

# The Associations between Apolipoprotein E Gene Epsilon2/Epsilon3/Epsilon4 Polymorphisms and the Risk of Coronary Artery Disease in Patients with Type 2 Diabetes Mellitus

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**Background and Objective:** Apolipoprotein E (APOE) plays important roles in lipoprotein metabolism and cardiovascular disease. Evidence suggests the *APOE* gene epsilon2/epsilon3/epsilon4 ( $\epsilon 2/\epsilon 3/\epsilon 4$ ) polymorphisms might be associated with the susceptibility of coronary artery disease (CAD) in patients with type 2 diabetes mellitus (T2DM). However, no clear consensus has yet been established. Therefore, the aim of this meta-analysis is to provide a precise conclusion on the potential association between *APOE*  $\epsilon 2/\epsilon 3/\epsilon 4$  polymorphisms and the risk of CAD in patients with T2DM based on case-control studies.

**Methods:** Pubmed, Embase, Chinese National Knowledge Infrastructure (CNKI), and Wanfang databases were searched for all relevant studies prior to August 2017 in English and Chinese language. The pooled odds ratios (ORs) and their corresponding 95% confidence intervals (CIs) were used to assess the strength of the relationships. The between-study heterogeneity was evaluated by Cochran's Q-test and the I<sup>2</sup> index to adopt fixed- or random- effect models.

**Results:** A total of 13 studies were eligible for inclusion. There was evidence for significant associations between *APOE*  $\varepsilon$ 4 mutation and the risk of CAD in patients with T2DM (for  $\varepsilon_3/\varepsilon_4$  vs.  $\varepsilon_3/\varepsilon_3$ : OR = 1.69, 95% CI = 1.38–2.08, *P* < 0.001; for  $\varepsilon_4/\varepsilon_4$  vs.  $\varepsilon_3/\varepsilon_3$ : OR = 2.72, 95% CI = 1.61–4.60, *P* < 0.001; for  $\varepsilon_4/\varepsilon_4+\varepsilon_3/\varepsilon_4$  vs.  $\varepsilon_3/\varepsilon_3$ : OR = 1.83, 95% CI = 1.52–2.22, *P* < 0.001; for  $\varepsilon_4$  allele vs.  $\varepsilon_3$  allele: OR = 1.64, 95% CI = 1.40–1.94, *P* < 0.001). In contrast, no significant associations were found in genetic model of *APOE*  $\varepsilon_2$  mutation (for  $\varepsilon_2/\varepsilon_2$  vs.  $\varepsilon_3/\varepsilon_3$ : OR = 1.67, 95% CI = 0.90–3.09, *P* = 0.104; for  $\varepsilon_2/\varepsilon_3$  vs.  $\varepsilon_3/\varepsilon_3$ : OR = 1.18, 95% CI = 0.93–1.51, *P* = 0.175; for  $\varepsilon_2/\varepsilon_2+\varepsilon_2/\varepsilon_3$  vs.  $\varepsilon_3/\varepsilon_3$ : OR = 1.26, 95% CI = 0.88–1.82, *P* = 0.212; for  $\varepsilon_2$  allele vs.  $\varepsilon_3$  allele: OR = 1.34, 95% CI = 0.98–1.84, *P* = 0.07).

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1

**Conclusions:** The *APOE* gene  $\epsilon$ 4 mutation is associated with an increased risk of CAD in patients with T2DM, while the  $\epsilon$ 2 variation has null association with this disease.

Keywords: coronary artery disease, type 2 diabetes mellitus, apolipoprotein E, epsilon2, epsilon3, epsilon4, genetic polymorphism

# INTRODUCTION

Type 2 diabetes mellitus (T2DM) is a long-term metabolic disease with a high incidence and prevalence in the world. T2DM is often accompanied by various complications such as hypertension, dyslipidemia and coronary artery disease (CAD) (Naito and Miyauchi, 2017). As the disease progresses, patients with T2DM have a 2 to 4-fold increased risk for developing CAD compared with non-diabetic individuals (Mohan et al., 2001; Emerging Risk Factors et al., 2010). In addition, cardiovascular disease including CAD in patients with T2DM is associated with significant mortality (Zhang et al., 2014b; Freitas Lima et al., 2015). Therefore, early prevention and vigorous control of T2DM and its complications are becoming an ever-increasing global health priority. A better understanding of the etiology of CAD in patients with T2DM will result in better clinical management.

Dyslipidemia, hypertension, obesity, and smoking status are well-established risk factors for T2DM (Paneni et al., 2013; Wang et al., 2015a). Additionally, human genetic association studies have revealed that numerous genetic mutations and polymorphisms also play a critical role (Wei et al., 2014; Raj et al., 2015; Sumi et al., 2017). Among the previous studies, apolipoprotein E (APOE) gene has been regarded as one of the most likely candidate genes which may be associated with CAD in T2DM patients.

APOE is a class of plasma apolipoprotein totaling 299 amino acids, and it is involved in lipoprotein metabolism and the development of cardiovascular diseases (Zheng et al., 1998). The APOE gene is mapped to chromosome 19q13.2 in a cluster with apolipoprotein C1 and C2 gene, and it consists of three introns and four exons. APOE is a polymorphic gene and the most commonly studied alleles/isoforms are: epsilon2 (ɛ2), epsilon3 (£3), and epsilon4 (£4). The differences between the three isoforms are the location of 112 and 158 in the amino acid chain where cysteine or arginine is present. These three APOE alleles are determined by the rs7412 and rs429358 singlenucleotide polymorphisms. The three alleles, APOE-E2 (cys112 and cys158), APOE-ɛ3 (cys112 and arg158) and APOE-ɛ4 (arg112 and arg158), yield six different genotypes for the APOE gene:  $\epsilon 2/\epsilon 2$ ,  $\epsilon 2/\epsilon 3$ ,  $\epsilon 2/\epsilon 4$ ,  $\epsilon 3/\epsilon 3$ ,  $\epsilon 3/\epsilon 4$ , and  $\epsilon 4/\epsilon 4$ . Because the  $\epsilon 3$  allele or  $\varepsilon 3/\varepsilon 3$  genotype is the most common allele or genotype among the population, they are well accepted as the "wild-type" and used as the "reference" in the genetic models (Zhang et al., 2000; Guo et al., 2007; Izar et al., 2009; Chaudhary et al., 2012; Hong et al., 2017).

The role of *APOE*  $\epsilon 2/\epsilon 3/\epsilon 4$  polymorphisms in the development of CAD in patients with T2DM is widely studied, but the results are still controversial and conflicting. In 1998,

Zheng et al. firstly investigated the association between *APOE* gene polymorphism and T2DM complicated with CAD in the Chinese population. The results showed that *APOE*- $\epsilon$ 4 allele increased the risk of CAD in T2DM (Zheng et al., 1998). Other studies have also confirmed Zheng's findings (Chaaba et al., 2008; Hong et al., 2017). However, *APOE*- $\epsilon$ 2 allele was also found to be associated with the risk of CAD in T2DM (Halim et al., 2012). In addition, no significant association between *APOE*  $\epsilon$ 2/ $\epsilon$ 3/ $\epsilon$ 4 polymorphisms and the risk of CAD in T2DM was reported in some studies (Zhang et al., 2000; Guo et al., 2007; Izar et al., 2009). To demonstrate with certainty the associations between the *APOE*  $\epsilon$ 2/ $\epsilon$ 3/ $\epsilon$ 4 polymorphisms and the risk of CAD in patients with T2DM, we conducted a systematic review and meta-analysis on published case-control studies.

# MATERIALS AND METHODS

#### Literature Search

This study was undertaken according to the methodology of MOOSE (Meta-analysis of Observational Studies in Epidemiology) statement (Stroup et al., 2000). We thoroughly searched all published studies in the Embase, PubMed, China National Knowledge Infrastructure (CNKI) and Wanfang databases up to August 2, 2017. The included articles were limited to Chinese and English language. The following keywords were used for searching: "apolipoprotein E" OR "APOE" AND "polymorphism" OR "single nucleotide polymorphism" OR "SNP" OR "variant" OR "variation" AND "coronary artery disease" OR "coronary heart disease" OR "CAD" OR "CHD" OR "atherosclerosis" OR "myocardial infarction" OR "myocardial infarct" OR "heart attack" OR "MI" AND "type 2 diabetes" OR "non-insulin dependent diabetes mellitus" OR "diabetes mellitus, type 2" OR "diabetes, type 2" OR "diabetes mellitus, non-insulin dependent" The Chinese databases were searched using the equivalent Chinese terms. In addition, hand searches for all related articles were performed. The detailed search strategies are presented in Supplementary Table 1.

# **Inclusion and Exclusion Criteria**

The first two investigators independently accessed the eligibility of the studies by screening the title, abstract and full-text, based on the inclusion and exclusion criteria. The inclusion criteria for all studies were as follows: (1) study on the associations between *APOE*  $\varepsilon 2/\varepsilon 3/\varepsilon 4$  polymorphisms and CAD in patients with T2DM, regardless of sample size. (2) case-control design. (3) detailed data for the *APOE* alleles or genotype distribution in case and control groups to estimate odds ratio (OR) with 95% confidence interval (CI). Exclusion criteria: (1) duplication of previous data; (2) review, comment and editorial; (3) no sufficient genotype data. Any dispute about the eligibility of an article was resolved by discussion.

#### **Data Extraction**

The data was drawn out based on a standard protocol. The following information was carefully extracted from all eligible publications independently by two authors (JQL and HR) using a standardized form: last name of first author, year of publication, study country, sample size in cases and controls, methods of genotyping, number genotypes and alleles. If similar data sets presented in different articles by the same research group, the data would be adopted only once. The collected data were compared, and possible disagreements were discussed by the authors and resolved with consensus.

#### **Quality Score Assessment**

The study quality was independently assessed by two reviewers. Quality assessment of genetic associations between *APOE* 

 $\epsilon 2/\epsilon 3/\epsilon 4$  polymorphisms and CAD in patients with T2DM is described in the Supplementary Table 2. The scores were adjusted according to the criteria developed for meta-analysis of molecular association studies by Thakkinstian et al. (2005). The total scores ranged from 0 to 13, with 13 representing the highest quality.

### **Statistics Analysis**

All the statistical analysis in this study was performed using Stata 12.0 (StataCorp, College Station, TX). Hardy-Weinberg equilibrium was performed in control groups using the chisquare test. The combined OR and 95%CI were calculated to evaluate the strength of the association between the *APOE*  $\epsilon 2/\epsilon 3/\epsilon 4$  polymorphisms and risk of CAD in T2DM subjects. The pooled ORs were, respectively, performed for nine genetic models ( $\epsilon 2/\epsilon 2$  vs.  $\epsilon 3/\epsilon 3$ ;  $\epsilon 2/\epsilon 3$  vs.  $\epsilon 3/\epsilon 3$ ;  $\epsilon 2/\epsilon 4$  vs.  $\epsilon 3/\epsilon 3$ ;  $\epsilon 3/\epsilon 4$ vs.  $\epsilon 3/\epsilon 3$ ;  $\epsilon 4/\epsilon 4$  vs.  $\epsilon 3/\epsilon 3$ ;  $\epsilon 2$  allele vs.  $\epsilon 3$  allele;  $\epsilon 4$  allele vs.  $\epsilon 3$  allele;  $\epsilon 2/\epsilon 2+\epsilon 2/\epsilon 3$  vs.  $\epsilon 3/\epsilon 3$ ;  $\epsilon 4/\epsilon 4+\epsilon 3/\epsilon 4$  vs.  $\epsilon 3/\epsilon 3$ ). The statistically significant level of the combined OR was determined



TABLE 1	Characteristics	of the included	studies for this	meta-analysis.
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First-author	Year	Country	Genotyping methods <sup>a</sup>	Quality score	Samp	e size			А	POE ge	enotype	es distr	ibution (	case/con	trol)		
					Case	Control	ε2/ε2	ε2/ε3	ε <b>3/ε</b> 3	ε3/ε4	ε4/ε4	ε <b>2</b> /ε4	ε2	ε3	ε4	ε2/ε2+ ε2/ε3	ε3/ε4+ ε4/ε4
Hong	2017	China	RT-PCR	10	114	106	1/1	14/11	61/72	31/20	2/0	5/2	21/15	167/175	40/22	15/12	33/20
Chaudhary	2012	Thailand	PCR-RFLP	12	147	155	1/1	11/2	88/117	46/30	1/4	0/1	13/5	233/266	48/39	12/3	47/34
Halim	2012	Egypt	PCR-RFLP	5	35	35	6/0	5/2	18/31	6/2	0/0	0/0	17/2	47/66	6/2	11/2	6/2
Al-Majed	2011	Kuwaiti	PCR-RFLP	9	41	105	3/7	1/2	21/73	4/6	12/15	0/2	7/18	47/154	28/38	4/9	16/21
Vaisi-Raygani	2010	Iran	PCR-RFLP	12	172	118	4/0	31/26	91/69	31/20	12/3	3/0	42/26	244/184	58/26	35/26	43/23
Shi	2009	China	PCR-RFLP	9	98	110	0/0	4/3	44/71	36/27	2/0	12/9	16/12	128/172	52/36	4/3	38/27
Izar	2009	Brazil	PCR-RFLP	11	386	604	3/7	60/86	241/388	57/81	9/4	14/31	80/131	599/943	89/120	63/93	66/85
Chaaba	2008	Tunisia	PCR-RFLP	9	71	86	0/0	3/9	57/68	NA	NA	0/1	NA	NA	NA	3/9	11/8
Zhang L	2008	China	PCR-RFLP	10	100	100	2/4	12/15	54/67	30/13	2/1	0/0	16/23	150/162	34/15	14/19	32/14
Guo	2007	China	Multi-ARMS- PCR	11	40	40	0/0	2/1	22/29	13/7	1/0	2/3	4/4	59/66	17/10	2/1	14/7
Pan	2002	China	PCR-RFLP	11	24	63	0/1	4/7	12/45	6/8	0/0	2/2	6/11	34/105	8/10	4/8	6/8
Zhang WH	2000	China	PCR-RFLP	9	61	63	1/0	2/7	46/50	10/6	1/0	1/0	5/7	104/113	13/6	3/7	11/6
Zheng	1998	China	PCR-RFLP	8	33	78	NA	NA	22/59	NA	NA	NA	NA	NA	NA	3/15	8/4

<sup>a</sup>Multi-ARMS-PCR: multiplex amplification refractory mutation system-polymerase chain reaction; PCR-RFLP: polymerase chain reaction restriction fragment length polymorphism; RT-PCR: real-time polymerase chain reaction.

NA: not available.

by the Z-test with P < 0.05. Heterogeneity between studies was calculated by using the Cochran's Q-test and Higgins I<sup>2</sup> index. In the absence of between-study heterogeneity ( $I^2 < 50\%$ ), the fixed effect model (Mantel-Haenszel method) was chosen to calculate the pooled estimates. Otherwise, random effect model (DerSimonian and Laird method) would be adopted if the  $I^2$ > 50% (Higgins et al., 2003). Subgroup analysis was performed according to the source of patients (Chinese and non-Chinese). Galbraith plot analysis and sensitivity analysis were conducted to detect whether there were outliers that could be the potential sources of heterogeneity between studies when heterogeneity was moderately large. Publication bias was evaluated by Begg's funnel plot and Egger's regression test (Begg and Mazumdar, 1994; Egger et al., 1997). If there is evidence of significant publication bias, the trim and fill method was performed to assess the potential influence of publication bias (Duval and Tweedie, 2000).

#### RESULTS

#### The Characteristics of the Included Studies

As depicted in **Figure 1**, a total of 222 articles were obtained by online search, and 2 articles were included by manual search. After removing duplicates, 175 articles were included. After screening title and abstract, 115 articles were excluded. As a result, 13 articles (Zheng et al., 1998; Zhang et al., 2000, 2008; Pan et al., 2002; Guo et al., 2007; Chaaba et al., 2008; Izar et al., 2009; Shi et al., 2009; Vaisi-Raygani et al., 2010; Al-Majed et al., 2011; Chaudhary et al., 2012; Halim et al., 2012; Hong et al., 2017) were eligible for the meta-analysis. The characteristics of the included articles are summarized in **Table 1**. The included studies were conducted in several countries including China,

Brazil, Thailand, Egypt, Iran, Kuwait, and Tunisia. All studies were performed in a case-control design and the sample sizes varied from 70 to 990. The quality score of the included studies ranged from 5 to 12 (mean: 9.69) out of a maximal score of 13.

#### **Quantitative Synthesis**

The main results of this meta-analysis for the association between APOE  $\varepsilon 2/\varepsilon 3/\varepsilon 4$  polymorphisms and the risk of CAD in T2DM patients are presented in Table 2. There was significant association in three genetic models which demonstrate, ɛ4 mutation contributed to an increased risk of CAD in patients with T2DM (Figure 2). The pooled results for the three genetic models in the overall analysis were as follows: for  $\varepsilon 3/\varepsilon 4$  vs.  $\varepsilon 3/\varepsilon 3$ : OR = 1.69, 95% CI = 1.38–2.08, P < 0.001; for  $\varepsilon 4/\varepsilon 4$  vs.  $\varepsilon 3/\varepsilon 3$ : OR = 2.72, 95% CI = 1.61–4.60, P < 0.001; for  $\epsilon 4/\epsilon 4 + \epsilon 3/\epsilon 4$  vs.  $\epsilon_3/\epsilon_3$ : OR = 1.83, 95% CI = 1.52–2.22, P < 0.001. In contrast, the ε2 variation had null association with this disease (Figure 3). No significant association in the overall analysis was found in genetic model of  $\varepsilon 2/\varepsilon 2$  vs.  $\varepsilon 3/\varepsilon 3$  (OR = 1.67, 95% CI = 0.90-3.09, P = 0.104);  $\varepsilon 2/\varepsilon 3$  vs.  $\varepsilon 3/\varepsilon 3$  (OR = 1.18, 95% CI = 0.93-1.51, P = 0.175);  $\varepsilon 2/\varepsilon 4$  vs.  $\varepsilon 3/\varepsilon 3$  (OR = 1.20, 95% CI = 0.78-1.84, P = 0.405;  $\epsilon 2/\epsilon^2 + \epsilon^2/\epsilon^3$  vs.  $\epsilon^3/\epsilon^3$  (OR = 1.26, 95%) CI = 0.88-1.82, P = 0.212). In addition, the genetic models of allele-based contrasts in the overall analysis also revealed a statistically significant pooled estimates for £4 allele vs. £3 allele (OR = 1.64, 95% CI = 1.40–1.94, P < 0.001) but not for  $\epsilon^2$  allele vs.  $\epsilon^3$  allele (OR = 1.34, 95% CI = 0.98-1.84, P = 0.07).

In the subgroup analysis according to the source of patients (Chinese and non-Chinese), the pooled ORs of all genetic models except the  $\epsilon 2/\epsilon 4$  vs.  $\epsilon 3/\epsilon 3$  model were consistent with the results

TABLE 2	Meta-analysis results of the associations between APC	OF £2/£3/£4 polymorphisms and risk of coronary	artery diseases in type 2 diabetes patients

Genetic model	Pooled OR (95%CI)	Z-value	P <sub>meta-analysis</sub>	NO. of studies	Model <sup>a</sup>	P <sub>heterogeneity<sup>b</sup></sub>	l <sup>2</sup> %
ε2/ε2 vs. ε3/ε3	1.67(0.90–3.09)	1.62	0.104	9	F	0.532	0.00
Chinese	2.03(0.98-4.21)	0.02	0.984	4	F	0.841	0.00
Non-Chinese	1.01(0.31–3.32)	1.90	0.058	5	F	0.208	32.00
ε2/ε3 vs. ε3/ε3	1.18(0.93–1.51)	1.36	0.175	12	F	0.151	30.10
Chinese	1.21(0.76-1.95)	0.80	0.423	6	F	0.450	0.00
Non-Chinese	1.32(0.72-2.42)	0.89	0.374	6	R	0.053	54.30
ε2/ε4 vs. ε3/ε3	1.20(0.78-1.84)	0.83	0.405	10	F	0.493	0.00
Chinese	2.17(1.10-4.28)	2.22	0.026	5	F	0.852	0.00
Non-Chinese	0.79(0.44-1.41)	0.79	0.428	5	F	0.746	0.00
ε3/ε4 vs. ε3/ε3	1.69(1.38-2.08)	4.99	< 0.001	11	F	0.312	13.90
Chinese	2.22(1.59-3.09)	4.71	< 0.001	6	F	0.954	0.00
Non-Chinese	1.42(1.09-1.85)	2.57	0.010	5	F	0.186	35.30
ε4/ε4 vs. ε3/ε3	2.72(1.61-4.60)	3.72	< 0.001	9	F	0.807	0.00
Chinese	4.26(1.16-15.61)	2.18	0.029	5	F	0.980	0.00
Non-Chinese	2.45(1.37-4.37)	3.03	0.002	4	F	0.291	19.70
ε2/ε2+ε2/ε3 vs. ε3/ε3	1.26(0.88-1.82)	1.25	0.212	13	R	0.071	39.50
Chinese	1.08(0.71-1.65)	0.34	0.734	7	F	0.538	0.00
Non-Chinese	1.52(0.81-2.85)	1.30	0.193	6	R	0.012	66.00
ε4/ε4+ε3/ε4 vs. ε3/ε3	1.83(1.52-2.22)	6.24	< 0.001	13	F	0.384	6.20
Chinese	2.44(1.78-3.36)	5.51	< 0.001	7	F	0.890	0.00
Non-Chinese	1.55(1.22-1.97)	3.60	< 0.001	6	F	0.360	8.80
ε2 allele vs. ε3 allele	1.34(0.98-1.84)	1.81	0.070	11	R	0.054	44.70
Chinese	1.19(0.84-1.69)	0.99	0.324	6	F	0.536	0.00
Non-Chinese	1.67(0.93–3.03)	1.71	0.088	5	R	0.007	71.50
$\epsilon$ 4 allele vs. $\epsilon$ 3 allele	1.64(1.40-1.94)	5.97	< 0.001	11	F	0.284	16.80
Chinese	2.08(1.58-2.74)	5.21	< 0.001	6	F	0.987	0.00
Non-Chinese	1.44(1.17–1.77)	3.50	< 0.001	5	F	0.138	42.60

OR, odd ratio; CI, confidence interval.

<sup>a</sup>F: fixed random effect model; R: random effect model.

<sup>b</sup>P<sub>heterogeneity</sub> value for between-study heterogeneity based on Cochran's Q test.

in the overall population. In the Chinese population, the  $\varepsilon 2/\varepsilon 4$  genotype increased the risk of CAD in patients with T2DM (OR = 2.17, 95% CI = 1.10-4.28, P = 0.026).

#### **Heterogeneity Analysis**

As shown in **Table 2**, there was moderate between-study heterogeneity in the genetic model of  $\varepsilon 2$  allele vs.  $\varepsilon 3$  allele  $(P_{\text{heterogeneity}} = 0.054, I^2 = 44.70\%)$  and  $\varepsilon 2/\varepsilon 2+\varepsilon 2/\varepsilon 3$  vs.  $\varepsilon 3/\varepsilon 3$  $(P_{\text{heterogeneity}} = 0.071, I^2 = 39.50\%)$  in the overall analysis. However, no significant heterogeneity was found in other genetic models (for  $\varepsilon 2/\varepsilon 2$  vs.  $\varepsilon 3/\varepsilon 3$ :  $P_{\text{heterogeneity}} = 0.532$ ,  $I^2 = 0\%$ ; for  $\varepsilon 2/\varepsilon 3$  vs.  $\varepsilon 3/\varepsilon 3$ :  $P_{\text{heterogeneity}} = 0.151$ ,  $I^2 = 30.10\%$ ; for  $\varepsilon 2/\varepsilon 4$  vs.  $\varepsilon 3/\varepsilon 3$ :  $P_{\text{heterogeneity}} = 0.493$ ,  $I^2 = 0\%$ ; for  $\varepsilon 3/\varepsilon 4$ vs.  $\varepsilon 3/\varepsilon 3$ :  $P_{\text{heterogeneity}} = 0.312$ ,  $I^2 = 13.90\%$ ; for  $\varepsilon 4/\varepsilon 4$  vs.  $\varepsilon 3/\varepsilon 3$ :  $P_{\text{heterogeneity}} = 0.284$ ,  $I^2 = 16.80\%$ ; for  $\varepsilon 4/\varepsilon 4+\varepsilon 3/\varepsilon 4$  vs.  $\varepsilon 3/\varepsilon 3$ :  $P_{\text{heterogeneity}} = 0.384$ ,  $I^2 = 6.20\%$ ). The heterogeneity analysis results indicated that the pooled results of this metaanalysis in most genetic models were statistically steady and robust. In addition, subgroup analysis indicated that there was no heterogeneity under all nine genetic models in the Chinese population.

#### Galbraith Plot Analysis and Sensitivity Analysis

There was evidence of moderately large between-study heterogeneity in the genetic model of  $\varepsilon 2$  allele vs.  $\varepsilon 3$  allele ( $P_{\text{heterogeneity}} = 0.054$ ,  $I^2 = 44.70\%$ ) and  $\varepsilon 2/\varepsilon 2+\varepsilon 2/\varepsilon 3$  vs.  $\varepsilon 3/\varepsilon 3$  ( $P_{\text{heterogeneity}} = 0.071$ ,  $I^2 = 39.50\%$ ), so Galbraith plot analysis and sensitivity analysis were performed to detect the possible sources of heterogeneity. Under the genetic model of  $\varepsilon 2$  allele vs.  $\varepsilon 3$  allele, the Galbraith plot analysis (**Figure 4A**) showed that the Halim et al. study was the outlier, which is consistent with the results of sensitivity analysis (**Figure 4B**). No heterogeneity existed after this outlier study was omitted ( $P_{\text{heterogeneity}} = 0.460$ ,  $I^2 = 0\%$ ). Thus, the study by Halim et al. may be the source of heterogeneity in the meta-analysis for the  $\varepsilon 2$  allele vs.  $\varepsilon 3$  genetic model.

Similarly, under the genetic model of  $\epsilon 2/\epsilon 2 + \epsilon 2/\epsilon 3$  vs.  $\epsilon 3/\epsilon 3$ , the Galbraith plot analysis (**Figure 4C**) and sensitivity analysis (**Figure 4D**) indicated that Halim and Chaudhary's

Study			0/_
ID		OR (95% Cl)	Weight
			Treight
Chinese			
Hong (2017)		1.95 (1.01, 3.74)	8.44
Shi (2009)		2.27 (1.22, 4.22)	8.49
Zhang L (2008)		2.84 (1.38, 5.84)	5.82
Guo (2007)		2.64 (0.91, 7.63)	2.75
Pan (2002)		2.81 (0.82, 9.67)	1.74
Zhang WH (2000) -		1.99 (0.68, 5.82)	3.14
Zheng (1998)		<ul> <li>5.36 (1.47, 19.61</li> </ul>	)1.22
Subtotal (I-squared = 0.0%, p = 0.890)	$\diamond$	2.44 (1.78, 3.36)	31.61
1			
Non-Chinese			
Chaudhary (2012)		1.84 (1.09, 3.09)	13.46
Halim (2012)		→ 5.17 (0.94, 28.35	0.81
Al-Majed (2011)		2.65 (1.18, 5.96)	4.33
Valsi-Raygani (2010)		1.42 (0.78, 2.57)	11.92
Zar(2009)		1.25 (0.87, 1.79)	33.79
Chaaba (2008) $-$		1.04 (0.02, 4.33)	4.07
Subiolal (I-squared = $8.8\%$ , p = $0.300$ )		1.55 (1.22, 1.97)	00.39
. Overall (Lequared = $6.2\%$ p = $0.284$ )		1 92 /1 52 2 22)	100.00
Overall (I-squared = $0.2\%$ , p = $0.504$ )	Ϋ́	1.03 (1.32, 2.22)	100.00
.0353	1	28.4	
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Forest plot for the association between APOE gene polymorphist $c_{4}+c_{3}/c_{4}$ vs. $c_{3}/c_{3}$ . The center of each square represents the OE	the area of the square	is for the weight of stuc	ie ∠ alabetes lies and the h
E4+E0/E4 VS. E0/E0. THE CENTER OF EACH SQUARE REPRESENTS THE OF	, une area or une square	is for the weight of stud	iies, ai iu ti le l

study were the outliers. When the two outlier studies were omitted, no heterogeneity existed in the remaining studies  $(P_{\text{heterogeneity}} = 0.681, I^2 = 0\%)$ . Therefore, the studies of Halim et al. and Chaudhary et al. may be the main contributors to the source of heterogeneity in the meta-analysis for the  $\varepsilon 2/\varepsilon 2 + \varepsilon 2/\varepsilon 3$ vs.  $\varepsilon 3/\varepsilon 3$  genetic model.

#### **Publication Bias**

No obvious asymmetry was observed in the shape of the funnel plot for the following genetic models:  $\epsilon 2/\epsilon 2$  vs.  $\epsilon 3/\epsilon 3$  (Figure 5A);  $\epsilon 2/\epsilon 3$  vs.  $\epsilon 3/\epsilon 3$  (Figure 5B);  $\epsilon 2/\epsilon 4$  vs.  $\epsilon 3/\epsilon 3$  (Figure 5C);  $\epsilon 4/\epsilon 4$  vs.  $\epsilon 3/\epsilon 3$  (Figure 5D);  $\epsilon 2$  allele vs.  $\epsilon 3$  allele (Figure 5E);  $\epsilon 2/\epsilon 2 + \epsilon 2/\epsilon 3$  vs.  $\epsilon 3/\epsilon 3$  (Figure 5F). In addition, the Begg's test and Egger's test also did not show any evidence of publication bias ( $P_{\text{Begg}} = 0.251$  and  $P_{\text{Egger}} = 0.08$  for  $\epsilon 2/\epsilon 2$  vs.  $\epsilon 3/\epsilon 3$ ,  $P_{\text{Begg}} = 0.373$  and  $P_{\text{Egger}} = 0.320$  for  $\epsilon 2/\epsilon 3$  vs.  $\epsilon 3/\epsilon 3$ ,  $P_{\text{Begg}} = 0.283$  and  $P_{\text{Egger}} = 0.403$  for  $\epsilon 2/\epsilon 4$  vs.  $\epsilon 3/\epsilon 3$ ,  $P_{\text{Begg}} = 0.466$  and  $P_{\text{Egger}} = 0.988$  for  $\epsilon 4/\epsilon 4$  vs.  $\epsilon 3/\epsilon 3$ ,  $P_{\text{Begg}} = 0.300$  and  $P_{\text{Egger}} = 0.331$  for  $\epsilon 2/\epsilon 2 + \epsilon 2/\epsilon 3$  vs.  $\epsilon 3/\epsilon 3$ ).

The results from the following three genetic models  $\varepsilon_3/\varepsilon_4$  vs.  $\varepsilon_3/\varepsilon_3$ ;  $\varepsilon_3/\varepsilon_4 + \varepsilon_4/\varepsilon_4$  vs.  $\varepsilon_3/\varepsilon_3$  and  $\varepsilon_4$  allele vs.  $\varepsilon_3$  allele performed by Begg's test ( $P_{\text{Begg}} = 0.213$ ,  $P_{\text{Begg}} = 0.033$ , and  $P_{\text{Begg}} = 0.043$ , respectively) or Egger's test ( $P_{\text{Egger}} = 0.013$ ;  $P_{\text{Egger}} = 0.001$  and  $P_{\text{Egger}} = 0.001$ , respectively) revealed publication bias.

Nevertheless, by using the trim and fill method, the recalculated estimates (OR = 1.50, 95%CI = 1.24–1.82; OR = 1.59, 95%CI = 1.34–1.89 and OR = 1.40, 95%CI = 1.22–1.62, respectively) remained statistically significant, which indicated that our meta-analysis results were steady and not influenced by publication bias. **Figure 6** shows the funnel plot of trim and fill method in the genetic model of  $\varepsilon 3/\varepsilon 4$  vs.  $\varepsilon 3/\varepsilon 3$  (**Figure 6A**),  $\varepsilon 3/\varepsilon 4 + \varepsilon 4/\varepsilon 4$  vs.  $\varepsilon 3/\varepsilon 3$  (**Figure 6B**),  $\varepsilon 4$  allele vs.  $\varepsilon 3$  allele (**Figure 6C**).

# DISCUSSION

T2DM is a well-established risk factor for the development of CAD. The management of CAD in patients with T2DM poses great challenges to the medical profession (Wei et al., 2015). The identification of susceptibility genes would be very helpful for the management of CAD in patients with T2DM. The link between *APOE*  $\varepsilon 2/\varepsilon 3/\varepsilon 4$  polymorphisms and CAD in diabetic patients has been highlighted in our study. This meta-analysis provides evidence for the significant associations between *APOE*  $\varepsilon 3/\varepsilon 3$ ;  $\varepsilon 4$  allele vs.  $\varepsilon 3/\varepsilon 3$ ;  $\varepsilon 4/\varepsilon 4 + \varepsilon 3/\varepsilon 4$  vs.  $\varepsilon 3/\varepsilon 3$ ;  $\varepsilon 4$  allele vs.  $\varepsilon 3$  allele) and an elevated risk of CAD in patients with T2DM. In contrast, no significant association was found in genetic model of *APOE*  $\varepsilon 2$  variation ( $\varepsilon 2/\varepsilon 2$  vs.  $\varepsilon 3/\varepsilon 3$ ;  $\varepsilon 2/\varepsilon 3$  vs.  $\varepsilon 3/\varepsilon 3$ ;  $\varepsilon 2/\varepsilon 2 + \varepsilon 2/\varepsilon 3$  vs.  $\varepsilon 3/\varepsilon 3$ ;  $\varepsilon 2$  allele vs.  $\varepsilon 3$ 



allele). However, CAD in patients with T2DM is believed to be multifactorial and involved in many susceptibility genes with small individual effects. Therefore, the integration of information derived from several polymorphisms in multiple susceptibility genes may become clinically useful.

It has been reported that lipoprotein-related mechanisms are associated with the impairment of the cardiovascular system among patients with diabetes (Jenkins et al., 2004). For example, serum low-density lipoprotein cholesterol (LDL-C) level was identified as an independent risk factor for CAD in T2DM patients (Jayashankar et al., 2016). APOE is initially recognized for its important role in plasma lipid metabolism and thus affects the serum lipid profiles in the body. The three APOE alleles ( $\varepsilon_2$ ,  $\varepsilon$ 3,  $\varepsilon$ 4) differ from each other by only one or two amino acids at positions 112 and 158, but these slight differences alter the structure and function of APOE. In general, the APOE-ε4 allele is associated with higher and the APOE-E2 allele with lower total plasma cholesterol and LDL-C concentrations compared with the APOE-ε3 allele (Bennet et al., 2007; Larifla et al., 2017). Therefore, abnormalities of lipoprotein metabolism may explain, at least in part, the associations between APOE  $\varepsilon 2/\varepsilon 3/\varepsilon 4$  polymorphisms and the risk of CAD in patients with T2DM.

Several meta-analysis studies have been conducted to assess the association between *APOE*  $\epsilon 2/\epsilon 3/\epsilon 4$  polymorphisms and risk of CAD in the general population. In 2004, Song et al

firstly found that carriers of the APOE-E4 allele had a 42% increased risk for CAD (OR = 1.42, 95% CI = 1.26-1.61) compared with the  $\varepsilon 3/\varepsilon 3$  genotypes (Song et al., 2004). Xu et al. found similar results which showed that the ɛ4 allele had a 46% higher risk of CAD (OR = 1.46, 95% CI = 1.28-1.66) (Xu et al., 2016). Similar findings were also observed in other meta-analysis (Yin et al., 2013; Xu et al., 2014, 2016; Zhang et al., 2014a, 2015; Wang et al., 2015b). Interestingly, the role of APOE-E2 allele in the risk of CAD may be dependent on the patient ethnicity (Xu et al., 2016). In addition, the association between APOE  $\varepsilon 2/\varepsilon 3/\varepsilon 4$  polymorphisms and the risk of T2DM in the general population was also well explored in previous meta-analysis (Anthopoulos et al., 2010; Yin et al., 2014). The results indicated that both APOE  $\varepsilon$ 2 and  $\varepsilon$ 4 alleles were associated with an increased risk of T2DM in the general population. In 2015, Wu et al. performed a meta-analysis on the association between APOE £2/£3/£4 polymorphisms and T2DM patients with CAD among Chinese Han population. They found that APOE-E4 allele resulted in an increased risk of T2DM patients with CAD in China (Wu et al., 2015). However, only five individual studies were included in their meta-analysis. To our knowledge, our meta-analysis represents the largest study to investigate the association between APOE ε2/ε3/ε4 polymorphisms and risk of CAD in the T2DM patients.



horizontal lines represent the 95% CIs, given named study is omitted.

Heterogeneity across studies is common in meta-analysis of genetic association study (Munafo and Flint, 2004). Heterogeneity should be taken into consideration in the interpretation of the meta-analysis results. However, one of the strengths in this meta-analysis was the lack of significant heterogeneity in all genetic models except the genetic model of  $\varepsilon 2$ allele vs. ɛ3 allele. Between-study heterogeneity can be attributed to the potential differences such as the definition of disease, ethnicity, genotyping methods and sample size in the included studies. To explore the potential sources of heterogeneity under the genetic model of ɛ2 allele vs. ɛ3 allele, Galbraith plot analysis and sensitivity analysis were employed to detect whether there were outliers that could be the potential sources of heterogeneity between studies. The study conducted by Halim et al was considered as the main contributors to between-study heterogeneity. The heterogeneity was effectively decreased after omitting the study. The frequency of APOE-E3 allele was nearly 95% in Halim's study, whereas lower than 90% in other studies (Zhang et al., 2008; Izar et al., 2009; Chaudhary et al., 2012; Hong et al., 2017). Consequently, the heterogeneity can be due to the distinct frequency of APOE  $\varepsilon 2/\varepsilon 3/\varepsilon 4$  polymorphisms among the included studies. Although Halim's study caused the substantial heterogeneity in the genetic model of  $\epsilon 2$  allele vs.  $\epsilon 3$  allele, the pooled effect was still insignificant after removing it.

There are several limitations in this meta-analysis that should be noted. First, the included studies were limited to only English or Chinese languages in our research and some eligible studies may be published in other languages, which would cause bias of the results. Second, all the included studies in this meta-analysis were the type of retrospective case-control studies, which may result in some selection bias. Third, publication bias existed in the following three genetic models:  $\varepsilon_3/\varepsilon_4$  vs.  $\varepsilon_3/\varepsilon_3$ ;  $\varepsilon_3/\varepsilon_4+\varepsilon_4/\varepsilon_4$ vs.  $\varepsilon 3/\varepsilon 3$ ;  $\varepsilon 4$  allele vs.  $\varepsilon 3$  allele. However, by using the trim and fill method, the recalculated ORs and their 95% CIs did not change, which indicated the stability and robustness of meta-analysis results. Last but not the least, T2DM complicated with CAD is a multifactorial disease caused by both genetic and environmental factors. The APOE-environment interactions should be considered. For example, the study by Talmud et al. has found that the impact of the APOE-E4 on the risk of CAD appeared to be restricted to smokers (Talmud et al., 2004).



**FIGURE 5** | Begg's funnel plot for the association between *APOE* gene polymorphism and the risk of coronary artery diseases in type 2 diabetes patients under the genetic model of ε2/ε2 vs. ε3/ε3 (**A**), ε2/ε3 vs. ε3/ε3 (**B**), ε2/ε4 vs. ε3/ε3 (**C**), ε4/ε4 vs. ε3/ε3 (**D**), ε2 allele vs. ε3 allele (**E**), and ε2/ε2+ε2/ε3 vs. ε3/ε3 (**F**). Size of the open circles is proportional to the weight of studies.



In conclusion, we observed a significant association between the *APOE* gene  $\varepsilon$ 4 mutation and an increased risk of CAD in patients with T2DM, while the  $\varepsilon$ 2 variation had null association with this disease. Taking into account the above limitations, more studies with larger sample size and incorporated with geneenvironment interactions are needed to definitively determine the association between the *APOE* gene  $\varepsilon$ 2/ $\varepsilon$ 3/ $\varepsilon$ 4 polymorphisms and the risk of CAD in patients with T2DM.

# **AUTHOR CONTRIBUTIONS**

Conceived and designed the study: J-QL and HR. Performed the search: J-QL, HR, M-ZL. Analyzed the data: J-QL and HR. Contributed reagents/material/analysis tools: J-QL, HR, M-ZL,

PX, P-FF, and D-XX. Wrote and review the manuscript: J-QL, HR and HB. Reference collection, data management, statistical analyses, paper writing, and study design: J-QL and HR.

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# SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fphys. 2017.01031/full#supplementary-material

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**Conflict of Interest Statement:** The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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