

Change in left atrial function predicts incident atrial fibrillation: the Multi-Ethnic Study of Atherosclerosis

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Aims	Longitudinal change in left atrial (LA) structure and function could be helpful in predicting risk for incident atrial fibril- lation (AF). We used cardiac magnetic resonance (CMR) imaging to explore the relationship between change in LA structure and function and incident AF in a multi-ethnic population free of clinical cardiovascular disease at baseline.
Methods and results	In the Multi-Ethnic Study of Atherosclerosis (MESA), 2338 participants, free at baseline of clinically recognized AF and cardiovascular disease, had LA volume and function assessed with CMR imaging, at baseline (2000–02), and at Exam 4 (2005–07) or 5 (2010–12). Free of AF, 124 participants developed AF over 3.8 ± 0.9 years (2015) following the second imaging. In adjusted Cox regression models, an average annualized change in all LA parameters were significantly associated with an increased risk of AF. An annual decrease of 1-SD unit in total LA emptying fractions (LAEF) was most strongly associated with risk of AF after adjusting for clinical risk factors for AF, baseline LA parameters, and left ventricular mass-to-volume ratio (hazard ratio per SD = 1.91, 95% confidence interval = 1.53–2.38, $P < 0.001$). The addition of change in total LAEF to an AF risk score improved model discrimination and reclassification (net reclassification improvement = 0.107, $P = 0.017$; integrative discrimination index = 0.049, $P < 0.001$).
Conclusion	In this multi-ethnic study population free of clinical cardiovascular disease at baseline, a greater increase in LA vol- umes and decrease in LA function were associated with incident AF. The addition of change in total LAEF to risk prediction models for AF improved model discrimination and reclassification of AF risk.
Keywords	atrial fibrillation • left atrium

Introduction

Atrial fibrillation (AF), the most common sustained cardiac arrhythmia, is increasing in prevalence as populations age.¹ It affects up to 9– 10% of those aged over 80,^{1,2} and contributes to significant morbidity, such as increased stroke rates,^{1,2} thrombo-embolic events,³ cognitive decline,⁴ and left ventricular (LV) dysfunction.⁵ Previous studies have shown that greater volumes and impaired function of the left atrium (LAm) are modestly associated with incident AF, independent of traditional risk factors.^{6–9} Adverse remodelling of the LAm has been proposed to facilitate both initiation and maintenance of AF by promoting ectopic triggers, and altering the wavelength of the re-entrant circuit,¹⁰ while recent studies have shown left atrial (LA) parameters to be useful in stratifying the risk of

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incident stroke in patients with AF.^{11,12} Thus, change in LA parameters may be helpful in risk stratification of incident AF in asymptomatic individuals.¹³

To date, few prospective studies have investigated the association between longitudinal change of the LAm with the development of AF, neither have they shown incremental benefit to existing AF risk prediction models.^{7,8} Furthermore, most studies utilized speckletracking echocardiography to analyse the LAm, which has proven challenging given the anatomic locations of the LAm, pulmonary veins and thin atrial wall.¹⁴ We used multimodality tissue tracking (MTT) on cardiac magnetic resonance (CMR) imaging, the gold-standard reference in atrial and ventricular volume measurement, and a wellestablished method for the assessment of cardiac deformation,^{15–17} to explore the relationship between change in LA parameters and incident AF. It is described as being superior to echocardiography in delineating static LA volumes, given its excellent ability to define the spatial resolution of endocardial and epicardial borders.^{15–17} In a population free of clinical cardiovascular disease at baseline, we hypothesized that an increase in LA volumes and decrease in LA function are associated with a greater risk of incident AF, after adjusting for known risk factors.

Methods

Study design

The Multi-Ethnic Study of Atherosclerosis (MESA) is a prospective, population-based multi-ethnic (White, African-American, Chinese, and Hispanic) cohort study of subclinical cardiovascular disease. The study design of MESA has been described in detail previously.¹⁸ In summary, between 2000–02 (Exam 1), 6814 men and women aged 45–84 years, free of clinical cardiovascular disease at enrolment, were recruited from six US field centres (Baltimore, MD; Chicago, IL, USA; Forsyth County, NC, USA; Los Angeles County, CA, USA; Northern Manhattan, NY; St Paul, MN, USA). Exam 1 was followed by Exam 2 (2002-04); Exam 3 (2004-05); Exam 4 (2005–07) and Exam 5 (2010–12). Approximately every 9 months, each participant was contacted by a telephone interviewer, as follow-up, to inquire about hospital admissions, cardiovascular outpatient diagnoses and mortality. Medical records and information were successfully obtained in 98% of reported hospitalized cardiovascular events and 95% of reported outpatient cardiovascular diagnostic encounters. The methodology of risk factor and outcome collection is detailed in Supplementary data online. All participants provided written informed consent. All study protocols were approved by the institutional review boards of each participating field centre.

The flowchart of the MESA population investigated in this study is illustrated in *Figure 1*. A baseline CMR study (Exam 1) was conducted in 4859 participants, measuring LA volumes, EF, and global peak longitudinal strain. Of these, 2338 eligible participants returned for a second CMR study (Exam 4/5) on average 9.4 ± 0.5 years later. Exam 5 measures were used in participants who had CMR at both Exams 4 and (n = 457), unless they developed AF in between Exams 4 and 5 (dotted line of *Figure 1*), whereby Exam 4 measures were used (n = 19). Participants were excluded if (i) they did not have at least two CMR studies (baseline and Exam 4/5) or (ii) due to unavailability or poor quality of images (n = 2421), or (iii) if they developed AF before the second study (Exam 4: n = 8; Exam 5: n = 92). Over a mean follow-up period of 3.8 ± 0.9 years after the second study, 132 participants developed incident AF (Exam 4: n = 19; Exam 5: n = 113). However, due to unavailability or poor quality of images (n = 8), only 124 AF cases were included in the event analysis.



Figure I Flowchart of study. (1): Baseline to second CMR study mean time: 9.4 ± 0.5 years (2): AF follow-up mean time: 3.8 ± 0.9 years after second CMR study *Exclusion: (1) did not have at least 2 CMR studies (baseline and Exam 4/5) or (2) images were unavailable or of poor quality (n = 2421), or (3) developed clinically recognized AF before the second CMR study (Exam 4, n = 8; Exam 5, n = 92). #Exclusion due to unavailability or poor quality of images (n = 8). AF, atrial fibrillation; CMR, cardiac magnetic resonance.

Identification of AF cases

Incident cases of AF during the follow-up period (2007/2012 (Exam 4/5)–2015) were identified through MESA surveillance and, for participants enrolled in fee-for-service Medicare, from inpatient and outpatient Medicare claims data. As a part of standard event surveillance procedures, all hospitalizations were identified during follow-up calls to study participants or a proxy. Discharge diagnosis and procedure codes from those hospitalizations were abstracted. AF was documented as present if an *International Classification of Diseases*, Ninth Revision diagnosis code 427.31 (AF) or 427.32 (atrial flutter) was present. If the first AF claim occurred before the baseline study, the participant was considered to have prevalent AF and therefore was excluded from the analysis.

CMR and MTT image analysis

Baseline CMR images were acquired using 1.5 T MR scanners: Signa LX or CVi (GE Medical Systems, Waukesha, WI, USA) or Symphony or

Sonata (Siemens Medical Systems, Erlangen, Germany). Long axis cine images were obtained from 2-chamber and 4-chamber views, using electrocardiogram-gated fast gradient-echo pulse sequence. All cine images were acquired with a temporal resolution of \sim 50 ms. A stack of shortaxis images recorded at end diastole was obtained for the assessment of LV mass. A detailed description of the MESA MRI protocol has been published.¹⁹

Multimodality Tissue Tracking software (MTT, v6.0, Toshiba, Japan) was used to quantify LA volume, EF, and strain from 2- and 4- chamber cine CMR images. This method has been validated previously with good-excellent [intraclass correlation (ICC); 0.88–0.98, P < 0.001] intra-and inter-reader reproducibility, and good (ICC; 0.44–0.82, P < 0.05– 0.001) inter-study reproducibility.^{20,21}

A single experienced operator blinded to the case status of the participant defined endocardial and epicardial borders of the LAm at end systole. Using the marked points, the software creates endocardial and epicardial borders, then tracks LA tissue in subsequent frames. The endocardial and epicardial contours generated by the software are then followed by the operator during the cardiac cycle for quality control. Maximum, pre-atrial, and minimum contraction LA volumes were extracted from volume curves that were created using the area-length method from apical 2- and 4-chamber views, using the following formula for Biplane calculation: Volume = $\frac{(0.848 \times \operatorname{areach} \times \operatorname{areach})}{([\operatorname{length}_{2Ch} + |\operatorname{ength}_{4Ch}]/2)}$. All LA volumes were indexed to body surface area (mL/m²), while LA emptying fractions (LAEF) were derived from LA volumes (Supplementary data online). Biplanar volume and function assessment on MTT had a strong positive linear correlation and concordance to other manual methods (e.g. Simpson's method).²¹

Strain measurement

MTT software calculates global longitudinal atrial strain by averaging longitudinal strain of all LA segments, determined by the software's automatic division of the LA wall into equal segments, in 2- and 4-chamber views during each cardiac cycle. Peak longitudinal LA strain (LASmax) was extracted from the global longitudinal strain curve (*Figure 2*). Calibration for strain measurements was accounted for in the MTT software as Exam 1 used Fast Gradient Recalled Echo (FGRE) images, while Exams 4 and 5 used Steady State Free Precision (SSFP) cine images.²²

Statistical analysis

Baseline characteristics of study participants are presented as mean \pm SD for continuous variables in *Table 1*. Annual change in LA variables was determined by averaging the change over time (years) between both studies, presented as mean \pm SD in *Table 2*. We used Cox proportional hazards regression models to study associations between annual change in LA variables and incident AF. The assumption of proportionality of hazards was confirmed for each model. The AF risk prediction score used was the CHARGE-AF (Cohort for Heart and Aging Research in Genomic Epidemiology-atrial fibrillation) risk model, previously validated using the Age, Gene and Environment-Reykjavik Study (AGES) and Rotterdam Study (RS).²³

Three models were generated to examine the associations between annual LA change and AF. In model 1, we adjusted for CHARGE-AF risk factors at the second CMR study (Exam 4/5): age, race, height, weight, systolic and diastolic blood pressure, antihypertensive medication use, smoking status, diabetes, and the development of myocardial infarction and congestive heart failure.²³ In model 2, we added baseline LA measures to account for baseline differences when measuring change, and potential measurement error bias.²⁴ Model 3 additionally adjusted for LV mass-to-volume ratio (MVR), a measure of LV diastolic dysfunction. The cumulative risk of AF over the follow-up years for the cohort, stratified by tertiles of LA parameters, was determined using the Kaplan–Meier curves, censoring at last follow-up. Differences across tertiles were compared using the log-rank test.

We compared model discrimination using the C-statistic for postsurvival analysis. The additional predictive value of annual LA change was calculated by the increment in the C-statistic, the categorical net reclassification improvement (NRI), and the integrative discrimination index (IDI).²⁵ Risk categories for NRI were defined a priori as <2.5%, 2.5–5.0%, and >5.0%, as estimated by the CHARGE-AF risk prediction model.²³ We assessed model calibration using Grønnesby and Borgan's modified Hosmer–Lemeshow x^2 statistic for survival analysis.²⁶ A *P*-value of <0.05 in two-tailed tests was considered statistically significant. All statistical analyses were performed using Stata software (version 12.0; Stata Corp, TX, USA).

Results

Population demographics of eligible participants (n = 2338) at baseline and the second CMR (9.4 ± 0.5 years after baseline), divided into those who developed AF (n = 132), and those who did not (n = 2206), are summarized in *Table 1*.

Annual change in LA structure and function

The annual change in LA variables over 9.4 ± 0.5 years are summarized in *Table 2*. Participants who developed incident AF had a greater increase in LA volumes (Δ LAVImax 1.06 ± 1.51 vs. 0.58 ± 1.14 mL/m²/year; P < 0.001; Δ LAVImin 1.05 ± 1.23 vs. 0.46 ± 0.77 mL/m²/year; P < 0.001) (*Table 2*). Similarly, peak LA strain and all LAEFs (total, passive, and active) decreased to a greater extent in AF cases (all P < 0.001; Δ Passive LAEF P = 0.015).

Association of change in LA variables with incident AF

The association of change in LA variables with incident AF over 3.8 ± 0.9 years of follow-up, analysed with multivariable Cox regression models, is summarized in *Table 3*. In the unadjusted model, all LA variables were significantly associated with incident AF, whereby an annual increase of 1-SD unit of Δ LAVImin was associated with the greatest increase in risk of incident AF [hazard ratio (HR) per SD = 1.47, 95% confidence interval (Cl) (1.31–1.64); *P* < 0.001] (Supplementary data online, Table S4).

After adjusting for CHARGE-AF risk factors in model 1, the association remained significant for all variables except for passive LAEF. After additionally adjusting for the baseline value of each LA variable in model 2, and LV MVR in model 3, all LA variables were significantly associated with incident AF (P < 0.001). We calculated the decrease in 1-SD unit of total LAEF(- Δ Total LAEF) and peak LA strain (- Δ Peak LA strain) to compare the increase in hazards per 1-SD unit for incident AF across the categories of volume (Δ LAVImin), EF and strain, as they showed the greatest magnitude of association with AF risk. An annual decrease of 1-SD unit in total LAEF was associated with the greatest increase in risk of incident AF after adjusting for clinical risk factors, baseline LA measures and LV MVR (HR per SD = 1.91, 95% CI 1.53–2.38, P < 0.001) (*Table 3C*). The Kaplan–Meier curves were plotted, stratified by tertiles of unadjusted Δ Total LAEF, statistically significant on the log-rank test (all P < 0.05) (*Figure 3*).





Improvement in risk prediction with addition of annual change in LA variables

The CHARGE-AF score demonstrated excellent performance for predicting AF in our population (C-statistic = 0.757, 95% CI 0.721– 0.794). We added Δ LAVImin, Δ Total LAEF, and Δ Peak LA strain separately to the CHARGE-AF model, as model 1, and additionally adjusted for baseline LA values in model 2, to investigate the additional discriminatory value. In model 2, Δ LAVImin had the highest discriminatory value (C-statistic = 0.787, 95% CI 0.747–0.824). Model 2 with Δ Total LAEF showed significant improvement to model discrimination and reclassification compared to model 2 alone (NRI = 0.107, P = 0.017; IDI = 0.049, P < 0.001) (*Table 4*).

Discussion

This study showed a relationship between total LAEF and incident AF: an average annual change in total LAEF showed the greatest association with incident AF risk in a multi-ethnic cohort, and improved model discrimination and reclassification in predicting risk of incident AF, after adjusting for clinical risk factors and baseline total LAEF. To our knowledge, the incremental model discrimination of LAEF to clinical AF risk scores has not been reported before. Our findings suggest that total LAEF may detect subtle and earlier structural, as well as functional, LA derangements in high-risk individuals, making it a more valuable risk predictor compared to structural LA remodeling (volume) alone (*Figure 4*). Validation of these findings were seen in a

Table I Population demographics

	Baseline (Exam 1)	Second study (Exam 4/5)) after 9.4 \pm 0.5 years
	n = 2338	No AF n = 2206	Incident AF n = 132
Age	59.3 ± 9.22	68.2 ± 8.97	75.1 ± 7.13
Male gender	1096 (47%)	1028 (47%)	68 (52%)
Race			
White, Caucasian	1005 (43%)	935 (42%)	70 (53%)
Chinese-American	275 (12%)	263 (12%)	12 (9%)
African-American	575 (24%)	545 (25%)	30 (23%)
Hispanic	483 (21%)	463 (21%)	20 (15%)
BMI	27.8 ± 4.95	28.1 ± 5.16	28.2 ± 5.12
Systolic blood pressure (mmHg)	123 ± 20.1	122.4 ± 19.8	125.2 ± 20.0
Diastolic blood pressure (mmHg)	71.8 ± 10.2	68.3 ± 9.8	67.5 ± 10.5
Antihypertensive medications	722 (31%)	1123 (51%)	94 (71%)
Glycaemic status			
Normal	1821 (79%)	1360 (62%)	75 (57%)
Impaired fasting glucose	261 (11%)	453 (21%)	28 (21%)
Diabetes mellitus	232 (10%)	376 (17%)	29 (22%)
Smoking			
Never	1221 (53%)	1035 (47%)	49 (37%)
Former	832 (36%)	992 (45%)	74 (56%)
Current	267 (11%)	166 (8%)	9 (7%)
Events			
Myocardial infarction	0	35 (2%)	4 (3%)
Congestive heart failure	0	19 (1%)	3 (2%)

AF, atrial fibrillation; BMI, body mass index.

recent investigation from the Framingham Offspring Study by Sardana *et al.* on LA Functional Index, a composite echocardiographic measure of LA structure and function, demonstrating a strong risk association with incident AF even in patients with normal LA volume; however, not significantly improving model discrimination on addition to AF prediction models.⁷

The association of structural and functional LA remodelling and incident AF has been extensively researched previously.^{6–9} Our findings of increasing LA volumes and decreasing LA function being associated with an increased risk of AF in community-based samples are consistent with current literature, in particular, showing minimum LA volume to have a higher risk association with AF compared to maximum LA volume, as seen in the study by Fatema *et al.*, which found minimum LA volume to be marginally superior to maximum LA volume in predictive ability of AF over a mean follow-up of 1.9 years.⁹

A previous investigation involving participants in the original Framingham Heart Study cohort, showed that a 5-mm increase in LA diameter was associated with a 39% increase in the HR for incident AF over a median follow-up of 7.2 years.⁸ Our study shows that an annual decrease of 1-SD unit in total LAEF was associated with a 91% increase in incident AF risk, after adjusting for AF risk factors, baseline LA values, and LV remodelling, differing from existing studies by providing an annualized rate of LA change and HRs per 1-SD unit. Annual LA change was averaged across 9 years, thus assumed to be linear over time, which may not have fully captured the noise in

	Annual cha (over 9.4 \pm		
LA variable	No AF n = 2206	Incident AF n = 132	P-value
ΔLAVImax (mL/m ² /year)	0.58 ± 1.14	1.06 ± 1.51	<0.001
Δ LAVIpreA (mL/m ² /year)	0.56 ± 0.96	0.97 ± 1.26	<0.001
Δ LAVImin (mL/m ² /year)	0.46 ± 0.77	1.05 ± 1.23	<0.001
Δ Total LAEF (%/year)	-0.69 ± 1.33	-1.32 ± 1.56	<0.001
Δ Passive LAEF (%/year)	-0.36 ± 1.09	-0.59 ± 1.07	0.015
Δ Active LAEF (%/year)	-0.65 ± 1.56	-1.22 ± 1.71	<0.001
Δ Peak LA strain (%/year)	-0.59 ± 1.66	-1.13 ± 1.59	<0.001

Table 2 Annual change in I A variables

 $\Delta,$ Annual change; AF, atrial fibrillation; EF, emptying fractions; Indexed volumes: maximum (VImax), pre-atrial (VIpreA), minimum (VImin); LA, left atrial.

year-to-year measurements, providing precedence for further investigations. The concept of dynamic change in risk profile, as patients age and accumulate exposure to risk factors, has been explored using other prediction models like the CHA₂DS₂-VASc, and HAS-BLED risk score, showing risk profile change to be superior to single baseline determinations.^{27–29} In this study, we investigated the predictive power of changes in parameters of LA structure and function to detect incident AF over and above their baseline magnitude.

	A. Model 1		B. Model 2		C. Model 3	
	Clinical risk (CHARGE-AF) ^a + ∆LA variable		Model 1 + Baseline LA variable		Model 2 + LV MVR	
Variable	HR (CI)	P-value	HR (CI)	P-value	HR (CI)	P-value
ΔLAVImax (mL/m ² /year)	1.24	0.003	1.36	<0.001	1.39	<0.001
	(1.08–1.42)		(1.19–1.56)		(1.21–1.60)	
Δ LAVIpreA (mL/m ² /year)	1.26	0.009	1.41	<0.001	1.42	< 0.001
	(1.06–1.50)		(1.19–1.67)		(1.20–1.69)	
Δ LAVImin (mL/m ² /year)	1.54	<0.001	1.64	<0.001	1.64	<0.001
	(1.30–1.83)		(1.40–1.91)		(1.40–1.92)	
Δ Total LAEF (%/year)	0.84	0.008	0.63	<0.001	0.62	<0.001
	(0.75–0.96)		(0.54-0.74)		(0.53–0.73)	
Δ Passive LAEF (%/year)	0.96	0.632	0.63	0.001	0.60	<0.001
	(0.79–1.15)		(0.48-0.81)		(0.45–0.79)	
Δ Active LAEF (%/year)	0.87	0.018	0.67	<0.001	0.67	<0.001
	(0.77–0.98)		(0.57-0.79)		(0.57-0.79)	
Δ Peak LA strain (%/year)	0.90	0.048	0.72	<0.001	0.70	<0.001
	(0.81–0.99)		(0.62-0.85)		(0.59–0.83)	
Δ LAVImin (per 1-SD)	1.42	<0.001	1.49	<0.001	1.49	<0.001
	(1.24–1.63)		(1.31–1.69)		(1.31–1.69)	
- Δ Total LAEF (per 1-SD)	1.26	0.008	1.86	<0.001	1.91	<0.001
	(1.06–1.49)		(1.50–2.31)		(1.53–2.38)	
- Δ Peak LA strain (per 1-SD)	1.20	0.048	1.71	<0.001	1.81	<0.001
	(1.01–1.43)		(1.31–2.24)		(1.36–2.41)	

Table 3 Association of annual LA change with incident AF in Cox regression models

Δ, Annual change; AF, atrial fibrillation; EF, emptying fractions; Indexed volumes: maximum (VImax), pre-atrial (VIpreA), minimum (VImin); LA, left atrial. ^aCHARGE-AF risk model: age, race, height, weight, systolic and diastolic blood pressure, antihypertensive medication, smoking status, diabetes, myocardial infarction, and con-

gestive heart failure by the second study.²⁴





Table 4	Model discrimination,	calibration	, NRI	, and IDI

Model 2: CHARGE-AF risk factors + Baseline LA variable					
	CHARGE-AF	Δ LAVI min (mL/m²/year)	Δ Total LAEF (%/year)	$\Delta \mathbf{P}$ eak LA strain (%/year)	
C-statistic (95% Cl)	0.757 (0.721–0.794)	0.787 (0.749–0.824)	0.779 (0.737–0.820)	0.770 (0.732–0.808)	
NRI ^a (P-value)		0.000 (0.99)	0.107 (0.017)	0.017 (0.63)	
IDI (P-value)		0.049 (<0.001)	0.049 (<0.001)	0.018 (<0.001)	
Calibration x^2 (<i>P</i> -value)	19.3 (0.02)	11.68 (0.23)	5.751 (0.77)	3.673 (0.93)	

Δ, Annual change; AF, atrial fibrillation; EF, emptying fractions; IDI, integrative discrimination index; LA, left atrial; NRI, net reclassification improvement; Vlmin, minimum indexed volume.

^aCategories of NRI:<2.5%, 2.5–5.0%, and >5.0%.²⁴



Figure 4 Change in total left atrial emptying fractions, derived by multimodality tissue tracking on CMR imaging, showed incremental risk prediction and reclassification of incident AF on addition to an AF risk prediction model.

The observation that annual LA change precedes the development of AF, warrants further investigation into the pathophysiology of AF. It has previously been suggested that LA structural and functional remodelling both lead to, and are caused by, AF. The resultant atrial deformation and fibrosis plays a role in altering electrical conduction properties, which may predispose to the development of arrhythmogenic micro-re-entrant circuits and triggered ectopic activity.¹⁰ However, this is but one of the many pathophysiological factors contributing to AF. Electrical remodelling, autonomic nervous system changes, and calcium-handling abnormalities all result in a re-entry prone substrate and atrial ectopic activity, contributing to the initiation and maintenance of AF.¹⁰ While it is possible that impaired total LAEF may be an indicator of any of the changes listed above, evidence that it precedes the development of AF does not necessarily imply causality. Abnormalities in LA volume or total LAEF may simply be markers for unidentified factors that are causally related to the

development of AF. How functional and structural LA remodelling contributes to these known mechanistic processes, or confers directionality to the development of AF, will require further studies for elucidation.

Limitations

We used linear, instead of 3D, methods to measure LA volumes, which may underestimate true volumes by 11.5–20%.³⁰ However, this method has been widely used and validated in research studies.^{20,21} Diagnosing incident AF based on hospital discharge codes meant a possible underestimation of subclinical AF that did not result in hospitalization, while subclinical episodes of AF between both CMR measures may have contributed to LA remodelling prior to clinical AF detection. However, a validation sub-study of 45 MESA participants with the diagnosis of AF based on hospital discharge codes showed that AF was confirmed in 93% on review of Medicare, implying a high specificity for the diagnosis.³¹ We also acknowledge that missing data from participants who did not participate in Exams 4 or 5 could have introduced bias in the study, however, as they tended to be older and had more risk factors, we hypothesize that their inclusion would be more likely to increase the strength of our associations.

Conclusion

In this multi-ethnic study population free of clinical cardiovascular disease at baseline, a greater annual increase in LA volume and annual decrease in LAEFs and strain, measured on CMR imaging, were associated with increased incident AF risk during follow-up. The addition of change in total LAEF to risk prediction models for AF showed improvement to model discrimination and reclassification of AF risk. Future studies should validate our findings to better understand this contributory role to the pathophysiology of AF.

Supplementary data

Supplementary data are available at European Heart Journal - Cardiovascular Imaging online.

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