knowledge, one of the few studies from an LMIC setting. Although this is a single-center study with a relatively small sample, our geographic catchment area is large and is representative of children with CKD in our region. Despite not reaching the target sample size, the pooled analyses revealed highly plausible findings that reflect the day-to-day reality in our clinic, including late presentation and loss to follow-up. Given the small numbers, subgroup analysis of comorbidities associated with CKD and other factors such as low birthweight, prematurity, medication use, and diet could not be analyzed. Owing to resource limitations, we could not measure albumin-tocreatinine ratio. However, with a significant association between proteinuria and progression, a spot urine protein-to-creatinine ratio remains a valid and affordable target for treatment in our cohort. Better data on the burden of hypertension using ambulatory blood pressure measurement would have been helpful; however, this was not routinely available in our setting. The significant association with blood pressure as measured with CKD progression also provides a reproducible target for intervention.

CONCLUSION

This prospective, longitudinal study shows that the profile of children with CKD stages II to IV in an LMIC is different from those in developed countries. The lower median baseline GFR suggests delay in diagnosis of CKD. The rate of progression of CKD was high, particularly in those with glomerular disease and associated with proteinuria and uncontrolled blood pressure, suggesting that the risk factors associated with progression of CKD must be identified early and treated adequately. The QoL analysis emphasizes the impact of CKD on affected children and highlights the role of socioeconomic status, which influences the ability to seek and to use medical care. There is an urgent need to increase awareness of pediatric CKD in LMICs to facilitate early diagnosis and intervention.

DISCLOSURE

All the authors declared no competing interests.

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SUPPLEMENTARY MATERIAL

Supplementary File (Word) Supplementary Methods. Supplementary References.

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Identification of ESRD in Cardiovascular Procedural Databases

Check for updates

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• ardiovascular revascularization procedural data are U routinely collected to assess procedural quality, create risk adjustment tools, and assess outcomes in populations not well-studied in clinical trials. The National Cardiovascular Data Registry (NCDR) and the Society for Thoracic Surgeons (STS) data sets are widely used for these purposes,^{1,2} and their data collection instruments are utilized as "off-the shelf" tools to facilitate standardized data capture.

Correct identification of dialysis-dependent endstage renal disease (ESRD) is particularly important because ESRD is a potent risk factor for cardiovascular mortality and procedural complications.³⁻⁷ Incorrect identification could impact risk-adjusted quality reporting for cardiac procedures as well as the retrospective analyses widely used to assess revascularization outcomes in dialysis patients. However, to our knowledge, the STS and NCDR instruments for identification of dialysis patients have not been validated. We assessed accuracy of dialysis identification by linking United States Renal Data System (USRDS) data to Massachusetts Data Analysis Center statewide data collected using the STS and NCDR instruments under a legal mandate requiring universal data capture on all patients undergoing coronary artery bypass graft (CABG) and percutaneous coronary intervention (PCI).^{8,9}

RESULTS

Study Characteristics

We identified 26,317 individuals undergoing CABG, and 99,848 undergoing PCI. The mean age was 66.8 \pm 10.5 years in the CABG group, and 64.7 \pm 12.6 years in the PCI group. Subjects were primarily White (CABG 89.9%; PCI 88.9%), with 55.1% of CABG and 71.7% of PCI patients admitted with acute coronary syndrome. Many procedures were performed urgently (CABG 62.7%; PCI 45.3%). Emergent or salvage procedures were rare for CABG (2.9%) but not PCI patients (23.9%). Diabetes, heart failure, and hypertension were common (Table 1).

Dialysis Identification

After excluding patients with kidney transplants (N = 49 for CABG; N = 193 for PCI), 295 of 26,268 (1.1%) CABG patients were identified by USRDS as having dialysis-dependent ESRD (Table 2). Of these, 278 (94.2%) were correctly identified by the STS instrument, and 17 (5.8%) were not. Conversely, 147 of

25,973 (0.6%) non-dialysis patients were incorrectly flagged as dialysis patients at the time of their procedure. Sensitivity for identification of dialysisdependent ESRD was 94.2%, specificity 99.4%, positive predictive value 65.4%, and negative predictive value 99.9%.

There were 950 of 99,655 (0.95%) PCI patients on dialysis identified by the USRDS. Of these, 876 (92.2%) were correctly identified, and 74 (7.8%) were not identified by the NCDR instruments. Of 98,705 individuals not on dialysis, 544 (0.6%) were incorrectly flagged as receiving chronic dialysis at the time of their procedure. Sensitivity was 92.2%, specificity 99.5%, positive predictive value 61.7%, and negative predictive value 99.9%. A supplementary analysis (Supplementary Figure S1 and Supplementary Results) identified dialysis type, hospital procedural volume, and procedural urgency as characteristics common to both data sets that differed in the number of falsenegative or false-positive patients compared with true-positive patients. Results for both PCI and CABG were similar following adjustment for hospital (Supplementary Table S1).

Change by Instrument Version and Calendar Year

Analyses stratified by STS version (CABG data) did not demonstrate significant variability in sensitivity according to instrument version (Figure 1), but specificity differed (sensitivity $P_{\text{trend}} = 0.84$; specificity $P_{\text{trend}} = 0.01$). However, differences were marginal, with an overall change in specificity of <0.4%. For PCI, sensitivity and specificity did not vary significantly by NCDR version (Figure 2; sensitivity $P_{\text{trend}} = 0.86$; specificity $P_{\text{trend}} = 0.64$). Trends across calendar years were qualitatively similar for the 2 data sets (not shown).

Impact on Risk-Adjusted Outcomes

For PCI, the area under the curve (AUC) was significantly lower (P = 0.02) for models incorporating the USRDS variable (0.899, 95% confidence interval [CI]: 0.891, 907) compared with the Massachusetts Data Analysis Center variable (0.900, 95% CI: 0.892, 0.908), but differences were marginal (Figure 3). Similarly, the continuous net reclassification index (-0.031, 95% CI: -0.058, -0.003) was consistent with weak effects on risk discrimination. For CABG, AUCs using the USRDS variable (0.765, 95% CI: 0.741,

Table 1. Overall population

CABG ^a		3G ^a	PCI ^b (N = 99,848)	
	(N = 26,317)			
Variable	N	%	N	%
Demographics and insurance				
Age (mean \pm SD), yr	66.8 ± 10.5		64.7 ± 12.6	
Male	20,018	76.06	69,090	69.20
Race				
White	23,666	89.93	88,733	88.87
Black	623	2.37	2770	2.77
Other	2028	7.71	8345	8.36
Insurance payor ^c				
Private	11,440	43.47	50,145	50.22
Government	14,146	53.75	46,482	46.55
Other	644	2.45	3221	3.23
Dialysis status				
Dialysis identified in the Mass-DAC instruments	431	1.64	1432	1.43
Dialysis identified through USRDS data	295	1.12	950	0.95
Baseline medical conditions				
Acute coronary syndrome	14,493	55.09	71,596	71.70
Diabetes	10,412	39.56	29,084	29.13
Heart failure	4631	17.60	11,446	11.46
Hypertension	22,145	84.15	75,354	75.47
Hypercholesterolemia	22,950	87.21	78,101	78.22
Peripheral vascular disease	4291	16.31	11,428	11.45
Prior myocardial infarction	13,183	50.09	21,615	21.65
Prior PCI	3179	12.08	15,341	15.36
Prior CABG	408	1.55	10,613	10.63
CABG or PCI performed at a teaching hospital	22,073	83.87	77,472	77.59
Urgent status	16,507	62.72	45,210	45.28
Emergent or salvage status	753	2.86	23,885	23.92
Cardiogenic shock	217	0.82	2082	2.09
Hospital characteristics				
Hospital procedural volume				
Low	5000	19.0	5043	5.05
Medium	9414	35.8	26,546	26.6
High	11,903	45.2	68,259	56.18

CABG, coronary artery bypass grafting; Mass-DAC, Massachusetts Data Analysis Center; PCI, percutaneous coronary intervention; USRDS, US Renal Data System. ^aFor CABG, hospital procedural volume is defined as low \leq 1418, medium >1418 to \leq 2175, high >2175.

^bHospital procedural volume for PCI is defined as low \leq 914 cases in total during the study period, medium >914 to \leq 4625, high \geq 4625.

^cEighty-seven (0.33%) CABG patients were missing payor.

0.790) and the Massachusetts Data Analysis Center variable (0.770, 95% CI: 0.745, 0.795) were not different (P = 0.06). The net reclassification index -0.045 (95% CI: -0.112, 0.022) was also consistent with only weak effects on reclassification. Lastly, differences in predicted risk were minimal, regardless of hospital procedural volume, for the vast majority of procedures (Supplementary Tables S2–S5 and Supplementary Figure S2).

DISCUSSION

We assessed dialysis identification by the NCDR and STS data instruments for 2003–2012 Massachusetts

 Table 2. Sensitivity and specificity for identification of chronic dialysis patients

Identification of dialysis in the USRDS and Massachusetts state revascularization data									
Dialysis according to Mass-DAC coronary artery			Dialysis identified in USRDS						
bypass-STS data		.,,	No	Yes	Total				
No			25,826	17	25,843				
Yes			147	278	425				
Total			25,973	295	26,268				
Sensitivity (%)	Specificity (%)	Positive value	Positive predictive value (%)		Negative predictive value (%)				
94.2	99.4	65	65.4		99.9				
Dialysis according	Dialysis identified in USRDS								
coronary intervention-NCDR data			No	Yes	Total				
No			98,161	74	98,235				
Yes			544	876	1420				
Total			98,705	950	99,655				
Sensitivity (%)	Specificity (%)	Positive value	predictive (%)	edictive Negative predict %) value (%)					
92.2	99.5	61	61.7		99.9				

Mass-DAC, Massachusetts Data Analysis Center; NCDR, National Cardiovascular Data Registry; STS, Society of Thoracic Surgeons; USRDS, US Renal Data System.

patients receiving CABG or PCI, by linking our data to the USRDS. Specificity and negative predictive values for identification of chronic dialysis patients were high, and false-positive rates were low. However, the proportion of individuals receiving maintenance dialysis was small, and positive predictive values were low (62% for PCI; 65% for CABG). The impact on overall prediction of procedural risk was small, suggesting that use of these tools to compare facility outcomes is reasonable despite the misidentification of an important risk factor like chronic dialysis status.

Prognostic risk scores derived from STS (the STS score) and NCDR data sets and instruments have been widely used to assess risk-adjusted outcomes, compare procedural results across providers, analyze outcomes of cardiac surgery, assess the impact of kidney disease and dialysis status on practice patterns, and assess postsurgical, postmyocardial infarction, and post-PCI outcomes.^{8,9,S1-S11} However, to our knowledge, the current investigation is the first to assess the precision of the dialysis variables, and our results suggest that their accuracy is suboptimal. Although sensitivity and specificity are high, the overall prevalence of maintenance dialysis patients was <1.5% in each cohort. Consequently, positive predictive values were low, with more than one third of patients identified by the STS instrument, and nearly 40% of those identified by the NCDR, not actually receiving maintenance dialysis. This raises questions about use of data based on the NCDR and STS instruments to assess cardiac procedures. Although the exclusion of patients with chronic kidney disease and ESRD from cardiovascular trials^{\$12}



Figure 1. Sensitivity and specificity according to the version of the Society for Thoracic Surgeons data collection instrument. False-negative compared with true-positive identification (a) and true-negative compared with false-positive identification (b) of dialysis patients undergoing coronary artery bypass grafting (CABG), according to the version of the Society for Thoracic Surgeons (STS) data instrument. *P* values are for tests to assess differences between versions and trends across versions. The underlying data are provided at the top of the table. Mass-DAC, Massachussetts Data Analysis Center; USRDS, United States Renal Data System.

makes use of these data sets to investigate cardiac treatment strategies attractive, our results suggest that identification of chronic dialysis patients within NCDRand STS-based data sets is not sufficiently accurate to provide reliable guidance for the care of dialysis patients. Misspecification of the dialysis variable could negatively influence adjustment for confounding and reduce the accuracy of public reporting of PCI and cardiac surgery outcomes. To avoid over-reliance on any single metric for assessing prognostic value, we examined changes in AUC and net reclassification index, and compared predicted and actual risk. We detected marginal effects on AUC, and predicted risk differed significantly in a minority of individual



Figure 2. Sensitivity and specificity according to the version of the National Cardiovascular Data Repository (NCDR) instrument used. Falsenegative compared with true-positive identification (a) and true-negative compared with false-positive identification (b) of dialysis patients undergoing percutaneous coronary intervention (PCI) according to the version of the instrument. *P* values are provided for tests to assess differences between versions and trends across versions. The underlying data are provided at the top of the table. Mass-DAC, Massachussetts Data Analysis Center; USRDS, United States Renal Data System.



Figure 3. Receiver operating curves for prediction equations incorporating Massachussetts Data Analysis Center (Mass-DAC) and United States Renal Data System (USRDS) dialysis variables. Plots show receiver operating curves for the regression equation incorporating the USRDS variable (blue) or the Mass-DAC variable (red) for percutaneous coronary intervention data (a) and coronary artery bypass graft data (b).

cases. Thus, in aggregate, our data suggest that the overall impact of dialysis status misspecification is small and unlikely to significantly compromise analyses of procedural risk and benefits or comparative hospital scorecards, although the net impact could be important in hospitals with a combination of low procedural volume and unusually high rates of misspecification.

Determining the underlying reasons for and best response to the inaccuracies we identified is necessary. Our results suggest that including explicit variables for peritoneal dialysis, targeting training efforts at low volume centers, and considering enhanced validation of data gathered during emergent or urgent procedures are steps with potential utility. Although we lacked the data needed to investigate misidentification of individuals with dialysis-dependent acute kidney injury as patients with dialysis-dependent ESRD, we also believe that clarification of the instrument fields to better discriminate between acute and chronic kidney disease should be considered.

Our analysis had several limitations. We analyzed data from a single state, and our results may not be fully generalizable. However, data collection using the NCDR and STS instruments is mandated in Massachusetts and is performed by trained staff; selected fields were audited to ensure fidelity. Nevertheless, better performance in the national data sets is theoretically possible. Additionally, state privacy regulations precluded use of social security numbers during matching to the USRDS, although we were able to utilize name, date of birth, and last known alive dates. Significant numbers of patients sharing these identifiers, within Massachusetts during the study period, is unlikely.

In conclusion, we matched Massachusetts PCI and CABG patients from the USRDS to identification of chronic dialysis patients by the NCDR and STS. Neither accurately identified individuals with dialysisdependent ESRD, suggesting that data collected using these instruments may not be useful for informing therapeutic choices in individuals requiring chronic dialysis and that efforts to improve these instruments are warranted.

DISCLOSURE

DMC received consulting fees from Amgen, Medtronic, and Lilly-Boehringer, and fees related to service on DSMBs (AstraZeneca, Allena Pharmaceuticals) and trial steering committees (Jannsen Pharmaceuticals-CREDECE Trial, Zoll Medical-WEDHEAD trial). All the other authors declared no competing interests.

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SUPPLEMENTARY MATERIAL

Supplementary File (PDF)

Supplementary Methods.

Supplementary Results.

Figure S1. Characteristics associated with misidentification of dialysis patients.

Figure S2. Predicted risk using US RDS and MASS-DAC variables.

Table S1. Sensitivity, specificity, positive and negativepredictive values in analyses corrected for hospital.

Table S2. Results of mixed-effects model using MASS-DACdialysis variable for prediction of 30-day mortality in thePCI data sets. Estimates are log-odds ratios.

Table S3. Results of mixed-effects model using USRD dialysis variable for prediction of 30-day mortality in the PCI data sets. Estimates are log-odds ratios.

Table S4. Results of mixed-effects model using USRDS dialysis variable for prediction of 30-day mortality in the CABG data sets. Estimates are log-odds ratios.

Table S5. Results of mixed-effects model using Mass-DACdialysis variable for prediction of 30-day mortality in theCABG data sets. Estimates are log-odds ratios.

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Community Screening for Diabetes, Hypertension, Nutrition, and Kidney Disease Among Kenyans



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