

Hospital mortality after hip fracture surgery in relation to length of stay by care delivery factors

A database study

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Abstract

Two hypotheses were offered for the effect of shorter hospital stays on mortality after hip fracture surgery: worsening the quality of care and shifting death occurrence to postacute settings.

We tested whether the risk of hospital death after hip fracture surgery differed across years when postoperative stays shortened, and whether care factors moderated the association.

Analysis of acute hospital discharge abstracts for subgroups defined by hospital type, bed capacity, surgical volume, and admission time.

153,917 patients 65 years or older surgically treated for first hip fracture.

Risk of hospital death.

We found a decrease in the 30-day risk of hospital death from 7.0% (95%CI: 6.6–7.5) in 2004 to 5.4% (95%CI: 5.0–5.7) in 2012, with an adjusted odds ratio [OR] 0.71 (95%CI: 0.63–0.80). In subgroup analysis, only large community hospitals showed the reduction of ORs by calendar year. No trend was observed in teaching and medium community hospitals. By 2012, the risk of death in large higher volume community hospitals was 34% lower for weekend admissions, OR=0.66 (95%CI: 0.46–0.95) and 39% lower for weekday admissions, OR=0.61 (95%CI: 0.40–0.91), compared to 2004. In large lower volume community hospitals, the 2012 risk was 56% lower for weekend admissions, OR=0.44 (95%CI: 0.26–0.75), compared to 2004.

The risk of hospital death after hip fracture surgery decreased only in large community hospitals, despite universal shortening of hospital stays. This supports the concern of worsening the quality of hip fracture care due to shorter stays.

Abbreviations: CI = confidence interval, CIHI = Canadian Institute of Health Information, OR = odds ratio.

Keywords: admission time, hip fracture, hospital type, mortality, trend, volume

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1. Introduction

Several studies show excess mortality after hip fracture compared with the general population.^[1] One in 10 extra deaths occur during hospitalization for hip fracture surgery.^[2] The risk of hospital death is associated with characteristics of patients, their fracture, and treatment.^[3,4] Where patients undergo surgical treatment for hip fracture may also influence the risk of postoperative hospital death.^[5–11] Factors of care delivery include hospital type and size, surgical volume, prioritization policy, care standards, transfers, care intensity, the availability of operating rooms and intensive care teams during the week, and management of non-medical delays. Recently, some argued that shortening of hospital stays after hip fracture surgery could affect hospital mortality either by worsening the quality of hip fracture care or by shifting the occurrence of death to postacute settings.^[11–13]

In Canada, hospitalization after hip fracture surgery shortened markedly in the 2000s as hospitals started introducing early discharge programs to improve efficiency.^[14] However, whether the risk of hospital death also changed during that period has not been reported. This study sought to test whether the risk of hospital death within 30 days after hip fracture surgery was different between 2004 and 2012 in Canada, and whether the time trend was moderated by factors of care delivery. In particular, we examine the time trends in subgroups of patients defined by a combination of hospital type, bed capacity, surgical

volume, and admission time, that is, by factors that influence access and outcomes of surgery.^[5,9,10,15]

2. Methods

2.1. Study group

We examined discharge abstracts of 153,917 patients 65 years or older who underwent surgery for nonpathological first hip fracture between January 1, 2004 and December 31, 2012 in all Canadian hospitals, except for the province of Quebec. The abstracts were selected from a database maintained by the Canadian Institute for Health Information (CIHI) using procedure codes for hip fracture surgery (CCI: 1VA74^^, 1VA53^^, 1VC74^^, 1SQ53^^, or CCP: 9054, 9114, 9134, 9351, 9359, 9361, 9362, 9363, 9364, 9369). All selected patients stayed at least one day after surgery. These did not include 25 patients with invalid procedure or discharge dates. The University of British Columbia Behavioural Research Ethics Board approved this study (H11-02611).

2.2. Primary outcome

The primary outcome was hospital death. Live discharge was treated as a competing event. Live discharges were identified by CIHI discharge destination codes: discharged home, discharge to home with support services, transferred to a long term care facility, or transferred to another nonacute facility (palliative care/hospice, addiction treatment centre). Deaths after surgery were ascertained within 30 inpatient days, on the premise that longer stays reflected nonacute hospitalization.^[16]

2.3. Subgroups

We studied the time trends in hospital mortality separately in 15 subgroups defined by factors of care delivery, namely, hospital type, bed capacity, annual volume of hip fracture surgery, and weekday and time of admission. Hospital type here serves as a proxy for standards of anesthesia, surgery and intensive care, adequacy of facilities and staffing levels, and attitudes to training, which may affect mortality outcomes.^[15] We used the CIHI classification that aggregates hospitals by type and the total number of beds in four groups: teaching, community-large, community-medium and community-small.^[17] Hospital surgical volume is often linked to quality of care and access to resources.^[6] We dichotomized hospitals into higher and lower volume categories by comparing their annual volumes in the year of index surgery with the median of average annual volumes among hospitals of the same type (174 surgeries for teaching hospitals, 141 for community large hospitals, 37 for community medium hospitals).^[18,19] Admission time was previously linked to access to resources with fewer support services from late Friday to Monday morning.^[9] In our analysis, weekday admissions between 8 am and 5 pm were classified as during working hours, and between 5 pm to 8 am as after hours, and weekend admissions were between 5 pm Friday and 8 am Monday.

2.4. Statistical analysis

We reported the distribution of patients by variable levels in each calendar year and compared the distributions across years using the chi-square test. Daily death rates were estimated by dividing the number of deaths found in the discharge abstracts by the total

number of inpatient days during the 30-day follow-up, overall and by calendar year.

We estimated the cumulative incidence function of death for each of 30 inpatient days after day of surgery, accounting for discharge rates, on the premise that patients remain at risk of hospital death only while they remain in hospital.^[20] We treated postoperative stays ended by transfers to another acute care facility, or by discharges occurring on the same day of surgery, as well as stays exceeding 30 days after surgery as the right-censored observations of time duration.^[21] We compared the cumulative incidence functions between years using Pepe-Mori 2-sample test.^[22]

We used proportional odds models and the pseudovalues method to test whether the cumulative incidences of death were different between each year and 2004.^[23] In the overall analysis, differences between the years were adjusted for age, sex, preadmission residence, diagnostic code for heart failure (ICD-10-CA I50, J81), chronic obstructive pulmonary disease (ICD-10-CA J41, J42, J43, J44, J47), ischemic heart disease (ICD-10-CA I20, I21, I22, I24, I25), cardiac dysrhythmias (ICD-10-CA I47, I48, I49), hypertension (ICD-10-CA I10.0, I10.1, I11), and diabetes (ICD-10-CA E10.0 –E10.7, E11.0–11.7, E13.0–13.7, E14.0–14.7) from all hospitalizations in 1 year prior to index admission, and fracture type, procedure type, weekday and time of admission, hospital type and size, hospital volume of hip fracture surgeries, demand at index admission, time to surgery, and province of residence.^[3–5,9,10,15] In the subgroup analysis, the differences were adjusted for age, sex, preadmission residence, comorbidity, fracture type at surgery, procedure type, demand at index admission, and time to surgery. Within each subgroup, we performed the nonparametric rank-order test for trend in adjusted ORs ordered by a calendar year.^[24] The competing-risk analysis was conducted with R packages *cmprsk*,^[25] *prodlm*,^[26] and *geepack*.^[27]

3. Results

3.1. Patient characteristics

In total, 153,917 hip fracture patients underwent surgery between 2004 and 2012 (Table 1). The majority were women (73.4%), and almost half were 85 years or older (45.6%). In half of the patients, fracture type was transcervical (52.0%). Major comorbidity was reported for 27.0%, with cardiac dysrhythmias being the most prevalent (9.4%).

The proportion of patients who underwent surgery in large community hospitals (44.7%) was higher than in teaching (38.5%) or medium community (14.9%) hospitals (Table 1). Less patients underwent surgery at small community and unassigned type (1.8%) hospitals. More patients underwent surgery at higher volume teaching (76.7%) and community large hospitals (68.4%) when compared to lower volume hospitals of each type. More patients were admitted after working hours (39.1%) or on weekends (37.5%) than during working hours (23.4%) at higher volume teaching hospitals. Similar distributions were seen for higher volume community large hospitals: after hours (36.0%), weekends (36.8%), and working hours (27.2%), and for lower volume community large hospitals: after hours (34.0%), weekends (36.1%), and working hours (29.8%). For lower volume teaching hospitals, more patients were admitted after working hours (37.9%) or on weekends (37.9%) than during working hours (24.2%). Similar proportions of patients were admitted during working hours, after hours and on weekends at community medium hospitals.

Table 1
Characteristics of 153,917 patients surgically treated for first hip fracture, by calendar year.

Characteristics	All years n = 153917	2004 n = 17222	2005 n = 17371	2006 n = 16808	2007 n = 16942	2008 n = 16826	2009 n = 16989	2010 n = 16955	2011 n = 17036	2012 n = 17768
Age, years										
65–74	23273 (15.1)	2553 (14.8)	2639 (15.2)	2466 (14.7)	2525 (14.9)	2460 (14.6)	2568 (15.1)	2507 (14.8)	2646 (15.5)	2909 (16.4)
75–84	60500 (39.3)	7343 (42.6)	7174 (41.3)	6821 (40.6)	6811 (40.2)	6612 (39.3)	6499 (38.3)	6392 (37.7)	6329 (37.2)	6519 (36.7)
85–94	62079 (40.3)	6512 (37.8)	6724 (38.7)	6713 (39.9)	6766 (39.9)	6852 (40.7)	6978 (41.1)	7127 (42.0)	7048 (41.4)	7359 (41.4)
≥95	8065 (5.2)	814 (4.7)	834 (4.8)	808 (4.8)	840 (5.0)	902 (5.4)	944 (5.6)	929 (5.5)	1013 (5.9)	981 (5.5)
Sex*										
Women	112965 (73.4)	12947 (75.2)	12869 (74.1)	12420 (73.9)	12546 (74.1)	12251 (72.8)	12397 (73.0)	12383 (73.0)	12332 (72.4)	12820 (72.2)
Men	40934 (26.6)	4274 (24.8)	4494 (25.9)	4388 (26.1)	4394 (25.9)	4572 (27.2)	4589 (27.0)	4571 (27.0)	4704 (27.6)	4948 (27.8)
Fracture type										
Transcervical	80027 (52.0)	9068 (52.7)	8982 (51.7)	8711 (51.8)	8722 (51.5)	8666 (51.5)	8889 (52.3)	8881 (52.4)	8777 (51.5)	9331 (52.5)
Pertrochanteric	66378 (43.1)	7455 (43.3)	7652 (44.1)	7295 (43.4)	7361 (43.4)	7323 (43.5)	7281 (42.9)	7176 (42.3)	7355 (43.2)	7480 (42.1)
Subtrochanteric	7512 (4.9)	699 (4.1)	737 (4.2)	802 (4.8)	859 (5.1)	837 (5.0)	819 (4.8)	898 (5.3)	904 (5.3)	957 (5.4)
Comorbidity†										
Heart failure	12088 (7.9)	1356 (7.9)	1322 (7.6)	1250 (7.4)	1216 (7.2)	1192 (7.1)	1416 (8.3)	1420 (8.4)	1448 (8.5)	1468 (8.3)
COPD	8026 (5.2)	903 (5.2)	943 (5.4)	923 (5.5)	827 (4.9)	898 (5.3)	848 (5.0)	876 (5.2)	922 (5.4)	886 (5.0)
IHD (acute)	9009 (5.9)	1008 (5.9)	1032 (5.9)	996 (5.9)	1051 (6.2)	1011 (6.0)	926 (5.5)	961 (5.7)	1050 (6.2)	974 (5.5)
Cardiac dysrhythmias	14425 (9.4)	1375 (8.0)	1409 (8.1)	1352 (8.0)	1306 (7.7)	1464 (8.7)	1702 (10.0)	1767 (10.4)	1960 (11.5)	2090 (11.8)
IHD (chronic)	2534 (1.6)	331 (1.9)	323 (1.9)	289 (1.7)	255 (1.5)	251 (1.5)	283 (1.7)	249 (1.5)	298 (1.7)	255 (1.4)
Hypertension	9105 (5.9)	1044 (6.1)	1058 (6.1)	971 (5.8)	837 (4.9)	927 (5.5)	982 (5.8)	1034 (6.1)	1092 (6.4)	1160 (6.5)
Diabetes	7029 (4.6)	313 (1.8)	379 (2.2)	803 (4.8)	1124 (6.6)	1240 (7.4)	938 (5.5)	740 (4.4)	734 (4.3)	758 (4.3)
Preadmission residence										
Home	91933 (59.7)	10540 (61.2)	10639 (61.2)	10209 (60.7)	10144 (59.9)	10063 (59.8)	9959 (58.6)	9751 (57.5)	10074 (59.1)	10554 (59.4)
Other	61984 (40.3)	6682 (38.8)	6732 (38.8)	6599 (39.3)	6798 (40.1)	6763 (40.2)	7030 (41.4)	7204 (42.5)	6962 (40.9)	7214 (40.6)
Admission time‡										
Working hours	42031 (27.3)	4832 (28.1)	4736 (27.3)	4628 (27.5)	4682 (27.6)	4574 (27.2)	4674 (27.5)	4652 (27.4)	4609 (27.1)	4644 (26.1)
After hours	55404 (36.0)	6057 (35.2)	6321 (36.4)	5926 (35.3)	6118 (36.1)	6089 (36.2)	6078 (35.8)	6109 (36.0)	6210 (36.5)	6496 (36.6)
Weekend	56414 (36.7)	6333 (36.8)	6314 (36.3)	6254 (37.2)	6141 (36.2)	6163 (36.6)	6236 (36.7)	6193 (36.5)	6176 (36.3)	6604 (37.2)
Transfer history										
No	140584 (91.3)	15809 (91.8)	15802 (91.0)	15266 (90.8)	15386 (90.8)	15354 (91.3)	15486 (91.2)	15451 (91.1)	15533 (91.2)	16497 (92.8)
Yes	13333 (8.7)	1413 (8.2)	1569 (9.0)	1542 (9.2)	1556 (9.2)	1472 (8.7)	1503 (8.8)	1504 (8.9)	1503 (8.8)	1271 (7.2)
Procedure type										
Fixation	92226 (59.9)	10438 (60.6)	10732 (61.8)	10140 (60.3)	10257 (60.5)	10067 (59.8)	10069 (59.3)	10108 (59.6)	10094 (59.3)	10321 (58.1)
Other	61691 (40.1)	6784 (39.4)	6639 (38.2)	6668 (39.7)	6685 (39.5)	6759 (40.2)	6920 (40.7)	6847 (40.4)	6942 (40.7)	7447 (41.9)
Timing of surgery										
First or next day	98393 (63.9)	11501 (66.8)	11325 (65.2)	10595 (63.0)	10529 (62.1)	10450 (62.1)	10657 (62.7)	10823 (63.8)	10811 (63.5)	11702 (65.9)
2 or more days	55524 (36.1)	5721 (33.2)	6046 (34.8)	6213 (37.0)	6413 (37.9)	6376 (37.9)	6332 (37.3)	6132 (36.2)	6225 (36.5)	6066 (34.1)
Hospital type										
Teaching	59301 (38.5)	6729 (39.1)	6739 (38.8)	6505 (38.7)	6574 (38.8)	6465 (38.4)	6595 (38.8)	6460 (38.1)	6535 (38.4)	6699 (37.7)
Community-Large	68871 (44.7)	7361 (42.7)	7496 (43.2)	7214 (42.9)	7338 (43.3)	7631 (45.4)	7776 (45.8)	7808 (46.1)	7895 (46.3)	8352 (47.0)
Community-Medium	22966 (14.9)	2492 (14.5)	2572 (14.8)	2551 (15.2)	2641 (15.6)	2500 (14.9)	2472 (14.6)	2546 (15.0)	2527 (14.8)	2665 (15.0)
Community-Small§	2779 (1.8)	640 (3.7)	564 (3.2)	538 (3.2)	389 (2.3)	230 (1.4)	146 (0.9)	141 (0.8)	79 (0.5)	52 (0.3)
Hospital volume										
Lower in its type	38977 (25.3)	5223 (30.3)	5035 (29.0)	4311 (25.6)	3939 (23.2)	3956 (23.5)	4300 (25.3)	4032 (23.8)	4760 (27.9)	3421 (19.3)
Province¶										
British Columbia	29158 (18.9)	3169 (18.4)	3243 (18.7)	3320 (19.8)	3152 (18.6)	3256 (19.4)	3228 (19.0)	3277 (19.3)	3161 (18.6)	3352 (18.9)
Alberta	16630 (10.8)	1848 (10.7)	1865 (10.7)	1792 (10.7)	1813 (10.7)	1800 (10.7)	1867 (11.0)	1856 (10.9)	1839 (10.8)	1950 (11.0)
Saskatchewan	8030 (5.2)	964 (5.6)	941 (5.4)	861 (5.1)	890 (5.3)	856 (5.1)	883 (5.2)	831 (4.9)	899 (5.3)	905 (5.1)
Manitoba	8644 (5.6)	880 (5.1)	999 (5.8)	963 (5.7)	928 (5.5)	969 (5.8)	996 (5.9)	943 (5.6)	980 (5.8)	986 (5.5)
Ontario	74732 (48.6)	8470 (49.2)	8417 (48.5)	8115 (48.3)	8258 (48.7)	8147 (48.4)	8094 (47.6)	8180 (48.2)	8338 (48.9)	8713 (49.0)
New Brunswick	5358 (3.5)	616 (3.6)	625 (3.6)	603 (3.6)	629 (3.7)	564 (3.4)	575 (3.4)	592 (3.5)	572 (3.4)	582 (3.3)
Nova Scotia	6534 (4.2)	728 (4.2)	743 (4.3)	670 (4.0)	714 (4.2)	706 (4.2)	782 (4.6)	727 (4.3)	722 (4.2)	742 (4.2)
Prince Edward Island	1101 (0.7)	114 (0.7)	132 (0.8)	111 (0.7)	142 (0.8)	117 (0.7)	119 (0.7)	120 (0.7)	128 (0.8)	118 (0.7)
Newfoundland	3510 (2.3)	412 (2.4)	383 (2.2)	350 (2.1)	391 (2.3)	383 (2.3)	419 (2.5)	406 (2.4)	367 (2.2)	399 (2.2)

COPD=chronic obstructive pulmonary disease, IHD=ischemic heart disease.

* 18 patients with unknown sex.

† By diagnostic codes from all hospitalisations in one year prior to index admission.

‡ 68 patients with unknown admission time.

§ Includes unassigned.

|| Below 174 surgeries for teaching, 141 surgeries for community large, 37 surgeries community medium.

¶ Not shown 220 patients from Territories.

In all subgroups, similar proportions of patients were over the age of 85, men, with at least 1 major comorbidity, with transcervical fracture type, and underwent fixation (see Table, Supplemental Digital Content 1–3, <http://links.lww.com/MD/B663>). Higher proportion of patients were admitted from home for lower volume teaching hospitals (65.2%) and community large hospitals (63.0%) than for higher volume teaching hospitals (54.2%) or community medium hospitals (54.6%) (see Table, Supplemental Digital Content 1–3, <http://links.lww.com/MD/B663>). Higher proportion of patients waited for surgery 2 or more days after admission for teaching hospitals (44.4%) than for community hospitals (30.8%) (see Table, Supplemental Digital Content 1–3, <http://links.lww.com/MD/B663>). The proportion of patients for higher volume teaching hospitals, higher volume community large hospitals, and community medium hospitals were distributed evenly over the years (not shown in Tables). The proportion of patients for lower volume teaching hospitals decreased from 14.9% of patients in 2004 to 6.8% in 2012. The proportion of patients for lower volume community large hospitals varied across calendar years with 12.3% in 2004, 9.0% in 2007, and 13.5% in 2011.

3.2. Death by calendar year

In this synthetic cohort, 8,032 (5.2%) patients died and 97,259 (63.2%) were discharged within 30 inpatient days of surgery. The average follow-up time was 13 days with 48,626 (31.6%) patients followed until censoring events or 30 days after surgery. Overall, the average death rate was 4.0 (95% confidence interval [CI] 3.9 to 4.1) per 1000 patient-days, with the highest rate of 4.3 (95% CI 4.0 to 4.6) per 1000 patient-days in 2004 and 2005 and the lowest rate of 3.7 (95% CI 3.5 to 4.0) per 1000 patient-days in 2012 (Table 2). When compared to 2004, the cumulative incidence functions of death were not different in 2005 ($P = .82$) or 2006 ($P = .34$), but were different in 2007 and 2008 ($P < .05$), and in 2009, 2010, 2011, and 2012 ($P < .01$). At 30 days after surgery, the cumulative incidence of death was highest among patients treated in 2004 at 7.0% (95% CI 6.6 to 7.5) and lowest in 2012 at 5.4% (95% CI 5.0 to 5.7) (Table 2).

Compared to 2004, the adjusted odds of death were 17% lower (odds ratio [OR] = 0.83, 95% CI 0.74 to 0.93), 22% lower (OR = 0.78, 95% CI 0.70 to 0.88), and 29% lower (OR = 0.71, 95% CI 0.63 to 0.80) in 2006, 2009, and 2012 respectively (Fig. 1). The trend test indicates a consistent reduction of the adjusted ORs ordered by calendar year ($z = -3.9, P < .001$). This

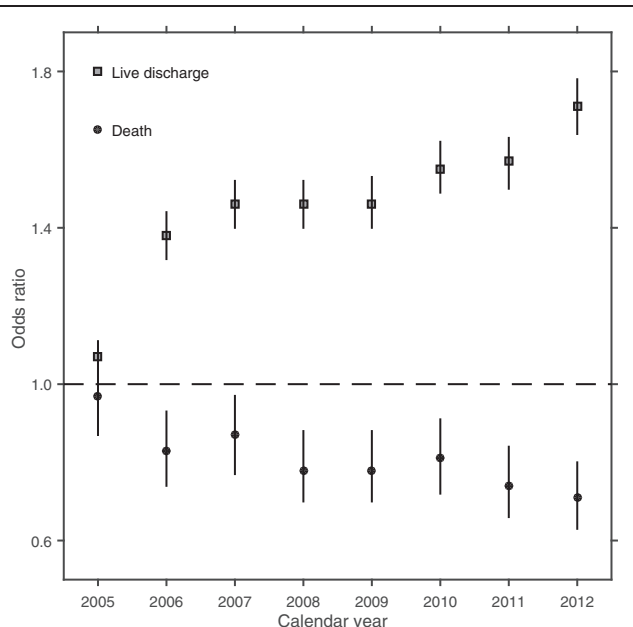


Figure 1. Risk of death and live discharge by calendar year, relative to 2004 (dashed line). Year-specific odds ratios and 95% confidence intervals were adjusted for age (85+ vs <85), sex, preadmission residence (residence vs home), comorbidity (at least 1 acute admission for heart failure, chronic obstructive pulmonary disease, ischemic heart disease, hypertension, or diabetes in the year prior to admission for hip fracture), fracture type (transcervical vs other), procedure type (fixation vs implant), weekday and time of admission, hospital type and size (teaching, large, medium, or small community), and hospital volume of hip fracture surgery (higher vs lower), demand at index admission (number of admissions in the week of initial hospitalization, larger vs smaller than hospital weekly capacity), time to surgery (2 days or more days vs less), and province of admission. Trend test for odds ratios of death $z = -3.9, P < .001$. Trend test for odds ratios of discharge $z = 5.5, P < .001$.

was in contrast to increasing trends for live discharges ($z = 5.5, P < .001$).

3.3. Death by calendar year in subgroups

Table 3 summarizes the rates and risks for death and for discharge in the subgroups. The lowest death rate was 3.3 (95% CI 2.8 to 3.7) per 1000 patient-days in teaching hospitals with lower volume of hip fracture surgeries for weekend admissions

Table 2
Hospital deaths and live discharges within 30 days after surgery by calendar year.

Year	No of patients	Follow-up time, days	No of deaths	Death rate (95% CI)*	% Died (95% CI)†	Discharge rate (95% CI)*	% discharged alive (95% CI)†
2004	17222	217121	934	4.3 (4.0–4.6)	7.0 (6.6–7.5)	39.5 (38.7–40.3)	67.7 (66.8–68.6)
2005	17371	219649	940	4.3 (4.0–4.6)	6.8 (6.4–7.3)	41.7 (40.9–42.6)	69.7 (68.9–70.6)
2006	16808	210902	889	4.2 (3.9–4.5)	6.2 (5.8–6.6)	48.9 (48.0–49.8)	74.5 (73.7–75.2)
2007	16942	215886	891	4.1 (3.9–4.4)	6.1 (5.7–6.4)	50.3 (49.3–51.2)	75.1 (74.4–75.9)
2008	16826	223673	897	4.0 (3.7–4.3)	5.9 (5.5–6.3)	50.0 (49.0–50.9)	74.8 (74.1–75.5)
2009	16989	228265	863	3.8 (3.5–4.0)	5.6 (5.2–6.0)	50.0 (49.1–50.9)	75.3 (74.6–76.0)
2010	16955	224666	890	4.0 (3.7–4.2)	5.8 (5.4–6.1)	51.4 (50.4–52.3)	75.6 (74.9–76.3)
2011	17036	226395	858	3.8 (3.5–4.0)	5.5 (5.1–5.9)	51.9 (51.0–52.9)	76.3 (75.6–76.9)
2012	17768	232398	870	3.7 (3.5–4.0)	5.4 (5.0–5.7)	53.6 (52.7–54.6)	77.0 (76.3–77.6)

CI = confidence interval.
* Per 1000 patient-days.
† At 30 days.

Table 3**Hospital deaths and live discharges in subgroups defined by hospital type, bed capacity, annual volume of hip fracture surgery, and time of admission***

Time of admission	Discharge rate, per 1000 patient-days	% Discharge alive at 30days	Death rate, per 1000 patient-days	% Deaths at 30 days
Teaching > 174 surgeries				
Working hours	39.8 (38.8–40.8)	68.8 (67.8–69.8)	3.6 (3.3–3.9)	6.1 (5.6–6.6)
After hours	38.1 (37.4–38.9)	67.1 (66.3–67.9)	3.5 (3.3–3.7)	6.0 (5.6–6.4)
Weekend	38.2 (37.4–39.0)	67.2 (66.4–68.0)	3.5 (3.3–3.7)	5.9 (5.6–6.3)
Teaching ≤ 174 surgeries				
Working hours	48.7 (46.7–50.7)	75.0 (73.4–76.6)	3.5 (3.0–4.1)	5.3 (4.5–6.1)
After hours	46.5 (44.9–48.0)	74.1 (72.8–75.3)	3.5 (3.1–3.9)	5.5 (4.8–6.1)
Weekend	47.8 (46.2–49.4)	74.4 (73.1–75.7)	3.3 (2.8–3.7)	5.0 (4.4–5.6)
Community large > 141 surgeries				
Working hours	56.4 (55.2–57.6)	78.1 (77.3–78.9)	4.5 (4.1–4.8)	6.1 (5.7–6.5)
After hours	57.7 (56.6–58.7)	79.2 (78.5–79.9)	4.1 (3.8–4.4)	5.6 (5.2–6.0)
Weekend	57.9 (56.9–59.0)	78.9 (78.3–79.6)	4.4 (4.1–4.7)	6.0 (5.6–6.4)
Community large ≤ 141 surgeries				
Working hours	49.1 (47.6–50.6)	73.8 (72.7–75.0)	4.3 (3.8–4.7)	6.3 (5.7–6.9)
After hours	49.6 (48.2–51.0)	75.2 (74.2–76.3)	4.1 (3.7–4.5)	6.1 (5.5–6.7)
Weekend	49.3 (48.0–50.7)	74.3 (73.2–75.3)	4.3 (3.9–4.7)	6.3 (5.7–6.8)
Community medium				
Working hours	56.2 (54.6–57.8)	77.3 (76.2–78.4)	5.0 (4.5–5.4)	6.6 (6.0–7.2)
After hours	56.1 (54.4–57.7)	78.4 (77.3–79.5)	4.5 (4.0–4.9)	6.1 (5.5–6.7)
Weekend	55.5 (53.9–57.1)	76.8 (75.7–77.9)	4.9 (4.4–5.4)	6.7 (6.0–7.3)

* Not included 1634 patients with unassigned hospital type, 1145 patients with community small hospital type, 657 patients treated in hospitals with annual volume below 12, and 68 patients with unknown admission time.

and the highest death rate was 5.0 (95% CI 4.5 to 5.4) per 1,000 patient-days in medium size community hospitals for working-hours admission. The 30-day cumulative incidence of death was lowest for weekend admissions to low-volume teaching hospitals 5.0% (95% CI 4.4 to 5.6) and highest for weekend admissions to medium community hospitals 6.7% (95% CI 6.0 to 7.3). The lowest discharge rate and 30-day cumulative incidence of discharge were observed in teaching higher volume hospitals for after-hours admissions at 38.1 (95% CI 37.4 to 38.9) per 1000 patient-days and 67.1% (95% CI 66.3% to 67.9%), respectively. The highest discharge rate was 57.9 (95% CI 56.9 to 59.0) per 1,000 patient-days in community large higher volume hospitals for weekend admissions. The highest 30-day cumulative incidence of discharge was in community large higher volume hospitals for after-hours admissions at 79.2% (95% CI 78.5% to 79.9%).

Figure 2 shows the year-specific adjusted ORs for death in each subgroup. For working-hours admissions (Fig. 2, row 1), the odds of death were 0.68 (95% CI 0.43 to 1.09) at higher volume teaching hospitals, 0.61 (95% CI 0.40 to 0.91) at higher volume large community hospitals, 0.99 (95% CI 0.40 to 2.42) at lower volume teaching hospitals, 0.75 (95% CI 0.42 to 1.35) at lower volume large community hospitals, and 0.53 (95% CI 0.32 to 0.85) at medium community hospitals in 2012 compared to 2004. For after-hours admissions (Fig. 2, row 2), the odds of death were 0.64 (95% CI 0.45 to 0.91) at higher volume teaching hospitals, 0.56 (95% CI 0.37 to 0.83) at higher volume large community hospitals, 0.90 (95% CI 0.40 to 2.06) at lower volume teaching hospitals, 0.80 (95% CI 0.46 to 1.40) at lower volume large community hospitals, and 1.10 (95% CI 0.65 to 1.87) at medium community hospitals in 2012 compared to 2004. For weekend admissions (Fig. 2, rows 3), the odds of death were 1.02 (95% CI 0.73 to 1.43) at higher volume teaching hospitals, 0.66 (95% CI 0.46 to 0.95) at higher volume large community hospitals, 0.61 (95% CI 0.25 to 1.48) at lower

volume teaching hospitals, 0.44 (95% CI 0.26 to 0.75) at lower volume large community hospitals, and 0.67 (95% CI 0.40 to 1.12) at medium community hospitals in 2012 compared to 2004.

The trend tests indicate consistent reduction in the adjusted ORs ordered by calendar year in large community hospitals with lower volumes for working-hours admissions ($z = -2.1$, $P < .05$), and with higher volumes for weekend admissions ($z = -2.0$, $P < .05$) and for after-hours admissions ($z = -2.5$, $P < .01$). No trend for the adjusted ORs ordered by calendar year was seen in other subgroups.

4. Discussion

In this study, we tested whether hospital mortality changed after shortening in postoperative stays among patients undergoing hip fracture surgery. We estimated the risk of hospital death for each calendar year between 2004 and 2012, when postoperative stays markedly shortened in Canada. Using the cumulative incidence function for death to account for the rate of live discharge, we found a gradual decrease in the 30-day risk of death from 7.0% (95% CI 6.6 to 7.5) in 2004 to 5.4% (95% CI 5.0 to 5.7) in 2012. After adjustment for characteristics of patients, their fracture, treatment, and care delivery, the risk of death was 29% lower in 2012 than in 2004 (OR=0.71, 95% CI 0.63 to 0.80), with a significant trend in reduction of the adjusted ORs ordered by calendar year ($P < .001$).

We further found that where patients undergo surgery moderated the association between calendar year and hospital mortality after hip fracture surgery during the study period. In particular, our results demonstrate the time trend differed in subgroups of patients defined by a combination of hospital type, bed capacity, annual volume of hip fracture surgeries, and weekday and time of admission. Only large community hospitals showed reduction of the adjusted ORs by calendar year.

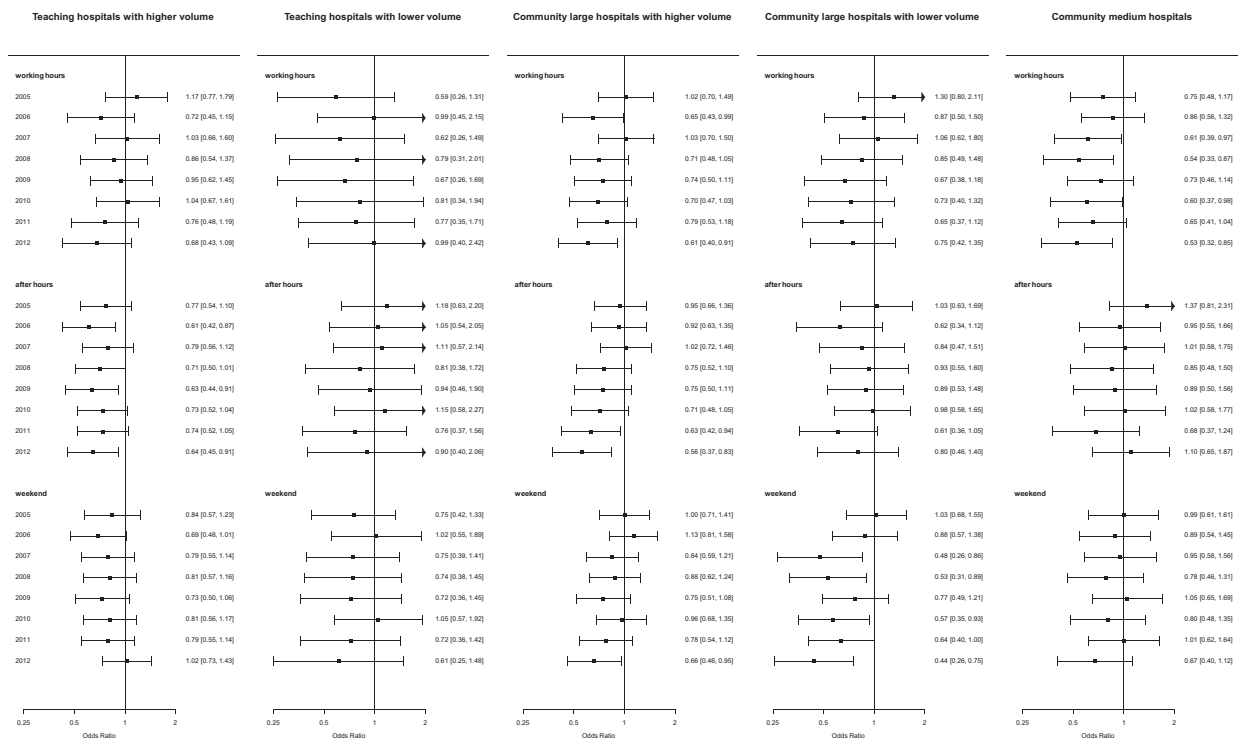


Figure 2. Risk of death after surgery by calendar year as compared to 2004 within subgroups defined by hospital type, bed capacity, annual volume of hip fracture surgery, and weekday and time of admission. Differences between years were adjusted for age (85+ vs < 85), sex, preadmission residence (residence vs home), comorbidity (at least 1 acute admission for heart failure, chronic obstructive pulmonary disease, ischemic heart disease, hypertension, or diabetes in the year prior to admission for hip fracture), fracture type (transcervical vs other), procedure type (fixation vs implant), demand at index admission (number of admissions in the week of initial hospitalization, larger vs smaller than hospital weekly capacity), and time to surgery (2 days or more days vs less).

Compared to 2004, the 2012 risk of death was 34% lower for weekend admissions to higher volume hospitals (OR=0.66, 95% CI 0.46 to 0.95), 39% lower for weekday admissions to higher volume hospitals (OR=0.61, 95% CI 0.40 to 0.91), and 56% lower for weekend admissions to lower volume hospitals (OR=0.44, 95% CI 0.26 to 0.75). No trend for the adjusted ORs ordered by calendar year was present in teaching hospitals or medium community hospitals.

Some suggested that early discharge programs shorten the exposure to the risk of death in hospital after hip fracture surgery^[28] and shift the occurrence of death to postacute settings.^[11] Shortening of postoperative stays also caused a concern of worsening the quality of hip fracture care.^[12] Following Andersen et al,^[29] we may expect fewer deaths in the intervention group than in the control group when an intervention increases the discharge rate even with no change in the death rate, simply because the number of event-free patients will decrease more quickly in the intervention group. Similarly, Wobblers et al reasoned that a decrease in the death rate even with no change in the rate of discharge in the intervention group leaves more patients exposed to the competing risk of discharge, resulting in more discharges in the intervention group.^[30] Thus, in the presence of competing risk of discharge, a policy question (“Does the stay shortening lower the risk of death?”) and an etiological question (“Does the stay shortening lower the death rate among patients still hospitalized?”) may have differing answers. We demonstrate that a time-trend in increasing the probability of discharge was not followed by a reduction in the risk of death in many care

settings. This points to a new research agenda aimed at understanding the effects of shortening of postoperative hospital stays on the standards of hip fracture care.

This is the first study to examine the occurrence of hospital death across calendar years within subgroups defined by the factors of care delivery, but there are some limitations. In particular, hospital type was based on the peer group assignments available only after 2010 and therefore there was a possibility for misclassification.^[31] Due to the observational nature of the study, there was a limited number of variables for adjustment. In particular, the presence of renal disease, prior cerebrovascular accident, dementia, or Parkinson’s disease may influence the occurrence of hospital death after hip fracture across calendar years. Within each subgroup we were able to control for only age, sex, preadmission residence, comorbidity, fracture subtype, type of surgical procedure, and time to surgery. We examined data for all Canadian hospitals outside the province of Quebec. Sirois et al^[32] reported an annual decrease in hospital death after hip fracture surgery of 4% over 20 years in a level 1 trauma center in Quebec, which is consistent with our results.

5. Conclusions

Despite universal shortening hospital stays in Canadian hospitals between 2004 and 2012, the risk of hospital death after hip fracture surgery decreased only in large community hospitals, but not in teaching and medium-size community hospitals. This supports the concern of worsening the quality of hip fracture care resulting from shorter stays in some care settings.

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References

- [1] Panula J, Pihlajamaki H, Mattila VM, et al. Mortality and cause of death in hip fracture patients aged 65 or older: a population-based study. *BMC Musculoskelet Disord* 2011;12:105.
- [2] Uzoigwe CE, Burnand HG, Cheesman CL, et al. Early and ultra-early surgery in hip fracture patients improves survival. *Injury* 2013;44:726–9.
- [3] Penrod JD, Litke A, Hawkes WG, et al. The association of race, gender, and comorbidity with mortality and function after hip fracture. *J Gerontol A Biol Sci Med Sci* 2008;63:867–72.
- [4] Vestergaard P, Rejnmark L, Mosekilde L. Loss of life years after a hip fracture. *Acta Orthop* 2009;80:525–30.
- [5] Forte ML, Virnig BA, Swiontkowski MF, et al. Ninety-day mortality after intertrochanteric hip fracture: does provider volume matter? *J Bone Joint Surg Am* 2010;92:799–806.
- [6] Kristensen PK, Thillemann TM, Johnsen SP. Is bigger always better? A nationwide study of hip fracture unit volume, 30-day mortality, quality of in-hospital care, and length of hospital stay. *Med Care* 2014;52:1023–9.
- [7] Metcalfe D, Olufajo OA, Zogg CK, et al. Are older adults with hip fractures disadvantaged in level 1 trauma centers? *Med Care* 2016;54:e16–22.
- [8] Wiggers JK, Guitton TG, Smith RM, et al. Observed and expected outcomes in transfer and nontransfer patients with a hip fracture. *J Orthop Trauma* 2011;25:666–9.
- [9] Freemantle N, Ray D, McNulty D, et al. Increased mortality associated with weekend hospital admission: a case for expanded seven day services? *BMJ* 2015;351:h4596.
- [10] Khan SK, Jameson SS, Avery PJ, et al. Does the timing of presentation of neck of femur fractures affect the outcome of surgical intervention. *Eur J Emerg Med* 2013;20:178–81.
- [11] Nordstrom P, Gustafson Y, Michaelsson K, et al. Length of hospital stay after hip fracture and short term risk of death after discharge: a total cohort study in Sweden. *BMJ* 2015;350:h696.
- [12] Cram P, Rush RP. Length of hospital stay after hip fracture: how low can we go before patients are at risk? *BMJ* 2015;350:h823.
- [13] FitzGerald JD, Boscardin WJ, Hahn BH, et al. Impact of the Medicare short stay transfer policy on patients undergoing major orthopedic surgery. *Health Serv Res* 2007;42(1 pt 1):25–44.
- [14] Sobolev B, Guy P, Sheehan KJ, et al. Time trends in hospital stay after hip fracture in Canada, 2004–2012: database study. *Arch Osteoporos* 2016;11:13.
- [15] Weller I, Wai EK, Jaglal S, et al. The effect of hospital type and surgical delay on mortality after surgery for hip fracture. *J Bone Joint Surg Br* 2005;87:361–6.
- [16] Kaboli PJ, Go JT, Hockenberry J, et al. Associations between reduced hospital length of stay and 30-day readmission rate and mortality: 14-year experience in 129 Veterans Affairs hospitals. *Ann Intern Med* 2012;157:837–45.
- [17] Canadian Institute for Health Information (CIHI) Peer Groups in the Electronic Discharge Abstract Database Reports, 2015. Ottawa: CIHI; 2015.
- [18] Urbach DR, Baxter NN. Does it matter what a hospital is “high volume” for? Specificity of hospital volume-outcome associations for surgical procedures: analysis of administrative data. *Qual Saf Health Care* 2004;13:379–83.
- [19] Ghaferi AA, Birkmeyer JD, Dimick JB. Hospital volume and failure to rescue with high-risk surgery. *Med Care* 2011;49:1076–81.
- [20] Pintilie M. *Competing Risks, A Practical Perspective*. New York: John Wiley & Sons; 2006;53–70.
- [21] Austin PC, Lee DS, Fine JP. Introduction to the analysis of survival data in the presence of competing risks. *Circulation* 2016;133:601–9.
- [22] Pepe MS, Mori M. Kaplan–Meier, marginal or conditional probability curves in summarizing competing risks failure time data? *Stat Med* 1993;12:737–51.
- [23] Klein JP, Andersen PK. Regression modeling of competing risks data based on pseudovalues of the cumulative incidence function. *Biometrics* 2005;61:223–9.
- [24] Cuzick J. A Wilcoxon-type test for trend. *Stat Med* 1985;4:87–90.
- [25] Gray B. *cmprsk: Subdistribution analysis of competing risks*. 2014. Available at: <http://CRAN.R-project.org/package=cmprsk> Accessed June 19, 2016.
- [26] Gerds T. *prodlm: Product-limit estimation for censored event history analysis*. 2014. Available at: <http://CRAN.R-project.org/package=prodlm> Accessed June 19, 2016.
- [27] Hojsgaard S, Halekoh U, Yan J. The R package *geepack* for generalized estimating equations. *J Stat Software* 2006;15:1–1.
- [28] Holvik K, Ranhoff AH, Martinsen MI, et al. Predictors of mortality in older hip fracture inpatients admitted to an orthogeriatric unit in oslo, norway. *J Aging Health* 2010;22:1114–31.
- [29] Andersen PK, Geskus RB, de WT, et al. Competing risks in epidemiology: possibilities and pitfalls. *Int J Epidemiol* 2012;41:861–70.
- [30] Wolbers M, Koller MT, Stel VS, et al. Competing risks analyses: objectives and approaches. *Eur Heart J* 2014;35:2936–41.
- [31] Canadian Institute for Health Information (CIHI) CIHI Portal Release Noted: Release 9.14.2015. Ottawa: CIHI; 2015.
- [32] Sirois MJ, Cote M, Pelet S. The burden of hospitalized hip fractures: patterns of admissions in a level I trauma center over 20 years. *J Trauma* 2009;66:1402–10.