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RESEARCH ARTICLE

Determinants of short birth interval among ever married reproductive age women: A community based unmatched case control study at Dessie city administration, Northern Ethiopia

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Abstract

Background

Short birth interval is a universal public health problem resulting in adverse fetal, neonatal, child and maternal outcomes. In Ethiopia, more than 50% of the overall inter birth spacing is short. However, prior scientific evidence on its determinants is limited and even then findings are inconsistent.

Methods

A community -based unmatched case-control study was employed on 218 cases and 436 controls. Cases were ever married reproductive age women whose last delivery has been in the past five years with birth interval of less than 3 years between the latest two successive live births whereas those women with birth interval of 3–5 years were taken as controls. A multistage sampling technique was employed on 30% of the kebeles in Dessie city administration. A pre-tested interviewer based questionnaire was used to collect data by 16 trained diploma nurses and 8 health extension workers supervised by 4 BSc nurses. The collected data were cleaned, coded and double entered into Epi-data version 4.2 and exported to SPSS version 22. Binary logistic regression model was considered and those variables with P<0.25 in the bivariable analysis were entered in to final model after which statistical significance was declared at P< 0.05 using adjusted odds ratio at 95% CI.

Result

In this study, contraceptive use (AOR = 11.2, 95% CI: 5.95-21.15), optimal breast feeding for at least 2 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth <25 years (AOR = 0.098, 95% CI:0.047-0.208), 95% CI:0.047-0.208, 95%

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Abbreviations: AOR, Adjusted odds ratio; EDHS, Ethiopian Demographics and Health Surveillance; FMOH, Federal Ministry of health; HEWs, Health Extension workers; LBW, Low birth weight; PROM, Premature rupture of membrane; WHO, World Health Organizations. 0.36, 95% CI: 0.282–0.761), having male preceding child (AOR = 0.46, 95% CI: 0.166– 0.793) and knowing the duration of optimum birth interval correctly (AOR = 0.45, 95% CI: 0.245–0.811) were significant determinants of short birth interval.

Conclusion

Contraceptive use, duration of breast feeding, age at first birth, preceding child sex and correct understanding of the duration of birth interval were significant determinants of short birth interval. Fortunately, all these significant factors are likely modifiable. Thus, the existing efforts of optimizing birth interval should be enhanced through proper designation and implementation of different strategies on safe breastfeeding practice, modern contraceptive use and maternal awareness about the health merits of optimum birth interval.

Background

Birthinterval refers to the time gap between two consecutive live births [1]. In 2005, World Health Organization consultation meeting on pregnancy intervals recommended a minimum inter pregnancy interval of at least 24 months to reduce the risk of adverse maternal, perinatal, and infant outcomes [2]. Moreover, the Ethiopian national family planning guideline recommends spacing childbirth at intervals of three to five years to reduce adverse fetomaternal and neonatal complications [3].

Short birth interval is a universal public health problem having association with adverse maternal, fetal, neonataland child outcomes such as low birth weight and perinatal death [4, 5], preterm delivery, small for gestational age [6], admission to neonatal intensive care unit [7], stillbirth, abortion, neonatal mortality [8], infant and under-5 mortality [8, 9], infant/child malnutrition including underweight, wasting, stunting [8, 10], neurodevelopmental and intellectual delay, autism, cerebral palsy [11], gestational diabetes [8, 12], precipitous labor [7], anemia [8, 13], uterine rupture, premature rupture of membrane, preeclampsia and chronic hypertension[8, 14, 15]. Most of these studies [4–7, 9–15] don't show causal association between short birth interval and the aforementioned pregnancy outcomes. Furthermore, the reported associations might have been largely attributed to confounding effects by genetically heritable familial factors [16, 17]. On the contrary, a systematic review of the available literature about the effects of birth spacing on maternal, perinatal, infant and child health witnessed the presence of causal mechanisms of association between short interbirth interval and its predictors [8].

Ethiopia had high population size as it was projected to reach more than 100 million and 4.0 total fertility rates in 2015. The country had also higher estimated pregnancy-related mortality ratio (PRM) of 412 deaths per 100,000 live births. Moreover, 1 in every 35 children dies within the first month; 1 in every 21 children dies before celebrating the first birthday; and 1 of every 15 children dies before reaching the fifth birthday (16). Therefore, the Ethiopian Federal Ministry of Health (FOMH) recommends spacing of childbirth at intervals of three to five years to reduce maternal, perinatal and infant mortality by optimizing the fertility rate in the country. However, in Ethiopia, more than 50% of the pregnancies occur within 3 years of their prior birth [18] which is shorter than the national recommendation of at least 3 years. Though initiatives like comprehensive implementation of family planning has been undertaken by the federal ministry of health at all levels of the health care system [3], the problem is of still greatest concern. This is so because birth intervals vary from society to society and within society itself within a country population [19, 20].

Since short birth interval is a potentially modifiable problem, a better knowledge and understanding of its determinants is imperative and essential to improve maternal health by designing and applying specifically targeted interventions thereby decreasing catastrophic pregnancy outcomes [9, 16]. However, evidence on the determinants of short birth interval in the study area is limited and even the nationally available data are inconsistent. Therefore, this study was aimed at identifying factors that have significant odds of association with short inter-birth interval among a community-based sample of Ethiopian women in Dessie city administration, 2019.

Methods

Study setting and period

Dessie city administration is located in northern part of Ethiopia at a distance of 401 km from Addis Ababa, capital of the country. It has an altitude of 2470 meters above sea level, situated between Tosa and Azewa mountains at11° 05′ North latitude and 39° 40′ East longitude. The city administration has 5 sub cities. Besides, for administrative sake, the city is categorized into 18 urban and 8 rural kebeles (the lowest administrative levels in the study area). Based on the 2014 Ethiopian population projection, Dessie district had a total population of 212,436 of whom 83.6% (177,688) lived in urban areas [19]. The study was held from 5/1/2019-12/5/2019.

Study design and participants' characteristics

A community based unmatched case-control study was conducted on a sample of eligible cases and controls. All the ever married reproductive age women who had at least two consecutive live births and whose last delivery within the past five years before the survey were eligible for the study.

The eligible women who had history of less than 3 years birth interval between their two successive live births were considered as cases. Besides, controls were considered to be those eligible women with birth interval of 3–5 years (including 3 and 5) between their two successive live births.

Sample size determination and sampling procedure

Taking several exposure variables into account, we calculated the respective sample size just by considering the assumption of case to control ratio of 1: 2; CI: 95%; Power: 80%; minimum detectable AOR = 2; design effect of 1.5 and 5% non-respondent rate. Among the given factors, we selected 'contraceptive use' because it yielded the maximum sample size as given in the following table (Table 1). Therefore, the final sample size was 678 (226 cases and 452 controls).

Then, multi stage sampling technique was employed to select the cases and controls. At first, 30% of the overall '*kebeles*' (three rural and five urban kebelles), were selected by simple random sampling technique. For those rural *kebeles*, the authors first checked family folder from health extension workers. We reviewed the family folder of permanently residing women in each kebele that fulfilled the inclusion criteria (less than 3 years birth intervalfor cases and 3–5 years' birth interval (including 3 and 5 years for controls)) by registering the birth date of the last two successive children in a family with their corresponding household identify permanently residing women that fulfilled the inclusion criteria (cases and controls) by registering the birth date of the last two successive children in a family with their corresponding household identify permanently residing women that fulfilled the inclusion criteria (cases and controls) by registering the birth date of the last two successive children in a family with their corresponding household identify permanently residing women that fulfilled the inclusion criteria (cases and controls) by registering the birth date of the last two successive children in a family with their corresponding household identify permanently residing women that fulfilled the inclusion criteria (cases and controls) by registering the birth date of the last two successive children in a family with their corresponding household identify permanently residing women that fulfilled the inclusion criteria (cases and controls) by registering the birth date of the last two successive children in a family with their corresponding household identification house the birth date of the last two successive children in a family with their corresponding household identification house the birth date of the last two successive children in a family with their corresponding household househ

Factors	Assumption	Total sample size	References
Contraceptive user	P of exposure in controls = 66.7%	678	(Hailu and Gulte, 2016)
Residence/urban	P of exposure in controls = 52.1%	540	(Yohannes <i>et al.</i> , 2011)
Husbands' occupation /Employee	P of exposure in controls = 51.7%	537	(Yohannes <i>et al.</i> , 2011)
Mothers' education /Has formal education	P of exposure in controls = 48.3%	524	(Hailu and Gulte, 2016)
Parity /> = 5 children	P of exposure in controls = 49.2%	524	(Begna Z. et al., 2013)
Sex of the index child /male	P of exposure in controls = 64.2%	638	(Begna Z. <i>et al.</i> , 2013)
Age of the mother/ 25–29	P of exposure in controls = 24.9%	576	(Begna Z. et al., 2013)
Status of index child /Alive	P of exposure in controls = 41.3%	509	(Tsegaye Dereje et al., 2017)
Wealth index/ Richest	P of exposure in controls = 25.2%	509	(Hailu and Gulte, 2016)

Table 1. Sample size determination involving different factors in the literature and the respective assumptions using open EPI INFO version 7 software.

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household identification number. Using the respective household identification number, a sampling frame of the households containing cases and controls was prepared for each kebele. Then, proportional allocation of sample size was employed to determine the study participants from each kebele. Finally, cases and controls were selected by simple random sampling technique from the existing sampling frame. Whenever more than one eligible woman was found in same selected household, only one woman was chosen by lottery method. Thus, a sample of 678 women (226 cases and 452 controls) was recruited from the sampling frame for the study (Fig 1).

Measurement and data collection procedure

Using interviewer based questionnaire, eight heath extension workers and sixteen diploma nurses underwent the data collection process including the weekend. During data collection, out of 654 eligible women (218 cases and 436 controls), 24 eligible women (8 cases and 16 controls) weren't accessed even after 2 different return visits. Therefore, these 24 absentees were replaced by other 24 randomly selected eligible mothers. The replaced mothers weren't systematically different from the original mothers because the replaced mothers were randomly selected from the already prepared sampling frame of eligible mothers (i.e. volunteers weren't included). Then, all the selected cases and controls were approached to be interviewed about factors related to their socio-demography, obstetrics, breastfeeding practice and modern contraception. Besides, the respondents were asked about their knowledge and attitude of birth interval. To determine children's birth dates, birth certificate or immunization cards were used. For those who were not immunized, health extension workers or mother's memory was consulted.

Data quality control

A structured English version interviewer based questionnaire (S1 Questionnaire) was first adapted from different literatures [1, 16, 20–22] and then translated to Amharic version (local language) for data collection purpose. The questionnaire was pretested just two weeks prior to the actual data collection using 33 eligible women (5% of the sample size) at the study area based on which some modifications were made to the originally prepared tool. Data collectors were closely monitored and guided by four BSC nurse supervisors. There was no missing information for any of the covariates in this study. This was because incomplete questionnaires were returned to the data collectors for completion by referring to the respective household identification number on a daily basis of checking all the questionnaires.



Fig 1. A flow diagram of sampling procedure.

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Data processing and analysis

Data were coded and double entered into Epi-Data software version 4.2 and then exported to SPSS version 22 for further processing and analysis. Descriptive statistics of different variables was done by cross tabulation. Binary logistic regression model using bivariable [crude odds ratio, [COR] and multivariable analyses [adjusted odds ratio, AOR] with 95% Confidence interval [CI] was employed. During bivariable analysis, variables whose p<0.25 were reserved for inclusion into the multivariable analysis in the final model after which statistical significance was declared at P<0.05 using adjusted odds ratio. Both Hosmer-Lemeshow's test (p = 0.753) and Omnibus Tests (p = .000) were used to check model fitness. Multi-collinearity was checked to see the linear correlation among the independent variables by using variance inflation factor and standard error. It was tried to minimize bias from intra-cluster correlation effect (dependencies) by considering only one of the eligible women in a selected household.

Besides, standard error was used during multivariate regressions and there was no any factor whose standard error greater than two indicating no dependency between mothers regarding the considered factors.

Estimation of household wealth index. Wealth index of the studied households were given scores based on the number and kinds of consumer goods they own including chairs, tables, chicken, transport (vehicles) and household characteristics like source of drinking water, toilet facilities, wall, roof and flooring materials. Among the nine characteristics, eight of them were extracted.

SPSS version 22 software was used to perform principal component analysis (PCA). Finally, wealth status was categorized into five groups and ranked from poorest to wealthiest quintile. Kaiser-Meyer-Olkin Measure of Sampling Adequacy was 0.751 and Bartlett's Test of Sphericity was significant.

Ethical approval and consent to participate. Ethical approval with ethics approval number of HU-CHMS-001 was obtained from Haramaya University, College of Health and Medical Sciences, Institutional Health Research Ethics Review Committee (IHRERC). An informed and voluntarily signed written consent (thumb print for those unable to write) was obtained from all the eligible mothers. Parental consent wasn't required because all the respondent mothers were above 16 years of old.

Results

Socio-demographic characteristics

From the overall sample of 678 mothers, 654 women (218 cases and 436 controls) agreed to be interviewed, thus making a response rate of 96.5%. Median age of the respondents at last delivery was 32 years. Twenty four (11%) of the cases and 82 (18.8%) of the controls were married at their age of 18 or less years. One hundred thirty five (61.9%) of the cases and 287 (65.8%) of the controls were within the age of25–34 years. Regarding their residence, 130 (59.6%) of the cases and 273 (62.6%) of the controls were urban residents. Nearly one fourth of the cases 58 (26.6%) and controls 110 (25.2.0%) had college and above level of education. One hundred twenty five (57.3%) of the cases and 232 (53.2%) of the controls were house wives. Moreover, 37(17.0%) of the cases and 107(24.5%) of the controls had the richest wealth index (Table 2).

Knowledge and attitude on birth interval

One hundred sixty five (75.7%) of the cases and 352 (80.7%) of the controls had ever heard about optimal birth interval. One hundred thirty four (61.5%) of the cases and 291 (66.7%) of the controls agreed that a minimum of 3 years birth spacing is essential between two successive births. Regarding husbands' perception of birth spacing, 120 (55%) of the cases and 246 (56.4%) of the controls had encouraging perception to birth spacing. One hundred forty four (66.1%) of the cases and 298(68.3%) of the controls had nobody to influence them to give birth with short interval. Two hundred and four (93.6%) of the cases and 404(92.7%) of the controls perceived that short birth interval have disadvantages on both maternal and child health. Regarding respondents' knowledge of the optimum birth interval, 130(78.8%) of the cases and 280(79.5%) of the cases 112(67.9%) and controls 289(82.1%) were health workers (Table 3).

Obstetrics related factors

The mean maternal age at first birth was $23(\pm 3.47)$ years. The median length of time from marriage to first birth was 24 months. Equal proportion (12.4%) of the cases and controls had

Factors	Category	Case (%)	Control(%)	P value	
Rsidence	Urban	130(59.7%)	273(62.65)	0.460	
	Rural	88(40.3%)	163(37.4%)		
Marital status	Married	186(85.3%)	364(83.5%)	0.759	
	Divorced	21(9.6%)	44(10.1%)		
	Widowed	11(5.1%)	28(6.4%)		
Religion	Orthodox	92(42.2%)	173(39.7%)	0.287	
	Muslim	124(56.9%)	249(57.1%)		
	Protestant	2(0.9%)	14(3.2%)		
Ethinicity	Amhara	200(91.7%)	399(91.5%)	0.926	
	Tgrai	7(3.2%)	11(2.5%)		
	Oromo	6(2.7%)	14(3.2%)		
	Others ¹	5(2.3%)	12(2.8%)		
Aother's education	No formal education	45(20.6%)	70(16.1%)	0.546	
	read and write	42(19.3%)	86(19.7%)		
	Elementary	34(15.6%)	81(18.6%)		
	Secondary	39(17.9%)	89(20.4%)		
	Collage and above	58(26.6%)	110(25.2%)		
Husband education	No formal education	50(22.9%)	69(15.8%)	0.104	
	read and write	32(14.7%)	69(15.8%)		
	Elementary	13(5.9%)	42(9.6%)		
	Secondary	41(18.8%)	72(16.5%)		
	College and above	82(37.6%)	184(42.2%)		
Mothers' occupation	employee(GO/NGO)	43(19.7%)	91(20.9%)	0.730	
-	house wife	125(57.3%)	232(53.2%)		
	Merchant	28(12.8%)	53(12.2%)		
	Student	9(4.1%)	29(6.7%)		
	Farmer	10(4.6%)	19(4.4%)		
	daily workers	3(1.4%)	11(2.5%)		
	Others ²	0(0%)	1(0.2%)		
 Husband occupation	employee(GO/NGO)	84(38.5%)	164(37.6%)	0.086	
-	Merchant	66(30.3%)	129(29.6%)		
	Student	0(0%)	2(0.5%)		
	Farmer	63(28.9%)	107(24.5%)		
	daily workers	4(1.8%)	23(5.3%)		
	Others ³	1(0.5%)	11(2.5%)		
Number of wives wealth index	One	216(99.1%)	434(99.5%)	0.478	
	More than one	2(0.9%)	2(0.5%)		
	Poorest	57(26.1%)	84(19.3%)	0.096	
	Second	35(16.1%)	80(18.3%)		
	Middle	47(26.6%)	83(19.0%)		
	Fourth	42(19.3%)	82(18.8%)		
	Richest	37(17.0%)	107(24.5%)		

Table 2. Socio-demographic characteristics on short birth interval among ever married mothers (case = 218, control = 436) in Dessie city administration, Dessie,
Ethiopia 2019.

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³Religious leader

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Factors	Category	Case (%)	Control (%)	P value
Heard about optimal birth interval	Yes	165(75.7%)	352(80.7%)	0.336
	No	53(24.3%)	84(19.3%)	
Optimum number of years between two successive births	Below three years	19(11.5)	46(13.1%)	0.70
	Three to five years	130(78.8%)	280(79.5%)	
	Above five years	13(7.8%)	23(6.5%)	
	I am not sure	3(1.8%)	3(0.8%)	0.562
minimum of 3 years of birth interval is essential between two successive births	Strongly agree	81(37.2%)	139(31.9%)	
	Agree	134(61.5%)	291(66.7%)	
	no idea	2(0.9%)	3(0.7%)	
	Disagree	1(0.5%)	3(0.7%)	
Husband's perception regarding birth spacing	Disagree strongly	28(12.8%)	27(6.2%)	0.00
	don't mind	57(26.1%)	152(34.9%)	
	Encouraging	120(55.04%)	246(56.4%)	
	Unknown	13(5.96%)	11(2.5%)	
External influences to give birth in short interval	My family	37(16.97%)	61(13.99%)	0.258
	Mother in law	21(9.63%)	60(13.76%)	
	Father in law	7(3.2%)	12(2.75%)	
	Societies norm	9(4.1%)	5(1.1%)	
	None	144(66.1%)	298(68.4%)	
Perceived advantages of optimum birth spacing	Yes	205(94.04%)	406(93.1%)	0.65
	No	13(5.96%)	30(6.9%)	
Perceived disadvantages of short birth interval	Yes	204(93.6%)	404(92.7%)	0.66
	No	14(6.4%)	32(7.3%)	

Table 3. Knowledge and attitude of birth interval among ever married reproductive age mothers (case = 218, control = 436) in Dessie city administration, Dessie,
Ethiopia 2019.

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bad fetal outcome at first delivery. Among the respondents, 5% of the cases and 2.5% of the controls experienced neonatal mortality. Besides, 3.7% of the cases and 1.8% of the controls had experienced stillbirth in their life time. Twenty five (5.7%) of the cases and 9(4.1%) of the controls had high birth order of their preceding child. From the overall respondents, 38 (17.4%) of the cases and 34 (7.8%) of the controls reported that their previous pregnancy was unplanned.

Forty six (21.1%) of the cases and 49(11.2%) of the controls had not ANC follow up for their previous pregnancy. Twenty five (11.5%) of the cases and 39(8.9%) of the controls had home delivery of their previous and last children. Majority of the cases 197(90.4%) and controls 397(89.9%) had spontaneous vaginal delivery of their previous child. Twenty six (11.9%) of the cases and 61(13.9%) of the controls ever had history of postpartum complications during their previous to last deliveries. From these complications, bleeding was reported among 6 (23.1%) of the cases and 30(49.2%) of the controls. The median duration of resuming postpartum sexual activity was 45 days. From the total respondents, 16 (7.3%) of the cases and 44 (10.1%) of the controls ever had chronic diseases like hypertension and diabetic mellitus before their last childbirth. The median ages of last and preceding child were 17and 60 months respectively (Table 4).

Breastfeeding and modern contraception related factors

Most of the cases 198(90.9%) breast fed their children for less than 24 months whereas 178 (40.8%) of the controls breastfed for at least 24 months. Moreover, more than half of the cases

Factors	Category	Case (%)	Control (%)	P value	
Fetal outcome of first delivery	Live birth	191(87.6%)	382(87.62%)	0.352	
	still birth	11(5.04%)	13(2.98%)		
	Abortion	3(1.4%)	13(2.98%)		
	Neonatal mortality	13(5.96%)	28(6.42%)		
Prior history of infertility	Yes	4(1.83%)	3(0.69%)	0.279	
	No	214(98.17%)	433(99.31%)		
Ever given birth to any child who died	Yes	31(14.2%) 58(13.3%)		0.723	
	No	187(85.8%)	378(86.7%)		
Male to female ratio of living children	More than one			0.355	
	One	63(28.89%)	135(30.96%)		
	Less than one	49(22.48%)	74(16.97%)		
	Males only	15(6.9%)	36(8.26%)		
	Females only	20(9.17%)	31(7.11%)		
Previous to last pregnancy is planned	Yes	180(82.6%)	402(92.2%)	0.001	
	No	38(17.4%)	34(7.8%)		
Practice postpartum abstinence before the last child	Yes	161(73.85%)	359(82.3%)	0.011	
	No	57(26.15%)	77(17.7%)		
Mode of delivery of previous to last birth	Vaginal delivery	197(90.4%)	392(89.9%)	0.981	
	Cesarean section	14(6.4%)	29(6.7%)		
	Instrumental delivery	7(3.2%)	15(3.4%)		
ANC follow up in preceding pregnancy	Yes	172(78.9%)	387(88.8%)	0.009	
	No	46(21.1%)	49(11.2%)		
Place of delivery of previous to last birth	Home	25(11.5%)	39(8.9%)	0.308	
	Health institution	193(88.5%)	397(91.1%)		
Pattern of menstruation in previous to last deliveries	Regular	Regular 185(84.9%) 362		0.550	
	Irregular	33(15.1%)	74(16.97%)		
Ever had chronic diseases (HTN, DM, others) before the last child	Yes	16(7.3%)	44(10.1%)	0.255	
	No	202(92.7%)	392(89.9%)		
Ever had history of postpartum complications in previous to last deliveries	Yes	26(11.9%)	61(13.99%)	0.464	
	No	192(88.1%)	375(86.01%)		
Last child sex	Male	121(55.5%)	238(54.6%)	0.824	
	Female	97(44.5%)	198(45.4%)		
Is last child alive	Yes	217(99.5%)	434(99.5%)	0.741	
	No	1(0.5%)	2(0.5%)		
previous to last child sex	Male	72(33%)	235(53.9%)	0.001	
-	Female	116(53.2%)	201(46.1%)		
Is previous to last child alive	Yes	215(98.6%)	434(99.5%)	0.254	
-	No	3(1.4%)	2(0.5%)		
Parity	<5	180 (82.5%)	370(84.8%)	0.450	
/	> = 5	38(17.5%)	66(15.2%)		

Table 4. Obstetrics related factors of short birth interval among ever married reproductive age mothers (case = 218, control = 436) in Dessie city administration, Dessie, Ethiopia 2019.

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80 (52.6%) and three fourth of the controls 295(73.8%) practiced exclusive breastfeeding to their preceding child. Ninety eight (44.9%) of the cases and 411 (94.3%) of the controls have utilized modern contraceptive methods after delivering their preceding child. Nearly all of the cases 213 (97.7%) and 434(99.5%) of the controls knew at least one type of modern contraceptive. One hundred eighty three (83.9%) of the cases and 428(98.2%) of the controls agreed that

family planning method is necessary for birth spacing. Regarding decision making about family planning in the house hold, ninety seven (44.5%) of the cases and 227(52.1%) of the controls decided based on couple agreement (Table 5).

Concerning the practice of modern contraceptive methods, forty three (43.9%) of the cases and 183(44.5%) of the controls utilized injectable type after delivering their preceding child (Fig 2).

Determinants of short birth interval

From the total fourteen variables that were entered to the multivariable logistic regression analysis, only five of them namely contraceptive use (AOR = 11.2, 95% CI: 5.95-21.15), optimal breast feeding for at least 2 years (AOR = 0.098, 95% CI:0.047-0.208), age at first birth<25 years (AOR = 0.36, 95% CI: 0.282-0.761), having male preceding child (AOR = 0.46, 95% CI: 0.166-0.793) and knowing the duration of optimum birth interval correctly (AOR = 0.45, 95% CI: 0.245-0.811) had significant odds of association with short birth interval. We used backward stepwise method to identify variables which had the largest contribution to the regression model. The result in forward or a stepwise variable selection method was similar on significance of the variables, but little change in adjusted odds ratio, p value and confidence interval were observed.

The odds of short birth interval among mothers who breastfed their prior child for at least 24 months were 90.2% lower (AOR = 0.098, 95% CI: 0.047–0.208) as compared to those having less than 12 months of breastfeeding duration. The odds of short birth interval among mothers having male preceding child was 54.0% lower than those whose child was female (AOR = 0.46, 95% CI: 0.166–0.793). Besides, the odds of short birth interval among those who didn't use modern contraceptives was11.2 times higher as compared to the users (AOR = 11.22, 95% CI: 5.95–21.15). Concerning maternal knowledge about the duration of birth interval, those mothers who knew the duration correctly had 55% lower odds of association with short birth

Factors	Category	Case (%)	Control (%)	P value	
Did you breast feed previous to last child	Yes	152(69.7%)	400(91.7%)	0.001	
	No	66(30.3%)	36(8.3%)		
Did you exclusively breastfeed previous to last child	Yes	80(52.6%)	295(73.8%)	0.001	
	No	72(47.4%)	105(26.2)		
Breast feeding duration	0-11	134(61.5%)	61(13.99%)	0.001	
	12-23	64(29.4%)	197(45.18%)		
	> = 24	20(9.2%)	178(40.83%)		
Jsing any of the modern methods before the conception of your last child	Yes	98(44.95%)	411(94.3%)	0.001	
	No	120(55.05%)	25(5.7%)		
Decision maker about Family planning	Self	104(47.7%)	190(43.58)		
	Both husband and wife	97(44.5%)	227(52.06%)	0.261	
	Husband only	3(1.4%)	13(2.98%)		
	No one	14(6.4%)	6(1.38%)		
Perception of family planning method	Agree	183(83.9%)	428(98.2%)	0.001	
	Disagree	34(15.6%)	4(0.9%)		
	Neutral	1(0.5%)	4(0.9%)		
Distance from health institution	Less than 30 minutes	93(42.7%)	197(45.2%)	0.799	
	30-1hrs	123(56.4%)	236(54.1%)		
	Greater than 1 hr	2(0.9%)	3(0.7%)		

Table 5. Breast feeding duration and contraceptive use among ever married reproductive age mothers in Dessie city administration, Dessie, Ethiopia 2019.

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interval (AOR = 0.45, 95% CI: 0.245-0.811) as compared to those who didn't know the duration correctly. Lastly, it was found that mothers who gave their first birth at the age of less than 28 years had 64% lower odds of association with short birth interval when compared to their counterparts (AOR = 0.46, 95% CI: 0.282-0.761) (Table 6).

Despite no statistical significance in the adjusted analysis, the crude odds of short birth interval was lower among mothers who had planned preceding pregnancy (COR = 0.401, 95% CI: 0.244–0.657) than those whose pregnancy wasn't planned. Besides, mothers who abstained in the post partum period had lower crude odds of short birth interval (COR = 0.61, 95% CI: 0.410–0.8941) than those who didn't abstain. Similarly, mothers who had ANC follow up [COR = 0.547, 95% CI: 0.348–0.858), mothers whose age at first marriage >25 years (COR = 5.256, 95% CI: 2.68–10.286), husband's encouraging perception of birth spacing (COR = 0.470, 95% CI: 0.266–0.833) and those mothers having the poorest wealth index (COR = 1.962, 95% CI: 1.187–3.245) were crudely associated with short birth interval (Table 6).

Discussion

This study was employed to investigate the determinants of short birth interval among ever married reproductive age mothers at Dessie city administration. Thus, from the adjusted analysis, it was found that contraceptive use, breast feeding duration, age at first birth, preceding child sex and knowing the appropriate duration of optimum birth interval correctly were significant determinants of short birth interval.

In this study, not using modern contraceptive method before getting pregnant of the last child was positively associated with short birth interval as compared to the users. This finding is similar to studies in Kassala, Eastern Sudan [23] and other prior Ethiopian studies [1, 20, 22, 24]. The consistency could be due to the fact that contraceptive use contributes to birth spacing

Factors	Case	Controls	Crude OR(95% CI)	p-value	AOR(95%CI)	p-value
Preceding pregnancy was planned						
yes	180	402	0.401(0.244-0.657)	.001	0.800(.348-1.839)	.599
no	38	34	1		1	
practice of postpartum abstinence in preceding child						
yes	161	359	0.606(0.410-0.8941)	.012	0.875(0.482-1.587)	.659
no	57	77	1		1	
ANC follow up in preceding pregnancy	1					
Yes	177	387	0.547(0.348-0.858)	.009	0.895(0.400-2.003)	0.787
No	41	49	1		1	
breast fed duration from previous to last child						
0-11	134	61	1		1	
12-23	64	197	0.148(0.098-0.224)	.001	0.291(0.154-0.550)	.001*
> = 24	20	178	0.051(0.029-0.089)	.001	0.098(0.047-0.208)	.001*
previous to Last child sex						
male	72	235	0.422(0.300-0.592)	0.01	0.463(0.282-0.761)	.002*
female	146	201	1		1	
using any of the modern methods before the conception of your last child		201	-		1	
yes	98	411	1	.001	11.221(5.953-21.151)	.001*
no	120	25	20.1(12.407-32.662)	1001		
knowledge to appropriate duration of birth interval	120					
correctly know	130	280	0.823(.589-1.149)	0.253	0.446(0.245-0.811)	.008*
not correctly know	88	156	1	0.235	1	
Husband education		150			-	
No formal education	82	138	1.302(0.926-1.830)	0.129	1.236 (0.633–2.416)	.535
Had formal education	136	298	1	01125	1	
age at first marriage		250				
less than 18	24	82	1		1	
18-25	154	328	1.604(0.979–2.628)	0.061	1.148(0.550-2.398)	.713
Greater than 25	40	26	5.256(2.68-10.286)	.001	0.478(0.113-2.024)	.316
age at first birth (years)		20	5.250(2.00 10.200)	.001	0.470(0.113 2.024)	.510
less than 28	160	413	0.154(0.092-0.257)	0.001	0.363(0.166-0.793)	0.011*
> = 28	58	23	1	0.001	1	0.011
no of living children	- 50	23	1		1	
0-2	55	90	1		1	
3-4	125	280	0.731(0.492–1.086)	0.120	.617(0.338–1.124)	.115
>=5	38	66	0.942(0.559-1.587)	0.120	1.109(0.489–2.514)	.696
Husband perception to birth spacing		00	0.942(0.559-1.587)	0.823	1.109(0.409-2.514)	.090
Disagree strongly	28	27	1		1	
Disagree strongly Dont mind				0.001		050
	57	152	0.362(0.196-0.666)		0.376(0.136-1.036)	.059
Encouraging	120	246	0.470(0.266-0.833)	0.010	0.557(0.221-1.401)	.214
Unknown	13	11	1.140(0.436-2.980)	0.790	0.873(0.195-3.908)	.859
Wealth index		0.4	1.0(2(1.107	0.000	2.012(0.052.1.515)	101
Poorest	57	84	1.962(1.187-3.245)	.009.	2.012(0.872-4.645)	.101
Second	35	80	1.265(0.733-2.183)	0.398	1.486(0.606-3.647)	.387
	477					
Middle Fourth	47 42	83 82	1.638(0.976-2.747) 1.481(0.874-2.510)	.062	2.378(1.086-5.210 1.823(0.780-4.262)	.030 0.166

Table 6. Multivariable analysis on the determinants of short birth interval among ever married reproductive age mothers in Dessie city administration, Dessie, Ethiopia, 2019.

 $^{*} \mathrm{for}$ Significant association at p<0.05)

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thereby reducing the total fertility rate by different mechanisms on normal reproductive process [25].

Mothers who breastfed their preceding child for at least 24 months had lower odds of short birth interval than those who breastfed for less than 12 months. This finding was supported by different studies which revealed lengths of birth interval to be influenced by duration of breastfeeding [26–28].Moreover, studies in Arba Minch District [20] and four disadvantaged regions of Ethiopia [21] showed similar finding which may be attributed to the fact that breast feeding has contraceptive effect due to the negative hormonal feedback mechanism of the hypothalamic-pitutary-ovarian axis. On the contrary, according to a community based cross sectional study in Southern Ethiopia, longer duration of breast feeding was significantly associated with increased incidence of short birth interval [22].The discrepancy might be due to differences in breast-feeding practices (exclusive breastfeeding, duration and frequency of breast feeding per 24 hours) and maternal factors (age, parity, nutritional status) [12] between the two studies. Besides, methodological and other socio-cultural differences between the two study populations might have contributed for the discrepancy.

Age at first birth was an important determinant of short birth interval. Hence, the odds of short birth interval among ever married reproductive age mothers who gave their first birth at the age of \geq 28 years were higher as regarded to those who gave their first birth at less than 28 years. This finding was consistent with evidences from a study in the United States [29]. The consistence might be due to the reason that elderly primiparity is often considered as a possible risk factor for limited fertility and hence elderly primiparous mothers rush to complete birthing of all their children as narrow spaced as possible [26]. But, this study was contrary to cross-sectional studies in Bangladesh [30, 31] which revealed that mothers having first birth at higher age usually have higher birth interval. The discrepancy could be attributed to the socio cultural and methodological variations among the two study population.

The study also showed that mothers who gave male child birth had lower odds to experience short birth interval than those whose child was female. This phenomenon was in line with evidences from case control studies in Arba Minch District [20] and rural pastoral communities of Southern Ethiopia [1]. The likely explanation of the congruence might be due to the fact that sex preference is a common culture in some communities so that giving son can be considered as a pride. Therefore, mothers who got female child from their prior birth become eager to be pregnant in short duration until they have the desired number of sons.

Mothers who knew the duration of optimum birth interval correctly had lower odds of short birth interval than those who didn't know. This finding was congruent with a case control study in Arba Minch District that showed lack of information about optimal birth spacing to be an indicated reason of short birth interval [20]. The likely explanation could be due to the fact that knowledge about the optimum inter birth interval is an important factor in motivating mothers to utilize family planning methods and practice safe breast feeding principles thereby preventing bad obstetric outcomes of short birth interval.

Based on our findings, local health care providers (physicians, midwives, nurses and health extension workers), the city health department and policy makers should focus on different strategies for creating parental awareness about the importance of modern contraceptive use, being primiparous before 28 years old and maternal knowledge of birth spacing. Moreover, we strongly recommend that mothers should prolong their breastfeeding practice for at least two years because its effect for optimizing birth interval has been witnessed by many other studies, WHO and UNICEF [32]. However, encouraging breast feeding up to two years may not warrant a reduction of birth interval because increasing breast feeding duration merely does not increase period of amenorrhea. This could in turn be due to differences among maternal breastfeeding practices, maternal age and parity. Women who are partially breast-feeding are

at higher risk of conceiving than women who are fully breast-feeding. The period of lactational amenorrhoea tends to be longer for older and multiparous than for younger and primiparous women. Besides, regardless of their breastfeeding practices, the other possible independent factor that may affect lactational infertility is maternal nutritional status. Therefore, despite the aforementioned confounders, maternal practice of optimal breastfeeding helps them optimize not only their health but also feto-neonatal and childhood survival.

Strength and limitation of the study

Using community based unmatched case control study design, high response rate and inclusion of both urban and rural communities could be considered as strengths of the study.

However, mothers' failure to recall of some important determinants like their own and children's age might have introduced recall bias into the study. Besides, accessing their socially desirable answers to some questions such as history of neonatal death would have caused social desirability bias. The recall bias was dealt with enabling mothers attach their children's birth dates to unforgettable Ethiopian holidays and calendar days. Besides, it was tried to minimize social desirability bias by conducting probed maternal interviews of the events (factors) by the trained data collectors. Some factors like husbands' perception of birth spacing may not have been measured appropriately. The study lacks support of qualitative data. Moreover, the results may not be representative of the ever married women of reproductive age group in Ethiopia due to smaller sample size in this study. Besides, the association of breastfeeding duration with inter-birth interval wasn't shown by subgroups of age, parity, breast feeding practices and nutritional status of the mothers, which can be considered as a limitation of the study. All the aforementioned limitations might have attributed for less precise measurement of some factors in the study.

Conclusion

From this study, contraceptive use, two and above years of breast feeding duration, less than 28 years of age at first birth, having male preceding child and knowing the duration of optimum birth interval correctly had significant negative odds of association with shortbirth interval.

Supporting information

S1 Questionnaire. Questionnaire used for data collection. (DOCX)

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References

- Begna Z., Sahilu Assegid, Wondwosen Kassahun and Mulusew Gerbaba. Determinants of inter birth interval among married women living in rural pastoral communities of southern Ethiopia: a case control study: BMC Pregnancy and Childbirth, 2013; 13:116. https://doi.org/10.1186/1471-2393-13-116 PMID: 23688144
- 2. WHO. Report of a WHO Technical Consultation on Birth Spacing Geneva, Switzerland. June 13–15, 2005. WHO, Department of Making Pregnancy Safer and Department of Reproductive Health and Research. 2006.
- 3. FMOH. National Guideline For Family Planning Services In Ethiopia. 2011.
- 4. Thapaliya R, Rai BK, Bhandari R, Rijal P and Gupta PP. The effect of birth interval on fetal outcomes. Health Renaissance; 2015; 13(3): 169–76.
- Mahande M. and Obure J. Effect of interpregnancy interval on adverse pregnancy outcomes in northern Tanzania: a registry-based retrospective cohort study. BMC Pregnancy Childbirth. 2016; 16: 140. https://doi.org/10.1186/s12884-016-0929-5 PMID: 27268015
- Kozuki Naoko, Lee Anne CC, Silveira Mariangela F, et al. The associations of birth intervals with smallfor gestational-age, preterm, and neonatal and infant mortality: a meta-analysis. BMC Public Health. 2013; 13 (Suppl 3):S3 2–9.
- Appareddy S, Pryor J and Bailey B. Inter-pregnancy interval and adverse outcomes: Evidence for an additional risk in health disparate populations. J Matern Fetal Neonatal Med. 2017; 30: 2640–4. https:// doi.org/10.1080/14767058.2016.1260115 PMID: 27903080
- Conde-Agudelo A., Rosas-Bermudez A., Castaño Fabio and Norton Maureen H.. Effects of Birth Spacing on Maternal, Perinatal, Infant, and Child Health: A Systematic Review of Causal Mechanisms. Studies in Family Planning 2012; Volume 43 93–114. https://doi.org/10.1111/j.1728-4465.2012.00308.x PMID: 23175949
- Dadi A. A systematic review and meta-analysis of the effect of short birth interval on infant mortality in Ethiopia. PLoS One. 2015; 10: e0126759. https://doi.org/10.1371/journal.pone.0126759 PMID: 26001079
- Shahjada A, Sharma BK, Sharma S, Mahashabde P and Bachhotiya A. Effects of birth interval on nutritional status of under five children in periurban area of Madhya Pradesh, India. International Journal of Medical Science and Public Health. 2014; 3: 723.
- Durkin MS, DuBois LA and Maenner MJ. Inter-Pregnancy Intervals and the Risk of Autism Spectrum Disorder: Results of a Population-Based Study. J Autism Dev Disord. 2015; 45: 2056–66. <u>https://doi.org/10.1007/s10803-015-2368-y</u> PMID: 25636677
- Hanley G. E., Hutcheon J. A., Kinniburgh B. A. and Lee L. Interpregnancy Interval and Adverse Pregnancy Outcomes: An Analysis of Successive Pregnancies. Obstet Gynecol. 2017; 129: 408–15. https://doi.org/10.1097/AOG.00000000001891 PMID: 28178044

- Ad Geidam, Inusa A and Isa B. Birth Interval and Maternal Outcome in University Of Maiduguri Teaching Hospital A Cross Sectional Study Borno Medical Journal 2016; Vol. 13: issue 2: 132–40.
- Wendt A, Gibbs CM, Peters S and Hogue CJ. Impact of increasing inter-pregnancy interval on maternal and infant health. Paediatr Perinat Epidemiol. 2012; 26 Suppl 1: 239–58.
- Bener A., Najah Mohammed Saleh, Khalil Salameh Khalil Mohd, et al. The impact of the interpregnancy interval on birth weight and other pregnancy outcomes. Rev Bras SaúdeMatern Infant, Recife, 2012; 12 (3): 233–41
- Class Q.A., Rickert M.E., Larsson H., Oberg A.S., Sujan A.C, Almqvist C., et al. (2018). Outcomedependent associations between short interpregnancy interval and offspring psychological and educational problems: A population-based quasi experimental study. International Journal of Epidemiology. https://doi.org/10.1093/ije/dyy042 PMID: 29566153
- Class Q.A., Rickert M.E., Oberg A.S., Sujan A.C., Almqvist C., Larsson H., et al. (2017). Within family analysis of inter-pregnancy interval and adverse birth outcomes. Obstetrics and Gynecology, 130 (6), 1304–1311. https://doi.org/10.1097/AOG.00000000002358 PMID: 29112654
- **18.** Central Statistical Agency (CSA) [Ethiopia] and ICF. Ethiopia Demographic and Health Survey 2016: Key Indicators Report. Addis Ababa, Ethiopia, and Rockville, Maryland, USA. CSA and ICF. 2016.
- **19.** CSA. Population and Housing Census of Ethiopia: Administrative Report. Addis Ababa. pp: 1–117. 2012.
- Hailu D and Gulte T. Determinants of Short Interbirth Interval among Reproductive Age Mothers in Arba Minch District, Ethiopia. Int J Reprod Med. 2016; 2016: 17. <u>https://doi.org/10.1155/2016/6072437</u> PMID: 27239553
- Dereje Tsegaye, Muluneh Shuremu and Kebebe Bidira. Practice of child spacing and its associated factors among women of child bearing age (15 to 49 years) in Illubabor zone, South West Ethiopia. International Journal of Nursing and Midwifery. 2017; 9: 102–8.
- Yohannes S, Wondafrash M, Abera M and Girma E. Duration and determinants of birth interval among women of child bearing age in Southern Ethiopia. BMC Pregnancy Childbirth. 2011; 11:38. <u>https://doi.org/10.1186/1471-2393-11-38 PMID: 21599927</u>
- 23. Ali A., Yassin Khalid and Ramadan Nagla. Determinant of Inter-Pregnancy Birth Interval in Kassala, Eastern Sudan Current Women's Health Reviews, 2014; 10 5–8.
- 24. Dereje Tesfaye, Shibru Temesgen and Zinabu Teka. Analysis of Determinants of Birth Interval in Four Disadvantaged Regions of Ethiopia. TurkiyeKlinikleri Journal of Biostatistics. 2015; 7: 63–76.
- 25. WHO. World Health Organization, Contraception Issues in Adolescent Health and Development. 2004.
- 26. Japheth OsotsiAwiti. Preceding Birth Interval Length and Maternal Health in Kenya. 2013.
- Owonikoko K, Adeniji O, Oke O, Fawole A and Adeniji A. Determinants of inter-pregnancy interval in Ogbomoso: an unmet need for contraceptive usage. International Journal of Reproduction, Contraception, Obstetrics and Gynecology. 2015; 4: 316.
- 28. Saadehl R. and Benbouzid D.. Breast-feeding and child-spacing: importance of information collection for public health policy. WHO Bulletin OMS Vol 68 1990.
- 29. Gemmill A. and Lindberg L. D. Short interpregnancy intervals in the United States. Obstet Gynecol. 2013; 122: 64–71. https://doi.org/10.1097/AOG.0b013e3182955e58 PMID: 23743455
- **30.** De Jonge H., Azad K., Seward N., et al. Determinants and consequences of short birth interval in rural Bangladesh: a cross-sectional study. BMC Pregnancy Childbirth. 2014; 14: 427. <u>https://doi.org/10.1186/s12884-014-0427-6 PMID: 25539669</u>
- **31.** FazleRabbi Ahbab Mohammad, Karmaker Shamal Chandra, Mallick Shahadat Ali and Sharmin Sayema. Determinants of Birth Spacing and Effect of Birth Spacing on Fertility in Bangladesh. Dhaka Univ J Sci). 2013; 61(1:105–10.
- UNICEF, 2015. Breastfeeding | Nutrition | https://www.unicef.org/nutrition/index_24824.html accessed on June 01/2020 at 2:00 pm