# The avoidable disease burden associated with overweight and obesity in Kenya: A modelling study 

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## Summary

Background Globally, there is a rising burden of non-communicable diseases related to high body mass index (BMI). Estimation of the magnitude of the avoidable disease burden related to high BMI in Kenya could inform priority setting in health.

Methods Using a proportional multistate life table model, we estimated the impact of the elimination of exposure to high BMI ( $>22.5 \mathrm{~kg} / \mathrm{m}^{2}$ ) on health adjusted life years, health adjusted life expectancy, and burden of 27 obesityrelated diseases. Participants were the 2019 Kenyan population modelled over their remaining lifetime.

Findings Elimination of high BMI could save approximately 83.5 million health-adjusted life years and increase the health-adjusted life expectancy by $2 \cdot 3$ ( $95 \%$ UI $2 \cdot 0-2 \cdot 8$ ) years for females and I. 0 ( $95 \%$ UI $0 \cdot 8-\mathrm{I} \cdot \mathrm{I}$ ) years for males. Over the first 25 years, over $7 \cdot 4$ million new cases of BMI-related diseases could be avoided and approximately half a million BMI related deaths postponed. The cumulative number of new cases of type 2 diabetes could reduce by approximately I. 6 million, cardiovascular diseases by over I. 3 million, chronic kidney disease by 850,473 and cancer would reduce by 55,624 estimated cases. In 2044, an estimated 867,664 prevalent cases of musculoskeletal disease would be prevented.

Interpretation The magnitude of avoidable high BMI-related disease burden in Kenya underscores the need to prioritise the control and prevention of overweight and obesity globally, especially in low- and middle-income settings, where obesity rates are rising rapidly. Reducing population BMI is challenging, but sustained and well-enforced sys-tem-wide approaches could be a great starting point.

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## Introduction

Globally, high body mass was responsible for over five million deaths and I 60 million disability-adjusted life years (DALYs) in 2019. ${ }^{\text {I }}$ Between 1975 and 2016, the age-standardised mean body mass index (BMI) increased in both men and women across all regions, including Africa. ${ }^{2,3}$ The latest published cross-sectional

[^0]household survey in Kenya indicates that $27 \%$ of the adult population have overweight or obesity ( $38.5 \%$ of women and $I 7 \cdot 5 \%$ of men), ${ }^{4}$ using the WHO definition of overweight as a BMI of $25 \cdot 0-29 \cdot 9 \mathrm{~kg} / \mathrm{m}^{2}$ and obesity as $\geq 30 \cdot 0 \mathrm{~kg} / \mathrm{m}^{2}$. ${ }^{5}$ This is coupled with a burden of underweight which although in decline, ir. $9 \%$ of the adult population have a BMI $<18.5 \mathrm{~kg} / \mathrm{m}^{2} .^{4}$ An 'obesity transition' is taking shape in Kenya. ${ }^{6,7}$ High body mass is ranked $7^{\text {th }}$ among the top io risk factors contributing to total DALYs in 2019 for all ages combined in Kenya and, it registered the highest increase between 2009 and 2019, at $+66 \cdot 2 \% .^{8}$ The rising BMI-related burden of non-communicable diseases (NCDs) forms part of a double burden of disease as the country continues to
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## Research in context

## Evidence before this study

We conducted a review of literature and found that overweight and obesity are a leading health risk and have one of the highest rates of increase globally and in Sub-Saharan Africa (SSA). Key data sources and studies that informed the global estimates were the 2017 and 2019 Global Burden of Disease (GBD) studies, Global health observatory data and the NCD Risk Factor Collaboration (NCD-RisC) study from 1975 to 2016 and the 2015 Kenya WHO STEPwise approach to Surveillance of NCD risk factors (STEPS) survey.

The evidence showed that Kenya is experiencing an epidemiological transition where the rising BMI-related burden of non-communicable diseases (NCDs) forms part of a double burden of disease as the country continues to fight infectious diseases, which although in decline, still dominate the country's overall disease burden.

## Added value of this study

Our findings add value to the existing evidence on the growing burden related to overweight and obesity in Kenya and similar settings. We estimate health-adjusted life years (HALYs) that could be gained from achieving optimal BMI levels, factoring in the rising trend in BMI and population ageing.

Our study quantifies the future avoidable burden which is relevant for prevention policy as we determine how much of the future burden can be prevented, and what the consequences of inaction are likely to be.

## Implications of all the available evidence

Our findings show a significant magnitude of avoidable BMI-related disease burden, making a strong case for policy makers to prioritise public health efforts for the prevention and control of obesity globally, especially in low- and middle-income settings where obesity rates are rising rapidly, and health systems are already under strain. Our outputs may provide additional measurable impact indicators that can be used in future to monitor progress of the national NCD control strategies in Kenya and similar settings.
fight infectious diseases, which although in decline, still dominate Kenya's disease burden. ${ }^{\text {I }}$

According to the World Health Organization (WHO) Global Action Plan for the Prevention and Control of NCDs 2013-2020, ${ }^{9}$ one of the nine voluntary global NCD targets is to halt the rise ( $0 \%$ increase) of global obesity by 2025. In Kenya, the National Health Policy 2014-2030 and the National Strategy for the prevention and control of NCDs (2015-2020), ${ }^{\text {IO,II }}$ have adopted the global target set for obesity. However, currently the health sector heavily directs resources towards fighting infectious diseases and less towards NCDs. Estimation
of the magnitude of the avoidable BMI-related disease burden in Kenya may empower policymakers to prioritise overweight and obesity alongside the responses to persisting communicable diseases.

In this study, we quantified the avoidable burden of BMI-related diseases in Kenya by modelling the potential health impact of changes in BMI for the 2019 Kenyan adult population over their remaining lifetime.

## Methods

## Current BMI distribution and trend

We used BMI data from the 2015 Kenya 'STEPwise approach to Surveillance of NCD risk factors' (STEPS) survey. ${ }^{4}$ Height and weight measures from the survey were used to compute the weighted mean BMI and standard deviation by five-year age group and sex, which fed into the risk factor section of the model (Supplementary file [SF] Table i). To derive the trend, we used age- and sex-specific mean BMI data for Kenya from the NCD Risk Factor Collaboration (NCD-RisC) study from I975 to 2016 (SF p. 3, Table 2 and Figures 1,2). ${ }^{3}$ We used the linear trend for our main analysis as it allows us to easily subject the modelled trend to changes (halve and double trend) in our sensitivity analysis. We also model the second order polynomial trend in sensitivity analysis.

In modelling the BMI trend, both the mean and standard deviation shift in equal proportion (i.e., a ratio of I:I is assumed). This is consistent with shifts observed in the work done by Fogel. ${ }^{\text {I2 }}$ We modelled BMI as lognormally distributed. A comparison of fitted gamma, Weibull, normal and lognormal distributions has shown the lognormal distribution to be the best fit for BMI. ${ }^{\text {I3,I4 }}$

## Disease epidemiology

We used age- and sex- specific mean estimates for incidence, prevalence, and cause-specific mortality rates, all-cause mortality rates, years lived with disability and the 2019 Kenya population data estimates from the Global Burden of Disease (GBD) 20I9 study to populate our model (SF Tables 3-7). ${ }^{\text {I,15 }}$ The GBD study presents a peer reviewed systematic analysis for the global burden of disease in 204 countries and territories. The authors use best available national data from primary sources in Kenya to compute national epidemiological estimates (SF Table 3). Several GBD studies have provided subnational estimates for Kenya, an indication of availability of significant primary data from the country. A detailed list of primary data sources is available through the GBD 20I9 sources. ${ }^{\text {I }}$ We used DisMod II software ${ }^{16}$ to enforce internal consistency in the Kenya epidemiological estimates obtained, while deriving remission and case fatality rates that are not provided in the GBD data (SF Figure 3). We calculated disability
weights (DW) using disease specific prevalence and years lived with disability (YLDs) estimates from the 2019 GBD study (SF Table 8). ${ }^{\text {I,17 }}$

## Modelled diseases and associations with BMI

We identified evidence on diseases whose association with BMI was found to satisfy the criteria for causal relationships. ${ }^{8,18}$ We model a total of 27 diseases related to high BMI as included in the most recent GBD 2019 study. ${ }^{8}$ Inclusion of risk-outcome pairs in the GBD study is preceded by a thorough evaluation process (SF p. 39). Only the risk- outcome pairs that were assessed as meeting the World Cancer Research Fund (WCRF) grades of 'convincing' or 'probable' evidence were included in the GBD study. ${ }^{8,19}$ The 27 modelled diseases are: type 2 diabetes, six cardiovascular conditions (atrial fibrillation and flutter, hypertensive heart disease, ischaemic heart disease, ischaemic stroke, intracerebral haemorrhage, subarachnoid haemorrhage), i8 cancers (breast [female], colon and rectum, gallbladder and biliary tract, kidney, leukaemia, liver, multiple myeloma, oesophageal, ovarian, pancreatic, thyroid, uterine), chronic kidney diseases, gout, low back pain, osteoarthritis (hip and knee), Alzheimer's disease and other dementias, asthma, cataract, gallbladder and biliary diseases. Consistent with the GBD study, specific types of diseases were modelled.

We used the GBD 2019 relative risk estimates (mean, lower and upper levels) by age and sex, reported as the relative risk of morbidity or mortality from a high-BMI-related disease, per five BMI-unit ( $5 \mathrm{~kg} / \mathrm{m}^{2}$ ) increase (SF Table 9). ${ }^{\circ}{ }^{\circ}$

In this study, we set the theoretical minimum risk exposure level (TMREL) ${ }^{2 \mathrm{I}}$ as our counterfactual level of risk exposure for the population. Prospective cohort studies ${ }^{8,18,22}$ have shown that the optimal BMI (TMREL) associated with the lowest risk of all-cause mortality is in the range of BMI $20-25 \mathrm{~kg} / \mathrm{m}^{2}$ for adults (aged $20+$ years). We use the term 'high BMI' to refer to BMI above this TMREL (optimal level). For modelling purposes, we used the mid-point (mean $22.5 \mathrm{~kg} / \mathrm{m}^{2}$ and a standard deviation of one) as the new BMI for the population where exposure to high BMI is eliminated. Further, in the potential impact fraction (PIF) ${ }^{23}$ calculation, we used the mid-point $\left(22.5 \mathrm{~kg} / \mathrm{m}^{2}\right)$ as the point at which the risk starts to rise (optimal BMI level) and varied this to 20 and $25 \mathrm{~kg} / \mathrm{m}^{2}$ in sensitivity analysis. The PIF calculation is described in SF pp. 46-47. We modelled the entire 2019 population in Kenya, and assumed the risks start rising only after age 20 and no burden among children. The available evidence shows that high BMI is associated with NCDs that are commonly seen in adults (apart from asthma) and respective relative risk measures are available for adults only.

## Specification of modelled scenario

We estimated the avoidable burden of BMI-related diseases in Kenya if the entire population attained an
optimal level of BMI (elimination of exposure to BMI $>22.5 \mathrm{~kg} / \mathrm{m}^{2}$ ). For this, we modelled the difference between health outcomes for a reference population with 'business as usual' where the current BMI levels and trend continue unabated for 25 years, and a comparator (an identical 'intervention' population) in which the entire population has an optimal BMI, operationalised as 'no exposure to risk factor'. Health benefits were not discounted.

## Multistate lifetable modelling

We conducted a modelling study using a proportional multi-state lifetable (pMSLT) model, ${ }^{24}$ to estimate the avoidable burden of BMI-related diseases in Kenya. We created the Kenya obesity model, an adaptation of the Assessing Cost-Effectiveness (ACE) model originally created for Australia. ${ }^{14,25-27}$ The pMSLT is a macrosimulation model that simulates multiple age cohorts of the population into the future. It allows for the inclusion of multiple diseases simultaneously, while accounting for comorbidity. ${ }^{24}$ The model is divided into a general section, that is a standard cause elimination life table (main lifetable), and a separate section for each disease with an independent illness-death (Markov) process (SF Figure 4). ${ }^{24}$ In our study, we generated a total of 37 disease specific sections. The diseases were modelled applying a set of differential equations to describe the transition between four states (healthy, diseased, dead from the disease, and dead from all other causes). ${ }^{28}$ Transition probabilities among the four states reflect rates of incidence, remission, case fatality, and mortality from all other causes. For every modelled disease, the potential impact fraction ${ }^{23}$ calculates the proportional change in incidence after a change in exposure to the risk factor, high BMI (SF pp. 46-47). The changes in incidence of the disease subsequently lead to changes in mortality and morbidity rates. These new 'post-intervention/ elimination of high BMI' rates are fed into the lifetable.

The lifetables were populated with a closed cohort of the entire 2019 population disaggregated by sex and five-year age group. The lifetables integrated all-cause mortality rates and years of life lived in poor health due to disease (SF Tables 5 [p. 32] and 6). The latter were modelled as the average years lived with disability (YLD) for a given age and sex ('prevalent YLD', pYLD). The mortality rate from 'all other causes' (those not explicitly modelled) was calculated by subtracting the sum of dis-ease-specific 'pre-intervention' mortality rates from the all-cause mortality rate. ${ }^{24}$ In our model, all-cause mortality rates and pYLDs for 'all other' causes are assumed to remain constant into the future. Disease-specific 'postintervention' mortality rates generated from each diseasespecific section of the model were used to update the mortality from 'all other' causes to derive 'post-intervention' all-cause mortality rates. Similarly, pYLDs for 'all other' causes were updated based on the sum of revised disease-
specific pYLDs to derive 'post-intervention' pYLDs. Dis-ease-specific pYLDs were the disability weights of the disease multiplied by disease prevalence.

All the disease specific changes feed into the lifetable to calculate the number of new health-adjusted life years (HALYs) and life expectancies. HALYs are summary measures of population health that allow morbidity and mortality to be simultaneously described within a single number. ${ }^{29,30}$ In our study, HALYs are the total number of years lived (life years) adjusted for disability (poor health due to the disease). The lifetable calculations for the stratified cohorts (by sex and five-year age group) are simulated with one-year cycle lengths until everyone dies or reaches the age of ioo years for both the reference and comparator populations. The differences in outcomes (change in incidence, prevalence, mortality, life years lived, HALYs, life expectancy (LE), health adjusted life expectancy [HALE]) between these two populations are estimated to determine the avoidable BMIrelated burden. To show the impact at the level of individuals, we also estimated the average LE gain for each hypothetical five-year age cohorts.

To run these analyses, we used Microsoft Excel for Microsoft 365 and two software Add-ins: the EpiGearXL 5.0 add-in for calculation of the potential impact fraction and Ersatz (Version I.35) for the uncertainty and sensitivity analysis. ${ }^{3 I}$ We used the GATHER checklist to guide the documenting of our data and methods. ${ }^{32}$

## Uncertainty analyses

The simultaneous and combined effect of the uncertainty in model inputs on our outcomes was quantified using Ersatz software that performed a Monte Carlo simulation with bootstrapping (2000 iterations) while incorporating probabilistic uncertainty from model inputs (relative risks). For all outputs, the $95 \%$ uncertainty intervals (UIs) were calculated. These are 2.5 and 97.5 percentiles capturing sampling error with input data.

## Sensitivity analyses

In one-way sensitivity analyses, we examined the implications of varying the 25 years of linear trend applied to reference populations to $0,5,10,15$ and 20 years. We modelled changes to the trend, halving and doubling the linear trend and application of a second order polynomial trend instead of linear. We also varied the point at which the risk starts to rise (optimal BMI level) in the PIF calculation from BMI $22.5 \mathrm{~kg} / \mathrm{m}^{2}$ used in the main analysis to $20 \mathrm{~kg} / \mathrm{m}^{2}$ and $25 \mathrm{~kg} / \mathrm{m}^{2}$.

## Ethics approval

Ethics approval was not required for this analysis. However, this study was carried out as part of a larger study that has ethical approval from the Griffith University

Human Research Ethics Committee (GU Ref No: 2019/ 707) and the Kenyatta National Hospital/University of Nairobi Ethics \& Research Committee (P8I/O2/202I).

## Patient consent

Patient consent was not required for this study.

## Role of the funding source

The funder had no role in the study design, collection, analysis, and interpretation of data; in the writing of the report; and in the decision to submit the paper for publication. All authors had access to the data and took the decision to submit the manuscript for publication.

## Results

## Changes in prevalence of overweight and obesity

We compared a Kenyan adult ( $\geq 20$ years) population with current BMI levels and trend applied, against an adult population that had no exposure to high BMI. We found that in 2019, an estimated 2.69 million cases of obesity would be avoided, of which $78 \%$ were among females. An estimated 5.82 million cases of overweight would be avoided, $6 \mathrm{I} \%$ of which were among females.

## Changes in health-adjusted life years and overall deaths

Over the lifetime of the 2019 Kenyan population, our model projected that elimination of exposure to high BMI could save approximately 83.5 million ( $95 \%$ uncertainty intervals [UIs] $73 \cdot 0$ to $96 \cdot \mathrm{I}$ ) HALYs (Table I). Greater gains would be seen in females than males. By the year 2025, our model predicted that nearly 210 thousand HALYs would be saved due to elimination of exposure to high BMI. By 2030, these gains are expected to increase to almost 783 thousand HALYs, while by 2044, approximately 5.7 million HALYs could be gained. Elimination of high BMI could save 22,735 lives by 2025 (or rather, postpone deaths). By 2030, the number of overall deaths postponed was projected to rise to 69,257 . For a 25 -year period (from 2019 to 2044), a total of 342,336 deaths could be postponed among the 2019 Kenyan population with more deaths postponed in females than in males.

## Changes in life expectancy and health adjusted life expectancy

Over the lifetime of the 2019 Kenyan population, our model predicted that elimination of exposure to high BMI could increase life expectancy (LE) by $2.0(95 \%$ UI $\mathrm{I} \cdot 5-2 \cdot 5$ ) years for females and 0.7 ( $95 \%$ UI $0.6-0.9$ ) years for males and increase health-adjusted life expectancy (HALE) by 2.3 ( $95 \%$ UI $2 \cdot 0-2 \cdot 8$ ) years for females and I.O ( $95 \%$ UI $0 \cdot 8-\mathrm{I} \cdot \mathrm{I}$ ) years for males (SF Table io).

| Variable | Male, mean (95\% UI) | Female, mean (95\% UI) | Total, mean (95\% UI) |
| :---: | :---: | :---: | :---: |
| HALYs gained over the lifetime of the 2019 population | $\begin{aligned} & 24,488,764 \\ & (21,170,666-28,175,394) \end{aligned}$ | $\begin{aligned} & 59,025,238 \\ & (49,553,478-70,930,081) \end{aligned}$ | $\begin{aligned} & 83,514,002 \\ & (73,043,515-96,061,248) \end{aligned}$ |
| HALYs gained over different time periods |  |  |  |
| Year 00-06 (2019-2025) | $\begin{aligned} & 78,209 \\ & (70,141-86,344) \end{aligned}$ | $\begin{aligned} & 131,804 \\ & (118,416-145,313) \end{aligned}$ | $\begin{aligned} & 210,014 \\ & (194,602-226,328) \end{aligned}$ |
| Year 00-11 (2019-2030) | $\begin{aligned} & 289,812 \\ & (260,600-320,095) \end{aligned}$ | $\begin{aligned} & 492,773 \\ & (443,362-543,752) \end{aligned}$ | $\begin{aligned} & 782,586 \\ & (723,979-844,629) \end{aligned}$ |
| Year 00-25 (2019-2044) | $\begin{aligned} & 2,053,948 \\ & (1,821,110-2,301,963) \end{aligned}$ | $\begin{aligned} & 3,646,534 \\ & (3,249,011-4,076,079) \end{aligned}$ | $\begin{aligned} & 5,700,482 \\ & (5,217,355-6,220,782) \end{aligned}$ |
| Deaths avoided over different time periods* |  |  |  |
| Year 00-06 (2019-2025) | $\begin{aligned} & 7437 \\ & (6,209-8,718) \end{aligned}$ | $\begin{aligned} & 15,298 \\ & (12,196-18,464) \end{aligned}$ | $\begin{aligned} & 22,735 \\ & (19,458-26,187) \end{aligned}$ |
| Year 00-11 (2019-2030) | $\begin{aligned} & 23,115 \\ & (19,331-27,019) \end{aligned}$ | $\begin{aligned} & 46,143 \\ & (37,334-55,236) \end{aligned}$ | $\begin{aligned} & 69,257 \\ & (60,033-79,096) \end{aligned}$ |
| Year 00-25 (2019-2044) | $\begin{aligned} & 112,194 \\ & (92,876-132,509) \end{aligned}$ | $\begin{aligned} & 230,142 \\ & (187,471-278,074) \end{aligned}$ | $\begin{aligned} & 342,336 \\ & (295,696-395,276) \end{aligned}$ |

Table 1: Effects of the elimination of exposure to high BMI on HALYs and deaths, in Kenya.
HALYs: Health adjusted life years, BMI: Body mass index, UI: Uncertainty interval.
*This is the overall mortality number from lifetable where we count both the reduction in mortality from BMI-related diseases and the increase in mortality due to other causes.

Figure I illustrates the estimated increase in HALE and LE for every five-year age cohort from birth to 94 years.

## Changes in disease burden

Type 2 diabetes. Over the 25 years between 2019 and 2044, our model predicts that elimination of exposure to high BMI in the Kenyan population would reduce the cumulative number of new cases of type 2 diabetes mellitus (T2DM) by approximately I. 6 ( $95 \%$ UI I•3-I.9) million (Table 2). In 2044, approximately I•I(95\% UI 0.9-I-3) million prevalent cases of $\mathrm{T}_{2} \mathrm{DM}$ due to high BMI
could be avoided (Table 3). This corresponds to a $59 \%$ ( $95 \%$ UI $53-65$ ) reduction in T2DM prevalence in 2044 (Figure 2, SF Table ir). The mean estimates show that the avoidable prevalent burden predicted is greater for females ( $615,377,95 \%$ UI $48 \mathrm{I}, 784-75 \mathrm{I}, 098$ ) compared with males (491,377, 95\%UI 372,682-615,592).

Cardiovascular disease. Over 25 years, the estimated reductions in cumulative numbers of new cases of ischemic heart disease (IHD), hypertensive heart disease (HHD), ischaemic stroke (IS), intracerebral haemorrhage (ICH), sub-arachnoid haemorrhage (SAH) and


Figure 1. Mean sex specific increases in HALE and LE over the lifetime for each 5-year age cohort. LE: life expectancy, HALE: healthadjusted life expectancy. Only mean estimates of increases in LE and HALE are plotted. In the SF table 10, the mean and uncertainty intervals of the LE and HALE gained for the youngest age group ( $0-4$ years) are reported.

| High BMI related disease | Numbers of new cases averted |  |  |
| :---: | :---: | :---: | :---: |
|  | Male, n ( $95 \%$ UI) | Female, $\mathbf{n}$ (95\% UI) | Total, n ( $95 \%$ UI) |
| Diabetes mellitus type 2 | 651,426 | 933,527 | 1,584,953 |
|  | (495,858-814,609) | (719,822-1,149,966) | (1,319,597-1,857,381) |
| Cardiovascular diseases |  |  |  |
| Ischemic heart disease | 189,192 | 250,088 | 439,280 |
|  | (136,850-242,694) | (180,986-326,201) | (350,152-531,832) |
| Ischemic stroke | 58,562 | 139,603 | 198,165 |
|  | (42,396-75,128) | (100,909-181,530) | (155,607-242,027) |
| Intracerebral haemorrhage | 113,756 | 175,876 | 289,632 |
|  | (75,000-159,416) | (115,533-240,745) | (215,950-367,584) |
| Subarachnoid haemorrhage | 19,183 | 32,409 | 51,592 |
|  | (12,467-27,022) | (21,350-44,145) | (38,118-65,966) |
| Hypertensive heart disease | 78,000 | 173,983 | 251,983 |
|  | (33,426-132,629) | (64,072-297,935) | (134,279-382,905) |
| Atrial fibrillation and flutter | 11,763 | 22,086 | 33,849 |
|  | (7,726-16,196) | (13,824-30,063) | (24,930-43,260) |
| Cancers |  |  |  |
| Oesophageal cancer |  | 18,038 | 26,666 |
|  | (1776-17,153) | (528-36,898) | (8027-47,381) |
| Colon cancer | 2275 | 1532 | 3807 |
|  | (1813-2730) | (499-2597) | (2674-4935) |
| Liver cancer due to alcohol use | 535 | 332 | 867 |
|  | (181-927) | (34-654) | (406-1345) |
| Liver cancer due to hepatitis B | 581 | 551 | 1,132 |
|  | (208-1,002) | (52-1051) | (545-1,766) |
| Liver cancer due to hepatitis $C$ | 231 | 518 | 749 |
|  | (71-404) | (41-1056) | (234-1323) |
| Gallbladder and biliary cancer | 177 | 2,518 | 2,694 |
|  | (8-357) | (1586-3566) | (1,716-3750) |
| Pancreatic cancer | 241 | 1,061 | 1,302 |
|  | (-84-632) | (288-1,910) | (446-2,242) |
| Breast cancer |  | 6,370 | 6,370 |
|  |  | (2,467-10,533) | (2467-10,533) |
| Uterine cancer |  | 6498 | 6,498 |
|  |  | (5830-7181) | (5830-7181) |
| Ovarian cancer |  | 815 | 815 |
|  |  | (-290-1,969) | (-290-1,969) |
| Kidney cancer | 404 | 1338 | 1742 |
|  | (276-529) | (1050-1631) | (1428-2055) |
| Thyroid cancer | 192 | 785 | 978 |
|  | (53-339) | (538-1035) | (695-1266) |
| Acute lymphoid leukaemia | 51 | 62 | 113 |
|  | (30-72) | (28-97) | (75-151) |
| Acute myeloid leukaemia | 102 | 221 | 323 |
|  | (56-147) | (95-339) | (191-455) |
| Chronic lymphoid leukaemia | 53 | 334 | 387 |
|  | (24-81) | (97-567) | (140-624) |
| Chronic myeloid leukaemia | 60 | 95 | 155 |
|  | (34-87) | (42-146) | (96-212) |
| Other Leukaemia | 185 | 175 | 360 |
|  | (96-278) | (75-272) | (228-492) |
| Multiple myeloma | 183 | 482 | 665 |
|  | (16-359) | (68-903) | (235-1120) |


| High BMI related disease | Numbers of new cases averted |  |  |
| :---: | :---: | :---: | :---: |
|  | Male, n ( $95 \%$ UI) | Female, $\mathbf{n}$ (95\% UI) | Total, n ( $95 \%$ UI) |
| Chronic kidney disease |  |  |  |
| CKD due to diabetes mellitus type 2 | 72,631 | 148,923 | 221,554 |
|  | (9177-155,566) | (22,230-292,402) | $(76,388-386,353)$ |
| CKD due to glomerulonephritis | 5254 |  | 13,848 |
|  | (869-11,619) | (1,098-18,100) | (4,843-24,908) |
| CKD due to hypertension | 25,929 | 51,358 | 77,287 |
|  | (5,485-54,712) | (9,993-103,501) | (28,635-136,559) |
| CKD due to other and unspecified causes | 144,060 | 393,723 | 537,784 |
|  | (21,976-307,130) | $(53,716-805,259)$ | (182,783-979,804) |
| Gallbladder and biliary diseases | 40,452 | 267,746 | 308,198 |
|  | (25,575-55,728) | (213,237-322,271) | (238,846-378,019) |
| Asthma | 126,948 | 310,603 | 437,551 |
|  | (90,055-166,572) | (219,833-406,617) | (309,912-572,806) |
| Alzheimer's disease and other dementias | 9,643 | 42,492 | 52,135 |
|  | (401-19,930) | (-856-91,078) | (-535-111,007) |
| Cataract | 17,639 | 55,668 | 73,307 |
|  | (6531-28,764) | (22,532-89,639) | (29,281-118,260) |
| Musculoskeletal diseases |  |  |  |
| Osteoarthritis hip | 6099 | 10,003 | 16,102 |
|  | (3299-8993) | $(5,579-14,469)$ | $(10,929-21,401)$ |
| Osteoarthritis knee | 239,724 | 588,860 | 828,584 |
|  | (135,576-349,888) | (317,606-885,289) | (532,395-1,138,651) |
| Low back pain | 559,631 | 1,019,277 | 1,578,908 |
|  | (405,662-722,318) | (732,348-1,302,013) | (1,256,131-1,898,529) |
| Gout | 208,690 | 115,815 | 324,505 |
|  | (114,971-318,690) | $(75,163-155,315)$ |  |
| Table 2: Estimated avoidable disease burden (incident cases) following elimination of exposure to high BMI in Kenya, 2019 - 2044. <br> BMI: Body mass index, UI: uncertainty interval. <br> The negative values reported could either be due to extra cases that arise in added years of life since a reduction in high BMI causes people to live longer and they are still at risk of these diseases, or the confidence intervals or boundaries of the respective relative risk values cross one. |  |  |  |

atrial fibrillation and flutter due to high BMI are illustrated Table 2. The greatest reduction in cumulative number of new cases was seen in IHD, while the least reduction is seen in atrial fibrillation and flutter. For each of the cardiovascular diseases, estimates of the avoidable prevalent cases attributed to high BMI are reported in Table 3.

Table 4 shows the avoidable number of deaths from cardiovascular diseases due to high BMI. The largest proportional reduction in mortality in 2044 was seen in HHD, projected to reduce by $57 \%$, followed by SAH ( $44 \%$ reduction) (Figure 2). Atrial fibrillation and flutter and ischaemic stroke recorded comparatively smallest relative reductions in mortality, $17 \%$ and $15 \%$ respectively.

Cancer. The top 5 cancers with greatest reduction of incident cases were: oesophageal $(26,666)$, uterine $(6,498)$, breast $(6,370)$, colon $(3,807)$, gallbladder and biliary $(2,694)$ cancers (Table 2). In 2044, the largest estimated proportional reductions in avoidable burden (incidence, prevalence, mortality) are seen in uterine
cancer, followed by oesophageal, gallbladder and biliary tract, and kidney cancers (Figure 2, SF Table Ii). Estimated avoidable prevalent cases and postponed deaths from cancer are reported in Tables 3 and 4 respectively.

Chronic kidney disease (CKD). If high BMI was eliminated in Kenya, our model projects that in 2044, there would be a $45 \%$ reduction of new cases of CKD due to hypertension, a $43 \%$ for CKD due to T2DM, a $38 \%$ for CKD due to other and unspecified causes, and $36 \%$ reduction for CKD due to glomerulonephritis (Table 2, Figure 2, SF Table ir). A total of 27,25I deaths from CKD attributable to high BMI could be prevented over the next 25 years. Of the four CKD types modelled, CKD due to hypertension is estimated to have the highest proportion of deaths averted (34\%) (Figure 2, SF Table iI). Additional results on the estimated cumulative number of avoidable new cases, prevalent cases and postponed deaths for respective CKD conditions attributable to high BMI are given in Tables 2,3 and 4 .

| High BMI related disease | Numbers of avoidable prevalence cases |  |  |
| :---: | :---: | :---: | :---: |
|  | Male, $\mathbf{n}$ ( $\mathbf{9 5 \%}$ UI) | Female, $\mathbf{n}$ (95\% UI) | Total, $\mathbf{n}$ ( $\mathbf{9 5 \%} \mathbf{~ U I )}$ |
| Diabetes mellitus type 2 | 491,377 | 615,377 | 1,106,755 |
|  | $(372,682-615,592)$ | (481,784-751,098) | (928,981-1,288,955) |
| Cardiovascular diseases |  |  |  |
| Ischemic heart disease | 114,212 | 151,122 | 265,334 |
|  | (80,442-149,087) | (107,070-199,664) | (208,335-324,512) |
| Ischemic stroke | 38,099 | 95,624 | 133,723 |
|  | (26,771-49,929) | (68,094-125,774) | (103,011-165,287) |
| Intracerebral haemorrhage | 50,627 | 81,376 | 132,004 |
|  | (32,358-72,255) | (52,274-112,522) | (97,026-169,426) |
| Subarachnoid haemorrhage | 8155 | 14,946 | 23,101 |
|  | (5083-11,805) | (9471-20,884) | (16,556-30,168) |
| Hypertensive heart disease | 37,613 | 62,850 | 100,464 |
|  | (16,100-64,285) | (24,420-107,519) | (56,556-149,701) |
| Atrial fibrillation and flutter | 6739 | 12,228 | 18,966 |
|  | (4150-9541) | (7256-17,088) | (13,488-24,768) |
| Cancers |  |  |  |
| Oesophageal cancer | 869 |  | 2637 |
|  | (160-1768) | (9-3727) | (759-4783) |
| Colon cancer | 447 | 263 | 710 |
|  | (348-546) | (53-483) | (474-944) |
| Liver cancer due to alcohol use | 38 | 23 | 61 |
|  | (12-68) | (1-47) | (27-98) |
| Liver cancer due to hepatitis B | 44 | 41 | 85 |
|  | $(15-78)$ | (3-80) | (40-135) |
| Liver cancer due to hepatitis $C$ | 16 | 34 | 50 |
|  | (4-29) | (1-73) | (14-91) |
| Gallbladder and biliary cancer | 13 | 192 | 205 |
|  | (-1-28) | (118-277) | (128-291) |
| Pancreatic cancer | 12 | 53 |  |
|  | (-8-35) | (9-99) | (16-116) |
| Breast cancer |  | 1685 | 1685 |
|  |  | (311-3162) | (311-3162) |
| Uterine cancer |  | 2119 | 2119 |
|  |  | (1896-2,347) | (1896-2347) |
| Ovarian cancer |  | 184 | 184 |
|  |  | (-87-463) | (-87-463) |
| Kidney cancer | 90 | 284 | 374 |
|  | (61-118) | (223-346) | (306-442) |
| Thyroid cancer | 62 | 283 | 345 |
|  | (17-110) | (193-374) | (245-448) |
| Acute lymphoid leukaemia | 5 | 6 | 11 |
|  | (3-7) | (3-10) | (7-15) |
| Acute myeloid leukaemia | 7 |  | 35 |
|  | (4-11) | (12-43) | (19-51) |
| Chronic lymphoid leukaemia | 9 |  | 62 |
|  | (3-14) | (13-93) | (19-103) |
| Chronic myeloid leukaemia | 5 | 8 | 14 |
|  | (3-8) | (4-13) | (8-19) |
| Other Leukaemia | 46 | 50 | 96 |
|  | (21-72) | (21-79) | (58-135) |
| Multiple myeloma | 23 | 61 |  |
|  | (0-47) | (5-119) |  |


| High BMI related disease | Numbers of avoidable prevalence cases |  |  |
| :---: | :---: | :---: | :---: |
|  | Male, n (95\% UI) | Female, $\mathbf{n}$ (95\% UI) | Total, n (95\% UI) |
| Chronic kidney disease |  |  |  |
| CKD due to diabetes mellitus type 2 | $\begin{aligned} & 50,584 \\ & (3370-112,513) \end{aligned}$ | $\begin{aligned} & 108,689 \\ & (10,243-220,915) \end{aligned}$ | $\begin{aligned} & 159,273 \\ & (47,564-286,636) \end{aligned}$ |
| CKD due to glomerulonephritis | $\begin{aligned} & 2152 \\ & (122-5038) \end{aligned}$ | $\begin{aligned} & 4269 \\ & (310-9212) \end{aligned}$ | $\begin{aligned} & 6421 \\ & (1963-12,012) \end{aligned}$ |
| CKD due to hypertension | $\begin{aligned} & 13,726 \\ & (2238-29,624) \end{aligned}$ | $\begin{aligned} & 29,056 \\ & (4642-59,014) \end{aligned}$ | $\begin{aligned} & 42,782 \\ & (15,021-76,421) \end{aligned}$ |
| CKD due to other and unspecified causes | $\begin{aligned} & 96,340 \\ & (6114-216,466) \end{aligned}$ | $\begin{aligned} & 270,847 \\ & (14,412-582,371) \end{aligned}$ | $\begin{aligned} & 367,187 \\ & (99,735-694,544) \end{aligned}$ |
| Gallbladder and biliary diseases | $\begin{aligned} & 12,990 \\ & (8086-18,047) \end{aligned}$ | $\begin{aligned} & 92,404 \\ & (72,844-112,281) \end{aligned}$ | $\begin{aligned} & 105,393 \\ & (80,874-130,347) \end{aligned}$ |
| Asthma | $\begin{aligned} & 57,086 \\ & (39,792-75,952) \end{aligned}$ | $\begin{aligned} & 145,153 \\ & (101,232-191,877) \end{aligned}$ | $\begin{aligned} & 202,239 \\ & (141,172-267,567) \end{aligned}$ |
| Alzheimer's disease and other dementias | $\begin{aligned} & 3809 \\ & (-1033-9163) \end{aligned}$ | $\begin{aligned} & 16,630 \\ & (-7006-43,045) \end{aligned}$ | $\begin{aligned} & 20,439 \\ & (-8197-52,185) \end{aligned}$ |
| Cataract | $\begin{aligned} & 7592 \\ & (-33-15,531) \end{aligned}$ | $\begin{aligned} & 24,120 \\ & (-944-49,050) \end{aligned}$ | $\begin{aligned} & 31,712 \\ & (-383-64,654) \end{aligned}$ |
| Musculoskeletal diseases |  |  |  |
| Osteoarthritis hip | $\begin{aligned} & 3452 \\ & (1252-5,756) \end{aligned}$ | $\begin{aligned} & 5947 \\ & (2215-9,715) \end{aligned}$ | $\begin{aligned} & 9399 \\ & (5086-13,720) \end{aligned}$ |
| Osteoarthritis knee | $\begin{aligned} & 143,569 \\ & (77,345-214,203) \end{aligned}$ | $\begin{aligned} & 369,987 \\ & (184,176-570,060) \end{aligned}$ | $\begin{aligned} & 513,556 \\ & (315,501-722,313) \end{aligned}$ |
| Low back pain | $\begin{aligned} & 85,001 \\ & (58,681-112,809) \end{aligned}$ | $\begin{aligned} & 154,527 \\ & (104,709-203,681) \end{aligned}$ | $\begin{aligned} & 239,528 \\ & (184,056-294,832) \end{aligned}$ |
| Gout | $\begin{aligned} & 68,001 \\ & (36,394-105,388) \end{aligned}$ | $\begin{aligned} & 37,180 \\ & (23,454-50,646) \end{aligned}$ | $\begin{aligned} & 105,181 \\ & (71,122-145,992) \end{aligned}$ |
| Table 3: Estimated avoidable disease burden (prevalent cases) following elimination of exposure to high BMI in Kenya, 2019-2044. <br> BMI: Body mass index, UI: uncertainty interval. <br> The negative values reported could be due to extra cases that arise in added years of life since a reduction of high BMI causes people to live longer, and they are still at risk of these diseases. Also, for the confidence intervals or boundaries of the respective relative risk values cross one. |  |  |  |

Osteoarthritis, low back pain and gout. For the musculoskeletal group of health outcomes related to high BMI, the greatest number of avoidable prevalent cases in the year 2044 was seen in osteoarthritis of the knee (Table 3).

Alzheimer's disease and other dementias, asthma, cataract, gallbladder and biliary diseases. Our model estimated that over the next 25 years, 52,135 cumulative new cases of Alzheimer's disease and other dementias related to high BMI could be averted, 437,55I new cases for asthma, 73,307 for cataract and 308,198 for gallbladder and biliary diseases (Table 2). Avoidable prevalent and mortality cases attributable to high BMI are illustrated in Tables 3 and 4, respectively and, proportional reductions given in Figure 2 and SF Table II).

Sensitivity analyses. Compared to the base case ( 25 years of trend), applying 20 years of trend reduced
the total number of HALYs expected to be gained due to elimination of high BMI over the lifetime of the Kenyan population by $ぃ \%$. Greater relative reductions of the expected gain in HALYs were seen when we applied I5 years ( $22 \%$ ), io years ( $32 \%$ ) and five years ( $42 \%$ ). When no trend was applied (o years of trend), a $5 \mathrm{I} \%$ reduction of expected gain in HALYs was seen. This shows that over half the HALY loss expected is due to future rises in BMI (Figure 3 and SF Table i2).

Notably, all scenarios still resulted in substantial numbers of expected gain in HALYs (Figure 3 and SF Table 12-I4). As expected, the lowest number of expected gains in HALYs (4I million) over the lifetime of the population was recorded for the scenario where no trend is applied.

Compared to main analysis (lowest BMI boundary of $22.5 \mathrm{~kg} / \mathrm{m} 2$ in PIF calculation), applying a boundary of $20 \mathrm{~kg} / \mathrm{m}^{2}$ resulted in a $25.5 \%$ increase in expected gain in HALYs over the lifetime of the population. Applying a boundary of $25 \mathrm{~kg} / \mathrm{m}^{2}$ led to a $25.4 \%$ decrease in expected gains in HALYs over the lifetime of the


Figure 2. Top 20 diseases with largest proportional reductions in incidence, prevalence and mortality following elimination of exposure to high BMI in Kenya, 2044. The horizontal bars show the mean. The whiskers show 95\% uncertainty intervals (Uls), the 2.5 and 97.5 percentiles capturing sampling error with input data. T2DM: type 2 diabetes, HHD: hypertensive heart disease, IHD: ischaemic heart disease, ICH: intracerebral haemorrhage, SAH: subarachnoid haemorrhage, dses: diseases, CKD hypertension: chronic kidney disease (CKD) due to hypertension, CKD DM: CKD due to diabetes mellitus, CKD other: CKD due to other causes, CKD glomerulonephritis: CKD due to glomerulonephritis, OA: Osteoarthritis, Ca: cancer, Liver Ca HepB: Liver cancer due to hepatitis B, Liver Ca Alcohol: liver cancer due to alcohol use, Liver CA HepC: Liver cancer due to Hepatitis C, M: male, F: female.

| High BMI related disease | Numbers of avoidable deaths |  |  |
| :---: | :---: | :---: | :---: |
|  | Male, n (95\% UI) | Female, n (95\% UI) | Total, n ( $95 \%$ UI) |
| Diabetes mellitus type 2 | 33,666 | 32,412 | 66,077 |
|  | (25,699-41,677) | (24,750-39,924) | (55,173-76,969) |
| Cardiovascular diseases |  |  |  |
| Ischemic heart disease | 31,923 | $45,427$ |  |
|  | (22,862-41,106) | (31,764-60,012) | (59,879-94,218) |
| Ischemic stroke | 6,284 | 12,716 | 18,999 |
|  | $(4,497-8,089)$ | (8,113-17,528) | (13,933-24,131) |
| Intracerebral haemorrhage | 51,289 | 79,786 | 131,075 |
|  | (34,477-70,889) | (53,032-108,308) | (98,672-165,349) |
| Subarachnoid haemorrhage | 3,966 | 4,913 | 8,879 |
|  | (2,598-5,557) | (3,253-6,646) | (6,672-11,272) |
| Hypertensive heart disease | 21,900 | 82,493 | 104,393 |
|  | (9,666-36,441) | (29,708-140,453) | (51,183-162,920) |
| Atrial fibrillation and flutter | 410 | 2,099 | 2,509 |
|  | (240-592) | (1,095-3,087) | (1,497-3,535) |
| Cancers |  |  |  |
| Oesophageal cancer | 7,355 | $15,512$ | $22,867$ |
|  | (1,469-14,643) | (333-31,740) | (6,728-40,706) |
| Colon cancer | 1,535 | 938 | $2,473$ |
|  | (1,205-1,860) | (183-1,726) | $(1,647-3,300)$ |
| Liver cancer due to alcohol use | 473 | 297 | 769 |
|  | (157-821) | (27-586) | (361-1,195) |
| Liver cancer due to hepatitis B | 515 | 494 | 1,009 |
|  | (182-890) | (43-945) | (483-1,580) |
| Liver cancer due to hepatitis C | 203 | 459 | 662 |
|  | (61-355) | (30-943) | (201-1,177) |
| Gallbladder and biliary cancer | 150 | 2,219 | 2,370 |
|  | (4-307) | (1,394-3,146) | (1,504-3,304) |
| Pancreatic cancer | 214 | 958 | 1,172 |
|  | (-83-572) | (243-1,739) | (383-2,037) |
| Breast cancer |  | 2,554 | 2,554 |
|  |  | (688-4,537) | (688-4,537) |
| Uterine cancer |  | $2,904$ | $2,904$ |
|  |  | $(2,604-3,213)$ | $(2,604-3,213)$ |
| Ovarian cancer |  | 423 | 423 |
|  |  | $(-311-1,180)$ | $(-311-1,180)$ |
| Kidney cancer | 236 | 788 | 1,024 |
|  | (159-312) | (616-963) | (836-1,212) |
| Thyroid cancer | 58 | 144 | 202 |
|  | (13-105) | (91-199) | (134-272) |
| Acute lymphoid leukaemia | 39 |  |  |
|  | (22-55) | (20-72) | (56-113) |
| Acute myeloid leukaemia | 78 | 187 |  |
|  | (42-114) | (80-288) | (153-376) |
| Chronic lymphoid leukaemia | 34 |  |  |
|  | (13-54) | (55-411) | (83-447) |
| Chronic myeloid leukaemia | 51 | $81$ | $132$ |
|  | (28-74) | (35-126) | (82-182) |
| Other Leukaemia | 100 | 99 | 199 |
|  | $(45-157)$ | (41-156) | (120-279) |
| Multiple myeloma | 138 | 373 | 512 |
|  | (4-279) | (35-720) | (159-881) |


| High BMI related disease | Numbers of avoidable deaths |  |  |
| :---: | :---: | :---: | :---: |
|  | Male, n (95\% UI) | Female, $\mathbf{n}$ (95\% UI) | Total, n (95\% UI) |
| Chronic kidney disease |  |  |  |
| CKD due to diabetes mellitus type 2 | $\begin{aligned} & 2,173 \\ & (179-4,730) \end{aligned}$ | $\begin{aligned} & 4,214 \\ & (252-8,447) \end{aligned}$ | $\begin{aligned} & 6,387 \\ & (1,842-11,323) \end{aligned}$ |
| CKD due to glomerulonephritis | $\begin{aligned} & 1,998 \\ & (309-4,468) \end{aligned}$ | $\begin{aligned} & 2,833 \\ & (250-6,137) \end{aligned}$ | $\begin{aligned} & 4,831 \\ & (1,523-8,834) \end{aligned}$ |
| CKD due to hypertension | $\begin{aligned} & 4,971 \\ & (989-10,492) \end{aligned}$ | $\begin{aligned} & 8,156 \\ & (1,233-16,665) \end{aligned}$ | $\begin{aligned} & 13,127 \\ & (4,653-23,381) \end{aligned}$ |
| CKD due to other and unspecified causes | $\begin{aligned} & 941 \\ & (70-2,101) \end{aligned}$ | $\begin{aligned} & 1,964 \\ & (6-4,261) \end{aligned}$ | $\begin{aligned} & 2,905 \\ & (799-5,417) \end{aligned}$ |
| Gallbladder and biliary diseases | $\begin{aligned} & 2,115 \\ & (1,279-2,978) \end{aligned}$ | $\begin{aligned} & 8,503 \\ & (6,662-10,378) \end{aligned}$ | $\begin{aligned} & 10,618 \\ & (7,945-13,350) \end{aligned}$ |
| Asthma | $\begin{aligned} & 2,889 \\ & (1,982-3,875) \end{aligned}$ | $\begin{aligned} & 8,787 \\ & (5,935-11,793) \end{aligned}$ | $\begin{aligned} & 11,676 \\ & (7,921-15,648) \end{aligned}$ |

Table 4: Estimated avoidable disease burden (postponed deaths) following elimination of exposure to high BMI in Kenya, 2019-2044.
BMI: Body mass index, UI: uncertainty interval.
The negative values reported could be due to extra cases that arise in added years of life since a reduction of high BMI causes people to live longer, and they are still at risk of these diseases. Also, for the confidence intervals or boundaries of the respective relative risk values cross one.
population. Both scenarios recorded substantial expected gain in HALYs over the lifetime (Figure 3 and SF Table 15). When we varied the trend applied, compared to the main analysis (linear trend applied), expected patterns were seen when the linear trend was halved and doubled (Figure 3, SF Table 16). Compared to base case (man analysis), applying the second order polynomial trend resulted in a decrease of the expected gains in HALYs over the lifetime of the population and, a reduced number of the avoidable deaths over various time periods (Figure 3, SF Table 17). In SF Tables I2-I7, we present additional results showing how the various sensitivity scenarios impacted the expected lifetime gains in LE and HALE, HALYs, and avoidable deaths.

## Discussion

Over the next 25 years, elimination of high BMI is estimated to postpone 342,336 overall deaths and save 5.7 million HALYs for the Kenya 2019 population. Over their lifetime, an estimated 83.5 million HALYs could be gained. Over the first 25 years, over $7 \cdot 4$ million new cases of BMI-related diseases could be avoided and approximately half a million BMI related deaths postponed. The cumulative number of new cases of type 2 diabetes could reduce by approximately i. 6 million, cardiovascular diseases by over I-3 million, chronic kidney disease by 850,473 and cancer would reduce by 55,624 cases. In 2044, 867,664 prevalent cases of musculoskeletal disease would be prevented. Over 25 years, the leading contributors of the total avoidable incident BMIrelated disease were musculoskeletal diseases (37\%), $\mathrm{T}_{2} \mathrm{DM}(2 \mathrm{I} \%)$ and cardiovascular diseases ( $\mathrm{t} 7 \%$ ). Cardiovascular diseases were the largest ( $69 \%$ ) contributors to
the total postponed mortality in the same period. Our results are in line with global findings where cardiovascular diseases, diabetes, kidney diseases and cancers were found to be the leading causes of high BMI related DALYs. ${ }^{1,3}$

The avoidable burden is higher for females than males. This is largely because a greater percentage of women had overweight or obesity at baseline ( $38.5 \%$ women, $\mathrm{I} 7.5 \%$ men). ${ }^{4}$ Also, this may be due to a significant avoidable burden from musculoskeletal conditions that are more prevalent in females than in males (particularly LBP and OA), and the averted incident cases and postponed deaths from cancers specific to females (breast, uterine and ovarian cancers).

Results from our sensitivity analysis where no trend was applied indicate that half of the burden could be avoided by achieving the national and global target of halting the rise of overweight and obesity, if a continuation of recent BMI trends is assumed as comparator.

In the underlying model structure, the Markovian assumption applies, i.e., individuals have no memory of (are independent of) previous disease states. ${ }^{24,33}$ This can be addressed using transitional tunnel states to incorporate memory and make the model more realistic, but it increases complexity. ${ }^{33}$ Second, the pMSLT model, has an assumption that modelled diseases are independent. ${ }^{24,34,35}$ This assumption has been found to have limited impact in most settings on overall estimates. ${ }^{34}$ However, as a cautious approach, since excess mortality caused by diabetes is the result of other diseases, such as cardiovascular disease, we excluded the contribution of 'post-intervention' T2DM-specific mortality rates to in the overall mortality rates in the lifetable (see section on multistate lifetable modelling).


Figure 3. Effects of the elimination of exposure to high BMI with varied trend years, lower BMI boundary, and trends applied. This figure shows the number of HALYs gained and deaths avoided (mean and $95 \%$ uncertainty intervals) on elimination of exposure to high BMI with 1) varying the 25 years of linear trend applied to reference populations to $0,5,10,15$ and 20 years 2 ) varying the point at which the risk starts to rise (optimal BMI level) in the PIF calculation from BMI $22.5 \mathrm{~kg} / \mathrm{m}^{2}$ used in the main analysis (base case) to $20 \mathrm{~kg} / \mathrm{m}^{2}$ and $25 \mathrm{~kg} / \mathrm{m}^{2}$ and 3) varying the trend, halving and doubling the linear trend and application of a second order polynomial trend instead of linear. The mean is shown in bars. The whiskers show $95 \%$ uncertainty intervals, the 2.5 and 97.5 percentiles capturing sampling error with input data. No.: number, HALYS: Health adjusted life years, Yrs.: Years.

Our model did not incorporate new birth cohorts or migration. Nonetheless, the life table modelling includes the effects of population aging. Also, our uncertainty intervals reflect only uncertainty in some of the input parameters, and certainly not all uncertainty in the analysis.

Regarding data inputs, several Kenyan studies ${ }^{36-42}$ presented incidence and prevalence data for a few of our modelled diseases, but the GBD 2019 study provided the most recent national estimates for all our modelled diseases. ${ }^{1,15}$ It is one of the data sources commonly used in health policy. A recent example is the
just released National Non-Communicable Diseases Strategic Plan: 202I/22-2025/2643 where data from the GBD 2019 study are used to describe the current disease outlook in Kenya.

A further limitation of the study is that we used BMI as the measure of overweight and obesity. While BMI correlates well with fat accumulation and metabolic health in large populations, it does not differentiate between lean and adipose tissue, or fat distribution which varies across individuals, ethnicities, athletes, and throughout the lifespan. ${ }^{22,44-47}$

Measures considered the gold standard for quantitative assessment of visceral fat tissue such as computerized axial tomography (CT) and magnetic resonance imaging (MRI) ${ }^{48}$ are not feasible for population-based studies due to radiation exposure, equipment size and costs. BMI is reported widely and has been endorsed as a useful measure for monitoring populations due to its simplicity of assessment, robust nature of the measurements, higher precision and accuracy than, for example, other anthropometric measures for adiposity such as waist circumference. ${ }^{44,49}$

The evidence on the association of BMI with health outcomes is taken mainly from pooled analyses of primary studies conducted in high income countries. ${ }^{8,20,50}$ In principle, the TMREL for BMI may vary by age, sex, height, length of follow up, and location if supported by clear evidence. ${ }^{\text {I2,22 }}$ Studies investigating the optimal BMI for African populations would be insightful. In our study we incorporate the available evidence by varying the TMREL of BMI $22.5 \mathrm{~kg} / \mathrm{m}^{2}$ used in the PIF calculation to $20 \mathrm{~kg} / \mathrm{m}^{2}$ and $25 \mathrm{~kg} / \mathrm{m}^{2}$.

Similarly, we did not find published evidence of relative risk estimates specific for the Kenyan population apart from two studies that found obesity/overweight was associated with greater odds of diabetes compared to normal BMI levels. ${ }^{41,5 I}$ The reported risk estimates are largely comparable to the age- and sex-specific risk measures that we use in our model for diabetes (SF p. 38). Robust cohort studies from Kenya are needed. Nevertheless, evidence from individual-participant-data meta-analysis of 239 prospective studies from four continents suggests that risk measures are broadly similar in different populations. ${ }^{22}$ Also, studies that include additional measures of adiposity such as waist circumference and waist-to-hip ratio, may be more informative as proportions of body fat may be different for various ethnic groups at the same value of BMI. ${ }^{18,22}$

A strength of this study is that we use the proportional multi-state life table Markov model which is a well-established method that allows for long term estimation of health outcomes. ${ }^{33}$ We explicitly model changes in risk factor distribution, disease incidence, prevalence and mortality, including future life years while adjusting for disease-specific disability and comorbidity. ${ }^{24,33}$ Since the lifetable approach provides estimates for changes in life expectancy and, given the
time horizon (lifetime) used, this study provides a comprehensive picture of the avoidable burden related to high BMI. Furthermore, we express health changes in terms of summary measures of population health (HALYs), which enables comparison of our findings across other domains. Additional strengths of this study are that we include an extensive set of high BMI related diseases in our model and, the BMI data for the Kenya population was sourced from the 2015 Kenya WHO STEPS survey where standard procedures were used to measure height and weight. ${ }^{4}$ Additionally, only population-based studies that measured height and weight in Kenya were included in the NCD Risk Factor Collaboration (NCD-RisC) pooled data that we used to estimate the trend in mean BMI. ${ }^{3}$

This is the first study to quantify the avoidable burden related to overweight and obesity in Kenya in future years. A related study is the GBD which provides estimates of the deaths and disability adjusted life years attributable to high BMI for a single year period. ${ }^{52}$ Unlike our study that quantifies the future avoidable burden, the GBD estimates are due to past exposures. They are relevant when estimating future health care needs. But past and present exposures can no longer be changed, so for prevention policy, it is more relevant to how much of the future burden can be prevented, and what the consequences of inaction are likely to be. In our study, we estimate HALYs that could be gained from achieving optimal BMI levels, factoring in the rising trend in BMI and population ageing. Also, GBD studies estimate DALYs as the sum of years of life lost (YLL) and years lost due to disability. The mortality rates used in GBD studies to estimate YLL are for a hypothetical population that has the lowest observed mortality rate at every age, whereas we used the country specific mortality rates for the life years lived component. However, the two studies are similar in that our study also includes an extensive set of high BMI related diseases and both studies find a significant burden associated with high BMI. Globally, a few studies have investigated the avoidable disease burden related to high BMI. ${ }^{53}$ Notably, these studies are conducted in high income countries, estimates of the avoidable burden are limited to a number of diseases and projected to a limited number of future years. Our study focused on a lower middle-income country, included an extensive set of diseases related to high BMI and estimated the avoidable burden over the entire lifetime of the 2019 Kenyan population.

The guiding health policy and strategies in Kenya have adopted the global target set to halt the rising burden of overweight and obesity. ${ }^{\text {Io,II, } 43}$ Quantifying the avoidable BMI-related disease burden in Kenya may give impetus for policy makers to prioritise population level measures that reduce overweight \& obesity and related NCDs in the country despite the persisting communicable disease burden. Our
outputs may provide additional measurable impact indicators that can be used in future to monitor progress of the national NCD control strategies in Kenya and similar settings.

In Kenya, BMI status in the adult population was higher in the older aged, in women, those in urban residence, those with higher education and greater wealth. ${ }^{4,54}$ Nearly everyone (99.8\%) had insufficient fruits and vegetable intakes and $80.3 \%$ did not engage in sufficient physical activity. ${ }^{55}$

Strategies that encourage physical activity and healthy diets at the individual level may encourage maintenance of healthy lifestyles especially for those residing in urban areas and wealthier individuals where high BMI is more prevalent. In low- and middle-income countries (LMICs) such as Kenya, it is estimated that by 2030, NCDs will contribute to three times as many disability adjusted life years (DALYs) and nearly five times as many deaths as communicable diseases, maternal, perinatal and nutritional conditions combined. ${ }^{56}$ To avoid significant disease burden in the future, strategies that seek to create healthier environments that decrease the stimulus for over-consumption and, encourage physical activity are more effective than those that only target individuals. ${ }^{6,57,58}$

Research on the nutritional value of traditional diets in Kenya may be helpful in advocating for the retention of healthy traditional diets. ${ }^{59}$ Research on how city and transport planning in Kenya can be used to create healthy neighbourhoods would guide formulation of policies that create a healthy environment for all. Since Kenya is still battling undernutrition in some populations, studies investigating how stunting impacts BMI levels in adulthood are key.

Future modelling studies could estimate avoidable high BMI related disease burden at subnational level and by socioeconomic factors such as education level, wealth quintiles, those in urban versus rural residence. Finally, to inform the government's priority areas, further research is needed to identify the most impactful and cost-effective context-specific strategies for NCD prevention and control in Kenya. ${ }^{60}$

Our findings quantify the future disease burden that could be avoided if high BMI was eliminated as a risk factor in Kenya. The magnitude of avoidable high BMI related disease burden underscores the need to prioritise the control and prevention of overweight and obesity globally, not only high-income countries but also in Kenya and other low- and middle-income settings, where obesity rates are rising rapidly. Reducing population BMI is challenging, but sustained and wellenforced system-wide approaches could be a great starting point towards this goal.

Our results may provide additional measurable impact indicators that can be used in future to monitor progress of the national NCD control strategies in Kenya and similar settings.

## Contributors

MNW conceived the study, developed the study protocol under the supervision of JLV and LNA. MNW did the modelling analysis and wrote the first version of the manuscript. LNA and JLV contributed to analysis, interpretation of findings and reviewed successive versions of the manuscript. All authors critically reviewed the manuscript and approved the final version for publication.

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## Data sharing statement

All data generated or analysed during this study are included in this published article and its supplementary information file.

## Declaration of interests

The authors declare that they have no conflicts of interest.

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## Supplementary materials

Supplementary material associated with this article can be found in the online version at doi:Io.ior6/j. eclinm.2022.IOI522.

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