BMJ Open Opioid-sparing effects of medical cannabis or cannabinoids for chronic pain: a systematic review and meta-analysis of randomised and observational studies

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ABSTRACT

Objective To assess the efficacy and harms of adding medical cannabis to prescription opioids among people living with chronic pain.

Design Systematic review.

Data sources CENTRAL, EMBASE and MEDLINE.

Main outcomes and measures Opioid dose reduction, pain relief, sleep disturbance, physical and emotional functioning and adverse events.

Study selection criteria and methods We included studies that enrolled patients with chronic pain receiving prescription opioids and explored the impact of adding medical cannabis. We used Grading of Recommendations Assessment, Development and Evaluation to assess the certainty of evidence for each outcome.

Results Eligible studies included five randomised trials (all enrolling chronic cancer-pain patients) and 12 observational studies. All randomised trials instructed participants to maintain their opioid dose, which resulted in a very low certainty evidence that adding cannabis has little or no impact on opioid use (weighted mean difference (WMD) -3.4 milligram morphine equivalent (MME); 95% CI (CI) -12.7 to 5.8). Randomised trials provided high certainty evidence that cannabis addition had little or no effect on pain relief (WMD -0.18 cm; 95% Cl -0.38 to 0.02; on a 10 cm Visual Analogue Scale (VAS) for pain) or sleep disturbance (WMD -0.22 cm; 95% Cl -0.4 to -0.06; on a 10 cm VAS for sleep disturbance; minimally important difference is 1 cm) among chronic cancer pain patients. Addition of cannabis likely increases nausea (relative risk (RR) 1.43; 95% CI 1.04 to 1.96; risk difference (RD) 4%, 95% CI 0% to 7%) and vomiting (RR 1.5; 95% CI 1.01 to 2.24; RD 3%; 95% CI 0% to 6%) (both moderate certainty) and may have no effect on constipation (RR 0.85: 95% Cl 0.54 to 1.35; RD -1%; 95% CI -4% to 2%) (low certainty). Eight observational studies provided very low certainty evidence that adding cannabis reduced opioid use (WMD -22.5 MME; 95% CI -43.06 to -1.97).

Conclusion Opioid-sparing effects of medical cannabis for chronic pain remain uncertain due to very low certainty evidence.

PROSPERO registration number CRD42018091098.

Strengths and limitations of this study

- This is the first meta-analysis to pool the results of randomised controlled trials and observational studies exploring the opioid-sparing effects of medical cannabis among people living with chronic pain.
- We conducted a comprehensive search for eligible studies, appraised the risk of bias of included studies and evaluated the certainty of evidence using the Grading of Recommendations Assessment, Development and Evaluation approach.
- Most observational studies incorporated inadequate adjustment for confounding, and all randomised trials, despite reporting this outcome, were not designed to address the effect of medical cannabis on opioid use.

INTRODUCTION

Chronic pain affects approximately one in five adults and is a common reason for seeking medical care.^{1 2} Opioids are commonly prescribed for this condition, particularly in North America;³ however, they only provide benefit to a minority of patients. A 2018 systematic review of 96 trials found high certainty evidence that, versus placebo, opioids provide important pain relief (≥1 cm improvement on a 10 cm Visual Analogue Scale (VAS) for pain) to 12% of patients for whom they are prescribed.⁴ Moreover, opioids are associated with harms such as overdose and death,⁵⁶ which are dose dependent.⁷⁻¹⁰ As a result, there is considerable interest in therapies that may allow patients with chronic pain using opioid therapy to reduce their opioid intake.

One promising approach is adding cannabis therapy, which low certainty evidence suggests may be similarly effective to opioids for reducing pain and improving physical functioning among people living with chronic pain.⁴ Experimental studies have shown that opioids and cannabis have similar signal transduction systems,¹¹ and observational studies in the USA demonstrated that the rates of opioid-related mortality reduced after cannabis was legalised.¹²⁻¹⁴ Between 64% and 77% of patients with chronic pain responding to cross-sectional surveys reported a reduction in long-term opioid use after adding medical cannabis to their treatment.^{15 16} A 2017 systematic review concluded that preclinical studies provided robust evidence for the opioid-sparing effects of cannabis.¹⁷ To clarify the issue, we undertook a systematic review of randomised controlled trials (RCT) and observational studies to explore the impact of adding medical cannabis on opioid dose, other patient-important outcomes and related harms in patients with chronic pain using prescribed opioid therapy.

This systematic review is part of the BMJ Rapid Recommendations project, a collaborative effort from the MAGIC Evidence Ecosystem Foundation (www.magicevidnece.org) and BMJ. This systematic review informed a parallel guideline published on BMJ.com¹⁸ and MAGICapp (https://app.magicapp.org/#/guideline/ jMMYPj).

METHODS

We followed standards for Meta-analysis Of Observational Studies in Epidemiology¹⁹ and Preferred Reporting Items for Systematic Reviews and Meta-Analyses guidelines.²⁰

Eligibility criteria

We included RCTs and observational studies, including cohort studies and case-control studies, in any language, that explored the impact of adding medical cannabis (ie, phytocannabinoids, endocannabinoids or synthetic cannabinoids) on the use of prescription opioids among people living with chronic pain. We defined pain as chronic if patients reported that symptoms had persisted for \geq 3 months.²¹ We excluded editorials, letters to the editor, preclinical studies, conference abstracts, case reports, case series, cross-sectional studies and studies with less than 2 weeks follow-up. We also excluded studies of recreational cannabis use as these products typically contain much higher amounts of the psychotropic cannabinoid tetrahydrocannabinol (THC) than would be administered for therapeutic purposes.^{22 23} We classified observational study designs according to recommendations by the Cochrane Observational Studies Methods Group.²⁴

Literature search and study selection

We searched the Cochrane Central Register of Controlled Trials (CENTRAL), EMBASE and MEDLINE from inception to March 2020 with no restriction on language of publication. An experienced medical librarian (RJC) developed our database-specific search strategies (online supplemental appendix A). We also searched the ClinicalTrials.gov registry to identify ongoing trials, and reference lists of all eligible studies and related systematic reviews for additional eligible studies. Two teams of paired reviewers independently screened titles, abstracts and full-text studies for eligibility using online systematic review software (Rayyan QCRI, Qatar Computing Research Institute). Reviewers resolved disagreements through discussion.

Data collection

Using standardised forms and a detailed instruction manual, pairs of reviewers independently abstracted data from each eligible study, including study and patient characteristics, and details of treatment (eg, dose, formulation and duration of cannabis add-on therapy). Our primary outcome was opioid dose. We also captured all patient-important outcomes, as guided by the Initiative on Methods, Measurement, and Pain Assessment in Clinical Trials,²⁵ including pain relief, sleep disturbance, physical and emotional functioning. Regarding adverse events, we focused on vomiting, nausea and constipation as a systematic review of values and preferences²⁶ demonstrated that patients living with chronic pain experience gastrointestinal complaints as the most important opioidinduced adverse events. We contacted authors to obtain unpublished data.

Risk of bias assessment

Following training and calibration exercises two independent reviewers used a modified Cochrane risk of bias tool^{27 28} to assess the risk of bias among eligible RCTs according to the following domains: allocation concealment, blinding of participants, study personnel, outcome assessors and data analyst, and lost to follow-up ($\geq 20\%$) missing data were assigned high risk of bias). Response options for each item were 'definitely or probably yes' (assigned a low risk of bias) and 'definitely or probably no' (assigned a high risk of bias) (online supplemental table 1). We used criteria suggested by the CLARITY group²⁹ to assess the risk of bias of observational studies including selection bias, confidence that all patients had the condition of interest, control for confounding variables, validity of outcome assessment(s), and infrequent missing data (<20%) (details available at www.evidencepartners.com/resources/methodological-resources/). (online supplemental tables 2-3).

Data analysis

We calculated inter-rater agreement regarding the eligibility of full-text studies using an adjusted kappa (κ) statistic.³⁰ We conducted separate analyses for RCTs and observational studies. All continuous measures for pain intensity and sleep disturbance were converted to a 10 cm VAS; the minimally important difference (MID) for both was 1 cm.^{31 32} All continuous outcomes that were reported by more than one study were pooled to derive the weighted mean difference (WMD) and associated 95% CI. We pooled binary outcomes (adverse events) as relative risks (RRs) and risk differences (RDs) and their

associated 95% CIs. We conducted all meta-analyses with random-effects models and the DerSimonian-Laird method. 33

When studies reported effects on continuous outcomes as the median and IOR, we derived the mean and SD using the method presented by Wan et al.³⁴ We also converted medians to means using the approach recommended by the Cochrane Handbook as a sensitivity analysis. When authors failed to report a measure of precision associated with mean differences, we imputed the SD from eligible studies that reported these measures (online supplemental technical appendix).³⁵ We included each comparison reported by multiarm studies and calculated a correction factor to account for the unit of analysis error (ie, when information from a treatment arm is used more than once in the same meta-analysis).³⁶ We explored the consistency of association between our pooled results and studies reporting the same outcome domains that were not possible to pool. We used Stata (StataCorp, Release V.15.1) for all analyses.Comparisons were two tailed using a threshold of $p \le 0.05$.

Subgroup analyses and meta-regression

We examined heterogeneity among pooled RCTs using the I^2 statistic, and through visual inspection of forest plots for pooled observational data, because statistical tests of heterogeneity can be misleading when sample sizes are large and associated confidence intervals are therefore narrow.³⁷ When we had at least two studies in each subgroup, we explored sources of heterogeneity with five prespecified subgroup hypotheses, assuming greater benefits with: (1) shorter versus longer duration of follow-up; (2) higher versus lower risk of bias; (3) enriched versus non-enriched study design; (4) chronic non-cancer versus chronic cancer-related pain and (5) higher versus lower THC content. We assumed similar directions of subgroup effects for harms, except for study design and THC content in which we expected greater harms with non-enriched trials and higher THC content. However, apart from item two (risk of bias), studies did not report sufficient data to undertake subgroup analyses.

The certainty of the evidence

We used the Grading of Recommendations Assessment, Development and Evaluation (GRADE) approach to assess the certainty of evidence on an outcome-byoutcome basis as high, moderate, low or very low.³⁸ With GRADE, RCTs begin as high-certainty evidence, but can be rated down because of risk of bias, imprecision, inconsistency, indirectness or publication bias. We rated down for imprecision if the 95% CI associated with a pooled continuous outcome included half the MID, or if the estimate of precision associated with the RR for binary outcomes included no effect. We considered an I² value between 75% and 100% to represent considerable inconsistency.³⁹ We rated down the certainty of evidence for indirectness if there were important differences between our research question and the patients enrolled, intervention tested

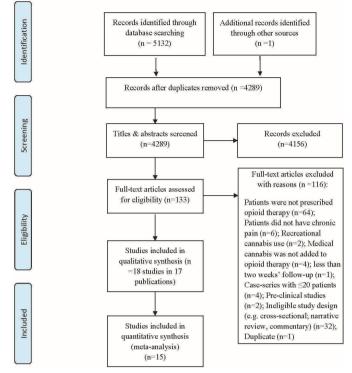


Figure 1 Study selection process in review of opioidsparing effects of cannabis in chronic pain.

or outcomes reported among studies contributing to our meta-analyses. 40

Using GRADE, observational studies begin as low certainty evidence, and while they can be rated down further for the same reasons as RCTs, they can also be rated up in the presence of a large magnitude of the effect, a dose–response gradient or consideration of plausible confounders or other biases that increase confidence in the estimated effect.⁴¹ We only reported the pooling results of observational studies when they resulted in the same or higher certainty of evidence than evidence from RCTs. When there were at least 10 studies for meta-analysis, we explored for small-study effects by visual assessment of funnel plot asymmetry and Egger's statistical test.⁴²

Patients and public involvement

Patients and public were not involved in this research.

RESULTS

Of 5133 records identified, we reviewed 133 articles in full text, and 18 studies reported in 17 publications proved eligible (figure 1, online supplemental appendix B); five RCTs in four publications^{43–46} and 13 observational studies.^{47–59} One study enrolled a mixed group of opioid and non-opioid users⁵⁰; however, our attempts to contact the authors to acquire pain intensity data for the subgroup of patients prescribed opioids proved unsuccessful. All five RCTs^{43–46} and three observational studies^{51 54 55} enrolled patients with chronic cancer-related pain; the remaining 10 observational studies explored adding

cannabis to opioids for patients with chronic non-cancer pain (eg, chronic low back pain, fibromyalgia, painful chronic pancreatitis),^{47 52 53 57-59} or a mix of cancer and non-cancer pain (table 1).^{48-50 56}

Among the 18 included studies, the percentage of female participants was 48% (median of individual trials 48%, IQR 43%-58%), and the median of the mean age was 56.3 (IOR 51.2-59.9). Follow-up ranged from 2 to 5 weeks among RCTs, and from 4 weeks to 6.4 years for observational studies. Only one RCT43 used an enrichment design (following the open-label phase, patients with at least 15% improvement in pain were randomised to the intervention and control groups) and all RCTs advised patients to maintain stable doses of all other prescribed pain medications, including opioids, during the study period (table 1). All included RCTs, and three of the observational studies^{48 51 52} administered synthetic cannabis products (ie, nabilone, dronabinol and nabiximole), five observational studies49 50 56 58 59 reported different combinations of THC:Cannabidiol (CBD) products, and six other observational studies^{47 53-55 57} did not provide details on the type of cannabis or cannabinoids provided (table 1, online supplemental table 4). Ten studies reported receiving industry funding, 43-46 49 51 52 57 58 five studies^{50 53-56} reported no-industry funding and three studies⁴⁷⁴⁸⁵⁹ did not report funding information (table 1).

Risk of bias of included studies

All included RCTs reported adequate allocation concealment and blinding of patients and healthcare providers; however, three trials^{43 45 46} were at risk of bias due to high lost to follow-up (online supplemental table 5). Each RCT specified that they employed an intention-to-treat analysis. All observational studies were at high risk of bias, typically due to lack of confidence in the assessment of exposure, non-representative samples and insufficient control for confounding (online supplemental file 6–7).

Outcomes for medical cannabis add-on therapy Opioid dose reduction

The primary limitation of RCTs was that all investigators instructed patients to not alter their dose of opioids. This represents a very serious indirectness of the findings regarding the research question, warranting rating down two levels, and was the primary reason for very low certainty evidence from the 1176 patients.^{43–45} Their results raised the possibility that adding medical cannabis may not be associated with a reduction in opioid use among patients living with chronic cancer pain (WMD –3.4 milligram morphine equivalent (MME); 95% CI –12.7 to 5.9; table 2; online supplemental figure 1). There were no differences in effect based on the lost to follow-up (online supplemental figure 2); test of interaction p=0.758).

Very-low certainty evidence from eight observational studies (seven of which enrolled people with chronic non-cancer pain)⁴⁷⁴⁸⁵⁰⁵¹⁵³⁻⁵⁵⁵⁸ raised the possibility that adding medical cannabis may reduce the use of opioids among

patients with predominantly chronic non-cancer pain (WMD -22.5 MME; 95% CI -43.06 to -1.97; table 2; online supplemental figure 3). Three observational studies that could not be pooled, as they only reported opioid reduction as a percentage, also found that providing medical cannabis allowed patients to decrease their opioid dose. The first study assessed the impact of providing medical cannabis to 61 patients with chronic low back pain who were prescribed opioid therapy (median opioid dose was 21 mg MME/day) and reported that 52% of patients (32 of 61) stopped all use of opioids at a median follow-up of 6.4 years.⁵⁷ The second study⁴⁹ reported that of 94 patients with chronic pain (both cancer and non-cancer pain) who began using CBD hemp extract, 53% were able to decrease their use of prescription opioids at 8 weeks. A third study⁵⁶ included 600 patients with chronic pain who indicated willingness to taper their opioid dose and were administered 0.5g daily of medicinal cannabis for each 10% reduction in opioid dose. After 6 months follow-up, 55% of patients reported a 30% reduction in opioid dose on average and 26% of them discontinued opioid use.

Pain relief

High-certainty evidence from five RCTs^{43–46} demonstrated that adding medical cannabis to opioid therapy resulted in trivial or no difference in cancer-related pain (WMD -0.18 cm; 95% CI -0.38 to 0.02 on the 10 cm VAS for pain; MID 1 cm; table 2; online supplemental figure 4). Results did not differ depending on lost to follow-up (online supplemental figure 5, a test of interaction p=0.623). Very low certainty evidence from observational studies suggested a large decrease in pain when medical cannabis was added to opioids (online supplemental figure 6).

Sleep disturbance

Five RCTs^{43–46} provided high certainty evidence that adding medical cannabis to prescription opioids results in a trivial improvement in sleep disturbance in people living with cancer-related chronic pain (WMD –0.22 cm; 95% CI –0.4 to –0.06 on the 10 cm VAS for sleep disturbance; MID 1 cm; table 2; online supplemental figure 7). Results did not differ between trials reporting the low and high lost to follow-up (online supplemental figure 8, a test of interaction p=0.93). Very low certainty evidence from observational studies suggested an improvement in sleep disturbance when medical cannabis was added to opioids (online supplemental table 8).

Other reported outcomes

A single RCT^{44} reported moderate certainty evidence that adding cannabis likely has little or no effect on emotional and physical functioning (online supplemental tables 9-10).

Adverse events

Nausea, vomiting or constipation

Four RCTs^{43–46} provided moderate certainty evidence that adding medical cannabis to opioid therapy likely increases the incidence of nausea (RR 1.43, 95% CI 1.04 to 1.96; RD

Table 1 Chara	cteristics of inclu	Characteristics of included studies (n=18)								
Author-year (country)	Study design	No of participants (% prescribed opioids)	Type of chronic pain (specific condition)	Age mean (SD)	% Female	Baseline opioid dose	Follow-up duration	Medical cannabis dose	Medical Analgesic cannabis dose cointervention	Funding source
Fallon, 2017 study I (multicentre trial") ¹³	Parallel arm RCT	n=399; nabiximols (n=20), placebo (n=199) (100%)	100% chronic cancer pain	59.8 (10.9)	43%	Receiving opioid therapy of <500 MME/day (nabiximols group: 199 MME/day±131; placebo group: 207 MME/day±135)	5 weeks	THC 27 mg/mL; CBD 25 mg/ mL (maximum allowed daily dosage of 10 sprays)	Patients were excluded if they planned to undergo clinical interventions that would affect pain	Otsuka Pharmaceutical
Fallon, 2017 study II (multicentre trial") ⁴³	Parallel arm RCT	n=206; nabiximols (n=103), placebo=103 (100%)	100% chronic cancer pain	61.5 (11.3)	49%	Receiving opioid therapy of <500 MME/day (nabiximols: 212 MME/day±136; placebo: 209 MME/ day±121)	5 weeks	THC 27 mg/mL; Patients were CBD 25 mg/ excluded if mL (maximum they planned t allowed daily undergo clinic dosage of 10 interventions t sprays) would affect p	Patients were excluded if they planned to undergo clinical interventions that would affect pain	Otsuka Pharmaceutical
Johnson, 2010 (multicentre trial!) ⁴⁴	Parallel arm RCT	n=177; THC: CBD extract (n=60), THC extract (n=58), placebo (n=59) (100%)	100% chronic cancer pain	60.2 (12.3)	46%	Receiving opioid therapy for at least 1 week before enrolment (THC:CBD: 258MME/ 258MME/ 188 MME±234; placebo: 367±886)	2 weeks	One spray: 2.7mg THC/2.5mg CBD. TBD. The maximum permitted dose: eight actuations over 3 hours and 48 actuations over 24-hours	Patients were excluded if they planned to undergo clinical interventions that would affect pain	GW Pharmaceuticals
Lichtman, 2018 (multicentre*) ⁴⁵	Parallel arm RCT	n=398; nabiximol (n=199), placebo (n=198) (100%)	100% chronic cancer pain	60 (11.5)	46%	Receiving opioid therapy of <500 MME/day (nabiximols: 193 MME/day±130; placebo: 186 MME/ day±131)	5 weeks	THC 27 mg/mL; CBD 25 mg/ mL (maximum allowed daily dosage of 10 sprays per day)	Patients were excluded if they planned to undergo clinical interventions that would affect pain	Otsuka Pharmaceutical
Portenoy, 2012 (multicentre [*]) ⁴⁶	Parallel arm RCT	n=360; nabiximols low dose (1-4 sprays/day) (n=91), medium dose (6-10 sprays/ day) (n=88), high dose day) (n=90), placebo (n=91) (100%)	100% chronic cancer pain	58 (12.2)	48%	Receiving opioid therapy of <500 MME/day (median was 120 MME/day; range 3-16 660)	5 weeks	THC 27 mg/mL; CBD 25 mg/ mL (maximum allowed daily dosage of 10 sprays per day)	Patients were allowed to use breakthrough opioid analgesic as required	GW Pharmaceuticals; Otsuka Pharmaceutical
Barlowe, 2019 (USA) ⁴⁷	Retrospective chart review	Enrolled in MCP (n=34), not enrolled in MCP (n=19) (100%)	100% CNCP (chronic painful pancreatitis)	49.9 (10.5)	45%	Not enrolled in MCP Range 4–297 183 MME/day±284; weeks enrolled in MCP 190 MME/day±273	Range 4–297 weeks	N	NR	Л
Bellnier, 2018 (USA) ⁴⁸	One-arm observational study	n=29 (100%)	90% CNCP; 10% cancer pain	61 (10)	65%	Patients were receiving a median opioid dose of 79.94 MME/day	13 weeks	10 mg capsules NR of THC/ CBD in a 1:1 ratio 3-times daily	٣	R
										Continued

Table 1 Continued	ned									
Author-year (country)	Study design	No of participants (% prescribed opioids)	Type of chronic pain (specific condition)	Age mean (SD)	% Female	Baseline opioid dose	Follow-up duration	Medical cannabis dose	Medical Analgesic cannabis dose cointervention	Funding source
Capano, 2020 (USA) ⁴⁹	One-arm observational study	n=131 (100%)	100% chronic pain (cancer and non- cancer)	56.1 (range: 39– 70)	68%	Receiving at least 50 MME/day	8 weeks	30 mg CBD/1 mg THC	щ	Ananda Professional
Haroutounian, 2016 (Israel) ⁵⁰	One-arm observational study	n=73 (35%)	93.2% CNCP; 6.8% chronic cancer pain	51.2 (15.4)†	38%†	Heceiving a median opioid dose of 60 MME/day (range 45–90)	26 weeks	Cigarettes: 6% to 14% THC, 0.2% to 3.8% CBD; Oral: 11% to 19% THC, 0.5% to 5.5% CBD	All participants were encouraged to attempt gradual dose reduction and possible discontinuation of other analgesics	No-external funding
Maida, 2008 (Canada) ⁵¹	Prospective cohort	Enrolled in MCP (n=47), not enrolled in MCP (n=65) (100%)	100% chronic cancer pain	69.7 (10.1)	42%	nabilone treated:60 MME/day±64; nabilone untreated: 67 MME/day±101	4 weeks	On average 1.79 mg two times daily nabilone	Patients were permitted to use concomitant analgesics	Valeant Pharmaceuticals Canada
Narang, 2008 (USA) ⁵²	One-arm observational study	n=30 (100%)	100% CNCP	Median=43.5 (range=21–67)	53%	Receiving an average opioid dose of 68 MME∕ day±57	4 weeks	Flexible dose schedule, dronabinol 5–20 mg three times daily	۳	Solvay Pharmaceuticals
0'Connell, 2019 (USA) ⁵³	One-arm observational study	n=77 (100%)	100% CNCP	54.1 (range=26–76)	58%	Receiving a mean opioid dose of 140 MME/day±184	26 weeks	NR	NR	No industry funding
Pritchard, 2020 (USA) ⁵⁴	Retrospective cohort	cannabis and opioids couse (n=22), opioids only (n=61) (100%)	100% chronic cancer pain	53.1 (11.7)	23%	MCP enrolled: 144 MME/day±129; MCP not enrolled: 119 MME/day±100	26 weeks	RN	R	No industry funding
Pawasarat, 2020 (USA) ⁵⁵	Retrospective chart review	Enrolled in MCP (n=137), not enrolled in MCP (n=95) (100%)	100% chronic cancer pain	58 (IQR:14.7)	56%	MCP enrolled: median 45 MME/ day, IQR=135; MCP not enrolled: median 97.5 MME/ day, IQR=150	Between 39 and 52 weeks for MCP enrolled;<26 weeks for not enrolled	R	N	No industry funding
Rod, 2019 (Canada) ^{se}	One-arm observational study	000=u	100% chronic pain (cancer and non- cancer)	Ë	۴	Receiving a mean opioid dose of 120 MME/day (range 90-240 MME/day)	26 weeks	CBD and THC ranged between 4% and 6%. Doses related directly to the opioid taper.	All participants indicated readiness to reduce opioid dose and also received psychological supports (eg, CBT, mindfulness, relaxation)	No external funding
										Continued

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Table 1 Continued	nued									
Author-year (country)	Study design	No of participants (% prescribed opioids)	Type of chronic pain (specific condition)	Age mean (SD)	% Female	Baseline opioid dose	Follow-up duration	Medical cannabis dose	Medical Analgesic cannabis dose cointervention	Funding source
Takakuwa, 2020 (USA) ⁵⁷	One-arm observational study	n=61 (100%)	100% CNCP (back pain)	50 (11.4)	38%	Receiving a median opioid dose of 21 MME/day	Median of 6.4 years among patients who ceased opioids completely	R	R	The Society of Cannabis Clinicians
Vigil, 2017 (USA) ⁵⁶	Retrospective chart review	Enrolled in MCP (n=37), not enrolled (n=29)(100%)	100% CNCP (90% back pain)	56.3 (11.8)	36%	Maximum daily dosage of <200 MME/day (enrolled in MCP: mean 24 MME/ day±23; not enrolled in MCP: mean 16 MME/ day±14)	52 weeks	٣	۴	University of New Mexico Medical Cannabis Research Fund
Yassin, 2019 (Israel) ⁵⁹	One-arm observational study	n=31 (100%)	100% CNCP (fibromyalgia)	33.4 (12.3)	%06	Receiving oxycodone 5 mg three times/day	26 weeks	THC to CBD ratio was 1:4; 20 g/month for 3 months, increased up to 30 g/month at the end of 6 months	Patients were permitted to use standardised analgesic therapy (duloxetine 30 mg once daily and Targin 5/2.5 mg two times a day). All other optates and atypical analgesics were stopped	R
*In Belgium, Bulgaria, †Based on the whole I CBD, cannabidiol; CB	Czech Republic, Estor population including ol T, cognitive behavioura	¹ In Belgium, Bulgaria, Czech Republic, Estonia, Germany, Hungary, Latvia, Lithuania, Poland, Romania, the UK and the USA. Hassed on the whole population including opiold users and non-users. CBD, cannabidiol; CBT, cognitive behavioural therapy; CNCP, chronic non-cancer pain; FU, follow-up; MME, milligram morphine equivalent; NR, not reported; RCT, randomised controlled trial; THC, tetrahydrocannabinol.	uania, Poland, Rom ser pain; FU, follow·	lania, the UK and the -up; MME, milligram n	USA. norphine equ	uivalent; NR, not reported	d; RCT, randomis	ed controlled trial; T	THC, tetrahydrocannabir	o.

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Table 2 GR/	ADE evide	ince profile of	GRADE evidence profile of medical cannabis or can	cannabinoids for patien	nabinoids for patients with chronic pain prescribed long-term opioid therapy	prescribed long-ter	rm opioid therapy		
# of studies	# of Patients	FU duration (Weeks)	Risk of bias*	Inconsistency (I ² , p value)†	Indirectness‡	Imprecision§	Publication bias	Treatment association (95% CI)	Overall certainty of evidence
Opioid dose: mo	rphine millig	Opioid dose: morphine milligram equivalents (MME) per day	s (MME) per day						
4 RCTs ^{43–45}	1176	2-5	No serious risk of bias ¶	No serious inconsistency (40%, p=0.15)	Very serious indirectness **	Serious imprecision 11 Not detected	Not detected	WMD -3.4 MME (-12.7 to 5.8)	Very low
8 Observational studies ^{47 48 50 51} 53-55 58	453	4–297	Serious risk of bias ##	Serious inconsistency (visual inspection)	No serious indirectness	No serious imprecision Not detected	Not detected	WMD -22.5 MME (-43.06 to -1.97)	Very low
Pain: 10cm VAS	for pain; lov	Pain: 10cm VAS for pain; lower is better; the MID=1 cm	MID=1 cm						
5 RCTs ⁴³⁻⁴⁶	1536	2–5	No serious risk of bias	No serious inconsistency (28%, p=0.20)	No serious indirectness	No serious imprecision	Not detected	WMD -0.18 (-0.38 to 0.02)	High
Sleep disturband	ce: 10 cm VA	S for sleep distu	Sleep disturbance: 10 cm VAS for sleep disturbance; lower is better; the MI	he MID=1 cm					
5 RCTs ⁴³⁻⁴⁶	1536	2–5	No serious risk of bias	No serious inconsistency (0%, p=0.45)	No serious indirectness	No serious imprecision Not detected	Not detected	WMD -0.22 (-0.39 to -0.06)	High
Nausea									
4 RCTs ⁴³⁻⁴⁶	1330	2–5	Serious risk of bias	No serious inconsistency (0%, p=0.88)	No serious indirectness	No serious imprecision Not detected	Not detected	RR 1.43 (1.04 to 1.96)	Moderate
Vomiting									
4 RCTs ⁴³⁻⁴⁶	1330	2–5	Serious risk of bias	No serious inconsistency (0%, p=0.50)	No serious indirectness	No serious imprecision Not detected	Not detected	RR 1.5 (1.01 to 2.24)	Moderate
Constipation									
3 RCTs ^{43 45 46}	1153	ى ا	Serious risk of bias §§	No serious inconsistency (0%, p=0.92)	No serious indirectness	Serious imprecision 11 Not detected	Not detected	RR 0.85 (0.54 to 1.35)	Low
*We assessed risk of bias using a modified Cochra thnconsistency refers to unexplained heterogeneity thndirectness results if the intervention, control, pa Serious imprecision refers to situations in which the Serious imprecision refers to situations in which the Serious inprecision refers to situations in which the insising outcome data (test of interaction p=0.758 missing outcome data (test of interaction p=0.758 missing outcome data the WND includes no effect. TTh 95% CI around the WND includes no effect. TTS were act high risk of bias due to lost it SigMost RCIs were at high risk of bias due to lost it FU, follow-up; GRADE, Grading of Recommendatic	of bias using a rs to unexplair rs to unexplair si f the intervel and refers to situ and refers to situ ata (test of inter ata (test of inter ata (test of inter ata (test of inter at high risk of 1 DE, Grading of	We assessed risk of bias using a modified Cochrane risk of bias instru filnconsistency refers to unexplained heterogenetty of results. For RCI #Indirectness results if the intervention, control, patients or outcomes Serious imprecision refers to situations in which the Cl includes both RJSome of the included RCTs were at high risk of bias, due to loss to ft missing outcome data (test of interaction p=0.788 and p=0.823 for poil +TDowngraded twice due to indirectness since all trials instructed parti HTThe 95%Cl around the WMD includes no effect. SMORS RCTs were at high risk of bias due to loss to for SUP of around the WMD includes to effect.	We assessed risk of bias using a modified Cochrane risk of bias instrument. Theonsistency refers to unexplained heterogenetry of results. For RCTs an 1° of 75%–100% indicates that heter thindirectness results if the intervention, control, patients or outcomes are different from the research question u specious imprecision refers to situations in which the CI includes both benefit and harm (the 95% CI includes 11 ¶Some of the included RCTs were at high risk of bias, due to loss to follow-up (>20%); however, we did not rate missing outcome data (test of intraction p=0.788 and p=0.783 and p=0.823 for opioid dose reduction and pain respectively). "Downgraded twice due to indirectness since all trials instructed participants to maintain their opioid dose duri #17The 85%CI around the xMDI includes no effect. #15The 85%CI around has due to lost to follow-up (>20%).	We assessed risk of bias using a modified Cochrane risk of bias instrument. Theonsistency refers to unexplained heterogeneity of results. For RC1s an 1 ⁶ of 75%–100% indicates that heterogeneity may be considerable. We assessed heterogeneity of pooled observational studies through visual inspection of forest plots. Endoinectness results for the intervention, control, patients or outcomes are different from the research question under investigation. Sensions impressions impressions refers to structions in which the Clinculdes both benefit and harm (the 95% Clinculdes 1 MID). Some of the included RC1s were at high risk of bias to be took the post of bias to the includes to a hole of the includes the loss to follow-up (>20%); how-up (>20%); how to read down for risk of bias as subgroup analysis showed no difference in treatment effect between trials at high and low risk of bias for missing outcome data (erc in interaction p=0.758 and p=0.623) for option does reduction and pain respectively). The 5% cl around the WID includes no effect. The 5% cl around the WID includes no effect. The second the read from the clost to follow-up (>20%). FU, follow-up; GRADE, Grading of Recommendations Assessment, Development and Evaluation; MID, minimally important difference; RC1, randomised controlled trial; RR, vISA, Visual Analogue Scale; WMD, weighted mean difference.	ty may be considerable. We ε vrestigation. • for risk of bias as subgroup i study period.	issessed heterogeneity of pr anatysis showed no differenc ised controlled trial; RR, reli	ooled observational studi ce in treatment effect bet lative risk; VAS, Visual An	ies through visual inspect ween trials at high and lo alogue Scale; WMD, wei	on of forest plots. <i>w</i> risk of bias for jhted mean difference.

4%, 95% CI 0% to 7%; online supplemental figures 9–10) and vomiting (RR 1.50; 95% CI 1.01 to 2.24; RD 3%; 95% CI 0% to 6%; online supplemental figures 11–12) in patients with cancer-related chronic pain prescribed opioid therapy. Three RCTs^{43 45 46} provided low certainty evidence that adding medical cannabis to opioid therapy may not increase constipation (RR 0.85, 95% CI 0.54 to 1.35; RD –1%; 95% CI –4% to 2%; online supplemental figures 13–14). Online supplemental table 11 summarises adverse events reported in observational studies.

DISCUSSION

Very low certainty evidence from randomised trials and observational studies was conflicting and leaves uncertain whether the addition of medical cannabis affects the use of prescribed opioids among people living with chronic pain. Compared with long-term opioid therapy for chronic cancer pain without medical cannabis, high certainty evidence showed that adding medical cannabis had little or no effect on pain or sleep disturbance. Results provided moderate certainty evidence that adding cannabis therapy to opioids likely increases both nausea (RR 1.43, 95% CI 1.04 to 1.96) and vomiting (RR 1.50; 95% CI 1.01 to 2.24) and low certainty evidence suggested no effect on constipation (RR 0.85, 95% CI 0.54 to 1.35).

Strengths of our review include a comprehensive search for eligible randomised and observational studies, appraisal of the risk of bias among individual studies, and use of the GRADE approach to rate the certainty of evidence. Our review has limitations, primarily due to features of primary studies eligible for review, which failed to report all recommended outcomes that have been established as important for people living with chronic pain. Most observational studies incorporated inadequate adjustment for confounding. All randomised trials, despite reporting this outcome, were not designed to address the effect of medical cannabis on opioid use. All eligible RCTs enrolled patients with chronic cancerrelated pain, and the generalisability to non-cancer chronic pain is uncertain. Specifically, substitution effects of medical cannabis for prescription opioids may also differ between chronic cancer and non-cancer pain; however, lack of variability among studies eligible for our review precluded exploration of this subgroup effect. Studies included in our review administered different formulations of cannabis and cannabinoid products; however, pooled effects of outcomes reported in RCTs showed no important heterogeneity.

A meta-analysis of preclinical studies,¹⁷ a narrative systematic review,⁶⁰ and several cross-sectional and case studies have reported an apparent reduction in opioid use with addition of cannabis therapy.^{9 10 61–65} In a national US populationbased survey⁶⁶ of 2774 cannabis users (both medical and non-medical use) 36% of respondents reported substituting cannabis for prescription opioids (discontinued opioid use). In this survey, the 60% of participants who identified as medical cannabis users were much more likely to substitute cannabis for prescription drugs than recreational users (OR 4.59; 95% CI 3.87 to 5.43). Another US survey⁶⁷ that included 841 patients prescribed long-term opioid therapy for chronic pain reported that 61% used medical cannabis, and 97% of this subgroup reported coincident reduction of their opioid use. Consistent with these findings, very low certainty evidence from observational studies in our review also suggests that adding medical cannabis allows patients predominantly with chronic non-cancer pain to reduce their use of opioids. Although RCT results do not support reduction in opioid dose by adding medical cannabis for opioids, the evidence is also very low certainty, primarily because investigators instructed patients to maintain their current opioid dose. This is a critical limitation, despite the 2019 National Institute for Health and Care Excellence guideline having concluded that providing medical cannabis for chronic pain does not reduce opioid use on the basis of these trials.⁶⁸ Future trials should randomise chronic pain patients who voluntarily agree to engage in a trial of opioid tapering to receive medical cannabis or placebo and report all patient-important outcomes.⁶⁹ Forced opioid tapering is ineffective⁷⁰ and may cause harm.⁷¹

CONCLUSION

The opioid-sparing effects of medical cannabis for chronic pain remain uncertain. Based on moderate-to-high certainty evidence, adding medical cannabis to opioid therapy among chronic cancer pain patients had little or no effect on neither pain relief nor sleep disturbance and likely increases the risk of nausea and vomiting. The accompanying BMJ Rapid Recommendation¹⁸ provides contextualised guidance based on this evidence, as well as three other systematic reviews on benefits,⁷² harms⁷³ and patients' values and preferences.⁷⁴

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