RESEARCH ARTICLE



Progressive effects of single-nucleotide polymorphisms on 16 phenotypic traits based on longitudinal data

Donghe Li¹ · Hahn Kang² · Sanghun Lee³ · Sungho Won^{1,4,5}

Received: 16 November 2019 / Accepted: 2 December 2019 / Published online: 4 January 2020 © The Author(s) 2020

Abstract

Background There are many research studies have estimated the heritability of phenotypic traits, but few have considered longitudinal changes in several phenotypic traits together.

Objective To evaluate the progressive effect of single nucleotide polymorphisms (SNPs) on prominent health-related phenotypic traits by determining SNP-based heritability (h_{snp}^2) using longitudinal data.

Methods Sixteen phenotypic traits associated with major health indices were observed biennially for 6843 individuals with 10-year follow-up in a Korean community-based cohort. Average SNP heritability and longitudinal changes in the total period were estimated using a two-stage model. Average and periodic differences for each subject were considered responses to estimate SNP heritability. Furthermore, a genome-wide association study (GWAS) was performed for significant SNPs. **Results** Each SNP heritability for the phenotypic mean of all sixteen traits through 6 periods (baseline and five follow-ups) were significant. Gradually, the forced vital capacity in one second (FEV1) reflected the only significant SNP heritability among longitudinal changes at a false discovery rate (FDR)-adjusted 0.05 significance level ($h_{snp}^2 = 0.171$, FDR = 0.0012). On estimating chromosomal heritability, chromosome 2 displayed the highest heritability upon periodic changes in FEV1. SNPs including rs2272402 and rs7209788 displayed a genome-wide significant association with longitudinal changes in FEV1 ($P = 1.22 \times 10^{-8}$ for rs2272402 and $P = 3.36 \times 10^{-7}$ for rs7209788). De novo variants including rs4922117 (near *LPL*, $P = 2.13 \times 10^{-15}$) of log-transformed high-density lipoprotein (HDL) ratios and rs2335418 (near *HMGCR*, $P = 3.2 \times 10^{-9}$) of low-density lipoprotein were detected on GWAS.

Conclusion Significant genetic effects on longitudinal changes in FEV1 among the middle-aged general population and chromosome 2 account for most of the genetic variance.

Keywords Heritability · Phenotypic trait · Genomic restricted maximum likelihood · Longitudinal changes

Electronic supplementary material The online version of this article (https://doi.org/10.1007/s13258-019-00902-x) contains supplementary material, which is available to authorized users.

Sanghun Lee integrative@korea.com

- Sungho Won won1@snu.ac.kr
- ¹ Interdisciplinary Program in Bioinformatics, Seoul National University, Seoul, Republic of Korea
- ² Biology Department, Morrissey College of Arts and Sciences, Boston College, Boston, MA, USA

Introduction

Single nucleotide polymorphism (SNP)-based heritability (h_{snp}^2) indicates the relative proportion of genetic variance explained based on SNPs used for genome-wide association studies (GWASs). For h_{snp}^2 estimation, the genomic restricted maximum likelihood (GREML) for

- ³ Department of Medical Consilience, Graduate School, Dankook University, Yongin, Republic of Korea
- ⁴ Institute of Health and Environment, Seoul National University, Seoul, Republic of Korea
- ⁵ Department of Public Health Science, Graduate School of Public Health, Seoul National University, 1 Kwanak-ro Kwanak-gu, Seoul 151-742, Republic of Korea

linear mixed models (LMMs) is often implemented in the genetic complex trait analysis (GCTA) tool (Yang et al. 2011). GREML first calculates the genetic relatedness matrices (GRM), which are used as variance-covariance matrices for random effects. The significance of estimates obtained through GREML depends on the study design; if it is applied to family-based samples, it displays pedigreebased heritability, but for unrelated subjects, it estimates h_{snp}^2 (Kim et al. 2015; Yang et al. 2017). Estimation of h_{snp}^2 involves considerable differences across not only methodologies but also procedures requiring careful interpretation of results (Evans et al. 2018; Ni et al. 2018). Moreover, the estimated heritability is potentially biased and misleading owing to measurement errors at various degrees. To overcome these challenges, the heritability determined from longitudinal data is more reliable than that determined from cross-sectional data. While most studies on h_{snn}^2 focused on the primary effect of SNPs, significant effects of SNPs on the average annual differences indicate the SNP-by-age interaction. Numerous examples illustrate the importance of age on longitudinal changes (Genetic epidemiologic studies on age-specified traits. NIA Aging and Genetic Epidemiology Working Group 2000; Nishimura et al. 2012; van de Pol and Verhulst 2006). For instance, one study showed that an annual decline in lung function is associated with age (Kim et al. 2016), and another study reported a genetic influence on changes in both lipoprotein risk factors and systolic blood pressure over a decade (Friedlander et al. 1997). Therefore, the h_{snp}^2 should be estimated on the basis of not only the mean of observed traits but also changes in the sufficient period. Hence, we applied a two-stage approach, which is a convenient method of analyzing longitudinal data by combining linear regression models to investigate the effect of SNPs on both average and longitudinal differences in phenotypic traits.

In this study, we investigated the magnitude of SNPs effect on average and longitudinal differences by using both genomic data and 16 phenotypic traits associated with major health indices using a phenotype-genotype dataset of unrelated individuals in a community-based cohort and evaluated their importance. Except for baseline, each phenotype was objectively measured every 2 years for 10-year follow-up, and six repeated measurements (maximum) were obtained for each individual. For each subject, both the average phenotypic traits and their longitudinal changes were estimated via subject-specific regression analysis, using intercepts and coefficients of ages, respectively. Each h_{sup}^2 value was estimated using GCTA. Our results show that lung function has the only significant h_{snp}^2 for longitudinal changes, while all average phenotypes of the 16 traits yielded a significant h_{snp}^2 value. Furthermore, the GWAS revealed certain novel genome-wide significant SNPs associated with the phenotypes analyzed herein.

Materials and Methods

Ethics approval and consent to participate

The respective Institutional Review Board (IRB) of Seoul National University reviewed and approved the informed consent, the study protocols and other documents (Permit No. E1605/002-003). All methods were performed in accordance with the relevant guidelines and regulations.

Korea Associated Resource (KARE) cohort data

Korea Associated Resource (KARE) data are based on a community-based epidemiological study and comprises subjects residing in Ansan (urban area) and Ansung (rural area) in the Gyeonggi Province of South Korea (Cho et al. 2009). A baseline survey was completed in 2001–2002, and 10,030 participants aged 40-69 years were recruited. After that, biennial repeated surveys were conducted, and the last survey were completed in 2013-2014 (Kim et al. 2017). Six different surveys were conducted in total. These measurement periods are indicated as periods 1-6 throughout this study, each with a different number of subjects (e.g., period 1: 8543 subjects [4052 male, 4491 female]; period 6: 5391 subjects [2502 male, 2889 female]). The number of overlapping subjects throughout the six periods was 4306 (2009 male, 2297 female). Among these, subjects whose traits were measured at least three times were considered, and 6843 participants (3273 male, 3570 female) were assessed in total.

Many participant phenotypes were recorded by trained interviewers through questionnaires and clinical measurement, and we only considered the 16 quantitative traits that they were measured objectively and associated with major health indices; these were classified into four groups: anthropometric, biochemistry, cardiopulmonary, and red blood cell traits (Table 1). As glycated hemoglobin (HbA1c), fasting blood glucose (GLU0), high-density lipoprotein (HDL), triglycerides (TG), and systolic blood pressure (SBP) displayed skewed distributions, they were log-transformed and denoted by log(HbA1c), log(GLU0), log(HDL), log(TG), and log(SBP), respectively. The missing rate of HbA1c was larger than 0.5 at period 2 and was excluded from further analyses. For each trait, subjects with more than three measurement observations were assessed.

Genotypes

Genotype data for the KARE cohort were obtained using the Affymetrix Genome-Wide Human SNP array 5.0 (Cho et al. 2009). Quality control (QC) analysis of SNPs and subjects was conducted using PLINK (Purcell et al. 2007)

Anthropomorphic traits	Height, Waist, Weight, body-mass index (BMI)
Biochemistry traits	
Glucose	Glycated hemoglobin (HbA1c), Fasting blood glucose (GLU0)
Cholesterol	Low-density lipoprotein (LDL), high-density lipoprotein (HDL), total cholesterol (TCHL), triglyceride (TG)
Cardiopulmonary traits	
Blood pressure	Systolic blood pressure (SBP), diastolic blood pressure (DBP)
Lung capacity	Predicted forced vital capacity (FVC) %, predicted forced expiratory volume in one second (FEV1) %, predicted FEV1/FVC %
Red blood cell traits	Hemoglobin levels (Hb)

Table 1 Sixteen phenotypic traits associated with major health indices

and ONETOOL (Song et al. 2018). SNPs with *P*-values from Hardy–Weinberg equilibrium (HWE) analysis $< 10^{-5}$, minor allele frequencies (MAFs) < 0.05, and genotype call rates < 95% were excluded. Furthermore, subjects with missing genotype call rates > 5% or sex-based inconsistencies were excluded. The missing genotypes for typed SNPs were imputed on the basis of the 1000-genome sequence reference data. After QC analysis, 305,158 SNPs were analyzed for SNP heritability and GWAS (Fig. 1).

Data availability

The data used in this study underwent an application process according to the Korean Genome and Epidemiology study data access policy, and can be downloaded from following website: http://nih.go.kr/index.es?sid=a5#a.

Determination of phenotypic averages and longitudinal changes in each subject

The phenotypic averages and longitudinal changes for each subject were determined and used to estimate SNP heritability and for the GWAS. Significant differences in phenotypic variances were observed for each period, and such heteroscedasticity was considered for phenotypic averages and longitudinal changes for each subject as follows. First, the linear regression model for subjects of the same period was adjusted for traits. The effect of sex, age, and top 10 principal component (PC) scores estimated from the genetic relationship matrix were adjusted as covariates. Considering w_{jk} as the residual variances of trait k (k=1, ..., 16) at the period j, the following linear regression model was adjusted with repeated measures for each subject i as follows:



$$y_{ijk} = \beta_{ik0} + \beta_{ik1} \left(age_{ij} - \overline{age_i} \right) + \epsilon_{ijk}, \epsilon_{ijk} \sim N\left(0, \frac{1}{w_{jk}}\sigma_{ik}^2\right)$$
(1)

Here, *i* indicates the *i*th subject, and $\overline{age_i}$ indicates the mean of ages at the observed time points. In the regression model (1), y_{iik} is the *j*th period observed value of subject *i* for trait k, β_{ik0} indicates the expected phenotypic mean of subject *i* for trait k when he or she is $\overline{age_i}$ years old, and β_{ik1} is the average longitudinal change in trait k. The estimated values of β_{ik0} and β_{ik1} were used to estimate the heritability and for GWAS analysis. For convenience, both are denoted by B_0 and B_1 , respectively.

Estimation of heritability

 B_0 and B_1 were separately used to estimate SNP heritability. Analyses were conducted using GCTA (Yang et al. 2011), and the restricted maximum likelihood estimator was used. Effects of age and sex were adjusted as covariates.

GWAS analysis

 B_0 and B_1 were separately analyzed to identify disease susceptibility loci for 16 different traits. The $\overline{age_i}$, sex, and 10 PC scores estimated from the genetic relationship matrix were included as covariates to adjust the population stratification.

Results

Estimation of heritability

A schematic representation of the heritability analysis and genome-wide association study is shown in Fig. 1. For 16 different traits of 6843 subjects, the mean and standard deviation values of each trait at period 1 are shown in Table 2 (see Table 1 for detailed information). Some missing values resulted in differences in the total number of subjects depending on the phenotype, and the sample sizes of those traits and descriptive statistics including sex and age were summarized.

For those 6843 subjects, a multidimensional scaling (MDS) plot was generated (Fig. 2). As shown in Fig. 2, subjects from the 1000 Genomes Project were also included, and our analyses were not affected by population stratification.

We calculated the descriptive statistics for B_0 and B_1 (Table 3, see "Materials and methods" for details). B_0 in Eq. (1) indicates the means of the predicted traits at $\overline{age_i}$ years of age. B_1 stands for the longitudinal changes in the traits of each subject. Table 3 shows that the means of B_0 are similar to those of period 1. Means of B_1 were generally closer to 0. Figure 3 shows the estimates of heritability with B_0 as the response in the GREML model, and the estimated heritability of height for the data peaked at 0.318 (P = 1.665 $\times 10^{-16}$, FDR = 2.664 $\times 10^{-15}$). The subsequent three highest heritability traits were total cholesterol (TCHL), log(HDL), and low-density lipoprotein (LDL), with values of 0.265 $(P=3.895 \times 10^{-12}, FDR=3.116 \times 10^{-11}), 0.241 (P=8.911)$ $\times 10^{-10}$, FDR = 4.753 $\times 10^{-9}$), and 0.222 (P = 5.178 $\times 10^{-9}$,

Table 2 Descriptive statistics of 16 traits	Trait	Trait (bas	seline)	Total (N)	Female		Age	
To traits		Mean	SD		N	%	Mean	SD
	Height (cm)	160.11	8.63	6823	3557	52.13%	51.90	8.69
	Waist (cm)	82.63	8.70	6835	3567	52.19%	51.90	8.69
	Weight (kg)	63.24	10.10	6822	3556	52.13%	51.90	8.69
	BMI (kg/m ²)	24.62	3.10	6822	3556	52.13%	51.90	8.69
	HbA1c (%)	5.74	0.82	6329	3321	52.47%	51.87	8.62
	GLU0 (mg/dl)	86.73	19.41	6728	3514	52.23%	51.85	8.67
	TG (mg/dl)	161.47	103.19	6840	3568	52.16%	51.91	8.70
	LDL (mg/dl)	115.00	32.89	6840	3568	52.16%	51.91	8.70
	HDL (mg/dl)	44.69	9.91	6840	3568	52.16%	51.91	8.70
	TCHL (mg/dl)	191.92	35.09	6840	3568	52.16%	51.91	8.70
	SBP (mmHg)	121.12	18.10	6843	3570	52.17%	51.91	8.70
	DBP (mmHg)	80.19	11.33	6843	3570	52.17%	51.91	8.70
	Hb (g/dl)	13.61	1.57	6840	3568	52.16%	51.91	8.70
	FVC (%predicted)	104.76	14.17	4291	2135	49.76%	50.37	8.17
	FEV1 (%predicted)	112.27	16.62	4290	2134	49.74%	50.37	8.16
	FEV1/FVC (predicted)	74.89	1.77	4291	2135	49.76%	50.37	8.17



Fig. 2 Population structures identified via a multidimensional scaling (MDS) plot. This plot shows that our analyses (KARE) are not affected by population stratification. AFR, AMR, EAS, EUR, and SAS indicate African, Ad Mixed American, East Asian, European, and South Asian populations, respectively, from the 1000 Genomes Project

FDR = 1.657×10^{-8}), respectively. These three traits are cholesterol-related. The heritability of waist was 0.218 ($P = 5.016 \times 10^{-9}$, FDR = 1.657×10^{-8}) and that of weight was 0.196 ($P = 2.046 \times 10^{-7}$, FDR = 5.456×10^{-7}). The heritability was 0.195 for Hb ($P = 4.926 \times 10^{-7}$, FDR = 9.852×10^{-7}) and 0.192 for log(TG) ($P = 4.419 \times 10^{-7}$, FDR = 9.852×10^{-7}). The heritability of the other traits with an FDR larger than 1×10^{-6} were less than 0.19.

397

Figure 4 shows the heritability estimates for B_1 , which are generally less than those for B_0 , and we found that lung function traits FVC and FEV1, wasit, diastolic blood pressure (DBP), BMI, and log(SBP) are relatively high. The highest heritability estimate was observed for FEV1 (0.171) and its FDR-adjusted P value was 0.0189. The heritability estimates of other traits were less than 0.1. The second highest heritability was 0.0941 for FVC, and its FDR-adjusted P value was 0.166. The heritability of waist was also relatively higher than that of other traits. Waist heritability and the FDR-adjusted P values were 0.0082 and 0.0657, respectively. The higher heritability estimates for B_1 indicate that the decreasing/increasing rates were associated with genetic factors. Height displayed the highest heritability estimates for B_0 , but its estimate for B_1 was low (0.0297). Height does not usually change after the age of 20, and its lower value here is probably attributable to it. For the other traits including log(HbA1c), LDL, log(HDL), TCHL, and Hb levels, SNP heritability estimates tended towards 0.

Furthermore, we determined chromosomal heritability estimates of FEV1, which displayed the highest heritability in the B_1 model. Consequently, chromosome 2 accounted for the highest proportion of phenotypic variance $(h_{snp}^2 = 0.0397)$, albeit with a high standard error (Fig. 5). There was a significant positive correlation between chromosome length and heritability (r=0.58, P=0.0045) in FEV1 (Fig. 6).

Genome-wide association studies

 B_0 and B_1 were considered responses for GWAS. Tables 4 and 5 show genome-wide significant SNPs at a significance

Trait	B_0				B_1	<i>B</i> ₁					
	Mean	SD	Min	Max	Mean	SD	Min	Max			
Height	159.906	8.724	130.241	187.866	-0.060	0.139	-2.168	0.747			
Waist	83.743	8.480	58.333	121.591	0.184	0.692	-4.968	6.904			
Weight	62.860	9.931	30.532	105.355	-0.094	0.480	-3.739	2.657			
BMI	24.531	2.992	14.197	38.831	-0.019	0.185	-1.486	1.048			
log(HbA1c)	1.737	0.107	1.256	2.441	0.002	0.011	-0.093	0.157			
log(GLU0)	4.538	0.149	4.260	5.733	0.012	0.018	-0.166	0.171			
log(TG)	4.834	0.423	3.584	7.189	-0.012	0.057	-0.409	0.412			
LDL	120.111	25.757	11.833	281.590	0.193	4.065	-29.218	28.549			
log(HDL)	3.782	0.193	3.100	4.567	-0.001	0.024	-0.211	0.135			
TCHL	194.208	28.114	97.986	343.106	-0.120	4.468	- 34.599	29.468			
log(SBP)	4.768	0.114	4.461	5.156	-0.001	0.017	-0.122	0.086			
DBP	78.252	8.239	50.639	111.556	-0.259	1.358	-12.330	8.066			
Hb	13.695	1.370	7.764	18.889	0.022	0.147	-1.641	1.468			
FVC	104.541	13.466	46.629	162.844	-0.090	2.478	-13.065	13.685			
FEV1	111.128	16.295	38.951	184.532	-0.239	2.575	-16.022	15.620			
FEV1/FVC	73.945	1.809	67.654	78.000	-0.213	0.127	-1.246	1.244			

Table 3Summary of B_0 and B_1 of 16 traits



Fig. 3 Single-nucleotide polymorphism heritability estimates of 16 traits with B_0 as the response. Error bars correspond to standard error values. The values above the error bar are *P* values and false discovery rate (FDR; bold)



Fig. 4 Single-nucleotide polymorphism heritability estimates of 16 traits with B_1 as the response. Error bars correspond to standard error values. The values above the error bar are *P* values and the false discovery rate (FDR; bold), and "*" indicates significant findings at an FDR of 0.05

level of 1×10^{-7} . Subsequent findings are summarized in the Supplementary material (Tables S1 and S2). Table 4 shows that SNPs have relatively lower *P* values for log(TG) and log(HDL) than any other trait. The most significant variant for log(TG) was rs6589566 in *ZPR1* with a *P*-value of 7.9×10^{-38} , the lowest *P* value among all 16 traits. Furthermore, *ZPR1* is associated with TG (Coram et al. 2013). The most significant variant of log(HDL) was rs16940212 with a *P* value of 2.08×10^{-18} in *ALDH1A2*, which is associated with HDL (Spracklen et al. 2017). Certain other significant variants were significantly associated with proximal genes and with traits assessed herein. The variant rs180349 ($P = 8.86 \times 10^{-35}$) of log(TG) was proximal to *BUD13*, which is associated with TG (Hoffmann et al. 2018). The variant rs17482753 ($P = 3.199 \times 10^{-18}$) is proximal to *LPL*, which was strongly associated with HDL (Hoffmann et al. 2018). Herein, we also detected some de novo variants including rs4922117 ($P = 2.13 \times 10^{-15}$) of log(HDL) and rs2335418 ($P = 3.2 \times 10^{-9}$) of LDL, which were previously unknown; however, both their proximal genes *LPL* and *HMGCR* are significantly associated with each trait (Hoffmann et al. 2018). The Manhattan Plot and



Fig. 5 Single-nucleotide polymorphism heritability estimates of FEV1 based on chromosomes with B_1 as the response. Error bars correspond to standard error values. The values above the error bar are *P* values and false discovery rate (FDR; bold)

Fig. 6 Correlation between chromosome length and chromosome-specific heritability. Numbers near dots are the chromosome numbers. Black line is the estimated regression line



QQ plot for the model with B_0 as the response are provided in the Supplementary material (Figures S1 and S2).

Table 5 shows the results of GWAS for B_1 . Based on the results, rs2272402 (*SLC6A1*) was the most significant variant both in FEV1 ($P = 1.22 \times 10^{-8}$) and FVC ($P = 1.40 \times 10^{-9}$) (Regional plots are shown in Supplementary Figures S5 and S7), and the *SLC6A1* enhancer was associated with lung function. Other variants including rs7209788 (*NARF*, $P = 3.36 \times 10^{-7}$) for FEV1 and rs2668162 (*FAM19A1*, $P = 6.18 \times 10^{-7}$) for FVC had *P*-values less than the 1 × 10^{-6} threshold (S2 Table). From the regional plot (Supplementary Figure S6), we found that rs4789777(*HEXDC*, $P = 4.599 \times 10^{-6}$) is highly correlated with rs7209788 of FEV1. The Manhattan Plot and QQ plot for the model with B_1 as the response are provided in the Supplementary material (Figures S3 and S4).

Discussion

In this study, SNP-based heritability estimates of 16 phenotypic traits were estimated longitudinal data from a 10-year follow-up of the KARE cohort. The GCTA tool was used with a two-stage approach to determine the heritability estimate of phenotypic mean and longitudinal changes in each trait. Moreover, chromosomal heritability estimates were determined and GWAS analyses were performed using the same approach. Overall, heritability estimates within the population-based cohort including KARE are potentially lower than those of pedigree or twin studies for all 16 traits, regardless of whether the response is B_0 that phenotypic mean of traits or B_1 which stands for the changes by time of traits. For example, the heritability of height herein was estimated to be approximately 0.318

	0											
Trait	SNP	CHR	BPI	BP2 A	1	A2 G	Ь	GENE	MAF	HWE_P	BETA	Ρ
log(GLU0)	rs1799884	7	44,229,068	44,229,068 A	-	D D	pstream	GCK	0.1872	0.9051	0.01889	5.62E-09
log(GLU0)	rs7754840	9	20,661,250	20,661,250 C	Ū	ц Г	tronic	CDKALI	0.478	1	0.01488	6.25E-09
Чb	rs5756505	22	37,467,354	37,467,354 C	Ū	LI LI	tronic	TMPRSS6	0.4979	0.9229	0.1125	2.61E-13
Hb	rs3768751	7	46,346,716	46,346,716 C	-	A Ir	Itronic	PRKCE	0.1796	1	-0.113	1.79E - 08
log(HbA1c)	rs7754840	9	20,661,250	20,661,250 C		LI LI	Itronic	CDKAL1	0.478	1	0.01254	5.41E-11
log(HDL)	rs16940212	15	58,694,020	58,694,020 T		LI LI	Itergenic	ALDH1A2	0.3414	0.9786	0.02988	2.08E-18
log(HDL)	rs17482753	8	19,832,646	19,832,646 T		LI LI	Itergenic	LPL(dist = 7876),SLC18A1(dist = 169,720)	0.1241	0.1333	0.04232	3.20E-18
log(HDL)	rs4922117	8	19,852,586	19,852,586 G		4 Ir	Itergenic	LPL(dist=27,816),SLC18A1(dist=149,780)	0.2077	0.418	0.0316	2.13E-15
LDL	rs599839	-	109,822,166	109,822,166 C		D	ownstream	PSRC1	0.06456	0.1925	-5.886	1.53E-11
LDL	rs12654264	5	74,648,603	74,648,603 T		≜ Ir	tronic	HMGCR	0.4758	0.8085	- 2.625	1.41E - 09
LDL	rs2335418	5	74,603,479	74,603,479 C		≜ Ir	Itergenic	ANKRD31 (dist = 70, 776), HMGCR (dist = 29, 514)	0.4232	0.5689	-2.6	3.20E-09
LDL	rs4045166	5	74,909,446	74,909,446 C		Ч	tronic	ANKDD1B	0.3326	0.3137	2.727	3.55E-09
LDL	rs10942739	5	74,786,083	74,786,083 T		Ч	tronic	COL4A3BP	0.3325	0.276	2.709	4.61E-09
LDL	rs688	19	11,227,602	11,227,602 T	Ŭ	ш С	xonic	LDLR	0.136	0.797	3.402	6.63E-08
TCHL	rs599839	1	109,822,166	109,822,166 G	_	D	ownstream]	PSRC1	0.06456	0.1925	-6.822	1.19E-12
TCHL	rs780092	7	27,743,154	27,743,154 G	-	≜ Ir	tronic	GCKR	0.3248	0.2948	-3.365	4.63E-11
TCHL	rs17321515	8	126,486,409	126,486,409 T	Ŭ	Ч	tergenic '	TRIB1(dist=35,762),LINC00861(dist=448,358)	0.4425	0.04178	2.782	4.20E-09
TCHL	rs1881396	7	27,844,601	27,844,601 G		L L	TR3	ZNF512	0.3313	0.8699	-2.793	3.18E-08
TCHL	rs6861279	5	74,919,409	74,919,409 T	Ŭ	ц Г	tronic	ANKDD1B	0.3386	0.1772	2.784	3.99E-08
TCHL	rs6734059	7	27,808,154	27,808,154 C		L In	tronic	ZNF512	0.3357	0.8921	- 2.724	6.36E-08
log(TG)	rs6589566	11	116,652,423	116,652,423 C	<u> </u>	L L	tronic	ZPR1	0.2169	0.3545	0.1113	7.90E-38
log(TG)	rs180349	11	116,611,827	116,611,827 A	5	L L	tergenic	LINC00900(dist=980,909),BUD13(dist=7059)	0.2265	0.752	0.1065	8.86E-35
log(TG)	rs10503669	8	19,847,690	19,847,690 T	Ŭ	-I I	tergenic	LPL(dist=22,920),SLC18A1(dist=154,676)	0.1207	0.003444	-0.08902	1.63E-16
log(TG)	rs780094	7	27,741,237	27,741,237 C		L Ll	tronic	GCKR	0.4626	0.9806	-0.05656	2.60E-15
log(TG)	rs7115242	11	116,908,283	116,908,283 T	0	н П	tronic	SIK3	0.2796	0.133	0.05987	8.66E-14
Only the sign	ificant variants	s with P	values less than	1×10^{-7} in each	trait	are inc	luded. The mo	ore significant results are listed in the supplement ma	aterial (Ta	ble S1)		

Table 4 Results of the genome-wide association study with B_0 as the response

Table 5 Results of the genome-wide association study with B_1 as the response

Trait	SNP	CHR	BP1	BP2	A1	A2	GP	GENE	MAF	HWE_P	BETA	Р
FEV1	rs2272402	3	11,075,461	11,075,461	А	G	Intronic	SLC6A1	0.07363	0.1473	-0.5823	1.22E-08
FVC	rs2272402	3	11,075,461	11,075,461	А	G	Intronic	SLC6A1	0.07363	0.1473	- 0.595	1.40E-09

Only the variants with P values less than 1×10^{-7} are included. The more variants under suggestive threshold (P values less than 1×10^{-5}) are listed in the supplement material (Table S2)

with B_0 as the response, which is lower than the conventional heritability estimate of height of approximately 0.8 based on the assumption-free model (Visscher et al. 2006). In the case of TCHL and LDL, each heritability estimate was determined to be 0.265 and 0.22, respectively, which are also lower than the heritability estimates of 0.67 and 0.69 for TCHL and LDL, respectively, on familial and pedigree analysis (van Dongen et al. 2013). The underlying reason may be explained on the basis of the missing heritability, which describes the difference in values between heritability estimated via GWAS and via familial studies (Sandoval-Motta et al. 2017). However, systemic inflation of estimated heritability estimates of polygenic phenotypes in familial studies may be confounded owing to a shared environment or environment-dependent genetic effects (Robinson et al. 2017). Therefore, the populationbased design similar to that of the present study potentially represents the average genetic effects regardless of various confounding environmental factors.

Based on the present B_0 and B_1 model, the heritabilities of B_1 are markedly lower than those of B_0 , indicating that most of the genetic variance of traits are not temporally influenced. Here, B_0 was not determined from the baseline measurements of traits but rather the average values of repeated measurements to yield a more robust and reasonable result. If baseline measurement and longitudinal changes (B_1) calculated from those were considered responses during the estimation of heritability, the estimate may have been potentially inaccurate owing to the correlation between baseline and B_1 values. Thus, by applying a regression model to estimate the average B_0 and longitudinal changes B_1 , the effect of B_0 on B_1 in each subject could be removed.

On GWAS, the two-stage model elucidated significant variants associated with the traits and their changes in the longitudinal data. We confirmed several proven variants and identified some other significant unreported variants. In the case of the B_0 model, rs4922117 ($P = 2.13 \times 10^{-15}$) of log (HDL) and rs2335418 ($P = 3.2 \times 10^{-9}$) of LDL were both unreported; however, their proximal genes *LPL* and *HMGCR*, respectively, were significantly associated with each trait (Hoffmann et al. 2018). Furthermore, unreported genes, such as rs180349, including non-coding SNPs with a significant *P* value for TG are proximal to *BUD13*, which is strongly associated with TG (Hoffmann et al. 2018). Variants

including rs17482753 also had significant *P* values and was proximal to *LPL*, which is strongly associated with the HDL trait (Hoffmann et al. 2018). In the B_1 model, rs2272402 (*SLC6A1*, $P = 1.22 \times 10^{-8}$) was significant in both FEV1 and FVC lung function. The *SLC6A1* enhancer is associated with pulmonary function. Therefore, the present results are concurrent with previous findings regarding genes associated with each phenotype.

Among the 16 phenotypic traits in this study, only FEV1 displayed longitudinally significant heritability herein (Fig. 4), thus reliably reflecting the physiological state of the lungs and airways and acting as a predictor of morbidity and mortality in the general population; FEV1 is also widely used to define chronic obstructive pulmonary disease (COPD) (Young et al. 2007). Lung function develops in early life, peaks at a specific time point in early adulthood, and subsequently declines with age. Therefore, the decline of lung function in middle-aged and older individuals is suggested to be heritable in the general population (Gottlieb et al. 2001). However, longitudinal studies on FEV1 and FEV1/FVC have suggested several significant genetic regions that markedly differ from the numerous genetic variants associated with lung function, with FEV1 being estimated at a single time point (John et al. 2017; Tang et al. 2014). Hence, gene-environment interactions and significant genetic heterogeneity in lung function have been observed in diseases such as asthma or COPD (Hansel et al. 2013; Imboden et al. 2012). Accordingly, the present study included the middle-aged general population with similar environmental exposure without specific lung diseases, thus suggesting that intact FEV1 decreased due to aging. Therefore, the present results show that FEV1 has significant SNP heritability for longitudinal changes (FDR = 0.0012 for FEV1).

This study has several limitations. First, the analysis of new variants in the present GWAS was not replicated for other cohorts. Second, the two-stage approach is statistically inefficient even though it is computationally fast. However, the sample size was very large, which hopefully minimized this problem. Furthermore, we considered subjects with at least three or more measurements, which potentially minimize statistical power loss. Third, gene-environment interactions were not analyzed, although the estimation of random effects in the mixed model was elusive. Fourth, GCTA itself has limitations for reasons such as data overfitting and skewed singular values (Kumar et al. 2016). Though our study optimized parameters to attain accurate results using GCTA, our sample size might have resulted in certain variations in comparison with other large studies. Furthermore, the issue regarding missing heritability was inevitable to an extent because the Affymetrix genotypic array represents only common variants for SNPs, while rare genetic SNP variants were not included herein (Bandyopadhyay et al. 2017).

Despite the aforementioned limitations, our study elucidates heritability estimates via a two-stage approach using a mixed model in GCTA and a GWAS, which further determines longitudinal change effects independently with a linear model, followed by estimating heritability using regression coefficients. This approach provides a reasonable and easy method to estimate heritability in longitudinal data and potentially assess both heritability of the phenotypic mean and changes through several periods. Essentially, our results show that significant SNP heritability is objectively confirmed for longitudinal changes in lung function decline including FEV1 in comparison with other health-related indices. Therefore, genetic studies on longitudinal FEV1 decline among the middle-aged general population and chromosome 2, which attributes the most in genetic variance should be encouraged.

Acknowledgements The authors gratefully acknowledge the support given by J.G, Sangchul Park, and all the subjects who have participated in this study. This study was supported by the Bio-Synergy Research Project (NRF-2017M3A9C4065964) of the Ministry of Science, ICT and Future Planning through the National Research Foundation, by Basic Science Research Program through the NRF of Korea funded by the Ministry of Education (NRF-2016R1D1A1A09919610) and by the Industrial Core Technology Development Program (20000134) funded by the Ministry of Trade, Industry and Energy (MOTIE, Korea).

Author contributions DL contributed to analysis and drafting of the article. SW and DL contributed to the study conception and design. SW, HK and SL contributed to critical revision of the paper. All authors reviewed the manuscript.

Compliance with ethical standards

Conflict of interest Donghe Li, Hahn Kang, Sanghun Lee and Sungho Won declare that they have no conflict of interest.

Open Access This article is licensed under a Creative Commons Attribution 4.0 International License, which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons licence, and indicate if changes were made. The images or other third party material in this article are included in the article's Creative Commons licence, unless indicated otherwise in a credit line to the material. If material is not included in the article's Creative Commons licence and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this licence, visit http://creativecommons.org/licenses/by/4.0/.

References

- Bandyopadhyay B, Chanda V, Wang Y (2017) Finding the sources of missing heritability within rare variants through simulation. Bioinform Biol Insights 11:1177932217735096
- Cho YS, Go MJ, Kim YJ, Heo JY, Oh JH, Ban HJ, Yoon D, Lee MH, Kim DJ, Park M et al (2009) A large-scale genome-wide association study of Asian populations uncovers genetic factors influencing eight quantitative traits. Nat Genet 41:527–534
- Coram MA, Duan Q, Hoffmann TJ, Thornton T, Knowles JW, Johnson NA, Ochs-Balcom HM, Donlon TA, Martin LW, Eaton CB et al (2013) Genome-wide characterization of shared and distinct genetic components that influence blood lipid levels in ethnically diverse human populations. Am J Hum Genet 92:904–916
- Evans LM, Tahmasbi R, Vrieze SI, Abecasis GR, Das S, Gazal S, Bjelland DW, de Candia TR, Goddard ME, Neale BM et al (2018) Comparison of methods that use whole genome data to estimate the heritability and genetic architecture of complex traits. Nat Genet 50:737
- Friedlander Y, Austin MA, Newman B, Edwards K, Mayer-Davis EI, King MC (1997) Heritability of longitudinal changes in coronary-heart-disease risk factors in women twins. Am J Hum Genet 60:1502–1512
- Genetic epidemiologic studies on age-specified traits (2000) NIA aging and genetic epidemiology working group. Am J Epidemiol 152:1003–1008
- Gottlieb DJ, Wilk JB, Harmon M, Evans JC, Joost O, Levy D, O'Connor GT, Myers RH (2001) Heritability of longitudinal change in lung function. The Framingham study. Am J Respir Crit Care Med 164:1655–1659
- Hansel NN, Ruczinski I, Rafaels N, Sin DD, Daley D, Malinina A, Huang L, Sandford A, Murray T, Kim Y et al (2013) Genomewide study identifies two loci associated with lung function decline in mild to moderate COPD. Hum Genet 132:79–90
- Hoffmann TJ, Theusch E, Haldar T, Ranatunga DK, Jorgenson E, Medina MW, Kvale MN, Kwok PY, Schaefer C, Krauss RM et al (2018) A large electronic-health-record-based genome-wide study of serum lipids. Nat Genet 50:401–413
- Imboden M, Bouzigon E, Curjuric I, Ramasamy A, Kumar A, Hancock DB, Wilk JB, Vonk JM, Thun GA, Siroux V et al (2012) Genomewide association study of lung function decline in adults with and without asthma. J Allergy Clin Immunol 129:1218–1228
- John C, Soler Artigas M, Hui J, Nielsen SF, Rafaels N, Pare PD, Hansel NN, Shrine N, Kilty I, Malarstig A et al (2017) Genetic variants affecting cross-sectional lung function in adults show little or no effect on longitudinal lung function decline. Thorax 72:400–408
- Kim Y, Lee Y, Lee S, Kim NH, Lim J, Kim YJ, Oh JH, Min H, Lee M, Seo HJ et al (2015) On the estimation of heritability with family-based and population-based samples. Biomed Res Int 2015:671349
- Kim SJ, Lee J, Park YS, Lee CH, Yoon HI, Lee SM, Yim JJ, Kim YW, Han SK, Yoo CG (2016) Age-related annual decline of lung function in patients with COPD. Int J Chron Obstr Pulm Dis 11:51–60
- Kim Y, Han BG, KoGESg (2017) Cohort profile: The Korean Genome and Epidemiology Study (KoGES) Consortium. Int J Epidemiol 46:1350
- Kumar SK, Feldman MW, Rehkopf DH, Tuljapurkar S (2016) Correction for Krishna Kumar et al., Limitations of GCTA as a solution to the missing heritability problem. Proc Natl Acad Sci USA 113:813
- Ni G, Moser G, Schizophrenia Working Group of the Psychiatric Genomics CSH, Wray NR, Lee SH (2018) Estimation of genetic correlation via linkage disequilibrium score regression

and genomic restricted maximum likelihood. Am J Hum Genet 102:1185–1194

- Nishimura M, Makita H, Nagai K, Konno S, Nasuhara Y, Hasegawa M, Shimizu K, Betsuyaku T, Ito YM, Fuke S et al (2012) Annual change in pulmonary function and clinical phenotype in chronic obstructive pulmonary disease. Am J Respir Crit Care Med 185:44–52
- Purcell S, Neale B, Todd-Brown K, Thomas L, Ferreira MA, Bender D, Maller J, Sklar P, de Bakker PI, Daly MJ et al (2007) PLINK: a tool set for whole-genome association and population-based linkage analyses. Am J Hum Genet 81:559–575
- Robinson MR, English G, Moser G, Lloyd-Jones LR, Triplett MA, Zhu Z, Nolte IM, van Vliet-Ostaptchouk JV, Snieder H, LifeLines Cohort S et al (2017) Genotype-covariate interaction effects and the heritability of adult body mass index. Nat Genet 49:1174–1181
- Sandoval-Motta S, Aldana M, Martinez-Romero E, Frank A (2017) The human microbiome and the missing heritability problem. Front Genet 8:80
- Song YE, Lee S, Park K, Elston RC, Yang HJ, Won S (2018) ONE-TOOL for the analysis of family-based big data. Bioinformatics 34:2851–2853
- Spracklen CN, Chen P, Kim YJ, Wang X, Cai H, Li S, Long J, Wu Y, Wang YX, Takeuchi F et al (2017) Association analyses of East Asian individuals and trans-ancestry analyses with European individuals reveal new loci associated with cholesterol and triglyceride levels. Hum Mol Genet 26:1770–1784
- Tang W, Kowgier M, Loth DW, Soler Artigas M, Joubert BR, Hodge E, Gharib SA, Smith AV, Ruczinski I, Gudnason V et al (2014) Large-scale genome-wide association studies and meta-analyses of longitudinal change in adult lung function. PLoS One 9:e100776

- van de Pol M, Verhulst S (2006) Age-dependent traits: a new statistical model to separate within- and between-individual effects. Am Nat 167:766–773
- van Dongen J, Willemsen G, Chen WM, de Geus EJ, Boomsma DI (2013) Heritability of metabolic syndrome traits in a large population-based sample. J Lipid Res 54:2914–2923
- Visscher PM, Medland SE, Ferreira MA, Morley KI, Zhu G, Cornes BK, Montgomery GW, Martin NG (2006) Assumption-free estimation of heritability from genome-wide identity-by-descent sharing between full siblings. PLoS Genet 2:e41
- Yang J, Manolio TA, Pasquale LR, Boerwinkle E, Caporaso N, Cunningham JM, de Andrade M, Feenstra B, Feingold E, Hayes MG et al (2011) Genome partitioning of genetic variation for complex traits using common SNPs. Nat Genet 43:519–525
- Yang J, Zeng J, Goddard ME, Wray NR, Visscher PM (2017) Concepts, estimation and interpretation of SNP-based heritability. Nat Genet 49:1304–1310
- Young RP, Hopkins R, Eaton TE (2007) Forced expiratory volume in one second: not just a lung function test but a marker of premature death from all causes. Eur Respir J 30:616–622

Publisher's Note Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.