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Dietary advanced glycation end products and cancer risk in Japan: From the Takayama study

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Abstract

Few large epidemiological studies have evaluated the association between dietary advanced glycation end products (AGEs) and cancer risk. We evaluated the relationship between dietary AGE intake and the incidence of total cancer and site-specific cancers in a population-based prospective study in Japan. Participants were 14,173 men and 16,549 women who were 35 years of age or older in 1992. Dietary intake was assessed via a validated food frequency questionnaire. Intake of the AGE N^{ε} carboxymethyl-lysine (CML) was estimated using databases of CML content in foods determined using ultraperformance liquid chromatography-tandem mass spectrometry. Cancer incidence was confirmed through regional population-based cancer registries. During a mean follow-up period of 13.3 years, 1954 men and 1477 women developed cancer. We did not observe a significant association between CML intake and the risk of total cancer in men or women. In men, compared with the lowest quartile of CML intake, the hazard ratios of liver cancer for the second, third, and highest quartiles were 1.69 (95% CI: 0.92-3.10), 1.48 (95% CI: 0.77-2.84), and 2.10 (95% Cl: 1.10–3.98; trend p = 0.04). Conversely, a decreased relative risk of male stomach cancer was observed for the second and highest quartiles of CML intake versus the lowest quartile, with hazard ratios of 0.73 and 0.67, respectively (trend p = 0.08). Our finding on the potential harmfulness of consuming AGEs on liver cancer risk is intriguing and warrants further study.

KEYWORDS

AGEs, cancer, cohort studies, diet, epidemiology

Abbreviations: AGEs, advanced glycation end products; CI, confidence intervals; CML, N^e-carboxymethyl-lysine; FFQ, food frequency questionnaire; HMGB1, high mobility group box protein-1; MET, metabolic equivalent task; RAGE, receptor for advanced glycation end products; sRAGE, soluble receptor for advanced glycation end products; UPLC-MS/MS, ultraperformance liquid chromatography-tandem mass spectrometry.

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1 | INTRODUCTION

Advanced glycation end products are complex compounds formed by the nonenzymatic glycation between reducing sugars and free amino groups of proteins or lipids, termed the Maillard reaction. AGEs are produced and then accumulate endogenously in the body during the aging process or via glucose metabolism.¹ Other sources are exogenous AGEs derived from foods. The formation of dietary AGEs is accelerated by food preparation and processing steps that use a high temperature for a long period of time, including grilling, roasting, frying, and baking.^{2,3} High AGE levels are common in highly processed foods such as bread, biscuits, cooked meats, and foods containing high sugar levels. Of dietary AGEs consumed, 10%-30% are absorbed intestinally and enter the circulation.⁴ The ingestion of foods with high AGE content can increase the blood and tissue concentrations of AGEs, and exogenous AGE levels are usually much higher than those of endogenously produced AGEs,⁵ although the association between dietary intake and circulating levels of AGEs has shown conflicting results in human studies.^{6,7}

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AGEs have been associated with metabolic diseases such as diabetes and atherosclerosis as well as aging.^{1,5} AGEs and their receptor (RAGE) are implicated in oxidative stress and inflammation.^{1,5} Although the role of AGEs in cancer is unknown, oxidative stress related to AGEs may have critical consequences to cancer-associated DNA damage.⁸ The presence of AGEs in human tumors and higher circulating AGE levels among cancer patients have been demonstrated.^{9,10} AGEs promoted proliferation, migration, and invasion of cancer cell lines.^{11,12} Therefore, dietary AGEs, a major source of exogenous AGEs, may be associated with the development of cancer.

 N^{ε} -carboxymethyl-lysine has been the product most commonly used to estimate dietary AGEs, because CML is thought to be one of the most abundant AGEs in the body and has a well characterized chemical structure. To date, five prospective studies have evaluated the association between dietary AGEs and cancer risk. The NIH-AARP Diet and Health Study showed that CML intake was positively associated with the risk of pancreatic cancer in men¹³; it also reported a positive association between CML intake and postmenopausal breast cancer.¹⁴ Another US-based study, the Prostate, Lung, Colorectal, and Ovarian Cancer Screening Trial, reported that higher CML intake was associated with an increased risk of breast cancer.¹⁵ These studies estimated the CML content of foods using a database based on an ELISA. However, the accuracy and reliability of the contents based on this technique are reported to be questionable. CML values have been shown to be overestimated with ELISA for high-fat products such as butter, olive oil, and mayonnaise and underestimated for carbohydrate-rich foods such as cereals, biscuits, and cookies compared with a more accurate technique, namely UPLC-MS/MS.^{16,17} By contrast, the European Prospective Investigation into Cancer and Nutrition (EPIC) study estimated dietary CML and other two AGEs, N^{ε} -(1-carboxyethyl)lysine and N^{δ} -(5-hydro-5-methyl-4-imidazolon-2-yl)-ornithine, using databases based on UPLC-MS/MS. They observed that dietary CML was inversely associated with the risks of colorectal

cancer and hepatocellular carcinoma, and positively associated with gallbladder cancer. 18,19

Based on previous conflicting results, we evaluated the associations between dietary CML intake using databases based on UPLC-MS/MS and the incidence of total and site-specific cancers in a population-based prospective study in Japan. This cohort study is the first on dietary AGEs and cancer in Asians, whose dietary habits differ from those in the West.

2 | MATERIALS AND METHODS

2.1 | Participants and design

This study utilized data from the Takayama study in Japan, the details of which have been described previously.²⁰ The cohort has been followed to study the associations between lifestyle factors and mortality or cancer incidence. The subjects were 36,990 residents who were 35 years of age or older and were not hospitalized in September 1992. In total, 31,552 residents (85.3%) participated in the study and completed a self-administered questionnaire including an FFQ.²⁰ The questionnaire included demographic characteristics, medical history, physical activity, smoking status, and reproductive factors. This study was approved by the institutional review board of the Gifu University Graduate School of Medicine **(26-277)**.

2.2 | Dietary AGE intake and other exposure variables

The FFQ comprises questions on the consumption frequency of foods and dishes (169 items) and the usual portion size of meals during the previous year.²¹ The missing data for FFQ were assigned the mode categories by sex and 10-year age group. The intake of each nutrient and food group was estimated with this information using the Japanese Standard Table of Food Composition, 5th revised and enlarged edition.²² To guantify the intake of dietary AGEs, we estimated dietary CML intake using databases based on UPLC-MS/ MS. The details of CML intake estimation have been described elsewhere.²³ Briefly, we used a public repository at the University of Dresden based on UPLC-MS/MS.²⁴ A great contribution to this repository came from measurements made by Hull et al.²⁵ We also used another database of CML levels in foods by Scheijen et al., measured using UPLC-MS/MS.²⁶ CML values were assigned to each food with different cooking methods. Some CML contents varied between both databases even for the same foods with the same cooking methods. The databases of Scheijen et al. and Hull et al. contained CML values for 34.9% and 24.2% of the cumulative number of foods in our FFQ, respectively. So, the database of Scheijen et al. were given priority. Because the values of Hull et al. were an average of 21% higher than those of Scheijen et al. for the foods included in our FFQ, the values from Hull et al. were divided by 1.21. The CML values for powdered milk and almonds were derived from Drusch

et al. and Zhang et al., respectively (0.2% of the cumulative number of foods).^{27,28} When a CML value was not available for a food with a certain preparation method, we used values for similar foods according to the cooking methods. These assignments covered 56.3% of the cumulative number of foods, and these foods accounted for 79.3% and 82.1% of total energy in men and women, respectively. As for the values still missing by UPLC-MS/MS, we referred to the database based on ELISA by Uribarri et al.,² except for meat products and starchy foods. By comparing the values measured by Uribarri et al. with those of Scheijen et al. and Hull et al., conversion factors were calculated according to food group. After these processes, we assigned CML values to 61.3% of the cumulative number of foods, and these foods accounted for 82.7% and 86.1% of total energy in men and women, respectively. Finally, we assigned median CML values according to food group and cooking method for the remaining foods. In our FFQ, the food items with the highest CML values per weight were roasted peanuts, biscuits, cakes, and so on. Pork steak, hamburg steak, and fried rice had the highest CML values per serving. Fresh fruits and vegetables and alcohol beverages had almost zero CML values.

Smokers were defined as people who had smoked a total of at least 20 packs of cigarettes throughout their lifetime. The number of years of smoking was determined for ever smokers. Physical activity was assessed by asking the participants how much time on average they spent on various intensities of physical activities during the previous year. The period spent on each intensity-level activity (h/week) was multiplied by the corresponding energy expenditure, expressed as a MET, and the sum score of the product was regarded as the physical activity score (METs h/week). The details of the physical activity assessment including validity are described elsewhere.²⁹

2.3 | Outcomes and follow-up

After we excluded 830 people who were diagnosed with any cancer before baseline and/or reported a positive history of cancer in the baseline questionnaire, 14,173 men and 16,549 women aged 35-101 years were included in the analysis. Data on death and emigration were obtained from residential registers or family registers. Cause of death was identified from death certificates provided by the Legal Affairs Bureau. The incidence of cancer was confirmed mainly through two regional population-based cancer registries in Gifu. The causes of cancer were coded according to the International Classification of Diseases and Health Related Problems, 10th Revision. Cancer was defined as code C00-C97. Follow-up was conducted until the date of cancer diagnosis, date of death, date of moving out of the study area, or the end of March 2008, whichever came first. During the mean follow-up period of 13.3 years, 1954 men and 1477 women developed cancer. The mortality-to-incidence ratio was 0.49, and 23.1% of patients were ascertained by death certificate-only registration. For such cases, a backward tracking review of the incidence date was conducted based on the description given in the death certificate. As a result, only 7.7% of patients had

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an unknown incidence date before the date of death, indicating satisfactory completeness of cancer registration in this cohort. During the study period, 1535 persons (5.0%) moved out of the study area. Among 238 participants whose date of emigration was unknown (0.8%), their last date of residence in the study area was assigned as their censored date.

We further investigated the association between the intake of dietary AGEs and the incidence of site-specific cancers that had 40 or more cases each for men and women so that there would be 10 cases or more in each intake group; cancers of the esophagus (C15), stomach (C16), colon (C18), rectum (C19–20), liver (C22), gallbladder and biliary duct (C23–24), pancreas (C25), lung (C33–C34), bladder (C67), and prostate (C61) for men and cancers of the stomach, colon, rectum, liver, gallbladder and biliary duct, pancreas, lung, breast (C50), cervix (C53), endometrium (C54), and ovary (C56) for women were taken for analysis. When a person had developed a primary cancer at two or more sites, only the first cancer was applied. Lung cancer was analyzed separately for never and ever smokers.

2.4 | Statistical analysis

CML intake was controlled for energy intake by the energy density method (g/1000kcal of total energy). The study participants were categorized into quartile groups (Q1–Q4) according to their energyadjusted CML intake. The characteristics of the participants were calculated as the mean (standard deviation) or the percentage of each category according to the quartile groups of CML intake for each sex.

Hazard ratios and 95% CIs for total and site-specific cancers were estimated for the quartile groups of CML intake using the Cox proportional hazards model. The reference group was set as the lowest quartile of CML intake. After potential cancer risk factors were identified through a literature review, the confounders included age (years, continuous), body mass index (quartiles), education years (≤8, 9-11, 12-14, ≥15 years), history of diabetes (yes, no), physical activity score (METs h/week, continuous), smoking status (never, former, current smoker who had smoked for 30 years or less, current smoker who had smoked for 31 years or more), alcohol consumption (g/day, continuous), total energy intake (kcal/day, continuous), and menopausal status (premenopausal, postmenopausal).^{30,31} We additionally adjusted for salt intake (quartiles) for stomach cancer; height (quartiles), aspirin use (yes, no), and total fiber, red meat, and processed meat intake (quartiles) for colon and rectal cancers; coffee consumption (none, from once a month to six times per week, once per day or more) for liver cancer; and height (quartiles), age at menarche (≤12, 13-14, 15-16, ≥17 years), age at first delivery (no, ≤20, 20-25, 26-30, ≥31 years), age at menopause (≤49, ≥50 years), number of children (0, 1, 2, \geq 3), and hormone replacement therapy (yes, no) for cancers of the breast, cervix, endometrium, and ovary. Dummy variables were created for missing data for categorical covariates. Tests for a linear trend were performed using the median values of CML intake for each category.

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Sensitivity analysis was conducted when the median CML values based on UPLC-MS/MS according to food group and cooking method were assigned to the foods with missing values instead of using the CML values based on ELISA. In addition, the analysis was limited to those who had no history of diabetes because AGEs are closely associated with glucose metabolism. Last, cases who were diagnosed with cancer in the first 2 years were excluded from the analysis because they might have had latent cancers at baseline.

All analyses were performed using SAS version 9.4 (SAS Institute). *p*-values were calculated by a two-sided test. A *p*-value of less than 0.05 was considered statistically significant in all analyses.

3 | RESULTS

The characteristics of the participants are shown according to the quartile groups of energy-adjusted CML intake in Tables 1 and 2. Men with higher CML intake were more likely to be older, be more educated, have reported a history of diabetes, be physically less active, have never smoked, have lower alcohol consumption, and have higher total energy, protein, and fat intake. In women, similar

associations were observed, except that younger age and greater physical activity were associated with higher CML intake.

There were no significant multivariate-adjusted hazard ratios of total cancer between the quartiles of CML intake in men (Table 3). Compared with men in the lowest quartile of CML intake, men in the second, third, and highest quartiles had increased risks of liver cancer with a significant linear trend (hazard ratios: 1.69, 1.48, and 2.10, respectively, trend p = 0.04). Conversely, a decreased relative risk of male stomach cancer was observed for the second and highest quartiles of CML intake versus the lowest quartile, with hazard ratios of 0.73 and 0.67, respectively (trend p = 0.08). In women, no significant associations were observed between CML intake and total and any site-specific cancer (Table 4).

To evaluate whether modifying factors exist in the association between CML intake and male liver cancer, participants were stratified by history of diabetes (yes, no), age (\geq 55, <55 years), smoking (ever, never), and alcohol consumption (\geq 23, <23g of ethanol equivalent; Table S1). Because the number of male liver cancer was not sufficient to divide into quartiles after stratification, the tertiles of CML intake were applied for the stratified analyses. The significance for interaction was assessed using the product term

TABLE 1 Characteristics of study participants at baseline (men)

	No. of valid		Intake of N^{ε} -carboxymethyl-lysine (CML)			
	response	Mean (SD) or %	Q1	Q2	Q3	Q4
No. of participants	14,173		3544	3543	3543	3543
Age (years)	14,173	55.1 (12.3)	53.6 (11.3)	53.9 (11.8)	55.5 (12.6)	57.4 (13.2)
Body mass index (kg/m²)	13,367	22.5 (2.8)	22.5 (2.7)	22.5 (2.7)	22.5 (2.9)	22.5 (2.9)
Height (cm)	13,606	164.5 (6.9)	164.4 (6.8)	164.8 (6.7)	164.7 (7.0)	164.3 (6.9)
History of diabetes (yes, %)	14,173	6.3%	5.0%	4.6%	6.1%	9.5%
Education years (%)	14,006					
≤8 years		23.0%	22.1%	21.5%	23.8%	24.4%
9–11 years		35.2%	40.8%	36.1%	32.8%	31.1%
12–14 years		30.5%	29.3%	32.5%	30.8%	29.2%
≥15 years		11.4%	7.7%	10.0%	12.6%	15.4%
Physical activity score (METs h/week)	14,173	27.4 (41.1)	28.4 (42.2)	28.8 (42.3)	27.4 (40.1)	24.8 (39.4)
Smoking status (%)	13,766					
Never		16.6%	13.7%	15.1%	18.0%	19.8%
Former		29.6%	25.6%	27.6%	30.0%	35.2%
Current (for <30 years)		31.4%	35.7%	34.3%	30.9%	24.7%
Current (for ≥30 years)		22.4%	25.0%	23.0%	21.1%	20.3%
Alcohol consumption (g/day)	14,173	41.2 (41.2)	61.0 (50.3)	44.1 (39.4)	34.6 (34.5)	25.1 (28.2)
CML intake (mg/day)	14,173	2.74 (1.23)	1.80 (0.62)	2.49 (0.79)	2.98 (1.03)	3.71 (1.43)
Total energy intake (kcal/day)	14,173	2604 (874)	2421 (721)	2591 (801)	2665 (909)	2739 (1006)
Carbohydrate intake (g/day)	14,173	362 (117)	349 (98)	369 (111)	371 (124)	361 (131)
Protein intake (g/day)	14,173	92.9 (37.6)	72.0 (24.6)	88.6 (30.1)	98.6 (36.2)	112.3 (44.5)
Fat intake (g/day)	14,173	60.1 (28.8)	41.1 (17.1)	55.8 (21.5)	65.3 (26.5)	78.2 (33.7)

Abbreviations: SD: standard deviation; MET, metabolic equivalent of task; Q1-Q4: the quartile groups of energy-adjusted CML intake.

TABLE 2 Characteristics of study participants at baseline (women)

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	No. of valid response		Intake of N^{ε} -carboxymethyl-lysine (CML)			
		Mean (SD) or %	Q1	Q2	Q3	Q4
No. of participants	16,549		4138	4137	4137	4137
Age (years)	16,549	56.3 (13.2)	57.7 (13.8)	56.1 (13.3)	56.0 (12.9)	55.3 (12.6)
Body mass index (kg/m ²)	15,578	22.0 (2.9)	21.9 (3.1)	22.0 (3.0)	22.1 (2.9)	22.0 (2.8)
Height (cm)	15,832	151.9 (6.4)	151.3 (6.7)	151.8 (6.5)	152.2 (6.3)	152.4 (6.1)
History of diabetes (yes, %)	16,549	2.9%	2.1%	2.5%	3.7%	3.5%
Menopausal status (pre)	16,141	39.5%	36.3%	40.2%	39.5%	41.8%
Education years (%)	16,280					
≤8 years		28.1%	34.3%	28.4%	26.5%	23.0%
9–11 years		39.1%	41.2%	39.1%	38.7%	37.3%
12-14 years		28.3%	21.9%	28.3%	30.3%	32.8%
≥15 years		4.6%	2.6%	4.2%	4.5%	6.9%
Physical activity score (METs h/week)	16,549	19.0 (29.2)	16.1 (26.4)	19.0 (29.2)	20.1 (30.3)	20.7 (30.4)
Smoking status (%)	14,801					
Never		82.4%	77.2%	82.4%	85.3%	84.7%
Former		4.6%	5.0%	4.5%	3.9%	5.1%
Current (for <30 years)		11.1%	14.9%	11.2%	9.5%	8.7%
Current (for ≥30 years)		1.9%	3.0%	1.9%	1.3%	1.5%
Alcohol consumption (g/day)	16,549	7.7 (16.8)	11.7 (24.8)	7.9 (15.6)	6.0 (11.6)	5.1 (10.2)
CML intake (mg/day)	16,549	2.57 (1.18)	1.73 (0.71)	2.35 (0.88)	2.74 (1.00)	3.48 (1.29)
Total energy intake (kcal/day)	16,549	2127 (788)	1930 (713)	2090 (772)	2147 (779)	2340 (831)
Carbohydrate intake (g/day)	16,549	312 (111)	310 (108)	315 (117)	308 (111)	316 (108)
Protein intake (g/day)	16,549	81.7 (34.3)	63.5 (25.2)	77.6 (29.2)	85.6 (32.2)	100.3 (38.3)
Fat intake (g/day)	16,549	55.3 (27.1)	38.8 (19.4)	51.6 (22.0)	59.0 (24.6)	71.8 (30.1)

Abbreviations: SD, standard deviation; MET, metabolic equivalent of task; Q1-Q4, the quartile groups of energy-adjusted CML intake.

of the CML intake and the stratifying factor. There was no significant interaction by diabetes, age, smoking or alcohol consumption, although the hazard ratios of liver cancer for the highest tertile versus the lowest tertile of CML intake were 2.15 among ever smokers (trend p = 0.01) and 0.45 among never smokers (trend p = 0.40). Similarly, those were 1.86 among those with no history of diabetes (trend p = 0.03) and 1.18 among those with history of diabetes (trend p = 0.75).

When we performed sensitivity analysis without using the data based on ELISA, the results were not altered substantially; compared with the lowest quartile of CML intake, the hazard ratios of total cancer for the second, third, and highest quartiles were 0.98 (95% CI: 0.86–1.12), 1.01 (95% CI: 0.8–1.15), and 1.06 (95% CI: 0.93–1.22), respectively, in men (trend p = 0.34) and 0.90 (95% CI: 0.78–1.04), 0.94 (95% CI: 0.82–1.09), and 0.99 (95% CI: 0.85–1.15), respectively, in women (trend p = 0.93). The hazard ratio of male liver cancer for the highest quartile versus the lowest quartile of CML intake was 2.02 (95% CI: 1.07–3.81; trend p = 0.04). The hazard ratios of male stomach cancer for the second and highest quartiles versus the lowest quartile of CML intake were 0.80 (95% CI: 0.60–1.06) and 0.68 (95% CI: 0.49–0.96), respectively (trend p = 0.06).

The associations between CML intake and total and site-specific cancers were not substantially altered in men and women by excluding those who had a history of diabetes (899 men and 487 women; Tables S2 and S3).

When we re-analyzed the data after excluding cancer cases during the first 2 years of follow-up (205 men and 139 women), the nonsignificant associations between CML intake and total cancer did not change among men and women. The hazard ratio of male liver cancer for the highest quartile versus the lowest quartile of CML intake was 2.17 (95% CI: 1.07–4.38; trend p = 0.04). The hazard ratios of male stomach cancer for the second and highest quartiles versus the lowest quartile of CML intake were 0.73 (95% CI: 0.53–0.99) and 0.66 (95% CI: 0.46–0.96), respectively (trend p = 0.13).

4 | DISCUSSION

In this prospective study in Japan, we did not observe a significant association between CML intake and the risk of total cancer in men or women. Higher CML intake was significantly associated with an increased risk of male liver cancer. Conversely, we found decreased TABLE 3 Multivariate-adjusted hazard ratios (95% CI) for total cancer and site-specific cancer according to intake of N^{ε} -carboxymethyllysine among men

	Intake of N ^ɛ -carb	oxymethyl-lysine (CML)			Trend
	Q1	Q2	Q3	Q4	p
No. of subjects	3544	3543	3543	3543	
Person years	46,628	47,135	45,114	43,068	
Total cancer (C00-C97)					
No. of cancers	490	455	489	520	
Hazard ratio (95% CI)	1.00 (Ref.)	0.93 (0.82-1.07)	1.02 (0.89-1.16)	1.05 (0.91-1.20)	0.34
Esophageal cancer (C15)					
No. of cancers	26	14	12	14	
Hazard ratio (95% CI)	1.00 (Ref.)	0.70 (0.36-1.38)	0.74 (0.35-1.55)	0.97 (0.45-2.07)	0.86
Stomach cancer (C16)					
No. of cancers	119	91	122	98	
Hazard ratio (95% CI) ^a	1.00 (Ref.)	0.73 (0.54-0.98)	0.94 (0.69-1.27)	0.67 (0.47-0.96)	0.08
Colon cancer (C18)					
No. of cancers	70	49	65	65	
Hazard ratio (95% CI) ^b	1.00 (Ref.)	0.72 (0.49-1.07)	0.99 (0.66-1.48)	1.00 (0.64-1.58)	0.69
Rectal cancer (C19–C20)					
No. of cancers	30	43	25	42	
Hazard ratio (95% CI) ^b	1.00 (Ref.)	1.59 (0.96-2.64)	0.97 (0.52-1.79)	1.61 (0.86-3.00)	0.29
Liver cancer (C22)					
No. of cancers	18	28	23	35	
Hazard ratio (95% CI) ^c	1.00 (Ref.)	1.69 (0.92-3.10)	1.48 (0.77-2.84)	2.10 (1.10-3.98)	0.04
Gallbladder and biliary du	uct cancer (C23–24)				
No. of cancers	10	16	14	13	
Hazard ratio (95% CI)	1.00 (Ref.)	1.54 (0.69-3.47)	1.22 (0.51-2.87)	0.96 (0.38-2.42)	0.72
Pancreas cancer (C25)					
No. of cancers	21	19	13	25	
Hazard ratio (95% CI)	1.00 (Ref.)	0.86 (0.46-1.63)	0.58 (0.28-1.21)	1.03 (0.53–2.00)	0.98
Lung cancer (C33–C34)_e	ever smokers (<i>n</i> = 11	,478)			
No. of cancers	72	58	70	63	
Hazard ratio (95% CI)	1.00 (Ref.)	0.77 (0.54–1.09)	0.92 (0.65–1.31)	0.78 (0.54–1.15)	0.34
Lung cancer (C33–C34)_ı	never smokers (<i>n</i> = 2	288)			
No. of cancers	0	2	0	8	
Hazard ratio (95% CI)	-	-	-	-	-
Bladder cancer (C67)					
No. of cancers	35	21	29	33	
Hazard ratio (95% CI)	1.00 (Ref.)	0.58 (0.34-1.01)	0.80 (0.47-1.35)	0.84 (0.49-1.45)	0.77
Prostate cancer (C61)					
No. of cancers	32	37	52	54	
Hazard ratio (95% CI)	1.00 (Ref.)	1.08 (0.67-1.75)	1.47 (0.92-2.34)	1.37 (0.84-2.24)	0.13

Note: Hazard ratios adjusted for age (years), body mass index (quartiles), education years ($\leq 8, 9-11, 12-14, \geq 15$ years), history of diabetes (yes, no), physical activity score (METs h/week), smoking status (never, former, current smoker who had smoked for 30 years or less, current smoker who had smoked for 31 years or more), alcohol consumption (g/day), and intakes of total energy (kcal/day).

Abbreviations: CI: confidence interval; MET, metabolic equivalent of task; Q1-Q4: the quartile groups of energy-adjusted CML intake.

^aAdditionally adjusted for salt intake (g/day), in addition to the variables described above.

^bAdditionally adjusted for height (quartiles), aspirin use (yes, no) and intakes of red meat and processed meat (g/d), in addition to the variables described above.

^cAdditionally adjusted for coffee consumption (none, once a month to six times per week, ≥once per day), in addition to the variables described above.

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TABLE 4 Multivariate-adjusted hazard ratios (95% confidence interval) for total cancer and site-specific cancer according to intake of N^{ε} carboxymethyl-lysine among women

	Intake of <i>N^e</i> -carboxymethyl-lysine (CML)				
	Q1	Q2	Q3	Q4	Trend
No. of subjects	4138	4137	4137	4137	
Person years	55,180	56,569	56,878	56,625	
Total cancer (C00–C97)					
No. of cancers	404	354	360	359	
Hazard ratio (95% CI)	1.00 (Ref.)	0.92 (0.80-1.07)	0.94 (0.81-1.09)	0.97 (0.84-1.13)	0.72
Stomach cancer (C16)					
No. of cancers	73	56	56	48	
Hazard ratio (95% CI) ^a	1.00 (Ref.)	0.81 (0.56-1.17)	0.80 (0.54-1.18)	0.74 (0.49-1.13)	0.17
Colon cancer (C18)					
No. of cancers	59	67	48	57	
Hazard ratio (95% CI) ^b	1.00 (Ref.)	1.20 (0.83-1.74)	0.84 (0.55-1.28)	1.05 (0.68-1.62)	0.82
Rectal cancer (C19-C20)					
No. of cancers	25	30	15	22	
Hazard ratio (95% CI) ^b	1.00 (Ref.)	1.11 (0.63–1.96)	0.53 (0.26-1.07)	0.77 (0.39–1.54)	0.21
Liver cancer (C22)					
No. of cancers	19	15	20	16	
Hazard ratio (95% CI) ^c	1.00 (Ref.)	0.94 (0.47-1.86)	1.33 (0.69-2.54)	1.10 (0.54-2.21)	0.60
Gallbladder and biliary due	ct cancer (C23–C24)				
No. of cancers	21	11	19	19	
Hazard ratio (95% CI)	1.00 (Ref.)	0.57 (0.28–1.20)	0.95 (0.50–1.79)	1.02 (0.53-1.96)	0.78
Pancreas cancer (C25)					
No. of cancers	14	22	17	16	
Hazard ratio (95% CI)	1.00 (Ref.)	1.70 (0.86-3.33)	1.29 (0.63–2.64)	1.27 (0.61–2.67)	0.68
Lung cancer (C33–C34)_e	ver smokers ($n = 2607$)				
No. of cancers	15	12	5	11	
Hazard ratio (95% CI)	1.00 (Ref.)	1.24 (0.57-2.67)	0.58 (0.21-1.62)	1.33 (0.58–3.03)	0.83
Lung cancer (C33–C34)_n	ever smokers ($n = 12,1$	94)			
No. of cancers	17	11	12	9	
Hazard ratio (95% CI)	1.00 (Ref.)	0.65 (0.30-1.39)	0.65 (0.30–1.39)	0.50 (0.21-1.16)	0.10
Breast cancer (C50)					
No. of cancers	43	45	49	43	
Hazard ratio (95% CI) ^d	1.00 (Ref.)	0.96 (0.63-1.46)	1.02 (0.67–1.55)	0.88 (0.56–1.36)	0.63
Cervical cancer (C53)					
No. of cancers	11	15	8	14	
Hazard ratio (95% CI) ^d	1.00 (Ref.)	1.50 (0.67–3.38)	0.85 (0.33-2.19)	1.40 (0.59-3.29)	0.69
Endometrial cancer (C54)					
No. of cancers	6	9	14	14	
Hazard ratio (95% CI) ^d	1.00 (Ref.)	1.28 (0.45-3.63)	1.91 (0.72–5.03)	1.74 (0.65-4.66)	0.20
Ovarian cancer (C56)					
No. of cancers	13	7	23	10	

Note: Hazard ratios adjusted for age (years), body mass index (quartiles), education years (≤ 8 , 9–11, 12–14, ≥ 15 years), history of diabetes (yes, no), physical activity score (METs h/week), smoking status (never, former, current smoker who had smoked for 30 years or less, current smoker who had smoked for 31 years or more), alcohol consumption (g/day), and intakes of total energy (kcal/day).

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TABLE 4 (Continued)

Abbreviations: CI: confidence interval; MET, metabolic equivalent of task; Q1–Q4: the quartile groups of energy-adjusted CML intake. ^aAdditionally adjusted for salt intake (g/day), in addition to the variables described above.

^bAdditionally adjusted for height (quartiles), aspirin use (yes, no) and intakes of red meat and processed meat (g/day), in addition to the variables described above.

^cAdditionally adjusted for coffee consumption (none, once a month to six times per week, ≥once per day), in addition to the variables described above.

^dAdditionally adjusted for height (quartiles), age at menarche (\leq 12, 13–14, 15–16, \geq 17 year), age at first delivery (no, \leq 20, 20–25, 26–30, \geq 31 year), age at menopause (\leq 49, \geq 50 year), number of children (0, 1, 2, \geq 3), and hormone replacement therapy (yes, no), in addition to the variables described above.

risks of stomach cancer for the second and highest quartiles of CML intake in men. There were no significant associations between CML intake and any site-specific cancer in women.

We found that CML intake was positively associated with the risk of liver cancer in men. The liver is responsible for the clearance and catabolism of circulating AGEs.³² Absorbed AGEs preferentially accumulate in hepatic tissue.⁶ So, the detrimental effects of AGEs might be more likely to occur in the liver. Experimental studies have suggested that a high AGE diet plays a role initiating liver inflammation.^{33,34} Dietary AGEs enhance the accumulation of toxic AGE and upregulate RAGE mRNA expression leading to AGE-RAGE interactions, which releases proinflammatory molecules and induces oxidative stress in hepatocytes and hepatic stellate cells.³⁴ RAGE regulates tumorigenesis in inflammation-associated liver carcinogenesis, while sRAGE has a protective role in hepatocellular injury by neutralizing the effects mediated by the AGE-RAGE.^{35,36} One case-cohort study reported an inverse association between hepatocellular carcinoma and circulating CML and sRAGE among Finnish male smokers,³⁷ whereas the association between dietary intake and circulating levels of AGEs is unsure.^{6,7} The EPIC study, another study on dietary AGEs and liver cancer, observed an inverse association between dietary CML and hepatocellular carcinoma in men and women combined; this contradict our study that observed a positive association between dietary CML and liver cancer in men. The subjects of the EPIC study had a little younger average age (51 year) than those of this study (55 year). The foods contributing to CML intake in the EPIC study was cereals (35%), meat (19%), cakes and biscuits (14%) and dairy (11%). The main food sources of dietary CML in this study were cereals (27%), meat (15%), dairy (12%), and pulses (8%). Bread and whole grains are commonly consumed in Western countries, whereas rice and refined grains are widely consumed in Japan. The AGE content of bread was much higher than that of rice. Whole grain intake was reported to be associated with a reduced risk of hepatocellular carcinoma in Western populations.³⁸ The associations between dietary AGEs and diseases may be different between Western and Asian countries, where diets differ greatly.

The increased risks of liver cancer in relation to higher CML intake were observed among ever smokers but not among never smokers. Smoking may modify the association between dietary AGEs and liver cancer. AGE precursors are present in tobacco and AGEs accumulation in tissues has been demonstrated to be higher in smokers than in non-smokers.³⁹ Smoking, another determinant of exogenous AGEs, may contribute to the body's total AGE load together with dietary AGEs

and enhance the effects of AGEs on liver cancer. Conversely, the positive association between dietary AGEs and liver cancer seemed to be more pronounced in nondiabetic men than in those with diabetes. Increased endogenous AGEs due to diabetes might reduce the effects of dietary AGEs on liver cancer. However, neither smoking nor diabetes significantly modified the association between dietary AGEs and liver cancer, and the EPIC study reported no interaction by diabetes, age, smoking, or alcohol consumption. Future studies are needed to confirm whether these modifying effects exist and to clarify the role of exogenous and endogenous AGEs on liver cancer.

Positive association between dietary AGEs and liver cancer was not observed among women. The EPIC study has not reported the association by sex.¹⁹ Although the reason for the sex difference is unknown, it is possible that the dietary sources of CML are not exactly the same for men and women (Table S4). Whereas cereals and meat contributed more to CML in men, a variety of foods contributed more evenly to CML in women. Meat consumption was reported to not only contain AGEs, but also to be associated with greater accumulation of AGEs in the body.³⁹ As sex difference in lipoprotein or glucose metabolism in response to dietary intake has been suggested,⁴⁰ there might also be sex differences in the response of oxidative stress and inflammation to dietary AGEs.

Contrary to the hypothesis that dietary AGEs might influence cancer development, we observed a decreased risk of stomach cancer among men with high CML intake. The risk estimates for stomach cancer in women were similar to those for men, although they were not significant, perhaps due to the relatively small number of female stomach cancer. The explanation for this inverse association is not clear. RAGE was overexpressed in the mucosa with gastric cancer compared with the noncancerous tissues.^{41,42} Subjects with the variant genotypes associated with enhanced RAGE signaling (82Gly/Ser and 82Ser/Ser) had a significantly higher risk of gastric cancer compared with the wild-type carriers (82Gly/Gly).⁴³ A recent experiment has suggested that CML modification of HMGB1 enhances the cancer-promoting effect of HMGB1.44 Conversely, the final products of the Maillard reaction, called "melanoidins," have been demonstrated to exhibit biological activities including antioxidant activity, metal chelating ability, and lipid peroxidation inhibitory activity, particularly in the gastrointestinal tract.⁴⁵ It was also suggested that melanoidins derived from food proteins inhibit urease-gastric mucin adhesion and suppress Helicobacter pylori colonization.⁴⁶ So, these mechanisms in the gut might counteract the harmful effects of AGE-RAGE on stomach cancer. Another cohort

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study has suggested an inverse association between dietary CML and colorectal cancer risk.¹⁸

The previous two cohorts that estimated CML intake using the ELISA-based database have reported a positive association between dietary intake of AGEs and the risk of pancreatic¹³ and breast cancer.^{14,15} These results may not exactly be comparable with our study by the UPLC-MS/MS-based database that observed no significant association between dietary AGEs and pancreatic or breast cancer. The dietary sources that contributed the most to CML intake were fats and oils followed by meat products in those studies, different from those in our cohort (cereals followed by meat products). More prospective studies are warranted on the association between dietary AGEs and site-specific cancers in diverse populations.

The strengths of our study include its use of databases for the CML content of food based on HPLC-MS/MS as well as its prospective design. The good participation rate, long follow-up period, and the consideration of several confounding factors are also strengths. However, several limitations should be mentioned. Our FFQ was not validated for the assessment of CML intake. The Spearman correlation coefficients between the FFQ and 12 1-day diet records kept over a 1-year period for total energy, carbohydrate, protein, and fat intake were 0.44, 0.34, 0.38, and 0.24, respectively, for men and 0.53, 0.45, 0.63, and 0.52, respectively, for women. The CML values of foods on our FFQ were not completely covered by previous Western databases. In particular, the values for some foods typical to the Japanese diet were missing, although they were likely to be low in CML. For example, udon (a white wheat-based noodle) and tempura (a dish of deep-fried seafood or vegetables in light batter) were assigned CML values with reference to the values for boiled pasta and grilled cod, respectively. Sashimi (a dish of uncooked fish cut into small pieces) was assigned a CML of 0. In addition, the cooking methods, including the amount and duration of heat exposure, may differ between Japanese and Western diets, even for the same foods with the same cooking methods. However, measurement errors by such misclassifications are likely to have occurred independently of cancer incidence. We estimated only the CML intake in foods because CML has been used as a surrogate marker to estimate the AGE contents of foods and is commonly used in epidemiological studies, although other dietary AGEs might be differently associated with the risk of cancer. The information on exposure, including the FFQ, was collected only at baseline and changes in lifestyle habits were not evaluated during the follow-up period. We did not obtain information regarding infections with H. pylori and hepatitis virus, although these are major risk factors for stomach and liver cancer. The sample size was limited, and therefore precluded analyses of several cancers with a small number of cases. Dietary patterns or lifestyle habits might have changed due to preclinical signs or underlying disease. However, the exclusion of cancer cases during the first 2 years of follow-up did not change the results significantly.

In conclusion, this prospective study observed that high CML intake was significantly associated with an increased risk of liver cancer in Japanese men. Our finding suggests the potential harm of consuming AGEs on liver cancer risk. Furthermore, as dietary AGEs might have differential effects depending on the site of cancer, the associations between dietary AGEs and site-specific cancers need to be further investigated.

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DISCLOSURE

The authors have no conflict of interest. Dr. Chisato Nagata is a current Editorial Board Member of Cancer Science.

ETHICAL APPROVAL

Approval of the research protocol by an Institutional Reviewer Board: This study protocol was approved by the institutional review board of the Gifu University Graduate School of Medicine (26-277).

Informed Consent: Consent was waived due to the valid reasons for lack of consent at the time of the survey, the difficulty in obtaining consent now, and the public necessity of this research.

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SUPPORTING INFORMATION

Additional supporting information can be found online in the Supporting Information section at the end of this article.

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