



Prognostic utility of pretreatment neutrophil-lymphocyte ratio in survival outcomes in localized non-small cell lung cancer patients treated with stereotactic body radiotherapy: Selection of an ideal clinical cutoff point



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ABSTRACT

Background and purpose: Neutrophil-lymphocyte ratio (NLR) has been associated with overall survival (OS) in non-small cell lung cancer (NSCLC). We aimed to assess the utility of NLR as a predictor of lung cancer-specific survival (LCS) and identify an optimal, pretreatment cutoff point in patients with localized NSCLC treated with stereotactic body radiotherapy (SBRT) within the Veterans Affairs' (VA) national database.

Materials and methods: In the VA database, we identified patients with biopsy-proven, clinical stage I NSCLC treated with SBRT between 2006 and 2015. Cutoff points for NLR were calculated using Contal/O'Quigley's and Cox Wald methods. Primary outcomes of OS, LCS, and non-lung cancer survival (NCS) were evaluated in Cox and Fine-Gray models.

Results: In 389 patients, optimal NLR cutoff was identified as 4.0. In multivariable models, NLR > 4.0 was associated with decreased OS (HR 1.44, $p = 0.01$) and NCS (HR 1.68, $p = 0.01$) but not with LCS (HR 1.32, $p = 0.09$). In a subset analysis of 229 patients with pulmonary function tests, NLR > 4.0 remained associated with worse OS (HR 1.51, $p = 0.02$) and NCS (HR 2.18, $p = 0.01$) while the association with LCS decreased further (HR 1.22, $p = 0.39$).

Conclusion: NLR was associated with worse OS in patients with localized NSCLC treated with SBRT; however, NLR was only associated with NCS and not with LCS. Pretreatment NLR, with a cutoff of 4.0, offers potential as a marker of competing mortality risk which can aid in risk stratification in this typically frail and comorbid population. Further studies are needed to validate pretreatment NLR as a clinical tool in this setting.

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Abbreviations: NLR, Neutrophil-lymphocyte ratio; NSCLC, non-small cell lung cancer; SBRT, stereotactic body radiotherapy; VA, Veterans Affairs; OS, overall survival; LCS, lung cancer-specific survival; NCS, non-lung cancer survival.

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1. Introduction

Non-small cell lung cancer (NSCLC) continues to be a leading etiology of cancer-related deaths in the United States [1–2]. Approximately 20–30% of NSCLC patients present with localized disease at diagnosis [3]. Although the long-standing, primary treatment for early stage lung cancer is surgical resection, a majority of patients are deemed ineligible for surgery owing to an amalgamation of poor baseline medical fitness (e.g. age, performance status) and concomitant comorbid conditions (e.g. COPD, heart disease) [4–6].

For patients that are medically inoperable or unwilling to undergo surgery, stereotactic body radiotherapy (SBRT) is an effective treatment option [7–10]. However, distant recurrence is common (up to 33%) after SBRT [6,11]. Therefore, identification of patients at high risk for lung cancer recurrence and death can potentially enable targeted treatment intensification with other modalities. Furthermore, patient selection remains important as SBRT is typically well tolerated but not completely devoid of risks [5]. Many patients with localized NSCLC may have alternative drivers for their demise, and thus treatment may not necessarily improve outcomes. Thus there are multiple motivations to identify additional prognostic markers to guide risk stratification and decision management in this population. Although there are many established patient and disease related prognostic markers for NSCLC, the SBRT patient population is primarily comprised of those who are medically inoperable and therefore have similarly poor factors across the board, and therefore prognostic markers in this population are lacking [4,12–13].

Chronic inflammation in the lungs is an established precursor state to NSCLC and other malignancies [14–15]. One metric of cellular inflammation, readily obtained from routine labs indicated in the workup of numerous cancer types, is the neutrophil-lymphocyte ratio (NLR). However, previous studies have demonstrated that an elevated NLR is associated with poor outcomes in both non-oncologic, inflammatory conditions (e.g. coronary artery disease, end stage renal disease) and oncologic conditions (e.g. pancreatic cancer, mesothelioma) [16–18]. In the pulmonary realm, NLR has also been associated with lung function decline and inferior COPD outcomes [19].

Previous studies have demonstrated an association between NLR and overall survival (OS) in patients with localized NSCLC treated with SBRT [20–22]. However, they did not delineate the association of NLR with lung cancer-specific survival (LCS) and non-lung cancer survival (NCS); it remains unknown if NLR is predictive of oncologic (lung cancer risk) or non-oncologic (underlying lung disease, competing mortality risk) outcomes in this setting. Furthermore, these previous studies have been limited by small cohort sizes, lack of granular details (such as pulmonary function tests), and lack of time-dependent cutoff points. Here, we attempt to

overcome these limitations. In patients with localized NSCLC treated with SBRT within the national Veterans’ Affairs (VA) database, we identify an optimal cutoff point for pretreatment NLR and perform competing risks analyses to delineate the prognostic potential of NLR on OS, LCS, and NCS individually.

2. Methods

2.1. Data source

The VA Informatics and Computing Infrastructure (VINCI) represents a comprehensive informatics platform that enables access to the database comprised of patient-level electronic health records and administrative data throughout the national VA health care system. Tumor registry data is uploaded by trained registrars in accordance with protocols issued from the American College of Surgeons, thereby capturing an estimated 90% of incident cancers within the VA system [23–24]. Cause specific mortality (ICD-10 code C34 for lung cancer) information was obtained from the National Death Index (NDI) and then manually chart reviewed to fill in missing entries. Our protocol was approved by the San Diego VA Institutional Review Board.

2.2. Patient selection

Of the veterans with biopsy-proven NSCLC diagnosed between January 1, 2006 and December 31, 2015 who were treated definitively with radiation, only those with clinical T1 or T2a (<5 cm in greatest dimension), N0 (no regional lymph node metastasis), and M0 (no distant metastasis) stages were included. Tumor staging was denoted by the 7th American Joint Committee on Cancer guidelines. Patients were excluded if they had a history of synchronous malignancy, treatment > 6 months after diagnosis, or missing covariables of interest. The final cohort was comprised of 389 patients with localized NSCLC treated with SBRT. The following text and Fig. 1 details the complete inclusion and exclusion criteria of this study’s final cohort.

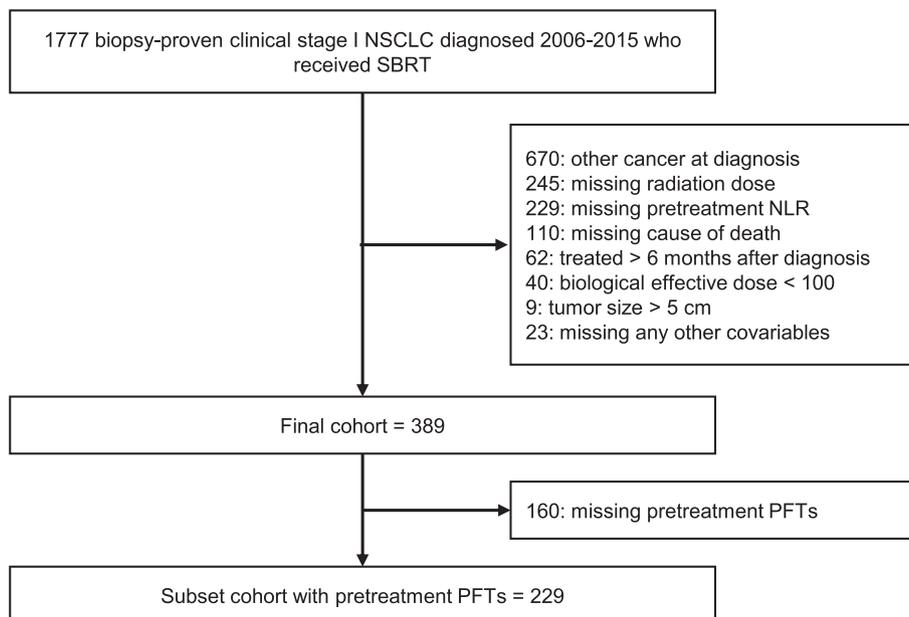


Fig. 1. Inclusion and exclusion criteria. (NSCLC = non-small cell lung cancer, SBRT = stereotactic body radiotherapy, PFTs = pulmonary function tests).

2.3. Patient, tumor, and treatment variables

SBRT was defined as delivery of one to five daily fractions of radiation directed at the lung. After identification of radiation receipt through the registry, we manually reviewed charts to extract radiation dose and fractionation information to ensure patients received definitive radiation directed at the lung. To account for the differences in dose and fractionation regimens, biologically equivalent dose was calculated for each patient with an α/β ratio of 10. Patients were excluded if they received a low biologically effective dose of radiation ($<100 \text{ Gy}_{10}$) [25].

The study covariables of interest were all derived from the VA cancer registry database. They include age at diagnosis, sex, race, tobacco history, histology, tumor size. Linked administrative data provided International Classification of Diseases–9 and 10 codes for comorbidities used to construct the Charlson comorbidity index score [26–27]. Forced expiratory volume in one second (FEV1) values were obtained by searching clinical notes within one year from the start of treatment. FEV1 values were converted to percent predicted based on patient characteristics at diagnosis using standard reference equations [28]. Laboratory data was used to identify complete blood count (CBC) with differential test results. From this, NLR is calculated as the absolute neutrophil count (ANC) divided by the absolute lymphocyte count (ALC). Pretreatment NLR was defined as any value available within six months prior to starting treatment. If multiple values were present for a patient, that patient's final pretreatment NLR was chosen to be the value closest to their treatment date. All patients were followed until death or last follow up with a VA provider before August 1, 2020.

2.4. Statistical analysis

In clinical data, transforming a continuous variable into a categorical variable to evaluate its predictive value on time dependent endpoints (such as survival) can make the model more interpretable and clinically useful. Although several techniques are commonly employed (e.g. median/quartiles, Receiver Operating Characteristic (ROC) tests), the appropriateness of such methods for time dependent endpoints have often been called into question [29–30]. Therefore we chose to employ the Contal and O'Quigley cutpoint method that uses the log-rank test statistic [29,31]. Using the methods described by Contal and O'Quigley and the SAS macro provided by Mandrekar et al., we were able to identify our continuous NLR variable to be eligible for dichotomization [31–32]. Then, the Contal and O'Quigley statistic and Cox Wald statistic were identified and compared for a panel of candidate cutoff points to identify the optimal cutoff point(s) of pretreatment NLR on each survival endpoint.

After identification of the optimal cutoff point of pretreatment NLR, the cohort was dichotomized based on this value. Patient characteristics between these two NLR cohorts were compared using Chi Square test and Wilcoxon's rank sum test for categorical and continuous variables, respectively. OS was assessed with Kaplan–Meier analysis for unadjusted models and with Cox proportional hazards analysis for multivariable models. Comparisons of LCS among groups were evaluated using a competing risk analysis framework to account for the competing risk of non-lung cancer mortality. Vice-versa logic was used to evaluate NCS. LCS and NCS were assessed with cumulative incidence analysis for unadjusted models and with Fine-Gray regression analysis for multivariable models. For all survival analysis, hazard ratios (HR) and 95% confidence intervals (CI) were reported. Throughout this study, all multivariable models incorporated variables chosen a priori – NLR batch, age at diagnosis, gender, race, Charlson score, tobacco history, histology, T stage, total biologically effective dose. All statistical analyses were performed using SAS version 9.4 (SAS

Institute, Cary, NC), with two-sided p-values < 0.05 considered statistically significant.

3. Results

3.1. Baseline characteristics and treatment exposure

The 389 patients in our final cohort were analyzed with a median post-treatment follow-up of 70 months. Median NLR of the cohort was 3.0 (range 0.4–42) obtained at a median of 41 days (range 1–174) before start of SBRT. The optimal cutoff pretreatment NLR value for our cohort was determined to be 4.0 using the Contal and O'Quigley cutpoint selection method ($p = 0.006$). This was validated to be an optimal cutoff point in the Cox Wald cutpoint selection method ($p = 0.0002$) as well. Therefore the cohort was split by this cutoff point into two batches – 275 patients with $\text{NLR} \leq 4.0$, 114 patients with $\text{NLR} > 4.0$.

Table 1 depicts the breakdown of demographic and clinical variables for the overall cohort and each NLR cohort. The NLR cohorts were broadly similar, with the only significant difference being a larger proportion of Black patients in the lower NLR cohort (17% vs 7%). The median (and range) total dose, dose per fraction, number of fractions, and BED received for the entire cohort were 50 Gy (45–75 Gy), 12 Gy (7.5–20 Gy), 5 fractions (3–8 fractions), and 112.5 Gy (100–187.5 Gy) respectively. These radiation specific variables did not differ between the NLR cohorts.

3.2. Survival

Median overall survival of the entire cohort was 31.5 months (95% CI 27.9–34.8 months). There were 304 deaths, of which 192 (63.2%) were attributed to lung cancer. The 5 year cumulative incidence of death from lung cancer and any non-lung cancer cause were 55.1% (95% CI 49.2–60.6%) and 38.2% (95% CI 31.8–44.6%), respectively. Kaplan–Meier analysis revealed a difference in OS ($p = 0.0002$) when stratified by the NLR cutoff, with median OS of 35.8 months (95% CI 31.3–43.8 months) in patients with $\text{NLR} \leq 4.0$ and 22.7 months (95% CI 17.6–27.2 months) in patients with $\text{NLR} > 4.0$. Gray's test found patients with $\text{NLR} > 4.0$ to have increased lung cancer mortality (LCM) ($p = 0.02$) and non-lung cancer mortality (NCM) ($p = 0.003$) compared to patients with $\text{NLR} \leq 4.0$. These unadjusted OS curves and cumulative incidence plots depicting LCM and NCM are seen in Fig. 2a–c.

In univariable Cox analysis, pretreatment $\text{NLR} > 4.0$ was associated with worse overall survival (HR = 1.59, 95% CI 1.24–2.04, $p = 0.0003$). In multivariable analysis, $\text{NLR} > 4.0$ continued to be an independent association with worse overall survival (HR 1.44, 95% CI 1.12–1.86, $p = 0.01$). For the other endpoints, similar multivariable analysis demonstrated $\text{NLR} > 4.0$ to be independently associated with decreased NCS (HR 1.68, 95% CI 1.11–2.53, $p = 0.01$) but only trend towards significance with LCS (HR 1.32, 95% CI 0.95–1.82, $p = 0.09$). These results are summarized in Table 2.

3.3. Subset analysis with PFT data

In a subset analysis of 229 patients with PFT data available within one year of starting SBRT, 162 patients had pretreatment $\text{NLR} \leq 4.0$ and 67 patients had pretreatment $\text{NLR} > 4.0$.

In multivariable analysis (including PFTs along with previously detailed covariables), $\text{NLR} > 4.0$ remained significantly associated with worse OS (HR 1.51, 95% CI 1.07–2.14, $p = 0.02$) and NCS (HR 2.18, 95% CI 1.24–3.84, $p = 0.01$) like in the overall cohort's models. However, NLR was not associated with LCS (HR 1.22, 95% CI 0.77–1.93, $p = 0.39$). Tables 1 and 3 detail the breakdown of PFT data in

Table 1
Baseline patient, tumor, and treatment characteristics of the overall cohort stratified by NLR ≤ 4.0 and NLR > 4.0 groups.

Variable	All Patients (N = 389)	NLR ≤ 4.0 (N = 275)	NLR > 4.0 (N = 114)	P value
Pretreatment NLR				<0.0001*
Median (range)	3.0 (0.4–42)	2.4 (0.4–4)	5.3 (4.1–42)	
Age				0.24
Median (range)	71 (49–93)	71 (49–93)	72 (53–90)	
Gender				0.99
Male	378 (97.2%)	267 (97.1%)	111 (97.4%)	
Female	11 (2.8%)	8 (2.9%)	3 (2.6%)	
Charlson Score				0.46
0	83 (21.3%)	61 (22.1%)	22 (19.3%)	
1	141 (36.2%)	103 (37.5%)	38 (33.3%)	
2+	165 (42.5%)	111 (40.4%)	54 (47.4%)	
Race				0.01*
White	335 (86.1%)	229 (83.3%)	106 (93.0%)	
Black	54 (13.9%)	46 (16.7%)	8 (7.0%)	
Smoker at Diagnosis				0.74
Yes	167 (42.9%)	120 (43.6%)	47 (41.2%)	
No	222 (57.1%)	155 (56.4%)	67 (58.8%)	
Histology				0.42
Adenocarcinoma	162 (41.6%)	120 (43.6%)	42 (36.9%)	
Squamous	156 (40.1%)	105 (38.2%)	51 (44.7%)	
Other/unknown	71 (18.3%)	50 (18.2%)	21 (18.4%)	
T stage				0.57
< 2 cm (T1a)	166 (42.7%)	118 (42.9%)	48 (42.1%)	
2–3 cm (T1b)	143 (36.8%)	104 (37.8%)	39 (34.2%)	
3–5 cm (T2a)	80 (20.5%)	53 (19.3%)	27 (23.7%)	
Total BED (Gy ₁₀)				0.96
Median (range)	112.5 (100–187.5)	112.5 (100–187.5)	112.5 (100–180)	
White Blood Cell Count (K/μL)				<0.0001*
Median (range)	8.0 (2.9–22.5)	7.5 (2.9–22.5)	9.1 (5–22.1)	
Absolute Neutrophil Count (K/μL)				<0.0001*
Median (range)	5.1 (0.8–19.6)	4.5 (0.8–12.1)	6.7 (2.6–19.6)	
Absolute Lymphocyte Count (K/μL)				<0.0001*
Median (range)	1.6 (0.1–13.5)	1.8 (0.4–13.5)	1.2 (0.1–12.5)	
Pretreatment FEV ₁				0.96
≤ 30%	(N = 229) 22 (9.6%)	(N = 162) 15 (9.3%)	(N = 67) 7 (10.5%)	
31–50%	71 (31.0%)	49 (30.2%)	22 (32.8%)	
51–80%	103 (45.0%)	74 (45.7%)	29 (43.3%)	
> 80%	33 (14.4%)	24 (14.8%)	9 (13.4%)	

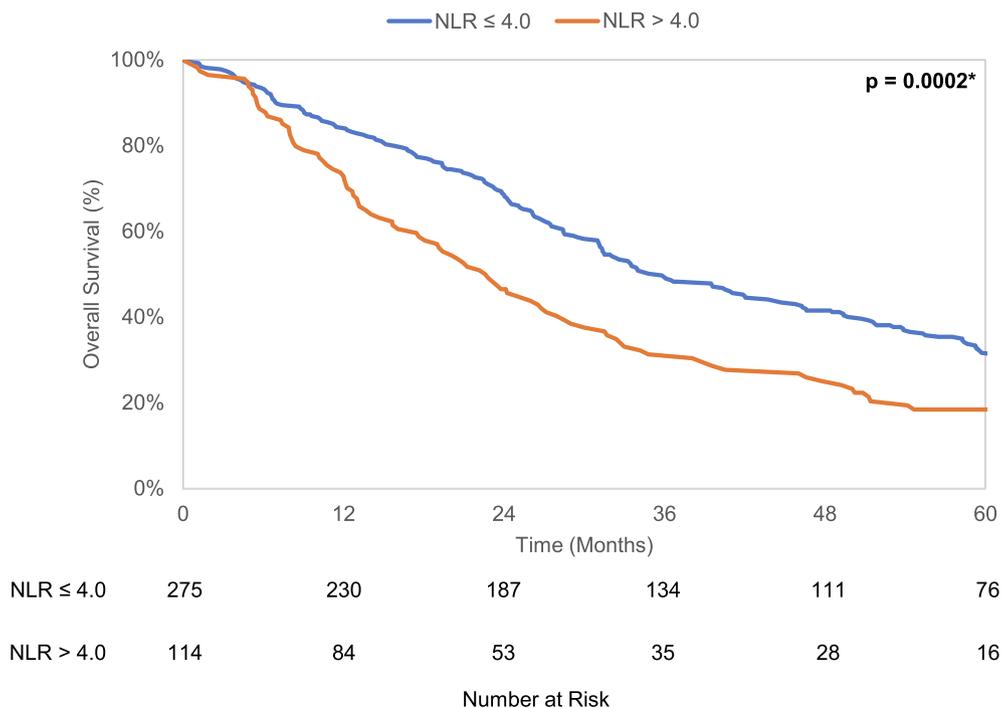


Fig. 2. a. Kaplan-Meier curves for overall survival stratified by NLR cutoff value 4.0. b. Cumulative incidence curves for lung cancer-specific mortality stratified by NLR cutoff value 4.0. c. Cumulative incidence curves for non-lung cancer mortality stratified by NLR cutoff value 4.0.

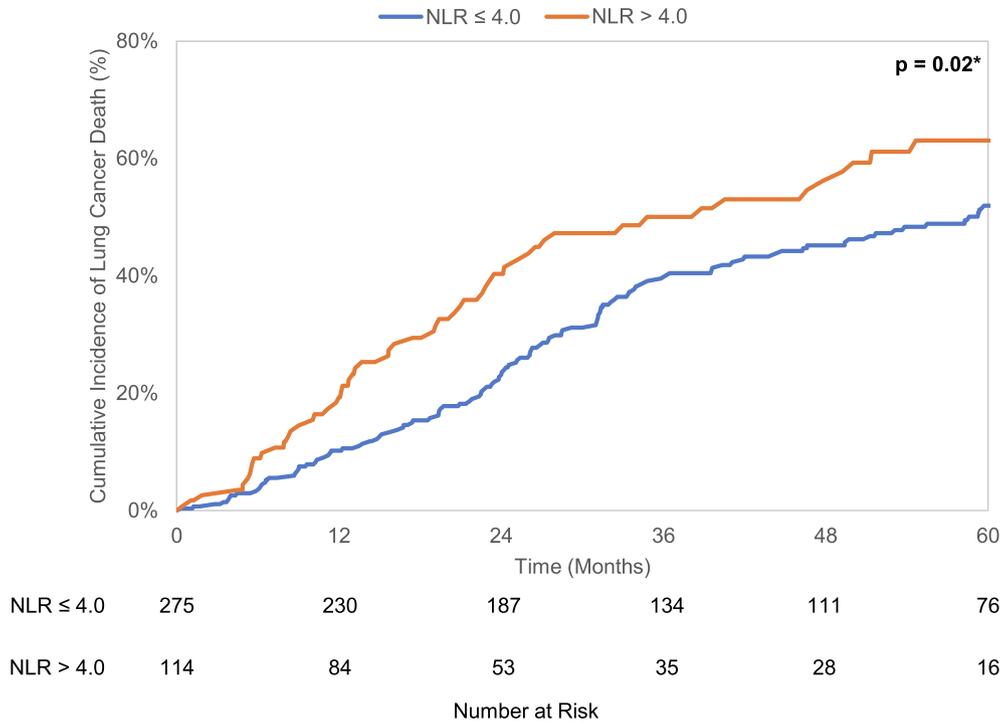


Fig. 2 (continued)

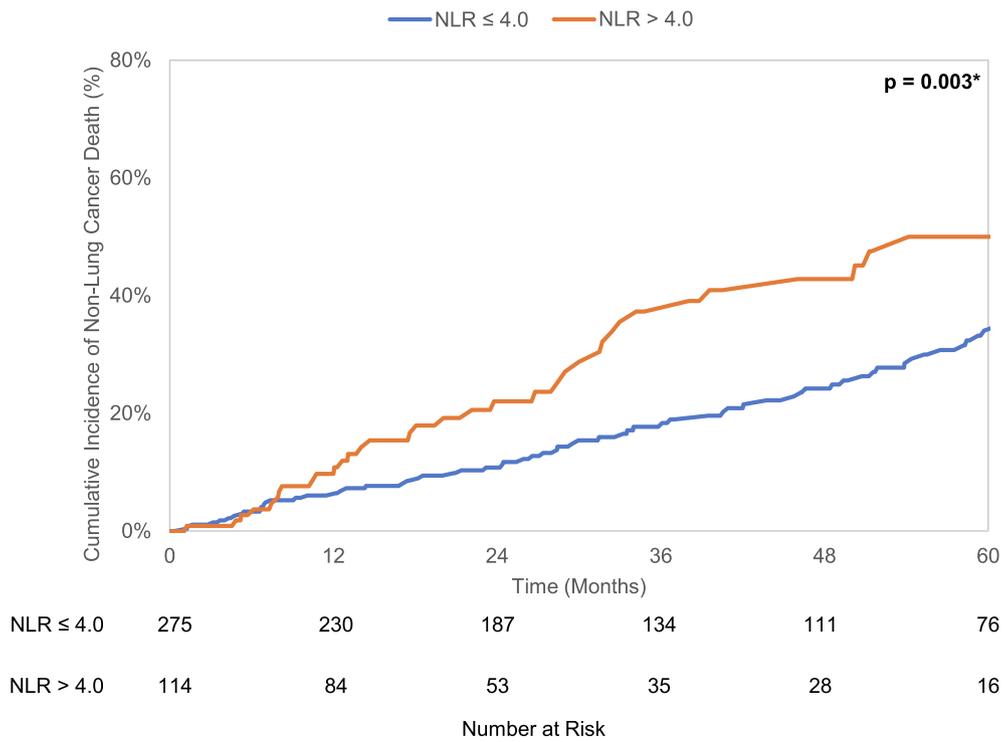


Fig. 2 (continued)

Table 2
Multivariable a priori regressions on overall survival (OS), lung cancer-specific survival (LCS), non-lung cancer survival (NCS) in the overall cohort.

Variable	OS		LCS		NCS	
	HR (95% CI)	P value	HR (95% CI)	P value	HR (95% CI)	P value
NLR						
≤ 4.0	1.0		1.0		1.0	
> 4.0	1.44 (1.12–1.86)	0.01*	1.32 (0.95–1.82)	0.09	1.68 (1.11–2.53)	0.01*
Age	1.03 (1.02–1.05)	0.0001*	1.03 (1.02–1.05)	0.0004*	1.03 (1.01–1.06)	0.01*
Gender						
Male	1.0		1.0		1.0	
Female	0.98 (0.59–1.62)	0.93	1.11 (0.53–2.32)	0.78	0.80 (0.31–2.08)	0.65
Charlson Score						
0	1.0		1.0		1.0	
1	1.35 (0.98–1.86)	0.07	1.56 (1.03–2.36)	0.04*	1.05 (0.62–1.78)	0.85
2+	1.56 (1.13–2.16)	0.01*	1.50 (0.97–2.32)	0.07	1.68 (1.02–2.78)	0.04*
Race						
White	1.0		1.0		1.0	
Black	0.73 (0.51–1.04)	0.08	0.84 (0.56–1.28)	0.43	0.57 (0.30–1.07)	0.08
Smoker at Diagnosis						
Yes	1.0		1.0		1.0	
No	0.80 (0.63–1.02)	0.07	0.89 (0.66–1.20)	0.46	0.70 (0.45–1.02)	0.06
Histology						
Adenocarcinoma	1.0		1.0		1.0	
Squamous	1.04 (0.80–1.33)	0.79	1.23 (0.89–1.70)	0.20	0.78 (0.52–1.18)	0.24
Other/unknown	0.86 (0.59–1.24)	0.41	1.05 (0.66–1.67)	0.84	0.61 (0.33–1.15)	0.13
T stage						
< 2 cm (T1a)	1.0		1.0		1.0	
2–3 cm (T1b)	1.32 (1.02–1.72)	0.04*	1.41 (1.01–1.96)	0.04*	1.20 (0.77–1.87)	0.42
3–5 cm (T2a)	1.25 (0.91–1.70)	0.16	1.23 (0.83–1.82)	0.31	1.29 (0.79–2.11)	0.31
Total BED	1.00 (0.99–1.00)	0.49	1.00 (0.99–1.00)	0.32	1.00 (0.99–1.01)	0.92

Table 3
Multivariable a priori (including pulmonary function tests) regressions on overall survival (OS), lung cancer-specific survival (LCS), non-lung cancer survival (NCS) in the subset cohort.

Variable	OS		LCS		NCS	
	HR (95% CI)	P value	HR (95% CI)	P value	HR (95% CI)	P value
NLR						
≤ 4.0	1.0		1.0		1.0	
> 4.0	1.51 (1.07–2.14)	0.02*	1.22 (0.77–1.93)	0.39	2.18 (1.24–3.84)	0.01*
Age	1.05 (1.03–1.07)	<0.0001*	1.05 (1.02–1.08)	0.0008*	1.05 (1.01–1.10)	0.03*
Gender						
Male	1.0		1.0		1.0	
Female	1.22 (0.67–2.22)	0.53	1.04 (0.37–1.93)	0.94	1.55 (0.53–4.50)	0.42
Charlson Score						
0	1.0		1.0		1.0	
1	1.41 (0.90–2.22)	0.13	1.45 (0.83–2.55)	0.20	1.39 (0.64–3.01)	0.40
2+	1.65 (1.05–2.59)	0.03*	1.40 (0.77–2.53)	0.27	2.30 (1.11–4.77)	0.03*
Race						
White	1.0		1.0		1.0	
Black	0.78 (0.49–1.24)	0.30	0.99 (0.58–1.68)	0.96	0.47 (0.20–1.12)	0.09
Smoker at Diagnosis						
Yes	1.0		1.0		1.0	
No	0.71 (0.50–0.99)	0.04*	0.76 (0.50–1.15)	0.19	0.63 (0.35–1.13)	0.12
Histology						
Adenocarcinoma	1.0		1.0		1.0	
Squamous	1.18 (0.83–1.68)	0.35	1.36 (0.85–2.15)	0.20	0.96 (0.55–1.70)	0.90
Other/unknown	0.85 (0.47–1.52)	0.58	1.04 (0.51–2.10)	0.92	0.62 (0.20–1.92)	0.41
T stage						
< 2 cm (T1a)	1.0		1.0		1.0	
2–3 cm (T1b)	1.31 (0.91–1.89)	0.14	1.40 (0.88–2.23)	0.15	1.15 (0.61–2.15)	0.67
3–5 cm (T2a)	1.21 (0.80–1.81)	0.37	1.25 (0.73–2.13)	0.42	1.15 (0.58–2.26)	0.69
Total BED	1.00 (0.99–1.00)	0.26	1.00 (0.99–1.00)	0.16	1.00 (0.99–1.01)	0.96
Pretreatment FEV1						
< 30%	1.0		1.0		1.0	
31–50%	0.53 (0.30–0.94)	0.03*	0.63 (0.29–1.33)	0.23	0.37 (0.15–0.93)	0.04*
51–80%	0.57 (0.33–0.99)	0.04*	0.60 (0.29–1.27)	0.18	0.51 (0.22–1.20)	0.12
> 80%	0.58 (0.30–1.11)	0.10	0.62 (0.27–1.44)	0.27	0.50 (0.18–1.38)	0.18

this subset cohort and the subsequent multivariable analysis on each survival endpoint.

3.4. Validation of previously proposed cutoff

Sebastian et al. proposed a pretreatment NLR cutoff of 3.6 identified by using the median pretreatment NLR value of their 156 patient cohort with localized NSCLC treated with SBRT [22]. They identified pretreatment NLR > 3.6 to be independently associated with overall mortality. Therefore we replicated our previous analysis using this previously proposed cutoff.

For our overall cohort of 389 patients, multivariable analysis revealed pretreatment NLR > 3.6 to not be associated with OS (HR 1.25, 95% CI 0.99–1.58, $p = 0.06$) or LCS (HR 1.12, 95% CI 0.84–1.51, $p = 0.44$). But it was associated with NCS (HR 1.52, 95% CI 1.03–2.24, $p = 0.04$). In our subset cohort of 229 patients with PFT data, multivariable analysis (incorporating PFTs) revealed pretreatment NLR > 3.6 to not be associated with OS (HR 1.31, 95% CI 0.65–1.47, $p = 0.91$), LCS (HR 0.98, 95% CI 0.65–1.47, $p = 0.91$), or NCS (HR 1.59, 95% CI 0.98–2.58, $p = 0.08$).

4. Discussion

In this hypothesis-generating study of patients with stage I NSCLC treated with SBRT in the VA system, we identified an optimal pretreatment NLR cutoff point of 4.0 and observed NLR to be independently associated with OS. However, this association was driven by the relationship between NLR and NCS rather than by the relationship between NLR and LCS. Subset analysis (incorporating PFT data which pose as a notable confounder) further evidenced our findings by demonstrating elevated NLR to have more pronounced associations with inferior OS and NCS. Although NLR's association with LCS trended towards significance ($p = 0.09$) in the overall cohort, it was notably insignificant ($p = 0.39$) in the subset analysis. Our analysis suggests that NLR may not have potential as a marker of lung cancer-specific survival in this setting. However, it does suggest that NLR has strong potential as a marker for competing risk mortality in this population.

Previous studies have explored the associations between pre-SBRT NLR and survival outcomes but have presented a mix of results. Of these studies that have proposed a pre-SBRT NLR cutoff point, Cannon et al. (cutoff 2.98) and Luo et al. (cutoff 2.06) utilized a ROC approach while Sebastian et al. (cutoff 3.6) chose the cohort's median as a cutoff point [20–22]. All of these cutoff points were evaluated to have no association with any survival outcomes in our dataset, but only data for the one with the largest cohort and strongest evidence (Sebastian et al.'s 3.6) was detailed in this paper. In the context of these previously proposed cutoff points yielding no associations in our dataset and given the time dependent nature of survival outcomes, we believe our study is unique in this domain to use statistically appropriate, time-dependent cutoff methods. Furthermore, our study is the largest cohort thus far with nearly 400 patients. Additionally, given that NLR has been associated with PFTs and that PFTs have been demonstrated to be an important factor in both lung function decline and NSCLC, we are able to account for a potentially notable confounding variable while most previous studies have not been able to [19,33]. Finally, unlike most previous studies, we are able to delineate NLR's impact on LCS and NCS individually. This enables a better understanding of the clinical utility NLR offers in this disease and population.

A large body of literature exists on the pathophysiology and interpretation of NLR in both oncologic and non-oncologic disease states. In various oncologic states, lymphocytes are hypothesized to primarily aid in the host response against cancer while neu-

trophils are hypothesized to primarily aid in the carcinogenesis response against the host [34–35]. Similar immunologic principles and theories have implicated NLR to be associated with outcomes in several non-oncologic disease states [16–17]. Therefore a major challenge to the use of NLR in the clinical setting is this hurdle of understanding what prognostic value it holds for each patient population in consideration. For localized NSCLC patients treated with SBRT, we hypothesize that NLR is predictive of competing mortality risk given our data. Giuliani et al. reported similar findings of no association with lung cancer-specific outcomes when analyzing NLR as a continuous variable [36]. This suggests an underlying propensity for mortality in the elevated NLR cohort. However, this propensity cannot be simply explained by a difference in classic variables of age, Charlson scores, etc. in the elevated NLR cohort as Table 1 demonstrates the similarity in host and tumor characteristics for both cohorts. This is in line with the unmet need for additional risk stratification in localized NSCLC treated with SBRT as most of these patients possess unfavorable, classic risk stratification variables that typically preclude them from surgery.

Although our study overcomes important limitations of prior studies in this realm, there are still notable limitations present. Alongside the inherent drawbacks of a retrospective database study, detailed records of pro-inflammatory states (e.g. rheumatologic disorders, infections, etc.) and specific medications (e.g. corticosteroids, immunomodulators, etc.) that affect a patient's immunologic state, and therefore likely their NLR, are not noted and raise concern for unaccounted confounding. Additionally, there is research demonstrating NLR to have differential prognostic capacity in different age and race populations [37]. Of note, Black patients in our cohort had higher NLR which possibly indicates an elevated, baseline inflammatory state in this population but this would need further research [38]. Although we hope the inclusion of these variables in multivariable models would account for this difference, the generalizability of our study to all populations consequently might be limited.

5. Conclusions

In summary, we identified an optimal pretreatment NLR cutoff of 4.0 and observed NLR to be independently associated with overall survival in patients with localized NSCLC who underwent SBRT. After accounting for clinical variables (notably including baseline lung function), this association was found to be driven by the relationship between NLR and non-lung cancer survival rather than by the statistically insignificant relationship between NLR and lung cancer-specific survival. Although there is a need for markers predictive of cancer-specific outcomes in this population, our analysis does not suggest NLR to have such potential. Instead, our hypothesis-generating study demonstrates strong potential for NLR as a marker for competing risk mortality in this setting. Given this population to typically be frailer and rife with comorbidities, this simple and cost-effective lab value with our proposed cutoff has potential to fill a need for further risk stratification and clinical decision making for these patients. Further studies are needed to validate these findings and investigate pretreatment NLR as a clinical tool in patients with localized NSCLC treated with SBRT.

Declaration of Competing Interest

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

Appendix A. Supplementary data

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.ctro.2021.03.010>.

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