Review

The accuracy of pain drawing in identifying psychological distress in low back pain—systematic review and meta-analysis of diagnostic studies

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Abstract. [Purpose] The aim of this systematic review and meta-analysis was to estimate the accuracy of qualitative pain drawings (PDs) in identifying psychological distress in subacute and chronic low back pain (LBP) patients. [Subjects and Methods] Data were obtained from searches of PubMed, EBSCO, Scopus, PsycINFO and ISI Web of Science from their inception to July 2014. Quality assessments of bias and applicability were conducted using the Quality of Diagnostic Accuracy Studies-2 (QUADAS-2). [Results] The summary estimates were: sensitivity=0.45 (95% CI 0.34, 0.61), specificity=0.66 (95% CI 0.53, 0.82), positive likelihood ratio=1.23 (95% CI 0.93, 1.62), negative likelihood ratio=0.84 (95% CI 0.70, 1.01), and diagnostic odds ratio=1.46 (95% CI 0.79, 2.68). The area under the curve was 78% (CI, 57 to 99%). [Conclusion] The results of this systematic review do not show broad and unqualified support for the accuracy of PDs in detecting psychological distress in subacute and chronic LBP. **Key words:** Pain drawing, Stress psychological, Low back pain

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INTRODUCTION

A pain drawing (PD) is a simple line drawing of the human body (front and back)¹⁾ on which patients can indicate their pain for both clinical and research purposes²⁾. PDs may differ from each other depending on whether the image is simple³⁾ or detailed^{4, 5)}, presents a frontal or sagittal perspective, or depicts a female or male figure⁶⁾. PDs also use different methods for making notations and scoring the drawing^{3, 4, 7)}. Although PDs are commonly used to identify pain site, magnitude, and regional extension, their use to identify psychological distress of patients has also been proposed⁸⁾. It is well documented that psychological factors can contribute to how a patient first reports low back pain (LBP) and influence the transition towards chronic pain^{9–11)}. Moreover, psychological distress has a relevant influence on patients' responses to a variety of therapeutic approaches to LBP, including conservative treatment^{12, 13}), chemonucleolysis and surgery⁴).

Following the scoring method devised by Ransford et al.⁷⁾, who concluded that PDs could be used to identify patients with elevated hysteria and hypochondriasis scores on the Minnesota Multiphasic Personality Inventory (MMPI), based on the conformity of the patients' markings of typical or expected pain patterns, other investigators have also proposed using PDs to evaluate an individuals' underlying psychological states, in order to predict outcomes and to avoid unnecessary procedures for LBP patients. Udén et al.⁸⁾ introduced an alternative scoring method that classified drawings as "normal" or "organic" versus "abnormal" or "non-organic".

The validity of PD scoring systems for measuring psychological status was systematically reviewed by Carnes et al.¹⁴⁾ who reported on the sensitivity, specificity and positive and negative predictive values of abnormal drawings. They concluded that the available data did not support the assumption that unusual pain drawings or extensive markings indicated a disturbed psychological state, and as a consequence, recommended against using PDs as psychological assessment tools. Given that Carnes' systematic review¹⁴⁾ included dysfunctions in different body structures, a precise conclusion about PDs cannot be drawn with respect

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to LBP. Furthermore, Carnes' analysis differentiated scoring methods and included reference tests covering a large spectrum of psychological disturbances. No meta-analysis was performed and only English-language articles were included, raising the risk of publication bias. Finally, it is important to emphasize that the database search in Carnes' review was performed in the first three months of 2003 and subsequently, different authors have expressed contrasting opinion about on the influence of psychosocial factors on PD in LBP^{15–18)}. In particular, Abbott¹⁶⁾ and Turk¹⁸⁾ have argued that for some patients, the drawing might be influenced by psychosomatic disorders, and could be used to alert medical and health professionals.

For all these reasons, another systematic review and specific meta-analysis would be a useful contribution to the assessment of the validity of PDs in identifying psychological distress among subacute and chronic LBP patients. This systematic review expands upon previous studies explicitly targeting a particular population of interest, and focuses on the specific characteristics of outcome measures as inclusion criteria.

SUBJECTS AND METHODS

Our literature search aimed to identify all available studies that evaluated the correlation between qualitative or semiqualitative PDs scored using only the methods described by Ransford or Udén and psychological distress (anxiety, depression or illness behavior). Records were identified by searching multiple literature databases, including PubMed, EBSCO, Scopus, PsycINFO and ISI Web of Science from their inception to July 2014. The search terms used were: 'pain drawing*' OR 'pain diagram*' OR 'pain measurement*' OR 'pain chart*' OR 'body chart*' OR 'body map*' OR 'mannequin*' OR 'manikin*' OR 'mannikin*' OR 'self-assessment*' combined with 'spinal pain' OR 'back pain' OR 'lumbar pain' OR 'lumbo-pelvic pain' OR 'lumbago' and with 'reliability' OR 'reproducibility of findings'. These keywords were identified after preliminary literature searches. Additional records were searched through other sources to complement the database findings. Two reviewers (AR, AR) independently applied the previously determined inclusion and exclusion criteria to select potentially relevant papers which were been initially identified based on title and abstract. Full-text copies of relevant trials were then obtained and independently evaluated by the reviewers. When a disagreement between reviewers occurred, it was resolved by a meeting held in consultation with another author (CV).

Types of studies – Published diagnostic studies without any restrictions on publication date or language were included.

Types of participants – The participants in selected studies had to be symptomatic adults, 18 years of age or older, with a diagnosis of subacute or chronic specific or non-specific LBP. Pain lasting between 1 and 3 months was categorized as "subacute," and as "chronic" when lasting for more than 3 months. In the absence of this explicit description, pain was considered subacute or chronic when the investigators themselves categorized a subject's pain in those terms. Trials were excluded if subjects with spinal pain without any distinction among cervical, thoracic or lumbar locations were included; if subjects had received surgery or if any of the participants had received a diagnosis such as myelopathy, fracture, infection, dystonia, tumor, inflammatory disease, or osteoporosis.

Index tests – Among all the types of PDs described in the literature, our analysis was limited to only those qualitative or semi-qualitative PDs that used the methods of Ransford or Udén, either exactly as originally described or with slight variation. These PDs have shown good intra- and inter-examiner reliability^{3, 8, 19–23}). Studies concerning other kinds of PDs (e.g. quantitative or descriptive) were excluded. When trials used both Ransford's or Udén's methods and different methods (i.e. quantitative PDs), only data from qualitative or semi-qualitative PDs were taken into consideration. Moreover, if localizing the site or the area of pain or formulating a prognosis were the main clinical purposes of the study, these trials were excluded.

Reference tests – A further criterion of inclusion was the comparison between PDs and measures of psychological distress. To be eligible for inclusion, studies had to assess anxiety or depression or the presence of an illness behavior according to Waddell's signs and symptoms²⁴⁾. Only studies using anxiety or depression scales which demonstrated good psychometric properties and Waddell's signs and symptoms were considered for inclusion. When studies used multidimensional scales, only subscales related to anxiety or depression were considered. Other outcome measures like performance tests, disability measures, medico-legal issues, or imaging were excluded.

Two authors (AR, AR) independently conducted data extraction. Three other authors (CV, LB, PP) were consulted in the case of persisting disagreement. Reviewers were not blinded to information regarding the authors, journal of origin, or the outcomes of r each paper reviewed. Using a standardized form, data extraction addressed participants, type of PD, reference tests, and the findings that were reported. The methodological quality of studies was assessed by two authors (JHV, CV) using the QUADAS-2 scale, which has been shown to be reliable and valid for rating the quality of diagnostic studies²⁵⁾. Trials were not excluded on the basis of quality.

Standard methods recommended for meta-analyses of diagnostic studies were used²⁶). ProMeta V.2.0 and Review Manager V.5.2 software were used for the statistical analyses. For each study, computed measures of test accuracy using standard methods were used: Sensitivity, specificity, positive likelihood ratio (LR⁺), negative likelihood ratio (LR⁻) and diagnostic odds ratio (DOR)²⁶⁾. These measures were pooled using a random effects model to forecast the heterogeneity of the comparisons. Since sensitivity and specificity are correlated, their joint distributions were summarized using a summary receiver operating characteristic curve (SROC). SROC plots display the results of individual studies in ROC space with each study plotted as a single sensitivity-specificity point. The size of the points depicted the precision of the estimate, typically scaled according to the inverse of the standard error of the logit (sensitivity) and logit (specificity), or according to their sample sizes. The area under the SROC curve is a global measure of overall performance. An area under the curve of 1 indicates perfect discriminatory ability^{26, 27)}. The diagnostic odds ratio (DOR) summarizes the diagnostic accuracy of the index test as a single number that describes how many times higher the odds are of obtaining a positive test result for a person with disease compared to a person without disease. Because the SROC curve and the DOR are not easy to interpret and use in clinical practice, likelihood ratios were also considered, since they are more clinically meaningful²⁷⁾. Likelihood ratios greater than 10 or less than 0.1 generate large and often conclusive changes from pre-test to post-test probability; likelihood ratios of 5 to 10 and 0.1 to 0.2 generate moderate shifts in pre-test to post-test probability; likelihood ratios of 2 to 5 and 0.5 to 0.2 generate small (but sometimes important) changes in probability; and likelihood ratios of 1 to 2 and 0.5 to 1 alter probability to a small (and rarely important) degree²⁸⁾. Heterogeneity was described using the Q and I-square statistics for pooling sensitivity, specificity, likelihood ratios and diagnostic odds ratio. A significant Q value indicates a lack of homogeneity of findings of studies, but considering the low number of studies for each outcome and knowing that this test is not very indicative in this case, a random-effects model was used. Publication bias was assessed using the Egger's weighted regression method²⁹ with precision (1/standard error) and log odds ratio plotted.

RESULTS

The database search identified 6,418 studies. Additional eligible studies (n=7) were identified through other sources. After removing duplicates, and then screening titles and abstracts of all the remaining unique articles, 15 full-text articles needed to be assessed to verify their eligibility for the inclusion in the present study. Ultimately, eight of these manuscripts were excluded for various reasons (surgical patients, pain not present at the time of the evaluation, general spinal pain, reference tests not evaluating psychological distress or illness behaviour, missing outcome data, or prognostic study). Thus, seven studies were finally selected for this review (Fig. 1)^{30–36}.

Overall, the seven papers reported on studies conducted in Europe (Germany, Sweden, UK), Asia (India), and North America (Canada, USA), and had been published between 1983 to 2005, with only 28.6% of them published after 2000. These studies involved a total of 1,622 patients who were enrolled and completed assessments (Table 1). Study samples ranged from 54 to 649 subjects with a mean sample size of 232 participants. The mean age of participants was specified only in four studies, and it was approximately 44.5 years (range: 41.9–47.0). The sex of the participants was indicated in 4 studies, in which a slight majority (56%) of the participants were male.

The quality of the eligible studies was assessed using the revised Quality Assessment for Studies of Diagnostic Accuracy (QUADAS-2) tools²⁷⁾. Overall, the quality of the studies was poor to moderate (Table 2). For the patient selection domain, three studies^{30, 33, 34)} had a high risk of bias because of inappropriate inclusion or exclusion criteria. Four studies^{30, 33–35)} raised substantial concerns regarding their applicability to our research question. For the index test



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Fig. 1. Flow diagram of studies through the different phases of the review

domain, three studies^{30, 34, 35)} had a high risk of bias. None of the studies specified the threshold used. For two studies, there were strong concerns regarding the applicability of a study^{30, 34)} because variations in execution or interpretation of the PD may have affected estimates of its diagnostic accuracy. With respect to the QUADAS-2 criterion of risk of bias relative to the reference standard, the reference standard for all studies was considered to be appropriate. However, one study³¹⁾ raised concerns regarding the applicability of the reference standard to the research question. The most common methodological concerns was the failure to report patient flow and timing which occurred in 87% of the studies^{30, 31, 33–36)}.

All seven studies, including 10 pair-wise comparisons, were included in the meta-analysis and their sensitivities and specificities are shown on the forest plot (Fig. 2). The forest plot shows the studies in order of sensitivity. Study-specific estimates of sensitivity and specificity are shown, with their 95% confidence intervals. Figure 2 shows the sensitivity and the specificity of each study with point estimates and confidence intervals. Sensitivity estimates were lower and more variable (range 4–86%) than those of specificity (range 40–100%). The area under the curve was 78% (CI, 57–99%), indicating that PDs demonstrated only a modestly acceptable discriminatory power to identify psychological distress in subacute and chronic LBP patients. Table 3 shows the summary measures of test accuracy with 95% CIs and tests for heterogeneity. Publication bias was not evident in the

Author	Participants	Outcome measures	Reported results
Von Baeyer CL (1983) ³⁵⁾	. 212 chronic LBP patients. 2 samples: California (111), and Saskatchewan (101) heterogeneous as to diagnosis and chronicity of back pain.	 Pain drawing by Ransford MMPI (Minnesota Multiphasic Pain Inventory) 	Not dichotomous data: higher correlation coefficient = 0.28 for the Hypochondriasis (Hs) scale ($p<0.01$) Dichotomous data: regarding Hs, 61% agreement of the cases overall, but only 44% of patients with elevated Hs score were correctly identified by their drawings scores.
Hildebrandt J. (1988) ³¹⁾	54 subjects (32 men, 22 women) randomly chosen among all chronic LBP patient visited in the authors' pain clinic within 1 month. Mean age: 46.4 years (men) and 44.5 years (women).	 Pain drawing slightly changed from Ransford scoring method. Erweiterte Revidierte Mehrdimensionale Schmerzs- kala (ERMSS) 	No correlation between PDs scores (> or <2) and affective score of ERMSS.
Lindal E. (1988) ³²⁾	Group 1: 54 subjects with LBP (22 men, 32 women) randomly contacted without any prior knowledge of their medical history. Mean age: 46.8 years. Group 2: 13 subjects with LBP scheduled for back surgery. Mean age: 51.1 years.	 Pain drawing by Udén. Rod-and-Frame Test (RFT) Metacontrast Technique (MCT) Eysenck Personality Inventory 	PDs classified as non-organic and possibly non-organic more common in group 2 (p <0.01) No significant correlation between the abnormality of PD and depression (MCT).
Chan CW. (1993) ³⁰⁾	49 subjects with chronic LBP (387 men, 264 women) classified as worker's compensation LBP patients (484) and as medicolegal LBP patients (167).	 Pain drawing by Udén with patients sorted by individually pooled pain categories Waddell score 	51.3% of patients with Waddell score <2 had nonorganic PD; 94.6% of patients with Waddell scores = 5 had nonor- ganic PD; 81.7% of patients with high Waddell scores had nonorganic PD.
Parker H. (1995) ²³⁾	Cohort 1: 100 subjects with LBP aged up to 65 years, without any major concurrent psychiatric illness, rheuma- tologic or neoplastic disease, or severe spinal structural abnormality. Cohort 2: 100 subjects with chronic LBP aged up to 65 years, without any major concurrent psychiatric illness, rheumatologic or neoplastic disease, or severe spinal struc- tural abnormality.	 Pain Drawing by Ransford, Pain Drawing by Margolis and Body Map Scoring System Modified Somatic Perception Questionnaire (MSPQ) Modified Zung (ZUNG) Modified Zung (ZUNG) Behavioral Signs (INSIGN) Behavioral Symptoms (INSYMPT) Distress and Risk Assessment Method (DRAM) 	The Modified Ransford scores did not show consistent cor- relations with the measures of psychological distress.
Pfingsten M. (2003) ³⁴⁾	109 chronic LBP patients (62 women 57 men), median age 41.9 years, mean duration of pain 49.9 months. Subsample composed of 36 patients: group 1 (clearly with-out psychological distress) = 22 patients; group 2 (clear psychological distress) = 14 patients.	 Pain drawing by Ransford Depression score (ADS) Affective score (SES/A) Other scales used: Numerical Rating Scale, Pain Disability Index 	Not dichotomous data: mild correlation (of a score >2) with ADS (r= 0.15) and SES (r=0.27). Dichotomous data: Ransford's method showed sensitivity = 86% and specificity = 46% .
Pande KC (2005) ³⁶⁾	331 patients (123 women, 208 men), with non-specific LBP, degenerative disc disease or lumbar stenosis. Mean age: 47.8 ± 15.1 (men) and 46.1 ± 13.4 (women)	 Pain drawing by Ransford Hospital Anxiety and Depression Scale (HADS) 	1) Performance of Pain Drawing respect in identifying patients with any degree of depression: sensitivity = 62% , specificity = 40% , positive predictive value = 69% 1) Performance of Pain Drawing respect in identifying patients with any degree of anxiety: sensitivity = 73% , specificity = 70% , positive predictive value = 78%

Table 1. Characteristics of included studies

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STUDY	RISK OF BIAS				APPLICABILITY CONCERNS		
	PATIENT SELECTION	INDEX TEST	REFERENCE STANDARD	FLOW AND TIMING	PATIENT SELECTION	INDEX TEST	REFERENCE STANDARD
Chan, 1993	8	\odot	\odot	\odot	\odot	$\overline{\otimes}$	\odot
Hildebrandt, 1987	\odot	\odot	\odot	$\overline{\otimes}$	\odot	\odot	$\overline{\otimes}$
Lindal, 1998	\odot	\odot	\odot	\odot	\odot	\odot	\odot
Pande, 2005	\odot	\odot	\odot	$\overline{\otimes}$	\odot	\odot	\odot
Parker, 1995	0	\odot	Ô	\odot	8	\odot	\odot
Pfingsten, 2003	8	$\overline{\otimes}$	\odot	$\overline{\otimes}$	$\overline{\otimes}$	8	\odot
Van Baeyer, 1983	?	$\overline{\otimes}$	\odot	$\overline{\odot}$	$\overline{\otimes}$?	?
	High Ris	k 20	nclear Risk				

Table 2. Methodological quality of included studies. Tabular presentation of OUADAS-2 results

0.23 [0.10, 0.42] Hildebrandt et al 1988 0.43 [0.24, 0.63] 0.58 [0.37, 0.77] 11 16 15 12 Pande et al (A) 2005 102 29 134 66 0.43 [0.37, 0.50] 0.69 [0.59, 0.79] Von Baeyer et al 1983 20 62 82 0.44 [0.34, 0.53] 0.80 [0.71, 0.88] 48 Pande et al (B) 2005 67 133 41 90 0.62 [0.52, 0.71] 0.40 [0.34, 0.47] 0.82 (0.74, 0.88)

0.86 [0.57, 0.98]

ensitivity (95% CI)

0.04 [0.00, 0.20]

0.08 [0.01, 0.27] 0.13 [0.06, 0.24]

Specificity (95% CI)

0.88 [0.69, 0.97]

0.92 [0.75, 0.99]

0.50 [0.23, 0.77]

1.00 [0.82, 1.00]

0.49 (0.44, 0.53)

0.45 [0.24, 0.68]

Sensitivity (95% CI)

0.2 0.4 0.6 0.8

Fig. 2. Forest plot of the estimates of sensitivity and specificity of PDs in identifying psychological distress in subacute and chronic LBP

Each solid square indicates the point estimate of sensitivity and specificity calculated for each study with 95% confidence intervals. Studies are presented in descending order of sensitivity

Table 3. Summary measures of test accuracy for all studies and tests of heterogeneity

Study

Lindal et al (A) 1988

Lindal et al (B) 1988

Parker et al (A) 1995

Parker et al (B) 1995

Chan et al 1993

Pfingsten et al 2003

ТΡ FP FN ΤN

2

8 7

7

103 268 23 254

12 12 2 10

3

2

n 23 19

24 22

22

54

24

All comparisons	Summary measure	Test for
(n=10)	of test accuracy*	heterogeneity [†]
	(95% CI)	(95% CI) p value
Sensitivity	0.45 (0.34, 0.61)	*
Specificity	0.66 (0.53, 0.82)	*
LR^+	1.23 (0.93, 1.62)	*
LR-	0.84 (0.70, 1.01)	*
DOR	1.46 (0.79, 2.68)	*

*Random effects model; CI: confidence interval; † Q test for heterogeneity:

LR+: positive likelihood ratio; LR-: negative likelihood ratio; DOR: diagnostic odds ratio

meta-analysis of the seven studies (t = -1.09; p=0.31).

DISCUSSION

This study aimed to search and review the scientific evidence for the validity of qualitative PDs in identifying psychological distress in subjects with subacute or chronic LBP. Seven diagnostic studies with large sample sizes, mostly involving adults with non-specific LBP, were selected and subjected to meta-analysis. The summaries of sensitivity and specificity, the LR⁺ and LR⁻ values, as well as the large range of the results (sensitivity range = 4-86%; specificity range = 40-100%) do not allow us to confirm the supposed discriminative power of PDs without reservation. Moreover, the area under the ROC curve of 78% suggests that there is only modest validity in using PDs to detect anxiety, depression or illness behaviour in LBP patients. Furthermore, on the basis of the OUADAS-2 results, all of the data were drawn from poor to moderate quality studies.

1 0 0.2 0.4 0.6 0.8 1

Although the index tests the selected studies were homogeneous (i.e., only qualitative PDs evaluated according to Ransford or Udén were selected), very different reference tests were employed to detect psychological disturbances. In fact, the Waddell's score, the Minnesota Multiphasic Pain Inventory, the Erweiterte Revidierte Mehrdimensionale Schmerzskala, the Metacontrast Technique, the Modified Somatic Perception Questionnaire, the Modified Zung Questionnaire, the Behavioral Signs, the Behavioral Symptoms, the Distress and Risk Assessment Method, the Depression score, and the Affective score were all used in the selected studies. These different reference standards document anxiety, depression, or illness behaviour across a wide conceptual swath.

One potential limitation of our study concerns our decision to analyze non-homogeneous outcome measures together. This option was preferred by our team because the specific relevance of any single one of these mood or behavioral disturbances in inducing a shift from acute to chronic LBP, or in obstructing recovery, has been elusive¹²). Furthermore, two selected studies^{32, 36)} analyzed both anxiety and depression within their samples, and failed to find substantial differences in the accuracy of PDs in identifying anxious or depressed subjects. This evidence strongly suggests that the diagnostic accuracy of PDs relative to psychological distress does not change whether the types of distress are assessed separately or together.

Our conclusion dawn from the few selected studies of PDs is similar to that of the previous systematic review conductedby Carnes et al.¹⁴, who also considered the quality of the studies that provided the data for analysis. In the light of our results, the current literature does not support the broad and unqualified use of qualitative PDs to identify psychological distress in subacute or chronic LBP, despite the value PDs may have in assessing symptoms, pain patterns or other clinical phenomena. However, future meta-analyses using higher quality studies may lead to different conclusions.

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