# Association of rs6265 and rs2030324 Polymorphisms in Brain-Derived Neurotrophic Factor Gene with Alzheimer's Disease: A Meta-Analysis 

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#### Abstract

Background: The association between polymorphisms rs6265 and rs2030324 in brain-derived neurotrophic factor (BDNF) and Alzheimer's disease (AD) has been widely reported, but the results remain controversial.

Methods: A comprehensive search of Pubmed, Web of Science, China National Knowledge Infrastructure (CNKI), Wanfang Med Online and China Biology Medical literature database (CBM) was performed. Pooled odds ratios (ORs) with $95 \%$ confidence intervals (Cls) were calculated using fixed or random-effects models. We excluded the studies with OR $>3.0$ or $O R<0.3$ for sensitive analysis. Subgroup analysis by ethnicity, form of AD and gender was carried out. Meta-regression was conducted to explore the potential sources of between-study heterogeneity.

Results: 29 articles with 7548 cases and 7334 controls concerning rs6265 and 22 articles with 5796 cases and 5706 controls concerning rs2030324 were included in this meta-analysis. The combined evidence suggested rs6265 contributing significantly to the increased risk of AD in females (codominant: fixed-effects model (FEM): OR=1.13, $95 \% \mathrm{Cl}=1.04-1.23$; dominant: $\mathrm{FEM}: \mathrm{OR}=1.17,95 \% \mathrm{CI}=1.05-1.31$ ), especially for Caucasian females (codominant: $\mathrm{FEM}: \mathrm{OR}=1.18,95 \% \mathrm{CI}=1.03-$ 1.34; dominant: $\mathrm{FEM}: \mathrm{OR}=1.18,95 \% \mathrm{Cl}=1.01-1.37$ ) and female late-onset Alzheimer's disease (LOAD) patients (codominant: FEM: $\mathrm{OR}=1.22,95 \% \mathrm{Cl}=1.05-1.41$; dominant: $\mathrm{FEM}: \mathrm{OR}=1.23,95 \% \mathrm{Cl}=1.03-1.46$ ). No evidence indicated an association between rs2030324 with AD in codominant (random-effects model (REM): $\mathrm{OR}=1.06,95 \% \mathrm{Cl}=0.89-1.26$ ) and dominant ( $\mathrm{REM}: \mathrm{OR}=1.05,95 \% \mathrm{CI}=0.86-1.27$ ) models.

Conclusion: This meta-analysis suggested A allele of rs6265 might increase the risk of AD in Caucasian females and female LOAD patients. In addition, no evidence indicated an association between rs2030324 with AD. Further studies are needed to confirm these results.


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## Introduction

Alzheimer's disease (AD) is an age-associated neurodegenerative disorder characterized by progressive decline in cognitive function, which typically begins with deterioration in memory [1]. The number of people with dementia worldwide in 2010 is estimated at 35.6 million and is projected to nearly double every 20 years to 65.7 million in 2030 and 115.4 million in 2050. AD is the most common form of dementia and possibly contributes to $60-70 \%$ of cases [2]. In the US alone, AD is related with an estimated healthcare cost of US\$172 billion per year [3]. The overwhelming number of AD patients, combined with the staggering economic burden, makes AD a public health problem.

The key pathological changes that observed in AD brain tissue are the accumulation of neuritic extracellular amyloid plaques and intracellular neurofibrillary tangles [4]. However, the neuropathological etiology of AD remains unclear, but are probably related
with the combined interaction between gene variants and environmental factors [5]. There have been many genetic polymorphisms reported to be associated with AD , such as amyloid precursor protein (APP), presenilin-1(PSEN1), presenilin2(PSEN2) [6], apolipoprotein E (APOE) [7,8] and sortilin-related receptor 1 (SORL1) [9]. However, as a complex disorder, the genes mentioned-above can not explain the overall genetic susceptibility and it is supposed that some other genes may participate in the development of AD.

Brain-derived neurotrophic factor (BDNF), as a member of the neurotrophic family, plays an important role in the growth, development, differentiation and regeneration of various types of neurons in the central nervous system [10]. Autopsy studies found reduced mRNA expression of BDNF in the hippocampus of patients with AD [11], which implicates the possible participation of BDNF in the pathogenesis of AD. There have been many
polymorphisms studied in BDNF gene, such as rs11030104, rs16917204, rs7103411, rs6265 and rs2030324. However, only the last two polymorphisms have been widely studied, with no linkage disequilibrium (LD) between them. What's more, these results are inconsistent and individual studies have relatively small power to confirm this association. For example, the $G$ allele of rs6265 confers risk effect for AD in subjects of Japanese ( $\mathrm{OR}=1.23,95 \% \mathrm{CI}=1.02-1.47$ ) [12], but no significant association was found in Italians [13]. Therefore, we performed a metaanalysis to identify the association of the two polymorphisms in $B D N F$ and AD susceptibility.

## Materials and Methods

## Search Strategy

A literature search was performed for available articles that were published in English or Chinese (up to November 2013) from the following databases: (1) Pubmed; (2) Web of Science; (3) China National Knowledge Infrastructure (CNKI); (4) Wanfang Med Online; (5) China Biology Medical literature database (CBM). The search used the following keywords: "Alzheimer's disease" or " AD " and "brain-derived neurotrophic factor" or " $B D \mathcal{N} F$ " and "polymorphism" or "mutation" or "variant". We also reviewed the references of included articles to identify additional articles not captured by our database searches.

## Inclusion Criteria

Two investigators reviewed all relevant studies independently to determine whether an individual study was eligible for inclusion. If the two investigators disagreed about the eligibility of a study, a senior researcher was invited to the discussion. The inclusion criteria were as follows: (1) case-control or cohort study published as an original study to evaluate the association between rs6265 and rs2030324 polymorphisms in $B D N F$ gene and AD susceptibility; (2) AD were diagnosed according to NINCDS-ADRDA criteria, DSM-IV criteria or CERAD criteria; (3) genotype frequencies were reported in the articles or could be calculated or ORs and $95 \%$ CIs can be obtained; (4) the genotype frequencies of controls are consistent with Hardy-Weinberg equilibrium (HWE); (5) if one study from the same population had been published more than once, we choose the most recent or complete one; (6) English or Chinese language articles were included.

## Data Extraction

Two investigators extracted the data independently and reached a consensus on all items. Information extracted from each study was as follows: first author, publication year, country, ethnicity of studied population, diagnostic criteria for cases, sample size, genotype distributions, mean age and the percentage of male.

## Statistical Analysis

The chi-square $\left(\chi^{2}\right)$ analysis was used to test deviation from Hardy-Weinberg equilibrium (HWE) for the rs6265 and rs2030324 genotype distribution of $B D N F$ gene in control groups, and $\mathrm{P}<0.05$ was considered as departure from HWE. Pooled measure was used as the inverse variance weighted mean of the logarithm of odds ratio (OR) with 95\% confidence intervals (CI) to evaluate the strength of the association of rs6265 and rs2030324 in $B D N F$ gene with risk of AD . We conducted analysis for each polymorphism considering dominant (AA+GA vs. GG for rs6265, TT+CT vs. CC for rs2030324), recessive (AA vs. GA+GG for rs6265, TT vs. CT+CC for rs2030324) and codominant (A vs. G for rs6265, T vs. C for rs2030324) models, respectively. $\mathrm{I}^{2}$ of Higgins and Thompson was used to describe heterogeneity among
studies [14]. The random- effects model (REM) was adopted if significant heterogeneity ( $\mathrm{I}^{2}>50 \%$ ) was found; otherwise, the fixed-effects model (FEM) was adopted, and Mantel-Haenszel was used to assess the fixed effects. For rs6265 and rs2030324, metaregression with restricted maximum likelihood estimation was performed to explore the potentially important covariates that might exert substantial impacts on between-study heterogeneity. Specific genetic variants causally associated with common diseases would have small effects (risk ratios mostly <2.0) [15,16]. Therefore, for sensitive analysis, we excluded the studies with $\mathrm{OR}>3.0$ or $\mathrm{OR}<0.3$ for both of the two polymorphisms (rs6265 and rs2030324) to control the impact of outlier values resulting from low cell counts within each single study on the pooled effect. Moreover, the 'leave one out' sensitive analysis was performed using $\mathrm{I}^{2}>50 \%$ as the criteria to evaluate the key studies with substantial impact on between-study heterogeneity [17]. When heterogeneity was observed, subgroup analysis was also carried out. Publication bias was evaluated by Harbord's test. An influence analysis was performed to describe how robust the pooled estimator is to removal of individual studies. If the point estimate of its omitted analysis lies outside the $95 \% \mathrm{CI}$ of the combined analysis, the individual study is suspected of excessive influence [18]. All the statistical analyses were conducted using the STATA version 10 (Stata Corporation, College Station, TX, USA). Two-tailed $\mathrm{P} \leq 0.05$ was considered as statistically significant.

## Results

## Literature Search and Study Characteristics

The search strategy identified 309 articles from English databases ( 177 articles from Web of Science, 132 articles from Pubmed), 176 articles from Chinese databases ( 161 articles from China National Knowledge Infrastructure, 9 articles from Wanfang Med Online, 6 articles from China Biology Medical literature database). 64 articles were reviewed in full-text. Furthermore, we excluded 30 articles according to the inclusion criteria, and obtained 1 additional article through references review. Finally, 35 articles were included in this meta-analysis. Figure 1 showed the flow diagram of literature search.

For rs6265, data from 29 published articles [12,13,19-45] with 32 studies were used including 7548 cases and 7334 controls. For rs2030324, 22 articles[12,13,22,25,26,29,31-36,38-40,43,46-51] with 23 studies including 5796 cases and 5706 controls were used. The ethnicity of the studies included Caucasian, Asian, African and others. All articles were case-control studies. The detailed characteristics of the two polymorphisms are showed in Table 1, Table 2, Table 3 and Table S1 in File S1.

## Influence Analysis and Publication Bias

For both of the two polymorphisms, no individual study has excessive influence on the pooled effect in any of dominant, recessive and codominant models (data not shown). Harbord's test showed no publication bias for both two polymorphisms.

## Quantitative Synthesis

Overall analysis for rs6265 and rs2030324. For overall analysis, no significant association was found between A allele and risk of AD in codominant ( REM : $\mathrm{OR}=1.03,95 \% \mathrm{CI}=0.95-$ 1.12), dominant (REM: $\mathrm{OR}=1.06,95 \% \mathrm{CI}=0.95-1.18$ ) and recessive model (FEM: $\mathrm{OR}=1.00,95 \% \mathrm{CI}=0.89-1.12$ ) for rs6265 polymorphism. No evidence indicated an association between rs2030324 with AD in codominant (REM: $\mathrm{OR}=1.06$, $95 \% \quad \mathrm{CI}=0.89-1.26$ ), dominant $\quad(\mathrm{REM}: \quad \mathrm{OR}=1.05, \quad 95 \%$


Figure 1. Flow diagram of literature search. doi:10.1371/journal.pone.0094961.g001
$\mathrm{CI}=0.86-1.27$ ) and recessive ( $\mathrm{FEM}: \mathrm{OR}=1.09,95 \% \mathrm{CI}=0.85-$ 1.39) models. The results for overall analysis are showed in Table 4. Figure 2 presented the forest plot of ORs in codominant model (A vs. G) in overall analysis for rs6265.

Evidence for heterogeneity ( $\mathrm{I}^{2}>50 \%$ ) was found in codominant and dominant models considering the association of rs6265 and rs2030324 polymorphisms with AD. Univariate meta-regression with the covariates of publication year, ethnicity, diagnostic criteria, form of AD, age (ratio of mean age in case group to that in control group) and gender (ratio of male percent in case group to that in control group) for the above-mentioned polymorphisms showed that no covariates have significant effect on between-study heterogeneity.

## Subgroup Analysis for Both Two Polymorphisms

In the stratified analysis by ethnicity, no evidence indicated the association between rs6265 and AD susceptibility both for Caucasians and Asians in codominant (Caucasian: REM: $\mathrm{OR}=1.03,95 \% \mathrm{CI}=0.90-1.17$; Asian: $\mathrm{FEM}: \mathrm{OR}=1.03,95 \%$ CI $=0.95-1.11$ ), dominant (Caucasian: $\mathrm{REM}: \mathrm{OR}=1.04,95 \%$ $\mathrm{CI}=0.88-1.22$; Asian: $\mathrm{FEM}: \mathrm{OR}=1.05,95 \% \mathrm{CI}=0.94-1.18$ ) and recessive (Caucasian: FEM : $\mathrm{OR}=1.01,95 \% \mathrm{CI}=0.81-1.25$; Asian: FEM: $\mathrm{OR}=1.01,95 \% \mathrm{CI}=0.88-1.15$ ) models. After stratified by gender, a statistical significant association between rs6265 and risk of AD was observed for females in codominant ( $\mathrm{FEM}: \mathrm{OR}=1.13,95 \% \mathrm{CI}=1.04-1.23$ ) and dominant model (FEM: OR $=1.17,95 \% \mathrm{CI}=1.05-1.31$ ), but not in recessive
Table 1. Main characteristics of BDNF gene rs6265 polymorphism genotype distributions in studies included in this meta-analysis.

| Author | Year | Country | Ethnicity | Diagnostic criteria | Mean age (case/ control) | Number (case/ control) | Genotypes GG/GA/AA |  | Allele frequency G/A |  | P for HWE |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  | AD | control | AD | control |  |
| Boiocchi et al. [13] | 2013 | Italy | Caucasian | NINCDS-ADRDA | 75/57 | 191/408 | 113/63/15 | 231/150/27 | 289/93 | 612/204 | 0.69 |
| Sonali et al. [19] | 2013 | India | Asian | DSM-IV | 64.9/64.9 | 57/63 | 3/32/22 | 12/23/28 | 38/76 | 47/79 | 0.10 |
| Ou et al. [20] | 2012 | China | Asian | NINCDS-ADRDA | 78.66/78.71 | 58/52 | 16/28/14 | 14/25/13 | 60/56 | 53/51 | 0.79 |
| Borroni et al. [21] | 2012 | Italy | Caucasian | DSM-IV | 77.6/64.2 | 234/162 | 128/87/19 | 89/63/10 | 343/125 | 241/83 | 1.00 |
| Fukumoto et al. [22] | 2010 | Japan | Asian | NINCDS-ADRDA | 73.5/67.1 | 657/525 | 218/319/120 | 197/249/79 | 755/559 | 643/407 | 1.00 |
| Qi et al. [23] | 2009 | China | Asian | NINCDS-ADRDA | 86.8/86.3 | 80/86 | 27/32/21 | 27/48/11 | 86/74 | 102/70 | 0.18 |
| Feher et al. [24] | 2009 | Hungary | Caucasian | NINCDS-ADRDA | 73.7/71.7 | 160/164 | 94/56/10 | 52/79/33 | 244/76 | 183/145 | 0.75 |
| Qian et al. [25] | 2008 | China | Asian | NINCDS-ADRDA | 72.89/72.77 | 105/105 | 28/52/25 | 25/56/24 | 108/102 | 106/104 | 0.56 |
| Cozza et al. [26] | 2008 | Italy | Caucasian | DSM-IV | $\mathrm{Na} / 64.2$ | 251/97 | 152/84/15 | 60/33/4 | 388/114 | 153/41 | 1.00 |
| Yu et al. [27] | 2008 | China | Asian | NINCDS-ADRDA | 73.25/68.82 | 99/99 | 31/41/27 | 28/51/20 | 103/95 | 107/91 | 0.84 |
| He et al. [28] | 2007 | China | Asian | NINCDS-ADRDA | 75.7/70.1 | 513/575 | 155/245/113 | 165/285/125 | 555/471 | 615/535 | 0.93 |
| Huang et al. [29] | 2007 | America | Caucasian | NINCDS-ADRDA | Na/72 | 220/128 | 150/66/4 | 98/25/5 | 366/74 | 221/35 | 0.06 |
| Forero et al. [30] | 2006 | Colombia | Mixed | NINCDS-ADRDA | 73.3/71.8 | 101/168 | 72/27/2 | 131/34/3 | 171/31 | 296/40 | 0.71 |
| Zhang et al. [31] | 2006 | America | Caucasian | NINCDS-ADRDA | 69.1/37.5 | 295/250 | 178/108/9 | 166/74/10 | 464/126 | 406/94 | 0.68 |
| Tsai et al. [32] | 2006 | China | Asian | NINCDS-ADRDA | 74.9/73.7 | 175/189 | 43/92/40 | 64/95/30 | 178/172 | 223/155 | 0.66 |
| Akatsu et al. [33] | 2006 | Japan | Asian | CERAD | 83.5/81.6 | 95/108 | 25/58/12 | 35/53/20 | 108/82 | 123/93 | 1.00 |
| Saarela et al. [34]* | 2006 | Finland | Caucasian | NINCDS-ADRDA | Na/79 | 97/101 | 62/32/3 | 81/17/3 | 156/38 | 179/23 | 0.11 |
| Lee et al. [35] | 2005 | America | Unknown | NINCDS-ADRDA | 80.4/75 | 95/70 | 45/47/3 | 32/30/8 | 137/53 | 94/46 | 0.79 |
| Desai et al. [36]* | 2005 | America | Caucasian | NINCDS-ADRDA | Na/75.7 | 995/671 | 662/299/34 | 456/197/18 | 1623/367 | 1109/233 | 0.69 |
| Desai et al. [36]* | 2005 | America | African | NINCDS-ADRDA | Na/70.49 | 64/45 | 59/5/0 | 42/3/0 | 123/5 | 87/3 | 1.00 |
| Bian et al. [37] | 2005 | China | Asian | NINCDS-ADRDA | Na/70.2 | 203/239 | 49/113/41 | 73/115/51 | 211/195 | 261/217 | 0.70 |
| Vepsalainen et al. [38] | 2005 | Finland | Caucasian | NINCDS-ADRDA | Na/70 | 375/460 | 280/87/8 | 342/109/9 | 647/103 | 793/127 | 0.85 |
| Matsushita et al. [12] | 2005 | Japan | Asian | NINCDS-ADRDA | 76.1/75.2 | 487/471 | 171/247/69 | 150/223/98 | 589/385 | 523/419 | 0.41 |
| Nishimura et al. [39] | 2005 | Japan | Asian | NINCDS-ADRDA | 77.4/68.4 | 172/275 | 61/85/26 | 88/140/47 | 207/137 | 316/234 | 0.54 |
| Bodner et al. [40] | 2005 | America | Caucasian | NINCDS-ADRDA | 83/81 | 256/194 | 163/85/8 | 126/62/6 | 411/101 | 314/74 | 0.82 |
| Li et al. [41]* | 2005 | England | Caucasian | NINCDS-ADRDA | $\mathrm{Na} / \mathrm{Na}$ | 359/396 | 239/105/15 | 269/114/13 | 583/135 | 652/140 | 0.86 |
| Li et al. [41]* | 2005 | America | Caucasian | NINCDS-ADRDA | $\mathrm{Na} / \mathrm{Na}$ | 188/361 | 109/73/6 | 235/110/16 | 291/85 | 580/142 | 0.51 |
| Li et al. [41]* | 2005 | America | Caucasian | NINCDS-ADRDA | $\mathrm{Na} / \mathrm{Na}$ | 388/349 | 251/126/11 | 237/105/7 | 628/148 | 579/119 | 0.34 |
| Nacmias et al. [42] | 2004 | Italy | Caucasian | NINCDS-ADRDA | 72.2/72.9 | 83/97 | 48/29/6 | 55/38/4 | 125/41 | 148/46 | 0.58 |
| Bagnoli et al. [43] | 2004 | Italy | Caucasian | NINCDS-ADRDA | 71.1/72.9 | 128/97 | 62/60/6 | 55/38/4 | 184/72 | 148/46 | 0.58 |
| Combarros et al. [44] | 2004 | Spain | Caucasian | NINCDS-ADRDA | 75.3/79.9 | 237/218 | 149/78/10 | 143/67/8 | 376/98 | 353/83 | 1.00 |
| Ventriglia et al. [45] | 2002 | Italy | Caucasian | NINCDS-ADRDA | 72/Na | 130/111 | 85/33/12 | 54/48/9 | 203/57 | 156/66 | 0.82 |

Table 2. Genotype and allele distribution for rs6265 polymorphism in female and other subgroups.

| Author | Year | Country | Ethnicity | Mean age (case/ control) | Genotypes GG/GA/AA |  | Allele frequency ( $\mathbf{G} / \mathbf{A}$ ) |  | Genotypes GG/GA/AA |  | Genotypes GG/GA/AA |  | Allele frequency (G/A) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | AD(female) | control | AD(female) | control | AD(male) | control | LOAD | control | LOAD | control |
| Boiocchi et al. [13] | 2013 | Italy | Caucasian | 75/57 | 69/42/9 | 130/81/14 | 180/60 | 341/109 | 109/33 | 271/95 | Na | Na | Na | Na |
| Ou et al. [20] | 2012 | China | Asian | 78.66/78.71 | 7/6/5 | 4/7/5 | 20/16 | 15/17 | 219/171 | 251/235 | Na | Na | Na | Na |
| Fukumoto et al. [22] | 2010 | Japan | Asian | 73.5/67.1 | 142/205/80 | 122/143/40 | 489/365 | 387/223 | 40/40 | 38/34 | Na | Na | Na | Na |
| Yu et al [27] | 2008 | China | Asian | 73.25/68.82 | 13/11/13 | 11/27/11 | 37/37 | 49/49 | 47/21 | 40/24 | Na | Na | Na | Na |
| He et al. [28] | 2007 | China | Asian | 75.7/70.1 | 92/152/74 | 97/170/65 | 336/300 | 364/300 | 34/16 | 48/24 | 131/189/93 | 128/208/94 | 451/375 | 464/396 |
| Forero et al. [30] | 2006 | Colombia | Mixed | 73.3/71.8 | 51/20/2 | 90/23/2 | 122/24 | 203/27 | 49/7 | 93/13 | Na | Na | Na | Na |
| Tsai et al. [32] | 2006 | China | Asian | 74.9/73.7 | 19/50/15 | 33/50/18 | 88/80 | 116/86 | 530/128 | 420/100 | Na | Na | Na | Na |
| Akatsu et al. [33] | 2006 | Japan | Asian | 83.5/81.6 | 16/36/6 | 30/42/14 | 68/48 | 102/70 | 35/1 | 23/1 | Na | Na | Na | Na |
| Saarela et al. [34]* | 2006 | Finland | Caucasian | Na/79 | 45/21/2 | 46/10/0 | 111/25 | 102/10 | 104/86 | 142/126 | 62/32/3 | 81/17/3 | 156/38 | 179/23 |
| Lee et al. [35] | 2005 | America | Unknown | 80.4/75 | 31/28/2 | 20/14/4 | 90/32 | 54/22 | 185/109 | 161/139 | Na | Na | Na | Na |
| Desai et al. [36]* | 2005 | America | Caucasian | $\mathrm{Na} / 75.7$ | 449/201/19 | 287/115/9 | 1099/239 | 689/133 | 90/92 | 107/69 | 662/299/34 | 456/197/18 | 1623/367 | 1109/233 |
| Desai et al. [36]* | 2005 | America | African | $\mathrm{Na} / 70.49$ | 42/4/0 | 31/2/0 | 88/4 | 64/2 | 40/34 | 21/23 | 59/5/0 | 42/3/0 | 123/5 | 87/3 |
| Bian et al. [37] | 2005 | China | Asian | Na/70.2 | 20/67/21 | 36/47/22 | 107/109 | 119/91 | 45/13 | 77/13 | 34/73/27 | 58/90/42 | 141/127 | 206/174 |
| Matsushita et al. [12] | 2005 | Japan | Asian | 76.1/75.2 | 117/170/53 | 104/154/63 | 404/276 | 362/280 | 66/58 | 58/42 | 137/195/54 | 150/223/98 | 469/303 | 523/419 |
| Li et al. [41]* | 2005 | England | Caucasian | $\mathrm{Na} / \mathrm{Na}$ | 178/73/14 | 192/73/5 | 429/101 | 457/83 | 266/194 | 256/184 | 239/105/15 | 269/114/13 | 583/135 | 652/140 |
| Li et al. [41]* | 2005 | America | Caucasian | $\mathrm{Na} / \mathrm{Na}$ | 51/32/4 | 150/67/9 | 134/40 | 367/85 | 115/37 | 99/27 | 109/73/6 | 235/110/16 | 291/85 | 580/142 |
| Li et al. [41]* | 2005 | America | Caucasian | $\mathrm{Na} / \mathrm{Na}$ | 163/81/4 | 150/60/5 | 407/89 | 360/70 | 118/26 | 140/38 | 251/126/11 | 237/105/7 | 628/148 | 579/119 |
| Nacmias et al. [42] | 2004 | Italy | Caucasian | 72.2/72.9 | 36/19/3 | 39/22/0 | 91/25 | 100/22 | 146/42 | 201/51 | Na | Na | Na | Na |
| Combarros et al. [44] | 2004 | Spain | Caucasian | 75.3/79.9 | 107/47/7 | 105/44/6 | 261/61 | 254/56 | 221/59 | 219/49 | Na | Na | Na | Na |

[^0]Table 3. Main characteristics of BDNF gene rs2030324 polymorphism genotype and allele distributions in studies included in this meta-analysis.

| Author | Year | Country | Ethnicity | Diagnostic criteria | Mean age (case) control) | Number (case/ control) | Genotypes CC/CT/TT |  | Allele frequency C/T |  | P for HWE |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  | AD | control | AD | control |  |
| Boiocchi et al. [13] | 2013 | Italy | Caucasian | NINCDS-ADRDA | 75/57 | 192/384 | 55/93/44 | 103/192/89 | 203/181 | 398/370 | 1.00 |
| Cousin et al. [46] | 2011 | France | Caucasian | NINCDS-ADRDA | Na/66.2 | 425/470 | 370/54/1 | 419/50/1 | 794/56 | 888/52 | 1.00 |
| Fukumoto et al. [22] | 2010 | Japan | Asian | NINCDS-ADRDA | 73.5/67.1 | 657/525 | 611/45/1 | 490/34/1 | 1267/47 | 1014/36 | 0.46 |
| Hou et al. [47]* | 2009 | China | Asian | NINCDS-ADRDA | 79.21/76.07 | 203/138 | 172/31/0 | 117/21/0 | 375/31 | 255/21 | 1.00 |
| Qian et al. [25] | 2008 | China | Asian | NINCDS-ADRDA | 72.89/72.77 | 105/105 | 104/1/0 | 96/8/1 | 209/1 | 200/10 | 0.20 |
| Cozza et al. [26] | 2008 | Italy | Caucasian | DSM-IV | Na/64.2 | 251/97 | 212/35/4 | 80/15/2 | 459/43 | 175/19 | 0.22 |
| Huang et al. [29] | 2007 | America | Caucasian | NINCDS-ADRDA | Na/72 | 220/128 | 202/16/2 | 113/15/0 | 420/20 | 241/15 | 1.00 |
| Zhang et al. [31] | 2006 | America | Caucasian | NINCDS-ADRDA | 69.1/37.5 | 295/250 | 271/22/2 | 220/30/0 | 564/26 | 470/30 | 1.00 |
| Tsai et al. [32] | 2006 | China | Asian | NINCDS-ADRDA | 74.9/73.7 | 175/189 | 151/24/0 | 167/20/2 | 326/24 | 354/24 | 0.16 |
| Akatsu et al. [33] | 2006 | Japan | Asian | CERAD | 83.5/81.6 | 95/108 | 89/6/0 | 101/7/0 | 184/6 | 209/7 | 1.00 |
| Saarela et al. [34]* | 2006 | Finland | Caucasian | NINCDS-ADRDA | Na/79 | 97/101 | 88/9/0 | 81/19/1 | 185/9 | 181/21 | 1.00 |
| Lee et al. [35] | 2005 | American | Caucasian | NINCDS-ADRDA | 80.4/75 | 106/73 | 102/4/0 | 66/7/0 | 208/4 | 139/7 | 1.00 |
| Desai et al. [36]* | 2005 | American | Caucasian | NINCDS-ADRDA | Na/75.7 | 719/523 | 629/86/4 | 454/69/0 | 1344/94 | 977/69 | 0.15 |
| Desai et al. [36]* | 2005 | American | African | NINCDS-ADRDA | Na/70.49 | 58/42 | 54/4/0 | 38/4/0 | 112/4 | 80/4 | 1.00 |
| Olin et al. [48] | 2005 | American | Caucasian | Na | 77.7/66 | 212/202 | 173/36/3 | 189/13/0 | 382/42 | 391/13 | 1.00 |
| Vepsalainen et al. [38] | 2005 | Finland | Caucasian | NINCDS-ADRDA | Na/70 | 375/460 | 90/199/86 | 124/239/97 | 379/371 | 487/433 | 0.40 |
| Matsushita et al. [12] | 2005 | Japan | Asian | NINCDS-ADRDA | 76.1/75.2 | 487/471 | 457/30/0 | 438/33/0 | 944/30 | 909/33 | 1.00 |
| Nishimura et al. [39] | 2005 | Japan | Asian | NINCDS-ADRDA | 77.4/68.4 | 172/275 | 154/18/0 | 264/11/0 | 326/18 | 539/11 | 1.00 |
| Bodner et al. [40] | 2005 | American | Caucasian | NINCDS-ADRDA | 83/81 | 256/194 | 230/26/0 | 175/19/0 | 486/26 | 369/19 | 1.00 |
| Nishimura et al. [49]* | 2004 | Brazil | Caucasian | NINCDS-ADRDA | 68.7172.3 | 188/188 | 175/13/0 | 170/17/1 | 363/13 | 357/19 | 0.38 |
| Bagnoli et al. [43] | 2004 | Italy | Caucasian | NINCDS-ADRDA | 71.1/72.9 | 128/97 | 113/14/1 | 83/14/0 | 240/16 | 180/14 | 1.00 |
| Riemenschneider et al. [50] | 2002 | German | Caucasian | NINCDS-ADRDA | 69.3/65.6 | 210/188 | 185/24/1 | 175/13/0 | 394/26 | 363/13 | 1.00 |
| Kunugi et al. [51] | 2001 | Japan | Asian | NINCDS-ADRDA | 74/55 | 170/498 | 150/19/1 | 477/21/0 | 319/21 | 975/21 | 1.00 |

Table 4. The overall pooled measure on the relation of BDNF gene rs6265 and rs2030324 polymorphism with AD.

| Loci | Inherited model | All included articles |  |  |  |  | After excluding articles with $\mathrm{OR}>3.0$ or $\mathrm{OR}<\mathbf{0 . 3}$ |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Number | REM Pooled OR (95\% CI) | FEM Pooled OR (95\% CI) | Q-value | $1^{2}$ (\%) | Number | REM Pooled OR (95\% CI) | FEM Pooled OR (95\% CI) | Q-value | $\begin{aligned} & \mathrm{I}^{2} \\ & (\%) \end{aligned}$ | Articles Excluded |
| rs6265 | Codominant | 7548/7334 | 1.03(0.95-1.12) | 1.03(0.97-1.08) | 63.14 | 50.9 | 7548/7334 | - | - | - | - | - |
|  | Dominant | 7548/7334 | 1.06(0.95-1.18) | 1.05(0.98-1.13) | 63.92 | 51.5 | 7491/7271 | 1.05(0.94-1.16) | 1.05(0.98-1.12) | 59.64 | 49.7 | [19] |
|  | Recessive ${ }^{\text {a }}$ | 7484/7289 | 1.00(0.86-1.17) | 1.00(0.89-1.12) | 42.74 | 29.8 | 7313/7100 | 1.04(0.93-1.17) | 1.04(0.93-1.17) | 28.00 | 0 | [24,35] |
| rs2030324 | Codominant | 5796/5706 | 1.06(0.89-1.26) | 1.06(0.96-1.17) | 52.13 | 57.8 | 5309/4901 | 0.99(0.88-1.12) | 1.01(0.91-1.12) | 22.49 | 15.5 | [ $25,48,51]$ |
|  | Dominant | 5796/5706 | 1.05(0.86-1.27) | 1.06(0.94-1.19) | 51.64 | 57.4 | 5309/4901 | 0.98(0.85-1.13) | 0.99(0.88-1.12) | 24.11 | 21.2 | [ $25,48,51]$ |
|  | Recessive ${ }^{\text {b }}$ | 4419/4405 | 1.09(0.85-1.39) | 1.09(0.85-1.39) | 15 | 0 | 2848/2743 | 1.05(0.82-1.34) | 1.05(0.82-1.34) | 10 | 0 | [31,32,36,48,51] |
| ${ }^{\text {a }}$ One study about African (Desaia et al)for recessive model was not sufficient to calculated pooled OR. <br> ${ }^{\text {b }}$ Seven studies (Hou et al. Akatsu et al. Lee et al. Desaia et al. Matsushita et al. Masataka et al. Bodner et al.) for recessive model were not sufficient to calculated pooled Abbreviations: AD, Alzheimer's disease; FEM, fixed-effects model; REM, random-effects model. <br> rs6265: Codominant model, A vs. G; Dominant model, AA+GA vs. GG; Recessive model, AA vs. GA+GG. rs2030324: Codominant model, T vs C; Dominant model, TT+CT vs. CC; Recessive model, $\Pi$ vs. CT+CC. doi:10.1371/journal.pone.0094961.t004 |  |  |  |  |  |  |  |  |  |  |  |  |

model (FEM: OR $=1.13,95 \%$ CI $=0.95-1.35$ ). However, no association was detected for males in codominant model (FEM: $\mathrm{OR}=0.98,95 \% \mathrm{CI}=0.88-1.08)$, dominant model $(\mathrm{OR}=1.01$, $95 \% \mathrm{CI}=0.88-1.16$ ) and recessive model (FEM: OR $=0.91,95 \%$ $\mathrm{CI}=0.74-1.12$ ). When stratified by ethnicity in females, for Caucasian females, the A allele was found contributing significantly to the increased risk of AD in codominant model (FEM: $\mathrm{OR}=1.18,95 \% \mathrm{CI}=1.03-1.34$ ) and dominant model (FEM: OR $=1.18,95 \% \mathrm{CI}=1.01-1.37$ ), not in recessive model (FEM: $\mathrm{OR}=1.40,95 \% \mathrm{CI}=0.94-2.10)$. With regard to Asian females, no association was detected in codominant model (FEM: $\mathrm{OR}=1.09, \quad 95 \% \mathrm{CI}=0.98-1.22$ ), dominant model (FEM: OR $=1.15,95 \% \mathrm{CI}=0.98-1.36$ ) and recessive model (FEM: $\mathrm{OR}=1.09,95 \% \mathrm{CI}=0.90-1.33$ ). When stratified by gender in LOAD patients, the A allele was observed significantly associated with AD in female LOAD patients in codominant model (FEM: $\mathrm{OR}=1.22,95 \% \mathrm{CI}=1.05-1.41$ ) and dominant model (FEM: $\mathrm{OR}=1.23,95 \% \mathrm{CI}=1.03-1.46$ ), but not in recessive model (FEM: OR $=1.47,95 \% \mathrm{CI}=0.88-2.44$ ). However, no association was found in male LOAD patients in any of the above-mentioned models. Figure 3 presented the forest plot of ORs in codominant model (A vs. G) in female group for rs6265.

For rs2030324, in the subgroup analysis, for Caucasian, no association was found in codominant model (REM: OR $=1.00$, $95 \% \mathrm{CI}=0.82-1.21$ ), dominant model ( $\mathrm{REM}: \mathrm{OR}=0.97,95 \%$ $\mathrm{CI}=0.78-1.22$ ) and recessive model (FEM: $\mathrm{OR}=1.10,95 \%$ $\mathrm{CI}=0.86-1.40$ ). With regard to Asians, there was no association between T allele and AD susceptibility in codominant model (REM: OR $=1.22,95 \% \mathrm{CI}=0.80-1.85$ ), dominant model (REM: $\mathrm{OR}=1.24,95 \% \mathrm{CI}=0.82-1.89$ ) and recessive model (FEM: $\mathrm{OR}=0.81,95 \% \mathrm{CI}=0.18-3.71)$. When stratified by gender, there was still no association between rs2030324 and AD both for females and males in codominant model (Female: FEM $\mathrm{OR}=0.97,95 \% \mathrm{CI}=0.75-1.26$; Male: $\mathrm{FEM} \mathrm{OR}=1.02,95 \%$ $\mathrm{CI}=0.74-1.40$ ), dominant model (Female: FEM: OR $=1.02,95 \%$ $\mathrm{CI}=0.74-1.40$; Male: $\mathrm{OR}=0.87,95 \% \mathrm{CI}=0.56-1.34)$. The pooled results are summarized in Table 5 and Table S2 and Table S3 in File S1.

## Sensitivity Analysis

After excluding articles with $\mathrm{OR}>3.0$ or $\mathrm{OR}<0.3$, low heterogeneity ( $\mathrm{I}^{2}<50 \%$ ) was found in the codominant, dominant and recessive models, and the results for both two polymorphisms are consistent with the ones without sensitive analysis. Furthermore, the "leave one out" analysis was carried out when significant heterogeneity ( $\mathrm{I}^{2}>50 \%$ ) was observed. After excluded key articles with substantial impact on between-study heterogeneity, the data showed rs6265 confers a risk effect for Caucasians in codominant model (FEM: $\mathrm{OR}=1.08,95 \% \mathrm{CI}=1.00-1.17$, $\mathrm{P}=0.042$ ). With respect to rs2030324, the result indicated T allele was associated with EOAD patients in codominant (FEM: $\mathrm{OR}=2.02,95 \% \mathrm{CI}=1.18-3.47$ ) and dominant model (FEM: $\mathrm{OR}=2.01,95 \% \mathrm{CI}=1.15-3.52$ ).

## Discussion

To our knowledge, many case-control studies have been carried out to investigate the role of $B D N F$ gene in the development of AD. However, these results remained controversial. 29 articles concerning rs6265 and 22 articles concerning rs2030324 were included in our meta-analysis. With respect to rs2030324, there was no evidence for an association with AD. What's more, the combined evidence suggested that rs6265 was not associated with AD for overall analysis. After stratified by ethnicity in females, our


Figure 2. Forest plots of relationship between BDNF gene rs6265 polymorphism and AD risk in codominant model (A vs. G) for overall analysis. White diamond donates the pooled OR. Black squares indicate the OR in each study, with square sizes inversely proportional to the standard error of the OR. Horizontal lines represent 95\% Cls. doi:10.1371/journal.pone.0094961.g002
data indicated rs6265 lead to the increased risk of AD in Caucasian females, but not for Asians. When stratified by gender in LOAD patients, the A allele was found contributing significantly to the increased risk of AD in female LOAD patients. The results of our study were not consistent with the previous meta-analysis conducted by Fukumoto et al. [22] in 2010 including 16 studies revealing a gender-related association between rs6265 polymorphism and AD susceptibility. The reason might be that the quantitative assessments in our study were based on a larger sample size and we have performed a detailed subgroup analysis.

The Met66-BDNF protein has been shown to be associated with reduced transport of BDNF from the Golgi region to appropriate secretory granules in neurons, compared with the

Val66-BDNF protein [52]. Moreover, the A allele of rs6265 was related with poorer episodic memory, abnormal hippocampal activation, and lower hippocampal n-acetyl aspartate (NAA) in human subjects [53]. Epidemiological studies showed higher incidence and prevalence of AD in women than in men $[3,5]$. Molecular mechanisms underlying this correlation had been considered as colocalization estrogen receptors with BDNFsynthesizing neurons in the forebrain [54] and induction BDNF expression by estrogen through the estrogen response element [55]. In addition, though the two forms of AD (EOAD and LOAD) have different patterns of genetic epidemiology [4], the results of our study revealed no difference between them.


Figure 3. Forest plots of relationship between BDNF gene rs6265 polymorphism and AD risk in codominant model (A vs. G) for female group. White diamond donates the pooled OR. Black squares indicate the OR in each study, with square sizes inversely proportional to the standard error of the OR. Horizontal lines represent 95\% Cls.
doi:10.1371/journal.pone.0094961.g003

Between-study heterogeneity is common in meta-analysis for genetic association studies [56] and it is essential to explore the potential sources of between-study heterogeneity [57]. This metaanalysis also showed significant between-study heterogeneity in dominant model and codominant model for both two polymorphisms. An indeterminate number of characteristics that varied among studies could be the sources of between-study heterogeneity, such as publication year, ethnicity, diagnostic criteria, form of AD , age, and gender etc. Therefore, in order to explore the potential sources of between-study heterogeneity for rs6265 and rs2030324, meta-regression was adopted. However, the aforementioned covariates were not important contributors to this disease-effect heterogeneity. Considering that our meta-analysis showed significant heterogeneity, subgroup analyses by ethnicity (Caucasian and Asian), form of AD (EOAD and LOAD) and gender (female and male) etc. were performed to explore the sources of heterogeneity. However, between-study heterogeneity still existed in subgroups, suggesting the presence of other unknown confounders. AD is a complex multi-factorial disease and is related with the combined effects between gene variants and
environmental factors. Therefore, other genetic and environment variables, as well as their possible interaction, may be potential contributors to this disease-effect unconformity. We further conducted a sensitivity analysis excluding articles with $\mathrm{OR}>3$ or $\mathrm{OR}<0.3$. After sensitivity analysis, low heterogeneity ( $\mathrm{I}^{2}<50 \%$ ) was found in the codominant, dominant and recessive models, and the results are consistent with the one before heterogeneity analysis for rs6265 and rs2030324, strongly identified the stability of our results. Moreover, no publication bias was found in any of the above-mentioned models for both two polymorphisms.

The major strength of our meta-analysis is that the results were based on the large number of participators, allowing a much greater possibility of reaching definitive conclusions. Additionally, we conducted subgroup analyses to explore the potential sources of heterogeneity, and sensitivity analysis was carried out to ensure the stability of our results. However, our meta-analysis also had some limitations. Firstly, lack of the original data of included articles made it impracticable to excluded potential confounders completely, especially the confounding of age. Secondly, different diagnostic criteria may have possible influence on the diagnosis of AD.
Table 5. Subgroup analysis on the relation of BDNF gene rs6265 and rs2030324polymorphism with AD in codominant model.

| Loci | Data | Type | Inherited model | All included articles |  |  |  |  | After excluding articles with $\mathrm{OR}>3.0$ or $\mathrm{OR}<\mathbf{0 . 3}$ |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Number | REM Pooled OR (95\% CI) | FEM Pooled OR (95\% CI) | Q-valu | $1^{2}$ (\%) | Number | REM Pooled OR (95\% CI) | FEM Pooled OR (95\% CI) | Qvalue | $I^{2} \text { (\%) }$ | Articles Excluded |
| rs6265 | Ethnicity | Caucasian | Codominant | 4587/4264 | 1.03(0.90-1.17) | 1.03(0.95-1.11) | 45.2 | 64.6 | - | - | - |  | - | - |
|  |  | Asian | Codominant | 2701/2787 | 1.04(0.94-1.14) | 1.03(0.95-1.11) | 15.67 | 29.8 | - | - | - | - | - | - |
|  | Gender | Female | Codominant | 3246/3120 | 1.13(1.04-1.23) | 1.13(1.04-1.23) | 18 | 0 | - | - | - | - | - | - |
|  |  | Male | Codominant | 1813/1976 | 0.98(0.88-1.08) | 0.98(0.88-1.08) | 18 | 0 | - | - | - | - | - | - |
|  | Form | EOAD | Codominant | 270/665 | 0.96(0.78-1.19) | 0.96(0.78-1.19) | 2 | 0 | - | - | - |  | - | - |
|  |  | LOAD | Codominant | 3024/3014 | 1.05(0.93-1.18) | 1.03(0.94-1.12) | 12.97 | 38.3 | - | - | - |  | - | - |
|  | Female | Caucasian | Codominant | 1676/1619 | 1.18(1.03-1.34) | 1.18(1.03-1.34) | 7 | 0 | - | - | - | - | - | - |
|  |  | Asian | Codominant | 1390/1315 | 1.10(0.97-1.24) | 1.09(0.98-1.22) | 8.37 | 16.4 | - | - | - | - | - | - |
|  | Male | Caucasian | Codominant | 836/936 | 1.04(0.88-1.23) | 1.04(0.88-1.23) | 7 | 0 | - | - | - | - | - | - |
|  |  | Asian | Codominant | 897/943 | 0.97(0.80-1.16) | 0.95(0.83-1.08) | 12.13 | 42.3 | - | - | - | - | - | - |
|  | LOAD | Female | Codominant | 1383/1211 | 1.22(1.05-1.41) | 1.22(1.05-1.41) | 5 | 0 | - | - | - | - | - | - |
|  |  | Male | Codominant | 682/666 | 1.07(0.88-1.30) | 1.07(0.88-1.30) | 5 | 0 |  | - | - | - | - | - |
| rs2030324 | Ethnicity | Caucasian | Codominant | 3674/3355 | 1.00(0.82-1.21) | 1.03(0.92-1.15) | 28.63 | 54.6 | 3462/3153 | 0.96(0.84-1.11) | 0.99(0.89-1.11) | 15.35 | 21.8 | [48] |
|  |  | Asian | Codominant | 2064/2309 | 1.22(0.80-1.85) | 1.22(0.98-1.52) | 21.41 | 67.3 | 1789/1706 | 1.11 (0.85-1.47) | 1.10(0.86-1.39) | 6.34 | 21.1 | [25,51] |
|  | Form | EOAD | Codominant | 303/1359 | 1.39(0.66-2.90) | 1.55(0.95-2.53) | 6.3 | 52.4 | - | - | - | - | - | - |
|  |  | LOAD | Codominant | 2038/2351 | 1.21(0.77-1.89) | 1.18(0.97-1.43) | 36.2 | 77.9 | 1760/1651 | 0.91(0.74-1.13) | 0.91(0.74-1.13) | 6 | 0 | [48,51] |
|  | Gender | Female | Codominant | 696/603 | 0.97(0.75-1.26) | 0.97(0.75-1.26) | 2 | 0 | - | - | - | - | - | - |
|  |  | Male | Codominant | 356/444 | 1.02(0.74-1.40) | 1.02(0.74-1.40) | 2 | 0 | - | - | - |  | - | - |

[^1]In conclusion, this meta-analysis suggested A allele of rs6265 might increase the risk of AD in Caucasian females and female LOAD patients. In addition, no evidence indicated an association between rs2030324 with AD. Since potential biases and confounders could not be ruled out completely in this study, further studies are needed to confirm these results.

## Supporting Information

File S1 This includes the files Search Strategy S1 and Tables S1 to S3. Search Strategy S1. Keywords of literature search for different database. Table S1. Genotype and allele distribution for rs2030324 polymorphism in female and other subgroups. Table S2. Subgroup analysis on the relation of BDNF

## References

1. Hashimoto R, Hirata Y, Asada T, Yamashita F, Nemoto K, et al. (2009) Effect of the brain-derived neurotrophic factor and the apolipoprotein E polymorphisms on disease progression in preclinical Alzheimer's disease. Genes Brain Behav 8: 43-52.
2. World Health Organization Website. Available: http://www.who.int/mental_ health/publications/dementia_report_2012/en/. Accessed 2014 Nov 15.
3. Pubmed Website. Available: http://www.ncbi.nlm.nih.gov/pubmed/20298981. Accessed 2014 Nov 15.
4. Reitz C, Brayne C, Mayeux R (2011) Epidemiology of Alzheimer disease. Nat Rev Neurol 7: 137-152.
5. Pubmed Website. Available: http://www.ncbi.nlm.nih.gov/pubmed/21702424. Accessed 2014 Nov 15.
6. Campion D, Dumanchin C, Hannequin D, Dubois B, Belliard S, et al. (1999) Early-onset autosomal dominant Alzheimer disease: prevalence, genetic heterogeneity, and mutation spectrum. Am J Hum Genet 65: 664-670.
7. Saunders AM, Strittmatter WJ, Schmechel D, George-Hyslop PH, PericakVance MA, et al. (1993) Association of apolipoprotein E allele epsilon 4 with late-onset familial and sporadic Alzheimer's disease. Neurology 43: 1467-1472.
8. Maezawa I, Jin LW, Woltjer RL, Maeda N, Martin GM, et al. (2004) Apolipoprotein E isoforms and apolipoprotein AI protect from amyloid precursor protein carboxy terminal fragment-associated cytotoxicity. J Neurochem 91: 1312-1321.
9. Rogaeva E, Meng Y, Lee JH, Gu Y, Kawarai T, et al. (2007) The neuronal sortilin-related receptor SORL1 is genetically associated with Alzheimer disease. Nat Genet 39: 168-177.
10. Fahnestock M, Garzon D, Holsinger RM, Michalski B (2002) Neurotrophic factors and Alzheimer's disease: are we focusing on the wrong molecule? J Neural Transm Suppl 241-252.
11. Connor B, Young D, Yan Q, Faull RL, Synek B, et al. (1997) Brain-derived neurotrophic factor is reduced in Alzheimer's disease. Brain Res Mol Brain Res 49: 71-81.
12. Matsushita S, Arai H, Matsui T, Yuzuriha T, Urakami K, et al. (2005) Brainderived neurotrophic factor gene polymorphisms and Alzheimer's disease. J Neural Transm 112: 703-711.
13. Boiocchi C, Maggioli E, Zorzetto M, Sinforiani E, Cereda C, et al. (2013) Brainderived neurotrophic factor gene variants and Alzheimer disease: an association study in an Alzheimer disease Italian population. Rejuvenation Res 16: 57-66.
14. Higgins JP, Thompson SG (2002) Quantifying heterogeneity in a meta-analysis. Stat Med 21: 1539-1558.
15. Ioannidis JP (2006) Commentary: grading the credibility of molecular evidence for complex diseases. Int J Epidemiol 35: 572-578; discussion 593-576.
16. Khoury MJ, Little J, Gwinn M,Ioannidis JP (2007) On the synthesis and interpretation of consistent but weak gene-disease associations in the era of genome-wide association studies. Int J Epidemiol 36: 439-445.
17. Kangas-Kontio T, Vavuli S, Kakko SJ, Penna J, Savolainen ER, et al. (2009) Polymorphism of the manganese superoxide dismutase gene but not of vascular endothelial growth factor gene is a risk factor for diabetic retinopathy. Br J Ophthalmol 93: 1401-1406.
18. Tobias A (1999) Assessing the influence of a single study in the meta-analysis estimate. Stata Technical Bulletin 8.
19. Sonali N, Tripathi M, Sagar R, Vivekanandhan S (2013) Val66Met polymorphism and BDNF levels in Alzheimer's disease patients in North Indian population. Int J Neurosci 123: 409-416.
20. Ou LY (2012) The assoiation research between BDNF gene G196A polymorphism and sporadic Alzheimer's Disease[dissertation]. Guangxi Medical University.
21. Borroni B, Bianchi M, Premi E, Alberici A, Archetti S, et al. (2012) The brainderived neurotrophic factor Val66Met polymorphism is associated with reduced hippocampus perfusion in frontotemporal lobar degeneration. J Alzheimers Dis 31: 243-251.
22. Fukumoto N, Fujii T, Combarros O, Kamboh MI, Tsai SJ, et al. (2010) Sexually dimorphic effect of the Val66Met polymorphism of BDNF on susceptibility to
gene rs6265 polymorphism with AD in dominant and recessive model. Table S3. Subgroup analysis on the relation of BDNF gene rs2030324 polymorphism with AD in dominant and recessive model.
(DOC)

## Checklist S1 PRISMA checklist. (DOC)

## Author Contributions

Conceived and designed the experiments: YL. Performed the experiments: YL SC ZTX DFZ. Analyzed the data: YL SC ZTX DFZ. Contributed reagents/materials/analysis tools: YL SC ZTX DFZ. Wrote the paper: YL SC.

Alzheimer's disease: New data and meta-analysis. Am J Med Genet B Neuropsychiatr Genet 153B: 235-242.
23. Qi HM, Cha JM, Sheng YL, Qian Y (2009) Association of brain-derived neurotrophic factor and apolipoprotein gene polymorphisms with sporadic Alzheimer's disease. Heilongjiang Medical journal 33: 724-726.
24. Feher A, Juhasz A, Rimanoczy A, Kalman J, Janka Z (2009) Association between BDNF Val66Met polymorphism and Alzheimer disease, dementia with Lewy bodies, and Pick disease. Alzheimer Dis Assoc Disord 23: 224-228.
25. Qian Y, Zhang ZJ, Zhang XB, Yong.Gui Y, Yu H, et al. (2008) Haplotype analysis of Alzheimer's disease BDNF gene G196A and C270T polymorphisms. Journal of Southeast University 27: 161-165.
26. Cozza A, Melissari E, Iacopetti P, Mariotti V, Tedde A, et al. (2008) SNPs in neurotrophin system genes and Alzheimer's disease in an Italian population. J Alzheimers Dis 15: 61-70.
27. Yu H, Zhang Z, Shi Y, Bai F, Xie C, et al. (2008) Association study of the decreased serum BDNF concentrations in amnestic mild cognitive impairment and the Val66Met polymorphism in Chinese Han. J Clin Psychiatry 69: 1104 1111.
28. He XM, Zhang ZX, Zhang JW, Zhou YT, Tang MN, et al. (2007) Lack of association between the BDNF gene Val66Met polymorphism and Alzheimer disease in a Chinese Han population. Neuropsychobiology 55: 151-155.
29. Huang R, Huang J, Cathcart H, Smith S, Poduslo SE (2007) Genetic variants in brain-derived neurotrophic factor associated with Alzheimer's disease. J Med Genet 44: e66.
30. Forero DA, Benitez B, Arboleda G, Yunis JJ, Pardo R, et al. (2006) Analysis of functional polymorphisms in three synaptic plasticity-related genes (BDNF, COMT AND UCHL1) in Alzheimer's disease in Colombia. Neurosci Res 55: 334-341.
31. Zhang H, Ozbay F, Lappalainen J, Kranzler HR, van Dyck CH, et al. (2006) Brain derived neurotrophic factor (BDNF) gene variants and Alzheimer's disease, affective disorders, posttraumatic stress disorder, schizophrenia, and substance dependence. Am J Med Genet B Neuropsychiatr Genet 141B: 387393.
32. Tsai SJ, Hong CJ, Liu HC, Liu TY, Liou YJ (2006) The brain-derived neurotrophic factor gene as a possible susceptibility candidate for Alzheimer's disease in a chinese population. Dement Geriatr Cogn Disord 21: 139-143.
33. Akatsu H, Yamagata HD, Kawamata J, Kamino K, Takeda M, et al. (2006) Variations in the BDNF gene in autopsy-confirmed Alzheimer's disease and dementia with Lewy bodies in Japan. Dement Geriatr Cogn Disord 22: 216222.
34. Saarela M LL, Juha O (2006) No association between the brain-derived neurotrophic factor $196 \mathrm{G}>\mathrm{A}$ or 270C $>\mathrm{T}$ polymorphisms and Alzheimer's or Parkinson's disease. Folia Neuropathol 44: 12-16.
35. Lee J, Fukumoto H, Orne J, Klucken J, Raju S, et al. (2005) Decreased levels of BDNF protein in Alzheimer temporal cortex are independent of BDNF polymorphisms. Exp Neurol 194: 91-96.
36. Desai P, Nebes R, DeKosky ST,Kamboh MI (2005) Investigation of the effect of brain-derived neurotrophic factor (BDNF) polymorphisms on the risk of lateonset Alzheimer's disease ( AD ) and quantitative measures of AD progression. Neurosci Lett 379: 229-234.
37. Bian JT, Zhang JW, Zhang ZX, Zhao HL (2005) Association analysis of brainderived neurotrophic factor (BDNF) gene 196 A/G polymorphism with Alzheimer's disease (AD) in mainland Chinese. Neurosci Lett 387: 11-16.
38. Vepsalainen S, Castren E, Helisalmi S, Iivonen S, Mannermaa A, et al. (2005) Genetic analysis of BDNF and TrkB gene polymorphisms in Alzheimer's disease. J Neurol 252: 423-428.
39. Nishimura M, Kuno S, Kaji R,Kawakami H (2005) Brain-derived neurotrophic factor gene polymorphisms in Japanese patients with sporadic Alzheimer's disease, Parkinson's disease, and multiple system atrophy. Mov Disord 20: 10311033.
40. Bodner SM, Berrettini W, van Deerlin V, Bennett DA, Wilson RS, et al. (2005) Genetic variation in the brain derived neurotrophic factor gene in Alzheimer's disease. Am J Med Genet B Neuropsychiatr Genet 134B: 1-5.
41. Li Y, Rowland C, Tacey K, Catanese J, Sninsky J, et al. (2005) The BDNF Val66Met polymorphism is not associated with late onset Alzheimer's disease in three case-control samples. Mol Psychiatry 10: 809-810.
42. Nacmias B, Piccini C, Bagnoli S, Tedde A, Cellini E, et al. (2004) Brain-derived neurotrophic factor, apolipoprotein E genetic variants and cognitive performance in Alzheimer's disease. Neurosci Lett 367: 379-383.
43. Bagnoli S, Nacmias B, Tedde A, Guarnieri BM, Cellini E, et al. (2004) Brainderived neurotrophic factor genetic variants are not susceptibility factors to Alzheimer's disease in Italy. Ann Neurol 55: 447-448.
44. Combarros O, Infante J, Llorca J, Berciano J (2004) Polymorphism at codon 66 of the brain-derived neurotrophic factor gene is not associated with sporadic Alzheimer's disease. Dement Geriatr Cogn Disord 18: 55-58.
45. Ventriglia M, Bocchio Chiavetto L, Benussi L, Binetti G, Zanetti O, et al. (2002) Association between the BDNF $196 \mathrm{~A} / \mathrm{G}$ polymorphism and sporadic Alzheimer's disease. Mol Psychiatry 7: 136-137.
46. Cousin E, Mace S, Rocher C, Dib C, Muzard G, et al. (2011) No replication of genetic association between candidate polymorphisms and Alzheimer's disease. Neurobiol Aging 32: 1443-1451.
47. Hou L (2009) Association study of BDNF and ER $\alpha$ gene polymorphisms with Late-oneset Alzheimer's disease in southern Chinese Han population [dissertation]. Guangzhou Medical College.
48. Olin D, MacMurray J, Comings DE (2005) Risk of late-onset Alzheimer's disease associated with BDNF C270T polymorphism. Neurosci Lett 381: 275278.
49. Nishimura AL, Oliveira JR, Mitne-Neto M, Guindalini C, Nitrini R, et al. (2004) Lack of association between the brain-derived neurotrophin factor (C-270T)
polymorphism and late-onset Alzheimer's disease (LOAD) in Brazilian patients. J Mol Neurosci 22: 257-260.
50. Riemenschneider M, Schwarz S, Wagenpfeil S, Diehl J, Muller U, et al. (2002) A polymorphism of the brain-derived neurotrophic factor (BDNF) is associated with Alzheimer's disease in patients lacking the Apolipoprotein E epsilon4 allele. Mol Psychiatry 7: 782-785.
51. Kunugi H, Ueki A, Otsuka M, Isse K, Hirasawa H, et al. (2001) A novel polymorphism of the brain-derived neurotrophic factor (BDNF) gene associated with late-onset Alzheimer's disease. Mol Psychiatry 6: 83-86.
52. del Toro D, Canals JM, Gines S, Kojima M, Egea G, et al. (2006) Mutant huntingtin impairs the post-Golgi trafficking of brain-derived neurotrophic factor but not its Val66Met polymorphism. J Neurosci 26: 12748-12757.
53. Egan MF, Kojima M, Callicott JH, Goldberg TE, Kolachana BS, et al. (2003) The BDNF val66met polymorphism affects activity-dependent secretion of BDNF and human memory and hippocampal function. Cell 112: 257-269.
54. Miranda RC, Sohrabji F, Toran-Allerand CD (1993) Neuronal colocalization of mRNAs for neurotrophins and their receptors in the developing central nervous system suggests a potential for autocrine interactions. Proc Natl Acad Sci U S A 90: 6439-6443.
55. Sohrabji F, Miranda RC, Toran-Allerand CD (1995) Identification of a putative estrogen response element in the gene encoding brain-derived neurotrophic factor. Proc Natl Acad Sci U S A 92: 11110-11114.
56. Munafo MR FJ, et al (2004) Meta-analysis of genetic association studies. Trends Genet 20: 439-444.
57. Lau J, Ioannidis JP, Schmid CH (1998) Summing up evidence: one answer is not always enough. Lancet 351: 123-127.


[^0]:    *The cases of these studies are late-onset Alzheimer's disease (LOAD). Abbreviations: Na , not available.
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[^1]:    Abbreviations: FEM, fixed-effects model; REM, random-effects model. EOAD, early-onset Alzheimer's Disease; LOAD, late-onset Alzheimer's Disease.
    rs6265: Codominant model, A vs. G; rs2030324: Codominant model, T vs C.
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