

## Child Weight Growth Trajectory and its Determinants in a Sample of Iranian Children from Birth until 2 Years of Age

Sayed-Mohsen Hosseini, Mohamad-Reza Maracy, Sheida Sarrafzade, Roya Kelishadi<sup>1</sup>

Department of Bio-statistics and Epidemiology,  
School of Health, Isfahan University of Medical  
Sciences, Isfahan, Iran, <sup>1</sup>Department of Pediatrics,  
Faculty of Medicine and Child Growth and  
Development Research Center, Isfahan University  
of Medical Sciences, Isfahan, Iran

### Correspondence to:

Prof. Roya Kelishadi,  
Department of Pediatrics,  
Faculty of Medicine and Child  
Growth and Development  
Research Center, Isfahan University  
of Medical Sciences, Isfahan, Iran.  
E-mail: kelishadi@med.mui.ac.ir

Date of Submission: May 11, 2013

Date of Acceptance: Oct 20, 2013

**How to cite this article:** Hosseini S, Maracy M, Sarrafzade S, Kelishadi R. Child Weight growth trajectory and its determinants in a sample of Iranian children from birth until 2 years of age. *Int J Prev Med* 2014;5:348-55.

### ABSTRACT

**Background:** Growth is one of the most important indices in child health. The best and most effective way to investigate child health is measuring the physical growth indices such as weight, height and head circumference. Among these measures, weight growth is the simplest and the most effective way to determine child growth status. Weight trend at a given age is the result of cumulative growth experience, whereas growth velocity represents what is happening at the time.

**Methods:** This longitudinal study was conducted among 606 children repeatedly measured from birth until 2 years of age. We used linear mixed model to analyze repeated measures and to determine factors affecting the growth trajectory. LOWESS smooth curve was used to draw velocity curves.

**Results:** Gender, child rank, birth status and feeding mode had a significant effect on weight trajectory. Boys had higher weight during the study. Infants with exclusive breast feeding had higher weight than other infants. Boys had higher growth velocity up to age 6 month. Breast fed infants had higher growth velocity up to 6 month, but thereafter the velocity was higher in other infants.

**Conclusions:** Many of the studies have investigated child growth, but most of them used cross-sectional design. In this study, we used longitudinal method to determine effective factors on weight trend in children from birth until 2-year-old. The effects of perinatal factors on further growth should be considered for prevention of growth disorders and their late complications.

**Keywords:** Growth, linear mixed model, velocity, weight

### INTRODUCTION

Child growth is one of the important health indicators from early infancy until puberty,<sup>[1,2]</sup> such that all policymakers and staffs involved in children's health issues should be familiar with growth concepts and its natural limits and hence that diagnoses unnatural cases, determine the causes and help child healthy.<sup>[1]</sup> Physical growth can determine the individual health status. The child with appropriate growth trajectory doesn't have

nutritional disorders, chronic diseases.<sup>[3]</sup> Weight and height growth curves investigation are the most appropriate way to evaluate nutrition status and child growth.<sup>[2,4]</sup> Weight, height and head circumference are common measurements to investigate child growth. Weight is simplest and the most effective measure to determine child health.<sup>[2]</sup>

Both underweight and overweight may affect individual and society health status.<sup>[5]</sup> Underweight can be caused by diseases or social or environmental factors.<sup>[3]</sup> Underweight people are more exposed to loss energy and are prone to injury and infection.<sup>[5]</sup>

Overweight and obesity are other problems in human societies.<sup>[5-7]</sup> Studies in developing countries, Middle East and Iran imply obesity increases in childhood.<sup>[8-11]</sup> As the obesity epidemic, the prevalence of hypertension increased in children. Childhood obesity can increase the risk of adulthood obesity.<sup>[5,6,8]</sup> Obesity itself can increase the risk of disabilities of type II diabetes, hypertension, hyperlipidemia and stroke and may cause some kinds of cancer, infertility and some other problems.<sup>[6,12,13]</sup> World Health Organization in MONICA study which conducted in 1988 has reported Iran as one of the 7<sup>th</sup> countries with highest childhood obesity prevalence.<sup>[8]</sup> Breastfeeding is one of the important strategies to obesity prevention. As the length period of breastfeeding increases, the risk of obesity decreases.<sup>[7]</sup>

There are two main approaches to detecting infants with growth disorder, (1) the comparison of attained growth at a specific age with a reference chart and (2) measuring the growth indices within a time interval expressed as growth velocity.<sup>[14]</sup> Attained weight represents the cumulative growth experience that happened in the past, whereas growth velocity represents what is happening currently.<sup>[14,15]</sup> The growth velocity curve, which modeled with appropriate mathematical function, represents the best way to describe the complex patterns of weight growth during infancy and childhood.<sup>[16]</sup>

Many studies were investigated the effects of different factors on birth weight, but few surveys investigate these effects on their growth trajectory. We studied the effect of different factors on weight growth trajectory in this paper. We also represent weight growth velocity here.

Many researchers used cross-sectional design to investigate child growth. The alternative study design is to use longitudinal approach. The important characteristic in longitudinal studies is that each

subject measured repeatedly over time.<sup>[17,18]</sup> The main advantage of a longitudinal study compared to a cross-sectional study is that the individual development of a certain outcome variable over time can be studied.<sup>[19,20]</sup> Longitudinal studies are also more effective and have more statistical power than cross-sectional studies.<sup>[21]</sup> When we are interested in to study changes over time, we must collect repeated measurements.<sup>[22,23]</sup> In longitudinal data that we measure subjects repeatedly over time, the observations are not independent.<sup>[20,22]</sup> Ignoring this correlation can negatively impact parameter estimation, hypothesis testing and efficiency of study design.<sup>[21]</sup> Therefore, we need special methods which can include this correlation to analysis.

## METHODS

### Population and setting

This longitudinal study was conducted in Isfahan, Iran, among 606 children born in 2007-2008 and followed from birth until the age of 2 years. Demographic information including gender, child rank, birth status (pre-term or full-term), parents' age, parents' education and occupation, exposure to second-hand smoke, family history of chronic diseases such as hypertension and diabetes, as well as parental obesity, were collected. The checklists were filled in by expert health professionals through the interview with one of the infant's parents. Birth weight was documented from the hospital records. Additional repeated measurements for each infant were measured and recorded by trained health staffs in 2, 4, 6, 8, 10, 12, 15, 18 and 24 month of life. Only newborns with Iranian ethnicity and without apparent congenital anomaly were enrolled to the study. Weight was measured in light clothing to the nearest 100 g. The infant's feeding method was recorded as exclusive breastfeeding or a mixed diet (including both breast and formula milk).

Sex, child rank, feeding mode, parent's age, education and occupation, number of family members, using supplement food, tobacco smoke exposure, family history of hypertension, diabetics and obesity, were investigated in this study.

### Statistical analysis

In the current study, we used longitudinal design to study child weight growth. There are many statistical methods to analyze longitudinal data, but

one of the most effective methods is linear mixed model (LMM)<sup>[24]</sup> effects. LMM is an extension of typical linear regression analysis which includes random effects in the structure of the mean rather than the fixed effects.<sup>[17,19,25,26]</sup> LMM was used to determine factors affecting growth trends over time.

To make weight growth velocity, we calculated weight velocities by use of the 2-pointed average weight model 2-pointed model (2-PM). 2-PM is calculated as net weight gain over the time interval (g/kg/d):

$$\text{Growth velocity} = \frac{1000 \times (W_{n+1} - W_n)}{(D_{n+1} - D_n) \times \left[ \frac{W_{n+1} + W_n}{2} \right]}$$

In which *n* is measurement occasion, *W<sub>n</sub>* is weight in grams on measurement “*n*”, and *W<sub>n+1</sub>* is weight in grams on measurement “*n* + 1”, *D<sub>n</sub>* is age in the day on measurement “*n*” and *D<sub>n+1</sub>* is age in the day on measurement “*n* + 1”.<sup>[27]</sup>

Then “LOWESS” method was used to draw weight velocity curves.<sup>[19]</sup>

We used SPSS statistical package for windows (SPSS Inc., Chicago, Illinois) version 18.0 and S-Pluse8 software (Insightful Corp, Seattle, Washington). The level of statistical significance was set at *P* < 0.05.

## RESULTS

Overall, 606 infants were followed from birth up to 2 years of age. They consisted of 307 girls and 299 boys. Of newborns studied, 598 were full-term. The mean (standard deviation [SD]) of mothers’ and fathers’ age was 32.29 (6.33) and 37.22 (6.60) years, respectively. The mean (SD) birth weight was 3090.94 (496.20) and 3226.99 (572.78) g for girls and boys respectively. Birth weight was significantly higher in boys (*P* = 0.002). The characteristics of the study population are summarized in Table 1.

Gender (*P* < 0.0001), child rank in family, gestational age (*P* = 0.005), the type of milk consumed (*P* = 0.022) and maternal education had a significant effect on weight growth trend.

Figure 1a implies the growth curve for boys and girls. Mean weight was higher in boys from birth up to 2-year-old. This difference existed from birth time and became greater over time until 2 years of age. It means that there is interaction between sex and age. The amount of age and sex interaction was significant (*P* < 0.001).

**Table 1:** Demographic characteristics of the study population (*n*=606)

|                               | Frequency | Percent |
|-------------------------------|-----------|---------|
| Gender                        |           |         |
| Boy                           | 299       | 49.3    |
| Girl                          | 307       | 50.7    |
| Child rank in family          |           |         |
| 1                             | 344       | 56.8    |
| 2                             | 189       | 31.2    |
| ≥3                            | 73        | 12.0    |
| Gestational age               |           |         |
| Full-term                     | 598       | 98.7    |
| Pre-term                      | 8         | 1.3     |
| Type of milk consumed         |           |         |
| Breastfeed                    | 528       | 87.1    |
| Mixed milks                   | 78        | 12.9    |
| Exposure to second hand smoke |           |         |
| Yes                           | 135       | 22.3    |
| No                            | 454       | 74.9    |
| ≤6                            | 17        | 2.8     |
| Maternal education (years)    |           |         |
| 6-12                          | 268       | 44.2    |
| 14                            | 250       | 41.3    |
| >14                           | 88        | 14.5    |

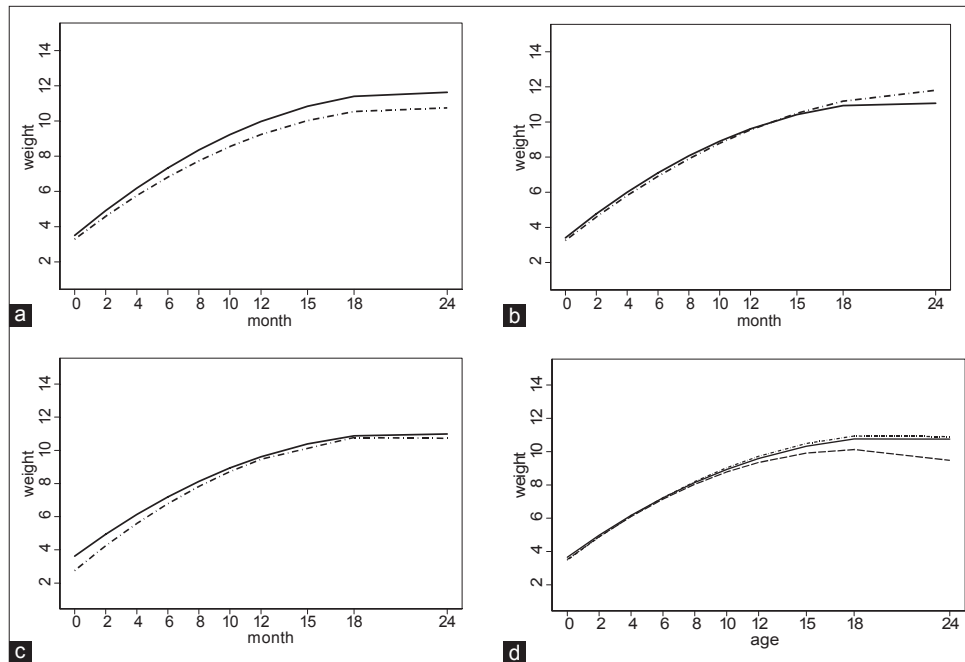
Exclusively breast fed infants had a little higher weight, up to 12 months. Thereafter, other infants had higher mean weight. This pattern implies interaction between age and feeding mode. The amount of interaction was significant (*P* = 0.044). The weight growth curve by feeding mode is shown in Figure 1b.

As depicted in Figure 1c, although pre-term infants had lower mean weight at birth time, but this difference became smaller over time. The interaction between age and birth status was not significant. However, due to the very small number of per-term subgroup, this finding should be regarded with caution.

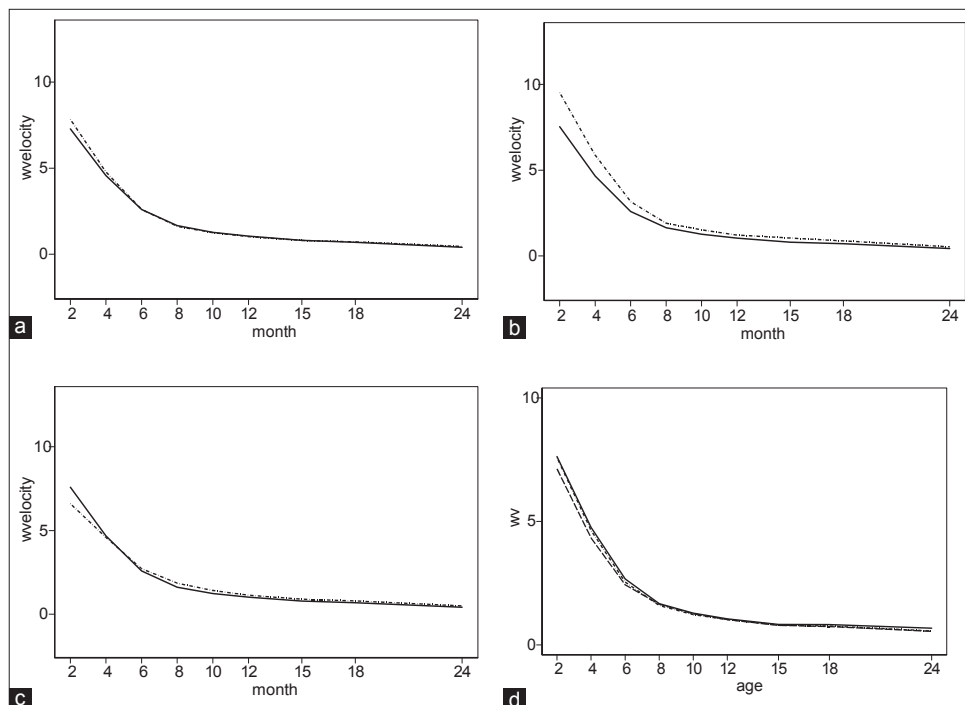
The interaction between age and child rank was not significant. The first child had significantly lower mean weight in comparison with the third rank and upper (*P* = 0.44). However, the difference was not significant between 1<sup>st</sup> and 2<sup>nd</sup> rank children [Figure 1d].

Maternal education was a significant predictor for the variables assessed in this study. Children of mothers with under diploma education had significantly lower mean weight than those with a diploma or higher education (*P* = 0.039).

In this study, parents’ age and occupation, father education, tobacco smoke exposure and family



**Figure 1:** Weight growth trajectory: (a) Weight growth curve by gender. Solid line implies boy curve and dashed line implies girls. (b) Child growth curve by feeding mode. The solid line implies exclusive breast fed infants' curve and dashed line implies mixed milks. (c) Weight growth curve by birth status. Solid line implies full-term infants and dashed line implies pre-term. (d) Weight growth curve by child rank in family. Solid line, dotted line and dashed line imply first, second and third and upper rank respectively



**Figure 2:** Weight velocity curves: (a) Weight velocity curve by sex. Girls curve is in solid and boys curve is in dashed line. (b) Weight velocity curve by birth status. Solid line implies full-term and dashed line implies pre-term infants. (c) Weight velocity curves by feeding mode. Solid line implies breast fed and dashed line implies mixed fed infants. (d) Weight velocity curves by child rank in family. The curves of first and second rank are matched and they can see as a solid line. Dashed line implies third rank and upper

history of hypertension, diabetes and obesity did not have statistically significant effect on the trend of weight gain.

Regression coefficients of the final model for weight are presented in Table 2.

To investigate weight growth velocity, we used “LOWESS” curve. Figure 2a implies weight velocity curves by gender. It shows that boys had higher growth velocity at the beginning and then their velocity came closer to that of girl’s. Figure 2b demonstrates that at birth time, pre-term infants had very higher velocity than their full-term counterparts; and it became close to full-term velocity curve up to 2 years of age.

Figure 2c implies weight growth velocity by feeding mode; It shows that breast fed infants had higher velocity early at birth time. Then it became close to mixed fed infants and after 6 months, they had lower velocity. Weight growth velocity by child rank is shown in Figure 2d.

## DISCUSSION

Growth is a very important event in childhood. Many researchers have investigated child growth,

but most of them used a cross-sectional design for their study. In the current study, we used longitudinal data to study weight trajectory from birth. Longitudinal studies have major advantages in comparison with cross-sectional surveys. We implemented LMM for longitudinal data, so we were able to examine a wide range of factors that might affect changes in weight-for-age during the first 2 years of the child’s life.

Infant gender, child rank, birth status (gestational age), maternal education and feeding mode (breast or mixed feeding), affected the weight changes in this period. Factors as parental age and occupation and father education were not associated with the trend of weight change over time.

Raum *et al.* in an article 2011, studied 1418 infants with both cross-sectional and longitudinal methods, and used LMM to analyze longitudinal data.<sup>[28]</sup> Karaolis-Danckert *et al.* in 2008 used mixed models to determine the relationship between weight velocity and body fat percent.<sup>[29]</sup> Regnault *et al.* has modeled the growth by using mixed model.<sup>[30]</sup> Heydari in 2009 used a longitudinal method to study weight and height growth in child under 2 years of age.<sup>[31]</sup>

**Table 2:** Model coefficients of weight growth

| Variable           | Estimate | Standard error | df   | t value | Lower   | Upper   | P value |
|--------------------|----------|----------------|------|---------|---------|---------|---------|
| Intercept          | 2.86     | 0.232          | 3366 | 12.31   | 2.41    | 3.32    | <0.0001 |
| Age                | 0.73     | 0.008          | 3366 | 96.02   | 0.71    | 0.74    | <0.0001 |
| I (age^2)          | -0.02    | 0.0003         | 3366 | -54.37  | -0.0167 | -0.0161 | <0.0001 |
| Gender             |          |                |      |         |         |         |         |
| Girl               | -0.26    | 0.056          | 598  | -4.54   | -3.67   | -0.15   | <0.0001 |
| Boy                | -        | -              | -    | -       | -       | -       | -       |
| Child rank         |          |                |      |         |         |         |         |
| Third and upper    | 0.23     | 0.087          | 598  | 2.63    | 0.06    | 0.40    | 0.009   |
| Second             | 0.15     | 0.061          | 598  | 2.51    | 0.03    | 0.27    | 0.013   |
| First              | -        | -              | -    | -       | -       | -       | -       |
| Gestational age    |          |                |      |         |         |         |         |
| Full-term          | 0.67     | 0.23           | 598  | 2.91    | 0.22    | 1.12    | 0.004   |
| Pre-term           | -        | -              | -    | -       | -       | -       | -       |
| Feeding mode       |          |                |      |         |         |         |         |
| Breast fed         | -        | -              | -    | -       | -       | -       | -       |
| Mixed diet         | -0.27    | 0.117          | 598  | -2.30   | -0.50   | -0.04   | 0.022   |
| Maternal education |          |                |      |         |         |         |         |
| 6-12               | -0.12    | 0.059          | 598  | -2.07   | -0.24   | -0.01   | 0.039   |
| 14                 | 0.03     | 0.081          | 598  | 0.33    | -0.13   | 0.19    | 0.741   |
| >14 years          | -        | -              | -    | -       | -       | -       | -       |
| Age: Gender        | -0.04    | 0.007          | 3366 | -5.13   | -0.05   | -0.02   | <0.0001 |
| Age: Feeding mode  | 0.04     | 0.015          | 3366 | 2.83    | 0.01    | 0.07    | 0.005   |

P value<0.05 is significant



Gender had a significant effect on weight growth trend, such that boys had higher mean weight from birth up to 2 years of age. In the study of Regnault *et al.* in 2010, boys were heavier than girls at birth time and they had higher weight velocity in 1 and 3 month of life.<sup>[30]</sup> Heydari *et al.* in 2005 studied 597 infants from birth to 2-year-old. In their study, boys mean weight was higher than girls, except for 1<sup>st</sup> month. But the mean weight difference between the boy and girl was not statistically significant before month 2 and at year 2.<sup>[32]</sup>

In our study, maternal education had a significant effect on weight growth trend; such that the birth weight of mothers with education lower than 14 years had significantly lower than children of mothers with higher education. In Raum *et al.* study, mother's education was significant before adjustment with other variables, but after adjustment it did not have a significant effect.<sup>[28]</sup> In another study, illiteracy of mothers was documented as a risk factor for child underweight.<sup>[33]</sup>

The family history of hypertension, diabetes and obesity did not have a significant effect on the child weight growth trend. In the study of Kalantari *et al.*, family history of overweight and obesity had a significant effect on obesity in children.<sup>[7]</sup> In the study of Regnault *et al.*, maternal obesity had a significant effect on birth weight, but this relationship was not significant at months 1 and 3 and paternal obesity did not have a significant effect on birth weight.<sup>[30]</sup>

Many studies documented maternal smoking during pregnancy as a risk factor of low birth weight and future obesity in their children. Raum *et al.* and Karaolis-Danckert *et al.* implied that smoking before and after pregnancy increases the risk of obesity in children.<sup>[28,29]</sup> In another study, the infants of a mother who smoked a cigarette during pregnancy had lower weight between births up to 3<sup>th</sup> month.<sup>[30]</sup> In our study, the effect of exposure to tobacco smoke was not significant on weight change trend. Likewise, it did not have any significant effect on birth weight.

In our study, exclusive breast fed infants had a little higher mean weight before the age of 12 months; however after this time, mixed diet infants had higher mean weight. This trend can implies the higher risk of childhood obesity in infants fed with other types of milk. Our findings are consistent with some previous studies.

Kelishadi *et al.* in 2005 observed that the frequency and duration of breast feed was lower in obese children.<sup>[34]</sup> Raum *et al.* showed that breast feeding decrease child obesity risk in 6-year-old children.<sup>[28]</sup>

In our study, the child rank in the family had a significant effect on the weight growth trend. This finding is in line with some previous studies in showing that firstborn children may be smaller than their counterparts.<sup>[30]</sup> Karaolis-Danckert *et al.* implied that being the first child of the family may increase the risk of childhood obesity.<sup>[29]</sup> However in the study of Raum *et al.*, the birth rank had no significant effect on future overweight.<sup>[28]</sup>

In the present study, the growth pattern of preterm newborns was significantly different from their full-term counterparts, but as described before, the result should be regarded with caution because the number of pre-term infants was much smaller than full-term babies.

We used the 2-pointed weight model to make velocity curves. Many studies used this model to make weight velocity.<sup>[14,27]</sup> As Patel *et al.* implied, 2-pointed model provided an accurate estimate of growth velocity.<sup>[27]</sup>

The effects of early life growth and incidence of diseases in later life have been studied in many studies.<sup>[35-39]</sup>

Determining factors affecting early growth can be useful in policy making for prevention of short-term and long-term complications of growth disorders.

## CONCLUSIONS

Given the long-term sequelae of growth disorders and the documented effects of early life factors on further growth, primordial preventive measures should be considered for the modifiable factors affecting the child growth.

## REFERENCES

1. Nasiri-Rine H, Salar-Kia N. Weight and height in children under 5 year in Tonekabon, Iran. *J Babol Univ of Med Sci* 2004;6:55-9.
2. Taghavi N, Ebrahimi H, Karimi A, Pourheydari M. Effects of formaldehyde on morphometric structure of testis in Balb/C mice. *Med Sci J Islamic Azad Univ Tehran Med Branch* 2007;17:95-101.
3. Vakili R. *Growth and its Disorders*. Mashhad: Mashhad University of Medical Sciences; 2003.

4. Emdadi-Fard M, Safarian M, Doosti H, Shakeri MT, Fazaeli M, Abasalti Z. Standardized percentile curves of height versus the age of Iranian children aged 25-60 months living in the northeast of Iran. *Ofogh-e-Danesh J* 2010;16:36-44.
5. Mahan LK, Escott-Stump S. *Krause's Food & Nutrition Therapy*. St. Louis, Mo: Saunders /Elsevier; 2008.
6. Jedrychowski W, Maugeri U, Kaim I, Budzyn-Mrozek D, Flak E, Mroz E, *et al.* Impact of excessive gestational weight gain in non-smoking mothers on body fatness in infancy and early childhood. Prospective prebirth cohort study in Cracow. *J Physiol Pharmacol* 2011;62:55-64.
7. Kalantari N, Shenavar R, Rashid-Khani B, Hoshar-Rad A, Knasihat-Kon A, Abdolazade SM. Relationship of obesity and overweight in school children in Shiraz with pattern of breastfeeding, birth weight and socioeconomic status. *O'loum-e taghziye Va Sanaye Ghazayi Iran* 2009;5:19-28.
8. Kelishadi R, Hashempour M, Sarafzadegan N, Sadri G, Ansari R, Alikhasi H, *et al.* Prevalence of overweight and obesity in adolescents and its relationship with environmental factors (Isfahan 1380). *J Med Sch Tehran Univ Med Sci* 2003;61:260-73.
9. Whitaker RC, Wright JA, Pepe MS, Seidel KD, Dietz WH. Predicting obesity in young adulthood from childhood and parental obesity. *N Engl J Med* 1997;337:869-73.
10. Willows ND, Johnson MS, Ball GD. Prevalence estimates of overweight and obesity in Cree preschool children in northern Quebec according to international and US reference criteria. *J Inf* 2007;97:311-6.
11. World Health Organization. Programme of Nutrition F, Health R. *Obesity: Preventing and Managing the Global Epidemic: Report of a WHO Consultation on Obesity, 3-5 June 1997*. Geneva: World Health Organization; 1998.
12. Ilyas M, Ellis EN. Management of childhood hypertension: A guide for primary care physicians. *J Ark Med Soc* 2006;103:137-40.
13. Asl MV, Nanbakhsh Z, Mehdinejad R. Incidence of obesity, hyperlipidemia, hyperglycemia in NIDDM patients in Urmia diabetic center. *J Urmia Nurs Midwifery Fac* 2006;4:112-6.
14. Olusanya BO, Renner JK. Predictors of growth velocity in early infancy in a resource-poor setting. *Early Hum Dev* 2011;87:647-52.
15. Botton J, Heude B, Maccario J, Ducimetière P, Charles MA, FLVS Study Group. Postnatal weight and height growth velocities at different ages between birth and 5 y and body composition in adolescent boys and girls. *Am J Clin Nutr* 2008;87:1760-8.
16. Bertino E, Coscia A, Boni L, Rossi C, Martano C, Giuliani F, *et al.* Weight growth velocity of very low birth weight infants: Role of gender, gestational age and major morbidities. *Early Hum Dev* 2009;85:339-47.
17. Diggle P. *Analysis of Longitudinal Data*. USA: Oxford University Press; 2002.
18. Lynn P. *Methodology of Longitudinal Surveys*. Hoboken, New Jersey, US: Wiley; 2009.
19. Fitzmaurice GM. *Longitudinal Data Analysis*. Boca Raton, Florida, US: Chapman and Hall/CRC; 2009.
20. Twisk JW. *Applied Longitudinal Data Analysis for Epidemiology: A Practical Guide*. Cambridge, England: Cambridge Univ. Press; 2003.
21. Edwards LJ. Modern statistical techniques for the analysis of longitudinal data in biomedical research. *Pediatr Pulmonol* 2000;30:330-44.
22. Nakai M, Ke W. Statistical models for longitudinal data analysis. *Appl Math Sci* 2009;3:1979-89.
23. Verbeke G, Molenberghs G. *Linear Mixed Models for Longitudinal Data*. New York City, US: Springer Verlag; 2009.
24. Norleans MX. *Statistical Methods for Clinical Trials*. Boca Raton, Florida, US: CRC; 2000.
25. Antonio K, Beirlant J. Actuarial statistics with generalized linear mixed models. *Insur Math Econ* 2007;40:58-76.
26. Kincaid C, editor. *Guidelines for Selecting the Covariance Structure in Mixed Model Analysis*. Vol. 30. SUGI; 2005.
27. Patel AL, Engstrom JL, Meier PP, Jegier BJ, Kimura RE. Calculating postnatal growth velocity in very low birth weight (VLBW) premature infants. *J Perinatol* 2009;29:618-22.
28. Raum E, Küpper-Nybelen J, Lamerz A, Hebebrand J, Herpertz-Dahlmann B, Brenner H. Tobacco smoke exposure before, during, and after pregnancy and risk of overweight at age 6. *Obesity (Silver Spring)* 2011;19:2411-7.
29. Karaolis-Danckert N, Buyken AE, Kulig M, Kroke A, Forster J, Kamin W, *et al.* How pre- and postnatal risk factors modify the effect of rapid weight gain in infancy and early childhood on subsequent fat mass development: Results from the Multicenter Allergy Study 90. *Am J Clin Nutr* 2008;87:1356-64.
30. Regnault N, Botton J, Forhan A, Hankard R, Thiebaugeorges O, Hillier TA, *et al.* Determinants of early ponderal and statural growth in full-term infants in the EDEN mother-child cohort study. *Am J Clin Nutr* 2010;92:594-602.
31. Heydari ST. Infants' growth charts in Jahrom, Iran. *Iran J Pediatr* 2009;19:25-34.
32. Heydari ST, Emamghoreishi F, Amini M. A comparative study on growth state of children less than two-years old in Jahrom, Southeastern Iran with NCHS measurements. *Ofogh-e-Danesh J* 2005;11:42-8.
33. Kabiri M, Parsinia M, Goodarzi M, Babayi GR. Relation

between physical growth of 0-2 year-old children and socioeconomic and educational situation of their parents in Karadj/Iran. *Iran J Pediatr* 2003;13:47-52.

34. Kelishadi R, Hashemipour M, Famouri F, Sabet B, Sanei M. The impact of breast feeding in prevention of obesith in children. *J Qazvin Univ Med Sci* 2005;35:88-93.
35. Amiri M. Early life conditions and trends in mortality at later life: Is there any relationship? *Int J Prev Med* 2011;2:53-5.
36. Barker DJ. Maternal nutrition, fetal nutrition, and disease in later life. *Nutrition* 1997;13:807-13.
37. Ben-Shlomo Y, Kuh D. A life course approach to chronic disease epidemiology: Conceptual models, empirical challenges and interdisciplinary perspectives. *Int J Epidemiol* 2002;31:285-93.
38. Hanson M, Gluckman P, editors. *Fetal Matrix: Evolution, Development and Disease*. Cambridge, England: Cambridge University Press; 2004.
39. Lawlor DA, Smith GD. Early life determinants of adult blood pressure. *Curr Opin Nephrol Hypertens* 2005;14:259-64.

**Source of Support:** Study was conducted as a thesis supported by Vice-chancellery for Research, Isfahan University of Medical Sciences, Isfahan, Iran, **Conflict of Interest:** None declared.

### Staying in touch with the journal

#### 1) Table of Contents (TOC) email alert

Receive an email alert containing the TOC when a new complete issue of the journal is made available online. To register for TOC alerts go to [www.ijpm.in/signup.asp](http://www.ijpm.in/signup.asp).

#### 2) RSS feeds

Really Simple Syndication (RSS) helps you to get alerts on new publication right on your desktop without going to the journal's website. You need a software (e.g. RSSReader, Feed Demon, FeedReader, My Yahoo!, NewsGator and NewzCrawler) to get advantage of this tool. RSS feeds can also be read through FireFox or Microsoft Outlook 2007. Once any of these small (and mostly free) software is installed, add [www.ijpm.in/rssfeed.asp](http://www.ijpm.in/rssfeed.asp) as one of the feeds.