

Effect of white-coat hypertension on arterial stiffness

A meta-analysis

Peng Cai, MD^a, Yan Peng, MD^a, Yan Wang, MD^{b,*}, Xukai Wang, MD^{a,*}

Abstract

Background: White-coat hypertension (WCH) is a debatable risk factor of cardio-cerebrovascular diseases and the current study results on the association between WCH and arterial stiffness are inconsistent. The aim was to investigate the effect of WCH on arterial stiffness using meta-analysis.

Methods: Based on prespecified search strategies and inclusion criteria, Medline, Embase, Web Of Science, Cochrane Library, and BioSciences Information Service Preview databases were reviewed. A total of 20 studies involving 1538 WCH patients and 3582 normotensives (NT) were included. Literatures were screened for data extraction and quality assessment. Overall analysis and subgroup analysis were conducted in RevMan version 5.3 and Stata version 14.0 software.

Results: Overall analysis showed that carotid-femoral pulse wave velocity (cf-PWV) was significantly higher in WCH group than in the NT group (P < .00001, 95% CI: 0.79–3.26). Subgroup analysis showed that in adults, cf-PWV was significantly higher in the WCH patients than in the NT subjects (P < .001, 95% CI: 0.46–0.87), while in juveniles, cf-PWV was comparable between the WCH group and the NT group (P = .25, 95% CI: -0.39 to 0.61).

Conclusion: This meta-analysis showed that WCH may increase arterial stiffness in adult population.

Abbreviations: cf-PWV = carotid-femoral pulse wave velocity, NT = normotensive, WCH = white-coat hypertension.

Keywords: arterial stiffness, cardio-cerebrovascular disease, isolated clinic hypertension, pulse wave velocity, white-coat hypertension

1. Introduction

White-coat hypertension (WCH), also termed isolated clinic hypertension, is seen in the patients who show hypertension during the clinic visits.^[1] Currently, the diagnostic criteria of hypertension has been updated and the diagnostic criteria of WCH vary by guidelines.^[2,3] The widely used traditional criteria defines WCH as:

Clinic systolic blood pressure \geq 140 mm Hg and/or diastolic pressure \geq 90 mm Hg, and mean ambulatory blood pressure <135/85 mm Hg daytime or home blood pressure <135/85 mm Hg. WCH was once considered a benign phenomenon, but

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Received: 28 May 2018 / Accepted: 26 September 2018 http://dx.doi.org/10.1097/MD.000000000012888 several studies have established its relationship with multiple metabolic disorders such as impaired glucose tolerance, insulin resistance, and metabolic syndrome.^[4,5] Ongoing studies have been directed to clarify the role of WCH in cardio-cerebrovascular impairments.^[6]

Arterial stiffness examination is a noninvasive tool to evaluate cardio-cerebrovascular risks. Many clinical studies and basic researches have revealed arterial stiffness as a risk factor of cardiocerebrovascular diseases. With the popularization of arterial stiffness examination, some indicators such as pulse wave velocity (PWV), ambulatory arterial stiffness index (AASI), and augmentation index have been developed. Of note, both American Heart Association scientific statement and European expert consensus have recommended PWV as the golden standard for arterial stiffness with consideration to its high accuracy and applicability.^[7] To identify the target organs of WCH in cardio-cerebrovascular impairments, several clinical studies have attempted to investigate the relationship between WCH and arterial stiffness. However, their results vary due to confounding factors such as small sample size, racial difference, inconsistent methods, and discrepant inclusion criteria.^[8] In light of the inconsistencies of relationship between WCH and arterial stiffness, this systematic review and meta-analysis were conducted to evaluate the relationship between WCH and arterial stiffness.

2. Methods

2.1. Search strategies

Medline, Embase, Web Of Science, Cochrane Library, and BioSciences Information Service (BIOSIS) Preview databases were searched using the combination of text words and keywords

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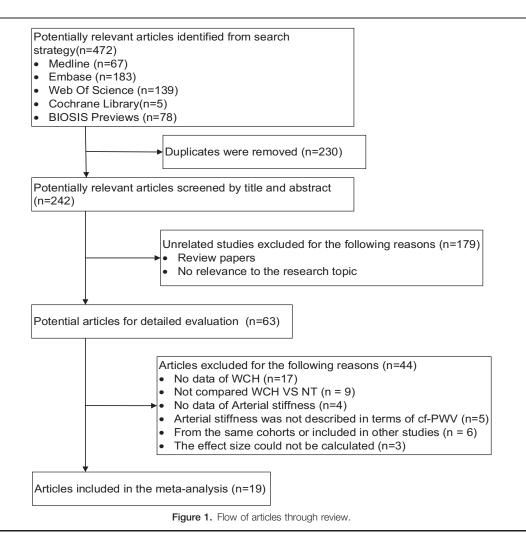
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of the following terms: "clinic hypertension, "office hypertension," "white-coat," "PWV," "pulse wave velocity," "arterial stiffness," "aortic stiffness," and "vascular stiffness." Publication date was limited to December 23, 2017.

2.1.1. Inclusion criteria.

- 1. Arterial stiffness measured by cf-PWV;
- 2. Case-control studies including WCH group and NT group;
- 3. WCH was defined as an office BP ≥140/90 mm Hg with day ABPM <135/85 mm Hg.

Literatures of the same study population, poor research quality, and incomplete data reporting were excluded. If a paper included several independent case-control groups, they were screened and the eligible ones were included in the meta-analysis. Figure 1 shows the flowchart of study design.

2.2. Data extraction and quality assessment

Two investigators (PC and YP) independently searched literature, screened studies, and extracted data on the basis of searched strategies, and inclusion criteria. The quality of studies was assessed by population selection, comparability between cases and controls, and exposure measurement in accordance with the Newcastle-Ottawa Quality Assessment Scale (NOS). The NOS contains 8 items with a maximum score of 9 points. All studies were classified as low quality (0–3 points), medium quality (4–6 points), or high quality (7–9 points) based on NOS.^[9]

2.3. Statistical analysis

The cf-PWV was compared between WCH group and NT group. All statistical analyses were conducted in RevMan software version 5.3 (The Cochrane Collaboration, Copenhagen, Denmark) and Stata version 14.0 (Stata Corp LP, College Station, TX). All the data were calculated for their 95% confidence intervals (95% CI). Statistical difference was defined as a 2-sided *P* value equal to or smaller than .05.

All the data were transformed into mean±standard deviation format by either RevMan version 5.3 software or manual calculation. Publication bias analysis, sensitivity analysis, heterogeneity analysis, data synthesis, Z test, meta-regression analysis, and subgroup analysis were performed. Publication bias was analyzed with Begg and Egger tests and visually examined by funnel plot. Sensitivity analysis was performed with Cohen test and graphical methods. Twelve was used to quantitatively assess heterogeneity. When significant heterogeneity was indicated by I2>50%, the random-effects model was used to calculate effect size; otherwise, fixed-effects model was used, followed by Z test. Subgroup analysis was performed for age, blood pressure, Table 1

| Reference | Cohort | Sample size (% WCH) | Age, y (WCH) | PWV, m/s (WCH) | PWV, m/s (NT) | Instrument for inspecting PWV | Antihypertensive drug users | NOS quality |
|--|-----------|------------------------|-----------------|-------------------|------------------|--|--------------------------------|----------------|
| Ribeiro et al ^[10] | Portugal | 47 (59.6) | 44±5 | 9.3 ± 1.68 | 9.2±1 | Complior device (Artech Medical, Pantin, France) | No | High |
| Ribeiro et al [10] | Portugal | 34 (52.9) | 45±6 | 11.6±0.87 | 9.6±1.16 | Complior device (Artech Medical, Pantin, France) | No | High |
| Silva et al ^[11] | Portugal | 219 (39.7) | 48 | 9.9±1.45 | 8.9±1.45 | Complior device (Artech Medical, Pantin, France) | Mixed | High |
| Stolarz-Skrzypek et al ^[12] | Poland | 222 (9.1) | 32.8±12.9 | 9.39±1.23 | 8.56 ± 1.45 | Complior device (Artech Medical, Pantin, France) | No | High |
| Andrikou et al ^[13] | Greece | 125 (64.8) | 52±8 | 7.5±1.2 | 6.8 ± 0.5 | Complior device (Artech Medical, Pantin, France) | Mixed | High |
| Schillaci et al ^[14] | Italy | 204 (65.2) | 49±12 | 9.3±2 | 8.5±2 | SphygmoCor device (AtCor Sydney, Australia), | No | High |
| Martin et al ^[15] | Australia | 65 (44.6) | 55.8 ± 8.3 | 7.8 ± 0.72 | 8±0.72 | Millar Mikro-tip (Millar Instruments, Houston, USA) | No | High |
| Sozeri et al ^[16] | Turkey | 108 (7.4) | 5-18 | 5.6 ± 0.61 | 5.3 ± 0.7 | Vicorder (Skidmore Medical Limited, Bristol, UK) | No | High |
| Hopkins et al ^[17] | UK | 35 (25.7) | 50 ± 1.7 | 10.2 ± 1.03 | 8.7 ± 1.03 | Unknown | Mixed | Medium |
| Protogerou et al ^[18] | UK | 134 (27.6) | 59.2±12.3 | 8.9 ± 2.7 | 7.8±1.7 | GD Konstantonis | Mixed | High |
| Sung et al ^[19] | China | 403 (38) | 58±13 | 8.9 ± 2.1 | 8.1±1.6 | Parks model 802 (Parks Medical Electronics, Inc) | No | High |
| Jurko et al ^[20] | Slovakia | 56 (50) | 17-18 | 7.7±2.5 | 7.9±2 | system VaSera 1500 (Japan) | No | Medium |
| Chatzistamatiou et al [21] | Greece | 273 (47.6) | 56 | 8.1 ± 1.49 | 7.5±1.49 | SphygmoCor device (AtCor Sydney, Australia), | Mixed | High |
| Afsar et al ^[22] | Turkey | 120 (9.2) | 57.2±14.0 | 8.52 ± 2.5 | 7.23±1.7 | MPX5050 (Freescale Inc, Tempe, AZ). | Mixed | High |
| Almeida et al ^[23] | Portugal | 490 (64.3) | 48±15 | 9.7 ± 2.4 | 9.5±2 | Complior device (Artech Medical, Pantin, France) | No | High |
| Scuteri et al [24] | USA | 1908 (9.8) | 58.8±12.5 | 7.9 ± 2 | 6.4 ± 1.7 | 3500-ATL Ultramark Inc | Mixed | High |
| Wojciechowska et al ^[25] | Poland | 135 (14.8) | 49±15.3 | 8.14 ± 1.5 | 6.47±1 | SphygmoCor device (AtCor Sydney, Australia), | Mixed | High |
| Barochiner et al [26] | Argentina | 71 (32.4) | 71.6±9.4 | 9.2±3.4 | 8±2.3 | SphygmoCor device (AtCor Sydney, Australia), | Yes | High |
| Nemcsik et al [27] | Hungary | 84 (20.2) | adult | 10.49±2.76 | 8.06 ± 1.61 | PulsePen | Mixed | Medium |
| Androulakis et al ^[28] | Greece | 387 (52.7) | 54.3±0.9 | 8.6 ± 0.98 | 7.6±0.88 | Complior device (Artech Medical, Pantin, France) | No | High |

NOS=Newcastle-Ottawa Quality Assessment Scale, NT=normotensives, PWV=pulse wave velocity, UK=United Kingdom, USA=United States of America, WCH=white-coat hypertension.

instrument for inspecting PWV, history of diabetes mellitus and/ or cardiovascular diseases, and quality score. For patients without history of diabetes mellitus and cardiovascular diseases, we further conducted subgroup analysis by antihypertensive treatments, meta-regression analysis was conducted in Stata version 14.0 to identify the sources of heterogeneity. All analyses were based on previous published studies, thus no ethical approval and patient consent are required.

3. Results

3.1. Studies retrieved and characteristics

A total of 472 articles were retrieved from Medline, Embase, Web Of Science, Cochrane Library, and BIOSIS Preview databases. After duplicate removal, the articles were screened by title, abstract and then full-text, thus 19 articles were finally included. The eligible articles included 5120 subjects (WCH group: 1538, NT group: 3582) from 20 studies and 12 countries. Baseline characteristics varied by study. Two studies included juveniles, while the remaining studies included adults. Only 1 study specifically included antihypertensive drug users, 10 studies specifically included nonantihypertensive drug users, and the remaining studies included mixed users. Regarding comorbidities, 9 studies excluded patients with diabetes mellitus or cardiovascular diseases. NOS score was medium and high in 3 and 17 studies, respectively. Table 1 shows the baseline characteristics.^[10–28]

3.2. Relationship between WCH and PWV

3.2.1. Overall analysis. Meta-analysis of 20 eligible studies showed cf-PWV was significantly higher in WCH group than in NT group (Z=6.57, P < .00001, 95% CI: 0.79–3.26; Fig. 2), but the heterogeneity was noticeable (I2=82%). Egger test and Begg test revealed neither publication bias nor small-study effects (Egger test, P=.751; Begg test, P=.626), and Fig. 3 visually reflected the publication bias. For random-effects model, sensitivity analysis revealed no significant changes of effect size (Fig. 4). Meta-regression analysis indicated the heterogeneity was partly attributed to comorbidities including diabetes mellitus and cardiovascular diseases (P < .05, R-squared=21.24%).

3.2.2. Subgroup analysis. The studies were stratified by the history of antihypertensive drug use, age, instrument for

| | WCH | | | NT | | | | Mean Difference | Mean Difference |
|---|----------------------|-------|----------|----------|--------|------------------------|-----------|------------------------|---|
| Study or Subgroup | Mean | SD | Total | Mean | SD | Total | Weight | IV. Random, 95% CI | Year IV, Random, 95% Cl |
| Ribeiro, L 1 2000 | 9.3 | 1.68 | 28 | 9.2 | 1 | 19 | 4.4% | 0.10 [-0.67, 0.87] | 2000 |
| Ribeiro, L 2 2000 | 11.6 | 0.87 | 18 | 9.6 | 1.16 | 16 | 4.8% | 2.00 [1.30, 2.70] 2 | 2000 |
| Silva, J. A. 2004 | 9.9 | 1.45 | 87 | 8.9 | 1.45 | 132 | 6.2% | 1.00 [0.61, 1.39] | 2004 |
| Stolarz-Skrzypek, K 2008 | 9.39 | 1.23 | 20 | 8.56 | 1.45 | 202 | 5.3% | 0.83 [0.26, 1.40] | 2008 |
| Andrikou, I 2011 | 7.5 | 1.2 | 81 | 6.8 | 0.5 | 44 | 6.6% | 0.70 [0.40, 1.00] | 2011 |
| Schillaci, G 2011 | 9.3 | 2 | 133 | 8.5 | 2 | 71 | 5.3% | 0.80 [0.22, 1.38] | 2011 |
| Martin, C. A 2011 | 7.8 | 0.72 | 29 | 8 | 0.72 | 36 | 6.4% | -0.20 [-0.55, 0.15] 2 | 2011+ |
| Sozeri, B 2012 | 5.6 | 0.61 | 8 | 5.3 | 0.7 | 100 | 6.0% | 0.30 [-0.14, 0.74] 2 | 2012 + |
| Hopkins, S. 2013 | 10.2 | 1.03 | 9 | 8.7 | 1.03 | 26 | 4.4% | 1.50 [0.72, 2.28] | 2013 |
| Protogerou, A. D 2013 | 8.9 | 2.7 | 37 | 7.8 | 1.7 | 97 | 3.7% | 1.10 [0.17, 2.03] | 2013 |
| Sung, S. H 2013 | 8.9 | 2.1 | 153 | 8.1 | 1.6 | 250 | 6.3% | 0.80 [0.41, 1.19] | 2013 |
| Jurko, A 2014 | 7.7 | 2.5 | 28 | 7.9 | 2 | 28 | 2.9% | -0.20 [-1.39, 0.99] 2 | 2014 |
| Chatzistamatiou, E 2015 | 8.1 | 1.49 | 130 | 7.5 | 1.49 | 143 | 6.4% | 0.60 [0.25, 0.95] | 2015 |
| Afsar, B 2015 | 8.52 | 2.5 | 11 | 7.23 | 1.7 | 109 | 2.1% | 1.29 [-0.22, 2.80] 2 | 2015 |
| Almeida, J 2016 | 9.7 | 2.4 | 315 | 9.5 | 2 | 175 | 6.2% | 0.20 [-0.20, 0.60] 2 | 2016 + |
| Scuteri, A 2016 | 7.9 | 2 | 187 | 6.4 | 1.7 | 1721 | 6.7% | 1.50 [1.20, 1.80] 2 | - The restored and the second s |
| Wojciechowska, W 2016 | 8.14 | 1.5 | 20 | 6.47 | 1 | 115 | 4.8% | 1.67 [0.99, 2.35] | 2016 |
| Barochiner, J 2017 | 9.2 | 3.4 | 23 | 8 | 2.3 | 48 | 2.0% | 1.20 [-0.33, 2.73] | 2017 |
| Nemcsik, J 2017 | 10.49 | 2.76 | 17 | 8.06 | 1.61 | 67 | 2.4% | 2.43 [1.06, 3.80] | |
| Androulakis, E 2017 | 8.6 | 0.98 | 204 | 7.6 | 0.88 | 183 | 7.0% | 1.00 [0.81, 1.19] | |
| Total (95% CI) | | | 1538 | | | 3582 | 100.0% | 0.86 [0.60, 1.12] | • |
| Heterogeneity: Tau ² = 0.24; | Chi ² = 1 | 05.80 | , df = 1 | 9 (P < 0 | .0000 | 1); ² = (| 82% | 0 00 | |
| Test for overall effect: Z = 6 | | | | | | 1100 | | | -2 -1 0 1 2 |
| | • | | , | t plot c | of the | comp | arison: w | hite-coat hypertensior | Favours [NT] Favours [WCH] n versus normotension. |

inspecting PWV and study quality, and Table 2 shows all subgroup analysis results. For adults, PWV was significantly higher in WCH group than in NT group (P<.001, 95% CI: 0.46–0.87), but PWV was not different between WCH group and NT

group in juveniles (P=.253, 95% CI: -0.39 to 0.61). In the subgroup analysis of 9 studies excluding patients with diabetes mellitus or cardiovascular diseases, heterogeneity was significantly reduced (I2=45%; Fig. 5), and PWV differed between

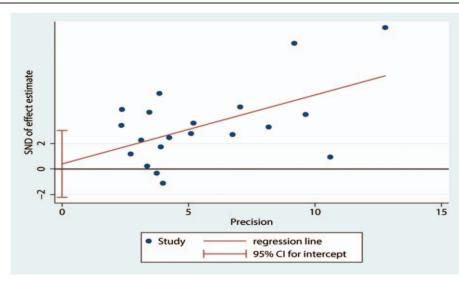
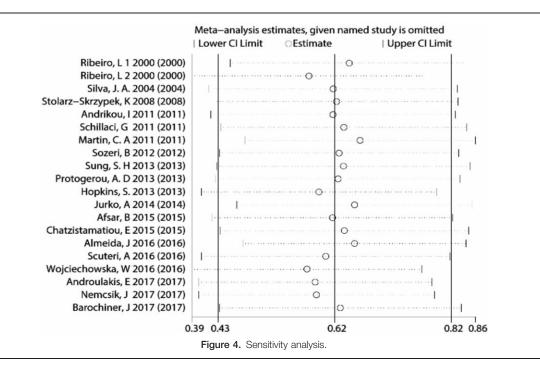


Figure 3. Publication bias. SND=standard normal deviation.



WCH group and NT group (P < .00001, 95% CI: 0.43–0.73; Fig. 5). When these 9 studies were further divided by history of antihypertensive drug use, untreated group and mixed group showed significantly reduced heterogeneity (I2=2%; I2=0%; Fig. 5), and PWV differed between WCH group and NT group (P=.01, 95% CI: 0.07–0.55; P < .00001, 95% CI: 0.56–0.95; Fig. 5).

4. Discussion

Table 2

Meta-analysis evaluated the relationship between WCH and arterial stiffness. It was found that adult WCH patients had significantly higher cf-PWV than normal population, indicating higher risks of cardio-cerebrovascular diseases in these patients. However, juveniles did not show the phenomenon, probably attributable to short duration of WCH and a low degree of arterial stiffness. Moreover, only 2 studies containing 164 juveniles were analysed, which might not have fully represented the real situations of juveniles. More studies are needed to reveal the relationship between WCH and arterial stiffness in the juvenile population.

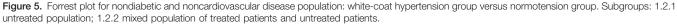
During literature screening, some studies were identified which used AASI and augmentation index to quantify arterial

stiffness.^[29,30] These accessory examinations have been accepted by clinical practitioners. In particular, AASI calculated from ambulatory blood pressure monitoring is easy to use. Nevertheless, PWV, as the golden standard of arterial stiffness, has a markedly higher diagnostic accuracy than other indexes. Metaanalysis included clinical studies which had employed PWV as an examination method to best show the relationship between WCH and arterial stiffness. Recently, Upala et al^[31] published another meta-analysis about the relationship between WCH and arterial stiffness, but they reported no significant association between WCH and arterial hypertension on the basis of 4 eligible observational studies containing persistent hypertension group, WCH group and normal control group. In our opinion, due to their inclusion methods, they might have excluded many casecontrol studies which only contained WCH group and normal control group thus the study did not sufficiently reveal the relationship between WCH and arterial stiffness. Based on prespecified search strategies and inclusion criteria, Medline, Embase, Web Of Science, Cochrane Library, and BIOSIS Preview databases were reviewed. A total of 20 studies involving 1538 WCH patients and 3582 normotensives were included in our study, which would better reflect the effect of WCH on arterial stiffness.

| | Age (y) | | Antihypertens | ive drug users | Instrum | ent for inspect | ting PWV | Study | Quality | History of DM and (or) CVD | |
|-------------------|--------------------|-----------------|---------------|----------------|--------------------|----------------------|-----------|-----------|-----------|----------------------------|-----------|
| Subgroups | Juveniles (<18) | Adults (≥18) | No | Mixed | Complior device | SphygmoCor device | Others | Medium | High | No | Yes |
| Number of studies | 2 | 18 | 10 | 9 | 7 | 4 | 9 | 3 | 17 | 9 | 11 |
| P value | .253 | <.001 | <.001 | .001 | <.001 | .001 | <.001 | <.001 | <.001 | <.001 | <.001 |
| P value | 23.4% | 85.4% | 87.9% | 70.2% | 90.3% | 82.4% | 81.0% | 85% | 87.4% | 45% | 87% |
| 95% <i>Cl</i> | -0.39 to 0.61 | 0.46-0.87 | 0.12-0.74 | 0.61 - 1.05 | 0.29-1.08 | 0.23 - 1.09 | 0.27-0.86 | 0.35-1.05 | 0.52-0.66 | 0.33-0.78 | 0.76-1.57 |

CVD = cardiovascular disease, DM = diabetes mellitus, NT = normotensives, PWV = pulse wave velocity, WCH = white-coat hypertension.

| WCH | | | | | NT | | | Mean Difference | | Mean Difference | | |
|--------------------------------------|------------|--------|---------|-------------|--------|-------|--------|---------------------|------|----------------------------|--|--|
| Study or Subgroup | Mean | SD | Total | Mean | SD | Total | Weight | IV, Fixed, 95% C | Year | IV, Fixed, 95% CI | | |
| 1.2.1 Untreated popula | ation | | | | | | | | | | | |
| Ribeiro 1 2000 | 9.3 | 1.68 | 28 | 9.2 | 1 | 19 | 4.0% | 0.10 [-0.67, 0.87] | 2000 | | | |
| Schillaci 2011 | 9.3 | 2 | 133 | 8.5 | 2 | 71 | 7.0% | 0.80 [0.22, 1.38] | 2011 | | | |
| Sozeri 2012 | 5.6 | 0.61 | 8 | 5.3 | 0.7 | 100 | 11.8% | 0.30 [-0.14, 0.74] | 2012 | + | | |
| Jurko 2014 | 7.7 | 2.5 | 28 | 7.9 | 2 | 28 | 1.7% | -0.20 [-1.39, 0.99] | 2014 | | | |
| Almeida 2016 | 9.7 | 2.4 | 315 | 9.5 | 2 | 175 | 14.8% | 0.20 [-0.20, 0.60] | 2016 | | | |
| Subtotal (95% CI) | | | 512 | | | 393 | 39.3% | 0.31 [0.07, 0.55] | | • | | |
| Heterogeneity: Chi ² = 4. | 07, df = | 4 (P = | 0.40); | l² = 2% | | | | | | | | |
| Test for overall effect: Z | = 2.50 (| P = 0. | 01) | | | | | | | | | |
| 4.0.0 Mins I and a latin | | | | | | | | | | | | |
| 1.2.2 Mixed population | | | | | | | | | | | | |
| Silva 2004 | 9.9 | | 87 | | 1.45 | 132 | 15.2% | 1.00 [0.61, 1.39] | | | | |
| Andrikou 2011 | 7.5 | 1.2 | 81 | 6.8 | | 44 | 25.9% | 0.70 [0.40, 1.00] | | | | |
| Chatzistamatiou 2015 | 8.1 | 1.49 | 130 | | 1.49 | 143 | 18.6% | 0.60 [0.25, 0.95] | | | | |
| Afsar 2015 | 8.52 | 2.5 | 11 | 7.23 | 1.7 | 109 | 1.0% | 1.29 [-0.22, 2.80] | 2015 | | | |
| Subtotal (95% CI) | | | 309 | | | 428 | 60.7% | 0.75 [0.56, 0.95] | | • | | |
| Heterogeneity: Chi ² = 2. | 84, df = | 3 (P = | 0.42); | $ ^2 = 0\%$ | | | | | | | | |
| Test for overall effect: Z | = 7.54 (| P < 0. | 00001) | | | | | | | | | |
| Total (95% CI) | | | 821 | | | 821 | 100.0% | 0.58 [0.43, 0.73] | | • | | |
| Heterogeneity: Chi ² = 14 | 4.63, df = | = 8 (P | = 0.07) | : ² = 45 | % | | | | | | | |
| Test for overall effect: Z | | | | | 10,000 | | | | | -2 -1 0 1 2 | | |
| | | | | | | | | | | Favours [NT] Favours [WCH] | | |



A limitation of this meta-analysis is that we had no individual patient data, only the literature data can be combined and analyzed. A further limitation was that the heterogeneity for overall analysis was noticeable (I2 = 82%), so the meta-regression analysis and subgroup analysis were performed. Subgroup analysis is the highlight of meta-analysis, especially that of the patients without diabetes mellitus or cardiovascular diseases. Maine-Syracuse case-control study has demonstrated the significant relationship between type-2 diabetes mellitus (especially uncontrolled type-2 diabetes mellitus) and arterial stiffness. Previous studies have proven the close relationship between cardiovascular diseases (e.g., coronary artery disease) and arterial stiffness.^[32,33] Therefore, subgroup analysis for the patients without diabetes mellitus or cardiovascular diseases was conducted. The results showed significantly reduced heterogeneity in the eligible studies, which was further reduced by the secondary subgroup analysis stratified by history of antihypertensive drug use. In this way, subgroup analyses identified the relationship between WCH and arterial stiffness. By stepwise subgroup analyses, the eligible criteria was gradually narrowed to reduce the heterogeneity and to enhance the reliability of study results. Meta-regression analysis also identified diabetes mellitus and cardiovascular diseases as important sources of overall heterogeneity.

This study showed that WCH may cause arterial stiffness in adult population. This kind of mechanisms may help uncover the multiple target organ damages in the future. WCH is common in clinical practice, but its pathophysiological mechanisms and target organ damages remain unclear. As a result, many clinicians are confused about its diagnosis and treatments. Based on these study findings, more attention is to be given to the role of WCH in cardio-cerebrovascular target organ damages, and reasonable diagnostic and therapeutic standards of WCH should be further explored.

Author contributions

Formal analysis: Yan Wang. Funding acquisition: Xukai Wang. Software: Peng Cai, Yan Peng. Writing – review & editing: Peng Cai, Xukai Wang.

References

- Huang YQ, Jie LI, Chen JY, et al. The relationship between soluble CD40 ligand level and atherosclerosis in white-coat hypertension. J Hum Hypertens 2017;32:40–5.
- [2] Filipovsky J. [White-coat hypertension and masked hypertension]. Vnitr Lek 2015;61:401–5.
- [3] Cuspidi C, Tadic M, Mancia G, et al. White-coat hypertension: the neglected subgroup in hypertension. Korean Circ J 2018;48:552–64.
- [4] Litwin M. Why should we screen for arterial hypertension in children and adolescents? Pediatr Nephrol 2018;33:83–92.
- [5] Kent ST, Burkholder GA, Tajeu GS, et al. Mechanisms influencing circadian blood pressure patterns among individuals with HIV. Curr Hypertens Rep 2015;17:88.
- [6] Huang YL, Huang WJ, Mai WY, et al. White-coat hypertension is a risk factor for cardiovascular diseases and total mortality. J Hypertens 2017;35:677–88.
- [7] Loehr LR, Meyer ML, Poon AK, et al. Prediabetes and diabetes are associated with arterial stiffness in older adults: the ARIC study. Am J Hypertens 2016;29:1038–45.
- [8] Courand PY, Harbaoui B, Serraille M, et al. Ruling out white coat hypertension with NT-proBNP: a new paradigm away from blood pressure assessment. Int J Cardiol 2016;207:57–8.
- [9] Rissanen R, Berg HY, Hasselberg M. Quality of life following road traffic injury: a systematic literature review. Accid Anal Prev 2017;108:308–20.
- [10] Ribeiro L, Gama G, Santos A, et al. Arterial distensibility in subjects with white-coat hypertension with and without diabetes or dyslipidaemia:

comparison with normotensives and sustained hypertensives. Blood Press Monit 2000;5:11-7.

- [11] Silva JA, Barbosa L, Bertoquini S, et al. Relationship between aortic stiffness and cardiovascular risk factors in a population of normotensives, white-coat normotensives, white-coat hypertensives, sustained hypertensives and diabetic patients. Rev Port Cardiol 2004; 23:1533–47.
- [12] Stolarz-Skrzypek K, Olszanecka A, Lubaszewski W, et al. Left ventricular mass and pulse wave velocity in patients with masked hypertension and white-coat hypertension. Nadcisnienie Tetnicze 2008;12:80–6.
- [13] Andrikou I, Tsioufis C, Dimitriadis K, et al. Similar levels of low-grade inflammation and arterial stiffness in masked and white-coat hypertension: comparisons with sustained hypertension and normotension. Blood Press Monit 2011;16:218–23.
- [14] Schillaci G, Pucci G, Pirro M, et al. Combined effects of office and 24-h blood pressure on aortic stiffness in human hypertension. J Hypertens 2011;29:869–75.
- [15] Martin CA, Cameron JD, Chen SS, et al. Two hour glucose post loading: a biomarker of cardiovascular risk in isolated clinic hypertension. J Hypertens 2011;29:749–57.
- [16] Sozeri B, Ozdemir Y, Deveci M, et al. The effects of white coat hypertension in children. Pediatr Nephrol 2012;27:1735.
- [17] Hopkins S, Ramsay I, Williams S, et al. White coat hypertension is common in HIV-positive individuals and not associated with markers of increased vascular aging. HIV Med 2013;14:40–1.
- [18] Protogerou AD, Panagiotakos DB, Zampeli E, et al. Arterial hypertension assessed " out-of-office" in a contemporary cohort of rheumatoid arthritis patients free of cardiovascular disease is characterized by high prevalence, low awareness, poor control and increased vascular damageassociated " white coat" phenomenon. Arthritis Res Ther 2013;15: R142.
- [19] Sung SH, Cheng HM, Wang KL, et al. White coat hypertension is more risky than prehypertension: important role of arterial wave reflections. Hypertension 2013;61:1346–53.
- [20] Jurko A, Tonhajzerova I, Jurko T, et al. Arterial stiffness in children with essential hypertension and white coat hypertension. European J Preventive Cardiol 2014;21:S120.
- [21] Chatzistamatiou E, Moustakas G, Trachanas C, et al. Aortic stiffness and essential hypertension phenotypes. Euro Heart J 2015;36:854.

- [22] Afsar B, Elsurer R, Kirkpantur A, et al. Urinary sodium excretion and ambulatory blood pressure findings in patients with hypertension. J Clin Hypertens (Greenwich) 2015;17:200–6.
- [23] Almeida J, Monteiro J, Silva JA, et al. Central pressures and central hemodynamic values in white coat hypertensives are closer to those of normotensives than to those of controlled hypertensives for similar age, gender, and 24-h and nocturnal blood pressures. Rev Port Cardiol 2016;35:559–67.
- [24] Scuteri A, Morrell CH, Orru M, et al. Gender specific profiles of white coat and masked hypertension impacts on arterial structure and function in the SardiNIA study. Int J Cardiol 2016;217:92–8.
- [25] Wojciechowska W, Stolarz-Skrzypek K, Olszanecka A, et al. Subclinical arterial and cardiac damage in white-coat and masked hypertension. Blood Press 2016;25:249–56.
- [26] Barochiner J, Aparicio LS, Alfie J, et al. Arterial stiffness in treated hypertensive patients with white-coat hypertension. J Clin Hypertens 2017;19:6–10.
- [27] Nemcsik J, Korosi B, Batta D, et al. Evaluation of affective temperaments, depression and anxiety in white-coat, well-treated and resistant hypertension and in healthy controls. J Hypertens 2017;35:e209.
- [28] Androulakis E, Papageorgiou N, Lioudaki E, et al. Subclinical organ damage in white-coat hypertension: the possible role of cystatin C. J Clin Hypertens (Greenwich) 2017;19:190–7.
- [29] Mestanik M, Jurko A, Mestanikova A, et al. Arterial stiffness evaluated by cardio-ankle vascular index (CAVI) in adolescent hypertension. Can J Physiol Pharmacol 2016;94:112–6.
- [30] Fukushima T, Eguchi K, Ohkuchi A, et al. Changes in central hemodynamics in women with hypertensive pregnancy between before and after delivery. J Clin Hypertens (Greenwich) 2016;18:329–36.
- [31] Upala S, Sanguankeo A. Association between white-coat hypertension and arterial stiffness: a systematic review and meta-analysis. Hypertension 2016;68:216.
- [32] Vriz O, Bertin N, Ius A, et al. Carotid artery stiffness and development of hypertension in people with paraplegia and no overt cardiovascular disease: a 7-year follow-up study. J Cardiovasc Echogr 2017;27:132–40.
- [33] Cai P, Peng Y, Li L, et al. Fibroblast growth factor 23 (FGF23) gene polymorphisms are associated with essential hypertension risk and blood pressure levels in Chinese Han population. Clin Exp Hypertens 2018;680–5.