

Reconciling heterogeneous dengue virus infection risk estimates from different study designs

Authors and affiliations

Angkana T. Huang^{1,2,3}, Darunee Buddhari², Surachai Kaewhiran⁴, Sophon Iamsirithaworn⁴, Direk Khampaen⁴, Aaron Farmer², Stefan Fernandez², Stephen J. Thomas⁵, Isabel Rodriguez Barraquer⁶, Taweewun Hunsawong², Anon Srikiatkachorn^{2,7}, Gabriel Ribeiro dos Santos¹, Megan O'Driscoll¹, Marco Hamins-Puertolas⁶, Timothy Endy⁸, Alan L. Rothman⁷, Derek A. T. Cummings³, Kathryn Anderson⁵, Henrik Salje¹

1 University of Cambridge, Cambridge, UK

2 Armed Forces Research Institute of Medical Sciences, Bangkok, Thailand

3 University of Florida, Florida, USA

4 Ministry of Public Health, Nonthaburi, Thailand

5 State University of New York Upstate Medical University, USA,

6 University of California, San Francisco, USA

7 University of Rhode Island, USA

8 Coalition for Epidemic Preparedness Innovations, USA

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Abstract

Uncovering rates at which susceptible individuals become infected with a pathogen, i.e. the force of infection (FOI), is essential for assessing transmission risk and reconstructing distribution of immunity in a population. For dengue, reconstructing exposure and susceptibility statuses from the measured FOI is of particular significance as prior exposure is a strong risk factor for severe disease. FOI can be measured via many study designs. Longitudinal serology are considered gold standard measurements, as they directly track the transition of seronegative individuals to seropositive due to incident infections (seroincidence). Cross-sectional serology can provide estimates of FOI by contrasting seroprevalence across ages. Age of reported cases can also be used to infer FOI. Agreement of these measurements, however, have not been assessed. Using 26 years of data from cohort studies and hospital-attended cases from Kamphaeng Phet province, Thailand, we found FOI estimates from the three sources to be highly inconsistent. Annual FOI estimates from seroincidence was 2.46 to 4.33-times higher than case-derived FOI. Correlation between seroprevalence-derived and case-derived FOI was moderate (correlation coefficient=0.46) and no systematic bias. Through extensive simulations and theoretical analysis, we show that incongruences between methods can result from failing to account for dengue antibody kinetics, assay noise, and heterogeneity in FOI across ages. Extending standard inference models to include these processes reconciled the FOI and susceptibility estimates. Our results highlight the importance of comparing inferences across multiple data types to uncover additional insights not attainable through a single data type/analysis.

Significance statement

Dengue virus infections are surging globally. Knowing who, where, and how many people are at risk of infection is crucial in determining means to protect them. Here, we compare three current approaches in measuring risk (two involving blood samples and one involving case counts) to estimate the risk of infection. Estimates derived from each method differed greatly. By accounting for rise and falls of antibodies following infections, noise in the antibody titer measurements, and heterogeneity in infection risk across ages, we reconciled the measurements. As measurements from blood samples and case counts are pillars in uncovering risk of most infectious diseases, our results signifies integrating these processes into risk measurements of pathogens beyond dengue virus.

1 Introduction

2 Quantifying historical infection intensity of pathogens is essential to assess infection burden and
3 susceptibility of populations through time, insights that are pivotal in predicting future
4 transmission potentials and shaping effective intervention strategies (1–3). Infection intensity is
5 often quantified as the rate at which susceptible individuals become infected, a concept known
6 as the Force of Infection (FOI). For dengue virus (DENV) infections, quantifying infection risk
7 through FOI is of particular significance, as infection burden is non-linearly linked to the
8 observable disease burden. First infection by one of the four DENV serotypes is primarily
9 subclinical but the generated immune response is the most widely recognized risk factor for
10 severe disease following a second infection by a different serotype (4). The FOI can also be
11 used to estimate how immunity is distributed in the population (by age, for example) to identify
12 who is at risk of infections having already acquired immunity (5–7). Information on infection risk
13 in populations and the distribution of immunity are integral to optimizing the impact of the two
14 currently licensed vaccines and avoiding deleterious outcomes (8–14).

15 Typically, two main sources of data are employed to estimate historical infection intensity, or
16 FOI, in populations: serological data and case count data. In parallel, two different study designs
17 have been used to estimate forces of infection: longitudinal and cross-sectional. Longitudinal
18 serological studies are often considered the gold standard, as they directly track the transition of
19 seronegative individuals to seropositive (seroincidence) (7, 15). Cross-sectional serological
20 data, which includes individuals of different ages, can provide estimates of FOI by drawing upon
21 differences in exposure histories across birth cohorts (16–20). Similarly, age-stratified case
22 count data can extract information from age distribution of cases over time which reflects the
23 variation in exposure histories among different age groups (21–24). To infer FOI from age-
24 stratified case counts, models are employed to link the infection process with the generation of
25 reported cases. The model typically accounts for reporting rates, but can also include processes
26 that influence illness manifestations (23).

27 These approaches rely on different assumptions about antibody responses following infection,
28 age-specific differences in infection risk, the role of cross-reactivity from infection or vaccination
29 from related viruses, accuracy of the serological assay, and how immunity preceding infections
30 affects the risk of symptoms. However, the importance of these different assumptions on the
31 resulting FOI estimates is largely unknown. Further, little is known about the consistency in
32 estimates derived from the different approaches. In this study, we leverage 26 years of data
33 from a single location to compare FOI estimates obtained from various data types. In this single
34 location both serological and clinical case data is available from longitudinal cohorts and from a
35 passive surveillance system. We compare estimates derived from different subsets of the
36 available data, identify the sources of discrepancies and develop methods to improve estimates
37 through joint inference when multiple data types are available.

38 Results

39 Dengue data in Kamphaeng Phet

40 Kamphaeng Phet province, Thailand (KPP) represents a dengue hyper-endemic region with
41 four consecutive longitudinal cohort studies conducted from 1998 to the present: Kamphaeng
42 Phet Prospective Study 1 (KPS1, 1998-2002), KPS2 (2004-2007), KPS3 (2010), and
43 Kamphaeng Phet Family Cohort Study (KFCS, 2015-ongoing) (25–28), **Figure 1** and **Table S1**.
44 KPS1 and KPS2 were school children cohorts while KPS3 was a one-year cohort of children in
45 the community. KFCS is a community cohort focused on multi-generational households (28).
46 Individuals were bled every 3, 6, 6, and 12 months in these cohorts, respectively, and tested for
47 anti-DENV antibodies via hemagglutination inhibition assay (HAI). Percentages of seropositive
48 samples (geometric mean titer (GMT) ≥ 10) increased with age except for samples obtained at
49 very young ages, attributable to the presence of maternally-derived antibodies and cross-
50 reactive antibodies from Japanese Encephalitis vaccination (29, 30), **Figure 1b**. Among
51 participants aged nine, 75%, 57%, 53%, and 49% have GMT ≥ 10 , respectively. All individuals
52 in KFCS have GMT ≥ 10 after age 30 (97% with GMT ≥ 20).

53 Within Mueng, the capital district, the Kamphaeng Phet Provincial Hospital (KPPH) serves as
54 the sole tertiary care facility in the province. Between 1994 and 2020, KPPH reported a total of
55 17,773 cases suspected of dengue among KPP residents (of which 12,819 were lab confirmed),
56 representing an annual incidence of 0.5 to 3.3 cases per thousand population (**Figure 1c**).
57 Mueng residents accounted for 55% of these cases.

58 Inferred force of infection (FOI) differs across data sources

59 Considering the cohorts as both longitudinal measures (multiple samples per individual) and
60 cross-sectional data (single sample per individual), we estimated the annual per-serotype FOI
61 between 1998 and 2019 using standard models for each data type (16–20), **Figure S1**. Bleeds
62 taken before age three were excluded to avoid interference from maternally-derived antibodies
63 and/or cross-reactive antibodies from Japanese Encephalitis vaccination. We derived case-
64 based FOI by fitting a model which takes into account differences in symptomatic rates across
65 the four possible infections of individuals (one by each serotype) and variations in time and age
66 for DENV-infected individuals to seek care at KPPH (23). We excluded cases under age one as
67 their symptomatic rate upon first DENV infection differs from the others due to maternally-
68 derived immune-enhancement (31). All models assumed that infection risks in the excluded
69 ages remained similar to the rest of the population despite differences in test positive
70 tendencies or clinical presentations.

71 Applying a standard geometric mean titer (GMT) threshold of 10 to define seropositivity of
72 serological samples, we found that the seroincidence-derived annual FOIs were consistently
73 higher than the case-derived annual FOIs (3.40-fold on average, 95%CI: 2.46 to 4.33, **Figure**
74 **2a**) and seroprevalence-derived annual FOIs (95%CI: 1.70, 3.96-folds). Ratios between cross-
75 sectional seroprevalence-derived annual FOIs and case-derived annual FOIs did not appear to
76 vary systematically (95%CI: 0.72, 1.23-folds), with moderate correlation between the two
77 (correlation coefficient=0.46). The estimates derived from both serological sources exhibited

78 wide uncertainty. Raising the GMT threshold to 20 to mitigate false positives from e.g.,
79 individuals seropositive from JEV vaccination, did not lower FOI estimates from seroincidence
80 (95%CI: 1.95, 3.55-folds). However, it led to systematically lower seroprevalence-derived FOI
81 relative to case-derived FOI (0.43 to 0.81-fold, **Figure S2a**). The discrepancy patterns remained
82 similar when compared to case-derived FOIs inferred from lab-confirmed cases (**Figure S3**).

83 These discrepancies and uncertainties in FOI values resulted in notable differences in the
84 reconstructed susceptibility fractions across the different approaches (see **Figure 2b**). For
85 instance, in the most recent year of the study (2019), case-derived reconstructions suggested
86 56% of 9yrs old remained DENV-naive (95%CI: 49%, 64%) while seroprevalence-derived and
87 seroincidence-derived reconstructions suggested 40% (95%CI: 30%, 51%) and 24% (95%CI:
88 12%, 37%) of 9yrs old remained DENV-naive, respectively.

89 **Simulations to study effects of violated model assumptions on inferred FOI**

90 To identify sources of the FOI discordance, we performed an extensive suite of simulations in
91 which data generation and true infection rates were known. We analyzed simulated data using
92 our different approaches described above. Our simulations incorporated varying assumptions of
93 the effects of waning antibody titers, measurement error in assay readouts, and titers against
94 cross-reactive pathogens. Prior research has demonstrated that following primary DENV
95 infections, antibody titers rise rapidly but then wane exponentially to a steady titer approximately
96 5-times lower within a year (7). After a subsequent infection by a different DENV serotype, titers
97 increase to levels that are robust to detection. Measurement error in assay readouts can lead to
98 titers falling below seropositivity thresholds, while individuals without prior exposure to DENV
99 may exhibit seropositivity due to titers against other flaviviruses (29, 32), **Figure 3a**.
100 Additionally, variations in infection risk across different age groups are possible (23). We
101 simulated infection timings of 500,000 individuals with defined FOI by year to eliminate
102 imprecisions in estimates resulting from insufficient statistical power to study effects of these
103 processes on the inferred FOI.

104 We found that in the absence of random measurement error and when the seropositivity
105 threshold is low enough to correctly discriminate DENV-exposed individuals from naives,
106 waning monotypic titers do not lead to biased FOI estimates from either serological data types
107 (**Figure S4a-b** and **Figure S5a**). However, when the simulation included random measurement
108 error, using a low threshold led to false positives which inflated seroincidence-derived FOI
109 (**Figure S4c**). The inflation was exacerbated by the presence of cross-reactive titers (**Figure 3d**,
110 **Figure S4d**). While raising the threshold to define seropositivity helped mitigate inflation if titers
111 of exposed individuals remained high (**Figure S5d**), the trade-off for false negatives when
112 monotypic titers did wane led to even more pronounced over-estimates of FOIs (**Figure S5e-f**).
113 The over-estimation arose from the greater chance of testing positive in the follow-up bleed in
114 DENV-exposed individuals that tested falsely negative at pre-interval compared to truly DENV-
115 naive individuals. In fact, the over-estimation can be severe even at lower thresholds where
116 fewer false negative individuals were expected (**Figure S4e-f**, see also **Supplementary**
117 **Mathematical Analysis**).

118 In the absence of waning titers, seroprevalence-derived FOI appeared robust to assay noise
119 and cross-reactive titers at both seropositivity thresholds (**Figure S4c-d, Figure S5c-d**).
120 Expectedly, false negativity due to waning monotypic titers led to underestimation of FOI, which
121 became more extreme with higher seropositivity thresholds (**Figure S4e-f, Figure S5e-f**).

122 We found that age-specific differences in risk of infection could also cause discrepancies in FOI
123 estimates. Even when the susceptibility status of individuals could be perfectly ascertained,
124 seroincidence-derived FOI estimates were systematically different from seroprevalence-derived
125 FOI. (**Figure 4e, Figure S4g, Figure S5g**). When age-specific risk was present in conjunction
126 with waning titers and assay noise, the difference was exacerbated (**Figure 3f**).

127 **Correcting for violated assumptions recovers temporal FOI of simulation ground** 128 **truth at varying efficiencies**

129 In highly powered datasets, when only assay noise was present, we found that inflation in
130 estimated seroincidence could be perfectly mitigated at both seropositivity thresholds (10 and
131 20, **Figure S6a, Figure S7a**) by correcting for (presumed known) test positive probabilities
132 given the susceptibility status of individuals and the distribution in infection risk by age. In the
133 presence of both assay noise and waning monotypic titers, this simple correction, which did not
134 take into account exact infection timings of individuals and the variable amounts of titers waned
135 across individuals, reduced but did not eliminate the inflation in estimated seroincidence-derived
136 FOI (**Figure S6b, Figure S7b**). Importantly, the new estimates at a threshold of 20 showed
137 greater discrepancies from the ground truth than at a threshold of 10 (1.35 to 1.63-fold
138 difference as compared to the ground truth vs 0.81 to 0.98-fold difference). The correction
139 efficiencies remained similar when additional corrections for cross-reactive titers in DENV-
140 naives and non-uniform risk in age were needed (**Figure S6c-d, Figure S7c-d**). In contrast, the
141 simple correction was able to efficiently correct for biases in seroprevalence-derived FOI in all
142 cases (**Figure S6a-e, Figure S7a-e**).

143 When we reduced the size of the simulated datasets to match the number and time points of
144 samples available in our cohort studies, we found that correlation with the ground truth for both
145 serological data types was lower despite the same adjustments to correct for assay noise,
146 waning monotypic titers and age-specific differences in risk (**Figure S6g vs Figure S6d**).
147 Uncertainties in the estimates increased substantially (**Figure S6f-g**).

148 **Reconciled infection risk in KPP is non-uniform across ages**

149 Taking into account sources of discrepancies in estimation of FOI learnt from the simulation
150 studies, namely, age-specific infection risk, antibody kinetics, and assay variability, we
151 developed a model that is jointly informed by both serological data types to estimate the shared
152 underlying infection risk in KPP and a separate case-based model with age-specific risk
153 extension.

154 Infection risk estimates from the 'joint serology model' and the extended case-based model
155 showed good agreement, **Figure 4a-b**. Both models suggested elevated infection risk in KPP
156 between ages 6-17yrs compared to the reference class (age 0-5yrs). Correlation between the

157 temporal FOIs was moderate (cor. coef.=0.48) but without systematic differences in magnitudes
158 (ratio between serology-based to case-based FOIs of 0.72 to 1.34). Reconstructions of
159 susceptibility by age and year from the two models were highly congruent (cor. coef. \geq 0.97
160 without signs of systematic differences, **Figure 4c**).

161 **Test positive probabilities are key to the reconciliation**

162 Using multiple serological data sources with shared underlying processes, we were able to
163 characterize factors governing probabilities of falsely testing positive at various thresholds when
164 DENV-naive, and testing positive when DENV-exposed, in tandem with the infection risks. The
165 factors are namely the post-infection antibody titers captured in the serological samples and
166 variability in the assay measurements, **Figure S10**. We estimated that the captured titer rises
167 between bleeding intervals of the cohort study participants in response to primary DENV
168 infections that occurred during the intervals were comparable across studies: an average rise of
169 7.89 log₂ (95%CI: 4.83, 13.10). The titers then declined to a steady level of 2.76 log₂ (95%CI:
170 2.53, 2.96). Standard deviation of assay readouts was estimated to be 0.51 (95%CI: 0.36, 0.64)
171 which corresponds to 99.9% to 100% of monotypic titers above a threshold of 10 over the long-
172 term (i.e., cross-sectional seroprevalence studies) or 92.7% to 93.7% at a threshold of 20.
173 Averaged across the studies, DENV-naive individuals have a 6.2% to 7.3% chance of testing
174 positive for DENV at threshold of 10 and <0.2% at threshold of 20.

175 We re-inferred temporal FOIs from each of the serological data sources presuming various
176 other sets of test positive probabilities and found that congruence with case-derived FOIs were
177 reduced (**Figure 4d**, **Figure S11**). Importantly, FOIs from seroincidence varied greatly across
178 test positive probabilities leading to pronounced changes in congruence compared to
179 seroprevalence-derived FOIs.

180 **Discussion**

181 Leveraging a unique opportunity where over two decades of longitudinal serological data and
182 hospital case count data are available from the same community, we assessed the congruence
183 in FOI estimated from different data types. We found large discrepancies between the FOI
184 estimates. Consequently, susceptibility in the population inferred from the estimates were
185 drastically different. Our investigations revealed causes of these discrepancies as a lack of
186 accounting for antibody kinetics and assay noise (affecting serological data types), and age-
187 specific infection risk (affecting all data types).

188 Longitudinal serology is crucial to track infections of individuals and ascertain their evolving
189 exposure statuses (5, 7). However, we found identifying DENV infections based on
190 seroconversion from negative to positive at a specified threshold was highly sensitive to the
191 interplay between antibody kinetics and assay noise. Importantly, the heterogeneous
192 seroconversion tendencies across sera pairs could not be efficiently corrected for by applying
193 average test positive probabilities across sera pairs. Our findings support the use of individual-
194 based titer reconstructions, mechanistically taking into account sources of bias, to detect
195 infections (7, 33, 34).

196 Cross-sectional serology is the most common data source used to infer dengue burden (17, 19,
197 35–38). While this approach bypasses biases in case reporting, our findings highlight that the
198 processes that affect a test coming back positive or not should be considered (39). Applying a
199 single seroreversion rate to all exposed individuals may be sufficient to account for waning
200 antibodies in non-endemic settings. However, given the increased durability of antibodies in
201 multitypically exposed individuals (40), settings with multiple serotypes co-circulating will need
202 to account for heterogeneity in antibody kinetics across individuals with different infection
203 histories. The likely exposure to cross-reacting viruses (e.g., JEV or ZIKV) due to shared vector
204 ecology or vaccination will also require appropriate corrections. Our analyses assumed that
205 cross-reactive titers were evenly present in DENV-naives aged ≥ 3 yrs. This may be true for
206 Japanese Encephalitis vaccine-induced titers as children were vaccinated widely at young ages
207 (≤ 2 yr (41)) but would only be true for co-circulating viruses such as ZIKV (29, 42) if FOI of
208 these viruses were extremely high such that all children were exposed before age 3. Further
209 explorations are needed to assess how heterogeneous presence of cross-reactive titers would
210 impact estimation of FOI of DENV, especially if these exposures alter antibody kinetics to
211 subsequent DENV infections (43).

212 Age-stratified case data provides an alternative means to estimate past dengue burdens (21–
213 23). As the method involves adjusting for multiple processes to uncover the true age distribution
214 of infections (which is informative of the infection burden) from the observed age distribution,
215 long periods of surveillance are necessary to reliably estimate FOI. With 26 years of data fine-
216 scale age strata, our FOI estimates from case data tracked closely with estimates from our joint
217 model (our best approximation of the underlying dengue burden). In settings with less data, the
218 ability to disentangle observation processes from infection risk would be more limited making
219 simplifying assumptions necessary. Analytical studies and simulations are needed to assess the
220 impact of those assumptions on the inferred FOI.

221 Our inferences revealed evidence of age-specific infection risk, a characteristic often neglected
222 in dengue epidemiology. Whether the variation reflects behavioral, immunological, and/or
223 physiological differences (44–46), the heterogeneity challenges generalizations of risk
224 measured in a sample to the general population. Uncovering processes leading to these
225 differences is key to overcoming this challenge and will facilitate the development and
226 management of interventions.

227 The importance of characterizing antibody kinetics, assay variability, and distribution of risk in
228 the population demonstrated in our study applies broadly beyond dengue as serological and
229 case data are pillars in quantifying infection burdens in most diseases. Our results highlight the
230 need to compare inferences across multiple data types and analysis methods both to flag
231 blindspots in each of the inferences to mitigate misconclusions and to uncover additional
232 insights not attainable through a single data type/analysis.

233 Material and methods

234 Ethics statement

235 Use of data from the Kamphaeng Phet Hospital were reviewed and approved by Walter Reed
236 Army Institute of Research Institutional Review Board (protocol number 1313 and 1957). The
237 study protocol for KPS1 was approved by the Office of the Army Surgeon General, University of
238 the Massachusetts Medical School, and the Ministry of Public Health, Thailand. The protocol for
239 KPS2 was additionally approved by the University of California–Davis and San Diego State
240 University (protocol number 654 and 1042). The KPS 3 and KFCS cohort study was approved
241 by the Thailand Ministry of Public Health Ethical Research Committee; Siriraj Ethics Committee
242 on Research Involving Human Subjects; Institutional Review Board for the Protection of Human
243 Subjects, State University of New York Upstate Medical University; and Walter Reed Army
244 Institute of Research Institutional Review Board (protocol number 1552 and 2119).

245 Empirical data for serological models

246 Longitudinal samples from four longitudinal cohort studies in Kamphaeng Phet province were
247 included in this study: KPS1 (1998-2002), KPS2 (2004-2008), KPS3 (2010), and KFCS (2016-
248 2019) (27). KPS1, 2 and 3 were cohorts of school children while KFCS focused on multi-
249 generational households. To generate cross-sectional seroprevalence data from the longitudinal
250 samples, one sample was randomly selected per individual. In the family cohort study, KFCS,
251 only one randomly selected individual was selected per family to avoid reported
252 interdependence between family members (27). Inferences involving seroprevalence data were
253 repeated using three independent random samples.

254 Antibody titers were measured using hemagglutination inhibition assay (HAI) against DENV1
255 (Hawaii strain), DENV2 (New Guinea C strain), DENV3 (H87 strain), and DENV4 (814669 strain
256 in KPS1-3 and H241 strain in KFCS) as described elsewhere (46, 47). Measurements were
257 done in 2-fold serial dilutions between 1:10 and 1:10240. Titers <10 and >10240 were imputed
258 as 5 and 20480, respectively. For each sample, geometric mean titers (GMT) were computed
259 from the four serotype-specific titers. The linear scale GMT, A_{linear} , were log transformed via
260 equation $\log_2(A_{\text{linear}} / 10) + 1$ so that samples with linear scale titers of <10 to all four
261 serotypes were zero.

262 Empirical data for case-based models

263 Age-annotated cases suspected for dengue that sought care at the Kamphaeng Phet Provincial
264 Hospital (KPPH) between 1994 and 2020 were included in this study. Cases were considered
265 lab-confirmed when acute samples from the patients were tested positive for DENV via
266 polymerase chain reaction (PCR) or virus isolation, or enzyme-linked immunosorbent assays
267 (ELISAs) as per criteria described elsewhere (48–50). Inferences were restricted to Mueng
268 residents to match the spatial coverage of the cohort studies. Population age censuses for
269 1994-2020 were acquired from the Department of Provincial Administration, Ministry of the
270 Interior through the Official Statistics Registration Systems (51).

271 Simulated data

272 We simulated one million observations (two bleeds of three months apart per individual for
273 500,000 individuals) to eliminate imprecisions of estimates from insufficient power. Observations
274 were made between ages 5-15 years where the occurrence of first infections was expected to
275 be concentrated. First, we simulated ages at which the individuals acquired their first, second,
276 third, and fourth DENV infections under defined annual (and age-specific) force of infections.
277 We then used anti-DENV antibody kinetics reported in Salje, 2018 (6) to generate true
278 underlying titers of individuals at the observation time points. For scenarios where assays were
279 imperfect, the observed titers were drawn from normal distributions with means equal to the true
280 underlying titers and standard deviation $\sigma = 0.49$. In scenarios where non-zero titers in DENV-
281 naives due to the presence of cross-reactive titers were considered, titers in DENV-naive
282 individuals were drawn from a normal distribution with mean 0.266 (20% of the long-term titer in
283 monotypic sera) and standard deviation $\sigma_0 = \sigma = 0.49$.

284 To simulate data with power equivalent to empirical serological data, we performed the same
285 procedures but with the number of individuals, observation time points and ages matching those
286 in the cohort studies.

287 Standard models in FOI inferences

288 Inferring force of infection (FOI, λ) of dengue from non-serotype-specific datasets typically
289 assumes equivalent FOI across all serotypes in circulation, long-lived protection against the
290 infecting serotype, and no cross protection against other serotypes, **Figure S1**. Hence,
291 probability that an individual birth cohort h has escaped a particular serotype up to age a is
292 $p_{esc} = \exp(-\sum_{t=h}^{h+a} \bar{\lambda}(t))$ where $\bar{\lambda}(t)$ is the average per-serotype force of infection at time t . It
293 follows that the probability that the individual has acquired i infections is

$$294 \binom{4}{i} p_{esc}^{4-i} (1 - p_{esc})^i$$

295 Assuming that antibodies in infected individuals are robustly above a chosen positivity
296 threshold, FOI can be linked to **cross-sectional serology** via probability of testing positive
297 written as $1 - p_{esc}^4$. In **longitudinal serological studies**, the probability that a seronaive
298 individual at time t tests positive at $t + \Delta t$ is similarly $1 - \exp(-4 \sum_t^{t+\Delta t} \bar{\lambda}(t))$.

299 Increase in the proportion of individuals in birth cohort h that have acquired at least i infections
300 between time t and $t + \Delta t$ indicates occurrence of the i -th infection during the time interval,
301 $I_{\Delta t}(i, h, t)$. Let proportion $Q(i)$ of the i -th infections of individuals result in clinical presentations
302 that are severe enough to trigger care seeking and proportion $\phi(a, t)$ of those severe infections
303 go on to report to KPPH, a being the age of cohort h at time t . We express $\phi(a, t)$ as $\phi(a) \cdot \phi(t)$
304 where $\phi(a)$ is the piecewise constant age-specific reporting rate and $\phi(t)$ is the piecewise
305 constant time-specific reporting rate. Considering $Pop(h, t)$ the population size of birth cohort h
306 at time t , we would expect the number of dengue cases from this birth cohort in this time
307 interval who reported to KPPH to be

$$308 \sum_{i \in \{1, 2, 3, 4\}} I_{\Delta t}(i, h, t) \cdot Q(i) \cdot \phi(a) \cdot \phi(t) \cdot Pop(h, t)$$

309 For **case counts aggregated by age**, the expected counts is the sum of expected cases of
310 those birth cohorts contributing to the respective age bins.

311 **Extended models for FOI inferences**

312 Standard models for all data types were extended to estimate age-specific infection risks
313 relative to the youngest age group. Age groups were defined with consideration of data points
314 available to inform estimates for the groups and consistency between data types to ease
315 comparison, see **Table S5**.

316 The **joint serology model** estimates parameters characterizing antibody kinetics and assay
317 noise in tandem with the infection risks. Probabilities of testing positive given susceptibility
318 states of individuals were then derived from those parameters. Due to known significant waning
319 in monotypic titers, we allowed for differing probabilities of testing positive in pre- vs post-
320 interval bleeds for individuals who have been infected once to reflect their difference in expected
321 time since infection. To derive these probabilities, we estimate the level of long-lasting titers
322 Ω_{long} present in both bleeds and the additional short-lived titers Ω_{short} present only in post-
323 interval bleeds (see **Figure S10**). Cross-reactive titers in DENV-naives, shared across studies,
324 were estimated relative to Ω_{long} . Following Salje et al (6), we formulated the relationship
325 between true underlying titers of individuals A , standard deviation of assay measurement
326 around the true titer σ , and probability of testing positive $P(\oplus | A, \sigma)$ as $1 - \Phi((A - v)/\sigma)$ where
327 Φ is the cumulative density function of a standard normal distribution. Because bleeding
328 intervals differed across cohort studies, we allowed Ω_{short} to differ across studies. The test
329 positive probability for seroprevalence data was assumed to be the same as pre-interval bleeds.
330 Likelihood was evaluated against seroprevalence and seroincidence data at two seropositivity
331 thresholds (10 and 20) to better inform this relationship.

332 **Model fitting**

333 In all models, we estimate annual FOI from ten years prior to the first observation to the year of
334 the last data point. FOI prior to the ten years were assumed to be constant. We used Bernoulli
335 likelihood to fit to serological data and negative binomial likelihood to fit to case data with priors
336 as defined in **Table S2**.

337 Parameters of all models were estimated from the data using Rstan v2.21.2 (52) with five
338 independent chains, each of length 2,000 (200 discarded as warm-up). Posteriors of all chains
339 combined were considered converged when R-hat < 1.1 and effective sample size > 300 for all
340 parameters. Where inferences were done for three repeated random samples (i.e. models
341 involving seroprevalence data), reported parameter estimates were from posterior draws pooled
342 across the repeats. Convergence was assessed prior to the pooling.

343 **Quantifying congruence**

344 We quantify congruence between any two sets of estimates via Pearson's correlation and an
345 average ratio between their posterior medians. The average ratio was obtained by fitting a linear

346 regression with zero intercept between the posterior medians and 95% confidence interval of
347 the ratio was calculated as the point estimate +/- 1.96 * standard error.

348 **Data and code availability**

349 Data and code used to generate all results are available at
350 <https://zenodo.org/doi/10.5281/zenodo.11635046>.

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355 **Disclaimer**

356 Material has been reviewed by the Walter Reed Army Institute of Research. There is no
357 objection to its presentation and/or publication. The opinions or assertions contained herein are
358 the private views of the author, and are not to be construed as official, or as reflecting true views
359 of the Department of the Army or the Department of Defense. The investigators have adhered to
360 the policies for protection of human subjects as prescribed in AR 70–25.

Main Text Figures

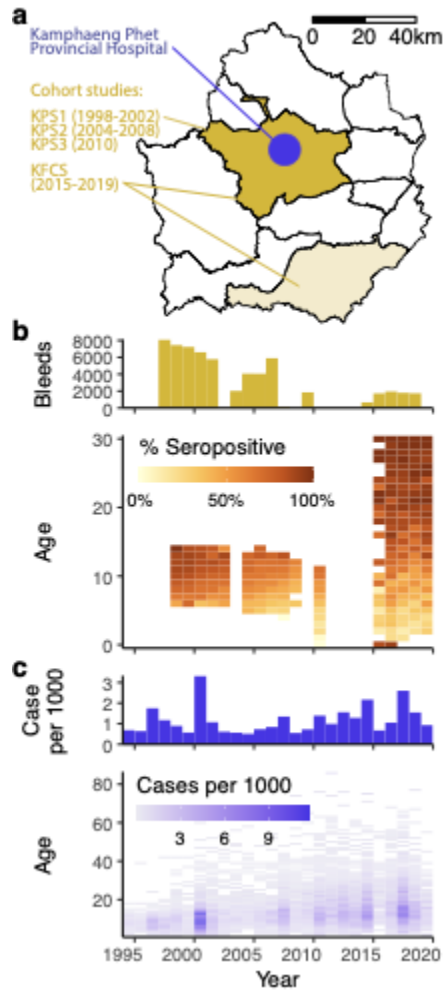


Figure 1. Study data. a) Map of Kamphaeng Phet province showing spatial coverage of cohort studies (colored) and location of Kamphaeng Phet Provincial Hospital (KPPH, blue point). **b)** Number of bleeds by year (top) and percentages of with GMT ≥ 10 by age and year of collection (bottom). **c)** Number of dengue cases reported at KPPH per thousand population by year (top), and by year and age (bottom).

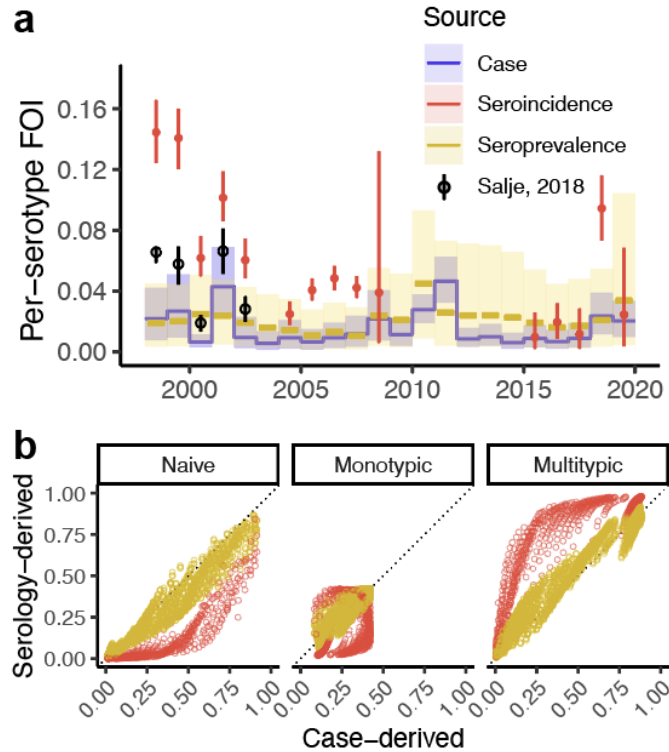


Figure 2. Estimates from standard force of infection (FOI) inference models. a) Annual FOI estimated from each of the data sources: sero-incidence data (red) and seroprevalence data (yellow) using seropositivity threshold of $\text{GMT} \geq 10$, and case data (blue). Annual FOI estimated from longitudinal samples of KPS1 (black, (7)) are included for comparison. **b)** Relationship between serology-derived (y-axis, respective colors) and case-derived susceptibility reconstructions (x-axis). Each point in the reconstruction represents the proportion in each age-year that has not been infected with DENV (naive), has been infected by one serotype (monotypic) or more than one serotype (multitypic).

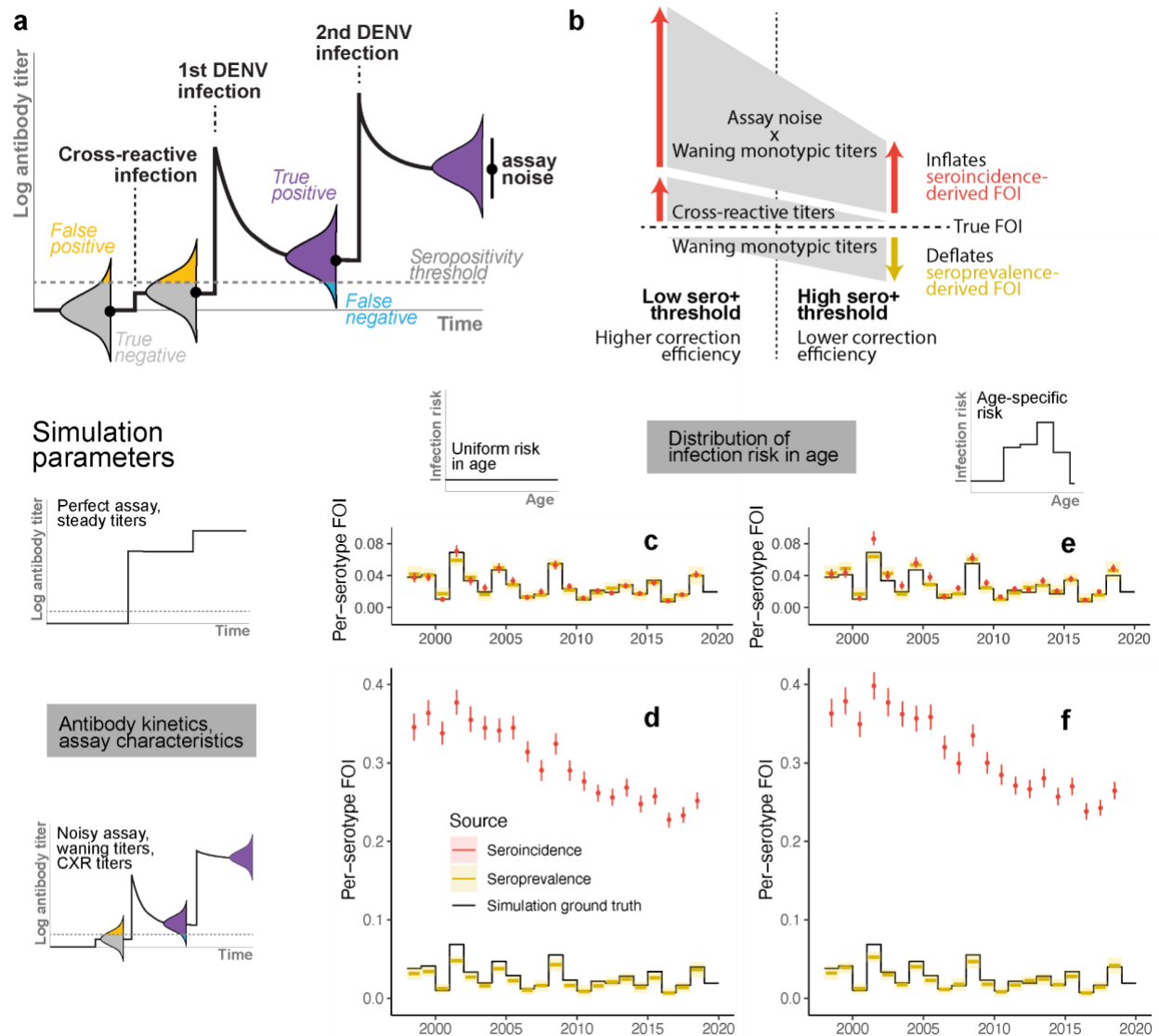


Figure 3. Biases in serology-derived force of infection (FOI) using simulated data with known true parameters. a) Illustration of anti-DENV antibody kinetics as an individual acquires a cross-reactive (CXR) virus infection or vaccination (i.e., not DENV), one DENV infection, and >1 DENV infections. Measured titers distribute around the true underlying titers with variability depending on the assay characteristics. **b)** Schematic of biases in serology-derived FOI and their correction efficiencies at low and high seropositivity thresholds. **c-f)** Antibody kinetics, assay characteristics (rows), and distribution of infection risk in age among susceptible individuals (columns) used to generate observed titer measurements and FOI inferred from those respective simulations using standard models for seroincidence (red) and seroprevalence (yellow).

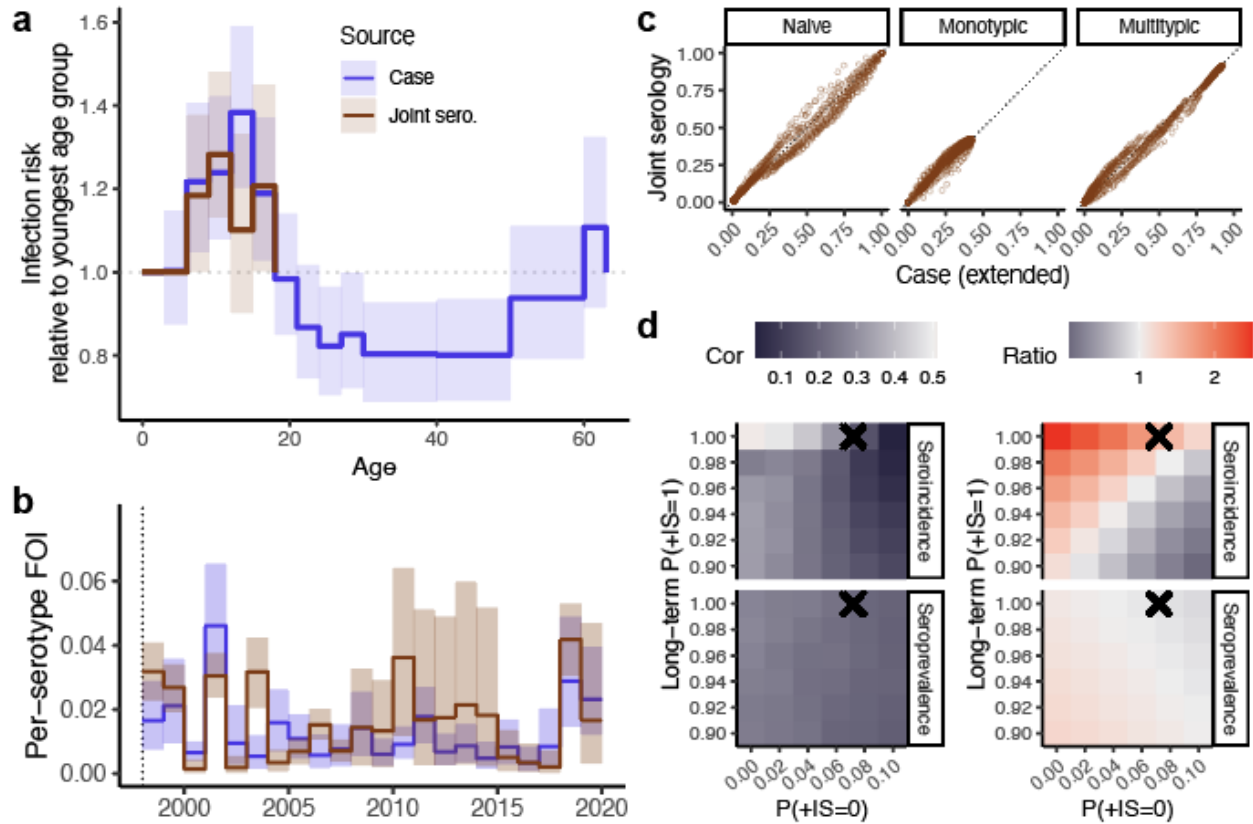


Figure 4. Infection risks in Kamphaeng Phet. **a)** Distribution of infection risk by age, **b)** annual per-serotype FOI inferred from the joint serology model (brown) and the extended case-based model (blue), and **c)** relationships between the susceptibility reconstructions. Each point in the reconstruction represents the proportion in each age-year that has not been infected with DENV (naive), has been infected by one serotype (monotypic) or more than one serotype (multitypic). **d)** Effects of presumed test positive probabilities in DENV-naives (x-axis) and long-term test positive probabilities in monotypically-infected individuals (y-axis) on the correlation and ratio between temporal FOIs inferred from the extended case-based model and temporal FOIs inferred from a single data source (either seroincidence or seroprevalence at seropositivity threshold of 10) imposed with age-specific risk inferred from the extended case-based model. Test positive probabilities estimated from the joint serology model are annotated as crosses for comparison.

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