



Body Mass Index Trajectories in the First 5 Years and Associated Antenatal Factors

Molly Mattsson^{1*}, Deirdre M. Murray², Colin P. Hawkes³, Mairead Kiely⁴, Carol Ní Chaoimh⁵, Fergus P. McCarthy⁵, Regien Biesma^{6†} and Fiona Boland^{7†}

¹ Division of Population Health Sciences, Royal College of Surgeons in Ireland, Dublin, Ireland, ² Department of Paediatrics and Child Health, University College Cork, Cork, Ireland, ³ Division of Endocrinology and Diabetes, The Children's Hospital of Philadelphia, Philadelphia, PA, United States, ⁴ Cork Centre for Vitamin D and Nutrition Research, School of Food and Nutritional Sciences, University College Cork, Cork, Ireland, ⁵ Irish Centre for Maternal and Child Health Research, Cork University Maternity Hospital, University College Cork, Cork, Ireland, ⁶ University Medical Center Groningen, University of Groningen, Groningen, Netherlands, ⁷ Data Science Centre, Royal College of Surgeons in Ireland, Dublin, Ireland

OPEN ACCESS

Edited by:

Radhika Muzumdar, University of Pittsburgh, United States

Reviewed by:

Andrew Whatmore, The University of Manchester, United Kingdom Brigid Gregg, University of Michigan, United States Barbara Heude, Institut National de la Santé et de la Recherche Médicale (INSERM), France

*Correspondence:

Molly Mattsson mollymattsson@rcsi.ie

[†]These authors have contributed equally to this work

Specialty section:

This article was submitted to Pediatric Endocrinology, a section of the journal Frontiers in Pediatrics

Received: 28 October 2020 Accepted: 25 January 2021 Published: 19 February 2021

Citation:

Mattsson M, Murray DM, Hawkes CP, Kiely M, Ní Chaoimh C, McCarthy FP, Biesma R and Boland F (2021) Body Mass Index Trajectories in the First 5 Years and Associated Antenatal Factors. Front. Pediatr. 9:622381. doi: 10.3389/fped.2021.622381 **Background:** The increasing prevalence of childhood obesity is an important public health issue and the development of obesity in early life and associated risk factors need to be better understood. The aim of this study was to identify distinct body mass index trajectories in the first 5 years of life and to examine their associations with factors identified in pregnancy, including metabolic parameters.

Methods: BMI measurements from 2,172 children in Ireland enrolled in the BASELINE cohort study with BMI assessments at birth, 2, 6, and 12 months, and 2 and 5 years were analyzed. Growth mixture modeling was used to identify distinct BMI trajectories, and multivariate multinomial logistic regression was used to assess the association between these trajectories and antenatal factors.

Results: Three distinct BMI trajectories were identified: normal (89.6%); rapid gain in the first 6 months (7.8%); and rapid BMI after 12 months (2.6%). Male sex and higher maternal age increased the likelihood of belonging to the rapid gain in the first 6 months trajectory. Raised maternal BMI at 15 weeks of pregnancy and lower cord blood IGF-2 were associated with rapid gain after 1 year.

Conclusion: Sex, maternal age and BMI, and IGF-2 levels were found to be associated with BMI trajectories in early childhood departing from normal growth. Further research and extended follow-up to examine the effects of childhood growth patterns are required to understand their relationship with health outcomes.

Keywords: growth, childhood obesity, epidemiology, growth mixture modeling, IGF

INTRODUCTION

Obesity is a growing public health issue globally and a risk factor for multiple non-communicable diseases (NCD) (1-6). The worldwide prevalence of obesity has increased from 3.2% in 1975 to 10.8% in 2014 in men, and from 6.4 to 14.9% in women. The prevalence in the Irish population surpasses the global mean and has been estimated as 26% both men and women (7). Children and adolescents with obesity have been found to be around five times more likely to have obesity in

1

adulthood compared to children without obesity (8). In Ireland, the prevalence of obesity at 5 years of age increased from 1.2% in 1975 to 11.3% in 2016 in boys, and from 1.2% in 1975 to 10.1% in 2016 in girls.

Tracking children's growth patterns longitudinally allows for the assessment of dynamic changes in size. This may permit a more accurate identification of young children at higher risk when compared with assessments at a single time point, and provide insight into early life determinants of childhood and adult overweight and obesity (9). Body mass index (BMI) for assessment of weight status is generally not used under the age of 2 (10) and instead the American Academy of Pediatrics recommend weight-for-length (WFL) for this age group (11). However, previous studies have shown that BMI z-scores in infancy to have a significantly higher positive predictive value for early childhood obesity than WFL-Z (12) and that raw BMI may be a better indicator of body composition at 1 months of age compared to WFL (13).

Longitudinal random-effects and latent growth curve models are commonly used to explore growth trajectories in childhood. These methods allow for individual variability; however they assess the average pattern of change and assume that individuals belong to the same underlying population, represented by a single growth curve. As an alternative, a latent class approach may be adopted, allowing researchers to identify and describe underlying subgroups within a population based on different trajectories (14). Latent class growth analysis (LCGA) estimates a mean growth curve for each class, but no individual variation around the mean growth curve is allowed. Growth mixture modeling (GMM) combines the features of the random effects model and LCGA by estimating both mean growth curves for each class and individual variation around these growth curves (15).

The development of obesogenic growth trajectories and childhood obesity is complex. From a life course model of chronic disease epidemiology perspective, childhood obesity has previously been conceptualized in a framework where individuallevel factors, including biological, social, and behavioral risks, are acting within the influence of the child's family environment, which is, in turn imbedded in the context of the community environment (16). Additionally, it is helpful to consider critical periods of biological and behavioral plasticity for obesity risk, beginning as early as fetal life (17). In this study, we have aimed to investigate BMI trajectories in early life and factors identified during pregnancy. Sociodemographic, lifestyle and metabolic health factors in pregnancy, including low socioeconomic status (SES), smoking, maternal pre-pregnancy BMI, gestational weight gain (GWG), and gestational diabetes mellitus (GDM), have been consistently associated with multiple aspects of child growth, such as birthweight, growth velocity, BMI, and childhood obesity (18-21).

We previously conducted a systematic review of group-based trajectory modeling for BMI trajectories in childhood, and found that trajectories of excessive rapid gain were associated with several predictors, including high maternal pre-pregnancy BMI, GWG and smoking during pregnancy (22). The association of maternal and infant metabolic parameters, beyond prepregnancy BMI, GWG and GDM, in relation to offspring's longitudinal BMI trajectories does however remain largely unexplored. Insulin-like growth factor (IGF)-1 and IGF-2 are peptides primarily secreted by the liver. Small studies have previously demonstrated a correlation between cord blood IGF-I concentration and birthweight (23, 24) as well as weight at 6 months (25). Cord blood concentrations of IGF-2 have been found to be related to IGF-2 levels at age five, which in turn has been related to fat mass (24). Further, low IGF-2 levels have been documented in pre-puberal children with obesity (26). Leptin is a hormone secreted primarily by adipocytes and evidence suggests that low fetal leptin concentrations may mediate weight gain during infancy and play a role in the development of obesity (27– 29). A recent study suggests that cord blood leptin may play a modest mediating role in early postnatal catch-up or catch-down in weight (30).

The aim of this study was to identify distinct BMI trajectories in the first 5 years of life using GMM methods and to examine the associations between these trajectories and factors identified during pregnancy, including maternal and fetal metabolic parameters.

METHODS

Study Population and Design

Study subjects were participants of the Cork BASELINE (Babies after SCOPE: Evaluating the Longitudinal Impact using Neurological and Nutritional Endpoints) Birth Cohort Study (31), a mother-infant prospective birth cohort study based in Cork, Ireland. It was initiated in 2008 as a follow-up to the SCOPE (Screening for Pregnancy Endpoints) Ireland study (32), a major multi-center prospective pregnancy study involving primiparous low-risk women. Exclusion criteria included women recognized as high risk of pre-eclampsia, small for gestational age baby, or spontaneous preterm birth, as these were the primary outcomes of interest in the study. A total of 1,768 women were recruited to SCOPE Ireland at Cork University Maternity Hospital (CUMH). Each participant was interviewed by a research midwife at 14-16 and 19-21 weeks of gestation and detailed clinical information, maternal anthropometry measurements, ultrasound data, and blood specimens were collected. Umbilical cord blood was collected at birth. One thousand, five hundred thirty-seven SCOPE participants consented for their infants to participate in the BASELINE study and during a second stream of recruitment, a further 646 infants were recruited after delivery from the postnatal wards of CUMH, with a singleton pregnancy being the main inclusion criterion (31). Pediatric follow-up with in-person assessments were conducted at birth, 2, 6, and 12 months and at 2 and 5 years. Data on the child's early-life environment, diet, health, and development were recorded at each assessment. In total, 2,172 infants were followed up after birth.

Maternal Data

Maternal age, smoking and alcohol use during pregnancy, and BMI were obtained during the SCOPE study interviews at 14– 16 and 19–21 weeks, and are only available for the mothers who participated in both SCOPE and BASELINE. GDM diagnosis was established post-delivery from medical records. Education level, income, and marital status were determined at the 2 month assessment of BASELINE.

Infant Data

Umbilical cord samples (including leptin and IGF-1 and -2 concentrations) were obtained at birth for infants in the SCOPE study and z-score variables were created using the *z*-score command in Stata. Leptin was only collected in a subset of children (n = 405). Gestational age (GA) at delivery was obtained from medical records, with preterm birth defined as <37 weeks gestation.

Child Anthropometric Measurements

Measurements of child weight and length/height were obtained at birth 2, 6, and 12 months and 2 years and 5 years of age. The available BMI data at each time point is outlined in **Supplementary Table 1**. Naked weight was measured using digital scales at birth and at 2 months correct to the nearest 0.01 kg and at 6 and 12 months and at 2 and 5 years correct to the nearest 0.1 kg. Supine length correct to the nearest 0.1 cm was measured at birth 2, 6 and 12 months. At 2 and 5 years, standing height was measured using a wall mounted stadiometer (33). Waist circumference was measured at each assessment. All measures were performed by trained research staff.

Statistical Analysis

BMI trajectories in the first 5 years of life were analyzed using GMM. A longitudinal change model was assumed where each growth pattern was characterized by random intercept and linear, quadratic, and cubic terms by age, allowing for curved developmental patterns. Cubic variance was fixed, thus not allowing for individual variance for cubic terms. The residual variance for BMI at 5 years was fixed to zero, as a small and not significant negative residual was identified for this variable. We used the maximum likelihood robust estimator to account for missing data by full information maximum likelihood (FIML). This process approximates missing data by estimating a likelihood function for each individual based on variables that are present, such that all the available data points are used (34). The optimal number of latent trajectories was identified based on four model-fit indices: Sample-size adjusted Bayesian information criterion (BIC), adjusted Bootstrap likelihood ratio test (BLRT), Lo-Mendell Rubin test (LMRT), entropy, and interpretability of the trajectories. A lower BIC value indicates a better model fit, while the BLRT and LMRT provide a *p*-value indicating whether a model with one less trajectory group (k-1 model) should be rejected in favor of a model with k trajectories (35). Entropy is a statistic that ranges from 0 to 1 with high values (>0.8)indicating that individuals are classified with confidence (14). Distinct trajectories were coded as a categorical variable (with k number of categories) and were named based on their visual appearance. The selected model was reproduced in children with no missing BMI data (n = 915).

Associations between factors identified in pregnancy and at delivery (maternal age, education level, income, marital status, smoking and alcohol use during pregnancy, BMI at 15 weeks, child sex, and cord blood leptin, IGF-1, and IGF-2) and BMI trajectories were examined using multivariate multinomial logistic regression, with the most commonly occurring trajectory chosen as the reference category. GDM and preterm delivery were not included due to their low prevalence in the study. **Supplementary Figure 1** includes a flow chart for the number of children included at each stage of the study. Analysis was conducted using Mplus version 8 (36) and Stata version 14 (37). The GroLTS (Guidelines for Reporting on Latent Trajectory Studies) Checklist was used as a guide for completing analysis and manuscript preparation (38).

RESULTS

BMI Trajectories in the First 5 Years

Table 1 describes the distribution of demographic and clinical characteristics for the whole sample as well as according to BMI trajectories. Based on the fit indices in Table 2 the three-class model was selected. Sample-size adjusted BIC and BLRT indicated a better fit for a four trajectory model, however LMRT indicated that a four trajectory model was not significantly superior to the three trajectory one. Furthermore, the four trajectory model identified two very small subgroups (2.0 and 1.3%), which may reduce interpretability. Supplementary Table 2 outlines the average BMI and standard deviation (SD) at each time point by sex and trajectory membership. The majority of the children (89.6%, n = 1,947) exhibited a BMI trajectory corresponding to the 50th to 75th percentile according to WHO growth standards for both boys and girls at each time point (39). Secondly, 7.8% (n = 169) of children had a higher BMI at birth compared to the other two groups (75th percentile), which increased to the 95th percentile in the first 6 months, remained at 95th percentile between six and 24 months, and decreased to the 85th percentile by 5 years. Finally, 2.6% (n = 56) had a BMI at the 60th percentile at birth, a BMI between the 75th and 85th percentile between two and 12 months, followed by the 95th percentile at 2 years and 99th percentile at 5 years (Figure 1). Sensitivity analyses using subjects with no missing BMI data in the first 5 years (n = 915) showed similar trajectory patterns.

Determinants of BMI Trajectories in the First 5 Years

Sex and maternal age were associated with membership of class 2, the trajectory characterized by early rapid gain. Girls were less likely than boys to belong to this trajectory [Relative risk ratio (RRR) (95% Confidence Interval (CI)): 0.17 (0.06–0.49)], as were children of older mothers [RRR (95% CI): 0.88 (0.79–0.99) for a maternal age increase of 1 year]. Maternal BMI and cord blood IGF-2 were associated with membership of class 3, the rapid gain after 12 months group. Children of mothers with higher BMI were more likely to belong to this trajectory [RRR (95% CI): 2.27 (1.15–4.48) for an increase of 1 kg/m² in maternal BMI at 15 weeks]. Further, an increase of 1SD of cord blood IGF-2 was found to be inversely associated with membership [RRR (95% CI): 0.40 (0.16–0.98)]. **Table 3** provides a full outline of the multivariate multinomial logistic regression analysis.

TABLE 1 | Distribution of demographic and clinical characteristics according to BMI trajectory.

Average age \pm standard deviation (SD), years ($r = 1,572$) 30.0 ± 4.4 30.0 ± 4.4 30.0 ± 4.5 29.7 ± 4.9 0.88 n (%) Ethnicity ($n = 1,950$)		All trajectories	Class 1	Class 2	Class 3	P-value ^a
$\begin{array}{ c c c c c c c c c c c c c c c c c c c$	Average age \pm standard deviation (SD), years ($n = 1,572$)	30.0 ± 4.4	30.0 ± 4.4	30.0 ± 4.5	29.7 ± 4.9	0.88
$\begin{array}{c c c c c c c c c c c c c c c c c c c $		<i>n</i> (%) ^b	n (%)	n (%)	n (%)	
Irish or other Caucasian 1.911 (98.0) 1.691 (98.0) 165 (97.6) 55 (98.2) Other 39 (2.0) 34 (2.0) 4 (2.4) 1 (1.8) Educational attainment (n = 1,949) V 152 (90.6) 16 (9.5) 010 (17.9) No college/university education 1,741 (89.3) 1,542 (89.4) 153 (90.5) 46 (82.1) Income level per year (n = 1,873) V 20 (38.5) 444 (26.8) 20 (38.6) €43-63 k 458 (23.3) 380 (22.9) 42 (25.6) 14 (26.9) €64+ 922 (49.2) 826 (49.9) 78 (47.6) 18 (34.6) Married/de-facto relationship 104 (5.3) 86 (5.0) 105 (94.1) 48 (85.7) Smoking during pregnancy (n = 1,572) V V V V No 1,166 (74.2) 1,050 (74.5) 57 (71.3) 92 (72.5) Yes 1278 (81.3) 1,144 (81.1) 103 (84.4) 31 (77.5) No 294 (18.7) 266 (18.9) 19 (15.6) 9 (22.5) Yes Yes 1278 (81.3) 1,144 (81.1)	Ethnicity ($n = 1,950$)					0.79
Other 39 (2.0) 34 (2.0) 4 (2.4) 1 (1.8) Educational attainment (r = 1.949)	Irish or other Caucasian	1.911 (98.0)	1,691 (98.0)	165 (97.6)	55 (98.2)	
Educational attainment (n = 1,949) 0.89 (10.7) ^h 182 (10.6) 16 (9.5) 10 (17.9) Any college/university education 1.08 (10.7) ^h 15.2 (8.9.4) 15.3 (0.5.5) 4.6 (17.9) Income level per year (n = 1,87.3) 20.8 (10.7) ^h 4.5 (15.2 (17.2) 4.4 (26.8) 2.0 (38.5) €43-63.k 515 (27.5) 4.51 (27.2) 4.4 (26.8) 2.0 (38.5) €64.h 2.02 (49.2) 2.02 (49.2) 2.02 (49.2) 1.4 (26.9) €64.h 2.02 (49.2) 2.02 (49.2) 2.02 (49.2) 2.02 (49.2) Marital status (n = 1,950) 5.00 (17.6) 8.6 (5.0) 10.0 (5.9) 8.14 (3.5) Single 1.04 (6.3) 8.6 (6.0) 10.0 (5.9) 8.16 (3.7) 2.00 (17.5) Marited/de-facto relationship 1.846 (94.7) 1.630 (95.0) 159 (94.1) 4.8 (85.7) 2.00 (17.5) No 1.666 (74.2) 1.050 (74.5) 8.7 (71.3) 2.07 (25.5) 2.00 (17.5) No 1.016 (74.2) 1.050 (74.5) 8.7 (71.3) 2.07 (25.5) 2.00 (17.5) No 1.016 (74.2)	Other	39 (2.0)	34 (2.0)	4 (2.4)	1 (1.8)	
No college/university education 208 (10.7) ^b 182 (10.6) 16 (9.5) 10 (17.9) Any college/university education 1,741 (83.3) 1,542 (89.4) 153 (90.5) 46 (82.1) Income level per year (n = 1,873) - 0.23 - - 0.23 < <e43 k<="" td=""> 515 (27.5) 44 (26.8) 20 (38.5) 14 (26.9) 18 (34.6) <e43 -="" e43="" k<="" td=""> 436 (23.3) 380 (22.9) 78 (47.6) 18 (34.6) - <e64 +<="" td=""> 922 (49.2) 826 (49.9) 78 (47.6) 18 (34.6) - Marial status (n = 1,950) - - 0.009 - - Single 104 (5.3) 86 (5.0) 10 (5.9) 8 (14.3) - Married/de-facto relationship 1046 (74.2) 1,050 (74.5) 87 (71.3) 29 (72.5) - Yes 406 (25.8) 360 (25.5) 35 (28.7) 11 (27.5) - - No 1,278 (81.3) 1,144 (81.1) 103 (84.4) 31 (77.5) - - No 2,924 (82.7) 2.</e64></e43></e43>	Educational attainment ($n = 1,949$)					0.19
Any college/university education 1,741 (89.3) 1,542 (89.4) 153 (80.5) 46 (82.1) Income level per year (n = 1,873) 515 (27.5) 451 (27.2) 44 (26.8) 20 (38.5) €43-631k 515 (27.5) 380 (22.9) 42 (25.6) 14 (26.8) 50 (27.8) €64+ 92 (49.2) 380 (22.9) 48 (47.8) 18 (34.5) 500 (27.8) Martial status (n = 1,950) 50 (46.9) 10 (6.9) 8 (14.3) 68 (14.3) 68 (15.0) 10 (5.9) 8 (14.3) Single 104 (6.3) 86 (5.0) 10 (5.9) 8 (14.3) 500 (27.5)	No college/university education	208 (10.7) ^b	182 (10.6)	16 (9.5)	10 (17.9)	
Income level per year (n = 1,873) 0.23 ≪€43 k 515 (27.5) 451 (27.2) 44 (26.8) 20 (38.5) €64- 0.20 (40.2) 826 (49.3) 380 (22.9) 42 (25.6) 14 (26.9) €64+ 0.20 (40.2) 826 (49.3) 78 (47.6) 18 (34.6) Martial status (n = 1,950) 0.009 8 (14.3) Single 104 (5.3) 86 (5.0) 10 (5.9) 8 (14.3) Martied/de-facto relationship 1.846 (94.7) 1.639 (95.0) 159 (94.1) 48 (85.7) No 1.666 (74.2) 1.050 (74.5) 87 (71.3) 2.9 (72.5) Yes 360 (25.8) 360 (25.8) 361 (73.5) 2.9 (72.5) Yes 1.666 (74.2) 1.050 (74.5) 87 (71.3) 2.9 (72.5) Yes 2.94 (18.7) 2.66 (18.9) 19 (15.6) 9 (22.5) No 2.94 (18.7) 2.66 (18.9) 19 (15.6) 9 (22.5) Yes 1.278 (81.3) 1.144 (81.1) 103 (84.4) 31 (77.5) Average bellytt at 15 weeks ± SD, kg /m² (n = 1,572) 2.8 ± 1.9 2.8 ± 1.9 2.9 ± 2.1 0.82 Qib	Any college/university education	1,741 (89.3)	1,542 (89.4)	153 (90.5)	46 (82.1)	
<€43 k	Income level per year ($n = 1,873$)					0.23
€43-63 k 436 (23.3) 380 (22.9) 42 (25.6) 14 (26.9) €64+ 922 (49.2) 826 (49.9) 78 (47.6) 18 (34.6) Marital status (n = 1,950) v 90000 81 (4.3) 86 (5.0) 10 (5.9) 8 (14.3) Marital status (n = 1,950) v 90000 81 (6.0) 10 (5.9) 8 (14.3) Marital status (n = 1,950) 10 (46.9) 10 (46.9) 86 (5.0) 10 (5.9) 8 (14.3) Marital status (n = 1,572) 10 (5.6) 87 (71.3) 29 (72.5) 700 No 1166 (74.2) 1,050 (74.5) 87 (71.3) 29 (72.5) 700 Ves 1,166 (74.2) 1,050 (74.5) 87 (71.3) 29 (72.5) 700 No 29 (25.5) 35 (28.7) 11 (27.5) 700 90 (22.5) 70.5 Ves 1,278 (81.3) 1,144 (81.1) 103 (84.4) 31 (77.5) 40002 Average weight change between 15 and 20 weeks ± SD, kg (n = 1,572) 24.9 ± 1 24.8 ± 1.9 2.9 ± 1.7 2.9 ± 2.1 0.30 No	<€43 k	515 (27.5)	451 (27.2)	44 (26.8)	20 (38.5)	
€64+ 922 (9.2) 826 (9.9) 78 (47.6) 18 (34.6) Marital status (n = 1,950)	€43-63 k	436 (23.3)	380 (22.9)	42 (25.6)	14 (26.9)	
Marital status (n = 1,950) 0.009 Single 104 (5.3) 86 (5.0) 10 (5.9) 8 (14.3) Married/de-facto relationship 1,846 (94.7) 1,639 (95.0) 159 (94.1) 48 (85.7) Somoking during pregnancy (n = 1,572) 1,166 (74.2) 1,050 (74.5) 87 (71.3) 29 (72.5) No 1,166 (74.2) 1,050 (74.5) 87 (71.3) 29 (72.5) Yes 406 (25.8) 36 (25.5) 35 (28.7) 11 (27.5) Alcohol in pregnancy (n = 1,572) 294 (18.7) 266 (18.9) 19 (15.6) 9 (22.5) Yes 1,278 (81.3) 1,144 (81.1) 103 (84.4) 31 (77.5) Average BMI at 15 weeks ± SD, kg/m ² (n = 1,572) 2.49 ± 4.1 24.8 ± 4.1 25.3 ± 4.5 2.7.7 ± 3.9 <0.001	€64+	922 (49.2)	826 (49.9)	78 (47.6)	18 (34.6)	
Single 104 (5.3) 86 (5.0) 10 (5.9) 8 (14.3) Married/de-facto relationship 1,846 (94.7) 1,639 (95.0) 159 (94.1) 48 (85.7) Smoking during pregnancy (n = 1,572) - 0.73 29 (72.5) Yes 406 (25.8) 360 (25.5) 35 (28.7) 11 (27.5) Alcohol in pregnancy (n = 1,572) - 0.55 No 294 (18.7) 266 (18.9) 19 (15.6) 9 (22.5) Yes 1,278 (81.3) 1,144 (81.1) 103 (84.4) 31 (77.5) Average BMI at 15 weeks ± SD, kg/m² (n = 1,572) 24.9 ± 4.1 24.8 ± 4.1 25.3 ± 4.5 27.7 ± 3.9 <0.001	Marital status ($n = 1,950$)					0.009
Maried/de-facto relationship 1,846 (94.7) 1,639 (95.0) 159 (94.1) 48 (85.7) Smoking during pregnancy (n = 1,572) 1,166 (74.2) 1,050 (74.5) 87 (71.3) 29 (72.5) Yes 406 (25.8) 360 (25.5) 35 (28.7) 11 (27.5) Alcohol in pregnancy (n = 1,572) 294 (18.7) 266 (18.9) 19 (15.6) 9 (22.5) Yes 1,278 (81.3) 1,144 (81.1) 103 (84.4) 31 (77.5) Average BMI at 15 weeks ± SD, kg/m² (n = 1,572) 24.9 ± 4.1 24.8 ± 4.1 25.3 ± 4.5 27.7 ± 3.9 <0.001	Single	104 (5.3)	86 (5.0)	10 (5.9)	8 (14.3)	
Smoking during pregnancy (n = 1,572) 0,166 (74.2) 1,050 (74.5) 87 (71.3) 29 (72.5) Yes 406 (25.8) 360 (25.5) 35 (28.7) 11 (27.5) Alcohol in pregnancy (n = 1,572) 294 (18.7) 266 (18.9) 19 (15.6) 9 (22.5) Yes 1,278 (81.3) 1,144 (81.1) 103 (84.4) 31 (77.5) Average BMI at 15 weeks ± SD, kg/m² (n = 1,572) 24.9 ± 4.1 24.8 ± 4.1 25.3 ± 4.5 27.7 ± 3.9 <0.001	Married/de-facto relationship	1,846 (94.7)	1,639 (95.0)	159 (94.1)	48 (85.7)	
No1,166 (74.2)1,050 (74.5)87 (71.3)29 (72.5)Yes406 (25.8)360 (25.5)35 (28.7)11 (27.5)Alcohol in pregnancy $(n = 1,572)$ 294 (18.7)266 (18.9)19 (15.6)9 (22.5)Yes1,278 (81.3)1,144 (81.1)103 (84.4)31 (77.5)Average BMI at 15 weeks \pm SD, kg/m² $(n = 1,572)$ 24.9 \pm 4.124.8 \pm 4.125.3 \pm 4.527.7 \pm 3.9<0.001	Smoking during pregnancy ($n = 1,572$)					0.73
Yes $406 (25.8)$ $360 (25.5)$ $35 (28.7)$ $11 (27.5)$ Alcohol in pregnancy $(n = 1,572)$ $294 (18.7)$ $266 (18.9)$ $19 (15.6)$ $9 (22.5)$ Yes $1,278 (81.3)$ $1,144 (81.1)$ $103 (84.4)$ $31 (77.5)$ Average BMI at 15 weeks \pm SD, kg/m² $(n = 1,572)$ 24.9 ± 4.1 24.8 ± 4.1 25.3 ± 4.5 27.7 ± 3.9 <0.001 Average weight change between 15 and 20 weeks \pm SD, kg $(n = 1,572)$ 2.8 ± 1.9 2.8 ± 1.9 2.9 ± 1.7 2.9 ± 2.1 0.82 GDM $(n = 1,572)$ T T T T 0.021 0.021 0.021 No $1,536 (97.7)$ $1,382 (98.0)$ $117 (95.9)$ $37 (92.5)$ T Yes $36 (2.3)$ $28 (2.0)$ $5 (4.1)$ $3 (7.5)$ T Male $1,095 (50.4)$ $950 (48.8)$ $121 (71.6)$ $24 (42.9)$ T Female $1,077 (49.6)$ $997 (51.2)$ $48 (28.4)$ $32 (57.1)$ Pre-term delivery (<37 weeks) $(n = 2,172)$ T T T No $2,087 (96.1)$ $1,872 (96.2)$ $164 (97.0)$ $51 (91.1)$ Yes $5(3.9)$ $75 (3.9)$ $5 (3.0)$ $5 (8.9)$	No	1,166 (74.2)	1,050 (74.5)	87 (71.3)	29 (72.5)	
$ \begin{array}{c c c c c c c c c c c c c c c c c c c $	Yes	406 (25.8)	360 (25.5)	35 (28.7)	11 (27.5)	
No $294 (18.7)$ $266 (18.9)$ $19 (15.6)$ $9 (22.5)$ Yes $1,278 (81.3)$ $1,144 (81.1)$ $103 (84.4)$ $31 (77.5)$ Average BMI at 15 weeks \pm SD, kg/m² ($n = 1,572$) 24.9 ± 4.1 24.8 ± 4.1 25.3 ± 4.5 27.7 ± 3.9 <0.001 Average weight change between 15 and 20 weeks \pm SD, kg ($n = 1,572$) 2.8 ± 1.9 2.8 ± 1.9 2.9 ± 1.7 2.9 ± 2.1 0.82 GDM ($n = 1,572$) V V V V V V V V No $1,536 (97.7)$ $1,382 (98.0)$ $117 (95.9)$ $37 (92.5)$ V Yes $36 (2.3)$ $28 (2.0)$ $5 (4.1)$ $3 (7.5)$ V Child sex ($n = 2,172$) V V V V V Male $1,095 (50.4)$ $950 (48.8)$ $121 (71.6)$ $24 (42.9)$ V Female $1,077 (49.6)$ $997 (51.2)$ $48 (28.4)$ $32 (57.1)$ 0.12 No $2,087 (96.1)$ $1,872 (96.2)$ $164 (97.0)$ $51 (91.1)$ 0.12 No $2,087 (96.1)$ $1,872 (96.2)$ $164 (97.0)$ $51 (91.1)$ Yes $85 (3.9)$ $75 (3.9)$ $5 (3.0)$ $5 (8.9)$	Alcohol in pregnancy ($n = 1,572$)					0.55
Yes $1,278 (81.3)$ $1,144 (81.1)$ $103 (84.4)$ $31 (77.5)$ Average BMI at 15 weeks \pm SD, kg/m² ($n = 1,572$) 24.9 ± 4.1 24.8 ± 4.1 25.3 ± 4.5 27.7 ± 3.9 <0.001 Average weight change between 15 and 20 weeks \pm SD, kg ($n = 1,572$) 2.8 ± 1.9 2.8 ± 1.9 2.9 ± 1.7 2.9 ± 2.1 0.82 GDM ($n = 1,572$) V V V V V V V V V No $1,536 (97.7)$ $1,382 (98.0)$ $117 (95.9)$ $37 (92.5)$ V <td>No</td> <td>294 (18.7)</td> <td>266 (18.9)</td> <td>19 (15.6)</td> <td>9 (22.5)</td> <td></td>	No	294 (18.7)	266 (18.9)	19 (15.6)	9 (22.5)	
$ \begin{array}{cccccccccccccccccccccccccccccccccccc$	Yes	1,278 (81.3)	1,144 (81.1)	103 (84.4)	31 (77.5)	
Average weight change between 15 and 20 weeks \pm SD, kg $(n = 1,572)$ 2.8 ± 1.9 2.8 ± 1.9 2.9 ± 1.7 2.9 ± 2.1 0.82 GDM $(n = 1,572)$ $$	Average BMI at 15 weeks \pm SD, kg/m ² ($n = 1,572$)	24.9 ± 4.1	24.8 ± 4.1	25.3 ± 4.5	27.7 ± 3.9	<0.001
GDM (n = 1,572) 0.03 No 1,536 (97.7) 1,382 (98.0) 117 (95.9) 37 (92.5) Yes 36 (2.3) 28 (2.0) 5 (4.1) 3 (7.5) Child sex (n = 2,172)	Average weight change between 15 and 20 weeks \pm SD, kg ($n = 1,572$)	2.8 ± 1.9	2.8 ± 1.9	2.9 ± 1.7	2.9 ± 2.1	0.82
No1,536 (97.7)1,382 (98.0)117 (95.9)37 (92.5)Yes36 (2.3)28 (2.0)5 (4.1)3 (7.5)Child sex (n = 2,172)	GDM ($n = 1,572$)					0.03
Yes 36 (2.3) 28 (2.0) 5 (4.1) 3 (7.5) Child sex (n = 2,172)	No	1,536 (97.7)	1,382 (98.0)	117 (95.9)	37 (92.5)	
Child sex (n = 2,172) Male 1,095 (50.4) 950 (48.8) 121 (71.6) 24 (42.9) Female 1,077 (49.6) 997 (51.2) 48 (28.4) 32 (57.1) Pre-term delivery (<37 weeks) (n = 2,172)	Yes	36 (2.3)	28 (2.0)	5 (4.1)	3 (7.5)	
Male 1,095 (50.4) 950 (48.8) 121 (71.6) 24 (42.9) Female 1,077 (49.6) 997 (51.2) 48 (28.4) 32 (57.1) Pre-term delivery (<37 weeks) (n = 2,172)	Child sex ($n = 2,172$)					<0.001
Female 1,077 (49.6) 997 (51.2) 48 (28.4) 32 (57.1) Pre-term delivery (<37 weeks) (n = 2,172)	Male	1,095 (50.4)	950 (48.8)	121 (71.6)	24 (42.9)	
Pre-term delivery (<37 weeks) (n = 2,172) 0.12 No 2,087 (96.1) 1,872 (96.2) 164 (97.0) 51 (91.1) Yes 85 (3.9) 75 (3.9) 5 (3.0) 5 (8.9)	Female	1,077 (49.6)	997 (51.2)	48 (28.4)	32 (57.1)	
No2,087 (96.1)1,872 (96.2)164 (97.0)51 (91.1)Yes85 (3.9)75 (3.9)5 (3.0)5 (8.9)	Pre-term delivery (<37 weeks) ($n = 2,172$)					0.12
Yes 85 (3.9) 75 (3.9) 5 (3.0) 5 (8.9)	No	2,087 (96.1)	1,872 (96.2)	164 (97.0)	51 (91.1)	
	Yes	85 (3.9)	75 (3.9)	5 (3.0)	5 (8.9)	

^aBased on one-way ANOVA (for continuous variables) or chi-square test (for categorical variables).

^bn (%); indicates column percentages.

Significant at p < 0.05 level.

DISCUSSION

We have identified three distinct BMI trajectories in the first 5 years of life: normal, rapid BMI gain in the first 6 months, and rapid gain after 12 months. The trajectories were found to be associated with several determinants identified during pregnancy and labor, including sex, maternal age, maternal BMI, and cord blood IGF-2.

The trajectory patterns observed in this study are consistent with those identified in a previously published systematic review. Similarly to previous studies, we identified a normal trajectory and trajectories characterized by rapid gain at different time points. We did not identify a consistently high or low trajectory, which have been identified by some studies in the past (22). When using BMI as a measure of growth in longitudinal studies; raw BMI values or BMI z-scores may be used. BMI z-scores are standardized and indicate a child's position relative to same-age, same-sex children (40). Thus, BMI z-score trajectories indicate how children's BMI relative to peers changes over time, while raw BMI growth trajectories describe children's BMI change over time. While both may be used, some researchers have recommended raw BMI over z-scores for use in longitudinal analyses, as the within-child variability over time depends on the child's level of adiposity and findings presented in BMI units are more interpretable (41, 42)

In our study, boys were found to be more likely to belong to class 2, the trajectory characterized by early rapid gain, while no differences were identified for class 3. Further research is warranted to explain the potential sex difference in early BMI trajectories.

TABLE 2 | Fit indices.

No. of classes	Class proportions	Sample-size adjusted BIC	LMRT for k-1 vs. k classes	BLRT for k-1 vs. k classes	Entropy
		Value	P-value	P-value	
1	n/a	33,011	n/a	n/a	n/a
2	C1: 2,115 (97.4%) C2: 57 (2.6%)	32,784	0.005	<0.001	0.84
3	C1: 169 (7.8%) C2: 1,947 (89.6%) C3: 56 (2.6%)	32,717	0.009	<0.001	0.69
4	C1: 1,895 (87.2%) C2: 44 (2.0%) C3: 205 (9.4%) C4: 28 (1.3%)	32,681	0.288	<0.001	0.71



Adiposity, as measured by BMI, increases during the first year of life and then decreases. The adiposity rebound is the second rise in adiposity, which occurs between 3 and 7 years of age in individual children. It corresponds to fat cells starting to increase in number after an earlier phase of increasing then decreasing in size. The age the rebound occurs has been shown to predict obesity in later life, with an earlier AR (usually before 5 years of age) being associated with an increased risk of obesity (43, 44) and the timing of AR having been shown to be related to features of BMI trajectories (45). In this study early AR may be identified in class 3, as children in this class had a higher BMI at 5 years compared to 2 years of age (Supplementary Table 2 and Figure 1), indicating an AR at some point between these two measurements. As measurements of child weight and length were not made between 2 and 5 years of age, a more nuanced analysis of this time period was therefore not possible and as there was no follow-up past 5 years the timing of AR could not be identified in class 1 and 2.

The concept of fetal programming originates in a hypothesis that exposure to certain environmental influences during critical periods of development and growth may have significant consequences on an individual's long-term health. Developed by David Barker; it was originally called the Barker Hypothesis, then known as the Fetal Origin of Adult Disease, and now the Developmental Origins of Health and Disease (DOHaD) (46). According to this hypothesis, the fetus responds to a hostile uterine environment by developing adaptations that not only foster its immediate viability, but also its survival if a similar environment is encountered later in life (47). In terms of obesity development, there is a growing body of evidence that suggests that the origins of obesity and metabolic dysfunction can be traced back to the developing fetus responding to suboptimal conditions during critical periods of cellular proliferation, differentiation, and maturation by producing structural and functional changes in cells, tissues and organ systems, thus increasing the offspring's risk of developing a range of complex disorders, including obesity and metabolic

TABLE 3 Multinomia	l logistic r	egression fo	or trajectory	membership	(n =	= 402)
----------------------	--------------	--------------	---------------	------------	------	--------

	Class 2 RRR (95% Cl) <i>p</i> -value	Class 3 RRR (95% CI) <i>p</i> -value
Maternal age, y	0.88 (0.79-0.99) ^a 0.04	0.98 (0.84–1.15) 0.84
Educational attainment (vs. no third level)		
Any third level	1.15 (0.27–4.85) 0.85	1.06 (0.09–11.54) 0.96
Income level per year (vs. < €43 k)		
€43–63 k	0.56 (0.18–1.70) 0.31	3.74 (0.56–25.10) 0.18
€64k+	0.85 (0.30–2.42) 0.76	1.45 (0.21–10.25) 0.71
Marital status (vs. single)		
Married/de-facto relationship	0.92, 0.99	0.27 (0.01–5.03) 0.38
Smoking during pregnancy (vs. not smoking)	1.01 (0.40–2.57) 0.98	0.64 (0.11–3.78) 0.62
Alcohol in pregnancy (vs. no alcohol)	4.20 (0.91–19.40) 0.07	0.29 (0.06–1.28) 0.06
BMI at 15 weeks, kg/m ²	1.04 (0.66–1.65) 0.85	2.27 (1.15–4.48) 0.018
Child sex (vs. male)		
Female	0.17 (0.06–0.49) 0.001	4.06 (0.92–17.93) 0.065
Cord blood leptin (per 1SD increase)	1.08 (0.71–1.68) 0.73	1.13 (0.60–2.13) 0.71
Cord blood IGF-1 (per 1SD increase)	1.09 (0.71–1.68) 0.70	0.53 (0.23–1.22) 0.14
Cord blood IGF-2 (per 1SD increase)	1.22 (0.84–1.79) 0.29	0.40 (0.16–0.98) 0.046

^aReference category: class 1. Significant at p < 0.05 level.

dysfunction (48-50). In addition to fetal programming, the risk of childhood obesity and metabolic dysfunction may be further increased by a multitude of life course exposures, including SES, food production and marketing, and obesogenic environments. The effect of the exposure of these complex factors may be intergenerational; thus if mothers were exposed, contributing to their own obesity development, then their children are at higher of being exposed to all or some of the same factors (51, 52). High maternal pre-pregnancy BMI has previously been identified as the most frequently identified risk factor for membership of a rapid gain trajectory (22). In this study, BMI at 15 weeks gestation was associated with the trajectory characterized by early stable-high BMI followed by later rapid gain. How maternal BMI influences the BMI trajectory of the offspring is still not well-understood, however previous studies have found that overnutrition in utero may lead to high-risk BMI trajectories during early childhood, and that this may occur through increased transfer of maternal energy substrates, such as glucose, lipids and amino acids to the fetus (53). Additionally, mothers with overweight or obesity may be more likely to experience placental dysfunction (54) and may be at higher risk of micronutrient deficiency (55). In this study we did not identify any associations between BMI trajectories or socioeconomic status, including maternal education, income, or marital status.

To our knowledge, this is the first study to examine latent BMI trajectories in childhood in association with cord blood leptin, IGF-1 and -2 levels. While no associations were identified for leptin or IGF-1, a higher IGF-2 concentration in the cord blood was found to be negatively associated with class 3, i.e., those experiencing rapid gain between 1 and 5 years. IGF-2 has been previously linked to intrauterine programming predisposing to cardiovascular risk in postnatal life (56). The "Dutch hunger winter" studies of the period of famine induced by the Germanimposed food embargo in the western part of The Netherlands toward the end of World War II in the winter of 1944-1945 has provided support for the DOHaD concept, due to its unique nature. Although the food embargo was immediately lifted after liberation in May 1945, and children exposed to famine in utero during the hunger winter were well-nourished in childhood and had accelerated weight gain, decades later, they still experienced a higher incidence of cardiovascular disease (57). A cohort of individuals prenatally exposed to the Dutch Hunger Winter were tested six decades later, with results showing that the periconceptional exposure to famine was associated with reduced DNA methylation of the imprinted IGF-2 gene (58). The degree of IGF-2 methylation at birth has previously been linked to the development of childhood overweight and obesity (59). Cord blood IGF-2 concentrations have previously been found to be related to IGF-2 levels at age five, which in turn have been related to fat mass (24), however the relationship between cord blood IGF and growth in childhood remains largely unexplored and further research is warranted.

Strengths and Limitations

This study has several strengths. The large sample size and prospective design allowed the collection of repeated weight and height measurements and the evaluation of longitudinal childhood BMI trajectories. All weight and height data were obtained by trained staff using standardized instruments using strict Standard Operating Procedures. As the sample was recruited from the only maternity hospital in Cork, Ireland it included mothers and children from a broad range of social circumstances. Mean maternal age (30.9 years) was comparable with that reported nationally for the same year (31.5) (60) and the high prevalence of Caucasian women (98%) mirrors the Irish Census of 2006 (95%) (61). Income was normally distributed across the categories, with 45% of the participants reporting household incomes between 43,000 and 84,000 per year, and the average household income in Ireland in 2011 53,000 per year.

However, limitations remain. The comparative strength and limitations of GMM over other alternative approaches has been the subject of some debate (62). A more straightforward and commonly used method is following a child's BMI or z-score over time, with excessive growth indicated by crossing major percentile lines on a standard growth chart (63). Compared to the relatively complex and computer-intensive nature of GMM, this method is simple to implement. Furthermore, the assignment of children to a distinct developmental pattern is based on their highest estimated group-membership probability to the identified pattern. Thus, these latent patterns should not be considered as the actual developmental patterns but, rather, as approximations of more complex ones. Consequently, the findings will reflect associations of determinants with a model-derived class variable based on modeled BMI patterns. While GMM approaches remove the constraint on withinclass variation in the LCGA method, disadvantages with such approaches have been highlighted with respect to interpretability and the identification of non-existent subclasses (62). However, while choosing the correct model and number of classes in GMM is not straightforward, the crossing of percentiles method assumes growth to be a linear function of size at different ages. Conversely, GMM is capable of modeling non-linear growth curves, estimating individual trajectories and identifying distinctive subgroups in the population, and we therefore believe this method provides an important dimension for consideration and that these advantages may outweigh its complexity.

Lastly, in terms of the study sample and data some limitations must be noted. Firstly, some variables of interest could not be included due to the lack of availability, including parity and gestational weight gain. While the study sample was comparable to the general population in several aspects as detailed above, the women participating in the study were better educated than the general population, with 89% having obtained post-secondary school education, compared to 45% in the 2006 census (61). Further, the issue of missing data must be acknowledged. Of the 2,172 children with data collected after birth, 1,135 (52%) were followed up at 5 years. However, as FIML was used to account for missing data, all children were included in the trajectory analysis. Sensitivity analyses using subjects with no missing BMI data in the first 5 years (n = 915) showed similar trajectory patterns, thus illustrating the robustness of the extracted BMI trajectories. For the multinomial regression analysis, only 402 children were included in the final model. The low number is a result of leptin only being collected in a subset of children (n = 405), and the missing data may therefore be deemed missing at random. This may be further demonstrated in Supplementary Table 3, which shows population characteristics for participants with and without leptin data. Further research with larger numbers of children is required.

CONCLUSION

We identified three distinct BMI trajectories in the first 5 years of life which were associated with maternal age and

REFERENCES

- Berrington de Gonzalez A, Hartge P, Cerhan JR, Flint AJ, Hannan L, MacInnis RJ, et al. Body-mass index and mortality among 1.46 million white adults. N Engl J Med. (2010) 363:2211–9. doi: 10.1056/NEJMoa10 00367
- 2. Emerging Risk Factors C, Wormser D, Kaptoge S, Di Angelantonio E, Wood AM, Pennells L, et al. Separate and combined associations of

BMI, sex, and cord blood concentration of IGF-2. The potential public health and clinical implications of these findings are important. First, identification of BMI trajectories in early childhood may be helpful in identifying high-risk groups. Second, the assessment of prenatal determinants of BMI trajectories may enhance the understanding of etiologic pathways of childhood obesity and potentially target interventions. Further research and extended followup to examine the effects of early childhood growth patterns are required to understand this complex public health issue.

DATA AVAILABILITY STATEMENT

The datasets presented in this article are not readily available and access to the data must be applied for through the principal investigators. Requests to access the datasets should be directed to d.murray@ucc.ie.

ETHICS STATEMENT

The studies involving human participants were reviewed and approved by Clinical Research Ethics Committee of the Cork Teaching Hospitals. Written informed consent to participate in this study was provided by the participants' legal guardian/next of kin.

AUTHOR CONTRIBUTIONS

MM was involved in the study conception and design, requested the data sets, performed statistical analysis of the data, and drafted the manuscript. RB and DM were involved in the study conception and design and critically reviewed the manuscript. FB was involved in the study conception and design, critically reviewed the manuscript, and provided statistical support. FM and CH critically reviewed the draft manuscript. MK and CN analyzed and supplied the leptin data. All authors approved the final manuscript as submitted and agree to be accountable for all aspects of the work.

SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fped. 2021.622381/full#supplementary-material

body-mass index and abdominal adiposity with cardiovascular disease: collaborative analysis of 58 prospective studies. *Lancet*. (2011) 377:1085–95. doi: 10.1016/S0140-6736(11)60105-0

 Lu Y, Hajifathalian K, Ezzati M, Woodward M, Rimm EB, Danaei G. Metabolic mediators of the effects of body-mass index, overweight, and obesity on coronary heart disease and stroke: a pooled analysis of 97 prospective cohorts with 1.8 million participants. *Lancet.* (2014) 383:970–83. doi: 10.1016/s0140-6736(13)61836-x

- Ni Mhurchu C, Rodgers A, Pan WH, Gu DF, Woodward M, Asia Pacific Cohort Studies C. Body mass index and cardiovascular disease in the Asia-Pacific region: an overview of 33 cohorts involving 310 000 participants. *Int J Epidemiol.* (2004) 33:751–8. doi: 10.1093/ije/dyh163
- Whitlock G, Lewington S, Sherliker P, Clarke R, Emberson J, Halsey J, et al. Body-mass index and cause-specific mortality in 900 000 adults: collaborative analyses of 57 prospective studies. *Lancet. Mar 28.* (2009) 373:1083–96. doi: 10.1016/s0140-6736(09)60318-4
- Zheng W, McLerran DF, Rolland B, Zhang X, Inoue M, Matsuo K, et al. Association between body-mass index and risk of death in more than 1 million Asians. N Engl J Med. (2011) 364:719–29. doi: 10.1056/NEJMoa1010679
- NCD Risk Factor Collaboration. Trends in adult body-mass index in 200 countries from 1975 to 2014: a pooled analysis of 1698 population-based measurement studies with 19.2 million participants. *Lancet.* (2016) 387:1377– 1396. doi: 10.1016/S0140-6736(16)30054-X
- Simmonds M, Llewellyn A, Owen CG, Woolacott N. Predicting adult obesity from childhood obesity: a systematic review and meta-analysis. *Obes Rev.* (2016) 17:95–107. doi: 10.1111/obr.12334
- Ziyab AH, Karmaus W, Kurukulaaratchy RJ, Zhang H, Arshad SH. Developmental trajectories of Body Mass Index from infancy to 18 years of age: prenatal determinants and health consequences. *J Epidemiology Commun Health*. (2014) 68:934–41. doi: 10.1136/jech-2014-203808
- Cole TJ, Lobstein T. Extended international (IOTF) body mass index cutoffs for thinness, overweight and obesity. *Pediatr Obes.* (2012) 7:284–94. doi: 10.1111/j.2047-6310.2012.00064.x
- Daniels SR, Hassink SG. The role of the pediatrician in primary prevention of obesity. *Pediatrics*. (2015) 136:e275–92. doi: 10.1542/peds.2015-1558
- Roy SM, Spivack JG, Faith MS, Chesi A, Mitchell JA, Kelly A, et al. Infant BMI or Weight-For-Length And Obesity Risk In Early Childhood. *Pediatrics*. (2016) 137:e20153492. doi: 10.1542/peds.2015-3492
- Roy SM, Fields DA, Mitchell JA, Hawkes CP, Kelly A, Wu GD, et al. Body mass index is a better indicator of body composition than weight-for-length at age 1 month. J Pediatr. (2019) 204:77–83.e1. doi: 10.1016/j.jpeds.2018.08.007
- Ram N, Grimm KJ. Growth mixture modeling: a method for identifying differences in longitudinal change among unobserved groups. *Int J Behav Dev.* (2009) 33:565–576. doi: 10.1177/0165025409343765
- Jung T, Wickrama KAS. An introduction to latent class growth analysis and growth mixture modeling. Sock Person Psychol Com. (2008) 2:302–17. doi: 10.1111/j.1751-9004.2007.00054.x
- Davison KK, Birch LL. Childhood overweight: a contextual model and recommendations for future research. *Obes Rev.* (2001) 2:159–71. doi: 10.1046/j.1467-789x.2001.00036.x
- Hanson MA, Gluckman PD. Early developmental conditioning of later health and disease: physiology or pathophysiology? *Physiol Rev.* (2014) 94:1027–76. doi: 10.1152/physrev.00029.2013
- Godoy GA, Korevaar TI, Peeters RP, Hofman A, de Rijke YB, Bongers-Schokking JJ, et al. Maternal thyroid hormones during pregnancy, childhood adiposity and cardiovascular risk factors: the generation R study. *Clin Endocrinol.* (2014) 81:117–25. doi: 10.1111/cen.12399
- Heude B, Thiebaugeorges O, Goua V, Forhan A, Kaminski M, Foliguet B, et al. Pre-pregnancy body mass index and weight gain during pregnancy: relations with gestational diabetes and hypertension, and birth outcomes. *Matern Child Health J.* (2012). 16:355–63. doi: 10.1007/s10995-011-0741-9
- Pizzi C, Cole TJ, Richiardi L, dos-Santos-Silva I, Corvalan C, De Stavola B. Prenatal influences on size, velocity and tempo of infant growth: findings from three contemporary cohorts. *PLoS ONE.* (2014) 9:e90291. doi: 10.1371/journal.pone.0090291
- Woo Baidal JA, Locks LM, Cheng ER, Blake-Lamb TL, Perkins ME, Taveras EM. Risk factors for childhood obesity in the first 1,000 days: a systematic review. *Am J Prevent Med.* (2016) 50:761–79. doi: 10.1016/j.amepre.2015.11.012
- Mattsson M, Maher GM, Boland F, Fitzgerald AP, Murray DM, Biesma R. Group-based trajectory modelling for BMI trajectories in childhood: a systematic review. *Obes Rev.* (2019) 20:998–1015. doi: 10.1111/obr.12842
- Rohrmann S, Sutcliffe CG, Bienstock JL, Monsegue D, Akereyeni F, Bradwin G, et al. Racial variation in sex steroid hormones and the insulin-like growth factor axis in umbilical cord blood of male neonates. *Cancer Epidemiol Biomark Prevent.* (2009) 18:1484–91. doi: 10.1158/1055-9965.epi-08-0817

- Ong K, Kratzsch J, Kiess W, Dunger D. Circulating IGF-I levels in childhood are related to both current body composition and early postnatal growth rate. *J Clin Endocrinol Metab.* (2002) 87:1041–4. doi: 10.1210/jcem.87.3.8342
- Patel N, Hellmuth C, Uhl O, Godfrey K, Briley A, Welsh P, et al. Cord metabolic profiles in obese pregnant women: insights into offspring growth and body composition. J Clin Endocrinol Metab. (2018) 103:346–55. doi: 10.1210/jc.2017-00876
- Street ME, Smerieri A, Montanini L, Predieri B, Iughetti L, Valenzise M, et al. Interactions among pro-inflammatory cytokines, IGF system and thyroid function in pre-pubertal obese subjects. *J biol Regul Homeost Agents*. (2013) 27:259–66.
- Alexe DM, Syridou G, Petridou ET. Determinants of early life leptin levels and later life degenerative outcomes. *Clin Med Res.* (2006) 4:326–35. doi: 10.3121/cmr.4.4.326
- Ong KK, Ahmed ML, Sherriff A, Woods KA, Watts A, Golding J, et al. Cord blood leptin is associated with size at birth and predicts infancy weight gain in humans. *J Clin Endocrinol Metab.* (1999) 84:1145–8. doi: 10.1210/jcem.84.3.5657
- Parker M, Rifas-Shiman SL, Belfort MB, Taveras EM, Oken E, Mantzoros C, et al. Gestational glucose tolerance and cord blood leptin levels predict slower weight gain in early infancy. *J Pediatr.* (2011) 158:227–33. doi: 10.1016/j.jpeds.2010.07.052
- 30. Taine M, Khalfallah O, Forhan A, Glaichenhaus N, Charles MA, Heude B. Does cord blood leptin level mediate the association between neonatal body size and postnatal growth? Results from the EDEN mother-child cohort study. *Ann Hum Biol.* (2020) 47:159–65. doi: 10.1080/03014460.2020.1748712
- O'Donovan SM, Murray DM, Hourihane JO, Kenny LC, Irvine AD, Kiely M. Cohort profile: the cork BASELINE birth cohort study: babies after SCOPE: evaluating the longitudinal impact on neurological and nutritional endpoints. *Int J Epidemiol.* (2015) 44:764–75. doi: 10.1093/ije/dyu157
- North RA, McCowan LME, Dekker GA, Poston L, Chan EHY, Stewart AW, et al. Clinical risk prediction for pre-eclampsia in nulliparous women: development of model in international prospective cohort. *BMJ.* (2011) 342:d1875. doi: 10.1136/bmj.d1875
- Chaoimh CN, Murray DM, Kenny LC, Irvine AD, Hourihane JO, Kiely M. Cord blood leptin and gains in body weight and fat mass during infancy. *Eur J Endocrinol.* (2016) 175:403–10. doi: 10.1530/eje-16-0431
- Little TD, Jorgensen TD, Lang KM, Moore EW. On the joys of missing data. J Pediatr Psychol. (2014) 39:151–62. doi: 10.1093/jpepsy/jst048
- Tein JY, Coxe S, Cham H. Statistical power to detect the correct number of classes in latent profile analysis. *Struct Equ Model.* (2013) 20:640–57. doi: 10.1080/10705511.2013.824781
- Muthén LK, Muthén BO. Mplus User's Guide. Eighth Edition. Los Angeles, CA: Muthén & Muthén (1998–2017).
- 37. StataCorp. Stata Statistical Software: Release 14. College Station, TX: StataCorp LP (2015).
- van de Schoot R, Sijbrandij M, Winter SD, Depaoli S, Vermunt JK. The GRoLTS-checklist: guidelines for reporting on latent trajectory studies. *Struct Equ Model Multidiscipl J.* (2017) 24:451–67. doi: 10.1080/10705511.2016.1247646
- 39. Who Multicentre Growth Reference Study Group, de Onis M. WHO child growth standards based on length/height, weight and age. Acta Paediatr. (2006) 95:76–85. doi: 10.1111/j.1651-2227.2006. tb02378.x
- 40. Kuczmarski RJ, Ogden CL, Guo SS, Grummer-Strawn LM, Flegal KM, Mei Z, et al. (2000). CDC growth charts for the united states: methods and development. In: *Vital and Health Statistics Series 11, Data From the National Health Survey*. Hyattsville: National Center for Health Statistics. (2002).
- Berkey CS, Colditz GA. Adiposity in adolescents: change in actual BMI works better than change in BMI z score for longitudinal studies. *Ann Epidemiol.* (2007) 17:44–50. doi: 10.1016/j.annepidem.2006.07.014
- Cole TJ, Faith MS, Pietrobelli A, Heo M. What is the best measure of adiposity change in growing children: BMI, BMI %, BMI z-score or BMI centile? *Eur J Clin Nutr.* (2005) 59:419–25. doi: 10.1038/sj.ejcn.1602090
- Rolland-Cachera MF, Deheeger M, Bellisle F, Sempe M, Guilloud-Bataille M, Patois E. Adiposity rebound in children: a simple indicator for predicting obesity. *Am J Clin Nutr.* (1984) 39:129–35. doi: 10.1093/ajcn/39. 1.129

- Whitaker RC, Pepe MS, Wright JA, Seidel KD, Dietz WH. Early adiposity rebound and the risk of adult obesity. *Pediatrics*. (1998) 101:E5.
- 45. Wen X, Kleinman K, Gillman MW, Rifas-Shiman SL, Taveras EM. Childhood body mass index trajectories: modeling, characterizing, pairwise correlations and socio-demographic predictors of trajectory characteristics. *BMC Med Res Methodol.* (2012) 12:38. doi: 10.1186/1471-2288-12-38
- Barker DJ. The origins of the developmental origins theory. J Intern Med. (2007) 261:412–7. doi: 10.1111/j.1365-2796.2007.01809.x
- Gluckman PD, Hanson MA, Spencer HG. Predictive adaptive responses and human evolution. *Trends Ecol Evol.* (2005) 20:527–33. doi: 10.1016/j.tree.2005.08.001
- Godfrey KM, Reynolds RM, Prescott SL, Nyirenda M, Jaddoe VWV, Eriksson JG, et al. Influence of maternal obesity on the longterm health of offspring. *Lancet Diabet Endocrinol.* (2017) 5:53–64. doi: 10.1016/S2213-8587(16)30107-3
- Entringer S, Buss C, Swanson JM, Cooper DM, Wing DA, Waffarn F, et al. Fetal programming of body composition, obesity, and metabolic function: the role of intrauterine stress and stress biology. *J Nutr Metab.* (2012) 2012:632548. doi: 10.1155/2012/632548
- Gluckman PD, Hanson MA. Living with the past: evolution, development, and patterns of disease. Science. (2004) 305:1733–6. doi: 10.1126/science.1095292
- Llewellyn CH. Genetic susceptibility to the "obesogenic" environment: the role of eating behavior in obesity and an appetite for change. *Am J Clin Nutr.* (2018) 108:429–430. doi: 10.1093/ajcn/nqy210
- Bush NR, Allison AL, Miller AL, Deardorff J, Adler NE, Boyce WT. Socioeconomic disparities in childhood obesity risk: association with an oxytocin receptor polymorphism. *JAMA Pediatr.* (2017) 171:61–7. doi: 10.1001/jamapediatrics.2016.2332
- Aris IM, Bernard JY, Chen L-W, Tint MT, Pang WW, Lim WY, et al. Infant body mass index peak and early childhood cardio-metabolic risk markers in a multi-ethnic Asian birth cohort. *Int J Epidemiol.* (2017) 46:513–25. doi: 10.1093/ije/dyw232
- Myatt L, Maloyan A. Obesity and placental function. Semin Reprod Med. (2016) 34:42–9. doi: 10.1055/s-0035-1570027
- Scholing JM, Olthof MR, Jonker FAM, Vrijkotte TGM. Association between pre-pregnancy weight status and maternal micronutrient status in early pregnancy. *Public Health Nutr.* (2018) 21:2046–55. doi: 10.1017/S1368980018000459

- Cianfarani S. Insulin-like growth factor-II: new roles for an old actor. Front Endocrinol. (2012) 3:118. doi: 10.3389/fendo.2012.00118
- 57. Painter RC, de Rooij SR, Bossuyt PM, Simmers TA, Osmond C, Barker DJ, et al. Early onset of coronary artery disease after prenatal exposure to the dutch famine. *Am J Clin Nutr.* (2006) 84:322–7; quiz 466–7. doi: 10.1093/ajcn/84.1.322
- Heijmans BT, Tobi EW, Stein AD, Putter H, Blauw GJ, Susser ES, et al. Persistent epigenetic differences associated with prenatal exposure to famine in humans. *Proc Natl Acad Sci. USA*. (2008) 105:17046–9. doi: 10.1073/pnas.0806560105
- Perkins E, Murphy SK, Murtha AP, Schildkraut J, Jirtle RL, Demark-Wahnefried W, et al. Insulin-like growth factor 2/H19 methylation at birth and risk of overweight and obesity in children. J Pediatr. (2012) 161:31–9. doi: 10.1016/j.jpeds.2012.01.015
- 60. Perinatal Statistics Report. (2010). *ESRI Survey and Statistical Report Series* 41.26.06. Dublin: Economic and Social Research Institute (2012).
- 61. Census 2006. Volume 10: Education and Qualifications. Dublin: Stationery Office (2006)
- Tu YK, Tilling K, Sterne JA, Gilthorpe MS. A critical evaluation of statistical approaches to examining the role of growth trajectories in the developmental origins of health and disease. *Int J Epidemiol.* (2013) 42:1327– 39. doi: 10.1093/ije/dyt157
- Taveras EM, Rifas-Shiman SL, Sherry B, Oken E, Haines J, Kleinman K, et al. Crossing growth percentiles in infancy and risk of obesity in childhood. *Arch Pediatr Adolesc Med.* (2011) 165:993–8. doi: 10.1001/archpediatrics. 2011.167

Conflict of Interest: The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

Copyright © 2021 Mattsson, Murray, Hawkes, Kiely, Ní Chaoimh, McCarthy, Biesma and Boland. This is an open-access article distributed under the terms of the Creative Commons Attribution License (CC BY). The use, distribution or reproduction in other forums is permitted, provided the original author(s) and the copyright owner(s) are credited and that the original publication in this journal is cited, in accordance with accepted academic practice. No use, distribution or reproduction is permitted which does not comply with these terms.