LETTER

Multiple myeloma gammopathies



Clinical features and survival outcomes in IgD myeloma: a study by Asia Myeloma Network (AMN)

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To the Editor:

Immunoglobulin D (IgD) myeloma is a rare isotype that comprises 1–2% of multiple myeloma (MM) patients [1–3], which has significantly inferior survival for a median overall survival (OS) between 13 and 21 months [4–6].

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Given the lack of large cohort with comprehensive clinical and cytogenetic assessment, knowledge about IgD myeloma is obtained mostly from a limited sample size [7]. Therefore, we carried out a multicenter retrospective study to evaluate the prevalence, clinical features, prognosis, and to develop and validate a prognostic model, including 356 patients with IgD myeloma from 14 centers of Asian Myeloma Network (AMN).

Data were collected from China, Korea, and Singapore diagnosed from 2002 to 2019 (Supplementary Table 1). Ethical committee approvals were obtained and study protocol was approved by the Institutional Review Board of each institution. To avoid clinical information leak, and get

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Variable	IgD MM	Non-IgD MM	P value	** ' 11			
variable	(N = 356, %)	Non-IgD MM $(N = 712, \%)$	P value	Variable	IgD MM (N = 356, %)	Non-IgD MM $(N = 712, \%)$	P value
Sex				≤0.01, ≥100	138 (38.8)	356 (50)	
Male	241 (67.7)	429 (60.3)	0.018	Data missing	81 (22.7)	0 (0)	
Female	115 (32.3)	283 (39.7)		Del (13q) in FISH			
Age at diagnosis, year	ars			Yes	77 (21.7)	261 (36.7)	0.001
Median (range)	56 (32–85)	62 (23–96)	< 0.001	No	224 (62.9)	451 (63.3)	
<65	286 (80.3)	468 (65.7)		Data missing	55 (15.4)	0 (0)	
≥65	70 (19.7)	244 (34.3)		Del (17p) in FISH			
DS Stage				Yes	35 (9.9)	75 (10.5)	0.609
I	8 (2.2)	16 (2.2)	0.812	No	266 (74.7)	637 (89.5)	
II	23 (6.5)	54 (7.6)		Data missing	55 (15.4)	0 (0)	
III	323 (91.3)	642 (90.2)		1q21 gains in FISH			
ISS stage				Yes	91 (25.6)	368 (51.7)	< 0.001
I	70 (19.7)	179 (25.1)	< 0.001	No	179 (50.2)	344 (48.3)	
II	64 (17.9)	262 (36.8)		Data missing	86 (24.2)	0 (0)	
III	222 (62.4)	271 (38.1)		t (11;14) in FISH			
Plasma cells of BM	` '	,		Yes	88 (24.7)	96 (13.5)	< 0.001
≥50	149 (41.9)	143 (20.1)	< 0.001	No	213 (59.9)	616 (86.5)	
<50	207 (58.1)	569 (79.9)		Data missing	55 (15.4)	0 (0)	
Hemoglobin level (g		(,,,,,		t (4;14) in FISH	,	. ,	
<100	231 (64.9)	420 (59)	0.063	Yes	4 (1.1)	136 (19.1)	< 0.001
≥100	125 (35.1)	292 (41)		No	297 (83.5)	576 (80.9)	
Platelet count (10 ⁹ /L	` '	- > - (11)		Data missing	55 (15.4)	0 (0)	
<100	76 (22.3)	91 (12.8)	< 0.001	t (14;16) in FISH	00 (101.)	0 (0)	
≥100	280 (78.7)	621 (87.2)	10.001	Yes	3 (0.9)	8 (1.1)	0.859
Serum LDH (U/L)	200 (70.7)	021 (07.2)		No	298 (83.7)	704 (98.9)	0.009
≥245	136 (38.2)	177 (24.9)	< 0.001	Data missing	55 (15.4)	0 (0)	
<245	220 (61.8)	535 (75.1)	VO.001	Double hit ^a	33 (13.1)	0 (0)	
Serum creatinine lev		333 (73.1)		yes	20 (5.6)	116 (16.3)	< 0.001
≥2	137 (38.5)	131 (18.4)	< 0.001	no	251(70.5)	596 (83.7)	10.001
<2	219 (61.5)	581 (81.6)	VO.001	Data missing	85 (23.9)	0(0)	
Serum calcium level		301 (01.0)		Triple hit ^b	03 (23.7)	0(0)	
≥2.65	85 (23.9)	105 (14.7)	< 0.001	Yes	1 (0.3)	11 (1.5)	0.135
<2.65	271 (76.1)	607 (85.3)	<0.001	No	269 (75.6)	701 (98.5)	0.133
Light chain restriction		007 (03.3)		Data missing	86 (24.1)	0 (0)	
Kappa	40 (11.2)	407 (57.2)	< 0.001	t (11;14) and Del (1	. ,	0 (0)	
Lambda	316 (88.8)	305 (42.8)	<0.001	Yes	28 (7.9)	28 (3.9)	0.001
Extramedullary plass		303 (42.8)		No	273 (76.7)	684 (96.1)	0.001
Yes	68 (19.1)	106 (14.9)	0.079	Data missing	55 (15.4)	0 (0)	
No	288 (80.9)	606 (85.1)	0.079	t (11;14) and Del (1		0 (0)	
R-ISS stage	200 (00.9)	000 (05.1)		Yes	10 (2.8)	5 (0.7)	0.002
I	42 (11.8)	114 (16)	< 0.001	No	291 (81.8)	707 (99.3)	0.002
I II	42 (11.8) 175 (49.2)	449 (63.1)	\U.UU1	Data missing	55 (15.4)	0 (0)	
III				t (11;14) and 1q21 g		0 (0)	
	115 (32.3) 24 (6.7)	149 (20.9) 0 (0)		Yes	27 (7.6)	39 (5.5)	0.012
Data missing	24 (0.7)	0 (0)		No	243 (68.2)	673 (94.5)	0.012
FLCR	127 (29.5)	256 (50)	0.050		` ′	` ′	
0.01–100	137 (38.5)	356 (50)	0.959	Data missing	86 (24.2)	0 (0)	

Table 1 (continued)

Variable	IgD MM	Non-IgD MM	P value	
	(N = 356, %)	(N = 712, %)		
t (11;14) and t (4;14	1) in FISH			
Yes	1 (0.3)	0 (0)	0.124	
No	300 (84.3)	712 (100)		
Data missing	55 (15.4)	0 (0)		
t (11;14) and t (14;1	16) in FISH			
Yes	0 (0)	0 (0)	NA	
No	301 (84.6)	712 (100)		
Data missing	55 (15.4)	0 (0)		
t (11;14) and double	e hit in FISH			
Yes	6 (1.7)	2 (0.3)	0.003	
No	265 (74.4)	710 (99.7)		
Data missing	85 (23.9)	0 (0)		
t (11;14) and triple	hit in FISH			
Yes	0 (0)	0 (0)	NA	
No	270 (75.8)	712 (100)		
Data missing	86 (24.2)	0 (0)		

MM multiple myeloma, Ig immunoglobulin, DS Durie Salmon, ISS international staging system, R-ISS revised ISS, LDH Lactate dehydrogenase, BM bone marrow, FISH fluorescence in situ hybridization, Del deletion, FLCR free light chains ratio, NA not avaliable.

a real sense of the accurate model's outcomes, we split existing 356 IgD MM to three parts, namely training cohort (one center from Shanghai, n=212), validation cohort 1 (two centers from Beijing, n=81), and validation cohort 2 (centers from Korea and Singapore, n=63). The Least Absolute Shrinkage and Selector Operation (LASSO) Cox regression model to determine prognostic factors from the variables with P < 0.05 in the log-rank tests was performed as described [8, 9]. The quality of the prediction model was measured using the concordance index (C-index) and areas under the time-dependent receiver-operating characteristics (ROC) curves (AUCs). A bootstrap with 1000 re-samples was used for internal validation. SAS 9.4 and R 3.5.1 were used for the statistical analysis.

A total of 356 patients with IgD myeloma represented 2–8.8% of all myeloma patients, especially over 5% IgD myeloma prevalence in Chinese centers. We compared the clinical characteristics of IgD myeloma with 712 (1:2) non-IgD myeloma patients random selected as control matched for year of diagnosis and systemic therapy from Shanghai Changzheng Hospital. Baseline characteristics of total cohort are listed on Table 1 and different centers are shown in Supplementary Table 2. IgD myeloma patients had a higher frequency in male, younger than 65 years, advanced

R-ISS stage III, hypercalcemia, elevated creatinine levels, and elevated LDH. Cytogenetic information was available for 301 patients (84.6%), while the 1q21 probe was only performed in 75.8% patients. Notably, 29.2% frequency of t(11;14) was predominantly higher compared to those in non-IgD subtypes (P < 0.001). Among the 88 IgD patients harbored t(11;14), the most frequent chromosome abnormalities (CA) coupled with t(11;14) were 13q-(31.8%), 1q21 + (30.7%), and followed 17p- (11.4%). 'Double-hit' or 'triple-hit' [10, 11] only occurred 5.6% and 0.3% patients, respectively.

And then, we compared IgD myeloma patients with IgG, IgA, and light chain patients random selected as matched control (Supplementary Table 3). The median age was younger in IgD compared with others myeloma subtypes. Notably, the frequency of t(11;14) was significantly higher than non-IgD subtypes (IgD 29.2% vs IgG 10.6% vs IgA 8.4%, P < 0.001), but was a slight higher than light chain subtype (29.2 vs 24.9%). 'Double-hit' phenotype was significant lower in IgD myelomas than others subtypes.

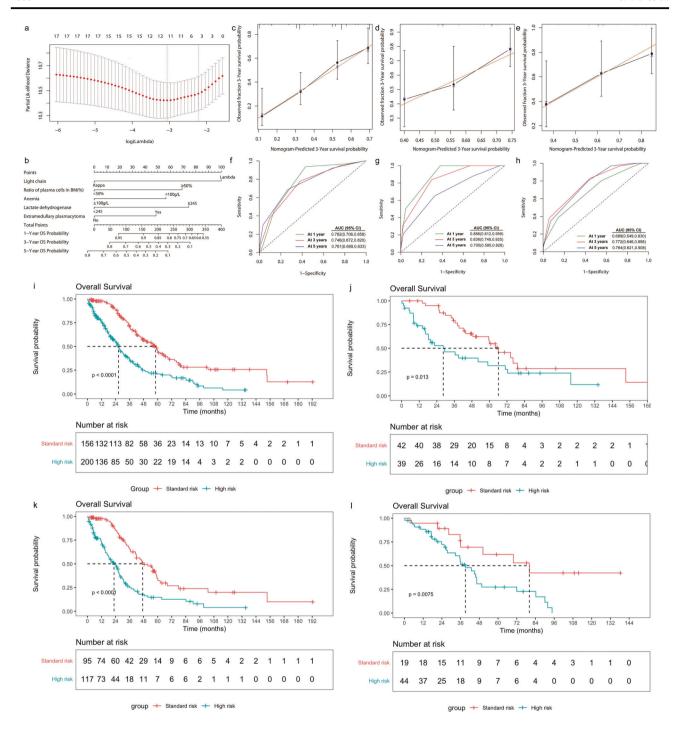
Frontline treatment modalities used are shown in Supplementary Table 4. The overall response rate (ORR) was 88.8%, and very good partial response or better was 58.6% (Supplementary Table 5). After a median follow-up of 8.2, 7.3, and 4.9 years for the three cohorts, the median OS were 36.5 months for the total cohort and 31.2 months in training cohort, 52.2 months in validation cohort 1, and 45.7 months in validation cohort 2 (Supplementary Fig. 1 and Supplementary Table 6). Patients received IMiDs showed a relatively longer median OS than others regimens, however, which untranslated into a significant survival benefit (P =0.17, Supplementary Fig. 2a), and might be the subgroups limitation. Patients received ASCT had a median OS of 45.7 months, which was a slightly longer than 35 months for non-ASCT patients (P = 0.4, Supplementary Fig. 2b). We subsequently investigated whether cytogenetic aberration was a prognostic factor [12], which showed that CA did not have an impact on OS, suggesting other molecular events overcome initial CA risk features and impact prognosis.

Subsequently, the LASSO Cox regression model to determine prognostic factors from the univariate analysis was performed (Supplementary Table 7). Five clinical parameters with statistically relevant, including lambda light chain, plasma cells in BM \geq 50%, hemoglobin < 100 g/L, LDH \geq 245 U/L, and extramedullary plasmacytoma, were integrated into multivariate LASSO regression model (Supplementary Table 8). A nomogram was developed and the risk score was computed as follows: 0.9215 × lambda light chain + 0.6376 × plasma cells in BM (\geq 50%) + 0.5203 × anemia (<100 g/L) + 0.6864 × LDH (\geq 245 U/L) + 0.4484 × extramedullary plasmacytoma (variable present = 1, absent = 0, Fig. 1a, b). The predictive accuracy for OS calculated using the C-index was 0.705 (95% CI, 0.663–0.747). In the internal

^aThe cooccurrence of any 2 of the following: t(4;14), t(14;16), gain (1q), del(17p).

^bThe cooccurrence of 3 or more of the following: t(4;14), t(14;16), gain(1q), del(17p).

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validation, the corrected C-index of OS was 0.696. Similarly, the C-index for OS in validation cohort 1 was 0.690 (95% CI, 0.612–0.768) and 0.703 (95% CI, 0.608–0.798) in validation cohort 2. The calibration curves of the alternative nomogram to predict the 3-year OS presented in Fig. 1c–e suggested a good fit for the observed nomogram, when compared with the ideal nomogram. The panel displayed an AUC value at 1-year, 3-year, and 5-year OS, and the validation sets had a similar high AUC values at these timepoints (Fig. 1f–h).

On the basis of the distribution of the risk scores and the 3-year survival probability, two categories of risk were created with the cut-off point at 1.56: standard risk (risk score \leq 1.56, n=156) and high-risk subgroup (risk score > 1.56, n=200). The clinical characteristics between derivations were presented in Supplementary Table 6. Patients with IgD myeloma at standard risk were significantly better than high-risk subgroup (Fig. 1i). Similar results were obtained in training and validation cohorts respectively (Fig. 1j–l). Notably we

◆ Fig. 1 Development of a predictive model and validation. LASSO Cox regression model to determine prognostic factors from the variables with P < 0.05 in the log-rank tests was performed. The prediction model was established on the basis of variables selected from the LASSO Cox regression model and weighted using the Cox regression coefficient. Cross-validation for tuning parameter selection in the LASSO model (a). The nomogram based on data from training patients to predict individual prognosis (b). The calibration curves of an alternative nomogram to predict 3-year OS of IgD myeloma in training patients (c), validation cohort 1 (d) and validation cohort 2 (e). The X-axis represents the predicted survival probability calculated using the nomogram, while the Y-axis represents the actual survival probability for patients. The gold 45-degree line represents the ideal nomogram, while the black line represents the observed nomogram. The AUC values were 0.746-0.782 (f), 0.705-0.886 (g), and 0.689-0.772 (h) in the training cohort, validation cohort 1 and validation cohort 2, respectively. On the basis of the distribution of the risk scores and the 3-year survival probability, two categories of risk were created with the cut-off point at 1.56: standard risk (risk score ≤ 1.56. n = 156) and high-risk subgroup (risk score > 1.56, n = 200). The 3year OS in patients with IgD myeloma at standard risk were significantly better than those of patients at high-risk subgroup (70.6 ± 4.1% vs $36.1 \pm 3.8\%$, P < 0.0001, 1i). Similar results were obtained in the training cohort $(65.8 \pm 5.7\% \text{ vs } 25.1 \pm 4.6\%, P < 0.0001, 1k)$, validation cohort 1 (79.4 \pm 6.5% vs 46.2 \pm 8.6%, P = 0.013, 1j), and validation cohort 2 (69.5 \pm 11.5% vs 54.4 \pm 8.3%, P = 0.0075, 11).

identified that ASCT showed a survival advantage in highrisk group, while it did not improve the survival in standard risk group (Supplementary Fig. 3). Moreover the prediction value of the model was independent of induction modalities (Supplementary Fig. 4).

The clinical features of IgD myeloma in this study indicated a potentially high-risk population. We identified a higher prevalence difference may be related to the difference in ethnicity, an efficient algorithm diagnosis, and assembled patients. We found 40.6% of patients had abnormal karyotypes. Translocation (11;14) was predominantly high frequency, which showed the similar to others studies [13, 14]. By far, the most clinically relevant feature of myeloma with t (11;14) is increased expression of the anti-apoptotic protein BCL-2 [15], and venetoclax of Bcl-2 inhibitor provides a new option for t(11;14), which may have a special significances for IgD patients with t(11;14).

The median OS with 36.5 months in IgD myeloma showed an inferior survival than patients with more common myeloma subtypes [16, 17]. Although the ORR (88.8%) has not improved rapidly in this AMN study, patients received IMiD with lenalidomide showed a better trend median OS than patients received PIs, conventional chemotherapy regimens, even PI + thalidomide as PI + IMiD subgroup, while PI + IMiD overcame improvement due to high peripheral neuropathy to limit the application. However, the patients received IMiD is too small to draw the definite conclusion, which might provide a clue to optimize treatment. Meanwhile, we demonstrated ASCT would benefit for IgD myeloma patients with high-risk score.

Previous studies concentrated on the prognostic value of single variable [1, 3, 5, 18, 19], as it is a challenge to construct a comprehensive prognostic model due to the rarity of IgD myeloma. We used 356 IgD myeloma patients to develop a prognostic model based on the 17 variables of the multivariate LASSO model to estimate the 3-year OS. Five clinical parameters were identified as clinically relevant to compute the risk score, which a high-risk score was consistently associated with a poor survival outcome. Moreover, the prediction value of the model was independent of induction modalities. This risk model improves the classification of IgD myeloma and may facilitate the development of risk-adapted treatment strategies. Although this retrospective design, presence of missing data, and heterogeneous treatment could have resulted in the introduction of selection bias and underestimate or overestimate of some of results, we believe that the large sample size and the multicenter study performed in such a rare disease could balance out some of these weaknesses.

In conclusion, we described the clinical features of 356 IgD myeloma patients, as well as developed and validated a predictive model containing five baseline clinical variables that could group the IgD patients into standard risk and high risk. Meanwhile, we demonstrated that IMiDs therapy might be a trend to benefit for the patient's outcome, and ASCT could benefit patient with high risk within the predictive model. These findings may provide guidance for management of IgD myeloma and better prognostic stratification for development of risk-adapted treatment strategies.

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Author contributions JL, XH and YJ collected and analyzed the data, and wrote the first draft, and approved the final version of the paper; JL, CS, MKK, FZ, and WC performed patient management and approved the final version of the paper; YJ, AL, LY and ZY, performed patients' follow-up, participated in final data analysis and approval of the final version of the paper; QC analyzed the data and approval of the final version of the paper; JHL, KK, and WJC, performed patient management revised the paper critically, and approved the final version of the paper; SKK and BD revised the paper critically, and approved the final version of the paper, WF, and JH designed the study, performed patient management, and approved the final version of the paper, and; JD designed the study, performed patient management, analyzed the data, wrote the first draft, approved the final version of the paper.

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Compliance with ethical standards

Conflict of interest The authors declare that they have no conflict of interest.

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References

- Jancelewicz Z, Takatsuki K, Sugai S, Pruzanski W. IgD multiple myeloma. Review of 133 cases. Arch Intern Med. 1975;135:87–93.
- Morris C, Drake M, Apperley J, Iacobelli S, van Biezen A, Bjorkstrand B, et al. Efficacy and outcome of autologous transplantation in rare myelomas. Haematologica. 2010;95:2126–33.
- Blade J, Lust JA, Kyle RA. Immunoglobulin D multiple myeloma: presenting features, response to therapy, and survival in a series of 53 cases. J Clin Oncol. 1994;12:2398–404.
- Zagouri F, Kastritis E, Symeonidis AS, Giannakoulas N, Katodritou E, Delimpasi S, et al. Immunoglobulin D myeloma: clinical features and outcome in the era of novel agents. Eur J Haematol. 2014;92:308–12.
- Wechalekar A, Amato D, Chen C, Keith Stewart A, Reece D. IgD multiple myeloma-a clinical profile and outcome with chemotherapy and autologous stem cell transplantation. Ann Hematol. 2005; 84:115–7.
- Kuliszkiewicz-Janus M, Zimny A, Sokolska V, Sasiadek M, Kuliczkowski K. Immunoglobulin D myeloma-problems with diagnosing and staging (own experience and literature review). Leuk Lymphoma. 2005;46:1029–37.

- Selene II, Jose JA, Khalil MJ, Faisal MS, Malik MN. Presentation patterns, diagnostic markers, management strategies, and outcomes of IgD multiple myeloma: a systematic review of literature. Cureus. 2019;11:e4011.
- 8. Ternes N, Rotolo F, Michiels S. Empirical extensions of the lasso penalty to reduce the false discovery rate in high-dimensional Cox regression models. Stat Med. 2016;35:2561–73.
- Liu Y, Huang A, Chen Q, Chen X, Fei Y, Zhao X, et al. A distinct glycerophospholipid metabolism signature of acute graft versus host disease with predictive value. JCI Insight. 2019;5:e129494.
- Boyd KD, Ross FM, Chiecchio L, Dagrada GP, Konn ZJ, Tapper WJ, et al. A novel prognostic model in myeloma based on co-segregating adverse FISH lesions and the ISS: analysis of patients treated in the MRC Myeloma IX trial. Leukemia. 2012;26:349-55.
- Shah V, Sherborne AL, Walker BA, Johnson DC, Boyle EM, Ellis S, et al. Prediction of outcome in newly diagnosed myeloma: a meta-analysis of the molecular profiles of 1905 trial patients. Leukemia. 2018;32:102–10.
- Kumar SK, Rajkumar SV. The multiple myelomas current concepts in cytogenetic classification and therapy. Nat Rev Clin Oncol. 2018;15:409–21.
- An G, Xu Y, Shi L, Zou D, Deng S, Sui W, et al. t(11;14) multiple myeloma: a subtype associated with distinct immunological features, immunophenotypic characteristics but divergent outcome. Leuk Res. 2013;37:1251–7.
- Chen L, Fan F, Deng J, Xu J, Xu A, Sun C, et al. Clinical characteristics and prognosis of immunoglobulin D myeloma in the novel agent era. Ann Hematol. 2019;98:963–70.
- Swan D, Delaney C, Natoni A, O'Dwyer M, Krawczyk J. Successful venetoclax salvage in the setting of refractory, dialysisdependent multiple myeloma with t(11;14). Haematologica. 2019; 105:e141–e143.
- Smith D, Yong K. Advances in understanding prognosis in myeloma. Br J Haematol. 2016;175:367–80.
- Goyal G, Rajkumar SV, Lacy MQ, Gertz MA, Buadi FK, Dispenzieri A, et al. Impact of prior diagnosis of monoclonal gammopathy on outcomes in newly diagnosed multiple myeloma. Leukemia. 2019;33:1273–7.
- Fibbe WE, Jansen J. Prognostic factors in IgD myeloma: a study of 21 cases. Scand J Haematol. 1984;33:471–5.
- Kim MK, Suh C, Lee DH, Min CK, Kim SJ, Kim K, et al. Immunoglobulin D multiple myeloma: response to therapy, survival, and prognostic factors in 75 patients. Ann Oncol. 2011; 22:411–6.