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OPEN The role of the systemic inflammatory response in predicting outcomes in patients with operable cancer: Systematic review and meta-analysis

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Cancer remains a leading causes of death worldwide and an elevated systemic inflammatory response (SIR) is associated with reduced survival in patients with operable cancer. This review aims to examine the evidence for the role of systemic inflammation based prognostic scores in patients with operable cancers. A wide-ranging literature review using targeted medical subject headings for human studies in English was carried out in the MEDLINE, EMBASE, and CDSR databases until the end of 2016. The SIR has independent prognostic value, across tumour types and geographical locations. In particular neutrophil lymphocyte ratio (NLR) (n = 158), platelet lymphocyte ratio (PLR) (n = 68), lymphocyte monocyte ratio (LMR) (n = 21) and Glasqow Prognostic Score/modified Glasqow Prognostic Score (GPS/mGPS) (n = 60) were consistently validated. On meta-analysis there was a significant relationship between elevated NLR and overall survival (OS) (p < 0.00001)/ cancer specific survival (CSS) (p < 0.00001), between elevated LMR and OS (p < 0.00001)/CSS (p < 0.00001), and elevated PLR and OS (p < 0.00001)/CSS (p = 0.005). There was also a significant relationship between elevated GPS/mGPS and OS (p < 0.00001)/CSS (p < 0.00001). These results consolidate the prognostic value of the NLR, PLR, LMR and GPS/mGPS in patients with resectable cancers. This is particularly true for the NLR/GPS/ mGPS which should form part of the routine preoperative and postoperative workup.

Cancer remains one of the leading causes of mortality worldwide and is responsible for 8.8 million deaths per year¹. Overall, it has been estimated that one in three people will develop cancer in their lifetime, and one in four will die from it^{2,3}. Indeed, in the UK alone it is estimated that 150,000 people die because of cancer each year^{1,3}. Such a large burden of disease accounts for a significant proportion of the healthcare budgets of the UK, US and worldwide medical care^{1,3,4}.

Four cancers: lung, colorectal, breast and prostate, account for approximately half of all new cases and deaths². For a range of solid organ malignancies including colorectal, lung, breast and prostate cancers, definitive local therapy in the form of surgical resection remains the cornerstone of treatment².

The genetic composition of many different types of cancer has been widely reported, however there is also increasing evidence that the host inflammatory response plays an important role in the development and progression of cancer^{3,5-7}. In 2010 Roxburgh and McMillan published the first comprehensive review of the role of the systemic inflammatory response in predicting survival in patients with primary operable cancer². They identified 80 studies where the systemic inflammatory response was related to either overall, and cancer specific survival². However the majority of studies used singular markers of the inflammatory response such as CRP, albumin neutrophil, lymphocyte and platelet counts, indeed just 18 studies reported combined prognostic scores to improve prediction of survival². These included eight that reported the prognostic value of the GPS, and nine studies that reported the prognostic value of NLR. While these studies reported a significant relationship between the systemic inflammatory response and survival there were variable thresholds used for the single or combined markers resulting in considerable variability in the magnitude of the effect reported².

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However, since this review there has been a marked increase in the number of studies reporting the prognostic value of combined scoring systems based on the systemic inflammatory response. The majority reported have principally been ratios of components of the white cell count such as the neutrophil lymphocyte ratio (NLR), platelet lymphocyte ratio (PLR), lymphocyte monocyte ratio (LMR) but also acute phase proteins such as C-reactive protein/albumin ratio (CAR). Another approach is to combine scores of the acute phase proteins such as GPS/mGPS^{3,8,9}. The presence of an elevated systemic inflammatory response as shown by the presence of circulating white cells and acute phase proteins is an important unifying host characteristic in patients with cancer. The prognostic ability of the combined scores has been widely reported and there have been reviews of NLR^{9,10} and mGPS¹¹ and in advanced cancer¹². The present review is the first since 2010 to focus on primarily operable cancer and to include all recognised systemic inflammation based prognostic scores. This will rationalise the evidence for the role of systemic inflammation based prognostic scores in patients with primary operable cancers.

Methods

This systematic review and meta-analysis of published literature was undertaken according to a pre-defined protocol described in the PRISMA-P statement and in a similar fashion to that recently reported with advanced inoperable cancer. The primary outcome was to assess the prognostic value of the validated combined scores of the systemic inflammatory response (NLR, PLR, LMR, GPS and mGPS) in patients with primary operable cancer.

This was carried out by a wide-ranging literature search to identify studies carried out up to December 2016. Medical subject heading (MeSH) terms(Cancer, GPS, Glasgow Prognostic Score, mGPS, modified Glasgow Prognostic Score, NLR, Neutrophil Lymphocyte Ratio, LMR, Leucocyte Monocyte Ratio, PLR, Platelet Lymphocyte Ratio), were used in the US National Library of Medicine (MEDLINE), the Excerpta Medica database (EMBASE) and the Cochrane Database of Systematic Reviews (CDSR) to identify articles.

On completion of the online search, the title and abstract of each identified study was examined for relevance. Studies not in cancer patients, studies not available in English and those published in abstract form only were excluded. Where there were multiple publications from the same cohort the most recent paper was included. Full texts were obtained for all studies deemed potentially relevant. Once further exclusions outlined below were carried out, the bibliographies of all included articles were subsequently hand searched to identify any additional studies.

Only articles that reported survival analysis and gave hazard ratios (HR) with associated confidence intervals were included in the final meta-analysis. Articles reporting survival analysis in relative risk (RR) and odds ratio (OR) were also included but not in the meta-analysis. Studies that did not follow the majority of other studies in terms of score or ratio direction interpretation were excluded from the final meta-analysis. Studies with patients who had chemotherapy and/or radiotherapy before or after surgery were also included.

Statistics. The HRs and 95% CIs were directly retrieved from the article. If several estimates were reported for the same marker, the multivariate estimate was used in preference to the univariate analysis. Data was assessed for heterogeneity using the I² statistic and χ^2 test interpreted using the guidance from the Cochrane Handbook for Systematic Reviews of Interventions¹³. The degrees of heterogeneity were defined as minimal between 0% and 30%, moderate between 30% and 50%, substantial between 50% and 80% and considerable between 80% and 100%. Given the likely differences in methodology of the studies included, meta-analysis was performed using the random- effects (DerSimonian – Laird method) model. The Z test was used to assess the overall impact of systemic inflammation based scores on overall and cancer specific survival. All P values were 2-sided and P < 0.05 were considered statistical significant. Evidence of publication bias was evaluated using visual inspection of funnel plots. All analyses were performed using Review Manager (RevMan) [Computer program]. Version 5.3. Copenhagen: The Nordic Cochrane Centre, the Cochrane Collaboration, 2014.

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Study selection process. The study selection process is summarised in Fig. 1. Initial search strategy identified 4780 articles whose titles and abstracts were reviewed. Articles were excluded if the treatment regime was chemotherapy/radiotherapy only (n = 659), where survival was not the primary outcome measure (n = 2811), full articles were not available (n = 372), and those that were a systematic review/meta-analysis (n = 374).

This led to a review of the full text of 564 articles. A further 351 articles were excluded if progression free survival (PFS) was the only outcome measured (n = 112), if the treatment regime was chemotherapy/radiotherapy only (n = 58) and if survival was not expressed as HR/OR/RR (95% CI; n = 181). The remaining 213 articles, had their bibliographies reviewed in a systematic manner and this identified a further 31 articles to be included in the final analysis leading to final figure of 244 articles considered in the present systematic review and meta-analysis.

Studies of the prognostic value of Glasgow Prognostic Score (GPS) or modified Glasgow Prognostic Score (mGPS) in patients with primary operable cancer. Eighty articles with both overall survival (OS) and/or cancer specific survival (CSS) as their primary outcome measures were identified (Supplementary Table). This comprised data on 25,207 patients (9,361 deaths) reporting the significant prognostic value of GPS/mGPS in cohorts of patients with primary operable cancer (Supplementary Table). Seventy two studies were carried out in a retrospective manner while eight were prospective (Supplementary Table). Seventy two studies used multivariate and eight used univariate survival analysis (Supplementary Table).

After exclusion forty eight studies examined the relationship with overall survival including 16,160 patients (6,051 deaths), as the primary outcome measure. On meta-analysis there was a significant association between GPS/mGPS and overall survival (HR 1.86 95% CI 1.68–2.07, p < 0.00001) with a substantial degree of

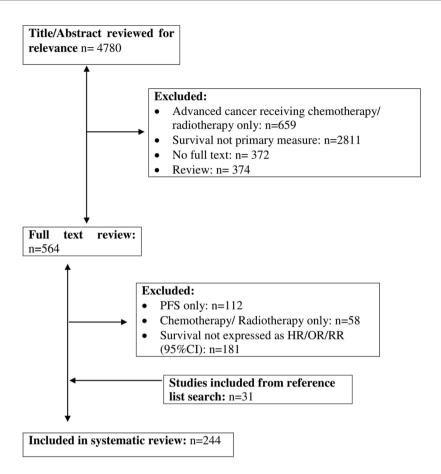


Figure 1. PRISMA flowchart demonstrating study selection.

heterogeneity ($1^2 = 61\%$, Fig. 2). These included studies on colorectal (n = 12), oesophageal (n = 7), liver (n = 6), gastric (n = 6), pancreatic (n = 5), lung (n = 4), gallbladder (n = 2), colorectal liver metastases (n = 1), renal (n = 1), bladder (n = 1), cholangiocarcinoma (n = 1), oral (n = 1) and vulval cancers (n = 1).

On meta-analysis of those studies carried out in colorectal cancer (n=12), including 4,739 patients (1,883 deaths), there was a significant association between elevated GPS/mGPS and overall survival (HR: 1.62 95% CI 1.42–1.84, p < 0.00001) with a substantial degree of heterogeneity ($I^2 = 51\%$, Fig. 3). These included studies carried out in the UK (n=8), Japan (n=2), Korea (n=1) and Australia (n=1). The proportion of patients who had an elevated GPS/mGPS was 60% in Australia, 39% in Japan, 37% in the UK and 21% in Korea.

On meta-analysis of studies involving oesophageal cancer (n=7), including 1,918 patients (669 deaths), there was a significant association between GPS/mGPS and overall survival (HR: 1.73 95% CI 1.31–2.29, p < 0.0001) with a minimal degree of heterogeneity ($I^2 = 34\%$, Fig. 4). These included studies carried out in Japan (n=4), Germany (n=1), China (n=1) and Ireland (n=1). The proportion of patients who had an elevated GPS/mGPS was 19% in Japan, 46% in Germany, 28% in China and 22% in Ireland.

On meta-analysis of studies involving liver cancer (n = 6), including 2,142 patients (801 deaths), there was a significant association between GPS/mGPS and overall survival (HR: $2.87\,95\%$ CI 1.79-4.60, p < 0.0001) with a substantial degree of heterogeneity ($I^2 = 71\%$, Fig. 5). These included studies carried out in Japan (n = 3) and China (n = 3). The proportion of patients who had an elevated GPS/mGPS was 20% in Japan and 12% in China.

On meta-analysis of studies involving gastric cancer (n=6), including 2,471 patients (753 deaths), there was a significant association between GPS/mGPS and overall survival (HR: 1.95 95% CI 1.36–2.79, p=0.0003) with a substantial degree of heterogeneity ($I^2=70\%$, Fig. 6). These included studies carried out in Japan (n=4), China (n=1) and Italy (n=1). The proportion of patients who had an elevated GPS/mGPS was, 30% in Japan, 23% in China and 52% in Italy.

On meta-analysis those studies carried out in pancreatic cancer (n=5), including 549 patients (501 deaths), there was a significant association between GPS/mGPS and overall survival (HR: 1.70 95% CI 1.21–2.38, p=0.002) with a substantial degree of heterogeneity ($I^2=60\%$, Fig. 7). These included studies carried out in the UK (n=2), Japan (n=1), Italy (n=1) and Austria (n=1). The proportion of patients who had an elevated GPS/mGPS was 45% in the UK, 23% in Japan, 68% in Italy and 34% in Austria.

After exclusion twenty nine studies examined cancer specific survival (CSS) including 9,053 patients (2,686 deaths), as its primary outcome measure. On meta-analysis there was a significant association between GPS/mGPS and cancer specific survival (HR 2.08 95% CI 1.82–2.39, p < 0.00001) with a substantial degree of heterogeneity ($I^2 = 68\%$, Fig. 8). These included studies on colorectal (n = 16), oesophageal (n = 4), oesophago-gastric (n = 2), gastric (n = 2), renal cell (n = 2), colorectal liver metastases (n = 1), oral (n = 1) and bladder cancers (n = 1).

			Favours [Experimental]	Favours [Control]		Hazard Ratio		Hazard Ratio
Study or Subgroup	log[Hazard Ratio]	SE	Total	Total	Weight	IV, Random, 95% CI	Year	IV, Random, 95% CI
Leitch EF 2007 (51)	0.732	0.232	45	149	2.5%	2.08 [1.32, 3.28]	2007	-
Richards CH 2010 (57)	0.47	0.12	136	320	3.7%	1.60 [1.26, 2.02]	2010	-
Roxburgh CS 2010 (56)	0.548	0.165	125	287	3.2%	1.73 [1.25, 2.39]	2010	-
Kobayashi T 2010 (60)	1.122	0.488	30	63	1.0%	3.07 [1.18, 7.99]	2010	-
Hefler-Frischmuth K 2010 (58)	0.0953	0.4	27	93	1.3%	1.10 [0.50, 2.41]	2010	-
Roxburgh CS 2011 (66)	0.47	0.126	135	302	3.6%	1.60 [1.25, 2.05]	2011	-
Roxburgh CS 2011 (69)	1.172	0.395	33	76	1.3%	3.23 [1.49, 7.00]	2011	_
Vashist YK 2011 (67)	0.916	0.191	71	495	2.9%	2.50 [1.72, 3.63]	2011	-
Jamieson NB 2011 (65)	0.815	0.233	109	135	2.4%	2.26 [1.43, 3.57]	2011	
Moug SJ 2011 (61)	0.445	0.145	63	206	3.4%	1.56 [1.17, 2.07]	2011	-
Nozoe T 2011 (68)	1.231	0.531	184	232	0.8%	3.42 [1.21, 9.70]	2011	
Wang DS 2012 (79)	0.334	0.136	162	324	3.5%	1.40 [1.07, 1.82]	2012	-
Jamieson NB 2012 (83)	0.571	0.201	173	173	2.8%	1.77 [1.19, 2.62]	2012	-
Kubota T 2012 (75)	1.654	0.384	92	1017	1.4%	5.23 [2.46, 11.10]	2012	
Lamb GW 2012 (80)	1.428	0.266	59	169	2.1%	4.17 [2.48, 7.02]	2012	
La Torre M 2012 (77)	0.574	0.205	84	101	2.7%	1.78 [1.19, 2.65]	2012	-
Oshiro Y 2013 (86)	1.025	0.45	46	62	1.1%	2.79 [1.15, 6.73]	2013	
Stotz M 2013 (84)	0.0908	0.176	110	110	3.0%	1.10 [0.78, 1.55]	2013	+
Son HJ 2013 (89)	0.796	0.577	55	624	0.7%	2.22 [0.72, 6.87]	2013	+
Shiba H 2013 (85)	2.43	1.231	25	30	0.2%	11.36 [1.02, 126.81]	2013	· · · · · ·
Horino K 2013 (87)	1.334	0.314	128	352	1.8%	3.80 [2.05, 7.02]	2013	
Wu XS 2014 (95)	2.387	0.751	75	85	0.5%	10.88 [2.50, 47.41]	2014	
Pinato DJ 2014 (91)	0.405	0.177	61	220	3.0%	1.50 [1.06, 2.12]	2014	
Takeno S 2014 (42)	0.214	0.154	215	552	3.3%	1.24 [0.92, 1.68]	2014	 -
Huang J 2014 (92)	0.49	0.146	153	349	3.4%	1.63 [1.23, 2.17]	2014	-
Nakamura M 2014 (38)	1.0028	0.495	44	168	0.9%	2.73 [1.03, 7.19]	2014	
Aurello P 2014 (32)	0.647	0.213	62	102	2.6%	1.91 [1.26, 2.90]	2014	
Hirashima K 2014 (40)	0.815	0.372	38	244	1.4%	2.26 [1.09, 4.68]	2014	
Xu XL 2015 (22)	0.604	0.226	259	468	2.5%	1.83 [1.17, 2.85]	2015	
Watt DG 2015 (106)	0.278	0.0845	292	508	4.0%	1.32 [1.12, 1.56]	2015	-
Okamura Y 2015 (107)	0.536	0.315	86	256	1.8%	1.71 [0.92, 3.17]	2015	
Shibutani M 2015 (104)	1.979	0.926	69	254	0.3%	7.24 [1.18, 44.43]	2015	-
Shiba H 2015 (105)	1.33	0.621	16	51	0.6%	3.78 [1.12, 12.77]	2015	-
Kawashima M 2015 (41)	0.365	0.3007	227	1043	1.9%	1.44 [0.80, 2.60]	2015	+-
Ni XC 2015 (102)	1.472	0.284	40	367	2.0%	4.36 [2.50, 7.60]	2015	
Miyazaki T 2015 (98)	0.756	0.369	44	97	1.4%	2.13 [1.03, 4.39]	2015	-
Arigami T 2015 (39)	0.07696	0.382	98	238	1.4%	1.08 [0.51, 2.28]	2015	
Hirahara N 2015 (103)	0.715	0.341	16	141	1.6%	2.04 [1.05, 3.99]	2015	
Farhan-Alanie OM 2015 (100)	0.525	0.161	56	178	3.2%	1.69 [1.23, 2.32]	2015	-
Ferro M 2015 (101)	0.223	0.267	430	1037	2.1%	1.25 [0.74, 2.11]	2015	
Walsh SM 2016 (113)	0.215	0.298	104	223	1.9%	1.24 [0.69, 2.22]	2016	
Park JH 2016 (109)	0.247	0.0636	435	1000	4.2%	1.28 [1.13, 1.45]	2016	-
Toyokawa T 2016 (37)	0.0208	0.402	77	185	1.3%	1.02 [0.46, 2.25]	2016	
Ishizuka M 2016 (21)	0.593	0.218	142	627	2.6%	1.81 [1.18, 2.77]	2016	-
Abe T 2016 (108)	2.0436	0.769	17	46	0.4%	7.72 [1.71, 34.84]	2016	
Fu YP 2016 (110)	1.255	0.474	377	772	1.0%	3.51 [1.39, 8.88]	2016	
Chan JC 2016 (112)	0.795	0.214	353	386	2.6%	2.21 [1.46, 3.37]	2016	-
Fan H 2016 (111)	0.801	0.22	373	1243	2.6%	2.23 [1.45, 3.43]	2016	
Total (95% CI)			6051	16160	100.0%	1.86 [1.68, 2.07]		. •
Heterogeneity: Tau ² = 0.06; Chi ²	= 121.67, df = 47 (P	< 0.0000	1); I ² = 61%					0.01 0.1 1 10 100
Test for overall effect: Z = 11.55	(P < 0.00001)							Favours [Experimental] Favours [Control]

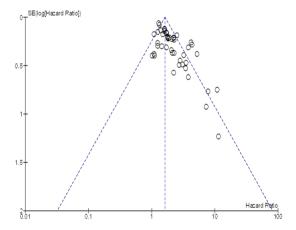
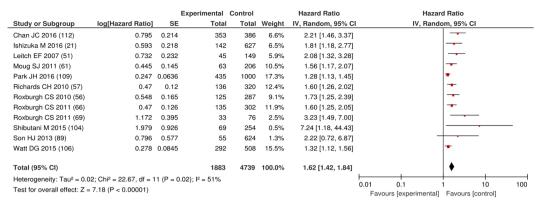


Figure 2. Forrest and Funnel Plot of Studies investigating the prognostic value of GPS/mGPS in terms of OS in an unselected cohort of patients with operable cancer.

On meta-analysis of studies involving colorectal cancer (n = 16), including 5121 patients (1300 deaths), there was a significant association between GPS/mGPS and cancer specific survival (HR: 1.75 95% CI 1.55–1.98, p < 0.00001) with a moderate degree of heterogeneity ($I^2 = 42\%$, Fig. 9). These included studies carried out in the UK (n = 15) and Japan (n = 1). The proportion of patients who had an elevated GPS/mGPS was 39% in the UK and 8% in Japan.

Studies of the prognostic value of Neutrophil Lymphocyte Ratio (NLR) in patients with primary operable cancer. One hundred and fifty eight articles with both overall survival (OS) and/or cancer specific survival (CSS) as their primary outcome measures were identified (Supplementary Table). This comprised data



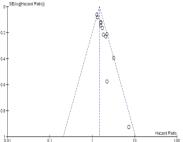
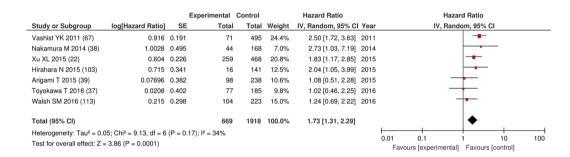


Figure 3. Forrest and Funnel Plot of Studies investigating the prognostic value of GPS/mGPS in terms of OS in patients with operable colorectal cancer.



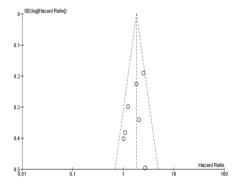


Figure 4. Forrest and Funnel Plot of Studies investigating the prognostic value of GPS/mGPS in terms of OS in patients with operable oesophageal cancer.

on 63,837 patients (22,681 deaths) reporting the significant prognostic value of NLR in cohorts of patients with primary operable cancer. All one hundred and fifty eight studies were carried out in a retrospective manner (Supplementary Table). One hundred and twenty eight studies used multivariate and thirty used univariate survival analysis (Supplementary Table).

After exclusion one hundred and nineteen studies examined the relationship with overall survival including 49,664 patients (18,542 deaths), as the primary outcome measure. On meta-analysis there was a significant association between NLR and overall survival (HR 1.73 95% CI 1.56–1.91, p < 0.00001) with a considerable degree of

			Experimental	Control	Hazard Ratio			Hazard Ratio
Study or Subgroup	log[Hazard Ratio]	SE	Total	Total	Weight	IV, Random, 95% CI	Year	IV, Random, 95% CI
Horino K 2013 (87)	1.334	0.314	128	352	18.2%	3.80 [2.05, 7.02]	2013	
Huang J 2014 (92)	0.49	0.146	153	349	24.1%	1.63 [1.23, 2.17]	2014	
Okamura Y 2015 (107)	0.536	0.315	86	256	18.2%	1.71 [0.92, 3.17]	2015	
Ni XC 2015 (102)	1.472	0.284	40	367	19.3%	4.36 [2.50, 7.60]	2015	
Fu YP 2016 (110)	1.255	0.474	377	772	13.1%	3.51 [1.39, 8.88]	2016	
Abe T 2016 (108)	2.0436	0.769	17	46	7.2%	7.72 [1.71, 34.84]	2016	
Total (95% CI)			801	2142	100.0%	2.87 [1.79, 4.60]		•
Heterogeneity: Tau ² = 0.2 Test for overall effect: Z =).004); I ² = 71%					0.01 0.1 10 100 Favours [experimental] Favours [control]	

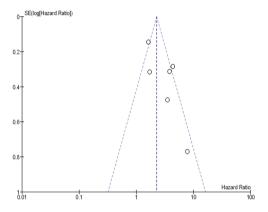
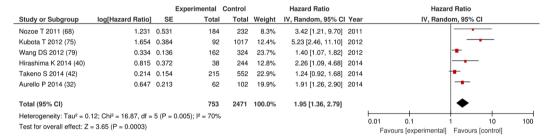


Figure 5. Forrest and Funnel Plot of Studies investigating the prognostic value of GPS/mGPS in terms of OS in patients with operable liver cancer.



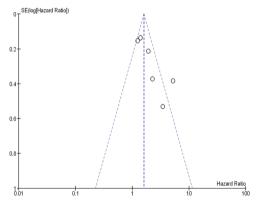


Figure 6. Forrest and Funnel Plot of Studies investigating the prognostic value of GPS/mGPS in terms of OS in patients with operable gastric cancer.

heterogeneity ($I^2 = 98\%$, Fig. 10). The most common NLR threshold examined was ≥ 5 (n = 29). Other thresholds were ≥ 3 (n = 9), ≥ 2.5 (n = 7), NLR as continuous variable (n = 7), ≥ 4 (n = 7) and ≥ 2 (n = 5). Other thresholds were used in < 5 studies and thus, meta-analysis was not carried out (n = 55).

On meta-analysis of those studies with a threshold of ≥ 5 (n = 29), including 9,997 patients (4,012 deaths) there was a significant association between elevated NLR and overall survival (HR: 1.92 95% CI 1.67–2.20, p < 0.00001) with a moderate degree of heterogeneity ($I^2 = 47\%$, Fig. 11). These included colorectal (n = 8), lung (n = 4), colorectal liver metastases (n = 4), oesophageal (n = 3), gastric (n = 2), soft tissue sarcoma (n = 2), liver (n = 2), pancreatic (n = 1), renal (n = 1), pleural mesothelioma (n = 1) and hepato-pancreatico-biliary cancers (n = 1).

			Experimental	Control		Hazard Ratio		Hazard Ratio
Study or Subgroup	log[Hazard Ratio]	SE	Total	Total	Weight	IV, Random, 95% CI	Year	r IV, Random, 95% CI
Jamieson NB 2011 (65)	0.815	0.233	109	135	22.2%	2.26 [1.43, 3.57]	2011	-
La Torre M 2012 (77)	0.574	0.205	84	101	24.4%	1.78 [1.19, 2.65]	2012	·
Jamieson NB 2012 (83)	0.571	0.201	173	173	24.7%	1.77 [1.19, 2.62]	2012	·
Shiba H 2013 (85)	2.43	1.231	25	30	1.9%	11.36 [1.02, 126.81]	2013	3
Stotz M 2013 (84)	0.0908	0.176	110	110	26.9%	1.10 [0.78, 1.55]	2013	· •
Total (95% CI)			501	549	100.0%	1.70 [1.21, 2.38]		•
Heterogeneity: Tau ² = 0.08	; Chi² = 9.88, df = 4 (P = 0.04	4); I² = 60%					0.01 0.1 1 10 100
Test for overall effect: Z = 3						Favours [experimental] Favours [control]		

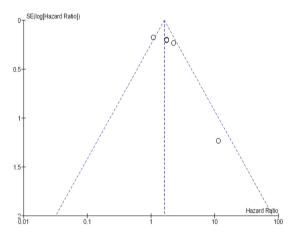
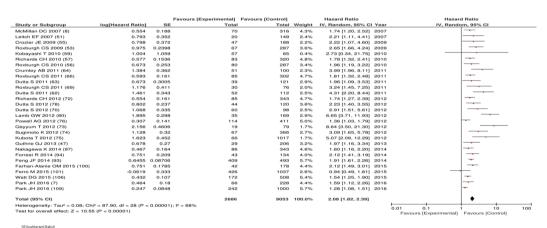


Figure 7. Forrest and Funnel Plot of Studies investigating the prognostic value of GPS/mGPS in terms of OS in patients with operable pancreatic cancer.



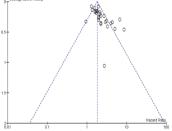
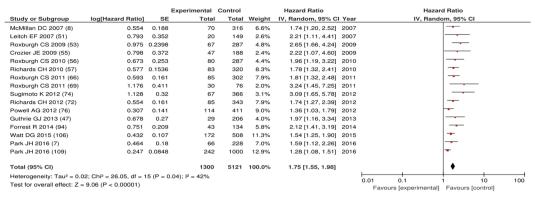


Figure 8. Forrest and Funnel Plot of Studies investigating the prognostic value of GPS/mGPS in terms of CSS in an unselected cohort of patients with operable cancer.

On meta-analysis of those studies with a threshold of ≥ 5 and colorectal cancer (n = 8), including 3,379 patients (825 deaths) there was a significant association between an NLR ≥ 5 and overall survival (HR: 1.80 95% CI 1.37–2.37, p < 0.0001) with moderate heterogeneity (I² = 45%, Fig. 12). In these eight studies, there was a variation in their geographical locations including the UK (n = 2), Korea (n = 2), Taiwan (n = 1), Austria (n = 1), US (n = 1) and Australia (n = 1). The proportion of patients who had an NLR ≥ 5 with colorectal cancer was 25% in the UK, 5% in Korea, 25% in Taiwan, 11% in US and 30% in Australia. 29% in Korea and 20% in Japan. No country had more than 4 studies and therefore no further meta-analysis was carried out.



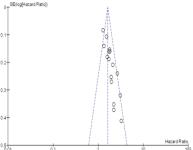


Figure 9. Forrest and Funnel Plot of Studies investigating the prognostic value of GPS/mGPS in terms of CSS in patients with operable colorectal cancer.

On meta-analysis of those studies with a threshold of \geq 3 (n = 9), including 2,638 patients (835 deaths) there was a significant association between elevated NLR and overall survival (HR: 1.83 95% CI 1.48–2.27, p < 0.00001) with a moderate degree of heterogeneity (I²=44%, Fig. 13). These included gastric (n = 2), liver (n = 1), biliary tract (n = 1), bladder (n = 1), breast (n = 1), colorectal (n = 1), pleural mesothelioma (n = 1) and endometrial cancers (n = 1). In these nine studies, there was a variation in their geographical locations including Japan (n = 4), Canada (n = 2), China (n = 1), Belgium (n = 1) and Australia (n = 1). The proportion of patients who had an NLR \geq 3 was 28% in Japan, 47% in Canada, 33% in China, 31% in Belgium and 52% in Australia. No tumour site had more than four studies and therefore no further meta-analysis was carried out.

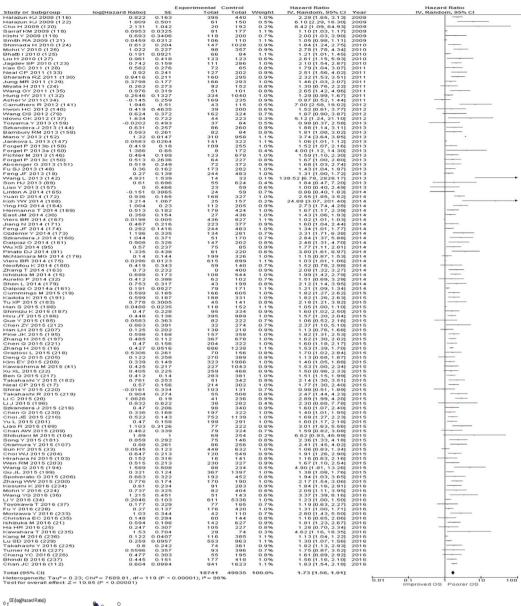
On meta-analysis of those studies with a threshold of \geq 2.5 (n = 7), including 1,888 patients (475 deaths) there was a significant association between elevated NLR and overall survival (HR: 1.78 95% CI 1.29–2.44, p = 0.0004) with a moderate degree of heterogeneity (I² = 42%, Fig. 14). These included lung (n = 3), oesophageal (n = 1), colorectal (n = 1), soft tissue sarcoma (n = 1) and liver cancers (n = 1). In these seven studies, there was a variation in their geographical locations including Japan (n = 5), China (n = 1) and US (n = 1). The proportion of patients who had an NLR \geq 2.5 was 30% in Japan, 28% in China and 50% in US. No tumour site had more than four studies and therefore no further meta-analysis was carried out.

On meta-analysis those studies with NLR as continuous variable (n=7), including 2,472 patients (1,466 deaths) there was a moderate association between elevated NLR and overall survival (HR: 1.05 95% CI 1.02–1.08, p=0.001) with a substantial degree of heterogeneity ($I^2=63\%$, Fig. 15). These included pancreatic (n=2), renal (n=2), colorectal (n=1), lung (n=1) and bladder cancers (n=1). In these seven studies, there was a variation in their geographical locations including the UK (n=2), US (n=2), China (n=1), Austria (n=1) and Australia (n=1). No tumour site had more than four studies and therefore no further meta-analysis was carried out.

On meta-analysis those studies with a threshold of \geq 4 (n = 7), including 2,195 patients (697 deaths) there was a significant association between elevated NLR and overall survival (HR: 1.36 95% CI 1.01–1.84, p = 0.04) with a substantial degree of heterogeneity (I²=73%, Fig. 16). These included glioblastoma (n = 2), gastric (n = 1), oesophageal (n = 1), ovarian (n = 1), breast (n = 1) and colon cancers (n = 1). In these seven studies, there was a variation in their geographical locations including Japan (n = 2), China (n = 1), the UK (n = 1), Belgium (n = 1), Austria (n = 1) and Ireland (n = 1). The proportion of patients who had an NLR \geq 4 was 15% in Japan, 32% in China, 22% in Belgium and 36% in Ireland. No tumour site had more than four studies and therefore no further meta-analysis was carried out.

On meta-analysis those studies with a threshold of ≥ 2 (n = 5), including 3,065 patients (1,068 deaths) there was a significant association between elevated NLR and overall survival (HR: 1.48 95% CI 1.28–1.72, p < 0.00001) with minimal heterogeneity ($I^2 = 0\%$, Fig. 17). These cancers included gastric (n = 2), colorectal (n = 1), liver (n = 1) and pancreatic (n = 1). In these five studies, there was a variation in their geographical locations including China (n = 3) and Korea (n = 2). The proportion of patients who had an NLR ≥ 2 was 60% in China and 39% in Korea. No tumour site had more than four studies and therefore no further meta-analysis was carried out.

After exclusion forty one studies examined the relationship with cancer specific survival including 17,539 patients (4,617 deaths), as its primary outcome measure. On meta-analysis there was a significant association



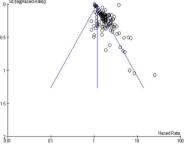


Figure 10. Forrest and Funnel Plot of Studies investigating the prognostic value of NLR in terms of OS in an unselected cohort of patients with operable cancer.

between NLR and cancer specific survival (HR 1.32 95% CI 1.24–1.41, p < 0.00001) with a considerable degree of heterogeneity ($I^2 = 81\%$, Fig. 18). The most common NLR thresholds used was ≥ 5 (n = 7), ≥ 3 (n = 6) and NLR as continuous variable (n = 5). Other thresholds did not have more than four studies and therefore meta-analysis was not carried out (n = 19).

On meta-analysis those studies with a threshold of \geq 5 (n=7), including 1,283 patients (531 deaths) there was a significant association between elevated NLR and cancer specific survival (HR: 1.89 95% CI 1.53–2.34, p < 0.00001) with minimal heterogeneity (I²=0%, Fig. 19). These included colorectal (n=2), liver only colorectal metastases (n=1) and soft tissue sarcoma (n=1), adrenal (n=1), pancreatic (n=1) and renal cancers (n=1). In these seven

			Experimental	Control		Hazard Ratio		Hazard Ratio
Study or Subgroup	log[Hazard Ratio]	SE	Total	Total	Weight	IV, Random, 95% CI	Year	IV, Random, 95% CI
Halazun KJ 2008 (116)	0.822	0.163	395	440	5.9%	2.28 [1.65, 3.13]	2008	-
Kishi Y 2009 (119)	0.693	0.3406	118	200	2.9%	2.00 [1.03, 3.90]	2009	-
Halazun KJ 2009 (122)	1.809	0.501	61	150	1.6%	6.10 [2.29, 16.30]	2009	
Hung HY 2011 (132)	0.2546	0.1327	334	1040	6.7%	1.29 [0.99, 1.67]	2011	-
Sharaiha RZ 2011 (130)	0.8416	0.211	160	295	4.9%	2.32 [1.53, 3.51]	2011	-
Neal CP 2011 (133)	0.92	0.241	127	202	4.3%	2.51 [1.56, 4.02]	2011	
Wang DS 2012 (79)	0.624	0.372	162	324	2.5%	1.87 [0.90, 3.87]	2012	
Idowu OK 2012 (137)	1.634	0.722	44	223	0.9%	5.12 [1.24, 21.10]	2012	
Kwon HC 2012 (140)	0.419	0.4635	39	200	1.8%	1.52 [0.61, 3.77]	2012	 •
Carruthers R 2012 (141)	1.946	0.51	43	115	1.6%	7.00 [2.58, 19.02]	2012	
Forget P 2013c (150)	0.513	0.2636	64	227	3.9%	1.67 [1.00, 2.80]	2013	-
Wang L 2013 (142)	4.931	1.539	14	33	0.2%	138.52 [6.78, 2828.17]	2013	
Forget P 2013b (150)	0.419	0.18	109	255	5.6%	1.52 [1.07, 2.16]	2013	-
Son HJ 2013 (89)	0.61	0.696	55	624	0.9%	1.84 [0.47, 7.20]	2013	
Absenger G 2013 (151)	0.519	0.249	72	372	4.2%	1.68 [1.03, 2.74]	2013	-
Szkandera J 2014 (160)	1.044	0.371	51	170	2.5%	2.84 [1.37, 5.88]	2014	_
Yuan D 2014 (172)	0.936	0.165	168	327	5.9%	2.55 [1.85, 3.52]	2014	_
Aurello P 2014 (32)	0.412	0.398	62	102	2.3%	1.51 [0.69, 3.29]	2014	 • • • • • • • • • •
Linton A 2014 (165)	-0.151	0.3865	24	59	2.4%	0.86 [0.40, 1.83]	2014	
Pinato DJ 2014 (91)	1.335	0.438	61	220	2.0%	3.80 [1.61, 8.97]	2014	
Pine JK 2015 (195)	0.598	0.168	157	358	5.8%	1.82 [1.31, 2.53]	2015	-
Shirai Y 2015 (220)	-0.0161	0.334	103	131	2.9%	0.98 [0.51, 1.89]	2015	
Neal CP 2015 (17)	0.57	0.156	214	302	6.1%	1.77 [1.30, 2.40]	2015	-
Chan AW 2015 (209)	0.462	0.339	79	324	2.9%	1.59 [0.82, 3.08]	2015	 •
Spolverato G 2015 (206)	0.663	0.322	192	452	3.1%	1.94 [1.03, 3.65]	2015	•
Choi JE 2015 (210)	0.522	0.143	752	1139	6.4%	1.69 [1.27, 2.23]	2015	-
Chen ZY 2015 (212)	0.863	0.391	32	274	2.4%	2.37 [1.10, 5.10]	2015	
Kawashima M 2015 (41)	0.425	0.217	227	1043	4.8%	1.53 [1.00, 2.34]	2015	·
Turner N 2016 (227)	0.5596	0.357	93	396	2.7%	1.75 [0.87, 3.52]	2016	 •
Total (95% CI)			4012	9997	100.0%	1.92 [1.67, 2.20]		
Heterogeneity: Tau ² = 0.06	Chi ² = 53.32, df = 2	8 (P = 0.0	003); I ² = 47%					
Test for overall effect: Z = 9	.25 (P < 0.00001)							0.01 0.1 1 10 100 Favours [experimental] Favours [control]
								ravours [experimental] Favours [control]

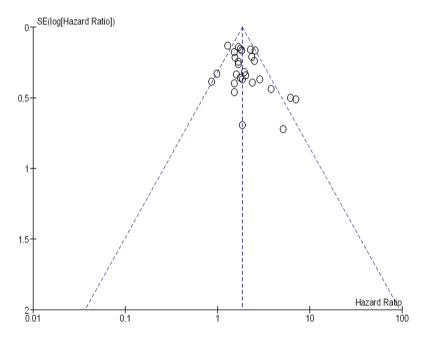
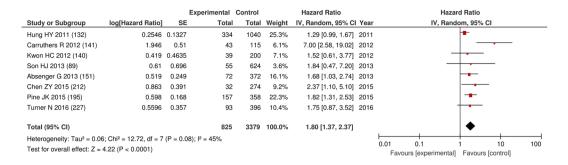


Figure 11. Forrest and Funnel Plot of Studies investigating the prognostic value of NLR \geq 5 in terms of OS in an unselected cohort of patients with operable cancer.

studies, there was a variation in their geographical locations including the UK (n=3), Austria (n=2), US (n=1) and South Korea (n=1). The proportion of patients who had an NLR ≥ 5 was 19% in the UK, 35% in US and 7% in South Korea. No tumour site had more than four studies and therefore no further meta-analysis was carried out.

On meta-analysis those studies with a threshold of ≥ 3 (n = 6), including 2,367 patients (525 deaths) there was a significant association between elevated NLR and cancer specific survival (HR: 1.81 95% CI 1.42–2.30, p < 0.00001) with a moderate degree of heterogeneity (I² = 32%, Fig. 20). These included renal (n = 2), bladder (n = 1), colorectal (n = 1), oesophageal (n = 1) and gastric cancers (n = 1). In these six studies, there was a variation in their geographical locations including Japan (n = 2), Korea (n = 1), China (n = 1), Taiwan (n = 1) and Canada (n = 1). The proportion of patients who had an NLR ≥ 3 was 25% in Japan, 20% in Korea, 20% in China, 40% in Taiwan and 51% in Canada. No tumour site had more than four studies and therefore no further meta-analysis was carried out.



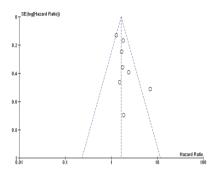


Figure 12. Forrest and Funnel Plot of Studies investigating the prognostic value of NLR \geq 5 in terms of OS in patients with operable colorectal cancer.

			Experimental	Control		Hazard Ratio		Hazard Ratio	
Study or Subgroup	log[Hazard Ratio]	SE	Total	Total	Weight	IV, Random, 95% CI	Year	IV, Random, 95%	Cl
Wang GY 2011 (135)	0.976	0.319	51	101	8.2%	2.65 [1.42, 4.96]	2011		_
Kao SC 2011 (128)	0.582	0.276	72	85	10.0%	1.79 [1.04, 3.07]	2011	-	
Forget P 2013a (150)	1.39	0.65	8	162	2.6%	4.01 [1.12, 14.35]	2013		•
Bambury RM 2013 (158)	0.593	0.261	82	84	10.7%	1.81 [1.08, 3.02]	2013	-	
McNamara MG 2014 (176)	0.14	0.144	199	326	18.6%	1.15 [0.87, 1.53]	2014	+	
Ishizuka M 2014 (15)	0.688	0.173	108	544	16.3%	1.99 [1.42, 2.79]	2014		
Hermanns T 2014 (169)	0.513	0.182	178	424	15.6%	1.67 [1.17, 2.39]	2014		
Takahashi R 2015 (219)	0.904	0.274	55	508	10.1%	2.47 [1.44, 4.23]	2015		_
Mohri Y 2016 (224)	0.737	0.325	82	404	8.0%	2.09 [1.11, 3.95]	2016		_
Total (95% CI)			835	2638	100.0%	1.83 [1.48, 2.27]		•	
Heterogeneity: Tau ² = 0.04; C	hi ² = 14.24, df = 8 (F	P = 0.08); I ² = 44%						40 400
Test for overall effect: Z = 5.5	5 (P < 0.00001)							0.01 0.1 1 Favours [experimental] Favou	10 100 irs [control]

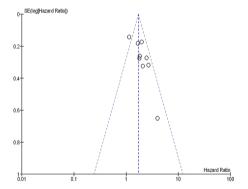
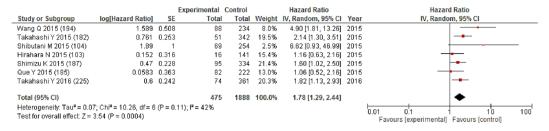


Figure 13. Forrest and Funnel Plot of Studies investigating the prognostic value of NLR \geq 3 in terms of OS in an unselected cohort of patients with operable cancer.

On meta-analysis those studies with NLR as continuous variable (n = 5), including 3,686 patients (1,312 deaths) there was a significant association between elevated NLR and cancer specific survival (HR: 1.06 95% CI 1.01–1.10, p = 0.008) with a substantial degree of heterogeneity (I 2 =80%, Fig. 21). These included renal (n = 1), bladder (n = 1), colorectal (n = 1), liver only colorectal metastases (n = 1) and gastric cancers (n = 1). In these six studies, there was a variation in their geographical locations including the US (n = 3), the UK (n = 1) and Australia (n = 1). No tumour site had more than four studies and therefore no further meta-analysis was carried out.



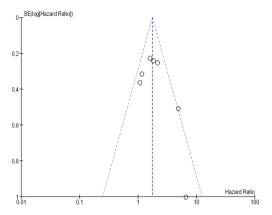
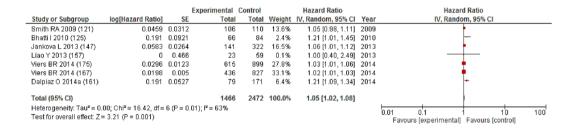


Figure 14. Forrest and Funnel Plot of Studies investigating the prognostic value of NLR \geq 2.5 in terms of OS in an unselected cohort of patients with operable cancer.



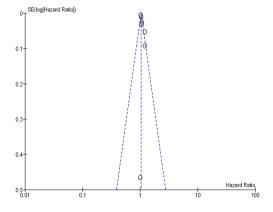


Figure 15. Forrest and Funnel Plot of Studies investigating the prognostic value of NLR as a continuous variable in terms of OS in an unselected cohort of patients with operable cancer.

Studies of the prognostic value of platelet lymphocyte ratio (PLR) in patients with primary operable cancer. Sixty eight articles with both OS and/or CSS as their primary outcome measures were identified (Supplementary Table). This comprised data on 29,273 patients (10,729 deaths) reporting the significant prognostic value of PLR in cohorts of patients with primary operable cancer (Supplementary Table). All sixty eight studies were conducted in a retrospective manner. Forty three studies were conducted in a multivariate and twenty five in a univariate manner (Supplementary Table).

After exclusions fifty five studies examined the relationship with overall survival including 25,601 patients (9,258 deaths), as the primary outcome measure. On meta-analysis there was a significant association between

			Experimental	Control		Hazard Ratio		Hazard Ratio
Study or Subgroup	log[Hazard Ratio]	SE	Total	Total	Weight	IV, Random, 95% CI	Year	IV, Random, 95% CI
Shimada H 2010 (124)	0.612	0.204	147	1028	16.8%	1.84 [1.24, 2.75]	2010	-
Miyata H 2011 (24)	0.262	0.273	92	152	13.6%	1.30 [0.76, 2.22]	2011	+-
Asher V 2011(134)	-0.145	0.259	169	235	14.2%	0.87 [0.52, 1.44]	2011	
Bambury RM 2013 (158)	0.593	0.261	82	84	14.1%	1.81 [1.08, 3.02]	2013	-
Absenger G 2013 (151)	0.798	0.25	72	372	14.6%	2.22 [1.36, 3.63]	2013	_ -
Forget P 2013a (150)	-0.673	0.816	17	172	3.1%	0.51 [0.10, 2.53]	2013	
Han S 2015 (188)	0.0488	0.0235	118	152	23.7%	1.05 [1.00, 1.10]	2015	<u>†</u>
Total (95% CI)			697	2195	100.0%	1.36 [1.01, 1.84]		•
Heterogeneity: Tau ² = 0.10;	1); I² = 73%					0.01 0.1 10 100		
Test for overall effect: Z = 2	.01 (P = 0.04)							Favours [experimental] Favours [control]

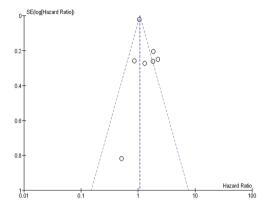
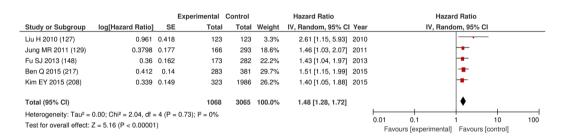


Figure 16. Forrest and Funnel Plot of Studies investigating the prognostic value of NLR \geq 4 in terms of OS in an unselected cohort of patients with operable cancer.



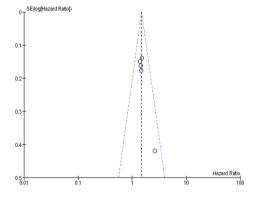


Figure 17. Forrest and Funnel Plot of Studies investigating the prognostic value of NLR \geq 2 in terms of OS in an unselected cohort of patients with operable cancer.

an elevated PLR and overall survival (HR 1.09 95% CI 1.06–1.11, p < 0.00001) with a substantial degree of heterogeneity ($I^2 = 80\%$, Fig. 22). The most common PLR thresholds examined were ≥ 300 (n = 10) and ≥ 150 (n = 7). Other thresholds did not have more than four studies and therefore meta-analysis was not carried out (n = 58).

On meta-analysis those studies with a threshold of \geq 300 (n = 10), including 3,713 patients (HR: 1.61 95% CI 1.20–2.18, p = 0.002) with a substantial degree of heterogeneity (I²=75%, Fig. 23). These included colorectal (n = 3), lung (n = 2), gastric (n = 2), colorectal liver metastases (n = 1), oesophageal (n = 1) and ovarian cancers (n = 1). In these ten studies, there was a variation in their geographical locations including the UK (n = 3), Korea (n = 2), China (n = 2), Hungary (n = 1), Italy (n = 1) and Japan (n = 1). The proportion of patients who had a PLR \geq 300 was 20% in the UK, 4% in Korea, 10% in China, 13% in Italy and 5% in Japan. No tumour site had more than four studies and therefore no further meta-analysis was carried out.

			Experimental	Control		Hazard Ratio		Hazard Ratio
Study or Subgroup	log[Hazard Ratio]	SE	Total	Total	Weight	IV, Random, 95% CI	Year	IV, Random, 95% CI
Jagdev SP 2010 (123)	1.435	0.492	63	286	0.4%	4.20 [1.60, 11.02]	2010	
Dutta S 2011 (62)	0.07696	0.187	52	112	2.1%	1.08 [0.75, 1.56]	2011	+
Gondo T 2012 (139)	0.666	0.322	54	189	0.9%	1.95 [1.04, 3.66]	2012	-
Dutta S 2012 (78)	0.174	0.2297	44	120	1.6%	1.19 [0.76, 1.87]	2012	+
Stotz M 2013 (84)	0.477	0.231	110	110	1.5%	1.61 [1.02, 2.53]	2013	-
Azuma T 2013 (153)	1.118	0.397	54	137	0.6%	3.06 [1.40, 6.66]	2013	
Pichler M 2013 (146)	0.464	0.324	59	678	0.9%	1.59 [0.84, 3.00]	2013	+-
Noh H 2013 (156)	1.406	0.471	25	442	0.4%	4.08 [1.62, 10.27]	2013	
Shibutani M 2013 (149)	0.476	0.186	136	674	2.1%	1.61 [1.12, 2.32]	2013	-
Guthrie GJ 2013 (47)	1.122	0.4656	29	206	0.4%	3.07 [1.23, 7.65]	2013	
Perisanidis C 2013 (155)	2.339	1.068	17	97	0.1%	10.37 [1.28, 84.12]	2013	
Jankova L 2013 (147)	0.00995	0.0502	86	322	6.7%	1.01 [0.92, 1.11]	2013	†
Luo HL 2014 (162)	1.853	0.661	24	234	0.2%	6.38 [1.75, 23.30]	2014	
Viers BR 2014 (175)	0.0392	0.0171	345	899	7.8%	1.04 [1.01, 1.08]	2014	<u> </u>
Forrest R 2014 (94)	0.82	0.423	43	134	0.5%	2.27 [0.99, 5.20]	2014	
Kubo T 2014 (166)	0.536	0.262	74	524	1.2%	1.71 [1.02, 2.86]	2014	
Dalpiaz O 2014 (161)	0.9999	0.398	58	202	0.6%	2.72 [1.25, 5.93]	2014	
Tanaka N 2014 (170)	0.385	0.183	129	665	2.2%	1.47 [1.03, 2.10]	2014	-
Szkandera J 2014 (160)	0.683	0.481	22	170	0.4%	1.98 [0.77, 5.08]	2014	+
Ying HQ 2014 (164)	1.019	0.243	100	205	1.4%	2.77 [1.72, 4.46]	2014	-
Hermanns T 2014 (169)	0.631	0.154	110	424	2.8%	1.88 [1.39, 2.54]	2014	-
Dalpiaz O 2014a (161)	0.148	0.074	54	171	5.6%	1.16 [1.00, 1.34]	2014	+
Viers BR 2014 (167)	0.0198	0.00747	233	827	8.0%	1.02 [1.01, 1.04]	2014	<u> </u>
Neofytou K 2015 (192)	0.182	0.0636	60	140	6.1%	1.20 [1.06, 1.36]	2015	*
Lee SK 2015 (211)	0.0862	0.0747	300	3116	5.6%	1.09 [0.94, 1.26]	2015	+
Duan H 2015 (202)	0.464	0.174	192	371	2.4%	1.59 [1.13, 2.24]	2015	-
Bagante F 2015 (193)	0.793	0.331	50	84	0.8%	2.21 [1.16, 4.23]	2015	
Deng Q 2015 (205)	0.425	0.164	235	389	2.6%	1.53 [1.11, 2.11]	2015	-
Shin JS 2015 (184)	1.823	0.913	5	269	0.1%	6.19 [1.03, 37.06]	2015	
Cummings M 2015 (19)	0.519	0.251	96	605	1.3%	1.68 [1.03, 2.75]	2015	
Kim M 2015 (215)	0.165	0.426	73	277	0.5%	1.18 [0.51, 2.72]	2015	
Neal CP 2015 (17)	0.656	0.164	204	302	2.6%	1.93 [1.40, 2.66]	2015	-
Fu Y 2016 (228)	0.351	0.15	171	420	2.9%	1.42 [1.06, 1.91]	2016	-
Morizawa Y 2016 (233)	0.956	0.257	32	110	1.3%	2.60 [1.57, 4.30]	2016	
Kosumi K 2016 (234)	0.61	0.28	65	283	1.1%	1.84 [1.06, 3.19]	2016	
Wang SC 2016 (231)	0.0953	0.0187	588	1498	7.8%	1.10 [1.06, 1.14]	2016	
Chen PC 2016 (214)	0.0488	0.178	221	323	2.3%	1.05 [0.74, 1.49]	2016	+
Mohri Y 2016 (224)	0.678	0.306	65	404	1.0%	1.97 [1.08, 3.59]	2016	
Bhindi B 2016 (237)	0.385	0.103	107	418	4.4%	1.47 [1.20, 1.80]	2016	-
Xie X 2016 (223)	0.179	0.185	147	317	2.2%	1.20 [0.83, 1.72]	2016	+
Kang M 2016 (236)	0.148	0.0481	85	385	6.8%	1.16 [1.06, 1.27]	2016	•
								1.
Total (95% CI)			4617	17539	100.0%	1.32 [1.24, 1.41]		[•
Heterogeneity: Tau ² = 0.01	; Chi ² = 208.77, df = 4	10 (P < 0.0	00001); I ² = 81%					0.01 0.1 1 10 100
Test for overall effect: Z = 8	3.83 (P < 0.00001)							0.01 0.1 1 10 100 Improved CSS Poorer CSS
								improved 000 Fuorer 000

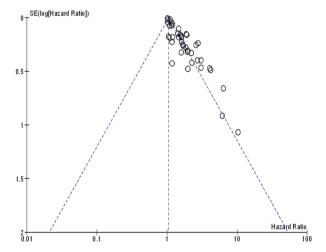
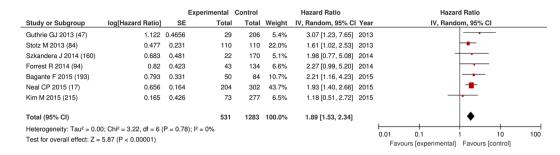


Figure 18. Forrest and Funnel Plot of Studies investigating the prognostic value of NLR in terms of CSS in an unselected cohort of patients with operable cancer.

On meta-analysis those studies with a threshold of \geq 150 (n=7), including 1,315 patients (667 deaths) there was a significant association between elevated PLR and overall survival (HR: 1.59 95% CI 1.29–1.97, p < 0.0001) with a minimal degree of heterogeneity (I²=29%, Fig. 24). These included oesophageal (n=2), pancreatic (n=2), liver (n=1), colorectal liver metastases (n=1) and colorectal cancers (n=1). In these seven studies, there was a variation in their geographical locations including China (n=2), Japan (n=2), the UK (n=1), Hong Kong (n=1) and Australia (n=1). The proportion of patients who had a PLR \geq 150 was 43% in China, 49% in Japan, 41% in the UK, 27% in Hong Kong and 75% in Australia. No tumour site had more than four studies and therefore no further meta-analysis was carried out.



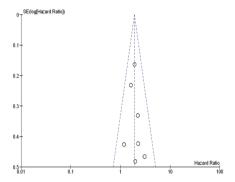
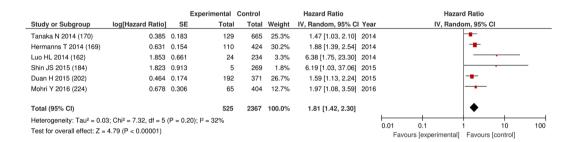


Figure 19. Forrest and Funnel Plot of Studies investigating the prognostic value of NLR \geq 5 in terms of CSS in an unselected cohort of patients with operable cancer.



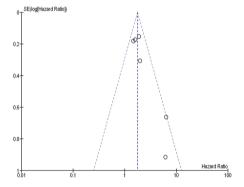


Figure 20. Forrest and Funnel Plot of Studies investigating the prognostic value of NLR \geq 3 in terms of CSS in an unselected cohort of patients with operable cancer.

After exclusions fifteen studies examined the relationship with cancer specific survival including 4,489 patients (1,769 deaths), as the primary outcome measure. On meta-analysis there was a significant association between an elevated PLR and cancer specific survival (HR 1.21 95% CI 1.06–1.38, p = 0.005) with a substantial degree of heterogeneity ($I^2 = 63\%$, Fig. 25). The most common PLR threshold examined was ≥ 300 (n = 4). Other thresholds used were >150 (n = 1), ≥ 25.4 (n = 1), >103 (n = 1), ≥ 132 (n = 1), ≥ 176 (n = 1), >190 (n = 1), ≥ 200 (n = 1), ≥ 240 (n = 1), ≥ 292 (n = 1), PLR as continuous variable (n = 1) and PLR per 100 units (n = 1). These included studies on oesophageal (n = 3), colorectal (n = 3), gastric (n = 2), colorectal liver metastases (n = 1), adrenal (n = 1), renal (n = 1), endometrial (n = 1), bladder (n = 1), soft tissue sarcoma (n = 1) and breast cancers (n = 1). Geographically studies were located in the UK (n = 5), China (n = 4), Austria (n = 2), Japan (n = 1), US

			Experimental	Control		Hazard Ratio			Haz	ard Ra	itio	
Study or Subgroup	log[Hazard Ratio]	SE	Total	Total	Weight	IV, Random, 95% Cl	Year		IV, Ran	dom,	95% CI	
Jankova L 2013 (147)	0.00995	0.0502	86	322	11.3%	1.01 [0.92, 1.11]	2013			+		
Viers BR 2014 (175)	0.0392	0.0171	345	899	25.8%	1.04 [1.01, 1.08]	2014			•		
Viers BR 2014 (167)	0.0198	0.00747	233	827	29.8%	1.02 [1.01, 1.04]	2014			٠		
Neofytou K 2015 (192)	0.182	0.0636	60	140	8.2%	1.20 [1.06, 1.36]	2015			-		
Wang SC 2016 (231)	0.0953	0.0187	588	1498	24.9%	1.10 [1.06, 1.14]	2016			•		
Total (95% CI)			1312	3686	100.0%	1.06 [1.01, 1.10]						
Heterogeneity: Tau ² = 0	.00; Chi ² = 20.05, df =	4 (P = 0.0	005); I ² = 80%					0.01	0.1	+	10	100
Test for overall effect: Z	= 2.63 (P = 0.008)							0.01 Favour	s [experimenta	l] Fa		

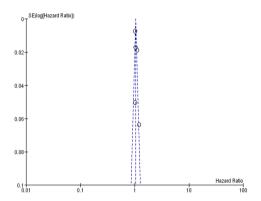


Figure 21. Forrest and Funnel Plot of Studies investigating the prognostic value of NLR as a continuous variable in terms of CSS in an unselected cohort of patients with operable cancer.

(n = 1), South Korea (n = 1) and Canada (n = 1). The proportion of patients who had an elevated PLR was 12% in the UK, 55% in China, 23% in Japan, 38% in US and 3% in South Korea. No specific PLR thresholds had more than four studies and therefore no further meta-analysis was carried out.

Studies of the prognostic value of lymphocyte monocyte ratio (LMR) in patients with primary operable cancer. Twenty one articles with both OS and/or CSS as their primary outcome measures were identified (Supplementary Table). This comprised data on 15,386 patients (4,298 deaths) reporting the significant prognostic value of LMR in cohorts of patients with primary operable cancer (Supplementary Table). All 21 studies were retrospective. Nineteen studies used multivariate and two used univariate survival analysis (Supplementary Table).

After exclusion twelve studies examined the relationship with overall survival including 11,913 patients (3,106 deaths), as the primary outcome measure. On meta-analysis there was a significant association between a elevated LMR and overall survival (HR 0.69 95% CI 0.63–0.74, p < 0.00001) with a substantial degree of heterogeneity ($I^2 = 61\%$, Fig. 26). There was a variety of LMR cut-offs used in each study including ≥ 2 , (n = 1), ≥ 2.14 (n = 1), ≥ 2.35 (n = 1), ≥ 2.38 (n = 1), ≥ 2.83 (n = 1), ≥ 2.85 (n = 1), ≥ 2.87 (n = 1), ≥ 3.23 (n = 1), ≥ 3.80 (n = 1), ≥ 4.32 (n = 1) and ≥ 4.95 (n = 1). These included studies on colorectal (n = 3), bladder (n = 2), liver only colorectal metastases (n = 1), gastric (n = 1), renal (n = 1), liver (n = 1), breast (n = 1), soft tissue sarcoma (n = 1) and cervical cancers (n = 1). Geographically the studies were carried out in China (n = 6), Austria (n = 3), the UK (n = 1), Canada (n = 1) and Australia (n = 1). The proportion of patients who had high LMRs was 71% in China, 68% in Japan, 64% in the UK, 49% in Australia and 48% in Austria. No specific LMR thresholds had more than four studies and therefore no further meta-analysis was carried out.

After exclusion five studies examined the relationship with cancer specific survival including 1,627 patients (697 deaths), as the primary outcome measure. On meta-analysis there was a significant association between a elevated LMR and cancer specific survival (HR 0.70 95% CI 0.60–0.82, p < 0.00001) with a moderate degree of heterogeneity ($I^2 = 47\%$, Fig. 27). There was a variety of LMR cut-offs used in each study including >2.35 (n = 1), >2.85 (n = 1), >2.93 (n = 1) and \geq 4.95 (n = 1). One study expressed LMR in terms of log. These included studies on liver only colorectal metastases (n = 1), gastric cancer (n = 1), oesophageal cancer (n = 1), bladder cancer (n = 1) and soft tissue sarcoma (n = 1). Geographically the studies were carried out in the China (n = 2), UK (n = 1), Austria (n = 1), and Canada (n = 1). The proportion of patients who had high LMRs was 68% in Japan, 64% in the UK, 50% in Austria and 40% in China. No specific LMR thresholds had more than four studies and therefore no further meta-analysis was carried out.

Studies of the prognostic value of other scores of the systemic inflammatory response in patients with primary operable cancer. Thirty five articles reported a variety of other scores reported in less than 10 studies each. These included the PNI (Prognostic Nutritional Index), COP-NLR (combined platelet count and NLR), NLR/PLR combination, CAR (CRP/albumin ratio), SI (systemic inflammatory score), SII (systemic inflammatory index), NLR/CRP combination, (HALP) haemoglobin, albumin, lymphocyte and platelet, NLR/ESR (erythrocyte sedimentation rate) combination, (WLR) white cell count to lymphocyte count ratio,

			Experimental	Control		Hazard Ratio		Hazard Ratio
Study or Subgroup	log[Hazard Ratio]	SE	Total	Total	Weight	IV, Random, 95% CI	Year	IV, Random, 95% CI
Smith RA 2009 (121)	0.00399	0.001016	93	110	32.5%	1.00 [1.00, 1.01]	2009	•
Bhatti I 2010 (125)	-0.0222	0.0456	66	84	6.7%		2010	+
Asher V 2011(134)	0.529	0.2546	169	235	0.3%		2011	
Kwon HC 2012 (140)	0.669	0.265	39	200	0.2%	1.95 [1.16, 3.28]		
Carruthers R 2012 (141)	0.405	0.31	43	115	0.2%		2012	+
Wang DS 2012 (79)	-0.143	0.136	162	324	0.9%	0.87 [0.66, 1.13]		+
Raungkaewmanee S 2012 (238)	0.344	0.306	50	166	0.2%		2012	+-
Feng JF 2013 (18)	0.56	0.135	244	483	0.9%		2013	-
Feng JF 2013 (239)	0.821	0.401	35	43	0.1%		2013	
Stotz M 2013 (84)	0.125	0.168	110	110	0.6%	1.13 [0.82, 1.58]	2013	+
Son HJ 2013 (89)	0.696	0.679	55	624	0.0%	2.01 [0.53, 7.59]		
Toiyama Y 2013 (159)	0.775	0.448	37	84	0.1%		2013	 -
Ying HQ 2014 (164)	0.14	0.2065	112	205	0.4%	1.15 [0.77, 1.72]		+
Jiang N 2014 (171)	0.0658	0.153	223	377	0.7%	1.07 [0.79, 1.44]		+
Pinato DJ 2014 (91)	0.47	0.57	61	220	0.1%	1.60 [0.52, 4.89]		· · · · · · · · · · · · · · · · · · ·
Szkandera J 2014 (242)	0.399	0.245	91	372	0.3%	1.49 [0.92, 2.41]		
Neofytou K 2014 (180)	0.775	0.351	59	140	0.1%		2014	
Szkandera J 2014 (160)	-0.494	0.364	51	170	0.1%		2014	
Feng JF 2014 (174)	0.61	0.137	244	483	0.9%		2014	· -
Baranyai Z 2014 (240)	1.253	0.238	335	336	0.3%	3.50 [2.20, 5.58]		· ·
Aurello P 2014 (32)	0.122	0.465	62	102	0.1%	1.13 [0.45, 2.81]		
Zhang T 2014 (163)	0.686	0.228	129	400	0.3%		2014	
Krenn-Pilko S 2014 (241)	0.652	0.329	136	793	0.2%	1.92 [1.01, 3.66]		-
Yuan D 2014 (172)	0.335	0.241	185	327	0.3%		2014	
Hsu JT 2015 (186)	-0.108	0.13	395	989	1.0%	0.90 [0.70, 1.16]		+
Li J 2015 (196)	-0.00702	0.621	38	282	0.0%		2015	
Spolverato G 2015 (206)	0.582	0.271	192	452	0.2%	1.79 [1.05, 3.04]		
Zhang WW 2015 (200)	0.769	0.196	170	190	0.5%	2.16 [1.47, 3.17]		
Zhang H 2015 (197)	-0.0346	0.122	367	678	1.1%	0.97 [0.76, 1.23]		+
Sun KY 2015 (33)	0.174	0.1096	448	632	1.4%	1.19 [0.96, 1.48]		-
Pang Q 2015 (245)	0.704	0.212	254	316	0.4%		2015	 -
Que Y 2015 (185)	0.956	0.406	82	222	0.1%		2015	
Kawashima M 2015 (41)	0.854	0.247	227	1043	0.3%	2.35 [1.45, 3.81]		
Shirai Y 2015 (220)	0.524	0.245	103	131	0.3%	1.69 [1.04, 2.73]		-
Kim EY 2015 (208)	0.0344	0.128	323	1986	1.0%	1.03 [0.81, 1.33]		+
Cummings M 2015 (19)	0.637	0.191	166	605	0.5%		2015	
Qu JL 2015 (199)	0.566	0.128	60	274	1.0%	1.76 [1.37, 2.26]	2015	-
Han LH 2015 (207)	0.0139	0.284	138	218	0.2%		2015	+
Neal CP 2015 (17)	0.218	0.108	214	302	1.4%	1.24 [1.01, 1.54]	2015	-
Chen Q 2015 (230)	0.344	0.162	197	322	0.7%	1.41 [1.03, 1.94]	2015	-
Xu XL 2015 (22)	0.113	0.127	259	468	1.1%			+
Wang Q 2015 (194)	0.47	0.502	88	234	0.1%	1.60 [0.60, 4.28]	2015	+
Chan AW 2015 (209)	0.206	0.248	79	324	0.3%	1.23 [0.76, 2.00]	2015	+-
Choi WJ 2015 (204)	0.593	0.27	120	549	0.2%	1.81 [1.07, 3.07]	2015	 -
Yoshida T 2015 (254)	0.524	0.245	103	131	0.3%	1.69 [1.04, 2.73]	2015	
Zhang GM 2015 (243)	0.149	0.332	55	124	0.2%	1.16 [0.61, 2.22]	2015	
Han S 2015 (188)	0.002996	0.00203	118	152	32.3%	1.00 [1.00, 1.01]		•
Deng Q 2015 (205)	0.0296	0.14	270	389	0.9%	1.03 [0.78, 1.36]		+
Messager M 2015 (244)	0.904	0.362	39	153	0.1%	2.47 [1.21, 5.02]		
Bhindi B 2016 (237)	0.148	0.0677	177	418	3.4%	1.16 [1.02, 1.32]	2016	+
Toyokawa T 2016 (37)	0.193	0.284	77	185	0.2%	1.21 [0.70, 2.12]	2016	+-
Chan JC 2016 (112)	0.465	0.0866	941	1623	2.2%		2016	
Li Y 2016 (34)	0.161	0.111	611	5336	1.4%	1.17 [0.95, 1.46]	2016	 -
Wang YQ 2016 (36)	0.565	0.28	51	143	0.2%	1.76 [1.02, 3.05]		
Ha HR 2016 (25)	-0.377	0.342	105	227	0.2%		2016	-+
Total (95% CI)			9258	25601	100.0%	1.09 [1.06, 1.11])
Heterogeneity: Tau ² = 0.00; Chi ² =	266.16, df = 54 (P < 0	0.00001); F	2 = 80%					0.01 0.1 1 10 100
Test for overall effect: Z = 6.20 (P <	< 0.00001)							0.01 0.1 1 10 100 Improved OS Poorer OS
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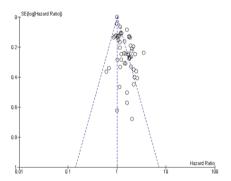
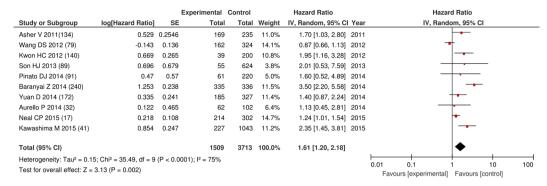


Figure 22. Forrest and Funnel Plot of Studies investigating the prognostic value of PLR in terms of OS in an unselected cohort of patients with operable cancer.

(APRI) AST-platelet ratio index, PI/CRP/WCC combination, Canton score, (AGR) albumin/globulin ratio, CRP/ Neutrophil combination, (PIS) Prognostic Inflammation Score, and the CONUT score.

Eight articles with both overall survival (OS) and/or cancer specific survival (CSS) as their primary outcome measures were identified (Supplementary Table). This comprised data on 2,666 patients (1,387 deaths) reporting the significant prognostic value of PNI in cohorts of patients with primary operable cancer. All eight studies were carried out in a retrospective manner (Supplementary Table). Six studies used multivariate and two used univariate survival analysis (Supplementary Table).



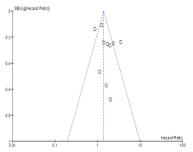
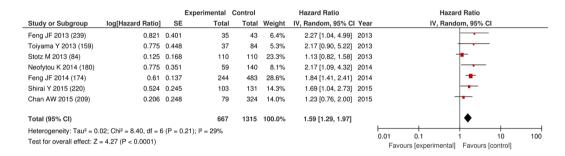


Figure 23. Forrest and Funnel Plot of Studies investigating the prognostic value of PLR \geq 300 in terms of OS in an unselected cohort of patients with operable cancer.



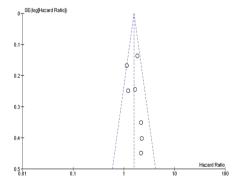
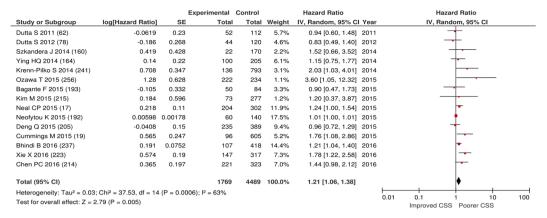


Figure 24. Forrest and Funnel Plot of Studies investigating the prognostic value of PLR \geq 150 in terms of OS in an unselected cohort of patients with operable cancer.

After exclusion seven studies examined the relationship with overall survival including 2,087 patients (1,087 deaths), as the primary outcome measure. On meta-analysis there was a significant association between PNI and overall survival (HR 1.76 95% CI 1.52–2.04, p < 0.00001) with minimal heterogeneity ($I^2 = 0\%$, Fig. 28). The most common PNI threshold examined was ≤ 45 (n = 3), ≤ 50 (n = 1), ≤ 50.5 (n = 1), 48.5 (n = 1), 48.2 (n = 1). These included hepatocellular (n = 3), gastric (n = 2), lung (n = 1) and colorectal liver metastases (n = 1). In these eight studies, there was a variation in their geographical locations including Japan (n = 2), UK (n = 1), Hong Kong (n = 1), China (n = 1), US (n = 1) and Italy (n = 1). The proportion of patients who with an elevated PNI was 74% in Hong Kong, 59% in Japan, 59% in Italy, 52% in China and 17% in the UK. No tumour site had more than four



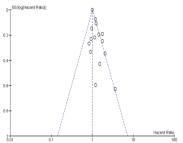


Figure 25. Forrest and Funnel Plot of Studies investigating the prognostic value of PLR in terms of CSS in an unselected cohort of patients with operable cancer.

			Experimental	Control		Hazard Ratio		Hazard Rat	io
Study or Subgroup	log[Hazard Ratio]	SE	Total	Total	Weight	IV, Random, 95% CI Yea	ar	IV, Random, 9	5% CI
Zhou X 2014 (250)	-0.374	0.142	250	426	9.9%	0.69 [0.52, 0.91] 201	4	-	
Stotz M 2014 (248)	-0.673	0.251	72	372	5.5%	0.51 [0.31, 0.83] 201	4		
Szkandera J 2014 (160)	-1.049	0.379	51	170	3.0%	0.35 [0.17, 0.74] 201	4		
Chen L 2015 (259)	-0.875	0.274	64	485	4.8%	0.42 [0.24, 0.71] 201	5		
Neal CP 2015 (17)	-0.449	0.153	214	302	9.3%	0.64 [0.47, 0.86] 201	5	-	
Deng Q 2015 (205)	0	0.157	270	389	9.1%	1.00 [0.74, 1.36] 201	5	+	
Hutterer GC 2015 (257)	-0.58	0.247	82	182	5.6%	0.56 [0.35, 0.91] 201	5		
Lin ZX 2015 (253)	-0.921	0.305	48	210	4.2%	0.40 [0.22, 0.72] 201	5		
Wen J 2015 (252)	-0.174	0.089	326	2000	12.8%	0.84 [0.71, 1.00] 201	5	•	
Li Y 2016 (34)	-0.273	0.104	611	5336	12.0%	0.76 [0.62, 0.93] 201	6	-	
Chan JC 2016 (112)	-0.564	0.088	941	1623	12.9%	0.57 [0.48, 0.68] 201	6	•	
Bhindi B 2016 (237)	-0.357	0.119	177	418	11.1%	0.70 [0.55, 0.88] 201	6	-	
Total (95% CI)			3106	11913	100.0%	0.65 [0.57, 0.75]		•	
Heterogeneity: Tau ² = 0.0	3; Chi ² = 28.43, df = 1	1 (P = 0	.003); I ² = 61%				-	- 	
Test for overall effect: Z =	5.89 (P < 0.00001)						0.01	0.1 1 Improved OS Poo	10 100 orer OS

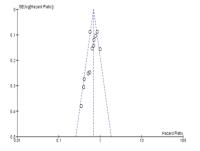


Figure 26. Forrest and Funnel Plot of Studies investigating the prognostic value of LMR in terms of OS in an unselected cohort of patients with operable cancer.

studies and therefore no further meta-analysis was carried out. Two studies examined the relationship with cancer specific survival including 579 patients (300 deaths), as the primary outcome measure. Both of these studies used a PNI threshold of \leq 45. No threshold was used in \geq 4 studies and thus, meta-analysis was not carried out.

Four studies reported the COP-NLR score. The first such study was by Ishizuka and coworkers ¹⁴ from Japan. In this multivariate survival analysis on patients with colorectal cancer, low COP-NLR was shown to be related to a statistically better cancer specific survival (OR: 0.46495% CI 0.267-0.807 p = 0.007). The second such study

			Experimental	Control		Hazard Ratio		Hazar	d Ratio	
Study or Subgroup	log[Hazard Ratio]	SE	Total	Total	Weight	IV, Random, 95% CI Ye	ar	IV, Rando	om, 95% CI	
Szkandera J 2014 (160)	-1.108	0.514	22	170	4.7%	0.33 [0.12, 0.90] 20	14	-		
Huang Y 2015 (258)	-0.511	0.198	129	348	19.9%	0.60 [0.41, 0.88] 20	15	-		
Deng Q 2015 (205)	0	0.173	235	389	22.9%	1.00 [0.71, 1.40] 20	15	-	•	
Neal CP 2015 (17)	-0.472	0.161	204	302	24.5%	0.62 [0.45, 0.86] 20	15	-		
Bhindi B 2016 (237)	-0.371	0.138	107	418	27.9%	0.69 [0.53, 0.90] 20	16	•		
Total (95% CI)			697	1627	100.0%	0.69 [0.55, 0.87]		♦		
Heterogeneity: Tau ² = 0.03	3; Chi ² = 7.52, df = 4	(P = 0.1	1); I ² = 47%				-		1 10	100
Test for overall effect: Z =					0.01	0.1 Improved CSS	1 10 Poorer CSS	100		

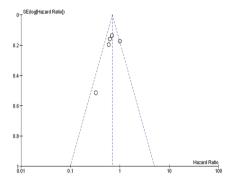
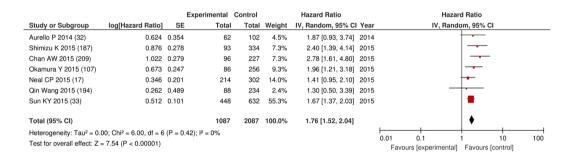


Figure 27. Forrest and Funnel Plot of Studies investigating the prognostic value of LMR in terms of CSS in an unselected cohort of patients with operable cancer.



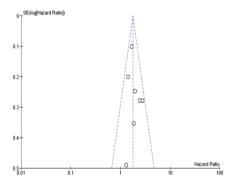


Figure 28. Forrest and Funnel Plot of Studies investigating the prognostic value of PNI in terms of OS in an unselected cohort of patients with operable cancer.

was also by Ishizuka and coworkers¹⁵ from Japan. In this multivariate survival analysis on patients with gastric cancer, elevated COP-NLR was shown to be related to a statistically significant worse overall survival (HR: 1.781 95% CI 1.094–2.899 p = 0.020). The third such study was by Zhang and coworkers¹⁶ from China. In this multivariate survival analysis on patients with lung cancer, elevated COP-NLR was shown to be related to a statistically significant worse overall survival (HR: 1.810 95% CI 1.587–2.056 p < 0.001). The fourth such study was by Neal and coworkers¹⁷ from the UK. In this univariate survival analysis on patients with colorectal liver metastases, elevated COP-NLR was shown to be related to a statistically significant worse overall survival (HR: 1.230 95% CI 1.005–1.505 p = 0.045) and worse cancer specific survival (HR: 1.243 95% CI 1.003–1.541 p = 0.047).

Three studies reported the combination of the NLR and PLR. The first such study was by Feng and coworkers from China. The combination of NLR and PLR is collectively named the CNP. The CNP was calculated based on data obtained on the day of admission, where patients with both elevated NLR (>3.45) and PLR (>166.5) were allocated a score of 2, and patients showing one or neither were allocated a score of 1 or 0, respectively. In this multivariate survival analysis on patients with oesophageal cancer, CNP 1 or 2 was shown to be related to a statistically worse overall survival (HR: 1.964 95% CI 1.371–2.814 p <0.001). The second such study was by Cummings and coworkers from the UK. In this multivariate survival analysis on patients with endometrial cancer, both high NLR and PLR was shown to be related to a statistically significant worse overall survival (HR: 2.54 95% CI 1.61–4.01 p <0.001) and worse cancer specific survival (HR: 2.26 95% CI 1.24–4.13 p =0.008). The third such study was by Chuan Li and coworkers from China. In this multivariate survival analysis on patients with liver cancer, elevated postoperative NLR-PLR was shown to be related to a statistically significant worse overall survival (HR: 2.894 95% CI 1.992–4.2 p <0.001).

Two studies reported the CAR. The first such study was by Ishizuka and coworkers²¹ from Japan. In this multivariate survival analysis on patients with colorectal cancer, CAR >0.038 was shown to be related to a statistically worse overall survival (HR: 2.613 95% CI 1.621–4.212 p <0.001). The second such study was by Xu and coworkers^{22,23} from China. In this multivariate survival analysis on patients with oesophageal cancer, CRP/Albumin ratio >0.50 was shown to be related to a statistically significant worse overall survival (HR: 2.44 95% CI 1.82–3.26 p <0.0001).

One study reported the SI a score involving leucocyte count, serum albumin and haemoglobin level. High leucocyte count ($>9,500\,\mu$ l), low serum albumin level ($3.5\,g/d$ l) and low haemoglobin level ($<12.5\,mg/d$ l) was each allocated a score of 1.The study was conducted by Miyata and coworkers²⁴ from Japan. In this multivariate survival analysis on patients with oesophageal cancer, SI score of 2/3 was shown to be related to a statistically significant worse overall survival (HR: $3.17\,95\%$ CI $1.74-5.78\,p=0.0002$).

One study reported on the SII which was determined as neutrophil \times platelet/lymphocyte. The study was conducted by Ha and coworkers^{25,26} from South Korea. In this multivariate survival analysis on patients with ampulla of vater cancer, SII \le 780 was shown to predict better overall survival (HR: 0.924 95% CI 0.44–1.93 p = 0.833).

One study reported on the combination of the NLR and CRP. The study was conducted by Tomita and coworkers²⁷ from Japan. In this multivariate survival analysis on patients with lung cancer, low NLR and low CRP (compared to both high) was shown to predict better overall survival (RR: 0.40395% CI 0.240-0.689 p = 0.0012).

One study reported on preoperative HALP. The study was conducted by Chen and coworkers²⁸ from China. In this multivariate survival analysis on patients with gastric cancer, HALP \geq 56.8 was shown to predict better overall survival (HR: 0.700 95% CI 0.496–0.987 p = 0.042).

One study reported on the combination of the NLR and ESR. The study was conducted by Hyun and coworkers from Korea. Patients were divided into three groups: those with ESR and NLR in the normal range (group 0), those with either elevated ESR or elevated NLR (group I), and those with both elevated ESR and elevated NLR (group II). In this multivariate survival analysis on patients with renal cancer, both elevated ESR and NLR was shown to predict worse overall survival (HR: $3.521\,95\%$ CI $1.888-6.567\,p<0.001$) and worse cancer specific survival (HR: $4.367\,95\%$ CI $1.987-9.597\,p<0.001$).

One study reported on the WLR. The study was conducted by East and coworkers 30 from the UK. In this multivariate survival analysis on patients with colon cancer, WLR \geq 3.4 was shown to predict worse overall survival (HR: 4.10 95% CI 3.13–7.42 p = 0.03).

One study reported on the APRI. The study was conducted by Shen and coworkers³¹ from China. In this multivariate survival analysis on patients with liver cancer, APRI \geq 0.62 was shown to predict worse overall survival (HR: 1.508 95% CI 1.127–2.016 p = 0.006).

One study reported on the combination of the PI, CRP and white cell count (0 if both low, 1 if either high, 2 if both high). The study was conducted by Aurello and coworkers³² from Italy. In this multivariate survival analysis on patients with gastric cancer, PI 2 was shown to predict worse overall survival (HR: 0.37 95% CI 0.16–0.82 p = 0.01).

One study reported on the Canton score involving PNI, NLR and platelet. The study was conducted by Sun and coworkers³³ from China. In this multivariate survival analysis on patients with gastric cancer, elevated Canton score was shown to predict worse overall survival (HR: 1.64395% CI 1.142-2.364 p = 0.007).

One study reported on the AGR. The study was conducted by Li and coworkers³⁴ from China. In this multivariate survival analysis on patients with colorectal cancer, AGR \geq 1.50 was shown to predict better overall survival (HR: 0.646 95% CI 0.543–0.767 p < 0.001).

One study reported on the combination of CRP and neutrophils. The study was conducted by Christina and coworkers³⁵ from Austria. In this multivariate survival analysis on patients with oral cancer, high CRP/ neutrophil was shown to predict worse overall survival (HR: 2.795% CI 0.68-10.75 p = 0.16).

One study reported on the PIS involving a combination of NLR and serum albumin. PIS was defined as follows: patients with increased NLR and decreased serum albumin were assigned score 0; patients with either increased NLR or decreased serum albumin were assigned score 1; patients with decreased NLR and increased serum albumin were assigned score 2. The study was conducted by Wang and coworkers from China. In this multivariate survival analysis on patients with ovarian cancer, PIS 2 was shown to predict better overall survival (HR: 0.1895% CI 0.09-0.38 p < 0.001).

Finally, the last study reported on the CONUT score involving serum albumin concentration, total lymphocyte count and total cholesterol concentration. The study was conducted by Toyokawa *and coworkers*³⁷ from Japan. In this multivariate survival analysis on patients with oesophageal cancer, high CONUT score was shown to predict worse overall survival (HR: 2.303 95% CI 1.191–4.455 p = 0.013).

Assessment of bias using funnel plot analysis of studies carried out in patients with primary operable cancer. Funnel plot analysis containing ten or more studies revealed bias towards studies reporting a relationship between an increased SIR as evidenced by the GPS/GPS (multiple tumour types Figs 2 and 8; colorectal cancer Figs 3 and 9), NLR (multiple tumour types Figs 10 and 18; NLR >5 Fig. 11), PLR (multiple tumour types Figs 22 and 25; PLR >300 Fig. 23), LMR (multiple tumour types Fig. 26) and poorer survival. The funnel plots also showed that a clear majority of studies had high patient numbers. This is particularly true for studies focusing on GPS/mGPS (Figs 2 and 8), NLR (Figs 10 and 18), PLR (Fig. 22) and LMR (Fig. 26).

Discussion

In the present review 244 reports of the prognostic value of systemic inflammation based prognostic scores were identified. This is in contrast to the initial review by Roxburgh and McMillan (2010) where 18 such studies were identified. In particular, those scores based on the ratio of components of a white cell count have been the subject of intense interest with, over the intervening 7 years, 158 studies reporting the value of the NLR, 68 reporting PLR and 21 reporting LMR. Also, the cumulative GPS/mGPS has been the subject of 80 reports. The majority of these studies have been carried out in lung and gastrointestinal cancer. For example, the GPS/mGPS had prognostic value in lung (5 studies), gastric cancer (7 studies), pancreatic (5 studies), and colon cancer (3 studies). A feature of this up to date review of systemic inflammation based prognostic scores is the identification of the proliferation of new scores derived from routinely available markers of the SIR. Most notable among these that have been validated in several studies are PINI (7 studies), COP-NLR (4 studies) and CNP (3 studies). It remains to be established whether any of the scores will have prognostic value in addition to the GPS/mGPS and NLR. Irrespective, there is increasing recognition and acceptance of the clinical utility of systemic inflammation based prognostic scores prior to surgery for cancer.

It is perhaps surprising that, given apparent the superior prognostic value of the GPS/ mGPS³ the relatively larger numbers of reports of the prognostic value of ratios based on components of the white cell count. However, the pre-operative differential white cell count is part of the standard pre-operative workup for the majority of cancer resections as it is used to help identify patients who may have an infection prior to surgery. Also, the white cell count is used to identify any pre-existing conditions that may affect the surgical procedure such as the hypercoagulability of thrombocytosis. Thus, these results are routinely available for retrospective studies. This might also explain the variety of prognostic thresholds reported for NLR, PLR and LMR. In contrast, reports on the prognostic value of the GPS/mGPS, not routinely assessed as part of the standard pre-operative workup, were more likely to be examined in prospective studies. This might explain the consistent adherence to the original thresholds reported for GPS/ mGPS. From the above there is a strong case for the GPS/mGPS to be incorporated into pre-operative workup of patients undergoing surgery for cancer.

It is of interest that while there is general uniformity of thresholds used in the GPS/mGPS studies, with most adhering to the original abnormal thresholds (CRP > 10 mg/l and albumin < 35 g/l), studies in East Asia particularly Japan have used thresholds of 7.5 mg/l³⁸, 5 mg/l^{39,40} and 3 mg/l^{41–43}. Such lower CRP thresholds are above the normal reference ranges in Japan/ East Asia cohorts and results in fewer patients breaching the CRP > 10 mg/l threshold. This observation of a greater proportion of patients with elevated systemic inflammation markers in Western countries compared with Eastern Asian countries is also apparent in white cell derived ratios. Given the objective and reproducible nature of systemic inflammation based prognostic scores it is likely that such observations are real. Indeed, there are recognized ethnic differences in the normal range of neutrophils and lymphocytes^{44–46}. For example, Azab and co-workers recently reported that, in more than 9,000 patients in the United States, there were ethnic differences in the NLR⁴⁶. Specifically, in the cohort as a whole the mean NLR was 2.15. In contrast, black Americans had a mean NLR of 1.76, Hispanic Americans had a mean NLR of 2.08 and white Americans had a mean NLR of 2.24⁴⁶. Also, within ethnicities, patients who had diabetes, cardiovascular disease, a high BMI and were smokers had a significantly higher NLR⁴⁶. Although, similar data for the GPS/mGPS has not yet appeared in the literature it is likely that there would be a similar effect on the GPS/mGPS. Therefore, given that the most common abnormal thresholds used for NLR are >5 and >3 it is likely that a combination of tumour and host genetic and environmental factors are responsible for such consistent East/West differences. These and the present results emphasise the importance of not only staging the tumour but also the host systemic inflammatory response in patients with operable disease⁷.

Recently, studies have directly compared the prognostic value of the two most common combined markers of the systemic inflammatory response, the NLR and the GPS/mGPS. Guthrie and coworkers⁴⁷ reported a comparison in both the preoperative and follow-up settings in patients with resectable colorectal cancer. In this study of 206 patients undergoing a surgical resection at a single institution it was reported that both preoperative mGPS (HR: 1.97, CI 1.16–3.34, p < 0.005) and NLR (HR: 3.07, CI 1.23–7.63, p < 0.05) were independently associated with cancer specific survival⁴⁷. However in the postoperative follow-up only mGPS (HR: 4.81, CI 2.13–10.83, p < 0.001) maintained its significance in terms of cancer specific survival⁴⁷. In contrast, Wang and coworkers (2012) reported that, in 177 patients with pancreatic cancer treated with surgery and palliative chemotherapy, although NLR and mGPS predicted overall survival only NLR was independently associated with overall survival (HR: 2.54 CI 1.31–4.90, p = 0.006)⁴⁸. Finally, Okuno and coworkers (2016) reported that, in 534 patients with perihilar cholangiocarcinoma, both the NLR and mGPS had prognostic value⁴⁹. However, on multivariate analysis, only the mGPS was independently associated with overall survival (HR: 1.58 CI 1.21–2.06, p = 0.001)⁴⁹.

The present review and meta-analysis has a number of limitations^{50–261}. For example funnel plot analysis, even after fixed effect analysis, showed that there was for all systemic inflammation based prognostic scores some asymmetry. This would suggest that there may be some reporting bias. The basis of this bias is not clear. Other than statistically significant results being more likely to be published other possible contributors may be that the studies included in the analysis were English language only publication, had small study size, included multiple

tumour types and included multiple thresholds. Nevertheless the consistency of prognostic value over a variety of systemic inflammation based prognostic scores and across larger studies, single tumour types and single thresholds would indicate that although there was evidence of bias in the meta-analysis, such scores do indeed have prognostic value. Similarly when only univariate analysis was available it was entered into the analysis. The majority of studies had HR derived from multivariate analysis (181 studies) and therefore harmonisation of HR results was not attempted. In the present meta-analysis there was considerable heterogeneity in the HR of some of the markers of the SIR. However, this was less when a consistent threshold for the marker was used. There are other potential contributors to such heterogeneity including geographical location. Such sub-analysis was limited by the number of studies available for meta-analysis. The strength of this present review is its comprehensive nature.

In summary, the results of this review consolidate the prognostic value of combined markers of the systemic inflammatory response including GPS/mGPS NLR, PLR and LMR in patients with resectable cancers. This is particularly true for the GPS/mGPS and NLR and in lung and GI cancers. These should form part of the routine preoperative workup and follow-up for all such patients undergoing resection for cancer.

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Author Contributions

Mr. Dolan and Mr. Lim both contributed equally to the compiling of papers and their statistical analysis as part of this systematic review and meta-analysis. They also both contributed equally to the writing of the final paper. As such they should be classed as joint first authors. Mr. McSorley provided day to day support and feedback for statistical analysis and the final paper. Professors McMillan and Horgan also provided day to day support and made significant input to the final paper.

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