were low-dose, “high” if both medications were high-dose, and “intermediate” if only one medication was high-dose.

We conducted 2 complementary analyses to disentangle whether differences in outcomes between FDC and multipill regimens might be related to improved adherence or rather, differences in the actual effects of medications between groups. First, we performed an on-treatment analysis, censoring patients when they first discontinued treatment. Therefore, outcomes were assessed only during active treatment, which removed the effect of adherence. Second, we conducted an intention-to-treat analysis that followed individuals irrespective of disruptions in treatment to quantify the difference in drug adherence between groups and assess its impact on clinical outcomes.

We present 2 measures to describe adherence to antihypertensive medications based on medication dispensing: the time to the first instance of discontinuation and the proportion of total days covered. Discontinuation of antihypertensive medication was defined as no repeat prescription within 150% of the previous days' supply. For example, a medication would be considered discontinued if the index prescription was dispensed for 30 days and a second prescription was not dispensed within 45 days. For the multipill combination group, repeat prescription of both medications within the 150% grace period was necessary to be considered continuous use. Medication class switching between ACEI or ARB or between different thiazides was not considered discontinuation. The on-treatment analysis followed individuals until the first instance of discontinuation. Because individuals may subsequently receive the index antihypertensive medications after an initial disruption in therapy, we also calculated the total duration of use of the index antihypertensive medications during follow-up. This was reported as the proportion of days of follow-up covered by the index antihypertensive medications in all periods of continuous use.

We hypothesized that, if the benefit of FDC therapy was related to improved medication adherence, the intention-to-treat analysis would demonstrate significant between-group differences in clinical outcomes whereas the on-treatment analysis would not.

Outcomes

The primary outcome was a composite of all-cause death and hospitalization for acute myocardial infarction (AMI), heart failure, or stroke. The administrative diagnostic codes (S6 Table) used to define the primary outcome have good sensitivity and positive predictive value, respectively: AMI, 89% and 87%; heart failure, 79% and 85%; and stroke, 76% and 97% [25,26]. Secondary outcomes included each individual component of the primary outcome, hospitalization for hypokalemia or hyponatremia to assess for safety, and the first instance of discontinuation of antihypertensive medication. Because we expected no association between antihypertensive treatment and cataract surgery, this was used as a “tracer outcome” to assess for residual confounding between the groups.

Statistical analysis

Propensity score matching resulted in balanced groups with no baseline covariates differing by more than 0.1 standardized difference except for index medication, which differed because ARBs were more common in FDC formulations than ACEI. Adjusting for covariates with less than 0.1 standardized difference imbalance provides negligible benefits in addressing residual confounding [27]. Thus, no additional covariates were included in the regression models. In the on-treatment analysis, individuals were censored when they first discontinued antihypertensive medications. The median time to the first instance of medication discontinuation was compared between the FDC and multipill combination groups using Kaplan-Meier estimators and the log-rank test. The proportion of days covered by antihypertensive medications in both groups was compared using the Wilcoxon signed rank sum test. Our primary analysis was a time-to-event analysis in the matched cohort using Cox proportional hazards regression including a robust variance estimator that accounted for clustering within matched sets [28]. When analyzing each component of the primary composite outcome separately, individuals were also censored for death. Proportional hazards assumptions were tested for the primary outcome using a time-dependent covariate by including an interaction term between the antihypertensive group and time in the Cox models. All of the above analyses were prespecified, with the exception of the calculation of the proportion of days covered by index antihypertensive medications, which was a post hoc analysis conducted to better describe medication use after we recognized that many participants continued to receive index medications after an instance of medication discontinuation.

The following additional analyses were provided in response to peer review comments. First, we conducted a sensitivity analysis by defining medication discontinuation as no repeat prescription within 300% of the previous days' supply. Second, to better understand the role that concomitant cardiovascular risk-modifying treatments may have played in influencing our findings, we described the receipt of other cardiovascular medications in the last 90 days of follow-up in the FDC and multipill groups. A standardized difference less than 0.1 between groups was considered to be well-balanced. Finally, to better understand patterns of medication use, we reported antihypertensive and other cardiovascular medications dispensed in the last 90 days of follow-up among individuals who did and did not discontinue their index antihypertensive medications, defined as having any break in therapy of greater than 150% of the previous days' supply.

All analyses were performed using SAS software (version 9.4).

Results

Study cohort and follow-up

After propensity score matching, we identified a cohort of 13,350 individuals (6,675 in each group) who were new users of combination antihypertensive therapy with an ACEI or ARB plus a thiazide diuretic. In the intention-to-treat analysis, median follow-up time was 1,826 days in both the FDC group (interquartile range 1,163–1,826) and the multipill group (IQR 1,142–1,826).

Baseline characteristics

The 2 groups were well-balanced on all baseline characteristics except medication class at index (Table 1) and were relatively similar before matching (S3 Table). The median age at index was 71 years (IQR 68–77). In both groups, 42.7% of individuals received low-dose medication, 43.0% received intermediate-dose, and 14.3% received high-dose. FDC users were more likely

Table 1. Baseline characteristics in propensity score-matched study cohort.

Characteristic	Multipill	FDC	Standardized Difference
	(N = 6,675)	(N = 6,675)	
Age, median (IQR), y	71 (68–77)	71 (68–77)	0
Female, n (%)	3,680 (55.1)	3,590 (53.8)	0.03
Neighborhood income quintile, n (%)			
1	1,345 (20.1)	1,403 (21.0)	0.02
2	1,388 (20.8)	1,371 (20.5)	0.01
3	1,303 (19.5)	1,268 (19.0)	0.01
4	1,244 (18.6)	1,231 (18.4)	0.01
5	1,342 (20.1)	1,350 (20.2)	0
Missing	53 (0.8)	52 (0.8)	0
Nursing home residence, n (%)	121 (1.8)	50 (0.7)	0.09
Rural residence, n (%)	951 (14.2)	785 (11.8)	0.07
Charlson comorbidity score, categorized			
No hospitalizations	5,917 (88.6)	5,923 (88.7)	0
0	471 (7.1)	439 (6.6)	0.02
1	136 (2.0)	163 (2.4)	0.03
2+	151 (2.3)	150 (2.2)	0
Healthcare utilization			
Hospitalizations in prior year, mean (SD)	0.09 (0.33)	0.09 (0.34)	0.01
Outpatient physician visits in prior year, median (IQR)	4 (2–8)	4 (2–8)	0
Visit to cardiologist in prior 3 months, n (%)	1,148 (17.2)	1,108 (16.6)	0.02
Cardiac catheterization in prior 5 years, n (%)	74 (1.1)	57 (0.9)	0.03
Total number of different prescription drugs in prior 100 days, mean (SD)	1.47 (2.17)	1.48 (2.13)	0
Medical comorbidities, n (%)			
Diabetes [†]	1,129 (16.9)	1,175 (17.6)	0.02
Stroke ^{††}	45 (0.7)	28 (0.4)	0.03
AMI ^{††}	17 (0.3)	13 (0.2)	0.01
Heart Failure ^{††}	40 (0.6)	31 (0.5)	0.02
Peripheral vascular disease ^{††}	32 (0.5)	33 (0.5)	0
Chronic kidney disease ^{††}	18 (0.3)	16 (0.2)	0.01
Cancer [†]	673 (10.1)	675 (10.1)	0
Chronic obstructive pulmonary disease [†]	352 (5.3)	348 (5.2)	0
Dementia ^{††}	349 (5.2)	320 (4.8)	0.02
Index medication use, n (%)			
ACEI	5,117 (76.7)	2,330 (34.9)	0.93
ARB	1,558 (23.3)	4,345 (65.1)	0.93
Hydrochlorothiazide	5,531 (82.9)	5,887 (88.2)	0.15
Chlorthalidone	73 (1.1)	0 (0.0)	0.15
Indapamide	1,071 (16.0)	788 (11.8)	0.12
Index medication dose category, n (%)			
Low	2,849 (42.7)	2,849 (42.7)	0
Medium	2,873 (43.0)	2,873 (43.0)	0
High	953 (14.3)	953 (14.3)	0
Other medications in prior 100 days, n (%)			
Noninsulin antihyperglycemic	673 (10.1)	698 (10.5)	0.01
Insulin	79 (1.2)	97 (1.5)	0.02
Statin	1,693 (25.4)	1,579 (23.7)	0.04

(Continued)

Table 1. (Continued)

Characteristic	Multipill	FDC	Standardized Difference
	(N = 6,675)	(N = 6,675)	
Warfarin	122 (1.8)	98 (1.5)	0.03
Direct oral anticoagulants	8 (0.1)	10 (0.1)	0.01
Digoxin	41 (0.6)	35 (0.5)	0.01
Clopidogrel	70 (1.0)	59 (0.9)	0.02

[†]Diagnosis occurred at any point in time.

^{††}Diagnosis occurred within 5 years of cohort entry. Index medication dose categorization is described in S1 Table.

Abbreviations: ACEI, angiotensin-converting enzyme inhibitor; AMI, acute myocardial infarction; ARB, angiotensin II-receptor blocker; FDC, single-pill fixed-dose combination; IQR, interquartile range; SD, standard deviation.

<https://doi.org/10.1371/journal.pmed.1002584.t001>

to receive an ARB (65.1% versus 23.3% in multipill group) than an ACEI and more likely to receive hydrochlorothiazide (88.2% versus 82.9% in multipill group) than other thiazides.

Medication use

The median time to the first instance of medication discontinuation was 191 days (IQR 45–741) in the FDC group and 150 days (IQR 45–446) in the multipill group ($P < 0.01$, Table 2). Medication discontinuation, defined as any break in therapy of greater than 150% of the previous days’ supply, occurred in 88.7% of individuals receiving multipill treatment and 83.1% in the FDC group (HR 0.80, 95% CI 0.77–0.83, $P < 0.01$). Individuals often resumed use of their index medications after a period of discontinuation (S4 Table). When examining use of the index antihypertensives over the entire study period, the proportion of days covered was 70% in the FDC group (IQR 19%–98%), which was significantly greater than 42% in the multipill group (IQR 11%–91%, $P < 0.01$). The proportion of days covered was similar in the sensitivity analysis using the less stringent definition of medication discontinuation (S5 Table).

Use of other cardiovascular risk-lowering medications was well-balanced between the FDC and multipill groups at baseline (Table 1) and in the last 90 days of follow-up (Table 3). Individuals in the FDC group were more likely to receive their index antihypertensive in the last 90 days of follow-up, but use of other antihypertensive medications was well-balanced (Table 3).

Clinical outcomes in the on-treatment analysis

In the on-treatment analysis following individuals until the first instance of medication discontinuation, there were no significant differences in primary or secondary outcomes between

Table 2. Medication use among individuals treated with multipill or FDC antihypertensive regimens.

Medication Use	Multipill	FDC
	N = 6,675	N = 6,675
Follow-up time, median (IQR), days	1,826 (1,142–1,826)	1,826 (1,163–1,826)
Time to first instance of discontinuation, median (IQR), days*	150 (45–446)	191 (45–741) [†]
Proportion of total days covered, median (IQR)	0.42 (0.11–0.91)	0.70 (0.19–0.98) [†]

*Indicates Kaplan-Meier estimate.

[†]Indicates $P < 0.01$ for between-group difference.

Time to first instance of discontinuation represents the first period of continuous medication use, defined as no disruption of greater than 150% of the previous days’ supply in receiving the index medications.

Abbreviations: FDC, single-pill fixed-dose combination; IQR, interquartile range.

<https://doi.org/10.1371/journal.pmed.1002584.t002>

Table 3. Use of antihypertensive and cardiovascular medications in the last 90 days of follow-up.

Medication Class	Multipill	FDC	Standardized Difference
	N = 6,675	N = 6,675	
	N (%)	N (%)	
Index antihypertensive	1,997 (29.9)	2,756 (41.3)	0.24
Other antihypertensive	1,925 (28.8)	1,908 (28.6)	0.01
Noninsulin antihyperglycemic	813 (12.2)	890 (13.3)	0.03
Insulin	149 (2.2)	170 (2.5)	0.02
Statin	2,209 (33.1)	2,148 (32.2)	0.02
Warfarin	231 (3.5)	201 (3.1)	0.02
Direct oral anticoagulants	86 (1.3)	85 (1.3)	0.00
Digoxin	NA*	NA*	<0.1*
Clopidogrel	215 (3.2)	201 (3.0)	0.01

*Data suppressed to comply with ICES privacy policies because calculations required the use of a cell involving 5 or fewer events.

This table reports the number and proportion of individuals who received each medication class in the last 90 days of follow-up. Standardized difference less than 0.1 was considered well-balanced.

Abbreviations: FDC, single-pill fixed-dose combination; ICES, Institute for Clinical Evaluative Sciences.

<https://doi.org/10.1371/journal.pmed.1002584.t003>

groups (Table 4). The composite primary outcome occurred at a rate of 2.4 events per 100 person-years in both the FDC and multipill groups (HR 1.06, 95% CI 0.86–1.31, $P = 0.60$).

There were also no significant differences between the groups with respect to hospitalizations for hypokalemia or hyponatremia, which occurred in fewer than 0.5% of cases (Table 2). There was no significant difference in the tracer outcome of cataract surgery between the groups (HR 0.99, 95% CI 0.85–1.14, $P = 0.83$).

In a sensitivity analysis using the less stringent definition of medication discontinuation, the results were similar, with the primary outcome occurring at a rate of 2.2 events and 2.1

Table 4. Clinical outcomes among individuals treated with multipill or FDC antihypertensive regimens, on-treatment analysis.

Outcome	Multipill	FDC	HR [†] (95% CI)	P value
	N = 6,675	N = 6,675		
	Event Rate* (Events/Years of Follow-up)	Event Rate* (Events/Years of Follow-up)		
Primary Outcome	2.4 (149/6,306)	2.4 (198/8,227)	1.06 (0.86–1.31)	0.60
Secondary Outcomes				
AMI	0.5 (34/6,322)	0.6 (46/8,258)	1.07 (0.69–1.68)	0.77
Heart failure	0.2 (11/6,330)	0.2 (19/8,261)	1.37 (0.66–2.99)	0.41
Stroke	0.4 (26/6,320)	0.5 (39/8,243)	1.26 (0.77–2.1)	0.37
Death	1.4 (86/6,333)	1.3 (108/8,267)	0.99 (0.75–1.32)	0.94
Safety Outcomes				
Hypokalemia	N/A ^{††}	N/A ^{††}	N/A ^{††}	N/A ^{††}
Hyponatremia	0.2 (11/6,332)	0.2 (14/8,265)	1.10 (0.50–2.49)	0.80
Tracer Outcome				
Cataract surgery	5.6 (331/5,946)	5.2 (397/7,663)	0.99 (0.85–1.14)	0.83

The primary outcome was a composite of death or hospitalization with AMI, heart failure, or stroke.

*Event rate per 100 person-years.

[†]HR was calculated with multipill group as the reference category.

^{††}Because there were fewer than 5 events, the data were suppressed to comply with ICES privacy policies, and a regression model was not fit.

Abbreviations: AMI, acute myocardial infarction; FDC, single-pill fixed-dose combination; HR, hazard ratio; ICES, Institute for Clinical Evaluative Sciences.

<https://doi.org/10.1371/journal.pmed.1002584.t004>

Table 5. Clinical outcomes among individuals treated with multipill or FDC antihypertensive regimens, primary intention-to-treat analysis.

Outcome	Multipill	FDC	HR [†] (95% CI)	P value
	N = 6,675	N = 6,675		
	Event Rate* (Events/Years of Follow-up)	Event Rate* (Events/Years of Follow-up)		
Primary Outcome	3.9 (1,008/25,967)	3.4 (904/26,226)	0.89 (0.81–0.97)	<0.01
Secondary Outcomes				
AMI	0.6 (158/26,376)	0.5 (142/26,569)	0.89 (0.71–1.12)	0.33
Heart failure	0.4 (97/26,526)	0.3 (91/26,605)	0.93 (0.70–1.24)	0.62
Stroke	0.5 (139/26,440)	0.6 (151/26,604)	1.08 (0.86–1.36)	0.51
Death	2.8 (755/26,699)	2.4 (646/26,854)	0.85 (0.77–0.94)	<0.01
Instance of drug discontinuation	93.4 (5,921/6,333)	67.0 (554/8,268)	0.80 (0.77–0.83)	<0.01
Safety Outcomes				
Hypokalemia	N/A ^{††}	N/A ^{††}	N/A ^{††}	N/A ^{††}
Hyponatremia	0.1 (35/26,626)	0.1 (30/26,790)	0.85 (0.52–1.39)	0.52
Tracer Outcome				
Cataract surgery	4.5 (1,072/24,027)	4.5 (1,089/24,118)	1.01 (0.93–1.10)	0.78

The primary outcome was a composite of death or hospitalization with AMI, heart failure, or stroke.

*Event rate per 100 person-years. An instance of drug discontinuation was defined as receiving no repeat medication within 150% of the previous days' supply of the index medications.

[†]HR was calculated with multipill group as the reference category.

^{††}Because there were fewer than 5 events, the data were suppressed to comply with ICES privacy policies, and a regression model was not fit.

Abbreviations: AMI, acute myocardial infarction; FDC, single-pill fixed-dose combination; HR, hazard ratio; ICES, Institute for Clinical Evaluative Sciences.

<https://doi.org/10.1371/journal.pmed.1002584.t005>

events per 100 person-years in the FDC group and the multipill group, respectively (HR 1.09, 95% CI 0.92–1.30, $P = 0.34$).

Clinical outcomes in the intention-to-treat analysis

The composite primary outcome in the intention-to-treat analysis occurred at a significantly lower rate in the FDC group than the multipill group (3.4 versus 3.9 events per 100 person-years; HR 0.89, 95% CI 0.81–0.97, $P < 0.01$; Table 5, Fig 2). In our analysis of secondary endpoints, the hazard of death was significantly lower among individuals in the FDC compared to the multipill group—2.4 versus 2.8 deaths per 100 person-years, respectively (HR 0.85, 95% CI 0.77–0.94, $P < 0.01$). No significant differences were observed in the hazards for other components of the composite endpoint (Table 5).

There were no significant differences between the groups with respect to hospitalizations for hypokalemia or hyponatremia, which occurred rarely (Table 5). There was no significant difference in the tracer outcome of cataract surgery between the groups (HR 1.01, 95% CI 0.93–1.10, $P = 0.78$).

Discussion

Among older adults initiating antihypertensive therapy, FDC treatment was associated with a significantly lower risk of composite clinical outcomes compared with multipill treatment, which may be related to better medication adherence. In the on-treatment analysis, outcomes were similar among adults who were actively receiving treatment. However, the intention-to-treat analysis revealed meaningful differences between groups with respect to adherence, with 70% of total days covered during follow-up in the FDC group compared with 42% in the multipill group. This was associated with a significantly lower risk of composite clinical outcomes

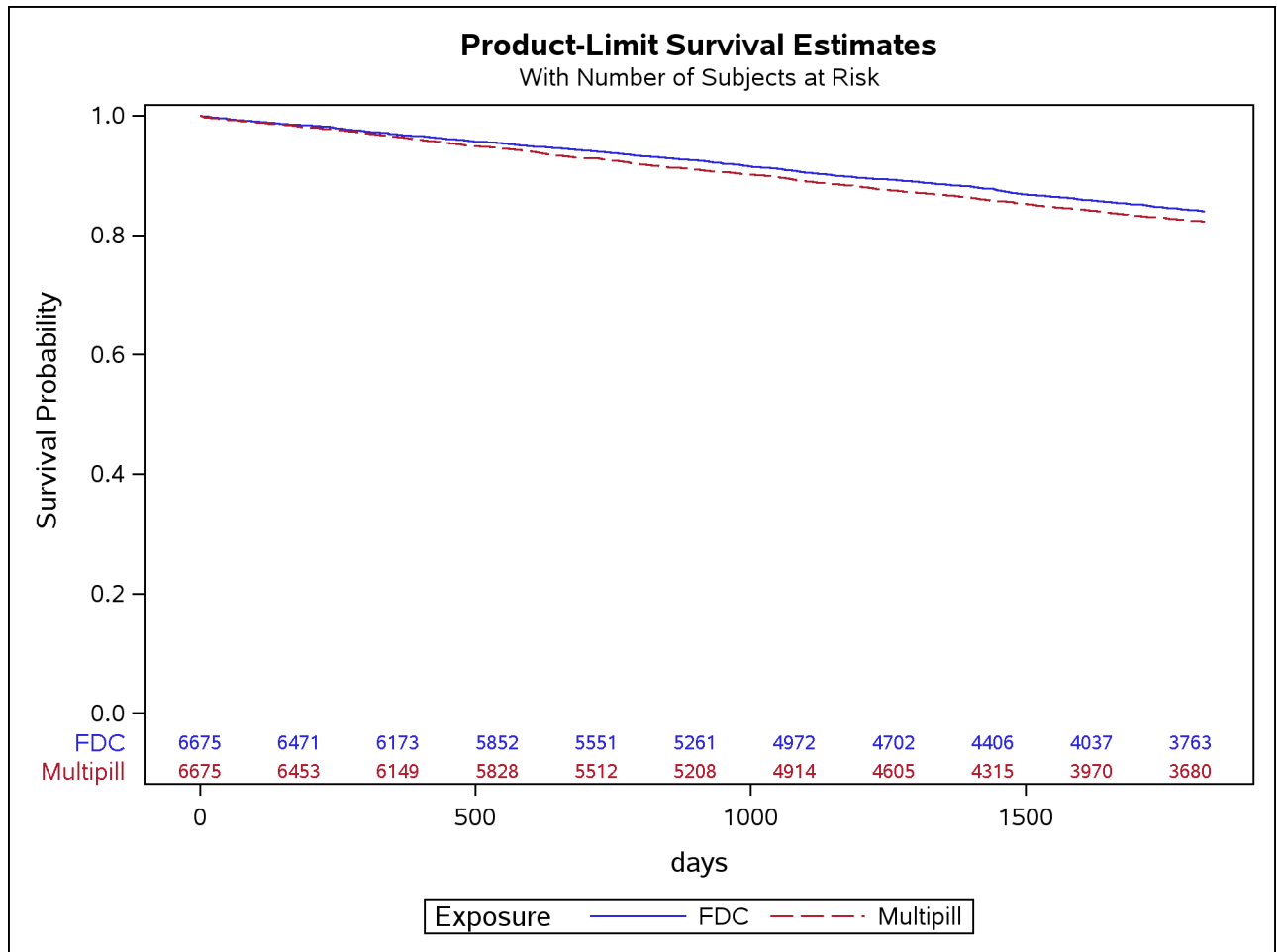


Fig 2. Survival estimates among individuals initiated on FDC versus multipill combination therapy. Legend: Kaplan-Meier estimates of survival probability. FDC, single-pill fixed-dose combination.

<https://doi.org/10.1371/journal.pmed.1002584.g002>

among the FDC treatment group, with an absolute difference of 0.5 fewer primary outcome events and 0.4 fewer deaths per 100 person-years of follow-up.

We used a novel application of a paired comparison between an intention-to-treat and on-treatment analysis to investigate whether adherence to treatment was related to better clinical outcomes. The on-treatment analysis only compared outcomes in patients up until the first instance of medication discontinuation and revealed that clinical outcomes were not significantly different when there were no medication adherence differences between groups. This contrasted with the intention-to-treat analysis, which followed patients despite disruptions in medication use and in which FDC therapy was associated with better medication adherence and subsequent clinical outcomes. This observation supports the hypothesis that improved medication adherence associated with FDC use confers important clinical benefits in a real-world setting.

The results of several sensitivity analyses support our findings. First, when the definition of medication discontinuation was made less stringent (permitting a 300% grace period instead of 150%), we observed similar results with respect to clinical outcomes and proportion of total days covered by index medications. This suggests that our findings are unlikely to have been biased by our chosen definition of medication discontinuation. Second, the use of other

cardiovascular risk-modifying medications was well-balanced between the FDC and multipill groups, both at baseline and in the last 90 days of follow-up. This suggests that our results are unlikely to have been biased by between-group differences in the use of other concomitant therapies.

Blood pressure control often requires multiple medications. More medications and more complex regimens reduce medication adherence [29], whereas simplifying regimens improves adherence [30]. FDC therapy offers an appealing solution by allowing more intensive treatment with a simpler regimen. Gupta and colleagues performed a systematic review and meta-analysis of randomized controlled trials and cohort studies of FDC compared with multipill therapy [10]. They identified 5 studies involving 17,999 individuals and found that FDC was associated with improved adherence (odds ratio 1.21, 95% CI 1.03–1.43). These findings are consistent with another meta-analysis, which included studies of hypertension and other conditions like HIV and which found that FDC therapy was associated with a 26% improvement in medication adherence [12]. A recent large United States study using claims data found that patients initiating FDC for hypertension were 13% more likely to be adherent to medications [8]. Our findings—that FDC use was associated with 28% more days covered with medications and an HR of 0.80 for the first instance of medication discontinuation (95% CI 0.77–0.83)—are very similar to the existing literature, which strengthens our confidence in the generalizability of our results. It is worth noting that provincial health insurance covered the cost of the medications in the study population, and therefore our findings may not be generalizable to settings in which there are substantial out-of-pocket cost differences between FDC and multipill regimens.

Randomized trials and observational studies have not rigorously examined whether FDC use is associated with better clinical outcomes. In an individual-patient data meta-analysis of 3 pragmatic trials, FDC therapy combining antihypertensive and other cardiovascular medications was associated with reduced blood pressure compared to usual care [31]. In their meta-analysis, Gupta and colleagues found a nonsignificant improvement in blood pressure control with FDC therapy [10], but clinical outcomes were not assessed in the included studies. Real-world evidence is contradictory and may reflect unmeasured bias [11,32]. One retrospective cohort study in the United Kingdom found fewer cardiovascular events among patients receiving FDC compared with multipill combinations (HR 0.74, 95% CI 0.70–0.77) [33]. However, this analysis was limited by substantial risk of confounding because matching was performed only on age, sex, and primary care practice, and groups remained unbalanced after matching. Moreover, groups were not matched based on medication class or other patient characteristics, and the analysis was not restricted to new users or to patients receiving only 2 antihypertensive medications at baseline, thus permitting important selection effects. Furthermore, adherence was not assessed in their study.

We addressed limitations in the existing literature and employed several methodologic approaches to improve the validity of comparisons in nonrandomized samples [34]. First, we used a similar active comparator. We compared FDC therapy with multipill combinations of the same medications, thus reducing potential confounding related to medication class. There was no compelling clinical reason to choose FDC or multipill combination therapy, and both are guideline-recommended, which suggests that there should be minimal indication bias. Second, we employed a new user study design and excluded patients with recent hospitalization or those who started combination medications on separate days to reduce the risk of selection bias. Third, we employed high-dimensional propensity score matching, which resulted in groups that were well-balanced on baseline characteristics. Finally, we used cataract surgery as a tracer outcome to assess for residual confounding and identified no difference between the

groups on this measure. Therefore, we were able to rigorously examine the association between FDC use and clinical outcomes in a real-world setting.

The only residual differences in our propensity score-matched groups were in the index medications. These were due to differences between the commonly used medications in FDC and multipill formulations. Individuals in the FDC group were more likely to receive an ARB (65.1% versus 23.3% in the multipill combination group). There was a smaller difference in the use of hydrochlorothiazide (88.2% in FDC group versus 82.9% in multipill combination group). This is unlikely to explain the difference in primary outcome. Although ARBs have been associated with a lower incidence of adverse effects than ACEI, the 2 medication classes have similar effectiveness for treating hypertension with no consistent differences in clinical outcomes [35,36]. Although there is controversy about the comparative effectiveness of chlorthalidone and hydrochlorothiazide [19], the differences in thiazide use between the 2 groups in our study were small and would have biased outcomes in favor of the multipill group where chlorthalidone use was more common. Adjusting for the differences in index medications would be inappropriate because this would adjust away the effect of the broader exposure group. Importantly, outcomes were not different between the groups in the on-treatment analysis, which suggests that the observed differences were attributable to adherence and not to the medications themselves.

There are several limitations to our study. First, the observed difference in the occurrence of the primary composite endpoint was driven by fewer deaths in the FDC group. We did not observe significantly fewer occurrences of the individual cardiovascular events. This may be explained by the relatively low event rate (individual cardiovascular events occurred in 1.4% to 2.4% of participants) and limitations in sample size. Our study included only new users over the age of 66 years, and treatment for hypertension is often initiated earlier in life. By excluding individuals with prior hospitalizations for cardiovascular events and by restricting our sample to new users of antihypertensive medications, we likely selected for a lower-risk population, which may have affected our ability to identify clinically important differences in individual cardiovascular endpoints or safety outcomes. Second, we were unable to identify cause of death and thus could not distinguish between cardiovascular death and all-cause mortality. Although this limits our ability to explain the observed reductions in mortality, our findings are consistent with previous literature demonstrating that hypertension control reduces both all-cause mortality and cardiovascular mortality [37]. Third, we did not have blood pressure measurements, which impaired our ability to adjust for the severity of baseline hypertension. We attempted to address this issue by matching based on initial medication dose, and the groups were well-balanced in this regard. Fourth, our measures of adherence were based on medication dispensing and assumed that medications were taken as prescribed. These are not direct measures of medication adherence but are considered acceptable measures in secondary analysis of clinical and administrative datasets [38]. Finally, despite employing multiple methodological approaches to address confounding, the possibility of residual confounding remains a limitation in this observational analysis.

Approximately 675 million people globally require combination antihypertensive therapy, and this number may grow as new guidelines call for more intensive blood pressure control [6]. Up to 40% of patients with hypertension in high-income countries are treated with multipill regimens [9], and this number may be higher in low- and middle-income countries [39]. Internationally, half of hypertension societies do not recommend FDC treatment [7]. Our study suggests that FDC formulations are associated with better medication adherence and clinical outcomes. Using FDC rather than multipill therapy represents a simple and potentially low-cost intervention that could substantially reduce the global burden of morbidity and mortality related to hypertension.

Supporting information

S1 Checklist. STROBE and RECORD checklist for “Fixed-dose combination antihypertensive medications, adherence, and clinical outcomes: A population-based retrospective cohort study”.

(DOCX)

S1 Table. Dose categorization for antihypertensive medication.

(DOCX)

S2 Table. Balance in the year of study enrollment in the FDC and multipill groups. FDC, single-pill fixed-dose combination.

(DOCX)

S3 Table. Baseline characteristics in study cohort before matching.

(DOCX)

S4 Table. Medications dispensed in the last 90 days of follow-up categorized by discontinuation status within each exposure group.

(DOCX)

S5 Table. Medication use among individuals treated with multipill or FDC antihypertensive regimens; medication discontinuation sensitivity analysis. FDC, single-pill fixed-dose combination.

(DOCX)

S6 Table. Administrative diagnostic codes for components of the primary outcome, based on the International Statistical Classification of Diseases and Related Health Problems, 9th revision and 10th revision, Canada (ICD-9 and ICD-10-CA).

(DOCX)

S1 Dataset Creation and Analysis Plan. Dataset creation and data analysis plan for “Fixed-dose combination antihypertensive medications, adherence, and clinical outcomes: A population-based retrospective cohort study,” edited for clarity.

(DOCX)

Acknowledgments

We thank Dr. Andreas Laupacis and Dr. Shaun Goodman for reviewing and commenting on an earlier version of this manuscript, and we thank Brogan Inc., Ottawa for use of their Drug Product and Therapeutic Class Database. Parts of this material are based on data and information compiled and provided by Cancer Care Ontario (CCO) and the Canadian Institute for Health Information (CIHI).

Disclaimer: The opinions, results, analyses, and conclusions reported in this paper are those of the authors, and not necessarily CCO or CIHI, and are independent from the funding sources. No endorsement by ICES, the SPOR Unit, or the Ontario MOHLTC is intended or should be inferred.

Author Contributions

Conceptualization: Amol A. Verma, Mina Tadrous, Tara Gomes, Muhammad M. Mamdani.

Formal analysis: Wayne Khuu.

Methodology: Amol A. Verma, Wayne Khuu, Mina Tadrous, Tara Gomes, Muhammad M. Mamdani.

Resources: Tara Gomes, Muhammad M. Mamdani.

Supervision: Muhammad M. Mamdani.

Writing – original draft: Amol A. Verma.

Writing – review & editing: Wayne Khuu, Mina Tadrous, Tara Gomes, Muhammad M. Mamdani.

References

1. Forouzanfar MH, Liu P, Roth GA, Ng M, Biryukov S, Marczak L, et al. Global Burden of Hypertension and Systolic Blood Pressure of at Least 110 to 115 mm Hg, 1990–2015. *JAMA*. 2017; 317: 165–182. <https://doi.org/10.1001/jama.2016.19043> PMID: 28097354
2. Forouzanfar MH, Afshin A, Alexander LT, Anderson HR, Bhutta ZA, Biryukov S, et al. Global, regional, and national comparative risk assessment of 79 behavioural, environmental and occupational, and metabolic risks or clusters of risks, 1990–2015: a systematic analysis for the Global Burden of Disease Study 2015. *Lancet*. 2016; 388: 1659–1724. [https://doi.org/10.1016/S0140-6736\(16\)31679-8](https://doi.org/10.1016/S0140-6736(16)31679-8) PMID: 27733284
3. Gradman AH, Basile JN, Carter BL, Bakris GL. Combination therapy in hypertension. *J Am Soc Hypertens*. Elsevier Ltd; 2010; 4: 90–98. <https://doi.org/10.1016/j.jash.2010.03.001> PMID: 20400053
4. James PA, Oparil S, Carter BL, Cushman WC, Dennison-Himmelfarb C, Handler J, et al. 2014 Evidence-Based Guideline for the Management of High Blood Pressure in Adults. *JAMA*. 2014; 311: 507. <https://doi.org/10.1001/jama.2013.284427> PMID: 24352797
5. Leung AA, Daskalopoulou SS, Dasgupta K, McBrien K, Butalia S, Zarnke KB, et al. Hypertension Canada's 2017 Guidelines for Diagnosis, Risk Assessment, Prevention, and Treatment of Hypertension in Adults. *Can J Cardiol*. Elsevier Inc.; 2017; 33: 557–576. <https://doi.org/10.1016/j.cjca.2017.03.005> PMID: 28449828
6. Whelton PK, Carey RM, Aronow WS, Casey DE, Collins KJ, Dennison Himmelfarb C, et al. 2017 ACC/AHA/AAPA/ABC/ACPM/AGS/APhA/ASH/ASPC/NMA/PCNA Guideline for the Prevention, Detection, Evaluation, and Management of High Blood Pressure in Adults. Hypertension. 2017; HYP.0000000000000065. <https://doi.org/10.1161/HYP.0000000000000065> PMID: 29133356
7. Chalmers J, Arima H, Harrap S, Touyz RM, Park JB. Global survey of current practice in management of hypertension as reported by societies affiliated with the International Society of Hypertension. *J Hypertens*. 2013; 31: 1043–8. <https://doi.org/10.1097/HJH.0b013e32835f7eef> PMID: 23429664
8. Lauffenburger JC, Landon JE, Fischer MA. Effect of Combination Therapy on Adherence Among US Patients Initiating Therapy for Hypertension: a Cohort Study. *J Gen Intern Med*. 2017; 32: 619–625. <https://doi.org/10.1007/s11606-016-3972-z> PMID: 28050754
9. Schaffer AL, Pearson S-A, Buckley NA. How does prescribing for antihypertensive products stack up against guideline recommendations? An Australian population-based study (2006–2014). *Br J Clin Pharmacol*. 2016; 82: 1134–1145. <https://doi.org/10.1111/bcp.13043> PMID: 27302475
10. Gupta AK, Arshad S, Poulter NR. Compliance, Safety, and Effectiveness of Fixed-Dose Combinations of Antihypertensive Agents: A Meta-Analysis. *Hypertension*. 2010; 55: 399–407. <https://doi.org/10.1161/HYPERTENSIONAHA.109.139816> PMID: 20026768
11. Buess D, Dieterle T, Leuppi JD, Zeller A, Martina B, Tschudi P, et al. [PP.39.07] Fixed-Dose Combinations of Antihypertensive Drugs May Not Improve Blood Pressure Control Compared to Free Drug Combinations—Findings from The Swiss Hypertension Cohort Study. *J Hypertens*. 2016; 34: e355. <https://doi.org/10.1097/01.hjh.0000492382.81220.e4>
12. Bangalore S, Kamalakkannan G, Parkar S, Messerli FH. Fixed-Dose Combinations Improve Medication Compliance: A Meta-Analysis. *Am J Med*. 2007; 120: 713–719. <https://doi.org/10.1016/j.amjmed.2006.08.033> PMID: 17679131
13. Baser O, Andrews LM, Wang L, Xie L. Comparison of real-world adherence, healthcare resource utilization and costs for newly initiated valsartan/amlodipine single-pill combination versus angiotensin receptor blocker/calcium channel blocker free-combination therapy. *J Med Econ*. 2011; 14: 576–583. <https://doi.org/10.3111/13696998.2011.596873> PMID: 21728914
14. Gautam CS, Saha L. Fixed dose drug combinations (FDCs): rational or irrational: a view point. *Br J Clin Pharmacol*. 2008; 65: 795–796. <https://doi.org/10.1111/j.1365-2125.2007.03089.x> PMID: 18294326

15. Angeli F, Reboldi G, Mazzotta G, Garofoli M, Ramundo E, Poltronieri C, et al. Fixed-Dose Combination Therapy in Hypertension. *High Blood Press Cardiovasc Prev*. 2012; 19: 51–54. <https://doi.org/10.2165/11632070-000000000-00000> PMID: 22867089
16. Andrade SE, Walker AM, Gottlieb LK, Hollenberg NK, Testa MA, Saperia GM, et al. Discontinuation of Anti-hyperlipidemic Drugs—Do Rates Reported in Clinical Trials Reflect Rates in Primary Care Settings? *N Engl J Med*. 1995; 332: 1125–1131. <https://doi.org/10.1056/NEJM199504273321703> PMID: 7700285
17. Levy AR, O'Brien BJ, Sellors C, Grootendorst P, Willison D. Coding accuracy of administrative drug claims in the Ontario Drug Benefit database. *Can J Clin Pharmacol*. 2003; 10: 67–71. PMID: 12879144
18. Park-Wyllie LY, Juurlink DN, Kopp A, Shah BR, Stukel TA, Stumpo C, et al. Outpatient gatifloxacin therapy and dysglycemia in older adults. *N Engl J Med*. 2006; 354: 1352–61. <https://doi.org/10.1056/NEJMoa055191> PMID: 16510739
19. Dhalla IA, Gomes T, Yao Z, Nagge J, Persaud N, Hellings C, et al. Chlorthalidone Versus Hydrochlorothiazide for the Treatment of Hypertension in Older Adults. *Ann Intern Med*. 2013; 158: 447. <https://doi.org/10.7326/0003-4819-158-6-201303190-00004> PMID: 23552325
20. Juurlink DN, Mamdani MM, Lee DS, Kopp A, Austin PC, Laupacis A, et al. Rates of Hyperkalemia after Publication of the Randomized Aldactone Evaluation Study. *N Engl J Med*. 2004; 351: 543–551. <https://doi.org/10.1056/NEJMoa040135> PMID: 15295047
21. PROGRESS Collaborative Group. Randomised trial of a perindopril-based blood-pressure-lowering regimen among 6105 individuals with previous stroke or transient ischaemic attack. *Lancet*. 2001; 358: 1033–1041. [https://doi.org/10.1016/S0140-6736\(01\)06178-5](https://doi.org/10.1016/S0140-6736(01)06178-5) PMID: 11589932
22. Schneeweiss S, Rassen JA, Glynn RJ, Avorn J, Mogun H, Brookhart MA. High-dimensional propensity score adjustment in studies of treatment effects using health care claims data. *Epidemiology*. 2009; 20: 512–22. <https://doi.org/10.1097/EDE.0b013e3181a663cc> PMID: 19487948
23. Division of Pharmacoepidemiology and Pharmacoeconomics. Pharmacoepidemiology Toolbox including High-dimensional Propensity Score (hd-PS) Adjustment version 2 [Internet]. Available from: <http://www.drugepi.org/dope-downloads/>. [cited 31 May 2017].
24. Weber MA, Schiffrin EL, White WB, Mann S, Lindholm LH, Kenerson JG, et al. Clinical Practice Guidelines for the Management of Hypertension in the Community. *J Clin Hypertens*. 2014; 16: 14–26. <https://doi.org/10.1111/jch.12237> PMID: 24341872
25. Juurlink D, Preyra C, Croxford R, Chong A, Austin P, Tu J, et al. Canadian Institute for Health Information Discharge Abstract Database: A validation study. *Inst Clin Eval Sci*. 2006; 1–77.
26. Porter J, Mondor L, Kapral MK, Fang J, Hall RE. How Reliable Are Administrative Data for Capturing Stroke Patients and Their Care. *Cerebrovasc Dis Extra*. 2016; 6: 96–106. <https://doi.org/10.1159/000449288> PMID: 27750249
27. Nguyen T-L, Collins GS, Spence J, Daurès J-P, Devereaux PJ, Landais P, et al. Double-adjustment in propensity score matching analysis: choosing a threshold for considering residual imbalance. *BMC Med Res Methodol*. 2017; 17: 78. <https://doi.org/10.1186/s12874-017-0338-0> PMID: 28454568
28. Austin PC. The use of propensity score methods with survival or time-to-event outcomes: Reporting measures of effect similar to those used in randomized experiments. *Stat Med*. 2014; 33: 1242–1258. <https://doi.org/10.1002/sim.5984> PMID: 24122911
29. Osterberg L, Blaschke T. Adherence to Medication. *N Engl J Med*. 2005; 353: 487–497. <https://doi.org/10.1056/NEJMra050100> PMID: 16079372
30. Schroeder K, Fahey T, Ebrahim S. How Can We Improve Adherence to Blood Pressure-Lowering Medication in Ambulatory Care? *Arch Intern Med*. 2004; 164: 722. <https://doi.org/10.1001/archinte.164.7.722> PMID: 15078641
31. Webster R, Patel A, Selak V, Billot L, Bots ML, Brown A, et al. Effectiveness of fixed dose combination medication ('polypills') compared with usual care in patients with cardiovascular disease or at high risk: A prospective, individual patient data meta-analysis of 3140 patients in six countries. *Int J Cardiol. Netherlands*; 2016; 205: 147–156. <https://doi.org/10.1016/j.ijcard.2015.12.015> PMID: 26736090
32. Mazza A, Lenti S, Schiavon L, Sacco AP, Dell'Avvocata F, Rigatelli G, et al. Fixed-Dose Triple Combination of Antihypertensive Drugs Improves Blood Pressure Control: From Clinical Trials to Clinical Practice. *Adv Ther*. 2017; 34: 975–985. <https://doi.org/10.1007/s12325-017-0511-1> PMID: 28299716
33. Belsey JD. Optimizing adherence in hypertension: a comparison of outcomes and costs using single tablet regimens vs individual component regimens. *J Med Econ*. 2012; 15: 897–905. <https://doi.org/10.3111/13696998.2012.689792> PMID: 22548677
34. Fralick M, Kesselheim AS, Avorn J, Schneeweiss S. Use of Health Care Databases to Support Supplemental Indications of Approved Medications. *JAMA Intern Med*. 2017; 2120: 1–9. <https://doi.org/10.1001/jamainternmed.2017.3919>

35. Matchar DB, McCrory DC, Orlando LA, Patel MR, Patel UD, Patwardhan MB, et al. Systematic review: comparative effectiveness of angiotensin-converting enzyme inhibitors and angiotensin II receptor blockers for treating essential hypertension. *Ann Intern Med.* 2008; 148: 16–29. Available from: <http://www.ncbi.nlm.nih.gov/pubmed/17984484>. PMID: 17984484
36. Li EC, Heran BS, Wright JM. Angiotensin converting enzyme (ACE) inhibitors versus angiotensin receptor blockers for primary hypertension. *Cochrane Database Syst Rev.* 2014; CD009096. <https://doi.org/10.1002/14651858.CD009096.pub2> PMID: 25148386
37. Van Vark LC, Bertrand M, Akkerhuis KM, Brugts JJ, Fox K, Mourad JJ, et al. Angiotensin-converting enzyme inhibitors reduce mortality in hypertension: A meta-analysis of randomized clinical trials of renin-angiotensin-aldosterone system inhibitors involving 158 998 patients. *Eur Heart J.* 2012; 33: 2088–2097. <https://doi.org/10.1093/eurheartj/ehs075> PMID: 22511654
38. Lam WY, Fresco P. Medication Adherence Measures: An Overview. *Biomed Res Int.* Hindawi Publishing Corporation; 2015;2015. <https://doi.org/10.1155/2015/217047> PMID: 26539470
39. Chow CK, Teo KK, Rangarajan S, Islam S, Gupta R, Avezum A, et al. Prevalence, awareness, treatment, and control of hypertension in rural and urban communities in high-, middle-, and low-income countries. *JAMA.* 2013; 310: 959–68. <https://doi.org/10.1001/jama.2013.184182> PMID: 24002282