# **ORIGINAL ARTICLE**

**Open Access** 





Takashi Nakagata<sup>1,2,3\*</sup>, Tsukasa Yoshida<sup>1,3</sup>, Daiki Watanabe<sup>1,3,4</sup>, Yukako Arishima-Hashii<sup>1</sup>, Yosuke Yamada<sup>1,3</sup>, Naomi Sawada<sup>5</sup>, Hidekazu Shimada<sup>6</sup>, Nobuo Nishi<sup>1</sup> and Motohiko Miyachi<sup>1,4</sup>

#### **Abstract**

**Background:** Weight misperception adversely affects health-related quality of life (HRQoI); however, few studies have evaluated the relationship between weight misperception and muscle mass. This study aimed to examine the relationship of weight misperception with low muscle mass using skeletal muscle index (SMI) estimated by multifrequency bioelectrical impedance analysis (MF-BIA) among community-dwelling Japanese.

**Methods:** Participants were 525 Japanese individuals aged 40–91 years old (male 89, female 436). Misperception was calculated by subtracting measured value from self-reported weight, presented as a percentage and categorized into tertiles based on sex (under-reporters, acceptable reporters, and over-reporters). Appendicular lean mass was estimated using MF-BIA, and low muscle mass was defined using SMI values of 7.0 and 5.7 kg/m² for males and females, respectively, based on the Asian Working Group for Sarcopenia 2019 consensus. We evaluated the association between prevalence of low muscle mass and weight misperception (under-reporters and over-reporters) using multivariate logistic regression including covariate.

**Results:** In total, 9.3% (49/525) of participants had low muscle mass. After adjusting for covariates, prevalence of low muscle mass was higher among over-reporters than acceptable-reporters (odds ratio [OR]; 2.37, 95% confidence interval [CI]; 1.03–5.44). Additionally, sensitivity analysis was performed on females, which confirmed that the prevalence of low muscle mass was higher in over-reporters than in acceptable-reporters (*OR*, 3.27; 95% *CI*, 1.18–9.12).

**Conclusion:** Weight misperception was significantly correlated with low muscle mass, especially in over-reporters.

Keywords: Aging, Health checkup, Sarcopenia, Self-weighing, Weight perception

# **Background**

Misperceptions of height and weight have been defined as the difference between one's self-perceived and objectively measured height and weight, respectively [1, 2]. Misperceptions of height and weight, which include under-reporting (objectively measured > self-reported) and over-reporting (objectively measured < self-reported), may be associated with variations in age, sex, body mass index (BMI), nationalities, ethnicity, and socioeconomic status [3, 4]. According to substantial evidence, weight misperception has adverse effects on physical and psychological health [5, 6] and may negatively

Full list of author information is available at the end of the article



© The Author(s) 2022. **Open Access** This article is licensed under a Creative Commons Attribution 4.0 International License, which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons licence, and indicate if changes were made. The images or other third partial in this article are included in the article's Creative Commons licence, unless indicated otherwise in a credit line to the material. If material is not included in the article's Creative Commons licence and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this licence, visit http://creativecommons.org/licenses/by/4.0/. The Creative Commons Public Domain Dedication waiver (http://creativecommons.org/publicdomain/zero/1.0/) applies to the data made available in this article, unless otherwise stated in a credit line to the data.

<sup>\*</sup>Correspondence: takashi.nakagata@gmail.com

<sup>&</sup>lt;sup>1</sup> National Institute of Health and Nutrition, National Institutes of Biomedical Innovation, Health and Nutrition, 1-23-1 Toyama, Tokyo 162-8636 Shinjuku-ku, Japan

affect health outcomes. A previous study examined the association between weight misperception and health-related quality of life (HRQoL) among adolescents and adult populations. Weight misperception, in particular, the underestimation or over-reporting of weight, has been shown to be significantly associated with low HRQoL, after adjusting for the influences of age, chronic disease, and socioeconomic status [2]. Furthermore, weight misperception is related to more unhealthy and fewer healthy weight control behaviors [7].

Age-related loss of skeletal muscle mass and strength, which are associated with functional impairment and physical disabilities in elderly people [8], are key components of sarcopenia and frailty. As previously mentioned, various studies examined the association between weight misperception and health outcomes in a wide range of population [2, 5, 6]. However, few studies have examined the association between weight misperception and health outcomes related to skeletal muscle mass and strength. A previous study investigated the relationship of perceived weight status and muscle strength measured by hand grip strength conducted among 12,727 adults (mean age,  $51.0 \pm 16.6$  years) using the Korea National Health and Nutrition Examination Survey (KNHANES) [9]. It demonstrated that underreporters of weight status had higher odds of sarcopenia than accurate-reporters and over-reporters of weight status. Although mechanisms underlying the observed relationships between weight-related behaviors and hand grip strength are unclear, having correct perceptions of one's own weight may not only be associated with HRQoL but also skeletal muscle mass.

Therefore, the present study aimed to examine the association between weight misperception and skeletal muscle mass, with the prevalence of low muscle mass measured using skeletal muscle index (SMI) values based on the Asian Working Group for Sarcopenia 2019 consensus.

#### **Methods**

# Study design and participants

From July 1 to July 9, 2019, we conducted face-to-face surveys with 539 individuals (female = 447, male = 92), who were aged  $\geq$  40 years and resided in Settsu city, Osaka prefecture. The participants were recruited either when they participated in a specific health examination conducted at the Settsu Health Center or via Settsu's public relations magazine and word of mouth. Details of study design and participants are described in our previous publication [10]. Of these participants, we excluded participants who lived outside Settsu ity (n = 10), missed skeletal muscle mass data (n = 2), refused to participate (n = 1), and were aged less than 40 (n = 1)

1). Ultimately, this cross-sectional study included 525 adults (436 females and 89 males) aged 40 to 91. Before the study began, participants provided written consent to participate after receiving information about the procedures and purpose of the study. The study protocol was approved by the Ethics Review Board of the National Institute of Health and Nutrition (Ikikenhatsu 178-1). This study was carried out in accordance with the principles outlined in the Declaration of Helsinki and was registered with the Japan Clinical Trials Registry (UMIN000036880).

#### **Procedures**

Each participant was asked to visit a public building in Settsu city once between 10 AM and 4 PM and complete questionnaires on daily life (e.g., about dwelling and physical activity) before collecting each one's objective weight and height measurements. To obtain self-reported body weight and height measurements, participants were asked, "What is your current body weight and height?" in self-administered questionnaires. They filled out self-reported height and weight using 0.1 cm and 0.1 kg units, respectively. Other health-related variables, which include smoking habit, drinking habit, exercise habit, self-rated health, self-rated physical fitness, and self-rated economic condition, were assessed using a self-administered questionnaire. Details of these questionnaires are described in our previous study [10]. The heights of the participants were then objectively measured to the closest 0.1 cm using an analog height meter. Body weights were measured, and appendicular lean mass (ALM) was estimated using the multifrequency bioelectrical impedance analysis (MF-BIA) (MC-780A-N, TANITA, Tokyo, Japan), which were then validated using dual-energy X-ray absorptiometry (DXA) [11]. Participants were evaluated in their underwear and were asked to stand barefoot on toe-and-heel electrodes while holding handgrips, with their arms hanging down a few centimeters from the hips. The details of the MF-BIA device used were previously described. All the procedures were conducted in July 2019.

# Skeletal muscle mass measurements and definition of low muscle mass

We calculated the skeletal muscle index (SMI) as follows:  $SMI = ALM/height^2$ . In this study, participants with lower SMI values for each sex were categorized as having "sarcopenia with low skeletal muscle mass." According to the Asian Working Group for Sarcopenia, the cutoff SMI values using the bioelectrical impedance analysis (BIA) method were 7.0 and 5.7 kg/m² in male and female, respectively [12].

# Weight misperception

Weight misperception was defined as the difference between self-reported and objectively measured values of weight and calculated as follows in our study. First, we calculated the difference of weight (kg) by subtracting objectively measured values from self-reported values of weight based on Ikeda's study [4]. Next, to calculate the difference (%) against the objectively measured value, the difference obtained was then divided by the objectively measured value and multiplied by 100 (represented as %). For example, weight misperception of an individual with self-reported weight = 65.0 kg and objectively measured weight = 66.0 kg can be calculated as follows:  $-1.0 \text{ kg} \div 66.0 \text{ kg} \times 100 \% = -1.51\%$ .

# Statistical analysis

For descriptive statistics, values of continuous variables were expressed as mean and standard deviation, or as median and interquartile range, while values of categorical variables were expressed as percentages (%). Continuous and ordinal participant characteristics were classified as follows: exercise habit ("Do you go walking or engage in other exercise at least once per week?" "yes" or "no"), smoking habit ("almost daily" or "sometimes" = yes, "used to, but quit," or "never" = no), drinking habit ("almost daily," or "sometimes," = yes, and "almost never" or "never" = no); self-reported health ("very healthy" or "somewhat healthy" = good self-reported health, and "not very healthy" or "unhealthy" = poor self-reported health); self-reported physical fitness ("extremely confident" or "somewhat confident" = good self-reported physical fitness, and "slightly anxious" or "very anxious" = poor self-reported physical fitness); and self-reported socioeconomic status ("easy" or "somewhat easy" = high socioeconomic status, and "somewhat hard" or "hard" = low socioeconomic status). These variables were classified with reference to covariates used in our previous study [13], and these models were decided with reference to covariates used in previous studies, which examined the association between weight misperception on HRQoL [2].

Second, participants were first classified into sex-specific tertiles based on extent of weight misperception (under-reporters, acceptable reporters, and over-reporters) [14]. In order to evaluate the relationship between weight misperception and prevalence of low muscle mass, we used a logistic regression model to calculate the sex, age, and multivariable adjusted ORs, as well as 95% CI. In the logistic regression model, sex and age were inputted as covariates to calculate the adjusted ORs. To calculate the multivariable-adjusted ORs, continuous variables (i.e., sex, age, and BMI) were inputted into the logistic regression model as covariates

(model 1). Then, exercise habit, smoking, drinking habit, self-reported health, self-reported physical fitness, and self-reported socioeconomic status were added to model 1 to calculate the multivariable-adjusted OR (model 2). Furthermore, linear regression analyses were conducted to obtain the correlation between the means of self-reported and objectively measured height, weight, and BMI. We used Microsoft Office Excel 2017 and PASW Statistics version 20.0 (SPSS, IBM Corp., Armonk, N.Y., USA) for data processing and statistical analyses, respectively. A two-tailed *p*-value lower than 5% was considered statistically significant.

#### Results

Table 1 shows the characteristics of the participants including skeletal muscle mass and health-related information. In our study, the median weight misperception (interquartile range) was 0.9 (-0.3, 2.0)%, while the prevalence of low muscle mass using SMI was 9.3% (49/525).

Figure 1 shows the scatter plots and correlation between the means of self-reported and objectively measured height, weight, and BMI. All variables showed a strong positive correlation between self-reported and objectively measured characteristics (r = 0.96-0.99); however, there were some outliers from the identical line.

Table 2 shows the characteristics of the participants with weight misperceptions. **Participants** divided into three groups based on sex-specific weight misperception tertiles: < 0.4%, 0.4–1.7%, and  $\ge 1.7\%$  in female and < -0.4%, -0.3-1.1%, and  $\ge 1.1\%$  in male, respectively. The median difference between selfreported weight and objectively-measured weight among under-reporters, accurate-reporters, and over-reporters were -0.8, 0.9, and 2.4%, respectively. We found no significant difference in age and height among underreporters, accurate-reporters, and over-reporters. Results show that over-reporters have lower body weight, BMI, and SMI.

Table 3 shows the multivariate-adjusted ORs of the prevalence of low muscle mass by tertiles of weight misperception. After adjusting for sex, age, and BMI (model 1), the OR for the prevalence of low muscle mass was higher (2.01 [95% CI, 0.90–4.47]) in the overreporting group than in the acceptable-reporting group. After adjusting for model 1 and other variables (model 2), the OR was higher (2.37 [95% CI, 1.03–5.44]) in the overreporting group than in the acceptable-reporting group. A total of 436 females were included in the sensitivity analysis after excluding males because of low statistical power. The results of this analysis confirmed the results from the main analysis (3.27 [95% CI, 1.18–9.12], Table 4).

**Table 1** Characteristics of participants

	•	•	
Variables	All (n = 525)	Female (n = 436)	Male (n = 89
Age, years	72 (67, 77)	72 (67, 77)	74 (69, 78)
Measured height, cm	$154.3 \pm 7.6$	$152.1 \pm 5.8$	$164.9 \pm 6.1$
Measured weight, kg	$53.7 \pm 9.2$	$51.9 \pm 8.3$	$62.3 \pm 8.3$
Measured BMI, kg/ m <sup>2</sup>	$22.5 \pm 3.2$	$22.4 \pm 3.3$	$22.9 \pm 2.7$
Weight misperception, kg <sup>a</sup>	0.5 (-0.2, 1.0)	0.5 (-0.1, 1.0)	0.3 (-0.6, 0.8)
Weight misperception, % <sup>a</sup>	0.9 (-0.3, 2.0)	1.0 (-0.1, 2.1)	0.5 -0.9, 1.3)
%fat, %	$29.6 \pm 7.9$	$31.2 \pm 7.2$	$21.6 \pm 5.8$
ALM, kg	$15.9 \pm 3.2$	$14.9 \pm 1.9$	$21.0 \pm 3.3$
SMI, kg/m <sup>2</sup>	$6.6 \pm 0.8$	$6.4 \pm 0.5$	$7.7 \pm 0.9$
Low muscle mass, n (%)	49 (9.3%)	33 (7.6%)	16 (18.0%)
Exercise habit, n (%)			
Yes	396 (75.4%)	329 (75.5%)	67 (75.3%)
No	129 (24.6%)	107 (24.5%)	22 (24.7%)
Smoking habit, n (%)			
No	507 (96.6%)	429 (98.4%)	78 (87.6%)
Yes	18 (3.4%)	7 (1.6%)	11 (12.4%)
Drinking habit, n (%)			
No	337 (64.2%)	310 (71.1%)	27 (30.3%)
Yes	118 (22.5%)	126 (28.9%)	62 (69.7%)
Self-reported health, r	n (%)		
Good	440 (83.8%)	367 (84.2%)	73 (82.0%)
Poor	85 (16.2%)	69 (15.8%)	16 (18.0%)
Self-reported physical	fitness, n (%)		
Good	269 (51.2%)	223 (51.1%)	46 (51.7%)
Poor	256 (48.8%)	213 (48.9%)	43 (48.3%)
Socioeconomic status	s, n (%)		
High	385 (73.5%)	324 (74.3%)	61 (69.3%)
Low	139 (26.5%)	112 (25.7%)	27 (30.7%)

Data are the means  $\pm\,\text{SD}$  or percentages. Age and weight misperceptions are median values (interquartile ranges)

BMI Body mass index, ALM Appendicular lean mass, SMI Skeletal muscle index

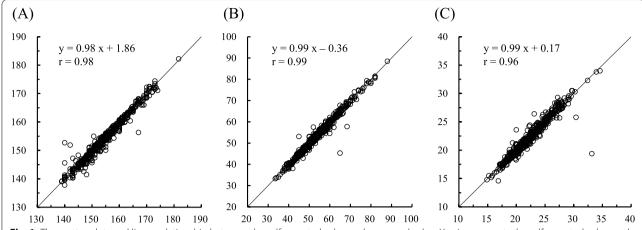
# Discussion

To our knowledge, this study is the first to evaluate the relationship between weight misperception and low muscle mass using MF-BIA in community-dwelling Japanese adults. Using MF-BIA-measured SMI, the present study indicates that participants who overreport body weight—self-reported weight was larger than objectively measured weight (median 0.9 %, [interquartile range (-0.3, 2.0)])—have higher odds of low muscle mass, as compared to the acceptable-reporting group (odds; 2.37 [95% CI, 1.03–5.44]).

Self-reported height and weight are often used for convenience in annual health checkups and large population surveys [15, 16], and previous studies have shown a high correlation between measured and selfreported height and weight among Japanese adults [4, 17-19]. Previous studies have been conducted focusing on the difference between measured and selfreported values and considered the associations among misperceptions of being "fat," HRQoL, and unhealthy weight control behaviors [2, 5-7]. For example, studies among Korean subjects who misperceived their weight were associated with significantly impaired HRQoL and sarcopenia based on hand grip strength [2, 9]. Weight misperception is also possibly associated with skeletal muscle mass; however, to the best of our knowledge, no relevant studies have been conducted. In our logistic analysis, misperception of body weight was significantly associated with low muscle mass using SMI; overreporting group had higher ORs for the prevalence of low muscle mass than the acceptable-reporting group (2.37 [95% CI, 1.03-5.44], Table 3). In addition, sensitivity analysis indicated that excluding males did not change that finding; the odds of low muscle mass in over-reporting groups was significantly higher than in acceptable-reporters (3.27 [95% CI, 1.18-9.12], Table 4).

Although the mechanism underlying misperception of weight-related low muscle mass is unclear in our crosssectional study, the possible factors for misperceiving body weight and their relationship to health outcomes have been discussed previously. One potential factor could be body status; BMI among over-reporters was lower than among acceptable and under-reporters in this study (under-reporters, 23.5  $\pm$  3.4; acc-reporters, 22.3  $\pm$ 2.9; and over-reports, 21.6  $\pm$  3.0). According to Ikeda's study on the characteristics of weight misperception among the general adult population in Japan based on the National Household Surveys, 1986, overweight and obese participants (BMI  $\geq$  25), both male and female, significantly under-reported their body weight as compared to their measured weight (range -5.9 kg to -1.3 kg) [4]. In contrast, underweight participants (BMI < 18.5) in that study over-reported their body weight (mean values, male; 2.4 kg, female; 1.1 kg) [4]. Previous studies showed that low BMI was associated with the prevalence of low muscle mass among Japanese populations [10, 20], and the association between overreporting body weight and low BMI might be related to low muscle mass in our study. Furthermore, low BMI may be associated with lower total energy intake per day. Body status has been linked to misreporting energy intake [21, 22], while being underweight ( $BMI < 18.5 \text{ kg/m}^2$ ) has been associated with over-reporting energy intake based on the ratio of energy intake per day to estimated energy

<sup>&</sup>lt;sup>a</sup> Calculated by subtracting measured values from self-reported values



**Fig. 1** The scatter plots and linear relationship between the self-reported value and measured value. X-axis represents the self-reported value, and Y axis represents the measured value. The solid black line represents the identity line. **A** height, **B** weight, and **C** body mass index

**Table 2** Physical characteristics by the results of weight misperception

Variables	Weight misperception		
	Under-reporters	acceptable-reporters	Over-reporters
Number (female/male)	174 (145/29)	176 (146/30)	175 (145/30)
Age, yrs	72 (67, 77)	72 (67, 76)	74 (68, 78)
Height, cm	$154.5 \pm 7.9$	$154.4 \pm 7.3$	$154.0 \pm 7.6$
Weight, kg	$56.2 \pm 9.4$	$53.4 \pm 8.7$	$51.4 \pm 8.9$
BMI, kg/m <sup>2</sup>	$23.5 \pm 3.4$	$22.3 \pm 2.9$	$21.6 \pm 3.0$
%fat, %	$31.5 \pm 8.3$	$28.9 \pm 7.7$	$28.3 \pm 7.3$
ALM, kg	$16.2 \pm 3.0$	$16.1 \pm 3.2$	$15.5 \pm 3.2$
SMI, kg/m <sup>2</sup>	$6.7 \pm 0.7$	$6.7 \pm 0.8$	$6.5 \pm 0.8$
Low muscle mass, n (%)	8 (4.6%)	12 (6.8%)	29 (16.6%)
Exercise habit, n (%)			
Yes	124 (71.3%)	134 (76.1%)	138 (78.9%)
No	50 (28.7%)	42 (23.9%)	37 (21.2%)
Smoking habit, n (%)			
No	166 (95.4%)	172 (97.7%)	169 (96.6%)
Yes	8 (4.6%)	4 (2.3%)	6 (3.4%)
Drinking habit, n (%)			
No	107 (61.5%)	119 (67.6%)	111 (63.4%)
Yes	67 (38.5%)	57 (32.4%)	64 (36.6%)
Self-reported health, n (%)			
Good	149 (85.6%)	143 (81.3%)	148 (84.6%)
Poor	25 (14.4%)	33 (18.8%)	27 (15.4%)
Self-reported physical fitness, n (%)			
Good	93 (53.4%)	87 (49.4%)	89 (50.9%)
Poor	81 (46.6%)	89 (50.6%)	86 (49.1%)
Socioeconomic status, n (%)			
High	121 (69.5%)	135 (76.7%)	129 (73.7%)
Low	52 (29.9%)	41 (23.3%)	46 (26.3%)

Data are the means  $\pm$  SD or percentages. Weight misperception and age are median (interquartile range). Sex-specific weight misperception tertiles in female: under-reporters, < 0.4 (%); acceptable-reporters, 0.4–1.7 (%); over-reporters,  $\geq$  1.7 (%). Sex-specific weight misperception tertiles in male: under-reporters, < 0.4 (%); acceptable-reporters, 0.3–1.1 (%); over-reporters,  $\geq$  1.1 (%)

BMI Body mass index, ALM Appendicular lean mass, SMI Skeletal muscle index

**Table 3** Multivariate-adjusted odds ratios and 95% confidence intervals for prevalence of low muscle mass by the results of weight misperception

	n	Number of low muscle mass	OR (95% CI)	Model 1 <i>OR</i> (95% <i>CI</i> )	Model 2 <i>OR</i> (95% <i>CI</i> )
Under-reporters	174	8 (4.6%)	0.66 (0.26–1.65)	0.76 (0.27–2.11)	0.97 (0.34–2.86)
Acceptable-reporters	176	12 (6.8%)	1.00 (reference)	1.00 (reference)	1.00 (reference)
Over-reporters	175	29 (16.6%)	2.72 (1.34-5.52)	2.01 (0.90-4.47)	2.37 (1.03-5.44)

Model 1, adjusted for sex, age, and BMI. Model 2, model 1 plus exercise habit, smoking habit, drinking habit, self-reported physical fitness, self-reported health, and self-reported socioeconomic status

CI Confidence interval, OR Odds ratio

**Table 4** Multivariate-adjusted odds ratios and 95% confidence intervals for prevalence of low muscle mass by the results of weight misperception in female

	n	Number of low muscle mass	OR (95% CI)	Model 1 <i>OR</i> (95% <i>Cl</i> )	Model 2 <i>OR</i> (95% <i>CI</i> )
Under-reporters	145	4 (2.8%)	0.56 (0.16–1.97)	0.84 (0.22–3.18)	1.25 (0.30–5.16)
Acceptable reporters	146	7 (4.8%)	1.00 (reference)	1.00 (reference)	1.00 (reference)
Over-reporters	145	22 (15.2%)	3.55 (1.47-8.60)	2.61 (1.00-6.81)	3.27 (1.18-9.12)

Model 1, adjusted for age and BMI. Model 2, model 1 plus exercise habit, smoking habit, drinking habit, self-reported physical fitness, self-reported health, and self-reported socioeconomic status

CI Confidence interval, OR Odds ratio

requirement in US adults using data from the National Health and Nutrition Examination Survey (NHANES) 2003–2012 [14]. Taking information from these previous studies [14, 21, 22], weight over-reporting may be associated with low muscle mass. Further studies are still needed to examine the relationship between weight misperception and energy intake. Second, individuals who misperceive their own body weight suggest that they may not have been weighed recently. Our study did not ask participants about whether they had concrete knowledge of their weight (e.g., do they own a scale, have a habit of self-weighing, or if they had been weighed at the doctor recently), although previous studies investigate the association between self-weighing frequency and healthy weigh-related behaviors. Houston et al. have examined the relationship of self-weighing frequency among the 4 quartiles (never, less than one time a week, one time a week, and several times a week or more) with BMI, physical activity, and other health-related measures among US adults and found that frequent self-weighing was associated with less sedentary time, more vigorous physical activity, and several healthier practices, including nutrition and healthy dietary habits in a cross-sectional study [23]. Furthermore, Zheng et al. examined the temporal patterns of self-weighing behavior and weight changes in free living over 52 weeks; participants who consistently self-weighed more than 6 days/week achieved higher weight loss compared to other participants [24]. Self-weighing or self-monitoring of weight, which are often recommended to prevent weight gain, to lose weight, or prevent weight regain [16, 25], could help increase their awareness of healthy weigh-related behaviors and may be linked to low muscle mass indirectly.

However, our study has several limitations. First, this was a cross-sectional study; therefore, further cohort studies are warranted to establish a relationship between low muscle mass and weight misperception. Second, the participants were not randomly selected from the city in our study; they may have been more health conscious than the general population, suggesting that selection bias might have been present. In our previous study, face-to-face participants had lower percentages of being an alcohol drinker or current smoker and a higher level of self-reported socioeconomic status than participants recruited by stratified random sampling in Settsu city [10]. Therefore, further research should be conducted with randomized sampling. Third, the sample size, especially for males, was small (n = 89); therefore, further studies with a larger sample size are required to confirm our data. Fourth, we investigated self-reported height and weight only once; therefore, we could not describe the reproducibility of weight misperception in our participants. Thus, further studies are needed to evaluate the reproducibility of the self-reported values. Finally, we did not measure muscle strength (e.g., handgrip strength) and gait speed; for screening and diagnosis of sarcopenia, both measurements were included.

# Conclusion

This study showed that individuals who over-report their body weight have higher odds of low muscle mass among community-dwelling Japanese adults aged 40 years and older. Our findings suggest that public health strategies should not only promote healthy weight but also focus on fostering accurate weight perception among the population.

#### **Abbreviations**

SMI: Skeletal muscle index; ALM: Appendicular lean mass; BMI: Body mass index; CI: Confidence interval; OR: Odds ratio; DXA: Dual-energy X-ray absorptiometry.

#### Acknowledgements

We thank all of the participants who so generously gave their time and effort for this work. Furthermore, we thank Editage (www.editage.jp) for English language editing.

#### Authors' contributions

Study concept and design: TN and TY. Investigation and data collecting: TN, TY, DW, AY, NS, HS, NN, and MM. Data analysis and writing — original draft: TN. Supervision: YY, NN, and MM. All authors read and approved the final manuscript.

#### Funding

This study was financially supported by a grant from the Health Disparity Resolution Program Promotion Project (frailty) in Osaka Prefecture.

# Availability of data and materials

The datasets used and/or analyzed during the study are available from the corresponding author on reasonable request.

# **Declarations**

# **Competing interests**

The authors declare that they have no competing interests.

#### Author details

<sup>1</sup>National Institute of Health and Nutrition, National Institutes of Biomedical Innovation, Health and Nutrition, 1-23-1 Toyama, Tokyo 162-8636 Shinjuku-ku, Japan. <sup>2</sup>Sportology Center, Juntendo University Graduate School of Medicine, 2-1-1 Hongo, Bunkyo-ku, Tokyo 113-8421, Japan. <sup>3</sup>Institute for Active Health, Kyoto University of Advanced Science, 1-1 Nanjo Otani, Sogabe-cho, Kameoka-city, Kyoto, Japan. <sup>4</sup>Faculty of Sport Sciences, Waseda University, 2-579-15 Mikajima, Tokorozawa, Saitama 359-1192, Japan. <sup>5</sup>Department of Health and Welfare (Current, Next Generation Development Department), Settsu City Local Government, 1-1-1 Mishima, Settsu-city, Osaka 566-8555, Japan. <sup>6</sup>Department of Public Health and Medical Affairs, Osaka Prefectural Government, 2-1-22 Otemae, Chuo-ku, Osaka-city, Osaka 540-8570, Japan.

# Received: 22 November 2021 Accepted: 26 April 2022 Published online: 05 May 2022

# References

 Duncan DT, Wolin KY, Scharoun-Lee M, Ding EL, Warner ET, Bennett GG. Does perception equal reality? Weight misperception in relation to

- weight-related attitudes and behaviors among overweight and obese US adults. Int J Behav Nutr Phys Act. 2011;8(1):20. https://doi.org/10. 1186/1479-5868-8-20.
- Park S, Lee S, Hwang J, Kwon J-W. The impact of weight misperception on health-related quality of life in Korean adults (KNHANES 2007–2014): a community-based cross-sectional study. BMJ Open. 2017;7(6):e016098. https://doi.org/10.1136/bmjopen-2017-016098.
- Großschädl F, Haditsch B, Stronegger WJ. Validity of self-reported weight and height in Austrian adults: sociodemographic determinants and consequences for the classification of BMI categories. Public Health Nutr. 2012;15(1):20–7. https://doi.org/10.1017/S1368980011001911.
- Ikeda N. Validity of self-reports of height and weight among the general adult population in Japan: findings from National Household Surveys, 1986. PLoS One. 2016;11(2):e0148297. https://doi.org/10.1371/journal. pone.0148297.
- Burns CM, Tijhuis MAR, Seidell JC. The relationship between quality of life and perceived body weight and dieting history in Dutch men and women. Int J Obes Relat Metab Disord. 2001;25(9):1386–92. https://doi. org/10.1038/sj.ijo.0801714.
- Herman KM, Hopman WM, Rosenberg MW. Self-rated health and life satisfaction among Canadian adults: associations of perceived weight status versus BMI. Qual Life Res. 2013;22(10):2693–705. https://doi.org/10. 1007/s11136-013-0394-9.
- Park B, Cho HN, Choi E, Seo DH, Kim N-S, Park E, et al. Weight control behaviors according to body weight status and accuracy of weight perceptions among Korean women: a nationwide populationbased survey. Sci Rep. 2019;9(1):9127. https://doi.org/10.1038/ s41598-019-45596-z.
- Cruz-Jentoft AJ, Bahat G, Bauer J, Boirie Y, Bruyere O, Cederholm T, et al. Sarcopenia: revised European consensus on definition and diagnosis. Age Ageing. 2019;48(1):16–31. https://doi.org/10.1093/ageing/afy169.
- Lee K. Weight underestimation and weight nonregulation behavior may be related to weak grip strength. Nutr Res. 2021;87:41–8. https://doi.org/ 10.1016/i.nutres.2020.12.016.
- Watanabe D, Yoshida T, Nakagata T, Sawada N, Yamada Y, Kurotani K, et al. Factors associated with sarcopenia screened by finger-circle test among middle-aged and older adults: a population-based multisite crosssectional survey in Japan. BMC Public Health. 2021;21(1):798. https://doi. org/10.1186/s12889-021-10844-3.
- Yamada Y, Nishizawa M, Uchiyama T, Kasahara Y, Shindo M, Miyachi M, et al. Developing and validating an age-independent equation using multi-frequency bioelectrical impedance analysis for estimation of appendicular skeletal muscle mass and establishing a cutoff for sarcopenia. Int J Environ Res Public Health. 2017;14(7). https://doi.org/10. 3390/ijerph14070809.
- 12. Chen LK, Woo J, Assantachai P, Auyeung TW, Chou MY, Iijima K, et al. Asian Working Group for Sarcopenia: 2019 consensus update on sarcopenia diagnosis and treatment. J Am Med Dir Assoc. 2020;21(3):300–307.e2. https://doi.org/10.1016/j.jamda.2019.12.012.
- Yoshida T, Watanabe D, Nakagata T, Yamada Y, Kurotani K, Sawada N, et al. Prevalence of frailty and its related factors in community-dwelling middle-aged and elderly adults in Settsu and Hannan cities in Osaka prefecture. Nihon Koshu Eisei Zasshi. 2021. https://doi.org/10.11236/jph. 20-111.
- Murakami K, Livingstone MB. Prevalence and characteristics of misreporting of energy intake in US adults: NHANES 2003-2012. Br J Nutr. 2015;114(8):1294–303. https://doi.org/10.1017/s0007114515002706.
- Preiss K, Brennan L, Clarke D. A systematic review of variables associated with the relationship between obesity and depression. Obes Rev. 2013;14(11):906–18. https://doi.org/10.1111/obr.12052.
- Pacanowski CR, Bertz F, Levitsky DA. Daily self-weighing to control body weight in adults:a critical review of the literature. SAGE Open. 2014;4(4):2158244014556992. https://doi.org/10.1177/2158244014 556992.
- Okamoto N, Hosono A, Shibata K, Tsujimura S, Oka K, Fujita H, et al. Accuracy of self-reported height, weight and waist circumference in a Japanese sample. Obes Sci Pract. 2017;3(4):417–24. https://doi.org/10. 1002/osp4.122.
- 18. Watanabe D, Yoshida T, Watanabe Y, Yamada Y, Kimura M, Kyoto-Kameoka Study G. A U-shaped relationship between the prevalence of frailty and body mass index in community-dwelling Japanese older adults: The

- Kyoto-Kameoka study. J Clin Med. 2020;9(5). https://doi.org/10.3390/icm9051367.
- Yong V, Saito Y. How accurate are self-reported height, weight, and BMI among community-dwelling elderly Japanese?: evidence from a national population-based study. Geriatr Gerontol Int. 2012;12(2):247–56. https:// doi.org/10.1111/j.1447-0594.2011.00759.x.
- 20. Kim H, Suzuki T, Kim M, Kojima N, Yoshida Y, Hirano H, et al. Incidence and predictors of sarcopenia onset in community-dwelling elderly Japanese women: 4-year follow-up study. J Am Med Dir Assoc. 2015;16(1):85.e1–8. https://doi.org/10.1016/j.jamda.2014.10.006.
- Johansson L, Solvoll K, Bjørneboe GE, Drevon CA. Under- and overreporting of energy intake related to weight status and lifestyle in a nationwide sample. Am J Clin Nutr. 1998;68(2):266–74. https://doi.org/10. 1093/ajcn/68.2.266.
- Mattisson I, Wirfält E, Aronsson CA, Wallström P, Sonestedt E, Gullberg B, et al. Misreporting of energy: prevalence, characteristics of misreporters and influence on observed risk estimates in the Malmö Diet and Cancer cohort. Br J Nutr. 2005;94(5):832–42. https://doi.org/10.1079/bjn20 051573.
- Houston M, vanDellen M, Cooper JA. Self-weighing frequency and its relationship with health measures. Am J Health Behav. 2019;43(5):975–93. https://doi.org/10.5993/ajhb.43.5.9.
- Zheng Y, Sereika SM, Burke LE, Olgin JE, Marcus GM, Aschbacher K, et al. Temporal patterns of self-weighing behavior and weight changes assessed by consumer purchased scales in the Health eHeart Study. J Behav Med. 2019;42(5):873–82. https://doi.org/10.1007/ s10865-018-00006-7
- Shieh C, Knisely MR, Clark D, Carpenter JS. Self-weighing in weight management interventions: a systematic review of literature. Obes Res Clin Pract. 2016;10(5):493–519. https://doi.org/10.1016/j.orcp.2016.01.004.

# **Publisher's Note**

Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.

# Ready to submit your research? Choose BMC and benefit from:

- fast, convenient online submission
- thorough peer review by experienced researchers in your field
- rapid publication on acceptance
- support for research data, including large and complex data types
- gold Open Access which fosters wider collaboration and increased citations
- maximum visibility for your research: over 100M website views per year

#### At BMC, research is always in progress.

**Learn more** biomedcentral.com/submissions

