The Effect of Migration Duration on Treatment Delay Among Rural-to-Urban Migrants After the Integration of Urban and Rural Health Insurance in China: A Cross-Sectional Study

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Abstract

Few researches have been focused on the treatment delay of rural-to-urban migrants in China. Our study aimed to investigate the effect of migration duration on treatment delay among rural-to-urban migrants in tertiary hospitals. A cross-sectional study was conducted based on a sample of 727 patients and surveyed factors including sociodemographics, medical costs, migration, treatment delay, and health cost-coping strategies. Totally, 727 patients were included, of which 61 delayed their treatment and 666 had no treatment delay. Statistically significant differences were found between different migration duration groups in marital status, education, insurance, family annual income, residency, payment before treatment, reported disease, and migration duration (P < .05). The results from multiple logistic regression showed that migration between I and 5 years (adjusted odds ratio [OR] = 7.24; 95% confidence interval [CI] = 1.59-32.87; P < .05) was considered the significant contributing risk factor for treatment delay tended to use less savings and borrow more money than those without. Rural-to-urban migrants with I to 5 years of migration were the most vulnerable group of having treatment delay. Migrants were more likely to borrow money to cope with the health expenditure. Targeted services should be provided to meet different needs of migrants according to migration duration.

Keywords

treatment delay, cost-coping strategy, migrants, migration duration, health insurance

What do we already know about this topic?

Previous studies showed that treatment delay was related mainly with sociodemographic factors, socioeconomic status, clinical variables, and psychological factors.

How does your research contribute to the field?

This study added some new knowledge to the effect of migration on treatment delay: Rural-to-urban migrants with 1 to 5 years of migration were more likely to have treatment delay.

What are your research's implications toward theory, practice, or policy?

Targeted services should be provided to meet different needs of migrants according to migration duration.

Introduction

In China, the proportion of urban population increased to 56% in 2015, and more than 30%, roughly 247 million who are living in urban areas, are migrants.¹ China is now initiating a new round of urbanization on an unprecedented scale boosted by significant reform on household registration.²

Urbanization not only can change environment, but also people's lifestyle, which can directly lead to various health ¹Jiujiang University, China ²Prince of Songkla University, Hat Yai, Thailand ³Jiujiang First People's Hospital, China

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Figure 1. The conceptual framework based on the Andersen Health Service Utilization Model including the selected variables in this study.

problems. The health consequences followed by urbanization are mixed. Current studies reveal the phenomenon called "healthy migrant paradox," that is, migrants have physical health advantages over the native-born at the initial stage of the immigration. However, the migrant health advantage diminishes significantly with increasing resident time in the host society.³⁻⁶ The huge scale of migration has generated particular difficulties for health service delivery in China.⁷

In 2016, China initiated the merge of urban and rural health insurance schemes to tackle the inequity caused by separated rural and urban health care schemes which vary in funding, deductibles, and so on.⁸ However, pilot experience showed that the merge varied from province to province. In many cities, pooling funds were still controlled in county level.⁹ Also, whether the merge of urban and rural health insurance results in improving migrants' health service use remains less explored.

Although the vast majority of patients have been protected by health insurance, many low-income and minority patients (many of them are migrants from rural areas) had to pay a certain amount, often the entire amount of money prior to treatment, given the fact that the profit-driven public hospital worries the potential risk of being in debt if they treat patients in advance.^{10,11} Many households experience financial hardship as a result of delayed necessary treatment.^{12,13}

Most of the studies focusing on treatment delay were under the concept of patient's delay—the length of delay between the onset/discovery/recognition of signs and symptoms and a patient's first visit to a health care/medical provider.¹⁴ Many studies showed that the patient's treatment delay was related mainly with sociodemographic factors, socioeconomic status (SES), clinical variables, and psychological factors.¹⁵⁻¹⁸ However, few studies looked into the effect of migration on treatment delay (the delay between being told to be hospitalized/receive a major treatment and a patient's hospitalization/getting a major treatment done). This study aims to investigate the relationship between rural-to-urban migration and treatment delay after the integration of urban and rural health insurance schemes in China.

Methods

Conceptual Framework

The conceptual framework of this study is on the basis of the revised and expanded Behavior Model for Vulnerable Populations originated from the well-known Andersen Health Service Utilization Model.^{19,20} Andersen^{21,22} introduced the model in 1968 and had undergone several phases of updates over time. According to the model for vulnerable populations, the factors that make some populations vulnerable might also affect health services utilization and health status. We assumed that migrants, as a vulnerable population, and their use of service might be affected by 3 sequential clusters of factors: predisposing, enabling, and need factors in 2 domains (traditional and vulnerable domains) (see Figure 1). In this study, predisposing factors (predisposition to use health services) include demographic characteristics (age, sex, marital status), social structures (education, occupation), and vulnerable social structure (immigration). The enabling factors (impede the use of health services) include personal resource (insurance and income), community resource (residency), and vulnerable personal resource (payment before treatment—as an ability to negotiate). Need factors include reported diseases and vulnerable population. The outcome variable is whether the patient uses inpatient treatment in time or not.

Region

Recently, it is noticeable that a fast growth occurred in the midwest China, which contributed up to a 43.6% increase in migrants.²³ In Western China, migrants accounted for 19.5% of the total migrants in 2015, a 5.3% increase compared with the year of 2014.²⁴ So, the study was conducted at Jiujiang city in Jiangxi province of China which is located in the

midwest region of China. The city was subdivided into 10 counties and had a population of 4 million. To sample patient with various diseases, we purposively chose 2 general tertiary hospitals out of a total of 6 tertiary hospitals. Both hospitals were the top hospitals with more than 1500 beds.

Data Management and Measurement

The design was a hospital-based cross-sectional study. The study population consisted of all patients with or without urban household registration who were hospitalized in the selected 2 tertiary general hospitals in Jiujiang City between January 2017 and December 2018, or their direct relatives who were willing to participate in this study with the following inclusion criteria: hospitalized ≥ 1 time and age ≥ 18 years, and only 1 exclusion criteria: diagnosed with mental illness.

Consecutive sampling was used. Each participant who met the inclusion criteria was included till the desired sample size is reached. Sample size calculation was based on the 5% difference in the proportion of having negative care-seeking behavior between migrants and nonmigrants with 2.5% precision at 95% confidence interval (CI) and 80% power. The calculation was completed using the "epicalc" package in R software (version 3.5.2). A total of 500 admitted cases are needed considering a 10% nonresponse rate. The details of the calculation are as follows: For testing 2 independent proportions (2-tailed test), the main outcome was the proportion of having delayed care-seeking behavior. Based on the statistics from the Mobile Population Social Integration and Mental Health Survey (MPSIMHS), the proportion for the 1- to 5-year migration group was estimated as $p_1 = 0.25$, and the <1 year and no migration group as $p_2 = 0.05$, ratio (r) = 10 (n_2/n_1) , $\alpha = 0.05$, Z(0.975) = 1.959964; $\beta = 0.20$, Z(0.800) = 0.841621. Sample size using a continuity correction: 1- to 5-year migration group $(n_1) = 25$, <1 year and no migration group $(n_2) = 250$.

The questionnaire used for the face-to-face interview consisted of 4 parts. The first part was general information on sociodemographic characteristics. The second part contained the care-seeking behavior and health care cost of the latest admission. The third part was the information on monthly and annual family income and expenditure. The fourth part contained medical record and financial bill of the latest admission information obtained from the hospital information system.

For the main independent variable—migration duration (length of residence after moving from the rural to the urban area), participants were asked a question how long they had moved from the rural to the urban area.

For the main outcome variable—treatment delay, patients who reported treatment delay were considered having a treatment delay.²⁵ Patients were asked a set of questions: Did you delay the hospitalization and major treatment for current inpatient admission; when were you told by your doctor that you needed to be hospitalized; and when are you told by your doctor that you need the major treatment? Then the information with the actual admission and treatment date was verified using the data extracted from the hospital information system.

Data Analysis

Both descriptive and analytical statistics were employed. Generalized linear models (GLMs) were used to test the relationship between treatment delay and migration, along with other independent variables, such as age group, sex, marital status, job type, residency, income provider (major contributors to household income), health insurance, and health expenditure. Inference was made based on the chi-square test in univariate analysis. A *P* value less than .05 suggests that there is a significant difference among different categories.

The best fitted multivariate model was met by stepwise variable selection based on Akaike information criterion (AIC). First, we select the possible variables based on the conclusions of previous similar studies and univariate analysis between 2 groups. Then we applied stepwise model selection to looking for a better fitted model. Last, we decided the final variables on the basis of statistical model selection based on AIC and suggestion from other studies. The 95% CI and *P* value were calculated in the logistic regression model. The likelihood ratio test and Wald test were used to test the statistical significance of the variable in the logistic regression model. A *P* value less than .05 suggests the statistical significance of a variable in the model. All statistical analyses were conducted using the R software version 3.5.2.

Ethical Consideration

This study was conducted according to the World Medical Association Declaration of Helsinki and conformed to the International Committee of Medical Journal Editors' (ICMJE) Recommendations for the Conduct, Reporting, Editing, and Publication of Scholarly Work in Medical Journals. Ethical approval was obtained from Institutional Review Ethics Committee of Prince of Songkla University (REC: 59-266-18-5***) and Ethics Committee of *** Jiujiang University (REC: JJU20160116***). Every participant's consent was obtained with a written consent form signed by the patient or their legally authorized representatives. Great care was taken to secure the confidentiality of participants. Data from the hospital information system were disaggregated. Informed consent was obtained from all participants in hospitals. The participants had the opportunity to quit the survey at any time and to review and correct their own answers.

Results

Table 1 shows the comparison of sociodemographic characteristics among different migration duration groups. There were statistically significant differences between different

	>5 years of migration (n = 87), n (%)	I-5 years of migration (n = 27), n (%)	<1 year and no migration $(n = 613)$, n (%)	Chi-square test	Р
Age group				23.80	<.001
18-30	5 (5.75)	4 (14.81)	29 (4.73)		
30-40	11 (12.64)	5 (18.52)	41 (6.69)		
40-50	12 (13.79)	9 (33.33)	96 (15.66)		
50-110	59 (67.82)	9 (33.33)	447 (72.92)		
Sex				0.14	.9319
Female	42 (48.28)	12 (44.44)	286 (46.66)		
Male	45 (51.72)	15 (55.56)	327 (53.34)		
Marital status				6.94	.1389
Divorced and widowed	10 (11.49)	l (3.70)	41 (6.69)		
Married	75 (86.21)	23 (85.19)	547 (89.23)		
Single	2 (2.30)	3 (11.11)	25 (4.08)		
Education	()	(16.92	.002
Primary school or lower	32 (36.78)	10 (37.04)	305 (49.76)		
Middle school	26 (29.89)	6 (22.22)	84 (13.70)		
High school or higher	29 (33.33)	11 (40.74)	224 (36.54)		
lob type				4.14	.1259
Formal worker	19 (21.84)	7 (25.93)	94 (15.33)		
Informal worker	68 (78.16)	20 (74.07)	519 (84.67)		
Residency				155.36	<.001
, Urban	84 (96.55)	27 (100)	210 (34.26)		
Rural	3 (3.45)	0 (0)	403 (65.74)		
Insurance				3.43	.1799
Urban and rural resident	54 (62.07)	15 (55.56)	422 (68.84)		
Other insurance	33 (37.93)	12 (44.44)	191 (31.16)		
Family annual income (Yuan)				2.63	.6209
0-10000	44 (50.57)	15 (55.56)	301 (49.10)		
10000-50000	7 (8.05)	0 (0.00)	53 (8.65)		
50 000-480 000	36 (41.38)	12 (44.44)	259 (42.25)		
Payment before treatment				8.25	.083
<50%	12 (13.79)	2 (7.41)	90 (14.68)		
50%-74%	26 (29.89)	5 (18.52)	110 (17.94)		
75%-100%	49 (56.32)	20 (74.07)	413 (67.37)		

Table 1. Sociodemographic Characteristics of the Studied Participants (N = 727).

migration groups in age, education, and residency (P < .05). And 66.67% of patients with 1 to 5 years of migration were below the age of 50.

Univariate analysis was performed to explore possible factors related to treatment delay (Table 2). Totally, 727 patients were included, of which 61 delayed their treatment and 666 had no treatment delay. There were statistically significant differences between different migration groups in marital status, education, insurance, family annual income, residency, payment before treatment, reported disease, and migration duration (P < .05). Results showed that 7 (23.33%) single patients had delayed treatment; 41 (11.82%) patients with primary and lower education reported delayed treatment; 28 (11.86%) patients not covered by integrated urban and rural insurance reported delayed treatment; 45 (11.08%) patients living in rural area reported delayed treatment; 34 (11.07%) patients with 10 000 to 50 000 Yuan family annual income reported delayed treatment; 49 (10.17%) patients paid more than 75% cost before treatment delay; 8 (34.78%) patients with cancer reported delayed treatment; among 1- to 5-year migrants, 6 (22.22%) had treatment delay, whereas 51 (8.32%) reported in <1 year migrants and non-migrants.

Odds ratios (ORs) for treatment delay were adjusted for age, sex, marital status, education, job type, migration, insurance, family annual income, family support, residency, payment before treatment, reported disease, and migration duration using multiple logistic regression.

Multiple logistic regression method was used to investigate the potential risk factors in the multivariable analyses (Table 2). After the model selection, the best model includes 13 independent variables (age, sex, marital status, education, job type, migration, insurance, family annual income, family

Table 2. Logistic Regression for the Factors Related to Treatment Delay.

	No delay (n = 666), n (%)	Delay (n = 61), n (%)	Crude OR (95% CI)	Adjusted OR (95% CI)
Predisposing factors				
Age group: ref. $=$ 18-30	32 (84.21)	6 (15.79)	1.00	1.00
30-40	50 (87.72)	7 (12.28)	0.75 (0.23-2.42)	2.3 (0.43-12.13)
40-50	107 (91.45)	10 (8.55)	0.5 (0.17-1.48)	1.1 (0.22-5.58)
50-110	477 (92.62)	38 (7.38)	0.42 (0.17-1.08)	0.75 (0.15-3.8)
Sex: ref. = female	311 (91.47)	29 (8.53)	Ì.00	1.00
Male	355 (91.73)	32 (8.27)	0.97 (0.57-1.63)	0.93 (0.5-1.73)
Marital status: ref. = divorced and widowed	49 (94.23)	3 (5.77)	1.00	1.00
Married	594 (92.09)	51 (7.91)	1.4 (0.42-4.66)	1.32 (0.36-4.94)
Single	23 (76.67)	7 (23.33)*	4.97 (1.18-20.99)	5.45 (0.77-38.59)
Education: ref. = primary school or lower	306 (88.18)	41 (11.82)	1.00	1.00
High school or higher	109 (93.97)	7 (6.03)	0.48 (0.21-1.1)	0.43 (0.13-1.38)
Middle school	251 (95.08)	13 (4.92)**	0.39 (0.2-0.74)	0.33 (0.14-0.76)**
Job type: ref. = formal worker	112 (93.33)	8 (6.67)	1.00	1.00
Informal worker	554 (91.27)	53 (8.73)	1.34 (0.62-2.89)	0.77 (0.28-2.15)
Migration: ref. $=$ migrant	112 (91.06)	(8.94)	1.00	1.00
Nonmigrant	554 (91.72)	50 (8.28)	0.92 (0.46-1.82)	3.18 (01232087.5)
Enabling factors				
Insurance: ref. = urban and rural resident	458 (93.28)	33 (6.72)*	1.00	1.00
Other insurance	208 (88.14)	28 (11.86)	1.87 (1.1-3.17)	1.5 (0.81-2.79)
Family annual income: ref. = 50 000- 480 000 Yuan	339 (94.17)	21 (5.83)	1.00	1.00
0-10 000 Yuan	54 (90.00)	6 (10.00)	1.79 (0.69-4.65)	1.21 (0.42-3.48)
10000-50000 Yuan	273 (88.93)	34 (11.07)*	2.01 (1.14-3.54)	2.29 (1.19-4.43)*
Family support: ref. = live alone	35 (89.74)	4 (10.26)	1.00	1.00
No migration	554 (91.72)	50 (8.28)	0.79 (0.27-2.31)	0.56 (0,185231.87)
Live with family	77 (91.67)	7 (8.33)	0.8 (0.22-2.89)	0.86 (0.19-3.81)
Residency: ref. $=$ urban	305 (95.02)	16 (4.98)	1.00	1.00
Rural	361 (88.92)	45 (11.08)**	2.38 (1.32-4.29)	3.26 (1.14-9.29)*
Payment before treatment: ref. $=$ $<$ 50%	101 (97.12)	3 (2.88)	1.00	1.00
50%-74%	132 (93.62)	9 (6.38)	2.3 (0.61-8.7)	2.45 (0.61-9.78)
75%-100%	433 (89.83)	49 (10.17)*	3.81 (1.16-12.47)	2.07 (0.6-7.16)
Need factors				
Reported disease: ref. = neurosurgery	36 (97.30)	I (2.70)	1.00	1.00
Hepatobiliary surgery	278 (94.24)	17 (5.76)	2.2 (0.28-17.04)	3.36 (0.38-29.93)
Orthopedics	236 (92.91)	18 (7.09)	2.75 (0.36-21.2)	4.63 (0.53-40.45)
Cancer	15 (65.22)	8 (34.78)***	19.2 (2.2-167.21)	34.95 (3.42-357.05)**
Digestive	49 (80.33)	12 (19.67)	8.82 (1.1-70.91)	20.72 (2.14-200.25)**
Cardiothoracic surgery	52 (91.23)	5 (8.77)	3.46 (0.39-30.89)	6.28 (0.6-66.09)
Duration of migration: ref. = >5 years	83 (95.40)	4 (4.60)	1.00	1.00
I-5 years	21 (77.78)	6 (22.22)*	5.93 (1.53-22.93)	7.24 (1.59-32.87)*
<i td="" year<=""><td>562 (91.68)</td><td>51 (8.32)</td><td>1.88 (0.66-5.35)</td><td>0.61 (0.04-8.49)</td></i>	562 (91.68)	51 (8.32)	1.88 (0.66-5.35)	0.61 (0.04-8.49)

Note. OR = odds ratio; CI = confidence interval.

*P < .05. **P < .01. ***P < .001.

support, residency, payment before treatment, reported disease, and migration duration).

Education, family annual income, residency, reported disease, and migration duration were the variables in multiple logistic regression to be significantly associated with treatment delay. Having middle school education was associated with a 67% decreased odds of treatment delay compared with having primary and lower education. Compared with patients with family annual income more than 50 000 Yuan, those with family annual income ranging from 10 000



Figure 2. Cost-coping strategy between patients with treatment delay and no delay.

to 50000 Yuan had higher odds of treatment delay (adjusted OR = 2.29; 95% CI = 1.19-4.43; P < .05). In the rural setting, there was a trend toward increased treatment delay that was statistically significant (adjusted OR = 3.26; 95% CI =1.14-9.29; P < .05). Compared with patients having neurosurgery, those with cancer (adjusted OR = 34.95; 95% CI =3.42-357.05; P < .01) and digestive system diseases (adjusted OR = 20.72; 95% CI = 2.14-200.25; P < .01) had higher odds of treatment delay. Patients with 1 to 5 years of migration had higher odds of treatment delay (adjusted OR = 7.24; 95% CI = 1.59-32.87; P < .05) compared with those with more than 5 years of migration. Among patients who paid more than 75% cost before treatment, there was a trend toward increased treatment delay that was not statistically significant (adjusted OR = 2.07; 95% CI = 0.6-7.16; P = .25).

Age, sex, marital status, job type, migration, insurance, and family support were not associated with treatment delay in multiple logistic regression analysis.

There were 2 significant differences between patients with and without treatment delay. In total, 33 (54.1%) in the delay group used savings to cope with their health expenditure, and 458 (68.8%) in the no delay group; 46 (75.4%) in the delay group borrowed money, and 301 (45.2%) in the no delay group (Figure 2).

Patients with different migration durations also showed various preferences when dealing with health expenditure.

Among patients with more than 5 years of migration, 48 (55.2%) used savings to cope with their health expenditure, 18 (66.7%) among patients with 1 to 5 years of migration, and 425 (69.3%) among patients with less than 1 year or no migration; 26 (29.9%) patients with more than 5 years of migration borrowed money, 12 (44.4%) among patients with 1 to 5 years of migration, and 309 (50.4%) among patients with less than 1 year or no migration 1 year or no migration (Figure 3).

Discussion

After the integration of urban and rural health insurance schemes, even the rural-to-urban migrants without urban household registry were entitled to get expanded the medication and services; moreover, patients now could get the reimbursement back right before they were discharged. The result in this study presented that the probability of having treatment delay was decreased among patients covered by integrated urban and rural insurance, which might result from the expanded benefit packages and instant reimbursement policy after the integration of urban and rural insurance.

Despite the benefits provided by the integrated insurance, our study found some evidence indicating that the rural-tourban migrants should receive special attention.

First, this study showed that most of the migrants, especially among migrants living in an urban area for 1 to 5 years, were young and middle-aged labors. The result also showed



Figure 3. Cost-coping strategy among different migration duration groups.

that migrants were more likely to face the risk of having treatment delay, especially those have lived in urban area for 1 to 5 years. The young and middle-aged migrants are usually the main income provider for their family, and their health-related financial risk can put an extra burden on patients and their family, who not only have to tolerate medical treatment cost but also suffer from the income loss due to absenteeism to work.²⁶ Under this circumstance, they will take more time to make decision whether to use inpatient care or major treatment that can lead to treatment delay. Our study also showed that with the increasing duration of migration, fewer patients used savings and borrowed money to cope with the health expenditure, which indicate that patients with longer migration are more likely to get a formal job and covered by employee health insurance. Targeted services should be provided to meet different needs of migrants according to migration duration.

It was also noticeable that there were more patients paid more than 75% cost before treatment in the delay group, and patients with treatment delay were less likely to use savings to cope with their health expenditure, most of which also needed to borrow money. For the low-income family, they may need to delay the treatment to raise money to pay for the cost. Reducing payment before treatment might serve as a vehicle to provide timely treatment for migrants. Early education about medical crowdfunding might help patients collect the money for treatment in time and prevent bankruptcy.²⁷⁻²⁹ Third, between the delay and no delay groups, there were many differences in marital status, education, insurance, family annual income, residency, payment before treatment, and reported disease, which were consistent with other studies^{15,30-32}: being single, having less education, living in rural area, and low family annual income were more likely to have treatment delay.

Finally, this study also indicated that patients with cancer and digestive system diseases might delay their treatment. Because the disease or the treatment was not as emergency as neurosurgery, patients were more likely to delay the treatment for alternative solution.

Limitations

Our study had some limitations. First, all cases were selected from 2 tertiary hospitals in 1 city, so the conclusion might be difficult to be generalized to a larger scale. Second, socioeconomic determinants related to treatment delay were dynamic factors, which might be different over time and from country to country. The interaction between migration and other factors might react differently under changing social context. Our findings may not be generalized to all migrants in China and in other countries. Third, social capital, social support, and social integration were also proved to be related to health-seeking behavior, but we could not collect the above information due to limited time and resources. However, this is the first study to test the relationship between migration duration and treatment delay.

Conclusions

This study added some new possible factors contributing to the treatment delay. Rural-to-urban migrants in different migration groups showed different probabilities of having treatment delay, and those who have 1 to 5 years of migration were the most vulnerable group; payment before treatment may also serve as an obstacle to timely treatment. Migrants were more likely to borrow money to cope with the health expenditure. Targeted services should be provided to meet different needs of migrants according to migration duration.

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Author Contributions

L.N. analyzed the results; Y.L. helped with the investigation; H.L. and J.C. provided technical support with hospital data analysis; and X.W. and H.S. reviewed the manuscript.

Data Availability Statement

The data that support the findings of this study are available from the corresponding author upon reasonable request.

Declaration of Conflicting Interests

The author(s) declared no potential conflicts of interest with respect to the research, authorship, and/or publication of this article.

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References

- 1. National Bureau of Statistics of China. *Statistical Bulletin* of the National Economic and Social Development. Beijing: National Bureau of Statistics of China; 2015.
- Opinions on Further Promoting the Reform of the Household Registration System. Beijing: The State Council of the People's Republic of China; 2014.
- Lin Y, Zhang Q, Chen W, et al. Association between Social Integration and Health among Internal Migrants in ZhongShan, China. *PLos One*. 2016;11(2):e0148397.
- Chen J. Internal migration and health: re-examining the healthy migrant phenomenon in China. *Soc Sci Med.* 2011;72(8): 1294-1301.
- Alvarez-Ude CF, Rebollo AP. [Psychological disturbances and deterioration of health-related quality of life of patients with stage 3-5 chronic kidney disease (not on dialysis)]. *Nefrologia*. 2007;28:57-62.

- Tong Y, Piotrowski M. Migration and health selectivity in the context of internal migration in China, 1997–2009. *Popul Res Policy Rev.* 2012;31(4):497-543.
- Gong P, Liang S, Carlton EJ, et al. Urbanisation and health in China. *Lancet*. 2012;379(9818):843-852.
- The State Council's Opinions on the Integration of the Basic Medical Insurance System for Urban and Rural Residents. Beijing: The State Council of People's Republic of China; 2016.
- Yinghui Q. Study on the Pilot Reform of Merging of China's Urban and Rural Residents Basic Medical Insurance System [Master]. Beijing: Beijing University of Chinese Medicine; 2015.
- Reynolds L, McKee M. Serve the people or close the sale? Profit-driven overuse of injections and infusions in China's market-based healthcare system. *Int J Health Plann Manage*. 2011;26(4):449-470.
- Williams P. How Health Care Delay and Avoidance Decisions Are Affected by Finances and Health Insurance [1450480]. Ann Arbor: ProQuest Dissertations and Theses, The University of Texas School of Public Health; 2008.
- van Doorslaer E, O'Donnell O, Rannan-Eliya RP, et al. Effect of payments for health care on poverty estimates in 11 countries in Asia: an analysis of household survey data. *Lancet*. 2006;368(9544):1357-1364.
- Xu K, Evans DB, Kawabata K, Zeramdini R, Klavus J, Murray CJ. Household catastrophic health expenditure: a multicountry analysis. *Lancet*. 2003;362(9378):111-117.
- Dubayova T, van Dijk JP, Nagyova I, et al. The impact of the intensity of fear on patient's delay regarding health care seeking behavior: a systematic review. *Int J Public Health*. 2010;55(5):459-468.
- Lefler LL, Bondy KN. Women's delay in seeking treatment with myocardial infarction: a meta-synthesis. J Cardiovasc Nurs. 2004;19(4):251-268.
- George SZ, Dover GC, Fillingim RB. Fear of pain influences outcomes after exercise-induced delayed onset muscle soreness at the shoulder. *Clin J Pain*. 2007;23(1):76-84.
- Li T, Zhang H, Shewade HD, Soe KT, Wang L, Du X. Patient and health system delays before registration among migrant patients with tuberculosis who were transferred out in China. *BMC Health Serv Res.* 2018;18(1):786.
- Wechkunanukul K, Grantham H, Damarell R, Clark RA. The association between ethnicity and delay in seeking medical care for chest pain: a systematic review. *JBI Database System Rev Implement Rep.* 2016;14(7):208-235.
- Babitsch B, Gohl D, von Lengerke T. Re-revisiting Andersen's Behavioral Model of Health Services Use: a systematic review of studies from 1998–2011. *Psychosoc Med.* 2012;9:1-15.
- Gelberg L, Andersen RM, Leake BD. The behavioral model for vulnerable populations: application to medical care use and outcomes for homeless people. *Health Serv Res.* 2000;34(6): 1273-1302.
- Andersen R. A behavioral model of families' use of health services. *Research Series No. 25. Center for Health Administration Studies, Chicago: University of Chicago*; 1968.
- Andersen RM. Revisiting the behavioral model and access to medical care: does it matter. *J Health Soc Behav.* 1995;36(1): 1-10.

- 23. Bi L, Fan Y, Gao M, Lee CL, Yin G. Spatial mismatch, enclave effects and employment outcomes for rural migrant workers: empirical evidence from Yunnan Province, China. *Habitat International*. 2019;86:48-60.
- Statistics NBo. Migrant Workers' Surveillance Report 2016. National Bureau of Statistics. http://www.stats.gov.cn/tjsj/ zxfb/201704/t20170428_1489334.html. Published 2016. Accessed April, 2017.
- Lechner A. Forgoing Medical Care due to Costs: the Causal Effect of Health Care Burdens on Health [M.P.P.]. Ann Arbor, MI: Georgetown University; 2011.
- Leive A, Xu K. Coping with out-of-pocket health payments: empirical evidence from 15 African countries. *Bull World Health Organ.* 2008;86(11):849-856.
- 27. Burtch G, Chan J. Reducing medical bankruptcy through crowdfunding: evidence from GiveForward. https://pdfs.

semanticscholar.org/c4c7/148df6f372a853ac6f9f7f56701ce16 5188c.pdf. Published 2014. Accessed March 8, 2020.

- Berliner LS, Kenworthy NJ. Producing a worthy illness: personal crowdfunding amidst financial crisis. *Soc Sci Med*. 2017;187:233-242.
- Yuan H, Lau RY, Xu W. The determinants of crowdfunding success: a semantic text analytics approach. *Decision Support Systems*. 2016;91:67-76.
- Facione NC, Miaskowski C, Dodd MJ, Paul SM. The selfreported likelihood of patient delay in breast cancer: new thoughts for early detection. *Prev Med*. 2002;34(4):397-407.
- Betti R, Vergani R, Tolomio E, Santambrogio R, Crosti C. Factors of delay in the diagnosis of melanoma. *Eur J Dermatol*. 2003;13(2):183-188.
- 32. Montella M, Crispo A, Botti G, et al. An assessment of delays in obtaining definitive breast cancer treatment in Southern Italy. *Breast Cancer Res Treat*. 2001;66(3):209-215.