Radiology: Cardiothoracic Imaging

Suboptimal Quality and High Risk of Bias in Diagnostic Test Accuracy Studies at Chest Radiography and CT in the Acute Setting of the COVID-19 Pandemic: A Systematic

Review

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Conflicts of interest are listed at the end of this article.

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Purpose: To synthesize the literature on diagnostic test accuracy of chest radiography, CT, and US for the diagnosis of coronavirus disease 2019 (COVID-19) in patients suspected of having COVID-19 in a hospital setting and evaluate the extent of suboptimal reporting and risk of bias.

Materials and Methods: A systematic search was performed (April 26, 2020) in EMBASE, PubMed, and Cochrane to identify chest radiographic, CT, or US studies in adult patients suspected of having COVID-19, using reverse-transcription polymerase chain reaction test or clinical consensus as the standard of reference. Two \times two contingency tables were reconstructed, and test sensitivity, specificity, positive predictive values, and negative predictive values were recalculated. Reporting quality was evaluated by adherence to the Standards for Reporting of Diagnostic Accuracy Studies (STARD), and risk of bias was evaluated by adherence to the Quality Assessment of Diagnostic Accuracy Studies-2 (QUADAS-2).

Results: Thirteen studies were eligible (CT = 12; chest radiography = 1; US = 0). Recalculated CT sensitivity and specificity ranged between 0.57 and 0.97, and 0.37 and 0.94, respectively, and positive predictive values and negative predictive values ranged between 0.59 and 0.92 and 0.57 and 0.96, respectively. On average, studies complied with only 35% of the STARD-guideline items. No study scored low risk of bias for all QUADAS-2 domains (patient selection, index test, reference test, and flow and timing). High risk of bias in more than one domain was scored in 10 of 13 studies (77%).

Condusion: Reported CT test accuracy for COVID-19 diagnosis varies substantially. The validity and generalizability of these findings is complicated by poor adherence to reporting guidelines and high risk of bias, which are most likely due to the need for urgent publication of findings in the first months of the COVID-19 pandemic.

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By May 11, 2020, there were 4006257 confirmed coronavirus disease 2019 (COVID-19) cases and 278892 deaths globally, of which there were 88891 cases and 4531 deaths in the last 24 hours, respectively (1). For efficient triage of patients suspected of having COVID-19, rapid diagnosis is desirable. Especially in the first months of the pandemic, when prompt clinical action was required, theoretical knowledge and clinical experience was limited. In addition, at the time of this writing, due to the shortage of prospective studies and trials, evidence-based guidelines for the management of patients with (suspected) COVID-19 are lacking. This has, among others, resulted in divergent opinions and recommendations for COVID-19 diagnosis, the latter mainly based on expert opinions and early publications (2–5).

Real-time reverse-transcription polymerase chain reaction (RT-PCR) is generally accepted as the reference test for COVID-19 diagnosis (6). Owing to its modest sensitivity (7,8), limited availability, and relative time-consuming analysis, complementary and/or replacement tests have been proposed, in particular CT imaging.

The value of CT as a screening instrument to rule out COVID-19 infection, or as a diagnostic tool for the confirmation of COVID-19, is reflected by its test characteristics such as sensitivity, specificity, and predictive values. Rather than being fixed values, these parameters strongly depend on patient characteristics (selection of patients and disease prevalence or pretest probability), imaging technique, and characteristics of the doctors interpreting these images (eg, clinical experience and subjective thresholds for decision making). Reported diagnostic accuracy may therefore vary substantially between studies and is prone to selected reporting, affecting the generalizability of published results. Furthermore, several types of bias such as incorporation bias and verification bias may be introduced which may result in the overestimation of diagnostic performance.

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Abbreviations

COVID-19 = coronavirus disease 2019, DTA = diagnostic test accuracy, GRADE = grading of recommendations, assessment, development and evaluations, NPV = negative predictive value, PCR = polymerase chain reaction, PPV = positive predictive value, QUADAS = Quality Assessment of Diagnostic Accuracy Studies, RT-PCR = reverse-transcription polymerase chain reaction, STARD = Standards for Reporting of Diagnostic Accuracy Studies

Summary

Diagnostic test accuracy imaging studies published in the first months of the COVID-19 pandemic show substantial variation in reported diagnostic accuracy and poor adherence to reporting guidelines and contain substantial risk of bias.

Key Points

- Recalculated CT sensitivity and specificity ranges in suspected COVID-19 in diagnostic test accuracy studies were 0.57–0.97 and 0.37–0.94, respectively. For positive and negative predictive values, these ranges were 0.59–0.92 and 0.57–0.96, respectively. The calculated disease prevalence range was 0.29–0.85.
- Adherence to reporting guidelines (STARD) was low, with on average only 35% (12/34) of items reported.
- High risk of bias was identified in 10 of 13 studies (77%), and high applicability concerns were found in eight of 13 studies (62%) in more than one QUADAS-2 domain, generating limited information for generalizability to clinical practice.

Since the outbreak of the virus, over 11000 articles on CO-VID-19 have been published (9). Authors and journals need to be commended for their efforts of generating scientific evidence in the midst of a pandemic and making these results available, often open access, in an expedited fashion. However, the time pressure and limited time for peer review may affect the quality of the published studies and increase the risk of bias and incomplete reporting.

The purpose of this systematic review was twofold: (*a*) to systematically search and synthesize the literature on diagnostic test accuracy of chest radiography, CT, and US in patients suspected of having COVID-19 in a hospital setting and (*b*) to evaluate the quality of reporting and risk of bias in studies reporting on diagnostic imaging tests in the acute setting of the COVID-19 pandemic.

Materials and Methods

The protocol of this systematic review and meta-analysis was prospectively published and registered online (PROSPERO, registration number CRD42020177432). This study was conducted according to the preferred reporting items for systematic reviews and meta-analyses diagnostic test accuracy guide-lines (10–12).

Eligibility Criteria

We included articles meeting the following criteria: (*a*) adults with suspected COVID-19 pneumonia presenting in a hospital setting, including emergency departments; (*b*) patients undergoing chest imaging including US, chest radiography, and/ or CT for diagnosis of COVID-19 infection; (*c*) COVID-19 diagnosis confirmed or ruled out by reference test (ie, RT-PCR

or clinical consensus). We included articles reporting on diagnostic test accuracy measures including sensitivity, specificity, positive predictive value (PPV) and negative predictive value (NPV), and/or area under the receiver operating characteristic curve analysis. Parameters beyond test accuracy were beyond the scope of this study. A second aim of this study was to assess risk of bias and quality and completeness of reporting and their potential implications for patient care and treatment decisions.

Search

The online libraries of PubMed, EMBASE, and Cochrane were systematically searched on March 30 and updated on April 26, 2020, using synonyms for COVID-19, chest radiography, CT, US, and imaging (Appendix E1 [supplement]). No limitations were applied to the search strategy. The online version of the journal *Radiology: Cardiothoracic Imaging* was separately searched for relevant studies, as this journal is not yet indexed by MEDLINE (PubMed) or EMBASE. Preprint articles were not included.

Study Selection

Titles and abstracts were screened based on predefined criteria by two reviewers (D.S. and R.W.v.H.) independently (Fig 1). Duplicate articles were excluded manually. Discordant judgments were resolved in a consensus meeting. Full-text screening for inclusion in the systematic review was performed by two reviewers (D.S. and R.W.v.H.) independently. Discrepancies were resolved by a third reviewer (A.F.v.d.H.). We excluded studies (a) including only patients with confirmed COVID-19 diagnosis, (b) including only children, (c) focusing only on artificial intelligence algorithms, (d) focusing on animals, (e) in a non-English language, (f) that did not allow for reconstruction of a (partial) 2×2 contingency table, and (g) that were case reports (n < 10), reviews, conference proceedings, and letters. Cross-referencing was performed. Considering the urgent nature, corresponding authors were not contacted to retrieve missing (outcome) data. A list of excluded studies is presented in Appendix E2 (supplement).

Data Extraction and Critical Appraisal

We extracted data on study design, study subject identification (number of subjects identified, number excluded, and final number included), participant demographics, symptoms, laboratory findings, and imaging features. Detailed information on index test and reference test protocols including definitions and the threshold for a positive test, time-interval analysis methods, and test results was extracted. For test results, the number of positive and negative index tests and reference tests as well as the true-positive (index test and reference test positive), false-positive (index test and reference test negative), true-negative (index test and reference test negative), and false-negative (index test negative and reference test positive) counts were extracted.

For each study, data were extracted by two researchers individually and cross-checked by a third researcher (D.S., R.W.v.H., and A.F.v.d.H.), except for results on index test and reference

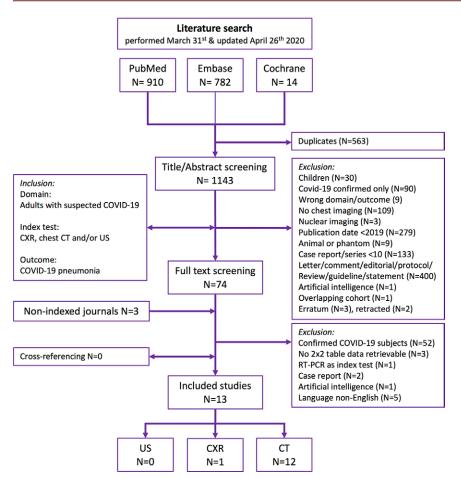


Figure 1: A flowchart of the systematic search results in the PubMed, EMBASE, and Cochrane databases with predefined selection criteria. Radiology: Cardiothoracic Imaging was screened for eligible articles as this novel journal is not yet indexed by MEDLINE. CXR = chest radiography.

test, which were obtained by three researchers individually. Discordances were resolved by consensus.

For each study, we assessed whether the author reported the purpose of imaging (screening, risk assessment, diagnosis, prognosis, staging, monitoring, or surveillance) and role of the imaging test (replacement, triage, add-on, parallel, or combined testing) (13).

Risk of bias and applicability were evaluated using the Quality Assessment of Diagnostic Accuracy Studies (QUADAS)-2 tool (36) based on 14 signaling questions including four riskof-bias domains (patient selection, index test, reference test, and flow and timing) and three applicability concern domains (patient selection, index test, and reference test) (Appendix E3 [supplement]). Applicability concerns the degree to which included patients and study setting (domain patient selection), the type of index test used, its conduct and interpretation (domain index test), and the target condition as defined by the standard of reference (domain reference standard) match the review question. Blinding for the index test result was considered not relevant in the assessment of reference test (RT-PCR) risk of bias, because this quantitative semiautomated method is unlikely to be affected. Risk of bias regarding applicability with regard to the target condition (COVID-19) was by default scored as low concern with RT-PCR as the reference test. Types of potential bias were assessed according to classifications as previously published by Whiting et al including bias on verification, incorporation, imperfect reference standard, spectrum, review, disease progression, and treatment paradox (14,15).

Reporting quality was rated according to the Standards for Reporting of Diagnostic Accuracy Studies (STARD) 2015 statement checklist (16). We evaluated all items on the checklist with the exception of (a) adverse events from performing the index test or reference standard, because this risk is considered negligible; (b) registration number of studies with a retrospective design, because we sympathize with the lack of retrospective study registration in this pandemic; (c) handling of missing data on the index test and reference standard for retrospective studies; and (d) rationale for choosing the reference standard when RT-PCR was used, because a polymerase chain reaction (PCR) is the unquestioned reference standard for the detection of a viral pneumonia. For those four items, we added "not applicable" to the scoring system. All QUADAS-2 and STARD items (Appendix E3 [supplement]) were rated by two individual readers and cross-checked by a third reader in case of discordance (D.S., R.W.v.H., and A.F.v.d.H.). For each study, overall risk of bias and applicability were evaluated according to the grading of

recommendations, assessment, development, and evaluations (GRADE)-21 framework (13,17). On the basis of the GRADE framework, certainty of evidence for the totality of the diagnostic test accuracy studies was summarized based on concerns regarding study design, risk of bias, indirectness and applicability, imprecision in diagnostic accuracy measure estimates (wide confidence intervals), inconsistency (large differences in estimates), and publication bias (13,17,18).

Data Analysis

Studies were categorized as "diagnostic test accuracy" (DTA) studies if measures of diagnostic accuracy were reported (at least test sensitivity or specificity); studies not reporting diagnostic accuracy measures were categorized as "non-DTA." Diagnostic test results were presented (*a*) as reported by authors and (*b*) as recalculated based on 2×2 frequency statistics as retrieved by our raters. The following diagnostic accuracy measures were recalculated based on 2×2 tables: sensitivity, specificity, PPV, NPV, accuracy, pretest probability, and positive and negative posttest probability (19). As an additional analysis, a predefined test interval of 3 days (ie, RT-PCR ≤ 3 days after chest imaging) was set as the cutoff for an appropriate time interval between the index test and reference test to limit the probability of interim infection or cross-transmission

Table 1: Cha	racteristics	of Studies Prov	viding Chest Ima	Table 1: Characteristics of Studies Providing Chest Imaging Accuracy Results in Patients with Suspected COVID-19 Pneumonia	sults in Patien	ts with Sus	pected CO	VID-19 Pner	umonia				
	Journal Im-	Journal Im- Date of Publi-			Inclusion of			Start Date	End Date	Uni- or Multi- Identified	Identified	Included	
Author	pact Factor* cation	* cation	Study Design	Data Collection	COVID-19	Setting	Country	Inclusion	Inclusion	center	Subjects	Subjects	
Ai et al (20)	7.608	26/02/20	Case series/cohc	Case series/cohort Retrospective	Suspected	Hospital	China	06/01/20	06/02/20	Unicenter	1049	1014	
Bai et al (21)	7.608	10/03/20	Case-control	Retrospective	Confirmed	Hospital	China	06/01/20	20/02/20	Multicenter	10 528	424	
Zhu et al (31) 2.049) 2.049	13/03/20	Case series/cohc	Case series/cohort Retrospective	Suspected	ED	China	24/01/20	20/02/20	Unicenter	NR	116	
Cheng et al (24)	3.161	14/03/20	Case series/coho	Case series/cohort Retrospective	Suspected	ED/FD	China	19/01/20	06/02/20	Unicenter	2445	38	
Long et al (27) 2.948	7) 2.948	25/03/20	Case series/cohc	Case series/cohort Retrospective	Suspected	Hospital	China	20/01/20	08/02/20	Unicenter	204	87	
Himoto et al 1.500 (26)	1.500	30/03/20	Case series/coho	Case series/cohort Retrospective	Suspected	Hospital	Japan	14/02/20	01/03/20	Unicenter	32	21	
Caruso et al (22)	7.608	03/04/20	Case series/cohort Prospective	ort Prospective	Suspected [†]	ED	Italy	04/03/20	19/03/20	Unicenter	158	158	
Choi et al (25) NA	5) NA	06/04/20	Case-control	Retrospective	Confirmed	Hospital	China and South Korea	l NR	NR	Multicenter	37	37	
Wen et al (29)) NA	06/04/20	Case series/cohc	Case series/cohort Retrospective	Suspected	Hospital	China	21/01/20	14/02/20	Multicenter	NR	103	
Yang et al (30) 5.099	() 5.099	12/04/20	Case series/cohc	Case series/cohort Retrospective	Suspected	Hospital	China	20/01/20	05/03/20	Unicenter	75	55	
Chen et al (23)	3.962	16/04/20	Case-control	Retrospective	Confirmed	Hospital	China	01/01/20	08/02/20	Multicenter	NR	136	
Miao et al (28)	1.651	19/04/20	Case series/coho	Case series/cohort Retrospective	Suspected	Hospital	China	12/01/20	13/02/20	Multicenter	166	130	
Dangis et al (32)	NA	21/04/20	Case series/cohc	Case series/cohort Retrospective	Suspected	ED	Belgium	14/03/20	24/03/20	Unicenter	192	192	
Note.—ED = * 2018 InCite † Previously pc	emergency d s Journal Citz ssitive test res	Note.—ED = emergency department, FD = fever departme * 2018 InCites Journal Citation Reports (Clarivate Analytic † Previously positive test result was part of inclusion criteria.	Note.—ED = emergency department, FD = fever department. * 2018 InCites Journal Citation Reports (Clarivate Analytics). † Previously positive test result was part of inclusion criteria.	Note.—ED = emergency department, FD = fever department, NA = not available, NR = not reported. * 2018 InCites Journal Citation Reports (Clarivate Analytics). † Previously positive test result was part of inclusion criteria.	NR = not repo	rted.							

between admitted patients. Depending on data availability, test accuracy results were recalculated restricted to this time interval. In the case that a study would only contain a subgroup of study participants relevant to the review, 2×2 data would be extracted for these patients only. QUADAS-2 results (counts for low risk, high risk, or unclear) were presented as overall, per domain, and per study. Adherence to STARD was analyzed by the number and percentage of STARD items reported and summarized by calculating the proportion of reported items to the total number of applicable STARD items. STARD in nature is designed for DTA studies, but non-DTA studies providing test results were also included in this review. Therefore, QUADAS-2 and STARD subgroup analyses were performed for DTA studies.

No formal analysis on small-study effects was performed. Categorical data were presented as number (frequency), and continuous data were presented as mean \pm standard deviation or median (interquartile range) based on data distribution. Analysis was done using IBM SPSS Statistics version 25 and Microsoft Excel for Mac version 16. Statistical significance was set at a level of P < .05 (assuming two-tailed tests).

Results

Search and Inclusion

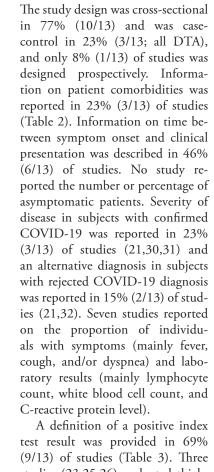
Our search yielded a total of 1706 articles (Fig 1). Thirteen studies on patients with suspected COVID-19 infection and available diagnostic accuracy data on chest CT and/or chest radiography performance as the index test and RT-PCR or clinical consensus as the reference test were included (8,20–32). No studies were found on chest US performance for COVID-19 diagnosis. No overlapping study populations were identified for the included studies.

Studies and Subjects

Ten studies (20–23,25–29,32) were categorized as DTA and three

Study					Demog	Demographics		Expo	Exposure His- tory*		Syı	Symptoms				Lab. Findings	ß
Author	Data†	Included Subjects	z	Age (y)	Male	BMI	Comorb.	Area (%)	Person (%)	Int. (d) Since Onset	Fever	Cough Dys.	Dys.	Other	Lymph Count ↓	CRP↑	Other
Ai et al (20)	Total	COVID-19 suspect	1014	51 ± 15	46%	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR
Bai et al (21)	Cases	COVID-19 confirmed and abnormal CT	l 219	45 ± 15	54%	NR	Rep.	36	Ŋ	4.9	65%	NR	NR	NR	84%	NR	WBC
	Controls	Viral pneumonia and abnormal CT	205	65 ± 19	50%	NR	Rep.	NA	NA	NR	56%	NR	NR	NR	56%	NR	WBC
Zhu et al (31)	Subgroup	COVID-19 confirmed	1 32	40 (27–53)) 46%	24.7 ± 3.2	Rep.	13	NR	5 (4-7)	84%	66%	NR	Rep.	59%	66%	WBC
		Non-COVID-19	84	35 (27–53)) 46%	22.9 + 3.2	Rep.	×	NR	4.0 (1-9)) 68%	62%	NR	Rep.	29%	48%	D-di- mer
Cheng et al (24)	Subgroup	COVID-19 confirmed and abnormal CT	1 11	50 ± 16	73%	NR	NR	NR	73	4.4	73%	64%	9%6	Rep.	NR	NR	NR
		Non-COVID-19 and abnormal CT	22	44 ± 16	32%	NR	NR	NR	32	5.5	77%	86%	18%	Rep.	NR	NR	NR
Long et al (27)	Subgroup	COVID-19 confirmed and abnormal CT	1 36	45 ± 18	56%	NR	NR	92*	NR	$2.6 \pm 1.7 \ 100\%$	7 100%	NR	NR	NR	64%	NR	WBC
		Non-COVID-19 and abnormal CT	51	47 ± 19	51%	NR	NR	57*	NR	3.2 ± 1.6	6 100%	NR	NR	NR	24%	NR	WBC
Himoto et al (26)	Subgroup	COVID-19 con- firmed, CT > 3 days after symptom onset	9	59 (45–81)	-81) 83%	NR	NR	NR	50	NR	NR	NR	NR	NR	NR	NR	NR
(-	Non-COVID-19 [‡]	15	66 (28–87)		NR	NR	0	0	NR	NR	NR	NR	NR	NR	NR	NR
Caruso et Total al (22)	: Total	COVID-19 suspect	158	57 ± 17	52%	NR	NR	NR	NR	NR	61%	56%	33%	NR	60%	88%	LDH
Choi et al (25)	Cases	COVID-19 confirmed and abnormal CT	I 17	45 ± 17	59%	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR
	Controls	Non–COVID-19 and negative chest radiograph	20	32 ± 14	45%	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR
Wen et al (29)	Total	COVID-19 suspect	103	46.0 ± 15	47%	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR
Yang et al (30)	Yang et al Subgroup (30)	COVID-19 con- firmed, pregnant, Chinese	13	30 ± 2	%0	20.4 ± 2.3	NR	NR	×	NR	15% of 65%"	15%	0	Rep.	mean	mean	WBC
															5	/T	

Table 2 (c	ontinued):	Table 2 (continued): Specification of Studied Subjects, Demographics, COVID-19 Exposure History, Patient Symptoms and Laboratory Findings	ied Sub	jects, Demo	graphic	s, covi	lD-19 Exp(osure H	istory, P	atient Sym	ptoms a	nd Labo	ratory F	inding	8		
Study					Demo	Demographics		Expc	Exposure His- tory*		Syı	Symptoms			La	Lab. Findings	sc
Author	Data†	Included Subjects	Z	Age (y)	Male	BMI	Comorb.	Area (%)	Person (%)	Int. (d) Since Onset	Fever	Cough Dys.		Other 0	Lymph Count ↓	CRP↑	Other
		Non-COVID-19 pregnant, Chinese	42	30 ± 3	%0	21.4 ± 2.8	NR	NR	0	NR	26% of 48%"	0	0	Rep.	mean	mean	WBC
Chen et al (23)	Cases	COVID-19 confirmed 70 and abnormal CT	02 P	43 ± 13	59%	NR	NR	NR	NR	NR	mean	69%	NR	Rep.	46%	mean	WBC
	Controls	Pneumonia and abnor- 66 mal CT	r- 66	47 ± 25	65%	NR	NR	NR	NR	NR	mean	85%	NR	Rep.	36%	mean	WBC
Miao et al (28)	Subgroup	COVID-19 confirmed 54 and abnormal CT	d 54	45 ± 13	52%	NR	Rep.	72	46	4.0 (1–6) NR	NR	NR	NR	NR	NR	NR	NR
		Non–COVID-19 and 76 abnormal CT	1 76	42 ± 14	65%	NR	Rep.	39	36	3.0 (1–5) NR	NR	NR	NR	NR	NR	NR	NR
Dangis et al (32)	Dangis et Subgroup al (32)	COVID-19 confirmed 83	d 83	67 ± 17	51%	NR	NR	NR	NR	7 (3–9)	%69	74%	53%	NR	NR	NR	NR
		Non-COVID-19	109	58 ± 18	41%	NR	NR	NR	NR	7 (3–9)	46%	67%	41%	NR	NR	NR	NR
Note.—Data are = dyspnea CRP = blood cell count.	ta are preser CRP = C-rea ount.	Note.—Data are presented as mean ± standard deviation, median (interquartile range), or median (range) unless otherwise specified. BMI = body mass index, Comorb. = comorbidities, Dys. = dyspnea CRP = C-reactive protein; Int. = interval, Lab. = laboratory, LDH = lactate dehydrogenase, Lymph count = lymphocyte count, NR = not reported, Rep. = reported, WBC = white blood cell count.	l deviatic :rval, Lał	n, median (in . = laboratory	terquart , LDH =	ile range = lactate	:), or media dehydroger	n (range) 1ase, Lyn) unless of 1ph count	therwise spec = lymphocy	cified. BM ⁄te count,	II = body NR = no	mass ind t reporte	lex, Con d, Rep.	norb. = co = reported	morbiditi I, WBC =	es, Dys. white
* Trip to cc † Per study ‡ Non-CO ^N	ntaminated data are pres /ID-19 base	* Trip to contaminated area and/or contact with COVID-19 suspect. † Per study data are presented stratified for cases and controls, or for patients with confirmed COVID-19 and without COVID-19. * Non-COVID-19 based on negative reverse transcription polymerase chain reaction test and/or clinical observation.	h COVII s and con unscriptic	D-19 suspect. Itrols, or for pain polymerase	atients w chain re	vith conf action to	îrmed CO ^v est and/or c	VID-19 ; linical ol	and witho bservation	ut COVID-	19.						



(24,30,31) as non-DTA (Table 1).

test result was provided in 69% (9/13) of studies (Table 3). Three studies (23,25,26) evaluated thickslice (3–5 mm) CT in that subset of patients and two (30,31) did not report slice thickness. Five studies dichotomized CT index test results into positive and negative, and a threshold was reported in 40% (4/10) of DTA studies (25,26,28,32). One study combined baseline and follow-up test results within subjects in the same primary test accuracy analysis (25).

A double PCR swab (nasopharynx and oropharynx) was taken in 8% (1/13) of studies; 46% (6/13) of studies took a single swab; and 46% (6/13) did not report on the sampling method for RT-PCR (Table 4). Repeated RT-PCR sampling was performed in 48% (6/13) of studies (20,21,27–30) and was not reported in 31% (4/13) of studies (23–26). Time from symptom onset to test and time between the index and reference test was reported in two and four studies, respectively.

Prenatal and postpartum.

Billateral malinies GGO Consol. Dili Halo ening Patring Lesions ai. sion Dist 90% 12% 46% 50% NR NR NR NR NR NR NR NR 75% Na 91% 60% 59% 31% 56% 5% NR NR NR NR 91% 0% 47% 80% 13% 7% NR 7% 8% NR 9% NR NR NR 10% 8% 10% 7% 8% NR 10% 7% 8% 10% 9% 10% <td< th=""></td<>
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100% 55% NR NR 82% NR 27% 0% 0% 86% 17% NR NR NR NR NR 3% 6% 86% 17% NR NR NR NR NR 3% 6% 100% 33% NR 0% 13% NR
86% 17% NR NR NR 3% 6% 100% 33% NR 0% NR 0% 33% 0% 0% 100% 33% NR 0% NR NR 0% 33% 0% 0% % 100% 72% 89% 12% NR NR 0% 33% 0% 0% % 100% 72% 89% 13% NR
100% 33% NR 0% NR 0% 33% 0% 0% % 100% 72% 89% 12% NR 0% 33% 0% 0% % 100% 72% 89% 12% 13% NR 39% 17% NR NR NA NA NA NA NA NA NA NA NA NR NR NA NA NA NA NA NA NA NA NR NR NR NR NR NR NR NR NR 99% 72% NR NR NR NR NR N6 10%
60 12% 13% NR 39% 17% NR NR NA 10% 10% 10% 10%
NA NA NR NR NR NR NR NR NR NR NR NR NR 99% 72% NR NR NR NR NR NR 0% 1% ⁴
NR NR 99% 72% NR NR NR NR NR 0% 1% [‡]
72% NR NR NR NR NR 0% 1% [‡]

Table 3	(contin	Table 3 (continued): Characteristics of CT and Chest Radiog	ıcteristic	s of CT	and Che	st Radiograph)	r (Inde	raphy (Index Tests)	~											
Study			Inc	Index Test			CT	CT Abnormalities	ıalities					CT	CT Findings*	*sc				
Author	Imaging	Author Imaging Thickness	Indi Rea Readers ing	Indiv. Read- ing	Result Dichot.	Positive Test Definition	Unilat- eral		No Abnor- Bilateral malities GGO	GGO	Vase Consol. Dil.	Vascular Dil.	(Re- versed) Halo	Septal Re- Thick- ticu ening Patt	Septal Re- Thick- ticular ening Pattern	Crazy Paving	Crazy Nodular Ly Paving Lesions ad.	Crazy Nodular Lymph- Paving Lesions ad.	Pleural Effu- sion	Periph- eral Distr.
Yang et al (30)	CT	X	NR	NR	Yes	Multiple patch- like shadows (early), GGO (middle), consolida- tion shadow (late stage)	NR	NR	8%	50%	NR	NR	NR	NR	NR	NR	NR	NR	42%	XX
Chen et al (23)	CT	Thick (5 mm) in most	7	No	No	Eight semantic imaging features*	NR	NR	10%	NR	NR	84%	NR	56%	NR	54%	NR	%0	%0	NR
Miao et al (28)	CT	Thin (< 3 mm)	7	Yes	No	Combination of findings, previous report**	26%	74%	%0	70%	22%	NR	NR	NR	NR	30%	15%	7%	%0	61%
Dangis et al (32)	CT	Thin (< 3 mm)	7	No	Yes	Six main fea- tures; GGO, distribu- tion ^{+†}	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR
Note.— applicab * Present † Present * In one % Likert : lesions: (and pleu * Total n tree-in-b * " Previou * Total n tree-in-b	Note.—AI = artificial int applicable, NR = not repv * Presented is the percent † Presented is the percent * In one case, likely relate \$ Likert scores with scores lesions; (<i>d</i>) bilateral GGC lesions; (<i>d</i>) bilateral GGC lesions; (<i>d</i>) bilateral the * Total number of mixed tree-in-bud, pleural thick tree-in-bud, pleural thick tree-in-bud, pleural thick tree-in-bud, pleural thick tree-in-bud, pleu	Note.—AI = artificial intelligence, Consol. = consolidation, Dil. = c applicable, NR = not reported. * Presented is the percentage of CT findings within COVID-19-co [†] Presented is the percentage of CT findings within the entire COV [*] In one case, likely related to nephropathy. § Likert scores with scores 4–5 positive for COVID-19. Likert scale lesions; (d) bilateral GGO- and peripheral-predominant lesions; (e) and pleural effusion (26). [*] Previous reports of radiologic findings in COVID-19 (38–41) usir [*] Toral number of mixed GGO in peripheral area, total number of c tree-in-bud, pleural thickening, and offending vessel augmentation. [*] Previous viral pneumonia report (43) and COVID-19 features: G monary atelectasis, pleural thickening, and lesion distribution.	nce, Con of CT find f CT find f CT find nephropa positive 1 d periphe c findings) in perip s, and offi port (43) skening, a ultifocal i ultifocal i	sol. = co and dings wit lings wit for COV for COV for COV and pred and COV and lesio involven	insolidati thin CO ^N thin the e 'TD-19. 1 ominant 'TD-19 (f as, total r 'essel aug WID-19 in distribu	Note.—AI = artificial intelligence, Consol. = consolidation, Dil. = dilatation, Distr. = distribution, GGO = ground-glass opacity, Indiv. = individual, Lymphad. = lymphadenopathy, NA = not applicable, NR = not reported. * Presented is the percentage of CT findings within COVID-19-confirmed patients with CT abnormalities. * Presented is the percentage of CT findings within the entire COVID-19-confirmed cohort. * Presented is the percentage of CT findings within the entire COVID-19-confirmed cohort. * In one case, likely related to nephropathy. * In one case, likely related to nephropathy. * In one case, likely related to methropathy. * Thereious reports of radiologic findings in COVID-19 (38–41) using the Fleischner Society lexicon as reference (42). * Total number of mixed GGO in peripheral-predominant lesions. (<i>b</i>) bilateral GGO- and peripheral effusion (26). * Total number of mixed GGO in peripheral area, total number of consolidations, total number of solid nodules with ground-glass opacities, interlobular septal thickening, crazy-paving pattern, tree-in-bud, pleural thickening, and offending vesel augmentation in lesions. * Total number of mixed GGO in peripheral area, total number of consolidation, crazy-paving, air bronchogram, cavitation, pulmonary nodule, lymphadenopathy, pleural effusion, pulmonary necectasis, pleural thickening, and lesion distribution. * Previous viral pneumonia report (43) and COVID-19 features: GGO, consolidation, crazy-paving, air bronchogram, cavitation, pulmonary nodule, lymphadenopathy, pleural effusion, under there area, pleural distribution. * Previous viral pneumonia report (43) and COVID-19 features: GGO, consolidation, crazy-paving, consolidation, and, at a later stage, crazy-paving,	n, Disi d patiet -confirt on (a) C al GGC al GGC al tations, nas. nasolida	tr. = dist nts with ned coh J- and p J- and p ter Socie ter Socie ter Socie at a late: at a late:	lilatation, Distr. = distribution, GGO = ground-g nfirmed patients with CT abnormalities. ID-19-confirmed cohort. based on (<i>a</i>) GGO-predominant lesic bilateral GGO- and peripheral-predominant lesic ag the Fleischner Society lexicon as reference (42). consolidations, total number of solid nodules with in lesions. GO, consolidation, crazy-paving, air bronchogran ribution, and, at a later stage, crazy-paving, conso	GGO = rmalities -predomi a srefert solid noc g, air brc -azy-pavii	ground- ;; (b) GC inant lesi inant lesi inant lesi inant lesi inant lesi inant lesi inant lesi inant lesi inant lesi nuchogra	 Iilatation, Distr. = distribution, GGO = ground-glass opacity, Indiv. = individual, Lymphad. = lymphadenopathy, NA = not nfrmed patients with CT abnormalities. ID-19-confirmed cohort. based on (a) GGO-predominant lesions; (b) GGO- and peripheral-predominant lesions; (c) bilateral GGO-predominant bilateral GGO- and peripheral-predominant lesions; (c) bilateral GGO-predominant bilateral GGO- and peripheral-predominant lesions; (b) consolidations; nodules, mediastinal lymphadenopathy, sg the Fleischner Society lexicon as reference (42). consolidations, total number of solid nodules with ground-glass opacities, interlobular septal thickening, crazy-paving patter in lesions. GO, consolidation, crazy-paving, air bronchogram, cavitation, pulmonary nodule, lymphadenopathy, pleural effusion, pulcion, and, at a later stage, crazy-paving, consolidation, and reversed halo sign. 	ity, Indiv eripheral out airwa) glass opa glass opa rion, pulr and rever	. = indiv -predon ~ abnorr ~ abnorr cities, ir cities, ir nonary	idual, L ninant le nalities, nterlobu nodule, sign.	ymphad sions; (c nodules lar septa lymphae	. = lympl) bilatera , mediast 1 thicken denopath	hadenopa ul GGO- _F tinal lymy ing, crazy	tthy, NA redomit phadenoj -paving l effusior	= not nant pathy, pattern, , pul-

Table 4: Characte	Table 4: Characteristics Reported for RT-PCR (Reference Test) in Suspected COVID-19 Diagnosis	r RT-PCR (I	Reference .	Test) in Suspected	COVID-19) Diagnosis					
				Reference Test	Test					Time Interval (d)	al (d)
Author	Test	Specified Certified	Certified	Tissue Sample	Second PCR*	Repeated Sampling (> 2)	Ref Test Diagnosis No. of Sub- Based on Repeated jects Sample > 1 PCR	s No. of Sub- d jects > 1 PCR	Symptom Onset-PCR	First-Sec- ond PCR	First-Sec- ond PCR Index Test-PCR
Ai et al (20)	RT-PCR	Yes	CFDA	Oropharynx	Yes	Yes	No, ≤ 3 days initial	258	NR	3 or ≥ 4	1 (0-7)
Bai et al (21)	RT-PCR	Yes	CFDA	NR	NR	Yes	Yes	NR	NR	NR	4.1
Zhu et al (31)	RT-PCR	NR	NR	NR	Yes	no	NA	NR	5 (2-7)	1	0-1
Cheng et al (24)	RT-PCR	NR	NR	Oropharynx	Yes	NR	NR	NR	NR	NR	NR
Long et al (27)	RT-PCR	NR	NR	NR	Yes	Yes	Yes	9	NR	2-8	NR
Himoto et al (26)	RT-PCR and/or clinical con- sensus	NR	NR	NR	NR	NR	NR	NR	14 (1–19)	NR	NR
Caruso et al (22)	RT-PCR	Yes	NR	Nasopharynx and Oropharynx	Yes	No	NA	158	NR	1	1
Choi et al (25)	CT and RT-PCR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR
Wen et al (29)	RT-PCR	NR	NR	Oropharynx, spu- tum, or BAL	Yes	Yes	Yes	NR	NR	1–3	NR
Yang et al (30)	RT-PCR	NR	NR	Oropharynx	Yes	Yes	Yes	NR	NR	NR	NR
Chen et al (23)	RT-PCR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR
Miao et al (28)	RT-PCR	NR	NR	Nasopharynx or sputum	Yes	Yes	Yes	NR	NR	~ 	NR
Dangis et al (32)	RT-PCR	Yes	NR	Nasopharynx	Yes	no	NA	NR	NR	1	NR
Note.—BAL = bronchoalveolar lav * If first (RT-)PCR negative result.	Note.—BAL = bronchoalveolar lavage, CFDA = China Food and Drug Administration, NR = not reported, RT-PCR = reverse-transcription polymerase chain reaction. * If first (RT-)PCR negative result.	CFDA = Chi	na Food anc	l Drug Administratio	on, NR = n	ot reported, R	(T-PCR = reverse-tr	anscription po	lymerase chain	reaction.	

Diagnostic Test Accuracy

Disease prevalence (pretest probability) was reported in 20% (2/10) of DTA studies, and measures of precision were given in 80% (8/10) of studies (Table 5). On the basis of the reconstructed 2 \times 2 tables, ranges of CT performance in the 10 DTA studies were as follows: sensitivity 0.57-0.97, specificity 0.37-0.94, PPV 0.59-0.92, and NPV 0.57-0.96. For non-DTA studies, this was 0.92-0.94, 0.05-0.33, 0.23-0.35, and 0.67-0.93, respectively. An additional analysis of CT performance for a CT-to-RT-PCR time interval \leq 3 days was abandoned as only one small cohort provided data for the preset time interval (27). One chest radiography study reported a sensitivity and specificity of 0.25 and 0.90, respectively, with an area under the receiver operating characteristic curve of 0.58. We were unable to reliably reconstruct the 2 \times 2 contingency table for this study.

Meta-Analysis

Pooling diagnostic test accuracy results and performing a meta-analysis were considered not justified given the study heterogeneity and QUA-DAS-2 and STARD results (Table 6).

Risk of Bias

The purpose of imaging, that is, diagnosis of patients with suspected COVID-19 infection was clearly described in 54% (7/13) of studies (20,22,23,26-29). Two studies clearly described the role of imaging as replacement (23) and parallel/combined (32). Overall, QUADAS-2-based signaling questions (Appendix E3 [supplement]) were scored as low risk or concern in 38% (69/182), unclear in 43% (78/182), and high risk or concern in 19% (35/182) of studies. Risk of bias was scored low risk in 17% (24/143), unclear in 48% (68/143) and high risk in 36% (51/143) of questions. Risk of bias was highest for patient selection and flow and timing (Fig 2). Bias for the index and reference test was unknown for most studies. Not one study scored low risk of

Table 5: Reported Test Accuracy for COVID-19 Diagnosis	orted Test	Accura	cy for CC	J VID-19 L	Jiagnosi	S												
Parameter	Model	N Pa- tients	Sen.	95% CI	Spec.	95% CI	PPV	95% CI NPV	٨PV	95% CI Acc.	Icc.	95% CI AUC		95% CI	Pretest Prob.	Posttest Prob.	LR+ % LR-	LR- %
CT																		
Ai et al (20)		1014	0.97	95–98	0.25	22-30	0.65	62–68	0.83	76–89	0.68	65-70	NR	NR	0.59	NR	NR	NR
Bai et al (21)	Rater 1	424	0.72	66–78	0.94	89–97	0.92	87–96	0.76	70–81	0.83	79–68	NR	NR	NR	NR	NR	NR
	Rater 2		0.72	65-78	0.88	83–92	0.87	81–91	0.74	69-80	0.80	76-83	NR	NR	NR	NR	NR	NR
	Rater 3		0.94	70-07	0.24	18 - 30	0.57	52-62	0.79	67–88	0.60	55-65	NR	NR	NR	NR	NR	NR
Long et al (27)		36	0.97	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR
Himoto et al (26)	"E"-rater 21 1	r 21	0.67	NR	0.93	NR	0.80	NR	0.88	NR	0.86	NR	0.88	0.67–0.98 NR	3 NR	NR	NR	NR
	"E"-rater 2	5	0.83	NR	0.80	NR	0.63	NR	0.92	NR	0.81	NR	0.92	0.71-0.99 NR) NR	NR	NR	NR
Caruso et al (22)	al	158	0.97	88–99	0.56	45–66	0.59	53-64	0.96	87–99	0.72	64–78	NR	NR	NR	NR	NR	NR
Wen et al (29)		103	0.93	85–97	0.53	27–77	0.92	83–96	0.42	18-70	NR	NR	NR	NR	NR	NR	NR	NR
Chen et al (23)		38	1.00	NR	0.37	NR	NR	NR	NR	NR	0.68	NR	0.81	0.67–0.95 NR	S NR	NR	NR	NR
Miao et al (28)	Highest sens.	130	0.57	NR	0.81	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR
	Highest spec.		0.20	NR	0.99	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR
Dangis et al (32)	la	192	0.87	80–94	0.94	89–98	0.91.	85-97	0.90	85-96	0.91	91–91	NR	NR	0.43	0.91	13.5	0.14
Chest radiog- raphy																		
Choi et al (25)*		37	0.25	20–26	0.90	89–96	NR	NR	NR	NR	NR	NR	0.58	(0.53- $0.73)^{\dagger}$	NR	NR	NR	NR
Note.—Acc. = accuracy, AUC = area under the curve, CI = confidence interval, LR = likelihood ratio, NPV = negative predictive value, NR = not reported, PPV = positive predictive value, Prob. = probability, Sen. = sensitivity, Spec = specificity. * Median of eight observers. * AUC range in eight observers.	= accuracy, <i>i</i> bility, Sen. = ight observe n eight obse	AUC = al = sensitiv Trs.	rea under ity, Spec =	the curve, = specificity	CI = con. y.	fidence into	erval, Lŀ	R = likelih	100d rat	io, NPV =	- negative	predictive	/alue, N	IR = not re	ported, PP	V = positive J	predictive v	'alue,

							2×2	2 data									Ductoot	Doctroot 1	Doctoot
Parameter N	Model	N Total	N Total N Calc. PCR+	PCR+	CT+	CT-	ΤΡ	FP]	TN F	FN PPV	V NPV	/ Sen.	Spec.	FP Rat	FP Rate FN RateAcc	eAcc	Prob.	Prob.	Prob.
CT diagnostic																			
Ai et al (20)		1014	1014	601	888	126	580	308	105	21 0.	0.65 0.83	3 0.97	7 0.25	0.75	0.03	0.68	0.59	0.56	0.12
Bai et al (21)	Rater 1	424	424	219	171	253	158	13	192	61 N	NA NA	A 0.72	2 0.94	0.06	0.28	0.83	NA	0.92	0.23
	Rater 2	424	424	219	181	243	157	24	181	62 N	NA NA	A 0.72	2 0.88	0.12	0.28	0.80	NA	0.86	0.24
	Rater 3	424	424	219	362	62	206	156	49	13 N	NA NA	1 0.94	4 0.24	0.76	0.06	0.60	NA	0.55	0.20
Long et al (27)		87	36	36	35	NR	35	NR	NR	1 N	NA NA	V 0.97	7 NA	NA	0.03	NA	NA	NA	NA
Himoto et al (26)	Rater 1 highest AUC	21	21	9	Ś	16	4		14	2 0.	0.80 0.88	8 0.67	7 0.93	0.07	0.33	0.86	0.29	0.91	0.26
	Rater 2 highest AUC	21	21	9	×	13	\sim	\tilde{c}	12	1 0.	0.63 0.92	2 0.83	3 0.80	0.20	0.17	0.81	0.29	0.81	0.17
Caruso et al (22)		158	158	62	102	56	60	42	54	2 0.	0.59 0.96	6 0.97	7 0.56	0.44	0.03	0.72	0.39	0.69	0.05
Wen et al (29)		103	103	88	89	14	82	7			0.92 0.57	7 0.93	3 0.53	0.47	0.07	0.87	0.85	0.67	0.11
Chen et al (23)	Valida- tion R- cohort	38	38	19	31	~	19	12	~	N 0	NA NA	1	0.37	0.63	0	0.68	NA	0.61	0
Miao et al (28)	Highest sens. model*	130	130	54	46	84	31	15	61	23 0.	0.67 0.73	3 0.57	7 0.80	0.20	0.43	0.71	0.42	0.74	0.35
Dangis et al (32)		192	192	83	79	113	72	~	102	11 0.	0.91 0.90	0 0.87	7 0.94	0.06	0.13	0.91	0.43	0.93	0.12
CT nondiagnostic																			
Zhu et al (31)		116	116	32	86	30	30	56	28	2 0.	0.35 0.93	3 0.94	4 0.33	0.67	0.06	0.50	0.28	0.58	0.16
Cheng et al (24)		38	38	11	33	Ś	11	22	5	0 0.	0.33 1.00	0 1.00	0 0.19	0.81	0.00	0.42	0.29	0.55	0.00
Yang et al (30) Chest radiogra-		55	55	13	52	ŝ	12	40	5	1 0.	0.23 0.67	7 0.92	2 0.05	0.95	0.08	0.25	0.24	0.49	0.62
phy Choi et al (25)*		37	23	20	*-	+-	\$	+-	18	Z	NA	NA	NA	NA	NA	ΝA	NA	NA	NA

11

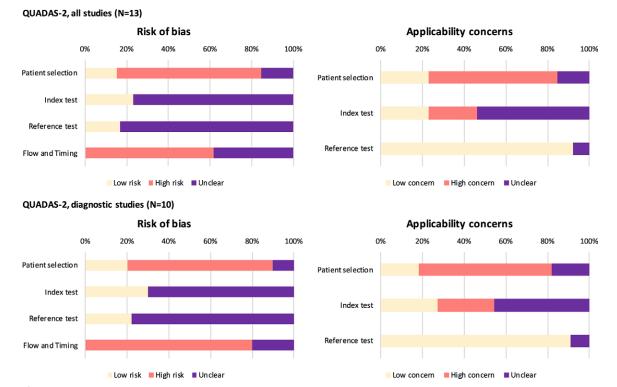
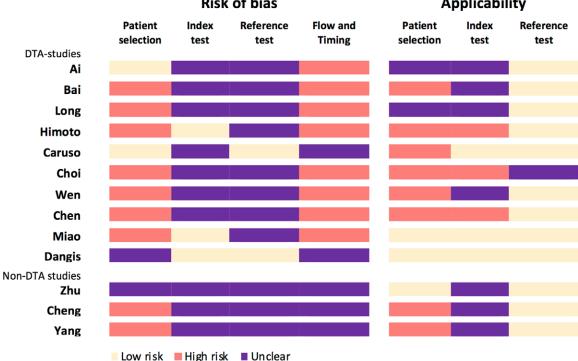


Figure 2: Quality Assessment of Diagnostic Accuracy Studies-2 (QUADAS-2) results per domain. QUADAS-2: results for all studies (upper part) and for diagnostic test accuracy studies (lower part) present per domain.



Risk of bias

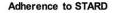
Applicability

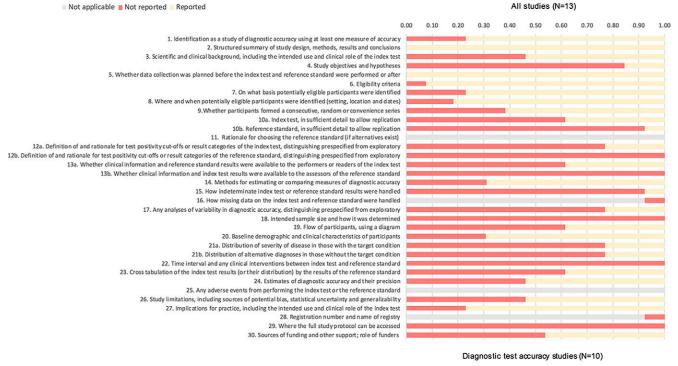
Figure 3: Quality Assessment of Diagnostic Accuracy Studies-2 (QUADAS-2) results per study. QUADAS-2: results for risk of bias and applicability presented as scored per study.

bias for all four QUADAS-2 domains, and three studies scored no high risk in any domain (Fig 3). High risk in at least one domain was scored in 77% (10/13) of studies, and high risk in at least two domains was scored in 54% (7/13) of studies.

Applicability

Applicability of studies was scored low concern in 38% (15/39), unclear in 26% (10/13), and high concern in 28% (11/13). Applicability concern was highest for patient selec-





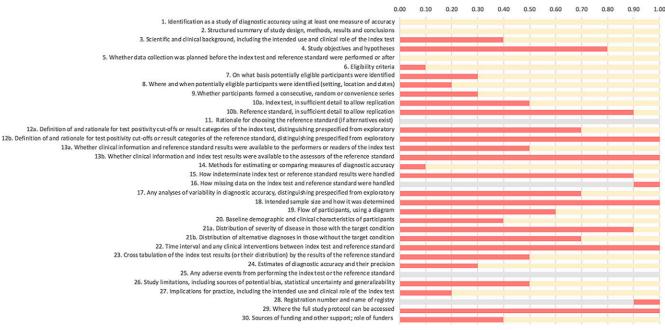


Figure 4: Adherence to Standards for Reporting of Diagnostic Accuracy Studies (STARD). Presented are the proportions of (non)reported items for each study according to the STARD guidelines; presented for all studies (top) and diagnostic test accuracy studies (bottom). The different STARD items concern the following sections in the reports: title or abstract (1), abstract (1,2), introduction (3), methods (4–18), results (19–25), discussion (26,27), and other (28–30).

tion, scored unclear in 15% (2/13) and high risk in 62% (8/13) of studies, respectively, followed by concern regarding index test applicability, scored unclear and high concern in 54% (7/13) and 23% (3/13) of studies, respectively (Fig 3). Two studies (15%) scored low concerns for applicability for all domains, and three studies (20%) scored only low or unclear concerns. High concern for applicability in at least one domain was scored in eight studies (62%), and high concern for applicability in at least two domains was scored in two studies (15%).

Quality of Reporting

The mean percentage of reported STARD items was 35% (12/34) for all studies (Fig 4; total STARD items = 34); for

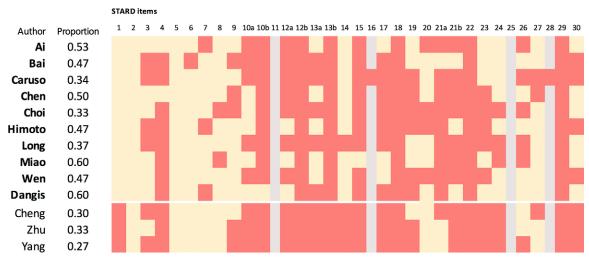


Figure 5: Standards for Reporting of Diagnostic Accuracy Studies (STARD) adherence per study. Graphical display of reported (green) and not reported (red) STARD items per study. The reported proportion is calculated by dividing the reported items by the total of reported and not reported items (not applicable items [gray] are not taken into account in this analysis) to applicable items. Upper 10 studies concern diagnostic test accuracy studies. The individual STARD items (presented as 1–30) are listed in Figure 4.

DTA studies, this was a mean of 41% (14/34) and for non-DTA, this was 26% (9/34). The mean proportion (to applicable items) of STARD adherence for all studies was 43% (range, 27%–60%), which was 47% (33%–60%) for DTA studies and 30% (27%–33%) for non-DTA studies (Fig 5).

The most unreported STARD included all items on the reference test, test positivity cutoffs of the reference test and index test, information on time interval and clinical intervention (eg, antiviral treatment) between index test and reference test, handling of indeterminate test results, intended sample size, study objectives, and hypothesis and severity of disease (Fig 4; Appendix E4 [supplement]). STARD items most often reported included identification of the diagnostic accuracy study, structured abstract, data collection prospective/retrospective, eligibility criteria, and methods for comparing measures of diagnostic accuracy (Appendix E4 [supplement]).

GRADE Framework

On the basis of the GRADE framework, the certainty of evidence for the totality of CT test accuracy studies for CO-VID-19 diagnosis was rated very low for both sensitivity and specificity estimates (Table 7). The certainty of evidence for chest radiography was not evaluated within GRADE as only one study was identified.

Discussion

Despite the vivid discussion on the use of imaging to screen and diagnose patients with suspected COVID-19 infection, only few studies addressed and reported diagnostic test accuracy of chest imaging. Reported diagnostic test accuracy of CT for diagnosis of suspected COVID-19 infection in individuals varies substantially, with ranges of 0.57–0.97 for sensitivity and 0.37–0.94 for specificity and recalculated PPV of 0.59–0.92 and NPV of 0.57–0.96, with a very poor level of evidence. The latter is not surprising, as these studies were conducted, written, and published in an extraordinary point in time, under time pressure and rather stressful conditions. However, poor adherence to reporting guidelines and substantial risk of bias may have substantial implications for the daily care of individuals suspected of having COVID-19 and the health care system (33).

The performance of a diagnostic test is subject to its setting: Tests may perform perfectly in a certain setting and achieve suboptimal results when applied in a different setting. For example, in clinical practice, PPV and NPV (the probability of diagnosing the disease with a positive test result and ruling out the disease with a negative test result, respectively) are the most important measures when estimating the risk of presence or absence of disease in patients with suspected infection. The predictive values, however, directly depend on disease prevalence (34). When studies only include patients with confirmed COVID-19, they will not include information on true-negative cases and are therefore prone to low thresholds for test positivity, thereby inflating test sensitivity. An example of such a case is the recently published meta-analysis on CT DTA studies for COVID-19 diagnosis (35), as the vast majority of subjects were patients with confirmed COVID-19. Hereby, the accuracy for discrimination between the diseased and nondiseased remains unknown. Other important sources of test variation include patient demographics, the severity of disease, index and reference test execution, and (predefined or chosen) thresholds but also those related to initial study design and risk of bias. QUADAS-2 and STARD are efficient and clear tools for authors and readers of diagnostic accuracy studies to evaluate the reported study setting and judge the bias, applicability, and expected test performance (16,36).

Complete reporting is needed to allow for the assessment of sources of variation and potential bias and judge reported and expected test performance. Bias in patient selection may result in overestimation of diagnostic accuracy. High pretest probability and strict selection of patients with severe disease or high risk

Table 7: GRADE Framework: Certainty of Evidence for the Totality of CT Test Accuracy Studies for COVID-19 Diagnosis	nework: Certainty	of Eviden	ce for the To	hality of CT Tes	st Accuracy Si	rudies for CO	VID-19 Diagne	osis			
		I]	Factors that May Decrease Certainty of Evidence	y Decrease Cer	tainty of Evider	JCE	Effect pe	Effect per 1000 Patients Tested	ts Tested	
Outcome	No. of Studies (No. of Patients)Study Design Risk of Bias	ly Design 1		Indirectness	Inconsistency Imprecision		Publication Bia	Pretest Prob- ability of Publication Bias ability of 29% 59%	Pretest Prob- ability of 59%	Pretest Prob-Pretest Prob- ability of ability of 59% 85%	Test Accuracy CoE
True-positives (patients with CO- VID-19) False-negative (patients incorrectly classi-	9 studies (2167 Cc patients)	phort and case-con- trol type studies	Very serious*	Very serious [†]	Serious	Not serious	None	165 to 290 0 to 125	336 to 590 0 to 254	336 to 590 484 to 850 0 to 254 0 to 366	000 Very low
fied as not having COVID-19) True-negative (patient: without CO-	fied as not having COVID-19) True-negative (patients 9 studies (2167 Cohort and Very serious* Very serious* without CO- patients) case-con-	hort and case-con-	Very serious*	Very serious [†]	Very serious Not serious	Not serious	None	170 to 667	98 to 385	36 to 141	⊕OOO Very low
VID-19) False-positives (pa- tients incorrectly classified as having COVID-19)		trol type studies						43 to 540	25 to 312	9 to 114	`
Note.—Sensitivity was 0.57 to 1.00, and specificity was 0.24 to 0.94. Prevalence was 29%, 59%, or 85%. * Multiple studies with high concern for bias in the QUADAS-2 domains: patient selection and flow and timing. † Multiple studies with high concern for applicability in the QUADAS-2 domains: patient selection and index test	0.57 to 1.00, and spe high concern for bias high concern for appli	cificity was in the QU icability in	0.24 to 0.94. ADAS-2 dom the QUADAS	Prevalence was ains: patient selk 3-2 domains: pa	29%, 59%, or ection and flow tient selection :	85%. ^ and timing. and index test.					

(COVID-19 exposure history) will typically result in spectrum bias with increased (or overestimated) test sensitivity. In this review, the majority of studies did not report on more demographics than age and sex, and no information on socioeconomic status was provided. Study subjects were typically patients presenting at the emergency department/hospital with fever and/or exposure history to COVID-19 (area or person) and with unknown severity of disease or the distribution of alternative diagnoses. Concerns with regard to generalizability rise with the inclusion of patients with very high suspicion and/or exposure history (29), a previous positive test (not further specified) (22), or the use of patients with another type of viral pneumonia as the control group (21,23); with the exclusion of patients with noninfectious lung disease (28), nonpregnant, or non-Chinese subjects (30); if CT was performed within 3 days after symptom onset (26); or CT findings were normal (21,23-25,27,28).

Six studies excluded patients with normal chest CT findings (or included abnormal CT only) (21,23-25,27,28) and/or did not perform RT-PCR in patients with suspected infection without CT abnormalities (106 of 204 excluded participants) (27). This selection may result in the overestimation of test performance. Test sensitivity and specificity may also be overestimated with an arbitrary choice of test threshold. As CT and chest radiography are multilevel tests, a definition of positive (or negative) test is required, but this was lacking in 69% (9/13) of included studies. One study defined a positive test cutoff when three of eight readers scored chest radiograph "positive" (25). For most, if not all, studies it was unclear how indeterminate results were handled. Higher sensitivity and (possibly) lower specificity may also be driven by verification bias (14,15,37). An example of partial verification bias occurred in the study in which RT-PCR was performed in patients with abnormal CT findings only (27). Differential verification bias may have occurred if the use of different (more invasive) swab specimens (eg, nasopharynx versus bronchoalveolar lavage) (28,29) or different reference test analysis methods or PCR kits (20,21) were driven by CT findings, although this is unclear. In addition, various rates of RT-PCR sensitivity have been reported, suggesting a potential imperfect standard of reference (7,8). Bias may be introduced if RT-PCR sensitivity is in fact lower, resulting in misdiagnosis and underestimation of index test performance.

Proper reporting of the reference test method and analysis is therefore required. However, studies typically did not provide information on the type of RT-PCR test kit, specification, or certification and/or whether single or multiple swabs were taken. In these cases, it was unclear whether study subjects received the same reference test. Since information on RT-PCR was insufficient, generalizability cannot be assessed. In addition, no report was made on treatment or other clinical intervention between index test and RT-PCR, thus the potential treatment paradox is unclear. Disease progression bias may potentially have been introduced considering the relatively large reported time intervals (up to 8 days) between CT and RT-PCR. Especially in early COVID-19, most patients with suspected COVID-19 were admitted to fever clinic departments and cross-infection may have occurred within this time interval. No study reported on blinding for the index test result, although with (semi)automated analyses, this may be considered irrelevant. Access to RT-PCR results when reading the index test will, however, induce unacceptable review bias and overestimate test performance. Blinding for RT-PCR results was unclear in four DTA studies (22,24,25,27). Other general concerns for generalizability of the index test not yet assessed include the use of thick-slice CT (23,25,26). CT performance also highly depends on readers' experience and rater reproducibility, although results on intrarater (32) or interrater (26,32) performance were rather scarce.

Our systematic review had several limitations. We attempted to identify all published articles by including a nonindexed journal, although more DTA studies may be reported in other nonindexed journals that we are unfamiliar with. Only a small number of studies were eligible for inclusion. Five non-English articles were excluded for language restrictions given the short time frame and urgency of this review, although likely similar bias and study weakness would have been encountered. A meta-analysis was not performed due to the low reporting quality. In the setting of prompt diagnosis, we sought to provide additional diagnostic test results for patients who underwent CT and RT-PCR tests within the predefined time interval of 3 days. Only one study allowed for recalculation of CT performance within the predefined time interval (27). Interpretation of the QUADAS-2 and STARD items is subjective; for this, multiple readers assessed the items individually.

To conclude, certainty of evidence was rated very low for both sensitivity and specificity estimates of CT for COVID-19 diagnosis in patients with clinical suspicion. Reported test accuracy of CT for diagnosis of COVID-19 infection varies substantially, from rather poor to excellent. Validity and generalizability of these findings is complicated by poor adherence to reporting guidelines and high risk of bias, which are most likely due to the need for urgent publication of findings in the first months of the COVID-19 pandemic.

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