

Article

Interrater and Intrarater Reliability of Cranial Anthropometric Measurements in Infants with Positional Plagiocephaly

Iñaki Pastor-Pons¹, María Orosia Lucha-López^{2,*}, Marta Barrau-Lalmolda¹, Iñaki Rodes-Pastor¹, Ángel Luis Rodríguez-Fernández³, César Hidalgo-García², and Jose Miguel Tricás-Moreno²

- ¹ Instituto de Terapias Integrativas, 50001 Zaragoza, Spain; inakipas@gmail.com (I.P.-P.); martabarraulalmolda@gmail.com (M.B.-L.); ikirodes95@gmail.com (I.R.-P.)
- ² Departamento de Fisiatría y Enfermería, Unidad de Investigación en Fisioterapia, Facultad de Ciencias de la Salud, Universidad de Zaragoza, 50009 Zaragoza, Spain; hidalgo@unizar.es (C.H.-G.); jmtricas@unizar.es (J.M.T.-M.)
- ³ Departamento de Fisioterapia, Facultad de Medicina, Universidad San Pablo CEU, 28925 Alcorcón, Spain; alrodfer@ceu.es
- * Correspondence: orolucha@unizar.es; Tel.: +34-626-480-131

Received: 26 November 2020; Accepted: 15 December 2020; Published: 17 December 2020



Abstract: (1) Background: anthropometric measurements with calipers are used to objectify cranial asymmetry in positional plagiocephaly but there is controversy regarding the reliability of different methodologies. Purpose: to analyze the interrater and intrarater reliability of direct anthropometric measurements with caliper on defined craniofacial references in infants with positional plagiocephaly. (2) Methods: 62 subjects (<28 weeks), with a difference of at least 5 mm between cranial diagonal diameters. Maximal cranial circumference, length and width and diagonal cranial diameters were measured. Intrarater (2 measurements) and interrater (2 raters) reliability was analyzed. (3) Results: intra- and interrater reliability of the maximal cranial length and width and right cranial diagonal was excellent: intraclass correlation coefficient (ICC) > 0.9. Intrarater and interrater reliability for the left cranial diagonal was excellent: ICC > 0.9 and difference in agreement in the Bland-Altman plot 0.0 mm, respectively. Intrarater and interrater reliability for the maximal cranial circumference was good: differences in agreement in Bland-Altman plots: intra: -0.03 cm; inter: -0.12 cm. (4) Conclusions: anthropometric measurements in a sample of infants with moderate positional plagiocephaly have shown excellent intra- and interrater reliability for maximal cranial length, maximal cranial width, and right and left cranial diagonals, and good intra- and interrater reliability in maximal cranial circumference measurement.

Keywords: positional plagiocephaly; deformational plagiocephaly; brachycephaly; head shape; anthropometry; head circumference

1. Introduction

Head and neck asymmetries are very common in healthy newborns [1]. Among these asymmetries, positional plagiocephaly (PP), a general term that describes the deformation of the skull and face, stands out, resulting from the application of prenatal or postnatal forces on the baby's head [2]. PP is characterized by asymmetric occipital flattening, accompanied by anterior displacement of the ear on the same side, contralateral parietal protrusion, and often ipsilateral frontal protrusion, with contralateral frontal flattening. These characteristics give the head a parallelogram shape when viewed from the top [3] and it can also be seen on the face with facial asymmetry [4]. Prevalence data are limited, depend on the geographical area, and are reported with a wide range. Data reported in the



literature varies from low (13–16%) [5–7] and median percentages (20–30%) of infants [8,9] to very high (61%) [1].

In addition to the classification of PP by visual estimation [10] which has only shown moderate reliability for clinical practice [11], there are different systems to objectify the cranial deformity: direct measurements with caliper, measurement of cranial asymmetry taking indirect references from photographs [12], 3D photographs [13], scanner images [14] or from a plastic modeling of the baby's head shape in a system called plagiocephalometry [15,16].

Anthropometric measurements with calipers have been used frequently to assess head shape but there is controversy regarding the reliability of the data [17] and lack of homogeneity regarding the anthropometric references used [18–22]. Craniometry with caliper is safe, fast, and low cost, which makes it an efficient method for clinical settings.

From the data obtained with anthropometry, data are extracted for the calculation of cranial indices or ratios. The cephalic index (CI) is calculated from the equation: cranial width/cranial length \times 100 [23] and determines the cranial morphology in terms of a more brachycephaly (CI > 85%) or dolichocephalic skull (CI < 75%) [23]. On the other hand, the cranial asymmetry indices or ratios require the diagonal diameters to be determined. The most used in the bibliography is the Cranial Vault Asymmetry Index (CVAI) [24,25]. The CVAI is calculated with the formula: cranial diagonal diameters difference/short diagonal diameter x 100 [23]. CVAI classifies plagiocephaly severity pursuant the Children's Healthcare of Atlanta scale in: level 1: <3.5%; level 2: 3.5 to 6.25%; level 3: 6.25 to 8.75%; level 4: 8.75 to 11.0%; level 5: >11.0% [26]. Classification of plagiocephaly severity may guide clinicians in the decision-making process regarding the treatment options of cranial asymmetry: repositioning, physical therapy, or cranial orthosis. Level 1 is considered within normal limits and no treatment is required. Level 2 requires repositioning at least. Level 3 calls for cranial remolding orthosis depending on age and history and levels 4 and 5 need cranial remolding orthosis [26]. Even with the use of cranial remolding orthosis, repositioning and physical therapy are recommended [27]. Early intervention translates into a significant improvement in PP regardless of the severity of the asymmetries [28]. Quantifying head shape is important for clinical management of PP and direct cranial anthropometric measurements provides an efficient solution for clinical settings.

The objective of the present study was to analyze the interrater and intrarater reliability of direct anthropometric measurements with caliper on defined craniofacial references, necessary for evaluation of PP in infants and for calculation of the most common indices used for the evaluation of cranial asymmetry.

2. Materials and Methods

2.1. Subjects

A cohort of 62 subjects under 28 weeks old with signs of PP were recruited. They were consecutively referred by pediatricians from sector III of the Aragonese Health Service.

According to the method of Walter et al. (1998), developed to calculate the required number of subjects for a reliability study, where reliability is measured using the intraclass correlation coefficient, with 0.8 being the minimum acceptable level of reliability and 0.9 the maximal expected level of reliability, 45.8 subjects are needed, admitting the following values for type I and type II errors: $\alpha = 0.05$ and $\beta = 0.20$ [29]. In this study, 62 subjects were recruited for the interrater and intrarater reliability study of anthropometric values, which is a sufficient sample to guarantee a good or almost perfect degree of agreement [30].

The inclusion criterion was to show a difference of at least 5 mm between cranial diagonal diameters [17]. Subjects with craniosynostosis, genetic, infectious, metabolic, or neurological diseases were excluded.

An informative document about the study was provided to the parents and an informed consent was signed after they had read the document and their questions about the study had been answered.

Regulations and guidelines regarding freedom, absence of coercion, disclosure of economic interests, understandable and complete information, confidentiality and acceptance were followed [31].

The Ethics Committee at the Aragon Health Sciences Institute approved the study (Registry No. C.P.–C.I. PI16/0275).

2.2. Measurements

The following anthropometric measurements were taken: maximal cranial circumference (MCC) (Martini et al., 2018); maximal cranial length (MCL): linear distance between glabella and opisthocranion; maximal cranial width (MCW): linear distance between the two euryon; diagonal cranial diameters (also called cranial diagonals) taken from the frontozygomatic suture to the contralateral lambdoid suture (Figures 1 and 2) [25].



Figure 1. Lateral view of diagonal cranial diameter evaluation, from the frontozygomatic point ipsilateral to the contralateral lambdoid suture, on the same horizontal plane.



Figure 2. Top view of diagonal cranial diameter evaluation, from the frontozygomatic point ipsilateral to the contralateral lambdoid suture, on the same horizontal plane.

From these data, the cranial vault asymmetry (CVA) (the difference between the cranial diagonal diameters) [32], the cephalic index (CI) and the CVAI were calculated. CI was calculated with the formula: cranial width/cranial length \times 100 [33], while CVAI was calculated using the formula: cranial diagonal diameters difference/short diagonal diameter \times 100 [23].

Measurements were made in a consultation of a clinical physiotherapy center, maintaining adequate lighting conditions for the procedure and as close as possible to the clinical daily routine. The subjects were evaluated by two trained raters in anthropometric evaluation of infants with plagiocephaly. Raters trained the measurements to agree on references search and measurement technique. They had 4 years of experience in the measurement of infants with plagiocephaly. An inextensible tape measure for the MCC and the caliper "mimos craniometer", manufactured by Think Pipe Line SLU, were used.

For the interrater reliability study, the measurements were performed first by rater 1 (measurement 1) and later by rater 2, without exchanging information, with no time interval between both measurements. For the intrarater reliability study, rater 1 measured the same parameters again next day (measurement 2). 24 h was considered sufficient time to not remember the data taken in the first measurement and to guarantee the reliability of the comparison of measurements.

In each measurement session, MCC, MCL, MCW were taken two times and cranial diagonal diameters were taken three times, by each examiner, non-consecutively, alternating the measurements of the different parameters and the mean of the three was recorded to carry out the statistical analysis of reliability.

2.3. Statistical Analyses

A descriptive analysis of qualitative variables, offering the absolute frequencies and the percentages in each category and of quantitative variables, offering the mean \pm standard deviation or median value (Q1–Q3) depending on whether the distribution was normal or non-normal, respectively, was carried out. Data distribution was analyzed with the Kolmogorov-Smirnov test with the Lilliefors correction, values of *p* < 0.05 were considered significant.

For intra- and interrater reliability analysis, if variables had a normal distribution, the intraclass correlation coefficient (ICC), one-factor model, random effects, was obtained. A 95% confidence interval was established for ICC. ICC has been interpreted according to the ranges established by Koo and Li [30]: values less than 0.5 are indicative of low reliability, values between 0.5 and 0.75 of moderate agreement, values between 0.75 and 0.9 of good agreement, and values greater than 0.90 of excellent reliability.

If any of the variables was not normally distributed, Bland-Altman plot was made to evaluate degree of agreement of the measurements. In Bland-Altman plots, three parallel lines were represented:

Upper limit of agreement: mean difference + $1.96 \times SD$.

Mean difference: mean value determined by one data series—mean value determined by the other data series. It reflects the systematic error.

Lower limit of agreement: mean difference – $1.96 \times SD$.

If two data series for which reliability is being studied obtain similar values on average, then the mean difference will be zero or close to zero. If it is far from this value, it would mean that the two methods produce different results.

The numerical analysis was performed using SPSS 22.0 for Windows and the Bland-Altman plots were performed with MedCalc for Windows.

3. Results

The interrater (two raters consecutively) and intrarater (same rater, on two consecutive days, at same time) reliability studies of the anthropometric measurements were carried out in 62 subjects.

Median age of the subjects at the time of measurement was 16 weeks (Table 1). 43.5% were female and 56.5% were male (Table 1). Descriptive values of cranial asymmetry indices can be consulted in Table 1.

Infants	with PP ($n = 62$)	Descriptive
Gender	Females $(n = 27)$	43.5%
Ma	ales ($n = 35$)	56.5%
А	ge (weeks)	16.0 (13.0; 19.0) **
	CI (%)	86.8 ± 7.4 *
CVA (mm)		8.2 (6.0; 11.6) **
CVAI (%)		6.4 (5.0; 9.4) **

Table 1. Sample description.

* Mean ± standard deviation. ** Median (Q1; Q3). CI: Cephalic index; CVA: Cranial vault asymmetry; CVAI: Cranial vault asymmetry index; PP: Positional plagiocephaly.

Descriptive values for the two MCC measurements of rater 1 are shown in Table 2.

Table 2. Data fo	r the intrarater	reliability stud	v of MCC.

Infants with PP ($n = 62$)	Rater 1	Descriptive	Kolmogorov-Smirnov. Sig.
МСС	Measurement 1	40.4 (39.4; 42.0) **	0.016
MCC	Measurement 2	40.8 ± 1.8 *	0.200

* Mean ± standard deviation. ** Median (Q1; Q3). MCC: Maximal cranial circumference.

Reliability was analyzed with the Bland-Altman Plot (Figure 3), since measurement 1 of rater 1 did not show a normal distribution, which would have allowed ICC calculation. However, the plot shows a good degree of agreement since the mean of the differences is very close to 0, being -0.03 cm.

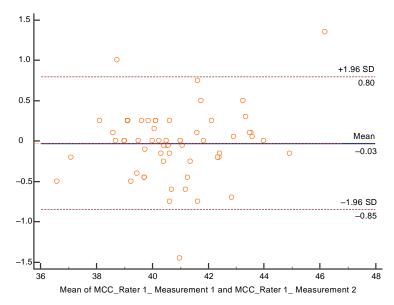


Figure 3. Bland-Altman plot of MCC intrarater reliability. SD: Standard deviation.

The intrarater reliability for the rest of the variables was analyzed with the ICC since all variables followed a normal distribution. Confidence intervals of the ICC were calculated, as well as its p values (Table 3).

Infants with	D.t.1	D	Kolmogorov-Smirnov.	100	95% Cor Inte	nfidence rval	p Value
PP $(n = 62)$	Rater 1	Descriptive	Sig.	ICC	Lower Limit	Upper Limit	
MCI	Measurement 1	132.6 ± 7.5 *	0.077	0.007	0.070	0.000	0.000
MCL	Measurement 2	132.8 ± 7.4 *	0.200	0.987	0.978	0.992	0.000
MCM	Measurement 1	$114.8 \pm 8.1 *$	0.200	0.077	0.0(2	0.007	0.000
MCW	Measurement 2	115.5 ± 7.6 *	0.200	0.977	0.962	0.986	0.000
Right cranial	Measurement 1	127.6 ± 8.2 *	0.200	0.000	0.000	0.004	0.000
diagonal	Measurement 2	128.1 ± 7.9 *	0.200	0.990	0.983	0.994	0.000
Left cranial	Measurement 1	130.5 ± 8.7 *	0.200	0.000	0.070		0.000
diagonal	Measurement 2	131.1 ± 8.2 *	0.200	0.983	0.972	0.990	0.000

Table 3. Intrarater reliability study of MCL, MCW, right cranial diagonal and left cranial diagonal.

* Mean ± standard deviation. ICC: Intraclass correlation coefficient; MCL: Maximal cranial length; MCW: Maximal cranial width.

Descriptive values of MCC and left cranial diagonal measurements of rater 1 and rater 2 are shown in Table 4.

	Table 4. Data	or the interrater reliabilit	v study of MCC and left	cranial diagonal.
--	---------------	------------------------------	-------------------------	-------------------

Infants with PP ($n = 62$)	Raters	Descriptive	Kolmogorov-Smirnov. Sig.
N/CC	Rater 1	40.4 (39.4; 42.0) **	0.016
MCC	Rater 2	40.9 ± 1.8 *	0.100
L off granial diagonal	Rater 1	130.5 ± 8.7 *	0.200
Left cranial diagonal	Rater 2	130.5 (125.1; 135.4) **	0.038

^{*} Mean ± standard deviation. ** Median (Q1; Q3).

Reliability was analyzed with Bland-Altman Plots (Figures 4 and 5), because MCC measurement of rater 1 and left cranial diagonal measurement of rater 2 did not show a normal distribution, which would have allowed ICC calculation. Plot referring to MCC (Figure 4) shows a good degree of agreement since the mean of the differences is very close to 0, being -0.12 cm.

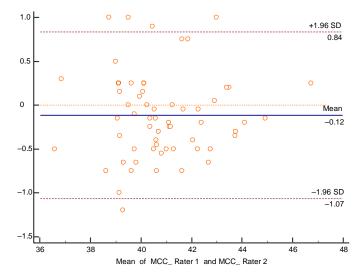


Figure 4. Bland-Altman plot of MCC interrater reliability.

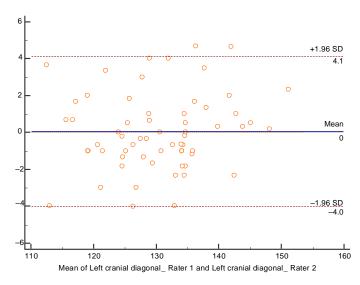


Figure 5. Bland-Altman plot of left cranial diagonal interrater reliability.

Plot referring to left cranial diagonal (Figure 5) shows an excellent degree of agreement, since mean difference is 0.0 cm.

The interrater reliability for the rest of the variables was analyzed with the ICC since all the variables followed a normal distribution. Confidence intervals of ICCs were calculated, as well as its *p* values (Table 5).

Infants with	D (Kolmogorov-Smirnov.	100		nfidence rval	
$PP (n = 62) \qquad \text{Raters}$	Descriptive	Sig.	ICC	Lower Limit	Upper Limit	<i>p</i> Value	
MCI	Rater 1	132.6 ± 7.5 *	0.077	0.092	0.070	0.000	0.000
MCL	Rater 2	133.4 ± 7.0 *	0.200	0.983	0.972	0.990	0.000
MCM	Rater 1	$114.8 \pm 8.1 *$	0.200	0.070	0.062	0.097	0.000
MCW Rater 2	Rater 2	116.2 ± 7.8 *	0.200	0.978	0.963	0.987	0.000
Right cranial	Rater 1	127.6 ± 8.2 *	0.200	0.007	0.077	0.000	0.000
diagonal	Rater 2	127.4 ± 8.2 *	0.200	0.986	0.977	0.992	0.000

Table 5. Interrater reliability study of maximal cranial length, maximal cranial width, right cranial diagonal.

* Mean ± standard deviation.

4. Discussion

The sample of this study consisted of 62 children with PP, with a difference of at least 5 mm between diagonal cranial diameters, i.e., children with at least moderate deformity [17]. 43.5% were female and 56.5% were male, so the slightly higher prevalence in male is according to data reported in the literature [34]. 75% of the sample was older than 13 weeks, so most of the sample had already exceeded three months, age below which PP could be found in almost half of infants [35].

Mean CI was 86.8%. Normal range described is between 75 and 85% [25], therefore, infants in the study had a tendency toward brachycephaly: they had the skull with a predominance of width over length.

Median CVA was 8.19 mm. According to Mortenson and Steinbok, who classify CVA into the following categories: normal CVA < 3 mm, mild/moderate CVA \geq 3 mm and CVA \leq 12 mm, moderate/severe CVA > 12 mm [17], the sample had a moderate PP.

Median CVAI was 6.43%. Plagiocephaly severity scale, according to CVAI is: level 1: <3.5%; level 2: 3.5 to 6.25%; level 3: 6.25 to 8.75%; level 4: 8.75 to 11.0%; level 5: >11.0% [26]. Therefore,

infants in this study presented a level 3, which depending on the age and medical history, may require cranial remodeling orthoses for their treatment.

Intrarater reliability was excellent for MCL and MCW, and for left and right cranial diagonal diameters. Good reliability was observed in MCC measurement.

Interrater reliability was excellent for MCL and MCW, and for left and right cranial diagonal diameters. Good reliability was observed in MCC measurement.

Slightly worse records in the reliability of the MCM were due to technical difficulty in the measurement. Exact references are not used for its recording. Maximal value of the fronto-occipital circumference is sought, which easily results in greater variability.

Our data confirm good results obtained by Mortenson and Steinbok (2006) regarding intrarater reliability and are superior to them regarding CVA interrater reliability, in infants referred for plagiocephaly or torticollis [17]. These authors took the distance between the frontozygomatic point (most medial point of the temporal crest of the frontal bone) and contralateral euryon (most lateral point of the neurocranium, it can be located in the parietal or in the temporal squama) as anthropometric references to establish the diagonals [17], while in this study the distance between the frontozygomatic point and the inner rim of the lambdoid suture of the contralateral side has been used, according to Wilbrand et al. [25], who reported good intra- and interrater reliability for measurements of circumference, length, width, and diagonal distances.

Skolnick proposed the distance between the contralateral frontozygomatic-euryon points as the best correlated with cranial perimeter [22,36], while in this study references proposed by Wilbrand were chosen, since the distance between the frontozygomatic-euryon points does not seem to show the global characteristic of the cranial deformation in its posterior part [22,36].

Skolnick et al. (2015) conducted a comparison study between direct anthropometric measurements and digital measurements. The study included caliper measurements of the length, width and diagonals and the measurement of the circumference by meter. In the results they found an excellent reproducibility of all the caliper measurements, and they appreciated a strong correlation between direct and digital measurements (R2 > 0.90). Caliper measurements were 1 to 4 mm shorter than digital with consistent variation [37]. Mendonca et al. (2013) in a previous study found less correlation between direct and digital measurements, with a significant difference of 6% in measurements of anteroposterior length and cranial width [38].

Direct anthropometric measurements on the skull using a caliper are a reliable tool for diagnosis and decision-making in plagiocephaly [25]. Results of our study with an analysis in a larger sample of children with PP contrast these previous results. It is an easy, effective, low cost and reproducible method if the examiners and assistants are well trained. These advantages favor that this measurement system seems to be the most used by the American Society of Maxillofacial Surgeons [39].

5. Conclusions

Anthropometric measurements taken in a sample of infants with moderate severity PP have shown excellent intra- and interrater reliability for MCL, MCW, and right and left cranial diagonals, and good intra- and interrater reliability in MCC measurement.

Author Contributions: Conceptualization: I.P.-P.; Methodology: I.P.-P.; Investigation: M.B.-L., I.R.-P.; Resources: M.B.-L.; Data Curation: I.R.-P.; Writing—Original Draft: I.P.-P.; Writing—Review—review & editing: M.O.L.-L., C.H.-G.; Formal Analysis: M.O.L.-L.; Visualization: Á.L.R.-F., J.M.T.-M.; Supervision: Á.L.R.-F., J.M.T.-M. All authors have read and agreed to the published version of the manuscript.

Funding: This research received no external funding.

Acknowledgments: The authors wish to thank the management and pediatricians in Section III of the Aragon Healthcare System (Zaragoza, Spain). Thanks to the Institute of Integrative Therapies (Instituto de Terapias Integrativas) in Zaragoza and its administrative staff. Thanks especially to the children and their families.

Conflicts of Interest: The authors declare no conflict of interest.

References

- 1. Stellwagen, L.; Hubbard, E.; Chambers, C.; Jones, K.L. Torticollis, facial asymmetry and plagiocephaly in normal newborns. *Arch. Dis. Child.* **2008**, *93*, 827–831. [CrossRef] [PubMed]
- 2. Kalra, R.; Walker, M.L. Posterior plagiocephaly. Child's Nerv. Syst. 2012, 28, 1389–1393. [CrossRef] [PubMed]
- 3. Robinson, S.; Proctor, M. Diagnosis and management of deformational plagiocephaly: A review. *J. Neurosurg. Pediatr.* **2009**, *3*, 284–295. [CrossRef] [PubMed]
- 4. Baumler, C.; Leboucq, N.; Captier, G. Étude de l'asymétrie mandibulaire dans les plagiocéphalies sans synostose. *Rev. Stomatol. Chir. Maxillofac.* **2007**, *108*, 424–430. [CrossRef]
- 5. Peitsch, W.K.; Keefer, C.H.; LaBrie, R.A.; Mulliken, J.B. Incidence of Cranial Asymmetry in Healthy Newborns. *Pediatrics* **2002**, *110*, e72. [CrossRef]
- 6. Littlefield, T.R.; Saba, N.M.; Kelly, K.M. On the current incidence of deformational plagiocephaly: An estimation based on prospective registration at a single center. In *Proceedings of the Seminars in Pediatric Neurology*; Elsevier: Amsterdam, The Netherlands, 2004; Volume 11, pp. 301–304.
- 7. Hutchison, B.L.; Hutchison, L.a.D.; Thompson, J.M.D.; Mitchell, E.A. Plagiocephaly and brachycephaly in the first two years of life: A prospective cohort study. *Pediatrics* **2004**, *114*, 970–980. [CrossRef]
- 8. Bialocerkowski, A. Physiotherapy reduces the risk of deformational plagiocephaly in infants who have a preferred head position when lying supine. *Aust. J. Physiother.* **2008**, *54*, 283. [CrossRef]
- 9. Koren, A.; Reece, S.M.; Kahn-D'angelo, L.; Medeiros, D. Parental Information and Behaviors and Provider Practices Related to Tummy Time and Back to Sleep. *J. Pediatr. Health Care* **2010**, *24*, 222–230. [CrossRef]
- Argenta, L.; David, L.; Thompson, J. Clinical classification of positional plagiocephaly. *J. Craniofacial Surg.* 2004, 15, 368–372. [CrossRef]
- 11. Spermon, J.; Spermon-Marijnen, R.; Scholten-Peeters, W. Clinical classification of deformational plagiocephaly according to Argenta: A reliability study. *J. Craniofacial Surg.* **2008**, *19*, 664–668. [CrossRef]
- Hutchison, B.L.; Stewart, A.W.; de Chalain, T.B.; Mitchell, E.A. A randomized controlled trial of positioning treatments in infants with positional head shape deformities. *Acta Pediatr. Int. J. Paediatr.* 2010, 99, 1556–1560. [CrossRef] [PubMed]
- Aarnivala, H.; Vuollo, V.; Heikkinen, T.; Harila, V.; Holmström, L.; Pirttiniemi, P.; Valkama, A.M. Accuracy of measurements used to quantify cranial asymmetry in deformational plagiocephaly. *J. Cranio Maxillofac. Surg.* 2017, 45, 1349–1356. [CrossRef] [PubMed]
- 14. Collett, B.R.; Gray, K.E.; Starr, J.R.; Heike, C.L.; Cunningham, M.L.; Speltz, M.L.; Heike, C.L.; Collett, B.R.; Cunningham, M.L.; Starr, J.R.; et al. Development at Age 36 Months in Children With Deformational Plagiocephaly. *Pediatrics* **2013**, *131*, e109–e115. [CrossRef] [PubMed]
- Van Wijk, R.M.; Van Til, J.A.; Groothuis-Oudshoorn, C.G.M.; L'Hoir, M.P.; Boere-Boonekamp, M.M.; Ijzerman, M.J. Parents' decision for helmet therapy in infants with skull deformation. *Child's Nerv. Syst.* 2014, *30*, 1225–1232. [CrossRef]
- Van Adrichem, L.L.N.A.; van Vlimmeren, L.L.A.; Čadanová, D.; Helders, P.J.M.; Engelbert, R.H.H.; Van Neck, H.J.W.; Koning, A.H.J. Validation of a simple method for measuring cranial deformities (plagiocephalometry). J. Craniofacial Surg. 2008, 19, 15–21. [CrossRef]
- 17. Mortenson, P.A.; Steinbok, P. Quantifying positional plagiocephaly: Reliability and validity of anthropometric measurements. *J. Craniofacial Surg.* **2006**, *17*, 413–419. [CrossRef]
- Mulliken, J.B. Analysis of posterior plagiocephaly: Deformational versus synostotic. *Plast. Reconstr. Surg.* 1999, 103, 371–380. [CrossRef]
- Lee, R.P.; Teichgraeber, J.F.; Baumgartner, J.E.; Waller, A.L.; English, J.D.; Lasky, R.E.; Miller, C.C.; Gateno, J.; Xia, J.J. Long-Term Treatment Effectiveness of Molding Helmet Therapy in the Correction of Posterior Deformational Plagiocephaly: A Five-Year Follow-Up. *Cleft Palate Craniofacial J.* 2008, 45, 240–245. [CrossRef]
- Rogers, G.F.; Miller, J.; Mulliken, J.B. Comparison of a Modifiable Cranial Cup versus Repositioning and Cervical Stretching for the Early Correction of Deformational Posterior Plagiocephaly. *Plast. Reconstr. Surg.* 2008, 121, 941–947. [CrossRef]
- 21. Wilbrand, J.-F.; Wilbrand, M.; Pons-Kuehnemann, J.; Blecher, J.-C.; Christophis, P.; Howaldt, H.-P.; Schaaf, H. Value and reliability of anthropometric measurements of cranial deformity in early childhood. *J. Cranio Maxillofac. Surg.* **2011**, *39*, 24–29. [CrossRef]

- 22. Skolnick, G.B.; Naidoo, S.D.; Nguyen, D.C.; Patel, K.B.; Woo, A.S. Evidence for Use of Frontozygomaticus and Contralateral Eurion as Hand-Caliper Landmarks for Assessment of Deformational Plagiocephaly. *J. Craniofacial Surg.* **2016**, *27*, 1498–1500. [CrossRef] [PubMed]
- 23. Loveday, B.P.T.; de Chalain, T.B. Active counterpositioning or orthotic device to treat positional plagiocephaly? *J. Craniofacial Surg.* **2001**, *12*, 308–313. [CrossRef] [PubMed]
- 24. Kluba, S.; Kraut, W.; Reinert, S.; Krimmel, M. What is the optimal time to start helmet therapy in positional plagiocephaly? *Plast. Reconstr. Surg.* **2011**, *128*, 492–498. [CrossRef] [PubMed]
- Wilbrand, J.F.; Schmidtberg, K.; Bierther, U.; Streckbein, P.; Pons-Kuehnemann, J.; Christophis, P.; Hahn, A.; Schaaf, H.; Howaldt, H.P. Clinical classification of infant nonsynostotic cranial deformity. *J. Pediatr.* 2012, 161, 1120–1125. [CrossRef]
- 26. Holowka, M.A.; Reisner, A.; Giavedoni, B.; Lombardo, J.R.; Coulter, C. Plagiocephaly severity scale to aid in clinical treatment recommendations. *J. Craniofacial Surg.* **2017**, *28*, 717–722. [CrossRef]
- 27. Steinberg, J.P.; Rawlani, R.; Humphries, L.S.; Rawlani, V.; Vicari, F.A. Effectiveness of Conservative Therapy and Helmet Therapy for Positional Cranial Deformation. *Plast. Reconstr. Surg.* **2015**, *135*, 833–842. [CrossRef]
- 28. Kelly, K.M.; Littlefield, T.R.; Pomatto, J.K.; Ripley, C.E.; Beals, S.P.; Joganic, E.F. Importance of Early Recognition and Treatment of Deformational Plagiocephaly with Orthotic Cranioplasty. *Cleft Palate Craniofacial J.* **1999**, *36*, 127–130. [CrossRef]
- 29. Walter, S.D.; Eliasziw, M.; Donner, A. Sample size and optimal designs for reliability studies. *Stat. Med.* **1998**, 17, 101–110. [CrossRef]
- 30. Koo, T.K.; Li, M.Y. A Guideline of Selecting and Reporting Intraclass Correlation Coefficients for Reliability Research. J. Chiropr. Med. 2016, 15, 155–163. [CrossRef]
- 31. Leibson, T.; Koren, G. Informed Consent in Pediatric Research. Pediatr. Drugs 2015, 17, 5–11. [CrossRef]
- 32. Cho, I.K.; Eom, J.R.; Lee, J.W.; Yang, J.D.; Chung, H.Y.; Cho, B.C.; Choi, K.Y. A Clinical Photographic Method to Evaluate the Need for Helmet Therapy in Deformational Plagiocephaly. *J. Craniofacial Surg.* **2018**, *29*, 149–152. [CrossRef] [PubMed]
- Roosenboom, J.; Lee, M.K.; Hecht, J.T.; Heike, C.L.; Wehby, G.L.; Christensen, K.; Feingold, E.; Marazita, M.L.; Maga, A.M.; Shaffer, J.R. Mapping genetic variants for cranial vault shape in humans. *PLoS ONE* 2018, 13, e0196148. [CrossRef] [PubMed]
- 34. Vernet, O.; de Ribaupierre, S.; Cavin, B.; Rilliet, B. Treatment of posterior positional plagiocephaly. *Arch. Pediatr. Organe Off. Soc. Fr. Pediatr.* **2008**, *15*, 1829. [CrossRef] [PubMed]
- 35. Mawji, A.; Vollman, A.R.; Hatfield, J.; McNeil, D.A.; Sauvé, R. The incidence of positional plagiocephaly: A cohort study. *Pediatrics* **2013**, *132*, 298–304. [CrossRef]
- 36. Skolnick, G.B.; Naidoo, S.D.; Patel, K.B.; Woo, A.S. Analysis of Digital Measures of Cranial Vault Asymmetry for Assessment of Plagiocephaly. *J. Craniofacial Surg.* **2014**, *25*, 1178–1182. [CrossRef]
- Skolnick, G.B.; Naidoo, S.D.; Nguyen, D.C.; Patel, K.B.; Woo, A.S. Comparison of Direct and Digital Measures of Cranial Vault Asymmetry for Assessment of Plagiocephaly. *J. Craniofacial Surg.* 2015, 26, 1900–1903. [CrossRef]
- 38. Mendonca, D.A.; Naidoo, S.D.; Skolnick, G.; Skladman, R.; Woo, A.S. Comparative Study of Cranial Anthropometric Measurement by Traditional Calipers to Computed Tomography and Three-dimensional Photogrammetry. *J. Craniofacial Surg.* **2013**, *24*, 1106–1110. [CrossRef]
- Purnell, C.A.; Benz, A.W.; Gosain, A.K. Assessment of Head Shape by Craniofacial Teams. J. Craniofacial Surg. 2015, 26, 1808–1811. [CrossRef]

Publisher's Note: MDPI stays neutral with regard to jurisdictional claims in published maps and institutional affiliations.



© 2020 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (http://creativecommons.org/licenses/by/4.0/).