# The Impact of 51 Risk Factors on Life Expectancy in Canada: Findings from a New Risk Prediction Model Based on Data from the Global Burden of Disease Study 

Jacek A. Kopec ${ }^{1,2, *}$, Eric C. Sayre ${ }^{2}$, Benajir Shams ${ }^{3}$, Linda C. Li ${ }^{2,4}$, Hui Xie ${ }^{2,5}$, Lynne M. Feehan ${ }^{4}$ and John M. Esdaile ${ }^{2,6}$<br>1 School of Population and Public Health, University of British Columbia, Vancouver, BC V6T 1Z3, Canada<br>2 Arthritis Research Canada, Vancouver, BC V5Y 3P2, Canada; esayre@arthritisresearch.ca (E.C.S.); lli@arthritisresearch.ca (L.C.L.); hxie@arthritisresearch.ca (H.X.); jesdaile@arthritisresearch.ca (J.M.E.)<br>3 Fraser Health Authority, Surrey, BC V3T 0H1, Canada; benajirshams86@gmail.com<br>4 Department of Physical Therapy, University of British Columbia, Vancouver, BC V6T 1Z3, Canada; lynnefeehan@gmail.com<br>5 Faculty of Health Sciences, Simon Fraser University, Burnaby, BC V5A 1S6, Canada<br>6 Department of Medicine, University of British Columbia, Vancouver, BC V5Z 1M9, Canada<br>* Correspondence: jkopec@arthritisresearch.ca

Citation: Kopec, J.A.; Sayre, E.C.; Shams, B.; Li, L.C.; Xie, H.; Feehan, L.M.; Esdaile, J.M. The Impact of 51 Risk Factors on Life Expectancy in Canada: Findings from a New Risk Prediction Model Based on Data from the Global Burden of Disease Study. Int. J. Environ. Res. Public Health 2022, 19, 8958. https://doi.org/10.3390/ ijerph19158958

Academic Editor: Paul B. Tchounwou

Received: 19 April 2022
Accepted: 19 July 2022
Published: 23 July 2022
Publisher's Note: MDPI stays neutral with regard to jurisdictional claims in published maps and institutional affiliations.

Copyright: © 2022 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (https:// creativecommons.org/licenses/by/ 4.0/).


#### Abstract

The aims of this study were (1) to develop a comprehensive risk-of-death and life expectancy (LE) model and (2) to provide data on the effects of multiple risk factors on LE. We used data for Canada from the Global Burden of Disease (GBD) Study. To create period life tables for males and females, we obtained age/sex-specific deaths rates for 270 diseases, population distributions for 51 risk factors, and relative risk functions for all disease-exposure pairs. We computed LE gains from eliminating each factor, LE values for different levels of exposure to each factor, and LE gains from simultaneous reductions in multiple risk factors at various ages. If all risk factors were eliminated, LE in Canada would increase by 6.26 years for males and 5.05 for females. The greatest benefit would come from eliminating smoking in males ( 2.45 years) and high blood pressure in females ( 1.42 years). For most risk factors, their dose-response relationships with LE were non-linear and depended on the presence of other factors. In individuals with high levels of risk, eliminating or reducing exposure to multiple factors could improve LE by several years, even at a relatively advanced age.


Keywords: life expectancy; risk factors; prediction models; Global Burden of Disease Study

## 1. Introduction

The health impact of risk factors at the individual level is often measured in terms of absolute risk of disease occurrence or mortality over a specified time period. Risk prediction models, such as the Framingham equations [1], have been employed for many years in the prevention of ischemic heart disease. Prediction models and online risk calculators for many other diseases have been developed, e.g., several cancers [2,3], diabetes [4], or osteoporotic fractures [5]. Despite these advances, and a growing emphasis on personalized medicine in the clinical setting [6], personalized prevention based on quantitative risk information is still an exception rather than a norm [7].

A less common but potentially useful alternative to disease-specific risk prediction is to assess the impact of risk factors in terms of all-cause mortality or life expectancy (LE) [8]. Trends in LE and the underlying causes are important for monitoring population health. In the USA and Canada, for example, LE has declined slightly in recent years [9]. The impact of risk factors on LE at the individual level can be assessed by analyzing data from cohort studies in which all-cause mortality is the outcome. Cohort-based LE effects for some risk factors, e.g., smoking [10], physical activity [11], some dietary factors [12], or metabolic
factors [13] have been published, but such data are difficult to compare across studies and have not been synthesized in the literature.

Another approach to assessing the impact of risk factors is to develop a "synthetic" model which utilizes data from many sources to combine information on mortality and relative risks associated with multiple diseases [8]. This approach has been taken by Lim et al. [14] in developing an all-cause mortality calculator using data from the Global Burden of Disease (GBD) Study and other sources [15]. However, a limitation of Lim's model is that it included only 12 risk factors (body mass index, systolic blood pressure, LDL cholesterol, fasting plasma glucose, seat belt use, smoking, alcohol, physical activity, fruits, vegetables, omega-3 fatty acids, and nuts), whereas GBD generates data for a much larger number of factors. The goal of the current study was to develop a new synthetic life expectancy model based on GBD data, hereafter referred to as the Comprehensive Health and Risk Manager (CHARM), and to estimate LE effects associated with a large number of risk factors.

## 2. Materials and Methods

### 2.1. Conceptual Model

CHARM is a computer program that estimates LE and risk of death from specific conditions based on a person-specific risk profile [16]. Model parameters comprise diseasespecific death rates by age and sex, age/sex-specific distributions of risk factors, and age/sex-specific relative risk functions for all disease-risk combinations included in the model. These parameters were used to create abridged period life-tables for males and females [17]. For each sex and 5-year age category, the death rate from each disease was obtained as the observed rate for the country of interest multiplied by the overall person-specific relative risk (RR) that depends on the values of all risk factors (risk profile).

Estimation of LE starts with the specification of the age and sex of the individual, and the values of the risk factors (e.g., systolic blood pressure, smoking, fruit consumption, etc.). For each risk factor, the exposure-specific RR was calculated relative to the country-specific mean exposure for continuous factors and prevalence-weighted relative risk for categorical factors. Dose-response functions included continuous exponential and more flexible interval risk functions for continuous exposures, and ordinal, nominal, and dichotomous functions for categorical exposures [18]. The overall age/sex-specific RR for each disease was calculated under two different models, a model assuming independent effects (to estimate the total effect of each risk factor) and a model accounting for mediation (to estimate the combined effect of multiple factors, including mediators) [19]. In both models, combined effects were computed according to an additive statistical model (an optional multiplicative model is also available) [19]. Cause-specific death rates from all diseases in a given age/sex category were then summed up to obtain the overall death rate, which was used in an abridged period life table to calculate LE [17]. Hence, estimated LE for a male or female with an average level of all factors would be equal to the observed LE for males or females, respectively, in the country of interest.

To obtain the effect of changes in a risk factor on LE at any age, we developed an (optional) lagged model, which assumes a gradual change in risk following a change in exposure. In this model, we assumed an exponential decay process, with the halflife parameter specific to each exposure-disease pair [20]. Technical details, including mathematical formulas and a description of the computer interface for CHARM, are provided as Supplementary Materials.

### 2.2. Data Sources

We used data from the 2016 and 2017 GBD studies (GBD 2016 and GBD 2017) to obtain age/sex-specific deaths rates for 270 diseases, including residual categories for comprehensiveness (Table S1, Supplementary Materials), population distributions (or means) for 60 exposures, and relative risk functions for 16,200 disease-exposure pairs. GBD is an international, collaborative research project whose goal is to provide comprehensive and
comparable mortality, morbidity and risk factor information for all countries over time [15]. The GBD methodology has been described in detail in numerous publications [21-27]. Data sources included peer-reviewed scientific publications, government reports, population surveys, administrative databases, vital registration, cancer registries, police reports, sales data, satellite measurements, and other sources. All data sources used in the study are listed in the publicly available Global Health Data Exchange database [28]. Additional details, including GBD data processing methods, are provided as Supplementary Materials.

### 2.3. Model Output

In this study, we used mortality and risk factor exposure data for Canada. We analyzed the effects of 51 out of 60 risk factors available in the model. Five factors that apply to children only and four factors that do not cause death were excluded. We computed LE gains from eliminating each factor, LE values for different levels of exposure to each factor, and provided an example of gains in LE from simultaneous reductions in multiple risk factors at various ages for a person with a complex risk factor profile.

### 2.4. Model Performance

To assess model performance, we compared LEs for 40 and 60-year-old males and females according to smoking and body mass index (BMI) between our model and the Mortality Population Risk Tool implemented in the Project Big Life calculator developed by Manuel et al. and based on a Canadian cohort (Canadian Community Health Survey) [8,29]. We also compared a 10-year risk of death from cardiovascular disease (CVD) for a 60-yearold male and female between CHARM and the SCORE chart for low-risk countries in Europe [30]. SCORE is a well-established cohort-based CVD risk-of-death calculator [31].

## 3. Results

### 3.1. Life Expectancy under Current vs Optimal Distribution of Risk Factors

Life expectancy at birth (LE0) in Canada in 2017 was 79.61 years for males and 83.74 for females. The difference in LE between males and females diminished with age and by age 90, it was less than one year. In Table 1 we show LE0 assuming the optimal level of exposure to each factor and the difference in LE0 between the optimal and current average levels of exposure in the population.

Table 1. Life expectancy at birth (LE0) assuming the optimal level of each factor and difference in LE0 between the optimal and current average exposure for 51 risk factors individually.

| Risk Factor (Units) | Life Expectancy <br> at Birth |  | Difference Relative <br> to Baseline | Optimal <br> Level |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females |  |
| Baseline (current average exposure) | 79.61 | 83.74 | 0 | 0 |  |
| Smoking (cigarettes/day since age 20) | 82.06 | 85.00 | 2.45 | 1.26 | 0 |
| Systolic blood pressure (mm Hg) | 81.12 | 85.16 | 1.51 | 1.42 | 110 |
| BMI (kg/m ${ }^{2}$ ) | 81.04 | 84.99 | 1.43 | 1.25 | 20 |
| Blood low-density lipoprotein(LDL) <br> level (mmol/L) | 80.59 | 84.58 | 0.98 | 0.84 | 0.7 |
| Sodium intake (g/day) | 80.41 | 84.42 | 0.80 | 0.68 | 1.0 |
| Whole grain consumption (g/day) | 80.23 | 84.25 | 0.62 | 0.51 | 150 |
| Nuts and seeds consumption (g/day) | 80.11 | 84.10 | 0.50 | 0.36 | 25 |
| Omega-3 fatty acids intake (mg/day) | 80.06 | 84.03 | 0.45 | 0.29 | 300 |

Table 1. Cont.

| Risk Factor (Units) | Life Expectancy at Birth |  | Difference Relative to Baseline |  | Optimal Level |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females |  |
| Fruit consumption (g/day) | 80.04 | 84.04 | 0.43 | 0.30 | 300 |
| Vegetable consumption (g/day) | 79.95 | 84.01 | 0.34 | 0.27 | 430 |
| Blood glucose (mmol/L) | 79.94 | 84.07 | 0.33 | 0.33 | 4.8 |
| Ambient particulate matter (PM 2.5) concentration $\left(\mu \mathrm{g} / \mathrm{m}^{3}\right)$ | 79.92 | 83.99 | 0.31 | 0.25 | 0 |
| Fiber intake (g/day) | 79.85 | 83.93 | 0.24 | 0.19 | 28 |
| Energy intake from polyunsaturated fatty acids (\%) | 79.83 | 83.90 | 0.22 | 0.16 | 13 |
| Alcohol consumption (g/day) | 79.82 | 83.72 | 0.21 | -0.02 | 0 |
| Energy intake from trans fatty acids (\%) | 79.79 | 83.89 | 0.18 | 0.15 | 0 |
| Physical activity (MET-min/week) | 79.78 | 83.95 | 0.17 | 0.21 | 4500 |
| Processed meat consumption (g/day) | 79.77 | 83.82 | 0.16 | 0.08 | 0 |
| Kidney function level | 79.71 | 83.94 | 0.10 | 0.20 | Cat. 5 |
| Bone lead concentration ( $\mu \mathrm{g} / \mathrm{g}$ ) | 79.71 | 83.81 | 0.10 | 0.07 | 0.02 |
| Calcium intake (g/day) | 79.70 | 83.82 | 0.09 | 0.08 | 1.5 |
| Exposure to second-hand smoking | 79.70 | 84.03 | 0.09 | 0.29 | No |
| Asbestos exposure at work | 79.67 | 83.74 | 0.06 | <0.01 | No/Low |
| Legume consumption (g/day) | 79.67 | 83.81 | 0.06 | 0.07 | 70 |
| Milk consumption (g/day) | 79.67 | 83.79 | 0.06 | 0.05 | 520 |
| Ambient ozone concentration (ppb) | 79.66 | 83.78 | 0.05 | 0.04 | 29.1 |
| Red meat consumption (g/day) | 79.66 | 83.76 | 0.05 | 0.02 | 0 |
| Sugar-sweetened beverage consumption (g/day) | 79.66 | 83.76 | 0.05 | 0.02 | 0 |
| Particulate matter, gases, and fumes exposure at work (2 categories) | 79.65 | 83.76 | 0.04 | 0.02 | No |
| Silica exposure at work | 79.64 | 83.75 | 0.03 | 0.01 | No |
| Water quality (12 categories) | 79.62 | 83.74 | 0.01 | $<0.01$ | Cat. 12 |
| Residential radon gas concentration $\left(\mathrm{Bq} / \mathrm{m}^{3}\right)$ | 79.62 | 83.74 | 0.01 | $<0.01$ | 10 |
| Diesel engine exhaust exposure at work (3 categories) | 79.62 | 83.74 | 0.01 | <0.01 | No |
| Chewing tobacco use (2 categories) | 79.62 | 83.74 | 0.01 | <0.01 | No |
| Intimate partner violence (2 categories) | 79.62 | 83.74 | 0.01 | <0.01 | No |
| Childhood sexual abuse (2 categories) | 79.62 | 83.74 | 0.01 | <0.01 | No |
| Sanitation facility (3 categories) | 79.61 | 83.74 | <0.01 | <0.01 | Cat. 3 |
| Handwashing facility (2 categories) | 79.61 | 83.74 | <0.01 | <0.01 | Cat. 2 |
| Household use of solid fuels (2 categories) | 79.61 | 83.74 | $<0.01$ | <0.01 | No |

Table 1. Cont.

| Risk Factor (Units) | Life Expectancy at Birth |  | Difference Relative to Baseline |  | Optimal Level |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females |  |
| Arsenic exposure at work (3 categories) | 79.61 | 83.74 | $<0.01$ | <0.01 | No |
| Benzene exposure at work (3 categories) | 79.61 | 83.74 | $<0.01$ | <0.01 | No |
| Beryllium exposure at work (3 categories) | 79.61 | 83.74 | $<0.01$ | <0.01 | No |
| Cadmium exposure at work (3 categories) | 79.61 | 83.74 | $<0.01$ | <0.01 | No |
| Chromium exposure at work (3 categories) | 79.61 | 83.74 | $<0.01$ | <0.01 | No |
| Formaldehyde exposure at work (3 categories) | 79.61 | 83.74 | $<0.01$ | <0.01 | No |
| Nickel exposure at work (3 categories) | 79.61 | 83.74 | $<0.01$ | $<0.01$ | No |
| Polycyclic aromatic hydrocarbon exposure at work (3 categories) | 79.61 | 83.74 | $<0.01$ | <0.01 | No |
| Sulfuric acid exposure at work (3 categories) | 79.61 | 83.74 | $<0.01$ | <0.01 | No |
| Trichloroethylene exposure at work (3 categories) | 79.61 | 83.74 | $<0.01$ | $<0.01$ | No |
| Occupation type for asthmagens exposure (9 categories) | 79.61 | 83.74 | $<0.01$ | <0.01 | Other |
| Blood hemoglobin level (g/dL) | 79.61 | 83.74 | N/A | <0.01 | 15 |
| All dietary factors | 81.80 | 85.58 | 2.19 | 1.84 | Optimal |
| All factors | 85.87 | 88.79 | 6.26 | 5.05 | Optimal |

LE0 values represent total effects (direct and mediated by other factors), except for all factors combined. Risk factors are ordered according to their effect on LE0, from largest to smallest (in males). Baseline LE0 assumes all risk factors at average values. Exposure categories for categorical risk factors are listed in Table S2 (Supplementary Materials). The optimal level is based on the Theoretical Minimum Risk Exposure Level (TMREL) used in GBD studies [27]. Dietary factors in the model are sodium, whole grain, nuts and seeds, omega-3 fatty acids, fruits, vegetables, fibre, processed meat, red meat, polyunsaturated fatty acids, trans fatty acids, calcium, legumes, milk, and sugar-sweetened beverages.

The greatest difference in LE0 for males was from eliminating smoking ( 2.45 years), followed by high systolic blood pressure (SBP) (1.51 years), high body mass index (BMI) (1.43 years), high low-density lipoproteins (LDL) cholesterol ( 0.98 years), and high sodium intake ( 0.80 years). For females, the greatest difference was for SBP (1.42 years) followed by smoking (1.26), BMI (1.25), LDL (0.84), and sodium (0.68). The effect of eliminating alcohol drinking was 0.21 years for males and -0.02 (slightly detrimental) for females. Increasing physical activity to 4500 MET-min/week would, on average, improve LE0 by 0.17 and 0.21 years for males and females, respectively. Many of the factors studied had a very small impact on LE0 ( 0.01 years or less), mainly due to relatively low population levels of exposure in Canada, such that the average person (not a person exposed) would not gain much from removing these factors. These include most chemical exposures at work and some environmental factors (Table 1).

The combined effect of improving all 51 risk factors from the current average to the optimal level (adjusted for mediation) was 6.26 years for males and 5.05 years for females (Table 1), resulting in LE0s of 85.87 years for males and 88.79 for females. The effect of optimizing all dietary factors was 2.19 years in males and 1.84 in females. The impact of
prevention can, of course, be much stronger in groups and individuals with higher than average levels of risk.

### 3.2. LE According to Risk Factor Levels

For most risk factors, the relationship between exposure and LE0 was non-linear (Tables 2-4). For example, for smoking, the effect of one unit of exposure was greater for low levels of smoking, whereas for SBP, BMI, alcohol, and some dietary factors, the opposite was true. Smoking 10 cigarettes a day since age 20 was associated with an LE0 of 78.21 for males and 79.58 for females, compared with 82.06 and 85.00 , respectively, for non-smokers (differences of 3.85 and 5.42 years), indicating a stronger impact of smoking on LE0 in females (Table 2). Among the dietary factors, there was a difference of 0.97 years for males and 0.88 years for females when comparing the optimal sodium intake ( 1 g a day) with a high intake ( 5.5 g a day). Eating 300 g (about 4 servings) of fruit a day compared to 75 g produced an LE0 difference of 0.66 and 0.55 years for males and females, respectively. For physical activity, the difference in LE0 between 0 and 1125 MET-min per week was 0.46 and 0.39 years for males and females, whereas the same absolute difference between 1125 and 2250 MET-min per week was only 0.13 and 0.12 years, respectively. With respect to alcohol, we considered $0 \mathrm{~g} /$ day to be optimal; however, drinking 18 g per day ( 1 drink is about 14 g ) had very little impact on LE0. Drinking 36 g of alcohol per day reduced LE0 by 0.51 and 0.49 years for males and females, respectively.

Table 2. Life expectancy at birth for selected levels of 21 behavioral risk factors and differences relative to optimal levels.

| Risk Factor Levels | Life Expectancy at Birth |  | LE Difference Relative to Best Level |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females |
| Smoking (cigarettes/day since age 20) |  |  |  |  |
| 0 | 82.06 | 85.00 | - | - |
| 10 | 78.21 | 79.58 | 3.85 | 5.42 |
| 20 | 76.68 | 76.99 | 5.38 | 8.01 |
| 30 | 75.34 | 75.03 | 6.72 | 9.97 |
| 40 | 74.29 | 73.84 | 7.77 | 11.16 |
| Physical activity (MET-min/week) |  |  |  |  |
| 4500 | 79.78 | 83.95 | - | - |
| 3375 | 79.69 | 83.88 | 0.09 | 0.07 |
| 2250 | 79.56 | 83.77 | 0.22 | 0.18 |
| 1125 | 79.43 | 83.65 | 0.35 | 0.30 |
| 0 | 78.97 | 83.26 | 0.81 | 0.69 |
| Alcohol consumption (g/day) |  |  |  |  |
| 0 | 79.82 | 83.72 | - | - |
| 18 | 79.79 | 83.75 | 0.03 | $-0.03$ |
| 36 | 79.31 | 83.23 | 0.51 | 0.49 |
| 54 | 78.57 | 82.58 | 1.25 | 1.14 |
| 72 | 77.60 | 81.82 | 2.22 | 1.90 |
| Sodium intake (g/day) |  |  |  |  |
| 1 | 80.41 | 84.42 | - | - |

Table 2. Cont.

| Risk Factor Levels | Life Expectancy at Birth |  | LE Difference Relative to Best Level |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females |
| 2.5 | 80.12 | 84.15 | 0.29 | 0.27 |
| 4 | 79.80 | 83.86 | 0.61 | 0.56 |
| 5.5 | 79.44 | 83.54 | 0.97 | 0.88 |
| 7 | 79.06 | 83.19 | 1.35 | 1.23 |
| Fruit consumption (g/day) |  |  |  |  |
| 300 | 80.04 | 84.04 | - | - |
| 225 | 79.83 | 83.87 | 0.21 | 0.17 |
| 150 | 79.61 | 83.69 | 0.43 | 0.35 |
| 75 | 79.38 | 83.49 | 0.66 | 0.55 |
| 0 | 79.12 | 83.28 | 0.92 | 0.76 |
| Vegetable consumption (g/day) |  |  |  |  |
| 430 | 79.95 | 84.01 | - | - |
| 322.5 | 79.8 | 83.9 | 0.15 | 0.11 |
| 215 | 79.65 | 83.78 | 0.30 | 0.23 |
| 107.5 | 79.48 | 83.66 | 0.47 | 0.35 |
| 0 | 79.30 | 83.53 | 0.65 | 0.48 |
| Whole grain consumption (g/day) |  |  |  |  |
| 150 | 80.23 | 84.25 | - | - |
| 112.5 | 80.06 | 84.11 | 0.17 | 0.14 |
| 75 | 79.87 | 83.96 | 0.36 | 0.29 |
| 37.5 | 79.67 | 83.80 | 0.56 | 0.45 |
| 0 | 79.44 | 83.63 | 0.79 | 0.62 |
| Red meat consumption (g/day) |  |  |  |  |
| 0 | 79.66 | 83.76 | - | - |
| 51.25 | 79.61 | 83.73 | 0.05 | 0.03 |
| 102.5 | 79.57 | 83.68 | 0.09 | 0.08 |
| 153.75 | 79.52 | 83.64 | 0.14 | 0.12 |
| 205 | 79.46 | 83.59 | 0.20 | 0.17 |
| Processed meat consumption (g/day) |  |  |  |  |
| 0 | 79.77 | 83.82 | - | - |
| 22.5 | 79.46 | 83.60 | 0.31 | 0.22 |
| 45 | 79.12 | 83.36 | 0.65 | 0.46 |
| 67.5 | 78.74 | 83.10 | 1.03 | 0.72 |
| 90 | 78.31 | 82.80 | 1.46 | 1.02 |
| Sugar-sweetened beverage consumption (g/day) |  |  |  |  |
| 0 | 79.66 | 83.76 | - | - |
| 250 | 79.42 | 83.61 | 0.24 | 0.15 |

Table 2. Cont.

| Risk Factor Levels | Life Expectancy at Birth |  | LE Difference Relative to Best Level |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females |
| 500 | 79.16 | 83.44 | 0.50 | 0.32 |
| 750 | 78.86 | 83.25 | 0.80 | 0.51 |
| 1000 | 78.52 | 83.04 | 1.14 | 0.72 |
| Milk consumption (g/day) |  |  |  |  |
| 520 | 79.67 | 83.79 | - | - |
| 390 | 79.65 | 83.77 | 0.02 | 0.02 |
| 260 | 79.63 | 83.75 | 0.04 | 0.04 |
| 130 | 79.60 | 83.73 | 0.07 | 0.06 |
| 0 | 79.58 | 83.71 | 0.09 | 0.08 |
| Legume consumption (g/day) |  |  |  |  |
| 70 | 79.67 | 83.81 | - | - |
| 52.5 | 79.59 | 83.76 | 0.08 | 0.05 |
| 35 | 79.52 | 83.71 | 0.15 | 0.10 |
| 17.5 | 79.43 | 83.65 | 0.24 | 0.16 |
| 0 | 79.35 | 83.60 | 0.32 | 0.21 |
| Nuts and seeds consumption (g/day) |  |  |  |  |
| 25 | 80.11 | 84.10 | - | - |
| 18.75 | 79.98 | 84.00 | 0.13 | 0.10 |
| 12.5 | 79.84 | 83.90 | 0.27 | 0.20 |
| 6.25 | 79.68 | 83.79 | 0.43 | 0.31 |
| 0 | 79.51 | 83.68 | 0.60 | 0.42 |
| Fiber intake (g/day) |  |  |  |  |
| 28 | 79.85 | 83.93 | - | - |
| 21 | 79.73 | 83.85 | 0.12 | 0.08 |
| 14 | 79.61 | 83.76 | 0.24 | 0.17 |
| 7 | 79.48 | 83.67 | 0.37 | 0.26 |
| 0 | 79.34 | 83.57 | 0.51 | 0.36 |
| Calcium intake (g/day) |  |  |  |  |
| 1.5 | 79.70 | 83.82 | - | - |
| 1.125 | 79.66 | 83.79 | 0.04 | 0.03 |
| 0.75 | 79.62 | 83.75 | 0.08 | 0.07 |
| 0.375 | 79.57 | 83.71 | 0.13 | 0.11 |
| 0 | 79.51 | 83.66 | 0.19 | 0.16 |
| Omega-3 fatty acids intake (mg/day) |  |  |  |  |
| 300 | 80.06 | 84.03 | - | - |
| 225 | 79.96 | 83.97 | 0.10 | 0.06 |
| 150 | 79.85 | 83.89 | 0.21 | 0.14 |
| 75 | 79.73 | 83.82 | 0.33 | 0.21 |
| 0 | 79.61 | 83.74 | 0.45 | 0.29 |

Table 2. Cont.

| Risk Factor Levels | Life Expectancy at Birth |  | LE Difference Relative to Best Level |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females |
| Energy intake from polyunsaturated fatty acids (\%) |  |  |  |  |
| 13 | 79.83 | 83.90 | - | - |
| 9.75 | 79.76 | 83.84 | 0.07 | 0.06 |
| 6.5 | 79.68 | 83.79 | 0.15 | 0.11 |
| 3.25 | 79.60 | 83.73 | 0.23 | 0.17 |
| 0 | 79.52 | 83.67 | 0.31 | 0.23 |
| Energy intake from trans fatty acids (\%) |  |  |  |  |
| 0 | 79.79 | 83.89 | - | - |
| 1.05 | 79.62 | 83.77 | 0.17 | 0.12 |
| 2.1 | 79.42 | 83.65 | 0.37 | 0.24 |
| 3.15 | 79.20 | 83.51 | 0.59 | 0.38 |
| 4.2 | 78.95 | 83.37 | 0.84 | 0.52 |
| Chewing tobacco use |  |  |  |  |
| No | 79.62 | 83.74 | - | - |
| Yes | 79.32 | 83.50 | 0.30 | 0.24 |
| Childhood sexual abuse |  |  |  |  |
| No | 79.62 | 83.74 | - | - |
| Yes | 79.58 | 83.73 | 0.04 | 0.01 |
| Exposure to intimate partner violence (ever) |  |  |  |  |
| No | 79.62 | 83.74 | - | - |
| Yes | 79.60 | 83.73 | 0.02 | 0.01 |

$\overline{\text { LE0 values are estimated for specific, selected levels of each factor, assumed to be constant over lifetime, and }}$ represent total effects (direct and mediated by other factors). For continuous exposures, the levels (5) are ordered from best to worst and evenly distributed. The best level is based on the Theoretical Minimum Risk Exposure Level (TMREL) in GBD studies [27] and the worst level is chosen based on exposure distribution, data availability, and clinical considerations (see Supplementary Materials). Exposure categories for categorical risk factors are listed in Table S2 in Supplementary Materials.

Table 3. Life expectancy at birth for selected levels of 7 metabolic risk factors and differences relative to optimal levels.

| Risk Factor Levels | Life Expectancy at Birth |  | LE Difference Relative to Best Level |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females |
| Body mass index (kg/m²) |  |  |  |  |
| 20 | 81.04 | 84.99 | - | - |
| 25 | 80.16 | 84.15 | 0.88 | 0.84 |
| 30 | 79.04 | 83.14 | 2.00 | 1.85 |
| 35 | 77.56 | 81.83 | 3.48 | 3.16 |
| 40 | 75.46 | 80.00 | 5.58 | 4.99 |
| Systolic blood pressure ( mm Hg ) |  |  |  |  |
| 110 | 81.12 | 85.16 | - | - |
| 127.5 | 80.01 | 84.13 | 1.11 | 1.03 |

Table 3. Cont.

| Risk Factor Levels | Life Expectancy at Birth |  | LE Difference Relative to Best Level |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females |
| 145 | 78.26 | 82.48 | 2.86 | 2.68 |
| 162.5 | 75.22 | 79.42 | 5.90 | 5.74 |
| 180 | 69.31 | 72.61 | 11.81 | 12.55 |

Blood low-density lipoprotein (LDL) cholesterol (mmol/L)

| 0.7 | 80.59 | 84.58 | - | - |
| :---: | :---: | :---: | :---: | :---: |
| 2.025 | 80.14 | 84.22 | 0.45 | 0.36 |
| 3.35 | 79.52 | 83.78 | 1.07 | 0.80 |
| 4.675 | 78.61 | 83.20 | 1.98 | 1.38 |
| 6 | 77.21 | 82.42 | 3.38 | 2.16 |
| 4.8 |  | Blood glucose (mmol/L) |  |  |
| 6.1 | 79.94 | 84.07 | - | - |
| 7.4 | 79.36 | 83.56 | 0.58 | 0.51 |
| 8.7 | 77.79 | 81.70 | 2.15 | 2.37 |
| 10 | 77.03 | 81.10 | 2.91 | 2.97 |
|  | 76.12 | 80.37 | 3.82 | 3.70 |


| Blood hemoglobin level (g/dL)* |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| 15 | 79.61 | 83.74 | - | - |
| 12.75 | 79.61 | 83.74 | - | $<0.01$ |
| 10.5 | 79.61 | 83.73 | - | 0.01 |
| 8.25 | 79.61 | 83.73 | - | 0.01 |
| 6 | 79.61 | 83.71 | - | 0.03 |

Bone lead concentration $(\mu \mathrm{g} / \mathrm{g})$

| 0.02 | 79.71 | 83.81 | - | - |
| :---: | :---: | :---: | :---: | :---: |
| 0.315 | 79.71 | 83.80 | $<0.01$ | 0.01 |
| 0.61 | 79.71 | 83.80 | $<0.01$ | 0.01 |
| 0.905 | 79.70 | 83.80 | 0.01 | 0.01 |
| 1.2 | 79.70 | 83.80 | 0.01 | 0.01 |


|  | Kidney function level |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Category 5 | 79.71 | 83.94 | - | - |
| Category 4 | 79.32 | 83.73 | 0.39 | 0.21 |
| Category 3 | 78.97 | 83.38 | 0.74 | 0.56 |
| Category 2 | 76.75 | 81.94 | 2.96 | 2.00 |
| Category 1 | 74.78 | 80.58 | 4.93 | 3.36 |

LE0 values are estimated for specific, selected levels of each factor, assumed to be constant over lifetime, and represent total effects (direct and mediated by other factors). For continuous exposures, the levels (5) are ordered from best to worst and evenly distributed. The best level is based on the Theoretical Minimum Risk Exposure Level (TMREL) in GBD studies [27] and the worst level is chosen based on exposure distribution, data availability, and clinical considerations (see Supplementary Materials). Exposure categories for categorical risk factors are listed in Table S2 in Supplementary Materials. * The effect of low blood hemoglobin is only estimated for women.

Table 4. Life expectancy at birth for selected levels of 23 environmental and occupational risk factors and differences relative to optimal levels.

| Risk Factor Levels | Life Expectancy at Birth |  | LE Difference Relative to Best Level |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females |
| Ambient particulate matter ( $\mathrm{PM}_{2.5}$ ) concentration ( $\mu \mathrm{g} / \mathrm{m}^{3}$ ) |  |  |  |  |
| 0 | 79.92 | 84.09 | - | - |
| 150 | 78.33 | 82.35 | 1.59 | 1.74 |
| 300 | 77.80 | 81.80 | 2.12 | 2.29 |
| 450 | 77.44 | 81.43 | 2.48 | 2.66 |
| 600 | 77.14 | 81.12 | 2.78 | 2.97 |
| Ambient ozone concentration (ppb) |  |  |  |  |
| 29.1 | 79.66 | 83.78 | - | - |
| 47.825 | 79.62 | 83.74 | 0.04 | 0.04 |
| 66.55 | 79.57 | 83.70 | 0.09 | 0.08 |
| 85.275 | 79.51 | 83.65 | 0.15 | 0.13 |
| 104 | 79.45 | 83.59 | 0.21 | 0.19 |
| Residential radon gas concentration ( $\mathrm{Bq} / \mathrm{m}^{3}$ ) |  |  |  |  |
| 10 | 79.62 | 83.74 | - | - |
| 207.5 | 79.19 | 83.34 | 0.43 | 0.4 |
| 405 | 78.59 | 82.76 | 1.03 | 0.98 |
| 602.5 | 77.75 | 81.94 | 1.87 | 1.80 |
| 800 | 76.59 | 80.78 | 3.03 | 2.96 |
| Exposure to second-hand smoke |  |  |  |  |
| No | 79.70 | 84.03 | - | - |
| Yes | 79.03 | 83.45 | 0.67 | 0.58 |
| Asbestos exposure at work |  |  |  |  |
| Low/No | 79.67 | 83.74 | - | - |
| High | 78.63 | 82.89 | 1.04 | 0.85 |
| Arsenic exposure at work |  |  |  |  |
| No | 79.61 | 83.74 | - | - |
| Low | 78.97 | 83.12 | 0.64 | 0.62 |
| High | 78.70 | 82.87 | 0.91 | 0.87 |
| Benzene exposure at work |  |  |  |  |
| No | 79.61 | 83.74 | - | - |
| Low | 79.53 | 83.67 | 0.08 | 0.07 |
| High | 79.39 | 83.57 | 0.22 | 0.17 |
| Beryllium exposure at work |  |  |  |  |
| No | 79.61 | 83.74 | - | - |
| Low | 79.61 | 83.74 | <0.01 | <0.01 |
| High | 79.46 | 83.60 | 0.15 | 0.14 |
| Cadmium exposure at work |  |  |  |  |
| No | 79.61 | 83.74 | - | - |

Table 4. Cont.

| Risk Factor Levels | Life Expectancy at Birth |  | LE Difference Relative to Best Level |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females |
| Low | 79.61 | 83.74 | <0.01 | <0.01 |
| High | 79.44 | 83.58 | 0.17 | 0.16 |
| Chromium exposure at work |  |  |  |  |
| No | 79.61 | 83.74 | - | - |
| Low | 79.61 | 83.74 | <0.01 | <0.01 |
| High | 79.45 | 83.59 | 0.16 | 0.15 |
| Diesel engine exhaust exposure at work |  |  |  |  |
| No | 79.62 | 83.74 | - | - |
| Low | 79.62 | 83.74 | <0.01 | <0.01 |
| High | 79.20 | 83.35 | 0.42 | 0.39 |
| Formaldehyde exposure at work |  |  |  |  |
| No | 79.61 | 83.74 | - | - |
| Low | 79.61 | 83.74 | <0.01 | <0.01 |
| High | 79.54 | 83.68 | 0.07 | 0.06 |
| Nickel exposure at work |  |  |  |  |
| No | 79.61 | 83.74 | - | - |
| Low | 79.15 | 83.29 | 0.46 | 0.45 |
| High | 78.63 | 82.80 | 0.98 | 0.94 |
| Polycyclic aromatic hydrocarbon exposure at work |  |  |  |  |
| No | 79.61 | 83.74 | - | - |
| Low | 79.61 | 83.74 | <0.01 | <0.01 |
| High | 79.34 | 83.48 | 0.27 | 0.26 |
| Silica exposure at work |  |  |  |  |
| No | 79.64 | 83.75 | - | - |
| Low | 79.18 | 83.31 | 0.46 | 0.44 |
| High | 79.04 | 83.18 | 0.60 | 0.57 |
| Sulfuric acid exposure at work |  |  |  |  |
| No | 79.61 | 83.74 | - | - |
| Low | 79.58 | 83.73 | 0.03 | 0.01 |
| High | 79.50 | 83.71 | 0.11 | 0.03 |
| Trichloroethylene exposure at work |  |  |  |  |
| No | 79.61 | 83.74 | - | - |
| Low | 79.61 | 83.74 | <0.01 | <0.01 |
| High | 79.58 | 83.72 | 0.03 | 0.02 |
| Particulate matter, gases, and fumes exposure at work |  |  |  |  |
| No | 79.65 | 83.89 | - | - |
| Low | 79.47 | 83.78 | 0.18 | 0.11 |
| High | 79.13 | 83.55 | 0.52 | 0.34 |

Table 4. Cont.

| Risk Factor Levels | Life Expectancy at Birth |  | LE Difference Relative to Best Level |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Males | Females | Males | Females |
| Occupation type for asthmagens exposure |  |  |  |  |
| Other | 79.61 | 83.74 | - | - |
| Administration | 79.61 | 83.74 | <0.01 | <0.01 |
| Technical | 79.61 | 83.74 | <0.01 | <0.01 |
| Sales | 79.61 | 83.73 | <0.01 | 0.01 |
| Agriculture | 79.61 | 83.73 | <0.01 | 0.01 |
| Mining | 79.61 | 83.74 | <0.01 | 0.01 |
| Transport | 79.61 | 83.74 | <0.01 | <0.01 |
| Manufacturing | 79.61 | 83.74 | <0.01 | <0.01 |
| Services | 79.61 | 83.74 | <0.01 | <0.01 |
| Household use of solid fuels causing air pollution |  |  |  |  |
| No | 79.61 | 83.74 | - | - |
| Yes | 78.73 | 82.76 | 0.88 | 0.98 |
| Water quality |  |  |  |  |
| Category 12 | 79.62 | 83.74 | - | - |
| Category 9 | 79.56 | 83.68 | 0.06 | 0.06 |
| Category 6 | 79.34 | 83.41 | 0.28 | 0.33 |
| Category 3 | 79.31 | 83.37 | 0.31 | 0.37 |
| Category 1 | 79.23 | 83.28 | 0.39 | 0.46 |
| Sanitation facility |  |  |  |  |
| Category 3 | 79.61 | 83.74 | - | - |
| Category 2 | 79.55 | 83.66 | 0.06 | 0.08 |
| Category 1 | 79.53 | 83.63 | 0.08 | 0.11 |
| Handwashing facility |  |  |  |  |
| Category 2 | 79.61 | 83.74 | - | - |
| Category 1 | 79.53 | 83.64 | 0.08 | 0.10 |

$\overline{\text { LE0 values are estimated for specific, selected levels of each factor, assumed to be constant over lifetime, and }}$ represent total effects (direct and mediated by other factors). For continuous exposures, the levels (5) are ordered from best to worst and evenly distributed. The best level is based on the Theoretical Minimum Risk Exposure Level (TMREL) in GBD studies [27] and the worst level is chosen based on exposure distribution, data availability, and clinical considerations (see Supplementary Materials). Exposure categories for categorical risk factors are listed in Table S2 in Supplementary Materials.

We evaluated the impact of 7 metabolic factors (Table 3). Of those, a BMI of 35 resulted in an LE0 of 77.56 for males and 81.83 for females, i.e., 3.48 and 3.16 years less than the LE0 for males and females with an optimal BMI of 20. Other metabolic factors with a potentially strong impact on LE0 were systolic blood pressure, LDL cholesterol, blood glucose, and kidney function.

Among the environmental and occupational factors, a high level of exposure to asbestos was associated with a 1.04-year reduction in LE0 in males and 0.85 years in females (Table 4). Work exposure to arsenic, nickel, silica, and particulate matter, gas, and fumes had a relatively strong impact on LE0. Air pollution was a potentially strong factor, with $\mathrm{PM}_{2.5}$ levels around $300 \mu \mathrm{~g} / \mathrm{m}^{3}$ associated with a reduction in LE0 of more than 2 years. Very high levels of radon exposure also had a significant impact on LE0.

### 3.3. Effect of Change in Individual Factors and Their Combinations

In Tables 5-7, we illustrate the mediation-adjusted LE impact of changing selected risk factors individually and in combination in a hypothetical person (male and female) with a specified risk profile. For this example, we selected 10 factors that we considered important for population health in Canada, were modifiable, and represented both behavioral and metabolic factors.

Table 5. An example of gains in LE50 for males and females 50 years of age with 10 risk factors, resulting from a reduction in exposure to each factor individually, according to a no-lag and lagged model, adjusted for mediation.

| Risk Factor | Initial <br> Level | Final <br> Level | Males |  | Females |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Lagged | No Lag | Lagged |  |  |
| Baseline life expectancy |  |  | 77.56 | 77.56 | 79.31 | 79.31 |
| Smoking (cig./day since age 20) | 10 | 0 | 3.18 | 2.63 | 4.23 | 3.63 |
| Systolic blood pressure (mm Hg) | 150 | 120 | 2.20 | 1.85 | 1.31 | 1.15 |
| Low-density lipoproteins <br> (mmol/L) | 5 | 2 | 1.73 | 1.42 | 0.87 | 0.77 |
| Body mass index (kg/m |  |  |  |  |  |  |
| Whole grain (g/day) | 35 | 25 | 1.07 | 0.86 | 0.75 | 0.62 |
| Fruit (g/day) | 10 | 150 | 0.70 | 0.57 | 0.42 | 0.36 |
| Processed meat (g/day) | 75 | 300 | 0.63 | 0.50 | 0.50 | 0.40 |
| Alcohol (g/day) | 42 | 0 | 0.52 | 0.42 | 0.24 | 0.21 |
| Physical activity <br> (MET-min/week) | 600 | 4000 | 0.32 | 0.29 | 0.22 | 0.20 |
| Sodium (g/day) | 7 | 1 | 0.11 | 0.09 | 0.09 | 0.07 |

The lagged model assumes a gradual change in risk following a change in exposure. LE50, life expectancy at age 50 years conditional on surviving until age 50; MET, metabolic equivalent of task. The risk factors are ordered by impact in males.

Table 6. An example of cumulative LE50 gains in males and females 50 years of age, resulting from a reduction in exposure to multiple risk factors, considering the levels of other factors and adjusting for mediation.

| Risk Factors Ordered from Strongest to Weakest |  |  |  | Risk Factors Ordered from Weakest to Strongest |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Males |  | Females |  | Males |  | Females |  |
| Risk Factors | Cumulative <br> LE50 Gain | Risk Factors | Cumulative <br> LE50 Gain | Risk Factors | Cumulative <br> LE50 Gain | Risk Factors | Cumulative <br> LE50 Gain |
| Baseline | 0.00 | Baseline | 0.00 | Baseline | 0.00 | Baseline | 0.00 |
| Smoking | 2.63 | Smoking | 3.63 | Sodium | 0.09 | Sodium | 0.07 |
| SBP | 4.73 | SBP | 5.22 | PA | 0.38 | PA | 0.27 |
| LDL | 5.58 | LDL | 5.79 | Alcohol | 0.70 | Meat | 0.47 |
| BMI | 6.43 | BMI | 6.76 | Meat | 1.13 | Alcohol | 0.78 |
| Grain | 6.61 | Fruit | 6.99 | Fruit | 1.60 | Grain | 1.10 |
| Fruit | 6.79 | Grain | 7.12 | Grain | 2.04 | Fruit | 1.44 |
| Meat | 7.07 | Alcohol | 7.63 | BMI | 2.83 | BMI | 2.07 |
| Alcohol | 7.50 | Meat | 7.86 | LDL | 3.72 | LDL | 2.58 |

Table 6. Cont.

| Risk Factors Ordered from Strongest to Weakest |  |  |  | Risk Factors Ordered from Weakest to Strongest |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Males |  | Females |  | Males |  | Females |  |
| Risk Factors | Cumulative <br> LE50 Gain | Risk Factors | Cumulative <br> LE50 Gain | Risk Factors | Cumulative LE50 Gain | Risk Factors | Cumulative LE50 Gain |
| PA | 7.72 | PA | 8.05 | SBP | 4.78 | SBP | 3.26 |
| Sodium | 7.85 | Sodium | 8.20 | Smoking | 7.85 | Smoking | 8.20 |

Ordering of risk factors is based on individual LE50 effects (Table 5). The initial and final levels of the risk factors are as follows: Smoking 10 and 0 cig./day since age 20; SBP: 150 and 120 mm Hg ; BMI: 35 and $25 \mathrm{~kg} / \mathrm{m}^{2}$; LDL: 5 and $2 \mathrm{mmol} / \mathrm{L}$; Whole grain: 10 and $150 \mathrm{~g} /$ day; Fruit: 75 and $300 \mathrm{~g} /$ day; Processed meat: 50 and $0 \mathrm{~g} /$ day; PA: 600 and 4000 MET-min/week; Alcohol: 42 and $0 \mathrm{~g} /$ day; Sodium: 7 and $1 \mathrm{~g} /$ day. Data have been obtained from a lagged model. LE50, life expectancy at age 50; SBP, systolic blood pressure; BMI, body mass index; LDL, low-density lipoproteins; PA, physical activity; MET, metabolic equivalent of task.

Table 7. An example of conditional LE and LE gains for males and females resulting from changing the level of exposure to 10 risk factors simultaneously, at different ages, adjusted for mediation.

|  | Males |  |  |  | Females |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age at <br> Change | LE <br> without <br> Change | LE with <br> Change | Difference | LE <br> without <br> Change | LE with <br> Change | Difference |  |
| 30 years | 75.81 | 85.32 | 9.51 | 77.91 | 87.55 | 9.64 |  |
| 50 years | 77.56 | 85.41 | 7.85 | 79.31 | 87.51 | 8.20 |  |
| 70 years | 82.39 | 86.28 | 3.89 | 83.07 | 87.49 | 4.42 |  |

The initial and final levels of the risk factors (Table 5) are as follows: Smoking 10 and 0 cig./day since age 20; SBP: 150 and 120 mm Hg ; BMI: 35 and $25 \mathrm{~kg} / \mathrm{m}^{2}$; LDL: 5 and $2 \mathrm{mmol} / \mathrm{L}$; Whole grain: 10 and $150 \mathrm{~g} /$ day Fruit: 75 and $300 \mathrm{~g} /$ day; Processed meat: 50 and $0 \mathrm{~g} /$ day; PA: 600 and 4000 MET-min/week; Alcohol: 42 and $0 \mathrm{~g} /$ day; Sodium: 7 and $1 \mathrm{~g} /$ day. Data have been obtained from a lagged model. Conditional LE is LE at a given age, conditional on surviving until that age. LE, life expectancy; SBP, systolic blood pressure; BMI, body mass index; LDL, low-density lipoproteins; PA, physical activity; MET, metabolic equivalent of task.

In the standard instantaneous (no lag) effect model, conditional LE at a given age following a change in exposure is the same as the LE of a person who has never been exposed. In the lagged model, we assume a gradual change in risk according to an exponential decay process, with a half-life parameter specific to each disease-exposure pair. In the no-lag model, quitting smoking at age 50 in our example (a person who has smoked 10 cigarettes/day since age 20) would produce gains of 3.18 and 4.23 years in males and females, respectively (Table 5). A more realistic effect of quitting would be obtained by assuming that the risk in those who quit will gradually approach the level of risk among lifetime non-smokers, resulting in gains of 2.63 and 3.63 years, as suggested by the lagged model.

The impact of changing the level of a risk factor may depend on the levels of other factors, as shown in Table 6 for the risk profile considered in Table 5. In this example, quitting smoking once other risk factors have been improved produces larger gains in LE50 than those observed previously, i.e., 3.07 years for males and 4.94 for females. For SBP, however, the interaction with other factors is in the opposite direction, such that the impact of reducing SBP is higher when the levels of other risks are also high.

In Table 7, we show the LE impact of modifying 10 risk factors at the same time, at age 30,50, or 70, for the hypothetical risk profile considered in Table 5. Differences in LE associated with improvements in all 10 factors ranged from 9.51 years in males and 9.64 in females at age 30 to 3.89 and 4.42 years, respectively, at age 70. Figure S1 (Supplementary Materials) illustrates the impact of these changes in risk on survival probability. It shows how, after a change in exposure at a specified age, survival probability departs from the survival curve of persons with a high level of exposure to all 10 factors and gradually approaches the survival curve of those who have never been exposed to such levels of risk.

### 3.4. Comparisons with Other Models

Comparisons of LEs according to the number of cigarettes and BMI between CHARM and the Big Life calculator are shown graphically in Figure S2 (Supplementary Materials). Among male non-smokers, the effect of BMI on LE was slightly lower in our model. Among female non-smokers, effects were similar, but LEs from our model tended to be slightly higher than those from Big Life in younger females and lower in older females (Figure S2C,D). For smoking, LEs from CHARM were higher for males, especially among light smokers, but lower for females, especially among heavy smokers (Figure S2G,H). The mean difference between the models was -0.027 years and the absolute mean difference was 1.96 years.

In Figure S3 (Supplementary Materials), we compare the risk of dying from CVD over 10 years in the European SCORE charts to the mediation-adjusted risk generated by CHARM for males and females age 60 according to smoking status, SBP and cholesterol level. Using a color scheme similar to the standard SCORE charts, the figure shows a high level of agreement between the two models. Mean absolute differences in predicted risk of dying from CVD over 10 years between the two models ranged from $0.44 \%$ for a female smoker to $1.52 \%$ for a male smoker. The maximum difference was $3.71 \%$ and was observed for a male non-smoker with an SBP of 180 and LDL of $5 \mathrm{mmol} / \mathrm{L}$.

## 4. Discussion

In this report we described a novel risk and LE model based on GBD data and provided estimates of the LE impact of 51 risk factors. For many of the factors studied, including most occupational and some dietary factors, these are the first published estimates of their impact on LE. When we set all risk factors to their optimal values, life expectancy at birth for Canada increased to 85.87 for males and 88.79 for females. In males, the greatest impact on LE0 could be gained from eliminating smoking, followed by high SBP, high BMI, high LDL cholesterol, and several dietary risks. In females, the greatest impact would be achieved from optimizing SBP, followed by smoking, BMI, LDL, and dietary factors. The effect of changing a given risk factor was often non-linear and depended on the levels of other factors. Reducing exposure to multiple factors resulted in potentially large gains in life expectancy, even at a relatively advanced age.

### 4.1. Validity of the Conceptual Model

Our model development process has been described briefly in the methods section. A detailed, technical description of the model and the interactive computer program we have developed are given in Supplementary Materials. It should be noted that our LE estimates are not expected to reflect past experience of any particular cohort. Rather, mortality rates used to calculate LE are based on country- and year-specific death rates. For a specific risk profile, the estimated LE should be interpreted as conditional on the assumption that age/sex-specific mortality rates will remain constant (standard assumption for period life tables) [17].

Our model generally assumes that risk factors do not interact in their effect on RR, except for mediation effects. We did not incorporate interactions between risk factors because the required data were not available from the GBD and the knowledge base to justify this is not sufficiently robust at this time [32]. However, our model does allow for RRs to vary according to age and sex. Furthermore, LE effects of risk factors depend in a complex way on the levels of other factors because of competing risks of death from different diseases. Our preference for the additive model of independence in the current study is supported by both theoretical considerations and empirical observations. From a theoretical perspective, independence defined as a lack of biological synergy or antagonism is consistent with a lack of interaction on the additive scale [19]. Empirically, our results for subjects with multiple risk factors suggest plausible effects on LE under the additive model. Moreover, all RRs that vary by age in the GBD data show a sub-multiplicative relationship with age [18,27].

To accommodate mediation effects, we applied two models. In estimating the individual (total) effect of each factor, without specifying other factors, we used a model without mediation. For example, the effect of BMI in Table 3 includes the direct effect as well as the effect mediated by SBP, LDL, fasting plasma glucose, and impaired kidney function. However, in estimating the combined effect of multiple factors (including mediators), we used the mediation-adjusted model and assumed that all other factors are kept at their specified or (if not specified) mean levels. For example, the BMI effect in Table 5 is the direct effect only and, therefore, is substantially smaller than the total effect.

### 4.2. Comparisons with Published Results

Our results can be compared with published LE effects of selected exposures from cohort-based models. In one of the first studies of the impact of smoking [33], LE of heavy smokers was 6-7 years lower than that of non-smokers. In subsequent studies, a range of values have been reported [34-36]. Doll et al. reported a 10-year difference in median survival age between heavy smokers and non-smokers in male British doctors [10]. Our results are consistent with data from cohort studies, except that the effect of smoking on LE in CHARM is stronger in females compared with males. A possible reason for the failure of cohort-based models to show this effect is a lack of adjustment for a lower average number of cigarettes smoked by female smokers compared with male smokers [37].

The association of diet with life expectancy and all-cause mortality has been controversial [38]. Implausible claims for the effects of some foods have been published, suggesting, for example, that consumption of nuts may increase life expectancy by several years [12]. Our results are more plausible. According to CHARM, given the current levels of consumption, removing all dietary risks would increase LE0 by about 2 years. However, consuming large amounts (5-10 times the average) of sugar or processed meat would reduce LE0 by more than a year, and 300 g of fruit a day (compared to none) would improve LE0 by close to a year. A similar effect would be expected from a very high intake of sodium ( $5-7 \mathrm{~g}$ per day), compared with the optimal intake. Given the well-known association between sodium and blood pressure, this estimate seems realistic and agrees with the literature [39].

Previous research on physical activity (PA) found a strong impact on all-cause mortality, but few studies provided data on LE [11,40]. Lim et al. reported LE gains of 1.27 years for males and 1.39 for females from eliminating low PA in the USA [14]. In our model, the effect of PA was smaller and was most pronounced at low levels of activity. It is possible that CHARM, which is based on a meta-analysis of published, confounder-adjusted associations of PA with 5 diseases (IHD, stroke, diabetes, colon cancer, and breast cancer) rather than all-cause mortality, underestimates the impact of PA if a larger number of diseases are associated with PA. However, randomized trials thus far do not support a strong association of PA with all-cause mortality [41].

### 4.3. Comparisons with Other Models

Additional empirical evidence of CHARM's validity is provided in Figures S2 and S3 (Supplementary Materials). These data show that the results from CHARM with respect to the effect of body mass index and smoking on LE broadly agree with those from the Big Life calculator. When discrepancies do occur, the results from CHARM seem equally or more plausible. For example, among males, smoking 40 cigarettes vs. none results in an LE40 difference of 7.83 years according to CHARM and approximately 7 years according to Big Life. However, the dose-response curve in Big Life is relatively flat beyond 5 cigarettes a day, whereas CHARM shows a continuous, albeit diminishing, increase in effect for heavier smokers. Furthermore, the Big Life calculator does not show a clear difference in the impact of smoking on LE between males and females, resulting in some discrepancy between the two models for heavy female smokers.

Differences in the estimates of risk of death from CVD between CHARM and SCORE were generally smaller than differences observed among established cohort-based mod-
els [42]. On average, CHARM estimates tended to be slightly higher for non-smokers, especially males with an extremely high SBP, and slightly lower for smokers. These differences could be due to methodological differences between the two models or actual differences in the joint effects of SBP, cholesterol, and smoking on CVD mortality between the GBD data and the European cohorts from which the SCORE model was derived.

### 4.4. Strengths and Limitations

CHARM is a comprehensive, interactive risk and life expectancy model that provides estimates for a large number of diseases (270) and risk factors (60). Previously, Lim and colleagues developed a multi-disease mortality risk calculator using GBD data, but their model included only 12 risk factors [14].

Unlike most published risk calculators based on statistical prediction models developed from cohort studies, CHARM is a synthetic model that uses all available, up-to-date data from a wide range of sources, collected, processed, and analyzed by the GBD. Cohortbased models are rarely fully representative of the underlying geographical population. Furthermore, parameters from a single cohort study are usually considered less reliable and less generalizable than those from a meta-analysis of many studies. Moreover, data from cohort studies are limited to a set of variables measured in those studies (may not include all important risk factors and confounders), may not collect data repeatedly over time, and may not be publicly available, which makes the model less transparent and verifiable. In addition, models developed from a cohort study may be difficult to update if the study has been discontinued or extensive new data collection and analyses are needed. In contrast, the relative risks and dose-response relationships underlying our estimates are publicly available, easily updatable, and based on meta-analyses of epidemiological studies carefully evaluated for potential biases and adjusted for confounding and mediation [18,27]. The model has been shown to provide plausible, internally consistent results that generally agree with published data and other models.

All results presented in this paper are based on mortality and exposure data for Canada. Therefore, our LE estimates are not strictly applicable to other countries. However, since the relative risk functions are considered by the GBD to be the same for all countries, the relative effects on LE should be generalizable to other populations, especially other high-income countries. Caution is required when applying our data to low and middle income countries because the impact of risk factors on mortality may depend on access to health care. Nonetheless, replacing current parameters with a parameter database for any other country would be straightforward.

A limitation of our model is that data for some factors that affect LE were not available. Specifically, we did not have risk functions for social determinants of health, genetic factors, and pre-existing conditions, as well as some environmental factors (e.g., temperature). Although these factors are generally unmodifiable, their inclusion in future versions of the model may improve the accuracy of predictions. A limitation of any model are potential inaccuracies in the data-derived parameters. This applies mainly to relative risk functions, derived from epidemiological literature, and the lag parameters. Some parameters derived from studies elsewhere may not apply to Canada. There may also be inaccuracies in exposure distribution data. Although uncertainty intervals for these parameters are available from the GBD, our model currently does not provide uncertainty intervals for LE or risk estimates. Therefore, similar to data from other published models, our estimates should be regarded as approximations and treated with caution. Finally, while LE is an important indicator of population health, a measure that reflects both quantity and quality of life, such as health-adjusted life expectancy (HALE) [26], would also be informative and could be considered in future studies.

### 4.5. Practical Applications

Prior studies of the impact of risk calculators on the uptake of preventive measures by individuals have produced mixed results [7,43-45]. This may be due to the kind of
information the calculator is providing and the way the results are presented [7,45]. When considering preventive interventions at the individual level, the results presented in this report can be useful. For example, data in Tables $2-4$ can be used to determine which factors are most important to consider for a given patient and to discuss the potential impact and feasibility of risk factor modifications. However, as our data show, the impact of an intervention on LE depends on the risk factor level, the amount of change, patient age, and other, coexisting risks. Usually, there are many realistic combinations of risk factor modifications that result in similar LE gains. A change in a strong factor may accomplish the same result as multiple changes in weaker factors. Therefore, the objective should be to find the intervention that is most effective and suitable given the patient's risk profile, preferences and goals. To this end, the best approach may be to use our model interactively and compare LE estimates for several changes in the risk profile. For example, for a female with a risk profile considered in Table 3, reducing smoking by 2 cigarettes a day (from 10 to 8), reducing BMI by 5 units ( 35 to 30 ), or reducing SBP by 13 mm Hg ( 150 to 137) result in a similar gain in LE of about 0.5 years.

## 5. Conclusions

Personalized prevention has been advocated for many years but despite significant progress in risk prediction for specific diseases, the data and tools needed to support large-scale applications of risk models across all diseases and risk factors have been lacking. In this report we estimated the effects of 51 risk factors on LE, derived from a new model that utilizes GBD data on mortality, risk factor distribution, and relative risk functions for 270 diseases. As such, the model is a synthesis of current epidemiological knowledge in a format suitable for immediate implementation. Limitations of the model notwithstanding, we hope this tool, and the data it generates, can assist patients and their doctors in making difficult choices between different options for risk factor modification.

Supplementary Materials: The following supporting information can be downloaded at: https: / /www.mdpi.com/article/10.3390/ijerph19158958/s1. Section S1: Sources of data; Section S2: User interface, Section S3: Mathematical description of the model; Table S1: List of diseases included in the model in alphabetical order; Table S2: Exposure categories for categorical risk factors; Figure S1: Survival probability (\%) for males (A) and females (B) with specified levels of 10 risk factors, according to age at which all risk factors are improved, adjusted for mediation; Figure S2: Comparison of the effect of smoking and body mass index on life expectancy in males and females 40 and 60 years of age, according to CHARM and the Big Life calculator (non-mediated model); and Figure S3: Comparison of CVD mortality over 10 years based on CHARM and the European SCORE chart for a 60-year-old male and female (mediated model).

Author Contributions: Conceptualization, J.A.K., E.C.S., H.X., L.C.L. and J.M.E.; Methodology, J.A.K., E.C.S. and H.X., Software: E.C.S.; Formal Analysis, E.C.S.; Data Curation, E.C.S. and B.S.; Writing-Original Draft Preparation, J.A.K. and E.C.S.; Writing—Review and Editing, J.A.K., E.C.S., B.S., L.C.L., H.X., L.M.F. and J.M.E.; Visualization, J.A.K. and E.C.S.; Supervision, J.A.K. and J.M.E.; Project Administration, J.A.K. and J.M.E.; Funding Acquisition, J.A.K., E.C.S., B.S., L.C.L., H.X., L.M.F. and J.M.E. All authors have read and agreed to the published version of the manuscript.

Funding: This work was supported by a grant from the Canadian Institutes of Health Research (grant number FRN 155556). The funders played no role in the design of the study, drafting the manuscript, or decision to publish.

Institutional Review Board Statement: Ethical review and approval were waived for this study because the study was based on existing, publicly available aggregate data, and did not involve any collection of new data on humans.

Informed Consent Statement: Informed consent was waived for this study because the study was based on existing, publicly available aggregate data, and did not involve any collection of new data on humans.

Data Availability Statement: GBD data used in this study are available from the Global Health Data Exchange, Institute for Health Metrics and Evaluation. Internet address: http:/ / ghdx.healthdata.org/, accessed on 10 January 2021.

Acknowledgments: The authors wish to thank Kara Estep for assistance in obtaining data from the GBD Study. All inferences, opinions, and conclusions drawn in this study are those of the authors and do not necessarily reflect the opinions or policies of the Institute for Health Metrics and Evaluation or the GBD Study.

Conflicts of Interest: The authors declare no conflict of interests.

## References

1. Wilson, P.W.; D'Agostino, R.B.; Levy, D.; Belanger, A.M.; Silbershatz, H.; Kannel, W.B. Prediction of coronary heart disease using risk factor categories. Circulation 1998, 97, 1837-1847. [CrossRef] [PubMed]
2. Gail, M.H.; Brinton, L.A.; Byar, D.P.; Corle, D.K.; Green, S.B.; Schairer, C.; Mulvihill, J.J. Projecting individualized probabilities of developing breast cancer for white females who are being examined annually. J. Natl. Cancer Inst. 1989, 81, 1879-1886. [CrossRef] [PubMed]
3. Gray, E.P.; Teare, M.D.; Stevens, J.; Archer, R. Risk prediction models for lung cancer: A systematic review. Clin. Lung Cancer 2016, 17,95-106. [CrossRef] [PubMed]
4. Stiglic, G.; Pajnkihar, M. Evaluation of major online diabetes risk calculators and computerized predictive models. PLoS ONE 2015, 10, e0142827. [CrossRef]
5. Ross, P.D.; Wasnich, R.D.; Davis, J.W.; Jameson, J.L.; Longo, D.L. Fracture prediction models for osteoporosis prevention. Bone 1990, 11, 327-331. [CrossRef]
6. Jameson, J.L.; Longo, D.L. Precision medicine—Personalized, problematic, and promising. N. Engl. J. Med. 2015, 372, 2229-2234. [CrossRef]
7. Müller-Riemenschneider, F.; Holmberg, C.; Rieckmann, N.; Kliems, H.; Rufer, V.; Müller-Nordhorn, J.; Willich, S.N. Barriers to routine risk-score use for healthy primary care patients: Survey and qualitative study. Arch. Intern. Med. 2010, 170, 719-724. [CrossRef]
8. Manuel, D.G.; Perez, R.; Sanmartin, C.; Taljaard, M.; Hennessy, D.; Wilson, K.; Tanuseputro, P.; Manson, H.; Bennett, C.; Tuna, M.; et al. Measuring burden of unhealthy behaviours using a multivariable predictive approach: Life expectancy lost in Canada attributable to smoking, alcohol, physical inactivity, and diet. PLoS Med. 2016, 13, e1002082. [CrossRef]
9. Global Burden of Disease (GBD); Institute for Health Metrics and Evaluation (IHME). GBD Compare. Viz Hub. Available online: https:/ /vizhub.healthdata.org/gbd-compare/ (accessed on 18 April 2022).
10. Doll, R.; Peto, R.; Boreham, J.; Sutherland, I. Mortality in relation to smoking: 50 years' observations on male British doctors. BMJ 2004, 328, 1519. [CrossRef]
11. Blond, K.; Brinkløv, C.F.; Ried-Larsen, M.; Crippa, A.; Grøntved, A. Association of high amounts of physical activity with mortality risk: A systematic review and meta-analysis. Br. J. Sports Med. 2020, 54, 1195-1201. [CrossRef]
12. Schwingshackl, L.; Schwedhelm, C.; Hoffmann, G.; Lampousi, A.M.; Knüppel, S.; Iqbal, K.; Bechthold, A.; Schlesinger, S.; Boeing, H. Food groups and risk of all-cause mortality: A systematic review and meta-analysis of prospective studies. Am. J. Clin. Nutr. 2017, 105, 1462-1473. [CrossRef]
13. Clarke, R.; Emberson, J.; Fletcher, A.; Breeze, E.; Marmot, M.; Shipley, M.J. Life expectancy in relation to cardiovascular risk factors: 38 year follow-up of 19,000 men in the Whitehall study. BMJ 2009, 339, b3513. [CrossRef] [PubMed]
14. Lim, S.S.; Carnahan, E.; Nelson, E.C.; Gillespie, C.W.; Mokdad, A.H.; Murray, C.J.; Fisher, E.S. Validation of a new predictive risk model: Measuring the impact of the major modifiable risks of death for patients and populations. Popul. Health Metr. $2015,13,27$. [CrossRef]
15. Institute of Health Metrics and Evaluation (IHME). About GBD. Available online: https:/ /www.healthdata.org/gbd/about.http: / /www.healthdata.org/gbd/about (accessed on 18 April 2022).
16. Kopec, J.; Sayre, E.; Shams, B.; Li, L.; Xie, H.; Feehan, L.; Esdaile, J. Development and preliminary validation of a novel risk and life expectancy calculator. In Proceedings of the 41st Conference of International Society for Clinical Biostatistics, Krakow, Poland, 23-27 August 2020.
17. Siegel, J.S.; Swanson, D.A. The Methods and Materials of Demography, 2nd ed.; Elsevier Academic Press: Amsterdam, The Netherlands, 2004.
18. GBD 2019 Risk Factors Collaborators. Global burden of 87 risk factors in 204 countries and territories, 1990-2019: A systematic analysis for the Global Burden of Disease Study 2019. Lancet 2020, 396, 1223-1249. [CrossRef]
19. Greenland, S.; Rothman, K. Concepts of Interaction. In Modern Epidemiology, 2nd ed.; Lippincot-Raven: New York, NY, USA, 1998; pp. 329-342.
20. Murray, C.J.L.; Lopez, A.D. (Eds.) Global Burden of Disease: A Comprehensive Assessment of Mortality and Disability from Diseases, Injuries, and Risk Factors in 1990 and Projected to 2020, 1st ed.; Global Burden of Disease and Injury Series; Harvard School of Public Health on behalf of the World Health Organization and the World Bank: Cambridge, MA, USA, 1996; Volume 1.
21. Murray, C.J.; Ezzati, M.; Flaxman, A.D.; Lim, S.; Lozano, R.; Michaud, C.; Naghavi, M.; Salomon, J.A.; Shibuya, K.; Vos, T.; et al. GBD 2010: Design, definitions, and metrics. Lancet 2012, 380, 2063-2066. [CrossRef]
22. GBD 2019 Diseases and Injuries Collaborators. Global burden of 369 diseases and injuries in 204 countries and territories, 1990-2019: A systematic analysis for the Global Burden of Disease Study 2019. Lancet 2020, 396, 1204-1222. [CrossRef]
23. GBD 2017 Causes of Death Collaborators. Global, regional, and national age-sex-specific mortality for 282 causes of death in 195 countries and territories, 1980-2017: A systematic analysis for the Global Burden of Disease Study 2017. Lancet 2018, 392, 1736-1788. [CrossRef]
24. GBD 2017 Mortality Collaborators. Global, regional, and national age-sex-specific mortality and life expectancy, 1950-2017: A systematic analysis for the Global Burden of Disease Study 2017. Lancet 2018, 392, 1684-1735. [CrossRef]
25. GBD 2017 Disease and Injury Incidence and Prevalence Collaborators. Global, regional, and national incidence, prevalence, and years lived with disability for 354 diseases and injuries for 195 countries and territories, 1990-2017: A systematic analysis for the Global Burden of Disease Study 2017. Lancet 2018, 392, 1789-1858. [CrossRef]
26. GBD 2017 DALYs and HALE Collaborators. Global, regional, and national disability-adjusted life-years (DALYs) for 359 diseases and injuries and healthy life expectancy (HALE) for 195 countries and territories, 1990-2017: A systematic analysis for the Global Burden of Disease Study 2017. Lancet 2018, 392, 1859-1922. [CrossRef]
27. GBD 2017 Risk Factor Collaborators. Global, regional, and national comparative risk assessment of 84 behavioural, environmental and occupational, and metabolic risks or clusters of risks for 195 countries and territories, 1990-2017: A systematic analysis for the Global Burden of Disease Study 2017. Lancet 2018, 392, 1923-1994.
28. Global Health Data Exchange-GHDx. Available online: http://ghdx.healthdata.org/ (accessed on 10 January 2021).
29. Project Big Life: Life Expectancy Calculator. Available online: https:/ /www.projectbiglife.ca/life-expectancy-calculator (accessed on 17 September 2020).
30. SCORE: European Low Risk Chart. Available online: https://www.escardio.org/static-file/Escardio/Subspecialty/EACPR/ Documents/score-charts.pdf (accessed on 17 September 2020).
31. Conroy, R.M.; Pyörälä, K.; Fitzgerald, A.P.; Sans, S.; Menotti, A.; De Backer, G.; De Bacquer, D.; Ducimetière, P.; Jousilahti, P.; Keil, U.; et al. Estimation of ten-year risk of fatal cardiovascular disease in Europe: The SCORE project. Eur. Heart J. 2003, 24, 987-1003. [CrossRef]
32. Lubin, J.H.; Couper, D.; Lutsey, P.L.; Yatsuya, H. Synergistic and non-synergistic associations for cigarette smoking and nontobacco risk factors for cardiovascular disease incidence in the Atherosclerosis Risk In Communities (ARIC) study. Nicotine Tob. Res. 2017, 19, 826-835. [CrossRef] [PubMed]
33. Haybittle, J.L. Cigarette smoking and life expectancy. Br. J. Prev. Soc. Med. 1966, 20, 101-104. [CrossRef] [PubMed]
34. Ozasa, K.; Katanoda, K.; Tamakoshi, A.; Sato, H.; Tajima, K.; Suzuki, T.; Tsugane, S.; Sobue, T. Reduced life expectancy due to smoking in large-scale cohort studies in Japan. J. Epidemiol. 2008, 18, 111-118. [CrossRef]
35. Brønnum-Hansen, H.; Juel, K. Abstention from smoking extends life and compresses morbidity: A population based study of health expectancy among smokers and never smokers in Denmark. Tob. Control 2001, 10, 273-278. [CrossRef]
36. Li, K.; Hüsing, A.; Kaaks, R. Lifestyle risk factors and residual life expectancy at age 40: A German cohort study. BMC Med. 2014, 12, 59. [CrossRef]
37. Huxley, R.R.; Woodward, M. Cigarette smoking as a risk factor for coronary heart disease in women compared with men: A systematic review and meta-analysis of prospective cohort studies. Lancet 2011, 378, 1297-1305. [CrossRef]
38. Ioannidis, J.P.A. The challenge of reforming nutritional epidemiologic research. JAMA 2018, 320, 969-970. [CrossRef]
39. Bibbins-Domingo, K.; Chertow, G.M.; Coxson, P.G.; Moran, A.; Lightwood, J.M.; Pletcher, M.J.; Goldman, L. Projected effect of dietary salt reductions on future cardiovascular disease. N. Engl. J. Med. 2010, 362, 590-599. [CrossRef]
40. Dhana, K.; Koolhaas, C.M.; Berghout, M.A.; Peeters, A.; Ikram, M.A.; Tiemeier, H.; Hofman, A.; Nusselder, W.; Franco, O.H. Physical activity types and life expectancy with and without cardiovascular disease: The Rotterdam Study. J. Public Health 2017, 39, e209-e218. [CrossRef] [PubMed]
41. Kujala, U.M. Is physical activity a cause of longevity? It is not as straightforward as some would believe. A critical analysis. Br. J. Sports Med. 2018, 52, 914-918. [CrossRef] [PubMed]
42. Allan, M.G.; Nouri, F.; Korownyk, C.; Kolber, M.R.; Vandermeer, B.; McCormack, J. Agreement among cardiovascular disease risk calculators. Circulation 2013, 127, 1948-1956. [CrossRef] [PubMed]
43. Karmali, K.N.; Persell, S.D.; Perel, P.; Lloyd-Jones, D.M.; Berendsen, M.A.; Huffman, M.D. Risk scoring for the primary prevention of cardiovascular disease. Cochrane Database Syst. Rev. 2017, 3, CD006887. [PubMed]
44. Lloyd-Jones, D.M.; Braun, L.T.; Ndumele, C.E.; Smith, S.C., Jr.; Sperling, L.S.; Virani, S.S.; Blumenthal, R.S. Use of risk assessment tools to guide decision-making in the primary prevention of atherosclerotic cardiovascular disease: A special report from the American Heart Association and American College of Cardiology. Circulation 2019, 139, e1162-e1177. [CrossRef] [PubMed]
45. Manuel, D.G.; Abdulaziz, K.E.; Perez, R.; Beach, S.; Bennett, C. Personalized risk communication for personalized risk assessment: Real world assessment of knowledge and motivation for six mortality risk measures from an online life expectancy calculator. Inform. Health Soc. Care. 2018, 43, 42-55. [CrossRef] [PubMed]
