# Sodium-glucose cotransporter 2 inhibitors and fracture risk in patients with type 2 diabetes mellitus: a meta-analysis of randomized controlled trials

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# Abstract

**Background:** Patients with type 2 diabetes mellitus (T2DM) have an increased risk of fracture compared with those without T2DM. Some oral glucose-lowering agents may increase the incidence of fracture. Whether sodium-glucose co-transporter 2 inhibitors (SGLT2is) are associated with increased risk of fracture remains unclear.

**Methods:** We retrieved articles from *PubMed, Embase, Cochrane Library* database, and other sources up to 24 October 2019. We included randomized controlled trials (RCTs) that reported fractures and analyzed the fracture incidence of SGLT2i, canagliflozin, dapagliflozin, and empagliflozin. Subgroup analysis was also performed based on baseline characteristics. **Results:** A total of 78 RCTs with 85,122 patients were included in our analysis. The overall SGLT2i fracture incidence was 2.56% *versus* 2.77% in the control group [odds ratio (OR), 1.03; 95% confidence interval (CI) (0.95, 1.12); p=0.49]. Compared with the control treatment, treatment with canagliflozin led to a higher rate of fractures [OR, 1.17; 95% CI (1.00, 1.37); p=0.05], but no significant difference was observed when compared with dapagliflozin [OR, 1.02; 95% CI (0.90, 1.15); p=0.79] or empagliflozin [OR, 0.89; 95% CI (0.73, 1.10); p=0.30]. Subgroup analysis showed that, in a follow-up of less than 52 weeks, SGLT2i decreased the incidence of fracture by 29% [OR, 0.71; 95% CI (0.55, 0.93); p=0.01], but this benefit was lost when the follow-up extended to more than 52 weeks [OR, 1.08; 95% CI (0.98, 1.18); p=0.12]. **Conclusion:** Canagliflozin seems to increase the risk of fracture, while other SGLT2is do not result in a higher incidence of fracture.

Keywords: canagliflozin, dapagliflozin, empagliflozin, fracture, SGLT2i

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## Introduction

Patients with type 2 diabetes mellitus (T2DM) have an increased risk of fracture compared with those without T2DM. In addition to common hip and spine fractures, femur, femoral neck, and pelvic fractures are often reported in observational studies.<sup>1–3</sup> This could be attributed to many factors, including bone mineral density (BMD), bone turnover, microarchitecture, and material properties.<sup>4,5</sup> Elderly individuals are more likely to experience both fractures and T2DM. Complications of T2DM can lead to delayed union or to non-union of fractures.<sup>6,7</sup>

Some oral glucose-lowering agents such as thiazolidinediones (especially rosiglitazone)<sup>8–10</sup> may increase the incidence of fracture, and further study into their effects in T2DM patients is warranted to mitigate this risk.

Sodium-glucose co-transporter 2 inhibitors (SGLT2is) are novel glucose-lowering agents that lower blood glucose by inhibiting glucose and sodium reabsorption in the proximal tubule of the kidney. In addition to lowering glucose, SGLT2is can also lead to weight loss, decreased blood pressure, and reduction in serum uric acid.<sup>11–13</sup> Recent

#### Meta-analysis

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clinical studies have demonstrated that SGLT2is can lower the risk of mortality, heart failure, renal failure, and cardiovascular events.14,15 Indeed, the 2019 European Society of Cardiology Guidelines recommend SGLT2i as a priority for patients with T2DM and cardiovascular diseases (CVDs), or at very high/high cardiovascular (CV) risk, to reduce CV events.<sup>16</sup> However, the CANVAS study found that the canagliflozin group had a higher incidence of bone fracture than the placebo group (15.4% versus 11.9% per year, p=0.02).<sup>17</sup> A subsequent study attempted to determine the reasons for this increase in fracture risk, but did not succeed.<sup>18</sup> A study with a follow-up period of 104 weeks found that SGLT2i may be related to an elevated incidence of fracture (7.7% versus 0%).<sup>19</sup> Another study indicated that the decrease in BMD in the canagliflozin group appears to be associated with increased fracture risk,<sup>20</sup> while other studies found that SGLT2is could lower the incidence of fracture.21,22 Given the controversial results of SGLT2is on fracture events, we sought to synthesize all available data to investigate the safety of SGLT2is.

# Materials and methods

# Search strategy

The following keywords were used in the literature search: "Sodium-Glucose Transporter 2 Inhibitors", "Dapagliflozin", "Canagliflozin", "Empagliflozin", "Ipragliflozin", "Sergliflozin", "Remogliflozin", "Tofogliflozin", "Luseogliflozin", "Sotagliflozin", and "Mizagliflozin". Databases such as *PubMed, Embase, Cochrane Central Register of Controlled Trials* (*CENTRAL*), and the clinical trial registration website https://www.clinicaltrials.gov/ were searched to identify randomized controlled trials (RCTs) whose comparators were SGLT2i and other treatments (including placebo) until 24 October 2019. We searched only articles published in English (further details are available in the Supplemental Material online).

# Eligibility criteria

Studies with the following criteria were eligible for inclusion:

- RCTs;
- Fractures reported in the Results or in the section of adverse events;

• Intervention group with a single medication (SGLT2i) or a mixture (containing SGLT2i and other hypoglycemic drugs), with a placebo control group or other active treatment. Trials were included irrespective of the dosage of SGLT2i and active treatment.

Animal experiments, case reports, cohort studies, pooled analyses, and studies with a sample size less than 50 were excluded.

# Data extraction and quality assessment

Four authors (YL, YY, JD, and SB) screened the retrieved citations and selected the potential references. In the case of disagreements, other authors (WL and XN) were consulted until a consensus was reached. All potential studies were further analyzed with full text. Baseline characteristics, follow-up period, outcome, and adverse events data were extracted by four authors (KNCS, ZX, YG, and YZ). If the data were incomplete or unclear, the study details were searched in clinitrialtrials.gov website or other published articles.

## Quality assessment

We evaluated the risk of bias for every study according to the *Cochrane handbook for systematic reviews of interventions* (version 5.1.0). Two authors independently examined the references and classified studies into low risk, unclear risk, and high risk through random sequence generation (selection bias), allocation concealment (selection bias), blinding of participants and personnel (performance bias), blinding of outcome assessment (detection bias), incomplete outcome data (attrition bias), selective reporting (reporting bias), and other bias.

## Outcomes

The primary endpoint was the overall fracture incidence of SGLT2i, and the secondary outcomes were fracture incidence with canagliflozin, dapagliflozin, and empagliflozin.

## Statistical analysis

All analyses were performed using Stata 15.1 software (StataCorp, College Station, TX, USA) and ReviewManager (RevMan) version 5.3 (The Cochrane Collaboration, Copenhagen, Denmark).

The study was reported according to the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) statement.23 The odds ratios (ORs) and 95% confidence intervals (CIs) were calculated by Mantel-Haenszel analysis to compare the safety of SGLT2is. To compare the real effects of SGLT2i on fractures, all doses of SGLT2i and all control groups were included. We also evaluated the fracture incidence of canagliflozin, dapagliflozin, and empagliflozin. Considering the influence of other factors on fractures, we performed subgroup analysis based on control type (placebo or other active comparators), follow-up period, ethnicity, age, data source, presence or absence of chronic kidney diseases (CKDs), history of CVD, and T2DM duration. We used the  $I^2$  statistic and chi-square test to evaluate heterogeneity across trials;  $I^2 > 50\%$  was considered to indicate substantial heterogeneity.24 The Mantel-Haenszel fixed-effects model was used where  $I^2 < 50\%$ ; otherwise, the Mantel-Haenszel random-effects model was used. By excluding each trial subsequently, we performed sensitivity analyses to evaluate the stability and reliability of the results. A visual funnel plot was used to evaluate publication bias.

#### Results

#### Study selection

Using the above-mentioned keywords, we identified 9750 potential references. A total of 4791 duplicates were excluded manually and using software, and the remaining 4959 manuscripts were screened by browsing titles and abstracts. Subsequently, the full text of 452 potentially eligible references was searched; of which, 374 articles were excluded as duplicate articles, not reporting fracture, and due to improper comparators. Finally, the remaining 78 RCTs of 85,122 patients were included in our meta-analysis (flowchart in Figure 1).

#### Characteristics of eligible studies

In total, 78 RCTs of 85,122 patients were included in our final analysis; 65 RCTs were selected from published articles and the other 13 trials were retrieved from the https://www.clinicaltrials.gov/ website. A total of 50,471 (59.3%) patients were treated with SGLT2is, and 34,651 (40.7%) were treated with other drugs or placebo. The shortest and the longest follow-up periods were 12 weeks



**Figure 1.** Flowchart of study selection. RCT, randomized controlled trial.

and 296 weeks, respectively, and the sample size in each study ranged from 121 to 17,160 patients. Canagliflozin was used in the treatment group in 16 studies, dapagliflozin in 33 studies, empagliflozin in 17 studies, bexagliflozin in two studies, ertugliflozin in six studies, ipragliflozin in two studies, remogliflozin in one study, and tofogliflozin in one study. Control group agents included exenatide, glimepiride, metformin, linagliptin, pioglitazone, sitagliptin, and placebo. The baseline characteristics and other data of the included studies are listed in the Supplemental Material.

#### Primary endpoint

The primary endpoint was the overall fracture incidence of SGLT2i. A total of 78 trials in 85,122 patients reported treatment with SGLT2i or control (including placebo). A total of 1294 fractures occurred in 50,471 patients treated with SGLT2i, and 961 fractures occurred in 34,651 patients treated with control agents (including placebo); SGLT2i did not increase the risk of fracture [2.56% *versus* 2.77%; OR, 1.03; 95% CI

(0.95, 1.12); p = 0.49]. Subgroup analysis showed that in a follow-up of no more than 52 weeks, patients treated with SGLT2i showed a 29% decrease in the incidence of fracture [OR, 0.71; 95% CI, (0.55, 0.93); p=0.01], but this benefit was lost when the follow-up extended to more than 1 year [OR, 1.08; 95% CI (0.98, 1.18; p=0.12]. There was no significant difference in the other subgroup analyses based on control group agents (active comparator versus placebo), ethnicity (White versus Asian), history of CVD, age (<60 years old or >60 years old), data source (published article versus website registration information), patients with or without CKD, and duration of T2DM (fracture risk is shown in Figure 2, subgroup analysis is listed in Table 1).

#### Secondary outcomes

The secondary outcomes were fracture incidences of canagliflozin, dapagliflozin, and empagliflozin. In the present meta-analysis, canagliflozin tended to increase the risk of fracture [3.51% versus 2.77%; OR, 1.17; 95% CI (1.00, 1.37); p=0.05]. With regard to subgroup analysis, canagliflozin increased the risk of fracture in patients with a follow-up  $\geq 104$  weeks, mean age >60 years old, duration of T2DM >10 years, or whose ethnicity is White (4.92% versus 3.48%, p=0.02; 4.82%versus 3.46%, p=0.03; 4.69% versus 3.39%, p=0.02; 3.78% versus 2.88%, p=0.03, respectively). Dapagliflozin and empagliflozin were not associated with a higher incidence of fracture (Figures 3-5 and Table 1).

#### Publication bias and quality assessment

We observed no obvious publication bias from the funnel plot (funnel plot is shown in the Supplemental Material).

## Sensitivity analysis

There was no obvious heterogeneity in the overall SGLT2i analysis or the dapagliflozin and empagliflozin group analysis ( $I^2 = 0$ ). In analysis of the canagliflozin group,  $I^2$  was 32%, but this dropped to 0 if the CANVAS study was omitted.

## Discussion

To the best of our knowledge, the present study is the largest meta-analysis to directly compare SGLT2is with other hypoglycemic agents. In our analysis of 78 RCTs, we found that canagliflozin seems to increase the risk of fracture, while other SGLT2is are not associated with a higher incidence of fracture.

It has been reported that there is increased risk of fracture in elderly T2DM patients that can be lowered by glycemic control.25 SGLT2i can achieve a 0.7% reduction in glycated hemoglobin,<sup>12</sup> which may explain why SGLT2i leads to a lower fracture incidence over a follow-up of less than 52 weeks. However, when this period is extended to more than 52 weeks, the impact of SGLT2i on phosphate metabolism may not be omitted. SGLT2is mediate their effects by inhibiting the reabsorption of glucose in the proximal tubule of the kidney;11 this increases the concentration of serum phosphate, likely through an increase in tubular reabsorption, which may result in a higher parathyroid hormone (PTH) concentration and increased fracture incidence. In addition, the change in phosphate concentration may provoke the secretion of fibroblast growth factor 23 (FGF23). Together, these factors may have a combined action on BMD and increase the fracture incidence.<sup>26,27</sup> What is more, a meta-analysis including 43 RCTs demonstrated that SGLT-2i can lower systolic blood pressure (BP) by 2.46 mmHg and diastolic BP by 1.46 mmHg.28 Scheen pointed out that SGLT-2is have a higher incidence of orthostatic hypotension compared with other hypoglycemic drugs in elderly populations, who have a higher incidence of fracture.<sup>29</sup> The antihypertensive effect of SGLT2 may play an important role in the occurrence of fracture. In summary, this combination effect of BP lowering and BMD change may neutralize the advantage of glycemic control and cause SGLT2i to have an effect on fracture incidence, comparable to that observed in other treatments.

Our results are consistent with those of previous meta-analyses<sup>30,31</sup> in that we find that SGLT2i may have a beneficial effect on fracture in a follow-up of less than 52 weeks, but when this period is extended to more than 52 weeks, this benefit disappears. However, we also found that canagliflozin may increase the incidence of fracture, which is different from the findings of the previous study; this is likely to be because we included the CANVAS study using the newest follow-up data,<sup>17,18</sup> which showed an obvious increase in fracture. Because the CANVAS study introduces heterogeneity to the subgroup

	SGLT	<b>'2i</b>	Cont	rol		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% C	M-H. Fixed, 95% Cl
Allegretti 2019	7	157	6	155	0.6%	1.16 [0.38, 3.53]	
Araki 2015 Araki 2016	6	1097	0	63	0.1%	0.76 [0.04, 13.57]	
Railey 2013	7	409	2	137	0.1%	1 18 [0.06, 37.22]	
Barnett 2014	5	403	12	319	1.3%	0.31 [0.11, 0.89]	
Bode 2015	17	477	5	237	0.6%	1.71 [0.62, 4.71]	<u>+</u>
Cefalu 2015	0	455	2	459	0.2%	0.20 [0.01, 4.20]	
Dagogo-Jack 2017	2	309	1	153	0.1%	0.99 [0.09, 11.01]	
Ferdinand 2019	0	78	0	72		Not estimable	
Ferrannini 2013	5	547	1	112	0.2%	1.02 [0.12, 8.85]	
Fioretto 2018	0	160	0	161	0.404	Not estimable	
Forst 2014	3	227	10	115	0.1%	3.60 [0.18, 70.31]	
Gallo 2019 Grupherger 2018	5	313	10	209	0.1%	1 08 0 22 17 87	
Halvorsen 2019	1	191	0	193	0.0%	3.05 [0.12, 75,27]	· · · · · · · · · · · · · · · · · · ·
Handelsman 2018	0	232	2	229	0.2%	0.20 [0.01, 4.10]	
Henry 2012	1	430	0	208	0.1%	1.46 [0.06, 35.90]	
Hollander 2019	10	889	2	437	0.3%	2.47 [0.54, 11.34]	
Inagaki 2012	2	308	0	75	0.1%	1.23 [0.06, 25.92]	
Inagaki 2014	0	178	2	93	0.3%	0.10 [0.00, 2.16]	
Inagaki 2016	0	76	1	70	0.2%	0.30 [0.01, 7.56]	
Jabbour 2014	0	223	1	224	0.1%	0.33 [0.01, 8.23]	
Jappour 2018	1	408	1	22/	0.1%	1.51 [0.06, 37, 26]	
Ji 2014	1	400	0	220	0.1%	1.51 [0.06, 37.26]	
Kaku 2013	1	225	0	54	0.1%	0.73 [0.03, 18 12]	
Kaku 2014	0	174	1	56	0.2%	0.11 [0.00. 2.64]	
Kaku,2014	1	174	o o	87	0.1%	1.51 [0.06, 37.52]	
Kashiwagi 2015	1	97	0	54	0.1%	1.69 [0.07, 42.31]	
Kawamori 2018	5	182	2	93	0.3%	1.29 [0.24, 6.75]	<del></del>
Kohan 2014	13	168	0	84	0.1%	14.67 [0.86, 249.89]	
Kovacs 2013	3	333	4	165	0.5%	0.37 [0.08, 1.65]	
Lavalle-González 2013	2	735	0	549	0.1%	3.75 [0.18, 78.18]	
Leiter 2014	5	482	8	483	0.8%	0.62 [0.20, 1.92]	
Ljunggren 2012	0	89	0	91	0.40/	Not estimable	
Lu 2015 Manaia 2016	0	552	1	271	0.1%	0.31 [0.01, 7.82]	
Mathieu 2015	0	160	2	160	0.2%	0.20 [0.01, 4.02]	
McMurray 2019	49	2373	50	2371	4.8%	0.98 [0.66, 1.46]	+
Miller 2018	0	194	1	97	0.2%	0.17 [0.01, 4.10]	
Müller-Wieland 2018	2	626	0	313	0.1%	2.51 [0.12, 52.44]	
Nauck 2014	6	400	9	401	0.9%	0.66 [0.23, 1.88]	
NCT 00643851	0	397	1	201	0.2%	0.17 [0.01, 4.15]	
NCT 01032629	265	2888	85	1442	10.1%	1.61 [1.25, 2.08]	-
NCT 01095653	1	261	0	132	0.1%	1.53 [0.06, 37.71]	
NCT 01106625	0	313	1	156	0.2%	0.17 [0.01, 4.08]	
NCT 01137812	2	377	1	378	0.1%	2.01 [0.18, 22.27]	
NGT 01159600	3	8/1	0	432	0.1%	3.49 [0.18, 67.64]	
NCT 01195062	2	308	1	188	0.1%	3.15 [0.15, 65.89]	
NCT 01422876	3	1074	0	267	0.2%	1 75 [0.09, 33, 93]	
NCT 01719003	1	995	Ő	332	0.1%	1.00 [0.04, 24.68]	
NCT 01989754	68	2907	78	2905	7.4%	0.87 [0.62, 1.21]	+
NCT 02564926	2	60	1	61	0.1%	2.07 [0.18, 23.44]	<del></del>
NCT 02589639	1	176	3	90	0.4%	0.17 [0.02, 1.62]	
NCT01137474	2	633	1	311	0.1%	0.98 [0.09, 10.88]	
Patel 2016	25	968	13	482	1.7%	0.96 [0.48, 1.89]	Ť
Perkovic 2019	67	2202	68	2199	6.4%	0.98 [0.70, 1.39]	<del>_</del>
Pollock 2019	1	293	2	155	0.3%	0.26 [0.02, 2.91]	
Riuderstrale 2018	31	107	33	/80	3.1%	0.96 [0.58, 1.58]	
Roden 2015	1	107	1	451	0.1%	3 03 [0.01, 8.12]	
Rosenstock 2012	2	281	0	139	0.0%	2.50 [0.12, 74.50]	<u> </u>
Rosenstock 2012	1	358	2	176	0.3%	0.24 [0.02. 2.71]	
Rosenstock 2015	2	324	1	170	0.1%	1.05 [0.09. 11.66]	
Rosenstock 2019	2	579	1	291	0.1%	1.01 [0.09, 11.13]	
Scott 2018	1	306	2	307	0.2%	0.50 [0.05, 5.54]	
Stenlof 2013	0	483	1	192	0.2%	0.13 [0.01, 3.26]	
Strojek 2011	0	447	1	145	0.2%	0.11 [0.00, 2.66]	
Sykes 2015	1	179	0	71	0.1%	1.20 [0.05, 29.85]	
Søfteland 2017	0	222	0	110		Not estimable	
Vilsboll 2019	2	324	0	319	0.0%	4.95 [0.24, 103.59]	
Winding 2013	3	607	1	193	0.1%	0.95 [0.10, 9.22]	
VVIVIOTE 2018	457	8582	440	8578	40.7%	1.04 [0.91, 1.19]	<b>T</b>
Tale 2014 Vana 2016	0	1/9	2	90 14F	0.3%	0.10 [0.00, 2.08]	
Yang 2018	2	130	1	133	0.1%	0.32 [0.03, 10.78]	
Zinman 2015	179	4687	91	2333	11.4%	0.98 [0.76, 1 27]	+
			01			0.00 [0.00, 1.27]	
Total (95% CI)		50471		34651	100.0%	1.03 [0.95, 1.12]	
Total events	1294		961				
Heterogeneity: Chi <sup>2</sup> = 68.	81, df = 73	B (P = 0.	.62); l <sup>2</sup> = (	0%			
Test for overall effect: Z =	= 0.70 (P =	0.49)					Favours [SGI T2i] Favours [Control]

**Figure 2.** Forest plot of fracture incidence between SGLT2is and other treatment. Cl, confidence interval; M-H, Mantel-Haenszel; SGLT2i, sodium-glucose co-transporter 2 inhibitor.

analysis of canagliflozin, the heterogeneity drops from 32% to 0 if the CANVAS study is excluded. Correspondingly, the OR of the canagliflozin group versus other hypoglycemic agents drops from 1.17 (1.00, 1.37) to 0.92 (0.75, 1.14) if the CANVAS study is excluded. Furthermore, it is 
 Table 1. Subgroup analysis of fracture incidence between SGLT2i and other treatment.

	SGLT2i	Canagliflozin	Dapagliflozin	Empagliflozin
Overall OR	1.03 (0.95, 1.12)	1.17 (1.00, 1.37)	1.02 (0.90, 1.15)	0.89 (0.73, 1.10)
Control type				
SGLT2i <i>versus</i> placebo	1.04 (0.95, 1.14)	1.17 (0.99, 1.38)	1.03 (0.91, 1.16)	0.87 (0.69, 1.10)
SGLT2i versus active comparator	0.94 (0.70, 1.26)	1.18 (0.64, 2.19)	0.81 (0.43, 1.53)	1.00 (0.63, 1.59)
Follow-up				
≪52 weeks	0.71 (0.55, 0.93)	0.78 (0.48, 1.28)	0.71 (0.45, 1.12)	0.53 (0.30, 0.92)
52-104 weeks	1.03 (0.77, 1.39)	-	1.04 (0.92, 1.18)	1.04 (0.21, 5.17)
≥104 weeks	1.08 (0.98, 1.19)	1.22 (1.04, 1.45)	1.07 (0.76, 1.51)	0.97 (0.77, 1.22)
Ethnicity				
White	1.03 (0.94, 1.12)	1.19 (1.01, 1.39)	1.00 (0.88, 1.13)	0.91 (0.73, 1.12)
Asian	1.03 (0.59, 1.79)	0.44 (0.12, 1.66)	2.26 (0.91, 5.65)	0.61 (0.18, 2.02)
Mean age				
<60 years	0.81 (0.64, 1.04)	0.86 (0.51, 1.43)	0.78 (0.49, 1.23)	0.89 (0.60, 1.33)
>60 years	1.07 (0.97, 1.17)	1.21 (1.02, 1.43)	1.03 (0.91, 1.17)	0.90 (0.70, 1.15)
Source of data				
Published studies	1.03 (0.95, 1.13)	1.18 (1.00, 1.38)	1.01 (0.90, 1.15)	0.91 (0.73, 1.12)
<b>Clinical registration</b>	0.86 (0.43, 1.73)	0.78 (0.14, 4.26)	1.18 (0.38, 3.70)	0.67 (0.24, 1.89)
Duration of DM				
<10 years	0.80 (0.57, 1.13)	0.87 (0.50, 1.52)	0.83 (0.44, 1.56)	1.29 (0.24, 6.75)
>10 years	1.11 (1.00, 1.23)	1.21 (1.03, 1.43)	1.05 (0.92, 1.20)	-
History of CKD				
Patients with CKD	0.84 (0.49, 1.43)	0.10 (0.00, 2.08)	2.05 (0.69, 6.04)	0.31 (0.11, 0.89)
Patients without CKD	1.04 (0.95, 1.13)	1.18 (1.01, 1.39)	1.01 (0.89, 1.14)	0.94 (0.76, 1.17)
History of CVD				
Patients with CVD	1.11 (0.97, 1.27)	1.21 (1.02, 1.44)	0.93 (0.64, 1.35)	0.98 (0.76, 1.27)
Patients without CVD	1.04 (0.91, 1.19)	-	1.04 (0.91, 1.19)	-

CKD, chronic kidney disease; CVD, cardiovascular disease; DM, diabetes mellitus; OR, odds ratio; SGLT2i, sodium-glucose co-transporter 2 inhibitor.

important to note that the CANVAS study had the longest mean follow-up period, of 296 weeks, with a large sample size of 4330, and the baseline risk for fracture is higher than that reported in other studies.<sup>32–34</sup> In addition, the primary outcome in CANVAS was low-trauma fractures as judged by the trial adjudication committee.<sup>17</sup> These factors may explain why patients treated

	Canaglif	lozin	Contr	ol	Odds Ratio		Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% Cl
Bode 2015	17	477	5	237	2.3%	1.71 [0.62, 4.71]	
Forst 2014	3	227	0	115	0.2%	3.60 [0.18, 70.31]	
Inagaki 2012	2	308	0	75	0.3%	1.23 [0.06, 25.92]	
Inagaki 2014	0	178	2	93	1.1%	0.10 [0.00, 2.16]	·
Inagaki 2016	0	76	1	70	0.5%	0.30 [0.01, 7.56]	· · · · · · · · · · · · · · · · · · ·
Ji 2014	1	450	0	226	0.2%	1.51 [0.06, 37.26]	
Lavalle-González 2013	2	735	0	549	0.2%	3.75 [0.18, 78.18]	
NCT 01032629	265	2888	85	1442	36.0%	1.61 [1.25, 2.08]	
NCT 01106625	0	313	1	156	0.7%	0.17 [0.01, 4.08]	• • • • • • • • • • • • • • • • • • • •
NCT 01137812	2	377	1	378	0.3%	2.01 [0.18, 22.27]	
NCT 01989754	68	2907	78	2905	26.6%	0.87 [0.62, 1.21]	
Patel 2016	25	968	13	482	5.9%	0.96 [0.48, 1.89]	
Perkovic 2019	67	2202	68	2199	23.1%	0.98 [0.70, 1.39]	+
Rodbard 2016	0	107	1	106	0.5%	0.33 [0.01, 8.12]	
Stenlof 2013	0	483	1	192	0.7%	0.13 [0.01, 3.26]	• • • • • • • • • • • • • • • • • • • •
Yale 2014	0	179	2	90	1.2%	0.10 [0.00, 2.08]	·
Total (95% CI)		12875		9315	100.0%	1.17 [1.00, 1.37]	◆
Total events	452		258				
Heterogeneity: Chi <sup>2</sup> = 22.03, df = 15 (P = 0.11); l <sup>2</sup> = 32%							
Test for overall effect: Z = 1.96 (P = 0.05) 0.01 0.1 1 Favours [Canagliflozin] Favours							Favours [Canagliflozin] Favours [Control]

**Figure 3.** Forest plot of fracture incidence between canagliflozin and other treatment. CI, confidence interval; M-H, Mantel-Haenszel.

	Dapaglif	lozin	Cont	rol		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% C	M-H, Fixed, 95% Cl
Araki 2016	1	122	0	60	0.1%	1.49 [0.06, 37.22]	· · · · · · · · · · · · · · · · · · ·
Bailey 2013	7	409	2	137	0.6%	1.18 [0.24, 5.73]	· · · · · · · · · · · · · · · · · · ·
Cefalu 2015	0	455	2	459	0.5%	0.20 [0.01, 4.20]	· · · · · · · · · · · · · · · · · · ·
Fioretto 2018	0	160	0	161		Not estimable	
Handelsman 2018	0	232	2	229	0.5%	0.20 [0.01, 4.10]	· · · · · ·
Henry 2012	1	430	0	208	0.1%	1.46 [0.06, 35.90]	
Jabbour 2014	0	223	1	224	0.3%	0.33 [0.01, 8.23]	
Jabbour 2018	1	458	1	227	0.3%	0.49 [0.03, 7.94]	· · · · · · · · · · · · · · · · · · ·
Kaku 2013	1	225	0	54	0.2%	0.73 [0.03, 18.12]	
Kaku,2014	1	174	0	87	0.1%	1.51 [0.06, 37.52]	
Kohan 2014	13	168	0	84	0.1%	14.67 [0.86, 249.89]	
Leiter 2014	5	482	8	483	1.5%	0.62 [0.20, 1.92]	
Ljunggren 2012	0	89	0	91		Not estimable	
Mathieu 2015	0	160	2	160	0.5%	0.20 [0.01, 4.15]	• • • •
McMurray 2019	49	2373	50	2371	9.4%	0.98 [0.66, 1.46]	+
Müller-Wieland 2018	2	626	0	313	0.1%	2.51 [0.12, 52.44]	
Nauck 2014	6	400	9	401	1.7%	0.66 [0.23, 1.88]	
NCT 00643851	0	397	1	201	0.4%	0.17 [0.01, 4.15]	• • •
NCT 01095653	1	261	0	132	0.1%	1.53 [0.06, 37.71]	
NCT 01195662	2	358	0	224	0.1%	3.15 [0.15, 65.89]	
NCT 02564926	2	60	1	61	0.2%	2.07 [0.18, 23.44]	
NCT01137474	2	633	1	311	0.3%	0.98 [0.09, 10.88]	
Pollock 2019	1	293	2	155	0.5%	0.26 [0.02, 2.91]	
Rosenstock 2012	2	281	0	139	0.1%	2.50 [0.12, 52.33]	
Rosenstock 2014	1	358	2	176	0.5%	0.24 [0.02, 2.71]	
Rosenstock 2019	2	579	1	291	0.3%	1.01 [0.09, 11.13]	
Scott 2018	1	306	2	307	0.4%	0.50 [0.05, 5.54]	
Strojek 2011	0	447	1	145	0.4%	0.11 [0.00, 2.66]	
Vilsboll 2019	2	324	0	319	0.1%	4.95 [0.24, 103.59]	
Wilding 2013	3	607	1	193	0.3%	0.95 [0.10, 9.22]	
Wiviott 2018	457	8582	440	8578	79.9%	1.04 [0.91, 1.19]	
Yang 2016	2	299	1	145	0.3%	0.97 [0.09, 10.78]	
Yang 2018	0	139	1	133	0.3%	0.32 [0.01, 7.84]	
Total (95% CI)		21110		17259	100.0%	1.02 [0.90, 1.15]	•
Total events	565		531				
Heterogeneity: Chi <sup>2</sup> = 18.37, df = 30 (P = 0.95); l <sup>2</sup> = 0%							
Test for overall effect: Z	z = 0.27 (P	= 0.79)		Favours [Dapagliflozin] Favours [Control]			

**Figure 4.** Forest plot of fracture incidence between dapagliflozin and other treatment. CI, confidence interval; M-H, Mantel-Haenszel.

with canagliflozin in the CANVAS study had a higher incidence of fracture.<sup>35</sup> A real-world study that enrolled 159,928 patients and compared

canagliflozin with glucagon-like peptide-1 (GLP-1) receptor agonist suggested that canagliflozin is not associated with a higher risk of fracture.<sup>32</sup>



**Figure 5.** Forest plot of fracture incidence between empagliflozin and other treatment. CI, confidence interval; M-H, Mantel-Haenszel.

Because we observed that the real-world study had a mean follow-up of 34 weeks, it is difficult to conclude the long-term safety. Similar to our results, a previous study that pooled the results of 10 trials showed that fracture risk in the canagliflozin group was increased.<sup>35</sup> Therefore, the long-term fracture incidence of canagliflozin deserves further attention.

The mechanism through which canagliflozin increases the risk of fracture is still unclear. SGLT-2i agents have different selectivity for SGLT-2 versus SGLT-1. Empagliflozin has the highest SGLT-2/SGLT-1 affinity ratio and canagliflozin the lowest.<sup>11</sup> Masiukiewicz and Ljunggren et al. suggest that ertugliflozin and dapagliflozin had no significant effect on BMD and other bone biomarkers,36,37 while Bilezikian et al. found that canagliflozin can lower the hip BMD and increase bone biomarkers.<sup>20</sup> Animal experiments also showed that canagliflozin can increase the concentration of phosphate, FGF23, and PTH, whereas tofogliflozin has been shown to have no clear effect on bone mass by microcomputed tomography.<sup>38</sup> Whether the difference in selectivity for SGLT-2/ SGLT-1 and different bone biomarkers lead canagliflozin to have a higher fracture incidence remains uninvestigated.

With regard to subgroup analysis, we found no obvious difference in SGLT2i and other treatment, with the exception of the canagliflozin group. Tang *et al.* found that SGLT2is had a

tendency to increase the risk of fracture in the Asian population,<sup>39</sup> but we did not observe that phenomenon in our subgroup analysis. Tang et al. included 2819 patients and found no significant difference [OR, 2.05; 95% CI (0.86, 4.87)], while we included 18 studies of 5279 patients (mainly Asians) and observed no significant difference [OR 1.03, 95% CI (0.59, 1.79)]. Furthermore, a study mainly including CKD patients showed that empagliflozin had a reduced incidence of fracture in stage 3 CKD<sup>21</sup>(1.6% versus 4.8%), but another study with a follow-up period of 104 weeks that mainly included moderate renal impairment CKD patients showed an obvious increase in fracture incidence for dapagliflozin (7.74% versus 0%).19 Our meta-analysis included eight trials of CKD patients and found that SGLT2is had no obvious effects on fracture incidence [OR, 0.84; 95% CI (0.49, 1.43), p=0.53]. However, because the heterogeneity across trials was high, we could not eliminate the influence on analysis; thus, further studies are needed to determine the fracture risk of SGLT2i.

#### Limitations

The present meta-analysis has some limitations in addition to the disadvantages in the original research. First, some data are acquired from the clinical registration website and not from the published article; this may introduce bias. However, there was no difference in our conclusion when we performed subgroup analysis based on data source and subgroup analysis. Second, some events data are not presented in the article, and, if after contact with the authors we were still unable to gain access to the raw data, the events data were transformed from the published data. Third, we were unable to access some baseline characteristics, which limited our ability to perform subgroup analysis.

In summary, canagliflozin seems to increase the risk of fracture, while other SGLT2is do not result in a higher incidence of fracture.

#### **Conflict of interest statement**

The authors declare that there is no conflict of interest.

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#### Supplemental material

Supplemental material for this article is available online.

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